Survey of Income and Program Participation

Transitions In and Out of Poverty: New Data from the Survey of Income and Program Participation

by

Patricia Ruggles, Urban Institute Roberton Williams, Congressional Budget Office

No. 8716

38

ACKNOWLEDGEMENT

The authors would like to thank Anne Bergsman for her invaluable programming help. Opinions expressed in this paper are those of the authors and do not necessarily represent the views of the Urban Institute or the Congressional Budget Office. This paper was prepared for presentation at the 1986 Annual Meetings of the Allied Social Science Association, New Orleans, Louisiana, December 1986. It is being made available in the SIPP Working Paper Series in order to broadly disseminate early findings from the Survey of Income and Program Participation.

TABLE OF CONTENTS

INTRO	DDUCTION	. 1
I.	Poverty Rates and Poverty Transitions as Measured in the SIPP	. 4
II.	Incidence of Demographic and Labor Market Events and Their Association with Poverty Status	. 9
III.	Poverty Transitions Made by Those in Families Experiencing Demographic and Employment-Related Events	.14
IV.	Poverty Transitions for Those in Female-Headed Families and in Families with Elderly Members	.17
	Female-headed Families Families with Elderly Members	.18
CONCL	.USION	.24
APPEN	IDIXES	
FOOTN	NOTES	

Transitions In and Out of Poverty:
New Data from the Survey of Income and Program Participation

INTRODUCTION

Factors associated with the onset of poverty spells and with exits from poverty have long been a major topic of interest to policy makers and analysts alike. From a policy perspective, information on the events likely to be associated with major changes in income is clearly helpful in designing anti-poverty initiatives. Further, a better understanding of the relative importance of various possible causes of poverty is potentially useful in improving the targeting of existing income assistance programs. And finally, if program participants and other members of the poverty population are to be encouraged to leave poverty and become self—sufficient, data on the major factors associated with exits from poverty may help us to assess the potential for success of alternative anti-poverty strategies.

Unfortunately, until recently it has been difficult to study specific transitions in and out of poverty in any detail. We have only a few major databases that provide information on individuals' incomes over a period of time, thereby allowing those who are newly poor or newly escaped from poverty to be distinguished from the rest of the poor or non-poor populations. The few longitudinal databases that exist, such as the Panel Study of Income Dynamics (PSID) and the National Longitudinal Survey (NLS), have typically provided information on annual income only, making it possible to identify only the year in which major income transitions occurred. Associating these annual income changes with specific events that occurred at some point during the year can be very difficult, since the event and the income change may actually have taken place at very different points in time and may not be causally related. Further, many income changes may be missed in such annual data, either because they

result in a poverty spell lasting less than a year or because they occur late enough in the year so that total annual income for that year does not fall below the poverty level. Finally, to the extent that income transitions and changes in household composition occur simultaneously, it can be difficult to identify the specific household members whose income is being reported for the year as a whole, as opposed to those whose income is included for only part of the year or possibly not at all. These problems can also make it difficult to identify newly poor or newly non-poor individuals and households.

Survey of Income and Program Participation (SIPP) allow us to avoid many of these problems in examining income transitions. The SIPP is a longitudinal panel survey that traces specific individuals over a period of more than two years. Interviews are conducted every four months, and information on income and family composition is collected on a month by month basis. Because data are collected on shorter intervals than in other longitudinal surveys, the specific points at which income transitions occur can be more precisely identified, and these transitions can be more closely matched with any associated changes in the composition of the household or in the employment status of household members.

This paper uses data from the SIPP to examine transitions into and out of poverty. The focus of the paper is on six specific demographic and employment-related events, and their association with becoming poor or leaving poverty. The events considered are the entrance of a new baby into the family, a death in the family, a marriage, a separation or divorce, the unemployment of a family member, and the entrance into employment of a family member. In all cases, the event has been considered in relation to the income status of all persons who were in the family in the month in

which the event occurred, and poverty statistics have all been calculated at the family level.

To assess the effects of these events on poverty transitions, their incidence among both the poor and non-poor populations is compared with that for the newly poor and the newly non-poor. The probability of becoming poor (or of becoming non-poor) if one is in a family where such an event occurs is then calculated for each of the events under investigation.

Examination of the relationship between these events and poverty transitions for the population as a whole is followed by similar analysis for people in four specific types of families. These family types are female—headed families, families headed by males or by married couples, families with elderly members (aged 65 and over), and families with no elderly members. Within each demographic group, both the overall incidence of each event and its association with changes in poverty status are investigated.

Although the overall aim of this paper is to examine the relationship between the various life events considered and transitions in and out of poverty, it should be noted that the data presented are primarily descriptive in nature, and do not imply a strict causative model of entry into or exit from poverty. Clearly, assessing the relationships between these events and poverty transitions in detail would require simultaneously taking into account the possible income effects of a wide array of other variables. In addition, we have by no means considered all possible event that could potentially affect income transitions. The data presented here are preliminary only, and represent a first attempt to identify some of variables that may be important in explaining poverty transitions.

Mevertheless, we believe that the results outlined below are highly suggestive, and point to some additional topics that deserve further investigation.

Before turning to those results, the next section of this paper briefly describes the specific SIPP sample used for this study, and presents some summary data on poverty rates and poverty transitions calculated from the SIPP. The third section of the paper provides data on the association of our six life events with poverty transitions for the poor and non-poor populations as a whole. The impacts of these events for those in the specific demographic groups outlined above are considered in section four, and the final section presents our conclusions and suggestions for further research.

I. Poverty Rates and Poverty Transitions as Measured in the SIPP

The estimates presented in this paper are based on data taken from the first five waves, or sets of interviews, of the SIPP. These data cover the 16 months from September 1983 through December 1984. Because data from each wave were released separately by the Census Bureau, individual records from each of the five waves had to be linked together to create files that spanned the entire 16 month period. There are records for approximately 60,000 individuals included in this linked file, although only about two-thirds of these people were in the SIPP sample in every month. Individuals have been assigned a zero weight for months in which they were not in the sample, cross-sectional weights assigned by the Census Bureau have been used in computing month by month statistics on income and demographic events. Data for about 50,000 individuals are available on average in any given month (specific counts for each month are shown in appendix table A.1).

One potential problem with using a file consisting of linked wave records involves continuity of the data across waves. In processing the raw SIPP data for release as public use samples, the Bureau of the Census performed a wide range of editing and imputation activities to ensure that

the data were complete and internally consistent. All edits and imputations were done on each wave of data individually, without reference to data from other waves, however. There was no attempt to assure continuity across waves, either in terms of information reported or that imputed by the Census Bureau.

Preliminary analysis of the linked files indicates that there is much greater variation in reported data for specific individuals across waves than there is within waves, particularly in terms of both amounts and types of income received. This increase in variation may arise from biases in reporting by respondents, who may tend to attribute current income levels to all previous months in the same wave, but it could also be associated with the fact that the Census Bureau's imputation procedures do not take into account income or other characteristics reported in other waves. Ideally, the longitudinally linked files should at some point receive the same type of careful editing for consistency that has been done for the cross-sectional files. Although the staggered interview schedule helps to smooth out the wave variations for aggregate values reported in the analysis, 3 this is not the case for transition events. Because both the events and movements into and out of poverty are more likely to show up in the data due to these variations in individual data, our findings should be interpreted with caution.

In spite of some possible suppression of within-wave variations in income, it is nonetheless clear that many individuals experience substantial variations in income on a month-to-month basis. Table 1 shows four alternative poverty measures as calculated from the SIPP for persons and families of several different types. These four measures, all of which are based on cash income only, include an annual poverty rate, based on family income over the year as a whole; an "ever poor" rate, showing the

proportion who were poor for at least one month during the year; an "always poor" rate, showing the proportion poor in every month during the year; and finally an average of the poverty rates calculated for each specific month during calendar year 1984. A fifth poverty rate, the official poverty rate as calculated from the Current Population Survey (CPS), has been included in Table 1 for comparison purposes as well.

Poverty rates under each of these alternative definitions are shown in Table 1 for each of five population subgroups and for the population as a whole. The subgroups shown are:

- Married couples with children—all those living in families headed by married couples living with their own children under age 18;
- Single parents with children—all those living in families headed by single parents living with their own children;
- O Unrelated Individuals—all persons not living in families; and
- Other—all persons living in families, but not in one of the types of families included in the first two categories above.

These four categories are mutually exclusive, and together account for the entire population. A fifth category, consisting of all persons aged 65 and over, regardless of family situation, is also shown in the table. It should be noted that all those in this category are also included in one of the other four, however.

Perhaps the most striking feature of the poverty rates shown in Table 1 is the large amount of within-year movement into and out of poverty that they imply for all five population subgroups, as well as for the population as a whole. For all subgroups, the proportion of people poor on average in any given month, based on their monthly income, is always higher than the proportion who are poor when their incomes over the year as a whole are taken into account. The ratio of monthly to annual poverty rates varies somewhat across these population subgroups, however, generally in ways that might be expected. For example, elderly persons and single parent

Table 1

ALTERNATIVE POVERTY RATES BY FAMILY TYPE, 1984 (in percents)

•		Survey of rogram Par				Current Population Survey
Family Type	Annual Rate	Poor All 12 Months	Poor In Any Month	Average of Monthly Rates		1984 Annual Rate
All Persons	11.0	5.9	26.2	13.7		14.4
Married Couples with Children	7.4	2.8	24.3	10.2		10.5
Single Parents with Children	39.9	25.8	60.8	42.7		44.7
Unrelated Individuals	17.7	11.0	35.9	21.9		21.8
Other Persons	4.5	2.0	14.3	6.3		5.3
Elderly Persons	10.3	6.8	18.5	12.1	•	12.4

SOURCE: Tabulations of data from the Survey of Income and Program Participation and the Current Population Survey.

families, who are most likely to rely on transfer incomes that typically vary relatively little from month to month, have annual poverty rates that are fairly close to their monthly rates, while families headed by married couples with children, who may be more likely to depend primarily on employment incomes, have annual and monthly poverty rates that diverge somewhat more.

The impacts of within-year movements into and out of poverty can be seen even more clearly by comparing those who are poor in at least one month with those who are poor over the year as a whole. For the full sample, the proportion poor at least one month is more than four times as high as the proportion poor in every month—about 26 percent as compared to about 6 percent. Again, the fluctuations seen differ across the five subgroups considered, with single-parent families having the most stable (and most consistently low) incomes, and with married couples with children experiencing the greatest fluctuations. These estimates imply, for example, that while married couples with children are almost as likely to be poor in at least one month as is an average member of the population as a whole, they are less than half as likely to be continuously poor for an entire year. Elderly persons, on the other hand, are substantially less likely than the average to experience at least one month of poverty, but those who are poor are more likely to stay poor over the year as a whole.

What causes the differences in the likelihood of becoming or remaining poor seen in Table 1? To what extent are these poverty transitions related to demographic events, such as a divorce or the birth of a child, and to what extent to labor market events, such as the loss of a job? The next section of this paper explores these questions for the population as a whole.

II. Incidence of Demographic and Labor Market Events and Their Association with Poverty Status

As discussed briefly in the introduction, this study considered the impacts of six specific demographic and economic events on transitions into and out of poverty. For the purposes of this analysis, poverty was defined on the basis of monthly cash income, and any family that received an income below one-twelfth of the relevant annual poverty threshold in any month was considered to be poor in that month. Individuals were considered poor in any month in which they were members of poor families. Thus, for example, individuals could enter poverty either by belonging to a family that lost income, or by joining a family that was already poor.

All of the analysis is on the basis of persons, not families, even though events, incomes, and poverty status are all determined at the family level. Thus, for example, a person is reported as having experienced a job loss in her or his family if any family member had earnings fall from above \$200 in one month to less than \$200 in the next, whether or not the individual's earnings changed. Note that this means that a family could have both a job loss and a job gain in a single month if different members experienced those two events. Some results that appear strange can be explained by this way of defining events.

number of persons who are counted as poor, as Table 1 showed, and many of those who are counted as becoming poor in this analysis may in fact have remained poor for only a short period of time. In future work on this topic, we hope to examine what determines the lengths of poverty spells as well as what transition points define spell beginnings and endings. Such an examination was beyond the scope of the present analysis, however.

The six specific demographic and economic events examined in this study were chosen largely because other studies based on annual longitudinal data, notably the PSID, have shown them to be important determinants of poverty transitions over longer analysis periods. We hypothesized that they would prove to be at least as important in explaining month-to-month poverty transitions. Because the SIPP, unlike databases using an annual income measure, is able to identify the specific point at which income transitions occur, these transitions can be matched more closely with particular life events. We believed that this might be particularly helpful in examining the impacts of family composition changes, whose timing (and even existance) can be difficult to determine in databases that rely on one yearly interview.

The six specific events examined in this study were:

- o Entrance of a New Baby into the Family. Births (and adoptions of infants) were measured by identifying all families that acquired a new sample member with age equal to zero. The birth was attributed to the month in which the new sample member was first reported.
- Death in the Family. A death was attributed to all families who lost a member whose reason for leaving the sample was coded as "death". However, this variable is as yet unedited on the SIFP, and preliminary analyses indicate that it may not, in fact, pick up all deaths that actually occurred in the sample. This event may therefore be somewhat under-represented in our analysis file.
- Occurrence of a Marriage. A marriage was attributed to all families containing at least one member who was coded as having moved from an unmarried status (including "never-married," "divorced," "separated," or "widowed") to "married spouse, present."
- Break-up of a Marriage. A marital break-up was attributed to all families including a member who was coded as moving from "married, spouse present" in one month to "divorced," "separated" or "widowed" in the next month. (In both this case and the previous one, the event was attributed to the family in which the person resided in the second month, and was treated as having occurred in that month.)

- Loss of a Job by a Family Member. A job loss was attributed to all families in which a member who had had earnings of more than \$200 in the previous month reported no earnings or earnings below \$200 in the current month. (An arbitrary level of \$200 was chosen to exclude earnings from casual employment; a person working 15 hours per week at minimum wage, for example, earns more than \$200 per month.)
- Employment of a Family Hember. An entry into employment was attributed to all families in which a member with no earnings or earnings under \$200 in the previous month reported earnings above \$200 in the current month.

In all cases, the event in question was attributed not only to the individual or individuals directly involved, but also to those in the same family in the month in which the event occurred. Because poverty status is measured at the family level and depends on family income, these events would affect the status not only of those for whom they occurred, but also of all others in the same family at the time of their occurrence.

Table 2 shows the overall incidence of each of these six events for the sample as a whole, for the poor and non-poor portions of the sample, and for those becoming poor or leaving poverty in the same month as the event. 6 The percentages shown in Table 2 represent the average percentage of those with the event across the sixteen months from September 1983 through December 1984. Thus, while sample sizes for some events may appear relatively small, particularly for the newly poor and newly non-poor populations (which each contain about 1000 sample persons in an average month) it should be noted that the figures presented here are actually monthly averages based on sixteen sets of observations of each event. For example, the data on new babies entering newly poor families imply that only about 0.07 percent of the sample would fall into this cell in an average month-or, since the total sample size is about 50,000, approximately 35 families. Over the sixteen months as a whole, however, there are 562 births occurring in such families—still not an enormous number, but enough for some reasonably reliable generalizations about their

Table 2

Percentage of Persons in Families
with Occurrences of a Major Life Event,
by Poverty Status

Event	All Persons	Poor Persons	Non-Poor Persons	Persons Becoming Poor This Month	Persons Leaving Poverty This Month
New Baby	0.5	1.0	. 0.5	3.3	0.4
Death	0.1	-	0.1		
Marriage	0.2	0.2	0.2	0.6	1.3
Marital Break-up	0.2	0.4	0.2	2.1	0.2
Job Loss	5.1	8.8	4.5	40.3	2.6
Job Gain	5.7	3.9	6.0	3.5	46.8
TOTAL NUMBER OF PERSONS IN CATEGORY (in Thousands):	232,254	34,423	197,831	4,912	4,561
Percentage of Population in Category:	100	14.8	85.2	2.1	2.0

Source: Calculated from the Survey of Income and Program Participation. Figures shown are averages for the months Sept. 1983 through Dec. 1984. See text for definitions of events and populations shown.

relative incidence. Appendix tables A.2 through A.5 show the month-by-month incidence of each event for the poor, non-poor, newly poor and newly non-poor populations respectively.

The most conspicuous finding illustrated by Table 2 is the very large role played by labor market events in determining entrances into and exits from poverty. More than 40 percent of those becoming poor in any given month had a family member lose a job in that month, while almost 47 percent of those leaving poverty had a family member gain a job. None of the demographic events, in contrast, was experienced by much more than 3 percent of those in families entering or leaving poverty.

Overall, it appears that about half of the families with poverty transitions experienced a relevant employment-related or demographic event in the month in which the transition occurred. Of course, not all the events tallied can be said to have directly caused the transitions in question—among other things, some families may have experienced more than one event in the relevant month. Further, some events clearly occur without affecting transitions in the expected direction—for example, some proportion of those becoming poor also have a family member gain a job in the same month. In some cases, this job gain may have been motivated by the loss of whatever other income source previously kept the family out of poverty, but there is presumably some random shifting of income and events in the population as well.

As might be expected, employment-related events were also much more common in the sample as a whole than were any of the demographic events.

Approximately 5 percent of the sample as a whole were in families that lost a job during an average month, while almost 6 percent were in families gaining jobs. In contrast, the most common demographic event was the arrival of a new baby, which occurred in one-half of one percent of all

families in an average month. Those in poor families experienced significantly higher rates of job loss, birth, and marital break-up than did the population as a whole, although their rates were still well below those seen for families becoming poor in any given month.

In summary, then, employment-related events are much more common than demographic events for the poor and non-poor alike, and they also appear to be associated with a relatively large proportion of all poverty transitions. On the other hand, those becoming poor and those leaving poverty in a given month do experience certain demographic events, as well as employment-related events, at a higher rate than those who do not make such transitions. Although these demographic events are much rarer than the employment-related events across the population as a whole, they may in some cases be just as likely to be associated with poverty transitions for those families in which they occur. The next section explores that issue.

III. Poverty Transitions Made by Those in Families Experiencing Demographic and Employment-Related Events

As seen in the last section, employment-related events, because they are so common in the population as a whole, are associated with a much higher proportion of all transitions into and out of poverty than are demographic events. For those who experience them, however, certain demographic events are also relatively likely to lead to a change in poverty status, as Table 3 shows. In fact, one demographic change—the break-up of a marriage—is significantly more likely to be associated with becoming poor than is the loss of a job. Of those in families experiencing a divorce (or other break-up), 23 percent became poor in the same month, compared to about 17 percent of those in families with a job

Table 3

Percentage of Persons in Families
with Occurrences of a Major Life Event
Who Experience a Change in Poverty Status

Event	Total Number of Persons with Event (in Thousands)	Persons with Event Who Become Poor This Month	Persons with Event Who Leave Poverty This Month
New Baby	1,258	12.9	1.3
Death	142	1.1	0.7
Marriage	444	6.4	13.7
Marital Break-up	447	23.0	2.3
Job Loss	11,904	16.6	1.0
Job Gain	13,192	1.3	16.2

Source: Calculated from the Survey of Income and Program Participation. Figures shown are averages for the months Sept. 1983 through Dec. 1984. See text for definitions of events and populations shown.

loser. Families gaining a new baby are also relatively likely to become poor—about 13 percent of those in families experiencing a birth entered poverty.

Some cautions should be observed in interpreting these results. In particular, we have not yet examined the duration of the poverty spells associated with these transitions, and some may be very short indeed. For example, some of the entrances into poverty experienced by families with new babies may result from temporary unpaid maternity leaves that reduce the family's total earnings in the very short run, but that have relatively little impact over the longer run. Similarly, many divorcees and job losers may be poor for a month or two, in the sense that they have little or no income over that period, but they may have savings or other resources

that are more than sufficient to tide them over until a new job is found or other arrangements for the support of the family are made. Conceivably, for example, even those who take a month's vacation without pay between giving up one job and starting another could be counted as poor if the family has little other income during that time. Such instances of "poverty" are clearly different in fundamental ways from longer-term spells of poverty, or even from short term spells experienced by those with few other resources. Although data on spell durations could not be tabulated in time for inclusion in this paper, we hope to pursue this issue further in our future work.

Just as certain events are relatively likely to lead to spells of poverty for those who experience them, others are relatively likely to result in transitions out of poverty. Not surprisingly, the two events most likely to be associated with an exit from poverty are job gains and marriage. About 16 percent of those in families with someone who gains a job, and almost 14 percent of those in families experiencing a marriage leave poverty in the same month. Because most of those who marry or even of those who gain jobs were not poor before the event occurred and thus could not have made a transition out of poverty, this represents a very high probability of leaving poverty for those among the poor who do experience each event. Overall, about 60 percent of those in previously poor families experiencing a marriage leave poverty in the same month, while about 65 percent of those in previously poor families with a job gainer do so. 10

In summary, the data presented in Tables 2 and 3 indicate that employment-related events are much more common than demographic events, both in the population as a whole and among those experiencing poverty transitions, and therefore they tend to be associated with a much larger

proportion of all transitions. Indeed, the count of events shown in Table 3 implies that employment-related events account for more than 90 percent of all the occurrences of events considered in this paper. Nevertheless, Table 3 also indicates that demographic events can also be important predictors of transitions for those who experience them. As seen above, those in families with a divorce are actually more likely to become poor than are those in families with a job loser, and those in families experiencing a birth are also relatively likely to fall into poverty. Demographic events may also have important consequences in terms of exits from poverty—for example, almost as large a proportion of those in families with a marriage as of those in families with a job gainer leave poverty in the same month.

So far, we have examined the impacts of employment-related and demographic events on poverty transitions taking place among the population as a whole. It is to be expected, however, that certain events are likely to be more important for subgroups within the population than they are for others. The next section of this paper considers the impacts of these events for two important subgroups—female—headed families and families with elderly members—who together account for a large proportion of the poverty population (and especially, of the long-term poor), and whose transitions into and out of poverty are therefore of particular policy interest.

IV. Poverty Transitions for Those in Female-Headed Families and in Families with Elderly Members

As seen above, events such as a divorce or the loss of a job can in many cases be associated with entries into poverty. Certain population subgroups, such a families headed by women, may be particularly likely to have suffered these events. In addition, both the data shown in Table 1

and results from other studies based on the PSID indicate that femaleheaded families and families with elderly members are more likely than other families to remain in poverty once they become poor. This section therefore examines the events likely to be associated with entries into and exits from poverty for persons in such families.

Female—headed Families. For the purposes of this study, a person was considered to be in a female—headed family in any month in which his or her family head was a woman rather than a married couple or a man. This determination was made at the family rather than household or sub-family level. As a result, for example, a young woman living with her own child and her parents, a married couple, would typically not have been counted as a member of a female—headed family in these data, since in most cases the parents would be counted as the family heads. The same young woman living with an unrelated married couple, however, would be counted as a female family head, even if the other couple were considered the household heads. 11

Table 4, which shows the impacts of the six life events discussed in the earlier part of this paper on poverty transitions experienced by those in female-headed and non-female headed families, indicates that those in female-headed families are indeed disproportionately likely to suffer certain adverse events—and are even more disproportionately likely to become poor. For example, although persons in female-headed families account for only about 20 percent of the sample in an average month, they experience 63 percent of the divorces and other marital disolutions. 12

If they do suffer a marital break-up, those in female-headed families are much more likely than those in other families to become poor: almost 31 percent of those in female-headed families experiencing a divorce or other break-up become poor in the same month, compared to only about 10 percent of those in other families. This discrepancy could occur for a

Table 4

Percentage of Persons in Families with Occurrences of a Major Life Event Who Experience a Change in Poverty Status, by Female Headship

	Total Number of Persons with Event (in Thousands)	Persons with Event Who Become Poor This Month	Persons with Event Who Leave Poverty This Month
Persons in Female-	Headed Families Wit	h:	
New Baby	253	24.6	1.7
Death	17	<u>.</u>	
Marriage	20	13.3	3.8
Marital Break-	up 282	30.6	1.7
Job Loss	1,826	26.3	0.8
Job Gain	2,067	2.5	23.5
Persons in All Other	er Families With:		
New Baby	1005	9.9	1.2
Death	125	1.3	0.8
Marriage	423	6.1	14.2
Marital Break-	. 165	10.1	3.2
Job Loss	10,078	14.9	1.0
Job Gain	11,125	1.1	14.8

Source: Calculated from the Survey of Income and Program Participation. Figures shown are averages for the months Sept. 1983 through Dec. 1984. See text for definitions of events and populations shown.

number of different reasons, some of which may have relatively short term impacts. For example, in one-earner households that experience a divorce, the earner is most likely to have been the husband, and his departure may substantially reduce family income for a month or two until the woman finds a job and/or arrangements for child support or alimony payments are made. (Of course, in some cases of this type the spell of poverty will last substantially longer—for example, if the woman cannot earn a enough to raise the family income above the poverty level, and her ex-spouse does not provide support payments sufficient to make up the difference.)

Although marital break-ups are somewhat more likely than job losses to be associated with entries into poverty for female-headed families who experience each of these events, job losses arestill much more common and account for a much larger share of all poverty entries for this subgroup. While members of female-headed families are actually somewhat less likely than the rest of the population to experience a job loss—they make up 20 percent of the total population but account for only about 15 percent of all job losses in an average month. Those that do experience such a loss are much more likely to become poor than are those in non-female-headed families in the same circumstances. About one-fourth of those in female-headed families enter poverty in this situation, compared to about 15 percent of those in other families. This difference probably stems in part from female-headed families being less likely than other families to have a second earner, and in part from their generally lower incomes from other sources.

The arrival of a new baby is also more likely to result in an entry into poverty for female-headed families than for other families—almost one-fourth of those in female-headed families with a new baby become poor, compared to one-tenth of those in other families, even though the overall

likelihood of a birth is about the same for both groups. Finally, even even ts not generally associated with becoming poor appear to be relatively likely to result in entrances into poverty for those in female-headed families. For example, 13 percent of those in female-headed families experiencing a marriage entered poverty in the same month. This apparently anomalous—and rare—result stems from the fact that such marriages cannot involve the family head and may thus add to needs without adding to resources. 14

Although female—headed families are more likely to enter poverty when suffering an adverse event, they are not in general more likely to leave poverty with a favorable event. The exception is job gains—more than 23 percent of those in female—headed families who find jobs leave poverty as a result, while only about 15 percent of those in other families do so. (Of course, female—headed families are much more likely to have been poor in the first place, and so have a greater opportunity to leave poverty under these circumstances.)

Unfortunately, the figures shown in Table 4 do not really indicate how many female—headed families leave poverty as the result of marriage, since as discussed above those who marry are in general no longer in female—headed families. In our future work on this topic, we hope to examine the impacts of marriage on those who were in female—headed families in the month before the marriage.

Families with Elderly Members. Families with elderly members account for about 16 percent of the population as a whole, but for a significantly smaller proportion of all events except deaths. Further, as Table 5 shows, those events that do occur are much less likely to move such families into or out of poverty (again with the exception of deaths). Job losses and gains are still the most likely events, even for families with elderly, but

only about 10 percent of those with such losses or gains enter or exit from poverty (as appropriate) in the same month, compared to about 17 percent of those in families without elderly. Marital break-ups (which include widowhood) are more likely to result in poverty for the elderly, with 19 percent of those experiencing such break-ups becoming poor. The absolute numbers of these break-ups are relatively small, however.

Although families with elderly members experiencing a death are more likely to enter powerty than those without a death, the proportion who do so is still very small—only about one and one—half percent. As discussed earlier, we suspect that there may be some problems with the coding of the variable indicating deaths, however, which is currently unedited. Among families with elderly members, a substantially larger proportion—over 4 percent—of those becoming poor in a given month were in families that experienced a death in either the current or immediately previous month. This was virtually the only case in which including those with the event in the previous month as well as those experiencing it in the current month increased the proportion becoming poor. In this case, it probably reflects the fact that income may be recorded for family members even in the month in which they die, so the full impacts of many deaths may not be felt until the next month.

Overall, the only event associated with leaving poverty for families with elderly members was employment. Although over 10 percent of those in families with elderly which included a job gainer left poverty in the same month, members of such families were only about half as likely to be job gainers as were those in families with no elderly. Overall, the data shown in Table 5 confirm the view that, although the elderly are somewhat less likely to enter poverty as the result of the types of events discussed in this paper, once they become poor they are much more likely to remain that way.

Table 5

Percentage of Persons in Families with Occurrences of a Major Life Event Who Experience a Change in Poverty Status, by Presence of an Elderly Person in the Family

	Total Number of Persons with Event (in Thousands)	Persons with Event Who Become Poor This Month	Persons with Event Who Leave Poverty This Month
Families with Elde	rly with:		
New Baby	24	6.7	0.2
Death	86	1.4	-
Marriage	11	3.3	3.7
Marital Break-	ար 63	19.4	
Job Loss	1,083	10.1	0.7
Job Gain	1,057	0.3	10.5
Families with No E	lderly with:		
New Baby	1,234	13.0	1.3
Death	56	0.7	1.7
Marriage	432	6.5	14.0
Marital Break-	up 384	23.6	2.6
Job Loss	10,821	17.3	1.0
Job Gain	12,135	1.4	16.7

Source: Calculated from the Survey of Income and Program Participation. Figures shown are averages for the months Sept. 1983 through Dec. 1984. See text for definitions of events and populations shown.

Conclusion

In conclusion, this paper has found significant correlations between the life events examined and transitions into and out of poverty. Overall, about half of the transitions observed occur in the same month as one of the six events examined. Job losses and job gains account for the vast majority of these transition-related events, but this is largely because such events are much more common, both among those making transitions and among those remaining in the same poverty status, than are the demographic events considered.

Certain demographic events are also very likely to be associated with poverty transitions. Perhaps most notably, 23 percent of those experiencing a marital break-up, and 13 percent of those in families with a new baby, become poor in the month in which the event occurred. By comparison, a job loss is associated with entry into poverty for about 17 percent of those in families with such losses. Job gains and marriages account for almost the same percentage of exits from poverty—16 percent and 14 percent, respectively—for those whose families experience them.

Although our findings so far indicate a strong association between experiencing one of these events and entering or leaving poverty, we believe that much further research in this area could fruitfully be done. For example, we have yet to explore the duration of poverty spells, and its association with specific entry and exit events. Further, it would be interesting to examine more closely those who experience important events and do not have a change in poverty status in the same month—it seems likely that at least in some cases the impacts of the event may be slightly delayed, resulting in a change in poverty status a few months later. We hope to consider some of these topics in our future research.

Appendix Table A1. MEIGHTED SIPP POPULATION CONTROL TOTALS BY MONTH (METH CELL COLUTE)

HONTH		YEAR	SIPP POP	NOT POOR	POOR	NEN POOR	New Not Poor
SEPTEMBER		1963	230030454.	194424225.	34406428.	0.	0.
	CELL	COUNTS	(52759)	(44366)	(6393)	(0)	(0)
	X OF	• • •		04.23	15.77	0.00	0.00
OCTOBER		1965	231043061.	1930344 37.	37208644.	5412797.	4415554.
		COUNTS	(\$2331)	(43863)	(8448)	(1207)	(1002)
	X OF			83.90	16.10	2.34	1.91
NOVEMBER		1963	231231553.	195286105.	35945449.	4399012.	6525 019.
		COUNTS	(52222)	(44051)	(8171)	(1004)	(1220) .
	X OF			84.45	15.55	1.90	2.39
DECEMBER		1965	231343321.	195254499.	34088822.	5905428.	5429795 .
		COUNTS	(\$2072)	(43891)	(8181)	(1317)	(1200)
	X OF			84.40	15.60	2.55	2.35
JANGUARY		1984	231742554.	196516729.	37225025.	6481072.	5023462.
		COUNTS	(51902)	(43535)	(8447)	(1524)	(1145)
	X OF	• • •		83.94	16.06	2.80	2.17
PEBRUARY		1904	231051114.	196685839.	35165275.	4275016.	5576873.
		COUNTS	(51884)	(43918)	1 7966)	(986)	(1263)
	X OF	• • •		84.83	15.17	1.84	2.41
MARCH		1904	232008746.	197840430.	34168116.	4632718.	B 215653.
		COUNTS	(51614)	(43%1)	(7653)	(1042)	(1190)
	X OF			85.27	14.73	2.00	2.25
PRIL		1984	232176787.	197426165.	34750622.	5179735 .	3861958.
		COUNTS	(51262)	(43529)	(7733)	(1164)	(877)
	X OF	POP		85.03	14.97	2.23	1.66
MAY		1904	23230%50.	199449836.	32859822.	3975543.	5413122.
		COUNTS	(508 03)	(43562)	1 7241)	1 9291	(1203)
-	X OF	POP		85.84	14.14	1.71	2.33
ME		1904	232556770.	199773321.	32783449.	5301410 .	4736287.
		COUNTS	(50599)	(43406)	(7193)	1 1200)	(1062)
	X OF	POP		85.90	14.10	2.28	2.04
JLY		1984	232652094.	199489671.	33162424.	5 491114.	4607323.
	CELL	COUNTS	(50350)	(43095)	(7255)	(1235)	(1009)
	X OF	POP		85.75	14.25	2.36	1.98
NUBUST		1904	232871996.	201262261.	31609735.	4930371.	5814777.
		COUNTS	(50104)	(43207)	(6897)	(1133)	(2269)
	% OF	• • •		84.43	13.57	2.12	2.50
BEPTEMB ER		1904	233102122.	199739853.	33362268.	6865826.	4442895.
		COUNTS	(49961)	(42734)	1 72471	F 25481	(996)
	X OF			85.69	14.31	2.95	1.91
OCTOBER		1904	233276048.	200285318.	32991530.	4811177.	4476152.
		COUNTS	(49903)	1 427791	(7124)	(1048)	(998)
	X OF	• •-•		85.86	14.14	2.06	1.92
NOVEMBER		1904	233470397.	200295157.	33175240.	\$411605.	4470604.
		COUNTS	(47709)	(400%)	(6815)	(1095)	(920)
	X OF			85.79	14.21	2.32	1.91
DECEMBER		1904	233601700.	199730527.	33071173.	550 7291.	3963660.
		COUNTS	(45479)	(30956)	(6673)	(1003)	(768)
	X OF	POP		85.50	14.50	2.36	1.70
AVERAGE H			232254337.	197830911.	34423426.	4911772.	4560046.
		POP		85.18	14.82	2.11	1.96

Table A2. Hetented count of pensons in factires exertencing each type of event, for factires which are

Column C		MA	HEN BABY	DEATH	Marriage	Marital Break-up	Job Gafn	Joh Toes
CHI COLMTS 1975 197 19								
Coll Coll	C. Lenkin		46756.	•	25752.	31080.	1307669	217744
Coll Courts	5	FLE COUNTS		<u>-</u>	-	7		
CELL COLORTS (1943) 6.0 (1947) 1007951. 1193246. Z V FOUR (1944) 2044 7.0 (1947) 119424. 119434.	~	2	0.13	9.0	0.07	•		
CELL COLUTY CELL		787	245335.	ď	2447			
Coli Colores Colores Coli Colores Colo	2	il courts	605				1333690	- WIRZ
The court The	*	2 5	3.					
CELL COLONTS C. F.	_	1965	250453	7016	70.00		7.5	62.0
The court The	8	_	(25)		30033	-50274	1154660.	X12070
THE COURTS (1989) (1987) (1989	*	_						2
CELL COUNTY TO FORD	JANEMEY		-			6.11	2.2	
THE COLOR STATE OF THE COLOR STATE S	8	Al captre			101192	19354	1106430.	377264.
THE COUNTY OF TH						121	. 247	1948
CELL COLANTS (FFEERINGS			6.19	0.27	0.13	2.X	
The court The		•	349003		51152.	133311.	1292389.	2468287.
CELL COLONTS 198555. 0.10 0.15 1.5 1.70740. Z	3.	•		(21)	2	(22	(301)	658
CELL COLORTS 1989 CELL COLORTS 1999 CELL				*.·	9.18	. S	3.8	7
Cold			MESSION.		45601.	£336.	1170748.	PIDDEPE
CTI COMPS	5,	it counts	2	2	=	191	20%	
CELL COLUMN		1	1.13	0.03	6.13	0.19	3.43	
COUNTY 1.17 1.51 2.61			407352.	FA359.	101027.	172156.	1123914	9618700
CELL COLMTS CENT CELL COLMTS	B :	•	-	181	192	39)	766	L KKK
CELL COMPTS CELL			1.17	6.13	0.29	95.		200
TOTAL COUNTY TO			405484	•	42181.	129417.	131767	9464749
CELL COLUMNS C. 753 C. 60 C. 13 C. 15 C. 1	B :	_	- 83	-		(22)	1318)	
CELL COLMTS 731 21330 76034 219245 185990 186900 1960	~		1.23	8.	0.13	0.39	6 . 9	
TOTAL COUNTY TO	¥		296479.	21330.	76039.	219265.	1539900.	2782469
THE COLOURS 1.00 0.07 0.23 0.67 0.70 TO SELL COLOURS 1.00 0.01 0.10 0.67 0.70 TO SELL COLOURS 1.00 0.01 0.10 0.64 0.70 TO SELL COLOURS 1.00 0.01 0.10 0.70 0.70 TO SELL COLOURS 1.00 0.10 0.10 0.70 0.70 TO SELL COLOURS 1.00 0.10 0.10 0.10 0.10 TO SELL COLOURS 1.00 0.00 0.10 0.10 0.10 TO SELL COLOURS 1.00 0.00 0.10 0.10 0.10 TO SELL COLOURS 1.00 0.00 0.10 0.10 TO SELL COLOURS 1.00 0.10 TO SELL COLOURS 1.00 0.10 TO SELL COLOURS 1.00	3,		(2)	; ;		(2)	346	(529)
CELL COLMTS 1 1 23 6 511 21915. 144999. ST OF POP 1.09 0.01 2.3 6 511 1.59 ST OF POP 1.09 0.01 2.3 6 511 4.75 CELL COLMTS 1.75 0.0 0.25 0.76 4.45 CELL COLMTS 1.75 0.0 0.25 0.76 4.45 CELL COLMTS 1.75 0.0 0.25 0.76 4.45 CELL COLMTS 1.17 0.0 0.25 0.76 4.45 CELL COLMTS 1.17 0.0 0.17 0.39 4.37 CELL COLMTS 1.17 0.0 0.3 4.37 CELL COLMTS 1.10 0.0 0.3 4.31 CELL COLMTS 1.10 0.0 0.3 4.31 CELL COLMTS 1.00 0.0 0.3 0.3 CELL COLMTS 1.00 0.0 0.3 0.3 CELL COLMTS 1.00 0.0 0.0 0.0 CELL COLMTS 1.0			8.	0.07	0.23	0.67	2.	8.2
TOTAL COUNTY TO			200002	3928.	61021.	221915.	1643030.	3207747
1960 1960	3 3	•			(23)		1698	(721)
CELL COMPTS 1327 0 78359. 233790. 1411144. Z OF POP 1.75 0.00 0.25 7.74 4.95 1.3031			D .1	10.0	. 10	0.67	4.4	19.6
Tot for 1.75 0.00 0.25 0.76 0.46 11075. 1437815. 1446. 11075. 1437815. 1446. 11075. 1437815. 1446. 11075. 1437815. 1446. 11076. 1446. 11076. 1446. 114787. 1446. 1457815. 1457815. 1446. 1457815. 1446. 1457815. 1446. 1457815. 1446. 1457816. 1446. 1457816. 1446. 1457816. 1446. 1457816. 1446. 1457816. 1446. 1		•		•	78359.	233990.	1411166.	3175437.
FELL COLMTS TO FOR T	i N		757	5 ((82		3041	700
CELL COLMTS (****) (*****) (*****) (*****) (*****) (*****) (******) (******) (******) (******) (******) (*******) (*******) (*******) (********) (********) (********) (********) (**********) (***********) (***********) (************) (*************** (******************* (******************** (*********************** (*************************** (********************************* (************************************	_				6.75	2. 7	***	2.8
X OF POP 1.17 0.01 0.17 0.39 4.31 1964 325073 11636 119570 114393 1532344 21 CELL COUNTS 6.69 2 1 35 1 22 1 341 X OF POP 6.99 6.64 6.36 6.35 4.64 1 70 6.99 6.90 6.35 6.35 4.64 1 70 6.9 6.9 6.35 6.35 4.64 2 6 FOP 1 101 6.1 6.1 3.3 6.64 4.23 2 7 FOP 1 101 6.1 6.1 3.3 6.26 6.25 1.28 6.28 1.28 6.28 1.28 6.28 1.28 6.28 1.28						151075.	1437015.	+1474.
CELL COLMTS (6.6) (2) (35) (22) (341) Z OF POP (6.9) (2) (35) (22) (341) Z OF POP (6.9) (2) (35) (22) (341) CELL COLMTS (101) (6.) (6.) (33) (209) Z OF POP (101) (6.) (6.) (33) (209) Z OF POP (101) (6.) (6.) (33) (209) Z OF POP (101) (6.) (6.) (33) (209) CELL COLMTS (77) (31) (23) (22) (204) Z OF POP (11) (6.) (6.) (32) (204) Z OF POP (11) (6.) (6.) (32) (204) Z OF POP (11) (6.)	×	_	1.17		117		(225)	110)
CELL COLMTS (64) (2) (35) (22) (341) Z OF POP (6.79 (6.36 (6.35 (6.24) 6.35 (6.44) 6.	OCTUBER	*	325073.	11676	11011	114101	4.51	12.42
Z OF POP 0.99 0.04 0.36 0.35 4.64 1344 453617. 26216. 23472. 161059. 1603990. CELL COLMTS 1.011 6.1 6.1 33 1.2899 Z OF POP 3771 1393. 0.09 0.07 0.49 4.23 CFLL COLMTS 1.771 3.1 2.3 1.2899 1.2899 CFLL COLMTS 1.771 3.1 2.3 1.25 1.2899 CFLL COLMTS 1.771 3.1 2.3 1.25 1.25 X OF POP 1.11 0.04 0.26 1.322445. 1326110. X OF POP 1.05 0.06 0.10 0.10 0.10 0.10 3.05	8	IL COUNTS	3	7	186		261	./27662
CELL COLMTS (101) (61) (61) (53) (200) (7.28	*	2 2		5	98'0	97.0		26.
CELL COLMTS (101) (61) (53) (289) 2 OF FOR 1.37 0.00 0.07 0.49 4.23 1904 877113. 13903. 0.6590. 17645: 1023904. CELL COLMTS (77) (31 25) (52) (204) X OF FOR 1.11 0.00 0.05 0.10 0.30 3.05 X OF FOR 1.03 0.06 0.10 0.30 3.05	MOVERER	Į	483617.	26210.	23472.	161859.	1467690	107874
X OF FOR 1.37 0.00 0.07 0.49 4.23 1904 1904 1905 1905 1905 1904 1905 1905 1905 1905 1905 1905 1905 1905	B :			;	3	33)	1 2891	
CELL COLATS 17713. 13703. 06570. 176463. 1025906. X OF POP 1.11 0.04 0.26 0.52 3.02 **SOF POP 1.03 0.05 0.10 0.36 3.05 **SOF POP 1326410.	2		1.37	3.	0.07	6.43	4.23	
FOR 1.11 0.04 0.26 0.52 3.02 1.02 1.02 1.02 1.02 1.02 1.02 1.02 1			577113.	13903.	96 576.	176463.	1023906.	P006172
38291. 18629. 60842. 132445. 1326110. FOP 1.05 0.05 0.18 0.38 3.05	B.	_	122		183	125	702	
POP 11.03 0.05 0.10 0.34 3.05 3.05		2	1.11	\$	9.36	0.52	3.05	\$
3.05	AVERAGE HONT	***************************************	382941.	10029.	60642	122446.	129431	
	×		1.03	90.0	• 7	5		
					In the second se)	•	

Table A3.

NETBITTED COUNT OF PERSONS IN FAMILIES EXPENIENCING EACH TYPE OF EVENT, FOR FAMILIES WITCH ARE NOT POOR

	HENULIN ALTON	HEN BABY	BEATH	Marriage	Break-up	Job Gain	Job Loss

Ę		79249.	17006.	51205.	24689.	. 1154669.	**************************************
_	are courts	=======================================	-	13)	3	1 26451	1 2062)
		\$	9 .0	0.03	1 .	2.5	4.67
		229769.	. Bette.	196534.	176972.	10691374.	7312E.
_	CELL COUNTS	(18)	121	(59	\$	143	19591
	\$ 15 ×	0.12	0.03	9.79	•	2.47	3.01
EVERER	1	293486.	110550.	. 10002	15042.	10267048.	9373046.
_	CELL COLATS	122	(82)	- 2	28	(2316)	(2135)
	204 20 X	9.15	3.0	0.12	8.	5.26	3.4
2	1	637421.	390243.	194262.	303171.	11236230.	9666799.
	CELL COLOUTS	(218)	-	?	(22)	. (2556)	(2205)
	X 07 70°	0.43	0.20	0.10	9.16	5.2	4.97
	***	A62334.	267503.	240382	280407.	11634760.	6740452.
	CELL COLOURS		9	102)	5	(2679)	1581)
	2 07 800	0.43	•17	0.12	0.15	5.42	3.43
2		1016277.	210152.	132479.	288993.	11271428.	6682247.
•	CPLL COLATS	(262)	-		2	(2550)	1533)
				D. 17	9.18	F. 20	3
E	1		077K2	Tenes.	240200	10145012	4597064
•		1961		1331		(2333)	15271
	2 OF POP	7		0.20	6.19	5.16	3.34
	100	1016723.	114934.	436066	. 269698.	13334359.	7210433.
	כבור כמשעצ	1242	92	156)	3	3048	(1617)
	X 05 700	0.51	90.0	0.22	0.13	6.6	3.62
	-	1051302.	96350.	5759M.	342056.	16329393.	10043684.
	CELL COURTS	(822)	102	(202)	(83)	1 36761	(\$230)
	% OF 70	9.52	90.0	0.23	0.17	0.17	5.03
>	***	091237.	272	. 13909	B06 272.	1300041.	10578097.
	CELL COUNTS		181	(902)	1011	(2902)	1648
	2 5 %	97.0	3.	0.30	97.0	6.52	5.30
15	**	125A79.	115120.	£63697.	279546.	13173171.	11307966.
	כבור כסושב	-	22	195	. (22)	2968	10093
	\$ 50 X	9.62	3.	27.0	•	• . E	29.5
TENERA	***	1273780.	101661.	61047.	.190439	14191410.	14334698.
	כבור כסחעב	(502)	(23)	(102)	157	1 3211)	(3213)
	2 2 2	3.0	90.0	0.31	0.31	7.10	7.18
Tabel	Ĭ	1062107.	#1%.	42524	46910.	11336231.	711221.
	CELL COUNTS	(822)	2	152)		1 2545)	
			20.0	12.0	22.0		2.5
	*	1502892.	. 90106.	411422	260482	7/48//1.	754020
	כבור כפראנצ	1 582					
		£.0	5	12:0	61.10 1.10 1.10 1.10 1.10 1.10 1.10 1.10		20.6
CEPTER		-0120421	.0216		3000/6		
	CELL COUNTS	1862					
		***	5	5.		3	
-			195176	- Atmost	11666	11046706	2763000

Table A4. Netwite court of persons in families experiencing each type of event, for families mitch are New Poor

Martin Martin				Maritral		•	
	MEN BARY	DEATH	Marriage	Break-un	Tob Oct	•	
OCTUBER 1965	04546				UTB: DOC	Job Loss	
CELL COLOURS		•	12918.	31000.	124.00		
404 40 X		-	S (2		.160010	
HOVETBER 1983	132790.		2	0.57	4. %	27.00	
CELL COLATS		•		41014.	42959.	1426	
DECEMBER 2 DECEMBER 2			21.0			(457)	
CELL CRAME		•	3991.	17411		43.06	
X 04 POP		0	2		. 22GD41		
JANEARY 1904			0.0	0.63	185 2. 66	(266	
CELL COLATS		<i>:</i>	32532.	3646.	101269.	2627105	
SOL SO Y	8.8				- 42	(603)	
Tell Contract		17972.	2667	3.5	8.8 8.8	1.2	
2 OF 900		-	121	23)	155281.	1655151.	
MARCH 1994		0.4 5		2.69	125		
CELL COLMES	(32)		29711.	43968.	133679.	34.72	
% OF POP		5 8		121 -	62	101	
1961 1964	232334.			. 3	8.09	26.63	
י פנ פני		-		106912.	200347.	1006548.	
PAY 1996	4.4		3.	. 70.2	(6E -	1395)	
CELL COUNTS		• • • • • • • • • • • • • • • • • • •	17523.	106628.	270753		
404 bo ×			3	321	(22)	399	
1961	103142.	•	K 20.5	2.69 2.69	6.01	43.01	
Z OF BOD		-	181	.620102	214926.	1650066.	
JARY 1986		8	1.8	2.2		161	
CELL COLATS	. 12)(CI	3928.	19227.	190696.	173319.	51.13	
404 PO X		-		43)	(24)	630)	
1861 ASTRON		•		2.47	3.16	43.39	
A GO BOB		•			2528A2.	2011055.	
SEPTEMBER 1904	98.57	8.	1.02	4.15		163	
CELL COUNTS			19196.	125436.	214772.	2623344	
CTUBER X OF FOR		8.0	9.24		(19)	(675)	
CELL CONTR	142260.	•	36368.	101270.	5.13	30.21	
204 20 X		= { -	131	191	\$.08/9201	
HOVENBER 1984	166073.		2.0	8.11	3.67	\$7.47	
CELL COURS	124			74301.	150445.	1996287.	
404 40 %	3.07			(91,	192	306	
TAIL COLUMN	198961.	•	52177.	199406	2.7	36.09	
2 06 BOS		•			PILL ,	175074.	
	3	8.		2.3	2.16	1 20	
AVERAGE PROFITS	162933.	1500.	20170				
	3.30	0.03	0.50	2000X	170254.	1900568.	
					ř	25.04	

Table A5. METONTED COUNT OF PERSONS IN FAULIES EXPERIENCING EACH TYPE OF EVENT, FOR FAULIES MICH ARE NEW NOT POOR

Columbia True Tru						Marital			
11500. 11500. 0. 0. 11500. 0. 0. 11500. 0. 0. 11500. 0. 0. 0. 0. 0. 0. 0.		3	HEM BABY	DEATH	Marriage	Break-up	Job Cath	Job Loss	
100 100	de la								
0.00 0.00 0.00 0.20 0.00 0.00 0.00 0.00			•	•	11507.		2264022	32725A.	
10027.								2	
1,000 1,00		. 5		3.	2.0	8.	***	7.41	
0.00	MANERAL EST		1007	•	244/4.	4978.	2182746.	1	
1.00	8 13		2		2	7	9		
10 10 10 10 10 10 10 10	2 45 %		8. D	8.8	¥.		39.51	6.73	
19065. 0.00 0.16 0.00 155.10 19065. 0.00 0.20 0.20 19065. 0.00 0.20 0.20 19065. 0.00 0.20 0.20 19065. 0.00 0.20 0.20 19065. 0.00 0.20 0.20 19065. 0.00 0.20 0.20 19065. 0.00 0.20 19065.	DECEMBER 19	2	•	•	29524.	•	1905617.	67221.	
100.00 10	כבוו כס	E	-	-	9	-	(413)	(12	
14045. 13034. 0. 13034. 0. 1402074. 0. 15651. 0. 0. 0. 0. 0. 0. 0. 0. 0. 0. 0. 0. 0.	2 6 %		•	8.	3.	60,0	15. 10	17.	
0.30	JANGARTY 19	8	14965.	•	13834.		PA12874	17109	
0.30	כנוו כס			3		;			
1, 1, 1, 1, 1, 1, 1, 1, 1, 1, 1, 1, 1,	20 %					5 8			
131 154	PERSONAL NAME AND ADDRESS OF THE PERSON NAME AND ADDRESS OF TH			}					
9.00 9.00 9.00 9.00 1.15 9.00 1.15 9.00 1.15 9.00 1.15 9.00 1.15 9.00 1.15 9.00 1.15 9.00 1.15 9.00 9.00 9.00 9.00 9.00 9.00 9.00 9.0							./400047		
737. 0. 60000. 1312. 256770. 251				5 8					
10		. 1		3			20.5	Z. 59	
6.09 0.00 1.15 0.26 45.26 1.51 0.0 22177. 10771. 1069343. 1.53 0.00 2.39 0.28 46.92 1.53 0.00 2.39 0.73 46.92 2.60 1.05 0.73 47.10 47.10 6.05 0.26 1.05 0.73 47.10 6.07 0.0 7.26 1.05 47.10 6.07 0.0 1.53 0.06 46.07 6.09 0.0 1.53 0.06 47.10 6.09 0.0 1.53 0.06 47.10 6.09 0.00 1.00 0.00 1.00 1.00 6.09 0.00 0.00 0.00 1.00 1.00 1.00 6.00 0.00 0.00 0.00 0.00 1.00 1.00 6.00 0.00 0.00 0.00 0.00 1.00 1.00 6.00 0.00 0.00 </td <td></td> <td></td> <td></td> <td>;</td> <td></td> <td>13012.</td> <td>.00/042</td> <td>118801.</td> <td></td>				;		13012.	.00/042	118801.	
1.55				5 6			1/29		
151 101 102					1.15	97.0	45.26	2.29	
2972. 14241. E702. 39724. E469712. 10.05 0.26 1.05 0.73 47.10 1.05 0.73 47.10 1.05 0.73 47.10 1.05 0.73 47.10 1.05 0.73 1.05 0.73 1.05 0.73 1.05 1.05 1.05 1.05 1.05 1.05 1.05 1.05		, [42177.	10771.	1009363.	66 137.	
2530. 1972.	מנו נפ					£ .	1624		
6.55		. 1	1.55		2.39	97.0	48.92	8.20	
0.55				19291	.020/6	39726.	249712.	67632.	
16037. 0.7 25.07. 3034. E172675. 0.06 0.34 0.00 1.53 0.06 2530. 0.00 1.53 0.06 2530. 0.00 1.00 0.00 0.00 0.23 47.10 0.00 0.00 0.00 2.03 47.10 0.00 0.00 0.00 2.03 47.10 0.00 0.00 0.00 2.03 47.10 0.00 0.00 0.00 2.03 27.20 0.00 0.00 0.00 2.03 27.20 0.00 0.00 0.00 2.03 27.20 0.00 0.00 0.00 2.03 27.20 0.00 0.00 0.00 2.03 27.10 0.00 0.00 0.00 2.03 1.00 0.00 0.00 0.00 2.03 1.00 0.00 0.00 0.00 0.00 2.03 0.00 0.00 0.00 0.00 2.03 0.00 0.00 0.00 0.00 0.00 2.03 0.00				3 (3 i	(22)	3	
251 11 1 4911 1		. 1	14017		1.09 3.6637	202		1.06	
2530. 2530.	כבור כס	ST-S	-	;	(52)			113677	
2530. 2510. 25	2 25 %		Z.0	8.0	1.53	***	45.07	2.39	
6.06 6.00 6.23 47.10 6.73 67.10 6.00 6.23 47.10 6.00 6.00 6.23 47.10 6.00 6.00 6.23 47.10 6.00 6.00 6.23 47.10 6.00 6.00 6.23 6.23 6.23 6.23 6.23 6.23 6.23 6.23	LL 13	\$	2530.	•	30508	10442.	2170009.	45054	
0.06 0.06 0.07 0.00 0.00 0.00 0.00 0.00			.(2	-	. 101	-	(473)	(21)	
0			8.0	8.	3.0	0.23	47.10	1.8	
0.00		5	•	•	16549.	11386.	2163119.	167972.	
16400. 0.0 0.20 37.20 1.20	8 % X X		5 (-	= ; _	(%) -	2	479)	9	
100				8.	Z . 9 .	0.20	37.20	8.89	
0.37					. 65756.	20007	2550107.	£21697.	
919. (11) (0.02				•					
11	OCTOBER 19	•		3			3441486	12050	
0.02 0.00 1.75 0.00 37.12	02 1130	STATE OF THE PERSON			(92	;	376)	127	
75101. 0. 66277. 6137. 1664746. 1 201 (231 (1) (352) (1) (352) (1) (352) (1) (352) (1) (352) (1) (352) (1) (352) (1) (352) (1) (1) (1) (1) (1) (1) (1) (2 25 ×		0.0	9.0	1.78	0.0	37.12	26.2	
1.66 (0.00	HOVEVEER 19	\$	75101.	•	88 277.	6137.	16676	\$693.	
1.64 0.00 1.91 0.14 37.73 1.92 1.93 0.14 37.73 1.93 1.93 1.93 1.93 1.93 1.93 1.93 1.9			102	3	1 23	1	1 352)	102 1	
13757. 0. 00579. 0. 1502430. 1 3111 0. 1502430. 0. 0. 0. 0. 0. 0. 0. 0. 0. 0. 0. 0. 0	2 6 7	ı.	3.7	8.	1.9	•.15	37.73	8.10	
16251. 449. 60016. 10065. 2133944. 0.36 0.02 46.79	MECENTER 19		15%?		6 1576.	•	1502430.	114461.	
16251. 949. 60010. 1006S. 2133964. 0.36 0.02 1.33 0.22 46.79	2 2 20 %		6.38	8	22.7		1111		
16251, 949, 66010, 10068, 2133966, 100 1.33 0.22 46.79									
0.36 0.02 1.35 0.22	AVERAGE PEDITY		16291.		6618.	10065.	2133964.	119038.	
	2 6 7		9.76	0.0	1.33	0.22	2:3	2.61	

FOOTNOTES

- 1/ Both having a new baby and experiencing a death are specific cases of changes in family size. All such cases cause poverty thresholds to rise or fall, and can therefore lead to changes in poverty status. A more general analysis might consider increases and decreases in family size as events that cause changes in needs—and possibly incomes. These broader events are not examined here, however.
- 2/ The two employment-related events are actually defined on the basis of monthly earnings rather than on reported employment status—a person has been defined as becoming unemployed if prior month's earnings were above \$200 and this month's earnings are below that level, while entry into employment involves moving from earnings of less than \$200 to earnings above that level. The threshold level of \$200 per month was chose somewhat arbitrarily to exclude small amounts of earnings from casual employment activities such as baby-setting. It should be noted that the unemployment variable, in particular, does not necessarily imply continued labor force participation, but rather includes persons who retire, who take maternity leave, and so forth.
- 3/ One-fourth of the sample was interviewed each month concerning events of the previous four months. As a result, only a quarter of the sample would experience a cross-wave transition for a given calendar month.
- 4/ Table 1 is taken from an earlier study by one of the authors of this paper, Roberton Williams, and is based on a slightly different SIPP sample than the rest of the results reported in this paper. In particular, this sample includes 12 months rather than 16 months of data, covering the calendar year 1984 period. It also excludes persons for whom a full year of data was not available. For more information on the specifics of this sample and of the methods used to calculate these four poverty measures, see Roberton Williams, "Poverty Rates and Program Participation in the SIPP and the CPS," paper prepared for presentation at the 1986 Joint Statistical Meetings, August 1986.
- 5/ See for example Mary Jo Bane and David Ellwood, "The Dynamics of Dependence and the Routes to Self-Sufficiency," Final Report to the U.S. Department of Health and Human Services, Harvard University, Cambridge, Mass. 1983, and Greg Duncan, Richard D. Coe, and Martha S. Hill, "Some Dynamic Aspects of Poverty," in Greg Duncan, ed., Years of Poverty, Years of Plenty, Institute for Social Research, Ann Arbor, Michigan 1984.
- 6/ Individuals were defined as "becoming poor" in a given month if their family incomes were below the relevant poverty threshold in that month, but above the thresholds in the previous month. Similarly, persons "leaving poverty" in a given month were defined as those who were poor in the previous month but not poor in the current month.
- 7/ Our data indicate that about 11 percent of the sample experience one of these six events in any given month, while about one-half of one percent experience two events. Less than one-tenth of one percent experience three events, and no one experienced four.

FOOTNOTES CONTINUED

- 8/ Families becoming poor in any given month are of course a subset of the poverty population as a whole for that month, so that the relatively high incidence of job loss, births, and marital break-ups seen for the newly poor may in fact explain the differences between the poor and non-poor populations as a whole.
- 9/ In some cases, marital break-ups will in effect cause job losses that will not appear as such in these tables, since they may result in an exployed person leaving the family. As defined here, job losses are attributed only to families that contain a person who was employed in the previous month.
- 10/ Because our poverty definition looks only at current month income, it is somewhat difficult to tell from our data (at least as now tabulated) exactly who was poor in the month prior to the occurrence of a given event. The estimates given here are therefore approximations only, based on the average number of transitions in both directions for those experiencing each event, as well as on the average number of persons experiencing each event and remaining poor. These averages do not take into accout month to month changesin family composition that could also affect these results, however.
- 11/ Family headship is in many cases a somewhat arbitrary designation, and extended-family households may be treated somewhat differently in the data depending on who in the household was interviewed.
- 12/ Since family composition is based on data for the month in which the event in question occurred, any woman who became a family head as the result of a divorce would be classified as a member of a female-headed family in that month, as would her children if they continued to live with her rather than her ex-spouse.
- 13/ In all likelihood, however, this lower percentage of job losses results from the fact that members of these families are significantly less likely to be employed in the first place.
- 14/ If the family head were to marry, the family would no longer be female-headed.

 Any marriage within female-headed families must thus have involved a dependent.

 Indeed, the very small number of marriages seen in such families may in itself explain this finding, since the sample of such marriages, even across all 16 months of data, cannot be very large.