

# Modeling Healthcare Expenditures

Anirban Basu, PhD  
The University of Chicago &  
The NBER

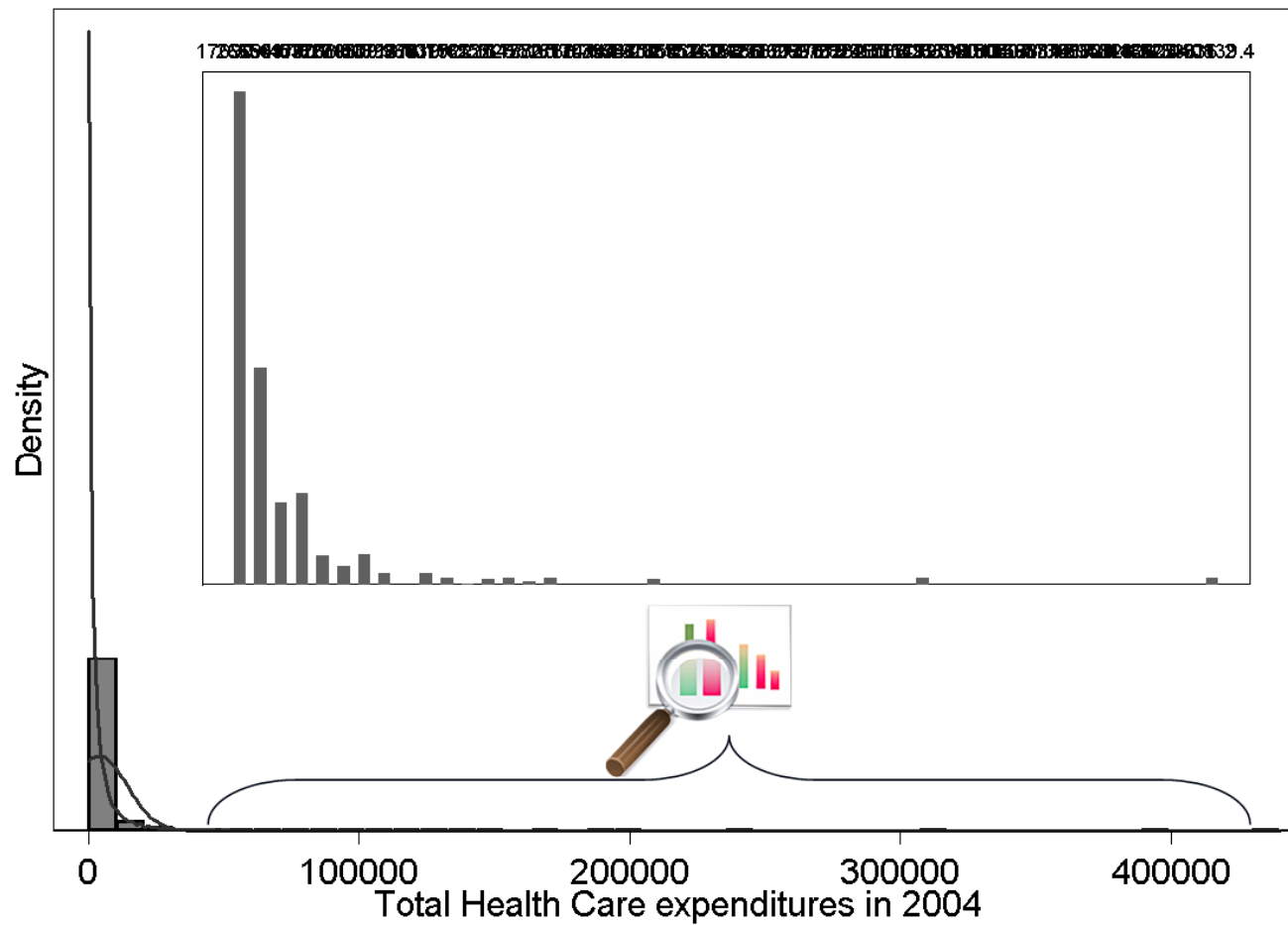
HERC CyberSeminar  
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# Goals

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- Review parametric/semi-parametric statistical models for health care expenditures
- Intent to study how mean costs responds to shifts in covariate levels
- Will assume “selection of observable” throughout
- Main focus is on bias, will look at some efficiency issues too.

# Healthcare costs in the U.S.



# Healthcare costs in the U.S.

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- A substantial fraction of people with no health care costs
- Handful of people with enormous yearly expenditures
- Often top 1% accounts for 25% of the total mean, sometimes the top 10% accounts for half of all costs
- Any attempt to study how the mean of such a distribution responds to a covariate would require careful attention to how different parts of the distribution moves in response to the changes in levels of that covariate

# Alternative Methods

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- **Classes of models:**
  - Single equation models
  - Multiple equation models
    - Two-part models
    - Multi-part models
    - Conditional density estimation
  - Mixture models
  - Propensity scores
  - Doubly robust models

# Outline

- Single-equation models: relate mean of costs to covariates using a single functional form
  - Linear models
  - Transformation models
    - Log-OLS
    - Box-Cox
  - Generalized Linear models
- Application to datasets of ~1000 to 100,000
- Discuss goodness-of-fit tests

# Notation

- $Y = \text{Costs}$
- $X = \text{Covariates} = (X_T, X_{-T})$
- $X_T = \text{A binary treatment variable}$
- $\beta = \text{regression coefficients}$
- $\mu = E(Y); \mu(\mathbf{x}) = E(Y | X=\mathbf{x})$
- $\text{Marginal effect: } d\mu(X)/dX$
- $\text{Incremental effect: } \mu(X=1) - \mu(X=0)$

## An illustrative health-economic study

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- Compare the 2-year cost of care under aggressive treatments (radiation or surgery) versus watchful waiting following diagnosis of prostate cancer
- Use SEER-Medicare linked database
- Data: diagnosis between 1995-2002
- Follow-up through 2004.
- No censoring issues
- Ignore issues with death



# The parameters of Interest

- Main parameters of interest:

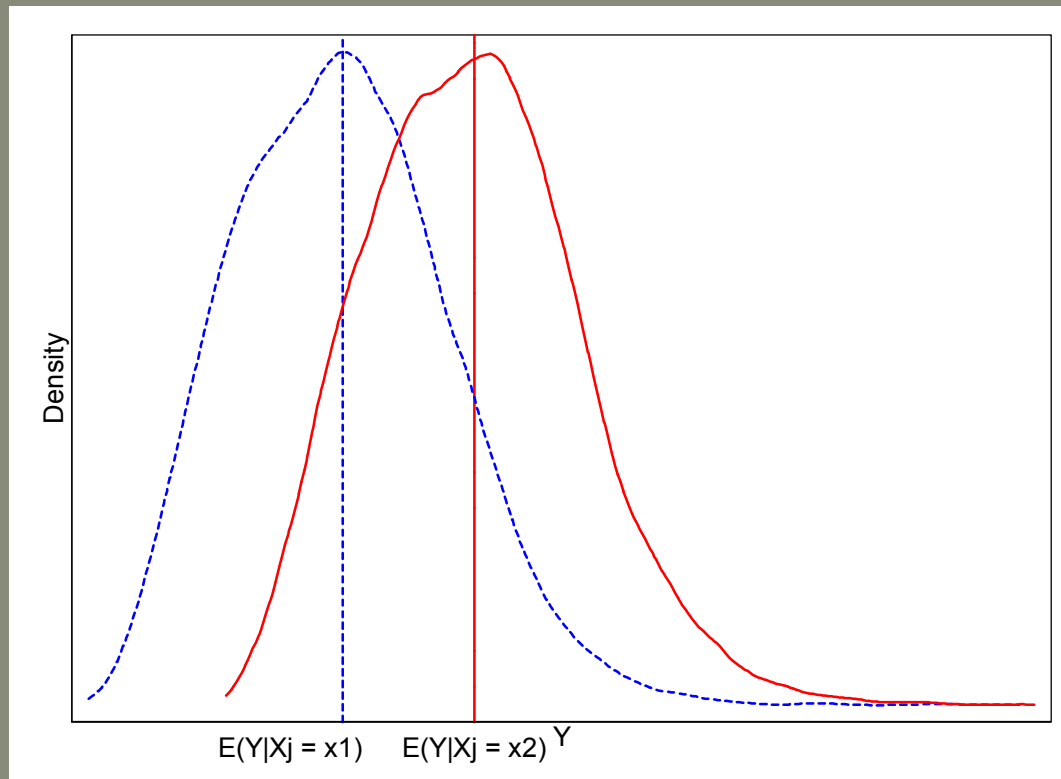
- $\mu(\mathbf{x}) = E(y | \mathbf{x})$   $y =$  Expenditures
- Incremental effect of aggressive treatment vs WW on total expenditures ( $\Delta_T$ )
- Avg. Costs for TRT – Avg. Costs for WW

$$\Delta_T = E_{(X_{-T}|X_T=1)} \left\{ \mu(X_T = 1, X_{-T}) - \mu(X_T = 0, X_{-T}) \right\}$$

- The interpretation of regression coefficients  $\beta$  is of secondary concern.

# Single equation models: Linear model

- $Y = X\beta + \varepsilon$



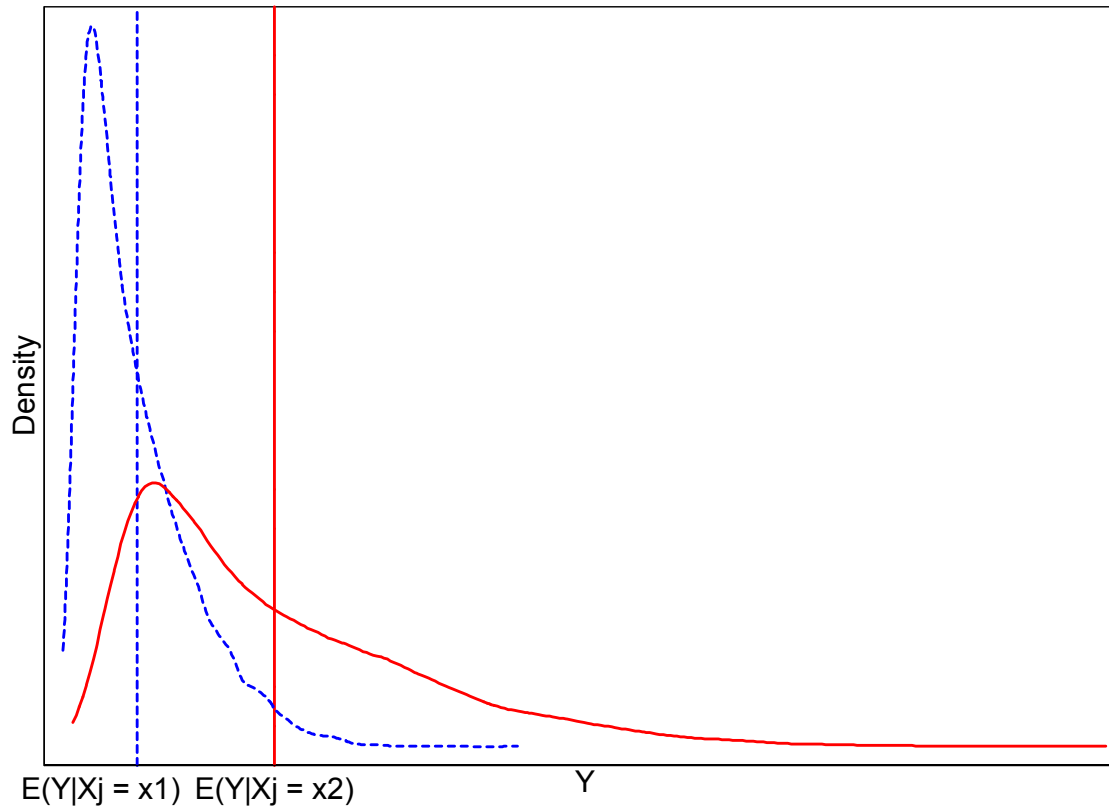
# Single equation models: Linear model

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- $Y = X\beta + \varepsilon$
- Assumes parallel shift for the entire distribution across levels of any one X, and for all X's
- Biased
  - the non-linearity in the response,
  - the instability caused by skewness and kurtosis,
- Inefficient
  - common failure to deal with the heteroscedasticity (variance increases with the mean)

# Single equation models: Transformation models – log OLS

- $\ln(Y) = X\beta + \varepsilon$



# Single equation models: Transformation models – log OLS

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- $\ln(Y) = X\beta + \varepsilon$
- Assumes proportional shift for the entire distribution across levels of any one X, and for all X's
- Advantages:
  - Usually overcomes skewness issues
  - Allows for additive effects in the log-scale, which translates to non-additive effects in the raw scale
  - Reduces problems with heteroscedasticity & kurtosis

# Transformation models

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- **Scale of estimation  $\neq$  scale of interest**
  - If modeling \$, results in log-\$
- **Model for geometric mean, not arithmetic mean**
  - Models  $E\{\log(y)\}$  and not  $\log(E\{y\})$

# Single equation models: Transformation models – log OLS

- $\ln(Y) = X\beta + \varepsilon$
- Disadvantages:
  - Creates a model for log-\$, not \$.
  - $E(\ln Y) \neq \ln(E(Y))$
  - Require retransformation,  $E(Y) = \exp(X\beta + \varepsilon)$   
 $= \exp(X\beta) * s$
  - Duan's smearing estimator provides an estimate for  $s$
  - $s\text{-hat} = \text{Mean}\{\exp(\ln Y - X \beta\text{-hat})\}$

# Single equation models: Transformation models – log OLS

- $\ln(Y) = X\beta + \varepsilon$
- Disadvantages:
  - $E(\ln Y) = X\beta$
  - $\ln E(Y) = X\beta + \ln\{s\}$
  - If  $s = s(X)$ ,  
$$d\ln E(Y)/dX = \beta + d\ln\{s(X)\}/dX$$
  - So, common interpretation of  $\beta$  no longer applies
  - If  $\varepsilon \sim \text{Normal}(0, \sigma^2)$ ,  $\ln(s) = 0.5 \sigma^2$
  - With log-scale heteroscedasticity,  
$$\text{Bias} = 0.5 d\sigma^2(X)/dX$$



# Single equation models: Transformation models – log OLS

Table 1

Univariate statistics on log medical expenditures for users\* health insurance experiment

Plan	Mean*	Variance**	Skewness	Kurtosis	Percentile					
					10%	25%	50%	75%	90%	95%
Free	6.349	2.083	0.202	3.385	4.541	5.469	6.289	7.166	8.290	8.976
25%	6.128	2.133	0.456	3.539	4.393	5.191	6.004	6.907	8.117	8.849
50%	6.039	2.003	0.513	4.067	4.265	5.134	5.969	6.766	7.906	8.628
95%	5.998	2.343	0.464	3.226	4.158	4.948	5.873	6.862	8.120	8.892
Individual deductible	6.153	2.384	0.305	2.993	4.237	5.068	6.039	7.093	8.322	8.876

\* In 1995 US\$, adjusted using the medical care component of the Consumer Price Index.

\*\* Plan differences are significant at  $F < 0.001$  for both mean and variance, by  $F$ -tests. Tests for variance based on Park Test.

Manning JHE, 1998

% change in costs between Free & 95%:

From log-OLS regression:  $6.349 - 5.998 = 0.351$  (or 35.1%)

Truth (assuming log-normality) =  $6.349 - 5.998 + 0.5 \cdot (2.083 - 2.343)$   
=  $0.221$  (or 22.1%)

# Single equation models: Transformation models – log OLS

- Plan specific smearing estimate can address bias
- Difficult to come up with appropriate smearing estimator when there are multiple X's that lead to heteroscedasticity
- Park Test – test to determine heteroscedasticity in log-scale
  - Use squared log-scale residuals as an estimator for variance
  - Study if X's predict this estimated variance
  - `reg lny x, robust`
  - `predict r, res`
  - `gen r2= r^2`
  - `glm r2 x, link(log) family(gamma) robust`

**Let's try out in Stata**

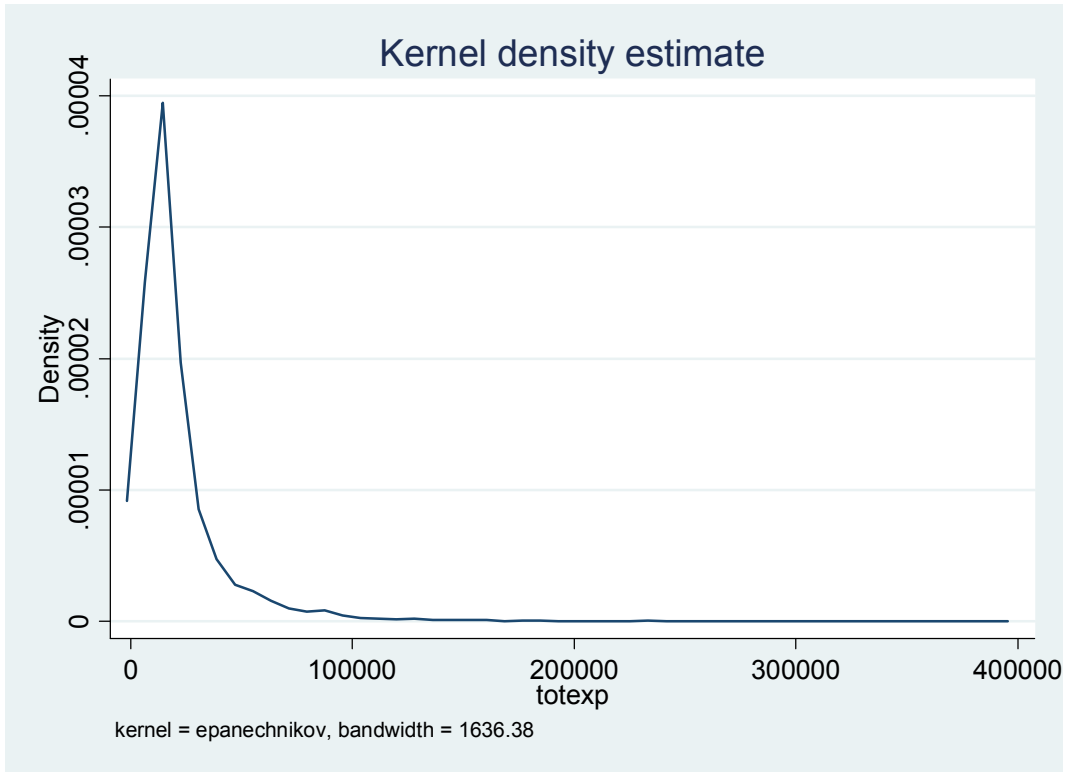
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MODELLING HEALTHCARE EXPENDITURES

```
-----
. use week2_data1, clear
. summ totexp,d
```

totexp				
-----				
	Percentiles	Smallest		
1%	78.32	2.78		
5%	907.92	3		
10%	2702.61	3.63	Obs	7129
25%	8912.18	4.18	Sum of Wgt.	7129
50%	14536.54		Mean	20265.17
		Largest	Std. Dev.	22937.29
75%	23375.21	279487.5		
90%	41139.5	307933.6	Variance	5.26e+08
95%	59268.49	316789.1	Skewness	4.537538
99%	112434.3	393908.4	Kurtosis	39.71721

```
. kdensity totexp, saving(grp1, replace)
```



```
. bysort aggr_trt: summ totexp, d
```

```
-> aggr_trt = 0
```

```
-----
```

Percentiles		Smallest		
1%	50.8	3.63		
5%	648.2	4.18		
10%	1423.23	4.57	Obs	2219
25%	4738.09	5	Sum of Wgt.	2219
50%	12317.91		Mean	19201.76
		Largest	Std. Dev.	24102.25
75%	23087.59	166289		
90%	43853.15	204529.9	Variance	5.81e+08
95%	65831.95	236276.2	Skewness	3.529187
99%	114372.3	316789.1	Kurtosis	24.20243

```
-----
```

-> aggr\_trt = 1

totexp

---

Percentiles		Smallest		
1%	99.18	2.78		
5%	1172.94	3		
10%	5141.575	4.37	Obs	4910
25%	10071	4.55	Sum of Wgt.	4910
50%	15343.07		Mean	20745.77
		Largest	Std. Dev.	22376.88
75%	23522.77	263867.3		
90%	39927.79	279487.5	Variance	5.01e+08
95%	57092.22	307933.6	Skewness	5.114373
99%	110871.9	393908.4	Kurtosis	48.88715

. di (20745.77 - 19201.76) / 19201.76  
.08040982

```
. /* OLS */
```

```
. reg totexp aggr_trt, robust
```

Linear regression

```
Number of obs =      7129  
F( 1, 7127) =      6.55  
Prob > F      =     0.0105  
R-squared     =     0.0010  
Root MSE     =     22928
```

totexp	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
<b>aggr_trt</b>	<b>1544.008</b>	<b>603.106</b>	<b>2.56</b>	<b>0.010</b>	<b>361.7408</b>	<b>2726.274</b>
_cons	19201.76	511.6135	37.53	0.000	18198.85	20204.67



```
. reg totemp aggr_trt agec white black single married well
mod stage1 stage2 lnc, robust
```

totemp	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
<b>aggr_trt</b>	<b>3490.596</b>	<b>636.6122</b>	<b>5.48</b>	<b>0.000</b>	<b>2242.646</b>	<b>4738.545</b>
agec	205.2538	52.4853	3.91	0.000	102.367	308.1406
white	1269.85	1027.947	1.24	0.217	-745.2315	3284.931
black	3706.038	1426.357	2.60	0.009	909.9544	6502.122
single	946.6497	1241.496	0.76	0.446	-1487.051	3380.351
married	-1048.995	696.5203	-1.51	0.132	-2414.382	316.3918
well	-4717.357	1261.522	-3.74	0.000	-7190.314	-2244.4
mod	-2995.753	745.5633	-4.02	0.000	-4457.279	-1534.227
stage1	-304.7167	602.6593	-0.51	0.613	-1486.108	876.6747
stage2	-1803.271	804.8815	-2.24	0.025	-3381.078	-225.4639
lnc	7871.044	502.2525	15.67	0.000	6886.48	8855.608
_cons	15268.15	1349.965	11.31	0.000	12621.82	17914.48

```

. /* Log-OLS */
. capture drop lny
. gen lny = ln(totexp)
. reg lny aggr_trt, robust

```

```

Linear regression                               Number of obs =      7129
                                                F( 1, 7127) =      93.48
                                                Prob > F          =      0.0000
                                                R-squared         =      0.0150
                                                Root MSE         =      1.2894

```

```

-----

```

lny	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
aggr_trt	.3431855	.035495	9.67	0.000	.2736047	.4127662
_cons	9.147605	.0310711	294.41	0.000	9.086697	9.208514

```

-----

```

```

. predict xb, xb
. gen mu = exp(xb)
. summ mu totexp

```

```

-----

```

Variable	Obs	Mean	Std. Dev.	Min	Max
mu	7129	12040.35	1780.561	9391.919	13237.27
totexp	7129	20265.17	22937.29	2.78	393908.4

```

-----

```

```

. **** Overall smearing
. drop mu
. gen smr = exp(lny - xb)
. summ smr

```

Variable	Obs	Mean	Std. Dev.	Min	Max
smr	7129	1.715782	2.016464	.00021	33.72997

```

. gen smear = r(mean)
. gen mu = exp(xb)*smear
. summ mu totexp

```

Variable	Obs	Mean	Std. Dev.	Min	Max
mu	7129	20658.62	3055.054	16114.49	22712.27
totexp	7129	20265.17	22937.29	2.78	393908.4

```
. bysort aggr_trt: summ mu totexp
```

```
-----  
-> aggr_trt = 0
```

Variable	Obs	Mean	Std. Dev.	Min	Max
mu	2219	16114.49	0	16114.49	16114.49
totexp	2219	19201.76	24102.25	3.63	316789.1

```
-----  
-> aggr_trt = 1
```

Variable	Obs	Mean	Std. Dev.	Min	Max
mu	4910	22712.27	0	22712.27	22712.27
totexp	4910	20745.77	22376.88	2.78	393908.4

```
/* Notice difference in log scale Variance */
```

```
. bysort aggr_trt: summ lny
```

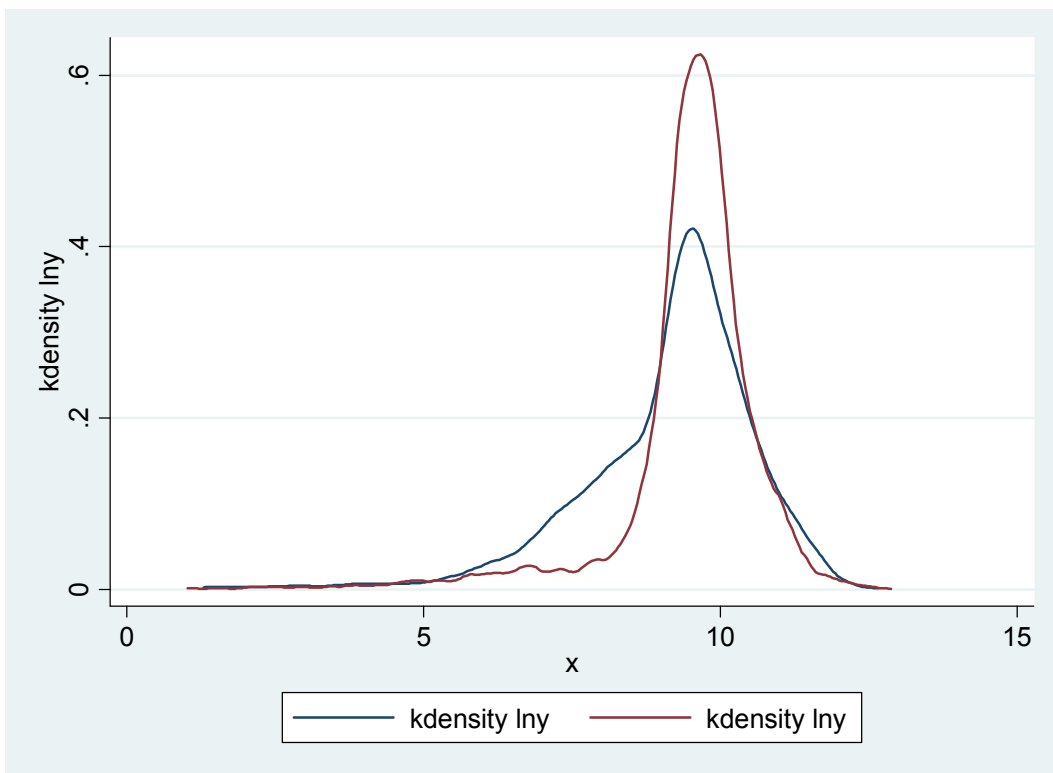
```
-----  
-> aggr_trt = 0
```

Variable	Obs	Mean	Std. Dev.	Min	Max
lny	2219	9.147605	1.463769	1.289233	12.66599

```
-----  
-> aggr_trt = 1
```

lny	4910	9.490791	1.202411	1.022451	12.88387
-----	------	----------	----------	----------	----------

```
. gr twoway (kdensity lny if aggr_trt==0) (kdensity lny if aggr_trt==1), saving
```



```
. **** Trt specific smearing
```

```
. drop mu
```

```
. summ smr if aggr_trt ==0
```

Variable	Obs	Mean	Std. Dev.	Min	Max
smr	2219	2.044498	2.566275	.0003865	33.72997

```
. gen smear0=r(mean)
```

```
. summ smr if aggr_trt ==1
```

Variable	Obs	Mean	Std. Dev.	Min	Max
smr	4910	1.567224	1.690446	.00021	29.75754

```
. gen smear1=r(mean)
```

```
. gen mu =exp(xb)*smear0 if aggr_trt ==0
```

```
. replace mu =exp(xb)*smear1 if aggr_trt ==1
```

```
. summ mu totexp
```

Variable	Obs	Mean	Std. Dev.	Min	Max
mu	7129	20265.17	714.9419	19201.76	20745.77
totexp	7129	20265.17	22937.29	2.78	393908.4

```
. bysort aggr_trt: summ mu totexp
```

```
----->  
aggr_trt = 0
```

Variable	Obs	Mean	Std. Dev.	Min	Max
mu	2219	19201.76	0	19201.76	19201.76
totexp	2219	19201.76	24102.25	3.63	316789.1

```
-----  
-> aggr_trt = 1
```

Variable	Obs	Mean	Std. Dev.	Min	Max
mu	4910	20745.77	0	20745.77	20745.77
totexp	4910	20745.77	22376.88	2.78	393908.4

```

. **** ln-OLS regression
. drop mu xb
. reg lny aggr_trt agec white black single married well mod
stage1 stage2 lnc, robust

```

lny	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
<b>aggr_trt</b>	<b>.4609178</b>	<b>.0372838</b>	<b>12.36</b>	<b>0.000</b>	<b>.3878303</b>	<b>.5340052</b>
agec	.0166473	.0027592	6.03	0.000	.0112385	.0220561
white	.2875066	.0703636	4.09	0.000	.1495731	.4254401
black	.239347	.0876897	2.73	0.006	.0674492	.4112448
single	.0702465	.066109	1.06	0.288	-.0593468	.1998397
married	-.0016466	.0381762	-0.04	0.966	-.0764833	.07319
well	-.3312179	.0710317	-4.66	0.000	-.4704613	-.1919746
mod	-.1428405	.0378877	-3.77	0.000	-.2171118	-.0685693
stage1	-.0147308	.0337002	-0.44	0.662	-.0807933	.0513317
stage2	-.0830997	.0480434	-1.73	0.084	-.177279	.0110797
lnc	.5280045	.0233166	22.65	0.000	.4822971	.5737119
_cons	8.627953	.0881616	97.87	0.000	8.455131	8.800776



```

. ** Park test - detect heteroscedasticity in log-scale
. predict xb, xb
. gen res2 = (lny- xb)^2
. glm res2 aggr_trt agec white black single married well
mod stage1 stage2 lnc, family(gamma) link(log) robust

```

---

res2	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
aggr_trt	-.5202051	.0652508	-7.97	0.000	-.6480943	-.3923159
agec	-.0253315	.0054215	-4.67	0.000	-.0359575	-.0147055
white	-.4874537	.1051322	-4.64	0.000	-.6935091	-.2813983
black	-.1256915	.1388546	-0.91	0.365	-.3978414	.1464585
single	-.0352572	.1358728	-0.26	0.795	-.3015629	.2310485
married	-.1367799	.073717	-1.86	0.064	-.2812624	.0077027
well	.100519	.1393567	0.72	0.471	-.172615	.373653
mod	-.0869734	.0785356	-1.11	0.268	-.2409005	.0669536
stage1	-.057802	.0720468	-0.80	0.422	-.1990111	.0834072
stage2	-.0547015	.1094026	-0.50	0.617	-.2691267	.1597236
lnc	-.5347174	.046929	-11.39	0.000	-.6266966	-.4427382
_cons	1.597288	.1376163	11.61	0.000	1.327565	1.867011

---

```

. test

```

```

      chi2( 11) = 283.67
Prob > chi2 = 0.0000

```

# Single equation models: Transformation models – Box-Cox

$$(Y^\lambda - 1)/\lambda = X\beta + \varepsilon$$

- $\lambda \rightarrow 0$ , Box-Cox  $\rightarrow$  log-OLS
- If  $\lambda$  is known apriori, can use Duan's smearing estimator
- If  $\lambda$  is estimated, use different estimator

$$E(Y | X) = (X\beta\lambda + 1)^{1/\lambda} \left\{ 1 + \frac{\sigma^2(1-\lambda)}{2(1 + X\beta\lambda)^2} \right\}$$

- All concerns about retransformation persist

# Single equation models: Generalized Linear Models

---

- $g\{E(Y | X)\} = X\beta$ ,  $g\{.\}$  is a link function
- No retransformation problem as link function applies to  $E(Y)$  and NOT  $Y$ .
- Estimation of GLM
  - Parametric – FIML
  - Semi-parametric – Quasi-likelihood
- FIML:
  - Gamma, Exponential, Inverse Gaussian, Poisson, Neg. Binomial
  - `glm y x, link(log) family(gamma) robust`

# Single equation models: Generalized Linear Models

- Quasi-likelihood approach

1. Relate  $\mu(\mathbf{X})$  to the linear predictor  $\mathbf{X}\beta$  using a link function.  $g(\mu(\mathbf{X})) = \mathbf{X}\beta$
2.  $V(Y) \sim h(\mu)$ ; e.g.  $V(Y) = \phi\mu^2 \Rightarrow$  Gamma Variance
3. Does not make any distributional assumption beyond 1st and 2nd moments
4. Estimate coefficients by solving quasi-score equations

$$\sum_{i=1}^N \mathbf{G}_{\beta_j}^i = \sum_{i=1}^N (y_i - \mu_i) V_i^{-1} (\partial \mu_i / \partial \beta_j) = 0$$

- `glm y x, link(log) fam(gamma) robust irls`

# Single equation models: Generalized Linear Models

## ● **Example: Gamma with log link**

- $\ln(\mu) = X\beta$       or       $\mu = \exp(X\beta)$
- $V = \theta_1 \mu^2$
- If mean and variance functions are correctly specified,  $\beta_1$  provides a consistent estimator of  $\partial \ln(E\{y | X\}) / \partial x_1$
- If the empirical distribution of  $Y$  is truly Gamma, then FIML will generate the MVUE estimator for  $\beta$
- If the empirical distribution of  $Y$  deviates from Gamma, but mean and variance functions are appropriately specified, quasi-likelihood will be a more robust estimator.

To Stata..

---

```
. /* Generalized Linear Models */
. glm totexp aggr_trt, link(log) family(gamma) robust
```

Iteration 3: log pseudolikelihood = -77820.337

Generalized linear models		No. of obs	=	7129
Optimization	: ML	Residual df	=	7127
		Scale parameter	=	1.291687
Deviance	= 7586.035219	1/df) Deviance	=	1.064408
Pearson	= 9205.852281	1/df) Pearson	=	1.291687

Variance function:  $V(u) = u^2$  [Gamma]  
 Link function :  $g(u) = \ln(u)$  [Log]

	AIC	=	21.83261
Log pseudolikelihood = -77820.3371	BIC	=	-55644.18

```
-----
```

totexp	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]
aggr_trt	.0773403	.0307692	2.51	0.012	.0170338 .1376469
_cons	9.862757	.0266422	370.19	0.000	9.810539 9.914975

```
-----
```

```
. predict mu, mu
. summ mu totexp
```

Variable	Obs	Mean	Std. Dev.	Min	Max
mu	7129	20265.17	714.9419	19201.76	20745.77
totexp	7129	20265.17	22937.29	2.78	393908.4

```
. bysort aggr_trt: summ mu totexp
```

```
-> aggr_trt = 0
```

Variable	Obs	Mean	Std. Dev.	Min	Max
mu	2219	19201.76	0	19201.76	19201.76
totexp	2219	19201.76	24102.25	3.63	316789.1

```
-> aggr_trt = 1
```

Variable	Obs	Mean	Std. Dev.	Min	Max
mu	4910	20745.77	0	20745.77	20745.77
totexp	4910	20745.77	22376.88	2.78	393908.4



. \*\* GAMMA FIML

. glm totexp aggr\_trt agec white black single married  
well mod stage1 stage2 lnc, link(log) family(gamma) robust

Iteration 3: log pseudolikelihood = -77563.068

---

	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
<b>aggr_trt</b>	<b>.1794719</b>	<b>.0320742</b>	<b>5.60</b>	<b>0.000</b>	<b>.1166076</b>	<b>.2423362</b>
agec	.010508	.0025216	4.17	0.000	.0055658	.0154501
white	.0854313	.0532399	1.60	0.109	-.0189169	.1897796
black	.1795773	.067756	2.65	0.008	.0467779	.3123767
single	.0333528	.0559583	0.60	0.551	-.0763235	.143029
married	-.0542054	.0347019	-1.56	0.118	-.1222198	.013809
well	-.2484179	.0642366	-3.87	0.000	-.3743194	-.1225165
mod	-.1447476	.0333845	-4.34	0.000	-.21018	-.0793151
stage1	-.0276048	.0291643	-0.95	0.344	-.0847657	.0295562
stage2	-.091044	.0429937	-2.12	0.034	-.1753101	-.0067779
lnc	.3698705	.0214508	17.24	0.000	.3278276	.4119133
_cons	9.628452	.0680018	141.59	0.000	9.495171	9.761733

---

```

. ** GAMMA QMLE
. glm totexp aggr_trt agec white black single married
well mod stage1 stage2 lnc, link(log) family(gamma) robust
irls

```

Iteration 6: deviance = 7071.498

totexp	Coef.	Semi-Robust Std. Err.	z	P> z	[95% Conf. Interval]	
<b>aggr_trt</b>	<b>.1794717</b>	<b>.0327351</b>	<b>5.48</b>	<b>0.000</b>	<b>.1153121</b>	<b>.2436313</b>
agec	.010508	.0025526	4.12	0.000	.005505	.015511
white	.0854311	.0535442	1.60	0.111	-.0195136	.1903758
black	.1795774	.0681181	2.64	0.008	.0460683	.3130864
single	.033353	.0559614	0.60	0.551	-.0763292	.1430353
married	-.0542054	.0349642	-1.55	0.121	-.1227339	.0143232
well	-.2484182	.0648256	-3.83	0.000	-.375474	-.1213623
mod	-.1447475	.0337717	-4.29	0.000	-.2109388	-.0785563
stage1	-.0276047	.0292326	-0.94	0.345	-.0848996	.0296902
stage2	-.0910438	.0436959	-2.08	0.037	-.1766862	-.0054015
lnc	.3698705	.0216682	17.07	0.000	.3274016	.4123394
_cons	9.628452	.0676074	142.42	0.000	9.495944	9.76096

# Method of recycled predictions

---

- In order to calculate effect in the additive scale
- Estimate parameters in
$$\exp(\beta_0 + \beta_1 * TRT + \beta_2 * X)$$
- Predict after turning on & off the TRT
- $\mu_1 = \exp(\beta_0 + \beta_1 * 1 + \beta_2 * X)$
- $\mu_0 = \exp(\beta_0 + \beta_1 * 0 + \beta_2 * X)$
- Summ  $\mu_1$  ,  $\mu_0$

```
. ** THE METHOD OF RECYCLED PREDICTIONS - TO ESTIMATE EFFECT IN  
ADDITIVE-SCALE
```

```
. gen trtold=aggr_trt
```

```
. replace aggr_trt =1
```

```
(2219 real changes made)
```

```
. predict mu1, mu
```

```
. replace aggr_trt=0
```

```
(7129 real changes made)
```

```
. predict mu0, mu
```

```
. replace aggr_trt = trtold
```

```
(4910 real changes made)
```

```
. summ mu1 mu0
```

Variable	Obs	Mean	Std. Dev.	Min	Max
mu1	7129	21447.39	6183.882	12332.27	59634.99
mu0	7129	17923.83	5167.942	10306.23	49837.65

# Single equation models: Generalized Linear Models

---

## ● Limitations

- Inefficient when empirical distribution of  $Y$  has fat tails (kurtotic)
- Log Link is the most common link function used, but assumption that  $\log(\mu) = X\beta$  may be violated
- Appropriate link or variance functions not known a-priori
- Various tests exist, not always sufficient to overcome these problems

# Single equation models: Modified Park Test following GLM

- **Proposed by Manning & Mullahy (2001)**

- $\text{Var}(Y | X) = \theta_1 \mu^{\theta_2}$
- For Gamma, we assume  $\theta_1 = 1, \theta_2 = 2$
- Test:  $\ln(\text{Var}(Y | X)) = \ln(\theta_1) + \theta_2 * \ln(\mu)$
- Use Residual  $^2$  as an estimate for variance.
- `glm r2 lnmu, link(log) family(gamma)  
robust`

```

/* Modified Park test */
. predict mu, mu
. gen lnmu = ln(mu)
. gen r2 = (totexp - mu)^2
. glm r2 lnmu, link(log) family(gamma) robust

```

---

r2	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
lnmu	1.468981	.2492923	5.89	0.000	.9803774	1.957585
_cons	5.419501	2.497738	2.17	0.030	.5240253	10.31498

---

```

. test lnmu=1
      chi2( 1) =      3.54
  Prob > chi2 =      0.0599

. test lnmu=2
      chi2( 1) =      4.54
  Prob > chi2 =      0.0332

. test lnmu=3
      chi2( 1) =     37.72
  Prob > chi2 =      0.0000

. drop mu lnmu r2

```

# Single equation models: Generalized Gamma Models

---

- Generalized Gamma FIML
- Three parameter distribution
  - Mean is a function of all three parameters
  - Only one of them is related to  $X\beta$ , implicit log-link
  - Special cases are Gamma, Exponential, Weibull, log-normal
  - Except for log-normal case, no retransformation problem
  - `stset y`
  - `gengam2 x`



```
. /* Generalized Gamma Regression */
```

```
. stset totexp
```

```
. ***** Download -gengam2- command from
```

```
http://home.uchicago.edu/~abasu/index\_files/Page317.htm AND place  
in C:\ado\Personal\  
.
```

```
. gengam2 aggr_trt agec white black single married well  
mod stage1 stage2 lnc, robust
```

_t	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
<b>aggr_trt</b>	<b>.193495</b>	<b>.0333561</b>	<b>5.80</b>	<b>0.000</b>	<b>.1281183</b>	<b>.2588717</b>
agec	.0104401	.002439	4.28	0.000	.0056598	.0152205
white	.0913009	.0521438	1.75	0.080	-.0108992	.1935009
black	.1798185	.0662471	2.71	0.007	.0499766	.3096605
single	.0355918	.0546193	0.65	0.515	-.07146	.1426436
married	-.0508413	.0332982	-1.53	0.127	-.1161046	.0144221
well	-.2535959	.0625437	-4.05	0.000	-.3761792	-.1310125
mod	-.1453943	.0321754	-4.52	0.000	-.2084568	-.0823318
stage1	-.0261988	.0281348	-0.93	0.352	-.0813419	.0289444
stage2	-.0900821	.0410532	-2.19	0.028	-.1705449	-.0096192
lnc	.3746245	.0203089	18.45	0.000	.3348198	.4144292
_cons	9.578895	.0675292	141.85	0.000	9.44654	9.711249

/ln_sig	-.053114	.0147078	-3.61	0.000	-.0819408	-.0242872
/kappa	.8766324	.0314132	27.91	0.000	.8150637	.9382012
-----						
sigma	.9482719	.013947			.9213265	.9760053
-----						

Tests for identifying distributions

Distributions	chi2	df	Prob>chi2
Std. Gamma (kappa = sigma)	3.27	1	0.0704
Log Normal (kappa = 0)	778.77	1	0.0000
Weibull (kappa = 1)	15.42	1	0.0001
Exponential (kappa = sigma = 1)	50.89	2	0.0000

New Variables for Schoenfield Residuals created: \_h\*, \_c\*

**. drop smr smear smear0 smear1 res2 \_\***

# Goodness-of-fit tests

---

- Pearson Correlation

- Correlation between raw-scale predictions and residuals

- Pregibon's Link Test / Reset Test

- Run same model with  $X\beta$  and  $X\beta^2$  as covariates

- Hosmer-Lemeshow Test

- Plot mean residuals across deciles of  $X\beta$ .
- Perform joint test that all means are zero

- Cope's Test

```

. /* GOODNESS OF FIT TESTS */
. ** For OLS
. reg totexp aggr_trt agec white black single married well
mod stage1 stage2 lnc, robust
. predict xb, xb
. gen mu_ols= xb
. gen res_ols = totexp - mu_ols
. ** Pearson Corr
. pwcorr res_ols mu_ols, sig

```

	res_ols	mu_ols
res_ols	1.0000	
mu_ols	-0.0000	1.0000

```

. ** Link Test
. gen xb2=xb^2
. reg totexp xb xb2, robust

```

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]
xb	.2645214	.3449794	0.77	0.443	-.4117406 .9407833
xb2	.0000171	8.43e-06	2.03	0.043	5.46e-07 .0000336
_cons	7353.987	3357.482	2.19	0.029	772.3251 13935.65

```

.      ** H-L test
.      xtile xbtile =xb, nq(10)
.      qui tab xbtile, gen(xbt)
.      reg res_ols xbt1 xbt2 xbt3 xbt4 xbt5 xbt6 xbt7
xbt8 xbt9 xbt10, nocons robust

```

res_ols	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
xbt1	857.5334	614.4044	1.40	0.163	-346.8819	2061.949
xbt2	681.1362	674.0399	1.01	0.312	-640.1824	2002.455
xbt3	-525.1787	642.1524	-0.82	0.413	-1783.988	733.631
xbt4	-142.7663	699.025	-0.20	0.838	-1513.063	1227.531
xbt5	938.962	933.5517	1.01	0.315	-891.0768	2769.001
xbt6	-638.9313	716.1985	-0.89	0.372	-2042.893	765.0306
xbt7	-1635.909	696.5204	-2.35	0.019	-3001.296	-270.5221
xbt8	-2250.963	710.2455	-3.17	0.002	-3643.255	-858.6706
xbt9	707.1719	1138.647	0.62	0.535	-1524.915	2939.259
xbt10	2007.334	1231.866	1.63	0.103	-407.489	4422.158

```

.test xbt1 xbt2 xbt3 xbt4 xbt5 xbt6 xbt7 xbt8 xbt9 xbt10

```

```

      F( 10, 7119) =      2.41
      Prob > F =      0.0074

```

```

. drop xb xb2 xbt* xbtile

```

```

. ** For lnOLS
. reg lny aggr_trt agec white black single married well
mod stage1 stage2 lnc, robust
. predict xb, xb
. gen smr = exp(lny- xb)
. qui summ smr
. gen smear = r(mean)
. gen mu_lols= exp(xb)*smear
. gen res_lols = totexp - mu_lols
. ** Pearson Corr
. pwcorr res_lols mu_lols, sig
| res_lols mu_lols
-----+-----
res_lols | 1.0000
mu_lols | -0.1776 1.0000
| 0.0000
. ** Link Test
. gen xb2=xb^2
. reg lny xb xb2, robust
|
| Coef. Robust
| lny | Std. Err. t P>|t| [95% Conf. Interval]
-----+-----
| xb | 7.586948 1.443073 5.26 0.000 4.758097 10.4158
| xb2 | -.3492898 .0757058 -4.61 0.000 -.4976956 -.2008839

```

```

    _cons |      -30.9968    6.869053    -4.51    0.000   -44.46218   -17.53141
-----+-----

```

```

.      ** H-L test
.      xtile xbtile =xb, nq(10)
.      qui tab xbtile, gen(xbt)
.      reg res_lols xbt1 xbt2 xbt3 xbt4 xbt5 xbt6 xbt7
xbt8 xbt9 xbt10, nocons robust

```

res_lols	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
xbt1	2479.074	635.594	3.90	0.000	1233.121	3725.027
xbt2	2585.675	755.5887	3.42	0.001	1104.497	4066.853
xbt3	608.5663	694.4084	0.88	0.381	-752.6807	1969.813
xbt4	573.4585	733.1452	0.78	0.434	-863.724	2010.641
xbt5	1438.745	775.6945	1.85	0.064	-81.84715	2959.336
xbt6	-215.099	826.1292	-0.26	0.795	-1834.558	1404.36
xbt7	-2386.189	757.4182	-3.15	0.002	-3870.954	-901.4244
xbt8	-3506.212	839.8621	-4.17	0.000	-5152.592	-1859.833
xbt9	-3966.24	1005.915	-3.94	0.000	-5938.133	-1994.347
xbt10	-10547.35	1207.501	-8.73	0.000	-12914.42	-8180.294

```

. test xbt1 xbt2 xbt3 xbt4 xbt5 xbt6 xbt7 xbt8 xbt9 xbt10
      F( 10, 7119) = 15.10
      Prob > F = 0.0000
. drop xb xb2 xbt* xbtile

```

```

. ** For log-GLM
. glm totexp aggr_trt agec white black single married
well mod stage1 stage2 lnc, link(log) family(gamma) robust
. predict xb, xb
. gen mu_glm = exp(xb)
. gen res_glm = totexp - mu_glm
.      ** Pearson Corr
.      pwcorr res_glm mu_glm, sig
      |   res_glm   mu_glm
-----+-----
res_glm |   1.0000
mu_glm  | -0.0062   1.0000
      |   0.6022
.      ** Link Test
.      gen xb2=xb^2
.      glm totexp xb xb2, link(log) family(gamma) robust
-----+-----
      |               Robust
totexp |   Coef.   Std. Err.      z    P>|z|    [95% Conf. Interval]
-----+-----
      |
      |   xb      2.601821    3.09824    0.84   0.401   -3.470618    8.67426
      |   xb2     -.0806185   .1555559   -0.52   0.604   -.3855025    .2242655
      |   _cons   -7.950714   15.41964   -0.52   0.606   -38.17266   22.27123
-----+-----

```



```

.      ** H-L test
.      xtile xbtile =xb, nq(10)
.      qui tab xbtile, gen(xbt)
.      reg res_glm xbt1 xbt2 xbt3 xbt4 xbt5 xbt6 xbt7
xbt8 xbt9 xbt10, nocons robust

```

---

res_glm	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
xbt1	-781.7378	554.103	-1.41	0.158	-1867.944	304.4688
xbt2	473.2616	736.601	0.64	0.521	-970.6954	1917.219
xbt3	-323.851	626.755	-0.52	0.605	-1552.477	904.7751
xbt4	49.60152	702.7822	0.07	0.944	-1328.061	1427.264
xbt5	1288.524	797.3756	1.62	0.106	-274.5689	2851.618
xbt6	207.3119	861.6822	0.24	0.810	-1481.841	1896.465
xbt7	-558.6503	690.9381	-0.81	0.419	-1913.094	795.7937
xbt8	-1955.442	718.1396	-2.72	0.006	-3363.209	-547.6744
xbt9	1364.908	1148.222	1.19	0.235	-885.9489	3615.764
xbt10	156.9733	1227.114	0.13	0.898	-2248.535	2562.482

---

```

. test xbt1 xbt2 xbt3 xbt4 xbt5 xbt6 xbt7 xbt8 xbt9 xbt10
      F( 10, 7119) = 1.48
      Prob > F = 0.1382

```

```

. drop xb xb2 xbt* xbtile

```

# Single equation models: Generalized Linear Models - EEE

---

## Flexible Mean Model

$$\frac{(\mu^\lambda - 1)}{\lambda} = \mathbf{X}\beta \quad \text{if } \lambda \neq 0$$

$$\log(\mu) = \mathbf{X}\beta \quad \text{if } \lambda = 0$$

## Flexible Variance Model

$$\mathbf{V}(\mathbf{Y}) = \phi\mu^\theta$$

$\theta = 1 \Rightarrow$  **Poisson**

$\theta = 2 \Rightarrow$  **Gamma**

$\theta = 3 \Rightarrow$  **Inv. Gauss**

# Single equation models: Generalized Linear Models - EEE

---

- **Goal: estimate  $\gamma = (\beta^T, \lambda, \phi, \theta)^T$**
- **EEE estimator estimates:**  
**Mean model parameters  $(\beta, \lambda)$  and**  
**Var. model parameters  $(\phi, \theta)$**   
**simultaneously using estimating equations.**
- **Extended Estimating Equations (EEE) ....**

# Single equation models: Generalized Linear Models - EEE

MEAN MODEL

$$\mathbf{G}_{\beta_j}^i = (\mathbf{y}_i - \mu_i) \mathbf{V}_i^{-1} (\partial \mu_i / \partial \beta_j)$$

$$\mathbf{G}_{\lambda}^i = (\mathbf{y}_i - \mu_i) \mathbf{V}_i^{-1} (\partial \mu_i / \partial \lambda)$$

VAR. MODEL

$$\mathbf{G}_{\phi}^i = [(\mathbf{y}_i - \mu_i)^2 - \mathbf{V}_i] \mathbf{V}_i^{-2} (\partial \mathbf{V}_i / \partial \phi)$$

$$\mathbf{G}_{\theta}^i = [(\mathbf{y}_i - \mu_i)^2 - \mathbf{V}_i] \mathbf{V}_i^{-2} (\partial \mathbf{V}_i / \partial \theta)$$

# Single equation models: Generalized Linear Models - EEE

---

- Latest Stata code available from  
<http://home.uchicago.edu/~abasu>

```
. summarize totcost, meanonly
. global scale=r(mean)
. generate y=totcost/$scale
. pglm y trt x
. pglm y trt x [pw=marsupwt]
. pglm predict y_hat, mu scale($sc)
. pglm predict ietrt, ie(trt) scale($sc)
. pglm predict me_x, me(x) scale($sc)
```

```
. /* EEE - EXTENDED ESTIMATING EQUATIONS */
. ***** Download -pglm- command from
http://home.uchicago.edu/~abasu/index\_files/Page317.htm AND place
in C:\ado\Personal\
```

```
. qui sum totexp, meanonly
. gen y = totexp/r(mean)
. global scale = r(mean)
. summ y
```

Variable	Obs	Mean	Std. Dev.	Min	Max
y	7129	1	1.131858	.0001372	19.4377

```
.
. pglm y aggr_trt agec white black single married well
mod stage1 stage2 lnc
```

```
Extended GEE with Power Variance Function          No of obs          =          7129
Optimization: Fisher's Scoring                    Residual df          =          7114
```

```
Variance:      (theta1*mu^theta2)
Link:          (mu^lambda - 1)/lambda
Std Errors:    Robust
```

	y	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
-----							
y							
	aggr_trt	.176795	.0319912	5.53	0.000	.1140935	.2394965
	agec	.010301	.0025097	4.10	0.000	.005382	.01522
	white	.0772856	.0524606	1.47	0.141	-.0255354	.1801065
	black	.1771791	.0671802	2.64	0.008	.0455084	.3088498
	single	.0369893	.0563494	0.66	0.512	-.0734535	.1474322
	married	-.053268	.0341707	-1.56	0.119	-.1202413	.0137054
	well	-.2421367	.063453	-3.82	0.000	-.3665023	-.1177711
	mod	-.144254	.0339966	-4.24	0.000	-.2108861	-.077622
	stage1	-.0237874	.0289403	-0.82	0.411	-.0805093	.0329345
	stage2	-.0913624	.0418396	-2.18	0.029	-.1733666	-.0093583
	lnc	.3719294	.0221651	16.78	0.000	.3284867	.4153721
	_cons	-.2769968	.0682525	-4.06	0.000	-.4107693	-.1432244
-----							
lambda							
	_cons	.1928841	.303307	0.64	0.525	-.4015866	.7873549
-----							
theta1							
	_cons	1.166449	.0810133	14.40	0.000	1.007666	1.325233
-----							
theta2							
	_cons	1.457585	.2222653	6.56	0.000	1.021953	1.893217
-----							

```

. predict xb, xb
. pglm predict mu_pglm, mu scale($scale)

```

```
. summ mu_pglm totexp
```

Variable	Obs	Mean	Std. Dev.	Min	Max
mu_pglm	7129	20265.42	5696.584	10770.24	51476.94
totexp	7129	20265.17	22937.29	2.78	393908.4

```
. pglm predict trteffect, ie(aggr_trt) scale($scale)
```

```
. summ trteffect
```

Variable	Obs	Mean	Std. Dev.	Min	Max
trteffect	7129	3470.224	790.525	2190.088	7521.893

```
. gen res_pglm = totexp - mu_pglm
```

### GOODNESS OF FIT

```
.  
. ** Pearson Corr  
. pwcorr res_pglm mu_pglm, sig
```

	res_pglm	mu_pglm
res_pglm	1.0000	
mu_pglm	-0.0013	1.0000
	0.9158	



```

.      ** H-L test
.      xtile xbtile =xb, nq(10)
.      qui tab xbtile, gen(xbt)
.      reg res_pglm xbt1 xbt2 xbt3 xbt4 xbt5 xbt6 xbt7
xbt8 xbt9 xbt10, nocons robust

```

res_pglm	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
xbt1	-609.5808	553.8617	-1.10	0.271	-1695.314	476.1529
xbt2	646.0331	742.0791	0.87	0.384	-808.6624	2100.729
xbt3	-471.0719	618.8141	-0.76	0.447	-1684.131	741.9877
xbt4	-84.50918	702.9003	-0.12	0.904	-1462.403	1293.384
xbt5	1116.715	791.1745	1.41	0.158	-434.2219	2667.652
xbt6	365.4751	872.6898	0.42	0.675	-1345.256	2076.207
xbt7	-1006.109	683.5575	-1.47	0.141	-2346.085	333.8666
xbt8	-1908.943	719.4691	-2.65	0.008	-3319.317	-498.5697
xbt9	1258.806	1145.722	1.10	0.272	-987.151	3504.762
xbt10	691.741	1227.973	0.56	0.573	-1715.451	3098.933

```

. test xbt1 xbt2 xbt3 xbt4 xbt5 xbt6 xbt7 xbt8 xbt9 xbt10
      F( 10, 7119) = 1.55
      Prob > F = 0.1164

```

```

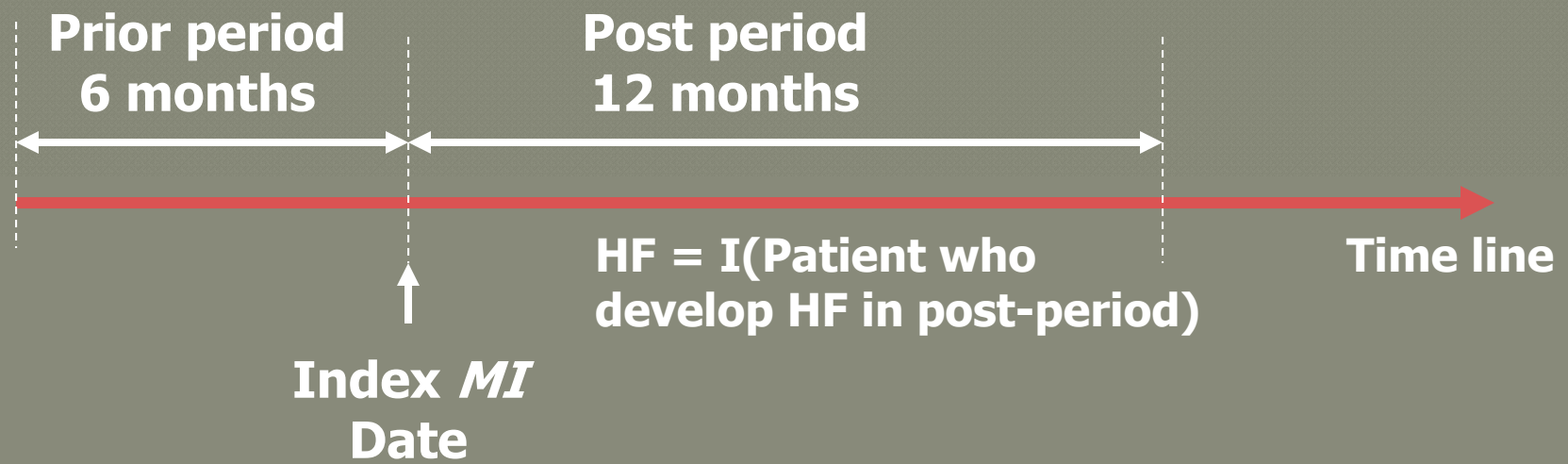
. drop xb xbt* xbtile

```

# Empirical Example: Cost of Post-MI HF

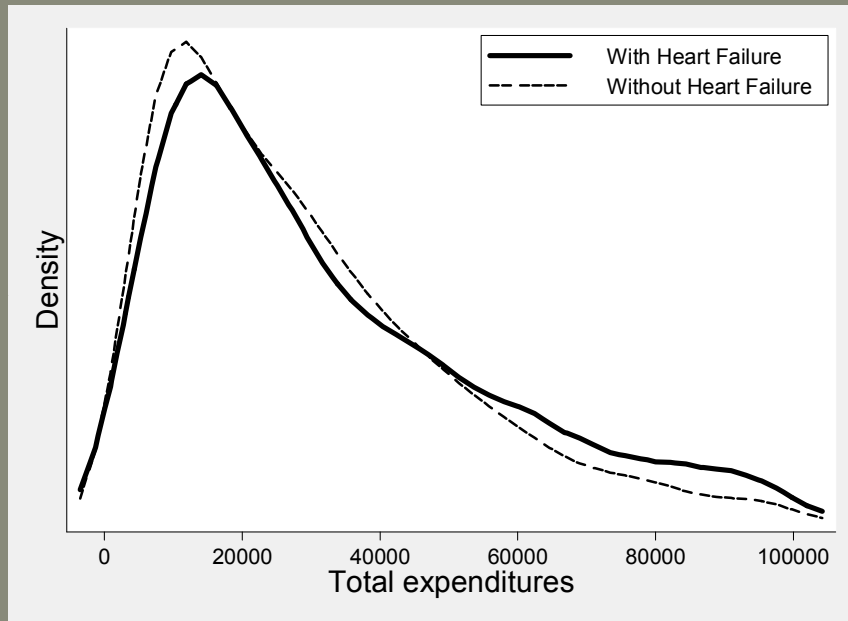
## ● Empirical example

- Cost attributable to Heart Failure (HF) post myocardial infarction (MI)
- Use Medstat data, extract all claims between July 1, 1997 and December 31, 2000



<b>Exclusion Criteria</b>	<b>Number of patients</b>
<b>Starting cohort</b>	<b>22091</b>
No claims in pre-MI or enrollment information	1279
Patients with pre-MI HF only	1696
Patients with pre-MI MI only	806
Patients with pre-MI HF and MI	269
Patients with post-MI MI	575
Post MI HF patients with intermediate coronary syndrome and coronary occlusion without MI	18
Patients with capitation type insurance plans	2258
Patients with VSD	16
Patients with mitral regurgitation	46
Patients with cost of index hospitalization < \$500	9
Patients with LOS>100 and cost < \$10,000	1
Patients with age<18 years or missing age	3
Patients with missing plan information	20
Patient with index date not in (1Dec1998 – 1Dec 2000)	4633
<b>Eligible cohort</b>	<b>10462</b>
Patients without full 1-year follow-up data.	3427
Observations with missing covariate values	3
<b>Final sample for this study</b>	<b>7428</b>

# Empirical Example: Cost of Post-MI HF



**With HF:** N=2259

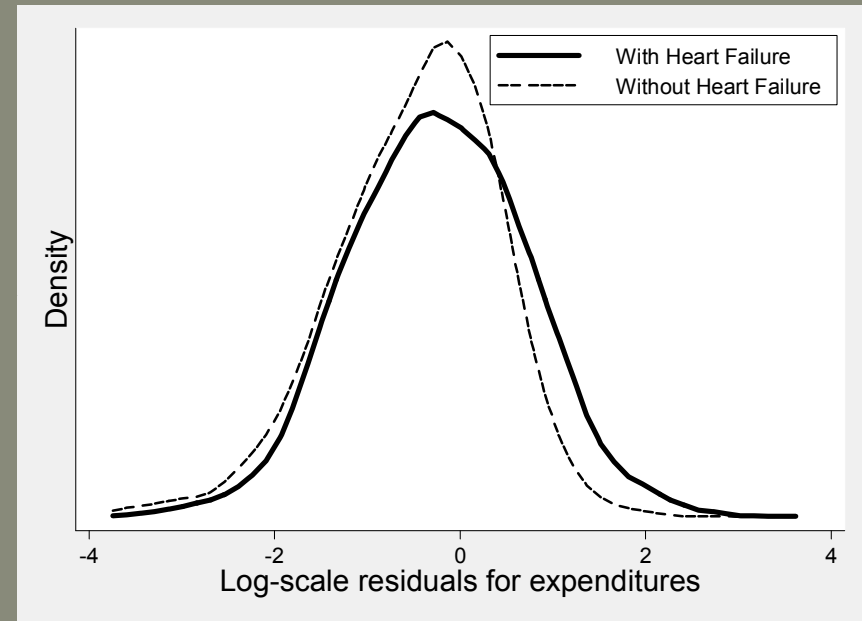
Mean (sd) = \$45,070 (57,366) ;

Sk. = 6.0; Kurt. = 73

**Without HF:** N=5169

Mean (sd) = \$31,708 (30,110) ;

Sk. = 4.4; Kurt. = 50



**With HF:** N = 2259

Sk. = 0.13; Kurt. = 3.2

**Without HF:** N = 5169

Sk. = -0.03; Kurt. = 3.4

# Empirical Example: Cost of Post-MI HF

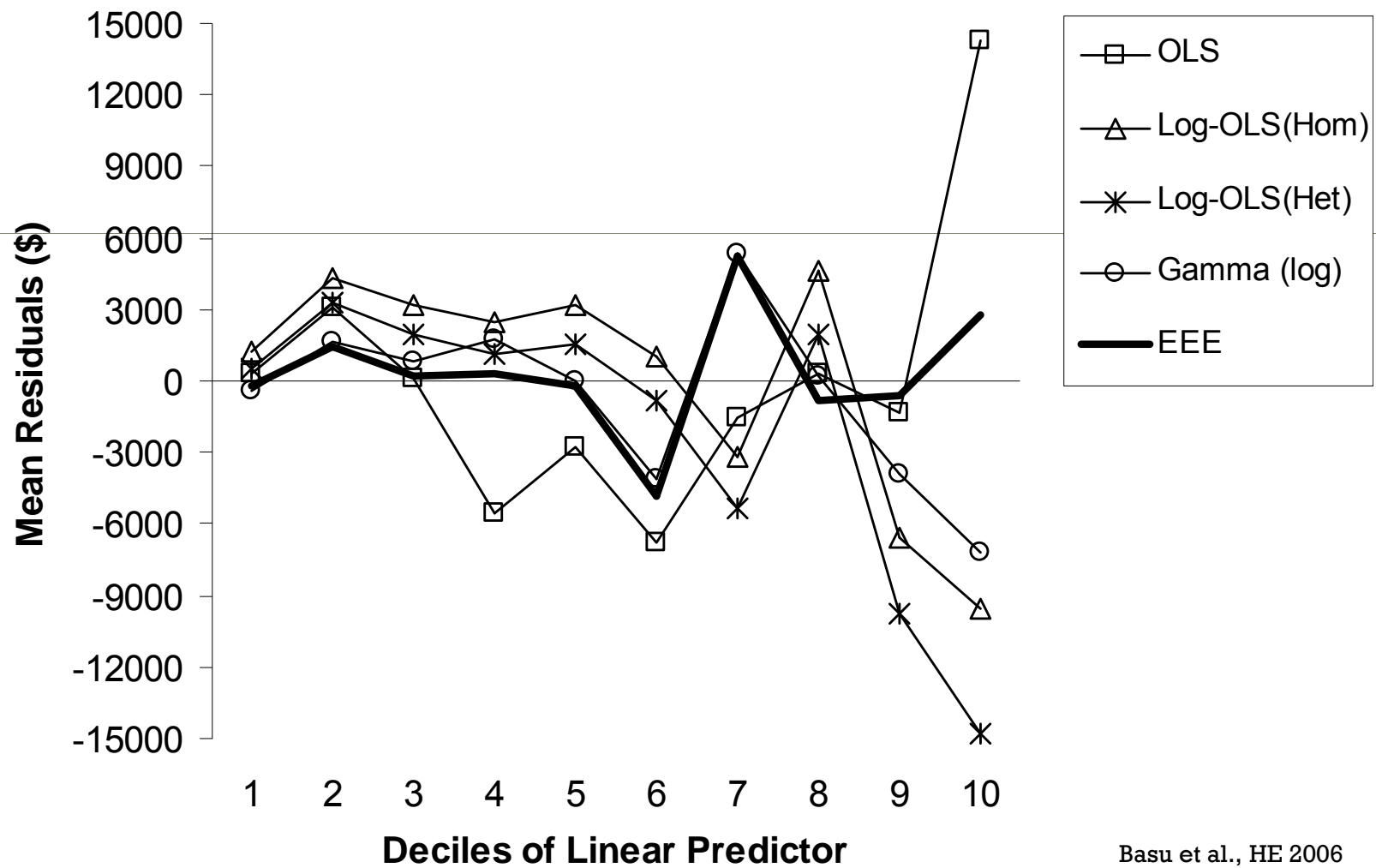
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- **Final Sample:** 7428 patients  
2259 (30.4%) with *HF*
- **Dependent variable:** Total expenditure
- **Independent variables:** HF, age, age-sq, insurance type, Medicare status, procedures performed, year of MI, type of *MI* and indicators for 30 categories of co-morbidities.
- **Compare:** OLS, log-OLS(with HF-specific smearing), GLM Gamma-Log link.

# Empirical Example: Cost of Post-MI HF

Estimator	$\hat{\pi}$
	Mean (std.err.)
OLS	15578 (1172)
[REDACTED]	[REDACTED]
Log-OLS (Het.)	17491 (1159)
Gamma (log link)	15890 (1104)
[REDACTED]	[REDACTED]

# Empirical Example: Cost of Post-MI HF



Basu et al., HE 2006

# Empirical Example: Cost of Post-MI HF

## ● Results from the EEE Model

- Mean Model (Link parameter)
- $\hat{\lambda} = 0.201$  (95% CI: 0.04, 0.37)
- Appropriate link  $\sim \mu^{0.2}$  or  $\mu^{0.25}$
- Variance Model:  $V(y) = \phi\mu^\theta$
- $\hat{\phi} = 0.80$  (95% CI: 0.72, 0.89);  $\hat{\theta} = 1.79$  (95% CI: 1.57, 2.01)
- Appropriate variance  $\sim \mu^2$

- Suggest fifth or fourth-root link with Gamma variance.



# Empirical Example: Cost of Post-MI HF

Estimator	$\hat{\pi}$ Mean (std.err.)	$(\hat{\pi}_{Row} - \hat{\pi}_{Col})$ Mean (std.err.)*			
			Log-OLS (Het.)	Gamma (log link)	EEE
OLS	15578 (1172)		<b>-1913</b> <b>(627)</b>	-312 (566)	856 (458)
Log-OLS (Het.)	17491 (1159)			<b>1601</b> <b>(239)</b>	<b>2769</b> <b>(560)</b>
Gamma (log link)	15890 (1104)				<b>1168</b> <b>(510)</b>
EEE	14722 (1135)				

# Other models

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- Multi-part models

- Two-part:           1<sup>st</sup> part model any costs (logit/probit)  
                          2<sup>nd</sup> part models costs given any
- Multi-part:        2-part models for each type of expenditures
- CDE:                1<sup>st</sup> part for multinomial categories of costs  
                          2<sup>nd</sup> parts for levels within each category

- Mixture models

# Comparison of Alternative Estimators

- Comparisons through simulations or empirical applications:
  - Duan et al., JBES 1983; Mullahy, JHE 1998; Lipscomb et al., SMDM 1999; Manning & Mullahy, JHE 2001; Deb & Trivedi, JHE 2002; Deb & Burgess, Hunter College Mimeo 2003; Basu et al., HE 2004; Buntin & Zaslavsky, JHE, 2004; Basu & Rathouz, Biostatistics 2005; Briggs et al., HE 2005; Manning et al., JHE 2006; Basu et al., HE 2006; Hill & Miller, HE 2009;
- Consistently found that no one estimator is appropriate for all expenditure data
- Theory has difficult time predicting signs of covariate effects – provides almost no guidance on functional form

# Other issues with cost models

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- Longitudinal costs – generalized estimating equations versus random effects model
- Right censoring
- Survival effects versus intensity effects
- Causal estimation – instrumental variables
- Predictions