## Appendix C. Chinook salmon bycatch-at-age methods and evaluation

### 1.1 Introduction

Currently, accurate in-season salmon abundance levels are unavailable and management must rely on analyses of historical data for developing alternatives. Developing regulations designed to reduce the impact of bycatch requires methods that appropriately relate these impacts to their respective salmon populations. A stochastic "adult equivalence" model was developed that accounts for sources of uncertainty. This extends from Witherell et al.'s (2002) evaluation and relaxes a number of assumptions. Such stochastic simulation approaches for evaluating management measures provide insight on the types of data required to better achieve objectives (e.g., Criddle 1996).

In 2007, the Council reviewed the methodology and encouraged refinements. In particular, these included:
a) Improving estimates of the salmon bycatch age composition,
b) Deriving realistic salmon maturation schedules which consider historical brood-year data,
c) Use of updated genetics information on stock origin,
d) Use of updated run size information, and
e) Refining the adult equivalent model to include a broader range of inputs (e.g., brood-year maturation rates and age specific natural mortality rates)

These updates and revisions were presented at the April 2008 Council meeting where further guidance for refinements was provided. This included explicit seasonal allocation of alternative cap levels and improved estimates of at sea survival. What follows is an update of the methods presented at the April 2008 Council meeting which describes the methods and data used to estimate AEQs and application to seasonal and sector allocations of cap levels currently proposed as alternative management actions.

### 1.2 Methods

Overall salmon bycatch levels are estimated based on extensive observer coverage. For the pollock fishery, the vast majority of tows are observed either directly at sea or based on offloading locations aboard motherships or shore-based processing plants. The observer data is used to allow inseason manager evaluate when to open and close all groundfish fisheries based on catch levels of prohibited species bycatch, such as salmon and halibut, and of target groundfish species. The process of applying observer data (in addition to other landings information) to evaluate fishery season length has relied on a pragmatic approach that expands the observed bycatch levels to extrapolate to unobserved fishing operations. More statistically rigorous estimators have been developed (Miller 2005) that can be applied to the North Pacific groundfish fisheries but these so far have not been implemented for inseason management purposes. Nonetheless, these estimators suggest that for the Eastern Bering Sea pollock fishery, the levels of salmon bycatch are precisely estimated with coefficients of variation of around $5 \%$. This indicates that, assuming that the observed fishing operations are unbiased relative to unobserved tows, the total salmon bycatch levels are precisely estimated for the fleet as a whole. For the purposes of this analysis, imprecision on the total annual salmon bycatch is considered negligible.

### 1.2.1 Salmon catch-at-age estimation methods

In order to appropriately account for the impact of salmon bycatch in the groundfish fisheries it is desirable to correct for the age composition of the bycatch. For example, the impact a bycatch level of 10,000 adult mature salmon would have is likely greater than the impact of 10,000 incidentally caught salmon that just emerged from rivers and expected to return for spawning in several years time. Hence, estimation of the age composition of the bycatch (and the measure of uncertainty) is critical.

[^0]Estimates of both length and age composition and their variance estimates were approximated using at two-stage bootstrap method. For a given year the first stage samples, with replacement, among all tows from which salmon were measured. Given this collection of tows, the individual fish measurements were resampled with replacement and all stratum-specific information was carried with each record. A separate process was carried out on the samples from which age data were collected following a similar two-stage approach. Once a sample of lengths and ages were obtained, age-length keys were constructed and applied to the catch-weighted length frequencies to compute age composition estimates. This process was repeated 100 times and the results stored to obtain a distribution of both length and age compositions.

Three years of length-at-age data were available from Myers et al. (2003). These data are based on salmon scale samples collected by the NMFS groundfish observer program from 1997-1999 and processed for age determination (and river of origin) by scientists at the University of Washington (Table 1). Extensive salmon bycatch length frequency data are available from the NMFS groundfish observer program since 1991 (Table 2). The age data were used to construct age length keys for nine spatiotemporal strata (one area for winter, two areas for summer-fall, for each of three fishery sectors). Each stratum was weighted by the NMFS Regional Office estimates of salmon bycatch (Table 3). To the extent possible, sex-specific age-length keys within each stratum were created and where cells were missing, a "global" sex-specific age-length key was used. The global key was simply computed over all strata within the same season. For years other than 1997-1999, a combined-year age-length key was used (based on all of the 1997-1999 data). This method was selected in favor of simple (but less objective) length frequency slicing based on evaluations of using the combined key on the individual years and comparing age-composition estimates with the estimates derived using annual age-length keys. The reason that the differences were minor are partially due to the fact that there are only a few age classes in the salmon bycatch and these are fairly well determined by their length at-age distribution (Figure 1).

### 1.2.1 Genetics sample composition

Scientists with Alaska Department of Fish and Game have developed a DNA baseline to resolve the stock composition mixtures of Chinook salmon in the Bering Sea (Templin et al. In prep.). This baseline includes 24,100 individuals sampled from over 176 rivers from the Kamchatka Peninsula, Russia, to the Central Valley in California (Table 4). The genetic stock identification (GSI) study used classification criteria whereby the accuracy of resolution to region-of-origin is must be greater than or equal to $90 \%$. This analysis identified 15 regional groups for reporting results. For this report, minor components in the bycatch are combined into the "other" category for clarity which results in a total of 9 stock units.

This study analyzed samples taken from the bycatch during the 2005 B season, both A and B seasons during 2006, and a sample from an excluder test fishery during the 2007 A season. Where possible, the genetics samples from the bycatch were segregated by major groundfish bycatch regions. Effectively, this entailed a single region for the entire fishery during winter (which is typically concentrated in space to the region east of $170^{\circ} \mathrm{W}$ ) and two regions during the summer, a NW region (west of $170^{\circ} \mathrm{W}$ ) and a southeast region (east of $170^{\circ} \mathrm{W}$ ). The genetic sampling distribution varies considerably by season and region compared to the level of bycatch (as reported by NMFS Regional Office; Table 3).

The samples used in the analysis were obtained during a feasibility study to evaluate using scales and other tissues as collected by the NMFS observer program for genetic sampling. Unfortunately, during this feasibility study, the collected samples failed to cover the bycatch in groundfish fisheries in a comprehensive manner. For example, in 2005 most sampling was completed prior to the month (October) when most of the bycatch occurred (Figure 2).

For the purposes of assigning the bycatch to region of origin, the level of uncertainty is important to characterize. While there are many approaches to implement assignment uncertainty, the method chosen
here assumes that the stratified stock composition estimates are unbiased and that the assignment uncertainty based on a classification algorithm (Seeb and Templin, In Prep; Table 6) adequately represents the uncertainty (i.e., the estimates and their standard errors are used to propagate this component of uncertainty). Inter-annual variability is also introduced in two ways: 1) by accounting for inter-annual variability in bycatch among strata; and 2) by using the point estimates (and errors) from the data (Table 6) over the different years (2005-2007) while weighting appropriately for the sampling intensity. The 2005 B-season results were given one third of the weight since sampling effort was low during October of that year (relative to the bycatch) while the 2006 B-season stock composition data was given two-thirds of the weight in simulating stock apportionments. For the A season, the 2007 data (collected from a limited number of tows) were given one fifth the weight while the 2006 was weighted 4 times that value.

The procedure for introducing variability in regional stock assignments of bycatch followed a Monte Carlo procedure with the point estimates and their variances used to simulate beta distributed random variables (which have the desirable property of being bounded by 0.0 and 1.0 ) and applied to the catch weightings (for the summer/fall (B) season) where areas are disaggregated. Areas were combined for the winter fishery since the period of bycatch by the fishery is shorter and from a more restricted area.

### 1.2.2 Estimating adult equivalence and impact rate

The impact of bycatch on salmon runs is the primary output statistic. This measure relates the historical bycatch levels relative to the subsequent returning salmon run $k$ in year $t$ as:
$u_{t, k}=\frac{C_{t, k}}{C_{t, k}+S_{t, k}}$
where $C_{t, k}$ and $S_{t, k}$ are the bycatch and stock size (run return) estimates of the salmon species in question. The calculation of $C_{t, k}$ includes the bycatch of salmon returning to spawn in year $t$ and the bycatch from previous years for the same brood year (i.e., at younger, immature ages). This latter component needs to be decremented by ocean survival rates and maturity schedules. This sum of catches (at earlier ages and years) can thus be represented as:

$$
\begin{equation*}
C_{t, k}=\sum_{a=1}^{A} c_{i, a, k} s_{a} \gamma_{a, k} \quad i=t-A+a \tag{2}
\end{equation*}
$$

where $c_{i, a, k}$ is the catch of age $a$ fish in year $i, A$ is the oldest age of their ocean phase, $s_{i, a, k}$ is the proportion of salmon surviving from age $a$ to $a+1$, and $\gamma_{a, k}$ is the proportion of salmon at sea that will return to spawn at age $a$. Maturation rates vary over time and among stocks detailed information on this is available from a wide variety of sources. For the purpose of this study, an average over putative stocks was developed based on a variety of studies (Table 7)

To carry out the computations in a straightforward manner, the numbers of salmon that remain in the ocean (i.e., they put off spawning for at least another year) are tracked through time until age 7 where for this model, all Chinook in the ocean at that age are considered mature and will spawn in that year.

Stochastic versions of the adult equivalence calculations acknowledge both run-size inter-annual variability and run size estimation error, as well as uncertainty in maturation rates, the natural mortality rates (oceanic), river-of-origin estimates, and age assignments. The variability in run size can be written as (with $\dot{S}_{t, k}$ representing the stochastic version of $S_{t, k}$ ):

$$
\begin{align*}
\dot{S}_{t, k}=\bar{S}_{k} e^{\varepsilon_{t}+\delta_{t}} & \varepsilon_{t} \sim N\left(0, \sigma_{1}^{2}\right),  \tag{3}\\
& \delta_{t} \sim N\left(0, \sigma_{2}^{2}\right)
\end{align*}
$$

where $\sigma_{1}^{2}, \sigma_{2}^{2}$ are specified levels of variability in inter-annual run sizes and run-size estimation variances, respectively.

The stochastic survival rates were simulated as:
$\dot{s}_{a, k}=1-\exp \left(-M_{a}+\delta\right), \quad \delta \sim N\left(0,0.1^{2}\right)$
whereas the maturity in a given year and age was drawn from beta-distributions:
$\dot{\gamma}_{a, k} \sim B\left(\alpha_{a}, \beta_{a}\right)$
with parameters $\alpha_{a}, \beta_{a}$ specified to satisfy the expected value of age at maturation (Table 7) and a prespecified coefficient of variation term (provided as model input).

Similarly, the parameter responsible for assigning bycatch to river-system of origin was modeled using a combination of years and "parametric bootstrap" approach, also with the beta distribution:
$\dot{p}_{k} \sim B\left(\alpha_{k}, \beta_{k}\right)$
again with $\alpha_{a}, \beta_{a}$ specified to satisfy the expected value the estimates and variances shown in Table 6. For the purposes of this study, the estimation uncertainty is considered as part of the inter-annual variability in this parameter. The steps (implemented in a spreadsheet) for the AEQ analysis can be outlined as follows:

1. Select a bootstrap sample of salmon bycatch-at-age $\left(\phi_{t, a}\right)$ for all years and strata;
2. Sum the bycatch-at-age for each year and proceed to account for year-of-return factors (e.g., stochastic maturation rates and ocean survival (Eqs. 2-5);
3. Partition the bycatch estimates to stock proportions (by year and area) drawn randomly from each parametric bootstrap;
4. Sum over all bycatch years and compare with run-size estimates for impact rate calculations;
5. Repeat 1-3 200 times;
6. Based on updated genetics results, assign to river of origin components ( $\dot{p}_{k}$, Eq. 6).
7. Compile results over all years and compute frequencies from which relative probabilities can be estimated;

Sensitivity analyses on maturation rates by brood year were conducted and contrasted with alternative assumptions about natural mortality schedules during their oceanic phase as follows:

| Model | 3 | 4 | 5 | 6 | 7 |
| :--- | ---: | ---: | ---: | ---: | ---: |
| 1 - None | 0.0 | 0.0 | 0.0 | 0.0 | 0.0 |
| 2 - Variable | 0.3 | 0.2 | 0.1 | 0.05 | 0.0 |
| 3 - Constant | 0.2 | 0.2 | 0.2 | 0.2 | 0.0 |

Evaluations of alternative Chinook salmon caps were done based on re-casting historical catch levels as if a cap proposal had been implemented. Since the alternatives all have specific values by season and sector, the effective limit on Chinook bycatch levels can vary for each alternative and over different years. This is caused by the distribution of the fleet relative to the resource and the variability of bycatch rates by season and years. To capture the effect of an alternative policy, the 2003-2007 mean "effective" cap for each alternative was computed and used as the seasonal limit for evaluation purposes (Table 8). These values were then used in the AEQ simulation model as season-specific caps. This means that the
minimum of the historical season-specific bycatch and the effective cap level given in Table 8 was applied for estimating the AEQ for each policy.

### 1.3 Results

### 1.3.1 Chinook salmon catch-at-age

The uncertainty in the distribution of seasonal length frequencies have improved over time (Figure 3). Applying these length frequencies (and associated uncertainty based on bootstrap sampling) results in annual totals of Chinook salmon bycatch by age as shown in Table 9 . When broken out by season there is some correlation between B season levels at one age and subsequent A season levels of the next age group (Table 10). Estimates of uncertainty due to age-specific bycatch sampling (for age and length) varied by season but showed some improvement (smaller values of coefficients of variation) for the main bycatch age groups in recent years (Table 11; Figure 4). For the evaluations of uncertainty in age assignments and impact analysis, the bootstrap samples of age composition were used and has the added advantage that the covariance structure is retained (e.g., Figure 5).

### 1.3.2 Chinook salmon bycatch stock composition

Application of GSI to estimate the composition of the bycatch by reporting region suggests that, if the goal is to provide estimates on the stock composition of the bycatch, there need is to adjust for the magnitude of bycatch occurring within substrata (e.g., east and west of $170^{\circ} \mathrm{W}$ during the B season, top panels of Figure 6). Applying the stock composition results presented in Table 6 over different years and weighted by catch gives stratified proportions that have similar characteristics to the raw genetics data (Table 12). Importantly, these stratified stock composition estimates can be applied to bycatch levels in other years which will result in overall annual differences in bycatch proportions by salmon stock region. This approach assumes that the salmon from early years were of similar stock composition, until planned investigations analyzing historical scale samples are complete, the degree of temporal variation in stock composition within season and spatial strata are unknown. These simulations can be characterized graphically in a way that shows the covariance structure among regional stock composition estimates (e.g., Figure 7).

Given the bycatch by strata estimates, it is possible to use the genetic composition data to estimate the historical expected stock proportions. However, this assumes the genetics data collected from 2005-2007 adequately represents the historical pattern. Clearly, it is preferable to have genetics samples for the historical period analyzed rather than assuming the stratumspecific stock composition estimates from the recent period reflect the past. That caveat stated, it is still interesting to note how historical annual bycatch composition varies depending on the locales of where Chinook are taken as bycatch (Figure 8) with median values presented in Table 13. To gain an appreciation of the impact, the Pacific Northwest group (PNW, also noted in some figures as $\mathrm{BC}+\mathrm{WA}+\mathrm{OR}$ ) and the Upper Yukon River annual proportions in the bycatch are strongly affected by the locales and seasons of where the bycatch occurred (Figure 9). Myers et al. (2003) found similar areaspecific patterns in their bycatch.

### 1.3.3 AEQ estimation

Using the weighted mean maturation schedule and the variable age-specific ocean mortality, the adult equivalents due to salmon mortality induced by the pollock fishery averaged about two thirds of the nominal (reported) annual bycatch in recent years (Figure 10). The AEQ model was shown to be sensitive to natural mortality assumptions but had little qualitative difference in the trend over time
(Figure 11). For the stochastic version, under Model 2 assumptions (decreasing mean age-specific natural mortality with age) results show a fair amount of uncertainty in the estimates of AEQ mortality (Figure 12).

Applying the stochastic (via the parametric bootstrap) time series of genetic stock components (see caveat above about extending stock composition estimates over an earlier period) to available runsize estimates allows computation of an impact or exploitation rate due to the pollock fishery bycatch. For the Upper Yukon River, this impact rate was well below $0.7 \%$ (Figure 13). For the "Coastal west Alaska" group, the impact rate estimates were considerably higher and have increased in recent years (Figure 14). Overall, from this analysis it appears that there is about a $10 \%$ chance that the coastal west Alaska group has experienced an exploitation rate greater than $3.5 \%$. However, the apparent increasing trend (consistent with increases in overall bycatch levels) warrants further monitoring.

For groups of Chinook stocks where run size information is incomplete it is possible to simply present the estimates of total adult equivalent mortality due to bycatch. For example, the estimates of Chinook mortalities that originated from stocks south of Alaska (Canada and the lower 48 states) range from around 3,000 fish during 2000, to as high as 13,000 fish in recent years (Figure 15).

### 1.3.4 Application to alternative cap scenarios

In Chapter 5 above, application to the subset of 36 bycatch alternatives for evaluation were presented. For each cap alternative and option, the hypothetical Chinook AEQ mortality totals under each cap and management option for 2003-2007 shows a fair amount of variability over different options and years (Table 14). For the western Alaska stocks, Myers' et al. (2003) scale pattern results were used to further break down these to river of origin (also presented in Chapter 5). Additionally, based on tables presented in Chapters 2 and 4, the savings in Chinook bycatch can be plotted relative to forgone pollock to show the trade-offs among alternatives (Figure 16).

### 1.4 Discussion

Myers' et al. (2003) recommended that NMFS estimate the variance of bycatch-at-age. Miller (2006) developed estimators on total salmon bycatch by the EBS trawl fleet and found that the CVs (coefficients of variation) of the estimates under the current sampling regime were on the order of $5 \%$ (assuming that hauls from unobserved vessels had the same bycatch pattern as that of observed vessels). This study provides an additional component of sampling variability attributed to length and age collections.

The samples from which Myers' et al. (2003) estimated ages were out of proportion relative to the bycatch. For example, in 1997 some $51 \%$ of the scale samples were from the A season whereas this represented only about $23 \%$ of the overall bycatch for that year (Table 15). Myers et al. corrected for the bycatch levels and achieved proportions at age similar to what was found in this study. However, during this period (1997-1999) the observers sampled over 41,500 Chinook salmon for lengths (compared to the estimated total Chinook bycatch over this period of 107,500 salmon). In this study, these length frequencies are combined with the age data to have a more complete sampling frame. An added benefit of including the length frequency samples is that scale sampling is impacted by the size of the fish. Fish that lose scales more easily are more often rejected for sample quality and scale loss tends to be higher for smaller fish. Having a complete length frequency set (where such sample rejection is unlikely to occur) should enhance the reliability of the age composition estimates. Having age structures read over more years would improve the estimates shown here and would help if further multi-stock models are constructed.

The time series of bycatch age composition estimates have only been briefly evaluated. Application extensions to these data can be explored with in-river brood year variability (e.g., Figure 17).

The stock composition estimates based on the genetics are qualitatively very similar to the scale-pattern study presented by Myers et al. (2003). The age composition, genetics, and modeling approach presented here should help to provide some foundation for evaluating the EIS that is being developed by NMFS and the Council and provide guidance for decisions on appropriate measures to reduce bycatch impacts. For example, it is possible to examine how a cap would have changed the impact rates historically. This can serve to illustrate the expected result of future cap regulation alternatives.

### 1.5 Literature Cited

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## FIGURES



2005

2006

Figure 1. Proportion of Chinook salmon samples collected for genetics compared to the proportion of bycatch by month for 2005 B-season only (top panel) and 2006 A and B season combined (bottom panel).

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Figure 2. Summary distribution of age samples by length collected by the NMFS groundfish observer program during 1997-1999 and analyzed by University of Washington scientists (Myers et al. (2003) for the A-season (top panel) and B season (bottom panel).


Figure 3. Length frequency by season and year of Chinook salmon occurring as bycatch in the pollock fishery. Error distributions based on two-stage bootstrap re-sampling procedure.


Figure 3. (continued) Length frequency by season and year of Chinook salmon occurring as bycatch in the pollock fishery. Error distributions based on two-stage bootstrap re-sampling procedure.


Figure 3. (continued) Length frequency by season and year of Chinook salmon occurring as bycatch in the pollock fishery. Error distributions based on two-stage bootstrap re-sampling procedure.


Figure 3. (continued) Length frequency by season and year of Chinook salmon occurring as bycatch in the pollock fishery. Error distributions based on two-stage bootstrap re-sampling procedure.

Chinook salmon bycatch A-Season



Figure 4. Chinook salmon bycatch age composition by year and A-season (top) and B-season (bottom). Vertical spread of blobs represent uncertainty as estimated from the two-stage bootstrap resampling procedure.


Figure 5. Bootstrap estimates of Chinook salmon bycatch example showing correlation of bycatch at different ages for the B-season in 1997 (top) and 1998 (bottom).


Figure 6. Chinook salmon bycatch results by reporting region for 2005 B season (top), 2006 B season (middle), and the 2006 and (partial sample) of 2007 A seasons (bottom). The top two panels include uncorrected results where bycatch differences between regions (east and west of $170^{\circ} \mathrm{W}$ ) are ignored (empty columns).


Figure 7. Simulated Chinook salmon stock proportion by region for the B season based on reported standard error values from ADFG analyses and assuming that the 2006 data has better coverage and is hence weighted $2: 1$ compared to the 2005 B-season data.


Figure 8. Chinook salmon bycatch results by genetics reporting regions for 2005 B season (top), 2006 B season (middle) and 2006 and (partial sample) of 207 season (bottom). The top two panels include uncorrected results where bycatch differences between regions (east and west of $170^{\circ} \mathrm{W}$ ) are ignored (empty columns).


Figure 9. Figure showing how the overall proportion of Upper Yukon River relates to the bycatch proportion that occurs in the NW region (west of $170^{\circ} \mathrm{W}$; top panel) and how the proportion of the BC-WA-OR (PNW) relates to the SE region (east of $170^{\circ} \mathrm{W}$; bottom panel) during the summer-fall pollock fishery, 1991-2007.


Figure 10. Time series of median Chinook adult equivalent bycatch from the pollock fishery, 1991-2007 compared to the annual totals. Dashed lines show the uncertainty due to the bootstrap age compositions of Chinook bycatch.


Figure 11. Time series of Chinook adult equivalent bycatch from the pollock fishery, 1991-2007 compared to the annual totals under different assumptions about ocean mortality rates.


Figure 12. Time series of Chinook adult equivalent bycatch from the pollock fishery, 1991-2007 compared to the annual totals with stochasticity in the bycatch age composition (via bootstrap samples), maturation rate ( $\mathrm{CV}=0.1$ ), natural mortality (Model 2, $\mathrm{CV}=0.1$ ).


Figure 13. Annual estimates of pollock fishery impacts on Upper Yukon returns, 1995-2006 (top panel) with stochasticity in natural mortality (Model $2, \mathrm{CV}=0.1$ ), maturation rate ( $\mathrm{CV}=0.1$ ), stock composition (as detailed above), and run size. The lower panel shows relative frequency of different impact levels given the simulations and bycatch history.


## Bycatch adult equivalents / Coast W AK Return

Figure 14. Annual estimates of pollock fishery impacts on Coastal west Alaska returns, 1994-2006 (top panel) with stochasticity in natural mortality (Model 2, CV=0.1), bycatch age composition (via bootstrap samples), maturation rate ( $\mathrm{CV}=0.1$ ), stock composition (as detailed above), and run size. The lower panel shows cumulative frequency of different impact levels given the simulations and bycatch history.


Figure 15. Annual estimated pollock fishery adult equivalent removals on stocks from the BC, WA, and Oregon returns, 1995-2007 with stochasticity in natural mortality (Model 2, CV=0.1), bycatch age composition (via bootstrap samples), maturation rate ( $\mathrm{CV}=0.1$ ), and stock composition (as detailed above).


Figure 16. Examples of trade-offs in hypothetical Chinook AEQ bycatch (horizontal axis) and forgone pollock (vertical axis) had the suite of 36 management options been in place for 2004 (upper left) through 2007 (lower right). The text plotted denote the sector split options and the symbols (and colors) represent A-B season splits: circle=50:50, square=58:42, diamond=70:30.


Figure 17. Chinook bycatch brood-year relative strength compared to the brood year variability observed in the Upper Yukon.

## TABLES

Table 1. Summary of Chinook salmon bycatch age data from Myers et al (2003) used to construct agelength keys for this analysis.

| Year | A | B | Total |
| ---: | ---: | ---: | ---: |
| 1997 | 842 | 756 | 1,598 |
| 1998 | 873 | 826 | 1,699 |
| 1999 | 645 | 566 | 1,211 |
| Total | 2,360 | 2,148 | 4,508 |

Table 2. The number of Chinook salmon measured for lengths in the pollock fishery by season (A and B), area ( $\mathrm{NW}=$ east of $170^{\circ} \mathrm{W}$; $\mathrm{SE}=$ west of $170^{\circ} \mathrm{W}$ ), and sector ( $\mathrm{S}=$ shorebased catcher vessels, $\mathrm{M}=$ mothership operations, $\mathrm{CP}=$ catcher-processors). Source: NMFS Alaska Fisheries Science Center observer data.

| Season | A | A | A | B | B | B | B | B | B |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Area | All | All | All | NW | NW | NW | SE | SE | SE |  |
| Sector | S | M | CP | S | M | CP | S | M | CP | Total |
| 1991 | 2,227 | 302 | 2,569 |  | 25 | 87 | 221 | 10 | 47 | 5,488 |
| 1992 | 2,305 | 733 | 889 | 2 | 4 | 14 | 1,314 | 21 | 673 | 5,955 |
| 1993 | 1,929 | 349 | 370 | 1 | 11 | 172 | 298 | 255 | 677 | 4,062 |
| 1994 | 4,756 | 408 | 986 | 3 | 93 | 276 | 781 | 203 | 275 | 7,781 |
| 1995 | 1,209 | 264 | 851 |  | 8 | 31 | 457 | 247 | 305 | 3,372 |
| 1996 | 9,447 | 976 | 2,798 |  | 17 | 161 | 5,658 | 1,721 | 493 | 21,271 |
| 1997 | 3,498 | 423 | 910 | 12 | 303 | 839 | 12,126 | 370 | 129 | 18,610 |
| 1998 | 3,124 | 451 | 1,329 |  | 38 | 191 | 8,277 | 2,446 | 1,277 | 17,133 |
| 1999 | 1,934 | 120 | 1,073 |  | 1 | 627 | 1,467 | 97 | 503 | 5,822 |
| 2000 | 608 | 17 | 1,388 | 4 | 40 | 179 | 564 | 3 | 120 | 2,923 |
| 2001 | 4,360 | 268 | 3,583 |  | 25 | 1,816 | 1,597 | 291 | 1,667 | 13,607 |
| 2002 | 5,587 | 850 | 3,011 |  | 23 | 114 | 5,353 | 520 | 494 | 15,952 |
| 2003 | 9,328 | 1,000 | 5,379 | 258 | 290 | 1,290 | 4,420 | 348 | 467 | 22,780 |
| 2004 | 7,247 | 594 | 3,514 | 1,352 | 557 | 1,153 | 8,884 | 137 | 606 | 24,044 |
| 2005 | 9,237 | 694 | 3,998 | 4,081 | 244 | 1,610 | 10,336 | 45 | 79 | 30,324 |
| 2006 | 17,875 | 1,574 | 5,716 | 685 | 66 | 480 | 12,757 | 3 | 82 | 39,238 |
| 2007 | 16,008 | 1,802 | 9,012 | 881 | 590 | 1,986 | 21,725 | 2 | 801 | 52,807 |

Table 3. Chinook salmon bycatch in the pollock fishery by season (A and B), area (NW=east of $170^{\circ} \mathrm{W}$; SE=west of $170^{\circ} \mathrm{W}$ ), and sector ( $\mathrm{S}=$ shorebased catcher vessels, $\mathrm{M}=$ mothership operations, CP=catcher-processors). Source: NMFS Regional Office, Juneau.

| Season | $\mathbf{A}$ <br> Area | $\mathbf{A}$ <br> All <br> All | $\mathbf{A}$ <br> $\mathbf{A l l}$ <br> $\mathbf{M}$ | $\mathbf{B}$ <br> $\mathbf{N W}$ | $\mathbf{B}$ <br> $\mathbf{N W}$ <br> $\mathbf{M}$ | $\mathbf{B}$ <br> $\mathbf{N W}$ <br> $\mathbf{C P}$ | $\mathbf{B}$ <br> $\mathbf{S E}$ <br> $\mathbf{S}$ | $\mathbf{B}$ <br> $\mathbf{S E}$ <br> $\mathbf{M}$ | $\mathbf{B}$ <br> $\mathbf{S E}$ <br> $\mathbf{C P}$ |  |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1991 | 10,192 | 9,001 | 17,645 | 0 | 48 | 318 | 1,667 | 103 | 79 | 39,054 |
| 1992 | 6,725 | 4,057 | 12,631 | 0 | 26 | 187 | 1,604 | 1,739 | 6,702 | 33,672 |
| 1993 | 3,017 | 3,529 | 8,869 | 29 | 157 | 7,158 | 2,585 | 6,500 | 4,775 | 36,619 |
| 1994 | 8,346 | 1,790 | 17,149 | 0 | 121 | 771 | 1,206 | 452 | 2,055 | 31,890 |
| 1995 | 2,040 | 971 | 5,971 |  | 35 | 77 | 781 | 632 | 2,896 | 13,403 |
| 1996 | 15,228 | 5,481 | 15,276 |  | 113 | 908 | 9,944 | 6,208 | 2,315 | 55,472 |
| 1997 | 4,954 | 1,561 | 3,832 | 43 | 2,143 | 4,172 | 22,508 | 3,559 | 1,549 | 44,320 |
| 1998 | 4,334 | 4,284 | 6,500 |  | 309 | 511 | 27,218 | 6,052 | 2,037 | 51,244 |
| 1999 | 3,103 | 554 | 2,694 | 13 | 12 | 1,284 | 2,649 | 362 | 1,306 | 11,978 |
| 2000 | 878 | 19 | 2,525 | 4 | 230 | 286 | 714 | 23 | 282 | 4,961 |
| 2001 | 8,555 | 1,664 | 8,264 | 0 | 162 | 5,346 | 3,779 | 1,157 | 4,517 | 3,444 |
| 2002 | 10,336 | 1,976 | 9,481 | 0 | 38 | 211 | 9,560 | 1,717 | 1,175 | 34,495 |
| 2003 | 16,488 | 2,892 | 14,428 | 764 | 864 | 2,962 | 6,437 | 1,076 | 1,081 | 46,993 |
| 2004 | 12,376 | 2,092 | 9,492 | 2,530 | 1,573 | 2,844 | 21,171 | 503 | 1,445 | 54,028 |
| 2005 | 14,097 | 2,111 | 11,421 | 8,873 | 744 | 4,175 | 26,113 | 144 | 168 | 67,847 |
| 2006 | 36,039 | 5,408 | 17,306 | 936 | 175 | 1,373 | 21,718 | 25 | 178 | 83,159 |
| 2007 | 35,458 | 5,860 | 27,943 | 1,672 | 3,494 | 4,923 | 40,079 | 50 | 2,225 | 121,704 |

Table 4. Table of Chinook baseline collections used in analysis of bycatch mixtures for genetics studies (from Templin et al. In Prep.).

| No. | Region | Location | Years | N |
| :---: | :---: | :---: | :---: | :---: |
| 1 | Russia | Bistraya River | 1998 | 94 |
| 2 |  | Bolshaya River | 1998, 2002 | 77 |
| 3 |  | Kamchatka River (Late) | 1997, 1998 | 119 |
| 4 |  | Pakhatcha River | 2002 | 50 |
| 5 | Norton Sound | Pilgrim River | 2005, 2006 | 82 |
| 6 |  | Unalakleet River | 2005 | 82 |
| 7 |  | Golsovia River | 2005, 2006 | 111 |
| 8 | Coast W AK (Lower Yukon) | Andreafsky River | 2002, 2003 | 236 |
| 9 |  | Anvik River | 2002 | 95 |
| 10 |  | Gisasa River | 2001 | 188 |
| 11 |  | Tozitna River | 2002, 2003 | 290 |
| 12 | Middle Yukon | Henshaw Creek | 2001 | 147 |
| 13 |  | S. Fork Koyuk | 2003 | 56 |
| 14 |  | Kantishna River | 2005 | 187 |
| 15 |  | Chena River | 2001 | 193 |
| 16 |  | Salcha River | 2005 | 188 |
| 17 |  | Beaver Creek | 1997 | 100 |
| 18 |  | Chandalar River | 2002, 2003, 2004 | 175 |
| 19 |  | Sheenjek River | 2002, 2004, 2006 | 51 |
| 20 | Upper Yukon | Chandindu River | 2000, 2001, 2003 | 247 |
| 21 |  | Klondike River | 1995, 2001, 2003 | 79 |
| 22 |  | Stewart River | 1997 | 99 |
| 23 |  | Mayo River | 1992, 1997, 2003 | 197 |
| 24 |  | Blind River | 2003 | 134 |
| 25 |  | Pelly River | 1996, 1997 | 140 |
| 26 |  | Little Salmon River | 1987, 1997 | 100 |
| 27 |  | Big Salmon River | 1987, 1997 | 117 |
| 28 |  | Tatchun Creek | 1987, 1996, 1997, 2002, 2003 | 369 |
| 29 |  | Nordenskiold River | 2003 | 55 |
| 30 |  | Nisutlin River | 19,871,997 | 56 |
| 31 |  | Takhini River | 1997, 2002, 2003 | 162 |
| 32 |  | Whitehorse Hatchery | 1985, 1987, 1997 | 242 |
| 33 | Coast W AK (Kuskokwim) | Goodnews River | 1993, 2005, 2006 | 368 |
| 34 |  | Arolik River | 2005 | 147 |
| 35 |  | Kanektok River | 1992, 1993, 2005 | 244 |
| 36 |  | Eek River | 2002, 2005 | 173 |
| 37 |  | Kwethluk River | 2001 | 96 |
| 38 |  | Kisaralik River | 2001, 2005 | 191 |
| 39 |  | Tuluksak River | 1993, 1994, 2005 | 195 |
| 40 |  | Aniak River | 2002, 2005, 2006 | 336 |
| 41 |  | George River | 2002, 2005 | 191 |
| 42 |  | Kogrukluk River | 1992, 1993, 2005 | 149 |
| 43 |  | Stony River | 1994 | 93 |
| 44 |  | Cheeneetnuk River | 2002, 2006 | 117 |
| 45 |  | Gagaryah River | 2006 | 190 |
| 46 |  | Takotna River | 1994, 2005 | 176 |
| 47 | Upper Kuskokwim | Tatlawiksuk River | 2002, 2005 | 191 |
| 48 |  | Salmon River (Pitka Fork) | 1995 | 96 |
| 49 | Coast W AK (Bristol Bay) | Togiak River | 1993, 1994 | 159 |
| 50 |  | Nushagak River | 1992, 1993 | 57 |
| 51 |  | Mulchatna River | 1994 | 97 |
| 52 |  | Stuyahok River | 1993, 1994 | 87 |
| 53 |  | Naknek River | 1995, 2004 | 110 |
| 54 |  | Big Creek | 2004 | 66 |
| 55 |  | King Salmon River | 2006 | 131 |
| 56 | N. AK Peninsula | Meshik River | 2006 | 42 |
| 57 |  | Milky River | 2006 | 67 |
| 58 |  | Nelson River | 2006 | 95 |
| 59 |  | Black Hills Creek | 2006 | 51 |
| 60 |  | Steelhead Creek | 2006 | 93 |
| 61 | S. AK Peninsula | Chignik River | 1995, 2006 | 75 |
| 62 |  | Ayakulik River | 1993, 2006 | 136 |
| 63 |  | Karluk River | 1993, 2006 | 140 |

Table 4. (continued) Table of Chinook baseline collections used in analysis of bycatch mixtures for genetics studies (from Templin et al. In Prep.).

| No. | Region | Location | Years | N |
| :---: | :---: | :---: | :---: | :---: |
| 64 | Cook Inlet | Deshka River | 1995, 2005 | 251 |
| 65 |  | Deception Creek | 1991 | 67 |
| 66 |  | Willow Creek | 2005 | 73 |
| 67 |  | Prairie Creek | 1995 | 52 |
| 68 |  | Talachulitna River | 1995 | 58 |
| 69 |  | Crescent Creek | 2006 | 164 |
| 70 |  | Juneau Creek | 2005, 2006 | 119 |
| 71 |  | Killey Creek | 2005, 2006 | 266 |
| 72 |  | Benjamin Creek | 2005, 2006 | 205 |
| 73 |  | Funny River | 2005, 2006 | 220 |
| 74 |  | Slikok Creek | 2005 | 95 |
| 75 |  | Kenai River (mainstem) | 2003, 2004, 2006 | 302 |
| 76 |  | Crooked Creek | 1992, 2005 | 306 |
| 77 |  | Kasilof River | 2005 | 321 |
| 78 |  | Anchor River | 2006 | 200 |
| 79 |  | Ninilchik River | 2006 | 162 |
| 80 | Upper Copper River | Indian River | 2004, 2005 | 50 |
| 81 |  | Bone Creek | 2004, 2005 | 78 |
| 82 |  | E. Fork Chistochina River | 2004 | 145 |
| 83 |  | Otter Creek | 2005 | 128 |
| 84 |  | Sinona Creek | 2004, 2005 | 157 |
| 85 | Lower Copper River | Gulkana River | 2004 | 211 |
| 86 |  | Mendeltna Creek | 2004 | 144 |
| 87 |  | Kiana Creek | 2004 | 75 |
| 88 |  | Manker Creek | 2004, 2005 | 62 |
| 89 |  | Tonsina River | 2004, 2005 | 75 |
| 90 |  | Tebay River | 2004, 2005, 2006 | 68 |
| 91 | Northern SE AK | Situk River | 1988, 1990, 1991, 1992 | 143 |
| 92 |  | Big Boulder Creek | 1992, 1993, 1995, 2004 | 178 |
| 93 |  | Tahini River | 1992, 2004 | 169 |
| 94 |  | Tahini River (LMH) Pullen Creek Hatchery | 2005 | 83 |
| 95 |  | Kelsall River | 2004 | 96 |
| 96 |  | King Salmon River | 1989, 1990, 1993 | 144 |
| 97 | Coast SE AK | King Creek | 2003 | 143 |
| 98 |  | Chickamin River | 1990, 2003 | 56 |
| 99 |  | Chickamin River - Little Port Walter | 1993, 2005 | 126 |
| 100 |  | Chickamin River - Whitman Lake Hatchery | 1992, 1998, 2005 | 331 |
| 101 |  | Humpy Creek | 2003 | 94 |
| 102 |  | Butler Creek | 2004 | 95 |
| 103 |  | Clear Creek | 1989, 2003, 2004 | 166 |
| 104 |  | Cripple Creek | 1988, 2003 | 143 |
| 105 |  | Genes Creek | 1989, 2003, 2004 | 95 |
| 106 |  | Kerr Creek | 2003, 2004 | 151 |
| 107 |  | Unuk River - Little Port Walter | 2005 | 150 |
| 108 |  | Unuk River - Deer Mountain Hatchery | 1992, 1994 | 147 |
| 109 |  | Keta River | 1989, 2003 | 144 |
| 110 |  | Blossom River | 2004 | 95 |
| 111 | Andrew Cr | Andrews Creek | 1989, 2004 | 152 |
| 112 |  | Crystal Lake Hatchery | 1992, 1994, 2005 | 397 |
| 113 |  | Medvejie Hatchery | 1998, 2005 | 273 |
| 114 |  | Hidden Falls Hatchery | 1994, 1998 | 155 |
| 115 |  | Macaulay Hatchery | 2005 | 94 |
| 116 | TBR Taku | Klukshu River | 1989, 1990 | 174 |
| 117 |  | Kowatua River | 1989, 1990 | 144 |
| 118 |  | Little Tatsemeanie River | 1989, 1990, 2005 | 144 |
| 119 |  | Upper Nahlin River | 1989, 1990 | 130 |
| 120 |  | Nakina River | 1989, 1990 | 141 |
| 121 |  | Dudidontu River | 2005 | 86 |
| 122 |  | Tahltan River | 1989 | 95 |

Table 4. (continued) Table of Chinook baseline collections used in analysis of bycatch mixtures for genetics studies (from Templin et al. In Prep.).

| No. | Region | Location | Years | N |
| :---: | :---: | :---: | :---: | :---: |
| 123 | BC/WA/OR | Kateen River | 2005 | 96 |
| 124 |  | Damdochax Creek | 1996 | 65 |
| 125 |  | Kincolith Creek | 1996 | 115 |
| 126 |  | Kwinageese Creek | 1996 | 73 |
| 127 |  | Oweegee Creek | 1996 | 81 |
| 128 |  | Babine Creek | 1996 | 167 |
| 129 |  | Bulkley River | 1999 | 91 |
| 130 |  | Sustut | 2001 | 130 |
| 131 |  | Ecstall River | 2001, 2002 | 86 |
| 132 |  | Lower Kalum | 2001 | 142 |
| 133 |  | Lower Atnarko | 1996 | 144 |
| 134 |  | Kitimat | 1997 | 141 |
| 135 |  | Wannock | 1996 | 144 |
| 136 |  | Klinaklini | 1997 | 83 |
| 137 |  | Nanaimo | 2002 | 95 |
| 138 |  | Porteau Cove | 2003 | 154 |
| 139 |  | Conuma River | 1997, 1998 | 110 |
| 140 |  | Marble Creek | 1996, 1999, 2000 | 144 |
| 141 |  | Nitinat River | 1996 | 104 |
| 142 |  | Robertson Creek | 1996, 2003 | 106 |
| 143 |  | Sarita | 1997, 2001 | 160 |
| 144 |  | Big Qualicum River | 1996 | 144 |
| 145 |  | Quinsam River | 1996 | 127 |
| 146 |  | Morkill River | 2001 | 154 |
| 147 |  | Salmon River | 1997 | 94 |
| 148 |  | Swift | 1996 | 163 |
| 149 |  | Torpy River | 2001 | 105 |
| 150 |  | Chilko | 1995, 1996, 1999, 2002 | 246 |
| 151 |  | Nechako River | 1996 | 121 |
| 152 |  | Quesnel River | 1996 | 144 |
| 153 |  | Stuart | 1997 | 161 |
| 154 |  | Clearwater River | 1997 | 153 |
| 155 |  | Louis Creek | 2001 | 179 |
| 156 |  | Lower Adams | 1996 | 46 |
| 157 |  | Lower Thompson River | 2001 | 100 |
| 158 |  | Middle Shuswap | 1986, 1997 | 144 |
| 159 |  | Birkenhead Creek | 1997, 1999, 2002, 2003 | 93 |
| 160 |  | Harrison | 2002 | 96 |
| 161 |  | Makah National Fish Hatchery | 2001, 2003 | 94 |
| 162 |  | Forks | 2005 | 150 |
| 163 |  | Upper Skagit River | 2006 | 93 |
| 164 |  | Soos Creek Hatchery | 2004 | 119 |
| 165 |  | Lyons Ferry Hatchery | 2002, 2003 | 191 |
| 166 |  | Hanford Reach | 2000, 2004, 2006 | 191 |
| 167 |  | Lower Deschutes River | 2002 | 96 |
| 168 |  | Lower Kalama | 2001 | 95 |
| 169 |  | Carson Stock - Mid and Upper Columbia spring | 2001 | 96 |
| 170 |  | McKenzie - Willamette River | 2004 | 95 |
| 171 |  | Alsea | 2004 | 93 |
| 172 |  | Siuslaw | 2001 | 95 |
| 173 |  | Klamath | 1990, 2006 | 52 |
| 174 |  | Butte Creek | 2003 | 96 |
| 175 |  | Eel River | 2000, 2001 | 88 |
| 176 |  | Sacramento River - winter run | 2005 | 95 |

Table 5. NMFS regional office estimates of Chinook salmon bycatch in the pollock fishery compared to genetics sampling levels by season and region, 2005-2007 (SE=east of $170^{\circ} \mathrm{W}$, NW=west of $170^{\circ} \mathrm{W}$ ).

|  |  | Area |  |  |  | Ara |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Season | SE | NW | Total | SE | NW |  |
|  | 2005 | B | 26,425 | 13,793 | 40,217 | $66 \%$ | $34 \%$ |  |
| Bycatch | 2006 | B | 21,922 | 2,484 | 24,405 | $90 \%$ | $10 \%$ |  |
|  | 2006 | A |  |  | 58,753 |  |  |  |
|  | 2007 | A |  |  | 69,261 |  |  |  |
|  | 2005 | B | 489 | 282 | 771 | $63 \%$ | $37 \%$ |  |
| Genetic | 2006 | B | 286 | 304 | 590 | $48 \%$ | $52 \%$ |  |
| Samples | 2006 | A |  |  | 801 |  |  |  |
|  | 2007 | A |  |  | 360 |  |  |  |

Table 6. ADFG estimates of stock composition based on genetic samples stratified by year, season, and region (SE = east of $170^{\circ} \mathrm{W}, \mathrm{NW}=$ west of $170^{\circ} \mathrm{W}$ ). Standard errors of the estimates are shown in parentheses and were used to evaluate uncertainty of stock composition. Source: $A D F G$ preliminary data.

|  |  | Coast | Cook | Middle | N AK |  | Upper |  |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Year / Season / Area | PNW | W AK | Inlet | Yukon | Penin | Russia | TBR | Yukon |
| 2005 B SE | $45.3 \%$ | $34.2 \%$ | $5.3 \%$ | $0.2 \%$ | $8.8 \%$ | $0.6 \%$ | $3.3 \%$ | $0.0 \%$ |
| N = 282 | $(0.032)$ | $(0.032)$ | $(0.019)$ | $(0.003)$ | $(0.021)$ | $(0.005)$ | $(0.016)$ | $(0.001)$ |
| 2005 B NW | $6.5 \%$ | $70.9 \%$ | $2.2 \%$ | $4.7 \%$ | $6.7 \%$ | $2.0 \%$ | $3.5 \%$ | $2.8 \%$ |
| N = 489 | $(0.012)$ | $(0.047)$ | $(0.011)$ | $(0.013)$ | $(0.042)$ | $(0.007)$ | $(0.012)$ | $(0.009)$ |
| 2006 B SE | $38.4 \%$ | $37.2 \%$ | $7.5 \%$ | $0.2 \%$ | $7.0 \%$ | $0.6 \%$ | $4.3 \%$ | $0.1 \%$ |
| $\mathrm{~N}=304$ | $(0.029)$ | $(0.032)$ | $(0.020)$ | $(0.004)$ | $(0.019)$ | $(0.005)$ | $(0.017)$ | $(0.002)$ |
| 2006 B NW | $6.4 \%$ | $67.3 \%$ | $3.0 \%$ | $8.0 \%$ | $2.1 \%$ | $3.3 \%$ | $0.5 \%$ | $8.0 \%$ |
| $\mathrm{~N}=286$ | $(0.016)$ | $(0.035)$ | $(0.020)$ | $(0.020)$ | $(0.016)$ | $(0.013)$ | $(0.007)$ | $(0.019)$ |
| 2006 A All | $22.9 \%$ | $38.2 \%$ | $0.2 \%$ | $1.1 \%$ | $31.2 \%$ | $1.1 \%$ | $1.1 \%$ | $2.3 \%$ |
| $\mathrm{~N}=801$ | $(0.015)$ | $(0.038)$ | $(0.004)$ | $(0.005)$ | $(0.039)$ | $(0.004)$ | $(0.007)$ | $(0.006)$ |
| 2007 A All | $9.4 \%$ | $75.2 \%$ | $0.1 \%$ | $0.5 \%$ | $12.0 \%$ | $0.2 \%$ | $0.1 \%$ | $0.1 \%$ |
| $\mathrm{~N}=360$ | $(0.016)$ | $(0.031)$ | $(0.004)$ | $(0.005)$ | $(0.025)$ | $(0.003)$ | $(0.002)$ | $(0.003)$ |

Table 7. Range of estimated mean age-specific maturation by brood year used to compute adult equivalents. The weighted mean value is based on the relative Chinook run sizes between the Nushagak and Yukon Rivers since 1997. Sources: Healey 1991, Dani Evenson (ADFG, pers. Comm.), Rishi Sharma (CRITFC, pers. Comm.).

|  | Weight | Age 3 | Age 4 | Age 5 | Age 6 | Age 7 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Yukon | 2.216 | $1 \%$ | $13 \%$ | $32 \%$ | $49 \%$ | $5 \%$ |
| Nushagak since 82 | 1.781 | $1 \%$ | $21 \%$ | $38 \%$ | $39 \%$ | $2 \%$ |
| Nushagak since 66 | 0 | $0 \%$ | $17 \%$ | $36 \%$ | $43 \%$ | $3 \%$ |
| Goodnews | 0 | $0 \%$ | $20 \%$ | $31 \%$ | $45 \%$ | $4 \%$ |
| SE Alaska (TBR) | 0.3 | $0 \%$ | $18 \%$ | $40 \%$ | $37 \%$ | $5 \%$ |
| BC, WA, OR, \& CA | 0.7 | $3 \%$ | $28 \%$ | $53 \%$ | $14 \%$ | $1 \%$ |
| Weighted mean |  | $1 \%$ | $18 \%$ | $37 \%$ | $40 \%$ | $3 \%$ |

Table 8. Chinook salmon effective bycatch "caps" in the pollock fishery by season (A and B) based on average values of the caps (if they occurred) had they been applied from 2003-2007.

| Cap, A/B, sector | A season | B season | Total |
| :--- | ---: | ---: | ---: |
| $87,50050 / 50$ opt1 | 31,099 | 24,339 | 55,438 |
| $87,50050 / 50$ opt2a | 31,950 | 32,844 | 64,793 |
| $87,50050 / 50$ opt2d | 36,899 | 28,791 | 65,690 |
| $87,50058 / 42$ opt1 | 44,118 | 20,321 | 64,439 |
| $87,50058 / 42$ opt2a | 41,653 | 30,463 | 72,116 |
| $87,50058 / 42$ opt2d | 42,234 | 24,258 | 66,492 |
| $87,50070 / 30$ opt1 | 49,368 | 16,277 | 65,644 |
| $87,50070 / 30$ opt2a | 44,665 | 18,427 | 63,092 |
| $87,50070 / 30$ opt2d | 55,376 | 17,815 | 73,191 |
| $68,10050 / 50$ opt1 | 27,784 | 18,272 | 46,056 |
| $68,10050 / 50$ opt2a | 26,459 | 28,264 | 54,723 |
| $68,10050 / 50$ opt2d | 25,196 | 24,258 | 49,455 |
| $68,10058 / 42$ opt1 | 29,569 | 17,581 | 47,150 |
| $68,10058 / 42$ opt2a | 28,587 | 21,247 | 49,834 |
| $68,10058 / 42$ opt2d | 32,676 | 19,997 | 52,674 |
| $68,10070 / 30$ opt1 | 41,021 | 13,253 | 54,274 |
| $68,10070 / 30$ opt2a | 35,980 | 15,495 | 51,475 |
| $68,10070 / 30$ opt2d | 42,234 | 14,640 | 56,874 |
| $48,70050 / 50$ opt1 | 19,292 | 16,196 | 35,488 |
| $48,70050 / 50$ opt2a | 18,053 | 17,439 | 35,493 |
| $48,70050 / 50$ opt2d | 21,242 | 16,725 | 37,966 |
| $48,70058 / 42$ opt1 | 21,142 | 13,253 | 34,394 |
| $48,70058 / 42$ opt2a | 19,592 | 15,495 | 35,087 |
| $48,70058 / 42$ opt2d | 23,610 | 14,640 | 38,250 |
| $48,70070 / 30$ opt1 | 27,784 | 10,225 | 38,009 |
| $48,70070 / 30$ opt2a | 26,459 | 12,262 | 38,721 |
| $48,70070 / 30$ opt2d | 25,196 | 11,612 | 36,809 |
| $29,30050 / 50$ opt1 | 9,761 | 10,225 | 19,985 |
| $29,30050 / 50$ opt2a | 10,637 | 12,262 | 22,900 |
| $29,30050 / 50$ opt2d | 10,070 | 11,612 | 21,682 |
| $29,30058 / 42$ opt1 | 12,725 | 8,740 | 21,465 |
| $29,30058 / 42$ opt2a | 12,177 | 10,520 | 22,697 |
| $29,30058 / 42$ opt2d | 12,031 | 10,634 | 22,665 |
| $29,30070 / 30$ opt1 | 15,120 | 6,885 | 22,005 |
| $29,30070 / 30$ opt2a | 17,010 | 7,065 | 24,074 |
| $29,30070 / 30$ opt2d | 14,859 | 6,775 | 21,634 |

Table 9. Calendar year age-specific Chinook salmon bycatch estimates based on the mean of 100 bootstrap samples of available length and age data. Age-length keys for 1997-1999 were based on Myers et al. (2003) data split by year while for all other years, a combined-year agelength key was used.

| Year | Age 3 | Age 4 | Age 5 | Age 6 | Age 7 | Total |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| $\mathbf{1 9 9 1}$ | 5,624 | 15,901 | 13,486 | 3,445 | 347 | 38,802 |
| $\mathbf{1 9 9 2}$ | 5,136 | 9,528 | 14,538 | 3,972 | 421 | 33,596 |
| $\mathbf{1 9 9 3}$ | 2,815 | 16,565 | 12,992 | 3,673 | 401 | 36,446 |
| $\mathbf{1 9 9 4}$ | 849 | 5,300 | 20,533 | 4,744 | 392 | 31,817 |
| $\mathbf{1 9 9 5}$ | 498 | 3,895 | 4,827 | 3,796 | 367 | 13,382 |
| $\mathbf{1 9 9 6}$ | 5,091 | 18,590 | 26,202 | 5,062 | 421 | 55,366 |
| $\mathbf{1 9 9 7}$ | 5,855 | 23,972 | 7,233 | 5,710 | 397 | 43,167 |
| $\mathbf{1 9 9 8}$ | 19,168 | 16,169 | 11,751 | 2,514 | 615 | 50,216 |
| $\mathbf{1 9 9 9}$ | 870 | 5,343 | 4,424 | 1,098 | 21 | 11,757 |
| $\mathbf{2 0 0 0}$ | 662 | 1,923 | 1,800 | 518 | 44 | 4,939 |
| $\mathbf{2 0 0 1}$ | 6,512 | 12,365 | 11,948 | 1,994 | 190 | 33,009 |
| $\mathbf{2 0 0 2}$ | 3,843 | 13,893 | 10,655 | 5,469 | 489 | 34,349 |
| $\mathbf{2 0 0 3}$ | 5,703 | 16,723 | 20,124 | 3,791 | 298 | 46,639 |
| $\mathbf{2 0 0 4}$ | 6,935 | 23,740 | 18,371 | 4,406 | 405 | 53,858 |
| $\mathbf{2 0 0 5}$ | 10,466 | 30,717 | 21,886 | 4,339 | 304 | 67,711 |
| $\mathbf{2 0 0 6}$ | 11,835 | 31,455 | 32,452 | 6,636 | 490 | 82,869 |
| $\mathbf{2 0 0 7}$ | 16,174 | 66,024 | 33,286 | 5,579 | 357 | 121,419 |

Table 10. Age specific Chinook salmon bycatch estimates by season and calendar age based on the mean of 100 bootstrap samples of available length and age data.

| Year/season | Age 3 | Age 4 | Age 5 | Age 6 | Age 7 | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1991 | 5,624 | 15,901 | 13,486 | 3,445 | 347 | 38,802 |
| A | 5,406 | 14,764 | 12,841 | 3,270 | 313 | 36,593 |
| B | 218 | 1,137 | 646 | 174 | 34 | 2,209 |
| 1992 | 5,136 | 9,528 | 14,538 | 3,972 | 421 | 33,596 |
| A | 1,017 | 4,633 | 13,498 | 3,798 | 408 | 23,355 |
| B | 4,119 | 4,895 | 1,040 | 174 | 13 | 10,241 |
| 1993 | 2,815 | 16,565 | 12,992 | 3,673 | 401 | 36,446 |
| A | 1,248 | 3,654 | 7,397 | 2,778 | 290 | 15,368 |
| B | 1,567 | 12,910 | 5,595 | 895 | 111 | 21,078 |
| 1994 | 849 | 5,300 | 20,533 | 4,744 | 392 | 31,817 |
| A | 436 | 3,519 | 18,726 | 4,211 | 326 | 27,218 |
| B | 413 | 1,781 | 1,807 | 533 | 66 | 4,599 |
| 1995 | 498 | 3,895 | 4,827 | 3,796 | 367 | 13,382 |
| A | 262 | 1,009 | 3,838 | 3,534 | 327 | 8,969 |
| B | 236 | 2,885 | 989 | 263 | 40 | 4,413 |
| 1996 | 5,091 | 18,590 | 26,202 | 5,062 | 421 | 55,366 |
| A | 863 | 7,187 | 23,118 | 4,431 | 349 | 35,947 |
| B | 4,228 | 11,403 | 3,085 | 632 | 71 | 19,418 |
| 1997 | 5,855 | 23,972 | 7,233 | 5,710 | 397 | 43,167 |
| A | 456 | 2,013 | 3,595 | 3,899 | 271 | 10,234 |
| B | 5,399 | 21,958 | 3,638 | 1,811 | 126 | 32,933 |
| 1998 | 19,168 | 16,169 | 11,751 | 2,514 | 615 | 50,216 |
| A | 1,466 | 2,254 | 8,639 | 2,079 | 512 | 14,950 |
| B | 17,703 | 13,915 | 3,112 | 435 | 103 | 35,266 |
| 1999 | 870 | 5,343 | 4,424 | 1,098 | 21 | 11,757 |
| A | 511 | 1,639 | 3,151 | 898 | 18 | 6,217 |
| B | 360 | 3,704 | 1,272 | 200 | 3 | 5,540 |
| 2000 | 662 | 1,923 | 1,800 | 518 | 34 | 4,939 |
| A | 365 | 1,167 | 1,406 | 453 | 26 | 3,416 |
| B | 298 | 757 | 395 | 66 | 8 | 1,522 |
| 2001 | 6,512 | 12,365 | 11,948 | 1,994 | 190 | 33,009 |
| A | 2,840 | 3,458 | 9,831 | 1,798 | 171 | 18,098 |
| B | 3,672 | 8,907 | 2,117 | 196 | 19 | 14,910 |
| 2002 | 3,843 | 13,893 | 10,655 | 5,469 | 489 | 34,349 |
| A | 1,580 | 5,063 | 9,234 | 5,328 | 478 | 21,683 |
| B | 2,263 | 8,830 | 1,421 | 141 | 11 | 12,666 |
| 2003 | 5,703 | 16,723 | 20,124 | 3,791 | 298 | 46,639 |
| A | 2,941 | 9,408 | 17,411 | 3,437 | 267 | 33,464 |
| B | 2,763 | 7,315 | 2,713 | 354 | 31 | 13,175 |
| 2004 | 6,935 | 23,740 | 18,371 | 4,406 | 405 | 53,858 |
| A | 1,111 | 5,520 | 13,090 | 3,763 | 354 | 23,838 |
| B | 5,824 | 18,220 | 5,282 | 643 | 51 | 30,020 |
| 2005 | 10,466 | 30,717 | 21,886 | 4,339 | 304 | 67,711 |
| A | 1,407 | 6,993 | 15,563 | 3,361 | 226 | 27,550 |
| B | 9,059 | 23,724 | 6,323 | 978 | 78 | 40,161 |
| 2006 | 11,835 | 31,455 | 32,452 | 6,636 | 490 | 82,869 |
| A | 3,604 | 17,574 | 30,447 | 6,404 | 465 | 58,494 |
| B | 8,231 | 13,881 | 2,005 | 232 | 25 | 24,374 |
| 2007 | 16,174 | 66,024 | 33,286 | 5,579 | 357 | 121,419 |
| A | 5,791 | 29,269 | 28,648 | 5,059 | 317 | 69,084 |
| B | 10,384 | 36,755 | 4,638 | 520 | 40 | 52,336 |

Table 11. Estimates of coefficients of variation of Chinook salmon bycatch estimates by season and calendar age based on the mean of 100 bootstrap samples of available length and age data.

| A season | Age 3 | Age 4 | Age 5 | Age 6 | Age 7 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1991 | 14\% | 6\% | 6\% | 10\% | 31\% |
| 1992 | 20\% | 9\% | 4\% | 9\% | 27\% |
| 1993 | 22\% | 9\% | 5\% | 10\% | 37\% |
| 1994 | 27\% | 12\% | 3\% | 10\% | 30\% |
| 1995 | 25\% | 12\% | 5\% | 6\% | 22\% |
| 1996 | 19\% | 6\% | 2\% | 9\% | 21\% |
| 1997 | 35\% | 12\% | 6\% | 7\% | 28\% |
| 1998 | 16\% | 9\% | 3\% | 10\% | 23\% |
| 1999 | 19\% | 10\% | 5\% | 11\% | 91\% |
| 2000 | 25\% | 9\% | 6\% | 9\% | 27\% |
| 2001 | 10\% | 6\% | 3\% | 7\% | 22\% |
| 2002 | 15\% | 6\% | 3\% | 4\% | 16\% |
| 2003 | 14\% | 6\% | 3\% | 8\% | 21\% |
| 2004 | 15\% | 6\% | 2\% | 5\% | 20\% |
| 2005 | 18\% | 6\% | 3\% | 7\% | 23\% |
| 2006 | 17\% | 5\% | 3\% | 7\% | 22\% |
| 2007 | 22\% | 5\% | 4\% | 8\% | 25\% |
| B season | Age 3 | Age 4 | Age 5 | Age 6 | Age 7 |
| 1991 | 23\% | 8\% | 12\% | 27\% | 67\% |
| 1992 | 9\% | 9\% | 25\% | 69\% | 87\% |
| 1993 | 19\% | 4\% | 9\% | 20\% | 65\% |
| 1994 | 17\% | 6\% | 6\% | 14\% | 27\% |
| 1995 | 21\% | 5\% | 12\% | 23\% | 48\% |
| 1996 | 6\% | 3\% | 7\% | 11\% | 29\% |
| 1997 | 12\% | 3\% | 10\% | 12\% | 39\% |
| 1998 | 5\% | 6\% | 9\% | 23\% | 36\% |
| 1999 | 16\% | 3\% | 8\% | 22\% | 149\% |
| 2000 | 9\% | 5\% | 8\% | 25\% | 49\% |
| 2001 | 7\% | 3\% | 8\% | 20\% | 52\% |
| 2002 | 6\% | $2 \%$ | 8\% | 17\% | 43\% |
| 2003 | 8\% | 3\% | 5\% | 15\% | 32\% |
| 2004 | 6\% | $2 \%$ | 5\% | 12\% | 30\% |
| 2005 | 5\% | $2 \%$ | 5\% | 10\% | 23\% |
| 2006 | 4\% | 3\% | 8\% | 15\% | 33\% |
| 2007 | 6\% | 2\% | 7\% | 13\% | 28\% |

Table 12. Mean values of catch-weighted stratified proportions of stock composition based on genetic sampling by season, and region (SE=east of $170^{\circ} \mathrm{W}$, NW=west of $170^{\circ} \mathrm{W}$ ). Standard errors of the estimates (in parentheses) were derived from 200 simulations based on the estimates from Table 6 and weighting annual results as explained in the text.

|  |  | Coast | Cook | Middle | N AK |  |  | Upper |  |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Season / Area | PNW | W AK | Inlet | Yukon | Penin | Russia | TBR | Yukon | Other |
| B SE | $45.0 \%$ | $34.7 \%$ | $5.1 \%$ | $0.1 \%$ | $8.6 \%$ | $0.6 \%$ | $3.4 \%$ | $0.0 \%$ | $2.4 \%$ |
|  | $(0.025)$ | $(0.024)$ | $(0.017)$ | $(0.002)$ | $(0.016)$ | $(0.004)$ | $(0.014)$ | $(0.001)$ | $(0.014)$ |
| B NW | $6.4 \%$ | $68.9 \%$ | $2.6 \%$ | $6.6 \%$ | $4.4 \%$ | $2.7 \%$ | $1.8 \%$ | $5.6 \%$ | $1.0 \%$ |
|  | $(0.010)$ | $(0.023)$ | $(0.012)$ | $(0.011)$ | $(0.019)$ | $(0.007)$ | $(0.006)$ | $(0.012)$ | $(0.008)$ |
| A All | $12.1 \%$ | $67.7 \%$ | $0.1 \%$ | $0.6 \%$ | $16.0 \%$ | $0.4 \%$ | $0.2 \%$ | $0.6 \%$ | $2.3 \%$ |
|  | $(0.012)$ | $(0.021)$ | $(0.003)$ | $(0.004)$ | $(0.019)$ | $(0.002)$ | $(0.002)$ | $(0.003)$ | $(0.010)$ |

Table 13. Median values of stochastic simulation results of AEQ Chinook mortality attributed to the pollock fishery by region, 1994-2007. These simulations include stochasticity in natural mortality (Model 2, $\mathrm{CV}=0.1$ ), bycatch age composition (via bootstrap samples), maturation rate ( $\mathrm{CV}=0.1$ ), and stock composition (as detailed above). NOTE: these results are based on the assumption that the genetics findings from the 2005-2007 data represent the historical pattern of bycatch stock composition (by strata).

|  | BC, WA, <br> OR, <br> and CA | Coastal <br> W. AK | Cook <br> Inlet | Middle <br> Yukon | N. Alaska <br> Peninsula | Other | Russia | Upper <br> Yukon | TBR <br> (SE | Total |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1994 | 5,198 | 21,518 | 242 | 201 | 4,898 | 714 | 147 | 194 | 198 | 33,310 |
| 1995 | 5,635 | 14,084 | 415 | 104 | 3,302 | 532 | 112 | 96 | 279 | 24,559 |
| 1996 | 6,974 | 17,025 | 520 | 154 | 3,939 | 632 | 142 | 137 | 364 | 29,886 |
| 1997 | 11,376 | 16,895 | 1,276 | 413 | 3,364 | 715 | 277 | 343 | 783 | 35,442 |
| 1998 | 10,967 | 14,218 | 1,110 | 103 | 3,382 | 696 | 165 | 87 | 711 | 31,439 |
| 1999 | 6,429 | 15,099 | 573 | 297 | 3,193 | 561 | 188 | 245 | 387 | 26,973 |
| 2000 | 2,815 | 9,383 | 219 | 167 | 2,106 | 330 | 99 | 147 | 152 | 15,418 |
| 2001 | 3,694 | 10,473 | 349 | 260 | 2,141 | 375 | 149 | 221 | 238 | 17,899 |
| 2002 | 6,236 | 14,516 | 509 | 106 | 3,467 | 609 | 117 | 96 | 341 | 25,997 |
| 2003 | 5,743 | 20,065 | 398 | 356 | 4,424 | 679 | 207 | 311 | 292 | 32,475 |
| 2004 | 10,164 | 21,904 | 1,018 | 466 | 4,592 | 859 | 305 | 393 | 685 | 40,386 |
| 2005 | 11,169 | 25,462 | 1,203 | 767 | 5,107 | 923 | 439 | 645 | 772 | 46,487 |
| 2006 | 12,719 | 36,337 | 892 | 363 | 8,355 | 1,348 | 290 | 339 | 633 | 61,275 |
| 2007 | 18,079 | 44,380 | 1,597 | 694 | 9,743 | 1,688 | 485 | 608 | 1,069 | 78,344 |

Table 14. Hypothetical adult equivalent Chinook salmon bycatch mortality totals under each cap and management option, 2003-2007. Numbers are based on the median AEQ values with the original estimates shown in the second row. Right-most column shows the mean over all years relative to the estimated AEQ bycatch. The shadings and the pies relate to the relative AEQ bycatch for each policy and year.

|  | 2003 | 2004 | 2005 | 2006 | 2007 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| No Cap | 33,215 | 41,047 | 47,268 | 61,737 | 78,814 |  |
| Cap, AB, sector |  |  |  |  |  |  |
| 87,500 70/30 opt2d | (1) 32,903 | 38,255 | 38,479 | 49,058 | 56,397 | 82\% |
| 87,500 70/30 opt2a | ( 33,081 | 38,485 | 38,753 | 49,986 | 54,164 | 82\% |
| 87,500 70/30 opt 1 | (-) 32,864 ( | 37,582 | 36,635 | 43,381 | 51,106 | 77\% |
| 87,500 58/42 opt2d | (1) 33,368 | 39,856 | 42,197 | 47,135 | 51,981 | 82\% |
| 87,500 58/42 opt2a | (-32,143 | 39,887 | 44,402 | 54,960 | 59,119 | 88\% |
| 87,500 58/42 opt1 | (-33,108 | 38,163 | 38,153 | 44,338 | 51,012 | 78\% |
| 87,500 50/50 opt2d | (1) 33,010 | 40,943 | 42,928 | 49,228 | 51,971 | 83\% |
| 87,500 50/50 opt2a | (-)30,747 | 38,967 | 43,140 | 47,977 | 53,212 | 82\% |
| 87,500 50/50 opt1 | (-) 33,151 | 39,747 | 41,912 | 43,139 | 43,599 | 77\% |
| 68,100 70/30 opt2d | () 33,162 ( | 36,866 | 36,314 | 40,583 | 45,112 | 73\% |
| 68,100 70/30 opt2a | (-)29,981 | 34,695 | 36,854 | 44,290 | 47,643 | 74\% |
| 68,100 70/30 opt 1 | (1) 32,948 ( | 36,791 | 35,507 | 39,891 | 42,666 | 72\% |
| 68,100 58/42 opt2d | () 32,364 ( | 37,417 | 37,704 | 40,948 | 43,194 | 73\% |
| 68,100 58/42 opt2a | (-)30,023 | 36,658 | 39,105 | 43,534 | 45,139 | 74\% |
| 68,100 58/42 opt1 | (-33,108 | 37,477 | 37,402 | 35,895 | 38,137 | 69\% |
| 68,100 50/50 opt2d | (-)30,769 | 37,607 | 41,249 | 38,952 | 38,063 | 71\% |
| 68,100 50/50 opt2a | (-)30,084 | 37,224 | 39,182 | 43,200 | 45,144 | 74\% |
| 68,100 50/50 opt 1 | (-) 32,342 | 37,659 | 38,203 | 36,334 | 35,679 | 69\% |
| 48,700 70/30 opt2d | (-) 29,249 | 33,665 | 33,408 | 30,077 | 28,277 | 59\% |
| 48,700 70/30 opt2a | (-)28,798 | 31,431 ( | 31,021 | 33,765 | 34,297 | 61\% |
| 48,700 70/30 opt 1 | (-)30,155 | 33,547 | 33,374 | 31,735 | 29,376 | 60\% |
| 48,700 58/42 opt2d | (-) 29,987 | 33,692 | 34,121 | 30,697 | 30,120 | 61\% |
| 48,700 58/42 opt2a | (-)27,722 | 31,175 | 32,007 | 28,025 | 27,065 | 56\% |
| 48,700 58/42 opt 1 | (-)28,349 | 33,201 | 33,788 | 30,543 | 25,454 | 58\% |
| 48,700 50/50 opt2d | (-) 28,797 | 33,773 | 33,600 | 30,876 | 29,647 | 60\% |
| 48,700 50/50 opt2a | (-)26,949 | 30,859 | 31,139 | 28,650 | 27,215 | 55\% |
| 48,700 50/50 opt 1 | (-)26,854 | 31,947 ( | 31,278 | 29,530 | 26,716 | 56\% |
| 29,300 70/30 opt2d | -19,200 | 22,679 | 23,095 $\bigcirc$ | 20,513 | 13,338 | 38\% |
| 29,300 70/30 opt2a | - 21,115 | 23,813 | 23,825 $\bigcirc$ | 20,612 $\bigcirc$ | 17,220 | 41\% |
| $\underline{29,30070 / 30 ~ o p t 1}$ | -19,252 | 22,524 $\bigcirc$ | 21,886 $\bigcirc$ | 19,101 $\bigcirc$ | 15,220 | $37 \%$ |
| 29,300 58/42 opt2d | -18,963 | 23,646 | 22,393 $\bigcirc$ | 20,476 | 15,041 | 38\% |
| 29,300 58/42 opt2a | -19,376 | 23,043 | 22,132 $\bigcirc$ | 20,827 $\bigcirc$ | 15,039 | 38\% |
| 29,300 58/42 opt 1 | $\bigcirc 18,259 \bigcirc$ | 21,267 $\bigcirc$ | 21,286 $\bigcirc$ | $18,331 \bigcirc$ | 14,924 | 36\% |
| 29,300 50/50 opt2d | 19,122 | 22,130 | 21,382 $\bigcirc$ | 18,665 | 14,048 | 36\% |
| 29,300 50/50 opt2a | 19,123 $\bigcirc$ | 21,927 $\bigcirc$ | 21,513 $\bigcirc$ | 20,925 $\bigcirc$ | 16,004 | 38\% |
| $\underline{\text { 29,300 50/50 opt1 }}$ | $\bigcirc 17,104 \bigcirc$ | 20,672 $\bigcirc$ | 19,676○ | 17,542 $\bigcirc$ | 13,161 | $34 \%$ |

Table 15. Comparison of sampling levels from Myers' et al. (2003) study and NMFS regional office estimates of Chinook bycatch levels from the pollock fishery, 1997-1999.

| Year | Area | Season | Myers' age <br> samples | Bycatch <br> Estimate | Myers' age <br> samples | Bycatch <br> Estimate |
| ---: | :---: | ---: | ---: | ---: | ---: | ---: |
| 1997 | All | A | 874 | 10,347 | $51 \%$ | $23 \%$ |
| 1997 | SE | B | 651 | 27,616 | $39 \%$ | $62 \%$ |
| 1997 | NW | B | 158 | 6,358 | $9 \%$ | $14 \%$ |
|  |  |  |  |  |  |  |
| 1998 | All | A | 906 | 15,118 | $51 \%$ | $30 \%$ |
| 1998 | SE | B | 730 | 35,307 | $41 \%$ | $69 \%$ |
| 1998 | NW | B | 138 | 820 | $8 \%$ | $2 \%$ |
|  |  |  |  |  |  |  |
| 1999 | All | A | 652 | 6,352 | $53 \%$ | $53 \%$ |
| 1999 | SE | B | 456 | 4,317 | $37 \%$ | $36 \%$ |
| 1999 | NW | B | 122 | 1,310 | $10 \%$ | $11 \%$ |


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