# 44th Northeast Regional Stock Assessment Workshop (44th SAW) 

## 44th SAW Stock Assessment Report

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07-03 44th Northeast Regional Stock Assessment Workshop (44th SAW). 44th SAW Assessment Summary Report. January 2007.

07-04 Estimated Bycatch of Loggerhead Sea Turtles (Caretta caretta) in U.S. Mid-Atlantic Scallop Trawl Gear, 2004-2005, and in Sea Scallop Dredge Gear, 2005, by KT Murray. February 2007.

07-05 Mortality and Serious Injury Determinations for Baleen Whale Stocks Along the United States Eastern Seaboard and Adjacent Canadian Maritimes, 2001-2005, by M Nelson, M Garron, RL Merrick, RM Pace III, and TVN Cole. February 2007.

07-06 The 2005 Assessment of Acadian Redfish, Sebastes fasciatus Storer, in the Gulf of Maine/Georges Bank region, by RK Mayo, JKT Brodziak, JM Burnett, ML Traver, and LA Col. April 2007.
07-07 Evaluation of a Modified Scallop Dredge's Ability to Reduce the Likelihood of Damage to Loggerhead Sea Turtle Carcasses, by HO Milliken, L Belskis, W DuPaul, J Gearhart, H Haas, J Mitchell, R Smolowitz, and W Teas. April 2007.

07-08 Estimates of Cetacean and Pinniped Bycatch in the 2005 Northeast Sink Gillnet and Mid-Atlantic Coastal Gillnet Fisheries, by D Belden. May 2007.
07-09 The Analytic Component to the Standardized Bycatch Reporting Methodology Omnibus Amendment: Sampling Design, and Estimation of Precision and Accuracy (2nd Edition), by SE Wigley, PJ Rago, KA Sosebee, and DL Palka. May 2007.

# 44th SAW Stock Assessment Report 

U.S. DEPARTMENT OF COMMERCE

National Oceanic and Atmospheric Administration
National Marine Fisheries Service
Northeast Fisheries Science Center
Woods Hole, Massachusetts
May 2007

## Northeast Fisheries Science Center Reference Documents

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This document's publication history is as follows: manuscript submitted for review -May 9, 2007; manuscript accepted through technical review -- May 9, 2007; manuscript accepted through policy review -- May 21, 2007; and final copy submitted for publication -- May 22, 2006. This document may be cited as:

44th Northeast Regional Stock Assessment Workshop (44th SAW): 44th SAW assessment report. US Dep. Commer., Northeast Fish. Sci. Cent. Ref. Doc. 07-10; 661 p. Available from: National Marine Fisheries Service, 166 Water Street, Woods Hole, MA 02543-1026.

The stock assessments which are the subject of this document were peer reviewed by a panel of assessment experts known as the Stock Assessment Review Committee (SARC). Panelists were provided by the Center for Independent Experts (CIE), University of Miami. Reports from the SARC panelists and a summary report from the SARC Chairman can be found at http://www.nefsc.noaa.gov/nefsc/saw.

## Assessment Report (44th SAW/SARC)

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## INTRODUCTION TO SAW-44 ASSESSMENT REPORT

The Northeast Regional Stock Assessment Workshop (SAW) process has three parts: preparation of stock assessments by the SAW Working Groups and/or by ASMFC Technical Committees / Assessment Committees; peer review of the assessments by a panel of outside experts who judge the adequacy of the assessment as a basis for providing scientific advice to managers; and a presentation of the results and reports to the Region's fishery management bodies.

Starting with SAW-39 (June 2004), the process was revised in two fundamental ways. First, the Stock Assessment Review Committee (SARC) is now a smaller panel with panelists provided by the University of Miami's Independent System for Peer Review (Center of Independent Experts, CIE). Second, the SARC no longer provides management advice. Instead, Council and Commission teams (e.g., Plan Development Teams, Monitoring and Technical Committees) formulate management advice, after an assessment has been accepted by the SARC.

Reports that are produced following SAW/SARC meetings include: an Assessment Summary Report - a brief summary of the assessment results in a format useful to managers; this Assessment Report - a detailed account of the assessments for each stock; and the SARC panelist report - a summary of the reviewer's opinions and recommendations as well as appendices consisting of a report from each panelist. SAW/SARC assessment reports are available online at http://www.nefsc.noaa.gov/nefsc/publication s/series/crdlist.htm. The CIE review reports and assessment reports can be found at http://www.nefsc.noaa.gov/nefsc/saw/.

The 44th SARC was convened in Woods Hole at the Northeast Fisheries Science Center, November 28 - December 4, 2006 to review three assessments (ocean quahog Arctica islandica, the northeast skate species complex [barndoor skate, Dipturus laevis; clearnose skate, Raja eglanteria; little skate, Leucoraja erinacea; rosette skate, Leucoraja garmani; smooth skate, Malacoraja senta; thorny skate, Amblyraja radiate; winter skate, Leucoraja ocellata)], and Atlantic surfclam Spisula solidissima. CIE reviews for SARC44 were based on detailed reports produced by the SAW Southern Demersal and Invertebrate Working Groups.

This Introduction contains a brief summary of the SARC comments, a list of SARC panelists, the meeting agenda, a list of working group meetings and a list of attendees (Tables 1 - 4). Maps of the Atlantic coast of the USA and Canada are also provided (Figures 1-5).

## Outcome of Stock Assessment Review Meeting

The ocean quahog assessment was accepted by the SARC. Current biomass appears well above the $\mathrm{B}_{\text {msy }}$ proxy and current F appears well below the $\mathrm{F}_{\text {msy }}$ proxy. The SARC was concerned with the biomass estimates from the main assessment model (KLAMZ) because the model did not link long-term average recruitment to virgin biomass. The reviewers also expressed concern about the accuracy and precision of the dredge efficiency estimate, the approach used to fill missing survey data cells, the appropriateness of proxies for $\mathrm{B}_{\text {msy }}$ and $\mathrm{F}_{\text {msy }}$, and the management of the entire offshore stock as a single unit.

Assessment results for the seven skate species were only partially accepted. The SARC rejected the estimates of the fishing mortality rate ( F ) as well as the proposed new Biological Reference Points (BRPs). The SARC felt that the absence of speciesspecific landings data made it extremely difficult to estimate F , and that estimates derived from the new model were too unreliable to accept at this time. The SARC felt that the existing BRPs were ad hoc and in need of improvement. The SARC felt that the proposed BRPs, derived from stockrecruit fits and length-based yield per recruit analysis, represented a positive step. However, the Committee did not feel that sufficient work had been done on the new BRPs to justify their use at this time. Accordingly, the assessment evaluated stock status with respect to the existing BRPs, and these results were accepted by the SARC.

No absolute estimates of total biomass or spawning stock biomass were made in the assessment. Finally, the SARC accepted work which examined the NEFSC Food Habits Database to estimate skate diets and skate consumptive demand in the ecosystem.

The Atlantic surfclam assessment was accepted by the SARC, although the Committee felt that the assessment could be improved by making better use of the available data on surfclam ages by developing a fully integrated age-structured model. Some of the concerns raised earlier about the ocean quahog assessment were also raised about the surfclam assessment. In addition, the Committee questioned whether the $\mathrm{B}_{\text {msy }}$ proxy (one half $\mathrm{B}_{1999}$ ) was appropriate, and suggested that this issue be reconsidered in a future assessment.

Table 1. 44th Stock Assessment Review Committee Panel.

# 44th Northeast Regional Stock Assessment Workshop (SAW 44) Stock Assessment Review Committee (SARC) Meeting 

November 28 - December 4, 2006
Woods Hole MA

## SARC Chairman (CIE):

Dr. Cynthia Jones, chair
Center for Quantitative Fisheries Ecology
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## SARC Panelists (CIE):

Dr. Vivian Haist, review panelist
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Mr. Patrick Cordue, review panelist
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Khandallah
Wellington 6035
New Zealand
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Email: plc@isl-solutions.co.nz

Table 2. Agenda, 44th Stock Assessment Review Committee Meeting.
44th Northeast Regional Stock Assessment Workshop (SAW 44)
Stock Assessment Review Committee (SARC) Meeting
Stephen H. Clark Conference Room - Northeast Fisheries Science Center Woods Hole, Massachusetts

November 28 - December 4, 2006

## AGENDA (11-27-06)

| TOPIC | PRESENTER | SARC LEADER | RAPPORTEUR |
| :--- | :--- | :--- | :--- |

Tuesday, 28 November (1:00-5:00 PM) $\qquad$
Opening
Welcome
James Weinberg, SAW Chairman
Introduction
Cynthia Jones, SARC Chairman
Agenda
Conduct of Meeting
Ocean quahog (A)
Larry Jacobson Vivian Haist Toni Chute
SARC Discussion
Cynthia Jones

Wednesday, 29 November (8:30 - Noon) $\qquad$
Skates (B)
Kathy Sosebee Patrick Cordue Michelle Traver
SARC Discussion
Cynthia Jones

Wednesday, 29 November (1:15-5:00 PM). $\qquad$
Atlantic surfclam (C)
Larry Jacobson Vivian Haist Laurel Col
SARC Discussion
Cynthia Jones

Table 2 continued.

Thursday, 30 November (8:30 - 5:00 PM) $\qquad$
Revisit Assessments ( $\mathrm{A}-\mathrm{C}$ ) with presenters, as needed.

Friday, 1 December (8:30 AM - )
Revisit Assessments (A - C) with presenters, if needed.
SARC Report writing. (closed)

Saturday, 2 December - Monday, 4 December .
SARC Report writing. (closed)

Table 3. 43rd Stock Assessment Workshop, list of working groups and meetings.
Assessment Group Chair Species Meeting Date/Place

SAW Southern Demersal Working Group
Paul Rago, NMFS NEFSC
Skate complex Oct. 24-26, 2006
Woods Hole

| Sondre Aanes | Institute of Marine Research, Bergen, Norway |
| :--- | :--- |
| Larry Alade | NEFSC |
| Laurel Col | NEFSC |
| Mike Fogarty | NEFSC |
| Mike Frisk | SUNY, Stony Brook |
| Todd Gedamke | VIMS |
| Dvora Hart | NEFSC |
| Fiona Hogan | UMass/SMAST |
| Chris Legault | NEFSC |
| Jason Link | NEFSC |
| Alyssa MacDonald | UMass/SMAST |
| Ralph Mayo | NEFSC |
| Hassan Moustahfid | NEFSC |
| Paul Nitschke | NEFSC |
| Mike Pennington | Institute of Marine Research, Bergen, Norway |
| Anne Richards | NEFSC |
| Gary Shepherd | NEFSC |
| Brian Smith | NEFSC |
| Katherine Sosebee | NEFSC |
| Michele Traver | NEFSC |
| Megan Tyrrell | NEFSC |
| Susan Wigley | NEFSC |



Invertebrate Working Group
M. Terceiro, NMFS NEFSC

Atl. Surfclam
Sept. 25-27,2006
Oct. 16-18, 2006
Oct. 30- Nov. 1, 2006
Woods Hole
T. Alspach (Sea Watch International, Ltd.)
A. Chute (NEFSC)
H. Dobby (Invited external participant, FRS Marine Laboratory, Aberdeen, Scotland)
C. Heaton (Mid-Atlantic Fishery Management Council, MAFMC)
J. Heifitz (Invited external participant, NMFS, AKFSC)
T. Hoff (MAFMC)
L. Jacobson (Northeast Fisheries Science Center, NEFSC) - assessment lead
C. Pickett (NEFSC)
E. Powell (Haskin Shellfish Laboratory, Rutgers University)
D. Wallace (Wallace \& Associates, Inc.)
J. Womack (Wallace \& Associates, Inc.)
J. Weinberg (NEFSC)

Table 4. 44th SAW/SARC, List of Attendees

| J. Womack | Wallace and Assoc. |
| :--- | :--- |
| T. Hoff | MAFMC |
| P. Nitschke | NEFSC |
| C. Pickett | NEFSC |
| D. Wallace | Wallace and Assoc. |
| L. Col | NEFSC |
| M. Terceiro | NEFSC |
| L. Jacobson | NEFSC |
| A. Applegate | NEFMC |
| F. Hogan | UMass/SMAST |
| A. MacDonald | UMass/SMAST |
| J. Moser | NEFSC |
| A. Richards | NEFSC |
| M. Traver | NEFSC |
| D. Hart | NEFSC |
| R. Brown | NEFSC |
| G. Shepherd | NEFSC |
| T. Alspach | Sea Watch International |
| L. O’Brien | NEFSC |



Figure 1. Offshore depth strata sampled during Northeast Fisheries Science Center bottom trawl research surveys.


Figure 2. Inshore depth strata sampled during Northeast Fisheries Science Center bottom trawl research surveys.


Figure 3. Statistical areas used for reporting commercial catches.


Figure 4. Catch reporting areas of the Northwest Atlantic Fisheries Organization (NAFO) for Subareas 3-6.


Figure 5. Clam strata for NEFSC resource surveys.

## A. ASSESSMENT OF OCEAN QUAHOGS ${ }^{1}$

### 1.0 TERMS OF REFERENCE (TOR)

1. Characterize the commercial and recreational catch including landings and discards.

Completed--Commercial landings were updated through 2005. Discards are negligible. However, a $5 \%$ allowance for incidental mortality due to contact with fishing gear is used in all assessment calculations.
2. Estimate fishing mortality, spawning stock biomass, and total stock biomass for the current year and characterize the uncertainty of those estimates. If possible, also include estimates for earlier years.

Completed--Fishing mortality, fishable and total stock biomass were estimated for 19782005. Confidence intervals were calculated to characterize uncertainty. Spawning biomass was calculated on an approximate basis after the SARC based on reviewers' suggestions.
3. Either update or re-estimate biological reference points (BRPs; proxies for $B_{M S Y}$ and $F_{M S Y}$ ), as appropriate. Comment on the scientific adequacy of existing and redefined BRPs.

Partially completed—Biomass reference points $B_{1978}$ (a proxy for virgin biomass), the management target $B_{M S Y}=1 / 2 B_{1978}$ and the management threshold $B_{\text {Threshold }}=1 / 4 B_{1978}$ were updated based on new information. Fishing mortality reference points ( $F_{\text {Target }}=F_{0.1}$ and $F_{\text {Threshold }}=F_{25 \%}$ ) were updated using new information about fishery selectivity and maturity in a length based per recruit model. Problems with the scientific adequacy of the current existing $F_{\text {Threshold }}$ proxy for $F_{\text {MSY }}$ are described. However, there was insufficient time to complete analyses required to recommend an optimum alternative. This work was deferred because fishing mortality rates are very low and there was no urgency.
4. Evaluate current stock status with respect to the existing BRPs, as well as with respect to new or re-estimated BRPs (from TOR 3).

Completed—Stock biomass and fishing mortality estimates for 2005 were compared to updated reference points.
5. Recommend what modeling approaches and data should be used for conducting single and multi-year stock projections, and for computing TACs or TALs.

Completed-A simple modeling approach and data were recommended for projecting biomass and fishing mortality of the ocean quahog stock through 2010.

[^0]6. If possible,
a) provide numerical examples of short term projections (2-3 years) of biomass and fishing mortality rate, and characterize their uncertainty, under various TAC/F strategies and
b) compare projected stock status to existing rebuilding schedules as appropriate.

Completed-Example calculations and projections through 2010 were carried out assuming three quota levels and at $F=F_{0.1}$.
7. Review, evaluate and report on the status of the SARC/Working Group Research Recommendations offered in recent SARC-reviewed assessments.

Completed-Several key research recommendations were accomplished in this assessment. In particular: 1) a survey was completed, reference points were calculated and biomass and fishing mortality were estimated for ocean quahog in Maine waters; 2) field data collected during 2002 and new data collected during 2005 were examined to determine if survey and commercial dredge efficiency depends on depth, sediment type or clam density; 3) survey selectivity and fishery selectivity curves were used to better interpret survey data; and 4) reference points were revised in this assessment using a new length based model and new fishery selectivity and maturity at length curves.

### 2.0 EXECUTIVE SUMMARY

A) This assessment for ocean quahog in the US EEZ is based on fishery data landings and LPUE data for 1978-2005 and NEFSC survey data for 1982-2005. Based on assessment results, the ocean quahog population is a relatively unproductive stock which is being fished down slowly towards its $B_{M S Y}$ reference point ( $1 / 2$ virgin biomass, estimated as $50 \%$ of biomass during 1978) gradually after about three decades of relatively low fishing mortality.
B) Ocean quahog in the US EEZ are not overfished and overfishing is not occurring. Stock biomass during 2005 was 3.039 million mt and above the revised management target of $1 / 2$ virgin biomass $=1.987$ million mt . The fishing mortality rate during 2005 for the exploitable region (all areas but GBK) was $F=0.0077 \mathrm{y}^{-1}$ and below the revised management target level $F_{0.1}=0.0278 \mathrm{y}^{-1}$.
C) Depletion experiments carried out during 1997-2005 on a cooperative basis with the fishing industry were used to estimate the efficiency of the NEFSC survey dredge, which is the basis for estimating biomass and fishing mortality. Based on all experiments to date, the NEFSC survey dredge has a capture efficiency of $16.5 \%$, which is less than values used in the earlier assessments ( $e=0.269$ in SARC38, and 0.346 in SARC31).
D) Biomass and fishing mortality estimates were improved in this assessment using new information about size selectivity of survey and commercial clam dredges.
E) The estimates of biomass and fishing mortality in this assessment do not include biomass or landings from Maine waters. However, stock biomass is small ( $\sim 1 \%$ ) relative to the rest of the EEZ and calculations would not change appreciably if Maine were included. As described below, the Maine fishery and stock component were assessed separately (Russell 2006). Highlights from the Maine assessment are presented here but interested persons should consult the Maine stock assessment report.
F) Biological reference points based on per recruit models ( $F_{0.1}$ and $F_{25 \%}$ ) were recalculated based on new length based per recruit model, and new fishery selectivity and maturity curves (see below).

| Reference Point | Old <br> (SARC- <br> $38)$ | New |
| :---: | :---: | :---: |
| $\mathrm{F}_{0.1}$ (target) | 0.0275 | 0.0278 |
| $\mathrm{~F}_{\text {MAX }}$ | 0.1810 | 0.0760 |
| $\mathrm{~F}_{25 \%}$ (threshold) | 0.0800 | 0.0517 |
| $\mathrm{~F}_{50 \%}$ | 0.0200 | 0.0180 |

G) From a technical perspective, the current threshold reference point for fishing mortality $F_{25 \%}=0.0517 \mathrm{y}^{-1}$ is a poor proxy for $F_{M S Y}$ in a long-lived species like ocean quahog with natural mortality rate $M=0.02 \mathrm{y}^{-1}$.
H) Proxies for virgin biomass and $B_{M S Y}$ in this assessment are substantially larger than in NEFSC (2003). In particular, the revised proxy in this assessment for $B_{M S Y}(1 / 2$ virgin biomass) was 1.987 million mt compared to 1.5 million mt for $B_{M S Y}$ in the last assessment. The new estimates are different primarily because revised survey dredge efficiency estimates are smaller ( $e=0.165$ instead of 0.269 0.346 ).
I) Biomass during 2005 was $76 \%$ of biomass during 1978 for the entire stock and $66 \%$ for the entire stock less GBK
J) Fishery LPUE, survey trends and assessment model estimates show substantial declines in stock biomass in southern regions (SVA, DMV and NJ) where the fishery has been continually active. In particular, biomass during 2005 was $5 \%$, $34 \%$ and $44 \%$ of biomass during 1978 for SVA, DMV and NJ. Biomass trends in northern regions which did not support the fishery until recently (LI, SNE and GBK) are relatively flat and stable. Biomass during 2005 was $94 \%, 75 \%$ and $100 \%$ of biomass during 1978 for LI, SNE and GBK.
K) An increasingly large fraction of the stock ( $83 \%$ during 2005 compared to $70 \%$ during 1978) is in northern regions (LI, SNE) where fishing is relatively recent and in the GBK region, which is not fished due to risk of PSP contamination.
L) Fishing mortality rates for southern areas where the fishery has been continually active (SVA, DMV and NJ) peaked in the late 1980's and early 1990's then declined as fishing effort shifted towards the north. Fishing mortality rates in northern areas were nearly zero before 1990 and increased substantially afterwards as fishing effort shifted towards the north. Fishing mortality rates for the entire stock increased from near zero in 1978 to average about $0.006 \mathrm{y}^{-1}$ ( $0.010 \mathrm{y}^{-1}$ for the entire stock less GBK) during early 1990 through 2005.
M) Recruitment events appear to be regional and sporadic (i.e. often separated by decades). Survey length composition data show that recruitment occurs throughout the resource sporadically and at an apparently low rate. Based on survey length composition data and published studies, at least some recent recruitment (small ocean quahog) is evident in DMV, NJ, LI, SNE and GBK during recent years. The potential contribution of recent recruitment to stock biomass and productivity is unknown.

## Maine waters

N) Ocean quahog in Maine waters are part of the unit stock covered by the FMP and support a small fishery that is managed under limited entry and quota systems that are separate from the individual transferable quota (ITQ) system used for ocean quahog in the rest of the EEZ.
O) The fishery and biological characteristics of ocean quahog in Maine waters are unique. In particular, the Maine fishery targets small ocean quahog for sale on the half shell market at prices roughly ten times the prices paid for larger ocean quahogs taken elsewhere in the EEZ. Management goals have for ocean quahog in Maine waters have not been described.
P) A survey and stock assessment were completed by the State of Maine for the portion of the ocean quahog stock occupying the major fishing grounds in Maine waters (Russell 2006). Most of the results presented here for the Maine fishery are from Russell (2006).
Q) Assessment results for Maine show relatively high levels of fishing effort and landings in recent years. LPUE levels have declined since the peak in 2002, but remain at relatively high levels overall.
R) Based on a per recruit model analysis, $F_{M A X}=0.0561, F_{0.1}=0.0247$ and $F_{50 \%}=$ $0.013 \mathrm{y}^{-1}$ for ocean quahog in the major fishing grounds of Maine waters only. These reference points are provided only for comparison and do not have any special status as targets or thresholds.
S) Based on survey results and dredge efficiency estimates for Maine, the biomass of ocean quahog during 2005 that was available to the fishery in Maine waters was

22,493 mt meats. In comparison, catch (landings plus a 5\% incidental mortality allowance) during 2005 was 505 mt meats.
T) Fishing mortality during 2005 in the areas surveyed and the principal fishing grounds in Maine waters was estimated to be $F=505 \div 22,493=0.022 \mathrm{y}^{-1}$, which is almost equal to $F_{0.1}=0.0247 \mathrm{y}^{-1}$, a reference point that would provide relatively high levels of yield while preserving some spawning stock.

### 3.0 INTRODUCTION

Ocean quahog (Arctica islandica) in the US Exclusive Economic Zone (EEZ) form a single stock for management purposes. With the exception of a relatively small component off the coast of Maine, the EEZ fishery is managed by under a single individual transferable quota (ITQ) system that was established for ocean quahog and Atlantic surfclam (Spisula sodidissma) in 1990. Murawski and Serchuk (1989) and Serchuk and Murawski (1997) provide detailed information about the history and operation of the fishery.

The ocean quahog fishery component off Maine is managed under a relatively small quota that is separate from the quota used to manage the ITQ fishery. The Maine component is of interest because of differences in biological, fishery, market and management characteristics. The ocean quahog assessment this year consists of two reports. The first (Russell 2006) estimates biomass, fishing mortality and per recruit reference points for the stock component in Maine waters based on a survey in 2005 and estimates of survey dredge efficiency. The second (this report) deals with the EEZ as a whole based on the NEFSC clam survey for 1982-2005 and summarized key aspects of the assessment for Maine waters.

Overfishing definitions and other management measures apply at the level of the entire stock although technical information is provided at the level of smaller stock assessment regions (Figure A1 and see below). Georges Bank (GBK) has been closed to ocean quahog harvesting since 1990 when Paralytic Shellfish Poison (PSP) was detected.

| Stock Assessment Region | Abbreviation |
| :---: | :---: |
| Maine | MNE |
| Georges Bank | GBK |
| Southern New England | SNE |
| Long Island | LI |
| New Jersey | NJ |
| Delmarva | DMV |
| Southern Virginia and North <br> Carolina | SVA |

Categories and units used in this assessment are defined below.

| Unit | Equivalent |
| :---: | :---: |
| Industry or Mid-Atlantic bushel (Industry bu) | $1.88 \mathrm{ft}^{3}$ |
| Maine (US standard) bushel (Maine bu) | $1.2448 \mathrm{ft}^{3}$ |
| Industry bushels x 10 | Pounds meat wt |
| Industry bushels x 4.5359 | Kilograms meat wt |
| Cage | 32 Industry bushels |
| Vessel ton class 1 | $1-4$ gross registered tons (GRT) |
| Vessel ton class 2 | $2-50 \mathrm{GRT}$ |
| Vessel ton class 3 | $51-150 \mathrm{GRT}$ |
| Vessel ton class 4 | $151-500 \mathrm{GRT}$ |
| Vessel ton class 5 | $501-1000 \mathrm{GRT}$ |

## Previous and current assessments

Stock assessments for ocean quahog in the EEZ were completed by NEFSC (1995; 1998; 2000; 2004). The last assessment (NEFSC 2004) concluded that the EEZ ocean quahog resource was not overfished and that overfishing was not occurring. This stock assessment arrives at the same conclusion.

The last assessment (NEFSC 2004) concluded that the qualitative condition of the stock off the coast of Maine was unknown and recommended that the Maine conduct a comprehensive survey and conduct experiments to estimate survey dredge efficiency. These recommendations were completed in this assessment and are presented in a separate report (Russell 2006).

## Biological characteristics ${ }^{2}$

Ocean quahog are common around Iceland, in the eastern Atlantic as far south as Spain, and in the western Atlantic as far south as Cape Hatteras (Theroux and Wigley 1983; Thorarinsdottir and Einarsson 1996; Lewis et al. 2001). They are found at depths of 10-400 m, depending on latitude (Theroux and Wigley 1983; Thompson et al. 1980). The US stock is almost completely within the EEZ outside of state waters at depths of about 20-80 m. In a study of the mitochondrial cytochrome $b$ gene, Dahlgren et al. (2000) did not find geographical differentiation between samples taken along the US coast from Maine to Virginia.

Ocean quahog are long-lived with some individuals aged at over 200 yrs (Jones 1983; Steingrimsson and Thorarinsdottir, 1995). Early studies of populations off New Jersey and Long Island (Thompson et al. 1980; Murawski et al. 1982) demonstrate that clams ranging in age from 50-100 years are common. In stock assessment work, adult ocean quahog are assumed to die from natural causes at the rate of about $2 \%$ annually (instantaneous rate of natural mortality $M=0.02 \mathrm{y}^{-1}$ ).

Ocean quahog grow slowly after the first years of life (Lewis et al. 2001, Figure A56). Maximum size is typically about 110 mm in shell length (SL) although larger specimens are common. Individuals large enough to recruit to the fishery grow only $0.51-0.77 \%$ per year in meat weight and $<1 \mathrm{~mm}$ per year in shell length (NEFSC 2004).

Size and age at maturity are variable. Off Long Island, the smallest mature quahog found was a male 36 mm long and 6 years old; the smallest and youngest mature female was 41 mm long and 6 yr old (Ropes et al. 1984). Some clams in this region are still sexually immature at ages of 8-14 years (Thompson et al. 1980; Ropes et al. 1984).

[^1]Females are more common than males among the oldest and largest individuals in the population (Ropes et al. 1984; Fritz 1991). Recruitment events are regional and infrequent in ocean quahog with decadal periods of little or no recruitment (Powell and Mann 2005).

### 4.0 COMMERCIAL AND RECREATIONAL CATCH (TOR-1)

Landings and quotas for the ITQ segment of the EEZ fishery are reported in different bushel units than landings and quotas for the fishery off Maine (Russell 2006). In particular, "ITQ" bushels are used for the ITQ component and "standard" bushels are used for the Maine component. Biomass and landings from both fishery components are reported in this assessment as meat weights (the weight of marketable product after removal from the shell), unless otherwise noted, because meat weights are directly comparable.

Total EEZ landings (including the ITQ and Maine fishery components) were relatively high during 1987-1996 with a peak of 22.5 thousand mt meats (Tables A1-A2 and Figure A2) or 4.9 million ITQ bushels (Table A3) during 1992. After 1996, landings declined to a low of about $15,000 \mathrm{mt}$ meats ( 3.3 million ITQ bushels) during 2000 and then increased to about $19,000 \mathrm{mt}$ meats ( 4.2 million ITQ bushels) during 2003. Landings declined after 2003 to about $14,000 \mathrm{mt}$ meats ( 3.2 million ITQ bushels) during 2005, which was the lowest level since 1981. Industry sources report that low landings during the most recent years were due to low market demand. The ITQ component accounted for almost all ( $\geq 98 \%$ ) of total EEZ landings during 1990-2005. Landings from Maine waters are minor in comparison to EEZ landings (Tables A2-A3 and Figure A2).

Landings from Maine waters increased steadily after 1990 to relatively high levels ( $\geq 326$ thousand mt meats annually) during 2000-2003 (Tables A2-A3). Landings in Maine waters decreased after 2003 to 294 thousand mt meats during 2005, which was the lowest level since 1999.

Landings by the ITQ component averaged $85 \%$ of the EEZ quota during 19902005 (Table A1). In contrast, the 100,000 Maine bushel quota allocated for ocean quahog in Maine waters was usually exhausted during 1999-2005 with vessels leasing ITQ shares in some years to harvest more than 100,000 mt meats from Maine waters (Tables A2-A3).

Landings of quahogs from state waters outside of Maine are near zero because ocean quahog are found offshore in relatively deep water. Landings in recreational fisheries are nil because commercial clam dredges are required to harvest ocean quahog and because ocean quahog are an industrial product with no recreational value.

### 4.1 Prices

Nominal exvessel prices for ITQ ocean quahog landings (expressed as dollars per ITQ bushel) decreased slightly during 2001-2004 (Table A4 and Figure A3). In real terms, prices during 2004 were about the average of real prices during 1994-2004. Prices for ocean quahog harvested in Maine waters (dollars per ITQ bushel) were roughly ten times higher than prices for ocean quahogs harvested in the rest of the EEZ (Table A4 and Figure A3).

### 4.2 Fishing effort

Total hours fished annually in the ITQ fishery component decreased from a peak of about $40,000 \mathrm{hr} \mathrm{y}^{-1}$ during 1991-1994 to about $30,000 \mathrm{hr} \mathrm{y}^{-1}$ during 1996-2004 and then decreased to about $20,000 \mathrm{hr} \mathrm{y}^{-1}$ during 2005 (Table A5 and Figure A4). The total number of trips in the ITQ fishery decreased steadily from about 3000 trip $\mathrm{y}^{-1}$ during 1991 to about 1000 trips y ${ }^{-1}$ during 2005 (Figure A5). In contrast, hours fished and trips increased in the Maine fishery component during 1991-2005. The number of active permits (vessels with landings) remained relatively constant during 1996-2004 but declined slightly during 2005 (Figure A6). Number of active permits, and fishing effort (hours fished and numbers of trips) is high in Maine waters relative to other stock assessment regions in the EEZ (Figure A4-A6).

### 4.3 Landings per unit effort (LPUE)

It is useful express trends in LPUE in terms of average catch rates for an actual vessel because industry sources report that fishing in the ITQ sector is profitable when LPUE is at least 110-120 bushels $\mathrm{h}^{-1}$ (D. Wallace, pers. comm.). The break-even LPUE reported in the last was assessment 80 bushels $\mathrm{h}^{-1}$ (NEFSC 2004). The new estimate is higher because of inflation, increased steaming time to relatively distant fishing grounds, operation of new larger vessels, and increased costs for food, fuel, insurance, etc. These estimates are not applicable to fishing in Maine waters.

LPUE (LPUE, bushels landed per hour fished) in the ocean quahog fishery may be a better measure of fishing success than a measure of stock abundance because changes in abundance or biomass for regions as a whole may be masked by concentration and movement of fishing effort between regions where ocean quahog density and catch rates are high (see below). In spite of these potential problems, LPUE and NEFSC clam survey data are highly correlated (see Section 5).

Trends in LPUE were not sensitive to the details of calculation (Table A6 and Figure A7). Three measures of LPUE were calculated for each stock assessment region based on vessel size classes 3-4 for the ITQ fishery and vessel size classes 1-2 for the Maine fishery. The size classes used in calculating LPUE accounted for almost all landings. "Nominal mean LPUE" was the average catch rates for individual trips in each region and year. "Total bushels/total hours" was the ratio of total landings and total hours fished. The "standardized index" for each region was calculated from the year effects estimated in a general linear model (described below).

General linear models (GLM) used to standardize LPUE data for ocean quahog were fit to trip-level log book data. A separate model was run for each stock assessment region because trends differed among regions. The dependent variable in GLM models was $\log$ LPUE (ITQ or Maine bushels per hour fished). There was no need to add a constant before taking logs because catch was greater than zero for all trips. The models included categorical year, month and vessel effects, which were statistically significant in every case. Other factors might have been included in GLM models but vessels and months were of special interest and other model formulations gave very similar trends in standardized LPUE.

The time series of standardized LPUE for each region was computed from the back-transformed year effects with adjustments so that the indices for each area were in units of LPUE for a single vessel that fished in each of the DMV, NJ, LI and SNE stock assessment regions. A different vessel was chosen for MNE.

GLM results show that standardized LPUE during 1985 declined in the DMV, NJ and LI stock assessment regions and fluctuated without trend in the SNE region (Table A6 and Figure A8). In the Maine fishery, standardized LPUE increased during 19912000, decreased afterwards but was still relatively high during 2005. Differences in trends among regions are discussed in detail below.

GLM results show that LPUE is slightly higher in the DMV, NJ, LI and SNE regions during February-April (Figure A9). LPUE in the Maine fishery peaks in June.

### 4.4 Spatial patterns in fishery data

Spatial patterns are important in interpreting fishery data and in managing fisheries for sessile and relatively unproductive organisms like ocean quahog. The ocean quahog stock is a complicated spatial mosaic with scattered productive and profitable fishing grounds where abundance is high and where fishing mortality tends to be concentrated. The size of productive fishing grounds for ocean quahog appears to be less than the size of ten minute squares (TNMS, $10^{\prime} \times 10^{\prime} \cong 100 \mathrm{~nm}^{2}$ ), which are the smallest spatial strata consistently reported on logbooks and used in this stock assessment.

As described in NEFSC (2004), spatial patterns in cumulative landings, cumulative effort and LPUE are related. The spatial distribution of landings and fishing effort in the ITQ fishery component changed markedly over time. During the 1980s, nearly all of the landings (Figure A2) and fishing effort (Figure A4-A5) were from the southern DMV and NJ stock assessment regions. As LPUE declined in the southern DMV and NH stock assessment regions (Figure A8), fishing effort and landings shifted offshore and north to the LI and SNE stock assessment regions. During 2005, in particular, the southern DMV and NJ stock assessment regions accounted for less than $20 \%$ of landings and fishing effort while the bulk of landings and effort (outside of Maine waters) were from LI (Figures A2 and A4-A6).

## Fishery data by ten-minute square (TNMS)

All vessels that fish for ocean quahog in the EEZ use logbooks to report landings and fishing effort by TNMS for each trip. TNMS are identified by six digit numbers. For example, TNMS 436523 is a ten-minute square that lies within the one-degree square with southeast corner at $43^{\circ} \mathrm{N}$ and $65^{\circ} \mathrm{E}$. TNMS are formed by dividing one-degree squares further into six columns and six rows that are $10^{\prime}$ wide. Columns are numbered 1-6 counting from west to east and the column number is given in the TNMS name before the row number. Rows are numbered 1-6 counting from north to south. Thus, TNMS 436523 is the ten-minute square whose southeast corner is at $43^{\circ} 30^{\prime} \mathrm{N}$ and $65^{\circ}$ 40' E.

Landings (Figure A10) during 1980-1990 were concentrated in relatively few TNMS that were primarily in the south and relatively inshore. Over time, TNMS with highest landings shifted offshore and north. Landings during 2001-2005 were concentrated in the LI stock assessment region.

Fishing effort (Figure A11) was concentrated in a few southern TNMS during 1980-1990 with three adjacent TNMS having effort levels higher than $1,000 \mathrm{~h} \mathrm{y}^{-1}$ and appreciable fishing effort south of $38^{\circ} \mathrm{N}$. Fishing effort spread into additional offshore and northern TNMS during 1991-1995 and 1996-2000. After 1995, there were few or no TNMS with effort levels above $1000 \mathrm{~h} \mathrm{y}^{-1}$. During 2001-2005, there was a no fishing effort south of $38^{\circ} \mathrm{N}$.

LPUE (Figure A12) was relatively high inshore and south during 1980-1990 with ten TNMS that had LPUE $\geq 161$ ITQ bushels $h^{-1}$. LPUE in the area below $40^{\circ} \mathrm{S}$ was generally high. LPUE declined in the south and fishing effort spread northward during 1991-1995 where LPUE was relatively high. During 1996-2000, LPUE declined in both the northern and southern areas. By 2001-2005, LPUE was often $\leq 80$ ITQ bushels $\mathrm{h}^{-1}$ below $40^{\circ} \mathrm{S}$.

## Trends

Trends in landings and LPUE during 1980-2005 were plotted for individual TNMS that were important in the fishery (Figures A13-A15). Important TNMS were selected by sorting TNMS according to total landings during 1980-1990, 1991-1995, 1996-2000 and 2001-2005 and then selecting the top 20 TNMS during each time period. All of the TNMS selected in this manner were combined to form a single unique set of 79 TNMS that were important to the fishery at some time during 1980-2005.

Trends in LPUE for individual TNMS tend to be relatively high in during the first years of exploitation and then to subsequently decline as effort, annual landings and cumulative landings increase over time (Figures A13-A15). Decreasing trends in LPUE appear strongest in southern areas such as TNMS 377422 to 397326 with the longest history of exploitation. LPUE does not appear to increase in a TNMS once fishing effort decreases.

Unlike LPUE which is highest in the first years of exploitation, landings and fishing effort tend to peak after 5-10 years of exploitation while LPUE is still relatively high and then to decrease over a 5-10 y period as grounds are fished down (Figures A13A15). In some TNMS with low recent LPUE levels (e.g. TNMS 387443-397316), fishing effort increased during 2001-2005 with some increase in landings.

### 4.5 Bycatch and discard

Landings and catch are almost equal in the ocean quahog fishery because discards are nil. Discard of ocean quahog in the ocean quahog fishery does not occur because undersize animals are automatically released by automatic sorting equipment. However, some incidental mortality occurs. Based on Murawski and Serchuk (1989), NEFSC (2004) assumed incidental mortality rates of $\leq 5 \%$ for ocean quahog damaged during fishing but not handled on deck. As in previous assessments, fishing mortality and other stock assessment calculations in this report assume $5 \%$ incidental mortality rates (i.e. landings x $1.05=$ assumed catch).

Bycatch of ocean quahog probably occurs in fishing for Atlantic surfclam but has not been quantified and is certainly minor. Off DMV and SVA in the southern end of the ocean quahog's range, survey catches including both surfclam and ocean quahog have become more common in recent years as surfclam have shifted towards deeper water in response to warm water conditions (Weinberg et al. 2005). However, mixed loads of surfclam and ocean quahog are not acceptable to processors and it is not practical to sort catches at sea so that vessels would tend to avoid areas where both species might be caught.

Bycatch and discard of ocean quahogs in other fisheries is nil. Ocean quahogs are not vulnerable to bottom trawls, scallop dredges (because they are too deep in sediments), or hook and line gear.

### 4.6 Commercial size-composition data

Commercial length composition data (shell lengths, SL) for ocean quahogs collected by port agents from landings indicate that the size composition of ocean quahog captured in the DMV stock assessment region differed during 1987-1994, 1995-2000 and 2001-2005 (Figure A16). Lengths for DMV during 1987-1994 and 2001-2005 were similar.

Commercial length composition data for NJ were stable during 1982-2002 with smaller ocean quahog landed during 2003-2005 (Figure A17). Length data for LI include relatively high proportions of large individuals ( $11-12 \mathrm{~cm} \mathrm{SL}$ ) during 1997-1999 (Figure A18). Length data for SNE during 1998-2005 were generally stable but with smaller ocean quahog landed during 1997-2000 (Figure A19). According to NEFSC (2004), smaller sizes landed from SNE during 1997-2000 were due to vessels targeting specific beds with relatively small ocean quahogs that had relatively high meat yield.

### 4.7 Fishery selectivity

Commercial fishery selectivity estimates used in this assessment for ocean quahog are from Thorarinsdottir and Jacobson (2005) who estimated selectivity of commercial dredges that harvest ocean quahog off Iceland. The selectivity curve $s_{L}=1 /\left(1+e^{7.63-0.105 L}\right)$, where $L$ is shell length in mm , indicates that about $10 \%, 50 \%$ and $90 \%$ of ocean quahog are available to the fishery at 51,72 , and $93 \mathrm{~mm} \mathrm{SL}(9,28$ and 86 y , based on the growth curve in Figure A59).

Dredges and towing speed in the US fishery are very similar to dredges and tow speed used in the selectivity experiments. The dredge used for selectivity experiments was $24 \mathrm{ft}(7.35 \mathrm{~m})$ in length, $5 \mathrm{ft}(1.5 \mathrm{~m})$ high and $12 \mathrm{ft}(3.65 \mathrm{~m})$ wide. The cutting blade was $10 \mathrm{ft}(3.05 \mathrm{~m})$ wide and set to penetrate sediments to a depth of 3 in $(8 \mathrm{~cm})$. The dredge was made of steel bars with intervening spaces of $11 / 4$ in ( 3.5 cm ) and was towed at about 2.1 knots ( $3.9 \mathrm{~km} \mathrm{~h}^{-1}$ ). Water pressure supplied to jets on the dredge from a pump on the ship was about 109 psi ( 7.5 bars). Water pressure levels in the US fishery are usually lower ( $\sim 80 \mathrm{psi}$ ) but water pressure probably has relatively little effect on size selectivity. Fishery selectivity curves are used in tracking trends in fishable biomass, estimating fishing mortality and in calculating biological reference points.

### 5.0 MORTALITY AND STOCK BIOMASS (TOR-2)

Mortality and stock biomass estimates for ocean quahog in the US EEZ are based on triennial NEFSC clam surveys, cooperative field studies used to measure survey dredge efficiency, and fishery data.

### 5.1 NEFSC Clam Surveys-Results

NEFSC clam surveys have been conducted since 1965 and are the main source of fishery-independent information about long term trends in abundance, biomass (Table A7, Figure A20), recruitment (Figure A21), stock distribution (Figures A22-A25 and Appendices A7-A8) and population length composition (Figure A26) for ocean quahog in the EEZ. The small area of coastal Maine waters is not covered by the NEFSC clam survey but it is minor in terms of stock biomass ( 20 vs. 2,700 thousand mt meats, Russell 2006) and landings ( $500 \mathrm{vs} .14,000 \mathrm{mt}$ meats).

Based on survey data and in general terms (see below for details), fishable abundance (mean number per tow), stock biomass (mean kg tow) and spawning biomass (mean kg/tow) declined during 1982-2005 in southern areas (SVA, DMV and NJ) where the bulk of fishing has occurred while fishable biomass in northern areas (LI, SNE and GBK) remained relatively high and stable (with the exception of GBK in the 1999 survey). LI is the only area with clear evidence of strong recruitment after 1982 based on survey length and recruit trend data. In particular, length data from LI show ocean quahog at 65 mm SL during 1978 that grew slowly over time and became indistinguishable from the rest of the LI stock by about 1994 (Figure A26). Recruitment trend data for LI are higher prior to 1994 than afterwards and variable in other regions (Figure A21). Trends in spawning and stock biomass were nearly the same.

## Survey methods

Survey data used in this assessment were from surveys during 1982-2005 by the $R / V$ Delaware II, which were carried out during the summer (June-July), using the standard NEFSC survey hydraulic dredge with a submersible pump, 152 cm (60 in) blade 5.08 cm and small $5.08 \mathrm{~cm}(2 \mathrm{in})$ mesh liner. The survey dredge differs from commercial dredges in being smaller, using the small mesh liner, and in having the pump mounted on the dredge, rather than the deck of the vessel. The survey dredge used since 1982 catches ocean quahog as small as 50 mm SL with some reliability.

Surveys prior to 1982 were not used in this assessment because they were carried out during different seasons, used other sampling equipment or, in the case of 1981, have not been integrated into the clam survey database (Table A7 in NEFSC 2004). The last stock assessment for ocean quahog (NEFSC 2004) used survey data for 1978-1980 assuming that catchability was different during than in later surveys. In effect, the data for 1979-1980 were treated as a short separate survey time series that had little or no effects on stock assessment estimates. Catchability coefficients for earlier surveys were much different than for surveys since 1981 (NEFSC 2004).

NEFSC clam surveys are organized around NEFSC shellfish strata which are combined to define stock assessment areas (Figure A1). Most of ocean quahog landings originate from areas covered by the survey. The survey did not cover GBK and SVA completely in all years and strata in other areas are occasionally missed (Table A8). Strata not sampled during a particular survey are filled by borrowing data from the same stratum in the previous and/or next survey, if data are available (NEFSC 2004). Survey data are never borrowed from surveys behind the previous or beyond the next survey.

Surveys follow a stratified random sampling design, allocating a pre-determined number of tows to each stratum. Stations used to measure trends in ocean quahog abundance are either random or nearly random. A few nearly random tows were added in previous surveys to ensure that important areas were sampled. Other non-random stations are occupied for a variety of purposes but not used to estimate relative trends in ocean quahog abundance.

A standard tow is nominally $0.125 \mathrm{~nm}(\mathrm{~m})$ in length (i.e. 5 minutes long at a speed of 1.5 knots). However, sensor data indicate that the actual tow lengths are greater (Weinberg et al. 2002 and see below).

Occasionally, randomly selected stations are found too rocky or rough to tow. In these cases during surveys since 1999, a search for fishable ground is made in the vicinity $(0.5 \mathrm{~nm})$ of the original station (NEFSC 2004). If no fishable ground is located, the station is given a special code ( $\mathrm{SHG}=151$ ) and the research vessel moves on to the next
station. The proportion of random stations that cannot be fished is used to estimate the proportion of habitat in a stratum or region that is suitable habitat for ocean quahog, which is used in calculation of ocean quahog biomass from survey data (see below).

Following most survey tows, all ocean quahog and Atlantic surfclam in the survey dredge are counted and shell length is measured to the nearest mm. A few very large catches may be subsampled. Mean meat weight ( kg ) per tow is computed with shell length-meat weight (SLMW) equations from NEFSC (2004).

SLMW relationships used with survey data to track trends in survey meat weight per tow are region-specific. SLMW relationships used for survey data in this analysis (Table A9) were the same as in the last assessment (NEFSC 2004). They were derived by averaging SLMW curves from the 1997 and 2002 surveys, which were based on fresh tissue minus shell weighed at sea. Samples from earlier surveys were from frozen meats.

NEFSC clam survey require a great deal of additional adjustments after extraction from the database and before they are used in trend or swept-area biomass calculations (e.g. adjustments for tow distance and fishery or survey selectivity). Clam survey database parameters that would be required to replicate each analysis are listed in Table A10).

## Survey gear selectivity

NEFSC (2004) estimated selectivity curves for ocean quahog in the NEFSC clam dredge based on catches by a commercial dredge with a small mesh liner during 2003 and survey catches in the same area during 2002. The selectivity curve $s_{L}=1 /\left(1+e^{8.122-0.119 L}\right)$ indicates that $50 \%$ of ocean quahog are fully available to the NEFSC clam dredge at about 68 mm SL, which can be compared to 73 mm for commercial dredges (Figure A27). The survey dredge tends to take smaller ocean quahogs than commercial dredges because of the relatively small 2 in liner in the survey dredge. Based on sizes retained by the survey dredge (NEFSC 2004), the survey dredge selectivity curve is reliable for ocean quahog $\geq 50 \mathrm{~mm}$ SL.

## Survey, stock and fishable abundance and biomass

Catch and length composition data for ocean quahog $\geq 50 \mathrm{~mm}$ SL from the NEFSC clam survey were used to estimate abundance and length composition for the stock as a whole. In particular, $N_{L}=n_{L} / s_{L}$ where $N_{L}$ is mean stock numbers or biomass per tow at length $L, n_{L}$ is survey catch and $s_{L}$ is survey selectivity.

Abundance and length composition for the fishable stock (i.e. available to the fishery) were estimated by correcting stock estimates for fishery selectivity. In particular, $\eta_{L}=\phi_{L} N_{L}$ where $\eta_{L}$ is fishable abundance and $\phi_{L}$ is fishery selectivity. Fishable abundance can be estimated directly from survey data for ocean quahog $\geq 50 \mathrm{~mm}$ SL using $\eta_{L}=n_{L} \phi_{L} / s_{L}$ (Figure A27).

Calculation of stock abundance and biomass occasionally produces very large estimates for small sizes where selectivity is small (near zero) when ratios $n_{L} / s_{L}$ become very large. Calculation of fishable abundance and biomass from survey data does not suffer from this problem because the adjustment of small sizes is relatively modest (Figure A27).

## Spawning stock biomass

Trends in spawning stock biomass for ocean quahog were estimated based on survey data by applying a maturity at length relationship for ocean quahog from Thorarinsdottir and Jacobson (2005) to survey length composition for the stock as a whole (i.e. after correction for survey dredge selectivity). In particular, $S_{L}=m_{L} N_{L} w_{L}$ where $S_{L}$ and $w_{L}$ are spawning biomass and mean body weight (from a length-weight relationship) See Section 6 for more information about the maturity curve.

## 2005 Survey

The 2005 NEFSC clam survey was carried out during late May to early June. There were three legs (stations 1-182 during May 24-June 2, stations 183-250 during June 9-June 17, and stations 251-433 during June 22-29). Four hundred and thirty three stations were occupied. Sensor data used to monitor dredge performance were collected at 399 stations. Two hundred and eighty random and nearly random stations were used to calculate trends in ocean quahog abundance. The set of strata covered during the 2005 survey was similar to strata covered during previous surveys except that no stations were occupied in the most northern (GBK) and southern (SVA) stock assessment regions (Table A8).

Trends in survey, stock and fishable mean kg per tow were calculated for ocean quahog $\geq 50 \mathrm{~mm}$ SL in each region (Table A7 and Figure A20). Smaller ocean quahog taken in surveys were not included because catches of small individuals is very low and because selectivity curves used to calculate stock and fishable abundance are not valid below 50 mm SL. Trends in survey, stock and fishable numbers and weight per tow for the same region were generally similar.

The precision of survey trend data from the 2005 survey was typical but results for DMV were relatively imprecise with high coefficients of variation (CV) due to a single large tow in stratum 15 (Table A7). CVs for trend data from surveys during 19822005 averaged about $0.3,0.2,0.2$ and 0.3 in the DMV, NJ, LI and SNE regions.

As described below, trends in NEFSC clam survey data are complicated by changes in survey dredge efficiency. ${ }^{3}$ In particular, survey data for 1994 were judged not comparable to survey data from other surveys because power to the dredge used to run the submersible pump during 1994 was set to 480 instead of 460 volts and dredge efficiency was artificially increased during 1994.

## Dredge performance

After the 1994 survey, sensors were used to monitor depth (ambient pressure), differential pressure, voltage, hertz and amperage of power supplied to the dredge, x -tilt (side to side), y-tilt (front to back) and ambient temperature during survey fishing operations. At the same time, sensors on board the ship monitor electrical frequency, GPS position, vessel bearing and vessel speed. All sensor data are recorded at 1 second intervals.

Good tows have characteristic sensor data patterns that are easy to interpret (Figure A28). Anomalous patterns indicate potential problems with the tow or sensors.

[^2]Differential pressure, amperage and y-tilt are particularly important. Differential pressure is one of the factors affecting the flow of water through the jets in front of the dredge blade. Amperage measures the work done by the pump in moving water through the jets. If water is blocked at the entrance to the pump, then both amperage and differential pressure will be low. If water is blocked downstream of the pump, then amperage will be low and differential pressure will be high. Y-tilt can be used to determine if the dredge is on the bottom with the blade in the sediment.

Differential pressure data collected during the 2005 clam survey show a spike early in the first leg (Figure A29) coinciding with a drop in amperage that was due to a faulty screen on the input to the dredge system that allowed rocks to enter and fill the manifold, which is downstream from the pump. The screen was repaired, rocks removed and the affected stations were reoccupied.

Differential pressure appeared to jump from about 40 to about 50 psi beginning at approximately station 221 during the second leg of the 2005 NEFSC clam survey at the same time that amperage might have declined (Figure A29). The timing of the change coincided with malfunction and repair of electrical equipment on the ship that supplies power to the pump on the dredge.

The apparent jump in differential pressure during the second leg of the 2005 survey triggered a careful analysis of survey sensor data and dredge performance (Appendix A1). The apparent problem with differential pressure was determined to stem from sensor drift. In particular, differential pressure measurements before and after the pump was turned on were generally biased high after station 220 to the same extent at each station. The difference between ambient measurements at the surface and during fishing for each tow (another way to estimate differential pressure) was usually about 40 psi and approximately equal to differential pressures measured in the normal manner during the first leg. The alternate estimates of differential pressure did show a slight but steady decline in differential pressure during the survey presumably due to wear on the pump (Appendix A1).

In the course of investigating the problems with differential pressure, a number of stations with poor dredge performance were identified based on problems with differential pressure, amperage, vessel speed, and y-tilt (Appendix A2). Four of the problematic stations (218, 225, 262 and 282) were in areas of typical ocean quahog habitat and would not have been omitted following standard survey procedures. ${ }^{4}$ Stations 218, 225, 262 and 282 from omitted from further analysis. Similar problems may have occurred in earlier surveys but can not be detected or removed for lack of sensor data. Analysis of sensor data from the 2002 survey will be analyzed to determine if similar problems occurred during 2002.

## Tow distance

Tow distance was estimated for each station in the 2005 NMFS clam survey based on speed over ground (SOG) data from the ship's GPS and dredge inclinometer data from the SSP. SOG was assumed to be the same for the ship and dredge.

Following NEFSC (2003), the dredge was assumed to be fishing effectively whenever the smoothed y-tilt was $\leq 5.16^{\circ}$ (see below). Based on the geometry of the

[^3]dredge, the blade penetrates the sediments to a depth of 1 inch when the $y$-tilt is $5.16^{\circ}$. Penetration increases as the y-tilt decreases.

Tow distance calculations for the 2005 survey were the same as in NEFSC (2003) except that missing values were interpolated as described below. The first step was to replace missing SOG and inclinometer data for each station with interpolated values from a cubic spline. The second step was to smooth the original plus interpolated SOG and inclinometer data with a centered seven point moving average (e.g. the smoothed value for $t=3$ was the average for $t=1$ to 7 ). ${ }^{5}$ The final step was to compute the effective tow distance for each tow $d_{j}$ using:

$$
d=\frac{\sum_{t} \delta_{t} s_{t}}{3600}
$$

where $t$ was a one-second interval, $\delta_{t}$ was a dummy variable equal to one when the dredge was fishing effectively (smooth y-tilt $\leq 5.16^{\circ}$ ) and zero otherwise, $s_{t}$ was SOG (knots) and 3600 is the number of seconds per hour. Tow distances calculated in this manner and used in this assessment for surveys during 1997-2002 (see below) were the same as in NEFSC (2003). The median tow distance for 2005 was consistent with median tow distances from the 1999 and 2002 surveys (see below). As pointed out in NEFSC (2003), the median tow distance for 1997 was $0.4-0.7 \mathrm{~nm}$ larger than median tow distances from other surveys because a slower winch was used to deploy the survey dredge (Table C7 in NEFSC 2003).

| Year | Median <br> Tow <br> Distance <br> $(\mathrm{NM})$ |
| :---: | :---: |
| 1997 | 0.26 |
| 1999 | 0.22 |
| 2002 | 0.19 |
| 2005 | 0.21 |

Tests showed that the new interpolation procedure had a negligible effect on tow distance estimates for the 2005 survey because missing values were rare. Similar results would likely be obtained for the 2002 survey, which also used the survey sensor package. Effects of interpolation on tow distance estimates were not investigated for 1997 and 1999 surveys but may be larger because sensor data from the 1997 and 1999 surveys were collected using less precise sensors with recording intervals that were sometimes longer than one second. This is a topic for future research.

[^4]
## Tow distance vs. depth

Tow distance is a key variable in estimating swept area biomass (see below). Weinberg et al. (2002) show that tow distance increases with depth for the NEFSC clam survey dredge when the dredge is deployed as in actual clam surveys. Regression analysis was used to determine if depth measurements could be used to infer tow length at survey stations when sensor data are not available. Based on graphical relationships (Figure A30), linear regression models were used, e.g. $d_{j}=\alpha+\beta D_{j}$ where $d_{j}$ was tow distance in nm (calculated from sensor data assuming the dredge was fishing when the smoothed $y$-tilt was $\leq 5.16^{\circ}$ ), and $D_{j}$ was average depth of the tow in meters as measured from the ship. Data used in the analysis were for random survey tows only (tows with database code RANDLIKE $>0$ ). Tows with sensor-based tow distances $<0.125 \mathrm{~nm}$ were omitted from the analysis because they were likely aborted or test tows.

A stepwise regression procedure was used to select the best model from a range of models based on the AIC statistic. In the Splus programming language, the simplest model considered was:

```
Smallest <- lm(d~1)
```

where " $\sim 1$ " indicates that the model consists of the mean for the entire data set. The most complicated model was:

```
Biggest<- lm(d ~ CRUISE + D / CRUISE)
```

which is equivalent to a separate regression models relating tow distance and depth in each of the 1997, 1999, 2002 and 2005 surveys (Figure A30).

The most complicated model was selected as the best model by the stepwise procedure based on AIC. The best model was statistically significant ( $p<0.0001$ ) and all parameters were statistically significant at the $p=0.1$ level (see below).

|  | Estimate | Standard Error | t-test | $\begin{gathered} \mathrm{p}- \\ \text { value } \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: |
| Survey effects (intercept parameters) |  |  |  |  |
| Intercept | 0.182 | 0.002 | 91.0098 | 0 |
| 1997 | -0.02 | 0.0028 | -7.2647 | 0 |
| 2002 | -0.0093 | 0.0015 | -6.1114 | 0 |
| 2005 | -0.0046 | 0.0013 | -3.6898 | 0.0002 |
| Depth effects (slope parameters) |  |  |  |  |
| Depth | 0.0009 | 0 | 20.0054 | 0 |
| 1997 | 0.0001 | 0.0001 | 1.8697 | 0.0618 |
| 2002 | -0.0001 | 0 | -2.7522 | 0.006 |
| 2005 | 0.0001 | 0 | 2.5433 | 0.0111 |
| Residual standard error: 0.02809 on 1179 degrees of freedom |  |  |  |  |
| Multiple R-Squared: 0.4634 |  |  |  |  |
| F-statistic: 145.4 on 7 and 1179 degrees of freedom, the p-value |  |  |  |  |

Residual plots indicated reasonably good model fit although distributions of residuals were skewed either to the left or right for some surveys. Based on the regression analysis, tow distance increases by an average of about $0.0009 \mathrm{~nm}(1.7 \mathrm{~m})$ per meter of depth.

Results show that missing tow distance data for NEFSC clam survey stations could be replaced with estimates based on depth from a survey-specific linear model. Unfortunately, differences among surveys were large enough to be important in estimating tow distance and should not be ignored. It does not appear that a single or average depth-tow distance relationship could be used to estimate tow distance for previous surveys with no sensor data for measurement of tow distances.

## Commercial and survey dredge efficiency

Dredge efficiency is defined for this assessment as the probability of capture (i.e. of being handled on deck) for an ocean quahog that is in the path of the dredge and large enough (e.g. $83+$ SL in a survey dredge or $90+\mathrm{mm}$ SL in commercial dredge, see below) to be fully selected by the dredge used in the experiment. Dredge efficiency for smaller ocean quahog is the product of the overall dredge efficiency for fully selected sizes and the selectivity for the particular size.

Collaborative "depletion" experiments were conducted following NEFSC clam surveys in 1997-2005 to estimate commercial and survey dredge efficiency (Figure A31). Commercial dredge efficiency estimates are of considerable interest but are most important in estimating efficiency of the survey dredge deployed from the $R / V$ Delaware II during NEFSC clam surveys. Commercial dredges are inherently more efficient than the survey dredge (due to higher pressure water jets) and tend to select larger ocean quahog. In this assessment differences in the size of catches are accommodated by restricting analysis to sizes large enough to be fully selected by survey and commercial gear used in the experiment (see below).

Considerable progress has been made since the last assessment, but efficiency estimates for ocean quahog are still more uncertain and difficult than for Atlantic surfclam (NEFSC 2003). Dredge efficiency is harder to estimate for ocean quahog because they are found in deeper water (which makes dredge position data less reliable) and because they burrow deeper into sediments (and are probably sampled less efficiently) to a degree that depends on environmental conditions.

All depletion experiments for ocean quahog involve fishing repeatedly in the same area, usually until a significant decline in catch per tow is noted. Sensors and GPS equipment are have been used since 1999 to track the performance of the dredge and position of the vessel during each tow (vessel position is used as a proxy for dredge position). Experiments during 1997-1998 used loran positions noted by hand. The accuracy of position information is an important consideration (see below). Catch and position data are used in a statistical analysis (see below) to estimate the efficiency of the dredge used in the experiment.

In a "Delaware II" depletion experiment, the $R / V$ Delaware $I I$ and NEFSC survey dredge are used to make depletion tows. The efficiency of the survey dredge is estimated from the depletion tow data directly using the "Patch" model (Rago et al., in press and see below). One Delaware II depletion experiment has been completed for ocean quahog (experiment OQ1999-01 DE2 in Table A11).

In "commercial" depletion experiments, a commercial vessel and dredge are used for depletion tows. The efficiency of the commercial dredge is estimated directly using the Patch model.

Commercial depletion experiments can be used to estimate survey dredge efficiency also if the $R / V$ Delaware $I I$ conducts setup tows prior to the commercial depletion experiment in the same or immediately adjacent area (see below). About five
non-overlapping setup tows are typically carried out. Sixteen commercial depletion experiments have been completed by commercial vessels of which thirteen included setup tows (Table A11 and Figure A31).

## Patch model

The Patch model was used exclusively to estimate depletion experiment data in this assessment. It has become a standard approach used in NEFSC stock assessment work for a variety of shell- and sedentary demurral finfish including Atlantic sea scallops NEFSC (2004b), ocean quahog (NEFSC 2004), Atlantic surfclam (NEFSC 2003) and goosefish (NEFSC 2005). Other estimators used for ocean quahog in previous assessments were either ad-hoc or based on estimators involving assumptions that are tenuous for ocean quahog (e.g. complete mixing after each depletion tow). Now that a sufficient number of depletion experiments have been completed, it is possible to use Patch model estimates exclusively.

The Patch model was used to estimate three parameters for each depletion experiment (initial ocean quahog density, dredge efficiency, and a measure of dispersion) by maximizing the likelihood of the observed catches under the assumptions that the dredge path is known and that the catches are sampled from a negative binomial distribution. The key point is that it is not necessary to assume ocean quahogs mix randomly (except in relatively small cells) after every depletion tow. Ideally, GPS is used to monitor the position of the ship (a proxy for position of the dredge) at one second intervals during each tow (see below). In computing the likelihood for the catch in each tow, the model considers the number of times each grid sampled during the tow had been swept by the dredge in previous tows. Likelihood profiles are used to compute confidence intervals for all model estimates and residual plots (observed - predicted catches) can be used to judge model fit.

## Revised estimators for survey dredge efficiency based on setup tows

Efficiency of the NEFSC clam survey dredge is estimated from commercial depletion experiment results by relating densities measured by the Delaware II in setup tows to initial density estimated from a commercial depletion experiment by the Patch model (Rago et al., in press). In particular:

$$
e=\frac{d}{D}
$$

where $e$ is estimated efficiency of the NEFSC survey dredge, $d$ is density (number $\mathrm{ft}^{-2}$ ) estimated from setup tows by survey dredge, and $D$ is density estimated by the Patch model. In this context, $d$ is understood to measure survey catch rates while $D$ is understood to measure the actual density of quahog on the bottom of the ocean within the boundaries of the depletion experiment site. Previous ocean quahog assessments (NEFSC 1998; NEFSC 2000; NEFSC 2004) used a different formula that is incorrect:

$$
e=\frac{d}{D} E
$$

where $E \leq 1$ is efficiency of the commercial dredge as estimated by the Patch model (note that this formula is correct if $E=1$, which is appropriate if $D$ is absolute initial density). For this assessment, all depletion experiments were reanalyzed using the correct formula and other changes described below. All other things being equal, the corrected formula increases research survey dredge efficiency estimates (and decreases swept-area biomass estimates) because $E<1$ so that $d / D \geq(d / D) E$.

## Revised assumptions about dredge selectivity

It is important that data used in the Patch model include only length groups that are (or are nearly) fully selected. For survey efficiency estimates from setup tows and commercial depletion experiments, size groups fully selected by both the survey and commercial gear should be used. This restriction is important for two reasons. Firstly, the estimator $e=d / D$ requires that $d$ and $D$ be for the same fully recruited size groups. Secondly, Patch model estimates of $E$ will be biased low if small size groups (with lower selectivity) are included.

Previous assessments (NEFSC 1998; NEFSC 2000; NEFSC 2004) assumed that Patch model estimates were valid as long as the survey dredge and commercial dredge used in the depletion experiment had "similar selectivity" for size groups included in the analysis. Commercial sampling equipment (dredge and shaker table) used in depletion experiments was usually adjusted prior to sampling so that the catch rates for small ocean quahog increased and the modified commercial and survey length composition data were made more similar. Decisions about which size groups to include in an analysis were made in previous assessment after experiments were completed based on length composition data from setup and depletion tows. In practice, length groups actually used in estimation varied from experiment to experiment (e.g. 71+ mm for the OQ2000-1, 76+ mm for the OQ2000-2, and all size groups for the OQ2002-1 to OQ2002-4 depletion studies). In experiments during 1997-1999 that used only one type of gear, all size groups were used.

## Revised depletion study catch data

For this assessment, all depletion experiments during 1997-2005 were analyzed or reanalyzed using depletion experiment catch data (numbers of ocean quahog per tow) for size groups that were at least $85 \%$ selected by all gear used in the experiment. In particular, catches for commercial depletion experiments and setup tows were for ocean quahog $90+\mathrm{mm}$ SL and catches for Delaware II depletion experiments were for ocean quahog $83+\mathrm{mm}$ SL. Based on selectivity curves (Figure A27), $87 \%$ and $93 \%$ of ocean quahog are selected by commercial and survey dredges at 90 mm SL . As mentioned above, commercial equipment was usually adjusted prior to use in depletion experiments so that commercial selectivity at 90 mm SL was likely higher than $90 \%$. Data analyzed from Delaware II depletion experiments were for ocean quahog $83+\mathrm{mm}$ SL because survey dredge selectivity is $85 \%$ at that size.

The decision to use the size at $85 \%$ selectivity as the cutoff was pragmatic. A higher selectivity cutoff level might be preferred on mathematical grounds but the variability of catch data decreased when fewer sizes were included. For example, data from the OQ2000-1 depletion experiment were used to estimate commercial dredge efficiency but could not be used to estimate survey dredge efficiency because relatively few ocean quahog $90+\mathrm{mm}$ were taken in setup tows. In OQ2000-1 setup tows, large ocean quahog comprised only $6 \%$ of the setup catch on average.

Calculation of catch of ocean quahog larger than a specified size (e.g. $90+\mathrm{mm}$ ) requires information about the catch in bushels in each tow, the number of clams per bushel ("bushel counts"), and the proportion of clams larger than $90+\mathrm{mm}$ (from length measurements. Ideally:

$$
n_{t, 90^{+}}=B_{t} n_{t} p_{t, 90^{+}}
$$

where $B_{t}$ is catch in bushels for tow $t, n_{t}$ is the number of ocean quahogs in a sample bushel and $p_{t, 90^{+}}$, is the proportion of the length sample that was at least 90 mm SL.

Bushel counts and length data measurements were not collected from every tow during depletion experiments. During most experiments, one bushel of ocean quahog was counted and one bushel was measured at intervals of 3-5 tows, and occasionally at longer intervals (Table A11). In some cases, the number of broken clams was recorded so that the number measured plus broken provided additional information about numbers per bushel.

A convention was developed to objectively calculate the number of ocean quahog above a specific size for tows without bushel counts or length data. For example, if an experiment consisted of 10 tows with samples taken on tows 2,6 and 9 , then $n_{2}$ was used for tows 1-2. The average of $n_{2}$ and $n_{6}$ was used for tows $3-5$. The average of $n_{6}$ and $n_{9}$ was used for tows 7-8. Finally, $n_{9}$ was used for tows $9-10$. In previous assessments, a variety of conventions (including the one used in this assessment) was employed for different tows and different depletion experiments.

In theory, bushel counts should increase and proportions of large individuals in catches should decrease as a depletion study is carried out and large ocean quahog are preferentially removed from the study site. This pattern was not, however, consistently observed.

Length and bushel count data from depletion and setup tows appears more important than recognized in previous assessments. More detailed length data (e.g. 1 bushel per tow) should therefore be collected during future depletion experiments. Lengths and bushel counts were likely under-sampled in depletion experiments to date (Table A11)

## Accuracy and precision of position data

Cell sizes used in Patch model runs for this assessment are 20-25 ft (Table A11). Previous assessments used 10-25 ft. Position data used in the Patch model for ocean quahog depletion experiments should be recorded at (or interpolated to) intervals $\leq$ 0.00001 degrees to avoid missing cells (see below). Position data recorded to 0.0001 degrees, for example, are too coarse, because the wrong cell would be assigned frequently due to imprecision in position measurements. This recommendation assumes that vessel position is an accurate proxy for dredge position. The accuracy of GPS data as information about dredge position likely deteriorates with depth. Problems with position information may be exaggerated to some extent for ocean quahog, which are found in relatively deep water. Potential effects of inaccurate position data should be evaluated by simulation analysis. Position data were smoothed prior to use in this assessment to account for imprecise position data from some depletion experiments (see below).

| Distance in feet for a change in   <br> latitude or longitude at $40^{\circ} \mathrm{N}$.   <br> Distance in Feet   <br> Degrees   |  |  |
| :---: | :---: | :---: |
| 1 | Latitude | Longitude |
| 0.1 | 364,560 | 279,269 |
| 0.01 | 3,646 | 27,927 |
| 0.001 | 365 | 2,793 |
| 0.0001 | 36.5 | 279 |
| 0.00001 | 4 | 27.9 |
| 0.000001 | 0.4 | 0.3 |

Position data used in the Patch model should be recorded at (or interpolated to) intervals $\leq 4$ second intervals to avoid skipping cells too frequently between position observations. The target tow speed for the $R / V$ Delaware II during depletion tows is 1.5 knots or $2.5 \mathrm{ft} \mathrm{sec}^{-1}$. Commercial vessels probably average about 2 knots or $3.4 \mathrm{ft} \mathrm{sec}^{-1}$ during commercial operations tows (D. Wallace, Wallace and Associates, pers. comm.) and about 3 knots or $5 \mathrm{ft} \mathrm{sec}^{-1}$ during depletion tows (E. Powell, Rutgers University, pers. comm..). Thus, sampling (or interplation) at intervals of 1-3 seconds is recommended because the $R / V$ Delaware $I I$ crosses a 20 ft cell in 8 seconds and a commercial vessel crosses a 20 ft cell in 4 seconds (see below). Smaller cell sizes require more frequent sampling or interpolation. Position data were interpolated in this assessment to account for relatively long sampling intervals in some depletion experiments (se below).

Time in seconds required to cross Patch model cells $15-25 \mathrm{ft}$ wide at vessel speeds of 1.5 and 2 knots.

Vessel speed
(knots)

| Feet | 1.5 | 3 |
| :---: | :---: | :---: |
| 15 | 5.9 | 2.9 |
| 20 | 7.9 | 3.9 |
| 25 | 9.9 | 4.9 |

## Smoothed position data for depletion experiments

Position data for 1997-2005 depletion experiments were from original Loran or GPS records. Start and stop times for GPS data were the same as used in the last assessment).

Position data from depletion studies during 2000-2005 were recorded to $10^{-6}$ degrees at one second intervals based on differential GPS or the equivalent (Table A11). However, position data from the 1999 Delaware II depletion study from GPS were recorded to only 0.0001 degrees and position data from loran readings in depletion studies during 1998-1998 were recorded to an accuracy of about 0.0001 degrees.

To avoid problems with erratic "stair pattern" tow tracks from coarse position data, original position data from all depletion experiments were smoothed prior to further analysis (Appendix A3). The smoother was a cubic spline when the number of observations $n \geq 15$, a quadratic polynomial when the number of observations was $5 \leq n$ $<15$ or a straight line when $2 \leq n<5$. Smooth lines were fit using latitude or longitude as the dependent variable and order of collection (a crude measure of time) as the
independent variable. Smoothed values were used in subsequent calculations, instead of the original data. Decisions about smoothing were ad-hoc but consistently applied and seemed to result in plausible tow paths for further analysis (Appendix A3). Fortunately, survey dredge efficiency estimates were from recent depletion studies with generally accurate position data sampled at relatively frequent intervals. With accurate data at frequent intervals, smoothing had very little effect of tow path data.

No position data were available for 2 out of 60 tows in the 1999 Delaware II depletion experiment. Crude estimates of the start and stop locations for these tows from previous assessments from a previous assessment were used instead.

Before analysis in the patch model, original or smoothed position data were interpolated along straight lines to a distance of 5 ft ( $\sim 1-2$ second intervals) to ensure that all cells that were crossed by the dredge would be recorded as "hits" in the Patch model program. This was apparently not done for all depletion experiments in previous assessments and it is possible that not all hits were included in previous estimates. In future assessments, interpolation should be based on the model (e.g. cubic spline) used to smooth the original position data, rather than by linear interpolation.

## Assumptions about cell size

All depletion studies were analyzed or reanalyzed using consistent and updated assumptions about cell size and indirect effects, which are closely related. Rago et al. (in press) suggested that the cell size be set at twice the width of the dredge used in the depletion experiment. They point out that decisions about cell size reflect a compromise between the accuracy of position data and the tenability of the assumption that animals mix within cells after each tow. Dredges used in depletion experiments were mostly $\geq 10$ ft wide with the exception of the commercial dredge in the OQ1997-1 commercial depletion experiment and the 5 ft dredge used in the OQ1999-1 (DE-2) Delaware II depletion experiment (Table A11).

In this assessment, the cell size in Patch model analyses was set at twice the dredge width or 20 ft , whichever was larger. This approach basically follows the advice in Rago et al. (in press) for all experiments during 2000-2005 while assuming that positional accuracy (particularly for experiments during 1997-2005) was never better than 20 ft . Patch model estimates for ocean quahog were moderately sensitive to the assumed cell size (Figure A32). In particular, efficiency estimates tend to increase and density estimates tend to decrease as the cell size assumed in the Patch model increase.

## Indirect effects

The "gamma" parameter in the Patch model is used to measure indirect effects (ocean quahog lost from the study site without being counted on deck). In this assessment gamma was fixed at the ratio of the dredge width and cell width ( $\gamma=0.5$ ) so that no indirect effects were assumed to occur. The gamma parameter is theoretically estimable but estimation has proven difficult in practice because the estimate for gamma is correlated with other estimates in the model and dependent on assumptions about cell size (Rago et al., in press). The previous assessment assumed indirect effects ( $\gamma=0.75$ ) in depletion experiments during 1997-2000 and no indirect effects ( $\gamma=0.5$ ) in depletion experiments during 2002. As shown in Rago et al. (in press) efficiency and density estimates from the Patch model tend to decrease as the assumed level of $\gamma$ increases.

## Sensitivity to initial parameter estimates

Patch model estimates were not sensitive to the starting values for parameter estimates. After an initial Patch model run for each experiment was completed, the model was rerun several times to determine if results were sensitive to starting parameter values. In particular, the model was rerun at least four times with HD/LE, LD/HE, HD/HE and LD/LE where HD, LD, HE and LE stand for higher and lower starting density values and higher and lower starting efficiency values. In general, higher starting values were 2-3 times higher than the initial estimate and lower starting values were onehalf to one-third of the initial estimate. The estimate providing the best fit to the catch data (smallest negative log-likelihood) was the best estimate.

## 2005 Depletion experiments

In 2005, five new commercial depletion experiments were completed with five setup tows and 17-21 depletion tows per site (Figures A33-A37). No Delaware II depletion studies were carried out for ocean quahog during 2005. Details about depletion studies during 2002 are described in NEFSC 2004, experiments during 1998 and 1999 are described in NEFSC (2000) and experiments during 1997-1998 are described in NEFSC (1998).

Survey sensor package equipment (with the exception of GPS and a backup depth sensor) did not function during ocean quahog depletion tows by the commercial vessel during 2005 due to battery failure, with the exception of initial tows at the OQ2005-6 depletion site.

The survey data that are available for 2005 commercial depletion tows (Figure A38) indicate that the commercial dredge was not always horizontal and hard on bottom at the OQ2005-06 depletion site due to the combined effect of low scope and choppy seas. The estimated efficiency for OQ2005-06 may have been reduced by these factors. The OQ2005-06 site was in the deepest water ( 65 m , Table A11) and conducted in choppy seas. The commercial dredge was deployed at this site with lower scope because the hose used to supply water to the dredge was relatively short. The sea was calmer and shallower at towing scope was greater at other relatively shallow depletion sites for ocean quahog during 2005. Although no sensor data are available, it is likely that the commercial dredge towed well at the other 2005 ocean quahog depletion sites.

As in previous years, commercial sampling equipment (dredge and shaker table) used in 2005 was adjusted to increase catch of relatively small ocean quahog. However, length composition data for the setup and depletion tows at each site during 2005 indicate that the selectivity of the two dredges differed (Figure A39). Confidence intervals and residual plots (Appendix A4) indicate that efficiency and density estimates from experiments during 2005 were reasonably precise.

## Depletion study results

For this assessment, all depletion experiments for ocean quahog during 19972005 were analyzed or reanalyzed using the Patch model based on revised data, assumptions and procedures described above. All of the underlying data, with the exception of the raw GPS position information collected during depletion studies during 1999-2005, were reevaluated. Residuals and confidence intervals for Patch model parameters are shown for each depletion experiment in Appendix A4. Estimates and model fit are summarized in Tables A11-A12. To build a bridge between new and old
results, differences between efficiency and density estimates in this and previous assessments are summarized in Table A13.

Estimates from commercial depletion experiments during 1997-1998 and the Delaware II depletion experiment during 1999 are probably less reliable than estimates from experiments during 2000-2005. Position data were relatively imprecise in depletion experiments prior to 2000 (Table A11). Goodness of fit to depletion catch data was poor for the OQ1998-1 and OQ1999-1 (DE-2) experiments (Appendix A4). Average annual commercial efficiency estimates from experiments during $1997(E=0.592)$ and 1998 ( $E=0.860$ ) were outside the range of average annual estimates for later years (i.e. $E=0.615,0.588$ and 0.559 during 2000-2005). The OQ1999-1 (DE-2) survey dredge efficiency estimate was anomalously high and the corresponding density estimate was anomalously low, relative to estimates from later commercial depletions with setup tows.

There were no clear relationships between dredge efficiency and density or depth (Figure A40). There is, however, a suggestion of a negative correlation between survey dredge efficiency and sediment size.

Revised Patch model estimates of commercial and survey dredge efficiency from historical depletion experiments were smaller than previous estimates with a few exceptions (Table A13). Revised density estimates were always smaller but the revised and previous density estimates are not comparable because they are for different size groups.

The seventeen commercial dredge efficiency estimates indicate that efficiency of commercial dredges is highly variable with $E=0.15$ to 1.00 (Tables A11-A12 and Figure A42). The average and median of estimates of commercial efficiency were 0.60 ( $\mathrm{CV}=24 \%$ ) and $0.66(\mathrm{CV}=14 \%)$.

Twelve survey dredge efficiency estimates were available, eleven from commercial depletion experiments with setup tows and one from a depletion study by the $R / V$ Delaware II (Tables A11-A12). Survey dredge efficiency estimates were also variable ( $e=0.098$ to 0.990 , Figure A43). Omitting the estimate from the OQ1999-1 (DE-2) experiment, which was anomalously high, survey dredge efficiency estimates ranged 0.098-0.297. The average and median of estimates of survey efficiency were $0.248(\mathrm{CV}=29 \%)$ and $0.165(\mathrm{CV}=18 \%)$. The ratio of median commercial efficiency and median survey dredge efficiency indicates that the NEFSC survey dredge is about onequarter as efficient as commercial dredges (Table A12). Survey dredge efficiency estimates did not appear correlated with commercial dredge efficiency estimates (Figure A41).

Density estimates for ocean quahog 90 mm SL (Table A11-A13 and Figure A42) ranged $0.007-0.295 \mathrm{ft}^{-2}$. The smallest density estimate ( $0.007 \mathrm{ft}^{-2}$ ) was from the OQ19991 (DE-2) survey depletion experiment, which gave an anomalously small survey dredge efficiency estimate. The highest density estimates ( $0.226-0.295 \mathrm{ft}^{-2}$ ) were the OQ2002-1 and OQ2002-2 depletion experiments.

## Best survey dredge efficiency estimate

The "best" estimates for survey dredge efficiency ( $e=0.165, \mathrm{CV}=18 \%$ ), commercial dredge efficiency ( $E=0.66, \mathrm{CV}=14 \%$ ) and ocean quahog density ( $D=0.082$ ocean quahog $\mathrm{ft}^{-2}, \mathrm{CV}=13 \%$ ) were the medians of all available estimates from ocean quahog depletion experiments during 1999-2005 (Table A12). Medians were used because they are robust to anomalous estimates, such as the high estimate for survey
dredge efficiency from the OQ1999-1 (DE-2) experiment and the low estimate of commercial dredge efficiency from the OQ1997-3 experiment (Table A11).

The new best estimate of survey dredge efficiency $(e=0.165)$ is smaller than the estimates used in the last assessment NEFSC (2004) for the 1997 survey ( $e=0.346$ ) and for the 1999-2000 surveys ( $e=0.269$ ).

Ideally, efficiency estimates would be survey specific because differences in sampling efficiency are possible. However it is not possible at present to estimate dredge efficiency for each survey with sufficient precision.

## Depletion experiments-building a bridge

As described above, factors that contribute to the differences between the previous and revised estimates are:

1) Revised computer programs
2) Corrected formula for survey dredge efficiency based on setup tows.
3) Cell size assumed in the Patch model set to the larger of 20 ft or twice the dredge width (affects OQ1997-01 and OQ1999-1 DE-2 only);
4) Depletion and setup catch data for ocean quahog $90+\mathrm{mm}$ SL (affects all depletion studies during 1997-2002);
5) Revised position data (new smoothing and interpolation, affects all studies during 1997-2002);
6) No indirect effects, i.e. $\gamma=$ ratio of dredge width and cell size (affects all depletion studies during 1997-2000);

Not all changes apply to each depletion experiment.
To build a bridge between old and new results, effects on efficiency and density estimates due to individual factors for the OQ1998-1 and OQ2002-1 depletion experiments are shown in Table A14. In the OQ2002-1 experiment, estimates were most sensitive to using the correct formula, revised position data, and revised catch data while the density estimate was most sensitive to using catch data for ocean quahog $90+\mathrm{mm} \mathrm{SL}$ only. In the OQ1998-1 experiment, estimates were most sensitive to using the revised position and catch data.

## Repeat stations

Stations from previous and the current survey are repeated during each survey to help detect potential changes in sampling efficiency. Catch data for stations sampled twice during the 2005 survey and during both the 2002 and 2005 surveys were analyzed for this assessment but results are not presented here because the repeat stations were in Atlantic surfclam habitat where ocean quahog catches were very low.

### 5.2 Efficiency corrected swept area biomass

Efficiency corrected swept area biomass (ESB) estimates were for years (1997, 1999, 2002 and 2005) when NEFSC clam surveys collected sensor data for each tow. Sensor data are important because ESB calculations require accurate measurements of tow distance. Differences in ESB estimates between this assessment and NEFSC (2004) for 1997-2002 are described in detail below under the heading "Building a bridge".

ESB estimates (Table A15) for ocean quahog were calculated:

$$
B=\frac{B^{\prime}}{e}
$$

where:

$$
B^{\prime}=\frac{\bar{\chi} A^{\prime}}{a}(1+\phi) u
$$

In ESB calculations, $e$ is the best estimate of survey dredge efficiency for ocean quahogs, $\bar{\chi}$ is mean catch of fishable ocean quahog per standard tow based on sensor data ( kg tow ${ }^{-1}$, see below), $A^{\prime}$ is habitat area $\left(\mathrm{nm}^{2}\right), a=0.0008225 \mathrm{~nm}^{2}$ tow ${ }^{-1}$ is the area that would be covered by the 5 ft wide survey dredge during a standard tow of 0.15 nm , and $u=10^{-6}$ converts kilograms to thousand metric tons. $B^{\prime}$ is the minimum swept-area biomass prior to correction for survey dredge efficiency.

The term $\phi$ used in ESB calculations is new in this assessment. It is the fraction of total biomass in deep water strata off LI (strata 32 and 36), SNE (strata 40, 44, 48) and GBK (strata 56, 58, 60 and 62) that were sampled only during 1999. According to NEFSC (2000), deep water strata accounted for $0 \%, 2 \%$ and $13 \%$ of total biomass in the LI, SNE and GBK regions during 2005. Data for deep water strata sampled only during 1999 are otherwise omitted in calculations and, in particular, calculation of mean catch per tow $\bar{\chi}$. NEFSC (2004) used a slightly different approach for GBK in the last assessment which gave essentially the same results.

Habitat area for ocean quahogs in each region was estimated:

$$
A^{\prime}=A u
$$

where $\underline{u}$ is the proportion of random tows in the region not precluded by rocky or rough ground (ocean quahogs occupy smooth sandy habitats), and $A$ is the total area computed by summing GIS area estimates for each survey stratum in the region. Mean catch per standard tow $(\bar{\chi})$ is the stratified mean catch of fishable ocean quahog for individual tows after adjustment to standard tow distance based on tow distance measurements from sensor data $\left(d_{s}\right)$ :

$$
\chi_{i}=\frac{C_{i} d}{d_{s}}
$$

Only random tows were used in calculations of ESB. Tows without sensor data, with gear damage or poor pump performance were excluded from ESB calculations.

Following NEFSC (2004), and as described above, tow distance was measured for each station assuming that the dredge was fishing when the blade penetrated the sediments to a depth of at least one inch. Thus, the tow distance at each station was the sum of the distance covered while the dredge angle was $\leq 5.2^{\circ}$.

ESB estimates for the entire ocean quahog stock during 1997-2005 (Table A15) were computed using a formula that facilitated variance calculations (see below):

$$
B_{\text {total }}=\frac{B_{\text {total }}^{\prime}}{e}=\frac{\sum_{r} B_{r}^{\prime}}{e}
$$

The $80 \%$ confidence intervals for efficiency corrected total fishable biomass during 1997, 1999, 2002 and 2005 overlapped suggesting that the estimates were not significantly different (Table A15).

## Catch-ESB Mortality estimates

Fishing mortality rates were estimated directly from the ratio of catch (landings plus an assumed $5 \%$ incidental mortality allowance) and ESB data for each region and year (Table A16). Biomass levels change slowly in ocean quahog, fishing and natural mortality rates are low for ocean quahog, and the survey during June provides a good approximation to average biomass. It was advantageous to use the ratio estimator because the surveys occur in June and because it was easy to include a wide range of uncertainties in variance calculations (see below).

## Uncertainty in ESB and mortality estimates

Variance estimates for ESB and related mortality estimates were important in using and interpreting results (Tables A15 and A16). Formulas for estimating ESB and mortality for a single stock assessment region are products and ratios of constants and random variables. Random variables in calculations are typically non-zero (or at least non-negative) and can be assumed to be approximately log normal. Therefore, we estimated uncertainty in ESB and related mortality estimates using a formula for independent log normal variables in products and ratios (Deming 1960):

$$
C V\left(\frac{a b}{c}\right)=\sqrt{C V^{2}(a)+C V^{2}(b)+C V^{2}(c)}
$$

where $\ln (a b / c), \ln (a), \ln (b)$ and $\ln (c)$ are normally distributed. The accuracy of Deming's formula for ESB estimates was checked by comparison to simulated estimates (NEFSC 2002). CV's by the two methods were similar as long as variables in the calculation were $\log$ normally distributed. In addition, distributions of the simulated products and ratios were skewed to the right and appeared lognormal.

CV estimates for terms used in ESB and related estimates (Tables A15-A16 and Figures A44-A45) were from a variety of sources and were sometimes just educated guesses. The CV for best estimate of survey dredge efficiency (e) was CV=0.177 calculated by bootstrapping the median ( 15,000 bootstrap iterations) (Table A12). For lack of better information, CVs for sensor tow distances (d), area swept per standard tow (a), total area of region $(A)$, percent suitable habitat $(u)$, and catch were all assumed to be $10 \%$. The CV for area swept $(a)$ is understood to include variance due to Doppler distance measurements and variability in fishing power during the tow due, for example, to rocky or muddy ground.

## Uncertainty in estimates for combined assessment regions

ESB for combined stock assessment areas was estimated as described above. Variance calculations accommodated covariance among regional estimates due to using a single estimate of survey dredge efficiency:

$$
C V^{2}\left(B_{\text {total }}\right)=C V^{2}(e)+C V^{2}\left(B_{\text {total }}^{\prime}\right)
$$

Previous assessments used the formula:

$$
\operatorname{Var}\left(B_{\text {total }}\right)=\sum_{r} \operatorname{Var}\left(B_{r}\right)
$$

where $\operatorname{Var}(x)$ is the variance of $x$. The formula used previously was incorrect because it assumed that efficiency and biomass estiamtes for each region were independent. The new formula makes the estimated confidence intervals for ESB and fishing mortality wider.

## Building a bridge

Efficiency corrected swept-area biomass estimates in this assessment are almost double the estimates in the previous assessment (Table A19). For example, total stock biomass during 2002 was 2.1 million mt in NEFSC (2004) while the revised estimate in this assessment is 3.8 million mt. Several factors are responsible for this change in the estimates for 2002: 1) changes to spreadsheet software used in computations, 2) an error in the survey data for 2002 (but not for other years); 3) accounting for ocean quahogs on GBK that are too deep to be taken in the survey ( $13 \%$ of total stock biomass); 4) use of fishable biomass rather than $70+\mathrm{mm}$ biomass, and 5) new estimates of survey dredge efficiency. Of all the factors, the revised survey dredge efficiency (followed by the corrected survey data for 2002) was the most important factor contributing to higher ESB estimates in this assessment (Table A19).

## 5.3 "VPA" estimates

VPA estimates of biomass and fishing mortality are useful for stock assessment regions where the KLAMZ model (see below) is not applicable. Assuming no recruitment and that growth exactly balances natural mortality, ocean quahog biomass on January $1^{\text {st }}$ and annual fishing mortality rates (Figure A46-A50) can be estimated for each stock assessment region using a simple virtual population analysis or "VPA" approach (NEFSC 2004). Efficiency corrected swept-area biomass estimates for 1999, 2002 and 2005 are averaged and used to anchor the calculations. Averages for 1999-2005 are used because the estimates for individual years are less precise (Table A15).

The VPA biomass estimate for January 1, 2002 is:

$$
b_{2002}=\frac{B_{1999}+B_{2002}+B_{2005}}{3}-\frac{C_{2002}}{2}
$$

where $b_{y}$ is the VPA biomass estimate for January 1 in year $y, B_{y}$ is the efficiency corrected swept area biomass for June in year $y, C_{2002}$ is total catch weight (landings plus a $5 \%$ allowance for incidental mortality). The first ratio on the right-hand side is average efficiency corrected swept-area biomass during 1999-2005 and used as an estimate of biomass in June of 2002. Catch for 2002 is divided by two prior to subtraction because NEFSC clam surveys occur during June, when the year is half over.

Biomass estimates for years prior to 2002 were calculated:

$$
b_{y<2002}=b_{2002}+\sum_{i=y}^{2001} C_{i}
$$

Biomass estimates for years after 2002 were calculated:

$$
b_{y>2002}=b_{2002}-\sum_{i=2002}^{y} C_{i}
$$

Fishing mortality rates from VPA estimates were calculated by solving the catch equation with instantaneous rates for natural mortality and somatic growth both zero.

### 5.4 KLAMZ Model

KLAMZ (see Appendix A5 for a complete technical description) is a forward projecting stock assessment model based on the Deriso-Schnute delay-difference equation (Deriso 1980; Schnute 1985; Quinn and Deriso 1999). The delay-difference equation is an implicitly age structured population dynamics model that is mathematically identical to explicitly age-structured models if fishery selectivity is "knife-edged", somatic growth follows the von Bertalanffy equation, and natural mortality is the same for all age groups in each year. Knife-edge selectivity means that all individuals alive in the model during the same year experience the same fishing mortality rate. Natural mortality rates and growth parameters can change from year to year in the KLAMZ model but are assumed to be the same for all individuals alive during the same year. The model is implemented in AD Model Builder and Excel but only the AD Model Builder version was used in this assessment.

The main assumptions in the KLAMZ model for ocean quahog are: recruitment is constant over time, fishery selectivity is knife-edged; the natural mortality rate is low or constant, and growth in weight can be described by a von Bertalanffy growth curve. Recruitment is assumed constant (at levels always estimated to be very low) because no recruitment index is available. The assumption of constant recruitment is used for ocean quahog because no reliable recruitment index current exists, recruitment levels are apparently very low, and trends in stock dynamics are appear due primarily to fishing mortality.

KLAMZ model runs for ocean quahog that linked virgin biomass calculations with estimated biomass during 1978 were explored during the SARC review for this assessment. NEFSC (2000) used an equvilent virgin biomass approach. NEFSC (2004) compared several approaches and ultimately rejected the virgin biomass approach due to poor fit to survey data. As shown during the review for this assessment, models for ocean quahog that linked initial and virgin biomass in this assessment did not yield plausible results in some cases and fit to survey data was substantially reduced.

Recruitment to the ocean quahog fishery is not knife-edged but occurs at sizes of $51-86 \mathrm{~mm}$ SL (Figure A27). Under these circumstances, KLAMZ is an approximate model can be use to track trends in fishable (instead of total) biomass. Fishable biomass is dominated by relatively large individual ocean quahogs that are readily captured (see research recommendations).

Despite the assumption of knife-edge selectivity, KLAMZ is a relatively robust model (i.e. with little or no retrospective bias) that has been used successfully in previous assessments for ocean quahog (NEFSC 2004) and other species. It provides useful estimates of long-term biomass and fishing mortality, performs relatively well with very limited information about age and growth and when explicitly age-structured models are difficult to apply. One of the chief reasons for the utility of the KLAMZ model is
statistical simplicity. The models used for ocean quahog in this assessment, for example, estimates only 2-3 parameters.

## Model configurations

Configurations of the KLAMZ model for ocean quahog in each region were similar to the "best" configurations identified in the last assessment (NEFSC 2004) following a thorough analysis of a wide range of alternate configurations. Changes are highlighted in the descriptions below.

KLAMZ model estimates were for ocean quahog in the DMV, NJ, LI and SNE regions during 1977-2005. The model was not used for SVA because survey data for SVA are noisy and incomplete. The KLAMZ model was fit to data for GBK for sensitivity analysis. Following NEFSC (2004), the KLAMZ model was not used to make best estimates for GBK because no fishing occurs there, the survey time series is short (1986-2002) and because apparent trends in stock biomass are not clear (see "GBK at virgin biomass?" below).

Data used in KLAMZ models for ocean quahog in this assessment were: NEFSC clam survey biomass trends and associated CV's for 1982-2005; efficiency corrected swept-area biomass estimates for 1997-2005 (see below); and catch during 1977-2005 (landings plus a $5 \%$ allowance for incidental mortality). LPUE data are included in the model but only for comparative purposes (i.e. they had nil effect on model estimates).

NEFSC (2004) chose to omit LPUE data entirely but the decision was unnecessary because it is useful to compare model trends with LPUE data and because the LPUE data have no effect on model estimates. LPUE data did not affect estimates in this assessment because the likelihood component for trends in LPUE data was set to a very low level $\left(10^{-6}\right)$ and the survey scaling parameter $Q$ for LPUE was calculated using a closed form maximum likelihood estimator (i.e. $Q$ was not estimated as a formal parameter). LPUE data did not affect variances estimates because LPUE data did not affect goodness of fit to other data.

Catch data for ocean quahog were assumed accurate and not estimated in the model. NEFSC clam survey data were used to measure trends in biomass. NEFSC clam survey data for 1994 were omitted because electrical voltage supplied to the pump on the survey dredge was set to 480 v , rather than 460 v , artificially increasing dredge efficiency during the 1994 survey (NEFSC 2004). Efficiency corrected swept-area biomass estimates for 1997-2005 are used to measure the scale of recent biomass levels but are not used to measure trends. Recruitment is assumed to be constant at some low level or zero. The natural mortality rate was $M=0.02 \mathrm{y}^{-1}$, except in DMV (see below).

As described above, the KLAMZ model in this assessment estimates trends in fishable biomass. In contrast NEFSC (2004) modeled biomass of ocean quahog $70+\mathrm{mm}$ SL. Survey data used in the model are trends in mean fishable biomass while survey data used by NEFSC (2004) were trends in ocean quahog 70+ mm SL. Based on the fishery selectivity curve for ocean quahog, $50 \%$ of ocean quahog are selected by commercial dredges at about 73 mm SL. Thus, the previous and current assumptions about recruitment to the fishable stock are reasonably compatible.

Assumptions about growth are the same as in the last assessment. In particular, the growth parameters $\rho=e^{K}$ (where $K=0.0176$ is the von Bertalanffy growth parameter for weight), $J_{t}=w_{k-1} / w_{k}=0.9693$ (where $w_{j}$ is predicted weight at age $j$ ) are constant and the same for all regions (NEFSC 2004). These growth parameters mean that quahogs in the model are slow growing, and that quahog recruit to the fishery (reach 70 mm SL ) at
age $k=26$ (Figure A59). Growth patterns differ among regions (Lewis et al. 2001 and Figure A56) but ocean quahog are difficult to age and there is too little information available to use region-specific growth curves (NEFSC 2000). The growth curve used in KLAMZ models for all areas but GBK was estimated from data collected in the MidAtlantic Bight where fishing occurs. Lewis et al.'s (2001) growth curve was used for GBK sensitivity analysis runs.

An assumed level of variance in instantaneous somatic growth rates (IGR) for old recruits is used to help estimate the initial age structure of ocean quahogs in the initial years of the model (Appendix A5). For ocean quahog in each region, IGR values during 1979-1980 were estimated assuming a lognormal distribution with arithmetic mean equal to the estimated IGR for 1981 and an arithmetic CV for years 1981-2005 estimated in a preliminary run. For ocean quahog, this constraint is unimportant because estimated age structures were stable due to assumptions about recruitment and low mortality rates.

ESB data are very important in KLAMZ models for ocean quahog as a source of information about biomass scale. Trends in ESB data during 1997-2005 were ignored in modeling because the time series is short (four years) and because information about trends from the NEFSC clam survey is already provided by the clam survey biomass index for 1982-2005. To use ESB data as a measure of scale while ignoring trend (see Appendix A5), the likelihood component for trends in ESB data were set to $10^{-6}$ so that the survey scaling parameter $Q$ was calculated but the trend was ignored. Information in ESB data about biomass scale is contained in the estimated survey scaling parameter $Q$.

As described in Appendix A5, the likelihood of the survey scaling factor is calculated assuming that estimates of $Q$ are from a lognormal prior distribution:

$$
L=0.5\left[\frac{\ln (Q)-\tau}{\varphi}\right]^{2}
$$

where $L$ is the negative $\log$ likelihood, $\varphi=\sqrt{\ln (1+C V)}$ and $\tau=\ln (\bar{q})-\frac{\varphi^{2}}{2}$ is the mean of the $\log$ normal distribution. For ocean quahog ESB data, the mean of the prior $\bar{q}=$ $\ln (1)=0$ if ESB data measure stock biomass accurately and $\mathrm{CV}=0.177$ is the bootstrap coefficient of variation (standard deviation / mean) for the median survey dredge efficiency used in calculating ESB (Table A12).

## Parameters estimated

KLAMZ models for ocean quahog in this assessment estimate either two or three parameters by maximum likelihood and numerical optimization. The three parameters potentially estimated are logarithms of: 1) biomass at the beginning of 1977, 2) escapement biomass (total biomass less biomass of new recruits) at the beginning of 1978, and 3) annual recruitment biomass (which is assumed constant over time for each region). In models where recruitment estimates were very low, recruitment was fixed at an assumed value that was nearly zero $\left(1 \mathrm{~kg} \mathrm{y}^{-1}\right)$ and the other two parameters were estimated.

Fishing mortality rates are calculated solving the catch equation numerically. Survey scaling parameters were calculated using a closed form maximum likelihood estimator.

## Variance estimates

Variances for biomass and fishing mortality estimates and for model parameters can be estimated by the delta method using exact derivatives calculated by AD Model Builder libraries or by bootstrapping (Appendix A5). Estimates in this assessment were from the delta method.

## KLAMZ Results-DMV

As in the previous assessment (NEFSC 2004), estimated recruitment was near zero and hard to estimate in preliminary runs for DMV. The annual recruitment level was therefore fixed at very low value ( $1 \mathrm{~kg} \mathrm{y}^{-1}$ ) in final runs.

The KLAMZ model for ocean quahog in the DMV area (Figure A48) fit NEFSC survey and LPUE data well (LPUE data did not affect model estimates). The CV of arithmetic scale residuals ( $26 \%$ ) for NEFSC survey data was smaller than the mean CV ( $32 \%$ ) for mean $\mathrm{kg} /$ tow survey data but within the range of observed values ( $21 \%-53 \%$ ). The estimated survey scaling parameter for ESB data was $Q=0.98$ indicating that the model was able to match the observed ESB biomass levels on average during 1995-2005 using the catch data and trends in NEFSC survey data.

Biomass estimates for DMV declined steadily after 1978. Estimated fishable biomass during 2005 was $34 \%$ of the estimate for 1978 (Figure A48). During 2005, fishable biomass was $101,000 \mathrm{mt}$ (CV 18\%) and mean fishing mortality was $0.0094 \mathrm{y}^{-1}$ (CV 18\%).

## KLAMZ Results-NJ

The KLAMZ model for ocean quahog in the NJ area (Figure A49) fit NEFSC survey and LPUE data well (LPUE data did not affect model estimates). The CV of arithmetic scale residuals (32\%) for NEFSC survey data was larger than the mean (19\%) and range ( $14 \%-24 \%$ ) of CV values for mean kg /tow survey data. The estimated survey scaling parameter for ESB data was $Q=0.95$ indicating that the model was able to match the observed ESB biomass levels on average during 1995-2005 using the catch data and trends in NEFSC survey data.

Biomass estimates for NJ declined steadily after 1978. Estimated fishable biomass in NJ during 2005 was $44 \%$ of the estimate for 1978. During 2005, fishable biomass was $401,000 \mathrm{mt}$ (CV 17\%) and mean fishing mortality was $0.0017 \mathrm{y}^{-1}$ (CV 17\%).

## KLAMZ Results-LI

The KLAMZ model for ocean quahog in the LI area (Figure A50) fit NEFSC survey data well. The model fit LPUE data well (Figure A50) except during early years (1986-1993) when the fishery was becoming established and LPUE was relatively high but falling rapidly reflecting, perhaps, fishing down on the very best ocean quahog beds (LPUE data did not affect model estimates). The CV of arithmetic scale residuals (28\%) for NEFSC survey data was larger than the mean (19\%) and at the upper bound of the range ( $14 \%-28 \%$ ) of CV values for mean $\mathrm{kg} /$ tow survey data. The estimated survey scaling parameter for ESB data was $Q=1.0$ indicating that the model was able to match the observed ESB biomass levels on average during 1995-2005 using the catch data and trends in NEFSC survey data.

Biomass estimates for LI increased steadily after 1978 until 1992 when fishing mortality increased to maximum levels. Estimated fishable biomass in LI during 2005
was $94 \%$ of the estimate for 1978 and $90 \%$ of the maximum estimated biomass during 1992. During 2005, fishable biomass was $678,000 \mathrm{mt}$ (CV 18\%) and mean fishing mortality was $0.016 \mathrm{y}^{-1}$ (CV 18\%).

## KLAMZ Results-SNE

The KLAMZ model for ocean quahog in the SNE area (Figure A51) did not fit NEFSC survey data or LPUE data as well as for other areas (LPUE data did not affect model estimates). Predicted survey values from the KLAMZ model decreased slowly in all years. Trends is fishable biomass based on mean survey kg/tow and LPUE data suggest an increasing trend in biomass before 1994 and a decreasing trend afterwards. These patterns are discussed in detail below.

The CV of arithmetic scale residuals (24\%) for NEFSC survey data was smaller than the mean $29 \%$ ) but within the range ( $18 \%-47 \%$ ) of CV values for mean $\mathrm{kg} /$ tow survey data. The estimated survey scaling parameter for ESB data was $Q=0.99$ indicating that the model was able to match the observed ESB biomass levels on average during 1995-2005 using the catch data and trends in NEFSC survey data.

Biomass estimates for SNE decreased steadily after 1978 until 1996 when landings and fishing mortality increased to peak levels. After 1996, biomass decreased at a slightly faster rate. Estimated fishable biomass in SNE during 2005 was $75 \%$ of the estimate for 1978. During 2005, fishable biomass was $595,000 \mathrm{mt}$ (CV 18\%) and mean fishing mortality was $0.003 \mathrm{y}^{-1}$ (CV 18\%).

## Uncertainty about historical estimates and hypotheses about lack of fit

The apparent lack of fit to survey trend and LPUE data for SNE contributes uncertainty to historical biomass estimates but has little effect on estimates for recent years which were anchored by efficiency corrected swept area biomass data. However, future assessments should consider more complicated models that address hypotheses described below that might explain upward trends in fishable biomass prior to 1994 and decreasing trends afterwards.

It is possible that the upward trend in LPUE during 1984-1993 reflects an exploration phase during which the fishery searched for and located prime fishing grounds. However, this explanation does not apply to survey trend data.

Changes in recruitment patterns and the assumption of constant recruitment in the KLAMZ model might explain the difference between trends in KLAMZ model estimates and survey trend and LPUE data. However, survey trends in fishable biomass are not consistent with survey length and recruit trend data. In particular, survey length data (Figure A26) and survey recruit abundance data (Figure A21) do not suggest strong recruitment prior to 1994 and weak recruitment afterwards. Survey length data for 19801994 do not show a mode of small ocean quahog recruiting to fishable size while survey trend data and LPUE were increasing. Survey length data after 1994 do not show reductions in recruits while survey trend and LPUE data were decreasing. Survey recruit abundance data seem, in particular, to suggest higher recruitment after 1994.

Changes in landings and fishing mortality may explain the trends in survey trend and LPUE data. Annual landings were low ( 0 to $1,000 \mathrm{mt}$ ) during 1978-1994 while the survey trend and LPUE data were increasing. After 1994, landings increased dramatically ( 2,000 to $9,000 \mathrm{mt}$ ) during while survey trend and LPUE data were decreasing.

## KLAMZ-methods for GBK trial and sensitivity runs

For the first time, the KLAMZ model was applied to GBK on a trial basis and to conduct sensitivity analyses. The trial run indicated increasing biomass in GBK since 1986. Rapidly increasing biomass estimates were due to the short and noisy survey trend data for GBK (Figure A20) and in particular the relatively low 1990 survey observation. The sensitivity analysis consisted of a run with the 1990 survey observation omitted.

The KLAMZ model for GBK covered 1986-2002 using NEFSC clam survey data for the same period when sampling was relatively consistent in all strata (Table A8). Survey data for 1994 were excluded due to problems with the pump voltage. Catches were zero in all years. In other respects, the configuration of the KLAMZ model for GBK was identical to the configuration used for ocean quahog in other stock assessment areas.

Based on Lewis et al. (2001), ocean quahog growth is faster on Georges Bank than in southern areas. A von Bertalanffy growth curve was therefore fit to weight at age information for ocean quahog in GBK to obtain growth parameters used in the KLAMZ model. The weight at age information was obtained by converting Lewis et al.'s (2001) growth curve for length to meat weight at age using length-weight parameters for GBK (Table A9). The resulting von Bertalanffy curve for growth in weight $\left(W_{a}=41.07\left(1-e^{-0.04525(a-0.3695)}\right)\right.$ where $W_{a}$ was meat weight $(\mathrm{g})$ at age $a$ years) closely approximated the weight at age information. The growth parameters used in the KLAMZ model were $\rho=\mathrm{e}^{-0.04525}=0.9558$ and $J=w_{k-1} / w_{k}=15.59 / 16.66=0.9362$ where $w_{k}$ was the meat weight at age 13 which is approximately when ocean quahog reach 70 mm SL and become available to fishing (if fishing occurs).

Confidence intervals for estimated biomass on GBK were computed assuming that errors were from a lognormal distribution. In particular, the $95 \%$ bounds for the biomass estimate $B$ were computed $B e^{ \pm 1.96 \sigma}$ where $\sigma=\sqrt{\left(1+C V^{2}\right)}$ and CV is the arithmetic scale coefficient of variation. The CV was the ratio of the biomass estimate and arithmetic standard deviation estimated in the KLAMZ model using AD-Model builder libraries and the delta method.

Recruitment and surplus production rates from the KLAMZ model for GBK were compared to results from the LI region where a strong recruitment event occurred and where biomass appears to have increased at least slightly during some years (Figure A50). Recruitment estimates (assumed constant) in the two regions were divided by the area $\left(\mathrm{nm}^{2}\right)$ of each region to make estimates for the two regions comparable on a per unit area basis. The annual instantaneous surplus production rate for each region is $\bar{P}=\bar{G}+\bar{r}-M$ where $\bar{G}$ and $\bar{r}$ are average rates for somatic growth and recruitment. The average growth rate is the mean of annual rates which are computed automatically in KLAMZ (Appendix A5). The average recruitment rate is the mean of annual recruitment rates which were computed $r_{t}=R_{t} / \bar{B}_{t}$ with the average biomass during each year $\bar{B}_{t}$ computed automatically in KLAMZ (Appendix A5).

## KLAMZ-results for GBK trial and sensitivity runs

The estimated trends from KLAMZ model runs for GBK (Figures A52-A53) were judged implausible and not used for GBK because of the short survey time series (six observations during 1986 to 2002), frequency of survey strata that were not sampled
(Table A8), lack of catch data due to no fishing on GBK, no contrast in biomass levels due to catch that are usually used in stock assessment modeling to measure stock productivity, interannual variability and lack of consistent trend in survey data over time, statistically insignificant trend in survey data (see below under the heading "GBK at virgin biomass?"), lack of LPUE data to serve as corroboration, lack of evidence for recruitment in survey length data, and lack of historical biomass estimates for 1978 that might be used to calculate historical biomass. In addition, KLAMZ model estimates for GBK seemed implausible because the average surplus production rate and average recruitment per unit area for GBK were substantially higher than estimates for LI where a strong recruitment trend occurred and where biomass levels may have increased.

The trial model fit NEFSC clam survey data after 1994 better than before 1994 (Figure A52). With the 1998 survey observation omitted, the model fit was much better (Figure A53). The estimated survey scaling parameter for ESB data was $Q=0.98$ in both runs indicating that the model was able to match the observed ESB biomass levels during 1995-2005.

In the trial run (Figure A52), estimated biomass increased by about 99\% from $735,000 \mathrm{mt}$ during 1985 to $1,466,000 \mathrm{mt}$ during 2002 ( $5 \%$ per year). Means for annual recruitment and surplus production rates on GBK during 1985-2002 were 2.3 and 8.8 times larger than for LI. Mean recruitment per unit area on GBK (Figure A52) was twice as high as on LI. The $95 \%$ confidence interval for trends in estimated biomass (Figure A52) was broad and, at the extremes, included scenarios with stable trends.

In the sensitivity run omitting the 1989 survey (Figure A53), the increasing trend in biomass was not as steep. In particular, estimated biomass increased by about 48\% from $940,000 \mathrm{mt}$ during 1985 to $1,389,000 \mathrm{mt}$ during 2002 (2.4\% per year). Means for annual recruitment and surplus production rates on GBK during 1985-2002 were 1.6 and 5 times larger than for LI. Mean recruitment per unit area on GBK (Figure A54b) was 1.5 times as high as on LI. The $95 \%$ confidence interval for trends in estimated biomass (Figure 56) was broad and largely compatible with scenarios with stable trend.

## "Best" Estimates

KLAMZ model estimates were used at the best source of information about DMV, NJ, LI, and SNE during 1977-2005. VPA estimates were used for SVA and efficiency correct swept area biomass estimates were used for GBK (VPA and efficiency corrected swept-area biomass estimates for GBK are the same because no fishing has occurred there). NEFSC (2004) used VPA estimates for LI instead of KLAMZ model estimates. However, KLAMZ model estimates appear useful with addition of the 2005 survey data.

Biomass of ocean quahog and the entire stock less GBK during 1978-2005 was estimated by summing best estimates for each stock assessment area. Fishing mortality in large areas was computed by solving the catch equation with total catch, total biomass and $M=0.02 \mathrm{y}^{-1}$. CV's were not calculated for whole stock biomass or fishing mortality estimates because of difficulties accommodating covariance in the estimates for individual area that was due to using the same survey efficiency estimates as prior information.

Best estimates (Table A20 and Figure A54) show declines in ocean quahog biomass for southern regions (SVA, DMV and NJ) where the fishery has been continually active. In particular, biomass during 2005 was $5 \%, 34 \%$ and $44 \%$ of biomass during 1978 for SVA, DMV and NJ (Table A21).

Best estimates of biomass in northern regions, which did not support the fishery until recently (LI, SNE and GBK), are relatively flat and stable. LI biomass actually increased during 1978-1992 before fishing occurred. Biomass during 2005 was $94 \%$, $75 \%$ and $100 \%$ of biomass during 1978 for LI, SNE and GBK (Table A21). Biomass during 2005 was $76 \%$ and $66 \%$ of biomass during 1978 for the entire stock and the entire stock less GBK (Table A21).

Best estimates of fishing mortality rates (Figure A55) for southern areas where the fishery has been continually active (SVA, DMV and NJ) peaked during the late 1980's and early 1990's then declined as fishing effort shifted towards the north (Figures A4-A6 and A11). Fishing mortality rates in northern areas (Figure A55) were nearly zero before 1990 and increased substantially in later years as fishing effort shifted towards the north. Fishing mortality rates for the entire stock increased from about $0.003 \mathrm{y}^{-1}$ during 1978 to an average of about $0.006 \mathrm{y}^{-1}\left(0.010 \mathrm{y}^{-1}\right.$ for the entire stock less GBK) during the early 1990s through 2005.

## Proportions of total fishable biomass at various density levels

Best biomass estimates and survey data were combined to partition best biomass estimates into components found in areas with a range of biomass density levels. Biomass density is important to profitability of the ocean quahog fishery because it determines commercial catch rates. Biomass density was measured as survey catch per tow (fishable kg/tow) because commercial catch rate data for random locations and the entire stock area were not available. The analysis used random NEFSC clam survey tows during 1980-2005 (1994 excluded) that were in areas deep enough ( $\geq 20 \mathrm{~m}$ ) to be ocean quahog habitat. All survey data was from random stations so that the survey data would measure survey catch rates across the study area on average.

Survey data for stock assessment regions other than GBK were grouped into tenyear time intervals to increase sample size. Five surveys during 1980-1989, three surveys during 1990-1999 (excluding 1994), and two surveys during 2000-2005 were used in the analysis. Survey data for GBK were grouped into two intervals 1966-1992 and 19972002 and analyzed as a single group (1966-2002) because GBK was covered in fewer surveys and sample size was lower. The 1994 survey was excluded from all analyses because of problems with survey dredge efficiency and electrical voltage of current supplied to the pump.

Survey tow data were grouped by $5 \mathrm{~kg} /$ tow biomass density categories (e.g. catches of $0-4.9 \mathrm{~kg} /$ tow were assigned to the same biomass density category). The grouped data were used to calculate the proportion of fishing grounds occupied by ocean quahog at each biomass density level, as well as the proportion of fishable biomass on fishing grounds at each biomass density level (see below).

Proportions of fishable biomass in one region during a single time period were calculated:

$$
X_{L}=\frac{p_{L} K_{L}}{\sum_{j} p_{j} K_{j}}
$$

where $p_{L}$ is the proportion of random survey tows in biomass density category $L, K_{L}$ is mean survey fishable $\mathrm{kg} /$ tow for random stations in the same biomass density category, and the summation in the denominator is over all biomass density categories. The percentage of random tows in each biomass density category $p_{L}$ is an estimate of the
proportion of fishing grounds in each biomass density category. Total biomass at each density level during 2005 was calculated by multiplying the proportions $X_{L}$ for each region by the best estimate of total biomass in each region.

Results (Table A17) show reductions in the proportions of areas with high catch rates $\left(p_{L}\right)$ and the proportion of total stock biomass in areas of high catch rates $\left(X_{L}\right)$ within the southern DMV and NJ stock assessment regions where the most of the fishing for ocean quahog occurred historically. Proportions were variable in LI and SNE where less fishing has occurred.

During 2005 (Table A18), the largest component ( $19 \%$ or 575 thousand mt meats) of total fishable stock biomass was on GBK in the highest ( $25+\mathrm{kg} /$ tow) biomass density category. In contrast, stock biomass levels in density categories larger than $10 \mathrm{~kg} /$ tow were low for other regions.

## Building a bridge

Best estimates in this assessment are higher than in the previous assessment (NEFSC 2004) due mostly to the change in estimated survey dredge efficiency (Table 21). As expected, the ratios between current and previous biomass estimates were similar to ratios for efficiency corrected swept area biomass levels (Table A19).

## GBK at virgin biomass?

This section describes a hypothesis that fishable biomass on GBK has increased substantially since 1978 due to relatively fast growth and recruitment. The hypothesis is new and untested for GBK which has never been fished and is usually assumed to be at a high "virgin" level. The hypothesis is important because it affects estimates of stock productivity, decisions about biomass reference points (i.e. virgin biomass) and stock status determinations. No fishing occurs on GBK due to potential for PSP contamination, but experimental ocean quahog fisheries in the area are planned. Reviewer's comments and suggestions are important and will be considered in the next assessment. However, they will not affect choice of the best biomass estimates for this assessment.

Best estimates for GBK in this and recent assessments assume a flat biomass trend since 1978 at an equilibrium "virgin" level (NEFSC 2000; NEFSC 2004). In particular, averages of efficiency corrected swept area biomass estimates during 19972002 were used as estimates of average biomass over longer time periods. As described above, preliminary KLAMZ model runs for GBK are not suitable for estimating long term trends in ocean quahog biomass at this time primarily due to limited prior to 1986.

Analysis of NEFSC survey data for GBK is complicated because survey coverage tends to be spotty on GBK (Table A8). During 1986-2002, survey coverage was relatively complete but $14 \%$ ( 18 out of 126) strata had no tows in a given year (Table A8). Only five strata (55, 57, 59, 71 and 73) were sampled during all seven years. As described above, the survey during 1994 is not comparable to other surveys during 19862002 because of voltage problems. Thus, only six survey observations are available for analyzing trends in ocean quahog recruitment and biomass on GBK.

Lewis et al. (2001) carried out a spatially detailed analysis of NEFSC survey data for GBK focusing on growth, spatial patterns in length composition and trends in abundance by size. The major finding was that small ocean quahog were present and that recruitment was apparently occurring on GBK during the 1990s. Lewis et al. (2001) noted that size distributions from the 1980s had a single mode and were dominated by large individuals, 75-90 mm SL. In contrast, bimodal size distributions were observed
and small individuals ( $<70 \mathrm{~mm} \mathrm{SL}$ ) often represented $20-50 \%$ of the catch in numbers at stations during the 1990s along the southeast flank of GBK. The small individuals were attributed to spawning during the 1980s. Lewis et al. (2001) did not evaluate the potential contribution of small ocean quahog to the fishable biomass for the stock as a whole.

Lewis et al. (2001) estimated a a von Bertalanffy growth curve for GBK that showed faster growth to maximum size than the growth curve for ocean quahog in the Mid-Atlantic Bight (Figure A56). Faster growth should result in higher productivity on GBK. Based on both growth curves, ocean quahog growth is relatively rapid during the first years of life and much slower in older individuals as they grow large enough to enter the fishery. The size at $50 \%$ selectivity to the commercial fishery ( 72 mm SL ) is a reference point that separates recruits and the fishable stock. At 72 mm SL , ocean quahog on GBK grow about 1.5 mm SL per year while ocean quahog in other areas grow about 0.8 mm SL per year (Figure A56). The corresponding percentage increase in meat weight growth at 72 mm is $6 \%$ per year for GBK and $3 \%$ per year for other areas (Figure A56).

## Survey length data

The survey length composition data presented in this assessment and used by Lewis et al. (2001) show that small ocean quahog and presumably recruitment occurs throughout the range of the ocean quahog stock (Figure A26 and see Section 7). The clearest example is in LI where length compositions during the 1970s and 1980s have an obvious mode due to recruitment of small individuals. As pointed out by Lewis et al. (2001), small ocean quahog were more common on GBK after 1990 and this pattern is evident in length composition data used in this assessment (Figure A26). Compared to other areas, however, length composition data for GBK are stable with relatively few small individuals and little apparent recruitment (Figure A26).

It is unlikely that ocean quahog in GBK too small to be taken in the survey ( $<50$ mm SL) are escaping detection by growing to fishable size during the time between surveys. Annual growth increments in GBK are 3 mm for ocean quahog 50 mm SL and increments decrease with size. Thus, a small 50 mm SL ocean quahog would be expected to growth to no more than 59 mm SL during the three year interval between surveys. Moreover, based on the growth curve for TBK, ocean quahog 50 mm SL are about age 4 y and recruits to the fishable stock at 70 mm SL are about age 14 y so that at least 10 y would be required to grow to fishable size from 50 mm SL.

## Trends

Survey trends were computed for 1986-2002 (excluding 1994) using data (uncorrected for survey gear selectivity, Table A23) for ocean quahog $<70 \mathrm{~mm} \mathrm{SH}$ (mean numbers per tow to measure recruitment) and $\geq 70+\mathrm{mm}$ (mean weight per tow to measure recruited stock biomass). Strata with no tows were filled by borrowing (see above), which is the standard procedure for ocean quahog.

The time series of mean weight per tow biomass indices for GBK are short (6 data points, Figure A57) but seem to suggest increasing trends. Regression lines fit to the two time series seem to indicate that biomass of ocean quahog 70+ increased rapidly and that biomass of smaller ocean quahog $<70 \mathrm{~mm}$ increased slowly during 1986-2002. Neither regression was statistically significant ( $p$-value $=0.43$ for ocean quahog $<70 \mathrm{~mm}$ SL and
$p$-value $=0.21$ for ocean quahog $70+\mathrm{mm}$ ). The apparently increasing trends were due largely to relatively low mean kg/tow in the 1989 survey (Figure A57).

### 6.0 BIOLOGICAL REFERENCE POINTS (TOR-3)

The Atlantic Surfclam and Ocean Quahog Fishery Management Plan (FMP, Amendment 12) defines biological reference points used as management targets and thresholds for stock biomass and fishing mortality. Targets are intended to represent desirable stock conditions. Thresholds are intended to identify overfishing (fishing mortality too high) and overfished (stock biomass too low) stock conditions.

Biological reference points used in managing US fisheries including the fishery for ocean quahog are linked in policy and law to maximum sustained yield (MSY) concepts. In particular, the overfishing threshold is meant to be smaller than or equal to $F_{M S Y}$, the fishing mortality rate that provides MSY. Fishing mortality levels higher than $F_{M S Y}$ constitute overfishing.

The biomass and fishing mortality targets specified in the FMP for ocean quahogs are $B_{\text {Target }}=B_{M S Y}$, which is assumed be one-half of the virgin biomass for the whole stock, and $F_{\text {Target }}=F_{0.1}$ for the exploited region (whole stock less GBK) The biomass and fishing mortality thresholds are $B_{\text {Threshold }}=1 / 2 B_{M S Y}$ and $F_{\text {Threshold }}=\mathrm{F}_{25 \%}$ (the fishing mortality rate that reduces life time egg production for an average female to $25 \%$ of the level with no fishing). The FMP does not specify whether the thresholds apply to the whole stock or exploited region only.

Biological reference points for ocean quahog defined in the FMP were recalculated for this assessment resulting in substantial changes to $F_{25 \%}$ and $F_{M A X}$ (the fishing mortality rate that maximizes yield per recruit). The new and old estimates for $F_{0.1}$ are similar (Table A24 and Figure A58). Sensitivity analysis indicates that assumptions about natural mortality had substantial effect on estimated reference points (Table A24).

In recalculating biological reference points, the Invertebrate Subcommittee noted that the current threshold reference point for fishing mortality (new estimate $F_{25 \%}=0.0517$ $\mathrm{y}^{-1}$, Table A24) is a poor proxy for $F_{M S Y}$ in a long-lived species like ocean quahog with natural mortality rate $M=0.02 \mathrm{y}^{-1}$ (Clark 2002; Thorarinsdottir and Jacobson 2005). From a purely technical perspective, it would be advantageous to reconsider biological reference points in the FMP for ocean quahog and their application to the entire or exploited portions of the stock.

Simulation analyses in Clark (2002) show that the highest sustainable catches for long lived stocks like ocean quahog are achieved when lower fishing mortality rates are applied at relatively high stock biomass levels. The same simulations show that fishing at $F_{25 \%}$ would eventually depress stock spawning stock biomass to less than $25 \%$ of the virgin level, a level likely far below $B_{M S Y}$. In the simulations, long-term yield from unproductive stocks was maximized at fishing mortality rates lower than $F_{50 \%}$ (Clark 2002). Fortunately, the ocean quahog fishery is currently managed under an individual ITQ system with a quota on landings that keeps fishing mortality rates lower than both $F_{0.01}$ and $F_{25 \%}$. The current quota is based on market demand and other economic factors.

## Revised biomass reference points (building a bridge)

New proxies for virgin biomass and $B_{M S Y}$ in this assessment are substantially larger than in NEFSC (2003). The proxy for virgin ocean quahog biomass was recalculated using the best estimates of stock biomass during 1978 for each region (3.973 million mt including GBK, Table A20). The proxy for $B_{M S Y}(1 / 2$ virgin biomass) in this assessment 1.987 million mt including GBK. Proxies for virgin biomass and $B_{M S Y}$ in NEFSC (2004) were smaller ( 3.3 and 1.5 million mt ). The new estimates are larger mainly because of changes in survey dredge efficiency estimates ( $e=0.165$ instead of $0.269-0.346)$. In addition, the new reference points are fishable biomass rather than biomass 70+ mm SL.

## Fishing mortality reference points (building a bridge)

Biological reference points for fishing mortality were calculated for ocean quahog in this assessment using a length-based per-recruit model that is part of the NEFSC Stock Assessment Toolbox. ${ }^{6}$ The length-based model is similar to the Thompson and Bell (1934) age-based model except that selectivity, maturity and growth are specified in terms of length, rather than age. The length-based approach is advantageous for ocean quahog because fishery selectivity and maturity are better known in terms of length than age (Figure A59).

Biological assumptions for reference point calculations in this assessment were generally comparable to assumptions in the last assessment (Figure A60). The ascending logistic fishery selectivity curve in per recruit model calculations was the same as in calculation of fishable survey biomass trends. The von Bertalanffy growth curve for length at age was the same as used earlier in this assessment for the MAB (Figure A59). Length-weight parameters $(\ln (\alpha)=-9.242, \beta=2.821)$ were averages for the stock as a whole.

Maturity at length was from Thorarinsdottir and Jacobson (2005) for ocean quahog in Icelandic waters with $10 \%, 50 \%$ and $90 \%$ of female ocean quahog mature at 40, 64, and 88 mm SL (2, 19, and 61 y , based on the growth curve in Figure A59). Based on the size range of samples (G. Thorarinsdottir, pers. comm..), the maturity curve is probably valid for ocean quahog in the size range used to estimate fishing mortality.

Maturity information for ocean quahog in the US EEZ is scant (see review in Cargnelli et al. 1999) but all available information and age-based per-recruit model calculations in the last assessment are compatible with the maturity at length estimates for ocean quahog in Icelandic waters (Figure A60).

### 7.0 STOCK STATUS (TOR-4)

Ocean quahog in the US EEZ are not overfished and overfishing is not occurring. Stock biomass during 2005 was 3.039 million mt (Table A20) and above the revised management target of $1 / 2$ virgin biomass $=1.987$ million mt (Figure A61). The fishing mortality rate during 2005 (all areas but GBK) was $F=0.0077 \mathrm{y}^{-1}$ (Table A20), which is below the revised management target level $F_{0.1}=0.0278 \mathrm{y}^{-1}$ (Figure A61)

[^5]
## Biological condition of the entire EEZ stock

The ocean quahog population is a relatively unproductive with total biomass gradually approaching the $B_{M S Y}$ reference point ( $1 / 2$ virgin biomass, estimated as $50 \%$ of biomass during 1978) gradually after about three decades of relatively low fishing mortality (Table A20 and Figures A54-A55).

Based on survey data (Figure A20), LPUE data (Figure A8) and best estimates for 1977-2005 (Figure A54), declines in stock biomass are most pronounced in southern regions (SVA, DMV and NJ) where the fishery has been active longest. In particular, stock biomass was below the $1 / 2$ virgin level during 2005 in SVA, DMV and NJ (Table A21).

An increasingly large fraction of the stock ( $42 \%$ during 2005 compared to $38 \%$ during 1978, Table A25) is in northern regions (LI and SNE) where fishing is relatively recent and in the GBK region, which is not fished due to risk of PSP contamination (Figure A54).

## Fishing effort and mortality

Fishing effort has shifted to offshore and northern grounds over time as catch rates and abundance in the south declined (Figures A2, A4, A8 and A54). Analysis of LPUE data for individual 10' squares indicates considerable fishing down on fishing grounds that historically supplied the bulk of landings (Figures A13-A15). There is no clear indication that LPUE increased on historical grounds after fishing effort was reduced.

Fishing mortality rates during 2005 are relatively low for the entire stock $\left(F=0.0045 \mathrm{y}^{-1}\right)$ and for the fishable stock ( $F=0.0077 \mathrm{y}^{-1}$ ), which excludes GBK (Figure A55). Fishing mortality rates in the south where biomass was relatively low during 2005 decreased substantially over the last decade to low levels ( $F=0.0,0.0094$ and $0.0017 \mathrm{y}^{-1}$ for SVA, DMV and NJ) during 2005. Fishing mortality rates for LI increased abruptly during 1992 as effort increased, declined and then increased to $F=0.0145 \mathrm{y}^{-1}$ in 2005. The fishing mortality rate in LI during 2005 is comparable to fishing mortality rates in southern areas as they were fished down to relatively low biomass levels.

## Productivity under fishing

Questions about the potential productivity of ocean quahog are becoming important as the stock is fished down from high virgin levels to $B_{M S Y}$. Uncertainties about productivity are close related to choice of an accurate $F_{M S Y}$ proxy and other decisions that affect sustainability and fishery profitability.

Ocean quahog in the EEZ do not currently show a clear increase in stock productivity, due to higher recruitment and increased growth rates, that would be expected as biomass declines to $B_{M S Y}$ levels. Given the long periods between settlement and recruitment and slow growth once ocean quahog reach fishable size, any increase in stock productivity may be delayed (Powell and Mann 2005).

Recruitment events appear to be regional and sporadic (i.e. often separated by decades). Survey length composition data show that recruitment occurs throughout the resource sporadically and at an apparently low rate. Based on survey length composition data, some recent recruitment is evident in DMV, NJ, LI, SNE and GBK during recent years (Figure A26). Lewis et al. (2001) describe recruitment on GBK during the 1990s. Powell and Mann (2005) used a lined commercial dredge on a directed survey during 2002 and detected recruitment in some regions across the Mid-Atlantic Bight. Slow
growth at sizes large enough to recruit to the fishery probably reduces the contribution of new recruits to fishery productivity (A62).

Information about growth of ocean quahog is sparse (Lewis et al. 2001). It is not possible to detect potential changes in growth at this time or to detect differences among regions (other than in GBK).

## Biological condition of ocean quahog in Maine waters

The State of Maine carried out a survey and a stock assessment was completed for a portion of the ocean quahog stock in Maine waters (Russell 2006). The survey and assessment cover the principal fishing grounds in Maine waters. The fishery and biological characteristics of ocean quahog in Maine coastal waters are unique. In particular, the fishery targets small ocean quahogs for sale on the half shell market at prices roughly ten times the price paid in the rest of the EEZ. Most of the information in this section is from the assessment report for Maine waters (Russell 2006).

Biological and fishery information for Maine waters were used in the length based per recruit model (also used for the rest of the EEZ, see Section 6) to estimate conventional biological reference points for Maine waters only. In particular, $F_{\text {MAX }}=$ $0.0561, F_{0.1}=0.0247$ and $F_{50 \%}=0.013 \mathrm{y}^{-1}$ for ocean quahog in Maine waters.

Assessment results for Maine show relatively high levels of fishing effort (Figure A4) and landings in recent years (Figure A2). LPUE levels have declined since the peak in 2002, but remain at relatively high levels overall (Figure A8).

Based on survey results and dredge efficiency estimates, stock biomass available to the fishery during 2005 was about $22,493 \mathrm{mt}$ meats. In comparison, catch (landings plus a $5 \%$ incidental mortality allowance) during 2005 was 505 mt meats. The biomass estimate and catch data are for the area surveyed which includes the main areas of commercial fishing in Maine waters. Biomass in Maine waters is underestimated to the extent that it excludes ocean quahog outside the area where fishing occurs and the survey was carried out.

Fishing mortality during 2005 the assessed was estimated to be $F=505 \div 22,493$ $=0.022 \mathrm{y}^{-1}$, which is almost equal to $F_{0.1}=0.0247^{-1}$ calculated from a per recruit model for ocean quahog in Maine waters. The $F_{0.1}$ estimate for Maine waters has no special significance in policy because, based on the FMP, biological reference points used in defining management targets and thresholds are estimated for and applied to the entire stock.

Management goals have not been described for ocean quahog in Maine waters but maximization of long term catch is a likely candidate. Based on simulation analyses for long-lived and unproductive fish species (Clark 2002), fishing mortality rates as low as $F_{50 \%}=0.013 \mathrm{y}^{-1}$ may be required if spawning stock must be conserved to maximize long term catch levels.

The importance of maintaining spawning stock in Maine waters may be low if the bulk of recruits originate in the EEZ outside of the relatively small Maine fishing grounds. In that case, $F_{0.1}=0.0247 \mathrm{y}^{-1}$ might be useful reference point for maximizing long term catch because it would probably provide relatively high levels of yield while preserving some spawning potential. If spawning biomass in Maine waters is completely irrelevant, then long term catch might be maximized by fishing at $F_{M A X}=0.0561 \mathrm{y}^{-1}$. However, $F_{M A X}$ is likely to require high levels of fishing effort and the estimate of $F_{M A X}$ is sensitive to small changes in growth and fishery selectivity parameters.

### 8.0 TAL and PROJECTIONS (TOR-5 \& 6)

Under current quota regulations, annual total allowable landings (TAL) for ocean quahog during 2007 is $24,190 \mathrm{mt}$ meats ( 5.333 million bushels). The quota and TAL will result in a fishing mortality rate of approximately $F=24,190 \div 1,775,000=0.014 \mathrm{y}^{-1}$ for the exploitable portion of the stock (excluding GBK) and $F=24,190 \div 3,039,000=0.008$ $\mathrm{y}^{-1}$ for the stock as a whole if biomass during 2007 is similar to biomass during 2005 ( 1,775 and 2,698 million mt ). TAL levels for longer time periods and for constant levels of fishing mortality can be calculated by projection, as described below.

## Projections

A simple method for making short term projections for ocean quahog biomass, catch and fishing mortality is demonstrated in this section with example calculations. Example calculations assume either: 1) constant regional catch at 4, 5.33 and 6 million bushels; 2) constant fishing mortality at the manager's target level, $F_{0.1}=0.0275 \mathrm{y}^{-1}$. In the calculations wit $F_{0.1}$, for example, predicted landings could be used as TAL.

All projection calculations use the following equations to represent biomass dynamics:

$$
\begin{aligned}
& X=G+r-M-F \\
& B_{t+1}=B_{t} e^{X} \\
& F=\frac{C}{B} \quad \text { or } \quad C=F B
\end{aligned}
$$

where $X$ is the net instantaneous annual rate of change, $G$ is the instantaneous rate for somatic growth in weight, $r$ is the rate for recruitment, $M=0.02 \mathrm{y}^{-1}$ is the natural mortality rate, $C$ is catch (e.g. quota for landings $+5 \%$ ), and $B$ is fishable biomass.

When catch is assumed known, the fishing mortality rate $F$ can be calculated iteratively (e.g. Solver in Excel). When $F$ is known, catch can be calculated directly.

Input data for projections are summarized in Table A26. Estimates of initial biomass (in 2005) and fishing mortality during 2005 were best estimates from Table A15. Catches (landings $+5 \%$ ) in 2006 are assumed to be the same as in 2005. In projections with constant $F=F_{0.01}=0.0278 \mathrm{y}^{-1}$ for exploited regions (excluding GBK) the proportions of catch in each region during 2006-2010 are assumed to be the same as in 2005. In projections for GBK, which is virgin and normally assumed to be at equilibrium carrying capacity in stock assessment work, rates for fishing mortality, natural mortality, growth and recruitment were zero so that stock biomass in GBK did not change over time. All of the projections suggest that the stock as a whole will continue to decline gradually over time (Table A27-A30). The decline is relatively rapid with $F=F_{0.01}$ (Table A31).

The method for ocean quahog is deterministic and does not consider natural variability in recruitment, growth or natural mortality. However, uncertainty in short term projections is primarily due to uncertainty in initial biomass estimates. Recruitment, natural mortality and growth of ocean quahog occur at low rates that have little effect on short term projections. Thus, CVs for efficiency corrected swept area biomass during 2005 (see below) can serve as reasonable measures of uncertainty in projections.

| CVs for projected biomass levels from Table A15. |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| SVA | DMV | NJ | LI | SNE | GBK | Total less <br> GBK | Total |
| $104 \%$ | $55 \%$ | $30 \%$ | $31 \%$ | $36 \%$ | $32 \%$ | $24 \%$ | $24 \%$ |

If uncertainty in short-term biomass projections is lognormal, then bounds for an asymmetric $95 \%$ confidence interval around projected biomass can be computed $B e^{ \pm 1.96 \sigma}$ where $\sigma=\sqrt{\ln \left(C V^{2}+1\right)}$.

### 9.0 RESEARCH RECOMMENDATIONS (TOR-7)

Recommendations from the previous assessment and new research recommendations are described sequentially.

## Recommendations from last assessment

- A complete survey and a valid survey dredge efficiency estimate are needed by the State of Maine to assess ocean quahogs off the coast of Maine.

A directed survey for ocean quahog that covered the main fishing grounds in Maine waters was completed by the Maine Department of Marine Resources during 2005 (Russell 2006). Data from box core and dredge sampling during 2006 were used to estimate survey dredge efficiency. The 2005 survey and efficiency estimate were used to estimate fishing mortality and biomass for ocean quahog in Maine waters (Russell 2006).

- Explore whether efficiency of the DE-II dredge and commercial dredges are affected by depth, sediment type, and clam density. This could be examined experimentally, or by having an efficient commercial dredge repeat stations sampled by the RV DEII. Also, evaluate non-extractive methods to estimate dredge efficiency and survey the resource.

Data collected during 2002 and new data collected during 2005 were examined in this assessment to determine if dredge efficiency depends on depth, sediment type or clam density. Additional data and analysis are required, however, to address this research recommendation. Non-extractive methods for estimating dredge efficiency were not investigated.

- Identify whether there are major differences in life histories and population dynamics between regions, and consider treating the EEZ stock as metapopulations.

A review of life history characteristics and analysis of population dynamics of ocean quahog in Maine waters was completed (Russell 2006). Alternate spatial based management approaches were not addressed in this assessment.

- Consider using ecological estimates of carrying capacity (based on available food, maximum size, predation, amount of suitable habitat) to evaluate/validate model estimates of virgin biomass.

Ecological estimates of carrying capacity were not addressed in this assessment. However, information suggesting that ocean quahog biomass on GBK (a virgin area) is increasing was examined and presented for review.

- Re-examine the rate of incidental mortality to ocean quahogs caused by commercial dredges.

No new field work or data analysis were carried out to address the research recommendation.

- Consider applying the relative selectivity function to the entire survey time series.

A survey selectivity curve was estimated for ocean quahog in the EEZ and a fishery selectivity curve estimated for ocean quahog off Iceland were used to better interpret survey data.

- Consider whether future stock assessment models should be based on age and abundance, rather than shell length and weight.

No progress.

- There is little information regarding $F_{M S Y}$ and $B_{M S Y}$ or suitable proxies for long lived species like ocean quahog. Traditional proxies (e.g., $F_{M S Y}=F_{25 \%}$ MSP $, F_{M S Y}=M, F_{M S Y}$ $=F_{0.1}$ and $B_{M S Y}$ at one-half virgin biomass) may be inappropriate for long lived organisms. The question of $\mathrm{F}_{\text {MSY }}$ and $\mathrm{B}_{\text {MSY }}$ proxies should be considered.

Traditional reference points from per recruit calculations were revised in this assessment using a new length based model and new estimates of fishery selectivity and maturity at length. Recent simulation work for long-lived rockfish and results for Icelandic ocean quahog were reviewed. The simulation results indicate that $F_{0.1}$ and $F_{25 \%}$ are likely poor proxies for $F_{\text {MSY }}$ in a long-lived organism like ocean quahog. Based on the simulations $F_{50 \%}$ may be a better proxy. These issues could be taken up the next time the fishery management plan is revised.

- Survey coverage of Georges Bank needs to be a priority in NMFS EEZ survey. Strata along the Hague line may need to be re-stratified and biomass estimates recalculated to include only US areas.

GBK was not surveyed during 2005 due to competing priorities for sampling in southern areas. However, this remains an important issue, particularly in view of hypotheses that stock biomass is increasing on GBK. Different stratification schemes were not investigated.

- If the management system requires accurate position information (e.g. VMS) from fishery vessels, evaluate the possible improvements to assessments using catch and location information from this source.

The working group discussed this topic but it is not mentioned in the report because the discussions were preliminary.

- Investigate the use of survey data collected prior to 1978.

No progress.

## New Recommendations (not prioritized)

- The $R / V$ Delaware II may not be available for use on NEFSC clam surveys after 1998 and it appears likely that the clam survey will become a cooperative effort with sampling from a commercial vessel. Both the $R / V$ Delaware II and commercial vessel should be used during 1998 so that catch rates, efficiency and selectivity patterns for the two vessels can be compared and calibrated. Planning should commence immediately.
- Fishing mortality and biomass reference points used as proxies for $F_{M S Y}$ and $B_{M S Y}$ should be reevaluated in the next assessment.
- Additional estimates of survey dredge efficiency from cooperative depletion studies are required.
- Develop a length (and possibly age) structured stock assessment model for ocean quahog that makes better use of survey and fishery length composition data which may provide better estimates of recruitment trends.
- Conduct further experimental work to determine the relationship between dredge efficiency, depth, substrate and clam density. A comprehensive study coincident with the next NEFSC clam survey would be most useful. The experimental design should include sufficient contrast in variables that may affect dredge efficiency.
- Cover GBK in the next NEFSC clam survey.
- Investigate the survey data from GBK during the 1989 survey to determine why it is low relative to survey observations during earlier years. This may be important in determining if biomass is increasing in GBK.
- Survey strata with no tows are a particular problem in the GBK region. The current procedure for filling holes in survey data involves borrowing data from adjacent surveys. This may not be optimal for ocean quahog surveys and GBK in particular. In the next assessment, consider filling holes in the GBK survey data using a model with stratum and year effects.
- Evaluate possible increasing trends in biomass for ocean quahog on GBK.
- Evaluate effects and contribution of recruitment to stock productivity.
- Improve estimates of biological parameters for age, growth (particularly of small individuals), and maturity for ocean quahog in both the EEZ and in Maine waters.
- Survey dredge and commercial dredge efficiency estimates should be reevaluated by field work during the next NEFSC clam survey. The next survey may be the last opportunity to estimate survey dredge selectivity. The commercial dredge selectivity curve was used in this assessment was estimated from field studies done off Iceland
where conditions may differ. Repeat tow experiments (i.e. survey stations reoccupied by commercial vessels) may be useful for this purpose.
- In the next assessment, projection calculations should be carried out using a model that is basically the same as the primary stock assessment model used to estimate biomass and fishing mortality (e.g. delay-difference population model in KLAMZ).
- Recommendations for future depletion studies.
- It was difficult to find areas with high concentrations of ocean quahog for depletion experiment sites during 2005. However, areas with lower densities of ocean quahog can be used if depletion tow distance is increased.
- Revised estimators for survey dredge efficiency based on commercial depletion experiments and setup tows use data for relatively large ocean quahog (i.e. $90+\mathrm{mm}$ ) only. Future depletion sites should contain reasonably high densities of large individuals.
- In future, every effort must be made to collect and record precise location data at short time intervals during depletion studies.
- Collect length and bushel count data from survey and depletion tows more frequently (e.g. every 1-2 tows). It might be advantageous to measure fewer individuals sampled from more tows.
- Analyze results from previous depletion studies to determine if differences between bushel counts and length composition data from different tows in the same depletion experiment are significantly different. Use the results to modify sampling protocols as appropriate.
- Changes in length composition during a depletion experiment might be incorporated into efficiency estimation by, for example, including selectivity parameters in the Patch model. Efficiency estimates (and commercial selectivity) might be more precise because more size groups would be included in catch data.
- It would be useful to analyze efficiency estimates in terms of season because ocean quahog are believed to change their depth in sediments on a seasonal basis.
- The next stock assessment should review the $M=0.02 \mathrm{y}-1$ assumption for ocean quahog.
- In the next assessment, KLAMZ model runs with two recruitment parameters should be explored for LI and SNE. Survey length composition show more recruitment prior to 1994 than afterwards. Model fit was not as good for SNE as other stock assessment regions.
- KLAMZ model runs for GBK should be explored further in the next assessment.


### 10.0 ACKNOWLEDGMENTS

The Invertebrate Subcommittee thanks all hands (ship's crew, scientific crew and skipper) on board the F/V Lisa Kim during the cooperative 2005 depletion experiments. The fishing industry funded and carried out a number of key research project that were important to the completion of this assessment. J. Weinberg, P. Rago and C. Pickett (Northeast Fishery Science Center) and R. Marzak (Haskin Shellfish Research Laboratory) made useful contributions.

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## OCEAN QUAHOG TABLES

Table A1. Landings (1,000 mt meats) for ocean quahog during 1967-2005 from dealer data (state + EEZ waters) and logbooks (EEZ only). Landings from state waters are calculated approximately by subtracting logbook landings from dealer landings. The EEZ quota and ratio of EEZ landings and EEZ quota are shown for comparison. Data for 2005 are preliminary and may be incomplete.

| Year | Dealer Database | EEZ <br> (Logbook) | State Waters (Logbook Dealer) | Percent Landings in EEZ | EEZ Quota | EEZ Landings / Quota (\%) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $1967{ }^{\text {a }}$ | 0.020 | 0.000 | 0.020 | 0.000 |  |  |
| 1968 | 0.102 | 0.000 | 0.102 | 0.000 |  |  |
| 1969 | 0.290 | 0.000 | 0.290 | 0.000 |  |  |
| 1970 | 0.792 | 0.000 | 0.792 | 0.000 |  |  |
| 1971 | 0.921 | 0.000 | 0.921 | 0.000 |  |  |
| 1972 | 0.634 | 0.000 | 0.634 | 0.000 |  |  |
| 1973 | 0.661 | 0.000 | 0.661 | 0.000 |  |  |
| 1974 | 0.365 | 0.000 | 0.365 | 0.000 |  |  |
| 1975 | 0.569 | 0.000 | 0.569 | 0.000 |  |  |
| 1976 | 2.510 | 1.854 | 0.656 | 0.739 |  |  |
| 1977 | 8.411 | 7.293 | 1.118 | 0.867 |  |  |
| 1978 | 10.415 | 9.197 | 1.218 | 0.883 |  |  |
| 1979 | 15.748 | 14.344 | 1.404 | 0.911 | 13.608 | 105\% |
| $1980^{\text {D, }}$ | 11.623 | 13.407 | -1.784 | 1.153 | 15.876 | 84\% |
| 1981 | 11.202 | 13.101 | -1.899 | 1.170 | 18.144 | 72\% |
| 1982 | 16.478 | 14.234 | 2.244 | 0.864 | 18.144 | 78\% |
| 1983 | 16.200 | 14.586 | 1.615 | 0.900 | 18.144 | 80\% |
| 1984 | 17.939 | 17.974 | -0.035 | 1.002 | 18.144 | 99\% |
| 1985 | 22.035 | 20.726 | 1.310 | 0.941 | 22.226 | 93\% |
| 1986 | 20.585 | 18.902 | 1.683 | 0.918 | 27.215 | 69\% |
| 1987 | 22.709 | 21.514 | 1.195 | 0.947 | 27.215 | 79\% |
| 1988 | 21.007 | 20.273 | 0.734 | 0.965 | 27.215 | 74\% |
| 1989 | 23.147 | 22.359 | 0.788 | 0.966 | 23.587 | 95\% |
| 1990 | 21.235 | 20.965 | 0.270 | 0.987 | 24.040 | 87\% |
| 1991 | 22.119 | 22.063 | 0.056 | 0.997 | 24.040 | 92\% |
| 1992 | 22.871 | 22.476 | 0.395 | 0.983 | 24.040 | 93\% |
| 1993 | 24.843 | 21.876 | 2.968 | 0.881 | 24.494 | 89\% |
| 1994 | 21.159 | 20.985 | 0.174 | 0.992 | 24.494 | 86\% |
| 1995 | 23.253 | 21.107 | 2.145 | 0.908 | 22.226 | 95\% |
| 1996 | 21.122 | 20.061 | 1.062 | 0.950 | 20.185 | 99\% |
| 1997 | 19.930 | 19.628 | 0.302 | 0.985 | 19.581 | 100\% |
| 1998 | 18.098 | 17.896 | 0.201 | 0.989 | 18.144 | 99\% |
| 1999 | 17.557 | 17.381 | 0.175 | 0.990 | 20.412 | 85\% |
| 2000 | 14.899 | 14.722 | 0.176 | 0.988 | 20.412 | 72\% |
| 2001 | 17.234 | 17.068 | 0.165 | 0.990 | 20.412 | 84\% |
| 2002 | 18.144 | 17.947 | 0.198 | 0.989 | 20.412 | 88\% |
| 2003 | 18.997 | 18.815 | 0.182 | 0.990 | 20.412 | 92\% |
| 2004 | 17.788 | 17.650 | 0.138 | 0.992 | 22.680 | 78\% |
| 2005 |  | 13.629 | -13.629 |  | 24.190 | 56\% |

${ }^{\text {a }}$ Landings for 1967-1979 are from NEFSC (1990)
${ }^{\mathrm{b}}$ Landings for 1980-1993 from NEFSC (2003).
${ }^{\text {c }}$ For 1980-2005, "Dealer Database Total" landings are from commercial landings databases (CFDETS or CFDERS), EEZ landings are from logbooks (Maine included), and "State Waters (Dealer-Logbook)" landings are the difference. Logbook landings are more accurate. In some years, logbook landings exceeded dealer database totals slightly.

Table A2. Ocean quahog landings (mt meats) by stock assessment region reported in logbooks for the US EEZ. Data for 1980-2003 are from logbooks and differ from the previous assessment (NEFSC 2004) because additional landings from other/unknown regions ("UNK") were allocated to regions in this assessment and because NEFSC (2004) treated Maine landings as other/unknown. Landings for 1978-1979 are not from logbooks and less reliable. Data for 2005 are preliminary and may be incomplete. Based on Maine reports, UNK amounts during 2002 were probably from Maine waters.

| YEAR | SVA | DMV | NJ | LI | SNE | GBK | MNE | UNK | Grand <br> Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1978 |  | 1,290 | 6,350 |  |  |  |  | 2,775 | 10,415 |
| 1979 |  | 5,450 | 6,030 |  |  |  |  | 4,268 | 15,748 |
| 1980 | 0 | 4,230 | 7,750 | 6 | 0 |  |  | 1,421 | 13,407 |
| 1981 | 56 | 3,637 | 8,402 | 3 | 0 |  |  | 1,003 | 13,101 |
| 1982 | 6 | 4,598 | 8,538 | 0 | 0 |  |  | 1,092 | 14,234 |
| 1983 | 0 | 5,396 | 8,249 | 21 | 629 | 0 | 0 | 291 | 14,586 |
| 1984 | 6 | 7,164 | 8,857 | 0 | 822 | 0 | 0 | 1,125 | 17,974 |
| 1985 | 160 | 7,200 | 10,676 | 40 | 693 | 0 | 0 | 1,956 | 20,726 |
| 1986 | 0 | 8,236 | 9,053 | 396 | 568 | 0 | 0 | 649 | 18,902 |
| 1987 | 0 | 10,533 | 9,077 | 1,180 | 696 | 0 | 0 | 27 | 21,514 |
| 1988 | 42 | 11,715 | 7,014 | 640 | 841 | 0 | 0 | 20 | 20,273 |
| 1989 | 0 | 6,439 | 14,100 | 605 | 1,196 | 0 | 0 | 20 | 22,359 |
| 1990 | 14 | 3,685 | 15,590 | 739 | 934 | 0 | 3 | 0 | 20,965 |
| 1991 | 0 | 4,839 | 14,575 | 1,674 | 865 | 0 | 110 | 0 | 22,063 |
| 1992 | 0 | 2,378 | 6,942 | 11,939 | 1,143 | 0 | 75 | 0 | 22,476 |
| 1993 | 0 | 1,953 | 10,205 | 8,642 | 1,020 | 0 | 56 | 0 | 21,876 |
| 1994 | 0 | 992 | 6,938 | 12,014 | 954 | 0 | 65 | 22 | 20,985 |
| 1995 | 0 | 699 | 5,356 | 9,526 | 5,412 | 0 | 114 | 0 | 21,107 |
| 1996 | 0 | 736 | 4,864 | 5,943 | 8,350 | 0 | 142 | 26 | 20,061 |
| 1997 | 0 | 1,072 | 4,229 | 5,141 | 8,968 | 0 | 218 | 0 | 19,628 |
| 1998 | 0 | 1,365 | 2,684 | 6,856 | 6,736 | 0 | 218 | 39 | 17,896 |
| 1999 | 0 | 1,090 | 3,038 | 6,329 | 6,618 | 0 | 279 | 27 | 17,381 |
| 2000 | 0 | 1,048 | 3,318 | 4,745 | 5,083 | 49 | 357 | 123 | 14,722 |
| 2001 | 0 | 894 | 4,560 | 5,692 | 4,694 | 13 | 326 | 889 | 17,068 |
| 2002 | 0 | 1,732 | 2,781 | 9,113 | 3,884 | 0 | 387 | 51 | 17,947 |
| 2003 | 0 | 896 | 3,692 | 11,617 | 2,177 | 0 | 359 | 73 | 18,815 |
| 2004 | 0 | 634 | 2,795 | 10,631 | 3,283 | 0 | 307 | 0 | 17,650 |
| 2005 | 0 | 932 | 664 | 9,688 | 2,015 | 0 | 294 | 35 | 13,629 |

Table A3. Ocean quahog landings by stock assessment region as reported in logbooks for the US EEZ. Figures are 1000 ITQ bushels except for Maine, which are reported as both ITQ and Maine bushels. Data for 2005 are preliminary and may be incomplete. Based on Maine reports, UNK amounts during 2002 were probably from Maine waters.

| YEAR | SVA | DMV | NJ | LI | SNE | GBK | MNE | MNE <br> (Maine <br> bushels) | UNK | Grand <br> Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1980 | 0 | 933 | 1,709 | 1 | 0 | 0 | 0 |  | 313 | 2,956 |
| 1981 | 12 | 802 | 1,852 | 1 | 0 | 0 | 0 |  | 221 | 2,888 |
| 1982 | 1 | 1,014 | 1,882 | 0 | 0 | 0 | 0 |  | 241 | 3,138 |
| 1983 | 0 | 1,190 | 1,819 | 5 | 139 | 64 | 0 | 0 | 64 | 3,280 |
| 1984 | 1 | 1,580 | 1,953 | 0 | 181 | 248 | 0 | 0 | 248 | 4,211 |
| 1985 | 35 | 1,587 | 2,354 | 9 | 153 | 431 | 0 | 0 | 431 | 5,001 |
| 1986 | 0 | 1,816 | 1,996 | 87 | 125 | 143 | 0 | 0 | 143 | 4,310 |
| 1987 | 0 | 2,322 | 2,001 | 260 | 153 | 6 | 0 | 0 | 6 | 4,749 |
| 1988 | 9 | 2,583 | 1,546 | 141 | 185 | 4 | 0 | 0 | 4 | 4,474 |
| 1989 | 0 | 1,420 | 3,108 | 133 | 264 | 4 | 0 | 0 | 4 | 4,934 |
| 1990 | 3 | 812 | 3,437 | 163 | 206 | 0 | 1 | 1 | 0 | 4,623 |
| 1991 | 0 | 1,067 | 3,213 | 369 | 191 | 0 | 24 | 37 | 0 | 4,901 |
| 1992 | 0 | 524 | 1,530 | 2,632 | 252 | 0 | 16 | 25 | 0 | 4,980 |
| 1993 | 0 | 431 | 2,250 | 1,905 | 225 | 0 | 12 | 19 | 0 | 4,841 |
| 1994 | 0 | 219 | 1,530 | 2,649 | 210 | 5 | 14 | 21 | 5 | 4,653 |
| 1995 | 0 | 154 | 1,181 | 2,100 | 1,193 | 0 | 25 | 38 | 0 | 4,691 |
| 1996 | 0 | 162 | 1,072 | 1,310 | 1,841 | 6 | 31 | 47 | 6 | 4,476 |
| 1997 | 0 | 236 | 932 | 1,133 | 1,977 | 0 | 48 | 73 | 0 | 4,400 |
| 1998 | 0 | 301 | 592 | 1,511 | 1,485 | 9 | 48 | 72 | 9 | 4,026 |
| 1999 | 0 | 240 | 670 | 1,395 | 1,459 | 6 | 62 | 93 | 6 | 3,931 |
| 2000 | 0 | 231 | 732 | 1,046 | 1,121 | 27 | 79 | 119 | 27 | 3,381 |
| 2001 | 0 | 197 | 1,005 | 1,255 | 1,035 | 196 | 72 | 109 | 196 | 4,065 |
| 2002 | 0 | 382 | 613 | 2,009 | 856 | 11 | 85 | 129 | 11 | 4,097 |
| 2003 | 0 | 198 | 814 | 2,561 | 480 | 16 | 79 | 120 | 16 | 4,284 |
| 2004 | 0 | 140 | 616 | 2,344 | 724 | 0 | 68 | 102 | 0 | 3,993 |
| 2005 | 0 | 206 | 146 | 2,136 | 444 | 8 | 65 | 98 | 8 | 3,110 |
|  |  |  |  |  |  |  |  |  |  |  |

Table A4. Real and nominal prices (dollars per ITQ bushel) for ocean quahogs landed by ITQ and Maine vessels. Real prices are 1991 dollars. Information for ITQ vessels from dealer data. Information for Maine vessels from MAFMC (2005). Price data for Maine vessels (originally prices for Maine bushel) were converted to prices per ITQ bushel). Adjustments for inflation from the US Bureau of Labor Statistics for unprocessed shellfish. ${ }^{\text {a }}$

| ITQ |  |  | Maine |  |
| :---: | :---: | :---: | :---: | :---: |
| Year | Nominal | Real | Nominal | Real |
| 1994 | $\$ 4.44$ | $\$ 4.20$ |  |  |
| 1995 | $\$ 4.30$ | $\$ 3.56$ |  |  |
| 1996 | $\$ 4.12$ | $\$ 3.40$ |  |  |
| 1997 | $\$ 4.13$ | $\$ 2.39$ |  |  |
| 1998 | $\$ 4.23$ | $\$ 2.41$ |  |  |
| 1999 | $\$ 4.24$ | $\$ 2.53$ |  |  |
| 2000 | $\$ 4.35$ | $\$ 2.55$ |  |  |
| 2001 | $\$ 5.54$ | $\$ 3.23$ |  |  |
| 2002 | $\$ 5.47$ | $\$ 3.33$ |  |  |
| 2003 | $\$ 5.37$ | $\$ 3.08$ | $\$ 61.73$ | $\$ 35.43$ |
| 2004 | $\$ 5.26$ | $\$ 3.02$ | $\$ 59.55$ | $\$ 34.17$ |

Table A5. Ocean quahog fishing effort (hours fished) by stock assessment region in the US EEZ based on logbook data. Figures for 1983-2003 differ from NEFSC (2003) because additional other/unknown ("UNK") trips were allocated to region and because data for subtrips (deliveries from the same trip to different dealers) were counted only once. Data for 2005 are preliminary and may be incomplete. Based on Maine reports, UNK amounts during 2002 were probably from Maine waters.

|  |  |  |  |  |  |  |  |  | SI |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| YEAR | SVA | DMV | NJ |  | SNE | GBK | MNE | UNK | Grand <br> Total |
| 1983 | 0 | 7,131 | 13,932 | 50 | 1,535 | 0 | 0 | 56 | 22,704 |
| 1984 | 15 | 11,096 | 15,488 | 0 | 2,523 | 0 | 0 | 1,231 | 30,353 |
| 1985 | 204 | 10,058 | 17,890 | 87 | 2,066 | 0 | 0 | 2,955 | 33,260 |
| 1986 | 0 | 12,260 | 14,350 | 361 | 1,145 | 0 | 0 | 1,012 | 29,127 |
| 1987 | 0 | 15,812 | 14,704 | 806 | 1,340 | 0 | 0 | 49 | 32,711 |
| 1988 | 64 | 19,100 | 11,598 | 615 | 1,639 | 0 | 0 | 64 | 33,079 |
| 1989 | 0 | 12,124 | 24,262 | 797 | 2,327 | 0 | 0 | 50 | 39,560 |
| 1990 | 25 | 8,166 | 29,327 | 1,283 | 1,838 | 0 | 286 | 0 | 40,924 |
| 1991 | 0 | 12,048 | 30,397 | 1,844 | 1,433 | 0 | 17,110 | 0 | 62,832 |
| 1992 | 0 | 5,513 | 15,998 | 13,148 | 1,964 | 0 | 13,424 | 0 | 50,047 |
| 1993 | 0 | 4,622 | 25,457 | 12,883 | 1,783 | 0 | 5,720 | 0 | 50,465 |
| 1994 | 0 | 2,260 | 20,543 | 19,165 | 2,082 | 0 | 5,056 | 57 | 49,162 |
| 1995 | 0 | 1,621 | 13,598 | 16,015 | 8,561 | 0 | 5,731 | 0 | 45,526 |
| 1996 | 0 | 1,521 | 9,340 | 10,238 | 11,866 | 0 | 8,404 | 54 | 41,422 |
| 1997 | 0 | 2,742 | 9,382 | 8,295 | 13,515 | 0 | 11,734 | 0 | 45,669 |
| 1998 | 0 | 3,225 | 6,983 | 10,509 | 10,639 | 0 | 11,631 | 79 | 43,066 |
| 1999 | 0 | 2,595 | 7,623 | 9,132 | 12,258 | 0 | 10,821 | 90 | 42,518 |
| 2000 | 0 | 2,517 | 7,966 | 7,071 | 10,542 | 63 | 12,215 | 612 | 40,986 |
| 2001 | 0 | 2,170 | 10,844 | 7,813 | 11,404 | 22 | 13,113 | 1,454 | 46,820 |
| 2002 | 0 | 4,290 | 6,683 | 11,605 | 7,797 | 0 | 16,779 | 85 | 47,240 |
| 2003 | 0 | 2,617 | 10,764 | 16,099 | 4,596 | 0 | 17,832 | 108 | 52,016 |
| 2004 | 0 | 2,476 | 7,953 | 14,478 | 6,665 | 0 | 19,013 | 0 | 50,586 |
| 2005 | 0 | 3,500 | 1,935 | 12,437 | 4,019 | 0 | 16,572 | 129 | 38,591 |

Table A6. Commercial landings per unit effort (LPUE) for ocean quahog by region.
Figures for Maine are for vessels in ton class groups 1-2 (1-50 GRT). Figures for all other regions are for vessels in ton class groups 3-4 (51-500 GRT). "Nominal Mean LPUE" is the simple average of LPUE for each trip in the region during the year. "Total Bushels / Total Hours" is total landings divided by total hours fished. "Standardized Index" is back-transformed year effects from a general linear model with year, month and vessel effects. The standardized indices are adjusted to the LPUE level of a single randomly chosen vessel (ton class 4 for the EEZ and ton class 1 for Maine) during June of each year. Data for 2005 are preliminary and may be incomplete.

| Year | DMV |  |  |  | NJ |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Nominal Mean LPUE | Total Bushels / Total Hours | Standardized Index | CV | Nominal Mean LPUE | Total Bushels / Total Hours | Standardized Index | CV |
| 1980 | 153 | 139 | 165 | 0.15 | 119 | 118 | 113 | 0.19 |
| 1981 | 149 | 140 | 159 | 0.15 | 122 | 118 | 113 | 0.19 |
| 1982 | 151 | 143 | 176 | 0.15 | 135 | 130 | 120 | 0.19 |
| 1983 | 175 | 167 | 201 | 0.15 | 138 | 131 | 124 | 0.19 |
| 1984 | 154 | 142 | 181 | 0.15 | 133 | 126 | 119 | 0.19 |
| 1985 | 167 | 158 | 192 | 0.15 | 140 | 132 | 124 | 0.19 |
| 1986 | 157 | 148 | 169 | 0.15 | 144 | 139 | 125 | 0.19 |
| 1987 | 159 | 147 | 158 | 0.15 | 136 | 136 | 116 | 0.19 |
| 1988 | 144 | 135 | 141 | 0.15 | 137 | 133 | 110 | 0.19 |
| 1989 | 127 | 117 | 131 | 0.15 | 133 | 128 | 105 | 0.19 |
| 1990 | 106 | 99 | 118 | 0.15 | 123 | 117 | 95 | 0.19 |
| 1991 | 94 | 89 | 102 | 0.15 | 110 | 106 | 82 | 0.19 |
| 1992 | 100 | 95 | 104 | 0.15 | 101 | 96 | 84 | 0.19 |
| 1993 | 105 | 93 | 105 | 0.15 | 95 | 88 | 75 | 0.19 |
| 1994 | 104 | 97 | 97 | 0.15 | 80 | 74 | 68 | 0.19 |
| 1995 | 102 | 95 | 91 | 0.16 | 93 | 87 | 79 | 0.19 |
| 1996 | 119 | 107 | 101 | 0.16 | 121 | 115 | 100 | 0.19 |
| 1997 | 93 | 86 | 90 | 0.15 | 105 | 99 | 86 | 0.19 |
| 1998 | 100 | 93 | 92 | 0.15 | 109 | 85 | 75 | 0.19 |
| 1999 | 96 | 93 | 88 | 0.15 | 95 | 88 | 80 | 0.19 |
| 2000 | 98 | 92 | 86 | 0.15 | 96 | 92 | 82 | 0.19 |
| 2001 | 90 | 91 | 76 | 0.16 | 98 | 93 | 80 | 0.19 |
| 2002 | 93 | 88 | 83 | 0.15 | 94 | 91 | 77 | 0.19 |
| 2003 | 77 | 74 | 68 | 0.15 | 79 | 74 | 63 | 0.19 |
| 2004 | 66 | 56 | 60 | 0.16 | 88 | 77 | 67 | 0.19 |
| 2005 | 61 | 59 | 56 | 0.15 | 80 | 76 | 64 | 0.18 |

Table A6 (continued).

| Year | LI |  |  |  | SNE |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Nominal Mean LPUE | Total Bushels / Total Hours | Standardized Index | CV | Nominal Mean LPUE | Total Bushels / Total Hours | Standardized Index | CV |
| 1980 |  |  |  |  |  |  |  |  |
| 1981 | 123 | 123 |  |  |  |  |  |  |
| 1982 |  |  |  |  |  |  |  |  |
| 1983 | 91 | 93 |  |  | 91 | 90 |  |  |
| 1984 |  |  |  |  | 73 | 72 | 73 | 0.17 |
| 1985 | 106 | 102 |  |  | 75 | 74 | 79 | 0.18 |
| 1986 | 262 | 242 | 267 | 0.23 | 115 | 109 | 114 | 0.17 |
| 1987 | 322 | 323 | 319 | 0.20 | 122 | 115 | 117 | 0.17 |
| 1988 | 232 | 230 | 210 | 0.22 | 114 | 113 | 113 | 0.17 |
| 1989 | 176 | 167 | 190 | 0.21 | 127 | 113 | 118 | 0.17 |
| 1990 | 180 | 127 | 221 | 0.23 | 129 | 112 | 136 | 0.17 |
| 1991 | 205 | 200 | 212 | 0.18 | 135 | 133 | 134 | 0.17 |
| 1992 | 207 | 200 | 227 | 0.15 | 119 | 128 | 164 | 0.17 |
| 1993 | 159 | 148 | 174 | 0.15 | 115 | 126 | 179 | 0.17 |
| 1994 | 152 | 138 | 161 | 0.15 | 100 | 101 | 142 | 0.17 |
| 1995 | 145 | 131 | 159 | 0.15 | 145 | 139 | 119 | 0.17 |
| 1996 | 136 | 128 | 149 | 0.16 | 164 | 155 | 137 | 0.17 |
| 1997 | 144 | 137 | 157 | 0.16 | 156 | 146 | 126 | 0.17 |
| 1998 | 155 | 144 | 160 | 0.16 | 147 | 140 | 120 | 0.17 |
| 1999 | 165 | 153 | 172 | 0.16 | 126 | 119 | 106 | 0.17 |
| 2000 | 156 | 148 | 163 | 0.16 | 109 | 106 | 99 | 0.17 |
| 2001 | 165 | 161 | 177 | 0.16 | 93 | 91 | 88 | 0.17 |
| 2002 | 182 | 173 | 178 | 0.15 | 122 | 110 | 122 | 0.17 |
| 2003 | 169 | 160 | 168 | 0.15 | 116 | 104 | 106 | 0.17 |
| 2004 | 179 | 162 | 166 | 0.15 | 115 | 109 | 106 | 0.17 |
| 2005 | 177 | 172 | 151 | 0.06 | 113 | 111 | 108 | 0.17 |

Table A6 (continued).

| Nominal <br> Mean <br> LPUE |  |  |  | Total <br> Bushels <br> ITotal <br> Hours |
| :--- | :--- | :--- | :--- | :--- | | Standardized |
| :---: |
| Index |$\quad$ CV

Table A7. Trends in survey, stock and fishable abundance and biomass for ocean quahog $\geq 50 \mathrm{~mm}$ SL during 1982-2005 based on

| Region | Year | Survey |  |  |  | Stock |  |  |  | Fishable |  |  |  | Number Tows | Number Positive Tows | Number Strata With Data |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | N/Tow |  | KG/Tow | CV | N/Tow |  | KG/Tow | CV | N/Tow |  | KG/Tow | CV |  |  |  |
| SVA | 1982 | 0.039 | 0.00 | 0.002 | 0.00 | 0.039 | 0.00 | 0.002 | 0.00 | 0.038 | 0.00 | 0.002 | 0.00 | 5 | 1 | 2 |
| SVA | 1983 | 1.892 | 0.58 | 0.099 | 0.58 | 1.917 | 0.58 | 0.101 | 0.58 | 1.854 | 0.58 | 0.097 | 0.58 | 10 | 3 | 2 |
| SVA | 1984 | 0.189 | 0.85 | 0.010 | 0.87 | 0.191 | 0.84 | 0.010 | 0.87 | 0.185 | 0.85 | 0.010 | 0.87 | 14 | 2 | 2 |
| SVA | 1986 | 0.285 | 0.00 | 0.013 | 0.00 | 0.294 | 0.00 | 0.013 | 0.00 | 0.275 | 0.00 | 0.012 | 0.00 | 9 | 1 | 2 |
| SVA | 1989 | 0.392 | 0.00 | 0.018 | 0.00 | 0.401 | 0.00 | 0.019 | 0.00 | 0.380 | 0.00 | 0.018 | 0.00 | 9 | 1 | 2 |
| SVA | 1992 | 0.000 | . | 0.000 | . | 0.000 | 0.00 | 0.000 | 0.00 | 0.000 | 0.00 | 0.000 | 0.00 | 9 | 0 | 2 |
| SVA | 1997 | 0.154 | 0.00 | 0.004 | 0.00 | 0.282 | 0.00 | 0.006 | 0.00 | 0.132 | 0.00 | 0.003 | 0.00 | 9 | 1 | 2 |
| SVA | 1999 | 0.081 | 0.55 | 0.002 | 0.61 | 0.182 | 0.50 | 0.003 | 0.54 | 0.069 | 0.56 | 0.002 | 0.61 | 19 | 2 | 2 |
| SVA | 2002 | 0.045 | 1.00 | 0.001 | 1.00 | 0.133 | 1.00 | 0.002 | 1.00 | 0.037 | 1.00 | 0.001 | 1.00 | 10 | 1 | 2 |
| SVA | 2005 | 0.000 | . | 0.000 |  | 0.000 | 0.00 | 0.000 | 0.00 | 0.000 | 0.00 | 0.000 | 0.00 | 9 | 0 | 2 |
| DMV | 1982 | 79.162 | 0.32 | 2.956 | 0.34 | 86.645 | 0.31 | 3.156 | 0.33 | 73.837 | 0.32 | 2.786 | 0.34 | 59 | 24 | 6 |
| DMV | 1983 | 86.228 | 0.49 | 2.549 | 0.42 | 106.611 | 0.52 | 2.988 | 0.45 | 76.158 | 0.48 | 2.301 | 0.41 | 54 | 28 | 6 |
| DMV | 1984 | 52.011 | 0.35 | 1.667 | 0.30 | 63.193 | 0.36 | 1.904 | 0.31 | 46.650 | 0.34 | 1.530 | 0.30 | 78 | 34 | 6 |
| DMV | 1986 | 75.681 | 0.23 | 2.532 | 0.22 | 86.737 | 0.24 | 2.800 | 0.22 | 68.939 | 0.23 | 2.342 | 0.22 | 61 | 28 | 6 |
| DMV | 1989 | 64.366 | 0.58 | 1.801 | 0.46 | 82.482 | 0.62 | 2.179 | 0.51 | 55.961 | 0.55 | 1.606 | 0.44 | 69 | 31 | 6 |
| DMV | 1992 | 71.982 | 0.36 | 2.285 | 0.31 | 85.405 | 0.40 | 2.589 | 0.33 | 64.676 | 0.35 | 2.093 | 0.30 | 69 | 25 | 6 |
| DMV | 1997 | 47.743 | 0.21 | 1.669 | 0.21 | 56.440 | 0.22 | 1.847 | 0.21 | 43.721 | 0.21 | 1.557 | 0.21 | 73 | 28 | 6 |
| DMV | 1999 | 28.359 | 0.29 | 0.948 | 0.27 | 33.388 | 0.29 | 1.056 | 0.27 | 25.821 | 0.29 | 0.878 | 0.26 | 70 | 23 | 6 |
| DMV | 2002 | 31.814 | 0.25 | 1.106 | 0.23 | 38.774 | 0.26 | 1.232 | 0.23 | 29.139 | 0.24 | 1.032 | 0.22 | 71 | 19 | 6 |
| DMV | 2005 | 19.407 | 0.49 | 0.694 | 0.53 | 24.842 | 0.45 | 0.776 | 0.50 | 17.906 | 0.50 | 0.652 | 0.53 | 66 | 21 | 6 |

Table A7 (cont.)

| Region | Year | Survey |  |  |  | Stock |  |  |  | Fishable |  |  |  | Number <br> Tows | Number Positive Tows | Number Strata With Data |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | N/Tow | CV | KG/Tow | CV | N/Tow | CV | KG/Tow | CV | N/Tow | CV | KG/Tow | CV |  |  |  |
| NJ | 1982 | 112.339 | 0.20 | 3.555 | 0.20 | 129.333 | 0.20 | 3.918 | 0.20 | 102.545 | 0.20 | 3.302 | 0.20 | 100 | 50 | 13 |
| NJ | 1983 | 86.092 | 0.21 | 2.832 | 0.21 | 98.417 | 0.21 | 3.090 | 0.21 | 79.201 | 0.21 | 2.649 | 0.21 | 98 | 55 | 13 |
| NJ | 1984 | 143.533 | 0.24 | 4.531 | 0.24 | 165.861 | 0.24 | 4.998 | 0.24 | 131.075 | 0.24 | 4.208 | 0.24 | 153 | 80 | 13 |
| NJ | 1986 | 142.520 | 0.23 | 4.847 | 0.23 | 158.243 | 0.24 | 5.213 | 0.23 | 132.170 | 0.23 | 4.555 | 0.22 | 103 | 52 | 13 |
| NJ | 1989 | 73.510 | 0.22 | 2.193 | 0.21 | 90.578 | 0.21 | 2.491 | 0.21 | 66.320 | 0.22 | 2.020 | 0.21 | 110 | 52 | 13 |
| NJ | 1992 | 88.043 | 0.18 | 3.023 | 0.17 | 97.822 | 0.18 | 3.246 | 0.17 | 81.725 | 0.18 | 2.843 | 0.17 | 110 | 52 | 13 |
| NJ | 1997 | 122.262 | 0.15 | 4.273 | 0.15 | 135.780 | 0.16 | 4.576 | 0.15 | 113.720 | 0.15 | 4.028 | 0.15 | 124 | 59 | 13 |
| NJ | 1999 | 59.480 | 0.15 | 2.019 | 0.14 | 72.266 | 0.15 | 2.221 | 0.14 | 54.889 | 0.15 | 1.900 | 0.14 | 132 | 61 | 13 |
| NJ | 2002 | 89.793 | 0.23 | 3.229 | 0.24 | 101.123 | 0.22 | 3.456 | 0.23 | 83.825 | 0.24 | 3.059 | 0.24 | 127 | 60 | 13 |
| NJ | 2005 | 47.076 | 0.16 | 1.568 | 0.15 | 62.364 | 0.15 | 1.769 | 0.15 | 43.117 | 0.15 | 1.473 | 0.14 | 103 | 54 | 13 |
| LI | 1982 | 278.856 | 0.15 | 7.021 | 0.16 | 434.976 | 0.16 | 9.325 | 0.15 | 239.652 | 0.15 | 6.258 | 0.16 | 43 | 37 | 9 |
| LI | 1983 | 185.877 | 0.21 | 5.232 | 0.21 | 253.508 | 0.22 | 6.355 | 0.21 | 163.619 | 0.21 | 4.742 | 0.21 | 38 | 36 | 9 |
| LI | 1984 | 235.154 | 0.17 | 6.536 | 0.16 | 318.987 | 0.18 | 7.967 | 0.17 | 206.330 | 0.17 | 5.906 | 0.16 | 71 | 63 | 9 |
| LI | 1986 | 311.430 | 0.22 | 8.625 | 0.21 | 416.390 | 0.23 | 10.480 | 0.21 | 273.066 | 0.22 | 7.782 | 0.21 | 36 | 31 | 9 |
| LI | 1989 | 226.213 | 0.34 | 5.062 | 0.29 | 367.492 | 0.38 | 7.152 | 0.33 | 190.104 | 0.33 | 4.384 | 0.28 | 40 | 36 | 9 |
| LI | 1992 | 323.335 | 0.18 | 8.313 | 0.16 | 465.234 | 0.20 | 10.625 | 0.17 | 279.032 | 0.17 | 7.401 | 0.16 | 42 | 36 | 9 |
| LI | 1997 | 401.643 | 0.16 | 11.156 | 0.16 | 518.847 | 0.17 | 13.351 | 0.16 | 353.149 | 0.16 | 10.049 | 0.16 | 42 | 35 | 9 |
| LI | 1999 | 232.273 | 0.17 | 6.280 | 0.15 | 310.519 | 0.19 | 7.671 | 0.16 | 202.716 | 0.17 | 5.628 | 0.14 | 45 | 41 | 9 |
| LI | 2002 | 253.059 | 0.21 | 6.969 | 0.20 | 330.414 | 0.21 | 8.385 | 0.20 | 222.209 | 0.21 | 6.268 | 0.20 | 43 | 40 | 9 |
| LI | 2005 | 151.233 | 0.18 | 4.122 | 0.19 | 218.396 | 0.19 | 5.121 | 0.18 | 132.758 | 0.18 | 3.717 | 0.19 | 45 | 39 | 9 |

Table A7 (cont.)

| Region | Year | Survey |  |  |  | Stock |  |  |  | Fishable |  |  |  | Number Tows | Number Positive Tows | Number Strata With Data |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | N/Tow | CV | KG/Tow | CV | N/Tow | CV | KG/Tow | CV | N/Tow | CV | KG/Tow | CV |  |  |  |
| SNE | 1982 | 277.607 | 0.27 | 6.981 | 0.25 | 345.845 | 0.28 | 8.222 | 0.26 | 245.458 | 0.27 | 6.283 | 0.25 | 48 | 30 | 10 |
| SNE | 1983 | 173.213 | 0.29 | 4.163 | 0.30 | 237.689 | 0.31 | 5.149 | 0.29 | 151.399 | 0.29 | 3.727 | 0.30 | 58 | 37 | 10 |
| SNE | 1984 | 188.458 | 0.27 | 4.753 | 0.29 | 234.355 | 0.26 | 5.588 | 0.28 | 166.802 | 0.27 | 4.280 | 0.29 | 69 | 38 | 10 |
| SNE | 1986 | 289.151 | 0.31 | 6.961 | 0.31 | 394.360 | 0.35 | 8.561 | 0.32 | 253.117 | 0.31 | 6.226 | 0.31 | 27 | 23 | 9 |
| SNE | 1989 | 274.664 | 0.19 | 6.707 | 0.18 | 353.181 | 0.21 | 8.050 | 0.19 | 241.358 | 0.19 | 6.003 | 0.18 | 34 | 29 | 10 |
| SNE | 1992 | 333.079 | 0.19 | 8.634 | 0.19 | 400.104 | 0.19 | 9.947 | 0.19 | 297.003 | 0.19 | 7.814 | 0.20 | 36 | 31 | 10 |
| SNE | 1997 | 292.893 | 0.54 | 6.128 | 0.45 | 447.963 | 0.61 | 8.405 | 0.52 | 246.944 | 0.52 | 5.335 | 0.43 | 39 | 27 | 10 |
| SNE | 1999 | 252.431 | 0.54 | 6.169 | 0.48 | 312.910 | 0.56 | 7.316 | 0.51 | 221.840 | 0.53 | 5.510 | 0.47 | 39 | 30 | 10 |
| SNE | 2002 | 180.674 | 0.22 | 5.103 | 0.22 | 206.737 | 0.22 | 5.663 | 0.22 | 164.245 | 0.22 | 4.697 | 0.22 | 29 | 28 | 9 |
| SNE | 2005 | 178.281 | 0.28 | 3.944 | 0.24 | 395.499 | 0.44 | 5.882 | 0.29 | 154.795 | 0.27 | 3.547 | 0.24 | 29 | 25 | 7 |
| GBK | 1986 | 276.488 | 0.19 | 6.207 | 0.18 | 427.632 | 0.23 | 8.633 | 0.20 | 232.206 | 0.19 | 5.313 | 0.18 | 48 | 21 | 16 |
| GBK | 1989 | 90.805 | 0.26 | 2.371 | 0.26 | 124.548 | 0.25 | 2.950 | 0.25 | 78.933 | 0.26 | 2.098 | 0.26 | 79 | 38 | 16 |
| GBK | 1992 | 346.253 | 0.21 | 9.225 | 0.21 | 485.713 | 0.19 | 11.427 | 0.20 | 302.841 | 0.21 | 8.208 | 0.21 | 74 | 41 | 16 |
| GBK | 1997 | 269.762 | 0.19 | 7.058 | 0.19 | 389.377 | 0.19 | 8.969 | 0.18 | 234.251 | 0.19 | 6.274 | 0.19 | 83 | 44 | 18 |
| GBK | 1999 | 273.398 | 0.17 | 7.806 | 0.19 | 365.971 | 0.16 | 9.391 | 0.18 | 241.903 | 0.17 | 7.060 | 0.19 | 77 | 47 | 18 |
| GBK | 2002 | 328.367 | 0.18 | 9.059 | 0.19 | 478.136 | 0.15 | 11.247 | 0.18 | 288.963 | 0.18 | 8.149 | 0.19 | 61 | 38 | 15 |

Table A8. Number of random and nearly random NEFSC survey tows used to estimate trends in abundance of ocean quahog. Figures in each cell are the number of tows in calculations for each combination of stratum and cruise. Figures in plain text are the number of original tows (without borrowing). Bold and outlined figures are for cells with zero tows originally that were filled by borrowing tows from the same strata during previous and/or subsequent cruises. Black cells are for cells with zero tows that could not be filled by borrowing. Note that there were too few tows in GBK during 1982-1984 and 2005 to calculate abundance indices for GBK during these years.

| Region | Stratum | Survey Year |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1982 | 1983 | 1984 | 1986 | 1989 | 1992 | 1994 | 1997 | 1999 | 2002 | 2005 |
| SVA | 5 | 4 | 9 | 13 | 8 | 8 | 8 | 8 | 8 | 16 | 8 | 8 |
|  | 6 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 3 | 2 | 2 |
| DMV | 9 | 30 | 26 | 35 | 29 | 37 | 37 | 39 | 39 | 38 | 39 | 39 |
|  | 10 | 2 | 2 | 3 | 3 | 3 | 3 | 3 | 3 | 3 | 3 | 3 |
|  | 11 | 2 | 2 | 4 | 2 | 2 | 2 | 2 | 2 | 2 | 2 | 2 |
|  | 13 | 19 | 18 | 25 | 20 | 20 | 20 | 21 | 22 | 19 | 20 | 20 |
|  | 14 | 2 | 2 | 3 | 3 | 3 | 3 | 5 | 3 | 3 | 3 | 3 |
|  | 15 | 4 | 4 | 8 | 4 | 4 | 4 | 5 | 4 | 5 | 4 | 4 |
| NJ | 17 | 11 | 11 | 18 | 12 | 12 | 12 | 12 | 14 | 12 | 12 | 12 |
|  | 18 | 3 | 3 | 6 | 3 | 3 | 3 | 3 | 3 | 3 | 3 | 3 |
|  | 19 | 3 | 3 | 6 | 3 | 3 | 3 | 3 | 3 | 3 | 3 | 3 |
|  | 21 | 18 | 18 | 22 | 19 | 20 | 20 | 23 | 26 | 39 | 29 | 29 |
|  | 22 | 3 | 3 | 6 | 3 | 3 | 3 | 5 | 3 | 3 | 3 | 3 |
|  | 23 | 7 | 6 | 11 | 5 | 4 | 5 | 5 | 5 | 5 | 5 | 5 |
|  | 25 | 9 | 9 | 13 | 8 | 9 | 9 | 9 | 12 | 8 | 9 | 9 |
|  | 26 | 2 | 2 | 5 | 3 | 3 | 3 | 3 | 3 | 3 | 3 | 3 |
|  | 27 | 4 | 4 | 8 | 4 | 4 | 4 | 4 | 4 | 4 | 4 | 4 |
|  | 87 | 8 | 7 | 10 | 9 | 9 | 9 | 9 | 9 | 9 | 16 | 16 |
|  | 88 | 15 | 15 | 24 | 17 | 20 | 20 | 20 | 21 | 22 | 20 | 20 |
|  | 89 | 15 | 15 | 21 | 15 | 18 | 17 | 17 | 19 | 18 | 18 | 18 |
|  | 90 | 2 | 2 | 3 | 2 | 2 | 2 | 2 | 2 | 2 | 2 | 2 |
| LI | 29 | 11 | 10 | 20 | 10 | 10 | 10 | 10 | 10 | 11 | 10 | 10 |
|  | 30 | 7 | 8 | 14 | 6 | 6 | 6 | 6 | 6 | 7 | 6 | 6 |
|  | 31 | 9 | 7 | 12 | 5 | 7 | 8 | 8 | 8 | 9 | 8 | 8 |
|  | 33 | 4 | 4 | 8 | 4 | 4 | 4 | 5 | 4 | 4 | 4 | 4 |
|  | 34 | 2 | 2 | 4 | 2 | 2 | 2 | 5 | 2 | 2 | 2 | 2 |
|  | 35 | 4 | 2 | 4 | 2 | 5 | 6 | 6 | 6 | 6 | 6 | 6 |
|  | 91 | 3 | 2 | 4 | 4 | 3 | 3 | 3 | 3 | 3 | 3 | 3 |
|  | 92 | 2 | 2 | 3 | 2 | 2 | 2 | 2 | 2 | 2 | 2 | 2 |
|  | 93 | 1 | 1 | 2 | 1 | 1 | 1 | 1 | 1 | 1 | 2 | 2 |

Table A8 (continued).

| Region | Stratum | Survey Year |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1982 | 1983 | 1984 | 1986 | 1989 | 1992 | 1994 | 1997 | 1999 | 2002 | 2005 |
| SNE | 37 | 7 | 4 | 7 | 3 | 6 | 3 | 5 | 4 | 4 | 3 | 3 |
|  | 38 | 3 | 2 | 5 | 3 | 3 | 3 | 5 | 3 | 3 | 3 | 3 |
|  | 39 | 6 | 4 | 6 | 2 | 5 | 5 | 5 | 5 | 5 | 5 | 5 |
|  | 41 | 6 | 5 | 7 | 5 | 6 | 6 | 6 | 6 | 5 | 6 | 6 |
|  | 45 | 3 | 7 | 9 | 4 | 4 | 4 | 4 | 4 | 4 | 3 | 3 |
|  | 46 | 2 | 5 | 5 | 3 | 2 | 3 | 5 | 3 | 3 | 2 | 2 |
|  | 47 | 4 | 3 | 4 | 2 | 2 | 4 | 5 | 4 | 3 | 1 | 1 |
|  | 94 | 1 | 2 | 2 |  | 1 | 1 | 2 | 2 | 4 | 2 | 2 |
|  | 95 | 4 | 14 | 11 | 4 | 4 | 4 | 4 | 4 | 4 | 4 | 4 |
|  | 96 | 12 | 12 | 13 | 1 | 1 | 3 | 2 | 4 | 4 |  |  |
| GBK | 54 |  | 3 | 3 | 3 | 6 | 3 | 3 | 3 | 3 | 0 | 0 |
|  | 55 | 3 | 3 | 3 | 3 | 1 | 3 | 3 | 3 | 2 | 2 | 2 |
|  | 56 |  |  |  |  |  |  |  | 4 | 4 | 4 |  |
|  | 57 |  |  | 2 | 2 | 1 | 2 | 5 | 2 | 2 | 2 | 2 |
|  | 58 |  |  |  |  |  |  |  | 5 | 5 | 5 |  |
|  | 59 | 1 | 4 | 5 | 1 | 2 | 6 | 5 | 5 | 4 | 5 | 5 |
|  | 60 |  |  | 2 | 2 | 2 | 4 | 2 | 5 | 5 | 5 | 5 |
|  | 61 | 8 | 1 | 6 | 5 | 12 | 7 | 6 | 6 | 6 | 6 | 6 |
|  | 62 |  |  | 1 | 1 | 1 | 4 | 4 | 4 | 4 | 4 | 4 |
|  | 65 |  |  | 3 | 3 | 5 | 2 | 2 | 3 | 4 | 1 | 1 |
|  | 67 |  | 5 | 5 | 5 | 7 | 7 | 7 | 7 | 7 |  |  |
|  | 68 | 1 | 8 | 7 | 3 | 6 | 6 | 5 | 5 | 5 |  |  |
|  | 69 | 2 | 5 | 11 | 6 | 6 | 6 | 7 | 6 | 7 | 7 |  |
|  | 70 | 1 | 2 | 6 | 4 | 8 | 4 | 4 | 4 | 3 | 2 | 2 |
|  | 71 |  | 2 | 2 | 3 | 1 | 2 | 3 | 3 | 1 | 2 | 2 |
|  | 72 | 2 | 10 | 8 | 1 | 8 | 8 | 8 | 8 | 6 | 6 |  |
|  | 73 | 1 | 1 | 4 | 3 | 6 | 6 | 6 | 6 | 5 | 6 | 6 |
|  | 74 | 3 | 4 | 1 | 3 | 7 | 4 | 4 | 4 | 3 | 3 | 3 |

Table A9. Parameter estimates for the relationship between shell length ( $L, \mathrm{~mm}$ ) and drained (fresh, not frozen) meat weight ( $W, \mathrm{~g}$ ) in ocean quahog (NEFSC 2004). The equation for the relationship is $W=e^{\alpha} L^{\beta}$.

| Region | Alpha | Beta |
| :---: | :---: | :---: |
| SVA | -9.042313 | 2.787987 |
| DMV | -9.042313 | 2.787987 |
| NJ | -9.847183 | 2.949540 |
| LI | -9.233646 | 2.822474 |
| SNE | -9.124283 | 2.774989 |
| GBK | -8.969073 | 2.767282 |

Table A10. Clam survey database parameters used to extract survey data for ocean quahog in this assessment. Parameters were the same for all regions. Negative parameter values are ignored in database calculations.

| Database Parameter | Survey length composition | $\begin{aligned} & \text { Trends } \\ & <70 \\ & \mathrm{~mm} \text { SL } \end{aligned}$ | Trends in survey, stock and fishable biomass | Efficiency corrected swept-area biomass |
| :---: | :---: | :---: | :---: | :---: |
| DISTANCE_TYPE | TREND | TREND | TREND | SENSORS |
| USEINCHESDOWN | 1 | 1 | 1 | 1 |
| LENGTH_BIN_SIZE_MM | 10 | 1000 | 1000 | 1000 |
| FIRST_LENGTH_MM | 1 | 0 | 50 | 50 |
| FIRST_BIN_IS_PLUSGROUP | -1 | -1 | -1 | -1 |
| LAST_LENGTH_MM | 250 | 69 | 250 | 250 |
| LAST_BIN_IS_PLUSGROUP | -1 | -1 | -1 | -1 |
| SVSPP_TO_USE | 409 | 409 | 409 | 409 |
| AREȦKIND | GIS | GIS | GIS | GIS |
| REV_DATE_FOR_AREAS | 2002 | 2002 | 2002 | 2002 |
| REV_DATE_FOR_LW | 2000 | 2000 | 2000 | 2000 |
| FIRST_JWSTCODE | -1 | -1 | -1 | -1 |
| LAST_JWSTCODE | -1 | -1 | -1 | -1 |
| FIRST__RANDLIKE | 1 | 1 | 1 | 1 |
| LAST_RANDLIKE | 2 | 2 | 2 | 2 |
| FIRST_STATION | -1 | -1 | -1 | -1 |
| LAST_STATION | -1 | -1 | -1 | -1 |
| FIRST_HAUL | 1 | 1 | 1 | 1 |
| LAST_HAUL | 3 | 3 | 3 | 3 |
| FIRST_GEARCOND | 1 | 1 | 1 | 1 |
| LAST_GEARCOND | 6 | 6 | 6 | 6 |
| FIRST_STRATUM | -1 | -1 | -1 | -1 |
| LAST_STRATUM | -96 | -96 | -96 | -96 |
| FIRST_REGION_CODE | 1 | 1 | 1 | 1 |
| LAST_REGION_CODE | 6 | 6 | 6 | 6 |
| WRITE_TOW_DATA | 1 | 1 | 1 | 1 |
| WRITE_STRATUM_DATA | 1 | 1 | 1 | 1 |
| FIRST_CRUISE | -199700 | -199700 | -199700 | 199700 |
| LAST_CRUISE | -200509 | -200509 | -200509 | 200509 |
| SurvSelxAlpha | 8.122 | 8.122 | 8.122 | 8.122 |
| SurvSelxBeta |  | -0.119 | -0.119 | -0.119 |
| FisherySelxAlpha |  | 7.63 | 7.63 | 7.63 |
| FisherySelxBeta |  | -0.105 | -0.105 | -0.105 |
| NOMINAL_TOW_DISTANCE_NM |  | 0.15 | 0.15 | 0.15 |
| MINVALIDDOPPLER ${ }^{-}$ |  | 0.04 | 0.04 | 0.04 |
| MAXVALIDDOPPLER |  | 0.3 | 0.3 | 0.3 |
| FILLHOLZ |  | 1 | 1 | 1 |

Table A11. Patch model estimates of efficiency for commercial and NEFSC survey clam dredges based on depletion experiments during 1997-2005. "NA" means not available.

| Study area |  |  |  |  |  | Depletion Tows |  |  |  |  |  | Patch Model |  |  |  |  |  |  | Setup Tows (if applicable) |  |  | NEFSC Dredge Efficiency | Foornotes |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Experiment | Region | Latitude (decimal degrees) | Longitude (decimal degrees) | Depth (m) | Mean Sediment Size (microns) | Depletion Study Vessel | $\begin{aligned} & \text { Depletion } \\ & \text { Date } \end{aligned}$ | Ship Position Data (source / nominal accuracy / time interval) | N tows used | N Bushel Counts / Length samples | $\begin{aligned} & \text { Depletion } \\ & \text { Vessel } \\ & \text { Blade } \\ & \text { Width (ft) } \end{aligned}$ | Cell Size <br> (ft) | Density $\left(N \mathrm{ft}^{-2}\right)$ | Depletion Vessel Efficiency | k | $\gamma$ | Neg. Log likelihoo d | Fit to Catch Data (R2s) | Setup Date | Setup or RV Depletion Stations | Density $\left(\mathrm{Nft}^{-2}\right)$ |  |  |
| OQ2005-1 | LI | 40.5190 | 72.0762 | 57 | 536 | $\begin{gathered} \hline \text { F/V Lisa } \\ \text { Kim } \end{gathered}$ | Sep-05 | $\begin{gathered} \text { GPS } / 6 \mathrm{ft} / 6 \\ \mathrm{sec} \end{gathered}$ | 20 | 4/4 | 10 | 20 | 0.073 | 0.183 | 1.97 | 0.50 | 127.0 | Ok | Jun-05 | $\begin{gathered} 165,231- \\ 234 \end{gathered}$ | 0.0120 | 0.165 | 1 |
| OQ2005-2 | LI | 40.3896 | 72.3895 | 53 | 438 | $\underset{\text { Kim }}{\text { F/V Lisa }}$ | Sep-05 | $\begin{gathered} \mathrm{GPS} / 6 \mathrm{ft} / 6 \\ \mathrm{sec} \end{gathered}$ | 21 | 4/4 | 10 | 20 | 0.047 | 0.402 | 8.57 | 0.50 | 131.8 | Ok | Jun-05 | $\begin{gathered} 162,235- \\ 238 \end{gathered}$ | 0.0080 | 0.169 | 1 |
| OQ2005-3 | LI | 40.6422 | 72.6517 | 35 | 267 | $\begin{aligned} & \text { F/V Lisa } \\ & \text { Kim } \end{aligned}$ | Sep-05 | $\begin{gathered} \mathrm{GPS} / 6 \mathrm{ft} / 6 \\ \mathrm{sec} \end{gathered}$ | 20 | 4/4 | 10 | 20 | 0.085 | 0.733 | 9.57 | 0.50 | 125.9 | Ok | Jun-05 | 3, 239-242 | 0.0101 | 0.119 | 1 |
| OQ2005-4 | LI | 40.6882 | 72.1815 | 46 | 308 | $\begin{aligned} & \text { F/V Lisa } \\ & \mathrm{Kim} \end{aligned}$ | Sep-05 | $\begin{gathered} \mathrm{GPS} / 6 \mathrm{ft} / 6 \\ \mathrm{sec} \end{gathered}$ | 17 | 4/4 | 10 | 20 | 0.027 | 0.815 | 12.31 | 0.50 | 89.4 | Ok | Jun-05 | $\begin{gathered} 168,243- \\ 246 \end{gathered}$ | 0.0042 | 0.154 | 1 |
| OQ2005-6 | LI | 40.0555 | 72.4167 | 65 | 554 | $\underset{\text { Kim }}{\text { F/V Lisa }}$ | Sep-05 | $\begin{gathered} \mathrm{GPS} / 6 \mathrm{ft} / 6 \\ \mathrm{sec} \end{gathered}$ | 20 | 4/4 | 10 | 20 | 0.137 | 0.660 | 2.55 | 0.50 | 146.3 | Ok | Jun-05 | 252-256 | 0.0210 | 0.153 | 1 |
| Mean CV for Mean |  |  |  | $\begin{gathered} 51 \\ 10 \% \end{gathered}$ | $\begin{aligned} & 421 \\ & 14 \% \end{aligned}$ |  |  |  | $\begin{gathered} 19.6 \\ 3 \% \end{gathered}$ |  |  |  | $\begin{gathered} 0.074 \\ 25 \% \end{gathered}$ | $\begin{gathered} 0.559 \\ 21 \% \end{gathered}$ | $\begin{aligned} & 6.99 \\ & 29 \% \end{aligned}$ |  |  |  |  |  | $\begin{gathered} 0.0110 \\ 25 \% \end{gathered}$ | $\begin{gathered} 0.152 \\ 6 \% \end{gathered}$ |  |
| $\begin{aligned} & \text { OQ2002-1 } \\ & \text { (LK-1) } \end{aligned}$ | LI | 40.7276 | 71.7373 | 60 | 331 | $\begin{aligned} & \hline \text { F/V Lisa } \\ & \text { Kim } \end{aligned}$ | Mar-02 | $\begin{gathered} \hline \mathrm{GPS} / 1 \mathrm{ft} / 6 \\ \mathrm{sec} \end{gathered}$ | 24 | 5/5 | 10 | 20 | 0.295 | 0.489 | 6.56 | 0.50 | 173.1 | Ok | Jun-02 | 5-9 | 0.0290 | 0.098 | 1,2,5 |
| $\begin{aligned} & \text { OQ2002-2 } \\ & \text { (LK-2) } \end{aligned}$ | LI | 40.1031 | 73.1911 | 48 | 277 | $\begin{aligned} & \text { F/V Lisa } \\ & \text { Kim } \end{aligned}$ | Mar-02 | $\begin{gathered} \mathrm{GPS} / 1 \mathrm{ft} / 6 \\ \mathrm{sec} \end{gathered}$ | 22 | 4/4 | 10 | 20 | 0.165 | 0.785 | 10.57 | 0.50 | 149.7 | Ok | Jun-02 | 25-29 | 0.0245 | 0.149 | 1,2 |
| $\begin{gathered} \text { OQ2002-3 } \\ (\mathrm{LK}-3) \end{gathered}$ | NJ | 38.8149 | 73.8133 | 50 | 195 | $\underset{\mathrm{Kim}}{\text { F/V Lisa }}$ | Mar-02 | $\begin{gathered} \text { GPS / } 1 \mathrm{ft} / 6 \\ \mathrm{sec} \end{gathered}$ | 20 | 4/4 | 10 | 20 | 0.081 | 0.777 | 11.57 | 0.50 | 133.4 | Ok | Jun-02 | 213-217 | 0.0239 | 0.297 | 1,2 |
| $\begin{gathered} \text { OQ2002-4 } \\ \text { (LK-4) } \end{gathered}$ | DMV | 37.8876 | 74.6449 | 48 | 135 | $\begin{aligned} & \text { F/V Lisa } \\ & \text { Kim } \end{aligned}$ | Mar-02 | $\begin{gathered} \mathrm{GPS} / 1 \mathrm{ft} / 6 \\ \mathrm{sec} \end{gathered}$ | 24 | 5/5 | 10 | 20 | 0.073 | 0.254 | 12.46 | 0.50 | 136.0 | Ok | Jun-02 | 272-276 | 0.0210 | 0.287 | 1, 2,9, 16 |
| Mean CV for Mean |  |  |  | $\begin{gathered} 52 \\ 6 \% \end{gathered}$ | $\begin{aligned} & 235 \\ & 18 \% \end{aligned}$ |  |  |  | $\begin{gathered} 22.5 \\ 4 \% \\ \hline \end{gathered}$ |  |  |  | $\begin{gathered} 0.153 \\ 34 \% \\ \hline \end{gathered}$ | $\begin{array}{r} 0.576 \\ 22 \% \\ \hline \end{array}$ | $\begin{gathered} 10.29 \\ 13 \% \end{gathered}$ |  |  |  |  |  | $\begin{gathered} 0.0704 \\ 65 \% \\ \hline \end{gathered}$ | $\begin{aligned} & 0.178 \\ & 0.274 \\ & \hline \end{aligned}$ |  |
| $\begin{gathered} \text { OQ2000-1 } \\ (\mathrm{JN}-1) \end{gathered}$ | LI | 40.6022 | 71.9875 | 58 | N/A | F/N John | Mar-00 | $\begin{gathered} \hline \text { GPS } / 1 \mathrm{ft} / \\ 30 \mathrm{sec} \end{gathered}$ | 22 | 5/5 | 12.5 | 25 | 0.100 | 0.730 | 5.55 | 0.50 | 157.4 | Ok | Jun-99 | 194-199 | NA | NA | $1,2,6$ |
| OQ2000-2 <br> (JN-2) | LI | 40.3945 | 72.5430 | 48 | N/A | $\begin{gathered} \text { FN John } \\ \mathrm{N} \end{gathered}$ | Mar-00 | $\begin{gathered} \mathrm{GPS} / 1 \mathrm{ft} / \\ 30 \mathrm{sec} \end{gathered}$ | 16 | 4/3 | 12.5 | 25 | 0.062 | 0.554 | 15.10 | 0.50 | 98.1 | Ok | Jun-99 | 178-180 | 0.0145 | 0.234 | $\begin{gathered} 1,2,7,11,12 \\ 17 \end{gathered}$ |
| $\begin{aligned} & \text { OQ2000-3 } \\ & (\mathrm{DM}-1) \end{aligned}$ | LI | 40.5830 | 72.7968 | 40 | N/A | F/V Danielle Maria | May-00 | $\begin{gathered} \mathrm{GPS} / 1 \mathrm{ft} / \\ 30 \mathrm{sec} \end{gathered}$ | 27 | 6/6 | 10 | 20 | 0.089 | 0.560 | 4.57 | 0.50 | 184.2 | Ok | Jun-99 | 3-8 | 0.0147 | 0.165 | $\begin{array}{\|c} 1,2,8,10,12 \\ 18 \end{array}$ |
| Mean CV for Mean |  |  |  |  |  |  |  |  | $\begin{aligned} & 21.7 \\ & 15 \% \\ & \hline \end{aligned}$ |  |  |  | $\begin{aligned} & 0.084 \\ & 14 \% \\ & \hline \end{aligned}$ | $\begin{gathered} 0.615 \\ 9 \% \\ \hline \end{gathered}$ | $\begin{aligned} & 8.405 \\ & 40 \% \\ & \hline \end{aligned}$ |  |  |  |  |  | $\begin{gathered} 0.0146 \\ 1 \% \\ \hline \end{gathered}$ | $\begin{aligned} & 0.199 \\ & 17 \% \\ & \hline \end{aligned}$ |  |
| $\begin{aligned} & \text { OQ1999-01 } \\ & \text { DE2 } \end{aligned}$ | LI | 40.6023 | 71.9848 | 57 | N/A | R/V Delaware II | Jun-99 | $\begin{gathered} \mathrm{GPS} / 36 \mathrm{ft} / \\ 1 \mathrm{sec} \end{gathered}$ | 60 | 8/8 | 5 | 10 | 0.007 | 0.990 | 4.05 | 0.25 | 253.1 | Poor |  | N/A |  | 0.990 | 14, 15 |
| $\begin{aligned} & \text { OQ1998-1 } \\ & (\mathrm{SH}-3) \end{aligned}$ | $\begin{gathered} \mathrm{LI} \\ \text { (Shinnecock) } \end{gathered}$ | 40.7665 | 72.1795 | 41 | N/A | F/N Cape Fear | $\begin{gathered} 3 / 1 / 11998 \\ 8 \end{gathered}$ | Loran / 40 ft / 30 sec . | 14 | 3/3 | 10 | 20 | 0.017 | 1.000 | 3.48 | 0.50 | 76.5 | Poor |  |  |  |  | 1,13 |
| $\begin{gathered} \text { OQ1998-2 } \\ (\mathrm{SH}-2) \end{gathered}$ | $\begin{gathered} \mathrm{Ll} \\ \text { (Shinnecock) } \end{gathered}$ | 40.7220 | 72.0075 | 45 | N/A | $\begin{aligned} & \text { F/N Cape } \\ & \text { Fear } \end{aligned}$ | Mar-98 | Loran / 40 ft / 30 sec . | 23 | 5/5 | 10 | 20 | 0.067 | 0.869 | 10.57 | 0.50 | 140.3 | Ok |  | NA |  | NA | 15 |
| $\begin{aligned} & \text { OQ1998-3 } \\ & \text { (NS-1) } \end{aligned}$ | SNE (Nantucket Shoals) | 40.4670 | 69.4830 | 63 | N/A | F/N Cape Fear | Apr-98 | $\begin{gathered} \text { Loran / } 40 \mathrm{ft} / \\ 30 \mathrm{sec} . \end{gathered}$ | 24 | 5/5 | 10 | 20 | 0.255 | 0.710 | 7.56 | 0.50 | 195.5 | Ok |  |  |  |  | 15 |
| $\begin{gathered} \text { Mean } \\ \text { CV for Mean } \\ \hline \end{gathered}$ |  |  |  |  |  |  |  |  | $\begin{aligned} & 20.3 \\ & 16 \% \\ & \hline \end{aligned}$ |  |  |  | $\begin{gathered} 0.113 \\ 64 \% \\ \hline \end{gathered}$ | $\begin{aligned} & 0.860 \\ & 10 \% \\ & \hline \end{aligned}$ | $\begin{gathered} 7.204 \\ 29 \% \\ \hline \end{gathered}$ |  |  |  |  |  |  |  |  |
| $\begin{aligned} & \text { OQ1997-1 } \\ & (\mathrm{SH}-1) \end{aligned}$ | $\begin{gathered} \mathrm{LI} \\ \text { (Shinnecock) } \end{gathered}$ | 40.2695 | 72.2985 | 58 | N/A | F/V Laura <br> Ann | Jul-97 | $\begin{gathered} \text { Loran / } 40 \mathrm{ft} \text { / } \\ 30 \mathrm{sec} . \end{gathered}$ | 28 | $7 / 7$ | 7.75 | 20 | 0.083 | 0.458 | 10.57 | 0.39 | 164.2 | Ok |  |  |  |  | 1,3 |
| OQ1997-2 <br> (WW-1) | NJ (Wildwood) | 38.5095 | 74.1115 | 49 | N/A | F/V Agitator | Aug-97 | $\begin{aligned} & \text { Loran / } 40 \mathrm{ft} \text { / } \\ & 30 \mathrm{sec} . \end{aligned}$ | 28 | $13 / 6$ | 10 | 20 | 0.084 | 0.150 | 2.37 | 0.50 | 176.0 | Ok |  | NA |  | NA | 1,4 |
| Mean CVfor Mean |  |  |  | $\begin{array}{r} 51 \\ 8 \% \\ \hline \end{array}$ |  |  |  |  | $\begin{aligned} & 19.6 \\ & 19 \% \\ & \hline \end{aligned}$ |  |  |  | $\begin{aligned} & 0.180 \\ & 45 \% \\ & \hline \end{aligned}$ | $\begin{gathered} 0.592 \\ 23 \% \\ \hline \end{gathered}$ | $\begin{array}{r} 6.01 \\ 25 \% \\ \hline \end{array}$ |  |  |  |  |  |  |  |  |

SEE FOOTNOTES ON NEXT PAGE


Table A12. Summary of new and revised density, commercial dredge efficiency, and survey dredge efficiency estimates for ocean quahog $90+\mathrm{mm}$ SL from the Patch model and setup tows.
\(\left.$$
\begin{array}{cccc}\hline & \begin{array}{c}\text { Density } \\
(\mathrm{N} \mathrm{ft}\end{array} \\
\text { Statistic }\end{array}
$$ \quad $$
\begin{array}{c}\text { Commercial } \\
\text { Vessel } \\
\text { Efficiency }\end{array}
$$ \quad \begin{array}{c}NEFSC <br>
Dredge <br>

Efficiency\end{array}\right]\)| N experiments | 18 | 17 | 12 |
| :---: | :---: | :---: | :---: |
| Minimum | 0.007 | 0.150 | 0.098 |
| Maximum | 0.295 | 1.000 | 0.990 |
| Median | 0.082 | 0.660 | 0.165 |
| Mean | 0.097 | 0.596 | 0.248 |
| Distribution of point estimates ${ }^{1}$ |  |  |  |
| sd | 0.141 | 0.267 | 0.241 |
| CV (sd/mean) | 1.453 | 0.448 | 0.972 |
| Lo 95\% | 0.000 | 0.073 | 0.000 |
| Hi 95\% | 0.373 | 1.000 | 0.722 |
| Distribution of average estimates ${ }^{1}$ |  |  |  |
| se | 0.033 | 0.065 | 0.070 |
| CV (se/mean) | 0.236 | 0.243 | 0.289 |
| Lo 95\% | 0.032 | 0.469 | 0.112 |
| Hi 95\% | 0.162 | 0.723 | 0.385 |
| Distribution of median estimates ${ }^{2}$ |  |  |  |
| se | 0.011 | 0.091 | 0.029 |
| Robust CV (se/median) | 0.132 | 0.138 | 0.177 |
| Lo 95\% | 0.047 | 0.402 | 0.136 |
| Hi 95\% | 0.089 | 0.733 | 0.261 |

[^6]Table A13. Original (used in the last assessment, NEFSC 2004) and revised ocean quahog density and efficiency estimates from the Patch model based on depletion experiments during 1997-2002. Percent change is (Revised-Previous).Previous x 100. "NA" means not available. Previous and revised density estimates are shown but are not directly comparable because they are based on different size groups.

| Experiment | Density ( $D, \mathrm{nft})^{-2}$ |  |  | Commercial efficiency ( $E$ ) |  |  | Setup Density ( $d, n \mathrm{ft}^{-2}$ ) |  |  | Survey efficiency (e) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Previous | Revised | Change | Previous | Revised | \% Change ${ }^{\text {c }}$ | Previous | Revised | \% Change ${ }^{\text {c }}$ | Previous | Revised | \% Change ${ }^{\text {c }}$ |
| OQ2002-1 | 0.550 | 0.295 | -46\% | 0.653 | 0.489 | -25\% | 0.068 | 0.029 | -57\% | 0.081 | 0.098 | 21\% |
| OQ2002-2 | 0.345 | 0.165 | -52\% | 0.810 | 0.785 | -3\% | 0.067 | 0.024 | -64\% | 0.158 | 0.149 | -6\% |
| OQ2002-3 | 0.111 | 0.081 | -27\% | 0.816 | 0.777 | -5\% | 0.037 | 0.024 | -36\% | 0.275 | 0.297 | 8\% |
| OQ2002-4 | 0.101 | 0.073 | -27\% | 0.599 | 0.254 | -58\% | 0.080 | 0.021 | -74\% | 0.474 | 0.287 | -39\% |
| OQ2000-1 ${ }^{\text {b }}$ | 0.413 | 0.100 | -76\% | 0.950 | 0.730 | -23\% | 0.169 | NA | NA | 0.389 | NA | NA |
| OQ2000-2 ${ }^{\text {b }}$ | 0.095 | 0.062 | -35\% | 0.922 | 0.554 | -40\% | 0.054 | 0.015 | -73\% | 0.524 | 0.234 | -55\% |
| OQ2000-3 ${ }^{\text {b }}$ | 0.180 | 0.089 | -51\% | 0.734 | 0.560 | -24\% | 0.053 | 0.015 | -72\% | 0.216 | 0.165 | -24\% |
| OQ1999-01 DE2 | 0.306 | 0.007 | NA | NA | NA | NA | NA | NA | NA | 0.470 | 0.990 | 111\% |
| OQ1998-1 | 0.105 | 0.017 | -84\% | 0.950 | 1.000 | 5\% | NA | NA | NA | NA | NA | NA |
| OQ1998-2 | 0.242 | 0.067 | -73\% | 0.401 | 0.869 | 117\% | NA | NA | NA | NA | NA | NA |
| OQ1998-3 | 0.570 | 0.255 | -55\% | 0.950 | 0.710 | -25\% | NA | NA | NA | NA | NA | NA |
| OQ1997-1 | 0.440 | 0.083 | -81\% | 0.488 | 0.458 | -6\% | NA | NA | NA | NA | NA | NA |
| OQ1997-2 | 0.060 | 0.084 | 39\% | 0.256 | 0.150 | -41\% | NA | NA | NA | NA | NA | NA |

Previous and revised density estimates are shown for completeness but are not comparable because they are based on different size groups.
${ }^{\mathrm{b}}$ Survey efficiencies calcuated based on information in NEFSC (2000, Tables C12 and C13) using e $=\mathrm{d} / \mathrm{D}^{*} \mathrm{E}$.

Table A14. Effects of new data and methods on efficiency and density estimates for ocean quahog from the Patch model and setup tows (where available).


Table A15. Efficiency corrected swept-area biomass estimates ( $1,000 \mathrm{mt}$ ) and CVs for the fishable stock of ocean quahog during 1997, 2000, 2002 and 2005 by stock assessment region. Figures for SVA and GBK during 2005 were taken from 2003 because no data were available for 2005.

|  | Estimate | CV |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| INPUT: Nominal tow distance ( $\mathrm{d}_{n}, \mathrm{~nm}$ ) INPUT: Dredge width (nm) | 0.15 |  |  |  |  |  |  |  |
|  | 0.0008225 |  |  |  |  |  |  |  |
| Area swept per standard tow ( $a, \mathrm{~nm}^{2}$ ) | $1.23375 \mathrm{E}-04$ | 10\% |  |  |  |  |  |  |
| Area of assessment region ( $A, \mathrm{~nm}^{2}$ ) - no correction for stations with unsuitable clam habitat |  |  |  |  |  |  |  |  |
| S. Virginia and N. Carolina (SVA) | 712 | 10\% |  |  |  |  |  |  |
| Delmarva (DMV) | 4,071 | 10\% |  |  |  |  |  |  |
| New Jersey ( NJ ) | 6,510 | 10\% |  |  |  |  |  |  |
| Long Island (LI) | 4,463 | 10\% |  |  |  |  |  |  |
| Southern New England (SNE) | 4,922 | 10\% |  |  |  |  |  |  |
| Georges Bank (GBK) | 7,821 | 10\% |  |  |  |  |  |  |
| Total | 28,499 |  |  |  |  |  |  |  |
| INPUT: Fraction suitable habitat ( $u$ ) |  |  |  |  |  |  |  |  |
| S. Virginia and N. Carolina (SVA) | 100\% | 10\% |  |  |  |  |  |  |
| Delmarva (DMV) | 100\% | 10\% |  |  |  |  |  |  |
| New Jersey ( NJ ) | 100\% | 10\% |  |  |  |  |  |  |
| Long Island (LI) | 100\% | 10\% |  |  |  |  |  |  |
| Southern New England (SNE) | 96\% | 10\% |  |  |  |  |  |  |
| Georges Bank (GBK) | 90\% | 10\% |  |  |  |  |  |  |
| Habitat area in assessment region ( $A^{\prime}$, nm2) |  |  | INPUT: Biomass fraction in unsurveyd deep water |  |  |  |  |  |
| S. Virginia and N. Carolina (SVA) | 712 | 14\% |  | S. Virginia a | Carolina (SVA) | 0\% | 10\% |  |
| Delmarva (DMV) | 4,071 | 14\% |  |  | elmarva (DMV) | 0\% | 10\% |  |
| New Jersey (NJ) | 6,510 | 14\% |  |  | ew Jersey (NJ) | 0\% | 10\% |  |
| Long Island (LI) | 4,463 | 14\% |  |  | ong Island (LI) | 0\% | 10\% |  |
| Southern New England (SNE) | 4,714 | 14\% |  | Southern | England (SNE) | 2\% | 10\% |  |
| Georges Bank (GBK) | 7,039 | 14\% |  |  | es Bank (GBK) | 13\% | 10\% |  |
| INPUT: Original survey mean catch from fishable stock (kg/tow, for tows adjusted to nominal tow distance using sensors) |  |  |  |  |  |  |  |  |
|  | Estimates for |  | Estimates for |  | Estimates for |  | Estimates for |  |
|  | 1997 | CV | 1999 | CV | 2002 | CV | 2005 | CV |
| S. Virginia and N. Carolina (SVA) | 0.0013 | 100\% | 0.0007 | 55\% | 0.0004 | 100\% | 0.0004 | 100\% |
| Delmarva (DMV) | 0.6528 | 23\% | 0.4449 | 26\% | 0.6863 | 24\% | 0.4221 | 48\% |
| New Jersey ( NJ ) | 1.7341 | 15\% | 0.9728 | 14\% | 1.8614 | 23\% | 1.0441 | 14\% |
| Long Island (LI) | 4.5648 | 17\% | 3.0065 | 14\% | 3.4414 | 17\% | 2.1812 | 16\% |
| Southern New England (SNE) | 2.2252 | 37\% | 2.6964 | 45\% | 3.2654 | 26\% | 2.2555 | 24\% |
| Georges Bank (GBK) | 2.6710 | 16\% | 3.1454 | 18\% | 3.8760 | 17\% | 3.8760 | 17\% |
| Swept-area biomass without efficiency correction (B', 1000 mt ): |  |  |  |  |  |  |  |  |
| S. Virginia and N. Carolina (SVA) | 0.0076 | 102\% | 0.0040 | 59\% | 0.0022 | 102\% | 0.0022 | 102\% |
| Delmarva (DMV) | 21.5388 | 30\% | 14.6803 | 33\% | 22.6452 | 31\% | 13.9280 | 52\% |
| New Jersey ( NJ ) | 91.4993 | 25\% | 51.3297 | 24\% | 98.2159 | 30\% | 55.0929 | 24\% |
| Long Island (LI) | 165.1265 | 26\% | 108.7572 | 24\% | 124.4894 | 26\% | 78.9022 | 26\% |
| Southern New England (SNE) | 86.7210 | 42\% | 105.0878 | 49\% | 127.2624 | 33\% | 87.9046 | 31\% |
| Georges Bank (GBK) | 172.2007 | 26\% | 202.7813 | 27\% | 249.8861 | 26\% | 249.8861 | 26\% |
| Total fishable biomass less GBK | 365 | 17\% | 280 | 21\% | 373 | 16\% | 236 | 16\% |
| Total fishable biomass | 537 | 14\% | 483 | 17\% | 623 | 14\% | 486 | 16\% |
| INPUT: Survey dredge efficiency (e) | 0.165 | 18\% | 0.165 | 18\% | 0.165 | 18\% | 0.165 | 18\% |
| Efficiency adjusted swept area fishable biomass (B, 1000 mt ) |  |  |  |  |  |  |  |  |
| S. Virginia and N. Carolina (SVA) | 0.046 | 104\% | 0.024 | 61\% | 0.013 | 104\% | 0.013 | 104\% |
| Delmarva (DMV) | 131 | 35\% | 89 | 37\% | 137 | 36\% | 84 | 55\% |
| New Jersey ( NJ ) | 555 | 31\% | 311 | 30\% | 596 | 35\% | 334 | 30\% |
| Long Island (LI) | 1,002 | 32\% | 660 | 30\% | 755 | 32\% | 479 | 31\% |
| Southern New England (SNE) | 526 | 46\% | 638 | 52\% | 772 | 37\% | 533 | 36\% |
| Georges Bank (GBK) | 1,045 | 31\% | 1,230 | 32\% | 1,516 | 32\% | 1,516 | 32\% |
| Total fishable biomass less GBK | 2,214 | 24\% | 1,698 | 28\% | 2,261 | 24\% | 1,431 | 24\% |
| Total fishable biomass | 3,258 | 23\% | 2,928 | 24\% | 3,776 | 23\% | 2,947 | 24\% |

Lower bound for $\mathbf{8 0 \%}$ confidence intervals on fishable biomass ( $\mathbf{1 0 0 0} \mathbf{~ m t}$, for lognormal distribution with no bias correction)

|  | $\begin{array}{\|c\|} \hline \text { Estimates for } \\ 1997 \end{array}$ | Estimates for 1999 | Estimates for 2002 | $\begin{aligned} & \text { Estimates for } \\ & 2005 \end{aligned}$ |
| :---: | :---: | :---: | :---: | :---: |
| S. Virginia and N. Carolina (SVA) | 0.015 | 0.012 | 0.004 | 0.004 |
| Delmarva (DMV) | 84 | 56 | 88 | 44 |
| New Jersey ( NJ ) | 378 | 213 | 385 | 229 |
| Long Island (LI) | 675 | 452 | 509 | 324 |
| Southern New England (SNE) | 302 | 340 | 487 | 342 |
| Georges Bank (GBK) | 708 | 823 | 1,021 | 1,021 |
| Total fishable biomass less GBK | 1,627 | 1,199 | 1,667 | 1,060 |
| Total fishable biomass | 2,448 | 2,153 | 2,830 | 2,189 |

Upperbound for $\mathbf{8 0 \%}$ confidence intervals on fishable biomass ( $\mathbf{1 0 0 0} \mathbf{~ m t , ~ f o r ~ l o g n o r m a l ~ d i s t r i b u t i o n ~ w i t h ~ n o ~ b i a s ~ c o r r e c t i o n ) ~}$

| S. Virginia and N. Carolina (SVA) | 0.137 | 0.050 | 0.040 | 0.040 |
| ---: | :---: | :---: | :---: | :---: |
| Delmarva (DMV) | 202 | 141 | 215 | 163 |
| New Jersey (NJ) | 814 | 454 | 923 | 488 |
| Long Island (LI) | 1,488 | 962 | 1,122 | 706 |
| Southern New England (SNE) | 918 | 1,197 | 1,225 | 833 |
| Georges Bank (GBK) | 1,542 | 1,839 | 2,251 | 2,251 |
| Total fishable biomass less GBK | 3,012 | 2,405 | 3,066 | 1,931 |
| Total fishable biomass | 4,336 | 3,982 | 5,039 | 3,967 |

Table A16. Ocean quahog fishing mortality estimates based on catch and efficiency corrected sweptarea biomass for fishable ocean quahog during 1997, 1999, 2002 and 2005. CV's are based on analytical variance calculations assuming log normality, and include uncertainty in catch, survey data, swept-area, amount of suitable habitat, and survey dredge efficiency.


Table A17. Proportions of total fishable ocean quahog biomass during 1980-2005 at a range of survey biomass density levels, by region.

| Fishable biomass density levels (kg/tow) from survey data |  |  |  |  |  |  | Sum of | Total | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Years | 0 to 4 | 5 to 9 | 10 to 14 | 15 to 19 | 20 to 24 | 25+ | Proportions (check) | Number Tows | Number of Surveys |
| Proportions of tows (and stock area) at each survey catch rate level: |  |  |  |  |  |  |  |  |  |
| Southern Virgina (SVA) |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 1.00 |  |  |  |  |  | 1.00 | 47 | 5 |
| 1990-1999 | 1.00 |  |  |  |  |  | 1.00 | 37 | 3 |
| 2000-2005 | 1.00 |  |  |  |  |  | 1.00 | 19 | 2 |
| Delmarva (DMV) |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 0.90 | 0.04 | 0.03 | 0.01 | 0.01 | 0.02 | 1.00 | 317 | 5 |
| 1990-1999 | 0.92 | 0.05 | 0.01 | 0.01 | 0.00 |  | 1.00 | 207 | 3 |
| 2000-2005 | 0.96 | 0.02 | 0.01 | 0.01 |  |  | 1.00 | 131 | 2 |
| New Jersy (NJ) |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 0.84 | 0.07 | 0.03 | 0.02 | 0.02 | 0.03 | 1.00 | 458 | 5 |
| 1990-1999 | 0.82 | 0.11 | 0.04 | 0.02 | 0.01 |  | 1.00 | 307 | 3 |
| 2000-2005 | 0.92 | 0.05 | 0.02 |  |  | 0.01 | 1.00 | 183 | 2 |
| Long Island (LI) |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 0.57 | 0.21 | 0.12 | 0.06 | 0.01 | 0.04 | 1.00 | 218 | 5 |
| 1990-1999 | 0.49 | 0.19 | 0.12 | 0.10 | 0.02 | 0.07 | 1.00 | 121 | 3 |
| 2000-2005 | 0.64 | 0.24 | 0.06 | 0.02 | 0.01 | 0.02 | 1.00 | 84 | 2 |
| Southern New England (SNE) |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 0.75 | 0.09 | 0.08 | 0.03 | 0.02 | 0.03 | 1.00 | 245 | 5 |
| 1990-1999 | 0.67 | 0.16 | 0.08 | 0.04 | 0.01 | 0.04 | 1.00 | 114 | 3 |
| 2000-2005 | 0.65 | 0.23 | 0.07 | 0.04 | 0.02 |  | 1.00 | 57 | 2 |
| Georges Bank (GBK) |  |  |  |  |  |  |  |  |  |
| 1986-1992 | 0.82 | 0.06 | 0.03 | 0.01 | 0.01 | 0.06 | 1.00 | 201 | 3 |
| 1997-2002 | 0.68 | 0.10 | 0.07 | 0.03 | 0.05 | 0.07 | 1.00 | 219 | 3 |
| All years | 0.75 | 0.08 | 0.05 | 0.02 | 0.03 | 0.07 | 1.00 | 420 | 6 |
| Mean survey catch rate (kg/tow) at each survey catch rate level ( $p_{L}$ ): |  |  |  |  |  |  |  |  |  |
| Southern Virgina (SVA) |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 0.054 |  |  |  |  |  |  |  |  |
| 1990-1999 | 0.007 |  |  |  |  |  |  |  |  |
| 2000-2005 | 0.002 |  |  |  |  |  |  |  |  |
| Delmarva (DMV) |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 0.490 | 5.856 | 11.604 | 18.761 | 21.994 | 31.082 |  |  |  |
| 1990-1999 | 0.413 | 7.133 | 13.556 | 17.734 | 21.847 |  |  |  |  |
| 2000-2005 | 0.307 | 7.888 | 11.960 | 15.524 |  |  |  |  |  |
| New Jersy (NJ) |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 0.848 | 7.115 | 12.577 | 17.033 | 20.956 | 35.668 |  |  |  |
| 1990-1999 | 0.647 | 6.845 | 11.748 | 17.546 | 23.198 |  |  |  |  |
| 2000-2005 | 0.938 | 6.166 | 12.707 |  |  | 29.972 |  |  |  |
| Long Island (LI) 20.972 |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 1.703 | 7.100 | 12.281 | 17.431 | 20.781 | 38.945 |  |  |  |
| 1990-1999 | 1.252 | 7.523 | 12.508 | 16.974 | 22.793 | 30.846 |  |  |  |
| 2000-2005 | 1.779 | 6.894 | 12.780 | 16.666 | 20.087 | 39.638 |  |  |  |
| Southern New England (SNE) |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 1.002 | 7.084 | 12.200 | 17.286 | 21.627 | 33.942 |  |  |  |
| 1990-1999 | 1.001 | 7.461 | 11.993 | 17.384 | 20.904 | 36.563 |  |  |  |
| 2000-2005 | 1.387 | 7.238 | 12.077 | 16.226 | 21.845 |  |  |  |  |
| Georges Bank (GBK) |  |  |  |  |  |  |  |  |  |
| 1986-1992 | 0.627 | 6.874 | 12.945 | 16.049 | 23.225 | 44.962 |  |  |  |
| 1997-2002 | 0.626 | 7.681 | 12.370 | 16.595 | 23.386 | 40.787 |  |  |  |
| All years | 0.627 | 7.381 | 12.535 | 16.413 | 23.349 | 42.576 |  |  |  |
| Proportions of stock biomass at each survey catch rate level ( $X_{L}$ ) : |  |  |  |  |  |  |  |  |  |
| Southern Virgina (SVA) |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 1.00 |  |  |  |  |  | 1.00 |  |  |
| 1990-1999 | 1.00 |  |  |  |  |  | 1.00 |  |  |
| 2000-2005 | 1.00 |  |  |  |  |  | 1.00 |  |  |
| Delmarva (DMV) |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 0.23 | 0.12 | 0.15 | 0.12 | 0.07 | 0.31 | 1.00 |  |  |
| 1990-1999 | 0.30 | 0.27 | 0.15 | 0.20 | 0.08 |  | 1.00 |  |  |
| 2000-2005 | 0.43 | 0.26 | 0.13 | 0.17 |  |  | 1.00 |  |  |
|  |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 0.22 | 0.15 | 0.14 | 0.09 | 0.10 | 0.29 | 1.00 |  |  |
| 1990-1999 | 0.23 | 0.34 | 0.20 | 0.17 | 0.07 |  | 1.00 |  |  |
| 2000-2005 | 0.49 | 0.17 | 0.16 |  |  | 0.19 | 1.00 |  |  |
| Long Island (LI) 0.10 |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 0.15 | 0.22 | 0.23 | 0.15 | 0.03 | 0.22 | 1.00 |  |  |
| 1990-1999 | 0.08 | 0.18 | 0.19 | 0.21 | 0.07 | 0.28 | 1.00 |  |  |
| 2000-2005 | 0.22 | 0.32 | 0.15 | 0.08 | 0.05 | 0.18 | 1.00 |  |  |
| Southern New England (SNE) |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 0.18 | 0.16 | 0.22 | 0.13 | 0.08 | 0.23 | 1.00 |  |  |
| 1990-1999 | 0.12 | 0.22 | 0.18 | 0.14 | 0.03 | 0.30 | 1.00 |  |  |
| 2000-2005 | 0.21 | 0.38 | 0.19 | 0.13 | 0.09 | 0.00 | 1.00 |  |  |
| Georges Bank (GBK) 0.10 |  |  |  |  |  |  |  |  |  |
| 1986-1992 | 0.11 | 0.10 | 0.08 | 0.05 | 0.08 | 0.58 | 1.00 |  |  |
| 1997-2002 | 0.07 | 0.12 | 0.13 | 0.07 | 0.16 | 0.45 | 1.00 |  |  |
| All years | 0.08 | 0.11 | 0.11 | 0.06 | 0.13 | 0.50 | 1.00 |  |  |

Table A18. Proportions of total 2005 stock biomass at a range of survey density levels, by region.

| Survey catch rate level (kg/tow) |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Region | 0 to 4 | 5 to 9 | 10 to 14 | 15 to 19 | 20 to 24 | 25+ | Total |
| Total 2005 biomass (mt meats) |  |  |  |  |  |  |  |
| Southern Virgina (SVA) | 17 | 0 | 0 | 0 | 0 | 0 | 17 |
| Delmarva (DMV) | 43,532 | 26,628 | 13,459 | 17,470 | 0 | 0 | 101,089 |
| New Jersy (NJ) | 195,400 | 68,833 | 63,047 | 0 | 0 | 74,354 | 401,634 |
| Long Island (LI) | 151,198 | 217,001 | 100,560 | 52,457 | 31,612 | 124,762 | 677,590 |
| Southern New England (SNE) | 123,098 | 225,647 | 115,846 | 77,824 | 52,388 | 0 | 594,802 |
| Georges Bank (GBK) | 82,714 | 148,850 | 163,456 | 87,709 | 206,009 | 574,872 | 1,263,610 |
| Total | 595,959 | 686,960 | 456,369 | 235,460 | 290,008 | 773,987 | 3,038,741 |
| Total 2005 biomass (bushels) |  |  |  |  |  |  |  |
| Southern Virgina (SVA) | 3,731 | 0 | 0 | 0 | 0 | 0 | 3,731 |
| Delmarva (DMV) | 9,597,036 | 5,870,504 | 2,967,208 | 3,851,373 | 0 | 0 | 22,286,120 |
| New Jersy (NJ) | 43,077,930 | 15,174,947 | 13,899,368 | 0 | 0 | 16,391,987 | 88,544,232 |
| Long Island (LI) | 33,333,071 | 47,840,106 | 22,169,510 | 11,564,629 | 6,969,113 | 27,504,966 | 149,381,395 |
| Southern New England (SNE) | 27,138,182 | 49,746,067 | 25,539,371 | 17,157,064 | 11,549,366 | 0 | 131,130,049 |
| Georges Bank (GBK) | 18,235,073 | 32,815,497 | 36,035,560 | 19,336,384 | 45,416,674 | 126,736,217 | 278,575,405 |
| Total | 131,385,021 | 151,447,120 | 100,611,016 | 51,909,450 | 63,935,154 | 170,633,170 | 669,920,932 |
| Percent of total 2005 biomass |  |  |  |  |  |  |  |
| Southern Virgina (SVA) | 0.001\% | 0.000\% | 0.000\% | 0.000\% | 0.000\% | 0.000\% | 0.001\% |
| Delmarva (DMV) | 1.43\% | 0.88\% | 0.44\% | 0.57\% | 0.00\% | 0.00\% | 3.33\% |
| New Jersy (NJ) | 6.43\% | 2.27\% | 2.07\% | 0.00\% | 0.00\% | 2.45\% | 13.22\% |
| Long Island (LI) | 4.98\% | 7.14\% | 3.31\% | 1.73\% | 1.04\% | 4.11\% | 22.30\% |
| Southern New England (SNE) | 4.05\% | 7.43\% | 3.81\% | 2.56\% | 1.72\% | 0.00\% | 19.57\% |
| Georges Bank (GBK) | 2.72\% | 4.90\% | 5.38\% | 2.89\% | 6.78\% | 18.92\% | 41.58\% |
| Total | 19.61\% | 22.61\% | 15.02\% | 7.75\% | 9.54\% | 25.47\% | 100.00\% |

Table A19. Calculations to build a bridge between efficiency corrected swept area biomass estimates for ocean quahog during 2002 in NEFSC (2004) and new estimates in this assessment. Columns show cumulative effects from each change in data and methods starting with NEFSC's (2004) estimates on the left and ending with the new estimates on the right.

| Region | $\begin{gathered} \text { NEFSC } \\ (2004) \end{gathered}$ | Step 1 (New spread sheet) | Step 2 (Correct survey data) | Step 3 (Add biomass in deep water) | Step 4 (Use fishable biomass) | This assessment (New efficiency estimate | Ratio (New / NEFSC(2004) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Data and configuration |  |  |  |  |  |  |  |
| Efficiency | 0.269 | 0.269 | 0.269 | 0.269 | 0.269 | 0.165 | 0.61 |
| Size groups in Patch model | 70+ | 70+ | 70+ | 70+ | Fishable | Fishable | NA |
| Deep water percentage | 0\% | 0\% | 0\% | 13\% | 13\% | 13\% | NA |
| Survey data | Erroneous | Erroneous | Correct | Correct | Correct | Correct | NA |
| 2002 efficiency corrected swept-area biomass estimates (1000 mt) |  |  |  |  |  |  |  |
| SVA | 0.01 | 0.01 | 0.01 | 0.01 | 0.01 | 0.01 | 1.93 |
| DMV | 71 | 71 | 89 | 89 | 84 | 137 | 1.93 |
| NJ | 330 | 330 | 383 | 383 | 365 | 596 | 1.81 |
| LI | 454 | 454 | 498 | 498 | 463 | 755 | 1.66 |
| SNE | 428 | 437 | 511 | 511 | 473 | 772 | 1.80 |
| GBK | 833 | 833 | 875 | 989 | 929 | 1,516 | 1.82 |
| Total less GBK | 1,283 | 1,292 | 1,481 | 1,481 | 1,385 | 2,261 | 1.76 |
| Total | 2,116 | 2,125 | 2,356 | 2,470 | 2,314 | 3,776 | 1.78 |

Table A20. "Best" fishable biomass and fishing mortality estimates for ocean quahog during 1978-2005, by stock assessment region and for the entire EEZ stock (with and without GBK).

|  | SVA | DMV |  | NJ |  | LI |  | SNE |  | GBK |  | Entire stock less GBK | Entire Stock |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Model | VPA | KLAMZ |  | KLAMZ |  | KLAMZ |  | KLAMZ |  | Mean ESB 1997-2002(Same as VPA) |  | NA | NA |
|  | Estimate | Estimate | cV | Estimate | cV | Estimate | cV | Estimate | cV | Estimate | cV | Estimate | Estimate |
| Scaling parameter for swept area biomass | $\begin{gathered} 1 \\ \text { (assumed) } \end{gathered}$ | 0.99 | NA | 0.95 | NA | 1.00 | NA | 0.99 | NA | $\begin{gathered} 1 \\ \text { (assumed) } \end{gathered}$ | NA | NA | NA |
| $\begin{aligned} & \text { Recruitment } \\ & (1000 \mathrm{mt} \text { ) } \end{aligned}$ | NA | 0 (assumed) |  | 0.541 | 0.43 | 9.860 | 0.34 | 4.799 | 1.06 | NA |  | $\begin{aligned} & 15.199 \\ & \text { (excludes } \\ & \text { SVA) } \\ & \hline \end{aligned}$ | 15.199 (excludes SVA \& GBK) |
| Fishable Stock Biomass (1000 mt) |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 1978 | 0.338 | 299 | 0.10 | 904 | 0.22 | 718 | 0.26 | 788 | 0.32 | 1,264 | 0.19 | 2,710 | 3,973 |
| 1979 | 0.338 | 292 | 0.10 | 879 | 0.22 | 721 | 0.25 | 782 | 0.31 | 1,264 | 0.19 | 2,674 | 3,938 |
| 1980 | 0.338 | 280 | 0.10 | 855 | 0.21 | 723 | 0.24 | 777 | 0.30 | 1,264 | 0.19 | 2,636 | 3,900 |
| 1981 | 0.338 | 270 | 0.11 | 831 | 0.21 | 726 | 0.24 | 771 | 0.29 | 1,264 | 0.19 | 2,599 | 3,862 |
| 1982 | 0.275 | 261 | 0.11 | 807 | 0.21 | 729 | 0.23 | 765 | 0.28 | 1,264 | 0.19 | 2,562 | 3,826 |
| 1983 | 0.268 | 251 | 0.11 | 783 | 0.20 | 732 | 0.22 | 759 | 0.27 | 1,264 | 0.19 | 2,525 | 3,789 |
| 1984 | 0.268 | 240 | 0.11 | 761 | 0.20 | 734 | 0.22 | 753 | 0.26 | 1,264 | 0.19 | 2,489 | 3,753 |
| 1985 | 0.261 | 228 | 0.12 | 738 | 0.20 | 737 | 0.21 | 747 | 0.25 | 1,264 | 0.19 | 2,450 | 3,714 |
| 1986 | 0.075 | 216 | 0.12 | 713 | 0.20 | 739 | 0.20 | 741 | 0.24 | 1,264 | 0.19 | 2,409 | 3,672 |
| 1987 | 0.075 | 203 | 0.12 | 691 | 0.19 | 742 | 0.20 | 735 | 0.24 | 1,264 | 0.19 | 2,370 | 3,634 |
| 1988 | 0.075 | 189 | 0.13 | 669 | 0.19 | 743 | 0.20 | 729 | 0.23 | 1,264 | 0.19 | 2,330 | 3,594 |
| 1989 | 0.031 | 173 | 0.14 | 650 | 0.19 | 745 | 0.19 | 723 | 0.22 | 1,264 | 0.19 | 2,291 | 3,555 |
| 1990 | 0.031 | 163 | 0.15 | 625 | 0.19 | 747 | 0.19 | 717 | 0.21 | 1,264 | 0.19 | 2,252 | 3,515 |
| 1991 | 0.017 | 157 | 0.15 | 598 | 0.19 | 749 | 0.18 | 711 | 0.21 | 1,264 | 0.19 | 2,214 | 3,478 |
| 1992 | 0.017 | 149 | 0.15 | 573 | 0.19 | 749 | 0.18 | 705 | 0.20 | 1,264 | 0.19 | 2,176 | 3,440 |
| 1993 | 0.017 | 143 | 0.16 | 556 | 0.18 | 740 | 0.18 | 699 | 0.20 | 1,264 | 0.19 | 2,139 | 3,402 |
| 1994 | 0.017 | 139 | 0.16 | 536 | 0.18 | 734 | 0.18 | 693 | 0.19 | 1,264 | 0.19 | 2,102 | 3,366 |
| 1995 | 0.017 | 135 | 0.16 | 520 | 0.18 | 724 | 0.18 | 688 | 0.18 | 1,264 | 0.19 | 2,067 | 3,331 |
| 1996 | 0.017 | 131 | 0.16 | 505 | 0.18 | 718 | 0.18 | 678 | 0.18 | 1,264 | 0.19 | 2,033 | 3,296 |
| 1997 | 0.017 | 128 | 0.16 | 492 | 0.18 | 714 | 0.18 | 665 | 0.18 | 1,264 | 0.19 | 2,000 | 3,263 |
| 1998 | 0.017 | 125 | 0.16 | 479 | 0.18 | 712 | 0.18 | 652 | 0.18 | 1,264 | 0.19 | 1,968 | 3,231 |
| 1999 | 0.017 | 121 | 0.17 | 468 | 0.17 | 708 | 0.18 | 642 | 0.18 | 1,264 | 0.19 | 1,938 | 3,202 |
| 2000 | 0.017 | 117 | 0.17 | 457 | 0.17 | 705 | 0.18 | 631 | 0.18 | 1,264 | 0.19 | 1,910 | 3,173 |
| 2001 | 0.017 | 114 | 0.17 | 445 | 0.17 | 703 | 0.18 | 622 | 0.18 | 1,264 | 0.19 | 1,884 | 3,148 |
| 2002 | 0.017 | 111 | 0.17 | 433 | 0.17 | 700 | 0.18 | 614 | 0.18 | 1,264 | 0.19 | 1,857 | 3,121 |
| 2003 | 0.017 | 107 | 0.17 | 422 | 0.17 | 694 | 0.18 | 607 | 0.18 | 1,264 | 0.19 | 1,830 | 3,093 |
| 2004 | 0.017 | 104 | 0.17 | 412 | 0.17 | 685 | 0.18 | 601 | 0.18 | 1,264 | 0.19 | 1,802 | 3,065 |
| 2005 | 0.017 | 101 | 0.18 | 402 | 0.17 | 678 | 0.18 | 595 | 0.18 | 1,264 | 0.19 | 1,775 | 3,039 |

Table A20 (continued).

| Fishing mortality rate |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1978 | 0.0000 | 0.0044 | 0.10 | 0.0097 | 0.22 | 0.00000 | NA | 0.0000 | NA | 0.0000 | 0.00 | 0.0039 | 0.0026 |
| 1979 | 0.0000 | 0.0191 | 0.10 | 0.0095 | 0.22 | 0.00000 | NA | 0.0000 | NA | 0.0000 | 0.00 | 0.0059 | 0.0040 |
| 1980 | 0.0000 | 0.0172 | 0.10 | 0.0103 | 0.21 | 0.00001 | 0.24 | 0.0000 | NA | 0.0000 | 0.00 | 0.0051 | 0.0034 |
| 1981 | 0.2085 | 0.0148 | 0.11 | 0.0111 | 0.21 | 0.00000 | 0.24 | 0.0000 | NA | 0.0000 | 0.00 | 0.0051 | 0.0034 |
| 1982 | 0.0252 | 0.0195 | 0.11 | 0.0116 | 0.21 | 0.00000 | NA | 0.0000 | NA | 0.0000 | 0.00 | 0.0056 | 0.0037 |
| 1983 | 0.0000 | 0.0224 | 0.11 | 0.0109 | 0.20 | 0.00003 | 0.22 | 0.0009 | 0.27 | 0.0000 | 0.00 | 0.0058 | 0.0039 |
| 1984 | 0.0257 | 0.0326 | 0.11 | 0.0126 | 0.20 | 0.00000 | NA | 0.0012 | 0.26 | 0.0000 | 0.00 | 0.0072 | 0.0048 |
| 1985 | 1.2454 | 0.0358 | 0.12 | 0.0163 | 0.20 | 0.0001 | 0.21 | 0.0010 | 0.25 | 0.0000 | 0.00 | 0.0085 | 0.0056 |
| 1986 | 0.0000 | 0.0407 | 0.12 | 0.0134 | 0.20 | 0.0006 | 0.20 | 0.0008 | 0.24 | 0.0000 | 0.00 | 0.0079 | 0.0052 |
| 1987 | 0.0000 | 0.0539 | 0.13 | 0.0134 | 0.19 | 0.0016 | 0.20 | 0.0010 | 0.24 | 0.0000 | 0.00 | 0.0091 | 0.0059 |
| 1988 | 0.8817 | 0.0649 | 0.14 | 0.0106 | 0.19 | 0.0009 | 0.20 | 0.0012 | 0.23 | 0.0000 | 0.00 | 0.0087 | 0.0057 |
| 1989 | 0.0000 | 0.0383 | 0.14 | 0.0221 | 0.19 | 0.0008 | 0.19 | 0.0017 | 0.22 | 0.0000 | 0.00 | 0.0098 | 0.0063 |
| 1990 | 0.6092 | 0.0230 | 0.15 | 0.0255 | 0.19 | 0.0010 | 0.19 | 0.0013 | 0.21 | 0.0000 | 0.00 | 0.0094 | 0.0060 |
| 1991 | 0.0000 | 0.0317 | 0.15 | 0.0249 | 0.19 | 0.0023 | 0.18 | 0.0012 | 0.21 | 0.0000 | 0.00 | 0.0100 | 0.0064 |
| 1992 | 0.0000 | 0.0163 | 0.16 | 0.0123 | 0.19 | 0.0161 | 0.18 | 0.0016 | 0.20 | 0.0000 | 0.00 | 0.0104 | 0.0066 |
| 1993 | 0.0000 | 0.0139 | 0.16 | 0.0187 | 0.19 | 0.0118 | 0.18 | 0.0015 | 0.20 | 0.0000 | 0.00 | 0.0103 | 0.0065 |
| 1994 | 0.0000 | 0.0073 | 0.16 | 0.0132 | 0.18 | 0.0166 | 0.18 | 0.0014 | 0.19 | 0.0000 | 0.00 | 0.0100 | 0.0063 |
| 1995 | 0.0000 | 0.0053 | 0.16 | 0.0105 | 0.18 | 0.0133 | 0.18 | 0.0080 | 0.19 | 0.0000 | 0.00 | 0.0103 | 0.0064 |
| 1996 | 0.0000 | 0.0057 | 0.16 | 0.0098 | 0.18 | 0.0084 | 0.18 | 0.0125 | 0.18 | 0.0000 | 0.00 | 0.0099 | 0.0061 |
| 1997 | 0.0000 | 0.0085 | 0.16 | 0.0087 | 0.18 | 0.0073 | 0.18 | 0.0137 | 0.18 | 0.0000 | 0.00 | 0.0099 | 0.0060 |
| 1998 | 0.0000 | 0.0112 | 0.16 | 0.0057 | 0.18 | 0.0097 | 0.18 | 0.0105 | 0.18 | 0.0000 | 0.00 | 0.0091 | 0.0056 |
| 1999 | 0.0000 | 0.0092 | 0.17 | 0.0066 | 0.17 | 0.0090 | 0.18 | 0.0105 | 0.18 | 0.0000 | 0.00 | 0.0090 | 0.0054 |
| 2000 | 0.0000 | 0.0091 | 0.17 | 0.0074 | 0.17 | 0.0068 | 0.18 | 0.0082 | 0.18 | 0.0000 | 0.00 | 0.0077 | 0.0047 |
| 2001 | 0.0000 | 0.0084 | 0.17 | 0.0110 | 0.17 | 0.0086 | 0.18 | 0.0080 | 0.18 | 0.0000 | 0.00 | 0.0091 | 0.0054 |
| 2002 | 0.0000 | 0.0160 | 0.17 | 0.0065 | 0.17 | 0.0132 | 0.18 | 0.0064 | 0.18 | 0.0000 | 0.00 | 0.0097 | 0.0058 |
| 2003 | 0.0000 | 0.0085 | 0.17 | 0.0089 | 0.17 | 0.0170 | 0.18 | 0.0036 | 0.18 | 0.0000 | 0.00 | 0.0103 | 0.0061 |
| 2004 | 0.0000 | 0.0062 | 0.17 | 0.0069 | 0.17 | 0.0157 | 0.18 | 0.0055 | 0.18 | 0.0000 | 0.00 | 0.0098 | 0.0058 |
| 2005 | 0.0000 | 0.0094 | 0.18 | 0.0017 | 0.17 | 0.0145 | 0.18 | 0.0034 | 0.18 | 0.0000 | 0.00 | 0.0077 | 0.0045 |

Table A21. Ocean quahog biomass in 2005 as a percentage of biomass in 1978, based on best estimates.

| SVA | DMV | NJ | LI | SNE | GBK | Entire <br> stock <br> less <br> GBK | Entire <br> Stock |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $5 \%$ | $34 \%$ | $44 \%$ | $94 \%$ | $75 \%$ | $100 \%$ | $66 \%$ | $76 \%$ |

Table A22. Comparison of best estimates for ocean quahog biomass during 2004 from the previous (NEFSC 2004) and current assessments.

|  | SVA | DMV | NJ | LI | SNE | GBK | Entire stock <br> less GBK | Entire <br> Stock |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Assessment | 1978 Biomass Estimates (Virgin Biomass) |  |  |  |  |  |  |  |
| This assessment | 0.338 | 299 | 904 | 718 | 788 | 1,264 | 2,710 | 3,973 |
| NEFSC (2004) | 0.297 | 298 | 455 | 534 | 386 | 655 | 1,674 | 2,329 |
| Ratio (new/old) | 1.1 | 1.0 | 2.0 | 1.3 | 2.0 | 1.9 | 1.6 | 1.7 |
| This assessment | 0.0169 | 103.8 | 411.5 | 685 | 601.3 | 1264 | 1801.603121 | 3065 |
| NEFSC (2004) | 0.013 | 91 | 284 | 478 | 349 | 655 | 1,201 | 1,856 |
| Ratio (new/old) | 1.3 | 1.1 | 1.5 | 1.4 | 1.7 | 1.9 | 1.5 | 1.7 |

Table A23. Mean numbers per tow for ocean quahog $<70 \mathrm{~mm} \mathrm{SL}$ and mean weight per tow for ocean quahog 70+ mm SL in NEFSC clam surveys on GBK during 1986-2002 (1994 omitted due to high pump voltage).

| Year | $<70 \mathrm{~mm}$ <br> SL <br> $\left(\mathrm{Ntw}^{-1}\right)$ | CV | $70+\mathrm{mm}$ <br> SL <br> $\left(\mathrm{KG} \mathrm{tow}^{-1}\right)$ | CV |
| :---: | :---: | :---: | :---: | :---: |
| 1986 | 40.5 | 0.60 | 5.7 | 0.17 |
| 1989 | 7.0 | 0.32 | 2.3 | 0.26 |
| 1992 | 31.7 | 0.35 | 9.0 | 0.21 |
| 1997 | 62.0 | 0.35 | 6.6 | 0.19 |
| 1999 | 35.3 | 0.34 | 7.5 | 0.19 |
| 2002 | 39.7 | 0.18 | 8.7 | 0.20 |

Table A25. Percentage of ocean quahog biomass in each stock assessment region during 1978 and 2005. Percentages for SVA, DMV, NJ, LI, SNE and GBK in the same row sum to $100 \%$.

| Year | SVA | DMV | NJ | LI | SNE | GBK | Entire <br> stock <br> less <br> GBK |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1978 | $0.009 \%$ | $8 \%$ | $23 \%$ | $18 \%$ | $20 \%$ | $32 \%$ | $68 \%$ |
| 2005 | $0.001 \%$ | $3 \%$ | $13 \%$ | $22 \%$ | $20 \%$ | $42 \%$ | $58 \%$ |

Table A26. Input data for ocean quahog projections.

| Year | SVA | DMV | NJ | LI | SNE | GBK | Total Less GBK | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Somatic growth rate ( $\mathrm{G}^{-1}$ ) |  |  |  |  |  |  |  |  |
| 2005 | 0.0045 | $1.0600 \mathrm{E}-$ 07 | 0.0013 | 0.0101 | 0.0066 | 0 | 0.0064 | 0.0037 |
| Recruitment rate ( $r=$ Recruitment / Average Biomass in $2005 \boldsymbol{y}^{-1}$ ) |  |  |  |  |  |  |  |  |
| 2005 | 0.0060 | $\begin{array}{r} 1.0038 \mathrm{E}- \\ 08 \end{array}$ | 0.0014 | 0.0146 | 0.0081 | 0.0000 | 0.0086 | 0.0050 |
| Natural mortality ( $M^{y^{-1}}$ ) |  |  |  |  |  |  |  |  |
| 2005 | 0.0200 | 0.0200 | 0.0200 | 0.0200 | 0.0200 | 0 | 0.0200 | 0.0117 |
| Initial Biomass |  |  |  |  |  |  |  |  |
| 2005 | 0.017 | 101 | 402 | 678 | 595 | 1,264 | 1,775 | 3,039 |
| Landings ( $\mathrm{mt}^{\text {c }} \boldsymbol{y}^{-1}$ ) |  |  |  |  |  |  |  |  |
| 2005 | 0.000 | 0.890 | 0.634 | 9.251 | 1.924 | 0 | 12.6990 | 12.6990 |
| Catch (landings +5\% allowance for incidental mortality, mt $\boldsymbol{y}^{-1}$ ) |  |  |  |  |  |  |  |  |
| 2005 | 0.000 | 0.935 | 0.665 | 9.713 | 2.021 | 0 | 13.3340 | 13.3340 |
| Fishing mortality ( $\mathrm{Fr}^{\boldsymbol{- 1} \text { ) }}$ |  |  |  |  |  |  |  |  |
| 2005 | 0.0000 | 0.0094 | 0.0017 | 0.0145 | 0.0034 | 0 | 0.0077 | 0.0045 |

Table A27. Projected biomass and fishing mortality for ocean quahog during 2005-2010 based on a 4 million bushel ( $18,144 \mathrm{mt}$ meats) annual quota during 2007-2010. Landings during 2006 are assumed the same as in 2005. Proportions of total catch in each year for each region are the same as in 2005.

| Year | SVA | DMV | NJ | LI | SNE | GBK | Total Less GBK | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Somatic growth rate ( $\mathrm{G}^{\text {-1 }}$ ) |  |  |  |  |  |  |  |  |
| 2005 | 0.0045 | 0.0000 | 0.0013 | 0.0101 | 0.0066 | 0.0000 | 0.0064 | 0.0037 |
| Recruitment rate ( $r=$ Recruitment / Average Biomass in $2002 \boldsymbol{y}^{-1}$ ) |  |  |  |  |  |  |  |  |
| 2005 | 0.0060 | 0.0000 | 0.0014 | 0.0146 | 0.0081 | 0.0000 | 0.0086 | 0.0050 |
| Natural mortality ( $\mathrm{M}_{\boldsymbol{y}}{ }^{-1}$ ) |  |  |  |  |  |  |  |  |
| 2005 | 0.0200 | 0.0200 | 0.0200 | 0.0200 | 0.0200 | 0.0000 | 0.0200 | 0.0117 |
| Net instantaneous rate of change, less fishing ( $X-F=G+r-M y^{-1}$ ) |  |  |  |  |  |  |  |  |
| 2005 | $0.0095$ | $0.0200$ | $0.0174$ | $0.0047$ | $0.0052$ | 0.0000 | -0.0050 | 0.0029 |
| Fishing mortality first year ( $F \boldsymbol{y}^{-1}$ ) |  |  |  |  |  |  |  |  |
| 2005 | 0.0000 | 0.0094 | 0.0017 | 0.0145 | 0.0034 | 0.0000 | 0.0077 | 0.0045 |
| Landings (mt meats $y^{-1}$ ) |  |  |  |  |  |  |  |  |
| $\begin{aligned} & 2005- \\ & 2006 \end{aligned}$ | 0 | 1 | 1 | 9 | 2 | 0 | 13 | 13 |
| 2007- | 0 | 1 | 1 | 13 | 3 | 0 | 18 | 18 |
| 2010 |  |  |  |  |  |  |  | 18 |
| Catch (mt meats $\boldsymbol{y}^{-1}$, landings+5\% allowance for incidental mortality) |  |  |  |  |  |  |  |  |
| $\begin{aligned} & 2005- \\ & 2006 \end{aligned}$ | 0 | 1 | 1 | 10 | 2 | 0 | 13 | 13 |
| 2007- | 0 | 1 | 1 | 14 | 3 | 0 | 19 | 19 |
| Initial Biomass |  |  |  |  |  |  |  |  |
| 2005- | 0 | 101 | 402 | 678 | 595 | 1,264 | 1,775 | 3,039 |
| 2006 |  |  |  |  |  |  |  |  |
| Projected biomass (mt meats) |  |  |  |  |  |  |  |  |
| 2006 | 0 | 98 | 394 | 671 | 590 | 1,264 | 1,753 | 3,016 |
| 2007 | 0 | 95 | 387 | 664 | 585 | 1,264 | 1,731 | 2,995 |
| 2008 | 0 | 92 | 379 | 654 | 579 | 1,264 | 1,703 | 2,967 |
| 2009 | 0 | 89 | 372 | 643 | 573 | 1,264 | 1,676 | 2,940 |
| 2010 | 0 | 86 | 364 | 632 | 567 | 1,264 | 1,649 | 2,912 |
| Projected fishing mortality rate ( $F \boldsymbol{y}^{-1}$ ) |  |  |  |  |  |  |  |  |
| 2006 | 0.000 | 0.010 | 0.002 | 0.015 | 0.003 | 0.000 | 0.008 | 0.004 |
| 2007 | 0.000 | 0.014 | 0.002 | 0.021 | 0.005 | 0.000 | 0.011 | 0.006 |
| 2008 | 0.000 | 0.015 | 0.003 | 0.021 | 0.005 | 0.000 | 0.011 | 0.006 |
| 2009 | 0.000 | 0.015 | 0.003 | 0.022 | 0.005 | 0.000 | 0.011 | 0.007 |
| 2010 | 0.000 | 0.016 | 0.003 | 0.022 | 0.005 | 0.000 | 0.012 | 0.007 |

Table A28. Projected biomass and fishing mortality for ocean quahog during 2005-2010 based on a 5.333 million bushel ( $24,189 \mathrm{mt}$ meats) annual quota during 2007-2010. Landings during 2006 are assumed the same as in 2005. Proportions of total catch in each year for each region are the same as in 2005.

| Year | SVA | DMV | NJ | LI | SNE | GBK | Total Less GBK | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Somatic growth rate ( $\mathrm{G} \mathrm{y}^{-1}$ ) |  |  |  |  |  |  |  |  |
| 2005 | 0.0045 | 0.0000 | 0.0013 | 0.0101 | 0.0066 | 0.0000 | 0.0064 | 0.0037 |
| Recruitment rate ( $r=$ Recruitment / Average Biomass in $2002 \boldsymbol{y}^{-1}$ ) |  |  |  |  |  |  |  |  |
| 2005 | 0.0060 | 0.0000 | 0.0014 | 0.0146 | 0.0081 | 0.0000 | 0.0086 | 0.0050 |
| Natural mortality ( $\mathrm{M}^{-1}$ ) |  |  |  |  |  |  |  |  |
| 2005 | 0.0200 | 0.0200 | 0.0200 | 0.0200 | 0.0200 | 0.0000 | 0.0200 | 0.0117 |
| Net instantaneous rate of change, less fishing ( $X-F=G+r-M y^{-1}$ ) |  |  |  |  |  |  |  |  |
| 2005 | -0.0095 | -0.0200 | -0.0174 | 0.0047 | -0.0052 | 0.0000 | -0.0050 | -0.0029 |
| Fishing mortality first year ( $\mathrm{F}^{\boldsymbol{y}}{ }^{-1}$ ) |  |  |  |  |  |  |  |  |
| 2005 | 0.0000 | 0.0094 | 0.0017 | 0.0145 | 0.0034 | 0.0000 | 0.0077 | 0.0045 |
| Landings (mt meats $y^{-1}$ ) |  |  |  |  |  |  |  |  |
| 2005-2006 | 0 | 1 | 1 | 9 | 2 | 0 | 13 | 13 |
| 2007-2010 | 0 | 2 | 1 | 18 | 4 | 0 | 24 | 24 |
| Catch (mt meats $\boldsymbol{y}^{-1}$, landings+ $5 \%$ allowance for incidental mortality) |  |  |  |  |  |  |  |  |
| 2005-2006 | 0 | 1 | 1 | 10 | 2 | 0 | 13 | 13 |
| 2007-2010 | 0 | 2 | 1 | 19 | 4 | 0 | 25 | 25 |
| Initial Biomass |  |  |  |  |  |  |  |  |
| 2005-2006 | 0 | 101 | 402 | 678 | 595 | 1,264 | 1,775 | 3,039 |
| Projected biomass (mt meats) |  |  |  |  |  |  |  |  |
| 2006 | 0 | 98 | 394 | 671 | 590 | 1,264 | 1,753 | 3,016 |
| 2007 | 0 | 95 | 387 | 664 | 585 | 1,264 | 1,731 | 2,995 |
| 2008 | 0 | 92 | 379 | 649 | 578 | 1,264 | 1,697 | 2,961 |
| 2009 | 0 | 88 | 371 | 633 | 571 | 1,264 | 1,663 | 2,927 |
| 2010 | 0 | 85 | 363 | 618 | 564 | 1,264 | 1,630 | 2,893 |
| Projected fishing mortality rate ( $\mathrm{F}_{\boldsymbol{y}}{ }^{-1}$ ) |  |  |  |  |  |  |  |  |
| 2006 | 0.000 | 0.010 | 0.002 | 0.015 | 0.003 | 0.000 | 0.008 | 0.004 |
| 2007 | 0.000 | 0.019 | 0.003 | 0.028 | 0.007 | 0.000 | 0.015 | 0.009 |
| 2008 | 0.000 | 0.020 | 0.003 | 0.029 | 0.007 | 0.000 | 0.015 | 0.009 |
| 2009 | 0.000 | 0.021 | 0.003 | 0.030 | 0.007 | 0.000 | 0.015 | 0.009 |
| 2010 | 0.000 | 0.021 | 0.004 | 0.030 | 0.007 | 0.000 | 0.016 | 0.009 |

Table A29. Projected biomass and fishing mortality for ocean quahog during 2005-2010 based on a 6 million bushel ( $27,215 \mathrm{mt}$ meats) annual quota during 2007-2010. Landings during 2006 are assumed the same as in 2005. Proportions of total catch in each year for each region are the same as in 2005.

| Year | SVA | DMV | NJ | LI | SNE | GBK | Total Less GBK | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Somatic growth rate ( $\mathrm{G}^{-1}$ ) |  |  |  |  |  |  |  |  |
| 2005 | 0.0045 | 0.0000 | 0.0013 | 0.0101 | 0.0066 | 0.0000 | 0.0064 | 0.0037 |
| Recruitment rate ( $r=$ Recruitment / Average Biomass in $2002 \boldsymbol{y}^{-1}$ ) |  |  |  |  |  |  |  |  |
| 2005 | 0.0060 | 0.0000 | 0.0014 | 0.0146 | 0.0081 | 0.0000 | 0.0086 | 0.0050 |
| Natural mortality ( $M^{\text {y }}$-1) |  |  |  |  |  |  |  |  |
| 2005 | 0.0200 | 0.0200 | 0.0200 | 0.0200 | 0.0200 | 0.0000 | 0.0200 | 0.0117 |
| Net instantaneous rate of change, less fishing ( $X-F=G+r-M y^{-1}$ ) |  |  |  |  |  |  |  |  |
| 2005 | -0.0095 | -0.0200 | -0.0174 | 0.0047 | -0.0052 | 0.0000 | -0.0050 | -0.0029 |
| Fishing mortality first year ( $\boldsymbol{F}^{\text {y }}{ }^{-1}$ ) |  |  |  |  |  |  |  |  |
| 2005 | 0.0000 | 0.0094 | 0.0017 | 0.0145 | 0.0034 | 0.0000 | 0.0077 | 0.0045 |
| Landings (mt meats $y^{-1}$ ) |  |  |  |  |  |  |  |  |
| 2005-2006 | 0 | 1 | 1 | 9 | 2 | 0 | 13 | 13 |
| 2007-2010 | 0 | 2 | 1 | 20 | 4 | 0 | 27 | 27 |
| Catch (mt meats $\boldsymbol{y}^{-1}$, landings+ $5 \%$ allowance for incidental mortality) |  |  |  |  |  |  |  |  |
| 2005-2006 | 0 | 1 | 1 | 10 | 2 | 0 | 13 | 13 |
| 2007-2010 | 0 | 2 | 1 | 21 | 4 | 0 | 29 | 29 |
| Initial Biomass |  |  |  |  |  |  |  |  |
| 2005-2006 | 0 | 101 | 402 | 678 | 595 | 1,264 | 1,775 | 3,039 |
| Projected biomass (mt meats) |  |  |  |  |  |  |  |  |
| 2006 | 0 | 98 | 394 | 671 | 590 | 1,264 | 1,753 | 3,016 |
| 2007 | 0 | 95 | 387 | 664 | 585 | 1,264 | 1,731 | 2,995 |
| 2008 | 0 | 91 | 379 | 647 | 577 | 1,264 | 1,694 | 2,957 |
| 2009 | 0 | 88 | 371 | 629 | 570 | 1,264 | 1,657 | 2,921 |
| 2010 | 0 | 84 | 363 | 611 | 563 | 1,264 | 1,620 | 2,884 |
| Projected fishing mortality rate ( $F \boldsymbol{y}^{-1}$ ) |  |  |  |  |  |  |  |  |
| 2006 | 0.000 | 0.010 | 0.002 | 0.015 | 0.003 | 0.000 | 0.008 | 0.004 |
| 2007 | 0.000 | 0.021 | 0.004 | 0.032 | 0.007 | 0.000 | 0.017 | 0.010 |
| 2008 | 0.000 | 0.022 | 0.004 | 0.033 | 0.008 | 0.000 | 0.017 | 0.010 |
| 2009 | 0.000 | 0.023 | 0.004 | 0.034 | 0.008 | 0.000 | 0.017 | 0.010 |
| 2010 | 0.000 | 0.024 | 0.004 | 0.035 | 0.008 | 0.000 | 0.018 | 0.010 |

Table A30. Projected biomass and fishing mortality for ocean quahog during 2005-2010 based on $F=F_{0.1}=0.0278 \mathrm{y}^{-1}$ for exploitable region (total area less GBK) during 2007-2010. Landings during 2006 are assumed the same as in 2005. Proportions of total catch in each year for each region are the same as in 2005.

| Year | SVA | DMV | NJ | LI | SNE | GBK | Total Less |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| GBK |  |  |  |  |  |  |  |$\quad$ Total

Table A31. Summary of example projections.

| Year | Biomass All Regions (1000 $\mathrm{mt})$ | Biomass less GBK (1000 mt) | Landings (1000 $\mathrm{mt})$ | F All Regions $\left(y^{-1}\right)$ | $\begin{gathered} \text { F less GBK } \\ \left(y^{-1}\right) \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Quota $=4$ million bushels (18,144 mt meats) |  |  |  |  |  |
| 2006 | 3,016 | 1,753 | 13 | 0.004 | 0.008 |
| 2007 | 2,995 | 1,731 | 18 | 0.006 | 0.011 |
| 2008 | 2,967 | 1,703 | 18 | 0.006 | 0.011 |
| 2009 | 2,940 | 1,676 | 18 | 0.007 | 0.011 |
| 2010 | 2,912 | 1,649 | 18 | 0.007 | 0.012 |
| Quota $=5.333$ million bushels ( $\mathbf{2 4 , 1 8 9}$ mt meats) |  |  |  |  |  |
| 2006 | 3,016 | 1,753 | 13 | 0.004 | 0.008 |
| 2007 | 2,995 | 1,731 | 24 | 0.009 | 0.015 |
| 2008 | 2,961 | 1,697 | 24 | 0.009 | 0.015 |
| 2009 | 2,927 | 1,663 | 24 | 0.009 | 0.015 |
| 2010 | 2,893 | 1,630 | 24 | 0.009 | 0.016 |
| Quota $=6$ million bushels (27,215 mt meats) |  |  |  |  |  |
| 2006 | 3,016 | 1,753 | 13 | 0.004 | 0.008 |
| 2007 | 2,995 | 1,731 | 27 | 0.010 | 0.017 |
| 2008 | 2,957 | 1,694 | 27 | 0.010 | 0.017 |
| 2009 | 2,921 | 1,657 | 27 | 0.010 | 0.017 |
| 2010 | 2,884 | 1,620 | 27 | 0.010 | 0.018 |
| $F=F_{0.1}=0.028 y^{-1}$ in exploited regions ( $F=0$ for $G B K$ ) |  |  |  |  |  |
| 2006 | 3,016 | 1,753 | 13 | 0.004 | 0.028 |
| 2007 | 2,960 | 1,696 | 44 | 0.016 | 0.028 |
| 2008 | 2,905 | 1,642 | 42 | 0.015 | 0.028 |
| 2009 | 2,853 | 1,589 | 40 | 0.015 | 0.028 |
| 2010 | 2,802 | 1,538 | 39 | 0.015 | 0.028 |

## OCEAN QUAHOG FIGURES



Figure A1. Stock assessment regions for ocean quahog in the US EEZ, with NEFSC shellfish survey strata numbers and boundaries.


Figure A2. Ocean quahog commercial landings (meat weights) from the US EEZ during 19782005. Data for 2005 are preliminary and may be incomplete.


Figure A3. Real and nominal exvessel prices for ocean quahog in the ITQ and Maine fishery components.


Figure A4. Hours fished for ocean quahog in the US EEZ during 1983-2005 based on logbook records.


Figure A5. Number of trips for ocean quahog in the US EEZ during 1991-2004 based on logbook records.


Figure A6. Number of active permits (fishing vessels) for ocean quahog in the US EEZ during 19910-2004 based on logbook records. The total number of permits in the graph for any year may exceed the total number of active permits in the fishery because some vessels fished in more than one area.

## DMV LPUE




Figure A7. Trends in three measures of LPUE for ocean quahog in the DMV (ITQ bushels per hour) and MNE (Maine bushels per hour) stock assessment regions.


Figure A8. Trends in standardized LPUE for ocean quahog during 1980-2005 by stock assessment region.


Figure A9. Trends in standardized LPUE month effects for ocean quahog during 1980-2005 by stock assessment region.


Figure A10. Spatial patterns in average annual landings (1000 ITQ bushels $\mathrm{y}^{-1}$ ) for ocean quahog from logbook records. Data in TNMS far offshore reflect errors in logbook data.


Figure A11. Spatial patterns in average annual fishing effort (hours fished $\mathrm{y}^{-1}$ ) for ocean quahog from logbook records. Data in TNMS far offshore reflect errors in logbook data.


Figure A12. Spatial patterns in average LPUE (ITQ bushels per hours fished) for ocean quahog from logbook records. Data in TNMS far offshore reflect errors in logbook data.


Figure A15. Trends in annual LPUE (ITQ bushels per hours fished, vessel ton class 3-4) for ocean quahog in important TNMS during 1980-2005.


Figure A16. Commercial length composition data for ocean quahog landed in the DMV stock assessment region.


Figure A17. Commercial length composition data for ocean quahog landed in the NJ stock assessment region.

## Long Island



Figure A18. Commercial length composition data for ocean quahog landed in the LI stock assessment region.

Southern New England


Figure A19. Commercial length composition data for ocean quahog landed in the SNE stock assessment region.


Figure A20. NEFSC clam survey trends for ocean quahog stock abundance (mean n/tow), biomass (mean kg/tow), and spawning biomass (mean kg/tow) during 1982-2005. Data for 1994 are omitted because of electrical problems with pump voltage that artificially increased dredge efficiency. Survey data shown in graphs were adjusted based on survey selectivity to estimate trends for the entire stock.


Figure A21. NEFSC clam survey trends for ocean quahog recruit ( $<70 \mathrm{~mm} \mathrm{SL}$ ) abundance (mean n/tow) during 1982-2005. Trends are shown with ("Stock") and without ("Survey") corrections for survey dredge selectivity. Data for 1994 are omitted because of electrical problems with pump voltage that artificially increased dredge efficiency. The apparent outlier for stock $n /$ tow in DMV during 1992 is due to a relatively large catch of small ocean quahog which was increased substantially when adjusted for survey dredge selectivity.


Figure A22. Location and size of recruit ocean quahog ( $<70 \mathrm{~mm}$ ) catches in 2005 NEFSC clam survey, between Long Island and Cape Hatteras.


Figure A23. Location and size of large ocean quahog $(70+\mathrm{mm})$ catches in 2005 NEFSC clam survey, between Long Island and Cape Hatteras.


Figure A24. Location and size of recruit ocean quahog ( $<70 \mathrm{~mm}$ ) catches in 2005 NEFSC clam survey, between Georges Banks and Long Island.


Figure A25. Location and size of large ocean quahog ( $70+\mathrm{mm}$ ) catches in 2005 NEFSC clam survey, between Georges Bank and Long Island.

Delmarva


Figure A26. Length composition for ocean quahog in NEFSC clam surveys, by region. Frequencies are proportional to mean numbers per tow at length, without adjustment for survey dredge selectivity.

New Jersey


Figure A26 (continued)

Long Island


Figure A26 (continued)


Figure A26 (continued)

## George's Bank



Figure A26 (continued)


Figure A27. Fishery and survey selectivity curves for ocean quahog. The ratio of the fishery and survey selectivity curves, which can be used to convert survey abundance at size directly to fishable abundance at size, is also shown.


SSP Ambient Temp.







Figure A28. Survey sensor package data for an NEFSC clam survey tow with acceptable dredge performance.


Figure A28. (continued)


Figure A29. Differential pressure and amperage measured by sensors on the survey dredge during the 2005 NEFSC clam survey. Vertical lines separate the first, second and third legs. Top: Mean values for each station. Bottom: Mean values for each station smoothed by a seven point moving average.


## Sensor tow distance and depth for NEFSC Clam Surveys



Figure A30. Tow distance measurements for NEFSC clam surveys from sensor data (top) and tow distance as a function of depth (bottom). Straight lines in the bottom panel show the best regression model. Curved lines are from loess regression and are intended to show trends.


Figure A31a. Locations of ocean quahog depletion experiments off the Long Island area, 19972005.


Figure A31b. Locations of ocean quahog depletion experiments off the New Jersey-Delmarva area, 1997-2005.


Figure A32. Sensitivity of Patch model estimates of ocean quahog density and dredge efficiency from depletion experiments and the Patch model. All of the experiments shown in the figure except OQ1999-1 (DE-2) were commercial experiments with a 10 ft dredge. The OQ1999-1 (DE-2) experiment was a Delaware II depletion experiment using a 5 ft dredge. The default cell size for Patch model analysis was 20 ft in all cases.


Figure A33. Setup and depletion tows for the OQ2005-1 ocean quahog depletion study. Setup tows by the $R / V$ Delaware II are identified by station numbers. Depletion tows by the F/V Lisa Kim are tightly clustered along parallel tracks. Tow paths appear straight because they are shown as straight lines between start and stop points.


Figure A34. Setup and depletion tows for the OQ2005-2 ocean quahog depletion study. Setup tows by the $R / V$ Delaware II are identified by station numbers. Depletion tows by the F/V Lisa Kim are tightly clustered along parallel tracks. Tow paths appear straight because they are shown as straight lines between start and stop points.


Figure A35. Setup and depletion tows for the OQ2005-3 ocean quahog depletion study. Setup tows by the $R / V$ Delaware II are identified by station numbers. Depletion tows by the F/V Lisa Kim are tightly clustered along parallel tracks. Tow paths appear straight because they are shown as straight lines between start and stop points.


Figure A36. Setup and depletion tows for the OQ2005-4 ocean quahog depletion study. Setup tows by the $R / V$ Delaware II are identified by station numbers. Depletion tows by the F/V Lisa Kim are tightly clustered along parallel tracks. Tow paths appear straight because they are shown as straight lines between start and stop points.


Figure A37. Setup and depletion tows for the OQ2005-6 ocean quahog depletion study. Setup tows by the $R / V$ Delaware II are identified by station numbers. Depletion tows by the F/V Lisa Kim are tightly clustered along parallel tracks. Tow paths appear straight because they are shown as straight lines between start and stop points.



Figure A38. Inclinometer data for selected tows by a commercial dredge at depletion experiment site OQ2005-06, which was carried
 because of low sensor batteries.


Figure A39. Length composition data from setup and depletion tows at a typical 2005 depletion site for ocean quahog (OQ2005-02).


Figure A40. Patch model dredge efficiency estimates vs. depth, estimated density from the Patch model and mean sediment size for ocean quahog in hydraulic dredges used on commercial vessels during depletion studies and the hydraulic dredge used during research surveys by the $F / V$ Delaware II. All data shown in plots on the left hand side are efficiency estimates for commercial vessels used in depletion studies. All data shown in plots on the right hand side are efficiency estimates for the R/V Delaware II based on commercial depletion estimates with setup tows by the Delaware II or, in the case of "DE-2 OQ1999-1", a depletion study carried out directly by the R/V Delaware II.


Figure A41. Survey dredge efficiency estimates for ocean quahog from depletions studies by commercial vessels and by the R/V Delaware II.


Figure A42. Distribution of survey dredge efficiency estimates for ocean quahog from depletion studies by commercial vessels and by the survey vessel ( $R / V$ Delaware II).


Figure A43. Distribution of ocean quahog density estimates $\left(\mathrm{n} \mathrm{ft}^{-2}\right)$ for ocean quahog $90+\mathrm{mm}$ SL from depletion studies by commercial vessels and by the survey vessel $(R / V$ Delaware II).


Figure A44. Uncertainty in efficiency corrected swept area biomass estimates for fishable ocean quahog during 2005. Note that the x -axis differs in the panel for SVA but is the same in all other panels to facilitate comparisons.


Long Island (LI)


Total fishable biomass less GBK


Fishing Mortality ( $\mathrm{F} \mathrm{y}^{-1}$ )



Total fishable biomass


Fishing Mortality ( $\mathrm{F} \mathrm{y}^{-1}$ )

Figure A45. Uncertainty in fishing mortality estimates for ocean quahog during 2005 based on catch data and efficiency corrected swept-area biomass. X-axes are scaled to the same maximum to facilitate comparisons.


Figure A46. Trends in fishable biomass for ocean quahog from the "VPA" model, by region.


Figure A47. Trends in fishable biomass and fishing mortality for ocean quahog from the "VPA" model.


Figure A48. KLAMZ model results for ocean quahog in the DMV stock assessment region. The bottom right panel shows population corrected swept-area biomass data used as prior information is shown in the bottom left panel. Trends in efficiency corrected swept area biomass and LPUE data did not affect model estimates and are shown for comparison only.


Figure A49. KLAMZ model results for ocean quahog in the NJ stock assessment region. The bottom right panel shows population estimates. Other panels show goodness of fit to trend data. The survey scaling parameter estimate for efficiency corrected swept-area biomass data used as prior information is shown in the bottom left panel. Trends in efficiency corrected swept area biomass and LPUE data did not affect model estimates and are shown for comparison only.



Figure A51. KLAMZ model results for ocean quahog in the SNE stock assessment region. The bottom right panel shows population estimates. Other panels show goodness of fit to trend data. The survey scaling parameter estimate for efficiency corrected swept-area biomass data used as prior information is shown in the bottom left panel. Trends in efficiency corrected swept area biomass and LPUE data did not affect model estimates and are shown for comparison only.

Figure A52. Results from a trial run of the KLAMZ model for ocean quahog in the GBK stock assessment region during 1986-2002


Figure A53. Results from a sensitivity run of the KLAMZ model for ocean quahog in the GBK stock assessment region with survey data for 1989 removed.
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Figure A54. Best biomass estimates for ocean quahog in the US EEZ.


Figure A55. Best fishing mortality estimates for the ocean quahog stock in the US EEZ and the total stock less GBK.


Annual growth increments


Annual percent growth in meat weight


Figure A56. Growth, annual growth increments and percent annual change in meat weights for ocean quahog in GBK and in the Mid-Atlantic Bight (MAB) based on von Bertalanffy growth curves. The growth curve for GBK is from Lewis et al. (2001). The growth curve for MAB is used in this assessment for the fishable ocean quahog stock (which excludes GBK).

## Recruits (<70 mm SL)



Figure A57. Trends in survey biomass (no correction for selectivity) for ocean quahog from NEFSC clam surveys during 1986-2002 (1994 omitted due to high pump voltage).


Figure A58. Per recruit model results from a new length based per recruit model and from NEFSC (2004).


Figure A59. Growth, maturity and fishery selectivity curves used in length-based per recruit model used to calculate biological reference points for ocean quahog. Maturity and selectivity (originally functions of length, middle panel) were expressed as functions of age (bottom panel) by inverting the growth curve.


Figure A60. New and old (SARC-38 in NEFSC 2004) maturity and fishery selectivity curves used in per recruit models for ocean
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Figure A61. Best estimates of fishable ocean quahog biomass for the entire ocean quahog stock (top) and fishing mortality for the exploitable stock (excluding GBK) during 2005, with confidence intervals and reference points. The confidence intervals are approximate and based on the CV for the efficiency corrected swept-area biomass estimates for 2005.

## OCEAN QUAHOG APPENDICES

APPENDIX A1. Survey sensor package data from the 2005 NEFSC clam survey. Differential pressure and other data were analyzed to determine if the pump on the survey dredge performed as expected.

## R/V Delaware II Clam Dredge Pump Performance ${ }^{7}$

## Introduction

From an initial review of the Survey Sensor Pack (SSP) data, the dredge pump manifold differential pressure showed a significant variation over the course of the survey's three cruise legs (See Figure 1). This variation was sporadic during the first survey leg with the pressure spikes being attributed to blocked manifold nozzles from visual inspections at the dredge's retrieval. This however, can not explain the consistent upward trend in the manifold differential pressure starting in the middle of the 2nd survey cruise leg which continued to the middle of the 3 rd leg with a then subsequent small falling trend towards the end of the survey. The numerous and sporadic pressure drop spikes that were also noted were not readily explainable by any events that occurred during the survey cruise.


Appendix A1. Figure 1-SSP Manifold Differential Pressure Figure 2-AC Pump Frequency
It was also noted that the frequency recorded also showed a large variation during the ends of the 1st and 2nd survey legs and was consistently higher than the 60 hertz that should have been expected (See Figure 2).

An overheated wire connection on the clam survey package's main breaker was discovered during station 217's tow and temporarily repaired for the remainder of the 2nd survey leg. The clam survey package's main breaker was replaced at the completion of the 2 nd survey leg.

To first investigate these anomalies, a visual inspection of the clam survey sensor data plots for all of the survey tows was done. In particular the Y-Tilt (dredge angle), Manifold Differential Pressure, Pump AC Amps/Volts/Frequency, and Vessel Speed were reviewed. Each tow was graded in an Excel worksheet to summarize the basic characteristics as noted below.

[^7]- Good/Bad Tow or Missing Sensor Data
- Approximate Manifold Differential Pressure
- Manifold Clogging or Pump Intake Blockage
- Erratic Dredge Angle (Y-Tilt); Front Middle, and End of Tow
- Dredge Pump Frequency; Front Middle, and End of Tow
- Tow Speed; Front Middle, and End of Tow
- Did a Low Speed Spike Occur (Tow speed $<\frac{1}{2}$ knot)?

The first discovery is the explanation the sporadic pressure drop spikes in the manifold differential pressure. These pressure drop spikes are likely being caused by a temporary blockage of the pumps intake or the pump ingesting the discharge from the dredge manifold which somehow disrupts the pump's intake flow.


Figure 3 shows a typical tow where this pump intake blockage has likely occurred. Note that there is a corresponding drop in the dredge pump's amps draw as the manifold pressure drops. This is typical for a centrifugal style pump such as is on the clam dredge. The drop in pressure could be minor as in Figure 3 or very substantial as shown in Figure 4. Figure 4 is likely an example of the pump ingesting the manifold discharge as it occurred when a very low speed spike, less than $1 / 2$ knots, also occurred.

The visual inspection of the senor plots also revealed the likely cause for the variation in the general trend of the pump manifold pressure. Using Figures 3 and 4, note that the differential pressures recorded before the pump was started were significantly different. For Figure 3 the starting value is about 5 PSI and for Figure 4 the value is about 15 PSI, a significant difference. Based on this, the following sensor values were graphed on a 10 station interval (those stations with obvious problems were ignored and the next nearest good station was selected, see Figure 5).

Manifold Differential Pressure Before Starting the Dredge Pump.
Manifold Differential Pressure After Starting the Dredge Pump.
Difference Between the After and Before Starting Values (Pump Pressure Rise)


Appendix A1. Figure 5
From Figure 5 the pressure rise in the dredge pump manifold is fairly steady with a consistent downward trend that is typical of a centrifugal pump becoming worn from sand/silt ingestion over the survey. The spikes at stations $49,153,171$, and 231 are likely due to minor clogging of the manifold nozzles as there is a corresponding drop in the amps draw from the pump. This is shown in Figure 6 which also graphs the amps draw, AC voltage, pump power, and tow depth.

Based on this the conclusion is the general performance of the clam dredge pump was fairly uniform over the entire survey and the previous noted variations in the manifold differential pressure are likely due to a calibration drift in the SSP sensor. Interestingly this drift starts to occur at about station 217 , which is when the problem with the main clam package breaker was noticed and repaired. How the breaker problem could cause a sensor drift is not known as the SSP package uses an internal DC battery completely separate from the AC system containing the clam package breaker.


Appendix A1. Figure 6
The variation that occurred in the recorded frequency remains a mystery even after the review of the sensor plots and conversations with the ship's engineer. The value should be very steady and between 59 and 61 hertz which is the output from the ship's generator. Figure 7 shows the typical variation in frequency that occurred during the survey.


The frequency was fairly steady at the start of the survey, and then started a gradual degradation during the last half of the survey's first leg. This degradation in recorded values was not consistent with wide variations between tows. Shortly after the start of the 3rd leg at about station 271, the problem appears to have cleared itself and the frequency was very steady for the remainder of the survey. While there is no direct explanation for this change, it does not to appear to have had any effect on the performance of the clam dredge. The hertz values seen by the pump during the survey are likely have to been the steady standard 59 to 61 hertz values shown on the ship's main switchboard. The changes are likely a problem is in the calibration of the sensor for the frequency not being at 60 hertz and some type of sensor interference for the variations experienced.

The last observation from the sensor plots and data is the occurrence of a rhythmic spike in the AC frequency and volts sensor plots. This occurred throughout the entire survey and a typical example is shown in Figure 8. As with the frequency variation discussed above this appears to be a sensor problem. First it is impossible for a generator to vary its speed as would be shown in the frequency plot. In addition there is no corresponding spikes in the amps or pump pressure that should occur if the volts were truly spiking.


Appendix A1. Figure 8

APPENDIX A2. Clam survey tows with poor performance. This appendix describes a proposal for using sensor data to identify NEFSC clam survey tows with poor performance. Current criteria for identifying tows with poor performance are based on data recorded on deck by the watch chief after each tow. In particular, the survey variable "HAUL" can be used to describe problems with tow duration, and the survey variable "GEARCOND" describes the condition of the dredge after a tow. The proposal described below uses sensor data collected on the dredge and on board the ship. Sensor based criteria could not be applied to data for surveys before 1997 because sensors were not used on the ship. The proposal is for discussion and review and does not represent a recommendation by the Invertebrate Subcommittee.

## NMFS R/V Delaware II Clam Survey Dredge Development of Good/Bad Tow Selection Criteria ${ }^{8}$

## Introduction

From a review of the Survey Sensor Pack (SSP) data from the NMFS 2005 Surf Clam and Ocean Quahog survey, the survey dredge's basic parameters showed a significant variation in the over the course of the survey's three cruise legs. This was primarily both a general upward trend in the manifold's differential pressure and sporadic pressure spikes over the survey (see figure 1). In addition there were occasionally tows that experienced significant variations in the dredge's fore and aft towing angle.


Appendix A2. Figure 1 - Average Survey Dredge Manifold Pressure vs. Survey Station Number
From a previous report (Appendix A2), these parameter variations were explored and their potential effect on the survey dredge's sampling efficiency reviewed. The general upward manifold pressure trend was attributed to a sensor calibration drift, not a true change in manifold pressure, and thus had no likely affect on the dredge's efficiency. The survey tows with manifold pressure spikes and the variations in the dredge's towing angle however were likely causing a significant change in the dredge's sampling efficiency, with the most extreme cases probably preventing the dredge from fishing at all.

Since these survey tows with the manifold pressure spikes and the towing angle variations have a significantly different, and unknown, sampling efficiency than the survey's overall efficiency determined by the depletion studies and other methods,

[^8]inclusion of them in the survey will likely create a bias in the final survey results. Because of this, those survey tows that have some of their key parameters that differ significantly from the normal values should be excluded from the survey as "bad" tows.

## Key Dredge Performance Parameters

The following general parameters are recorded from the SSP and onboard ship sensors for each of the NMFS clam dredge's survey tows.

Tilt-X - Side to side dredge angle.
Tilt-Y - Fore and aft dredge towing angle.
SSP Ambient Temperature - Sea water temperature at the dredge.
SSP Ambient Pressure - Ambient sea water pressure at the dredge (depth).
Differential Pressure - Dredge's water manifold deferential pressure.
AC Amps - Dredge pump's amperage draw.
AC Volts - Dredge pump's voltage.
AC Freq - Dredge pump's frequency.
Vessel Speed - Speed of the DEII
Of these parameters, the two key ones for the dredge's sampling efficiency are;

Tilt-Y - Fore and aft dredge towing angle.
Differential Pressure - Dredge's water manifold differential pressure.
Both of these are the parameters that are directly associated with how the dredge fishes. The Tilt-Y parameter will indicate if the dredge's knife is in sufficient contact with the sea bottom to be in a fishing position. The Differential Pressure indicates if sufficient water is being forced through the dredge's manifold to adequately liquefy the sea bottom.

The AC Amps, AC Volts, and AC Freq are not key parameters as any changes in them will be reflected in the manifold Differential Pressure values. Similarly, Vessel Speed is also not a key parameter in determining a good or bad tow. In this case any vessel speed variations (and thus the survey dredge) are handled in the standardization of each tow to a set "standard" tow distance. SSP Ambient Temperature and Pressure are not key parameters, as they have no effect on overall dredge performance.

| 2005 NMFSNcean Clam <br> Survey Average <br> Dredge Tow Angle <br> Average <br>  <br>  <br>  <br> Towing <br> Angle- |  |
| :---: | :---: |
| Station \# | Degrees |
| 20 | 2.56 |
| 29 | 2.14 |
| 39 | 2.39 |
| 50 | 2.71 |
| 59 | 2.53 |
| 69 | 2.03 |
| 78 | 1.94 |
| 90 | 2.52 |
| 103 | 2.22 |
| 114 | 2.47 |
| 124 | 2.52 |
| 134 | 2.89 |
| 143 | 2.23 |
| 152 | 2.24 |
| 159 | 2.29 |
| 162 | 2.28 |
| 173 | 2.47 |
| 262 | 2.21 |
| 270 | 2.13 |
| 280 | 2.11 |
| 291 | 1.72 |
| 303 | 2.29 |
| 313 | 2.24 |
| 322 | 2.32 |
| 335 | 2.54 |
| Average | 2.32 |
| Average | 0.19 |
| Deviation | 2.29 |
| Median | 2.2 |

The Tilt-Y and Manifold Pressure parameters will each be handled separately, but with a similar method, in determining a good or bad survey tow. A bad tow would then occur when either parameter varies by a specified difference from their normal values.

## Good/Bad Tow Tilt-Y Selection Criteria

The Tilt-Y parameter is a fixed fishing, not fishing (i.e. pass/fail) situation. From previous studies of the NMFS survey dredge the knife theoretically makes contact with the bottom at 4.4 degrees and is fully down at 0 degrees, referenced to the dredge side
runners. For the selection criteria the pass/fail cutout was set at the mid point of 2.2 degrees when the knife is at its half fishing depth in the sea bottom.

The dredge however does not tow with the side runners level as the aft end of the dredge will settle into the trough created in the ocean bottom by the water manifold while the forward dredge end rides on the bottom surface. From the table above this angle is approximately 2.3 degrees. This angle needs to be added to the 2.2 degree pass/fail point above to adjust for the dredge towing angle from the SSP data, which gives an adjusted pass/fail point of 4.5 degrees.

To use this set point, the SSP data will be evaluated by first calculating the total time the dredge Tilt-Y towing angle is above the 4.5 degree set point versus the total time the dredge was on the bottom. The tow will be deemed a bad tow if this time equals or exceeds $20 \%$ of the total towing time. For the four quahog strata survey stations deemed as a bad tow, the resultant time values using the 4.5 degree set point are tabulated below. Based on these Tilt-Y criteria, Station 218 is considered to be a bad tow and should be removed from the survey.

| Good/Bad Dredge Towing Angle Time Summary - Seconds |  |  |  |  |  |
| ---: | :---: | :---: | :---: | :---: | :---: |
| Station \# | 218 | 225 | 262 | 282 |  |
| Time Above 4.5 Degrees | 111 | 120 | 64 | 78 |  |
| Time Below 4.5 Degrees | 337 | 545 | 485 | 469 |  |
| Total Suvey Tow Time | 448 | 665 | 549 | 547 |  |
| Precent Time Above 4.5 Degrees | $24.8 \%$ | $18.0 \%$ | $11.7 \%$ | $14.3 \%$ |  |

## Good/Bad Tow Manifold Pressure Selection Criteria

While the Tilt-Y parameter could be handled as a "Knife Edged" pass/fail selection criteria, this will not work for the Manifold Pressure parameter. First there are two different problem modes that can occur, a manifold pressure above or below the normal value. In addition a linear variation in the pressure doesn't correspond into a linear variation in the water flow through the nozzles.

When the manifold pressure drops below the normal value (37-39 PSI), this is indicating a blocked pump intake which is restricting water flow through the manifold nozzles. A manifold pressure increase on the hand is indicating a blockage in the manifold and/or nozzles. This blockage though is also restricting the water flow through the manifold nozzles. These variations in water flow versus manifold pressure are shown in the graph below.


Because of this non-linearity, the good/bad selection criteria for the Manifold Pressure parameter will need to take into account the magnitude of the difference from normal values. That is the farther the Manifold Pressure value at a given time is from the normal value, the larger the influence that time period will have on the tow being declared a bad tow. This will allow for several different bad tow scenarios to be designated. They are.

1) A small increase or decrease in pressure over the entire tow period.
2) A large increase or decrease in pressure over a short portion of a tow.
3) A combination of small or large pressure variations during a tow.

The selection criteria time period weighting factor (WF) for the Manifold Pressure parameter will be formatted using the following formulas.

WF $=2 \times($ MP-40 $) / 40$ when the Manifold Pressure is Higher than Normal or
WF $=1$ when the Manifold Pressure is in the Normal range or
WF $=2 \times((35-M P) / 35 \times 0.83)$ when the Manifold Pressure is Lower than Normal where MP $=$ SSP measured Manifold Pressure in PSI.

The " 0.83 " is used to bring the potential below value range ( 0 to 35 PSI) into same magnitude as the potential above value range (40 to 69 PSI or 29 PSI range). An average normal Manifold Pressure value of 35-40 PSI was selected based on previous analysis of the 2005 SSP survey data in "R/V Delaware II Clam Dredge Pump Performance" which showed a range in manifold pressure from 39 PSI at the start to 36 PSI at the end of the survey. The doubling of the difference is used to account for the non-linearity by increasing the weighting factor disproportionably for Manifold Pressures farther from the normal value.

For the SSP data the weighting factor will be calculated for each data point which represents a one second time interval. The weighting factors for each second period will then be added to get a total weighted towing time. A bad tow will be declared when this weighted towing time exceeds the actual towing time that was within the normal range by more then $25 \%$. See sample table below for examples.

Based on these Manifold Pressure criteria, Stations 225, 262, and 282 are considered to be a bad tow and should be removed from the survey.

| Good/Bad Manifold Pressure Time Summary |  |  |  |  |  | Seconds |
| ---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Station \# | 218 | 225 | 262 | 282 |  |  |
| Weighted Time Above 40 PSI | 0.00 | 0.13 | 0.00 | 0.00 |  |  |
| Time in Normal Range | 14 | 337 | 190 | 159 |  |  |
| Weighted Time Below 35 PSI | 0.335 | 446.83 | 156.62 | 398.33 |  |  |
| Precent Time Outside Normal | $2.4 \%$ | $132.6 \%$ | $82.4 \%$ | $250.5 \%$ |  |  |






















































































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OQ2000-3 (DM-1)

$190$







Smoothed Position D




















































APPENDIX A5. Technical description of the KLAMZ stock assessment model.

## Larry Jacobson NEFSC, Woods Hole May 25, 2007

The KLAMZ assessment model is based on the Deriso-Schnute delay-difference equation (Deriso 1980; Schnute 1985; Quinn and Deriso 1999). The delay-difference equation is a relatively simple and implicitly age structured approach to counting fish in either numerical or biomass units. It gives the same results as explicitly age-structured models (e.g. Leslie matrix model) if fishery selectivity is "knife-edged", if somatic growth follows the von Bertalanffy equation, and if natural mortality is the same for all age groups in each year. Knife-edge selectivity means that all individuals alive in the model during the same year experience the same fishing mortality rate. ${ }^{9}$ Natural and fishing mortality rates, growth parameters and recruitment may change from year to year, but delay-difference calculations assume that all individuals share the same mortality and growth parameters within each year. The KLAMZ model includes simple numerical models (e.g. Conser 1995) as special cases because growth can be turned off so that all calculations are in numerical units (see below).

As in many other simple models, the delay difference equation explicitly distinguishes between two age groups. In KLAMZ, the two age groups are called "new" recruits ( $R_{t}$ in biomass or numerical units at the beginning of year $t$ ) and "old" recruits $\left(S_{t}\right)$ that together comprise the whole stock $\left(B_{t}\right)$. New recruits are individuals that recruited at the beginning of the current year (at nominal age $k$ ). ${ }^{10}$ Old recruits are all older individuals in the stock (nominal ages $k+1$ and older, survivors from the previous year). As described above, KLAMZ assumes that new and old recruits are fully vulnerable to the fishery. The most important differences between the delay-difference and other simple models (e.g. Prager 1994; Conser 1995; Jacobson et al. 1994) are that von Bertalanffy growth is used to calculate biomass dynamics and that the delaydifference model captures transient age structure effects due to variation in recruitment, growth and mortality exactly. Transient effects on population dynamics are captured exactly because, as described above, the delay-difference equation is algebraically equivalent to an explicitly age-structured model with von Bertalanffy growth.

[^9]The KLAMZ model incorporates a few extensions to Schnute's (1985) revision of Deriso's (1980) original delay difference model. Most of the extensions facilitate tuning to a wider variety of data that anticipated in Schnute (1985). The KLAMZ model is programmed in both Excel and in $\mathrm{C}++$ using AD Model Builder ${ }^{11}$ libraries. The AD Model Builder version is faster, more reliable and probably better for producing "official" stock assessment results. The Excel version is slower and implements fewer features, but the Excel version remains useful in developing prototype assessment models, teaching and for checking calculations.

The most significant disadvantage in using the KLAMZ model and other delaydifference approaches, beyond the assumption of knife-edge selectivity, is that age and length composition data are not used in tuning. However, one can argue that age composition data are used indirectly to the extent they are used to estimate growth parameters or if survey survival ratios (e.g. based on the Heinke method) are used in tuning (see below).

## Population dynamics

The assumed birth date and first day of the year are assumed the same in derivation of the delay-difference equation. It is therefore natural (but not strictly necessary) to tabulate catch and other data using annual accounting periods that start on the assumed biological birthday of cohorts.

## Biomass dynamics

As implemented in the KLAMZ model, Schnute's (1985) delay-difference equation is:

$$
\mathrm{B}_{\mathrm{t}+1}=(1+\rho) \tau_{\mathrm{t}} \mathrm{~B}_{\mathrm{t}}-\rho \tau_{\mathrm{t}} \tau_{\mathrm{t}-1} \mathrm{~B}_{\mathrm{t}-1}+\mathrm{R}_{\mathrm{t}+1}-\rho \tau_{\mathrm{t}} \mathrm{~J}_{\mathrm{t}} \mathrm{R}_{\mathrm{t}}
$$

where $B_{t}$ is total biomass of individuals at the beginning of year $t ; \rho$ is Ford's growth coefficient (see below); $\tau_{t}=\exp \left(-Z_{t}\right)=\exp \left[-\left(F_{t}+M_{t}\right)\right]$ is the fraction of the stock that survived in year $t, Z_{t}, F_{t}$, and $M_{t}$ are instantaneous rates for total, fishing and natural mortality; and $R_{t}$ is the biomass of new recruits (at age $k$ ) at the beginning of the year. The natural mortality rate $M_{t}$ may vary over time. Instantaneous mortality rates in KLAMZ model calculations are biomass-weighted averages if von Bertalanffy growth is turned on in the model. However, biomass-weighted mortality estimates in KLAMZ are the same as rates for numerical estimates under the assumption of knife-edge selectivity because all individuals are fully recruited. The growth parameter $J_{t}=w_{t-1, k-1} / w_{t, k}$ is the ratio of mean weight one year before recruitment (age $k-1$ in year $t-1$ ) and mean weight at recruitment (age $k$ in year $t$ ).

It is not necessary to specify body weights at and prior to recruitment in the KLAMZ model (parameters $v_{t-1}$ and $V_{t}$ in Schnute 1985) because the ratio $J_{t}$ and recruitment biomass contain the same information. Schnute's (1985) original delay difference equation is:

$$
\mathrm{B}_{\mathrm{t}+1}=(1+\rho) \tau_{\mathrm{t}} \mathrm{~B}_{\mathrm{t}}-\rho \tau_{\mathrm{t}} \tau_{\mathrm{t}-1} \mathrm{~B}_{\mathrm{t}-1}+w_{\mathrm{t}+1, \mathrm{k}} \mathrm{~N}_{\mathrm{t}+1}-\rho \tau_{\mathrm{t}} w_{\mathrm{t}-1, \mathrm{k}-1} \mathrm{~N}_{\mathrm{t}}
$$

[^10]To derive the equation used in KLAMZ, substitute recruitment biomass $R_{t+1}$ for the product $w_{t+1, k} N_{t+1, k}$ and adjusted recruitment biomass $J_{t} R_{t}=\left(w_{t-1, k-1} / w_{t, k}\right) w_{t, k} N_{t, k}=$ $w_{t-1, k-1} N_{t}$ in the last term on the right hand side. The advantage in using the alternate parameterization for biomass dynamic calculations in KLAMZ is that recruitment is estimated directly in units of biomass and the number of growth parameters is reduced. The disadvantage is that numbers of recruits are not estimated directly by the model. When required, numerical recruitments must be calculated externally as the ratio of estimated recruitment biomass and the average body weight for new recruits.

## Numerical population dynamics

Growth can be turned on off so that abundance, rather than biomass, is tracked in the KLAMZ model. Set $J_{t}=1$ and $\rho=0$ in the delay difference equation, and use $N_{t}$ (for numbers) in place of $B_{t}$ to get:

$$
\mathrm{N}_{\mathrm{t}+1}=\tau_{\mathrm{t}} \mathrm{~N}_{\mathrm{t}}+\mathrm{R}_{\mathrm{t}+1}
$$

Mathematically, the assumption $J_{i}=1$ means that no growth occurs the assumption $\rho=0$ means that the von Bertalanffy $K$ parameter is infinitely large (Schnute 1985). All tuning and population dynamics calculations in KLAMZ for biomass dynamics are also valid for numerical dynamics.

## Growth

As described in Schnute (1985), biomass calculations in the KLAMZ model are based on Schnute and Fournier's (1980) re-parameterization of the von Bertalanffy growth model:

$$
\mathrm{w}_{\mathrm{a}}=\mathrm{w}_{\mathrm{k}-1}+\left(\mathrm{w}_{\mathrm{k}}-\mathrm{w}_{\mathrm{k}-1}\right)\left(1+\rho^{1+\mathrm{a}-\mathrm{k}}\right) /(1-\rho)
$$

where $w_{k}=V$ and $w_{k-l}=v$. Schnute and Fournier's (1980) growth model is the same as the traditional von Bertalanffy growth model $\left\{W_{a}=W_{\max }\left[1-\exp \left(-K\left(a-t_{z e r o}\right)\right]\right.\right.$ where $W_{\max }, K$ and $t_{\text {zero }}$ are parameters $\}$. The two growth models are the same because $W_{\max }=\left(w_{k}-\rho\right.$ $\left.w_{k-1}\right) /(1-\rho), K=-\ln (\rho)$ and $t_{\text {zero }}=\ln \left[\left(w_{k}-w_{k-1}\right) /\left(w_{k}-\rho w_{k-1}\right)\right] / \ln (\rho)$.

In the KLAMZ model, the growth parameters $J_{t}$ can vary with time but $\rho$ is constant. Use of time-variable $J_{t}$ values with $\rho$ is constant is the same as assuming that the von Bertalanffy parameters $W_{\max }$ and $t_{\text {zero }}$ change over time. Many growth patterns can be mimicked by changing $W_{\max }$ and $t_{\text {zero }}$ (Overholtz et al., 2003). $K$ is a parameter in the $\mathrm{C}++$ version and, in principal, estimable. However, in most cases it is necessary to use external estimates of growth parameters as constants in KLAMZ.

## Instantaneous growth rates

Instantaneous growth rate (IGR) calculations in the KLAMZ model are an extension to the original Deriso-Schnute delay difference model. IGRs are used extensively in KLAMZ for calculating catch biomass and projecting stock biomass forward to the time at which surveys occur. The IGR for new recruits depends only on growth parameters:

$$
G_{t}^{N e w}=\ln \left(\frac{w_{k+1, t+1}}{w_{k, t}}\right)=\ln \left(1+\rho-\rho J_{t}\right)
$$

IGR for old recruits is a biomass-weighted average that depends on the current age structure and growth parameters. It can be calculated easily by projecting biomass of old recruits $S_{t}=B_{t}-R_{t}$ (escapement) forward one year with no mortality:

$$
S_{t}^{*}=(1+\rho) S_{t}-\rho \tau_{t-1} B_{t-1}
$$

where the asterisk $\left({ }^{*}\right)$ means just prior to the start of the subsequent year $t+1$. By definition, the IGR for old recruits in year $t$ is $G_{t}^{\text {Old }}=\ln \left(S_{t}^{*} / S_{t}\right)$. Dividing by $S_{t}$ gives:

$$
G_{t}^{\text {Old }}=\ln \left[(1+\rho)-\rho \tau_{t-1} \frac{B_{t-1}}{S_{t}}\right]
$$

IGR for the entire stock is the biomass weighted average of the IGR values for new and old recruits:

$$
G_{t}=\frac{R_{t} G_{t}^{\text {New }}+S_{t} G_{t}^{\text {Old }}}{B_{t}}
$$

All IGR values are zero if growth is turned off.

## Recruitment

In the Excel version of the KLAMZ model, annual recruitments are calculated $R_{t}=e^{\Omega_{t}}$ where $\Omega_{t}$ is a log transformed annual recruitment parameter, which is estimated in the model. In the $\mathrm{C}++$ version, recruitments are calculated based on $\log$ geometric mean recruitment $(\mu)$ and a set of annual $\log$ scale deviation parameters $\left(\omega_{t}\right)$ :

$$
\Omega_{t}=\mu+\omega_{t}
$$

The deviations $\omega_{t}$ are constrained to average zero. ${ }^{12}$ With the constraint, estimation of $\mu$ and the set of $\omega_{t}$ values ( $1+n$ years parameters) is equivalent to estimation of the smaller set ( $n$ years) of $\Omega_{t}$ values.

## Natural mortality

Natural mortality rates $\left(M_{t}\right)$ are assumed constant in the Excel version of the KLAMZ model. In the C++ version, natural mortality rates may be estimated as a constant value or as a set of values that vary with time. In the model:

$$
M_{t}=m e^{\sigma_{t}}
$$

where $m=\exp (\pi)$ is the geometric mean natural mortality rate, $\pi$ is a model parameter that may be estimated (in principal but not in practical terms), and $\omega_{t}$ is the log scale year-specific deviation. Deviations may be zero (turned off) so that $M_{t}$ is constant, may vary in a random fashion due to autocorrelated or independent process errors, or may be

[^11]based on a covariate. ${ }^{13}$ Model scenarios with zero recruitment may be initializing the parameter $\pi$ to a small value (e.g. $10^{-16}$ ) and not estimating it.

Random natural mortality process errors are effects due to predation, disease, parasitism, ocean conditions or other factors that may vary over time but are not included in the model. Calculations are basically the same as for survey process errors (see below).

Natural mortality rate covariate calculations are similar to survey covariate calculations (see below) except that the user should standardized covariates to average zero over the time period included in the model:

$$
\kappa_{t}=K_{t}-\bar{K}
$$

where $\kappa_{t}$ is the standardized covariate, $K_{t}$ is the original value, and $\bar{K}$ is the mean of the original covariate for the years in the model. Standardization to mean zero is important because otherwise $m$ is not the geometric mean natural mortality rate (the convention is important in some calculations, see text).

Log scale deviations that represent variability around the geometric mean are calculated:

$$
\varpi_{t}=\sum_{j=1}^{n} p_{j} \kappa_{t}
$$

where $n$ is the number of covariates and $p_{j}$ is the parameter for covariate $j$. These conventions mean that the units for the covariate parameter $p_{j}$ are $1 /$ units of the original covariate, the parameter $p_{j}$ measures the log scale effect of changing the covariate by one unit, and the parameter $m$ is the $\log$ scale geometric mean.

## Fishing mortality and catch

Fishing mortality rates $\left(F_{t}\right)$ are calculated so that predicted and observed catch data (landings plus estimated discards in units of weight) "agree" to the extent specified by the user. It is not necessary, however, to assume that catches are measured accurately (see "Observed and predicted catch").

Fishing mortality rate calculations in Schnute (1985) are exact but relating fishing mortality to catch in weight is complicated by continuous somatic growth throughout the year as fishing occurs. The KLAMZ model uses a generalized catch equation that incorporates continuous growth through the fishing season. By the definition of instantaneous rates, the catch equation expresses catch as the product:

$$
\hat{C}_{t}=F_{t} \bar{B}_{t}
$$

where $\hat{C}_{t}$ is predicted catch weight (landings plus discard) and $\bar{B}_{t}$ is average biomass.

[^12]Following Chapman (1971) and Zhang and Sullivan (1988), let $X_{t}=G_{t}-F_{t}-M_{t}$ be the net instantaneous rate of change for biomass. ${ }^{14}$ If the rates for growth and mortality are equal, then $X_{t}=0, \bar{B}_{t}=B_{t}$ and $C_{t}=F_{t} B_{t}$. If the growth rate $G_{t}$ exceeds the combined rates of natural and fishing mortality $\left(F_{t}+M_{t}\right)$, then $X_{t}>0$. If mortality exceeds growth, then $X_{t}<0$. In either case, with $X_{t} \neq 0$, average biomass is computed:

$$
\bar{B}_{t} \approx-\frac{\left(1-e^{X_{t}}\right) B_{t}}{X_{t}}
$$

When $X_{t} \neq 0$, the expression for $\bar{B}_{t}$ is an approximation because $G_{t}$ approximates the rate of change in mean body weight due to von Bertalanffy growth. However, the approximation is reasonably accurate and preferable to calculating catch biomass in the delay-difference model with the traditional catch equation that ignores growth during the fishing season. ${ }^{15}$ Average biomass can be calculated for new recruits, old recruits or for the whole stock by using either $G_{t}^{\text {New }}, G_{t}^{\text {Old }}$ or $G_{t}$.

In the KLAMZ model, the modified catch equation may be solved analytically for $F_{t}$ given $C_{t}, B_{t}, G_{t}$ and $M_{t}$ (see the "Calculating $F_{t}$ " section below). Alternatively, fishing mortality rates can be calculated using a log geometric mean parameter $(\Phi)$ and a set of annual $\log$ scale deviation parameters $\left(\psi_{t}\right)$ :

$$
F_{t}=e^{\Phi+\mu_{t}}
$$

where the deviations $\psi_{t}$ are constrained to average zero. When the catch equation is solved analytically, catches must be assumed known without error but the analytical option is useful when catch is zero or very near zero, or the range of fishing mortality rates is so large (e.g. minimum $\mathrm{F}=0.000001$ to maximum $\mathrm{F}=3$ ) that numerical problems occur with the alternative approach. The analytical approach is also useful if the user wants to reduce the number of parameters estimated by nonlinear optimization. In any case, the two methods should give the same results for catches known without error.

## Surplus production

Annual surplus production is calculated "exactly" by projecting biomass at the beginning of each year forward with no fishing mortality:

$$
\mathrm{B}_{\mathrm{t}}^{*}=(1+\rho) \mathrm{e}^{-\mathrm{M}} \mathrm{~B}_{\mathrm{t}}-\rho \mathrm{e}^{-\mathrm{M}} \mathrm{~L}_{\mathrm{t}-1} \mathrm{~B}_{\mathrm{t}-1}-\rho \mathrm{e}^{-\mathrm{M}} \mathrm{~J}_{\mathrm{t}} \mathrm{R}_{\mathrm{t}}
$$

By definition, surplus production $P_{t}=B^{*}{ }_{t}-B_{t}$ (Jacobson et al. 2002).

## Per recruit modeling

Per recruit model calculations in the Excel version of the KLAMZ simulate the life of a hypothetical cohort of arbitrary size (e.g. $R=1000$ ) starting at age $k$ with constant

[^13]$M_{t}, F$ (survival) and growth ( $\rho$ and $J$ ) in a population initially at zero biomass. In the first year:
$$
\mathrm{B}_{1}=\mathrm{R}
$$

In the second year:

$$
\mathrm{B}_{2}=(1+\rho) \tau \mathrm{B}_{1}-\rho \tau \mathrm{J} \mathrm{R}_{1}
$$

In the third and subsequent years:

$$
\mathrm{B}_{t+1}=(1+\rho) \tau \mathrm{B}_{\mathrm{t}}-\rho \tau^{2} \mathrm{~B}_{\mathrm{t}-1}
$$

This iterative calculation is carried out until the sum of lifetime cohort biomass from one iteration to the next changes by less than a small amount ( 0.0001 ). Total lifetime biomass, spawning biomass and yield in weight are calculated by summing biomass, spawning biomass and yield over the lifetime of the cohort. Lifetime biomass, spawning biomass and yield per recruit are calculated by dividing totals by initial recruitment $(R)$.

## Status determination variables

The user may specify a range of years (e.g. the last three years) to use in calculating recent average fishing mortality $\bar{F}_{\text {Recent }}$ and biomass $\bar{B}_{\text {Recent }}$ levels. These status determination variables are used in calculation of status ratios such as $\bar{F}_{\text {Recent }} / F_{\text {MSY }}$ and $\bar{B}_{\text {Re cent }} / \mathrm{B}_{\mathrm{MSY}}$.

## Goodness of Fit and Parameter Estimation

Parameters estimated in the KLAMZ model are chosen to minimize an objective function based on a sum of weighted negative log likelihood (NLL) components:

$$
\Xi=\sum_{v=1}^{N_{\Xi}} \lambda_{v} L_{v}
$$

where $N_{\Xi}$ is the number of NLL components $\left(L_{v}\right)$ and the $\lambda_{v}$ are emphasis factors used as weights. The objective function $\Xi$ may be viewed as a NLL or a negative $\log$ posterior (NLP) distribution, depending on the nature of the individual $L_{v}$ components and modeling approach. Except during sensitivity analyses, weighting factors for objective function components ( $\lambda_{v}$ ) are usually set to one. An arbitrarily large weighting factor (e.g. $\lambda_{v}=1000$ ) is used for "hard" constraints that must be satisfied in the model. Arbitrarily small weighting factors (e.g. $\lambda_{v}=0.0001$ ) can be used for "soft" model-based constraints. For example, an internally estimated spawner-recruit curve or surplus production curve might be estimated with a small weighting factor to summarize stockrecruit or surplus production results with minimal influence on biomass, fishing mortality and other estimates from the model. Use of a small weighting factor for an internally estimated surplus production or stock-recruit curve is equivalent to fitting a curve to model estimates of biomass and recruitment or surplus production in the output file, after the model is fit (Jacobson et al. 2002).

## Likelihood component weights vs. observation-specific weights

Likelihood component weights $\left(\lambda_{v}\right)$ apply to entire NLL components. Entire components are often computed as the sum of a number of individual NLL terms. The NLL for an entire survey, for example, is composed of NLL terms for each of the annual survey observations. In KLAMZ, observation-specific (for data) or instance-specific (for constraints or prior information) weights (usually $w_{j}$ for observation or instance $j$ ) can be specified as well. Observation-specific weights for a survey, for example, might be use to increase or decrease the importance of one or more observations in calculating goodness of fit.

## NLL kernels

NLL components in KLAMZ are generally programmed as "concentrated likelihoods" to avoid calculation of values that do not affect derivatives of the objective function. ${ }^{16}$ For $x \sim \mathrm{~N}\left(\mu, \sigma^{2}\right)$, the complete NLL for one observation is:

$$
L=\ln (\sigma)+\ln (\sqrt{2 \pi})+0.5\left(\frac{x-u}{\sigma}\right)^{2}
$$

The constant $\ln (\sqrt{2 \pi})$ can always be omitted because it does not affect derivatives. If the standard deviation is known or assumed known, then $\ln (\sigma)$ can be omitted as well because it is a constant that does not affect derivatives. In such cases, the concentrated negative log likelihood is:

$$
L=0.5\left(\frac{x-\mu}{\sigma}\right)^{2}
$$

If there are $N$ observations with possible different variances (known or assumed known) and possibly different expected values:

$$
L=0.5 \sum_{i=1}^{N}\left(\frac{x_{i}-\mu_{i}}{\sigma_{i}}\right)^{2}
$$

If the standard deviation for a normally distributed quantity is not known and is (in effect) estimated by the model, then one of two equivalent calculations is used. Both approaches assume that all observations have the same variance and standard deviation. The first approach is used when all observations have the same weight in the likelihood:

$$
L=0.5 N \ln \left[\sum_{i=1}^{N}\left(x_{i}-u\right)^{2}\right]
$$

where $N$ is the number of observations. The second approach is equivalent but used when the weights for each observation $\left(w_{i}\right)$ may differ:

$$
L=\sum_{i=1}^{N} w_{i}\left[\ln (\sigma)+0.5\left(\frac{x_{i}-u}{\sigma}\right)^{2}\right]
$$

In the latter case, the maximum likelihood estimator:

[^14]$$
\hat{\sigma}=\sqrt{\frac{\sum_{i=1}^{N}\left(x_{i}-\hat{x}\right)^{2}}{N}}
$$
(where $\hat{x}$ is the average or predicted value from the model) is used for $\sigma$. The maximum likelihood estimator is biased by $N /\left(N-d_{f}\right)$ where $d_{f}$ is degrees of freedom for the model. The bias may be significant for small sample sizes but $d_{f}$ is usually unknown.

## Landings, discards, catch

Discards are from external estimates $\left(d_{t}\right)$ supplied by the user. If $d_{t} \geq 0$, then the data are used as the ratio of discard to landed catch so that:

$$
D_{t}=L_{t} \Delta_{t}
$$

where $\Delta_{t}=D_{t} / L_{t}$ is the discard ratio. If $d_{t}<0$ then the data are treated as discard in units of weight:

$$
D_{t}=\operatorname{abs}\left(d_{t}\right) .
$$

In either case, total catch is the sum of discards and landed catch $\left(C_{t}=L_{t}+D_{t}\right)$. It is possible to use discards in weight $d_{t}<0$ for some years and discard as proportions $d_{t}>0$ for other years in the same model run. If catches are estimated (see below) so that the estimated catch $\hat{C}_{t}$ does not necessarily equal observed landings plus discard, then estimated landings are computed:

$$
\hat{L}_{t}=\frac{\hat{C}_{t}}{1+\Delta_{t}}
$$

and estimated discards are:

$$
\hat{D}_{t}=\Delta_{t} \hat{L}_{t} .
$$

## Calculating $\mathrm{F}_{\mathrm{t}}$

As described above, fishing mortality rates may be estimated based on the parameters $\Phi$ and $\psi_{t}$ to satisfy a NLL for observed and predicted catches:

$$
L=0.5 \sum_{t=0}^{N} w_{t}\left(\frac{\hat{C}_{t}-C_{t}}{\kappa_{t}}\right)^{2}
$$

where the standard error $\kappa_{t}=C V_{\text {catch }} \hat{C}_{t}$ with $C V_{\text {catch }}$ and weights are $w_{t}$ supplied by the user. The weights can be used, for example, if catch data in some years are less precise than in others. Using observation specific weights, any or every catch in the time series can potentially be estimated.

The other approach to calculating $F_{t}$ values is by solving the generalized catch equation (see above) iteratively. Subtracting predicted catch from the generalized catch equation gives:

$$
g\left(F_{t}\right)=C_{t}+\frac{F_{t}\left(1-e^{X_{t}}\right)}{X_{t}} B_{t}=0
$$

where $X_{t}=G_{t}-M_{t}-F_{t}$. If $X_{t}=0$, then $\bar{B}_{t}=B_{t}$ and $F_{t}=C_{t} / B_{t}$.

If $X_{t} \neq 0$, then the Newton-Raphson algorithm is used to solve for $F_{t}$ (Kennedy and Gentle 1980). At each iteration of the algorithm, the current estimate $F_{t}^{i}$ is updated using:

$$
F_{t}^{i+1}=F_{t}^{i}-\frac{g\left(F_{t}^{i}\right)}{g^{\prime}\left(F_{t}^{i}\right)}
$$

where $g^{\prime}\left(F_{t}^{i}\right)$ is the derivative $F_{t}^{i}$. Omitting subscripts, the derivative is:

$$
g^{\prime}(F)=-\frac{B e^{-F}\left[\left(e^{F}-e^{\gamma}\right) \gamma+e^{\gamma} F \gamma-e^{\gamma} F^{2}\right]}{X^{2}}
$$

where $\gamma=G-M_{t}$. Iterations continue until $g\left(F_{t}^{i}\right)$ and $\operatorname{abs}\left[g\left(F_{t}^{i+1}\right)-g\left(F_{t}^{i+1}\right)\right]$ are both less than a small number (e.g. $\leq 0.00001$ ).

Initial values are important in algorithms that solve the catch equation numerically (Sims 1982). If $M_{t}+F_{t}>G_{t}$ so that $X_{t}<0$, then the initial value $F_{t}^{0}$ is calculated according to Sims (1982). If $M_{t}+F_{t}<G_{t}$ so that $X_{t}>0$, then initial values are calculated based on a generalized version of Pope's cohort analysis (Zhang and Sullivan 1988):

$$
F_{t}^{0}=\gamma_{t}-\ln \left[\frac{\left(B_{t} e^{0.5 \gamma_{t}}-C_{t}\right) e^{0.5 \gamma_{t}}}{B_{t}}\right]
$$

## F for landings versus $F$ for discards

The total fishing mortality rate for each year can be partitioned into a component due to landed catch ${ }^{L} F_{t}=\frac{D_{t}}{C_{t}} F_{t}$, and a component due to discard ${ }^{D} F_{t}=\frac{L_{t}}{C_{t}} F_{t}$.

## Predator consumption as discard data

In modeling population dynamics of prey species, estimates of predator consumption can be treated like discard in the KLAMZ model as a means for introducing time dependent natural mortality. Consider a hypothetical example with consumption data $\left(\mathrm{mt} \mathrm{y}^{-1}\right)$ for three important predators. If the aggregate consumption data are included in the model as "discards", then the fishing mortality rate for discards ${ }^{d} F_{t}$ (see above) would be an estimate of the component of natural mortality due to the three predators. In using this approach, the average level of natural mortality $m$ would normally be reduced (e.g. so that $m_{\text {new }}+{ }^{d} \bar{F}=m_{\text {old }}$ ) or estimated to account for the portion of natural mortality attributed to bycatch.

Surplus production calculations are harder to interpret if predator consumption is treated as discard data because surplus production calculations assume that $F_{i}=0$ (see above) and because surplus production is defined as the change in biomass from one year to the next in the absence of fishing (i.e. no landings or bycatch). However, it may be useful to compare surplus production at a given level of biomass from runs with and without consumption data as a means of estimating maximum changes in potential fishery yield if the selected predators were eliminated (assuming no change in disease, growth rates, predation by other predators, etc.).

## Effort calculations

Fishing mortality rates can be tuned to fishing effort data for the "landed" catch (i.e. excluding discards). Years with non-zero fishing effort used in the model must also have landings greater than zero. Assuming that effort data are lognormally distributed, the NLL for fishing effort is:

$$
N L L=0.5 \sum_{y=1}^{n_{\text {eff }}} w_{y}\left[\frac{\ln \left(E_{y} / \hat{E}_{y}\right)}{\sigma}\right]^{2}
$$

where $w_{y}$ is an observation-specific weight, $n_{\text {eff }}$ is the number of active effort observations (i.e. with $w_{y}>0$ ), $E_{y}$ and $\bar{E}_{y}$ are observed and predicted fishing effort data, and the log scale variance $\sigma$ is a constant calculated from a user-specified CV.

Predicted fishing effort data are calculated:

$$
\hat{E}_{y}=\zeta F_{y}^{\vartheta}
$$

where $\zeta=e^{u}, \vartheta=e^{b}$, and $u$ and $b$ are parameters estimated by the model. If the parameter b is not estimated, then $\vartheta=1$ so that the relationship between fishing effort and fishing mortality is linear. If the parameter $b$ is estimated, then $\vartheta \neq 1$ and the relationship is a power function.

## Predator data as fishing effort

As described under "Predator consumption as discard data", predator consumption data can be treated as discard. If predator abundance data are available as well, and assuming that mortality due predators is a linear function of the predator-prey ratio, then both types of data may be used together to estimate natural mortality. The trick is to: 1) enter the predator abundance data as fishing effort; 2) enter the actual fishery landings as "discard"; 3) enter predator consumption estimates of the prey species as "landings" so that the fishing effort data in the refer to the predator consumption data; 4) use an option in the model to calculate the predator-prey ratio for use in place of the original predator abundance "fishing effort" data; and 5) tune fishing mortality rates for landings (a.k.a. predator consumption) to fishing effort (a.k.a. predator-prey ratio).

Given the predator abundance data $\kappa_{y}$, the model calculates the predator-prey ratio used in place of fishing effort data $\left(E_{y}\right)$ as:

$$
E_{y}=\frac{\kappa_{y}}{B_{y}}
$$

where $B_{y}$ is the model's current estimate of total (a.k.a "prey") biomass. Subsequent calculations with $E_{y}$ and the model's estimates of "fishing mortality" $\left(F_{y}\right.$, really a measure of natural mortality) are exactly as described above for effort data. In using this approach, it is probably advisable to reduce $m$ (the estimate of average mortality in the model) to account for the proportion of natural mortality due to predators included in the calculation. Based on experience to date, natural mortality due to consumption by the suite of predators can be estimated but only if $m$ is assumed known.

## Initial population age structure

In the KLAMZ model, old and new recruit biomass during the first year ( $R_{l}$ and $\left.S_{I}=B_{I}-R_{I}\right)$ and biomass prior to the first year $\left(B_{0}\right)$ are estimated as log scale parameters. Survival in the year prior to the first year ("year 0 ") is $\tau_{0}=e^{-F_{0}-M_{1}}$ with $F_{0}$ chosen to obtain catch $C_{0}$ (specified as data) from the estimated biomass $B_{0}$. IGRs during year 0 and year 1 are assumed equal $\left(G_{0}=G_{l}\right)$ in catch calculations.

Biomass in the second year of as series of delay-difference calculations depends on biomass $\left(B_{0}\right)$ and survival $\left(\tau_{0}\right)$ in year 0 :

$$
\mathrm{B}_{2}=(1+\rho) \tau_{1} \mathrm{~B}_{1}-\rho \tau_{1} \tau_{0} \mathrm{~B}_{0}+\mathrm{R}_{2}-\rho \tau_{1} \mathrm{~J}_{1} \mathrm{R}_{1}
$$

There is, however, there is no direct linkage between $B_{0}$ and escapement biomass ( $S_{l^{\prime}}=B_{1^{-}}$ $R_{I}$ ) at the beginning of the first year.

The missing link between $B_{0}, S_{1}$ and $B_{1}$ means that the parameter for $B_{0}$ tends to be relatively free and unconstrained by the underlying population dynamics model. In some cases, $B_{0}$ can be estimated to give good fit to survey and other data, while implying unreasonable initial age composition and surplus production levels. In other cases, $B_{0}$ estimates can be unrealistically high or low implying, for example, unreasonably high or low recruitment in the first year of the model $\left(R_{I}\right)$. Problems arise because many different combinations of values for $R_{l}, S_{l}$ and $B_{0}$ give similar results in terms of goodness of fit. This issue is common in stock assessment models that use forward simulation calculations because initial age composition is difficult to estimate. It may be exacerbated in delay-difference models because age composition data are not used.

The KLAMZ model uses two constraints to help estimate initial population biomass and initial age structure. ${ }^{17}$ The first constraint links IGRs for escapement ( $G^{\text {Old }}$ ) in the first years to a subsequent value. The purpose of the constraint is to ensure consistency in average growth rates (and implicit age structure) during the first few years. For example, if IGRs for the first $n_{G}$ years are constrained ${ }^{18}$, then the NLL for the penalty is:
$L_{G}=0.5 \sum_{t=1}^{n_{G}}\left[\frac{\ln \left(G_{t}^{\text {old }} / G_{n_{G}+1}^{\text {old }}\right)}{\sigma_{G}}\right]^{2}$
where the standard deviation $\sigma_{G}$ is supplied by the user. It is usually possible to use the standard deviation of $Q_{t}^{\text {old }}$ for later years from a preliminary run to estimate $\sigma_{G}$ for the first few years. The constraint on initial IGRs should probably be "soft" and non-binding $(\lambda \approx 1)$ because there is substantial natural variation in somatic growth rates due to variation in age composition.

The second constraint links $B_{0}$ to $S_{I}$ and ensures conservation of mass in population dynamics between years 0 and 1 . In other words, the parameter for escapement biomass in year 1 is constrained to match an approximate projection of the biomass in year 0 , accounting for growth, and natural and fishing mortality. The constraint is intended to be binding and satisfied exactly (e.g. $\lambda=1000$ ) because incompatible values of $S_{I}$ and $B_{0}$ are biologically impossible. In calculations:

[^15]$$
S_{1}^{p}=B_{0} e^{G_{1}-F_{0}-M_{1}}
$$
where $S_{1}^{p}$ is the projected escapement in year 1 and $B_{0}$ is the model's estimate of total biomass in year 0 . The instantaneous rates for growth and natural mortality from year 1 ( $G_{l}$ and $M_{l}$ ) are used in place of $G_{0}$ and $M_{0}$ because the latter are unavailable. The NLL for the constraint:
$$
L=\left[\ln \left(\frac{S_{1}^{p}}{S_{1}}\right)^{2}\right]^{2}+\left(S_{1}^{p}-S_{1}\right)^{2}
$$
uses a log scale sum of squares and an arithmetic sum of squares. The former is effective when $S_{l}$ is small while the latter is effective when $S_{l}$ is large. Constants and details in calculation of NLL for the constraint are not important because the constraint is binding (e.g. $\lambda=1000$ ).

## Equilibrium pristine biomass

It may be useful to constrain the biomass estimate for the first year in a model run towards an estimate of equilibrium pristine biomass if, for example, stock dynamics tend to be stable and catch data are available for the first years of the fishery, or as an alternative to the approach described above for initializing the age structure of the simulated population in the model. Equilibrium pristine biomass $\widetilde{B}_{0}$ is calculated based on the model's estimate of average recruitment and with no fishing mortality (calculations are similar to those described under "Per-recruit modeling" except that average recruitment is assumed in each year). ${ }^{19}$ The NLL term for the constraint is:

$$
L=\ln \left(\frac{\widetilde{B}_{0}}{B_{0}}\right)^{2}
$$

Pristine equilibrium biomass is used as a hard constraint with a high emphasis factor ( $\lambda$ ) so that the variance and constants normally used in NLL calculations are not important.

## Estimating natural mortality

As described above, natural mortality calculations involve a parameter for the geometric mean value ( $m$ ) and time dependent deviations ( $\omega_{t}$, which may or may not be turned on). Constraints on natural mortality process errors and natural mortality covariates can be used to help estimate the time dependent deviations and overall trend. The geometric mean natural mortality rate is usually difficult to estimate and best treated as a known constant. However, in the C++ version of the KLAMZ model, $m=e^{\pi}$ (where $\pi$ is an estimable parameter in the model) and estimates of $m$ can be conditioned on the constraint:

$$
L=0.5\left[\frac{\ln \left(w / w_{T \text { arget }}\right)}{\sigma_{\bar{\sigma}}}\right]^{2}
$$

[^16]where $w_{\text {Target }}$ is a user supplied mean or target value and $\sigma_{\bar{\pi}}$ is a log scale standard deviation. The standard deviation is calculated from an arithmetic scale CV supplied by the user. Upper and lower bounds for $m$ may be specified as well.

## Goodness of fit for trend data

Assuming lognormal errors ${ }^{20}$, the NLL used to measure goodness-of-fit to "survey" data that measure trends in abundance or biomass (or survival, see below) is:

$$
L=0.5 \sum_{j=1}^{N_{v}}\left[\frac{\ln \left(I_{v, j} / \hat{I}_{v, j}\right)}{\sigma_{v, j}}\right]^{2}
$$

where $I_{v, t}$ is an index datum from survey $v$, hats " "»" denote model estimates, $\sigma_{v, j}$ is a log scale standard error (see below), and $N_{v}$ is the number of observations. There are two approaches to calculating standard errors for $\log$ normal abundance index data in KLAMZ and it is possible to use different approaches for different types of abundance index data in the same model (see below).

## Standard errors for goodness of fit

In the first approach, all observations for one type of abundance index share the same standard error, which is calculated based on overall goodness of fit. This approach implicitly estimates the standard error based on goodness of fit, along with the rest of the parameters in the model (see "NLL kernels" above).

In the second approach, each observation has a potentially unique standard error that is calculated based on its CV. The second approach calculates log scale standard errors from arithmetic CVs supplied as data by the user (Jacobson et al. 1994):

$$
\sigma_{v, t}=\sqrt{\ln \left(1+C V_{v, t}^{2}\right)}
$$

Arithmetic CV's are usually available for abundance data. It may be convenient to use $C V_{v, t}=1.31$ to get $\sigma_{v, t}=1$.

There are advantages and disadvantages to both approaches. CV's carry information about the relative precision of abundance index observations. However, CV's usually overstate the precision of data as a measure of fish abundance ${ }^{21}$ and may be misleading in comparing the precision of one sort of data to another as a measure of trends in abundance (e.g. in contrasting standardized LPUE that measure fishing success, but not abundance, precisely with survey data that measure trends in fish abundance

[^17]directly, but not precisely). Standard errors estimated implicitly are often larger and more realistic, but assume that all observations in the same survey are equally reliable.

## Predicted values for abundance indices

Predicted values for abundance indices are calculated:

$$
\hat{I}_{v, t}=Q_{v} A_{v, t}
$$

where $Q_{v}$ is a survey scaling parameter (constant here but see below) that converts units of biomass to units of the abundance index. $A_{v, t}$ is available biomass at the time of the survey.

In the simplest case, available biomass is:

$$
A_{v, t}=s_{v, N e w} R_{t} e^{-X_{t}^{N e w} \Delta_{v, t}}+s_{v, O l d} S_{t} e^{-X_{t}^{o l d} \Delta_{v, t}}
$$

where $s_{v, \text { New }}$ and $s_{v, \text { Old }}$ are survey selectivity parameters for new recruits $\left(R_{t}\right)$ and old recruits $\left(S_{t}\right) ; X_{t}^{\text {New }}=G_{t}^{\text {New }}-F_{t}-M_{t}$ and $X_{t}^{\text {Old }}=G_{t}^{\text {Old }}-F_{t}-M_{t} ; j_{v, t}$ is the Julian date at the time of the survey, and $\Delta_{v, t}=j_{v, t} / 365$ is the fraction of the year elapsed at the time of the survey.

Survey selectivity parameter values ( $s_{v, N e w}$ and $s_{v, O l d}$ ) are specified by the user and must be set between zero and one. For example, a survey for new recruits would have $s_{v, \text { New }}=1$ and $s_{v, \text { Old }}=0$. A survey that measured abundance of the entire stock would have $s_{v, \text { New }}=1$ and $s_{v, \text { Old }}=1$.

Terms involving $\Delta_{v, t}$ are used to project beginning of year biomass forward to the time of the survey, making adjustments for mortality and somatic growth. ${ }^{22}$ As described below, available biomass $A_{v, t}$ is adjusted further for nonlinear surveys, surveys with covariates and surveys with time variable $Q_{v, t}$.

## Scaling parameters (Q) for log normal abundance data

Scaling parameters for surveys with lognormal statistical errors were computed using the maximum likelihood estimator:

$$
Q_{v}=e^{\frac{\sum_{i=1}^{N_{v}}\left[\ln \left(\frac{I_{v, i}}{A_{v, i}}\right) / \sigma_{v, j}^{2}\right]}{\sum_{j=1}^{N_{j}}\left(1 / \sigma_{v, j}^{2}\right)}}
$$

where $N_{v}$ is the number of observations with individual weights greater than zero. The closed form maximum likelihood estimator gives the same answer as if scaling parameters are estimated as free parameters in the assessment model assuming lognormal survey measurement errors.

[^18]
## Survey covariates

Survey scaling parameters may vary over time based on covariates in the KLAMZ model. The survey scaling parameter that measures the relationship between available biomass and survey data becomes time dependent:

$$
\hat{I}_{v, t}=Q_{v, t} A_{v, t}
$$

and

$$
Q_{v, t}=Q_{v} e^{\sum_{r=1}^{n_{n}} d_{r, t} \theta_{r}}
$$

with $n_{v}$ covariates for the survey and parameters $\theta_{r}$ estimated in the model. Covariate effects and available biomass are multiplied to compute an adjusted available biomass:

$$
A_{v, t}^{\prime}=A_{v, t} e^{\sum_{r=1}^{n_{i}} d_{r, t} \theta_{r}}
$$

The adjusted available biomass $A_{v, t}^{\prime}$ is used instead of the original value $A_{v, t}$ in the closed form maximum likelihood estimator described above.

Covariates might include, for example, a dummy variable that represents changes in survey bottom trawl doors or a continuous variable like average temperature data if environmental factors affect distribution and catchability of fish schools. Dummy variables are usually either 0 or 1 , depending on whether the effect is present in a particular year. With dummy variables, $Q_{v}$ is the value of the survey scaling parameter with no intervention $\left(d_{r, t}=0\right)$.

For ease in interpretation of parameter estimates for continuous covariates (e.g. temperature data), it is useful to center covariate data around the mean:

$$
d_{r, t}=d_{r, t}^{\prime}-\overline{d_{r}^{\prime}}
$$

where $d_{r, t}^{\prime}$ is the original covariate. When covariates are continuous and mean-centered, $Q_{\nu}$ is the value of the survey scaling parameter under average conditions ( $d_{r, t}=0$ ) and units for the covariate parameter are easy to interpret (for example, units for the parameter are $1 /{ }^{\circ} \mathrm{C}$ if the covariate is mean centered temperature in ${ }^{\circ} \mathrm{C}$ ).

It is possible to use a survey covariate to adjust for differences in relative stock size from year to year due to changes in the timing of a survey. However, this adjustment may be made more precisely by letting the model calculate $\Delta_{v, t}$ as described above, based on the actual timing data for the survey during each year.

## Nonlinear abundance indices

With nonlinear abundance indices, and following Methot (1990), the survey scaling parameter is a function of available biomass:

$$
Q_{v, t}=Q_{v} A_{v, t}^{\Gamma}
$$

so that:

$$
\hat{I}_{v, t}=\left(Q_{v} A_{v, t}^{\Gamma}\right) A_{v, t}
$$

Substituting $e^{\gamma}=\Gamma+1$ gives the equivalent expression:

$$
\hat{I}_{v, t}=Q_{v} A_{v, t}^{e^{v}}
$$

where $\gamma$ is a parameter estimated by the model and the survey scaling parameter is no longer time dependent. In calculations with nonlinear abundance indices, the adjusted available biomass:

$$
A_{v, t}^{\prime}=A_{v, t}^{e^{\prime}}
$$

is computed first and used in the closed form maximum likelihood estimator described above to calculate the survey scaling parameter. In cases where survey covariates are also applied to a nonlinear index, the adjustment for nonlinearity is carried out first.

## Survey $Q$ process errors

The C++ version of the KLAMZ model can be used to allow survey scaling parameters to change in a controlled fashion from year to year (NEFSC 2002):

$$
Q_{v, t}=Q_{v} e^{\varepsilon_{v, t}}
$$

where the deviations $\varepsilon_{v, t}$ are constrained to average zero. Variation in survey Q values is controlled by the NLL penalty:

$$
L=0.5 \sum_{j=1}^{N}\left[\frac{\varepsilon_{v, j}}{\sigma_{v}}\right]^{2}
$$

where the $\log$ scale standard deviation $\sigma_{\nu}$ based on an arithmetic CV supplied by the user (e.g. see NEFSC 2002). In practice, the user increases or decreases the amount of variability in Q by decreasing or increasing the assumed CV .

## Survival ratios as surveys

In the C++ version of KLAMZ, it is possible to use time series of survival data as "surveys". For example, an index of survival might be calculated using survey data and the Heinke method (Ricker 1975) as:

$$
A_{t}=\frac{I_{k+1, t+1}}{I_{k, t}}
$$

so that the time series of $A_{t}$ estimates are data that may potentially contain information about scale or trends in survival. Predicted values for a survival index are calculated:

$$
\hat{A}_{t}=e^{-Z_{t}}
$$

After predicted values are calculated, survival ratio data are treated in the same way as abundance data (in particular, measurement errors are assumed to be lognormal). Selectivity parameters are ignored for survival data but all other features (e.g. covariates, nonlinear scaling relationships and constraints on $Q$ ) are available.

## Recruitment models

Recruitment parameters in KLAMZ may be freely estimated or estimated around an internal recruitment model, possibly involving spawning biomass. An internally estimated recruitment model can be used to reduce variability in recruitment estimates (often necessary if data are limited), to summarize stock-recruit relationships, or to make use of information about recruitment in similar stocks. There are four types of internally estimated recruitment models in KLAMZ: 1) random variation around a constant mean; 2) random walk around a constant mean (autocorrelated variation); 3) random variation around a Beverton-Holt recruitment model; and 4) random variation around a Ricker recruitment model. The user must specify a type of recruitment model but the model is not active unless the likelihood component for the recruitment model is turned on ( $\lambda>0$ ).

The first step in recruit modeling is to calculate the expected log recruitment level $E\left[\ln \left(R_{t}\right)\right]$ given the recruitment model. For random variation around a constant mean, the expected $\log$ recruitment level is the log geometric mean recruitment:

$$
E\left[\ln \left(R_{t}\right)\right]=\sum_{j=1}^{N} \ln \left(R_{j}\right) / N
$$

For a random walk around a constant mean recruitment, the expected log recruitment level is the logarithm of recruitment during the previous year:

$$
E\left[\ln \left(R_{t}\right)\right]=\ln \left(R_{t-1}\right)
$$

with no constraint on recruitment during the first year $R_{l}$.
For the Beverton-Holt recruitment model, the expected log recruitment level is:

$$
E\left[\ln \left(R_{t}\right)\right]=\ln \left[e^{a} T_{t-\ell} /\left(e^{b}+T_{t-\ell}\right)\right]
$$

where $a=e^{\alpha}$ and $b=e^{\beta}$, the parameters $\alpha$ and $\beta$ are estimated in the model, $T_{t}$ is spawning biomass, and $\{$ is the lag between spawning and recruitment. Spawner-recruit parameters are estimated as log transformed values ( $e^{\alpha}$ and $e^{\beta}$ ) to enhance model stability and ensure the correct sign of values used in calculations. Spawning biomass is:

$$
T_{t}=m_{\text {new }} R_{t}+m_{\text {old }} S_{t}
$$

where $m_{\text {new }}$ and $m_{\text {old }}$ are maturity parameters for new and old recruits specified by the user. For the Ricker recruitment model, the expected log recruitment level is:

$$
E\left[\ln \left(R_{t}\right)\right]=\ln \left(S_{t-\ell} e^{a-b S_{t-\ell}}\right)
$$

where $a=e^{\alpha}$ and $b=e^{\beta}$, and the parameters $\alpha$ and $\beta$ are estimated in the model.
Given the expected $\log$ recruitment level, $\log$ scale residuals for the recruitment model are calculated:

$$
r_{t}=\ln \left(R_{t}\right)-E\left[\ln \left(R_{t}\right)\right]
$$

Assuming that residuals are $\log$ normal, the NLL for recruitment residuals is:

$$
L=\sum_{t=f_{\text {frsit }}}^{N} w_{t}\left[\ln \left(\sigma_{r}\right)+0.5\left(r_{t} / \sigma_{r}\right)^{2}\right]
$$

where $\lambda_{t}$ is an instance-specific weight usually set equal one. The additional term in the NLL $\left[\ln \left(\sigma_{r}\right)\right]$ is necessary because the variance $\sigma_{r}^{2}$ is estimated internally, rather than specified by the user.

The log scale variance for residuals is calculated using the maximum likelihood estimator:

$$
\sigma_{r}^{2}=\frac{\sum_{j=t_{\text {frst }}}^{N} r_{j}}{N}
$$

where $N$ is the number of residuals. For the recruitment model with constant variation around a mean value, $t_{\text {first }}=1$. For the random walk recruitment model, $t_{\text {first }}=2$. For the Beverton-Holt and Ricker models, $t_{\text {first }}=\ell+1$ and the recruit model imposes no constraint on variability of recruitment during years 1 to $\ell$ (see below). The biased maximum likelihood estimate for $\sigma^{2}$ (with $N$ in the divisor instead of the degrees of freedom) is used because actual degrees of freedom are unknown. The variance term $\sigma^{2}$ is calculated explicitly and stored because it is used below.

## Constraining the first few recruitments

It may be useful to constrain the first \{ years of recruitments when using either the Beverton-Holt or Ricker models if the unconstrained estimates for early years are erratic. In the KLAMZ model, this constraint is calculated:

$$
N L L=\sum_{t=1}^{t_{\text {mana }}-1} w_{t}\left\{\ln \left(\sigma_{r}+0.5\left[\frac{\ln \left(R_{t} / E\left(R_{t_{p_{r a x}}}\right)\right)}{\sigma_{r}}\right]^{2}\right\}\right.
$$

where $t_{\text {first }}$ is the first year for which expected recruitment $E\left(R_{l}\right)$ can be calculated with the spawner-recruit model. In effect, recruitments that not included in spawner-recruit calculations are constrained towards the first spawner-recruit prediction. The standard deviation is the same as used in calculating the NLL for the recruitment model.

Prior information about abundance index scaling parameters $(Q)$
A constraint on one or more scaling parameters $\left(Q_{v}\right)$ for abundance or survival indices may be useful if prior information is available (e.g. NEFSC 2000; NEFSC 2001; NEFSC 2002). In the Excel version, it is easy to program these (and other) constraints in an $a d-h o c$ fashion as they are needed. In the AD Model Builder version, log normal and beta distributions are preprogrammed for use in specifying prior information about $Q_{v}$ for any abundance or survival index.

The user must specify which surveys have prior distributions, minimum and maximum legal bounds ( $q_{\min }$ and $q_{\max }$ ), the arithmetic mean $(\bar{q})$ and the arithmetic CV for the prior the distribution. Goodness of fit for $Q_{v}$ values outside the bounds ( $q_{\text {min }}, q_{\max }$ ) are calculated:

$$
L=\left\lvert\, \begin{aligned}
& 10000\left(Q_{v}-q_{\max }\right)^{2} \text { if } Q_{v} \geq q_{\max } \\
& 10000\left(q_{\min }-Q_{v}\right)^{2} \text { if } Q_{v} \leq q_{\min }
\end{aligned}\right.
$$

Goodness of fit for $Q_{v}$ values inside the legal bounds depend on whether the distribution of potential values is log normal or follows a beta distribution.

## Lognormal case

Goodness of fit for lognormal $Q_{v}$ values within legal bounds is:

$$
L=0.5\left[\frac{\ln \left(Q_{v}\right)-\tau}{\varphi}\right]^{2}
$$

where the $\log$ scale standard deviation $\varphi=\sqrt{\ln (1+C V)}$ and $\tau=\ln (\bar{q})-\frac{\varphi^{2}}{2}$ is the mean of the corresponding log normal distribution.

## Beta distribution case

The first step in calculation goodness of fit for $Q_{v}$ values with beta distributions is to calculate the mean and variance of the corresponding "standardized" beta distribution:

$$
\bar{q}^{\prime}=\frac{\bar{q}-q_{\min }}{D}
$$

and

$$
\operatorname{Var}\left(q^{\prime}\right)=\left(\frac{\bar{q} C V}{D}\right)^{2}
$$

where the range of the standardized beta distribution is $D=q_{\max }-q_{\text {min }}$. Equating the mean and variance to the estimators for the mean and variance for the standardized beta distribution (the "method of moments") gives the simultaneous equations:

$$
\bar{q}^{\prime}=\frac{a}{a+b}
$$

and

$$
\operatorname{Var}\left(q^{\prime}\right)=\frac{a b}{(a+b)^{2}(a+b+1)}
$$

where $a$ and $b$ are parameters of the standardized beta distribution. ${ }^{23}$ Solving the simultaneous equations gives:

$$
b=\frac{\left(\bar{q}^{\prime}-1\right)\left[\operatorname{Var}\left(q^{\prime}\right)+\left(\bar{q}^{\prime}-1\right) \bar{q}^{\prime}\right]}{\operatorname{Var}\left(q^{\prime}\right)}
$$

and:

$$
a=\frac{b \bar{q}^{\prime}}{1-\bar{q}^{\prime}}
$$

Goodness of fit for beta $Q_{v}$ values within legal bounds is calculated with the NLL:

[^19]$$
L=(a-1) \ln \left(Q_{v}^{\prime}\right)+(b-1) \ln \left(1-Q_{v}^{\prime}\right)
$$
where $Q_{v}^{\prime}=Q_{v} /\left(Q_{v}-q_{\min }\right)$ is the standardized value of the survey scaling parameter $Q_{v}$.

## Surplus production modeling

Surplus production models can be fit internally to biomass and surplus production estimates in the model (Jacobson et al. 2002). Models fit internally can be used to constrain estimates of biomass and recruitment, to summarize results in terms of surplus production, or as a source of information in tuning the model. The NLL for goodness of fit assumes normally distributed process errors in the surplus production process:

$$
L=0.5 \sum_{j=1}^{N_{P}}\left(\frac{\widetilde{P}_{j}-P_{j}}{\sigma}\right)^{2}
$$

where $N_{p}$ is the number of surplus production estimates (number of years less one), $\widetilde{P}_{t}$ is a predicted value from the surplus production curve, $P_{t}$ is the assessment model estimate, and the standard deviation $\sigma$ is supplied by the user based, for example, on preliminary variances for surplus production estimates. ${ }^{24}$ Either the symmetrical Schaefer (1957) or asymmetric Fox (1970) surplus production curve may be used to calculate $\widetilde{P}_{t}$ (Quinn and Deriso 1999).

It may be important to use a surplus production curve that is compatible with recruitment patterns or assumptions about the underlying spawner-recruit relationship. More research is required, but the asymmetric shape of the Fox surplus production curve appears reasonably compatible with the assumption that recruitment follows a BevertonHolt spawner-recruit curve (Mohn and Black 1998). In contrast, the symmetric Schaefer surplus production model appears reasonably compatible with the assumption that recruitment follows a Ricker spawner-recruit curve.

The Schaefer model has two log transformed parameters that are estimated in KLAMZ:

$$
\widetilde{P}_{t}=e^{\alpha} B_{t}-e^{\beta} B_{t}^{2}
$$

The Fox model also has two log transformed parameters:

$$
\widetilde{P}_{t}=-e\left(e^{e^{\alpha}}\right) \frac{B_{t}}{e^{\beta}} \log \left(\frac{B_{t}}{e^{\beta}}\right)
$$

See Quinn and Deriso (1999) for formulas used to calculate reference points ( $F_{M S Y}, B_{M S Y}$, $M S Y$, and $K$ ) for both surplus production models.

## Catch/biomass

[^20]Forward simulation models like KLAMZ may tend to estimate absurdly high fishing mortality rates, particularly if data are limited. The likelihood constraint used to prevent this potential problem is:

$$
L=0.5 \sum_{t=0}^{N}\left(d_{t}^{2}+q^{2}\right)
$$

where:

$$
d_{t}=\left\lvert\, \begin{gathered}
F t-\Phi \text { if } F t>\Phi \\
0 \text { otherwise }
\end{gathered}\right.
$$

and
with the threshold value $\kappa$ normally set by the user to about 0.95 . Values for $\kappa$ can be linked to maximum F values using the modified catch equation described above. For example, to use a maximum fishing mortality rate of about $F \approx 4$ with $M=0.2$ and $G=0.1$ (maximum $X=4+0.2-0.1=4.1)$, set $\kappa \approx F / X\left(1-e^{-X}\right)=4 / 4.1\left(1-e^{-4}\right)=0.96$.

## Uncertainty

The AD Model Builder version of the KLAMZ model automatically calculates variances for parameters and quantities of interest (e.g. $R_{t}, F_{t}, B_{t}, F_{M S Y}, B_{M S Y}, \bar{F}_{\text {Recent }}$, $\bar{B}_{\text {Re cent }}, \bar{F}_{\text {Re cent }} / F_{M S Y}, \bar{B}_{\text {Recent }} / B_{M S Y}$, etc.) by the delta method using exact derivatives. If the objective function is the log of a proper posterior distribution, then Markov Chain Monte Carlo (MCMC) techniques implemented in AD Model Builder libraries can be used estimate posterior distributions representing uncertainty in the same parameters and quantities. ${ }^{25}$

## Bootstrapping

A FORTRAN program called BootADM can be used to bootstrap survey and survival index data in the KLAMZ model. Based on output files from a "basecase" model run, BootADM extracts standardized residuals:

$$
r_{v, j}=\frac{\ln \left(I_{v, j} / \hat{I}_{v, j}\right)}{\sigma_{v, j}}
$$

along with log scale standard deviations ( $\sigma_{v, j}$, originally from survey CV's or estimated from goodness of fit), and predicted values $\left(\hat{I}_{v, j}\right)$ for all active abundance and survival observations. The original standardized residuals are pooled and then resampled (with replacement) to form new sets of bootstrapped survey "data":

$$
{ }^{x} I_{v, j}=\hat{I}_{v, j} e^{r \sigma_{v, j}}
$$

where $r$ is a resampled residual. Residuals for abundance and survival data are combined in bootstrap calculations. BootADM builds new KLAMZ data files and runs the KLAMZ model repetitively, collecting the bootstrapped parameter and other estimates at each iteration and writing them to a comma separated text file that can be processed in

[^21]Excel to calculate bootstrap variances, confidence intervals, bias estimates, etc. for all parameters and quantities of interest (Efron 1982).

## Projections

Stochastic projections can be carried out using another FORTRAN program called SPROJDDF based on bootstrap output from BootADM. Basically, bootstrap estimates of biomass, recruitment, spawning biomass, natural and fishing mortality during the terminal years are used with recruit model parameters from each bootstrap run to start and carryout projections. ${ }^{26}$ Given a user-specified level of catch or fishing mortality, the delay-difference equation is used to project stock status for a user-specified number of years. Recruitment during each projected year is based on simulated spawning biomass, log normal random numbers, and spawner-recruit parameters (including the residual variance) estimated in the bootstrap run. This approach is similar to carrying out projections based on parameters and state variables sampled from a posterior distribution for the basecase model fit. It differs from most current approaches because the spawnerrecruit parameters vary from projection to projection.

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APPENDIX A6. Location and size of northern ocean quahog ( $70+\mathrm{mm} \mathrm{SL}$ ) survey catches between Georges Bank and Long Island,
1982-2005.



APPENDIX A7. Location and size of Mid-Atlantic ocean quahog ( $70+\mathrm{mm} \mathrm{SL}$ ) catches. Mid-Atlantic Bight, 1982-2005.


Appendix A7. (cont.)


Appendix A7. (cont.)


Appendix A7. (cont.)

# APPENDIX A8. Stock Assessment for Ocean Quahog in Maine Waters 

Prepared by Robert Russell (assessment lead, Maine Department of Marine Resources, BoothBay Harbor, ME) and the Invertebrate Subcommittee

## Executive Summary

The Maine ocean quahog resource is a unique segment of the quahog stock in Federal waters. As of 1999 under Amendment 10 to the Fishery Management Plan (FMP) for Atlantic Surfclam and Ocean Quahog, Maine was given a separate annual quota of 100,000 "Maine" bushels (bushels used to record landings in Maine are $66 \%$ as large as bushels used to report landings in the rest of the EEZ). Fishing is carried out using a "dry" dredge (with no water jets to loosen sediments).

Maine quahogs, often referred to as "mahogany" clams are a substitute for Mercinaria mercinaria in the half shell market. Maine quahogs are harvested at a much smaller size (38-64 mm shell length) than MidAtlantic quahogs (89-140 mm shell length).

Landings peaked in Maine in 2002 at 147,191 bushels and have fallen since to a level of 98,153 bushels in 2005. During this time period paralytic shellfish poisoning (PSP) kept many productive beds closed.

The State of Maine conducted a pilot survey for ocean quahogs in 2002 which provided useful information on abundance and distribution along with estimates of key biological parameters. Results from the pilot study were used to plan and narrow the focus of the 2005 survey.

Lacking from the pilot study was an estimate of dredge efficiency which is required to estimate biomass and mortality rates from landings and survey data. Based on data from boxcore samples and "follow on" survey tows during 2005-2006, the efficiency of the commercial dredge used during the 2005 survey was $16.1 \%$. In other words, $16.1 \%$ of relatively large (fully recruited) ocean quahogs in the path of the dredge are captured in each pass.

Based on survey density data and estimated dredge efficiency, the biomass of harvestable ocean quahogs during 2005 in the commercial fishing grounds ( $54 \mathrm{~nm}^{2}$ ) surveyed off Maine is $22,493 \mathrm{mt}$ meat weight. Based on the ratio of landings and biomass, the fishing mortality rate in the commercial fishing grounds surveyed off Maine is $F=0.022 \mathrm{y}^{-1}$.

Biological reference points have not been established for the Maine segment of the ocean quahog stock. However, a per recruit model analysis with parameters for the Maine segment of the stock was used to estimate reference points that are often used in fishery management. Based on per recruit modeling, $F_{\max }=0.0561, F_{0.1}=0.0247$ and $F_{50 \%}=0.013 \mathrm{y} \mathrm{y}^{-1}$.
$\mathrm{F}_{0.1}=0.0247 \mathrm{y}^{-1}$ (corresponding to a harvest rate of $2.5 \%$ per year) might be a reasonable reference point for managers if the goal is to maximize yield per recruit while preserving some spawning stock. Simulation analysis (Clark 2002) indicates that $F_{50 \%}=0.013$ ( $1.3 \%$ per year) might be a reasonable reference point for managers if the goal was to preserve enough spawning potential to maintain the resource in the long term. The estimated fishing mortality rate during $2005 \mathrm{~F}=0.022$ $\mathrm{y}^{-1}$ is nearly equal to $\mathrm{F}_{0.1}=0.0247 \mathrm{y}^{-1}$ and the assumed natural mortality rate $\mathrm{M}=0.02 \mathrm{y}^{-1}$ but higher than $F_{50 \%}=0.013$.

Survey size frequency distributions indicate differences in the size of quahogs between the "western" and "eastern" beds inside the commercial fishing grounds. Larger quahogs were found in eastern beds that had been closed to fishing for three year due to PSP.

Size frequency distributions from boxcores showed signs of recent settlement in the eastern bed (quahogs less than 5 mm SL ). However size classes between 5 and 35 mm SL were entirely missing throughout the survey indicating that recruitment is sporadic. Although growth is relatively rapid in Maine waters, it may be 3 decades or longer before these recruits become large enough to enter the fishery.

Stock assessment advice concerning ocean quahog in Maine waters would be easier to provide if management goals were formulated and if biological reference points for biomass and fishing mortality were defined.

## Introduction

The Maine fishery for Ocean quahogs, although harvesting the same species (Artica islandica), is persecuted in a different way and fills a different sector of the shellfish market than the rest of the EEZ fishery. The Maine "mahogany" quahog is harvested at a smaller size (38-64 mm or 1.5-2.5 in shell length, SL) than elsewhere in the EEZ fishery where ocean quahogs are harvested at $89-140 \mathrm{~mm}$ (3.5-5.5 in) SL.

Ocean quahog from Maine waters are marketed as a less expensive alternative for Mercenaria mercinaria (Maine DMR 2003). Harvesting takes place year round with the highest market demand during the summer holidays (Memorial Day through Labor Day). During this peak harvest period 30-40 out of a total of 57 license holders may land some volume of product.

The majority of the vessels in the Maine fleet are between 10.7-13.7 m (35-45 ft) and classified as "undertonnage" or "small" in issuing permits. All of the vessels use a "dry" dredge (with no hydraulic jets to loosen the sediments) with a cutter bar set by regulation at no more than 0.91 m (36 in). There are no restrictions on any other dimension of the dredge.

Quahog Fishing in Maine takes place in relatively few locations along the coast north of 43 degree 50 minute latitude (Figure 1). Historically the bulk of fishing activity has taken place between Mt. Desert Rock and Cross Island with two significant quahog beds south of Addison and Great Wass Island covering an area of approximately 60 square nautical miles.

The Maine fishery began to expand into Federal waters in the 1980's due in part to PSP closures within state waters. In 1990 it was determined that this fishing activity conflicted with the Magnuson-Stevens Fishery Management Conservation Act which calls for a stock to be managed as a unit throughout its range. The Maine fishery was granted "experimental" status from 1990-1997. In 1998, the Maine fishery was fully incorporated under Amendment 10 of the FMP and given an initial annual quota of 100,000 bushels based on historical landings data. There was no independent assessment of the resource available at that time. The State of Maine is responsible under Amendent 10 to certify harvest areas free of PSP and to conduct stock assessments.

In 2002 the State of Maine conducted a pilot survey to assess the distribution and abundance of quahogs along the Maine coast (MEDMR $2003{ }^{27}$ ). This survey was a critical first step in establishing distribution, size composition and relative abundance information for the Maine fishery and for directing the design of the current survey work. While this initial survey provided valuable

[^23]information it did not have the resources to estimate dredge efficiency and therefore was not able to estimate total biomass or biological reference points. The survey during 2005 focused effort on two issues: determining dredge efficiency, and mapping quahog densities in the region of highest commercial activity.

Estimates of biomass and mortality presented in this report are only for the commercial beds south of Addison and Jonesport/GreatWass Maine. This approach was chosen due to available resources and because it was conservative. Other quahog beds are known to exist along many parts of the Maine coast. If mortality targets could be met using the estimates from the primary fishing grounds then biomass outside the survey area can act as a defacto preserve.

## Fishery Data

Data throughout this report is presented in metric units. In some cases there are specialized terms and conversion factors which are listed below.

| "MidAtlantic" bushels of Ocean Quahogs x 10 | $=$ | lbs meat. |
| :--- | :--- | :--- |
| "MidAtlantic" bushels of ocean quahogs x 4.5359 | $=$ | kg meat |
| 1 "MidAtlantic" (= "industry") bushel | $=$ | 1.88 cubic feet |
| 1 "Maine" (= "US Standard") bushel | $=$ | 1.2448 cubic feet |
| "Undertonnage" vessel | $=$ | $1-4.9 \mathrm{GRT}$ |
| "Small" vessel | $=$ | $5-49.9 \mathrm{GRT}$ |
| 1 "Maine" bushel | $=$ | 0.0049 mt meat weight |

In 2005 there were 57 ocean quahog licenses in the State of Maine. Of these 57 licenses 30 reported landings. The number of active licenses has decreased each year since 2002 when 38 licenses had reported fishing activity.

Landings have also decreased steadily since 2002 when they were at a recorded high of 147,191 Maine bushels (TableZ 1). Landings for 2005 were 98, 153 Maine bushels. LPUE in recent years tracked downward with landings until the 2005 season when it showed a slight increase from 5.37 to 5.85 Maine bushels per hour towing (Figure 2). This increase may be an artifact of the open and closed status of parts of the main commercial beds due to PSP because the most productive quahog bed was reopened at the end of 2005 after a 3 year closure.

Incidental mortality in ocean quahog off Maine is an important topic for future research. Maine has a very high level of fishing activity relative to the size of the fleet. Approximately 16,766 hours of fishing took place during 2005 representing over 67,000 tows at 8 min per tow. Using standard industry dredge dimensions and tow speeds this level of fishing activity represents 28.68 nautical miles ${ }^{2}$ of bottom swept by commercial dredges.

All catches are tagged and vessel logbooks are submitted to track quota status. Marine Patrol has not had enough resources to check the validity of logbook entry or to confirm the vessels on purchased quota are reporting accurately.

## Research Surveys

With the limited funds dedicated for survey work on quahogs, it was decided to focus all of the 2005 survey effort on the primary commercial fishing grounds south of Addison and Great Wass. This decision is important in the interpretation of all following data as results because estimates pertain only to these two beds and not to the coast of Maine as a whole. Vessel logbooks and the 2002 independent survey abundance indices show that the majority of fishing activity and a sizable portion of the resource was in this region (Figure 3).

The first step in designing the survey was to establish a $1 \mathrm{~km}^{2}$ grid overlay using Arcveiw 3.2 over the known commercial beds. Based on number of days at sea, 260 sites (tows) could be completed. The centers of the $2601 \mathrm{~km}^{2}$ grids covering the commercial beds were selected as start points for survey tows (Figure 4). These points were transferred to The Cap'n Voyager Software for use on board the survey vessel.

The Quahog bed south of Addison, (referred to as "western") had been the only open fishing grounds for 3 years due to PSP issues in other beds. The quahog bed south of Great Wass Island, (referred to as "eastern") had been unfished for 3 years but had previously been one of the most productive fishing grounds.

## Survey gear and procedures

The commercial vessel F/V Promise Land is a 12.8 m ( 42 ft ) Novi Style dragger piloted by Capt. Michael Danforth that was contracted to perform all the survey drag operations. All survey tows were conducted using the same dredge with dimensions: cutter bar 0.91 m ( 36 in ), 2.44 m ( 8 $\mathrm{ft})$ long $\times 1.83 \mathrm{~m}(6 \mathrm{ft})$ wide $\times 1.22 \mathrm{~m}(4 \mathrm{ft})$ high, overall weight $1,361 \mathrm{~kg}(3,000 \mathrm{lbs})$, bar spacing
all grills $19.05 \mathrm{~mm}(3 / 4 \mathrm{in})$ (Figure 5 ). The survey dredge was the same dredge used by the F/V Promise Land during normal fishing activity.

As the vessel approached the start of a tow, bottom type and the feasibility of conducting a tow were assessed. If suitable bottom was not immediately present at the predetermined start point, the vessel would start crossing runs within the grid. If after 5 to 6 crosses no towable bottom or a tow path free of fixed lobster gear could not be found, then the grid location was deemed untowable, a note was made, and the captain continued on to the next site. When a suitable tow path was found within a grid the dredge was lowered to the bottom by free-spooling until the ratio of cable length to depth was $3: 1$. Once the desired cable length was reached the drum was locked, a two minute timer was started and a GPS point was taken.

Tows were made into the current at approximately $6.48 \mathrm{~km} / \mathrm{hr}$ ( 3.5 knots) speed over ground (average tow 214 m ). After two minutes elapsed, a second GPS point was taken and the dredge was brought to the surface.

Tow distances calculated using the start and stop GPS points are good estimates of the distance actually traveled by the dredge. The manner in which the dredge is set and retrieved does not create a situation in which the dredge continues to fish as it is retrieved or before the drum is locked. In particular, the weight of the dredge keeps it in place on the bottom when the drum is unlocked at the end of the tow. In addition, the practice of backing the vessel toward the stopping point at the end of each tow means that the dredge was unlikely to travel very far at the end of the tow as it is lifted into the water column.

After the dredge was retrieved and before it was brought on board the vessel, excess mud was cleaned from the dredge by steaming in tight circles with the dredge in the vessel's prop wash (Figure 6). Once on board, the dredge was emptied and photographed with a digital camera (Figure 7). The contents were placed on a shaker table (Figure 8), bycatch was noted and then all live quahogs were sorted out from the catch. From each tow a 5 L subsample of quahogs was taken at random (the entire catch was taken if catch was less than 5 L ). The subsample was used to estimate tow counts, volume, and size frequency of the catch. The remainder of the catch was placed in calibrated buckets to determine total catch volume.

All data collected on board during operations were entered into a Juniper Systems handheld Allegro field computer running Data Plus Professional Software. All GPS data were collected using a pair of Garmin Etrex handheld units and transmitted in real time to the Allegro and a laptop
running Cap'n Voyager Software. Data entry screens on the Allegro for the abundance survey consisted of: 1) trip information (date, time out, weather, sea state, time in, and comments); 2) site information (depth, bottom type, start tow GPS position, speed, end tow GPS position, and comments); 3) catch information (sample portion 5 L or all, volume, weight, count, photo id, size frequency 5 L or all, and comments); and 4) bycatch information (species, abundance).

The lengths (longest dimension) of all subsampled quahogs were measured to the nearest 0.01 mm and entered into the Allegro handheld using a Fowler Ultra-Cal IV digital caliper with an RS232 port. Estimated counts of quahogs were made by counting the number of clams in the 5 L sample and then expanding that value using the total volume of the catch. All data were analyzed using Excel with variances calculated using a bootstrap program (10,000 iterations) written by Dr. Yong Chen at the University of Maine, Orono.

Tow distances were determined by The Cap'n Software and were checked using ESRI ArcInfo software. All data from the tows were standardized to a 200 m tow prior to further analysis.

## Dredge efficiency

The Maine dry dredge is much less efficient ( $2-17 \%$, ME DMR 2003) than hydraulic dredges used in the rest of the EEZ which can be up to $95 \%$ efficient (Medcolf and Caddy, 1971). A reliable estimate of dredge efficiency is needed to convert survey densities to a biomass estimate (NEFSC 2004).

One method of estimating dredge efficiency is through depletion experiments which are used to measure survey dredge efficiency for NEFSC clam surveys in Federal waters. Depletion studies for ocean quahog involve sensor and data processing equipment that were not readily available. The dry dredge used in the Maine survey is relatively small compared to the depth of fishing. We hypothesized that it would be difficult to control the dredge precisely given the depth, size of dredge and strong currents in the region off Maine.

For the conditions off Maine is was determined that the best approach to estimating dredge efficiency would be through the use of a boxcore samples (to directly estimate quahog density) followed by survey tows in the same area. Considering only ocean quahog available to the fishery, the ratio of density measured by "follow on" dredge tows divided by boxcore density is an estimate of survey dredge efficiency (Thorarinsdottir and Jacobson 2005).

The $F / V$ Promise Land with its large A frame and winches was able to deploy the 544 kg $(1,200 \mathrm{lb})$ Ocean Instruments 610 boxcore with a core capacity of $0.062 \mathrm{~m}^{2}$ and maximum penetration up to 60 cm (Figure 9). Follow on tows were conducted using the same gear used during all previous portions of the survey.

Boxcore work was conducted at three locations during three separate trips, one in August of 2005, one in January of 2006 and the last in April 2006. In all three experiments, follow on survey tows were made the day after the cores had been taken. The locations sampled were in the eastern quahog bed in an area of relatively high abundance (Figure 10). This area was also selected because it was a closed fishing ground during the August 2005 trip which would eliminate the possibility of the boxcore sites being commercially towed before follow on tows could be made. In January and April 2006 the region had been reopened to commercial fishing. However, VHF radio announcements describing the type of work underway were broadcast to local fisherman who were very cooperative and stayed well away from the experimental areas until all follow on tows could be completed the next day. Data entered into the Juniper Systems Allegro field computer included information about: 1) the trip (date, start tow, end tow), core (core \#, core length, count, volume, weight, count of newly settled).

Each experiment began by establishing a single long towpath. To do this, the vessel was slowed to the standard tow speed of 3.5 kts and a GPS point was taken and plotted. After 2 min steaming along a fixed heading, a second GPS point was taken and plotted. These waypoints determined the endpoints for the follow on commercial tows and the path for boxcore sampling. Cores were then taken haphazardly along the tow path ( 60 for the August 2005 trip, 34 on the January 2006 trip and 30 on the April 2006 trip).

Once a core was brought on board it was measured for overall length and sieved through a large screen $\left(1 \mathrm{~cm}^{2}\right.$ mesh size). All quahogs were counted and their total volume and weight were measured.

During coring operations, it was noted that the upper 1-2 cm of very soft sediment contained recently settled quahogs ( $<5 \mathrm{~mm}$ length). The number of quahogs in this size range were recorded separately for all further cores and newly settled quahogs were retained to be preserved. During the January and April 2006 trips the top 5 cm of each core was removed and washed separately through a $300 \mu$ sieve and all quahogs $<5 \mathrm{~mm}$ SL were preserved.

It was noted during boxcore sampling during the August 2005 boxcore trip that there was a change in sediment type beginning around $12-15 \mathrm{~cm}$ from the surface of each core. At this transition the sediment turned to a matrix of solid clay and old quahog shell. None of the live quahogs found in the cores in 2005 were below this transition. To assess this, the maximum depth within the core of live quahogs was measured during the 2006 trips.

After the maximum number of cores had been completed for a given trip the commercial dredge was deployed at one of the endpoints of the established tow path. Standard commercial towing was conducted for 2 min along the same path as the cores had been taken allowing the dredge to tow from one endpoint to the next. After each round of coring, 6 tows were made along the same path, three in one direction and 3 opposing to help mitigate any effect from tide.

## Dredge survey results

A total of 2591 km 2 survey grids were selected for sampling (TableZ2). Out of the 259 there were 183 ( 121 in the western bed and 62 in the eastern bed) or $70.7 \%$ that were towable. Only two stations were untowable due to fixed lobster gear or other known obstructions. The remainder of the untowable sites were due to inappropriate substrate.

Calculations of fishable area were reduced by the area of the sites that were untowable. Total biomass calculations are based only on the towable area ( $183 \mathrm{~km}^{2}$ ). The site that had a known obstruction was not included as it is not fished by area harvesters because of the risk to their gear and the site with lobster gear was not included based on personal comments from Capt. Mike Danforth that it was an area of hard untowable substrate. Tow distance, catch volume and counts were all standardized to a 200 m tow. Actual tow distances averaged 214 m .

The density plot for the survey (Figure 11) shows the highest concentration of biomass in the eastern bed. The eastern section had been closed to quahog fishing for almost three years. Substrate data (Figure 12) from Kelly et al. (1998) show the complexity of the substrate in the eastern section with highest quahog densities found near the boundary of hard rocky substrate with gravels, sands or mud. Substrate data collected independently using sidescan imaging showed that Kelly et al.'s (1998) substrate information was relatively accurate. However, in some cases substrate labeled as "sand" or "gravel-sand mix" near our most productive tows may have been shell hash from old quahog beds that was seen in boxcores from the same area.

Size frequencies for all subsampled quahogs $(\mathrm{n}=20,737)$ taken during the survey are shown in Figure 13. Size frequencies were also plotted separately for quahogs sampled from the western and eastern beds (Figure 14). The western bed had a mean SL of $47.6 \mathrm{~mm} \pm 4.6 \mathrm{~mm}$ and the eastern bed had a mean SL of $52.4 \mathrm{~mm} \pm 5.1 \mathrm{~mm}$. Cumulative size frequency distributions and a Kolmogorov-Smirnov test were used to test the null hypothesis that the size frequency distributions in the eastern and western areas were the same (Zar 1999). The null hypothesis was rejected ( $\mathrm{p}=0.001$ )

Because the two beds have differing size compositions and abundance levels, it was decided to calculate abundance for the two beds separately before estimating combined abundance for the entire survey area. Abundance estimates (see below) assume a dredge efficiency of 0.161 (Table Z3 shows effects of different dredge efficiencies on abundance and bushel estimates).

To estimate the total biomass for the commercial fishing grounds the size frequency distributions were converted to proportion of the population in each 1 mm size bin. Shell length $(L)$ was converted to meat wet weight ( $W$ ) using $\mathrm{W}=4.97 \times 10-6 \times \mathrm{L}^{3.5696}$ (Maine DMR 2003). Meat weights were converted to total biomass (meats and shells) by applying the average meat yield from the pilot survey of $17.5 \%$ and combining the values for the separate beds.

| Variable | Bed | Estimate | CV |
| :--- | :--- | :--- | :--- |
| Abundance | Western | $1.7108 \times 10^{9}$ | $8 \%$ |
|  | Eastern | $2.4058 \times 10^{9}$ | $11 \%$ |
|  | Total | $4.1163 \times 10^{9}$ | $8 \%$ |
| Bushels | Western | $1.715 \times 10^{6}$ | $9 \%$ |
|  | Eastern | $2.787 \times 10^{6}$ | $11 \%$ |
|  | Total | $4.502 \times 10^{6}$ | $9 \%$ |
| Total Biomass (mt) | Western | 47,704 | $8 \%$ |
|  | Eastern | 94,977 | $13 \%$ |
|  | Total | 128,529 | $7 \%$ |
| Meat Weight (mt) | Western | 8,348 | $8 \%$ |
|  | Eastern | 16,621 | $11 \%$ |
|  | Total | 22,493 | $8 \%$ |

## Box core results

Efficiency estimates from box core experiments are presented based on sizes taken in the commercial fishery ( 35 mm SL and greater). The estimated dredge efficiency was $16.1 \%$ with a $95 \%$ bootstrap confidence interval of $11.4 \%-21.6 \%$.

Another important result from the boxcore work was that the average depth of live quahogs in the region sampled was no deeper than 9.55 cm (CV 20\%)._The standard commercial dry dredge has cutting teeth that are set to a depth of 7.62 cm . We did not see evidence of anaerobic quahogs located deep in the sediments as has been reported elsewhere (Chenowith and Dennison,1993; Taylor 1976). Based on these results, it would seem that the majority of quahogs in this region would be impacted after one pass of a dredge.

## Per recruit modeling

Biological and fishery parameters from a variety of sources were used to carry out a per recruit analysis for ocean quahog in Maine waters. Age at length and growth information was taken from Kraus et al. (1992). Von Bertalanffy growth parameters estimated from a sample of 663 quahogs from Machias Bay were: $L_{i n f}=59.470 \pm 2.089, K=0.055 \pm 0.006$, and $t_{o}=-0.235 \pm 0.483$. The growth curves from Maine indicate relatively fast growth the first few years of life in comparison to curves for other areas (Figure 19). Length-weight parameters were from the 2002 Maine Quahog survey: $\mathrm{W}=4.97 \times 10^{-6} * \mathrm{~L}^{3.5696}$. Length-weight curves for the Maine ocean quahogs and the rest of the EEZ stock were similar (Figure 20). Size at maturity data estimates were based on Rowell et al. (1990) who found that females became fully mature at an average size of 49.2 mm for a quahog stock in Nova Scotia, Canada.

Fishery selectivity was modeled as a linear ramp function that was zero at 37 mm SL and one at 47 mm . Following surveys, quahog of various sizes were pushed through the grates on the commercial dredge ( $19.05 \mathrm{~mm}, 3 / 4 \mathrm{in}$. bar spacing) to see what sizes might be retained. Clams from 34 mm to 38 mm generally passed through the grate with some getting caught. After 41 mm almost all clams were thick enough to be retained. The regression model for shell depth and shell length in Feindel (2003) shows that a $19.05 \mathrm{~mm}(3 / 4 \mathrm{in})$ bar spacing is the thickness of an ocean quahog with 38.7 mm SL.

The per recruit model used in this analysis was a length based approach which can be downloaded from the Northeast Fisheries Science Center as part of the NMFS Stock Assessment Toollbox. ${ }^{28}$ The length based per recruit model was also used by Thorarinsdottir and Jacobson (2005). The biological reference points estimated in per recruit modeling for ocean quahog were $F_{\max }=0.0561, F_{0.1}=0.0247$ and $F_{50 \%}=0.013 \mathrm{y}^{-1}$ (Figure 18).

Sensitivity analysis (Figure 21) shows biological reference points from the per recruit model for ocean quahog are most sensitive to fishery selectivity parameters and, in particular, the length at which ocean quahogs in Maine waters become fully recruited to the fishery.

## Fishing mortality rate

For this report fishing mortality is estimated as the catch in biomass/average biomass ${ }^{-1}$. The survey during 2005 took place over a period of two months and mortality rates are relatively low so that survey biomass is a good proxy for average biomass. Following NEFSC (2004), the catch for 2005 used in fishing mortality estimation was landings plus a $5 \%$ allowance for incidental mortality to account for clams that are killed during fishing activity but not harvested. Catch including the $5 \%$ incidental mortality allowance for 2005 was 505 mt and the biomass estimate was $22,493 \mathrm{mt}$ giving $\mathrm{F}=505 \div 22,493=0.022 \mathrm{y}^{-1}$. Thus, the estimated fishing mortality rate is roughly equal to $F_{0.1}$ but higher than $F_{50 \%}$.

## Stock Status

Ocean quahog biomass in Maine waters was 22,493 mt meat weight and 2.7 million mt meat weight for the EEZ stock as a whole during 2005. It is not necessary to evaluate stock status of ocean quahog in Maine waters relative overfishing definitions because the stock component off Maine is a relatively small part of the EEZ stock as a whole. Overfishing definitions apply to the EEZ stock as a whole.

It was not possible to evaluate current biomass levels relative to a biological reference points associated with maximum productivity, depleted stock or historical levels because no appropriate biological reference points or historical biomass estimates are available.

The fishing mortality rate during $2005 F=0.022 \mathrm{y}^{-1}$ was almost equal to $F_{0.1}=0.0247$ and the assumed natural mortality rate $M=0.02 \mathrm{y}^{-1}$ but almost double $F_{50 \%}=0.013 \mathrm{y}^{-1}$. F $\mathrm{F}_{0.1}$ might be a

[^24]reasonable reference point for managers if the goal is to maximize yield per recruit while preserving some spawning stock. Simulation analysis (Clark 2002) indicates that $F_{50 \%}(1.3 \%$ per year) might be a reasonable reference point for managers if the goal was to preserve enough spawning potential to maintain the resource in the long term. However, preservation of spawning potential may not be necessary if recruitment originates mostly outside of Maine waters.

There is evidence of recent recruitment (newly settled ocean quahog $<5 \mathrm{~mm} \mathrm{SL}$ ) in one of the beds that were surveyed. However, although growth is relatively rapid in Maine waters, it may be 3 decades or longer before these recruits become large enough to enter the fishery.

Stock assessment advice concerning ocean quahog in Maine waters would be easier to provide if management goals were formulated and if biological reference points for biomass and fishing mortality were defined.

## Research Recommendations

1. Impact on habitat and substrate should be investigated for the Maine Dredge along with good estimates of area swept by fishing activity,
2. More work needs to be done to determine age, growth rates and size/age at maturity for Maine ocean quahogs. New digitized methods may help in this process.
3. Need better estimates of gear selectivity.

## Acknowledgements

This report was completed with extensive help from the Captains and crews of the F/V Promise Land and F/V Whitney and Ashley. Additional assistance was provided by Scott Feindel, Larry Jacobson and the NEFSC Invertebrate (Clam) working group.

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## Maine Ocean Quahog Report -- Appendix - Paired Tows Experiment

## Survey design

The current (2005) survey for ocean quahogs was conducted using a substantially larger vessel (F/V Promise Land 12.8 m ) and drag than the 2002 survey vessel the F/V Whitney and Ashley (11m ). In order to link the data from the 2002 pilot survey with the 2005 survey we needed a correction factor between the two vessels and drags. One concern with the pilot survey from industry members had been that the drag on the Ashley and Whitney was to light to get a good sample of the quahogs on bottom and would tend to underestimate abundance. The State of Maine contracted the original vessel, captain and drag to conduct side by side tows with the current survey vessel on April 16, 2005. It was determined that the two vessels would steam to an area in the closed fishing grounds that had a relatively high abundance of quahogs and conduct 8 coordinated close side by side tows in three replicate areas, 24 tows in all.

## Survey gear

Each vessel was equipped with the same survey gear as had been used during their respective trips. Once a suitable tow path had been established both vessels in unison deployed there dredges and let out equal lengths of cable (Figure 22). The captain of the F/V Promise Land was responsible for setting the pace and path of towing and for radioing the precise start and stop times for a tow. Tow positions were recorded onboard the F/V Promise Land. Once both dredges had been recovered and washed in the vessel wake all live quahogs were removed and placed in graduated containers to determine total volume. Either a 5L subsample or the entire catch, which ever was greater, was taken for count estimates and size frequency measurements.

## Data collection

Both vessels were equipped with a Allegro handheld field computer and data was entered under the categories: trip information (date, vessel, weather, sea state), tow information (tow number, depth, bottom type, start tow gps, speed, end tow gps, weight 5L, count 5L, estimated total count), size information ( length). All tow locations were also entered into the Cap'n Voyager software. All data was analyzed in Excel and bootstrapped using Dr. Chen's program.

## Paired tows results

Results from the side by side tows indicate a 2.5:1 ratio between the F/V Promise Land and the F/V Whitney and Ashley. The data collected from the tows was bootstrapped 10,000 times to estimate the standard error and $95 \%$ CI (Figure 23) Mean number per tow from the F/V Promise land was 1452 (CV 14\%). Mean number per tow from F/V Whitney and Ashley was 583 (CV $13 \%)$.

The size frequency distribution from quahogs collected from subsamples during the tows (Figure 24) indicates a difference in selectivity between the two drags. A K/S test run on cumulative fractions shows a difference in the two distributions at the 0.02 level (Figure 25). The square mesh liner in the dredge on the F/V Whitney and Ashley was 19.05 mm on a side while the bar spacing on the F/V Promise Land is 19.05 mm . The smallest quahog present in both dredges subsamples is only 1 mm different at 35 mm and 36 mm SL respectively. Bar spacing may play a role in the selectivity difference since a square grid would have many more intersections to trap smaller animals or increase the likelihood of clogging the dredge with mud.

The size frequencies not only show that the lighter drag on the F/V Whitney and Ashley retained smaller quahogs it did not sample larger quahogs present in the area. This effect would not be caused by smaller openings but is an indication that the dredge may under sample larger quahogs. If smaller quahogs need to be closer to the surface because of siphon length or substrate availability than the lighter drag on the F/V Whitney and Ashley would have a bias to select a smaller quahog than a heavier dredge that can cut deeper into the substrate. Also the tow speeds set by the F/V Promise Land were faster than those regularly used by the F/V Whitney and Ashley. The lighter drag may not have been as effective at the slightly higher speeds used in the paired towing. The 2002 survey had two types of tows. Those conducted randomly through out the State and those done systematically based on distance from reported commercial catches. The systematic survey may be biased towards heavy catch areas so only the random sites that overlap the 2005 survey area were used for this rough comparison. Area biomass estimates from the 2002 pilot study are based on 25 completed tows.

The current estimate for the region which overlaps many of the same stations is based on 183 completed tows at a much finer scale. This may partly explain the differences between the two
estimates. Also three years of fishing has taken place since the initial survey in which nearly 467,000 Maine bushels have been landed from the same region.

The updated 2002 estimate for the current survey area is $5.99 \times 10^{6}$ bushels with a $95 \% \mathrm{CI}$ within $47 \%$ of the mean. The estimate from the 2005 survey is $4.502 \times 10^{6}$ bushels with a $95 \% \mathrm{CI}$ within $25.4 \%$ of the mean.
$\left.\begin{array}{|r|r|r|r|r|}\hline & & & & \\ \text { Year } & \begin{array}{c}\text { Landings(Maine bushels) } \\ \text { all vessel classes combined }\end{array} & \begin{array}{l}\text { Landings (only records with } \\ \text { both effort and catch>0) }\end{array} & \begin{array}{l}\text { Nominal } \\ \text { LPUE } \\ \text { (ME } \\ \text { (hrs fished) }\end{array} \\ \text { bushel/hr) }\end{array}\right]$

Appendix A8. Table 1. Landings data for 1990-2005 from vessel logbooks. LPUE is reported for those records with both catch and effort data.

| sizes selected by dredge(>34mm SL) |  |  |  |  | all sizes |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | lower 95\% | average | upper 95\% | lower95\% | average | upper 95\% |
| Efficiency \% |  | 11.4 | 16.1 | 21.6 | 3.9 | 5.4 | 7.1 |
| east | mean | 3.3977E+09 | $2.4058 \mathrm{E}+09$ | $1.7932 \mathrm{E}+09$ | $9.9317 \mathrm{E}+09$ | $7.1729 \mathrm{E}+09$ | 5.4554E+09 |
|  | se | $3.6358 \mathrm{E}+08$ | $2.5744 \mathrm{E}+08$ | $1.9189 \mathrm{E}+08$ | $1.0628 \mathrm{E}+09$ | $7.6757 \mathrm{E}+08$ | $5.8378 \mathrm{E}+08$ |
| west | mean | $2.4161 \mathrm{E}+09$ | $1.7108 \mathrm{E}+09$ | $1.2752 \mathrm{E}+09$ | $7.0625 \mathrm{E}+09$ | 5.1007E+09 | $3.8794 \mathrm{E}+09$ |
|  | se | $1.9464 \mathrm{E}+08$ | $1.3782 \mathrm{E}+08$ | $1.0272 \mathrm{E}+08$ | $5.6894 \mathrm{E}+08$ | $4.1090 \mathrm{E}+08$ | $3.1251 \mathrm{E}+08$ |
| all | mean | 5.8134E+09 | 4.1163E+09 | 3.0682E+09 | $1.6993 \mathrm{E}+10$ | $1.2273 \mathrm{E}+10$ | $9.3341 \mathrm{E}+09$ |
|  | se | $4.6013 \mathrm{E}+08$ | $3.2580 \mathrm{E}+08$ | $2.4284 \mathrm{E}+08$ | $1.3450 \mathrm{E}+09$ | $9.7138 \mathrm{E}+08$ | $7.3880 \mathrm{E}+08$ |


| Bushel Estimates <br> based on 10,000 bootstrap runs <br> Efficiency (\%) | 11.4 | 16.1 | 21.6 |  |
| :--- | :--- | ---: | ---: | ---: |
| east | mean | $3.936 \mathrm{E}+06$ | $2.787 \mathrm{E}+06$ | $2.078 \mathrm{E}+06$ |
|  | se | $4.156 \mathrm{E}+05$ | $2.943 \mathrm{E}+05$ | $2.193 \mathrm{E}+05$ |
|  |  |  |  |  |
| west | mean | $2.422 \mathrm{E}+06$ | $1.715 \mathrm{E}+06$ | $1.278 \mathrm{E}+06$ |
|  | se | $2.209 \mathrm{E}+05$ | $1.564 \mathrm{E}+05$ | $1.166 \mathrm{E}+05$ |
|  |  |  |  |  |
| all | mean | $2.160 \mathrm{E}+01$ | $4.502 \mathrm{E}+06$ | $3.356 \mathrm{E}+06$ |
|  | se | $1.793 \mathrm{E}+09$ | $3.872 \mathrm{E}+05$ | $2.886 \mathrm{E}+05$ |

Appendix A8. Table 2. Effects of efficiency estimates on count and bushel estimates.


Appendix A8. Figure 1. Under the current Surfclam/Ocean Quahog FMP, the Maine fishing area is defined as north of the $43^{\circ} 50^{\prime} \mathrm{N}$. This line roughly splits the Maine coast in two.


Appendix A8. Figure 2. Catch and effort trends in the Maine quahog fishery. In 2002 one of the primary quahog beds was closed due to PSP. It was reopened in the last quarter of 2005.


Appendix A8. Figure 3. Commercial harvest locations during 2003-2005. Point size represents total bushels reported to that location by all vessels.


Appendix A8. Figure 4. Spatial grids for abundance survey in relation to commercial activity.


Appendix A8. Figure 5. Commercial drag used in all surveys in 2005.


Appendix A8. Figure 6. Cleaning the catch before it is brought on board. This practice is used in commercial operations as well.


Appendix A8. Figure 7. Typical catch as it comes on board. Tow duration 2 minutes.


Appendix A8. Figure 8. The catch being processed on a standard shaker table.


Appendix A8. Figure 9. Ocean Instruments 610 Boxcore along with a typical core sampled.


Appendix A8. Figure 10. Locations of Boxcore samples.
Areas with high quahog density were chosen from the abundance survey results.


Appendix A8. Figure 11. Density Plot from towable 2005 survey locations.


Appendix A8. Figure 12. Survey tows overlay on substrate data from Joe Kelly.


Appendix A8. Figure 13. Size frequencies for all tows in the western and eastern beds.


Appendix A8. Figure 14. Size frequencies for western and eastern bed. Used as basis for K/S test


Appendix A8. Figure 15. Cumulative distributions for length composition in the western and eastern beds. The curves are significantly different at the $p=0.001$ level.



B

Appendix A8. Figure 16. Results from bootstrap runs on mean count per tow split by west (A) east (B) and on bushels per tow split west (C, next page) east ( $D$, next page).


Figure 16. (cont.)


Appendix A8. Figure 17. Size frequencies from boxcore and follow on tows.


Appendix A8. Figure 18. Per recruit model results for Maine ocean quahogs.


Appendix A8. Figure 19. Three growth curves for quahog. Data for the Krauss curve was from Maine.


Appendix A8. Figure 20. Meat weight shell length relationships for three quahog stocks. Data for the Kruass-Feindel curve was from Maine.

| Fully recruited <br> length <br> length$\|$ F-01 Fmax F50\%MSP <br> 30 0.0196 0.0348 0.0109 <br> 35 0.0215 0.0419 0.0116 <br> 40 0.0242 0.0543 0.0126 <br> 45 0.0275 0.0801 0.0143 <br> 50 0.0319 0.168 0.018 <br> 55 0.0376 -1 0.0309 |  |  |  |
| ---: | ---: | ---: | ---: |


| Fully <br> Mature |  |  |  |
| :--- | :--- | :--- | ---: |
| length | F-01 | Fmax | F50\%MSP |
| 30 | 0.0253 | 0.0604 | 0.0168 |
| 35 | 0.0253 | 0.0604 | 0.0164 |
| 40 | 0.0253 | 0.0604 | 0.0157 |
| 45 | 0.0253 | 0.0604 | 0.0146 |
| 50 | 0.0253 | 0.0604 | 0.013 |
| 55 | 0.0253 | 0.0604 | 0.0105 |
| 60 | 0.0253 | 0.0604 | -1 |
| 65 | 0.0253 | 0.0604 | -1 |

Appendix A8. Figure 21. Sensitivity of YPR to size at recruitment and maturity.


Appendix A8. Figure 22. Side by side towing operations underway.


Appendix A8. Figure 23. Results from both bootstrap runs for the paired tows between the F/V Promise Land and the F/V Whitney and Ashley. The F/V Promise Land has a catch ratio to the F/V Whitney and Ashley of 2.5:1


Appendix A8. Figure 24. Size frequencies for the two vessels in the paired tow experiments.


Appendix A8. Figure 25. Cumulative distribution plots for length data in paired tows.

# B. ASSESSMENT OF NORTHEAST SKATE SPECIES COMPLEX 

Report of the SAW Southern Demersal Working Group
(Members are listed at front of Report)
(EDITOR'S NOTE: In this skate assessment report, tables and figures are numbered according to Term of Reference, TOR. For example, Figure 3.1 would be the first figure for TOR 3.)

### 1.0 EXECUTIVE SUMMARY AND TERMS OF REFERENCE

## TOR 1. Characterize the commercial and recreational catch including landings and discards.

The principal commercial fishing method in the directed skate fishery is otter trawling. Skates are frequently taken as bycatch during groundfish trawling and scallop dredge operations and discarded. Recreational and foreign landings are currently insignificant. There are few regulations governing the harvesting of skates in U.S. waters. Skates have been reported in New England fishery landings since the late 1800s. Reported commercial fishery landings, primarily from off Rhode Island, however, never exceeded several hundred metric tons until the advent of distant-water fleets and the industrial fishery during the 1950s and 1960s. Skate landings reached $9,500 \mathrm{mt}$ in 1969 primarily from the distant water fleet, but declined quickly during the 1970s, falling to 800 mt in 1981. Since that time, landings have increased, partially in response to increased demand for lobster bait, and more significantly, to the increased export market for skate wings. Landings are not reported by species, with over $99 \%$ of the landings reported as "unclassified skates." Wings were likely taken from large-bodied skates (winter, thorny and barndoor), with winter and thorny skate currently known to be used for human consumption. Bait landings are presumed to be primarily from little skate, based on areas fished and known species distribution patterns. Landings increased to 12,900 mt in 1993 and then declined somewhat to $7,200 \mathrm{mt}$ in 1995. Landings increased again and the 2004 reported commercial landings of $16,073 \mathrm{mt}$ were the highest on record. Estimates of discards suggest they may be 2-4 times larger than the average landings. The commercial fishery discard mortality rates by species are unknown.

Aggregate recreational landings of the seven species in the skate complex are relatively insignificant when compared to the commercial landings, never exceeding 300 mt during the 1981-1998 time series of Marine Recreational Fishery Statistics Survey (MRFSS) estimates. The number of skates reported as released alive averages an order of magnitude higher than the reported landed number. Party/charter boats have historically been undersampled compared to the private/rental boat sector that accounts for most of the recreational catch, and may have a different discard rate. The recreational fishery release mortality rate of skates is unknown, but is likely comparable to that for flounders and other demersal species, which generally ranges from 10-15\%. Assuming a $10-15 \%$ release mortality rate would suggest that recreational fishery discard mortality is of about the same magnitude as the recreational landings.

TOR 2. Estimate fishing mortality, spawning stock biomass, and total stock biomass for the current year and characterize the uncertainty of those estimates. If possible, also include estimates for earlier years.

Fishing Mortality
(EDITOR'S NOTE: MODEL-BASED FISHING MORTALITY ESTIMATES WERE PROPOSED; BUT THEY WERE REJECTED BY THE REVIEW PANEL)

## Total Biomass

NEFSC survey data were the primary source of information to index biomass of skate species. Indices of winter skate abundance and biomass from the NEFSC autumn surveys were stable, but below the time series mean, during the late 1960s and 1970s. Winter skate indices increased to the time series mean by 1980, and then reached a peak during the mid 1980s. Winter skates indices began to decline in the late 1980s. Current NEFSC indices of winter skate abundance are below the time series mean, at about the same value as during the early 1970s. Current NEFSC indices of winter skate biomass are about $20 \%$ of the peak observed during the mid 1980s. Indices of little skate abundance and biomass from the NEFSC spring were stable, but below the time series mean, during the 1970s. Little skate spring survey indices began to increase in 1982, reached a peak in 1999, and declined thereafter. Indices of barndoor skate abundance and biomass from the NEFSC autumn surveys were at the highest values during early to late 1960s, and then declined to 0 fish per tow during the early 1980s. Since 1990, autumn survey indices have steadily increased, with the survey nearing the peak values found in the 1960s. NEFSC autumn survey indices for thorny skate have declined continuously over the last 40 years. NEFSC indices of thorny skate abundance have declined steadily since the late 1970s, reaching a historically low value in 2005 is less than $10 \%$ of the peak observed in the 1970s. Indices of smooth skate abundance and biomass from the NEFSC autumn survey were at a peak during the late 1970s. NEFSC survey indices declined during the 1980s, before stabilizing during the early 1990s at about $25 \%$ of the values of the 1970s. NEFSC spring and autumn survey indices for clearnose skate increased from the mid-1980s through 2000 and have since declined to about average values. Indices of rosette skate abundance and biomass from the NEFSC surveys were at a peak during 1975-1980, before declining through 1986. NEFSC survey indices for rosette skate increased from 1986 through 2001, declined slightly and recent indices are near the peak values of the late 1970s.

Spawning Stock Biomass:
Winter skate SSB generally follows the pattern of the autumn total biomass index with very low values in the 1970s followed by the large expansion of the size composition in the 1980s. The index of SSB declined in the mid- to late 1990s, increased slightly, and is currently at low values. Little skate SSB has been fairly stable through the time series with slightly higher values from 1999-2004 than in the 1980s and early 1990s. The pattern in barndoor skate SSB indices is much the same as that of total biomass with high values in the early 1960s, followed by very low to nonexistent values in the 1970s and

1980s, and then a consistent increase in the 1990s and 2000s. The decline in thorny skate SSB indices is more pronounced than for the total biomass index. Smooth skate SSB indices are very variable, but exhibit a slight decline over the time series. Clearnose skate SSB has increased over the time period. Rosette skate SSB has been variable but has generally increased.

TOR 3. Either update or redefine biological reference points (BRPs; proxies for BMSY and FMSY), as appropriate. Comment on the scientific adequacy of existing and redefined BRPs.

## Existing Reference Points:

Biomass reference points (Figure B2) are based entirely on survey data because commercial catches are not available by species. For all species except barndoor, the $B_{\text {msy }}$ proxy $\left(B_{\text {target }}\right)$ is estimated as the $75^{\text {th }}$ percentile of the appropriate survey series for that species (see Summary Status Table). For barndoor skate, the $\mathrm{B}_{\text {msy }}$ proxy is the average of the autumn survey biomass indices from a short period, 1963-1966. This period is used for barndoor skates because the survey captured few barndoor skates for a protracted period after these years. The stocks are declared to be overfished when the three-year moving average of the NMFS trawl survey index (mean weight per tow) is less than one half of the $75^{\text {th }}$ percentile of mean weight per tow of the reference survey series for that species ( $\mathrm{B}_{\text {threshold }}$ ).

The overfishing definition is based on changes in survey biomass indices. In any year, if the three-year moving average of the survey biomass index for a skate species declines by more than a critical percentage from the previous year's moving average, then fishing mortality is assumed to be greater than $\mathrm{F}_{\text {msy }}$ and overfishing is assumed to be occurring for that skate species. The critical percentages for each species are given in the Summary Status Table (below).

Proposed Reference Points:
(EDITOR'S NOTE: NEW REFERENCE POINTS WERE PROPOSED; HOWEVER THEY WERE NOT ACCEPTED BY THE REVIEW PANEL)

TOR 4. Evaluate current stock status with respect to the existing BRPs, as well as with respect to updated or redefined BRPs (from TOR 3).

Summary Status Table - Northeast Skate Species - Basis: Existing Reference Points

| Species | Series | Btarget | Bthresh | Current | Status | Target Percent | Current | Status |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Winter | GOM-MA Off | 6.46 | 3.23 | 3.34 | Not Overfished | -20 | -22.9 | Overfishing |
|  | Autumn 67-98 |  |  |  |  |  |  |  |
| Little | GOM-MA All | 6.54 | 3.27 | 4.59 | Not | -20 | -15.9 | No Overfishing |
|  | Spring 82-99 |  |  |  | Overfished |  |  |  |
| Barndoor | GOM-SNE Off | 1.62 | 0.81 | 0.96 | Not | -30 | 9.8 | No Overfishing |
|  | Autumn 63-66 |  |  |  | Overfished |  |  |  |
| Thorny | GOM-SNE Off | 4.41 | 2.20 | 0.56 | Overfished | -20 | -11.2 | No Overfishing |
|  | Autumn 63-98 |  |  |  |  |  |  |  |
| Smooth | GOM-SNE Off | 0.31 | 0.16 | 0.18 | Not | -30 | 3.7 | No Overfishing |
|  | Autumn 63-98 |  |  |  | Overfished |  |  |  |
| Clearnose | MA All | 0.56 | 0.28 | 0.63 | Not | -30 | -16.2 | Overfishing No Overfishing |
|  | Autumn 75-98 |  |  |  | Overfished |  |  |  |
| Rosette | MA Offshore | 0.029 | 0.015 | 0.049 | Not | -60 | 9.7 | No Overfishing |
|  | Autumn 67-98 |  |  |  | Overfished |  |  |  |

TOR 5.Review, evaluate and report on the status of the SARC/Working Group Research Recommendations offered in recent SARC-reviewed assessments.

Completed. See Section 5.

TOR 6.Examine the NEFSC Food Habits Database to estimate diet composition and annual consumptive demand for seven species of skates for as many years as feasible.

Most skates are benthivorous in their feeding habits. A clear prominence on Cancer crabs, other crabs, amphipods, polychaetes and similar benthic macrofauna and megafauna was apparent in the diets of these skates. Some of the larger skates- barndoor, thorny, and winter- can be piscivorous, particularly with ontogeny. The vast majority of fish (or fish-like) prey for these skates were small pelagic fishes and squids.

Save winter and little skates, overall consumption by most skate stocks is a relatively small amount of biomass flow. Most total consumption by any particular species of skate was scaled singularly by the abundance of that species. The vast majority of consumptive removals by all skates except little and winter was $<20 \mathrm{MT}$ per year.

As an aggregate group, skates consume a very small fraction of the total energy flow in the ecosystem. Skate consumptive removal is two to three orders of magnitude lower than biomass or production of skate prey. When abundance estimates are scaled by gear efficiency, it is possible that skates could consume a notable fraction of forage fish and squid biomass relative to what is removed by a fishery. Yet most of those forage fish stocks are at relatively high levels of abundance.

### 2.0 INTRODUCTION

The seven species in the Northeast Region (Maine to Virginia) skate complex are distributed along the coast of the northeast United States from near the tide line to depths exceeding 700 m ( 383 fathoms). The species are: little skate (Leucoraja erinacea), winter skate (L. ocellata), barndoor skate (Dipturus laevis), thorny skate (Amblyraja radiata), smooth skate (Malacoraja senta), clearnose skate (Raja eglanteria), and rosette skate (L. garmani).

In the Northeast region, the center of distribution for the little and winter skates is Georges Bank and Southern New England. The barndoor skate is most common in the Gulf of Maine, on Georges Bank, and in Southern New England. The thorny and smooth skates are commonly found in the Gulf of Maine. The clearnose and rosette skates have a more southern distribution, and are found primarily in Southern New England and the Chesapeake Bight. Skates are not known to undertake large-scale migrations, but they do move seasonally in response to changes in water temperature, moving offshore in summer and early autumn and returning inshore during winter and spring. Members of the skate family lay eggs that are enclosed in a hard, leathery case commonly called a mermaid's purse. Incubation time is 6 to 12 months, with the young having the adult form at the time of hatching (Bigelow and Schroeder 1953).

The last stock assessment for the skate complex was conducted in 1999 at SARC/SAW 30 (NEFSC 2000). At that time there was no Fishery Management Plan (FMP) in place. The National Marine Fisheries Service had been petitioned to list barndoor skate as endangered based on a paper published by Casey and Myers (1998) and was also asked to assess the other species in the complex. SARC 30 found no cause to list barndoor as endangered but recommended that the species remain on the candidate species list as well as to put thorny skate on the candidate species list. Biomass reference points were developed for all seven species and four were listed as overfished. Fishing mortality reference points were developed for winter and little skate and overfishing was occurring for winter skate.

Following SARC 30, an FMP was developed by the New England Fishery Management Council (NEFMC) when they were informed of the overfished status of thorny and barndoor (winter and smooth biomass increased in the 1999 autumn survey and were no longer considered overfished). The FMP was implemented in September of 2003 with a primary requirement for mandatory reporting of skate landings by species by both dealers and vessels. The FMP prohibited possession of barndoor and thorny skate, as well as smooth skate from the Gulf of Maine. A trip limit of $10,000 \mathrm{lbs}$ was implemented for winter skate with a Letter of Authorization for the bait fishery (little skate) to exceed the trip limit. Biomass reference points developed at SARC 30 were maintained, but new fishing mortality reference points were developed.

### 3.0 TOR 1. Characterize the commercial and recreational catch including landings and discards

### 3.1 Commercial Fishery Landings

Skates have been reported in New England fishery landings since the late 1800s. However, commercial fishery landings, primarily from off Rhode Island, never exceeded several hundred metric tons until the advent of distant-water fleets and the industrial fishery during the 1950s and 1960s. Skate landings reached $9,500 \mathrm{mt}$ in 1969, but declined quickly during the

1970s, falling to 800 mt in 1981 (Table B1.1, Figure B1.1). Landings then increased markedly, partially in response to increased demand for lobster bait, and more significantly, to the increased export market for skate wings. Landings increased to $12,900 \mathrm{mt}$ in 1993 and then declined somewhat to $7,200 \mathrm{mt}$ in 1995. Landings increased again and the 2004 reported commercial landings of $16,073 \mathrm{mt}$ were the highest on record (Table B1.1, Figure B1.1).

United States landings of skates are reported in all months (Table B1.2). There is a relatively even distribution of landings across months, but the summer months do show a slightly higher percentage, probably due to the increased demand for lobster bait during those months.

Skate landings are primarily from Massachusetts and Rhode Island (mainly New Bedford and Point Judith) with $85-95 \%$ of the landings occurring in those two states (Table B1.3). Landings from other states did occur back through time and the table somewhat reflects better reporting as more states reported in the NMFS database. Also, the difference in total landings between Table B1.1 and B1.3 is likely the result of landings from the industrial fishery not included in the Weighout database. These landings were sampled during the 1960s and 1970s for species composition and prorated. Skates accounted for about $10 \%$ of those landings.

Otter trawls are the primary gear used to land skates in the United States, with some landings coming from sink gill nets (Table B1.4). In the last couple of years, landings from longline gear have increased slightly in importance. The increase in other gear reflects the new reporting system implemented in 2004.

Landings are generally not reported by species, with over $99 \%$ of the landings reported as Aunclassified skates@ until the FMP was implemented in September of 2003 (Table B1.5). Wings are most likely taken from winter and thorny skates, the two species currently known to be used for human consumption. Bait landings are presumed to be primarily from little skate, based on areas fished and known species distribution patterns. Landings of barndoor and thorny skate are being reported by the dealers even though there is a possession prohibition for those two species. There are also wings reported for rosette, little and smooth which are known to be too small for wings. The distribution of skate landings by state and species also shows that some species are landed in areas that they do not occur (Table B1.6). For example, in 2004, barndoor were landed in Virginia which is too far south for barndoor skate.

### 3.2 Commercial Fishery Discards

Discard estimates from SAW/SARC 30 were revised in this assessment. The previous method, which employed primary species groups to bin the discard data, was found to be a biased estimator (NEFSC 2006). Instead, the ratio-estimator used in this assessment is based on the methodology described in Rago et al. (2005). It relies on a $\mathrm{d} / \mathrm{k}$ ratio where the kept component is defined as the total landings of all species within a "fishery". A fishery is defined as a homogeneous group of vessels with respect to gear type, season, and geographic region. Each of these attributes is an observable property and easily defined within existing data bases. Moreover, it is not dependent on ambiguous properties such as "target species" or imprecise selfreported attributes such as area fished.

The discard ratio for spiny dogfish in stratum $h$ is the sum of discard weight over all trips divided by sum of kept weights over all trips:

$$
\begin{equation*}
\hat{R}_{h}=\frac{\sum_{i=1}^{n_{h}} d_{i h}}{\sum_{i=1}^{n_{h}} k_{i h}} \tag{1}
\end{equation*}
$$

where $d_{i h}$ is the discards for dogfish within trip $i$ in stratum $h$ and $k_{i h}$ is the kept component of the catch for all species. $\mathrm{R}_{\mathrm{h}}$ is the discard rate in stratum h . The stratum weighted discard to kept ratio is obtained by weighted sum of discard ratios over all strata:

$$
\begin{equation*}
\hat{R}=\sum_{h=1}^{H}\left(\frac{N_{h}}{\sum_{h=1}^{H} N_{h}}\right) \hat{R}_{h} \tag{2}
\end{equation*}
$$

The total discard within a stratum is the product of the estimate discard ratio R and the total landings for the fishery in stratum h, i.e., $D_{h}=R_{h} K_{h}$.

Annual estimated discards by fishery for 1989-2005 are summarized in Table B1.7. Total discards in 1990 were estimated to be about $80,000 \mathrm{mt}$. Most of this came from the otter trawl fishery. However, in the first two years, there were no estimates of discards from the scallop dredge fishery, which represent a significant portion in later years. The peak in the estimates was in 1992 at almost $90,000 \mathrm{mt}$, almost half came from the scallop dredge fishery. Estimates have since declined except for 2002 which was inflated by one blue crab pot trip which is probably not representative of that fishery. Estimates in recent years are still higher than reported landings but are much lower than the estimates from the early 1990s. This is likely due to reduced effort in the multispecies groundfish fishery as well as the scallop dredge fishery. Sampling of the three main gear types (otter trawl, sink gill net, and scallop dredge) has improved in recent years (Tables B1.8-B1.10).

The discard estimates were not dis-aggregated to skate species because species identification is uncertain in the Domestic Observer Program. Catches of skates by species were mapped to determine if the data were potentially useful. Winter and little skate distributions look reasonable (Figures B1.2-B1.3). Barndoor distribution from the observer data shows fairly substantial amounts off Virginia and North Carolina (Figure B1.4). These are unlikely to be correctly identified. The distributions of thorny and smooth are also curious showing catches in the Mid-Atlantic (Figures B1.5-B1.6). The reverse is true for clearnose and rosette (Figures B1.7-B1.8). These two species have a southern distribution and the maps show considerable amounts of fish found in the Gulf of Maine. The length compositions of kept and discarded fish also show that there are identification problems (Figures B1.8-B1.15). In particular, the length frequency for kept little skate has fish that are 60 to 80 cm which is a larger size than this species can attain. The same thing occurs for smooth and rosette showing larger sizes than is possible.

### 3.3 Recreational Fishery Catch

Aggregate recreational landings of the seven species in the skate complex are relatively insignificant when compared to the commercial landings, never exceeding 300 mt during the 1981-1998 times series of Marine Recreational Fishery Statistics Survey (MRFSS) estimates. Little and clearnose skates are the most frequently landed species of the complex. For little skate, total landings varied between $<1000$ and 56,000 fish, equivalent to $<1$ to 15 mt , during 1981-1998. For clearnose skate, total landings varied between 2,000 and 145,000 fish, equivalent to 2 to 232 mt , during 1981-1998. The number of skates reported as released alive averages an order of magnitude higher than the reported landed number. Party/charter boats have historically been undersampled compared to the private/rental boat sector that accounts for most of the recreational catch, and may have a different discard rate. The recreational fishery release mortality rate of skates is unknown, but is likely comparable to that for flounders and other demersal species, which generally ranges from $10-15 \%$. Assuming a $10-15 \%$ release mortality rate would suggest that recreational fishery discard mortality is of about the same magnitude as the recreational landings. Data from 1999 through 2005 were similar in magnitude.

### 4.0 TOR 2. Estimate fishing mortality, spawning stock biomass, and total stock biomass for the current year and characterize the uncertainty of those estimates. If possible, also include estimates for earlier years.

### 4.1 Research survey data - Total Stock Biomass

Indices of relative abundance from NEFSC bottom trawl surveys form the basis for most of the conclusions about status of the seven species in the skate complex. The NEFSC trawl survey has been conducted in the autumn from the Gulf of Maine to Southern New England since 1963 (Azarovitz 1981) and the Mid-Atlantic was added in 1967 (Figure B2.1). A spring survey was started in 1968 with stations $<=27 \mathrm{~m}$ added in 1975 (Figures B2.2-2.4). All statistically significant NEFSC gear, door, and vessel conversion factors were applied to little, winter, and smooth skate indices when applicable (Sissenwine and Bowman, 1978; NEFC 1991). Juvenile little and winter skates are not readily distinguished in the field. The numbers of juveniles were split between the two species based on the abundance of the adults in the same tow.

For the aggregate skate complex, the spring survey index of biomass was relatively constant from 1968 to 1980, but then increased to peak levels in the mid to late 1980s. The index of skate complex biomass then declined steadily until 1994, but increased until 2000 and has since decreased (Figure B2.5A). If the species in the complex are divided into large (barndoor, winter, and thorny) and small sized skates (little, clearnose, rosette, and smooth), it is evident that the large increase in skate biomass in the mid to late 1980s was dominated by winter and little skate (Figure B2.5B,C). The biomass of large sized skates steadily declined from the mid-1980s to the mid-1990s and has since been stable (Figure B2.5B). The increase in aggregate skate biomass from the mid-1990s to 2000 was due to an increase in little skate and the subsequent decline is also due to little skate (Figure B2.5C).

Indices of relative abundance for some of the species have also been developed from MADMF and CTDEP research surveys.

The previous SARC computed variance estimates for the survey indices assuming a normal error distribution. A recommendation was made to explore alternate error distributions since this assumption may not hold at very low stock sizes and results in confidence intervals
which are below zero. Another alternative to assuming any error distribution is to use bootstrap methods. The bootstrap methodology of Smith (1997) was implemented using the Splus software written by Stephen Smith (DFO, Halifax). In order to bootstrap the NEFSC survey data, some strata had to be combined to ensure that at least two tows were made in each stratum during each year (Table B2.1). The second figure in each species section shows the stratified mean without combining strata, the mean combining strata and the bootstrapped mean.

## Winter skate

NEFSC spring and autumn bottom trawl surveys indicate that winter skate are most abundant in the Georges Bank (GBK) and Southern New England (SNE) offshore strata regions, with few fish caught in the Gulf of Maine (GOM), or Mid-Atlantic (MA) regions (NEFSC 2000; Figure B2.6). In the NEFSC spring survey offshore strata (1968-2006), the annual total catch of winter skate has ranged from 160 fish in 1976 to 1,891 fish in 1985. In the NEFSC autumn survey offshore strata (1963-2005), the annual total catch of winter skate has ranged from 115 fish in 1975 to 1,187 fish in 1984. Calculated on a per tow basis, these spring survey catches equate to maximum stratified mean number per tow indices for the GOM-MA offshore strata of about 7.9 fish, or 16.4 kg , per tow during 1985; autumn maximum catches equate to indices of 3.7 fish, or 13.3 kg per tow, in 1984 (Tables B2.2-B2.3).

The catchability of winter skate in the NEFSC winter bottom trawl survey (which substitutes a chain sweep with small cookies for the large rollers used in the spring and autumn surveys, to better target flatfish) is significantly higher than in the spring and autumn series, especially for smaller winter skates. NEFSC winter survey (1992-2006) annual catches of winter skate have ranged from 841 fish in 1993 to 4,055 fish in 1996, equating to a maximum stratified mean catch per tow of 43.5 fish or 25.2 kg per tow in 1996 (Table B2.4). The winter survey is focused in the Southern New England and Mid-Atlantic offshore regions, with a limited number of samples on Georges Bank, and no sampling in the Gulf of Maine (Figure B2.7). The NEFSC scallop dredge survey also catches winter skates mostly on Georges Bank (Figures B2.8-B2.9). The scallop survey also does not sample in the Gulf of Maine and on the very shallowest portions of Georges Bank.

Indices of winter skate abundance and biomass from the NEFSC spring and autumn surveys were stable, but below the time series mean, during the late 1960s and 1970s (Figure B2.10). Winter skate indices increased to the time series mean by 1980, and then reached a peak during the mid 1980s. Winter skate indices began to decline in the late 1980s. Current NEFSC indices of winter skate abundance are below the time series mean, at about the same value as during the early 1970s. Current NEFSC indices of winter skate biomass are about 20\% of the peak observed during the mid 1980s (Figures B2.10). The combining of strata did not have much impact on the stratified mean (Figures B2.11-B2.14).

The minimum length of winter skate caught in NEFSC surveys is 15 cm ( 6 in ), and the largest individual caught was 116 cm ( 46 in ) total length, during the 1985 spring survey on Georges Bank (Tables B2.2-B2.4). The median length of the survey catch has ranged from 28 cm in the 2003 winter survey to 79 cm in the 1978 spring survey and the 1985 autumn survey. The median length of the survey catch generally declined from 1979 to the mid-1990s in both the spring and autumn surveys, increased through 2002, and then declined slightly to currently remain about $45-52 \mathrm{~cm}$ (18-20 inches)(Figure B2.15). Length frequency distributions from the NEFSC spring and autumn surveys show several modes, most often at 40,60 , and 80 cm (Figures B2.16-B2.20). The spring survey length distributions show large modes at about 40 cm during the mid-1980s through the mid 1990s, suggesting strong recruitment during that period.

Truncation of the length distributions is evident in the NEFSC spring and autumn series since 1990.

The strata set used for bootstrapping the winter survey differed from the standard consistent strata set used for the information in Table 2.4. Given that the strata on Georges Bank were not sampled in some years, the set for bootstrapping was limited to Southern New England to the Mid-Atlantic (Table B2.1). This created more of a difference between the original mean, with usually a lower index when Georges Bank was included in the original (Figure B2.21B2.22). The indices of both abundance and biomass fluctuated without trend through the series.

The difference between the original mean and the combined strata mean in the scallop survey was due to the bootstrapped mean consisting of only strata which caught some winter skate (Figures B2.23-B2.24) while the original was the entire scallop survey strata set. There are no biomass estimates from 1985 through 2000 since no weights were taken at sea and the survey in 1999 was completed on a commercial scalloper and therefore the data are not comparable. Abundance was high in the mid-1980s, declined through the 1990s, increased through 2000 and then declined.

Indices of abundance for winter skate are available from the Massachusetts Division of Marine Fisheries (MADMF) spring and autumn research trawl surveys in the inshore waters of Massachusetts for the years 1978-2006. MADMF biomass indices of winter skate were moderate to high from 1981 through 1987. Thereafter, both spring and autumn indices declined to time series lows in 1989-1991. The spring index rebounded to moderate levels during 19921996 before dropping again to low values in the late 1990s and remaining low through 2006 (Figure B2.25). The autumn index is more erratic, but generally shows the same pattern.

Indices of abundance for winter skate are available from the Connecticut Department of Environmental Protection (CTDEP) spring and autumn finfish trawl surveys in Long Island Sound for the years 1984-2006 (1992 and later only for biomass). Annual CTDEP survey catches have ranged from 0 to 115 skates. CTDEP survey indices suggest that after increasing to a time series high from 1984 through 1989, winter skate in Long Island Sound has declined slightly (Figure B2.26).

## Little skate

NEFSC bottom trawl surveys indicate that little skate are abundant in the inshore and offshore strata in all regions of the northeast US coast, but are most abundant on Georges Bank and in Southern New England (NEFSC 2000, Figure B2.27). In the NEFSC spring surveys (1976-2006), the annual total catch of little skate has ranged from 3,512 fish in 1986 to 16,406 fish in 1999 (Table 2.5). In the NEFSC autumn surveys (1975-2005), the annual total catch of little skate has ranged from 1,124 fish in 1993 to 6,523 fish in 2003 (Table 2.6). Calculated on a per tow basis, these spring survey catches equate to maximum stratified mean number per tow indices for the GOM-MA inshore and offshore strata of about 28 fish, or 10 kg , per tow during 1999; autumn maximum catches equate to indices of 18 fish, or 7.7 kg , per tow in 2003 (Tables B2.5-B2.6).

The catchability of little skate in the NEFSC winter bottom trawl survey (which substitutes a chain sweep with small cookies for the large rollers used in the spring and autumn surveys, to better target flatfish) is significantly higher than in the spring and autumn series. NEFSC winter survey (1992-2006) annual catches of little skate have ranged from 8,870 fish in 2003 to 18,418 fish in 1992, equating to a maximum stratified mean catch per tow of 170 fish or 66 kg per tow in 1992 (Table B2.7). The winter survey is focused in the Southern New England and Mid-Atlantic offshore regions, with a limited number of samples on Georges Bank, and no
sampling in the Gulf of Maine (Figure B2.28). The NEFSC scallop dredge survey also catches little skates in all areas of sampling (Figures B2.29-B2.30). The scallop survey also does not sample in the Gulf of Maine, on the very shallowest portions of Georges Bank and parts of Southern New England.

Indices of little skate abundance and biomass from the NEFSC spring and autumn surveys were stable, but below the time series mean, during the 1970s. Little skate spring survey indices began to increase in 1982, reached a peak in 1999, and declined thereafter (Figure B2.31). Autumn survey indices have been relatively stable over the duration of the time series, with a slight increase in recent years (Figure B2.31). The application of the NEFSC gear conversion factors to spring survey indices decreased the indices in 1981 and earlier years by 75 percent. The combining of strata had slightly more impact for little skate than for winter skate, since many of the inshore strata were combined (Figures B2.32-B2.35).

The minimum length of little skate caught in NEFSC surveys is $6 \mathrm{~cm}(3 \mathrm{in})$, and the largest individual caught was $62 \mathrm{~cm}(24 \mathrm{in})$ total length, during the 1978 autumn survey on Georges Bank. The median length of the survey catch has ranged from 31 cm in the 1979 and 1987 spring surveys to 44 cm , most recently in the 2005 autumn survey. The median length of the survey catch has been generally stable over the duration of the spring and autumn surveys and is currently about 42 cm in the spring and 43 cm in the autumn (17 inches)(Figure B2.36). Length frequency distributions from the NEFSC spring and autumn surveys show several modes, most often at $10,20,30$, and 45 cm , which may represent ages $0,1,2$, and 3 and older little skate (Figures B2.37-B2.40).

The strata set used for bootstrapping the winter survey differed from the standard consistent strata set used for the information in Table 2.7. Given that the strata on Georges Bank were not sampled in some years, the set for bootstrapping was limited to Southern New England to the Mid-Atlantic (Table B2.1). This created more of a difference between the original mean, with usually a higher index when Georges Bank was included in the original (Figure B2.41B2.42). The indices of both abundance and biomass declined through 2000, increased for a few years and subsequently declined..

The difference between the original mean and the combined strata mean in the scallop survey was due to the bootstrapped mean consisting of only strata which caught some little skate (Figures B2.43-B2.44) while the original was the entire scallop survey strata set. There are only differences in the early part of the time series when more strata were sampled. There are no biomass estimates from 1985 through 2000 since no weights were taken at sea and the survey in 1999 was completed on a commercial scalloper and therefore the data are not comparable. Abundance indices increased to a peak in 2000 and have subsequently declined.

Indices of abundance for little skate are available from the Massachusetts Division of Marine Fisheries (MADMF) spring and autumn research trawl surveys in the inshore waters of Massachusetts for the years 1978-2006 (Figure B2.45). MADMF biomass indices of little skate declined through the 1980's to time series lows in 1989 (autumn) and 1991 (spring). Biomass indices quickly rose to high levels in the early 1990's, and have since fluctuated without trend.

Indices of abundance for little skate are available from the Connecticut Department of Environmental Protection (CTDEP) spring and autumn finfish trawl surveys in Long Island Sound for the years 1984-2006 (1992 and later only for biomass). Little skate are the most abundant species in the skate complex in Long Island Sound, with annual CTDEP survey catches ranging from 142 to 837 skates. CTDEP survey indices suggest an increase in abundance of little skate in Long Island Sound over the 1984-2006 time series followed by a decline (Figure B2.46).

## Barndoor skate

NEFSC bottom trawl surveys (Figure B2.47) indicate that barndoor skate are most abundant in the Gulf of Maine, Georges Bank, and Southern New England offshore strata regions, with very few fish caught in inshore ( $<27$ meters depth) or Mid-Atlantic regions. Bigelow and Schroeder (1953), however, noted that historically barndoor skate were found in inshore waters to the tide-line, and in depths as great as 400 meters off Nantucket. In the NEFSC spring surveys (1968-2006), the annual total catch of barndoor skate has ranged from 0 fish (several years during the 1970s and 1980s) to 196 fish in 2006 (Table B2.8). In the NEFSC autumn surveys (1963-2005), the annual total catch of barndoor skate has ranged from 0 fish (several years in the 1970s and 1980s) to 120 fish in 1963 (Table B2.9). Calculated on a per tow basis, the autumn survey catches equate to maximum stratified mean number per tow indices for the GOM-SNE offshore strata of about 0.8 fish, or 2.6 kg , per tow in 1963 (Tables B2.8-B2.9).

The catchability of barndoor skate in the NEFSC winter bottom trawl survey (which substitutes a chain sweep with small cookies for the large rollers used in the spring and autumn surveys, to better target flatfish) is significantly higher than in the spring and autumn series and may be particularly higher for smaller skates as in winter skates. NEFSC winter survey (19922006) annual catches of barndoor skate have ranged from 0 fish in 1992 to 355 in 2006, equating to a maximum stratified mean catch per tow of 3.2 fish or 3.0 kg per tow in 1999 (Table B2.10). The winter survey is focused in the Southern New England and Mid-Atlantic offshore regions, with a limited number of samples on Georges Bank, and no sampling in the Gulf of Maine (Figure B2.48). The NEFSC scallop dredge survey also catches barndoor skates primarily on Georges Bank (Figure B2.48). The scallop survey also does not sample in the Gulf of Maine, on the very shallowest portions of Georges Bank and parts of Southern New England.

Indices of barndoor skate abundance and biomass from the NEFSC spring and autumn surveys were at their highest values during early to late 1960s, and then declined to 0 fish per tow during the early 1980s. Since 1990, both spring and autumn survey indices have steadily increased, with the spring survey at the highest value and the autumn survey nearing the peak values found in the 1960s (Figure B2.49). The combining of strata did not have much impact on the stratified mean (Figures B2.50-B2.53).

The minimum length of barndoor skate caught in NEFSC surveys is 20 cm ( 8 inches), and the largest individual caught was 136 cm ( 54 in ) total length, during the 1963 autumn survey in the Gulf of Maine. The median length of the survey catch has ranged from 20 cm in the 1985 spring survey to 119 cm in the 1972 spring survey. The median length of the survey catch has been stable in recent years in both the spring and autumn surveys, and is currently $70-75 \mathrm{~cm}$ (2830 in ; Figure B2.54). Length frequency distributions from the NEFSC spring and autumn surveys illustrate the decline in abundance of barndoor skate to survey catches of zero during the 1980s (Figures B2.55-B2.59). Recent catches have included individuals as large as those recorded during the peak abundance of the 1960s, and the large number of fish between 40 and 80 cm evident during the 1960s is now apparent in recent surveys.

The strata set used for bootstrapping the winter survey differed from the standard consistent strata set used for the information in Table 2.10. Given that the strata on Georges Bank were not sampled in some years, the set for bootstrapping was limited to Southern New England to the Mid-Atlantic (Table B2.1). This created more of a difference between the original mean, with usually a lower index when Georges Bank was included in the original (Figure B2.60-B2.61). The indices of both abundance and biomass have increased substantially from 1993 to 2006. The NEFSC winter survey length frequency distributions for indicate a significant increase in the abundance of barndoor skate at lengths less than 80 cm (Figure B2.62).

The difference between the original mean and the combined strata mean in the scallop survey was due to the bootstrapped mean consisting of only strata which caught some barndoor skate (Figures B2.63-B2.64) while the original was the entire scallop survey strata set. There are no biomass estimates from 1985 through 2000 since no weights were taken at sea and the survey in 1999 was completed on a commercial scalloper and therefore the data are not comparable. Abundance indices increased consistently while the biomass indices have been more variable.

## Thorny skate

NEFSC bottom trawl surveys indicate that thorny skate are most abundant in the Gulf of Maine and Georges Bank offshore strata regions, with very few fish caught in inshore ( $<27$ meters depth), Southern New England, or Mid-Atlantic regions (Figure B2.65). In the NEFSC spring surveys (1968-2006), the annual total catch of thorny skate has ranged from 29 fish in 2006 to 574 fish in 1973 (Table 2.11). In the NEFSC autumn surveys (1963-2005), the annual total catch of thorny skate has ranged from 35 fish in 2005 to 874 fish in 1978 (Table 2.12). Calculated on a per tow basis, these spring and autumn survey catches equate to maximum stratified mean number per tow indices for the GOM-SNE offshore strata of about 2 to 3 fish, or about 6.0 kg , per tow during the early 1970s (Tables B2.11-2.12).

The NEFSC scallop dredge survey also catches thorny skates primarily on the edges of Georges Bank (Figure B2.66). The scallop survey also does not sample in the Gulf of Maine, on the very shallowest portions of Georges Bank and parts of Southern New England. A summer shrimp survey is conducted in the Gulf of Maine which also catches thorny skate (Figure B2.66). Indices from this survey have not been updated.

NEFSC spring and autumn survey indices for thorny skate have declined continuously over the last 40 years. Indices of thorny skate abundance and biomass from the NEFSC spring and autumn surveys were at a peak during the early 1970s, reaching 2.9 fish per tow ( 5.3 kg per tow) in the spring survey and1.8 fish per tow ( 5.9 kg per tow) in the autumn survey. Kulka and Mowbray (1998) indicated a similar period of high abundance for thorny skate in Canadian waters. NEFSC indices of thorny skate abundance have declined steadily since the late 1970s, reaching historically low values in 2005 and 2006 that are less than $10 \%$ of the peak observed in the 1970s (Figure B2.67). The combining of strata did not have much impact on the stratified mean (Figures B2.68-B2.71).

The minimum length of thorny skate caught in NEFSC surveys is about 10 cm ( 4 inches), and the largest individual caught was 111 cm ( 44 inches) total length, most recently during the 1977 spring survey on Georges Bank (Tables B2.11-B2.12). The median length of the survey catch has ranged from 23 cm in the 2003 autumn survey to 63 cm in the 1971 autumn survey. The median length of the survey catch has trended downward through most of the survey time series, but has been stable in recent years in autumn surveys, and is currently 40-50 cm (16-20 inches; Figure B2.72). Length frequency distributions from the NEFSC spring and autumn surveys show a pattern of decline in abundance of larger individuals consistent with an increase in total mortality over the survey time series (Figures B2.73-B2.77).

The difference between the original mean and the combined strata mean in the scallop survey was due to the bootstrapped mean consisting of only strata which caught some thorny skate (Figures B2.78-B2.79) while the original was the entire scallop survey strata set. There are no biomass estimates from 1985 through 2000 since no weights were taken at sea and the survey in 1999 was completed on a commercial scalloper and therefore the data are not comparable. Abundance indices declined from a peak in 1986 while the biomass indices declined since 2001.

Indices of abundance for thorny skate are available from the Massachusetts Division of Marine Fisheries (MADMF) spring and autumn research trawl surveys in the inshore waters of Massachusetts for the years 1978-2006. MADMF indices of thorny skate biomass have been variable over the time series, but there is a decreasing trend evident in both the spring and autumn time series. The spring index has stabilized around the median of $0.2 \mathrm{~kg} /$ tow throughout the 2000's, while the autumn index has been below the median of $0.6 \mathrm{~kg} / \mathrm{tow}$ since 1994 except for 2001 and 2002 (Figure B2.80).

## Smooth skate

NEFSC bottom trawl surveys indicate that smooth skate are most abundant in the Gulf of Maine and Georges Bank offshore strata regions, with very few fish caught in inshore ( $<27$ meters depth), Southern New England, or Mid-Atlantic regions (Figure B2.81). In the NEFSC spring surveys (1968-2006), the annual total catch of smooth skate has ranged from 12 fish in 1996 to 179 fish in 1973 (Table B2.13). In the NEFSC autumn surveys (1963-2005), the annual total catch of smooth skate has ranged from 10 fish in 1976 to 130 fish in 1978 (Table B2.14). Calculated on a per tow basis, these spring and autumn survey catches equate to maximum stratified mean number per tow indices for the GOM-MA offshore strata of 0.6 to 1.6 fish, or about 0.6 to 0.9 kg , per tow during the 1970s (Tables B2.13-B2.14).

The NEFSC scallop dredge survey also catches smooth skates primarily on the edges of Georges Bank (Figure B2.82). The scallop survey also does not sample in the Gulf of Maine, on the very shallowest portions of Georges Bank and parts of Southern New England. A summer shrimp survey is conducted in the Gulf of Maine which also catches smooth skate (Figure B2.82). Indices from this survey have not been updated.

Indices of smooth skate abundance and biomass from the NEFSC surveys were at a peak during the early 1970s for the spring series and the late 1970s for the autumn series (Figure B2.83). NEFSC survey indices declined during the 1980s, before stabilizing during the early 1990s at about $25 \%$ of the autumn and $50 \%$ of the spring survey index values of the 1970s. The combining of strata did not have much impact on the stratified mean (Figures B2.84-B2.87).

The minimum length of smooth skate caught in NEFSC surveys is about 8 cm ( 3 inches), and the largest individual caught was 73 cm ( 29 inches) total length, during the 2000 autumn survey on Georges Bank (Tables B2.13-B2.14). The median length of the survey catch has ranged from 26 cm in the 1993 autumn survey to 53 cm in the 1971 autumn survey. The median length of the survey catch in the GOM-SNE offshore region shows no trend over the full survey time series, and is currently at about 40 cm (16 in) (Figure B2.88). Length frequency distributions from the NEFSC spring and autumn surveys in the GOM offshore region show modes at 30 and 50 cm (Figures B2.89-B2.93). The relatively high abundances evident in the 1969-1983 spring surveys at the larger mode may represent the accumulated abundance at several older ages. Truncation of the larger mode is evident in the spring distributions during the 1980s and most of the 1990s. The 1999 spring survey length frequency distribution indicated strong recruitment in the region.

The difference between the original mean and the combined strata mean in the scallop survey was due to the bootstrapped mean consisting of only strata which caught some smooth skate (Figures B2.94-B2.95) while the original was the entire scallop survey strata set. There are no biomass estimates from 1985 through 2000 since no weights were taken at sea and the survey in 1999 was completed on a commercial scalloper and therefore the data are not comparable. Abundance indices were low at the beginning of the time series and have since increased.

## Clearnose skate

NEFSC bottom trawl surveys indicate that clearnose skate are most abundant in the MidAtlantic offshore and inshore strata regions, with very few fish caught in Southern New England and no fish caught in other survey regions (Figure B2.96). In the NEFSC spring surveys (19762006), the annual total catch of clearnose skate has ranged from 9 fish in 1979 to 136 fish in 1993 (Table B2.15). In the NEFSC autumn surveys (1975-2005), the annual total catch of clearnose skate has ranged from 19 fish in 1983 to 221 fish in 2001 (Table B2.16). Calculated on a per tow basis, these spring and autumn survey catches equate to maximum stratified mean number per tow indices for the Mid-Atlantic offshore and inshore strata set of 1.2-1.6 fish, or about 0.8-0.9 kg, per tow during the mid 1990s and 2000s (Tables B2.15-B2.16).

The catchability of clearnose skate in the NEFSC winter bottom trawl survey (which substitutes a chain sweep with small cookies for the large rollers used in the spring and autumn surveys, to better target flatfish) is significantly higher than in the spring and autumn series. NEFSC winter survey (1992-2006) annual catches of clearnose skate have ranged from 343 fish in 1999 to 3,086 fish in 1996, equating to a maximum stratified mean catch per tow of 12 fish or 15 kg per tow in 1996 (Table B2.17). The winter survey is focused in the Southern New England and Mid-Atlantic offshore regions, with a limited number of samples on Georges Bank, and no sampling in the Gulf of Maine (Figure B2.97).

NEFSC spring and autumn survey indices for clearnose skate have been increased from the mid-1980s through 2000 and have since declined to about average values (Figure B2.98). The combining of strata had more impact for clearnose skate than for other species, since many of the inshore strata were combined and the most southern strata were combined into one stratum (Figures B2.99-B2.102).

The minimum length of clearnose skate caught in NEFSC surveys is about 10 cm (4 inches), and the largest individual caught was 93 cm ( 33 in ) total length, during the 1992 and 2000 winter surveys in the Mid-Atlantic Bight region (Tables B2.15-B2.17). The median length of the survey catch has ranged from 41 cm in the 1980 spring survey to 67 cm in the 1995 spring survey. The median length of the spring survey catch has increased over the time series, from about 50 cm during the late 1970s to at about 60 cm in recent years ( 24 inches; Figure B2.103). The median length of the autumn survey catch has been stable over the time series, and is also at about 60 cm . Length frequency distributions from the NEFSC spring and autumn surveys show a consistent mode at $60-70 \mathrm{~cm}$ that may represent the accumulated abundance of several older ages (Figures B2.104-B2.107).

The strata set used for bootstrapping the winter survey differed from the standard consistent strata set used for the information in Table 2.17. Given that the strata on Georges Bank were not sampled in some years, the set for bootstrapping was limited to a few Southern New England strata and the Mid-Atlantic (Table B2.1). This created more of a difference between the original mean, with usually a lower index when Georges Bank was included in the original (Figure B2.108-B2.109). The indices of both abundance and biomass have generally fluctuated without trend.

Indices of abundance for clearnose skate are available from the Connecticut Department of Environmental Protection (CTDEP) spring and autumn finfish trawl surveys in Long Island Sound for the years 1984-1998 (1992 and later only for biomass). The CTDEP survey had caught very few clearnose skate, with annual catches ranging from 0 to 20 skates through 1998, but the indices have increased in Long Island Sound over the times series (Figure B2.110).

## Rosette skate

NEFSC bottom trawl surveys indicate that rosette skate are most abundant in the MidAtlantic offshore strata region, with very few fish caught in Southern New England and Georges Bank and no fish caught in the Gulf of Maine or inshore (Figure B2.111). In the NEFSC spring surveys (1968-2006), the annual total catch of rosette skate has ranged from 0 fish, in 1970 and1984, to 70 fish in 1977 (Table B2.18). In the NEFSC autumn surveys (1967-2005), the annual total catch of rosette skate has ranged from 1 fish, most recently in 1982, to 46 fish in 1999 (Table B2.19). Calculated on a per tow basis, these spring survey catches equate to maximum stratified mean number per tow indices for the Mid-Atlantic offshore strata set of about 0.6 fish, or about 0.1 kg , per tow during 1977 (Tables B2.18-B2.19).

The catchability of rosette skate in the NEFSC winter bottom trawl survey (which substitutes a chain sweep with small cookies for the large rollers used in the spring and autumn surveys, to better target flatfish) is significantly higher than in the spring and autumn series. NEFSC winter survey (1992-2006) annual catches of rosette skate have ranged from 143 fish in 1993 to 1029 fish in 2003, equating to a maximum stratified mean catch per tow of 2.8 fish or 0.7 kg per tow in 2003 (Table B2.20). The winter survey is focused in the Southern New England and Mid-Atlantic offshore regions, with a limited number of samples on Georges Bank, and no sampling in the Gulf of Maine (Figure B2.112).

Indices of rosette skate abundance and biomass from the NEFSC surveys were at a peak during 1975-1980, before declining through 1986. NEFSC survey indices for rosette skate increased from 1986 through 2001, declined slightly and recent indices are near the peak values of the late 1970s (Figure B2.113). The combining of strata had more impact for rosette skate than for other species, since the deep offshore strata were combined with the next deepest stratum and the most southern strata were combined into one stratum (Figures B2.114-B2.117).

The minimum length of rosette skate caught in NEFSC surveys is about 7 cm (3 inches), and the largest individual caught was 57 cm ( 22 inches) total length, during the 1971 spring survey in the Mid-Atlantic Bight region (Tables B2.18-B2.20). The median length of the survey catch has ranged from 18 cm in the 1985 spring survey to 57 cm in the 1971 spring survey, during which only 1 rosette skate was caught. The median length of the survey catch has been stable over the spring and autumn time series at about $36-37 \mathrm{~cm}$ (14 inches; Figure B2.118). Length frequency distributions from the NEFSC spring and autumn surveys show a consistent mode at $30-40 \mathrm{~cm}$ (Figures B2.119-B2.123).

The strata set used for bootstrapping the winter survey differed from the standard consistent strata set used for the information in Table 2.17. Given that the strata on Georges Bank were not sampled in some years and the deepwater strata which are important for rosette skate were not sampled until 1998, the set for bootstrapping was limited to a few Southern New England strata and the Mid-Atlantic from 1998 on (Table B2.1). This created more of a difference between the original mean, with usually a lower index when Georges Bank was included in the original (Figure B2.124-B2.125). The indices of both abundance and biomass increased through 2002 and have subsequently declined.

### 4.2 Research survey data - Spawning Stock Biomass

Maturity information was available in some form for all species to split the survey length information into mature and immature animals (Table 2.21). The series chosen for each species was the same as chosen for reference points at SARC30. There is a protracted spawning as females likely lay eggs year round so there is no need to pick a season based on spawning time. As it is generally the longest running series, the autumn survey was used for all species except
little skate. For little skate, the spring series from 1982 on was used; this date was chosen to avoid gear conversion issues.

Winter skate SSB generally follows the pattern of the autumn total biomass index with very low values in the 1970s followed by the large expansion of the size composition in the 1980s (Table B2.22; Figure B2.126). The index of SSB declined in the mid- to late 1990s, increased slightly, and is currently at low values. Little skate SSB has been fairly stable through the time series with slightly higher values from 1999-2004 than in the 1980s and early 1990s (Table B2.22; Figure B2.126). The pattern in barndoor skate SSB indices is much the same as that of total biomass with high values in the early 1960s, followed by very low to nonexistent values in the 1970s and 1980s, and then a consistent increase in the 1990s and 2000s (Table B 2.22 ; Figure B2.126). The decline in thorny skate SSB indices is more pronounced than for the total biomass index (Table B2.22; Figure B2.126). Smooth skate SSB indices are very variable, but exhibit a slight decline over the time series (Table B2.22; Figure B2.126). Clearnose skate SSB has increased over the time period (Table B2.22; Figure B2.126). Rosette skate SSB has been variable but has generally increased (Table B2.22; Figure B2.126).

### 4.3 Fishing mortality estimates

The length-based mortality estimators of Beverton and Holt (1956) and Hoenig (1987) were considered for the estimation of fishing mortality rates for winter, little, barndoor, thorny, and clearnose skates from NEFSC spring and autumn length frequency distributions. Only these five species were analyzed since age and growth information is available for these species and unavailable for rosette and smooth (Table 2.21).

## (EDITOR'S NOTE: MODEL-BASED FISHING MORTALITY ESTIMATES WERE PROPOSED; THEY ARE NOT SHOWN BECAUSE THEY WERE NOT ACCEPTED BY THE REVIEW PANEL)

### 4.3.1 Mortality from Mean Length Gedamke and Hoenig (2006) Method

Gedamke and Hoenig (2006) developed a method to estimate mortality from mean length data in nonequilibrium situations. It is an extension of the Beverton-Holt length-based mortality estimator that assumes constant recruitment throughout the time series and mortality at fixed levels for certain periods within the time series. The approach allows for the transitory changes in mean length to be modeled as a function of mortality rate changes. After an increase in mortality, mean length will gradually decrease due to larger animals being less prevalent in the population. After a decrease in mortality, mean length will increase slowly due to growth of the fish in the population. The rates of change in both cases depend on the von Bertalanffy growth parameters and the magnitude of change in the mortality rates. Since the method requires only a series of mean length above a user defined minimum size and the von Bertalanffy growth parameters, it can be applied in many data poor situations. Gedamke and Hoenig (2006) demonstrated the utility of this approach using both simulated data and an application to data for goosefish caught in the NEFSC fall groundfish survey.
(EDITOR'S NOTE: FISHING MORTALITY ESTIMATES WERE PROPOSED; THEY ARE NOT SHOWN BECAUSE THEY WERE NOT ACCEPTED BY THE REVIEW PANEL)

### 4.3.2 Thorny Skate Length Tuned Model (LTM)

## Introduction

A forward projecting length tuned model (LTM) was modified to fit only survey abundance indices and survey size information for the estimation of fishing mortality rates. Results from this analysis were compared to the Hoenig length based estimates to help determine the influences of assuming equilibrium conditions. The LTM model does not assume equilibrium conditions since fishing mortality estimates in year $n$ will influence the population size structure in year $n+1$. However the initial population in year one of the model is calculated assuming equilibrium conditions.

Herein we used a simple forward projecting age-based model tuned with age-3 recruitment (estimated from fish in the survey that were between 35 and 45 cm ), survey numbers of $40+\mathrm{cm}$ fish and length frequency of the $40+\mathrm{cm}$ fish. The Length Tuned Model was developed in the AD model builder framework. The model estimates fishing mortality and relative recruitment changes each year, fishing mortality to produce the initial population length frequency ( $\mathrm{F}_{\text {start }}$ ), and Qs for each survey index. Initial population abundance was fixed since no catch information can be used to scale the model in terms of abundance.
(EDITOR'S NOTE: RESULTS FROM THIS MODEL ARE NOT SHOWN BECAUSE THEY WERE NOT ACCEPTED BY THE REVIEW PANEL)

### 5.0 TOR 3. Either update or redefine biological reference points (BRPs; proxies for BMSY and FMSY), as appropriate. Comment on the scientific adequacy of existing and redefined BRPs.

### 5.1 Current Reference Points

The existing biomass reference points were developed at SARC 30 (NEFSC 2000) with $\mathrm{B}_{\mathrm{msy}}$ Proxy formulated as the $75^{\text {th }}$ percentile of the given time series of each species, except barndoor (Table B3.1) and half that value for $\mathrm{B}_{\text {threshold. }}$. It was assumed that all species had at some time passed through $\mathrm{B}_{\text {msy }}$ at some point in the time series. For barndoor skate, the mean of the first four years of the autumn survey were used instead, given that biomass had been extremely low during most of the time series. To reduce the variability in the survey estimates, a three-year moving average of the survey indices was proposed to evaluate stock status for all species.

The fishing mortality reference points developed at SARC 30 were not accepted by the NEFMC and a different method for evaluating fishing mortality was developed by the Plan Development Team (PDT). The thresholds for fishing mortality are based on annual percentage declines of the three-year average of the NEFSC trawl survey time series chosen for the biomass reference points. The percentages are specified for each species individually based on historical variation within the survey. The thresholds also include what is termed a precautionary "backstop" that indicates that overfishing is occurring if the trawl survey mean weight per tow declines for three consecutive years. The main part of the definition is that overfishing is occurring when the three-year moving average of the given survey biomass index declines by more than the average CV of the time series.

### 5.2 Alternative Reference Points

## (EDITOR'S NOTE: ALTERNATIVE REFERENCE POINTS WERE PRESENTED; THEY ARE NOT SHOWN BECAUSE THEY WERE NOT ACCEPTED BY THE REVIEW PANEL)

### 6.0 TOR 4. Evaluate current stock status with respect to the existing BRPs, as well as with respect to updated or redefined BRPs (from TOR 3).

### 6.1 Current Reference Points

For winter skate, the 2003-2005 NEFSC autumn survey biomass index average of 3.34 $\mathrm{kg} /$ tow is below the biomass target of $6.46 \mathrm{~kg} /$ tow but above the threshold reference point of 3.23 $\mathrm{kg} /$ tow (Figure B4.1). Winter skate is not overfished. The 2003-2005 average of $3.34 \mathrm{~kg} /$ tow was more than $20 \%$ below the 2002-2004 average of $4.34 \mathrm{~kg} /$ tow (Table B4.1), therefore overfishing is occurring for winter skate.

For little skate, the 2004-2006 NEFSC spring survey biomass index average of 4.59 $\mathrm{kg} /$ tow is below the biomass target of $6.54 \mathrm{~kg} /$ tow but above the threshold reference point of 3.27 $\mathrm{kg} /$ tow (Figure B4.1). Little skate is not overfished. The 2004-2006 average of $4.56 \mathrm{~kg} / \mathrm{tow}$ was less than $20 \%$ below the 2003-2005 average of $5.65 \mathrm{~kg} /$ tow (Table B4.1), therefore overfishing is not occurring for little skate.

For barndoor skate, the 2003-2005 NEFSC autumn survey biomass index average of 0.96 $\mathrm{kg} /$ tow is below the biomass target of $1.62 \mathrm{~kg} /$ tow but above the threshold reference points of $0.81 \mathrm{~kg} /$ tow(Figure B4.1). Barndoor skate is not overfished. The 2003-2005 average of 0.96 $\mathrm{kg} /$ tow was above the 2002-2004 average of $0.88 \mathrm{~kg} /$ tow (Table B4.1), therefore overfishing is not occurring for barndoor skate.

For thorny skate, the 2003-2005 NEFSC autumn survey biomass index average of 0.56 $\mathrm{kg} /$ tow is below the biomass target and threshold reference points of $4.41 \mathrm{~kg} / \mathrm{tow}$ and 2.20 $\mathrm{kg} /$ tow (Figure B4.1). Thorny skate is overfished. The 2003-2005 average of $0.56 \mathrm{~kg} / \mathrm{tow}$ was less than $20 \%$ below the 2002-2004 average of $0.63 \mathrm{~kg} /$ tow (Table B4.1), therefore overfishing is not occurring for thorny skate.

For smooth skate, the 2003-2005 NEFSC autumn survey biomass index average of 0.18 $\mathrm{kg} /$ tow is below the biomass target of $0.31 \mathrm{~kg} /$ tow but above the threshold reference point of 0.16 $\mathrm{kg} /$ tow (Figure B4.1) . Smooth skate is not overfished. The 2003-2005 average of $0.18 \mathrm{~kg} /$ tow was above the 2002-2004 average of $0.17 \mathrm{~kg} /$ tow (Table B4.1), therefore overfishing is not occurring for smooth skate.

For clearnose skate, the 2003-2005 NEFSC autumn survey biomass index average of 0.63 $\mathrm{kg} /$ tow is above the biomass target and threshold reference points of $0.56 \mathrm{~kg} /$ tow and 0.28 $\mathrm{kg} /$ tow (Figure B4.1). Clearnose skate is not overfished. The 2003-2005 average of $0.63 \mathrm{~kg} / \mathrm{tow}$ was less than $30 \%$ below the 2002-2004 average of $0.75 \mathrm{~kg} /$ tow (Table B4.1), therefore overfishing is not occurring for clearnose skate.

For rosette skate, the 2003-2005 NEFSC autumn survey biomass index average of 0.049 $\mathrm{kg} /$ tow is above the biomass target and threshold reference points of $0.029 \mathrm{~kg} /$ tow and 0.015 $\mathrm{kg} /$ tow (Figure B4.1) . Rosette skate is not overfished. The 2003-2005 average of $0.049 \mathrm{~kg} /$ tow was above the 2002-2004 average of $0.045 \mathrm{~kg} /$ tow (Table B4.1), therefore overfishing is not occurring for rosette skate.

### 7.0 TOR 5. Review, evaluate and report on the status of the SARC/Working Group Research Recommendations offered in recent SARC-reviewed assessments.

1) The commercial fishery statistics sampling programs should be adapted to report skates landings by species.

Since the implementation of the Skate Complex FMP, there is a requirement to report landings of skates by species. However, training is needed to improve the accuracy of the reporting.
2) Commercial fishery size composition data should be collected by species.

Observers are collecting landings and discarded size composition by species. However, more training is needed to improve the accuracy of the data.
3) Sea sampling of directed skate landings and skate bycatch should be increased, and the identification of the species composition of the skate catch improved.

Observer coverage was increased in 2004 and 2005 primarily for the multi-species groundfish fisheries which have a large bycatch of skates. Observer coverage of scallop fisheries has improved as well. More training is needed to improve the accuracy of the species identification.
4) Age and growth studies, for all seven species in the complex, are needed.

Studies have been conducted for five of the seven species (Frisk 2004,Gedamke 2006, Gedamke et al. 2005, Gelschleiter 1998, Sulikowski et al. 2005) and samples have been collected by NEFSC for the other two species.
5) Maturity and fecundity studies, for all seven species in the complex, are needed. Use of life history models requires these data, and may prove useful in establishing biological reference points for the skate species.

Maturity studies estimating $L_{50}$ have been conducted for barndoor (Gedamke 2005), winter and little (Frisk 2004), and thorny (Sulikowski et al. 2006). Sosebee (2005) estimated size at first maturity for all seven species.
6) Estimates of commercial and recreational fishery discard mortality rates, for different fishing gears and coastal regions and/or bottom types, for all seven species in the complex, are needed.

Not completed.
7) Studies of the stock structure of the species in the skate complex are needed to identify unit stocks. Stock identification studies, especially for barndoor, thorny, winter, and little skate, are needed.

Not completed.
8) Explore possible stock-recruit relationships by examination of NEFSC survey data. A simultaneous examination of the species in the complex may prove a useful first step.

Stock-recruit relationships have been examined for five of the species in the complex. The second method is not appropriate for skates.
9) Investigate trophic interactions between skate species in the complex, and between skates and other groundfish.

Considerable progress has been made.
10) Further consideration of the validity of NEFSC trawl survey catchability conversion factors for skate species is needed (diel, gear, vessel).

Not completed.
11) Investigate the influence of annual changes in water temperature or other environmental factors on shifts in the range and distribution of the species in the skate complex. Establish the bathymetric distribution of the species in the complex off the U.S. Northeast coast.

Work has been done on winter skate to explore the changes in abundance between the Scotian Shelf and Georges Bank (Frisk et al, in review).
12) Investigate the SEAMAP survey data for clearnose and rosette skate.

Not completed.
13) Investigate historical NEFSC survey data from the Albatross III cruises during 1948-1962 when they become readily accessible, as they may provide valuable historical context for long term trends in skate biomass.

Not completed.
14) Recalculate the error distributions of the survey indices using alternative distributions.

Instead of assuming an error distribution, confidence intervals were derived using the bootstrap methods of Smith (1997).

### 8.0 TOR 6. Examine the NEFSC Food Habits Database to estimate diet composition and annual consumptive demand for seven species of skates for as many years as feasible.

### 8.1 Introduction

Skate food habits were evaluated for all seven species in the skate complex. The total amount of food eaten and the type of food eaten were the primary food habits data examined. From these basic food habits data, diet composition, per capita consumption, total consumption, and the amount of prey removed by skates were calculated. Contrasts to total energy flows in the ecosystem and fishery removals of commercially targeted skate prey were conducted to fully address the Term of Reference.

### 8.2 Methods

Each skate was analyzed separately; emphasizing at least two if not three size classes as appropriate (Table B6.1). These size classes correspond to notable changes in diet and life history and also minimized low data density (i.e., number of stomachs sampled) for each size class. Each skate was analyzed for a particular bottom trawl survey strata set germane for each case (Table B6.1). For all the estimates, small winter skates ( $<30 \mathrm{~cm}$ ) were grouped with immature little skates. Estimates were analyzed on an annualized basis for each species, save instances were data density of stomach samples was too low. In those cases data were evaluated across 5 -year time blocks. Although the food habits data collections started quantitatively in 1973, not all species of skates were sampled during the full extent of this sampling program. For more details on the food habits sampling protocols and approaches, see Link and Almeida (2000). Where data are available, they are used except in the case of little skate (see above for discussion on why those estimates begin in 1982). This sampling program was a part of the NEFSC bottom trawl survey program; for background and context, further details of the survey program can be found in Azarovitz (1981) and NEFC (1988).

## Basic Food Habits

To estimate mean stomach contents $\left(S_{i}\right)$, each skate had the total amount of food eaten (as observed from food habits sampling) calculated for each size class, temporal and spatial scheme. The denominator in the mean stomach contents (i.e., the number of stomachs sampled) was inclusive of empty stomachs. These means were weighted by the number of tows in a temporal and spatial scheme as part of a two-stage cluster design. Further particulars of these estimators can be found in Link and Almeida (2000). Units for this estimate are in g.

To estimate diet composition $\left(D_{i j}\right)$, the amount of each prey item was summed across all skate stomachs. These estimates were then divided by the total amount of food eaten in a size class, temporal and spatial scheme, totaling $100 \%$. These estimates are proportions and were only presented for those major prey comprising $>85 \%$ of the total for each size class, temporal and spatial scheme. Further particulars of these estimators can be found in Link and Almeida (2000).

## Consumption Rates

To estimate per capita consumption, the gastric evacuation rate method was used (Eggers 1977, Elliott and Persson 1978). There are several approaches used for estimating consumption, but this approach was chosen as it was not overly simplistic (as compared to \% body weight; Bajkov 1935) or overly complex (as compared to highly parameterized bioenergetics models; Kitchell et al. 1977). There has been extensive use of these models (Durbin et al. 1983, Ursin et al. 1985, Pennington 1985, Overholtz et al. 1991, 1999, 2000, Tsou \& Collie 2001a, 2001b, Link \& Garrison 2002, Link et al. 2002, 2006, Overholtz \& Link 2007). Units are in g year ${ }^{-1}$.

Using the evacuation rate model to calculate consumption requires two variables and two parameters. The per capita consumption rate, $C_{i}$ is calculated as:

$$
C_{i}=24 \cdot E_{i} \cdot \bar{S}_{i}^{\gamma}
$$

where 24 is the number of hours in a day and the evacuation rate $E_{i}$ is:

$$
E_{i}=\alpha e^{\beta T}
$$

and is formulated such that estimates of mean stomach contents $\left(S_{i}\right)$ and ambient temperature ( $T$; here used as bottom temperature from the NEFSC bottom trawl surveys (Taylor and Bascunan 2000; Taylor et al. 2005) are the only data required. The parameters $\alpha$ and $\beta$ are set as values chosen from the literature (Tsou and Collie 2001a, 2001b, Overholtz et al. 1999, 2000). The parameter $\gamma$ is a shape function is almost always set to 1 (Gerking 1994).

To evaluate the performance of the evacuation rate method for calculating consumption, a simple sensitivity analysis was executed. The first phase of the sensitivity analysis fixed the two parameters and two variables, varying them one at a time. These varied across both the normal range from the data or literature and across proximal orders of magnitude to the normative range. The second phase varied all two pairs of values simultaneously, presented as surface plots to denote areas of rapid change and areas of relative stability (flat surfaces).

## Scaling Consumption

After per capita consumption rates were estimated for each skate in a size class, temporal and spatial scheme, those estimates were scaled up to an annual and stock wide basis, $C$ :

$$
C=365 \cdot C_{i} \cdot N_{i}
$$

where $N_{i}$ is the swept area estimate of abundance for each skate in each size class, temporal and spatial scheme and 365 is the number of days in a year.

This total consumption was partitioned for the major prey items of each skate by multiplying it by the diet composition of each prey $\left(D_{i j}\right)$ to provide an estimate of prey removals by each skate. Both the total consumption and the amount of prey removed by each skate are presented as metric tons year ${ }^{-1}$.

To evaluate the consumptive demands of a skate stock and the predatory removals of a skate stock in a broader ecosystem context, two contrasts were executed. First, comparisons of total consumption by each skate and by all skates combined were compared to the amount of energy flows for the entire ecosystem. These total energy flows were calculated in a recent energy budget (Link et al. 2006). Skate consumption is presented as a percentage of total energy flows in the ecosystem.

Second, the total amount of commercially targeted prey eaten by skates was treated as a removal and summed across all skates. These estimates were then compared to concurrently estimated fishery landings to provide an evaluation of potential competition between skates and fisheries on some of their major prey.

One concern of this approach is that the abundance estimates used to scale per capita skate consumption up to total population level consumption were not corrected for catchability or gear efficiency of the bottom trawl survey. To evaluate the potential effect of this factor, efficiencies of $100,50,25$ and $10 \%$ were applied to estimates of total prey removal by all skates.

### 8.3 Results

## Sensitivity analysis

The fixed values for all parameters were mean stomachs, $S_{i}=10$, mean bottom temperature, $T=10$, scaling coefficient $\alpha=0.02$, and exponent coefficient $\beta=0.111$. The parameters are consistent with literature values for other elasmobranchs (Tsou and Collie 2001a, 2001b).

Examining the sensitivity to mean stomach contents demonstrates a clear linear relationship to per capita consumption across the full range of observed skate stomachs (Figure B6.1a). This is obvious the one factor that most highly data driven and represents an intuitive relationship- the more food measured that a skate eats, the higher the annual per capita consumption. The range of food consumed can be anywhere from 50 g to 60 kg , consistent with observed food habits for this species complex.

Examining the sensitivity to mean bottom temperature demonstrates a curvilinear relationship with per capita consumption (Figure B6.1b). The upper tail of the range (i.e., $>$ $15^{\circ} \mathrm{C}$ ) represents an increase up to $10-20 \mathrm{~kg}$ consumed per year. However, the per capita consumption in the range of typical temperatures encountered by skates are on the order of 4-6 kg per year.

Examining the sensitivity to changes in $\alpha$ similarly demonstrates a curvilinear relationship with per capita consumption (Figure B6.2a), albeit with $\alpha$ presented on a logarithmic scale. This relationship is much more convex than with temperature, with consumption values where $\alpha \sim 0.1$ approaching 30 kg per year. However, within the range of $\alpha$ typically reported from the literature $(\alpha=0.01$ to 0.05$)$ results in a consumption on the order of $5-10 \mathrm{~kg}$ per year.

Examining the sensitivity to changes in $\beta$ also demonstrates a curvilinear relationship with per capita consumption (Figure. B6.2b). At the upper tail of the analysis with $>0.2$ results in a consumption estimate of $15-20 \mathrm{~kg}$ per year. However, within the range of $\beta$ typically reported from the literature ( $\beta=0.1$ to 0.12 ) results in a consumption on the order of $5-7 \mathrm{~kg}$ per year.

The most sensitive factor, when within normal range, is mean stomach contents of these skates.

Examining some salient pairs, one sees that categorically when looking at the upper end of mean stomach contents versus $\beta, \alpha$ or $T$ (Figures B6.3-B6.5) there is a clear spike at the upper range of any of those three factors with stomach contents. These peaks can result in per capita consumption estimates of over 300 kg per year. However, when one looks at the typical range of $\beta, \alpha$ or $T$ the surfaces are much flatter and more stable, even at the upper range of $S_{i}$. A similar pattern emerges when comparing $\beta$ and $\alpha$ (Figure B6.6). Yet even this maximum-maximum range is on the order of 120 kg per capita consumption per year, much less than when including $S_{i}$. This surface is also much flatter than the other ones that include $S_{i}$.

To put the sensitivity analysis in perspective, when both parameters were within the normal range, the change to per capita consumption was $<$ half to one order of magnitude. The temperature variable across the maximum possible range only changes the per capita consumption by $<$ an order of magnitude. Most observed temperature ranges are $\ll$ quarter of an order of magnitude.

An order of magnitude change in the amount of food eaten results in an order of magnitude change in per capita consumption. Variance about any particular species of skate has a CV of $\sim 50 \%$. Thus, within any given species for each size class, temporal and spatial scheme, the
variability of $S_{i}$ is likely to only influence per capita consumption by half an order of magnitude or less.

Estimates of abundance, and changes in estimates thereof, are likely going to dominate the scaling of total consumption by a broader range of magnitudes than the parameters and variables requisite for an evacuation method of estimating consumption.

## Winter Skate

The mean stomach contents for winter skate show a relatively stable amount of food eaten for both size classes (Figure B6.7a). Small winter skates ( $<30 \mathrm{~cm}$ ) were grouped with immature little skates. In instances with large error bars, there is an appearance of a major increase in food eaten during the early 1980s, yet this may be due to limited sample sizes during that period. Except the early 1980s, the number of empty stomachs has remained similar across the time period, averaging $\sim 20 \%$ and $\sim 25 \%$ for the medium and large size classes respectively (Figure B6.7b).

The mean length of skates sampled for stomach contents was consistent over time, averaging approximately 45 cm and 80 cm for medium and large size classes respectively (Figure B6.8a). There is a relationship between the size of skates and the amount of food eaten by skates, despite the wide variability in a few years (Figure B6.8b).

The temperature for these strata (and the environment which this skate was experiencing) ranged between 7 and $10^{\circ} \mathrm{C}$ (Figure B6.9a).

The per capita consumption of this skate (Figure B6.9b) generally tracks the amount of food eaten (Figure B6.7a). Values average approximately 2 kg per year for the medium size class and between 9 kg per year for the large size class.

Total minimal estimates of swept area abundance (Figure B6.10a) are generally comparable to estimates noted above. There was generally no trend for all three size classes over the entire time period except the large size. The large winter skates class exhibited a peak in the 1980s followed by a notable decline in the 1990s, with some recovery now apparent in more recent years. This is one of the more abundant skate species.

Total consumption when scaled to the population level generally tracks abundance more than any other contributing factor (Figure B6.10b). Both size classes show a peak in the 1980s, consistent with the observed peak in the abundance of the larger size class (Figure B6.10a). Estimates here for total consumptive demand by this skate range between 20,000 and 180,000 MT per year.

The diet composition of winter skate is reflective of the generally benthivorous diet of all skates and the piscivorous nature of particularly larger skates (Table B6.2). Major prey of this skate are primarily forage fishes (herrings, hakes) or benthic megafauna (crabs, shrimp). The category other fish refers to those species that are not primarily commercially targeted. The category other crabs refers to those crabs that are not in the genus Cancer or Paguroidean family.

When allocating total consumption of winter skate proportionally to each prey item, forage fish, squids, and benthic macrofauna are clearly the major amount of prey removed by this skate (Figures B6.11-B6.12). Up to 80,000 MT of a particular prey item can be removed by this skate in any given year.

## Little Skate

The mean stomach contents for Little Skate show an increasing amount of food eaten in the 1980s for the both size classes, followed by a more stable amount during the past 20 years (Figure B6.13a). The number of empty stomachs has remained mostly similar across the time
period, averaging $\sim 10 \%$ for both size classes (Figure B6.13b). Recall that small winter skates ( $<$ 30 cm ) are grouped in with the immature little skates.

The mean length of skates sampled for stomach contents was consistent over time, averaging approximately 20 cm and 45 cm for immature and mature size classes (Figure B6.14a). There is a clear relationship between the size of skates and the amount of food eaten by skates (Figure B6.14b).

The temperature for these strata (and the environment which this skate was experiencing) ranged between 7 and $11^{\circ} \mathrm{C}$ (Figure B6.15a).

The per capita consumption of this skate (Figure B6.15b) generally tracks the amount of food eaten (Figure B6.13a). Values average approximately 500 g per year for immatures and 2.5 kg per year for matures.

Total minimal estimates of swept area abundance (Figure B6.16a) are generally comparable to estimates noted above. There were some fluctuations during the later 1990s and early 2000 s, but these were centered about, and returned to, the long term average abundance. This was the most abundant skate species in the ecosystem.

Total consumption when scaled to the population level generally tracks abundance more than any other contributing factor (Figure B6.16b). Both size classes exhibit a reasonably stable amount of food eaten, but the total consumption is dominated by the mature size class (Figure B6.16a). Estimates here for total consumptive demand by this skate range between 100,000 and 350,000 MT per year.

The diet composition of little skate is reflective of the generally benthivorous nature of all skates (Table B6.3). Most of the major prey of this skate are comprised of benthic macrofauna (polychaetes, amphipods) or benthic megafauna (crabs, bivalves).

When allocating total consumption of little skate proportionally to each prey item, benthic invertebrates are clearly the major amount of prey removed by this skate (Figure B6.17). Up to $100,000 \mathrm{MT}$ of a particular prey item can be removed by this skate in any given year.

## Barndoor Skate

The mean stomach contents for barndoor skate show a relatively stable amount of food eaten for the immature size class (Figure B6.18a). In the larger size class there are instances with large error bars, giving an appearance of a major decline in food eaten circa 2002 to 2003. Yet this may be due to limited sample sizes during 2002. The number of empty stomachs has remained similar across the time period, averaging $\sim 25 \%$ for both size classes (Figure B6.18b).

The mean length of skates sampled for stomach contents was consistent over time, averaging slightly less than 60 cm and slightly over 100 cm for immature and mature size classes respectively (Figure B6.19a). There is a clear relationship between the size of skates and the amount of food eaten by skates, despite the wide variability in a few years (Figure B6.19b).

The temperature for these strata (and the environment which this skate was experiencing) ranged between 7 and $9^{\circ} \mathrm{C}$, declining slightly in more recent years (Figure B6.20a).

The per capita consumption of this skate (Figure B6.20b) generally tracks the amount of food eaten (Figure B6.18a). Values typically range approximately 5 kg per year for immatures and between 10 to 20 kg per year for matures.

Total minimal estimates of swept area abundance (Figure B6.21a) are generally comparable to estimates noted above. There was a generally increasing trend for both size classes over time, although numbers are still relatively low.

Total consumption when scaled to the population level generally tracks abundance more than any other contributing factor (Figure B6.21b). Both size classes show a peak in 2002,
consistent with the observed peak in mean stomach contents (Figure B6.18.a). Estimates here for total consumptive demand by this skate range between 4,000 and 16,000 MT per year.

The diet composition of barndoor skate is reflective of the generally benthivorous nature of all skates and the piscivorous nature of particularly larger skates (Table B.6.4). Most of the major prey of this skate are comprised of forage fishes (herrings, hakes) or benthic megafauna (crabs, shrimp). The category other fish refers to those species that are not primarily commercially targeted. The category other crabs refers to those crabs that are not in the genus Cancer or Paguroidean family.

When allocating total consumption of barndoor skate proportionally to each prey item, herrings, Pandalid shrimps, and Cancer crabs are clearly the major amount of prey removed by this skate (Figure B6.22). Up to 8,000 MT of a particular prey item can be removed by this skate in any given year.

## Thorny Skate

The mean stomach contents for thorny Skate show a relatively stable amount of food eaten for two of the three size classes, with medium skates exhibiting a slight increase (Figure B6.23a). Aside from the 1976 to 1980 time period (five year block), the number of empty stomachs has remained similar across the time period, averaging $\sim 15$ to $20 \%$ for all size classes (Figure B6.23b).

The mean length of skates sampled for stomach contents was consistent over time for all three size classes, averaging approximately $20 \mathrm{~cm}, 45 \mathrm{~cm}$, and slightly less than 80 cm for the small, medium and large size classes respectively (Figure B6.24a). There is a clear relationship between the size of skates and the amount of food eaten by skates (Figure B6.24b).

The temperature for these strata (and the environment which this skate was experiencing) ranged between 7 and $9^{\circ} \mathrm{C}$, declining slightly in more recent years (Figure B6.25a).

The per capita consumption of this skate (Figure B6.25b) generally tracks the amount of food eaten (Figure B6.23a). Values average approximately 500 g per year for the small size class, 1.5 kg per year for the medium size class, and 12 kg per year for the large size class.

Total minimal estimates of swept area abundance (Figure B6.26a) are generally comparable to estimates noted above. There was a clear declining trend for all size classes over time, although numbers are still relatively low.

Total consumption when scaled to the population level generally tracks abundance more than any other contributing factor (Figure B6.26b). All three size classes show a peak in the early 1980s, consistent with the observed peak in mean stomach contents (Figure B6.23a). Estimates here for total consumptive demand by this skate range between 10,000 and $40,000 \mathrm{MT}$ per year.

The diet composition of thorny skate is reflective of the generally benthivorous nature of all skates and the piscivorous nature of particularly larger skates (Table B6.5). Most of the major prey of this skate are comprised of forage fishes (herrings, hakes) or benthic megafauna (crabs, euphasiids). The category other fish refers to those species that are not primarily commercially targeted. The category other crabs refers to those crabs that are not in the genus Cancer or Paguroidean family.

When allocating total consumption of thorny skate proportionally to each prey item, herrings, squids, polychaetes, silver hake and other fish are the major amount of prey removed by this skate (Figures B6.27-B6.28). Up to $8,000 \mathrm{MT}$ of a particular prey item can be removed by this skate in any given year.

## Smooth Skate

The mean stomach contents for Smooth Skate show a relatively stable amount of food eaten for both size classes (Figure B6.29a). The number of empty stomachs has remained stationary across the time period, albeit with a wide range of variability (particularly for immatures), averaging $\sim 15$ to $20 \%$ for both size classes (Figure B6.29b). There were no empties for one part of the time series.

The mean length of skates sampled for stomach contents was consistent over time, averaging around $20-25 \mathrm{~cm}$ and 50 cm for immature and mature size classes respectively (Figure B6.30a). There is a clear relationship between the size of skates and the amount of food eaten by skates (Figure B6.30b).

The temperature for these strata (and the environment which this skate was experiencing) ranged between 7 and $10^{\circ} \mathrm{C}$ (Figure B6.31a).

The per capita consumption of this skate (Figure B6.31b) generally tracks the amount of food eaten (Figure B6.29a). Values typically range between 0.5 to 1 kg per year for immatures and 2 to 3 kg per year for matures. Because these stomachs were calculated in five year time blocks, these estimates reflect that periodicity.

Total minimal estimates of swept area abundance (Figure B6.32a) are generally comparable to estimates noted above. There was a lot of variability and the abundance of both size classes varied without trend.

Total consumption when scaled to the population level generally tracks abundance and amount of food consumed more than any other contributing factors (Figure B6.32b). Both size classes are highly variable, with the majority of the consumption for this population occurring in the mature size class.. Estimates for total consumptive demand by this skate range between 1,000 and 5,000 MT per year.

The diet composition of smooth skate is reflective of the generally benthivorous nature of all skates (Table B6.6). Most of the major prey of this skate are comprised of common benthic megafauna (pandalids, euphausiids).

When allocating total consumption of smooth skate proportionally to each prey item, pandalid shrimp and euphausiids are clearly the major amount of prey removed by this skate (Figure B6.33). Up to $2,000 \mathrm{MT}$ of a particular prey item can be removed by this skate in any given year, but values are typically on the order of 500 to $1,000 \mathrm{MT}$.

## Clearnose Skate

The mean stomach contents for Clearnose Skate show a relatively stable amount of food eaten for the immature size class (Figure B6.34a). The same is true for the larger size class. In the larger size class there may be a slightly increasing trend in the amount of food eaten. In the instance with large error bars there is an appearance of a major change in the amount of food eaten. Again this may be due to limited sample sizes during that 2005. The number of empty stomachs has remained stationary across the time period, albeit with a wide range of variability (particularly for immatures), averaging $\sim 25$ to $30 \%$ for both size classes (Figure B6.34b).

The mean length of skates sampled for stomach contents was consistent over time, averaging around $45-50 \mathrm{~cm}$ and $60-65 \mathrm{~cm}$ for immature and mature size classes respectively (Figure B6.35a). There is a clear relationship between the size of skates and the amount of food eaten by skates, despite the wide variability in one year (Figure B6.35b).

The temperature for these strata (and the environment which this skate was experiencing) ranged between 9 and $13^{\circ} \mathrm{C}$ (Figure B6.36a).

The per capita consumption of this skate (Figure B6.36b) generally tracks the amount of food eaten (Figure B6.34a). Values typically range approximately 1 to 2 kg per year for immatures and 5 kg per year for matures. Because these stomachs were calculated in five year time blocks, these estimates are similar in that periodicity.

Total minimal estimates of swept area abundance (Figure B6.37a) are generally comparable to estimates noted above. There was a generally increasing trend for both size classes over time, although numbers are still relatively low.

Total consumption when scaled to the population level generally tracks abundance and amount of food consumed more than any other contributing factors (Figure B6.37b). Both size classes show a peak in 2002, consistent with the observed peak in abundance and mean stomach contents during that five year period (Figures B6.37a and B6.34a). Estimates here for total consumptive demand by this skate range between 2,000 and 18,000 MT per year.

The diet composition of clearnose skate is reflective of the generally benthivorous nature of all skates (Table B6.7). Most of the major prey of this skate are comprised of common benthic megafauna (crabs, misc. crustaceans). The category other crabs refers to those crabs that are not in the genus Cancer or Paguroidean family.

When allocating total consumption of clearnose skate proportionally to each prey item, other crabs, Cancer crabs, squids are clearly the major amount of prey removed by this skate (Figure B6.38). Up to 8,000-10,000 MT of a particular prey item can be removed by this skate in any given year, but values are typically on the order of 2,000 to $4,000 \mathrm{MT}$.

## Rosette Skate

The mean stomach contents for Rosette Skate show a relatively stable amount of food eaten for both the immature and mature size classes (Figure B6.39a). The number of empty stomachs was again around $30 \%$, but increased slightly in more recent years (Figure B6.39b).

The mean length of skates sampled for stomach contents was consistent over time, averaging approximately 22 cm and 38 cm for immature and mature size classes respectively (Figure B6.40a). There is a clear relationship between the size of skates and the amount of food eaten by skates (Figure B6.40b).

The temperature for these strata (and the environment which this skate was experiencing) ranged between 9 and $12^{\circ} \mathrm{C}$ (Figure B6.41a).

The per capita consumption of this skate (Figure B6.41b) generally tracks the amount of food eaten (Figure B6.39a). Values average approximately 200 g per year for immatures and 800 g per year for matures.

Total minimal estimates of swept area abundance (Figure B6.42a) are generally comparable to estimates noted above. There was a peak in 2001 for matures and 2002 for immatures. No major trend for both size classes was evident.

Total consumption when scaled to the population level generally tracks abundance more than any other contributing factor (Figure B6.42b). The mature size classes shows a peak in 2001 and the immatures show a peak in 2002, consistent with the observed abundances (Figure B6.42a). Estimates here for total consumptive demand by this skate range between 50 and 500 MT per year.

The diet composition of rosette skate is reflective of the generally benthivorous nature of all skates (Table B6.8). Most of the major prey of this skate are comprised of some form of benthic macrofauna (amphipods, polychaetes) or megafauna (crabs, shrimp). The category other crabs refers to those crabs that are not in the genus Cancer or Paguroidean family.

When allocating total consumption of rosette skate proportionally to each prey item, benthic macrofauna are clearly the major prey removed by this skate. Pandalid shrimps, squids, and Cancer crabs are also removed by this skate but in lesser amounts (Figure B6.43). Up to 70 MT of a particular prey item can be removed by this skate in any given year, but more typically 10-30 MT.

## All Skates relative to the ecosystem and fisheries on major prey

The total amount of skate consumption across all skates has averaged around 230,000 MT over the past $25-30$ years (Figure B6.44). This represents a relatively small amount of the total energy flow in the ecosystem. There is $3.9 \times 10^{9}$ MT of total throughput through the ecosystem (Link et al. 2006) and skate consumption represents less than $0.006 \%$ of that total energy flow in the system. The total removal of most major skate prey relative to their standing stock biomass (B) or annual production (P) is small (Table B.6.9). Estimates of B and P tend to be at least two to three orders of magnitude greater than C by all skates for any particular prey item.

Those prey which are commercially important species and which are also important skate prey can be removed by skates at a rate comparable to their fisheries (Figure B.6.44; Table B.6.10). In the minimum swept area scenario, most skate prey are on the order of one quarter or less of what is landed for those prey, with the exception of red hake. When decreasing gear efficiencies are incorporated, the relative removal by skate consumption compared with fishery removals becomes much higher. With gear efficiencies of $50 \%$, about half of fishery removals are removed by skate consumption for the two squids and silver hake, with over double removed by skates relative to the fishery for red hake. The pattern continues with increasingly less efficient assumptions, with squids and silver hake removed by skates up to twice of what is removed by the fishery at the lowest assumed value ( $10 \%$ ), while red hake is up to 10 times what is removed by the fishery. The only exception is herrings, which although have a large amount of biomass removed by skates, remain a relatively small amount of removals compared to those fishery removals.

Finally, it is worth noting that some of the potential species interactions of interest- e.g. skates eating yellowtail flounder, winter flounder, sea scallops, etc.- were not of sufficient magnitude to analyze. In fact, each of the species just mentioned as examples only comprised a very small ( $\ll 0.1 \%$ of the diet) for only one or two skate species.

### 8.4 Summary

Most skates are benthivorous in their feeding habits. A clear prominence on Cancer crabs, other crabs, amphipods, polychaetes and similar benthic macrofauna and megafauna was apparent in the diets of these skates. Some of the larger skates- barndoor, thorny, and wintercan be piscivorous, particularly with ontogeny. The vast majority of fish (or fish-like) prey for these skates were small pelagic fishes and squids.

Save winter and little skates, overall consumption by most skate stocks is a relatively small amount of biomass flow. Most total consumption by any particular species of skate was scaled singularly by the abundance of that species. The vast majority of consumptive removals by all skates except little and winter was $<20$ MT per year.

As an aggregate group, skates consume a very small fraction of the total energy flow in the ecosystem. Skate consumptive removal is two to three orders of magnitude lower than biomass or production of skate prey. When abundance estimates are scaled by gear efficiency, it is possible that skates could consume a notable fraction of forage fish and squid biomass relative
to what is removed by a fishery. Yet most of those forage fish stocks are at relatively high levels of abundance.

### 9.0 SOURCES OF UNCERTAINTY FOR ASSESSMENT

1) The species composition and size structure of landings are generally unknown.
2) The true level of discards and the discard mortality rate are unknown.
3) A lack of information on the stock structure of the species in the skate complex has increased the uncertainty of conclusions about historical trends in abundance, and recommendations of appropriate biological reference points.
4) Life history data are from localized areas for barndoor, thorny, and clearnose and incomplete or totally lacking for two other species.
5) Mortality estimates based on equilibrium assumptions which are only partially met for these stocks. A preferable approach for future assessments would be an age-based method for determining mortality rates and estimates of longevity. This will require several years of future adequate length and age sampling, both from the commercial and research survey catches.
6) The proposed SFA biomass reference points are based on selected time periods of survey indices, but it is unknown how these relate to true estimates of $\mathrm{B}_{\text {MSY }}$.

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## SKATE TABLES

Table B1.1. Total commercial landings of skate (mt) in NAFO subareas 5 and 6 by country from 1960-2005. U.S. landings are from NAFO database from 1964-1988, weighout from 1989-2005.

| US |  | USSR | Others | Total |
| :---: | :---: | :---: | :---: | :---: |
| 1964 | 4081 | 0 | 2 | 4083 |
| 1965 | 2343 | 0 | 20 | 2363 |
| 1966 | 2738 | 0 | 106 | 2844 |
| 1967 | 2715 | 2121 | 62 | 4898 |
| 1968 | 2417 | 3974 | 92 | 6483 |
| 1969 | 3045 | 6410 | 7 | 9462 |
| 1970 | 1583 | 2544 | 1 | 4128 |
| 1971 | 900 | 5000 | 5 | 5905 |
| 1972 | 866 | 7957 | 0 | 8823 |
| 1973 | 1191 | 6754 | 18 | 7963 |
| 1974 | 2026 | 1623 | 2 | 3651 |
| 1975 | 752 | 3216 | 0 | 3968 |
| 1976 | 754 | 412 | 46 | 1212 |
| 1977 | 1143 | 240 | 35 | 1418 |
| 1978 | 1130 | 216 | 7 | 1353 |
| 1979 | 1280 | 79 | 1 | 1360 |
| 1980 | 1577 | 0 | 4 | 1581 |
| 1981 | 838 | 0 | 9 | 847 |
| 1982 | 878 | 0 | 0 | 878 |
| 1983 | 3603 | 0 | 0 | 3603 |
| 1984 | 4157 | 0 | 0 | 4157 |
| 1985 | 3984 | 0 | 0 | 3984 |
| 1986 | 4159 | 0 | 94 | 4253 |
| 1987 | 5078 | 0 | 0 | 5078 |
| 1988 | 7255 | 0 | 9 | 7264 |
| 1989 | 6707 | 0 | 0 | 6707 |
| 1990 | 11403 | 0 | 0 | 11403 |
| 1991 | 11332 | 0 | 0 | 11332 |
| 1992 | 12525 | 0 | 0 | 12525 |
| 1993 | 12904 | 0 | 0 | 12904 |
| 1994 | 8783 | 0 | 0 | 8783 |
| 1995 | 7217 | 0 | 0 | 7217 |
| 1996 | 14213 | 0 | 0 | 14213 |
| 1997 | 10945 | 0 | 0 | 10945 |
| 1998 | 13829 | 0 | 0 | 13829 |
| 1999 | 11684 | 0 | 0 | 11684 |
| 2000 | 13360 |  |  | 13360 |
| 2001 | 13120 |  |  | 13120 |
| 2002 | 13004 |  |  | 13004 |
| 2003 | 15005 |  |  | 15005 |
| 2004 | 16073 |  |  | 16073 |
| 2005 | 13885 |  |  | 13885 |

Table B1.2. U.S. commerical landings (mt, live wt) of skates (all species) by month from 1964-2005.

| Month |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| year | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | Total |
| 1964 | 4050.3 | 2.0 | 3.9 | 3.6 | 3.1 | 2.0 | 1.6 | 0.9 | 1.3 | 1.6 | 2.0 | 2.1 | 6.4 | 4081.0 |
| 1965 | 2304.4 | 5.4 | 7.2 | 7.5 | 4.3 | 2.4 | 0.4 | 0.6 | 1.2 | 0.6 | 2.3 | 2.6 | 4.2 | 2343.0 |
| 1966 | 2707.1 | 6.4 | 7.3 | 6.0 | 1.0 | 0.9 | 0.2 | 0.1 | 0.7 | 1.7 | 1.4 | 2.4 | 2.9 | 2738.0 |
| 1967 | 2643.3 | 15.1 | 7.3 | 18.1 | 7.7 | 3.0 | 1.6 | 0.6 | 0.4 | 1.8 | 6.1 | 2.9 | 7.1 | 2715.0 |
| 1968 | 2381.3 | 10.3 | 1.9 | 5.3 | 1.3 | 1.5 | 1.3 | 1.5 | 2.6 | 3.0 | 2.8 | 2.5 | 1.7 | 2417.0 |
| 1969 | 2993.4 | 4.1 | 6.2 | 5.7 | 6.2 | 2.5 | 2.3 | 3.1 | 3.2 | 3.0 | 5.0 | 5.7 | 4.6 | 3045.0 |
| 1970 | 1513.4 | 6.1 | 8.6 | 13.9 | 7.0 | 4.1 | 3.4 | 5.6 | 5.3 | 8.3 | 4.1 | 2.1 | 1.1 | 1583.0 |
| 1971 | 836.7 | 4.9 | 6.2 | 8.5 | 7.3 | 7.7 | 2.7 | 3.0 | 2.8 | 3.5 | 8.2 | 3.9 | 4.7 | 900.0 |
| 1972 | 780.1 | 7.2 | 6.9 | 12.1 | 12.3 | 9.1 | 4.9 | 5.7 | 7.8 | 4.3 | 4.2 | 5.9 | 5.5 | 866.0 |
| 1973 | 1104.1 | 8.3 | 3.9 | 10.4 | 12.4 | 7.1 | 6.7 | 7.1 | 7.0 | 8.1 | 7.1 | 4.7 | 4.1 | 1191.0 |
| 1974 | 1945.9 | 5.7 | 4.9 | 5.6 | 12.3 | 8.0 | 4.6 | 4.4 | 12.3 | 6.7 | 5.2 | 2.6 | 7.8 | 2026.0 |
| 1975 | 637.9 | 7.3 | 10.1 | 16.6 | 16.2 | 13.0 | 7.3 | 6.7 | 7.6 | 9.8 | 5.6 | 6.9 | 6.9 | 752.0 |
| 1976 | 641.8 | 8.4 | 12.5 | 19.2 | 22.4 | 9.6 | 4.3 | 8.1 | 4.7 | 6.9 | 3.1 | 6.3 | 6.8 | 754.0 |
| 1977 | 994.7 | 15.4 | 19.7 | 27.9 | 20.0 | 9.0 | 8.9 | 6.8 | 11.0 | 7.0 | 8.8 | 9.3 | 4.5 | 1143.0 |
| 1978 | 827.4 | 19.3 | 24.7 | 11.7 | 29.8 | 30.5 | 46.4 | 33.9 | 26.2 | 23.2 | 20.9 | 19.3 | 16.7 | 1130.0 |
| 1979 | 787.4 | 24.8 | 24.8 | 46.5 | 62.6 | 50.4 | 28.1 | 29.4 | 55.5 | 38.8 | 42.1 | 52.9 | 36.5 | 1279.6 |
| 1980 | 961.1 | 61.5 | 112.6 | 121.1 | 82.8 | 63.9 | 27.3 | 26.4 | 24.4 | 22.8 | 27.4 | 20.5 | 25.4 | 1577.2 |
| 1981 | 509.9 | 33.9 | 30.8 | 54.4 | 31.1 | 26.7 | 25.3 | 15.1 | 24.5 | 23.1 | 12.3 | 19.2 | 31.9 | 838.4 |
| 1982 | 449.5 | 30.4 | 23.3 | 54.0 | 47.5 | 58.2 | 18.9 | 25.3 | 35.1 | 32.3 | 34.4 | 31.3 | 38.2 | 878.1 |
| 1983 | 2720.3 | 84.1 | 95.9 | 134.0 | 95.4 | 102.3 | 76.3 | 44.1 | 66.1 | 53.3 | 37.0 | 56.6 | 37.5 | 3603.0 |
| 1984 | 3325.7 | 99.4 | 127.3 | 134.9 | 108.6 | 84.0 | 36.7 | 30.9 | 29.0 | 25.9 | 37.0 | 54.2 | 63.0 | 4156.5 |
| 1985 | 3220.7 | 85.4 | 85.5 | 150.6 | 142.7 | 31.6 | 29.9 | 33.2 | 29.9 | 28.8 | 37.7 | 59.3 | 48.6 | 3984.1 |
| 1986 | 3173.4 | 98.6 | 89.7 | 149.7 | 147.8 | 91.8 | 36.4 | 33.7 | 49.0 | 28.2 | 72.6 | 86.3 | 102.5 | 4159.5 |
| 1987 | 3638.7 | 83.8 | 114.3 | 207.7 | 227.0 | 245.3 | 106.2 | 40.3 | 53.0 | 33.8 | 87.6 | 101.5 | 139.1 | 5078.4 |
| 1988 | 5141.7 | 281.6 | 338.2 | 378.7 | 284.0 | 150.3 | 74.5 | 154.5 | 137.9 | 75.0 | 54.1 | 66.2 | 118.8 | 7255.5 |
| 1989 | 4157.8 | 240.1 | 150.3 | 227.1 | 454.3 | 292.6 | 102.6 | 142.2 | 272.3 | 221.9 | 174.8 | 173.0 | 98.4 | 6707.3 |
| 1990 | 4252.9 | 136.6 | 182.0 | 424.8 | 834.4 | 948.5 | 1174.9 | 763.8 | 818.7 | 624.4 | 265.9 | 542.3 | 433.4 | 11402.5 |
| 1991 | 4255.9 | 464.0 | 423.8 | 460.9 | 606.0 | 419.8 | 370.4 | 658.1 | 925.7 | 515.5 | 565.5 | 958.9 | 708.0 | 11332.3 |
| 1992 | 4782.2 | 517.3 | 457.7 | 510.1 | 567.1 | 564.3 | 816.2 | 764.4 | 718.2 | 862.3 | 639.1 | 771.1 | 555.4 | 12525.3 |
| 1993 | 4860.4 | 335.1 | 265.6 | 471.2 | 741.7 | 875.2 | 823.2 | 1005.6 | 859.1 | 712.4 | 535.5 | 864.0 | 555.0 | 12904.0 |
| 1994 | 175.5 | 338.2 | 309.8 | 291.7 | 501.5 | 855.1 | 1238.4 | 780.9 | 1263.7 | 960.6 | 937.7 | 787.3 | 342.9 | 8783.3 |
| 1995 | 1.0 | 183.7 | 285.7 | 413.6 | 515.5 | 752.0 | 915.7 | 768.4 | 752.2 | 557.7 | 724.8 | 897.2 | 449.7 | 7217.1 |
| 1996 | 2.3 | 224.6 | 229.3 | 206.5 | 360.1 | 1012.0 | 1389.7 | 1539.8 | 1577.6 | 1720.4 | 2440.4 | 2411.8 | 1098.4 | 14212.8 |
| 1997 |  | 530.8 | 469.9 | 597.5 | 395.5 | 969.4 | 1127.6 | 1181.8 | 1189.6 | 1062.3 | 1084.2 | 1305.2 | 1031.1 | 10944.8 |
| 1998 |  | 518.9 | 589.8 | 625.4 | 814.9 | 1403.4 | 1702.2 | 1643.9 | 1512.7 | 1551.5 | 1224.9 | 1277.1 | 964.5 | 13829.2 |
| 1999 |  | 511.2 | 401.0 | 591.8 | 678.6 | 1295.5 | 1436.2 | 1039.3 | 1137.7 | 1388.8 | 1055.8 | 1250.0 | 898.1 | 11683.9 |
| 2000 |  | 668.1 | 615.2 | 1024.2 | 826.2 | 1187.7 | 1594.2 | 1188.5 | 1534.6 | 1270.1 | 946.4 | 1583.6 | 921.1 | 13359.9 |
| 2001 |  | 802.4 | 588.6 | 956.2 | 967.3 | 984.0 | 1058.2 | 1150.5 | 1465.1 | 1197.3 | 1115.1 | 1692.1 | 1143.7 | 13120.5 |
| 2002 |  | 742.3 | 730.7 | 783.2 | 1093.9 | 773.5 | 1372.6 | 998.7 | 1488.6 | 1247.8 | 1352.1 | 1264.4 | 1156.3 | 13004.0 |
| 2003 |  | 548.3 | 447.6 | 857.4 | 1043.7 | 1006.6 | 1183.0 | 1632.9 | 1867.9 | 1889.1 | 1993.3 | 1563.3 | 971.9 | 15004.9 |
| 2004 |  | 538.3 | 1279.4 | 1305.0 | 1391.0 | 1155.1 | 1456.9 | 2008.7 | 1557.9 | 1573.6 | 1115.7 | 1541.6 | 1150.2 | 16073.4 |
| 2005 |  | 869.9 | 1201.7 | 1070.1 | 1187.4 | 1098.5 | 1289.7 | 1650.4 | 1585.9 | 1320.7 | 824.4 | 987.2 | 798.7 | 13884.6 |

Table B1.3. U.S. Commercial landings ( mt , live wt) of skates (all species) by state from 1964-2005. Data are from weighout database.


Table B1.4. U.S. Commercial landings (mt, live wt) of skates (all species) by gear type fromo 1964-2005. Landings are from weighout database.

| gear |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| year | longline | otter trawl | other | sink gillnet | Total |
| 1964 | 0.1 | 30.5 |  | 0.0 | 30.7 |
| 1965 | 0.3 | 38.2 |  | 0.0 | 38.6 |
| 1966 |  | 30.9 |  |  | 30.9 |
| 1967 |  | 71.7 |  |  | 71.7 |
| 1968 |  | 35.7 |  |  | 35.7 |
| 1969 |  | 51.5 |  | 0.0 | 51.6 |
| 1970 | 0.6 | 68.8 | 0.0 | 0.2 | 69.6 |
| 1971 | 1.1 | 62.0 |  | 0.1 | 63.3 |
| 1972 | 3.7 | 80.8 | 0.1 | 1.3 | 85.9 |
| 1973 | 7.0 | 77.9 | 1.9 | 0.2 | 86.9 |
| 1974 | 10.5 | 64.3 | 0.2 | 5.1 | 80.1 |
| 1975 | 11.7 | 101.4 | 0.1 | 0.8 | 114.1 |
| 1976 | 16.2 | 93.3 | 0.2 | 2.5 | 112.2 |
| 1977 | 13.4 | 126.8 | 0.9 | 7.2 | 148.3 |
| 1978 | 4.4 | 290.0 | 3.2 | 5.0 | 302.6 |
| 1979 | 18.4 | 456.0 | 5.8 | 12.0 | 492.2 |
| 1980 | 16.5 | 577.9 | 6.0 | 15.6 | 616.1 |
| 1981 | 5.1 | 311.7 | 1.2 | 10.4 | 328.4 |
| 1982 | 2.0 | 408.4 | 7.4 | 10.8 | 428.7 |
| 1983 | 3.4 | 846.2 | 22.5 | 10.6 | 882.7 |
| 1984 | 5.0 | 796.5 | 19.1 | 10.3 | 830.8 |
| 1985 | 3.7 | 721.5 | 17.8 | 20.3 | 763.3 |
| 1986 | 6.6 | 954.4 | 14.2 | 10.9 | 986.1 |
| 1987 | 22.4 | 1384.4 | 16.1 | 16.8 | 1439.7 |
| 1988 | 5.7 | 2070.7 | 22.2 | 15.2 | 2113.7 |
| 1989 | 30.6 | 6636.1 | 27.3 | 13.4 | 6707.3 |
| 1990 | 3.8 | 11339.6 | 47.7 | 11.5 | 11402.5 |
| 1991 | 24.3 | 11169.9 | 77.0 | 61.1 | 11332.3 |
| 1992 | 21.9 | 12242.5 | 35.1 | 225.8 | 12525.3 |
| 1993 | 63.4 | 11913.6 | 204.6 | 722.3 | 12904.0 |
| 1994 | 197.2 | 7194.4 | 357.4 | 1034.3 | 8783.3 |
| 1995 | 97.1 | 5777.2 | 400.7 | 942.1 | 7217.1 |
| 1996 | 51.8 | 12944.3 | 134.4 | 1082.3 | 14212.8 |
| 1997 | 47.7 | 8822.8 | 471.6 | 1602.8 | 10944.8 |
| 1998 | 53.2 | 11724.8 | 576.4 | 1474.8 | 13829.2 |
| 1999 | 48.5 | 10059.3 | 144.9 | 1431.3 | 11684.0 |
| 2000 | 34.9 | 11464.0 | 72.0 | 1789.0 | 13360.0 |
| 2001 | 12.0 | 10835.0 | 27.7 | 2245.9 | 13120.5 |
| 2002 | 32.8 | 9667.7 | 31.0 | 3272.4 | 13004.0 |
| 2003 | 97.1 | 10254.3 | 43.0 | 4610.6 | 15004.9 |
| 2004 | 136.9 | 10694.3 | 2217.0 | 3025.3 | 16073.4 |
| 2005 | 342.7 | 7744.3 | 2532.9 | 3264.7 | 13884.6 |

Table B1.5. U.S. landings (mt, live wt) of skates by species and markey category from 1964-2005. Landings are from weighout database.

| Species and Market Category |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| YEAR | Uncl. | Uncl. | Winter |  | Winter | Little | Little |  | Barndoor | Barndoor | Thorny |  | Thorny | Smooth |  | Smooth | Clearnose | Clearnose | Rose | Rose | Total |  |
|  | Whole | Wings | Whole |  | Wings | Whole | Wings |  | Whole | Wings | Whole |  | Wings | Whole |  | Wings | Whole | Wings | Whole | Wings | Whole | Wings |
| 1964 | 30.7 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 30.7 | 0.0 |
| 1965 | 38.6 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 38.6 | 0.0 |
| 1966 | 30.9 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 30.9 | 0.0 |
| 1967 | 71.7 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 71.7 | 0.0 |
| 1968 | 35.7 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 35.7 | 0.0 |
| 1969 | 51.6 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 51.6 | 0.0 |
| 1970 | 69.6 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 69.6 | 0.0 |
| 1971 | 63.3 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 63.3 | 0.0 |
| 1972 | 85.9 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 85.9 | 0.0 |
| 1973 | 86.9 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 86.9 | 0.0 |
| 1974 | 80.1 |  |  | 0.0 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 80.1 | 0.0 |
| 1975 | 114.1 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 114.1 | 0.0 |
| 1976 | 112.2 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 112.2 | 0.0 |
| 1977 | 148.3 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 148.3 | 0.0 |
| 1978 | 302.6 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 302.6 | 0.0 |
| 1979 | 492.2 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 492.2 | 0.0 |
| 1980 | 616.1 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 616.1 | 0.0 |
| 1981 | 328.4 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 328.4 | 0.0 |
| 1982 | 277.2 | 151.4 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 277.2 | 151.4 |
| 1983 | 169.6 | 713.0 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 169.6 | 713.0 |
| 1984 | 68.1 | 762.8 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 68.1 | 762.8 |
| 1985 | 68.3 | 695.0 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 68.3 | 695.0 |
| 1986 | 262.6 | 723.5 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 262.6 | 723.5 |
| 1987 | 87.5 | 1352.2 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 87.5 | 1352.2 |
| 1988 | 74.2 | 2039.6 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 74.2 | 2039.6 |
| 1989 | 4163.1 | 2544.2 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 4163.1 | 2544.2 |
| 1990 | 5002.9 | 6399.6 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 5002.9 | 6399.6 |
| 1991 | 5069.2 | 6262.5 |  |  |  | 0.6 |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 5069.7 | 6262.5 |
| 1992 | 5860.5 | 6664.7 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 5860.5 | 6664.7 |
| 1993 | 5526.6 | 7377.5 |  | 0.0 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 5526.6 | 7377.5 |
| 1994 | 703.4 | 8079.9 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 703.4 | 8079.9 |
| 1995 | 3095.1 | 3985.5 |  |  |  | 136.6 |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 3231.7 | 3985.5 |
| 1996 | 3981.5 | 10230.8 |  | 0.4 |  | 0.2 |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 3982.0 | 10230.8 |
| 1997 | 5369.1 | 5575.6 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 5369.1 | 5575.6 |
| 1998 | 5391.8 | 8437.4 |  |  |  | 0.0 |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 5391.8 | 8437.4 |
| 1999 | 5026.7 | 6655.2 |  |  |  |  |  |  | 2.1 |  |  |  |  |  |  |  |  |  |  |  | 5028.7 | 6655.2 |
| 2000 | 3633.4 | 8690.5 |  | 0.0 |  | 1036.0 |  | 0.1 | 0.0 |  |  |  |  |  |  |  |  |  |  |  | 4669.4 | 8690.6 |
| 2001 | 4399.5 | 8718.6 |  | 2.2 |  | 0.0 |  | 0.1 |  |  |  |  |  |  |  | 0.1 |  |  |  |  | 4401.7 | 8718.8 |
| 2002 | 4396.9 | 8606.9 |  |  |  |  |  | 0.1 |  | 0.1 |  |  |  |  |  |  |  |  |  |  | 4396.9 | 8607.1 |
| 2003 | 4327.8 | 10650.0 |  | 0.8 | 26.0 | 0.2 |  |  |  |  |  |  |  |  | 0.1 |  |  |  |  |  | 4328.8 | 10676.0 |
| 2004 | 998.5 | 8451.6 |  | 2.8 | 2697.5 | 2867.4 |  | 8.6 | 0.3 | 0.1 |  | 0.0 | 95.6 |  | 1.0 | 927.2 | 3.5 | 16.6 |  | 2.7 | 3873.6 | 12199.8 |
| 2005 | 142.2 | 6710.1 |  | 59.3 | 3301.4 | 3465.3 |  | 15.6 | 0.3 | 5.4 |  | 1.5 | 126.2 |  | 0.6 | 1.0 | 32.5 |  | 16.6 | 5.9 | 3718.3 | 10165.6 |

Table B1.6. U.S. landings (my, live wt) of skates by state, species and market category for 2004-2005.

|  | Species and Market Category |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | Total |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Uncl. Whole | Uncl. Wings | Winter | $\begin{aligned} & \hline \text { Winter } \\ & \hline \text { Wings } \\ & \hline \end{aligned}$ | Little Whole | Little Wings | Barndoor <br> Whole | Barndoor Wings | Thorny Whole | $\begin{aligned} & \hline \text { Thorny } \\ & \hline \text { Wings } \\ & \hline \end{aligned}$ | Smooth | Smooth | Clearnose Whole | Clearnose Wings | $\begin{array}{\|l\|} \hline \text { Rosette } \\ \hline \text { Whole } \\ \hline \end{array}$ | Rosette <br> Wings |  |  |
| YEAR | State |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | Whole | Wings |
| 2004 | CT | 369.9 | 71.8 |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 369.9 | 71.8 |
|  | DE | 0.0 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0.0 | 0.0 |
|  | ME | 0.0 | 12.2 |  | 1.2 |  |  |  |  |  |  |  |  |  |  |  |  | 0.0 | 13.3 |
|  | MD | 1.0 | 2.4 |  | 2.7 | 0.1 |  |  |  |  |  |  |  |  |  |  |  | 1.1 | 5.1 |
|  | MA | 17.7 | 6482.2 | 0.2 | 2467.9 | 97.5 |  |  |  | 0.0 | 83.4 | 0.1 | 926.8 |  |  |  | 0.1 | 115.5 | 9960.4 |
|  | NH |  | 5.1 |  | 5.4 |  |  |  |  |  | 0.1 |  |  |  |  |  |  | 0.0 | 10.6 |
|  | NJ | 1.5 | 131.2 | 0.3 | 135.5 | 103.0 | 2.7 |  |  |  | 0.1 |  |  |  |  |  |  | 104.8 | 269.5 |
|  | NY | 23.3 | 183.6 | 1.2 | 0.6 | 0.7 | 0.1 |  |  |  | 12.0 | 1.0 | 0.3 |  |  |  |  | 26.1 | 196.7 |
|  | NC |  | 0.5 |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0.0 | 0.5 |
|  | RI | 584.1 | 1538.6 | 1.2 | 84.2 | 2666.1 | 5.8 |  |  |  |  |  |  |  |  |  | 2.6 | 3251.3 | 1631.3 |
|  | VA | 1.1 | 24.1 |  |  |  |  | 0.3 | 0.1 |  |  |  |  | 3.5 | 16.6 |  |  | 4.9 | 40.8 |
|  | Total | 998.5 | 8451.6 | 2.8 | 2697.5 | 2867.4 | 8.6 | 0.3 | 0.1 | 0.0 | 95.6 | 1.0 | 927.2 | 3.5 | 16.6 |  | 2.7 | 3873.6 | 12199.8 |
| 2005 | CT | 0.1 | 47.5 |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0.1 | 47.5 |
|  | ME |  | 10.2 |  | 0.5 |  | 0.2 |  |  |  |  |  |  |  |  |  |  | 0.0 | 10.8 |
|  | MD | 2.3 | 6.1 |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 2.3 | 6.1 |
|  | MA | 60.2 | 5699.4 | 21.7 | 3071.7 | 21.1 |  |  | 1.3 | 1.5 | 111.6 | 0.0 | 0.7 |  |  |  |  | 104.5 | 8884.8 |
|  | NH | 0.0 | 9.4 |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0.0 | 9.4 |
|  | NJ | 0.4 | 120.1 | 24.4 | 110.7 | 45.0 | 1.1 | 0.1 |  |  |  | 0.4 |  | 32.5 |  |  |  | 102.9 | 231.9 |
|  | NY | 12.3 | 96.6 | 0.4 | 1.5 | 12.7 | 0.2 | 0.2 | 4.1 |  | 12.6 | 0.0 | 0.3 |  |  | 16.6 |  | 42.2 | 115.3 |
|  | NC |  | 0.5 |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0.0 | 0.5 |
|  | RI | 66.6 | 690.9 | 12.8 | 116.9 | 3386.5 | 14.1 |  |  |  | 2.0 | 0.2 | 0.1 |  |  |  | 5.9 | 3466.1 | 829.9 |
|  | VA | 0.3 | 29.3 |  |  |  |  |  |  |  |  |  |  |  |  |  |  | 0.3 | 29.3 |
|  | Total | 142.2 | 6710.1 |  | 3301.4 | 3465.3 | 15.6 |  |  |  | 126.2 | 0.6 | 1.0 | 32.5 |  | 16.6 |  | 3718.3 | 10165.6 |

Table B1.7. Discards (mt) of skates (all species) by gear type, empty cells not filled in. Dashes indicate no sampling.

| year | Line Trawl | Longline | Otter | Scallop Traw | Pair Traw | Shrimp Trawl | Sink Gili | Scallop Dredge | Mid-Water | Lobster Pot | Blue Crab Pots | Scotish Seine | Grand Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1989 |  |  | 56,622 |  |  | 85 | 127 |  |  | - |  |  | 56,834 |
| 1990 | - | 0 | 77,805 | - | - | 258 | 624 | - | - | - | - |  | 78,687 |
| 1991 | 865 | 0 | 45,775 | - | - | 283 | 289 | 6,391 | 0 | 0 |  |  | 53,602 |
| 1992 | 1,438 | 30 | 45,334 | - | 0 | 245 | 452 | 39,705 | 0 | 0 |  |  | 87,204 |
| 1993 | 45 | 0 | 28,388 | - | 188 | 36 | 375 | 22,866 | 0 | 31 |  |  | 51,929 |
| 1994 | 0 | 0 | 32,458 | - |  | 13 | 856 | 10,525 | 0 | 0 |  |  | 43,852 |
| 1995 | 0 | 0 | 37,564 | - | - | 9 | 767 | 18,074 | 0 | 0 |  |  | 56,414 |
| 1996 | - | - | 32,693 | - | - | 35 | 1,090 | 18,321 |  | - 0 |  |  | 52,139 |
| 1997 | - | 0 | 10,032 | - | - | 1 | 537 | 15,606 |  | - 0 |  |  | 26,176 |
| 1998 | - | - | 14,051 | - | - | - | 593 | 14,626 |  | - 0 |  |  | 29,270 |
| 1999 | - | - | 16,827 | - | - | - | 1,057 | 15,901 | 0 | 5 |  |  | 33,789 |
| 2000 | - | - | 29,121 | - | - | - | 1,130 | 12,099 | 0 | 29 |  |  | 42,379 |
| 2001 | - | - | 42,461 | 365 | - | 0 | 609 | 6,070 | 0 | - | - | 18 | 49,523 |
| 2002 | 39 | - | 43,740 | 268 | - |  | 2,015 | 15,651 | 0 | - | 12,375 | 7 | 74,095 |
| 2003 | 15 | - | 32,370 | - | - | 11 | 946 | 14,977 | 0 | 0 |  | 1 | 48,320 |
| 2004 | 29 |  | 27,341 | 161 | - | 0 | 803 | 4,970 | 1 |  |  | 0 | 33,306 |
| 2005 | 825 | - | 13,824 | 35 | - | 2 | 2,180 | 2,794 | 0 | 0 | - |  | 19,660 |

Table B1.8. Discards of skates (all species) by year, quarter, region in the otter trawl fishery.


Table B1.9. Discards of skates (all species) by year, quarter, region in the sink gill net fishery.


Table B1.10. Discards of skates (all species) by year, quarter, region in the scallop dredge fishery.

|  |  | Quarter | 1 |  |  |  | 2 |  | 3 |  |  |  | 4 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| year |  | Region | MA | NE | Total | MA | NE | Total | A | NE | Total | MA | E | otal |
| 1989 |  | ntrips <br> dkratio <br> mt kept <br> mt discard | $\begin{array}{r} 10086.6 \\ 0.0 \end{array}$ | $\begin{array}{r} 23291.0 \\ 0.0 \\ \hline \end{array}$ | $\begin{array}{r} 33377.5 \\ 0.0 \end{array}$ | $\begin{array}{r} 15880.9 \\ 0.0 \end{array}$ | $\begin{array}{r} 28652.0 \\ 0.0 \end{array}$ | $\begin{array}{r} 44532.8 \\ 0.0 \end{array}$ | $\begin{array}{r} 10428.4 \\ 0.0 \end{array}$ | $\begin{array}{r} 25176.9 \\ 0.0 \end{array}$ | $\begin{array}{r} 35605.4 \\ 0.0 \end{array}$ | $\begin{array}{r} 5278.9 \\ 0.0 \\ \hline \end{array}$ | $\begin{array}{r} 18667.2 \\ 0.0 \end{array}$ | $\begin{array}{r} 23946.0 \\ 0.0 \end{array}$ |
| 1990 |  | ntrips <br> dkratio <br> mt kept <br> mt discard | $\begin{array}{r} 10987.0 \\ 0.0 \\ \hline \end{array}$ | $\begin{array}{r} 17618.5 \\ 0.0 \\ \hline \end{array}$ | $\begin{array}{r} 28605.6 \\ 0.0 \\ \hline \end{array}$ | $\begin{array}{r} 14895.0 \\ 0.0 \\ \hline \end{array}$ | $\begin{array}{r} 30679.0 \\ 0.0 \\ \hline \end{array}$ | $\begin{array}{r} 45574.0 \\ 0.0 \\ \hline \end{array}$ | $\begin{array}{r} 14342.6 \\ 0.0 \\ \hline \end{array}$ | $\begin{array}{r} 30581.7 \\ 0.0 \\ \hline \end{array}$ | $\begin{array}{r} 44924.2 \\ 0.0 \\ \hline \end{array}$ | $\begin{array}{r} 7677.8 \\ 0.0 \\ \hline \end{array}$ | $\begin{array}{r} 19732.3 \\ 0.0 \\ \hline \end{array}$ | $\begin{array}{r} 27410.1 \\ 0.0 \\ \hline \end{array}$ |
| 1991 |  | ntrips <br> dkratio <br> mt kept <br> mt discard | $\begin{array}{r} 10896.2 \\ 0.0 \end{array}$ | $\begin{array}{r} 23586.6 \\ 0.0 \end{array}$ | $\begin{array}{r} 34482.8 \\ 0.0 \end{array}$ | $\begin{array}{r} 18918.4 \\ 0.0 \end{array}$ | $\begin{array}{r} 31037.2 \\ 0.0 \end{array}$ | $\begin{array}{r} 49955.5 \\ 0.0 \end{array}$ | $\begin{array}{r} 10741.8 \\ 0.0 \end{array}$ | 23977.9 0.0 | 34719.7 0.0 | $\begin{array}{r} 1 \\ 0.56 \\ 6046.7 \\ 3366.0 \\ \hline \end{array}$ | $\begin{array}{r} 1 \\ 0.18 \\ 16561.7 \\ 3024.7 \\ \hline \end{array}$ | $\begin{array}{r} 2 \\ 0.74 \\ 22608.4 \\ 6390.7 \\ \hline \end{array}$ |
| 1992 |  | ntrips | 1 | 2 | 3 | 1 | 3 | 4 | 1 | 2 | 3 | 1 | 4 | 5 |
|  |  | dkratio | 0.20 | 0.36 | 0.55 | 0.11 | 0.24 | 0.35 | 0.12 | 0.59 | 0.71 | 0.26 | 0.40 | 0.66 |
|  |  | mt kept | 7389.5 | 17974.8 | 25364.3 | 12121.3 | 25380.6 | 37501.9 | 11000.5 | 24564.0 | 35564.6 | 5325.4 | 18270.0 | 23595.4 |
|  |  | mt discard | 1452.4 | 6390.4 | 7842.8 | 1274.9 | 6192.9 | 7467.8 | 1322.4 | 14390.4 | 15712.7 | 1410.8 | 7270.6 | 8681.4 |
| 1993 |  | ntrips | 3 | 4 | 7 | 3 | 4 | 7 | 1 | 2 | 3 | 1 | 3 | 4 |
|  |  | dkratio | 0.45 | 0.20 | 0.65 | 0.52 | 0.14 | 0.66 | 0.53 | 0.18 | 0.71 | 0.76 | 0.52 | 1.28 |
|  |  | mt kept | 4536.8 | 13875.1 | 18412.0 | 6136.5 | 13124.9 | 19261.4 | 5650.6 | 11626.6 | 17277.2 | 3277.8 | 10498.7 | 13776.5 |
|  |  | mt discard | 2030.6 | 2758.9 | 4789.6 | 3188.3 | 1795.5 | 4983.8 | 2989.8 | 2145.0 | 5134.8 | 2506.4 | 5451.2 | 7957.7 |
| 1994 |  | ntrips | 4 | 3 | 7 | 3 | 1 | 4 |  | 4 | 4 | 3 | 5 | 8 |
|  |  | dkratio | 0.38 | 0.20 | 0.57 | 0.05 | 0.17 | 0.22 |  | 0.08 | 0.08 | 0.50 | 0.21 | 0.71 |
|  |  | mt kept | 5189.9 | 7542.7 | 12732.6 | 10500.5 | 9248.8 | 19749.4 | 9023.3 | 9236.0 | 18259.3 | 4719.4 | 8918.3 | 13637.7 |
|  |  | mt discard | 1958.8 | 1472.3 | 3431.1 | 551.3 | 1541.6 | 2092.9 | 0.0 | 765.9 | 765.9 | 2356.8 | 1878.8 | 4235.6 |
| 1995 |  | ntrips | 6 | 3 | 9 | 2 | 3 | 5 | 3 | 2 | 5 |  | 5 | 5 |
|  |  | dkratio | 0.26 | 0.32 | 0.59 | 0.39 | 0.04 | 0.44 | 0.07 | 0.26 | 0.33 |  | 0.83 | 0.83 |
|  |  | mt kept | 5765.1 | 7520.0 | 13285.1 | 11081.4 | 13823.0 | 24904.4 | 7007.7 | 10248.7 | 17256.4 | 2340.3 | 7278.6 | 9618.9 |
|  |  | mt discard | 1522.5 | 2424.8 | 3947.3 | 4348.7 | 605.9 | 4954.5 | 520.6 | 2619.9 | 3140.5 | 0.0 | 6031.6 | 6031.6 |
| 1996 |  | ntrips | 6 | 7 | 13 | 4 | 5 | 9 | 3 | 4 | 7 | 4 | 5 | 9 |
|  |  | dkratio | 0.24 | 0.13 | 0.38 | 0.46 | 0.10 | 0.56 | 0.23 | 0.14 | 0.38 | 1.11 | 0.41 | 1.52 |
|  |  | mt kept | 3368.3 | 5907.8 | 9276.1 | 10880.0 | 13675.2 | 24555.2 | 6904.9 | 12142.7 | 19047.6 | 2663.1 | 9855.3 | 12518.4 |
|  |  | mt discard | 823.5 | 782.0 | 1605.5 | 5022.2 | 1378.6 | 6400.8 | 1606.2 | 1738.1 | 3344.3 | 2959.9 | 4010.7 | 6970.6 |
| 1997 |  | ntrips | 6 | 6 | 12 | 5 | 2 | 7 | 4 | 3 | 7 | 1 | 2 | 3 |
|  |  | dkratio | 0.55 | 0.26 | 0.81 | 0.55 | 0.14 | 0.69 | 0.33 | 0.36 | 0.69 | 0.10 | 0.10 | 0.20 |
|  |  | mt kept | 3375.8 | 7265.0 | 10640.9 | 7523.7 | 11622.1 | 19145.8 | 5540.9 | 9175.7 | 14716.6 | 2206.1 | 7496.9 | 9703.0 |
|  |  | mt discard | 1840.2 | 1890.2 | 3730.5 | 4153.4 | 1620.5 | 5773.9 | 1803.5 | 3314.1 | 5117.6 | 228.2 | 755.8 | 984.0 |
| 1998 |  | ntrips | 1 |  | 1 | 6 | 2 | 8 | 3 | 2 | 5 | 6 | 6 | 12 |
|  |  | dkratio | 0.10 |  | 0.10 | 0.38 | 0.13 | 0.52 | 0.47 | 0.64 | 1.11 | 0.60 | 0.27 | 0.87 |
|  |  | mt kept | 3212.1 | 6498.3 | 9710.4 | 6420.8 | 9324.1 | 15744.9 | 4168.5 | 7997.0 | 12165.5 | 2778.4 | 6975.2 | 9753.6 |
|  |  | mt discard | 310.1 | 0.0 | 310.1 | 2455.6 | 1236.1 | 3691.7 | 1961.9 | 5089.6 | 7051.5 | 1656.4 | 1915.9 | 3572.2 |
| 1999 |  | ntrips |  |  |  | 1 | 2 | 3 | 4 | 1 | 5 | 2 | 5 | 7 |
|  |  | dkratio |  |  |  | 0.29 | 0.10 | 0.38 | 0.56 | 0.33 | 0.89 | 0.04 | 0.09 | 0.14 |
|  |  | mt kept | 3981.4 | 7393.9 | 11375.2 | 11211.7 | 16989.1 | 28200.8 | 6866.1 | 16967.2 | 23833.3 | 2229.0 | 15535.5 | 17764.5 |
|  |  | mt discard | 0.0 | 0.0 | 0.0 | 3198.7 | 1638.8 | 4837.5 | 3833.0 | 5673.7 | 9506.7 | 92.6 | 1464.1 | 1556.6 |
| 2000 |  | ntrips | 4 | 3 | 7 | 6 | 25 | 31 | 11 | 107 | 118 | 7 | 93 | 100 |
|  |  | dkratio | 0.05 | 0.22 | 0.26 | 0.15 | 0.18 | 0.33 | 0.03 | 0.06 | 0.09 | 0.14 | 0.03 | 0.17 |
|  |  | mt kept | 5085.8 | 9377.8 | 14463.5 | 19064.4 | 22542.1 | 41606.5 | 14563.1 | 19221.4 | 33784.5 | 5843.4 | 16750.7 | 22594.0 |
|  |  | mt discard | 232.5 | 2038.5 | 2271.0 | 2945.8 | 4008.4 | 6954.3 | 478.3 | 1117.1 | 1595.4 | 823.7 | 454.6 | 1278.3 |
| 2001 |  | ntrips |  | 17 | 17 | 22 | 18 | 40 | 8 | 17 | 25 | 12 | 11 | 23 |
|  |  | dkratio |  | 0.02 | 0.02 | 0.03 | 0.03 | 0.07 | 0.06 | 0.04 | 0.09 | 0.04 | 0.06 | 0.10 |
|  |  | mt kept | 7693.3 | 15218.8 | 22912.1 | 24272.2 | 31980.4 | 56252.7 | 22261.8 | 25588.2 | 47850.0 | 14665.1 | 19349.4 | 34014.4 |
|  |  | mt discard | 0.0 | 366.6 | 366.6 | 847.8 | 995.2 | 1843.1 | 1241.1 | 899.7 | 2140.8 | 555.7 | 1163.9 | 1719.5 |
| 2002 |  | ntrips | 7 | 4 | 11 | 1 | 22 | 23 | 12 | 22 | 34 | 7 | 20 | 27 |
|  |  | dkratio | 0.08 | 0.08 | 0.16 | 0.10 | 0.06 | 0.16 | 0.08 | 0.11 | 0.19 | 0.07 | 0.08 | 0.14 |
|  |  | mt kept | 11123.6 | 17851.7 | 28975.3 | 30540.0 | 34154.5 | 64694.5 | 28493.7 | 30490.7 | 58984.4 | 14310.0 | 19683.6 | 33993.6 |
|  |  | mt discard | 835.8 | 1509.2 | 2345.0 | 3015.8 | 2132.3 | 5148.1 | 2385.2 | 3304.9 | 5690.1 | 962.1 | 1506.2 | 2468.3 |
| 2003 |  | ntrips | 15 | 14 | 29 | 14 | 6 | 20 | 17 | 17 | 34 | 15 | 24 | 39 |
|  |  | dkratio | 0.11 | 0.07 | 0.18 | 0.05 | 0.10 | 0.15 | 0.05 | 0.09 | 0.14 | 0.06 | 0.08 | 0.13 |
|  |  | mt kept | 11318.7 | 16164.5 | 27483.3 | 35699.1 | 36028.7 | 71727.8 | 31001.4 | 30538.0 | 61539.3 | 19571.0 | 22027.6 | 41598.6 |
|  |  | mt discard | 1214.6 | 1111.0 | 2325.6 | 1739.3 | 3689.0 | 5428.2 | 1538.6 | 2863.9 | 4402.4 | 1149.6 | 1670.8 | 2820.4 |
| 2004 |  | ntrips | 9 | 13 | 22 | 27 | 28 | 55 | 56 | 26 | 82 | 35 | 54 | 89 |
|  |  | dkratio | 0.08 | 0.09 | 0.17 | 0.04 | 0.04 | 0.07 | 0.03 | 0.06 | 0.09 | 0.05 | 0.04 | 0.09 |
|  |  | mt kept | 16614.0 | 18777.6 | 35391.5 | 11961.7 | 16771.9 | 28733.6 | 2262.9 | 6101.8 | 8364.7 | 1616.5 | 9072.8 | 10689.3 |
|  |  | mt discard | 1353.9 | 1662.9 | 3016.8 | 447.6 | 619.9 | 1067.5 | 65.7 | 355.2 | 420.9 | 83.1 | 382.1 | 465.1 |
| 2005 |  | ntrips | 28 | 33 | 61 | 24 | 28 | 52 | 70 | 43 | 113 | 38 | 25 | 63 |
|  |  | dkratio | 0.06 | 0.05 | 0.11 | 0.03 | 0.06 | 0.09 | 0.05 | 0.05 | 0.10 | 0.07 | 0.04 | 0.11 |
|  |  | mt kept | 972.3 | 9753.4 | 10725.7 | 1958.8 | 17194.4 | 19153.2 | 2204.5 | 14651.3 | 16855.7 | 1129.5 | 6036.1 | 7165.6 |
|  |  | mt discard | 55.6 | 528.7 | 584.4 | 54.5 | 996.4 | 1050.9 | 101.6 | 733.4 | 835.0 | 76.8 | 246.5 | 323.2 |
| Total |  | ntrips | 90 | 109 | 199 | 119 | 149 | 268 | 193 | 252 | 445 | 133 | 263 | 396 |

Table B2.1. Strata from the NMFS spring/fall, winter, and scallop surveys which were combined for bootstrapping.

| Spring/Fall-Offshore | Spring/Fall-Inshore | Winter Survey | Winter Rosette | Scallop Survey |
| :---: | :---: | :---: | :---: | :---: |
| 1010 | $3020+3030+3040+3050$ | 1010 | 1020 | 6060 |
| 1020 | $3060+3070+3080$ | 1020 | 1030 | 6070 |
| $1030+1040$ | $3090+3100+3110$ | 1030 | 1100 | 6100 |
| 1050 | $3120+3130+3140$ | 1050 | 1110+1120 | 6110 |
| 1060 | $3150+3160+3170$ | 1060+1070 | 1610 | 6140 |
| $1070+1080$ | 3180+3190 | 1080 | $1620+1630+1640$ | 6150 |
| 1090 | 3200 | 1090 | 1650 | 6180 |
| 1100 | 3210-3220 | 1100 | 1660 | 6190 |
| 1110+1120 | 3230 | 1110 | 1670+1680 | 6220 |
| 1130 | $3240+3250+3260$ | 1610 | 1690 | 6230 |
| 1140 | $3270+3280+3290$ | 1620+1630 | 1700 | 6240 |
| 1140+1150 | $3300+3310$ | 1650 | 1710+1720 | 6250 |
| 1160 | 3320 | 1660+1670 | 1740 | 6260 |
| 1170 | 3330+3340 | 1690 | 1750+1760 | 6270 |
| 1170+1180 | 3350 | 1700+1710 |  | 6280+6290 |
| 1190 | 3360+3370 | 1730 |  | 6300 |
| 1200 | 3380 | 1740+1750 |  | 6310 |
| 1210 | $3390+3400$ |  |  | 6330 |
| 1220 | 3410 |  |  | 6340 |
| 1230 | $3420+3430$ |  |  | 6350 |
| 1240 | 3440 |  |  | 6460 |
| 1250 | $3450+3460$ |  |  | 6470 |
| 1260 | 3550 |  |  | 6490 |
| 1270 | 3550+3560 |  |  | 6500 |
| 1280 | $3580+3590+3600+3610+$ |  |  | 6510 |
| 1290+1300 | $3630+3640+3650+3660$ |  |  | 6520 |
| 1330+1340+1350 (1) |  |  |  | 6530 |
| 1360 |  |  |  | 6540 |
| 1370 |  |  |  | 6550 |
| 1380 |  |  |  | 6580 |
| 1390+1400 |  |  |  | 6590 |
| 1610+1620+1630+ |  |  |  | 6600 |
| 1640+1650 (clearnose/rosette) |  |  |  | 6610 |
| 1650+1660 (winter/little) |  |  |  | 6621+6622 |
| 1670 |  |  |  | $6631+6631+6640$ |
| 1670+1680 |  |  |  | 6651+6652 |
| 1690 |  |  |  | 6661+6662 |
| 1700 |  |  |  | 6710+6720 |
| 1710+1720 |  |  |  | 6740 |
| 1730 |  |  |  |  |
| 1740 |  |  |  |  |
| 1750+1760 |  |  |  |  |

Table B2.2. Abundance and biomass from NEFSC spring surveys for winter skate for the Gulf of Maine to Mid-Atlantic region (offshore strata $1-30,33-40,61-76)$. The mean index, $95 \%$ confidence intervals, individual fish weight, minimum, mean, and maximum length, 5 th, 50th, and 95th percentiles of length, number of nonzero tows, and number of fish caught are presented for 1968-2006.

|  | weight/tow |  |  | number/tow |  |  | ind wt | length |  |  |  |  |  | nonzero tows | no fish |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | mean | lower | upper | mean | lower | upper |  | min | 5\% | 50\% | mean | 95\% | max |  |  |
| 1968 | 2.171 | 1.640 | 2.978 | 0.854 | 0.530 | 1.178 | 2.542 | 32 | 42 | 56 | 58.6 | 79 | 112 | 36 | 232 |
| 1969 | 5.913 | 4.283 | 7.543 | 2.790 | 1.907 | 3.672 | 2.119 | 15 | 25 | 53 | 53.5 | 79 | 111 | 68 | 640 |
| 1970 | 2.645 | 1.627 | 3.663 | 0.971 | 0.626 | 1.317 | 2.723 | 37 | 43 | 59 | 61.0 | 83 | 103 | 44 | 275 |
| 1971 | 3.387 | 2.066 | 4.708 | 1.894 | 0.873 | 2.915 | 1.788 | 15 | 30 | 48 | 51.8 | 76 | 103 | 41 | 513 |
| 1972 | 4.620 | 3.033 | 6.207 | 2.602 | 1.253 | 3.951 | 1.776 | 15 | 24 | 48 | 49.5 | 74 | 97 | 63 | 634 |
| 1973 | 2.905 | 2.024 | 3.786 | 1.257 | 0.824 | 1.689 | 2.311 | 21 | 32 | 55 | 55.5 | 79 | 100 | 49 | 347 |
| 1974 | 2.091 | 1.352 | 2.830 | 0.943 | 0.505 | 1.381 | 2.218 | 29 | 34 | 53 | 55.6 | 76 | 101 | 46 | 222 |
| 1975 | 2.395 | 1.521 | 3.269 | 0.893 | 0.556 | 1.230 | 2.682 | 17 | 38 | 59 | 59.4 | 79 | 99 | 46 | 227 |
| 1976 | 2.153 | 1.075 | 3.231 | 0.628 | 0.279 | 0.978 | 3.428 | 22 | 38 | 64 | 63.1 | 86 | 97 | 29 | 160 |
| 1977 | 3.111 | 1.815 | 4.408 | 0.838 | 0.513 | 1.163 | 3.712 | 20 | 29 | 69 | 64.7 | 93 | 106 | 35 | 204 |
| 1978 | 8.275 | -0.327 | 16.877 | 1.355 | 0.121 | 2.589 | 6.108 | 43 | 62 | 79 | 78.5 | 89 | 96 | 41 | 395 |
| 1979 | 1.852 | 1.095 | 2.608 | 0.333 | 0.206 | 0.459 | 5.568 | 23 | 35 | 78 | 73.5 | 93 | 105 | 50 | 204 |
| 1980 | 2.990 | 1.751 | 4.229 | 0.538 | 0.331 | 0.745 | 5.559 | 22 | 45 | 78 | 74.8 | 97 | 104 | 49 | 187 |
| 1981 | 4.140 | 2.905 | 5.376 | 2.083 | 1.199 | 2.966 | 1.988 | 15 | 22 | 39 | 47.6 | 91 | 104 | 56 | 586 |
| 1982 | 5.773 | 3.876 | 7.670 | 2.137 | 1.195 | 3.080 | 2.701 | 15 | 26 | 46 | 54.9 | 95 | 109 | 64 | 707 |
| 1983 | 14.329 | 8.182 | 20.476 | 3.264 | 1.772 | 4.756 | 4.391 | 15 | 28 | 67 | 64.4 | 96 | 108 | 65 | 817 |
| 1984 | 10.480 | 6.816 | 14.144 | 2.948 | 1.694 | 4.201 | 3.555 | 15 | 22 | 60 | 59.0 | 94 | 106 | 59 | 753 |
| 1985 | 16.373 | 11.119 | 21.627 | 7.861 | 4.653 | 11.069 | 2.083 | 15 | 22 | 46 | 54.3 | 94 | 116 | 65 | 1891 |
| 1986 | 10.019 | 6.973 | 13.064 | 3.538 | 2.181 | 4.894 | 2.832 | 15 | 27 | 58 | 62.2 | 97 | 108 | 67 | 969 |
| 1987 | 13.126 | 8.428 | 17.824 | 4.821 | 2.926 | 6.716 | 2.723 | 15 | 29 | 56 | 60.8 | 97 | 108 | 69 | 1221 |
| 1988 | 14.543 | 10.508 | 18.577 | 7.409 | 4.736 | 10.082 | 1.963 | 15 | 25 | 43 | 53.4 | 95 | 107 | 73 | 1827 |
| 1989 | 10.141 | 7.736 | 12.546 | 4.252 | 3.095 | 5.409 | 2.385 | 15 | 25 | 59 | 61.4 | 94 | 109 | 74 | 1429 |
| 1990 | 7.183 | 5.184 | 9.183 | 5.087 | 2.657 | 7.517 | 1.412 | 15 | 27 | 41 | 49.9 | 91 | 105 | 67 | 1678 |
| 1991 | 6.965 | 4.012 | 9.918 | 3.239 | 1.979 | 4.499 | 2.150 | 17 | 29 | 54 | 58.6 | 93 | 107 | 57 | 1027 |
| 1992 | 5.988 | 3.369 | 8.607 | 5.208 | 0.635 | 9.780 | 1.150 | 15 | 23 | 42 | 46.2 | 82 | 106 | 51 | 1303 |
| 1993 | 4.761 | 3.392 | 6.131 | 4.305 | 2.561 | 6.049 | 1.106 | 15 | 25 | 42 | 46.5 | 82 | 103 | 62 | 1118 |
| 1994 | 1.421 | 0.990 | 1.852 | 1.673 | 1.150 | 2.196 | 0.849 | 20 | 32 | 43 | 46.5 | 69 | 99 | 49 | 519 |
| 1995 | 2.151 | 1.340 | 2.961 | 1.998 | 1.231 | 2.766 | 1.076 | 15 | 34 | 44 | 48.4 | 71 | 103 | 49 | 476 |
| 1996 | 4.547 | 2.499 | 6.594 | 4.470 | 2.384 | 6.556 | 1.017 | 15 | 34 | 46 | 49.0 | 68 | 96 | 56 | 1004 |
| 1997 | 3.065 | 1.325 | 4.806 | 1.834 | 0.987 | 2.680 | 1.672 | 15 | 23 | 51 | 53.5 | 78 | 93 | 39 | 458 |
| 1998 | 1.504 | 0.913 | 2.096 | 1.045 | 0.561 | 1.529 | 1.439 | 15 | 32 | 51 | 53.4 | 79 | 94 | 52 | 341 |
| 1999 | 2.968 | 1.303 | 4.632 | 1.876 | 0.870 | 2.883 | 1.582 | 16 | 27 | 54 | 54.9 | 79 | 100 | 52 | 482 |
| 2000 | 4.358 | 2.273 | 6.443 | 1.998 | 1.041 | 2.954 | 2.181 | 15 | 34 | 62 | 62.2 | 82 | 99 | 57 | 457 |
| 2001 | 3.496 | 1.889 | 5.103 | 2.350 | 0.912 | 3.787 | 1.488 | 20 | 27 | 44 | 52.1 | 82 | 100 | 48 | 556 |
| 2002 | 3.132 | 1.650 | 4.614 | 1.688 | 0.949 | 2.426 | 1.856 | 15 | 29 | 59 | 58.6 | 82 | 93 | 48 | 407 |
| 2003 | 2.799 | 1.471 | 4.127 | 2.047 | 1.164 | 2.931 | 1.367 | 15 | 29 | 49 | 53.4 | 82 | 100 | 61 | 606 |
| 2004 | 2.446 | 1.512 | 3.379 | 1.547 | 1.015 | 2.080 | 1.581 | 18 | 29 | 50 | 54.6 | 85 | 97 | 56 | 356 |
| 2005 | 1.757 | 0.869 | 2.645 | 1.672 | 0.470 | 2.874 | 1.051 | 15 | 30 | 45 | 48.6 | 75 | 97 | 52 | 375 |
| 2006 | 3.041 | 1.020 | 5.062 | 3.067 | 0.465 | 5.668 | 0.992 | 15 | 24 | 43 | 47.2 | 75 | 99 | 55 | 779 |

Table B2.3. Abundance and biomass from NEFSC autumn surveys for winter skate for the Gulf of Maine to Mid-Atlantic region (offshore strata 1-30,33-40,61-76). The mean index, $95 \%$ confidence intervals, individual fish weight, minimum, mean, and maximum length, 5 th, 50th, and 95th percentiles of length, number of nonzero tows, and number of fish caught are presented for 1967-2005.

|  | weight/tow |  |  | number/tow |  |  | ind wt | length |  |  |  |  |  | nonzero tows | no fish |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | mean | Iower | upper | mean | Iower | upper |  | min | 5\% | 50\% | mean | 95\% | max |  |  |
| 1967 | 2.159 | 1.248 | 3.070 | 0.825 | 0.544 | 1.106 | 2.617 | 15 | 32 | 56 | 57.0 | 83 | 107 | 35 | 213 |
| 1968 | 1.865 | 1.264 | 2.466 | 0.928 | 0.573 | 1.284 | 2.009 | 15 | 25 | 51 | 51.8 | 80 | 100 | 56 | 227 |
| 1969 | 1.315 | 0.856 | 1.774 | 0.540 | 0.351 | 0.730 | 2.435 | 16 | 37 | 58 | 58.3 | 78 | 90 | 36 | 161 |
| 1970 | 2.996 | 1.663 | 4.328 | 1.357 | 0.576 | 2.138 | 2.208 | 21 | 33 | 54 | 56.0 | 77 | 97 | 53 | 331 |
| 1971 | 1.078 | 0.542 | 1.615 | 0.588 | 0.238 | 0.938 | 1.833 | 18 | 27 | 50 | 50.5 | 77 | 93 | 35 | 163 |
| 1972 | 2.958 | 2.113 | 3.804 | 2.071 | 1.413 | 2.728 | 1.429 | 15 | 24 | 42 | 46.9 | 74 | 96 | 64 | 592 |
| 1973 | 4.686 | 3.348 | 6.024 | 2.238 | 1.510 | 2.967 | 2.093 | 21 | 32 | 54 | 55.1 | 78 | 101 | 48 | 662 |
| 1974 | 2.097 | 1.418 | 2.777 | 1.024 | 0.672 | 1.376 | 2.048 | 17 | 30 | 52 | 53.6 | 77 | 103 | 39 | 262 |
| 1975 | 1.315 | 0.682 | 1.948 | 0.420 | 0.260 | 0.580 | 3.130 | 16 | 24 | 62 | 60.9 | 84 | 103 | 31 | 115 |
| 1976 | 2.655 | 0.918 | 4.392 | 0.766 | 0.257 | 1.274 | 3.468 | 19 | 22 | 70 | 59.9 | 83 | 98 | 21 | 190 |
| 1977 | 4.095 | 2.814 | 5.376 | 1.617 | 1.049 | 2.185 | 2.533 | 15 | 25 | 47 | 54.8 | 87 | 100 | 51 | 662 |
| 1978 | 4.989 | 3.778 | 6.199 | 1.042 | 0.777 | 1.307 | 4.787 | 15 | 36 | 77 | 73.6 | 94 | 105 | 94 | 762 |
| 1979 | 5.121 | 3.768 | 6.475 | 1.290 | 0.976 | 1.603 | 3.971 | 20 | 31 | 75 | 66.0 | 93 | 113 | 89 | 975 |
| 1980 | 6.233 | 3.806 | 8.660 | 1.558 | 1.015 | 2.100 | 4.002 | 15 | 37 | 66 | 66.4 | 95 | 108 | 60 | 602 |
| 1981 | 5.668 | 3.726 | 7.610 | 1.505 | 0.916 | 2.094 | 3.766 | 15 | 25 | 61 | 62.3 | 99 | 110 | 54 | 516 |
| 1982 | 8.306 | 4.780 | 11.831 | 3.889 | 0.502 | 7.275 | 2.136 | 15 | 22 | 35 | 46.7 | 92 | 112 | 45 | 950 |
| 1983 | 12.852 | 5.693 | 20.012 | 2.590 | 1.447 | 3.733 | 4.962 | 16 | 28 | 78 | 70.5 | 95 | 108 | 42 | 843 |
| 1984 | 13.323 | 8.465 | 18.181 | 3.653 | 2.450 | 4.857 | 3.647 | 15 | 21 | 55 | 59.0 | 95 | 110 | 52 | 1187 |
| 1985 | 9.182 | 6.552 | 11.811 | 2.665 | 1.842 | 3.488 | 3.446 | 15 | 32 | 79 | 69.7 | 97 | 107 | 37 | 827 |
| 1986 | 15.800 | 7.184 | 24.415 | 4.196 | 2.496 | 5.895 | 3.766 | 15 | 34 | 75 | 71.5 | 97 | 110 | 46 | 1089 |
| 1987 | 11.063 | 8.200 | 13.925 | 4.291 | 2.783 | 5.800 | 2.578 | 15 | 25 | 58 | 60.1 | 97 | 109 | 49 | 1165 |
| 1988 | 7.564 | 4.961 | 10.167 | 3.126 | 2.223 | 4.028 | 2.420 | 15 | 23 | 49 | 57.4 | 97 | 110 | 45 | 888 |
| 1989 | 5.081 | 3.288 | 6.874 | 2.084 | 1.422 | 2.745 | 2.439 | 15 | 27 | 59 | 61.0 | 96 | 106 | 48 | 720 |
| 1990 | 7.145 | 4.658 | 9.632 | 2.451 | 1.397 | 3.505 | 2.915 | 22 | 33 | 68 | 66.5 | 97 | 107 | 44 | 895 |
| 1991 | 4.724 | 3.627 | 5.821 | 2.631 | 1.866 | 3.396 | 1.796 | 17 | 31 | 48 | 56.3 | 94 | 106 | 58 | 941 |
| 1992 | 3.582 | 2.140 | 5.024 | 1.862 | 1.116 | 2.608 | 1.923 | 22 | 33 | 51 | 57.4 | 91 | 103 | 39 | 509 |
| 1993 | 1.905 | 1.280 | 2.530 | 1.458 | 0.965 | 1.951 | 1.307 | 16 | 33 | 48 | 52.8 | 88 | 104 | 50 | 452 |
| 1994 | 2.120 | 1.432 | 2.808 | 1.925 | 1.217 | 2.633 | 1.101 | 15 | 26 | 44 | 47.6 | 84 | 106 | 52 | 503 |
| 1995 | 1.985 | 1.214 | 2.757 | 1.769 | 1.047 | 2.491 | 1.122 | 17 | 31 | 46 | 49.4 | 77 | 102 | 43 | 424 |
| 1996 | 2.276 | 1.615 | 2.937 | 1.426 | 0.985 | 1.867 | 1.596 | 17 | 35 | 51 | 54.9 | 83 | 104 | 44 | 370 |
| 1997 | 2.455 | 1.150 | 3.760 | 1.611 | 0.738 | 2.484 | 1.524 | 19 | 34 | 54 | 55.5 | 79 | 101 | 55 | 415 |
| 1998 | 3.753 | 2.488 | 5.018 | 2.140 | 1.438 | 2.843 | 1.753 | 19 | 27 | 55 | 56.8 | 83 | 101 | 50 | 609 |
| 1999 | 5.089 | 2.080 | 8.098 | 2.642 | 1.320 | 3.963 | 1.927 | 15 | 31 | 58 | 58.0 | 80 | 111 | 53 | 966 |
| 2000 | 4.378 | 2.390 | 6.366 | 2.535 | 1.351 | 3.718 | 1.727 | 18 | 25 | 56 | 55.5 | 82 | 99 | 45 | 756 |
| 2001 | 3.887 | 2.442 | 5.333 | 2.165 | 1.415 | 2.914 | 1.796 | 15 | 32 | 58 | 57.8 | 83 | 98 | 53 | 601 |
| 2002 | 5.600 | 3.417 | 7.782 | 2.323 | 1.535 | 3.111 | 2.411 | 16 | 33 | 66 | 63.9 | 87 | 101 | 55 | 743 |
| 2003 | 3.386 | 2.111 | 4.662 | 1.498 | 0.928 | 2.068 | 2.260 | 16 | 33 | 62 | 63.0 | 87 | 104 | 43 | 435 |
| 2004 | 4.031 | 2.632 | 5.430 | 1.942 | 1.343 | 2.542 | 2.075 | 15 | 33 | 62 | 60.4 | 87 | 102 | 50 | 611 |
| 2005 | 2.615 | 1.791 | 3.439 | 1.671 | 1.005 | 2.337 | 1.565 | 18 | 31 | 52 | 55.1 | 81 | 98 | 54 | 475 |

Table B2.4. Abundance and biomass from NEFSC winter surveys for winter skate for the Georges Bank to Mid-Atlantic region (offshore strata 1-3,5-7,9-11,13-14,16,61-63,65-67,69-71,73-75). The mean index, $95 \%$ confidence intervals, individual fish weight, minimum, mean, and maximum length, 5 th, 50 th, and 95 th percentiles of length, number of nonzero tows, and number of fish caught are presented for 1992-2006. Stratum 16 not sampled in 1993, 2000, 2002-2006. Strata 13 and 14 not sampled in 2003. Stratum 63 not sampled in 1993. Stratum 14 not sampled in 2005.

|  | weight/tow |  |  | number/tow |  |  | ind wt | min | length |  |  |  |  | nonzero tows | no fish |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | mean | lower | upper | mean | lower | upper |  |  | 5\% | 50\% | mean | 95\% |  |  |  |
| 1992 | 31.571 | 21.666 | 41.476 | 39.759 | 23.811 | 55.707 | 0.794 | 15 | 24 | 38 | 42.4 | 74 | 105 | 62 | 4042 |
| 1993 | 10.261 | 6.052 | 14.469 | 10.676 | 2.331 | 19.021 | 0.961 | 15 | 23 | 41 | 44.1 | 81 | 106 | 47 | 841 |
| 1994 | 14.439 | 10.586 | 18.293 | 14.216 | 8.465 | 19.966 | 1.016 | 15 | 29 | 40 | 45.4 | 81 | 102 | 33 | 1079 |
| 1995 | 23.268 | 14.507 | 32.029 | 35.528 | 18.060 | 52.996 | 0.655 | 15 | 27 | 40 | 42.2 | 59 | 104 | 53 | 3773 |
| 1996 | 25.239 | 7.110 | 43.369 | 43.515 | 7.434 | 79.596 | 0.580 | 15 | 25 | 40 | 41.2 | 56 | 99 | 59 | 4055 |
| 1997 | 11.643 | 7.287 | 15.999 | 12.565 | 7.109 | 18.022 | 0.927 | 15 | 27 | 45 | 46.9 | 71 | 98 | 46 | 1414 |
| 1998 | 22.464 | 15.878 | 29.050 | 19.950 | 13.556 | 26.344 | 1.126 | 15 | 26 | 48 | 49.4 | 74 | 105 | 60 | 2092 |
| 1999 | 21.089 | 13.628 | 28.549 | 18.380 | 10.899 | 25.860 | 1.147 | 15 | 24 | 49 | 49.0 | 74 | 101 | 52 | 1932 |
| 2000 | 11.315 | 4.814 | 17.815 | 5.697 | 2.799 | 8.596 | 1.986 | 18 | 27 | 56 | 57.6 | 88 | 101 | 33 | 486 |
| 2001 | 28.634 | 19.682 | 37.585 | 15.555 | 9.234 | 21.875 | 1.841 | 16 | 30 | 58 | 57.5 | 84 | 100 | 76 | 2025 |
| 2002 | 28.733 | 17.246 | 40.220 | 15.982 | 6.565 | 25.400 | 1.798 | 15 | 24 | 49 | 55.1 | 88 | 107 | 53 | 1849 |
| 2003 | 17.425 | 7.871 | 26.979 | 29.540 | -6.318 | 64.399 | 0.590 | 15 | 15 | 28 | 34.8 | 75 | 99 | 34 | 1662 |
| 2004 | 26.618 | 13.793 | 39.444 | 13.833 | 9.244 | 18.422 | 1.924 | 15 | 31 | 55 | 58.0 | 86 | 102 | 58 | 1342 |
| 2005 | 19.424 | 8.976 | 29.872 | 16.081 | 6.327 | 25.836 | 1.208 | 16 | 26 | 48 | 50.3 | 76 | 95 | 46 | 972 |
| 2006 | 32.411 | 12.125 | 52.697 | 18.233 | 9.593 | 26.874 | 1.778 | 15 | 30 | 56 | 57.4 | 86 | 102 | 60 | 1776 |

Table B2.5. Abundance and biomass from NEFSC spring surveys for little skate for the Gulf of Maine to Mid-Atlantic region (offshore strata $1-30,33-40,61-76$, and inshore strata 1-66). The mean index, $95 \%$ confidence intervals, individual fish weight, minimum, mean, and maximum length, 5 th, 50 th, and 95 th percentiles of length, number of nonzero tows, and number of fish caught are presented for 1976-2006.

|  | weight/tow |  |  | number/tow |  |  | ind wt | length |  |  |  |  |  | nonzero tows | no fish |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | mean | lower | upper | mean | lower | upper |  | min | 5\% | 50\% | mean | 95\% |  |  |  |
| 1976 | 1.308 | 0.861 | 1.755 | 3.218 | 2.136 | 4.301 | 0.406 | 8 | 12 | 40 | 36.9 | 48 | 58 | 172 | 4202 |
| 1977 | 1.347 | 0.882 | 1.811 | 3.336 | 2.177 | 4.494 | 0.404 | 6 | 19 | 41 | 38.7 | 48 | 57 | 160 | 4218 |
| 1978 | 1.391 | 0.962 | 1.821 | 3.286 | 2.363 | 4.209 | 0.423 | 8 | 11 | 42 | 37.5 | 48 | 62 | 160 | 3945 |
| 1979 | 0.650 | 0.501 | 0.799 | 2.182 | 1.429 | 2.934 | 0.298 | 4 | 12 | 31 | 32.7 | 48 | 56 | 204 | 5684 |
| 1980 | 2.206 | 1.705 | 2.707 | 5.898 | 4.384 | 7.413 | 0.374 | 8 | 12 | 37 | 36.0 | 48 | 57 | 224 | 9031 |
| 1981 | 1.501 | 1.200 | 1.803 | 3.426 | 2.714 | 4.137 | 0.438 | 6 | 15 | 41 | 38.3 | 49 | 55 | 175 | 4113 |
| 1982 | 3.627 | 2.644 | 4.611 | 7.214 | 5.351 | 9.076 | 0.503 | 9 | 18 | 43 | 40.7 | 49 | 55 | 153 | 3564 |
| 1983 | 5.718 | 4.017 | 7.420 | 13.024 | 9.215 | 16.832 | 0.439 | 6 | 16 | 42 | 37.9 | 48 | 57 | 167 | 6365 |
| 1984 | 4.094 | 2.615 | 5.574 | 10.023 | 6.787 | 13.258 | 0.409 | 7 | 11 | 40 | 35.8 | 48 | 55 | 139 | 4573 |
| 1985 | 6.265 | 4.628 | 7.901 | 15.175 | 10.575 | 19.775 | 0.413 | 8 | 11 | 40 | 36.8 | 48 | 57 | 148 | 6535 |
| 1986 | 2.753 | 1.712 | 3.795 | 8.554 | 3.399 | 13.709 | 0.322 | 6 | 14 | 33 | 34.5 | 48 | 57 | 153 | 3512 |
| 1987 | 4.625 | 3.149 | 6.102 | 16.031 | 10.222 | 21.839 | 0.289 | 8 | 12 | 32 | 33.1 | 47 | 55 | 145 | 9584 |
| 1988 | 5.083 | 3.444 | 6.721 | 14.593 | 9.688 | 19.498 | 0.348 | 8 | 11 | 36 | 34.5 | 48 | 55 | 130 | 4195 |
| 1989 | 6.634 | 3.434 | 9.834 | 21.643 | 9.844 | 33.441 | 0.307 | 8 | 13 | 34 | 33.4 | 46 | 55 | 144 | 10760 |
| 1990 | 4.993 | 2.397 | 7.589 | 14.979 | 5.250 | 24.708 | 0.333 | 8 | 11 | 37 | 34.7 | 47 | 56 | 132 | 7085 |
| 1991 | 5.990 | 4.672 | 7.308 | 18.731 | 14.059 | 23.403 | 0.320 | 8 | 13 | 34 | 34.2 | 47 | 58 | 178 | 11986 |
| 1992 | 5.297 | 2.477 | 8.118 | 16.793 | 5.234 | 28.352 | 0.315 | 8 | 16 | 33 | 34.1 | 46 | 57 | 136 | 6392 |
| 1993 | 7.524 | 5.187 | 9.862 | 22.361 | 15.110 | 29.611 | 0.336 | 9 | 12 | 36 | 35.0 | 47 | 54 | 160 | 9574 |
| 1994 | 3.622 | 2.425 | 4.819 | 9.365 | 6.297 | 12.434 | 0.387 | 9 | 19 | 39 | 37.3 | 46 | 54 | 154 | 8548 |
| 1995 | 2.872 | 2.024 | 3.720 | 7.574 | 5.215 | 9.933 | 0.379 | 8 | 10 | 39 | 36.1 | 47 | 59 | 148 | 3801 |
| 1996 | 7.574 | 5.522 | 9.626 | 18.185 | 12.647 | 23.722 | 0.417 | 7 | 17 | 41 | 38.3 | 48 | 58 | 168 | 9086 |
| 1997 | 2.708 | 2.231 | 3.184 | 6.671 | 5.504 | 7.837 | 0.406 | 9 | 13 | 40 | 37.8 | 48 | 54 | 151 | 4840 |
| 1998 | 7.471 | 6.156 | 8.787 | 20.938 | 16.232 | 25.644 | 0.357 | 7 | 17 | 37 | 35.8 | 47 | 56 | 195 | 15710 |
| 1999 | 9.978 | 7.688 | 12.267 | 28.377 | 20.345 | 36.409 | 0.352 | 8 | 12 | 38 | 35.4 | 47 | 56 | 157 | 16406 |
| 2000 | 8.596 | 6.647 | 10.545 | 19.677 | 15.270 | 24.083 | 0.437 | 9 | 21 | 41 | 38.9 | 47 | 57 | 179 | 15367 |
| 2001 | 6.835 | 4.297 | 9.372 | 15.347 | 9.900 | 20.794 | 0.445 | 8 | 18 | 42 | 39.5 | 48 | 58 | 154 | 6978 |
| 2002 | 6.444 | 4.546 | 8.341 | 16.280 | 11.306 | 21.254 | 0.396 | 8 | 11 | 42 | 37.7 | 48 | 57 | 154 | 11983 |
| 2003 | 6.486 | 4.505 | 8.486 | 15.116 | 10.195 | 20.036 | 0.429 | 9 | 22 | 42 | 40.1 | 48 | 55 | 169 | 6919 |
| 2004 | 7.219 | 5.374 | 9.064 | 17.039 | 11.917 | 22.162 | 0.424 | 7 | 25 | 42 | 39.9 | 47 | 57 | 147 | 9866 |
| 2005 | 3.241 | 2.305 | 4.177 | 7.328 | 5.515 | 9.141 | 0.442 | 8 | 13 | 43 | 38.9 | 48 | 53 | 138 | 3108 |
| 2006 | 3.323 | 1.892 | 4.753 | 7.878 | 4.544 | 11.211 | 0.422 | 7 | 11 | 42 | 38.4 | 48 | 55 | 138 | 2771 |

Table B2.6. Abundance and biomass from NEFSC autumn surveys for little skate for the Gulf of Maine to Mid-Atlantic region (offshore strata 1-30,33-40,61-76, and inshore strata 1-66). The mean index, $95 \%$ confidence intervals, individual fish weight, minimum, mean, and maximum length, 5th, 50th, and 95th percentiles of length, number of nonzero tows, and number of fish caught are presented for 1975-2005.

|  | weight/tow |  |  | number/tow |  |  | ind wt | min | length |  |  |  |  | nonzero tows | no fish |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | mean | Iower | upper | mean | Iower | upper |  |  | 5\% | 50\% | mean | 95\% | max |  |  |
| 1975 | 2.379 | 1.508 | 3.249 | 4.858 | 3.063 | 6.654 | 0.490 | 10 | 18 | 43 | 40.3 | 49 | 56 | 118 | 1386 |
| 1976 | 2.185 | 1.582 | 2.788 | 4.576 | 3.278 | 5.875 | 0.477 | 8 | 22 | 43 | 40.6 | 48 | 58 | 74 | 1421 |
| 1977 | 3.172 | 2.271 | 4.072 | 6.589 | 4.683 | 8.495 | 0.481 | 9 | 22 | 43 | 40.7 | 49 | 56 | 122 | 2438 |
| 1978 | 2.938 | 2.140 | 3.736 | 5.613 | 3.947 | 7.279 | 0.523 | 10 | 22 | 44 | 42.0 | 49 | 62 | 144 | 3171 |
| 1979 | 2.902 | 2.343 | 3.461 | 5.944 | 4.790 | 7.098 | 0.488 | 8 | 21 | 44 | 41.0 | 49 | 58 | 177 | 4597 |
| 1980 | 2.312 | 1.768 | 2.855 | 5.055 | 4.102 | 6.008 | 0.457 | 9 | 13 | 43 | 37.9 | 49 | 55 | 142 | 2451 |
| 1981 | 2.779 | 2.175 | 3.382 | 5.847 | 4.479 | 7.215 | 0.475 | 9 | 19 | 43 | 39.9 | 49 | 58 | 111 | 1728 |
| 1982 | 5.799 | 2.673 | 8.925 | 15.391 | 6.979 | 23.803 | 0.377 | 9 | 18 | 36 | 36.4 | 48 | 56 | 123 | 3848 |
| 1983 | 1.990 | 1.340 | 2.639 | 5.244 | 3.268 | 7.219 | 0.379 | 8 | 17 | 38 | 36.6 | 49 | 55 | 100 | 1313 |
| 1984 | 2.483 | 1.688 | 3.279 | 5.487 | 3.789 | 7.185 | 0.453 | 10 | 13 | 43 | 38.3 | 49 | 56 | 95 | 1350 |
| 1985 | 2.423 | 1.629 | 3.217 | 6.103 | 4.006 | 8.199 | 0.397 | 9 | 17 | 40 | 37.5 | 49 | 58 | 119 | 2761 |
| 1986 | 1.502 | 1.125 | 1.879 | 4.203 | 2.759 | 5.648 | 0.357 | 10 | 16 | 36 | 35.7 | 49 | 55 | 96 | 1240 |
| 1987 | 2.311 | 1.532 | 3.090 | 8.104 | 4.084 | 12.124 | 0.285 | 10 | 14 | 31 | 32.4 | 48 | 55 | 96 | 2093 |
| 1988 | 1.177 | 0.663 | 1.692 | 3.524 | 2.144 | 4.903 | 0.334 | 9 | 13 | 34 | 33.8 | 48 | 56 | 80 | 1128 |
| 1989 | 2.321 | 1.091 | 3.552 | 6.698 | 3.574 | 9.823 | 0.347 | 5 | 13 | 38 | 35.2 | 48 | 56 | 100 | 2288 |
| 1990 | 1.242 | 0.802 | 1.681 | 3.204 | 1.913 | 4.495 | 0.388 | 9 | 17 | 40 | 37.3 | 48 | 54 | 98 | 1183 |
| 1991 | 3.552 | 1.494 | 5.610 | 8.854 | 3.301 | 14.408 | 0.401 | 11 | 24 | 40 | 39.3 | 47 | 55 | 102 | 2866 |
| 1992 | 1.542 | 1.126 | 1.958 | 4.294 | 2.993 | 5.595 | 0.359 | 6 | 14 | 38 | 36.0 | 49 | 63 | 107 | 1460 |
| 1993 | 1.180 | 0.805 | 1.555 | 3.136 | 2.174 | 4.099 | 0.376 | 10 | 14 | 41 | 36.3 | 49 | 55 | 115 | 1124 |
| 1994 | 1.906 | 1.349 | 2.463 | 4.329 | 3.102 | 5.556 | 0.440 | 9 | 18 | 42 | 39.4 | 49 | 59 | 131 | 1729 |
| 1995 | 2.682 | 1.795 | 3.569 | 5.527 | 3.739 | 7.316 | 0.485 | 9 | 21 | 43 | 41.2 | 48 | 56 | 118 | 2058 |
| 1996 | 2.239 | 1.504 | 2.973 | 5.146 | 3.582 | 6.711 | 0.435 | 9 | 13 | 42 | 38.1 | 49 | 60 | 112 | 1878 |
| 1997 | 2.148 | 1.533 | 2.763 | 4.825 | 3.407 | 6.243 | 0.445 | 10 | 21 | 43 | 40.0 | 49 | 60 | 109 | 1757 |
| 1998 | 2.704 | 1.968 | 3.441 | 5.914 | 4.237 | 7.591 | 0.457 | 10 | 20 | 43 | 40.2 | 49 | 57 | 129 | 1713 |
| 1999 | 3.210 | 2.344 | 4.076 | 7.698 | 5.042 | 10.355 | 0.417 | 6 | 21 | 41 | 38.4 | 48 | 58 | 143 | 2289 |
| 2000 | 2.550 | 1.607 | 3.493 | 5.711 | 3.761 | 7.661 | 0.447 | 10 | 22 | 43 | 40.1 | 49 | 63 | 116 | 1759 |
| 2001 | 2.845 | 2.032 | 3.658 | 6.044 | 4.265 | 7.823 | 0.471 | 10 | 22 | 43 | 41.4 | 49 | 57 | 130 | 1985 |
| 2002 | 3.375 | 2.371 | 4.379 | 7.358 | 5.170 | 9.545 | 0.459 | 9 | 23 | 43 | 40.8 | 49 | 54 | 135 | 2515 |
| 2003 | 7.740 | 5.218 | 10.261 | 18.199 | 11.697 | 24.702 | 0.425 | 10 | 18 | 41 | 39.3 | 48 | 55 | 141 | 6523 |
| 2004 | 2.265 | 1.388 | 3.141 | 4.556 | 2.714 | 6.399 | 0.497 | 8 | 26 | 43 | 42.3 | 49 | 57 | 122 | 2270 |
| 2005 | 3.766 | 2.281 | 5.252 | 7.606 | 4.698 | 10.515 | 0.495 | 9 | 21 | 44 | 41.8 | 49 | 55 | 122 | 2437 |

Table B2.7. Abundance and biomass from NEFSC winter surveys for little skate for the Georges Bank to Mid-Atlantic region (offshore strata $1-3,5-7,9-11,13-14,16,61-63,65-67,69-71,73-75)$. The mean index, $95 \%$ confidence intervals, individual fish weight, minimum, mean, and maximum length, 5 th, 50 th, and 95 th percentiles of length, number of nonzero tows, and number of fish caught are presented for 1992-2006. Stratum 16 not sampled in 1993, 2000, 2002-2006. Strata 13 and 14 not sampled in 2003. Stratum 63 not sampled in 1993. Stratum 14 not sampled in 2005.

|  | weight/tow |  |  | number/tow |  |  | ind wt | min | length |  |  |  |  | nonzero tows | no fish |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | mean | Iower | upper | mean | Iower | upper |  |  | 5\% | 50\% | mean | 95\% | max |  |  |
| 1992 | 66.321 | 50.335 | 82.306 | 170.155 | 127.459 | 212.852 | 0.390 | 9 | 21 | 39 | 38.0 | 47 | 62 | 89 | 18418 |
| 1993 | 56.377 | 43.992 | 68.761 | 166.927 | 120.808 | 213.045 | 0.338 | 9 | 19 | 36 | 35.8 | 46 | 53 | 94 | 16026 |
| 1994 | 49.812 | 37.387 | 62.236 | 131.570 | 95.199 | 167.940 | 0.379 | 10 | 20 | 39 | 37.5 | 47 | 60 | 67 | 10113 |
| 1995 | 57.368 | 39.311 | 75.424 | 138.769 | 87.458 | 190.081 | 0.413 | 8 | 24 | 40 | 39.1 | 47 | 53 | 95 | 14530 |
| 1996 | 64.056 | 47.616 | 80.495 | 150.579 | 108.945 | 192.213 | 0.425 | 9 | 15 | 41 | 38.7 | 47 | 62 | 102 | 15701 |
| 1997 | 51.901 | 39.986 | 63.816 | 117.751 | 92.288 | 143.214 | 0.441 | 9 | 23 | 42 | 40.2 | 47 | 58 | 92 | 12084 |
| 1998 | 57.512 | 49.249 | 65.775 | 138.503 | 111.869 | 165.136 | 0.415 | 9 | 20 | 41 | 38.7 | 47 | 57 | 105 | 14492 |
| 1999 | 58.566 | 46.296 | 70.837 | 138.876 | 104.459 | 173.292 | 0.422 | 6 | 22 | 41 | 39.3 | 48 | 55 | 99 | 14740 |
| 2000 | 50.7247 | 37.806 | 63.643 | 115.572 | 87.597 | 143.547 | 0.439 | 8 | 20 | 42 | 39.5 | 47 | 53 | 92 | 10722 |
| 2001 | 47.429 | 38.584 | 56.274 | 105.749 | 85.050 | 126.447 | 0.449 | 8 | 11 | 42 | 39.7 | 48 | 63 | 120 | 12956 |
| 2002 | 63.3207 | 49.704 | 76.937 | 149.228 | 116.464 | 181.993 | 0.424 | 8 | 23 | 42 | 40.2 | 48 | 56 | 110 | 17329 |
| 2003 | 63.943 | 44.340 | 83.546 | 151.185 | 105.428 | 196.943 | 0.423 | 9 | 24 | 41 | 40.0 | 48 | 54 | 62 | 8870 |
| 2004 | 71.8027 | 50.398 | 87.208 | 162.456 | 128.807 | 196.106 | 0.442 | 10 | 25 | 41 | 40.5 | 47 | 54 | 94 | 13822 |
| 2005 | 64.149 | 45.820 | 82.478 | 140.444 | 93.239 | 187.648 | 0.457 | 9 | 25 | 42 | 40.9 | 47 | 54 | 68 | 9544 |
| 2006 | 59.2538 | 48.374 | 70.134 | 116.433 | 96.399 | 136.467 | 0.509 | 9 | 23 | 43 | 42.1 | 49 | 55 | 87 | 12687 |

Table B2.8. Abundance and biomass from NEFSC spring surveys for barndoor skate for the Gulf of Maine to Southern New England region (offshore strata 1-30, 33-40). The mean index, $95 \%$ confidence intervals, individual fish weight, minimum, mean, and maximum length, 5th, 50 th, and 95 th percentiles of length, number of nonzero tows, and number of fish caught are presented for 1968-2006.

|  | weight/tow |  |  | number/tow |  |  | ind wt | min | length |  |  |  |  | nonzero tows | no fish |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | mean | lower | upper | mean | Iower | upper |  |  | 5\% | 50\% | mean | 95\% | max |  |  |
| 1968 | 0.374 | 0.075 | 0.673 | 0.138 | 0.026 | 0.249 | 2.716 | 41 | 46 | 61 | 71.7 | 115 | 118 | 10 | 21 |
| 1969 | 0.658 | -0.364 | 1.681 | 0.145 | -0.011 | 0.301 | 4.539 | 33 | 42 | 70 | 83.1 | 119 | 120 | 8 | 22 |
| 1970 | 0.111 | 0.033 | 0.188 | 0.047 | 0.017 | 0.078 | 2.350 | 45 | 44 | 62 | 68.2 | 104 | 105 | 9 | 10 |
| 1971 | 0.116 | 0.018 | 0.214 | 0.102 | 0.021 | 0.183 | 1.134 | 26 | 31 | 59 | 57.1 | 69 | 80 | 8 | 20 |
| 1972 | 0.222 | 0.028 | 0.416 | 0.023 | 0.005 | 0.041 | 9.617 | 63 | 62 | 119 | 104.7 | 123 | 124 | 6 | 6 |
| 1973 | 0.010 | -0.001 | 0.022 | 0.017 | 0.000 | 0.034 | 0.621 | 51 | 51 | 51 | 54.1 | 59 | 60 | 3 | 3 |
| 1974 | 0.020 | -0.005 | 0.045 | 0.017 | -0.002 | 0.037 | 1.146 | 43 | 43 | 58 | 53.3 | 59 | 60 | 3 | 3 |
| 1975 | 0.001 | -0.001 | 0.003 | 0.001 | -0.001 | 0.003 | 0.900 | 60 | 60 | 60 | 60.0 | 60 | 60 | 1 | 1 |
| 1976 | 0.010 | -0.010 | 0.030 | 0.006 | -0.005 | 0.017 | 1.800 | 61 | 61 | 61 | 61.0 | 61 | 61 | 1 | 1 |
| 1977 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | - | - | - | - | - | - | - | 0 | 0 |
| 1978 | 0.015 | -0.009 | 0.040 | 0.016 | -0.006 | 0.039 | 0.933 | 51 | 50 | 55 | 56.3 | 61 | 62 | 2 | 3 |
| 1979 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | - | - | - | - | - | - | - | 0 | 0 |
| 1980 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | - | - | - | - | - | - | - | 0 | 0 |
| 1981 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | - | - | - | - | - | - | - | 0 | 0 |
| 1982 | 0.002 | -0.001 | 0.005 | 0.002 | -0.002 | 0.005 | 1.000 | 54 | 54 | 54 | 54.0 | 54 | 54 | 1 | 1 |
| 1983 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | - | - | - | - | - | - | - | 0 | 0 |
| 1984 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | - | - | - | - | - | - | - | 0 | 0 |
| 1985 | 0.001 | 0.000 | 0.002 | 0.007 | -0.004 | 0.017 | 0.076 | 20 | 20 | 20 | 24.6 | 37 | 38 | 2 | 2 |
| 1986 | 0.003 | -0.001 | 0.007 | 0.011 | -0.004 | 0.026 | 0.250 | 33 | 33 | 41 | 37.5 | 41 | 42 | 2 | 2 |
| 1987 | 0.002 | -0.002 | 0.006 | 0.007 | -0.006 | 0.020 | 0.300 | 37 | 37 | 37 | 37.0 | 37 | 37 | 1 | 1 |
| 1988 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | - | - | - | - | - | - | - | 0 | 0 |
| 1989 | 0.007 | -0.007 | 0.021 | 0.006 | -0.006 | 0.019 | 1.100 | 60 | 60 | 60 | 60.0 | 60 | 60 | 1 | 1 |
| 1990 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | - | - | - | - | - | - | - | 0 | 0 |
| 1991 | 0.002 | -0.002 | 0.006 | 0.007 | -0.006 | 0.020 | 0.300 | 38 | 38 | 38 | 38.0 | 38 | 38 | 1 | 1 |
| 1992 | 0.136 | -0.117 | 0.389 | 0.013 | -0.006 | 0.032 | 10.397 | 41 | 41 | 117 | 98.2 | 124 | 125 | 2 | 4 |
| 1993 | 0.032 | 0.024 | 0.039 | 0.028 | 0.005 | 0.051 | 1.147 | 31 | 31 | 37 | 45.3 | 89 | 90 | 5 | 5 |
| 1994 | 0.084 | -0.023 | 0.191 | 0.029 | -0.001 | 0.059 | 2.926 | 46 | 46 | 65 | 70.1 | 120 | 121 | 4 | 6 |
| 1995 | 0.015 | -0.007 | 0.037 | 0.012 | -0.005 | 0.029 | 1.254 | 55 | 55 | 63 | 59.6 | 63 | 64 | 2 | 2 |
| 1996 | 0.062 | -0.039 | 0.162 | 0.025 | -0.003 | 0.054 | 2.465 | 23 | 23 | 66 | 63.2 | 111 | 112 | 4 | 6 |
| 1997 | 0.077 | 0.006 | 0.148 | 0.035 | 0.007 | 0.063 | 2.216 | 39 | 39 | 67 | 68.7 | 89 | 90 | 6 | 7 |
| 1998 | 0.169 | -0.024 | 0.363 | 0.061 | 0.015 | 0.106 | 2.799 | 26 | 26 | 60 | 64.4 | 122 | 123 | 8 | 15 |
| 1999 | 0.279 | -0.102 | 0.660 | 0.052 | 0.011 | 0.094 | 5.343 | 28 | 28 | 74 | 80.9 | 125 | 126 | 8 | 11 |
| 2000 | 0.473 | 0.246 | 0.699 | 0.138 | 0.076 | 0.200 | 3.419 | 19 | 20 | 68 | 71.4 | 125 | 127 | 14 | 29 |
| 2001 | 0.170 | 0.032 | 0.307 | 0.141 | 0.048 | 0.234 | 1.200 | 20 | 20 | 52 | 54.8 | 77 | 115 | 13 | 30 |
| 2002 | 0.477 | 0.233 | 0.721 | 0.129 | 0.047 | 0.212 | 3.690 | 35 | 35 | 66 | 77.3 | 127 | 133 | 13 | 26 |
| 2003 | 0.885 | 0.341 | 1.429 | 0.302 | 0.172 | 0.432 | 2.928 | 19 | 19 | 54 | 64.0 | 126 | 132 | 23 | 64 |
| 2004 | 0.103 | 0.039 | 0.167 | 0.111 | 0.032 | 0.189 | 0.928 | 19 | 19 | 55 | 50.6 | 81 | 89 | 12 | 24 |
| 2005 | 0.670 | 0.120 | 1.221 | 0.319 | 0.073 | 0.565 | 2.101 | 26 | 33 | 68 | 68.1 | 109 | 122 | 15 | 59 |
| 2006 | 1.706 | -0.995 | 4.407 | 0.586 | -0.087 | 1.260 | 2.910 | 19 | 19 | 69 | 69.9 | 123 | 134 | 22 | 196 |

Table B2.9. Abundance and biomass from NEFSC autumn surveys for barndoor skate for the Gulf of Maine to Southern New England region (offshore strata 1-30, 33-40). The mean index, $95 \%$ confidence intervals, individual fish weight, minimum, mean, and maximum length, 5th, 50th, and 95th percentiles of length, number of nonzero tows, and number of fish caught are presented for 1963-2005.

|  | weight/tow |  |  | number/tow |  |  | ind wt | min | length |  |  |  |  | nonzero tows | no fish |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | mean | lower | upper | mean | lower | upper |  |  | 5\% | 50\% | mean | 95\% | max |  |  |
| 1963 | 2.633 | 1.604 | 3.663 | 0.762 | 0.468 | 1.056 | 3.458 | 28 | 44 | 69 | 74.6 | 121 | 136 | 47 | 120 |
| 1964 | 1.212 | 0.489 | 1.934 | 0.400 | 0.229 | 0.570 | 3.030 | 40 | 41 | 69 | 72.7 | 112 | 122 | 32 | 63 |
| 1965 | 1.822 | 1.115 | 2.528 | 0.695 | 0.441 | 0.949 | 2.622 | 27 | 42 | 67 | 69.9 | 111 | 134 | 36 | 95 |
| 1966 | 0.811 | 0.394 | 1.229 | 0.459 | 0.243 | 0.675 | 1.767 | 23 | 38 | 60 | 63.0 | 88 | 115 | 26 | 62 |
| 1967 | 0.438 | -0.025 | 0.901 | 0.064 | 0.017 | 0.111 | 6.844 | 45 | 52 | 65 | 81.0 | 119 | 120 | 10 | 14 |
| 1968 | 0.285 | 0.123 | 0.447 | 0.132 | 0.067 | 0.198 | 2.150 | 42 | 42 | 67 | 69.1 | 96 | 132 | 18 | 29 |
| 1969 | 0.054 | -0.003 | 0.111 | 0.035 | -0.006 | 0.076 | 1.551 | 51 | 51 | 62 | 62.0 | 73 | 74 | 5 | 8 |
| 1970 | 0.066 | -0.046 | 0.178 | 0.011 | -0.005 | 0.027 | 5.868 | 66 | 66 | 65 | 89.1 | 128 | 129 | 2 | 2 |
| 1971 | 0.170 | -0.051 | 0.392 | 0.117 | -0.077 | 0.311 | 1.455 | 35 | 35 | 53 | 54.6 | 63 | 120 | 6 | 19 |
| 1972 | 0.096 | -0.073 | 0.265 | 0.012 | -0.001 | 0.026 | 7.751 | 59 | 59 | 70 | 90.3 | 132 | 133 | 3 | 3 |
| 1973 | 0.004 | -0.001 | 0.009 | 0.008 | -0.003 | 0.019 | 0.474 | 41 | 41 | 47 | 48.7 | 52 | 53 | 2 | 3 |
| 1974 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | - | - | - | - | - | - | - | 0 | 0 |
| 1975 | 0.017 | -0.016 | 0.049 | 0.010 | -0.010 | 0.031 | 1.600 | 70 | 70 | 70 | 70.0 | 70 | 70 | 1 | 2 |
| 1976 | 0.047 | 0.002 | 0.091 | 0.058 | -0.003 | 0.119 | 0.810 | 50 | 50 | 51 | 54.6 | 61 | 62 | 7 | 10 |
| 1977 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | - | - | - | - | - | - | - | 0 | 0 |
| 1978 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | - | - | - | - | - | - | - | 0 | 0 |
| 1979 | 0.009 | -0.008 | 0.026 | 0.003 | -0.003 | 0.009 | 3.000 | 78 | 78 | 78 | 78.0 | 78 | 78 | 1 | 1 |
| 1980 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | - | - | - | - | - | - | - | 0 | 0 |
| 1981 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | - | - | - | - | - | - | - | 0 | 0 |
| 1982 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | - | - | - | - | - | - | - | 0 | 0 |
| 1983 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | - | - | - | - | - | - | - | 0 | 0 |
| 1984 | 0.010 | -0.004 | 0.024 | 0.003 | 0.000 | 0.007 | 2.900 | 61 | 61 | 84 | 73.0 | 84 | 85 | 2 | 2 |
| 1985 | 0.004 | -0.004 | 0.012 | 0.002 | -0.002 | 0.005 | 2.300 | 70 | 70 | 70 | 70.0 | 70 | 70 | 1 | 1 |
| 1986 | 0.029 | -0.018 | 0.077 | 0.015 | -0.002 | 0.032 | 2.008 | 22 | 22 | 52 | 51.0 | 90 | 91 | 3 | 3 |
| 1987 | 0.014 | -0.005 | 0.032 | 0.012 | -0.004 | 0.027 | 1.200 | 53 | 53 | 63 | 58.5 | 63 | 64 | 2 | 2 |
| 1988 | 0.007 | -0.005 | 0.020 | 0.009 | -0.005 | 0.022 | 0.850 | 34 | 34 | 33 | 44.8 | 76 | 77 | 2 | 2 |
| 1989 | 0.005 | -0.005 | 0.014 | 0.002 | -0.002 | 0.007 | 2.100 | 71 | 71 | 71 | 71.0 | 71 | 71 | 1 | 1 |
| 1990 | 0.028 | -0.022 | 0.078 | 0.010 | -0.005 | 0.024 | 2.964 | 60 | 60 | 66 | 76.3 | 95 | 96 | 2 | 3 |
| 1991 | 0.031 | 0.000 | 0.062 | 0.020 | 0.000 | 0.040 | 1.579 | 54 | 54 | 61 | 61.3 | 73 | 74 | 4 | 5 |
| 1992 | 0.002 | -0.002 | 0.007 | 0.004 | -0.004 | 0.013 | 0.550 | 46 | 46 | 51 | 49.0 | 51 | 52 | 1 | 2 |
| 1993 | 0.141 | -0.040 | 0.321 | 0.023 | 0.004 | 0.042 | 6.180 | 45 | 45 | 74 | 86.6 | 127 | 128 | 5 | 6 |
| 1994 | 0.035 | 0.001 | 0.069 | 0.044 | 0.006 | 0.082 | 0.790 | 33 | 33 | 47 | 49.4 | 75 | 76 | 6 | 9 |
| 1995 | 0.111 | -0.009 | 0.231 | 0.040 | -0.006 | 0.085 | 2.810 | 48 | 48 | 62 | 70.9 | 113 | 114 | 4 | 10 |
| 1996 | 0.042 | -0.020 | 0.104 | 0.023 | 0.000 | 0.046 | 1.841 | 25 | 25 | 61 | 59.8 | 92 | 93 | 4 | 5 |
| 1997 | 0.105 | -0.024 | 0.234 | 0.026 | 0.004 | 0.047 | 4.065 | 36 | 36 | 79 | 73.3 | 124 | 125 | 5 | 5 |
| 1998 | 0.089 | -0.036 | 0.214 | 0.026 | 0.002 | 0.050 | 3.453 | 48 | 48 | 71 | 73.9 | 120 | 121 | 4 | 5 |
| 1999 | 0.300 | 0.051 | 0.549 | 0.085 | 0.041 | 0.130 | 3.511 | 23 | 23 | 54 | 68.0 | 120 | 121 | 13 | 15 |
| 2000 | 0.288 | 0.054 | 0.521 | 0.054 | 0.023 | 0.085 | 5.360 | 29 | 29 | 89 | 85.5 | 121 | 122 | 12 | 15 |
| 2001 | 0.543 | 0.050 | 1.036 | 0.149 | 0.052 | 0.247 | 3.635 | 24 | 40 | 75 | 75.5 | 121 | 126 | 16 | 34 |
| 2002 | 0.778 | 0.351 | 1.205 | 0.269 | 0.130 | 0.407 | 2.893 | 26 | 27 | 59 | 68.0 | 119 | 129 | 24 | 59 |
| 2003 | 0.553 | 0.255 | 0.852 | 0.251 | 0.157 | 0.345 | 2.203 | 22 | 22 | 48 | 57.1 | 115 | 120 | 29 | 55 |
| 2004 | 1.295 | 0.677 | 1.913 | 0.229 | 0.122 | 0.336 | 5.662 | 42 | 47 | 80 | 90.1 | 124 | 128 | 23 | 58 |
| 2005 | 1.036 | 0.482 | 1.590 | 0.360 | 0.207 | 0.513 | 2.877 | 18 | 25 | 64 | 68.1 | 118 | 132 | 29 | 73 |

Table B2.10. Abundance and biomass from NEFSC winter surveys for barndoor skate for the Georges Bank to Mid-Atlantic region (offshore strata 1-3,5-7,9-11,13-14,16,61-63,65-67,69-71,73-75). The mean index, $95 \%$ confidence intervals, individual fish weight, minimum, mean, and maximum length, 5 th, 50 th, and 95 th percentiles of length, number of nonzero tows, and number of fish caught are presented for 1992-2006. Stratum 16 not sampled in 1993, 2000, 2002-2006. Strata 13 and 14 not sampled in 2003. Stratum 63 not sampled in 1993. Stratum 14 not sampled in 2005.

|  | weight/tow |  |  | number/tow |  |  | ind wt | min | length |  |  |  |  | nonzero tows | no fish |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | mean | lower | upper | mean | lower | upper |  |  | 5\% | 50\% | mean | 95\% | max |  |  |
| 1992 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | - | - |  |  | - |  | - | 0 | 0 |
| 1993 | 0.123 | -0.066 | 0.311 | 0.052 | 0.004 | 0.100 | 2.358 | 20 | 20 | 65 | 57.3 | 119 | 120 | 4 | 6 |
| 1994 | 0.185 | -0.027 | 0.397 | 0.080 | 0.011 | 0.148 | 2.328 | 21 | 21 | 60 | 63.5 | 102 | 103 | 5 | 7 |
| 1995 | 0.362 | 0.121 | 0.603 | 0.198 | 0.056 | 0.340 | 1.828 | 33 | 33 | 62 | 63.6 | 88 | 109 | 11 | 24 |
| 1996 | 0.291 | 0.079 | 0.503 | 0.203 | 0.054 | 0.352 | 1.434 | 19 | 20 | 61 | 56.4 | 85 | 92 | 12 | 23 |
| 1997 | 0.618 | 0.208 | 1.028 | 0.275 | 0.032 | 0.519 | 2.247 | 35 | 38 | 65 | 67.7 | 112 | 117 | 10 | 28 |
| 1998 | 0.455 | 0.146 | 0.765 | 0.464 | 0.092 | 0.837 | 0.980 | 20 | 26 | 41 | 46.8 | 83 | 123 | 12 | 57 |
| 1999 | 1.053 | 0.347 | 1.760 | 0.709 | 0.318 | 1.099 | 1.486 | 23 | 27 | 46 | 53.2 | 113 | 124 | 22 | 81 |
| 2000 | 2.718 | 0.153 | 5.284 | 1.081 | 0.518 | 1.643 | 2.515 | 19 | 19 | 56 | 62.78 | 122 | 126 | 12 | 69 |
| 2001 | 1.373 | 0.375 | 2.370 | 0.929 | 0.168 | 1.691 | 1.477 | 19 | 30 | 60 | 58.7 | 95 | 127 | 21 | 107 |
| 2002 | 2.126 | 0.506 | 3.746 | 0.950 | 0.441 | 1.459 | 2.238 | 18 | 29 | 58 | 63.9 | 119 | 126 | 24 | 123 |
| 2003 | 0.872 | 0.429 | 1.316 | 0.776 | 0.227 | 1.324 | 1.125 | 26 | 31 | 46 | 52.0 | 90 | 131 | 11 | 47 |
| 2004 | 3.397 | 1.214 | 5.581 | 1.786 | 0.972 | 2.601 | 1.902 | 18 | 30 | 53 | 60.9 | 116 | 130 | 23 | 247 |
| 2005 | 1.061 | 0.542 | 1.581 | 1.23101 | 0.703 | 1.759 | 0.862 | 18 | 19 | 44 | 47.8 | 84 | 102 | 21 | 103 |
| 2006 | 3.015 | 1.519 | 4.511 | 3.171 | 1.622 | 4.719 | 0.951 | 20 | 29 | 51 | 52.9 | 78 | 111 | 37 | 355 |

Table B2.11. Abundance and biomass from NEFSC spring surveys for thorny skate for the Gulf of Maine to Southern New England region (offshore strata 1-30,33-40). The mean index, $95 \%$ confidence intervals, individual fish weight, minimum, mean, and maximum length, 5 th, 50 th, and 95 th percentiles of length, number of nonzero tows, and number of fish caught are presented for 1968-2006.

|  | weight/tow |  |  | number/tow |  |  | ind wt | min | length |  |  |  |  | nonzero <br> tows | no fish |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | mean | lower | upper | mean | lower | upper |  |  | 5\% | 50\% | mean | 95\% | max |  |  |
| 1968 | 3.181 | 2.137 | 4.225 | 1.600 | 1.067 | 2.134 | 1.987 | 12 | 16 | 44 | 47.8 | 91 | 105 | 60 | 252 |
| 1969 | 4.526 | 3.186 | 5.865 | 1.680 | 1.161 | 2.199 | 2.694 | 12 | 13 | 47 | 51.1 | 98 | 109 | 64 | 294 |
| 1970 | 4.202 | 3.229 | 5.174 | 1.990 | 1.478 | 2.502 | 2.112 | 12 | 16 | 41 | 48.2 | 95 | 110 | 84 | 363 |
| 1971 | 3.683 | 2.475 | 4.891 | 1.974 | 1.473 | 2.475 | 1.866 | 12 | 15 | 44 | 47.8 | 95 | 116 | 81 | 424 |
| 1972 | 4.984 | 3.757 | 6.212 | 2.219 | 1.773 | 2.665 | 2.246 | 12 | 16 | 47 | 50.7 | 94 | 110 | 91 | 443 |
| 1973 | 6.622 | 4.867 | 8.377 | 3.562 | 2.640 | 4.483 | 1.859 | 12 | 15 | 44 | 47.9 | 91 | 108 | 75 | 574 |
| 1974 | 3.774 | 2.939 | 4.608 | 2.450 | 1.938 | 2.962 | 1.540 | 9 | 14 | 43 | 45.8 | 87 | 106 | 81 | 376 |
| 1975 | 3.189 | 2.222 | 4.157 | 1.360 | 0.990 | 1.731 | 2.344 | 10 | 15 | 46 | 50.5 | 95 | 102 | 62 | 192 |
| 1976 | 2.895 | 2.041 | 3.750 | 1.671 | 1.281 | 2.060 | 1.733 | 13 | 15 | 43 | 47.2 | 90 | 106 | 79 | 339 |
| 1977 | 1.623 | 1.175 | 2.070 | 0.942 | 0.675 | 1.209 | 1.722 | 12 | 15 | 42 | 48.1 | 89 | 111 | 74 | 213 |
| 1978 | 1.250 | 0.806 | 1.695 | 0.800 | 0.579 | 1.020 | 1.564 | 10 | 15 | 49 | 46.8 | 83 | 97 | 71 | 191 |
| 1979 | 1.079 | 0.729 | 1.429 | 0.582 | 0.410 | 0.754 | 1.853 | 12 | 17 | 51 | 50.5 | 84 | 102 | 68 | 163 |
| 1980 | 2.105 | 1.308 | 2.901 | 1.319 | 0.880 | 1.757 | 1.596 | 11 | 13 | 37 | 43.6 | 92 | 100 | 60 | 250 |
| 1981 | 2.700 | 2.065 | 3.335 | 1.535 | 1.139 | 1.930 | 1.760 | 9 | 13 | 47 | 48.1 | 87 | 100 | 60 | 255 |
| 1982 | 2.345 | 1.685 | 3.004 | 1.144 | 0.878 | 1.411 | 2.049 | 10 | 17 | 53 | 52.4 | 85 | 97 | 62 | 218 |
| 1983 | 2.142 | 1.398 | 2.886 | 0.968 | 0.728 | 1.209 | 2.212 | 12 | 15 | 52 | 52.3 | 91 | 103 | 55 | 156 |
| 1984 | 1.453 | 0.818 | 2.087 | 0.608 | 0.462 | 0.755 | 2.389 | 12 | 16 | 51 | 53.0 | 96 | 100 | 40 | 97 |
| 1985 | 3.074 | 2.124 | 4.024 | 1.413 | 1.060 | 1.766 | 2.175 | 11 | 14 | 44 | 48.4 | 95 | 102 | 59 | 209 |
| 1986 | 2.619 | 1.974 | 3.263 | 1.718 | 1.377 | 2.058 | 1.525 | 10 | 15 | 38 | 44.0 | 83 | 98 | 69 | 276 |
| 1987 | 1.469 | 0.805 | 2.133 | 0.852 | 0.646 | 1.058 | 1.724 | 14 | 16 | 42 | 46.6 | 87 | 109 | 53 | 141 |
| 1988 | 1.173 | 0.735 | 1.612 | 1.106 | 0.766 | 1.446 | 1.061 | 11 | 14 | 32 | 38.5 | 82 | 98 | 59 | 176 |
| 1989 | 1.481 | 0.793 | 2.169 | 1.221 | 0.801 | 1.640 | 1.213 | 11 | 15 | 34 | 40.0 | 84 | 101 | 57 | 175 |
| 1990 | 1.565 | 0.833 | 2.296 | 1.097 | 0.688 | 1.506 | 1.427 | 14 | 16 | 39 | 44.5 | 82 | 99 | 49 | 167 |
| 1991 | 1.542 | 0.945 | 2.139 | 0.858 | 0.569 | 1.147 | 1.797 | 11 | 13 | 47 | 48.5 | 89 | 99 | 47 | 132 |
| 1992 | 1.092 | 0.621 | 1.564 | 0.612 | 0.384 | 0.840 | 1.784 | 14 | 15 | 47 | 48.4 | 89 | 102 | 31 | 86 |
| 1993 | 0.700 | 0.366 | 1.034 | 0.486 | 0.327 | 0.646 | 1.440 | 13 | 13 | 36 | 42.0 | 91 | 105 | 37 | 79 |
| 1994 | 0.435 | 0.242 | 0.629 | 0.439 | 0.270 | 0.609 | 0.991 | 12 | 12 | 37 | 39.3 | 67 | 92 | 39 | 80 |
| 1995 | 0.564 | 0.307 | 0.821 | 0.384 | 0.236 | 0.533 | 1.467 | 9 | 12 | 42 | 45.8 | 84 | 92 | 31 | 66 |
| 1996 | 0.371 | 0.178 | 0.563 | 0.321 | 0.106 | 0.535 | 1.156 | 12 | 12 | 36 | 40.8 | 80 | 93 | 24 | 63 |
| 1997 | 0.422 | 0.117 | 0.727 | 0.270 | 0.153 | 0.387 | 1.560 | 15 | 20 | 47 | 47.9 | 82 | 87 | 25 | 47 |
| 1998 | 0.480 | 0.209 | 0.752 | 0.334 | 0.236 | 0.431 | 1.440 | 12 | 14 | 35 | 40.8 | 89 | 98 | 42 | 85 |
| 1999 | 0.369 | 0.093 | 0.646 | 0.255 | 0.163 | 0.347 | 1.448 | 11 | 17 | 40 | 46.2 | 83 | 89 | 26 | 44 |
| 2000 | 0.423 | 0.166 | 0.680 | 0.470 | 0.013 | 0.927 | 0.900 | 12 | 12 | 24 | 34.0 | 82 | 89 | 28 | 103 |
| 2001 | 0.493 | 0.217 | 0.769 | 0.221 | 0.080 | 0.362 | 2.234 | 14 | 33 | 56 | 57.7 | 80 | 92 | 16 | 35 |
| 2002 | 0.333 | 0.138 | 0.529 | 0.248 | 0.127 | 0.369 | 1.340 | 13 | 15 | 38 | 42.0 | 88 | 93 | 24 | 53 |
| 2003 | 0.594 | 0.268 | 0.920 | 0.332 | 0.203 | 0.461 | 1.790 | 19 | 19 | 50 | 50.9 | 86 | 102 | 30 | 57 |
| 2004 | 0.368 | 0.178 | 0.557 | 0.212 | 0.128 | 0.296 | 1.731 | 15 | 15 | 47 | 49.3 | 91 | 95 | 22 | 48 |
| 2005 | 0.435 | 0.154 | 0.716 | 0.371 | 0.167 | 0.576 | 1.171 | 16 | 17 | 44 | 44.4 | 76 | 89 | 19 | 62 |
| 2006 | 0.201 | 0.035 | 0.366 | 0.186 | 0.020 | 0.352 | 1.079 | 12 | 14 | 41 | 41.9 | 83 | 87 | 15 | 29 |

Table B2.12. Abundance and biomass from NEFSC autumn surveys for thorny skate for the Gulf of Maine to Southern New England region (offshore strata 1-30, 33-40). The mean index, $95 \%$ confidence intervals, individual fish weight, minimum, mean, and maximum length, 5 th, 50 th, and 95 th percentiles of length, number of nonzero tows, and number of fish caught are presented for 1963-2005.

|  | weight/tow |  |  | number/tow |  |  | ind wt | min | length |  |  |  |  | nonzero tows | no fish |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | mean | lower | upper | mean | lower | upper |  |  | 5\% | 50\% | mean | 95\% | max |  |  |
| 1963 | 5.371 | 3.788 | 6.954 | 1.672 | 1.305 | 2.039 | 3.213 | 10 | 15 | 60 | 60.4 | 99 | 107 | 65 | 297 |
| 1964 | 4.403 | 3.273 | 5.534 | 1.651 | 1.110 | 2.192 | 2.667 | 10 | 14 | 49 | 52.7 | 96 | 110 | 66 | 278 |
| 1965 | 4.474 | 3.268 | 5.681 | 1.825 | 1.243 | 2.408 | 2.451 | 10 | 14 | 45 | 49.6 | 95 | 107 | 55 | 352 |
| 1966 | 7.971 | 6.163 | 9.780 | 2.371 | 1.855 | 2.886 | 3.362 | 9 | 13 | 61 | 59.4 | 95 | 112 | 72 | 364 |
| 1967 | 2.712 | 1.422 | 4.001 | 0.982 | 0.383 | 1.580 | 2.763 | 12 | 14 | 49 | 52.5 | 95 | 100 | 54 | 165 |
| 1968 | 4.421 | 3.321 | 5.521 | 1.440 | 1.040 | 1.840 | 3.071 | 12 | 16 | 55 | 57.5 | 97 | 107 | 59 | 217 |
| 1969 | 5.715 | 4.320 | 7.110 | 1.833 | 1.359 | 2.307 | 3.117 | 12 | 14 | 55 | 56.7 | 97 | 106 | 72 | 289 |
| 1970 | 7.347 | 5.630 | 9.065 | 2.216 | 1.474 | 2.958 | 3.316 | 8 | 19 | 57 | 60.4 | 98 | 109 | 77 | 403 |
| 1971 | 5.357 | 4.149 | 6.565 | 1.434 | 1.095 | 1.774 | 3.735 | 12 | 18 | 63 | 64.1 | 99 | 111 | 69 | 284 |
| 1972 | 4.119 | 2.974 | 5.263 | 1.717 | 1.302 | 2.132 | 2.399 | 12 | 16 | 51 | 53.1 | 94 | 105 | 75 | 306 |
| 1973 | 4.564 | 3.227 | 5.902 | 1.536 | 1.134 | 1.939 | 2.971 | 12 | 17 | 59 | 61.2 | 95 | 111 | 72 | 274 |
| 1974 | 3.038 | 2.166 | 3.910 | 1.392 | 1.025 | 1.759 | 2.182 | 10 | 14 | 50 | 51.1 | 89 | 111 | 79 | 293 |
| 1975 | 2.474 | 1.483 | 3.464 | 1.027 | 0.716 | 1.338 | 2.409 | 10 | 12 | 47 | 50.0 | 94 | 106 | 70 | 232 |
| 1976 | 1.720 | 1.003 | 2.437 | 0.798 | 0.543 | 1.052 | 2.157 | 12 | 15 | 44 | 49.1 | 91 | 103 | 57 | 143 |
| 1977 | 3.221 | 2.513 | 3.928 | 1.548 | 1.223 | 1.874 | 2.080 | 10 | 13 | 49 | 50.7 | 89 | 107 | 108 | 446 |
| 1978 | 4.291 | 3.473 | 5.109 | 2.145 | 1.643 | 2.648 | 2.000 | 10 | 16 | 49 | 51.1 | 88 | 107 | 155 | 874 |
| 1979 | 3.612 | 2.750 | 4.474 | 1.283 | 0.864 | 1.702 | 2.815 | 11 | 21 | 59 | 59.5 | 89 | 101 | 134 | 486 |
| 1980 | 4.601 | 3.344 | 5.859 | 1.882 | 1.484 | 2.280 | 2.445 | 11 | 14 | 54 | 54.4 | 90 | 100 | 84 | 416 |
| 1981 | 3.339 | 2.551 | 4.127 | 1.305 | 0.957 | 1.653 | 2.559 | 12 | 15 | 55 | 57.1 | 90 | 103 | 71 | 223 |
| 1982 | 0.646 | 0.312 | 0.981 | 0.393 | 0.194 | 0.592 | 1.644 | 11 | 13 | 33 | 43.0 | 85 | 96 | 31 | 83 |
| 1983 | 2.409 | 1.553 | 3.266 | 0.833 | 0.589 | 1.077 | 2.892 | 15 | 20 | 56 | 58.8 | 93 | 108 | 49 | 121 |
| 1984 | 2.887 | 1.978 | 3.795 | 1.270 | 0.975 | 1.565 | 2.272 | 10 | 13 | 48 | 49.8 | 94 | 107 | 70 | 211 |
| 1985 | 2.877 | 1.765 | 3.988 | 1.438 | 1.094 | 1.783 | 2.000 | 12 | 16 | 49 | 49.6 | 87 | 103 | 66 | 260 |
| 1986 | 1.629 | 1.068 | 2.189 | 1.019 | 0.771 | 1.268 | 1.598 | 11 | 15 | 35 | 44.2 | 83 | 101 | 61 | 183 |
| 1987 | 0.944 | 0.590 | 1.297 | 0.841 | 0.600 | 1.082 | 1.123 | 12 | 14 | 36 | 40.2 | 78 | 92 | 49 | 143 |
| 1988 | 1.488 | 0.998 | 1.978 | 1.099 | 0.702 | 1.497 | 1.354 | 13 | 15 | 31 | 41.5 | 84 | 101 | 56 | 208 |
| 1989 | 1.883 | 0.980 | 2.786 | 1.129 | 0.787 | 1.471 | 1.668 | 12 | 14 | 40 | 46.2 | 85 | 101 | 63 | 198 |
| 1990 | 1.704 | 1.090 | 2.318 | 1.040 | 0.744 | 1.335 | 1.639 | 12 | 17 | 42 | 47.2 | 85 | 95 | 53 | 202 |
| 1991 | 1.632 | 0.519 | 2.745 | 0.921 | 0.591 | 1.251 | 1.772 | 13 | 15 | 47 | 49.5 | 86 | 108 | 54 | 153 |
| 1992 | 0.962 | 0.551 | 1.373 | 0.775 | 0.461 | 1.088 | 1.242 | 12 | 13 | 36 | 41.2 | 83 | 99 | 48 | 144 |
| 1993 | 1.658 | 0.639 | 2.676 | 0.901 | 0.440 | 1.361 | 1.840 | 12 | 13 | 47 | 47.8 | 91 | 101 | 50 | 157 |
| 1994 | 1.509 | 0.343 | 2.675 | 0.981 | 0.311 | 1.652 | 1.538 | 13 | 17 | 45 | 46.9 | 84 | 97 | 41 | 170 |
| 1995 | 0.783 | 0.331 | 1.235 | 0.639 | 0.183 | 1.095 | 1.226 | 13 | 14 | 39 | 42.2 | 72 | 99 | 37 | 107 |
| 1996 | 0.814 | 0.360 | 1.269 | 0.602 | 0.362 | 0.842 | 1.352 | 14 | 14 | 39 | 43.3 | 85 | 99 | 37 | 102 |
| 1997 | 0.849 | 0.405 | 1.293 | 0.404 | 0.241 | 0.567 | 2.101 | 12 | 20 | 50 | 52.3 | 83 | 99 | 33 | 79 |
| 1998 | 0.648 | 0.297 | 0.999 | 0.307 | 0.145 | 0.468 | 2.113 | 13 | 14 | 51 | 52.4 | 87 | 93 | 30 | 60 |
| 1999 | 0.479 | 0.249 | 0.710 | 0.326 | 0.195 | 0.457 | 1.469 | 13 | 14 | 41 | 46.3 | 87 | 94 | 38 | 72 |
| 2000 | 0.832 | 0.391 | 1.274 | 0.374 | 0.239 | 0.510 | 2.224 | 13 | 17 | 49 | 52.7 | 92 | 102 | 27 | 70 |
| 2001 | 0.332 | 0.087 | 0.577 | 0.294 | 0.157 | 0.430 | 1.129 | 16 | 17 | 44 | 44.1 | 74 | 82 | 23 | 60 |
| 2002 | 0.436 | 0.188 | 0.684 | 0.260 | 0.126 | 0.393 | 1.679 | 14 | 15 | 35 | 44.2 | 85 | 95 | 25 | 52 |
| 2003 | 0.742 | 0.450 | 1.035 | 0.930 | 0.168 | 1.691 | 0.798 | 12 | 14 | 23 | 34.2 | 74 | 89 | 34 | 175 |
| 2004 | 0.710 | 0.272 | 1.148 | 0.358 | 0.167 | 0.550 | 1.980 | 14 | 18 | 45 | 50.1 | 87 | 90 | 23 | 65 |
| 2005 | 0.224 | 0.092 | 0.357 | 0.205 | -0.034 | 0.443 | 1.096 | 13 | 18 | 39 | 42.6 | 76 | 90 | 17 | 36 |

Table B2.13. Abundance and biomass from NEFSC spring surveys for smooth skate for the Gulf of Maine to Southern New England region (offshore strata 1-30,33-40). The mean index, $95 \%$ confidence intervals, individual fish weight, minimum, mean, and maximum length, 5 th, 50 th, and 95 th percentiles of length, number of nonzero tows, and number of fish caught are presented for 1968-2006.

|  | weight/tow |  |  | number/tow |  |  | ind wt | length |  |  |  |  |  | nonzero tows | no fish |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | mean | Iower | upper | mean | lower | upper |  | min | 5\% | 50\% | mean | 95\% |  |  |  |
| 1968 | 0.211 | 0.080 | 0.342 | 0.484 | 0.129 | 0.838 | 0.436 | 12 | 24 | 41 | 42.1 | 58 | 64 | 17 | 41 |
| 1969 | 0.377 | 0.193 | 0.562 | 0.834 | 0.521 | 1.147 | 0.452 | 11 | 19 | 48 | 43.3 | 58 | 63 | 28 | 82 |
| 1970 | 0.346 | 0.134 | 0.557 | 0.702 | 0.376 | 1.028 | 0.492 | 9 | 14 | 47 | 40.9 | 57 | 61 | 25 | 68 |
| 1971 | 0.800 | 0.395 | 1.205 | 1.185 | 0.650 | 1.719 | 0.675 | 9 | 20 | 51 | 48.2 | 61 | 63 | 40 | 114 |
| 1972 | 0.621 | 0.355 | 0.886 | 1.016 | 0.582 | 1.450 | 0.611 | 14 | 20 | 47 | 44.3 | 59 | 64 | 34 | 122 |
| 1973 | 1.000 | 0.745 | 1.255 | 1.907 | 1.401 | 2.414 | 0.524 | 9 | 24 | 45 | 44.2 | 59 | 65 | 51 | 179 |
| 1974 | 1.092 | 0.594 | 1.590 | 2.003 | 1.109 | 2.896 | 0.545 | 9 | 9 | 47 | 42.7 | 59 | 63 | 47 | 172 |
| 1975 | 0.240 | 0.133 | 0.346 | 0.383 | 0.224 | 0.543 | 0.626 | 19 | 25 | 49 | 46.8 | 59 | 61 | 22 | 37 |
| 1976 | 0.534 | 0.413 | 0.655 | 1.150 | 0.870 | 1.429 | 0.464 | 12 | 16 | 43 | 39.8 | 57 | 60 | 49 | 134 |
| 1977 | 0.122 | 0.066 | 0.178 | 0.302 | 0.158 | 0.445 | 0.405 | 15 | 18 | 40 | 41.4 | 57 | 60 | 28 | 45 |
| 1978 | 0.251 | 0.144 | 0.358 | 0.413 | 0.258 | 0.567 | 0.609 | 24 | 26 | 50 | 46.7 | 58 | 61 | 33 | 56 |
| 1979 | 0.218 | 0.097 | 0.340 | 0.410 | 0.163 | 0.657 | 0.533 | 15 | 19 | 39 | 40.2 | 54 | 61 | 27 | 54 |
| 1980 | 0.484 | 0.316 | 0.651 | 0.948 | 0.625 | 1.271 | 0.510 | 16 | 20 | 42 | 41.9 | 56 | 60 | 42 | 84 |
| 1981 | 0.358 | 0.227 | 0.489 | 0.782 | 0.513 | 1.050 | 0.458 | 8 | 13 | 38 | 37.2 | 57 | 65 | 38 | 70 |
| 1982 | 0.152 | 0.057 | 0.247 | 0.225 | 0.092 | 0.357 | 0.677 | 11 | 10 | 52 | 45.6 | 57 | 64 | 14 | 23 |
| 1983 | 0.363 | 0.219 | 0.507 | 0.531 | 0.335 | 0.727 | 0.683 | 11 | 21 | 50 | 47.9 | 57 | 69 | 25 | 50 |
| 1984 | 0.065 | 0.010 | 0.120 | 0.124 | 0.026 | 0.221 | 0.523 | 19 | 20 | 48 | 39.8 | 59 | 60 | 9 | 13 |
| 1985 | 0.211 | 0.136 | 0.286 | 0.450 | 0.298 | 0.602 | 0.469 | 18 | 20 | 41 | 40.4 | 57 | 63 | 31 | 59 |
| 1986 | 0.250 | 0.137 | 0.362 | 0.466 | 0.256 | 0.677 | 0.536 | 20 | 24 | 48 | 46.7 | 59 | 65 | 30 | 93 |
| 1987 | 0.069 | 0.029 | 0.108 | 0.105 | 0.044 | 0.166 | 0.655 | 43 | 42 | 48 | 50.2 | 59 | 62 | 12 | 15 |
| 1988 | 0.115 | 0.044 | 0.186 | 0.328 | 0.175 | 0.480 | 0.350 | 11 | 13 | 36 | 36.3 | 57 | 60 | 24 | 49 |
| 1989 | 0.225 | 0.107 | 0.343 | 0.620 | 0.402 | 0.838 | 0.363 | 13 | 15 | 37 | 38.8 | 60 | 63 | 30 | 88 |
| 1990 | 0.152 | 0.010 | 0.294 | 0.294 | 0.080 | 0.509 | 0.515 | 11 | 16 | 46 | 44.0 | 57 | 62 | 18 | 40 |
| 1991 | 0.137 | 0.073 | 0.200 | 0.237 | 0.136 | 0.337 | 0.576 | 11 | 17 | 49 | 47.1 | 59 | 62 | 22 | 34 |
| 1992 | 0.063 | 0.025 | 0.101 | 0.104 | 0.035 | 0.172 | 0.608 | 22 | 40 | 49 | 48.5 | 56 | 57 | 12 | 16 |
| 1993 | 0.086 | 0.021 | 0.151 | 0.214 | 0.020 | 0.408 | 0.403 | 21 | 23 | 42 | 41.2 | 56 | 58 | 14 | 35 |
| 1994 | 0.098 | 0.043 | 0.153 | 0.176 | 0.082 | 0.269 | 0.558 | 29 | 29 | 47 | 47.1 | 56 | 58 | 15 | 30 |
| 1995 | 0.101 | 0.050 | 0.152 | 0.234 | 0.119 | 0.349 | 0.432 | 9 | 20 | 42 | 41.9 | 55 | 59 | 18 | 33 |
| 1996 | 0.036 | 0.014 | 0.058 | 0.084 | 0.038 | 0.129 | 0.429 | 20 | 19 | 48 | 43.8 | 53 | 59 | 10 | 12 |
| 1997 | 0.037 | 0.015 | 0.059 | 0.122 | 0.035 | 0.208 | 0.307 | 17 | 20 | 36 | 38.9 | 55 | 58 | 11 | 22 |
| 1998 | 0.200 | 0.089 | 0.311 | 0.410 | 0.206 | 0.613 | 0.489 | 9 | 19 | 46 | 44.6 | 56 | 60 | 28 | 77 |
| 1999 | 0.243 | 0.068 | 0.418 | 0.925 | -0.074 | 1.924 | 0.262 | 18 | 20 | 32 | 35.6 | 51 | 65 | 23 | 111 |
| 2000 | 0.060 | 0.025 | 0.095 | 0.220 | -0.021 | 0.460 | 0.272 | 10 | 10 | 27 | 30.9 | 59 | 62 | 13 | 30 |
| 2001 | 0.058 | 0.020 | 0.096 | 0.125 | 0.058 | 0.192 | 0.466 | 19 | 28 | 46 | 44.6 | 57 | 60 | 16 | 25 |
| 2002 | 0.184 | 0.096 | 0.271 | 0.482 | 0.297 | 0.667 | 0.381 | 10 | 13 | 45 | 40.4 | 55 | 61 | 26 | 78 |
| 2003 | 0.224 | 0.161 | 0.287 | 0.642 | 0.429 | 0.348 | 0.348 | 14 | 19 | 40 | 40.4 | 55 | 59 | 36 | 95 |
| 2004 | 0.262 | 0.141 | 0.383 | 0.650 | 0.278 | 1.022 | 0.403 | 12 | 19 | 43 | 42.3 | 56 | 60 | 32 | 125 |
| 2005 | 0.457 | 0.125 | 0.788 | 1.207 | 0.288 | 2.126 | 0.378 | 10 | 27 | 42 | 42.4 | 53 | 60 | 22 | 178 |
| 2006 | 0.203 | 0.005 | 0.401 | 0.531 | -0.009 | 1.072 | 0.382 | 19 | 21 | 41 | 41.3 | 56 | 62 | 22 | 71 |

Table B2.14. Abundance and biomass from NEFSC autumn surveys for smooth skate for the Gulf of Maine to Southern New England region (offshore strata 1-30,33-40). The mean index, $95 \%$ confidence intervals, individual fish weight, minimum, mean, and maximum length, 5 th, 50 th, and 95 th percentiles of length, number of nonzero tows, and number of fish caught are presented for 1963-2005.

|  | weight/tow |  |  | number/tow |  |  | ind wt | min | length |  |  |  |  | nonzero tows | no fish |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | mean | lower | upper | mean | lower | upper |  |  | 5\% | 50\% | mean | 95\% |  |  |  |
| 1963 | 0.498 | 0.306 | 0.689 | 0.543 | 0.282 | 0.804 | 0.917 | 9 | 20 | 48 | 43.9 | 58 | 62 | 26 | 53 |
| 1964 | 0.326 | 0.152 | 0.501 | 0.360 | 0.209 | 0.512 | 0.906 | 9 | 20 | 42 | 41.7 | 59 | 64 | 19 | 35 |
| 1965 | 0.475 | 0.140 | 0.811 | 1.221 | 0.440 | 2.001 | 0.389 | 11 | 16 | 35 | 38.1 | 56 | 64 | 27 | 94 |
| 1966 | 0.323 | 0.175 | 0.471 | 0.867 | 0.519 | 1.216 | 0.372 | 13 | 17 | 37 | 38.6 | 58 | 59 | 28 | 60 |
| 1967 | 0.152 | 0.036 | 0.268 | 0.293 | 0.118 | 0.469 | 0.518 | 22 | 24 | 48 | 46.5 | 62 | 69 | 16 | 27 |
| 1968 | 0.385 | 0.211 | 0.559 | 0.665 | 0.375 | 0.955 | 0.579 | 17 | 20 | 48 | 45.9 | 58 | 62 | 24 | 56 |
| 1969 | 0.290 | 0.131 | 0.449 | 0.604 | 0.282 | 0.925 | 0.481 | 12 | 16 | 41 | 39.6 | 58 | 64 | 21 | 50 |
| 1970 | 0.232 | 0.121 | 0.343 | 0.530 | 0.289 | 0.771 | 0.437 | 9 | 13 | 45 | 38.3 | 59 | 62 | 25 | 50 |
| 1971 | 0.157 | 0.077 | 0.238 | 0.250 | 0.120 | 0.379 | 0.631 | 17 | 36 | 53 | 51.0 | 57 | 59 | 18 | 27 |
| 1972 | 0.332 | 0.185 | 0.478 | 0.499 | 0.285 | 0.713 | 0.664 | 16 | 24 | 49 | 49.8 | 62 | 64 | 30 | 52 |
| 1973 | 0.311 | 0.199 | 0.423 | 0.506 | 0.344 | 0.667 | 0.614 | 17 | 22 | 48 | 46.9 | 58 | 60 | 32 | 56 |
| 1974 | 0.123 | 0.055 | 0.192 | 0.180 | 0.088 | 0.273 | 0.684 | 11 | 11 | 50 | 48.5 | 60 | 63 | 13 | 21 |
| 1975 | 0.076 | 0.029 | 0.123 | 0.104 | 0.043 | 0.165 | 0.727 | 21 | 30 | 49 | 46.7 | 56 | 57 | 12 | 15 |
| 1976 | 0.039 | 0.004 | 0.074 | 0.077 | 0.020 | 0.135 | 0.501 | 17 | 36 | 41 | 43.9 | 52 | 60 | 9 | 10 |
| 1977 | 0.376 | 0.274 | 0.478 | 0.600 | 0.443 | 0.757 | 0.627 | 19 | 24 | 48 | 44.9 | 56 | 61 | 50 | 84 |
| 1978 | 0.450 | 0.240 | 0.661 | 0.635 | 0.359 | 0.912 | 0.709 | 8 | 25 | 50 | 48.0 | 59 | 66 | 49 | 130 |
| 1979 | 0.182 | 0.075 | 0.288 | 0.239 | 0.116 | 0.362 | 0.761 | 9 | 29 | 50 | 48.7 | 60 | 62 | 31 | 60 |
| 1980 | 0.343 | 0.167 | 0.519 | 0.522 | 0.254 | 0.789 | 0.658 | 15 | 23 | 52 | 46.4 | 58 | 62 | 37 | 60 |
| 1981 | 0.119 | 0.039 | 0.199 | 0.167 | 0.069 | 0.264 | 0.715 | 23 | 26 | 49 | 48.1 | 60 | 61 | 13 | 18 |
| 1982 | 0.039 | 0.007 | 0.071 | 0.074 | 0.025 | 0.123 | 0.521 | 9 | 9 | 49 | 41.9 | 63 | 64 | 11 | 11 |
| 1983 | 0.146 | 0.056 | 0.236 | 0.255 | 0.085 | 0.426 | 0.573 | 14 | 14 | 46 | 40.9 | 57 | 59 | 12 | 24 |
| 1984 | 0.199 | 0.106 | 0.292 | 0.389 | 0.171 | 0.607 | 0.512 | 14 | 22 | 37 | 39.2 | 58 | 71 | 23 | 39 |
| 1985 | 0.210 | 0.088 | 0.332 | 0.340 | 0.180 | 0.500 | 0.617 | 12 | 15 | 51 | 45.2 | 59 | 63 | 28 | 64 |
| 1986 | 0.209 | 0.118 | 0.300 | 0.392 | 0.216 | 0.567 | 0.534 | 13 | 21 | 47 | 45.0 | 63 | 66 | 24 | 63 |
| 1987 | 0.095 | 0.045 | 0.145 | 0.164 | 0.081 | 0.247 | 0.581 | 15 | 15 | 48 | 44.8 | 60 | 61 | 19 | 28 |
| 1988 | 0.284 | 0.103 | 0.465 | 0.446 | 0.223 | 0.670 | 0.637 | 20 | 20 | 51 | 48.3 | 59 | 65 | 27 | 90 |
| 1989 | 0.128 | 0.072 | 0.185 | 0.336 | 0.194 | 0.478 | 0.382 | 13 | 16 | 33 | 36.8 | 59 | 62 | 27 | 52 |
| 1990 | 0.194 | 0.120 | 0.268 | 0.332 | 0.202 | 0.462 | 0.584 | 16 | 23 | 48 | 46.4 | 58 | 62 | 27 | 45 |
| 1991 | 0.167 | 0.070 | 0.265 | 0.335 | 0.188 | 0.482 | 0.500 | 18 | 20 | 46 | 43.9 | 57 | 62 | 25 | 59 |
| 1992 | 0.126 | 0.024 | 0.228 | 0.316 | 0.120 | 0.511 | 0.400 | 12 | 18 | 43 | 40.0 | 58 | 60 | 16 | 56 |
| 1993 | 0.227 | 0.107 | 0.346 | 0.818 | 0.273 | 1.362 | 0.277 | 13 | 13 | 26 | 32.6 | 56 | 62 | 29 | 123 |
| 1994 | 0.099 | 0.030 | 0.169 | 0.269 | 0.105 | 0.433 | 0.370 | 11 | 11 | 36 | 38.0 | 57 | 59 | 17 | 36 |
| 1995 | 0.189 | 0.115 | 0.263 | 0.764 | 0.315 | 1.214 | 0.247 | 10 | 13 | 30 | 32.6 | 56 | 59 | 29 | 119 |
| 1996 | 0.176 | 0.093 | 0.260 | 0.421 | 0.249 | 0.594 | 0.418 | 15 | 18 | 46 | 41.6 | 56 | 59 | 26 | 55 |
| 1997 | 0.232 | 0.117 | 0.347 | 0.449 | 0.232 | 0.665 | 0.517 | 16 | 21 | 47 | 45.2 | 60 | 64 | 20 | 59 |
| 1998 | 0.028 | 0.005 | 0.051 | 0.108 | 0.021 | 0.194 | 0.263 | 18 | 17 | 29 | 35.2 | 51 | 53 | 11 | 18 |
| 1999 | 0.070 | 0.032 | 0.109 | 0.110 | 0.050 | 0.171 | 0.638 | 22 | 22 | 50 | 48.7 | 60 | 62 | 16 | 22 |
| 2000 | 0.154 | 0.083 | 0.226 | 0.318 | 0.190 | 0.447 | 0.485 | 10 | 11 | 45 | 42.3 | 59 | 73 | 27 | 55 |
| 2001 | 0.287 | 0.169 | 0.405 | 0.565 | 0.349 | 0.781 | 0.507 | 17 | 23 | 49 | 46.5 | 58 | 62 | 29 | 84 |
| 2002 | 0.111 | 0.067 | 0.155 | 0.209 | 0.140 | 0.278 | 0.533 | 15 | 24 | 50 | 46.2 | 60 | 62 | 25 | 32 |
| 2003 | 0.190 | 0.076 | 0.304 | 0.646 | 0.248 | 1.045 | 0.294 | 10 | 14 | 39 | 36.3 | 52 | 62 | 30 | 84 |
| 2004 | 0.214 | 0.126 | 0.303 | 0.467 | 0.283 | 0.652 | 0.458 | 18 | 24 | 47 | 45.3 | 55 | 59 | 29 | 58 |
| 2005 | 0.131 | 0.039 | 0.224 | 0.291 | 0.143 | 0.439 | 0.451 | 15 | 17 | 47 | 43.1 | 59 | 62 | 18 | 44 |

Table B2.15. Abundance and biomass from NEFSC spring surveys for clearnose skate for the Mid-Atlantic region (offshore strata 61-76, inshore strata 15-44). The mean index, $95 \%$ confidence intervals, individual fish weight, minimum, mean, and maximum length, 5 th, 50 th, and 95th percentiles of length, number of nonzero tows, and number of fish caught are presented for 1976-2006.

|  | weight/tow |  |  | number/tow |  |  | ind wt | length |  |  |  |  |  | nonzero <br> tows | no fish |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | mean | lower | upper | mean | lower | upper |  | min | 5\% | 50\% | mean | 95\% |  |  |  |
| 1976 | 0.100 | 0.020 | 0.179 | 0.129 | 0.040 | 0.218 | 0.770 | 26 | 26 | 43 | 48.5 | 66 | 67 | 8 | 12 |
| 1977 | 0.509 | 0.297 | 0.722 | 0.500 | 0.260 | 0.741 | 1.017 | 23 | 23 | 56 | 52.5 | 63 | 64 | 17 | 41 |
| 1978 | 0.211 | -0.094 | 0.516 | 0.237 | -0.057 | 0.530 | 0.893 | 20 | 20 | 57 | 52.2 | 68 | 69 | 8 | 21 |
| 1979 | 0.109 | 0.010 | 0.209 | 0.125 | 0.004 | 0.247 | 0.875 | 25 | 25 | 42 | 50.3 | 77 | 78 | 6 | 9 |
| 1980 | 0.319 | 0.100 | 0.538 | 0.456 | 0.136 | 0.775 | 0.700 | 25 | 25 | 41 | 45.1 | 64 | 69 | 14 | 44 |
| 1981 | 0.891 | -0.141 | 1.923 | 0.606 | 0.106 | 1.107 | 1.469 | 24 | 26 | 60 | 55.9 | 67 | 72 | 10 | 44 |
| 1982 | 0.328 | 0.165 | 0.491 | 0.368 | 0.126 | 0.610 | 0.892 | 30 | 32 | 52 | 53.6 | 66 | 71 | 14 | 40 |
| 1983 | 0.138 | 0.005 | 0.270 | 0.127 | 0.003 | 0.252 | 1.081 | 13 | 13 | 58 | 51.3 | 65 | 66 | 7 | 11 |
| 1984 | 0.380 | 0.103 | 0.658 | 0.288 | 0.018 | 0.557 | 1.321 | 48 | 48 | 62 | 60.7 | 70 | 74 | 11 | 25 |
| 1985 | 0.493 | -0.166 | 1.151 | 0.436 | -0.203 | 1.076 | 1.129 | 48 | 48 | 58 | 59.3 | 69 | 72 | 10 | 37 |
| 1986 | 0.155 | 0.035 | 0.274 | 0.232 | 0.038 | 0.427 | 0.666 | 27 | 27 | 44 | 44.8 | 68 | 69 | 11 | 15 |
| 1987 | 0.306 | 0.150 | 0.463 | 0.202 | 0.109 | 0.204 | 1.519 | 49 | 51 | 63 | 61.9 | 69 | 72 | 16 | 20 |
| 1988 | 0.340 | 0.171 | 0.508 | 0.300 | 0.097 | 0.502 | 1.134 | 44 | 44 | 58 | 57.1 | 67 | 71 | 11 | 19 |
| 1989 | 0.424 | 0.258 | 0.590 | 0.415 | 0.275 | 0.554 | 1.023 | 25 | 25 | 58 | 52.3 | 68 | 72 | 14 | 40 |
| 1990 | 0.501 | 0.283 | 0.719 | 0.420 | 0.243 | 0.597 | 1.192 | 30 | 30 | 59 | 56.2 | 67 | 72 | 15 | 52 |
| 1991 | 0.690 | 0.463 | 0.918 | 0.543 | 0.354 | 0.731 | 1.272 | 27 | 27 | 62 | 58.8 | 68 | 71 | 23 | 59 |
| 1992 | 0.748 | 0.324 | 1.172 | 0.489 | 0.218 | 0.760 | 1.529 | 46 | 46 | 63 | 63.0 | 68 | 80 | 23 | 47 |
| 1993 | 0.856 | 0.479 | 1.233 | 0.656 | 0.216 | 1.096 | 1.305 | 21 | 33 | 63 | 58.6 | 70 | 74 | 12 | 136 |
| 1994 | 0.319 | 0.052 | 0.585 | 0.188 | 0.043 | 0.333 | 1.699 | 51 | 57 | 65 | 66.0 | 73 | 74 | 8 | 24 |
| 1995 | 0.669 | 0.361 | 0.977 | 0.464 | 0.261 | 0.666 | 1.443 | 46 | 46 | 67 | 62.4 | 68 | 74 | 18 | 32 |
| 1996 | 1.224 | 0.194 | 2.254 | 0.948 | 0.255 | 1.641 | 1.291 | 13 | 27 | 62 | 59.8 | 70 | 75 | 30 | 95 |
| 1997 | 1.290 | 0.885 | 1.695 | 0.972 | 0.542 | 1.403 | 1.326 | 33 | 39 | 63 | 61.3 | 71 | 78 | 22 | 80 |
| 1998 | 0.903 | 0.674 | 1.133 | 0.667 | 0.369 | 0.964 | 1.355 | 26 | 38 | 62 | 60.2 | 70 | 74 | 29 | 81 |
| 1999 | 0.943 | 0.647 | 1.238 | 0.862 | 0.470 | 1.255 | 1.093 | 26 | 28 | 59 | 57.3 | 67 | 72 | 19 | 54 |
| 2000 | 1.391 | 1.046 | 1.736 | 1.140 | 0.789 | 1.491 | 1.221 | 24 | 40 | 59 | 59.4 | 70 | 76 | 31 | 126 |
| 2001 | 1.380 | 0.674 | 2.087 | 1.097 | 0.456 | 1.738 | 1.258 | 42 | 49 | 62 | 60.8 | 68 | 72 | 19 | 74 |
| 2002 | 0.836 | 0.281 | 1.392 | 0.617 | 0.241 | 0.993 | 1.355 | 29 | 42 | 62 | 60.5 | 69 | 74 | 23 | 59 |
| 2003 | 0.622 | 0.366 | 0.879 | 0.448 | 0.265 | 0.631 | 1.389 | 49 | 49 | 62 | 62.7 | 75 | 76 | 16 | 35 |
| 2004 | 0.433 | 0.050 | 0.815 | 0.376 | 0.049 | 0.703 | 1.151 | 35 | 35 | 59 | 56.2 | 70 | 72 | 9 | 23 |
| 2005 | 0.569 | 0.030 | 1.109 | 0.414 | 0.008 | 0.820 | 1.374 | 42 | 42 | 61 | 61.2 | 70 | 73 | 11 | 27 |
| 2006 | 0.567 | 0.189 | 0.946 | 0.420 | 0.179 | 0.661 | 1.350 | 36 | 41 | 63 | 60.7 | 68 | 72 | 18 | 39 |

Table B2.16. Abundance and biomass from NEFSC autumn surveys for clearnose skate for the Mid-Atlantic region (offshore strata 61-76, inshore strata 15-44). The mean index, $95 \%$ confidence intervals, individual fish weight, minimum, mean, and maximum length, 5 th, 50 th, and 95th percentiles of length, number of nonzero tows, and number of fish caught are presented for 1975-2005.

|  | weight/tow |  |  | number/tow |  |  | ind wt | length |  |  |  |  |  | nonzero tows | no fish |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | mean | lower | upper | mean | lower | upper |  | min | 5\% | 50\% | mean | 95\% |  |  |  |
| 1975 | 0.237 | 0.086 | 0.388 | 0.246 | 0.133 | 0.360 | 0.961 | 21 | 21 | 53 | 50.3 | 63 | 66 | 31 | 49 |
| 1976 | 0.302 | 0.189 | 0.415 | 0.348 | 0.236 | 0.459 | 0.869 | 18 | 34 | 52 | 52.1 | 64 | 69 | 26 | 54 |
| 1977 | 0.768 | 0.288 | 1.248 | 0.742 | 0.281 | 1.203 | 1.035 | 15 | 37 | 57 | 55.4 | 65 | 68 | 32 | 106 |
| 1978 | 0.156 | 0.073 | 0.240 | 0.224 | 0.086 | 0.363 | 0.697 | 10 | 10 | 44 | 40.8 | 64 | 66 | 14 | 23 |
| 1979 | 0.419 | 0.116 | 0.721 | 0.346 | 0.146 | 0.545 | 1.211 | 22 | 24 | 56 | 55.4 | 67 | 71 | 27 | 46 |
| 1980 | 0.685 | 0.408 | 0.961 | 0.549 | 0.322 | 0.775 | 1.248 | 33 | 37 | 59 | 58.1 | 69 | 72 | 32 | 80 |
| 1981 | 0.171 | 0.081 | 0.260 | 0.179 | 0.087 | 0.271 | 0.954 | 27 | 27 | 55 | 51.5 | 65 | 68 | 19 | 28 |
| 1982 | 0.213 | 0.099 | 0.326 | 0.183 | 0.095 | 0.271 | 1.163 | 32 | 43 | 59 | 58.3 | 67 | 72 | 26 | 37 |
| 1983 | 0.141 | 0.027 | 0.254 | 0.127 | 0.043 | 0.210 | 1.110 | 16 | 16 | 57 | 52.2 | 64 | 70 | 15 | 19 |
| 1984 | 0.178 | 0.064 | 0.293 | 0.189 | 0.063 | 0.315 | 0.945 | 34 | 37 | 53 | 54.0 | 67 | 83 | 20 | 32 |
| 1985 | 0.306 | 0.173 | 0.439 | 0.315 | 0.182 | 0.447 | 0.974 | 32 | 41 | 56 | 54.9 | 66 | 71 | 23 | 42 |
| 1986 | 0.545 | -0.038 | 1.027 | 0.591 | 0.091 | 1.092 | 0.921 | 23 | 23 | 59 | 52.6 | 64 | 71 | 31 | 62 |
| 1987 | 0.320 | 0.176 | 0.465 | 0.289 | 0.167 | 0.412 | 1.107 | 15 | 41 | 56 | 55.5 | 69 | 70 | 23 | 42 |
| 1988 | 0.335 | 0.157 | 0.513 | 0.329 | 0.163 | 0.495 | 1.019 | 33 | 37 | 57 | 56.0 | 66 | 71 | 19 | 60 |
| 1989 | 0.273 | 0.075 | 0.471 | 0.324 | 0.064 | 0.584 | 0.843 | 37 | 37 | 52 | 52.7 | 63 | 70 | 20 | 39 |
| 1990 | 0.402 | 0.157 | 0.646 | 0.306 | 0.114 | 0.499 | 1.311 | 16 | 41 | 60 | 57.9 | 69 | 72 | 17 | 50 |
| 1991 | 0.922 | 0.279 | 1.566 | 0.816 | 0.339 | 1.294 | 1.130 | 35 | 39 | 58 | 57.1 | 69 | 71 | 35 | 119 |
| 1992 | 0.345 | 0.185 | 0.505 | 0.312 | 0.185 | 0.440 | 1.104 | 16 | 42 | 59 | 56.7 | 67 | 69 | 22 | 48 |
| 1993 | 0.495 | 0.145 | 0.844 | 0.474 | 0.188 | 0.759 | 1.044 | 35 | 40 | 57 | 56.8 | 66 | 73 | 27 | 104 |
| 1994 | 0.938 | 0.479 | 1.398 | 0.842 | 0.494 | 1.190 | 1.115 | 35 | 40 | 57 | 57.1 | 66 | 73 | 35 | 129 |
| 1995 | 0.331 | 0.189 | 0.473 | 0.426 | 0.233 | 0.618 | 0.777 | 14 | 14 | 51 | 45.5 | 66 | 72 | 25 | 63 |
| 1996 | 0.430 | 0.194 | 0.666 | 0.369 | 0.163 | 0.576 | 1.165 | 29 | 45 | 59 | 58.8 | 68 | 72 | 20 | 42 |
| 1997 | 0.614 | 0.296 | 0.932 | 0.484 | 0.281 | 0.688 | 1.269 | 43 | 43 | 61 | 60.2 | 69 | 77 | 27 | 60 |
| 1998 | 1.121 | 0.115 | 2.128 | 1.096 | 0.124 | 2.068 | 1.023 | 34 | 43 | 57 | 57.5 | 68 | 73 | 32 | 98 |
| 1999 | 1.053 | 0.536 | 1.570 | 0.928 | 0.525 | 1.332 | 1.134 | 15 | 32 | 61 | 57.8 | 69 | 71 | 41 | 84 |
| 2000 | 1.032 | 0.422 | 1.642 | 0.795 | 0.353 | 1.238 | 1.298 | 14 | 47 | 60 | 60.5 | 69 | 74 | 29 | 61 |
| 2001 | 1.614 | 1.092 | 2.136 | 1.494 | 0.984 | 2.004 | 1.081 | 13 | 15 | 59 | 55.2 | 68 | 73 | 41 | 221 |
| 2002 | 0.891 | 0.372 | 1.411 | 0.863 | 0.317 | 1.409 | 1.033 | 14 | 38 | 55 | 56.0 | 68 | 73 | 27 | 63 |
| 2003 | 0.661 | 0.417 | 0.906 | 0.640 | 0.456 | 0.823 | 1.034 | 15 | 30 | 54 | 54.5 | 71 | 78 | 38 | 81 |
| 2004 | 0.709 | 0.201 | 1.217 | 0.590 | 0.172 | 1.008 | 1.201 | 37 | 43 | 62 | 60.1 | 69 | 75 | 18 | 55 |
| 2005 | 0.524 | 0.192 | 0.855 | 0.452 | 0.207 | 0.697 | 1.159 | 26 | 37 | 62 | 59.6 | 71 | 74 | 30 | 71 |

Table B2.17. Abundance and biomass from NEFSC winter surveys for clearnose skate for the Georges Bank to Mid-Atlantic region (offshore strata 1-3,5-7,9-11,13-14,16,61-63,65-67,69-71,73-75). The mean index, $95 \%$ confidence intervals, individual fish weight, minimum, mean, and maximum length, 5 th, 50 th, and 95 th percentiles of length, number of nonzero tows, and number of fish caught are presented for 1992-2006. Stratum 16 not sampled in 1993, 2000, 2002-2006. Strata 13 and 14 not sampled in 2003. Stratum 63 not sampled in 1993. Stratum 14 not sampled in 2005.

|  | weight/tow |  |  | number/tow |  |  | ind wt | min | length |  |  |  |  | nonzero tows | no fish |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | mean | lower | upper | mean | lower | upper |  |  | 5\% | 50\% | mean | 95\% | max |  |  |
| 1992 | 5.622 | 3.247 | 7.997 | 5.247 | 2.974 | 7.519 | 1.072 | 23 | 26 | 59 | 54.7 | 67 | 93 | 22 | 551 |
| 1993 | 6.013 | 3.818 | 8.208 | 5.973 | 3.852 | 8.093 | 1.007 | 22 | 33 | 57 | 54.3 | 67 | 81 | 23 | 716 |
| 1994 | 8.854 | 4.037 | 13.672 | 7.692 | 2.152 | 13.233 | 1.151 | 27 | 33 | 60 | 57.5 | 69 | 77 | 16 | 639 |
| 1995 | 7.924 | 2.521 | 13.327 | 6.247 | 1.301 | 11.194 | 1.268 | 24 | 45 | 61 | 60.2 | 69 | 76 | 23 | 737 |
| 1996 | 14.725 | 8.266 | 21.183 | 11.555 | 6.347 | 16.762 | 1.274 | 22 | 40 | 61 | 60.0 | 69 | 77 | 32 | 3086 |
| 1997 | 5.522 | 3.154 | 7.890 | 5.069 | 2.158 | 7.980 | 1.089 | 22 | 35 | 59 | 56.2 | 70 | 76 | 32 | 682 |
| 1998 | 6.031 | 4.470 | 7.592 | 4.878 | 3.195 | 6.560 | 1.236 | 22 | 36 | 60 | 58.3 | 71 | 88 | 32 | 1091 |
| 1999 | 3.826 | 2.335 | 5.317 | 3.022 | 1.586 | 4.459 | 1.266 | 23 | 37 | 61 | 59.6 | 70 | 76 | 30 | 343 |
| 2000 | 10.102 | 5.693 | 14.510 | 8.864 | 4.579 | 13.150 | 1.140 | 25 | 42 | 59 | 58.2 | 69 | 93 | 43 | 1449 |
| 2001 | 8.316 | 5.624 | 11.008 | 6.599 | 4.240 | 8.957 | 1.260 | 25 | 43 | 61 | 60.6 | 69 | 86 | 41 | 1300 |
| 2002 | 12.223 | 8.343 | 16.102 | 8.864 | 5.886 | 11.843 | 1.379 | 23 | 39 | 63 | 61.6 | 70 | 74 | 51 | 1704 |
| 2003 | 19.637 | 13.819 | 25.455 | 15.769 | 10.902 | 20.635 | 1.245 | 23 | 39 | 62 | 59.1 | 70 | 81 | 36 | 2260 |
| 2004 | 11.566 | 7.743 | 15.389 | 10.162 | 6.344 | 13.979 | 1.138 | 20 | 35 | 60 | 58.1 | 70 | 80 | 38 | 1880 |
| 2005 | 6.036 | 3.837 | 8.235 | 5.078 | 2.425 | 7.731 | 1.189 | 24 | 44 | 60 | 59.1 | 70 | 82 | 26 | 1047 |
| 2006 | 11.723 | 4.862 | 18.585 | 11.085 | 4.693 | 17.477 | 1.058 | 23 | 35 | 57 | 56.7 | 70 | 77 | 41 | 1916 |

Table B2.18. Abundance and biomass from NEFSC spring surveys for rosette skate for the Mid-Atlantic region (offshore strata 61-76). The mean index, $95 \%$ confidence intervals, individual fish weight, minimum, mean, and maximum length, 5 th, 50 th, and 95 th percentiles of length, number of nonzero tows, and number of fish caught are presented for 1968-2006.

|  | weight/tow |  |  | number/tow |  |  | ind wt | length |  |  |  |  |  | nonzero tows | no fish |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | mean | lower | upper | mean | lower | upper |  | min | 5\% | 50\% | mean | 95\% | max |  |  |
| 1968 | 0.005 | -0.002 | 0.012 | 0.014 | 0.000 | 0.029 | 0.356 | 33 | 33 | 33 | 34.4 | 35 | 36 | 3 | 3 |
| 1969 | 0.001 | -0.001 | 0.002 | 0.003 | -0.003 | 0.010 | 0.200 | 37 | 37 | 37 | 37.0 | 37 | 37 | 1 | 1 |
| 1970 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | - | - | - | - | - | - | - | 0 | 0 |
| 1971 | 0.005 | -0.005 | 0.014 | 0.010 | -0.009 | 0.028 | 0.500 | 57 | 57 | 57 | 57.0 | 57 | 57 | 1 | 1 |
| 1972 | 0.000 | 0.000 | 0.001 | 0.003 | -0.003 | 0.010 | 0.100 | 35 | 35 | 35 | 35.0 | 35 | 35 | 1 | 1 |
| 1973 | 0.006 | -0.001 | 0.012 | 0.023 | -0.006 | 0.052 | 0.240 | 38 | 38 | 38 | 38.6 | 41 | 42 | 4 | 5 |
| 1974 | 0.005 | -0.005 | 0.015 | 0.025 | -0.024 | 0.074 | 0.200 | 41 | 41 | 41 | 41.0 | 41 | 41 | 1 | 1 |
| 1975 | 0.001 | -0.001 | 0.003 | 0.005 | -0.005 | 0.014 | 0.200 | 38 | 38 | 38 | 38.5 | 39 | 39 | 1 | 2 |
| 1976 | 0.007 | 0.000 | 0.015 | 0.035 | -0.003 | 0.073 | 0.208 | 31 | 31 | 36 | 36.9 | 44 | 45 | 4 | 6 |
| 1977 | 0.102 | 0.019 | 0.186 | 0.552 | 0.107 | 0.998 | 0.185 | 20 | 26 | 32 | 33.6 | 37 | 42 | 11 | 70 |
| 1978 | 0.010 | 0.001 | 0.019 | 0.041 | 0.008 | 0.074 | 0.232 | 12 | 25 | 35 | 35.3 | 40 | 41 | 7 | 10 |
| 1979 | 0.007 | 0.005 | 0.009 | 0.040 | 0.031 | 0.048 | 0.171 | 13 | 13 | 34 | 31.6 | 40 | 41 | 4 | 10 |
| 1980 | 0.072 | 0.030 | 0.115 | 0.373 | 0.167 | 0.580 | 0.194 | 26 | 27 | 34 | 35.3 | 41 | 42 | 15 | 47 |
| 1981 | 0.013 | 0.001 | 0.025 | 0.057 | 0.006 | 0.109 | 0.231 | 19 | 28 | 37 | 36.3 | 41 | 42 | 6 | 17 |
| 1982 | 0.025 | 0.010 | 0.040 | 0.108 | 0.043 | 0.174 | 0.234 | 22 | 25 | 37 | 37.4 | 43 | 44 | 11 | 20 |
| 1983 | 0.002 | -0.001 | 0.004 | 0.012 | -0.006 | 0.029 | 0.147 | 29 | 29 | 34 | 34.2 | 35 | 36 | 2 | 5 |
| 1984 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | - | - | - | - | - | - | - | 0 | 0 |
| 1985 | 0.005 | -0.001 | 0.011 | 0.059 | 0.040 | 0.079 | 0.080 | 17 | 17 | 18 | 21.0 | 29 | 42 | 3 | 9 |
| 1986 | 0.002 | -0.002 | 0.006 | 0.012 | -0.008 | 0.031 | 0.182 | 32 | 32 | 35 | 35.3 | 35 | 36 | 2 | 2 |
| 1987 | 0.003 | -0.002 | 0.009 | 0.017 | -0.012 | 0.046 | 0.200 | 35 | 35 | 36 | 36.7 | 36 | 37 | 2 | 2 |
| 1988 | 0.020 | -0.001 | 0.041 | 0.111 | -0.002 | 0.223 | 0.180 | 26 | 26 | 35 | 32.8 | 35 | 36 | 4 | 6 |
| 1989 | 0.010 | -0.004 | 0.025 | 0.051 | -0.036 | 0.137 | 0.200 | 28 | 28 | 34 | 34.6 | 40 | 41 | 2 | 15 |
| 1990 | 0.010 | -0.004 | 0.024 | 0.049 | -0.022 | 0.121 | 0.200 | 36 | 36 | 35 | 36.0 | 35 | 36 | 3 | 3 |
| 1991 | 0.036 | 0.014 | 0.058 | 0.143 | 0.057 | 0.228 | 0.253 | 19 | 33 | 37 | 37.2 | 40 | 42 | 7 | 19 |
| 1992 | 0.014 | -0.001 | 0.029 | 0.063 | 0.012 | 0.113 | 0.223 | 24 | 24 | 37 | 36.0 | 40 | 41 | 5 | 5 |
| 1993 | 0.009 | 0.007 | 0.011 | 0.037 | 0.030 | 0.043 | 0.255 | 38 | 38 | 37 | 38.6 | 39 | 40 | 2 | 5 |
| 1994 | 0.005 | 0.001 | 0.009 | 0.021 | 0.006 | 0.035 | 0.243 | 36 | 36 | 38 | 38.7 | 40 | 41 | 4 | 4 |
| 1995 | 0.010 | 0.000 | 0.020 | 0.056 | 0.003 | 0.110 | 0.173 | 19 | 19 | 35 | 32.9 | 36 | 37 | 3 | 5 |
| 1996 | 0.014 | -0.011 | 0.039 | 0.095 | -0.013 | 0.203 | 0.149 | 9 | 9 | 35 | 29.3 | 42 | 43 | 5 | 19 |
| 1997 | 0.028 | 0.022 | 0.033 | 0.138 | 0.091 | 0.186 | 0.200 | 30 | 30 | 34 | 35.6 | 41 | 42 | 4 | 25 |
| 1998 | 0.038 | 0.007 | 0.068 | 0.132 | 0.041 | 0.223 | 0.287 | 32 | 33 | 38 | 38.0 | 41 | 42 | 11 | 15 |
| 1999 | 0.043 | 0.003 | 0.083 | 0.206 | 0.012 | 0.399 | 0.211 | 15 | 29 | 37 | 36.7 | 42 | 43 | 9 | 16 |
| 2000 | 0.026 | 0.009 | 0.043 | 0.106 | 0.040 | 0.171 | 0.247 | 30 | 32 | 37 | 38.0 | 41 | 42 | 7 | 15 |
| 2001 | 0.010 | -0.005 | 0.025 | 0.041 | -0.012 | 0.095 | 0.244 | 21 | 21 | 40 | 38.2 | 40 | 41 | 4 | 4 |
| 2002 | 0.019 | -0.007 | 0.045 | 0.076 | -0.029 | 0.180 | 0.252 | 12 | 12 | 38 | 34.1 | 39 | 40 | 3 | 5 |
| 2003 | 0.028 | -0.002 | 0.057 | 0.115 | 0.003 | 0.226 | 0.241 | 9 | 24 | 38 | 37.0 | 39 | 41 | 5 | 17 |
| 2004 | 0.023 | -0.009 | 0.055 | 0.084 | -0.025 | 0.193 | 0.276 | 30 | 32 | 39 | 39.2 | 40 | 41 | 3 | 7 |
| 2005 | 0.050 | -0.029 | 0.128 | 0.216 | -0.131 | 0.564 | 0.229 | 13 | 31 | 37 | 36.7 | 40 | 41 | 5 | 21 |
| 2006 | 0.012 | 0.007 | 0.016 | 0.051 | 0.020 | 0.081 | 0.230 | 25 | 25 | 39 | 35.5 | 40 | 41 | 5 | 8 |

Table B2.19. Abundance and biomass from NEFSC autumn surveys for rosette skate for the Mid-Atlantic region (offshore strata 61-76). The mean index, $95 \%$ confidence intervals, individual fish weight, minimum, mean, and maximum length, 5 th, 50 th, and 95 th percentiles of length, number of nonzero tows, and number of fish caught are presented for 1967-2005.

|  | weight/tow |  |  | number/tow |  |  | ind wt | min | length |  |  |  |  | nonzero tows | no fish |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | mean | lower | upper | mean | lower | upper |  |  | 5\% | 50\% | mean | 95\% | max |  |  |
| 1967 | 0.019 | 0.002 | 0.037 | 0.117 | 0.010 | 0.224 | 0.166 | 10 | 18 | 34 | 34.3 | 39 | 42 | 7 | 17 |
| 1968 | 0.003 | -0.001 | 0.008 | 0.023 | -0.019 | 0.065 | 0.135 | 28 | 28 | 28 | 28.9 | 37 | 38 | 2 | 2 |
| 1969 | 0.002 | -0.002 | 0.006 | 0.010 | -0.009 | 0.028 | 0.200 | 38 | 38 | 38 | 38.0 | 38 | 38 | 1 | 1 |
| 1970 | 0.009 | -0.006 | 0.024 | 0.033 | -0.025 | 0.090 | 0.276 | 39 | 39 | 39 | 39.5 | 39 | 40 | 2 | 3 |
| 1971 | 0.001 | -0.001 | 0.004 | 0.006 | -0.005 | 0.016 | 0.250 | 40 | 40 | 40 | 40.5 | 40 | 41 | 1 | 2 |
| 1972 | 0.016 | 0.001 | 0.032 | 0.058 | 0.021 | 0.094 | 0.285 | 12 | 12 | 34 | 34.2 | 40 | 41 | 7 | 8 |
| 1973 | 0.012 | -0.008 | 0.032 | 0.053 | -0.016 | 0.122 | 0.224 | 16 | 16 | 28 | 29.0 | 40 | 41 | 3 | 5 |
| 1974 | 0.012 | -0.002 | 0.026 | 0.079 | -0.014 | 0.171 | 0.156 | 23 | 23 | 34 | 33.8 | 40 | 41 | 4 | 11 |
| 1975 | 0.004 | -0.001 | 0.009 | 0.034 | -0.001 | 0.070 | 0.122 | 25 | 25 | 34 | 33.6 | 38 | 39 | 4 | 8 |
| 1976 | 0.024 | 0.003 | 0.045 | 0.149 | 0.016 | 0.281 | 0.163 | 28 | 28 | 33 | 33.7 | 37 | 40 | 7 | 21 |
| 1977 | 0.020 | -0.002 | 0.043 | 0.087 | -0.011 | 0.185 | 0.231 | 31 | 31 | 33 | 35.2 | 40 | 41 | 5 | 8 |
| 1978 | 0.007 | -0.007 | 0.022 | 0.015 | -0.014 | 0.043 | 0.500 | 39 | 39 | 39 | 39.0 | 39 | 39 | 1 | 1 |
| 1979 | 0.010 | -0.004 | 0.025 | 0.043 | -0.016 | 0.101 | 0.242 | 22 | 22 | 35 | 36.1 | 39 | 40 | 3 | 6 |
| 1980 | 0.090 | 0.042 | 0.138 | 0.312 | 0.120 | 0.505 | 0.287 | 14 | 25 | 38 | 36.6 | 41 | 42 | 10 | 24 |
| 1981 | 0.079 | 0.011 | 0.148 | 0.296 | 0.052 | 0.539 | 0.268 | 27 | 28 | 37 | 37.5 | 41 | 43 | 10 | 45 |
| 1982 | 0.006 | -0.006 | 0.018 | 0.020 | -0.019 | 0.059 | 0.300 | 39 | 39 | 39 | 39.0 | 39 | 39 | 1 | 1 |
| 1983 | 0.001 | -0.001 | 0.003 | 0.010 | -0.010 | 0.030 | 0.100 | 12 | 12 | 12 | 20.7 | 36 | 37 | 1 | 3 |
| 1984 | 0.029 | 0.005 | 0.053 | 0.128 | 0.033 | 0.223 | 0.229 | 13 | 26 | 36 | 35.6 | 39 | 40 | 7 | 16 |
| 1985 | 0.005 | 0.004 | 0.007 | 0.036 | 0.019 | 0.054 | 0.146 | 14 | 14 | 25 | 28.0 | 35 | 36 | 5 | 6 |
| 1986 | 0.003 | 0.001 | 0.004 | 0.009 | 0.005 | 0.013 | 0.300 | 37 | 37 | 37 | 38.2 | 39 | 40 | 3 | 3 |
| 1987 | 0.028 | 0.006 | 0.050 | 0.112 | 0.040 | 0.183 | 0.253 | 11 | 15 | 38 | 32.7 | 41 | 42 | 7 | 10 |
| 1988 | 0.021 | 0.000 | 0.043 | 0.093 | -0.002 | 0.188 | 0.228 | 30 | 30 | 32 | 35.0 | 41 | 42 | 5 | 8 |
| 1989 | 0.018 | -0.005 | 0.041 | 0.046 | -0.012 | 0.105 | 0.378 | 33 | 33 | 33 | 33.5 | 36 | 37 | 3 | 4 |
| 1990 | 0.023 | -0.004 | 0.049 | 0.099 | 0.001 | 0.198 | 0.228 | 32 | 32 | 37 | 37.7 | 41 | 42 | 5 | 10 |
| 1991 | 0.005 | -0.004 | 0.014 | 0.021 | -0.009 | 0.051 | 0.237 | 15 | 15 | 34 | 31.4 | 34 | 35 | 3 | 3 |
| 1992 | 0.035 | 0.006 | 0.064 | 0.170 | 0.033 | 0.308 | 0.203 | 25 | 25 | 35 | 35.3 | 41 | 42 | 9 | 11 |
| 1993 | 0.021 | 0.005 | 0.037 | 0.102 | 0.033 | 0.170 | 0.211 | 25 | 25 | 37 | 35.1 | 40 | 41 | 4 | 8 |
| 1994 | 0.073 | 0.000 | 0.146 | 0.301 | 0.006 | 0.597 | 0.242 | 27 | 27 | 37 | 36.8 | 42 | 43 | 6 | 21 |
| 1995 | 0.039 | -0.005 | 0.084 | 0.174 | -0.009 | 0.358 | 0.227 | 19 | 24 | 35 | 35.1 | 38 | 39 | 7 | 13 |
| 1996 | 0.043 | -0.014 | 0.100 | 0.273 | -0.127 | 0.674 | 0.158 | 7 | 19 | 32 | 31.6 | 38 | 42 | 7 | 21 |
| 1997 | 0.013 | 0.000 | 0.026 | 0.074 | -0.014 | 0.162 | 0.176 | 31 | 31 | 33 | 34.0 | 42 | 43 | 4 | 6 |
| 1998 | 0.050 | -0.008 | 0.108 | 0.208 | -0.042 | 0.458 | 0.241 | 33 | 33 | 37 | 38.1 | 40 | 41 | 7 | 22 |
| 1999 | 0.067 | 0.038 | 0.096 | 0.380 | 0.182 | 0.578 | 0.177 | 12 | 18 | 34 | 32.6 | 41 | 42 | 8 | 46 |
| 2000 | 0.033 | -0.006 | 0.073 | 0.134 | -0.015 | 0.283 | 0.248 | 26 | 30 | 35 | 36.5 | 39 | 40 | 7 | 10 |
| 2001 | 0.121 | -0.007 | 0.249 | 0.472 | -0.016 | 0.961 | 0.257 | 11 | 34 | 39 | 38.6 | 43 | 44 | 10 | 28 |
| 2002 | 0.052 | 0.009 | 0.095 | 0.347 | 0.045 | 0.648 | 0.150 | 8 | 8 | 30 | 28.0 | 40 | 42 | 11 | 29 |
| 2003 | 0.033 | 0.016 | 0.051 | 0.136 | 0.071 | 0.200 | 0.247 | 33 | 33 | 36 | 37.4 | 39 | 41 | 7 | 18 |
| 2004 | 0.048 | 0.003 | 0.092 | 0.231 | 0.030 | 0.432 | 0.206 | 19 | 29 | 35 | 35.5 | 37 | 40 | 8 | 29 |
| 2005 | 0.065 | 0.001 | 0.129 | 0.286 | -0.004 | 0.575 | 0.227 | 30 | 30 | 35 | 36.4 | 39 | 40 | 7 | 24 |

Table B2.20. Abundance and biomass from NEFSC winter surveys for rosette skate for the Georges Bank to Mid-Atlantic region (offshore strata 1-3,5-7,9-11,13-14,16,61-63,65-67,69-71,73-75). The mean index, $95 \%$ confidence intervals, individual fish weight, minimum, mean, and maximum length, 5th, 50th, and 95 th percentiles of length, number of nonzero tows, and number of fish caught are presented for 1992-2006. Stratum 16 not sampled in 1993, 2000, 2002-2006. Strata 13 and 14 not sampled in 2003. Stratum 63 not sampled in 1993. Stratum 14 not sampled in 2005.

|  | weight/tow |  |  | number/tow |  |  | ind wt | length |  |  |  |  |  | nonzero tows | no fish |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | mean | Iower | upper | mean | lower | upper |  | min | 5\% | 50\% | mean | 95\% |  |  |  |
| 1992 | 0.264 | 0.138 | 0.390 | 1.125 | 0.619 | 1.632 | 0.235 | 16 | 27 | 36 | 36.4 | 41 | 45 | 15 | 230 |
| 1993 | 0.149 | 0.048 | 0.251 | 0.663 | 0.197 | 1.130 | 0.225 | 26 | 29 | 36 | 36.7 | 39 | 41 | 9 | 143 |
| 1994 | 0.199 | 0.148 | 0.249 | 0.761 | 0.608 | 0.914 | 0.261 | 16 | 28 | 37 | 36.8 | 40 | 44 | 15 | 162 |
| 1995 | 0.195 | 0.066 | 0.323 | 0.774 | 0.273 | 1.275 | 0.252 | 19 | 32 | 37 | 37.9 | 41 | 42 | 23 | 197 |
| 1996 | 0.324 | 0.121 | 0.526 | 1.410 | 0.443 | 2.376 | 0.230 | 19 | 28 | 36 | 36.3 | 40 | 46 | 23 | 899 |
| 1997 | 0.258 | -0.051 | 0.567 | 1.079 | -0.194 | 2.353 | 0.239 | 13 | 30 | 36 | 36.9 | 40 | 44 | 21 | 238 |
| 1998 | 0.160 | 0.102 | 0.219 | 0.664 | 0.421 | 0.907 | 0.241 | 15 | 30 | 36 | 36.5 | 40 | 45 | 21 | 350 |
| 1999 | 0.271 | 0.043 | 0.500 | 1.151 | 0.082 | 2.220 | 0.236 | 24 | 27 | 37 | 36.6 | 41 | 44 | 25 | 228 |
| 2000 | 0.344 | 0.198 | 0.491 | 1.357 | 0.725 | 1.989 | 0.254 | 8 | 28 | 37 | 37.5 | 43 | 47 | 34 | 740 |
| 2001 | 0.437 | 0.185 | 0.690 | 1.718 | 0.797 | 2.640 | 0.254 | 9 | 24 | 38 | 37.6 | 41 | 46 | 36 | 790 |
| 2002 | 0.723 | 0.140 | 1.307 | 2.655 | 0.603 | 4.708 | 0.272 | 8 | 29 | 38 | 38.3 | 42 | 47 | 34 | 913 |
| 2003 | 0.670 | 0.195 | 1.144 | 2.774 | 0.802 | 4.745 | 0.242 | 8 | 26 | 37 | 36.9 | 41 | 47 | 28 | 1029 |
| 2004 | 0.300 | 0.171 | 0.429 | 1.192 | 0.653 | 1.730 | 0.252 | 16 | 31 | 37 | 37.8 | 41 | 46 | 29 | 784 |
| 2005 | 0.189 | 0.090 | 0.289 | 0.716 | 0.357 | 1.076 | 0.264 | 12 | 30 | 38 | 38.2 | 43 | 45 | 19 | 281 |
| 2006 | 0.437 | 0.209 | 0.665 | 1.738 | 0.821 | 2.654 | 0.251 | 8 | 31 | 37 | 37.7 | 42 | 45 | 28 | 513 |

Table B2.21. Estimates of size at 50\% maturity, length-weight parameters (Wigley et al 2003) and Von Bertalanffy Parameter estimates used to estimate SSB and to calculate Hoenig mortality estimates. Clearnose data, in parentheses, refers to diak width.

| Species (Study) | L50 | $\ln (\mathrm{a})$ | b | Linf | K | t0 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Winter (Frisk 2004) | 76 | -13.1531 | 3.3199 | 122.1 | 0.07 | -2.06 |
| Little (Frisk 2004) | 44 | -12.4462 | 3.128 | 56.1 | 0.19 | -1.17 |
| Barndoor (Gedamke et al. 2005) | 116 | -13.3224 | 3.2919 | 166.3 | 0.14 | -1.2912 |
| Thorny (Sulikowski 2005, 2006) | 88 | -12.088 | 3.1197 | 124.0 | 0.12 | -0.35 |
| Smooth (Sosebee 2005) | 50 | -13.0139 | 3.1812 |  |  |  |
| Clearnose(Gelsleichter 1998; Sosebee 2005) | 66 | -13.8683 | $3.423594 .3(61.8)$ | 0.17 | -0.88 |  |
| Rosette (Sosebee 2005) | 34 | -12.5504 | 3.0718 |  |  |  |

Table B2.22 Estimates of spawning stock biomass indices from NEFSC surveys using sizes at $50 \%$ maturity as knife-edge cutpoints.

| Winter |  | Little | Barndoor |  | Thorny | Smooth | Clearnose Rosette |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 1963 |  |  | 0.796 | 3.934 | 0.202 |  |  |
| 1964 |  |  | 0.227 | 2.799 | 0.091 |  |  |
| 1965 |  |  | 0.135 | 2.848 | 0.297 |  |  |
| 1966 |  |  | 0.000 | 4.673 | 0.218 |  | 0.022 |
| 1967 | 0.553 |  | 0.063 | 1.411 | 0.126 |  | 0.001 |
| 1968 | 0.338 |  | 0.073 | 2.857 | 0.229 |  | 0.002 |
| 1969 | 0.183 |  | 0.000 | 3.668 | 0.190 |  | 0.009 |
| 1970 | 0.534 |  | 0.060 | 5.155 | 0.152 |  | 0.002 |
| 1971 | 0.151 |  | 0.047 | 3.921 | 0.134 |  | 0.010 |
| 1972 | 0.464 |  | 0.077 | 2.593 | 0.244 |  | 0.001 |
| 1973 | 0.892 |  | 0.000 | 2.987 | 0.189 |  | 0.013 |
| 1974 | 0.377 |  | 0.000 | 1.368 | 0.080 |  | 0.003 |
| 1975 | 0.327 |  | 0.000 | 1.344 | 0.039 | 0.003 | 0.005 |
| 1976 | 1.117 |  | 0.000 | 0.943 | 0.015 | 0.019 | 0.020 |
| 1977 | 1.863 |  | 0.000 | 1.450 | 0.201 | 0.076 | 0.015 |
| 1978 | 3.008 |  | 0.000 | 1.514 | 0.288 | 0.007 | 0.004 |
| 1979 | 3.400 |  | 0.000 | 1.569 | 0.112 | 0.073 | 0.009 |
| 1980 | 3.663 |  | 0.000 | 1.972 | 0.217 | 0.166 | 0.070 |
| 1981 | 3.513 |  | 0.000 | 1.312 | 0.079 | 0.016 | 0.070 |
| 1982 | 4.203 | 2.744 | 0.000 | 0.261 | 0.035 | 0.038 | 0.005 |
| 1983 | 7.598 | 4.058 | 0.000 | 1.065 | 0.073 | 0.006 | 0.001 |
| 1984 | 7.253 | 2.655 | 0.000 | 1.480 | 0.095 | 0.041 | 0.024 |
| 1985 | 8.514 | 4.184 | 0.000 | 1.077 | 0.169 | 0.069 | 0.003 |
| 1986 | 12.279 | 1.599 | 0.000 | 0.653 | 0.152 | 0.030 | 0.002 |
| 1987 | 7.768 | 2.168 | 0.000 | 0.209 | 0.062 | 0.085 | 0.021 |
| 1988 | 5.594 | 2.936 | 0.000 | 0.521 | 0.207 | 0.072 | 0.011 |
| 1989 | 3.753 | 2.832 | 0.000 | 0.709 | 0.073 | 0.028 | 0.002 |
| 1990 | 6.129 | 2.983 | 0.000 | 0.790 | 0.122 | 0.072 | 0.023 |
| 1991 | 3.499 | 2.854 | 0.000 | 0.734 | 0.116 | 0.341 | 0.003 |
| 1992 | 2.083 | 2.384 | 0.000 | 0.292 | 0.079 | 0.080 | 0.033 |
| 1993 | 1.012 | 3.875 | 0.134 | 0.700 | 0.146 | 0.110 | 0.018 |
| 1994 | 0.841 | 1.742 | 0.000 | 0.434 | 0.072 | 0.184 | 0.063 |
| 1995 | 0.536 | 1.706 | 0.000 | 0.189 | 0.081 | 0.097 | 0.033 |
| 1996 | 0.793 | 4.551 | 0.000 | 0.318 | 0.128 | 0.083 | 0.029 |
| 1997 | 0.664 | 1.601 | 0.052 | 0.333 | 0.167 | 0.269 | 0.009 |
| 1998 | 1.576 | 3.634 | 0.062 | 0.319 | 0.016 | 0.234 | 0.051 |
| 1999 | 1.331 | 5.078 | 0.118 | 0.145 | 0.062 | 0.442 | 0.055 |
| 2000 | 1.753 | 4.424 | 0.048 | 0.420 | 0.102 | 0.371 | 0.028 |
| 2001 | 1.397 | 4.783 | 0.250 | 0.066 | 0.226 | 0.376 | 0.129 |
| 2002 | 3.154 | 4.858 | 0.366 | 0.196 | 0.094 | 0.261 | 0.034 |
| 2003 | 1.912 | 4.401 | 0.161 | 0.233 | 0.106 | 0.353 | 0.032 |
| 2004 | 2.222 | 4.340 | 0.773 | 0.365 | 0.146 | 0.259 | 0.043 |
| 2005 | 1.005 | 2.455 | 0.285 | 0.047 | 0.082 | 0.253 | 0.057 |
| 2006 |  | 2.472 |  |  |  |  |  |
|  |  |  |  |  |  |  |  |

Table B.2.23.
(EDITOR'S NOTE: BASED ON THE REVIEWER'S COMMENTS, THIS TABLE WAS NOT INCLUDED IN THIS REPORT. THE TABLE HAD ESTIMATES OF FISHING MORTALITY RATE.)

Table B3.1. Current estimates of biomass-based reference points for skates. The estimates for barndoor skate are an average of 1963-1966 biomass estimates.

|  | $75^{\text {th }}$ percentile through $1998 / 1999$ |  |
| :--- | :--- | :--- |
|  | Bmsy | Bthreshold |
| Winter | 6.46 | 3.23 |
| Little | 6.54 | 3.27 |
| Barndoor | 1.62 | 0.81 |
| Thorny | 4.41 | 2.20 |
| Smooth | 0.31 | 0.16 |
| Clearnose | 0.56 | 0.28 |
| Rosette | 0.029 | 0.015 |

Tables B.3.2 - B.3.24.
(EDITOR'S NOTE: BASED ON THE REVIEWER'S COMMENTS, THESE TABLES WERE NOT INCLUDED IN THIS REPORT. THE TABLES HAD CALCULATIONS FOR ALTERNATIVE REFERENCE POINTS.)

Table B4.1. Fishing mortality overfishing definition for skates based on the average coefficient of variation in the survey. The percentages are percent change from one three-year moving average to the next.

| Winter |  | Little |  | Barndoor |  | Thorny |  |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
|  | $-20 \%$ | $-20 \%$ | $-30 \%$ | $-20 \%$ | $-30 \%$ | Clearnose Rosette |  |
|  | -8.8 | -7.6 | -3.8 | -17.6 | -0.4 | 4.5 | $-60 \%$ |
| 1992 | -8.8 | 15.6 | 180.7 | -1.1 | 6.7 | 5.6 | -2.0 |
| 1993 | -3.9 | -12.6 | 2.0 | -2.9 | -13.0 | 0.9 | 110.9 |
| 1994 | -25.5 | -14.8 | 61.3 | -4.3 | 13.8 | -0.8 | 3.8 |
| 1995 | -21.0 | -3.3 |  |  |  |  |  |
| 1996 | 6.2 | 0.4 | -34.3 | -21.4 | -9.8 | -3.6 | 16.4 |
| 1997 | 5.3 | -6.5 | 37.3 | -21.2 | 28.6 | -19.1 | -38.4 |
| 1998 | 26.3 | 35.0 | -8.6 | -5.5 | -26.9 | 57.5 | 11.1 |
| 1999 | 33.2 | 13.5 | 109.2 | -14.5 | -24.2 | 28.8 | 22.5 |
| 2000 | 17.0 | 29.2 | 37.1 | -0.9 | -23.6 | 15.0 | 15.3 |
| 2001 | 1.0 | -2.4 | 66.0 | -16.1 | 102.3 | 15.4 | 47.1 |
| 2002 | 3.8 | -13.9 | 42.5 | -2.6 | 8.1 | -4.4 | -6.9 |
| 2003 | -7.2 | -9.6 | 16.5 | -5.6 | 6.5 | -10.5 | 0.2 |
| 2004 | 1.1 | 1.9 | 40.7 | 25.0 | -12.4 | -28.6 | -35.4 |
| 2005 | -22.9 | -15.9 | 9.8 | -11.2 | 3.7 | -16.2 | 9.7 |
| 2006 |  | -18.7 |  |  |  |  |  |

Table B6.1. The size class, temporal, and spatial scheme for each species of skate analyzed for food habits and consumptive demand. $S=$ small, $I=$ immature, $M=$ medium if small and large used; $M=$ mature if immature used, $L=$ large. All size class cutoffs are in cm . * small winter skates were combined with immature little skates to account for potential identification concerns.

|  | Barndoor Skate | Clearnose Skate | Little Skate | Rosette Skate | Smooth Skate | Thorny Skate | Winter Skate |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| SVSPP <br> Code | 022 | 024 | 026 | 025 | 027 | 028 | 023 |
| Survey <br> Strata Set | $\begin{aligned} & \hline 01010- \\ & 01300, \\ & 01330- \\ & 01400, \\ & 01351 \end{aligned}$ | $\begin{aligned} & \hline 03150- \\ & 03440, \\ & 01610- \\ & 01760 \end{aligned}$ | $\begin{aligned} & \hline 01010- \\ & 01300, \\ & 01330- \\ & 01400, \\ & 01351, \\ & 01610- \\ & 01760, \\ & 03010- \\ & 03660 \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 01610- \\ & 01760 \end{aligned}$ | $\begin{aligned} & \text { 01010- } \\ & 01300, \\ & 01330- \\ & 01400, \\ & 01351 \end{aligned}$ | $\begin{aligned} & \hline 01010- \\ & 01300, \\ & 01330- \\ & 01400, \\ & 01351 \end{aligned}$ | $\begin{aligned} & \hline 01010- \\ & 01300, \\ & 01330- \\ & 01400, \\ & 01351, \\ & 01610- \\ & 01760 \end{aligned}$ |
| Temporal Resolution | $\begin{aligned} & \text { 2000-2005, } \\ & \text { annual } \end{aligned}$ | $\begin{aligned} & \text { 1977-2005, } \\ & 5 \text { year } \\ & \text { block } \end{aligned}$ | $\begin{aligned} & \text { 1982-2005, } \\ & \text { annual } \end{aligned}$ | $\begin{aligned} & 1999- \\ & 2005, \end{aligned}$ annual | $\begin{aligned} & 1977- \\ & 2005, \\ & 5 \text { year } \\ & \text { block } \\ & \hline \end{aligned}$ | $\begin{aligned} & \hline 1977- \\ & 2005, \\ & 5 \text { year } \\ & \text { block } \\ & \hline \end{aligned}$ | $\begin{aligned} & \text { 1977- } \\ & 2005, \\ & \text { annual } \end{aligned}$ |
| Size Class: |  |  |  |  |  |  |  |
| S or I | <80 | $>60$ | <30* | <30 | <30 | <30 | <30* |
| M | >80 | <60 | $>30$ | $>30$ | $>30$ | 30-60 | 30-60 |
| L |  |  |  |  |  | $>60$ | $>60$ |

Table B6.2. Diet composition of Winter Skate. All values are expressed as whole numbers rather than percentages. Relmsw $=$

Table B6.3. Diet composition of Little Skate. All values are expressed as whole numbers rather than percentages. Relmsw = relative mean stomach weight, on average for the size class and time period given. $A R=$ animal remains, a well-digested, highly

Table B6.4. Diet composition of Barndoor Skate. All values are expressed as whole numbers rather than percentages. Relmsw $=$ relative mean stomach weight, on average for the size class and time period given. $\mathrm{AR}=$ animal remains, a welldigested, highly unresolved category.


Table B6.5. Diet composition of Thorny Skate. All values are expressed as whole numbers rather than percentages. Relmsw = relative mean stomach weight, on average for the size class and time period given. AR = animal remains, a welldigested, highly unresolved category.

Table B6.6. Diet composition of Smooth Skate. All values are expressed as whole numbers rather than percentages. Relmsw $=$ relative mean stomach weight, on average for the size class and time period given. $\mathrm{AR}=$ animal remains, a welldigested, highly unresolved category.

| Average of relmsw |  |  |  |  | Cancer Crabs | Crangon sp. | Misc. Crustac $\epsilon$ Other Decapor Other Crabs |  |  | Decapod Shrir Euphasiids |  | GADFAM | MERBIL | MYSIDA | Pandalid Shrin OTHFIS |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| size | yr5block | AMPHIP | ANNELI | AR |  |  |  |  |  |  |  |  |  |  |  |  |
|  | 1980 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 1 | 1985 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 100 | 0 | 0 | 0 | 0 | 0 |
|  | 1990 | 25.5072 | 0 | 17.3913 | 0 | 0 | 11.014 | 0 | 0 | 0 | 38.8406 | 0 | 0 | 5.7971 | 0 | 0 |
|  | 1995 | 0 | 0 | 14.9425 | 0 | 0 | 21.7241 | 17.2414 | 0 | 0 | 0 | 40.3448 | 0 | 5.7471 | 0 | 0 |
|  | 2000 | 33.7185 | - 2.395 | 38.1688 | 0 | 3.7815 | 4.2017 | 0 | 0 | 0 | 6.8277 | 0 | 0 | 2.1008 | 5.8298 | 0 |
|  | 2005 | 9.9932 | - 2.435 | 44.1366 | 0 | 1.2784 | 19.298 | 0 | 1.8263 | 5.4788 | 7.0768 | 0 | 0 | 2.1306 | 4.8853 | 0 |
| ITotal |  | 13.84378 | 0.966 | 22.92784 | 0 | 1.01198 | 11.24756 | 3.44828 | 0.36526 | 1.09576 | 30.54902 | 8.06896 | 0 | 3.15512 | 2.14302 | 0 |
| M | 1980 | 0 | 0 | 0.974 | 3.8277 | 9.2047 | 7.288 | 16.192 | 0 | 16.6234 | 0 | 0 | 0 | 0.1065 | 42.4603 | 0.0572 |
|  | 1985 | 1.0017 | 11.5192 | 10.0167 | 0 | 4.0067 | 4.5075 | 0 | 0 | 0 | 34.5576 | 0 | 0 | 0 | 30.384 | 4.0067 |
|  | 1990 | 1.3291 | - 0.217 | 9.1141 | 0 | 1.9169 | 9.8085 | 0.2713 | 1.4955 | 3.0923 | 49.0786 | 0.7053 | 2.4955 | 0 | 4.8102 | 11.8357 |
|  | 1995 | 0.3283 | - 1.4304 | 6.8296 | 0 | 0.4072 | 13.2809 | 0 | 2.794 | 1.5244 | 7.963 | 0 | 22.8866 | 0.0693 | 26.0084 | 8.634 |
|  | 2000 | 2.2124 | - 0.6341 | 10.6603 | 2.4559 | 3.1884 | 3.1664 | 0 | 8.62 | 3.2415 | 5.4317 | 0 | 11.5312 | 0.2531 | 34.2183 | 4.923 |
|  | 2005 | 1.3192 | - 1.5846 | 13.3123 | 2.3746 | 1.9784 | 27.6493 | 0.0916 | 2.6112 | 3.7497 | 5.5229 | 0.0458 | 1.0535 | 0.3893 | 26.5792 | 5.5011 |
| M Total |  | 1.031783333 | 2.564216667 | 8.4845 | 1.443033333 | 3.450383333 | 10.9501 | 2.75915 | 2.586783333 | 4.705216667 | 17.0923 | 0.125183333 | 6.3278 | 0.136366667 | 27.41006667 | 5.826283333 |

[^25]| Average of relmsw |  | analcat |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| size | yr5block | AMMFAM | ANNELI | BIVALV |  | Cancer Crabs | Cephalopods | Misc. Crustac $\in$ DECAPO |  | Other Crabs | OPHFA2 | Other Fish | SERFA2 | SOLFAM |
| 1 | 1980 | 0 | 0.497 | 1.7597 | 0.3006 | 30.6196 | 3.2767 | 0.2182 | 2.0069 | 34.9254 | 0 | 12.5642 | 0 | 0 |
|  | 1985 | 0.80335 | 1.06525\| | 2.02465 \| | 1.0842\| | 22.74435 \| | 1.63835 | \| 8.00895 | 1.00345 | 33.2557 | 0 | 11.34335 | 0 | 0 |
|  | 1990 | 1.6067 | 1.6335 | 2.2896 | 1.8678 | 14.8691 | 0 | 15.7997 | 0 | 31.586 | 0 | 10.1225 | 0 | 0 |
|  | 1995 | 0 | 5.0256 | 0.6391 | 0 | 32.2353 | 0 | 12.5359 | 0 | 2.2386 | 33.5783 | 1.8737 | 6.0441 | 0 |
|  | 2000 | 0 | 5.8414 | 5.973 | 4.1183 | 9.3624 | 6.3996 | 5.5937 | 0.1422 | 32.9217 | 0.7901 | 12.5543 | 0.2844 | 0 |
|  | 2005 | 0 | 0 | 0 | 5.5127 | 14.8842 | 23.1533 | 0.3308 | 0 | 18.7431 | 4.9614 | 13.5612 | 0 | 0 |
| 1 Total |  | 0.32134 | 2.5995 | 2.13228 | 2.35988 | 20.39412 | 6.56592 | 6.89566 | 0.42982 | 24.08296 | 7.86596 | 10.13518 | 1.2657 | 0 |
| M | 1980 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 100 | 0 | 0 |
|  | 1985 | 0 | 0 | 0 | 16.667 | 33.3333 | 0 | 0 | 0 | 0 | 0 | 50 | 0 | 0 |
|  | 1990 | 5.9811 | 0.2723 | 0 | 0 | 27.3371 | 3.104 | 3.7212 | 0 | 26.9831 | 10.8913 | 7.5876 | 0 | 0 |
|  | 1995 | 0 | 0.4189 | 0.3491 | 0 | 13.5727 | 0 | 0 | 0 | 9.6008 | 5.1204 | 66.7183 | 0 | 0 |
|  | 2000 | 0.9146 | 0.3593 | 0.4717 | 0.7186 | 6.2759 | 0 | 0.8493 | 0 | 21.273 | 16.332 | 35.1313 | 1.5056 | 3.6487 |
|  | 2005 | 0 | 0.6081 | 0.827 | 1.4975 | 15.4281 | 55.3925 | 0.2737 | 0 | 8.1955 | 0 | 5.4731 | 0 | 0.608116667 |
| M Total |  | 1.149283333 | 0.276433333 | 0.274633333 | 3.147183333 | 15.99118333 | 9.749416667 | 0.807366667 | 0 | 11.00873333 | 5.390616667 | 44.15171667 | 0.250933333 | 0.3317 |

Table B6.8. Diet composition of Rosette Skate. All values are expressed as whole numbers rather than percentages. Relmsw $=$ relative mean stomach weight, on average for the size class and time period given. AR = animal remains, a welldigested, highly unresolved category.


Table B6.9. Comparison of total skate consumptive removal of major skate prey relative to standing biomass and production estimates of those prey (from Link et al. 2006); these estimates are integrated across the entire ecosystem for the period 1996-2000. All values are in MT. C = consumptive removal of the prey by skates, as averaged during the period 2000-2006; $\mathrm{B}=$ biomass, $\mathrm{P}=$ production.

|  | C | B | P |
| :--- | :---: | :---: | :---: |
| Polychaetes | $3.23 \times 10^{4}$ | $4.30 \times 10^{6}$ | $1.08 \times 10^{7}$ |
| Molluscs | $3.24 \times 10^{4}$ | $2.80 \times 10^{6}$ | $9.27 \times 10^{6}$ |
| Cephalopods | $5.91 \times 10^{3}$ | $3.13 \times 10^{5}$ | $3.03 \times 10^{5}$ |
|  <br> Mackerel | $5.09 \times 10^{3}$ | $2.04 \times 10^{6}$ | $7.55 \times 10^{5}$ |
| Euphasiids and <br> similar crustaceans | $2.12 \times 10^{3}$ | $1.89 \times 10^{6}$ | $2.69 \times 10^{7}$ |

Table B6.10. Comparison of fishery landings of major skate prey with total skate consumptive removal of major commercially targeted skate prey across different assumed gear efficiencies used to estimate skate abundance. All values represent an average from 2000-2005 and are in MT. The C/L ratio contrasts the consumption to the fishery landings as a unitless scalar; values $>1$ indicate more of the prey is consumed by skates than is removed by the fishery..

|  | Fishery <br> Landings | $100 \%$ <br> Efficiency | $50 \%$ | $25 \%$ | $10 \%$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Illex and <br> Loligo | $2.53 \times 10^{4}$ | $5.91 \times 10^{3}$ | $1.18 \times 10^{4}$ | $2.36 \times 10^{4}$ | $5.91 \times 10^{4}$ |
| C/L ratio | - | 0.23 | 0.47 | 0.93 | 2.33 |
| Silver Hake | $9.37 \times 10^{3}$ | $2.15 \times 10^{3}$ | $4.30 \times 10^{3}$ | $8.59 \times 10^{3}$ | $2.15 \times 10^{4}$ |
| C/L ratio | - | 0.23 | 0.46 | 0.92 | 2.29 |
| Red Hake | $9.95 \times 10^{2}$ | $1.15 \times 10^{3}$ | $2.29 \times 10^{3}$ | $4.58 \times 10^{3}$ | $1.15 \times 10^{4}$ |
| C/L ratio | - | 1.15 | 2.30 | 4.60 | 11.51 |
| Herrings | $1.16 \times 10^{5}$ | $5.09 \times 10^{3}$ | $1.02 \times 10^{4}$ | $2.04 \times 10^{4}$ | $5.09 \times 10^{4}$ |
| C/L ratio | - | 0.04 | 0.09 | 0.18 | 0.44 |

## SKATE FIGURES



Figure B1.1. Total landings of skates in NAFO subareas 5 and 6.


Figure B1.2. Distribution of winter skates from the Observer Program, 1989-2005.


Figure B1.3. Distribution of little skates from the Observer Program, 1989-2005.


Barndoor Skates from 1989-2005 NEFSC Observer Data

Figure B1.4. Distribution of barndoor skates from the Observer Program, 1989-2005.


Thorny Skates from 1989-2005 NEFSC Observer Data

Figure B1.5. Distribution of thorny skates from the Observer Program, 1989-2005.


Figure B1.6. Distribution of smooth skates from the Observer Program, 1989-2005.


Clearnose Skates from 1989-2005 NEFSC Observer Data

Figure B1.7. Distribution of clearnose skates from the Observer Program, 1989-2005.


Rosette Skates from 1989-2005 NEFSC Observer Data

Figure B1.8. Distribution of rosette skates from the Observer Program, 1989-2005.

Winter Skate
Observer Length Data


Figure B1.9. Winter skate length composition from the NEFSC observer program 1996-2005.

Little Skate
Observer Length Data


Figure B1.10. Little skate length composition from the NEFSC observer program 1996-2005.

Barndoor Skate
Observer Length Data


Figure B1.11. Barndoor skate length composition from the NEFSC observer program 1996-2005.

Thorny Skate
Observer Length Data


Figure B1.12. Thorny skate length composition from the NEFSC observer program 1996-2005.

## Smooth Skate

Observer Length Data


Figure B1.13. Smooth skate length composition from the NEFSC observer program 1996-2005.

## Clearnose Skate

## Observer Length Data



Figure B1.14. Clearnose skate length composition from the NEFSC observer program 1996-2005.

## Rosette Skate

Observer Length Data


Figure B1.15. Rosette skate length composition from the NEFSC observer program 1996-2005.


Figure B2.1. Map of offshore strata sampled in the NEFSC spring, autumn, and winter surveys.


Figure B2.2. Map of inshore strata sampled in the NEFSC spring and autumn surveys in the Gulf of Maine.


Figure B2.3. Map of inshore strata sampled in the NEFSC spring and autumn surveys in Southern New England.


Figure B2.4. Map of inshore strata sampled in the NEFSC spring and autumn surveys in the Mid-Atlantic.


Figure B2.5. Species composition of skates from the spring survey. Panel A shows the composition of all species, panel B shows the composition of large species ( $>100 \mathrm{~cm}$ maximum length), and panel C shows the composition of the small species (maximum length $<100 \mathrm{~cm}$ ).




Figure B2.6. Distribution of winter skate from the spring and autumn NEFSC surveys from 1998-2006.


Figure B2.7. Distribution of winter skate from the NEFSC winter surveys from 2000-2006.


Figure B2.8. Distribution of winter skate from the NEFSC scallop surveys from 1985-1996.


Figure B2.9. Distribution of winter skate from the NEFSC scallop surveys from 1997-2006.

## Winter Skate <br> GOM-MA Offshore Only



Figure B2.10. Abundance and biomass of winter skate from the NESFC spring (circles) and autumn (squares) bottom trawl surveys from 1967-2006 in the Gulf of Maine to Mid-Atlantic offshore region.

# Winter Skate <br> GOM-MA Offshore Only - Spring Survey 



Figure B2.11. Abundance and biomass of winter skate from the NESFC spring bottom trawl surveys from 1968-2006 in the Gulf of Maine to Mid-Atlantic offshore region. The circles represent the original stratified mean, the squares represent the mean combining strata for bootstrapping, and the triangles represent the bootstrapped mean.

## Winter Skate - Spring Survey <br> GOM-MA Offshore Only



Figure B2.12. Bootstrapped abundance and biomass of winter skate from the NESFC spring bottom trawl survey in the Gulf of Maine to Mid-Atlantic region, offshore strata only. Mean index in solid squares, $95 \%$ confidence interval in open squares.

## Winter Skate GOM-MA Offshore Only - Autumn Survey



Figure B2.13. Abundance and biomass of winter skate from the NESFC autumn bottom trawl surveys from 1967-2005 in the Gulf of Maine to Mid-Atlantic offshore region. The circles represent the original stratified mean, the squares represent the mean combining strata for bootstrapping, and the triangles represent the bootstrapped mean.

## Winter Skate - Autumn Survey GOM-MA Offshore Only



Figure B2.14. Bootstrapped abundance and biomass of winter skate from the NESFC autumn bottom trawl survey in the Gulf of Maine to Mid-Atlantic region, offshore strata only. Mean index in solid squares, $95 \%$ confidence interval in open squares.

## Winter Skate <br> Percentiles of Length Composition



Figure B2.15. Percentiles of length composition (5, 50, and 95) of winter skate from the NESFC spring and autumn bottom trawl surveys from 1967-2006 in the Gulf of Maine to Mid-Atlantic offshore region.


Figure B2.16. Winter skate length composition from the NEFSC spring and autumn trawl surveys in the Gulf of Maine to Mid-Atlantic offshore regions, 1967-1972.


Figure B2.17. Winter skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Gulf of Maine to Mid-Atlantic offshore regions, 1973-1982.


Figure B2.18. Winter skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Gulf of Maine to Mid-Atlantic offshore regions, 1983-1992.


Figure B2.19. Winter skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Gulf of Maine to Mid-Atlantic offshore regions, 1993-2002.


Figure B2.20. Winter skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Gulf of Maine to Mid-Atlantic offshore regions, 2003-2006.

Winter Skate
Winter Survey


Figure B2.21. Abundance and biomass of winter skate from the NESFC winter bottom trawl surveys from 1992-2006. The circles represent the original stratified mean, the squares represent the mean combining strata for bootstrapping, and the triangles represent the bootstrapped mean.

## Winter Skate <br> Winter Survey



Figure B2.22. Bootstrapped abundance and biomass of winter skate from the NESFC winter bottom trawl survey. Mean index in solid squares, $95 \%$ confidence interval in open squares.

# Winter Skate Scallop Survey 



Figure B2.23. Abundance and biomass of winter skate from the NESFC scallop surveys from 1985-2006. The circles represent the original stratified mean, the squares represent the mean combining strata for bootstrapping, and the triangles represent the bootstrapped mean.


Figure B2.24. Bootstrapped abundance and biomass of winter skate from the NESFC scallop survey. Mean index in solid squares, $95 \%$ confidence interval in open squares.

## Winter Skate - Massachusetts Trawl Survey



Figure B2.25. Abundance and biomass of winter skate from the Massachusetts spring and autumn finfish bottom trawl survey in state waters (strata 11-36).

## Winter Skate - CTDEP Finfish Survey



Figure B2.26. Abundance and biomass of winter skate from the CTDEP spring and autumn finfish bottom trawl survey in Connecticut state waters, 1984-2006.


Little Skates from 1998-2002 NEFSC Spring Surveys


Little Skates from 1998-2002 NEFSC Fall Surveys


Figure B2.27. Distribution of little skate from the spring and autumn NEFSC surveys from 1998-2006.


Figure B2.28. Distribution of little skate from the NEFSC winter surveys from 2000-2006.


Little Skates from 1991-1993 NEFSC Scallop Surveys


Little Skates from 1994-1996 NEFSC Scallop Surveys

Figure B2.29. Distribution of little skate from the NEFSC scallop surveys from 1985-1996.


Figure B2.30. Distribution of little skate from the NEFSC scallop surveys from 1997-2006.

## Little Skate <br> GOM-MA All Strata



Figure B2.31. Abundance and biomass of little skate from the NESFC spring (circles) and autumn (squares) bottom trawl surveys from 1975-2006 in the Gulf of Maine to Mid-Atlantic (all strata).

## Little Skate GOM-MA All Strata - Spring Survey



Figure B2.32. Abundance and biomass of little skate from the NESFC spring bottom trawl surveys from 1979-2006 in the Gulf of Maine to Mid-Atlantic (all strata). The circles represent the original stratified mean, the squares represent the mean combining strata for bootstrapping, and the triangles represent the bootstrapped mean.

## Little Skate - Spring Survey GOM-MA All Strata



Figure B2.33. Bootstrapped abundance and biomass of little skate from the NESFC spring bottom trawl survey in the Gulf of Maine to Mid-Atlantic region (all strata). Mean index in solid squares, $95 \%$ confidence interval in open squares.

## Little Skate <br> GOM-MA All Strata - Autumn Survey



Figure B2.34. Abundance and biomass of little skate from the NESFC autumn bottom trawl surveys from 1979-2005 in the Gulf of Maine to Mid-Atlantic (all strata). The circles represent the original stratified mean, the squares represent the mean combining strata for bootstrapping, and the triangles represent the bootstrapped mean.

## Little Skate - Autumn Survey GOM-MA All Strata



Figure B2.35. Bootstrapped abundance and biomass of little skate from the NESFC autumn bottom trawl survey in the Gulf of Maine to Mid-Atlantic region (all strata). Mean index in solid squares, $95 \%$ confidence interval in open squares.

## Little Skate: GOM-MA All strata Percentiles of Length Composition



Figure B2.36. Percentiles of length composition (5, 50, and 95) of little skate from the NESFC spring and autumn bottom trawl surveys from 1975-2006 in the Gulf of Maine to Mid-Atlantic region (all strata).


Figure B2.37. Little skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Gulf of Maine to Mid-Atlantic (all strata), 1975-1982.


Figure B2.38. Little skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Gulf of Maine to Mid-Atlantic (all strata), 1983-1992.


Figure B2.39. Little skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Gulf of Maine to Mid-Atlantic (all strata), 1993-2002.


Figure B2.40. Little skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Gulf of Maine to Mid-Atlantic (all strata), 2003-2006.

## Little Skate <br> Winter Survey



Year
Figure B2.41. Abundance and biomass of little skate from the NESFC winter bottom trawl surveys from 1992-2006. The circles represent the original stratified mean, the squares represent the mean combining strata for bootstrapping, and the triangles represent the bootstrapped mean.


Figure B2.42. Bootstrapped abundance and biomass of little skate from the NESFC winter bottom trawl survey. Mean index in solid squares, $95 \%$ confidence interval in open squares.


Figure B2.43. Abundance and biomass of little skate from the NESFC scallop surveys from 1985-2006. The circles represent the original stratified mean, the squares represent the mean combining strata for bootstrapping, and the triangles represent the bootstrapped mean.

## Little Skate <br> Scallop Survey



Figure B2.44. Bootstrapped abundance and biomass of little skate from the NESFC scallop survey. Mean index in solid squares, $95 \%$ confidence interval in open squares.

## Little Skate - Massachusetts Trawl Survey



Figure 2.45. Abundance and biomass of little skate from the Massachusetts spring and autumn finfish bottom trawl survey in state waters (strata 11-36).

## Little Skate - CTDEP Finfish Survey



Figure B2.46. Abundance and biomass of little skate from the CTDEP spring and autumn finfish bottom trawl survey in Connecticut state waters, 1984-2006.


Figure B2.47. Distribution of barndoor skate from the spring and autumn NEFSC surveys from 2000-2006.


Figure B2.48. Distribution of barndoor skate from the winter NEFSC surveys from 2000-2006 and the NEFSC scallop surveys from 1991-2006.

## Barndoor Skate GOM-SNE Offshore Only



Figure B2.49. Abundance and biomass of barndoor skate from the NESFC spring (circles) and autumn (squares) bottom trawl surveys from 1963-2006 in the Gulf of Maine to Southern New England offshore region.

## Barndoor Skate GOM-SNE Offshore Only - Spring Survey



Figure B2.50. Abundance and biomass of barndoor skate from the NESFC spring bottom trawl surveys from 1968-2006 in the Gulf of Maine to Southern New England offshore region. The circles represent the original stratified mean, the squares represent the mean combining strata for bootstrapping, and the triangles represent the bootstrapped mean.

## Barndoor Skate - Spring Survey <br> GOM-SNE Offshore Only



Figure B2.51. Bootstrapped abundance and biomass of barndoor skate from the NESFC spring bottom trawl survey in the Gulf of Maine to Southern New England offshore region. Mean index in solid squares, $95 \%$ confidence interval in open squares.

## Barndoor Skate GOM-SNE Offshore Only - Autumn Survey



Figure B2.52. Abundance and biomass of barndoor skate from the NESFC autumn bottom trawl surveys from 1963-2006 in the Gulf of Maine to Southern New England offshore region. The circles represent the original stratified mean, the squares represent the mean combining strata for bootstrapping, and the triangles represent the bootstrapped mean.

## Barndoor Skate - Autumn Survey GOM-SNE Offshore Only



Figure B2.53. Bootstrapped abundance and biomass of barndoor skate from the NESFC autumn bottom trawl survey in the Gulf of Maine to Southern New England offshore region. Mean index in solid squares, $95 \%$ confidence interval in open squares.

## Barndoor Skate Percentiles of Length Composition



Figure B2.54. Percentiles of length composition (5,50, and 95) of barndoor skate from the NESFC spring and autumn bottom trawl surveys from 1963-2006 in the Gulf of Maine to Southern New England offshore region.


Figure B2.55. Barndoor skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Gulf of Maine to Southern New England offshore region, 1963-1972.


Figure B2.56. Barndoor skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Gulf of Maine to Southern New England offshore region, 1973-1982.


Figure B2.57. Barndoor skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Gulf of Maine to Southern New England offshore region, 1983-1992.


Figure B2.58. Barndoor skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Gulf of Maine to Southern New England offshore region, 1993-2002.


Figure B2.59. Barndoor skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Gulf of Maine to Southern New England offshore region, 2003-2006.

## Barndoor Skate Winter Survey



Figure B2.60. Abundance and biomass of barndoor skate from the NESFC winter bottom trawl surveys from 1993-2006. The circles represent the original stratified mean, the squares represent the mean combining strata for bootstrapping, and the triangles represent the bootstrapped mean.


Figure B2.61. Bootstrapped abundance and biomass of barndoor skate from the NESFC winter bottom trawl survey. Mean index in solid squares, $95 \%$ confidence interval in open squares.

## Winter Survey




LENGTH (cm)

Figure B2.62. Barndoor skate length composition from the NEFSC winter flatfish surveys, 1993-2006.

Barndoor Skate
Scallop Survey


Figure B2.63. Abundance and biomass of barndoor skate from the NESFC scallop surveys from 1991-2006. The circles represent the original stratified mean, the squares represent the mean combining strata for bootstrapping, and the triangles represent the bootstrapped mean.

## Barndoor Skate - Scallop Survey <br> GOM-SNE Offshore Only



Figure B2.64. Bootstrapped abundance and biomass of barndoor skate from the NESFC scallop survey. Mean index in solid squares, $95 \%$ confidence interval in open squares.


Figure B2.65. Distribution of thorny skate from the spring and autumn NEFSC surveys from 1998-2006.



Figure B2.66. Distribution of thorny skate from the NEFSC scallop and shrimp surveys from 1985-2006.

## Thorny Skate GOM-SNE Offshore Only



Figure B2.67. Abundance and biomass of thorny skate from the NESFC spring (circles) and autumn (squares) bottom trawl surveys from 1963-2006 in the Gulf of Maine to Southern New England offshore region.

# Thorny Skate <br> GOM-SNE Offshore Only - Spring Survey 



Figure B2.68. Abundance and biomass of thorny skate from the NESFC spring bottom trawl surveys from 1968-2006 in the Gulf of Maine to Southern New England offshore region. The circles represent the original stratified mean, the squares represent the mean combining strata for bootstrapping, and the triangles represent the bootstrapped mean.

## Thorny Skate - Spring Survey <br> GOM-SNE Offshore Only



Figure B2.69. Bootstrapped abundance and biomass of thorny skate from the NESFC spring bottom trawl survey in the Gulf of Maine to Southern New England offshore region. Mean index in solid squares, $95 \%$ confidence interval in open squares.

# Thorny Skate GOM-SNE Offshore Only - Autumn Survey 



Figure B2.70. Abundance and biomass of thorny skate from the NESFC autumn bottom trawl surveys from 1968-2006 in the Gulf of Maine to Southern New England offshore region. The circles represent the original stratified mean, the squares represent the mean combining strata for bootstrapping, and the triangles represent the bootstrapped mean.

## Thorny Skate - Autumn Survey GOM-SNE Offshore Only



Figure B2.71. Bootstrapped abundance and biomass of thorny skate from the NESFC autumn bottom trawl survey in the Gulf of Maine to Southern New England offshore region. Mean index in solid squares, $95 \%$ confidence interval in open squares.

## Thorny Skate: GOM-SNE Offshore Percentiles of Length Composition



Figure B2.72. Percentiles of length composition (5,50, and 95) of thorny skate from the NESFC spring and autumn bottom trawl surveys from 1963-2006 in the Gulf of Maine to Southern New England offshore region.


Figure B2.73. Thorny skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Gulf of Maine to Southern New England offshore region, 1963-1972.


Figure B2.74. Thorny skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Gulf of Maine to Southern New England offshore region, 1973-1982.


Figure B2.75. Thorny skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Gulf of Maine to Southern New England offshore region, 1983-1992.


Figure B2.76. Thorny skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Gulf of Maine to Southern New England offshore region, 1993-2002.


Figure B2.77. Thorny skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Gulf of Maine to Southern New England offshore region, 2003-2006.

# Thorny Skate Scallop Survey 



Figure B2.78. Abundance and biomass of thorny skate from the NESFC scallop surveys from 1985-2006. The circles represent the original stratified mean, the squares represent the mean combining strata for bootstrapping, and the triangles represent the bootstrapped mean.

# Thorny Skate <br> Scallop Survey 



Figure B2.79. Bootstrapped abundance and biomass of thorny skate from the NESFC scallop survey. Mean index in solid squares, $95 \%$ confidence interval in open squares.

## Thorny Skate - Massachusetts Trawl Survey



Figure B2.80. Abundance and biomass of thorny skate from the Massachusetts spring and autumn finfish bottom trawl survey in state waters (strata 25-36).


Figure B2.81. Distribution of smooth skate from the spring and autumn NEFSC surveys from 2000-2006.



Figure B2.82. Distribution of smooth skate from the NEFSC scallop and shrimp surveys from 1985-2006.

## Smooth Skate <br> GOM-SNE Offshore Only



Figure B2.83. Abundance and biomass of smooth skate from the NESFC spring (circles) and autumn (squares) bottom trawl surveys from 1963-2006 in the Gulf of Maine to Southern New England offshore region.

# Smooth Skate GOM-SNE Offshore Only - Spring Survey 



Figure B2.84. Abundance and biomass of smooth skate from the NESFC spring bottom trawl surveys from 1968-2006 in the Gulf of Maine to Southern New England offshore region. The circles represent the original stratified mean, the squares represent the mean combining strata for bootstrapping, and the triangles represent the bootstrapped mean.

## Smooth Skate - Spring Survey GOM-SNE Offshore Only



Figure B2.85. Bootstrapped abundance and biomass of smooth skate from the NESFC spring bottom trawl survey in the Gulf of Maine to Southern New England offshore region. Mean index in solid squares, $95 \%$ confidence interval in open squares.

## Smooth Skate <br> GOM-SNE Offshore Only - Autumn Survey



Figure B2.86. Abundance and biomass of smooth skate from the NESFC autumn bottom trawl surveys from 1968-2006 in the Gulf of Maine to Southern New England offshore region. The circles represent the original stratified mean, the squares represent the mean combining strata for bootstrapping, and the triangles represent the bootstrapped mean.

## Smooth Skate - Autumn Survey GOM-SNE Offshore Only



Figure B2.87. Bootstrapped abundance and biomass of smooth skate from the NESFC autumn bottom trawl survey in the Gulf of Maine to Southern New England offshore region. Mean index in solid squares, $95 \%$ confidence interval in open squares.

## Smooth Skate: GOM-SNE Offshore Percentiles of Length Composition



Figure B2.88. Percentiles of length composition (5,50, and 95) of smooth skate from the NESFC spring and autumn bottom trawl surveys from 1963-2006 in the Gulf of Maine to Southern New England offshore region.


Figure B2.89. Smooth skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Gulf of Maine to Southern New England offshore region, 1963-1972.


Figure B2.90. Smooth skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Gulf of Maine to Southern New England offshore region, 1973-1982.


Figure B2.91. Smooth skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Gulf of Maine to Southern New England offshore region, 1983-1992.


Figure B2.92. Smooth skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Gulf of Maine to Southern New England offshore region, 1993-2002.


Figure B2.93. Smooth skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Gulf of Maine to Southern New England offshore region, 2003-2006.

Smooth Skate
Scallop Survey


Figure B2.94. Abundance and biomass of smooth skate from the NESFC scallop surveys from 1985-2006. The circles represent the original stratified mean, the squares represent the mean combining strata for bootstrapping, and the triangles represent the bootstrapped mean.

Smooth Skate Scallop Survey


Figure B2.95. Bootstrapped abundance and biomass of smooth skate from the NESFC scallop survey. Mean index in solid squares, $95 \%$ confidence interval in open squares.


Figure B2.96. Distribution of clearnose skate from the spring and autumn NEFSC surveys from 2000-2006.


Clearnose Skates from 2000-2006 NEFSC Winter Surveys

Figure B2.97. Distribution of clearnose skate from the winter NEFSC surveys from 2000-2006.

## Clearnose Skate Mid-Atlantic All strata



Figure B2.98. Abundance and biomass of clearnose skate from the NESFC spring (circles) and autumn (squares) bottom trawl surveys from 1975-2006 in the Mid-Atlantic (all strata).

## Clearnose Skate Mid-Atlantic All Strata - Spring Survey



Figure B2.99. Abundance and biomass of clearnose skate from the NESFC spring bottom trawl surveys from 1976-2006 in the Mid-Atlantic (all strata). The circles represent the original stratified mean, the squares represent the mean combining strata for bootstrapping, and the triangles represent the bootstrapped mean.

## Clearnose Skate - Spring Survey Mid-Atlantic All Strata



Figure B2.100. Bootstrapped abundance and biomass of clearnose skate from the NESFC spring bottom trawl survey in the Mid-Atlantic region (all strata). Mean index in solid squares, $95 \%$ confidence interval in open squares.

## Clearnose Skate

Mid-Atlantic All Strata - Autumn Survey


Figure B2.101. Abundance and biomass of clearnose skate from the NESFC autumn bottom trawl surveys from 1976-2006 in the Mid-Atlantic (all strata). The circles represent the original stratified mean, the squares represent the mean combining strata for bootstrapping, and the triangles represent the bootstrapped mean.

## Clearnose Skate - Autumn Survey Mid-Atlantic All Strata



Figure B2.102. Bootstrapped abundance and biomass of clearnose skate from the NESFC autumn bottom trawl survey in the Mid-Atlantic region (all strata). Mean index in solid squares, $95 \%$ confidence interval in open squares.

## Clearnose Skate Percentiles of Length Composition



Figure B2.103. Percentiles of length composition (5, 50, and 95) of clearnose skate from the NESFC spring and autumn bottom trawl surveys from 1975-2006 in the Mid-Atlantic region (all strata).

Consistent strata set not available prior to 1975/76


Figure B2.104. Clearnose skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Mid-Atlantic (all strata), 1975-1982.


Figure B2.105. Clearnose skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Mid-Atlantic (all strata), 1983-1992.


Figure B2.106. Clearnose skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Mid-Atlantic (all strata), 1993-2002.


Figure B2.107. Clearnose skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Mid-Atlantic (all strata), 2003-2006.

## Clearnose Skate Winter Survey



Figure B2.108. Abundance and biomass of clearnose skate from the NESFC winter bottom trawl surveys from 1992-2006. The circles represent the original stratified mean, the squares represent the mean combining strata for bootstrapping, and the triangles represent the bootstrapped mean.

## Clearnose Skate Winter Survey




Figure B2.109. Bootstrapped abundance and biomass of clearnose skate from the NESFC winter bottom trawl survey. Mean index in solid squares, $95 \%$ confidence interval in open squares.

## Clearnose Skate - CTDEP Finfish Survey



Figure B2.110. Abundance and biomass of clearnose skate from the CTDEP spring and autumn finfish bottom trawl survey in Connecticut state waters, 1984-2006.


Figure B2.111. Distribution of rosette skate from the spring and autumn NEFSC surveys from 2000-2006.


Rosette Skates from 2000-2006 NEFSC Winter Surveys

Figure B2.112. Distribution of rosette skate from the winter NEFSC surveys from 2000-2006.

## Rosette Skate Mid-Atlantic Offshore strata



Figure B2.113. Abundance and biomass of rosette skate from the NESFC spring (circles) and autumn (squares) bottom trawl surveys from 1967-2006 in the Mid-Atlantic offshore region.

# Rosette Skate Mid-Atlantic Offshore Only - Spring Survey 



Figure B2.114. Abundance and biomass of rosette skate from the NESFC spring bottom trawl surveys from 1968-2006 in the Mid-Atlantic offshore region. The circles represent the original stratified mean, the squares represent the mean combining strata for bootstrapping, and the triangles represent the bootstrapped mean.

## Rosette Skate - Spring Survey Mid-Atlantic Offshore Strata Only



Year


Figure B2.115. Bootstrapped abundance and biomass of rosette skate from the NESFC spring bottom trawl survey in the Mid-Atlantic offshore region. Mean index in solid squares, $95 \%$ confidence interval in open squares.

## Rosette Skate Mid-Atlantic Offshore Only - Autumn Survey



Figure B2.116. Abundance and biomass of rosette skate from the NESFC autumn bottom trawl surveys from 1967-2005 in the Mid-Atlantic offshore region. The circles represent the original stratified mean, the squares represent the mean combining strata for bootstrapping, and the triangles represent the bootstrapped mean.

## Rosette Skate - Autumn Survey Mid-Atlantic Offshore Strata Only



Figure B2.117. Bootstrapped abundance and biomass of rosette skate from the NESFC autumn bottom trawl survey in the Mid-Atlantic offshore region. Mean index in solid squares, $95 \%$ confidence interval in open squares.

## Rosette Skate <br> Percentiles of Length Composition



Figure B2.118. Percentiles of length composition (5, 50, and 95) of rosette skate from the NESFC spring and autumn bottom trawl surveys from 1967-2006 in the Mid-Atlantic offshore region.


Figure B2.119. Rosette skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Mid-Atlantic offshore region, 1967-1972.


Figure B2.120. Rosette skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Mid-Atlantic offshore region, 1973-1982.


Figure B2.121. Rosette skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Mid-Atlantic offshore region, 1983-1992.


Figure B2.122. Rosette skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Mid-Atlantic offshore region, 1993-2002.


Figure B2.123. Rosette skate length composition from the NEFSC spring and autumn bottom trawl surveys in the Mid-Atlantic offshore region, 2003-2006.

## Rosette Skate Winter Survey



Figure B2.124. Abundance and biomass of rosette skate from the NESFC winter bottom trawl surveys from 1998-2006. The circles represent the original stratified mean, the squares represent the mean combining strata for bootstrapping, and the triangles represent the bootstrapped mean.


Figure B2.125. Bootstrapped abundance and biomass of rosette skate from the NESFC winter bottom trawl survey. Mean index in solid squares, $95 \%$ confidence interval in open squares.

Skate Complex SSB Indices


Figure B2.126. Trends in spawning stock biomass indices for seven species of skates.

FIGURES B2.127-B2.141.
(EDITOR'S NOTE: BASED ON THE REVIEWER'S COMMENTS, THESE FIGURES WERE NOT INCLUDED IN THIS REPORT. THE FIGURES DEALT WITH ESTIMATES OF FISHING MORTALITY RATE.)


Figure B3.1. Relationship between spawning stock biomass indices ( $>=76 \mathrm{~cm}$ ) and recruitment indices (no/tow, $34-39 \mathrm{~cm}$ ) for winter skate. The time lag between SSB and recruitment accounts for the assumed age 3 at recruitment plus one year for hatching time.

## Winter Skate




Figure B3.2. Stock-recruitment plots for winter skate and barndoor skate with the Beverton-Holt function plotted.


Figure B.3.3. Relationship between spawning stock biomass indices ( $>=44 \mathrm{~cm}$ ) and recruitment indices (no/tow, $38-42 \mathrm{~cm}$ ) for little skate. The time lag between SSB and recruitment accounts for the assumed age 3 at recruitment plus one year for hatching time.


Figure B3.4. Relationship between spawning stock biomass indices ( $>=116 \mathrm{~cm}$ ) and recruitment indices (no/tow, $55-69 \mathrm{~cm}$ ) for barndoor skate. The time lag between SSB and recruitment accounts for the assumed age 2 at recruitment plus one year for hatching time.

# Thorny Skate <br> Relationship Between SSB Indices and Recruitment Indices 



Figure B3.5. Relationship between spawning stock biomass indices ( $>=80 \mathrm{~cm}$ ) and recruitment indices (no/tow, $25-35 \mathrm{~cm}$ ) for thorny skate. The time lag between SSB and recruitment accounts for the assumed age 2 at recruitment plus one year for hatching time.

## Thorny Skate



## Clearnose Skate



Figure B3.6. Stock-recruitment plots for thorny skate and clearnose skate with the Beverton-Holt function plotted.

# Clearnose Skate <br> Relationship Between SSB Indices and Recruitment Indices 



Figure B3.7. Relationship between spawning stock biomass indices ( $>=66 \mathrm{~cm}$ ) and recruitment indices (no/tow, $42-50 \mathrm{~cm}$ ) for clearnose skate. The time lag between SSB and recruitment accounts for the assumed age 3 at recruitment plus one year for hatching time.

## Skate Complex Biomass Indices



Figure B4.1. NEFSC survey biomass indices (kg/tow). Thin lines with symbols are annual indices, thick lines are 3-year moving averages, the thin horizontal lines are the current biomass targets and thresholds.

FIGURES B4.2-B4.21.
(EDITOR'S NOTE: BASED ON THE REVIEWER'S COMMENTS, THESE FIGURES WERE NOT INCLUDED IN THIS REPORT. THE FIGURES DEALT WITH ESTIMATES OF ALTERNATIVE BIOLOGICAL REFERENCE POINTS FOR SKATES. )



Figure B6.1. Sensitivity of Average per Capita Annual Consumption to a) mean stomach contents and b) temperature.


Figure B6.2. Sensitivity of Average per Capita Annual Consumption to the parameters a) alpha and b) beta.


Figure 6.3. Sensitivity of Annual per Capita Consumption variation in both temperature and mean stomach contents.


Figure 6.4. Sensitivity of Annual per Capita Consumption variation in both alpha and mean stomach contents.


Figure 6.5. Sensitivity of Annual per Capita Consumption variation in both beta and mean stomach contents.


Figure 6.6. Sensitivity of Annual per Capita Consumption variation in both beta and alpha.

Winter Skate
Mean Stomach Contents (g)


Figure B6.7a. The annual mean stomach contents $(0.1 \mathrm{~g})$ of winter skate for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.


Figure B6.7b. The percentage of stomachs that were empty (i.e., containing no prey) of Winter skate for the strata set and time period noted. Each size class is noted


Figure B6.8a. The mean length ( 1 cm ) of Winter skate from which stomach samples were collected, for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.

Winter Skate
Mean Length vs. Mean Stomach Contents


Figure B6.8b. The annual mean stomach contents $(0.1 \mathrm{~g})$ and the mean length $(1 \mathrm{~cm})$ of Winter skate for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.


Figure B6.9a. The annual mean bottom temperature $\left(0.1^{\circ} \mathrm{C}\right)$ for the selected strata set, as taken from the bottom trawl survey over the time period noted. Error bars are $\pm 1$ S.E.


Figure B6.9b. The annual per capita consumption (g yr-1) of Winter skate for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.


Figure B6.10a. The annual mean swept area abundance of winter skate for the strata set and time period noted. Each size class is noted.


Figure B6.10b. The annual total consumption (MT yr-1) of Winter skate for the strata set and time period noted.

WINTER SKATE PREY REMOVAL


Figure B6.11. The amount of prey consumed ( $\mathrm{MT} \mathrm{yr}^{-1}$ ) by Winter skate for the strata set and time period noted. These estimates represent the combination of total annual total consumption and the diet compositions of Winter skate. These prey were selected as some of the major prey ( $\gg 5 \%$ of diet composition) of Winter skate.


Figure B6.12. The amount of prey consumed $\left(\mathrm{MT} \mathrm{yr}^{-1}\right)$ by Winter skate for the strata set and time period noted. These estimates represent the combination of total annual total consumption and the diet compositions of Winter skate. These prey were selected as some of the major prey ( $\gg 5 \%$ of diet composition) of Winter skate.

## Little Skate <br> Mean Stomach Contents (g)



Figure B6.13a. The annual mean stomach contents ( 0.1 g ) of Little skate for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.


Figure B6.13b. The percentage of stomachs that were empty (i.e., containing no prey) of Little skate for the strata set and time period noted. Each size class is noted.


Figure B6.14a. The mean length ( 1 cm ) of Little skate from which stomach samples were collected, for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.

Little Skate
Mean Length vs. Mean Stomach Contents


Figure B6.14b. The annual mean stomach contents $(0.1 \mathrm{~g})$ and the mean length $(1 \mathrm{~cm})$ of Little skate for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.


Figure B6.15a. The annual mean bottom temperature $(0.1 \mathrm{oC})$ for the selected strata set, as taken from the bottom trawl survey over the time period noted. Error bars are $\pm 1$ S.E.


Figure B6.15b. The annual per capita consumption $\left(\mathrm{g} \mathrm{yr}^{-1}\right)$ of Little skate for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.


Figure B6.16a. The annual mean swept area abundance of Little skate for the strata set and time period noted. Each size class is noted.


Figure B6.16b. The annual total consumption ( $\mathrm{MT} \mathrm{yr}^{-1}$ ) of Little skate for the strata set and time period noted.


Figure B6.17. The amount of prey consumed (MT yr-1) by Little skate for the strata set and time period noted. These estimates represent the combination of total annual total consumption and the diet compositions of Little skate. These prey were selected as some of the major prey ( $\gg 5 \%$ of diet composition) of Little skate.


Figure B6.18a. The annual mean stomach contents ( 0.1 g ) of barndoor skate for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.

## Barndoor Skate <br> Empties



Figure B6.18b. The percent of barndoor skates that had empty stomachs, by year and size class.


Figure B6.19a. The mean length ( 1 cm ) of Barndoor skate from which stomach samples were collected, for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.

## Barndoor Skate <br> Mean Length vs. Mean Stomach Contents



Figure B6.19b. The annual mean stomach contents $(0.1 \mathrm{~g})$ and the mean length $(1 \mathrm{~cm})$ of Barndoor skate for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.


Figure B6.20a. The annual mean bottom temperature $\left(0.1^{\circ} \mathrm{C}\right)$ for the selected strata set, as taken from the bottom trawl survey over the time period noted. Error bars are $\pm 1$ S.E.


Figure B6.20b. The annual per capita consumption ( $\mathrm{g} \mathrm{yr}^{-1}$ ) of Barndoor skate for the strata set and time period noted. Each size class is noted.


Figure B6.21a. The annual mean swept area abundance of Barndoor skate for the strata set and time period noted. Each size class is noted.


Figure B6.21b. The annual total consumption $\left(\mathrm{MT} \mathrm{yr}^{-1}\right)$ of Barndoor skate for the strata set and time period noted.

## BARNDOOR SKATE PREY REMOVAL



Figure B6.22. The amount of prey consumed ( $\mathrm{MT} \mathrm{yr}^{-1}$ ) by Barndoor skate for the strata set and time period noted. These estimates represent the combination of total annual total consumption and the diet compositions of Barndoor skate. These prey were selected as some of the major prey ( $\gg 5 \%$ of diet composition) of Barndoor skate.


Figure B6.23a. The annual mean stomach contents $(0.1 \mathrm{~g})$ of Thorny skate for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.


Figure B6.23b. The percentage of stomachs that were empty (i.e., containing no prey) of Thorny skate for the strata set and time period noted. Each size class is noted


Figure B6.24a. The mean length ( 1 cm ) of Thorny skate from which stomach samples were collected, for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.


Figure B6.24b. The annual mean stomach contents $(0.1 \mathrm{~g})$ and the mean length $(1 \mathrm{~cm})$ of Thorny skate for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.


Figure B6.25a. The annual mean bottom temperature $\left(0.1^{\circ} \mathrm{C}\right)$ for the selected strata set, as taken from the bottom trawl survey over the time period noted. Error bars are $\pm 1$ S.E.


Figure B6.25b. The annual per capita consumption ( $\mathrm{g} \mathrm{yr}^{-1}$ ) of Thorny skate for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.

Thorny Skate
Abundance


Figure B6.26a. The annual mean swept area abundance of Thorny skate for the strata set and time period noted. Each size class is noted.


Figure B6.26b. The annual total consumption ( $\mathrm{MT} \mathrm{yr}^{-1}$ ) of Thorny skate for the strata set and time period noted.


Figure B6.27. The amount of prey consumed (MT yr-1) by Thorny skate for the strata set and time period noted. These estimates represent the combination of total annual total consumption and the diet compositions of Thorny skate. These prey were selected as some of the major prey ( $\gg 5 \%$ of diet composition) of Thorny skate.

THORNY SKATE PREY REMOVAL


Figure B6.28. The amount of prey consumed (MT yr-1) by Thorny skate for the strata set and time period noted. These estimates represent the combination of total annual total consumption and the diet compositions of Thorny skate. These prey were selected as some of the major prey ( $\gg 5 \%$ of diet composition) of Thorny skate.


Figure B6.29a. The annual mean stomach contents ( 0.1 g ) of Smooth skate for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.


Figure B6.29b. The percentage of stomachs that were empty (i.e., containing no prey) of smooth skate for the strata set and time period noted. Each size class is noted


Figure B6.30a. The mean length ( 1 cm ) of Smooth skate from which stomach samples were collected, for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.


Figure B6.30b. The annual mean stomach contents $(0.1 \mathrm{~g})$ and the mean length ( 1 cm ) of Smooth skate for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.


Figure B6.31a. The annual mean bottom temperature $\left(0.1^{\circ} \mathrm{C}\right)$ for the selected strata set, as taken from the bottom trawl survey over the time period noted. Error bars are $\pm 1$ S.E.


Figure B6.31b. The annual per capita consumption $\left(\mathrm{g} \mathrm{yr}^{-1}\right)$ of Smooth skate for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.


Figure B6.32a. The annual mean swept area abundance of Smooth skate for the strata set and time period noted. Each size class is noted.


Figure B6.32b. The annual total consumption ( $\mathrm{MT} \mathrm{yr}^{-1}$ ) of Smooth skate for the strata sel and time period noted.


Figure B6.33. The amount of prey consumed (MT yr-1) by Smooth skate for the strata set and time period noted. These estimates represent the combination of total annual total consumption and the diet compositions of Smooth skate. These prey were selected as some of the major prey ( $\gg 5 \%$ of diet com osition of Smooth skate.


Figure B6.34a. The annual mean stomach contents ( 0.1 g ) of Clearnose skate for the strata and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.

## Clearnose Skate <br> Empties



Figure B6.34b. The percentage of stomachs that were empty (i.e., containing no prey) of Clearnose skate for the strata set and time period noted. Each size class is noted


Figure B6.35a. The mean length ( 1 cm ) of Clearnose skate from which stomach sample were collected, for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.

Clearnose Skate
Mean Length vs. Mean Stomach Contents


Figure B6.35b. The annual mean stomach contents ( 0.1 g ) and the mean length $(1 \mathrm{~cm})$ of Clearnose skate for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.


Figure B6.36a. The annual mean bottom temperature $\left(0.1^{\circ} \mathrm{C}\right)$ for the selected strata set as taken from the bottom trawl survey over the time period noted. Error bars are $\pm 1$ S.E.

Clearnose Skate
Per Capita Consumption


Figure B6.36b. The annual per capita consumption $\left(\mathrm{g} \mathrm{yr}^{-1}\right)$ of Clearnose skate for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.


Figure B6.37a. The annual mean swept area abundance of Clearnose skate for the strata set and time period noted. Each size class is noted.


Figure B6.37b. The annual total consumption (MT yr ${ }^{-1}$ ) of Clearnose skate for the stra set and time period noted.


Figure B6.38. The amount of prey consumed (MT yr-1) by Clearnose skate for the strata set and time period noted. These estimates represent the combination of total annual total consumption and the diet compositions of Clearnose skate. These prey were selected as some of the major prey ( $\gg 5 \%$ of diet composition) of Clearnose skate.


Figure B6.39a. The annual mean stomach contents ( 0.1 g ) of Rosette skate for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.


Figure B6.39b. The percentage of stomachs that were empty (i.e., containing no prey) of Rosette skate for the strata set and time period noted. Each size class is noted


Figure B6.40a. The mean length ( 1 cm ) of Rosette skate from which stomach samples were collected, for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.

Rosette Skate
Mean Length vs. Mean Stomach Contents


Figure B6.40b. The annual mean stomach contents $(0.1 \mathrm{~g})$ and the mean length $(1 \mathrm{~cm})$ of Rosette skate for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E.


Figure B6.41a. The annual mean bottom temperature $\left(0.1^{\circ} \mathrm{C}\right)$ for the selected strata set, as taken from the bottom trawl survey over the time period noted. Error bars are $\pm 1$ S.E.

Rosette Skate Per Capita Consumption


Figure B6.41b. The annual per capita consumption $\left(\mathrm{g} \mathrm{yr}^{-1}\right)$ of Rosette skate for the strata set and time period noted. Each size class is noted. Error bars are $\pm 1$ S.E


Figure B6.42a. The annual mean swept area abundance of Rosette skate for the strata sel and time period noted. Each size class is noted.


Figure B6.42b. The annual total consumption ( $\mathrm{MT} \mathrm{yr}^{-1}$ ) of Rosette skate for the strata set and time period noted.

ROSETTE SKATE PREY REMOVAL


Figure B6.43. The amount of prey consumed ( $\mathrm{MT} \mathrm{yr}^{-1}$ ) by Rosette skate for the strata set and time period noted. These estimates represent the combination of total annual total consumption and the diet compositions of Rosette skate. These prey were selected as some of the major prey ( $\gg 5 \%$ of diet composition) of Rosette skate.



Figure B6.44. Average amount of major prey consumed by all skates from 2000-2005. A. fish prey. B. invertebrate prey.

## C. ASSESSMENT OF ATLANTIC SURFCLAM

Report of the Invertebrate Subcommittee (see Appendix C1 for membership)

### 1.0 TERMS OF REFERENCE (TOR) AND SUMMARY

1. Characterize the commercial and recreational catch including landings and discards. Completed, see Section C3.
2. Estimate fishing mortality, spawning stock biomass, and total stock biomass for the current year and characterize the uncertainty of those estimates. If possible, also include estimates for earlier years. Completed, see Section C5.
3. Either update or redefine biological reference points (BRPs; proxies for $B_{M S Y}$ and $F_{M S Y}$ ), as appropriate. Comment on the scientific adequacy of existing and redefined BRPs. Completed, see section C6. Biomass reference points were updated based on new estimates of historical biomass levels and criteria in the Surfclam and Ocean Quahog Fisheries Management Plan. Fishing mortality reference points did not require updating. Current reference points were adequate for this assessment because stock biomass is relatively high and fishing mortality rates are low. However, it was noted that implicit assumptions about $B_{M S Y}$ and biomass during 1999 may not be valid and should be reevaluated.
4. Evaluate current stock status with respect to the existing BRPs, as well as with respect to updated or redefined BRPs (from TOR 3). Completed, see section C7. The stock is not overfished and overfishing is not occurring.
5. Recommend what modeling approaches and data should be used for conducting single and multi-year stock projections, and for computing TACs or TALs. Completed, see Section C8. A consistent set of stock assessment modeling, integrated bootstrap and stochastic projection software is now available that can deal with auto correlated recruitment patterns in surfclam. It is not necessary to describe approaches for setting TAC or TAL levels because the fishery is managed with constant quota levels.
6. If possible:
a. Provide numerical examples of short term projections (2-3 years) of biomass and fishing mortality rate, and characterize their uncertainty, under various TAC/F strategies. Completed, see Section C9. Example projections under a wide range of scenarios indicate that surfclam biomass will decline over the next 2-3 years to levels near the $B_{M S Y}$ proxy level that used is used by managers as a target. The recent and expected declines are due to poor recruitment and slow growth. There is no indication that the stock will become overfished or that overfishing will occur. Uncertainty is very high, particularly for longer term projections.
b. Compare projected stock status to existing rebuilding or recovery schedules, as appropriate. Not relevant. surfclam are not overfished and no rebuilding schedule exists.
7. Review, evaluate and report on the status of the SARC/Working Group Research Recommendations offered in recent SARC reviewed assessments. Completed, see Section C10.

## Plain terms summary

1) The following abbreviations are used to identify stock assessment and fishing regions for surfclam (Figure C1).

| Region (south to north) | Abbreviation |
| :---: | :---: |
| Southern Virginia | SVA |
| Delmarva | DMV |
| New Jersey | NJ |
| Long Island | LI |
| Southern New England | SNE |
| Georges Bank | GBK |

2) Overall, total surfclam biomass has declined during recent years due to slow growth and poor recruitment, particularly in southern regions. Despite declines, total stock biomass is still at a relatively high level. Fishing mortality is low in all regions.
3) Stock conditions are relatively good in northern regions such as LI, SNE and GBK where the bulk of the stock was found during 2005 and little fishing occurs. Stock conditions are poorer in southern regions, DMV and SVA in particular, where fishing has occurred since the 1980's and a relatively small fraction of the stock was found during 2005. Conditions in NJ, where most of the fishing and a large fraction of the stock occur, are intermediate.
4) The surfclam stock is not overfished and overfishing is not occurring. Overfishing and overfished stock conditions are not likely to occur in the near future.
5) Total landings from the EEZ stock during 2005 were less than the quota due, based on industry sources, to market factors.
6) The majority of landings during recent years were from the NJ region although some landings were also taken from DMV in the south. Landings in the northern SNE and LI regions increased during recent years were minor. No fishing occurs on GBK due to risk of paralytic shellfish poisoning (PSP).
7) Over time, surfclam biomass has shifted towards the north. During 2005, the largest fraction of stock biomass was in GBK, rather than in NJ or DMV.
8) Fishing effort and catch have shifted north during recent years as catch rates in the south have declined.
9) Total fishing effort increased during recent years while landings per unit effort (LPUE) decreased for the fishery as a whole.
10) LPUE has declined in NJ and drastically in DMV. LPUE in the LI region appears to be increasing.
11) Growth rates for surfclam in NJ, and particularly in DMV, have slowed in recent years so that the age at recruitment to the fishery has increased by 1-2 years. Delayed recruitment and slower growth after reaching fishable size reduce potential fishery yield by a substantial amount. Slower growth is due to environmental factors.
12) Recruitment has declined during recent years for the stock as a whole and is at or near record low levels in most regions.
13) Stock biomass for the entire stock was at record high levels during the late 1990s. Since then stock biomass has declined. In 2005, total stock biomass was about the same as before the peak.
14) Biomass trends for NJ were similar to trends for the entire stock. Biomass trends for DMV indicate steeper and continuous declines since the record high levels for DMV during the early 1980s.
15) Recent declines in biomass are due to negative surplus production. This means that factors that increase stock biomass including growth and recruitment have not been large enough to offset natural (not related to fishing) losses.
16) Fishing mortality rates are low in all regions. The environment, rather than fishing, apparently caused the recent declines in biomass.

### 2.0 INTRODUCTION

This stock assessment for the offshore subspecies of Atlantic surfclam (Spisula solidissima solidissima) was prepared for SAW/SARC-44 along with a stock assessment for ocean quahog (Arctica islandica). No information is provided about the smaller coastal form (S. s. similis) that occupies relatively southern inshore habitats (Hare and Weinberg 2005). The geographic distributions of the two subspecies overlap to a limited extent in the south and in some inshore waters to the north. However, S. s. similis is reproductively isolated from S. s. solidissima and not important to the commercial fishery.

The same methods were used in the assessments for surfclam and ocean quahog although the surfclam assessment was completed after the ocean quahog assessment and incorporates a number of improvements. Interested persons and reviewers should read
the ocean quahog assessment (i.e., Assessment A in this volume) first because the methods used for both species are described there in detail. Improvements to methods for surfclam and other details relevant only to surfclam are described below.

## Distribution and biology

Atlantic surfclam is a relatively large fast growing bivalve distributed in the western North Atlantic Ocean, along the coast of North American from the southern Gulf of St. Lawrence to Cape Hatteras (Figure C1). Individuals larger than 16 cm shell length (SL) are relatively common in NEFSC surveys. Commercial concentrations are found primarily off New Jersey, the Delmarva Peninsula, and on Georges Bank. Surfclam are found from the intertidal zone to a depth of about 60 m but densities are low at depths greater than 40 m . See Cargnelli et al. (1999) for a complete review of life history and distributional information. The distribution of Atlantic surfclam and the distribution of a related species (S. similis) overlap in the south and some inshore waters to the north (Hare and Weinberg 2005).

It is likely that all Atlantic surfclam along the northeast coast belong to the same biological population. Surfclam are common in both inshore state ( $\leq 3 \mathrm{mi}$ from shore) and offshore federal waters. Federal waters consist of the Exclusive Economic Zone (EEZ), between 3 and 200 mi from shore. The stock assessment applies only to the EEZ segment of the surfclam population in federal waters, however, because the EEZ is the management unit specified in the Fishery Management Plan for the Atlantic Surfclam and Ocean Quahog Fisheries (FMP). Surfclam in New Jersey and New York state waters support valuable fisheries that are managed by state authorities.

Surfclam in the EEZ are managed as a unit stock but there is substantial regional variability in exploitation rates and biological characteristics. A variety of calculations and estimates in this assessment are presented for smaller stock assessment regions which are defined below (Figure C1). Previous assessments separated the New Jersey (NJ) region into Northern New Jersey (NNJ) and Southern New Jersey (SNJ) components. In this assessment, the NJ region is treated as a single entity. SNJ and NNJ were combined to simplify the assessment and because data for SNJ were too limited and variable to be analyzed separately.

There is uncertainty about the timing of annual mark (annulus) formation in surfclam chondrophores, which are cut from shells and used to age surfclams taken in NEFSC clam surveys. There is additional uncertainty about indentifying the first annual mark (Jacobson et al. 2006). Despite these questions of interpretation, surfclam annual rings are relatively easy to count. In this assessment, the number of annual marks and age are assumed to be the same and the assumed birth date is January 1 so that, for example, a member of the 2004 year class taken during the 2005 NEFSC clam survey would be age 1 at the time of capture and expected to show one ring. Ages for surfclams taken in the commercial fishery that operates year round are more uncertain. Surfclams age $20+$ are relatively common and the maximum observed age exceeds 35 . See Jacobson et al. (2006) for information about procedures used to estimate surfclam age.

Surfclams are capable of reproduction at age 1 , although full maturity may not be reached until age 2. Spawning occurs during late summer and early fall. Eggs and sperm are shed directly into the water column. Recruitment to the bottom occurs after a planktonic larval period of about three weeks.

Weinberg (1998) and Weinberg and Helser (1996) show that growth rates vary among regions, over time and in response to surfclam density levels. Based on NEFSC
clam survey data (Figure C2), growth rates appear to have declined for surfclams in the southern DMV region and to a lesser extent in the NJ region since 1993. Slower growth in surfclams in DMV during recent years coincides with mortality in near shore areas off DMV probably due to warm water (Weinberg 2005) and lower occurrence of surfclams with $25+$ annual marks in survey data (Figure C2).

Length-weight parameters used in this assessment to convert numbers of surfclams of different shell lengths in surveys to meat weight equivalents are region specific and based on fresh (unfrozen) material (Table C1). Length-weight parameters vary among locations and over time. Although length-weight data are collected periodically during NEFSC clam surveys, recent assessments used the same lengthweight relationship for the sake of simplicity and consistency (NEFSC 2003). A simple and consistent approach is used because length-weight data are not available for the commercial catch (which targets clams with high meat yield) and because length-weight information for early surveys was based on frozen material.

## Management

The fishery for Atlantic surfclams and ocean quahogs in the EEZ are unique in being the first US fishery managed under an individual transferable quota (ITQ) system. ITQ management was established during 1990 by the Mid-Atlantic Fishery Management Council under Amendment 8 to the Fishery Management Plan for the Atlantic Surfclam and Ocean Quahog Fisheries (FMP). Management measures include an annual quota for EEZ waters, which was 26.2 thousand mt meats per year during 2001-2005, and mandatory logbooks that describe each fishing trip. See Murawski and Serchuk (1989) and Serchuk and Murawski (1997) for detailed information about history, management and fishery operations. MAFMC (2006) describes recent fishery conditions and management for both surfclams and ocean quahogs.

## Previous assessments

Stock assessments are generally done after NMFS clam surveys, which are conducted every 2-3 years. In the most recent stock assessment for surfclam, NEFSC (2003) concluded that the stock was above the management target level (the stock was not overfished) and that fishing mortality was below the management threshold value (overfishing was not occurring). The stock was characterized as declining from a relatively high biomass level at the rate of about $5 \%$ per year due to negative surplus production and, in particular, relatively low recruitment. Conclusions from this stock assessment are similar. See NEFSC (1993; 1995; 1998; 2000) for earlier surfclam stock assessments.

Beginning with NEFSC (1998), the primary emphasis in surfclam stock assessments was: 1) use of sensors to evaluate survey dredge performance; 2) estimating survey dredge efficiency via cooperative "depletion studies"; and 3) calculation of efficiency corrected swept-area biomass. Previous stock assessments used stock assessment models with variable results. In this assessment, data from all available depletion studies are analyzed using consistent and improved methods. The updated information is used in a stock assessment model that is successfully applied to the stock as a whole and to the important DMV and NJ regions.

### 3.0 COMMERCIAL CATCH (TOR-1)

In using landings data for surfclams, 1 industry standard bushel ( $1.88 \mathrm{ft}^{3}$ ) was assumed to produce 17 lbs or 7.711 kg of useable meats. Fishery landings in this assessment are reported as meat weights for ease in comparison to survey data and in calculations but were originally recorded in units of cages ( 1 cage $=32$ industry bu). LPUE data, however, are reported in this assessment as landings in bushels per hour fished.

As in previous assessments (NEFSC 2003), catch in all stock assessment analyses is the sum of landings plus a $12 \%$ upper bound for incidental mortality that may occur during fishing operations (i.e. assumed catch = 1.12 times landings). It is important to realize that the $12 \%$ figure is an upper bound and that actual incidental mortality is likely to be lower. Incidental mortality in the surfclam and ocean quahog fisheries is likely lower than might be expected because the total area fished is modest. The total area fished is relatively low because fishermen operate efficiently under ITQ management and target only areas of highest density. Moreover, the ITQ fishery operates with little or no regulation induced inefficiency (e.g. inefficiency due to area closures, trip limits, size limits, etc.). Discard of small surfclams occurred during 1982-1990 when size limits were used to regulate the surfclam fishery (Table C2) but are currently near zero. Recreational catch is near zero.

Size selectivity of commercial clam dredges and harvesting equipment has not been characterized quantitatively in detail. Based on commercial length data and experimental results, NEFSC (2003) assumed that surfclams in NJ were fully available to the commercial fishery at 120 mm SL and that surfclams in other regions were fully available to the commercial fishery at 110 SL.

In this assessment, surfclams $120+\mathrm{mm}$ SL are assumed to be the fishable stock in all regions. In contrast, that NEFSC (2003) used $120+\mathrm{mm}$ for NJ and $110+\mathrm{mm}$ SL for other regions. Fishing mortality estimates in this assessment, for example, compare total catch (landings plus an assumed $12 \%$ upper bound for incidental mortality) to the fishable stock $120+$ SL. The bulk of the fishery and much of the stock occurs in NJ, where NEFSC (2003) assumed recruitment at 120 mm SL. Based on commercial length data in NEFSC (2003) and shown below, there is no strong evidence that size at recruitment differs among regions. Consistent use of 120 mm SL simplifies the assessment and makes biomass and fishing mortality estimates for combined regions easier to interpret.

## Age at recruitment

Age at recruitment to the surfclam fishery depends on growth rates and, in particular, the ages at which surfclams reach 120 mm SL. Growth curves used in stock assessment modeling (described later) fit to survey age data indicate that surfclam recruited to the DMV fishery at about age $51 / 2$ y during 1982-1992 and at about age $71 / 2$ y during 1994-2005. Growth curves for NJ show that surfclams reached 120 mm SL and recruited to the fishery at about age 5 y during 1982-1992 and at about age 6 y during 1994-2005. Changes in age at recruitment should have substantial effect on potential fishery yield. Assuming a natural mortality rate of $M=0.15 \mathrm{y}^{-1}$, for example, numbers of recruits to the fishery per surviving larvae would be decreased by about $26 \%$ due to natural mortality during the two additional years prior to recruitment. This effect is likely
compounded by other reductions in productivity due to slower growth after recruitment to the fishery occurs.

## Landings, fishing effort and prices

Landings and fishing effort data for 1982-2005 were from mandatory logbooks. Data for earlier years were from NEFSC (2003) and MAFMC (2006).

Landings data for surfclams are relatively accurate in comparison to other fisheries because of a comprehensive system for tracking landings in the ITQ fishery. Effort data are, however not reliable for 1985-1990, due to regulations that restricted the duration of fishing to 6 hr . Effort data are relatively reliable during later years.

Surfclam landings were primarily from the US EEZ during 1965-2002 (Table C3 and Figure C3). EEZ landings peaked during 1973-1974 at about 33 thousand mt. EEZ landings were relatively high during 2001-2005 and varied between 21 and 25 thousand mt . Landings reached the quota in most years but were less than the quota during 2005 because of limited markets (according to industry sources).

The bulk of EEZ landings were from DMV during 1979-1980 and from NJ during every year since 1981 (Table C4 and Figure C4). During 2001-2005 DMV landings were modest with relatively small amounts reported from the LI and SNE regions. Trends in fishing effort were similar (Table C5 and Figure C5).

Nominal exvessel prices for the inshore and EEZ fisheries increased from about $\$ 8 \mathrm{bu}^{-1}$ during 1982 to $\$ 10 \mathrm{bu}^{-1}$ during 1994 and then declined to about $\$ 9.50 \mathrm{bu}^{-1}$ during 2000-2005 (Figure C6). Using 1980-1982 as a basis, prices declined in real terms from about $\$ 9 \mathrm{bu}^{-1}$ during 1982 to about $\$ 5 \mathrm{bu}^{-1}$ during 2005. Based on industry sources (D. Wallace, pers. comm.), the "break-even" price for surfclams during 2005 (i.e. price necessary to cover variable costs such as fuel, crew shares, food, etc.) was about \$4-\$5 $\mathrm{bu}^{-1}$ (nominal, 2005 dollars).

## Landings per unit effort

Nominal landings per unit effort (LPUE) based on logbooks was computed as total landings divided by total fishing effort for all vessels and all trips (Table C6 and Figure C7). In addition, standardized LPUE indices (Table C7 and Figure C7) were computed from a log-linear GLM model with year, month and vessel effects for each region (see Assessment A. Ocean quahog, in this Report). GLM models were fit to tow by tow logbook data for vessels in size class 3 and 4 (51-150 and 151-500 GRT) which are the bulk of the EEZ fishery. There were no records with zero catch and it was not necessary to add a constant before applying the log transformation to the data. Year effects were used as the index of LPUE after they were adjusted to the average of June catch rates for a single vessel that fished in all regions.

For surfclams, year, vessel and month effects were statistically significant for all regions. Although month effects were statistically significant, they were small, of little practical importance and because they did not show meaningful seasonal trends.

Trends in nominal and standardized LPUE were similar (Figure C7). In particular, LPUE declined steadily from peak levels during 1994 to relatively low levels during 2005 in the DMV region. LPUE declined slowly but steadily in the NJ region during 1991-1995 and in LI after 2000. LPUE levels during 2005 were at or near record lows. In contrast to other regions, LPUE levels in SNE increased rapidly after 1998 as the small fishery in SNE developed.

LPUE is not an ideal measure of fishable biomass trends for sessile and patchy stocks like surfclam because fishermen target high density beds and change their operations to maintain relatively high catch rates as stock biomass declines (Hilborn and Walters 1992). However, trends in LPUE and fishable biomass based on the NEFSC clam survey were similar during recent years for DMV and NJ where fishing has been heaviest and fishing grounds are widespread (Figure C29). In contrast, LPUE and survey trends were not similar for LI and SNE where less fishing has occurred and the fishery is not as widespread. The correlation in trends for DMV and NJ was likely due to reduced surfclam densities in many habitat areas where significant densities occurred. Previous assessments noted that the fishery in DMV and NJ and surfclam stock overlap relatively completely.

Spatial patterns in fishery data
Average landings, fishing effort and LPUE per year from logbooks were calculated for ten-minute squares (TNMS) during 1981-1990, 1991-1995, 1996-2000, and 2001-2005. For plots, data for TNMS with very low levels of landings and data for TNMS outside the range of the fishery (obvious errors) were omitted.

Spatial patterns in fishery data (Figure C 8 to C 9 ) show relatively high landings and fishing effort in the south mostly offshore in DMV and SVA during 1981-1990 with some activity near shore in NJ and in northern regions of SNE south of Cape Cod. In later years, fishing activity was mostly in NJ. During 1991-1995, there were no landings or effort in SVA or SNE, reduced activity in DMV, and increased activity in NJ with expansion to offshore regions. During 1996-2000, activity in DMV decreased and the fishery moved north with some activity off southern LI. During 2001-2005, landings and effort increased in DMV and SNE with some activity SNE southeast of Cape Cod.

TNMS with relatively high LPUE levels (Figure C10) were mostly off NJ and DMV in all years. During 2001-2005, LPUE levels were high in offshore NJ, with several areas of high LPUE in DMV and SNE southeast of Cape Cod.

## Important TNMS

TNMS "important" to the fishery were identified by choosing the twenty TNMS with the highest mean landings per year during 1981-1990, 1991-1995, 1996-2000 and 2001-2005 (see Assessment A. Ocean quahog, in this Report). Trends in landings, effort and LPUE were plotted (Figures C11-C13) for each to show changes in conditions within individual TNMS. Compared to less productive ocean quahog, landings, effort and LPUE were relatively high for some TNMS after many years of fishing activity.

## Fishery length composition

Taken together, port sample length data for DMV and NJ in the south indicate that the surfclam stock consisted of a wider range of sizes during the early 1980s (Figure C14 to C3-15). As expected, the port sample data for both regions appear to reflect the relatively strong 1991 year class which would have recruited to the fishery during the early and mid-1990s (see below). Although sampling levels are low and the data are difficult to interpret, smaller surfclam in landings from DMV and NJ during 2005 might be due to recruitment of the 1998 year class at age 7 (see below).

Port samplers routinely collected shell length measurements for 30 randomly selected surfclams from landings after selected fishing trips. Numbers of trips sampled and numbers measured were low (Table C8), particularly during recent years and care is
required in interpreting trends. Numbers of trips sampled is probably the best measure of the potential information in port sample length data because lengths tend to be similar for individuals from the same trip (Pennington et al., 2002).

Commercial length composition data for DMV indicate that surfclams landed during 1982-2005 were mostly $120+\mathrm{mm}$ SL during most years although smaller individuals were evident during 1992-1994 and 2005 (Figure C14). The apparent reduction in shell length during 2005 is difficult to interpret due to modest sampling (Table C8). Relatively large surfclams were landed in DMV during 1982-1985 indicating that large surfclams were more common in DMV at that time.

There were more port samples from NJ than DMV during most years (Table C8). Commercial length composition data for NJ indicate that most of the surfclams landed during 1982-2005 were at least 120 mm SL , although smaller individuals were evident during 1982-1985, 1993-1998 and 2005 (Figure C15).

Port sample data for LI are limited to 1983, 1993 and 2005 (Figure C16) and samples sizes are modest (Table C8). The data for 2005 show substantial numbers of small individuals. However, the data suggest that most of the landings in LI are at least $120+\mathrm{mm}$ SL.

Port sample data for SNE are limited to 1982-1990 (Figure C17) and samples sizes are modest (Table C8). The data suggest that most of the landings in SNE are at least $120+\mathrm{mm}$ SL.

## Fishery age composition

Fishery age composition data for DMV and NJ during 2005 (Figure C18) from port sample lengths and survey age-length keys indicate that most of the 2005 landings were ages $5+\mathrm{y}$. The strong 1992 (age 13 y in 2005) and 1998 (age 7 in 2005) year classes were important to the fishery during 2005.

Apparently strong year classes in the fishery length and age composition data for DMV and NJ may have due to low port sampling in some years and lack of age data for the commercial catch. However, survey age composition data (described later) suggest the same recruitment patterns.

Fishery age composition data for DMV and NJ do not show evidence of strong incoming year classes that would recruit to the fishery prior to 2010 (Figure C18). However, small surfclam are not selected by commercial dredges.

### 4.0 NEFSC CLAM SURVEY TREND DATA

NEFSC survey strata used to track surfclam trends (Table C9) are different than used for ocean quahog because surfclams live in relatively shallow water where ocean quahog are usually not found. After borrowing to fill holes (survey strata with no tows, see Assessment A. Ocean quahog, in this Report) a few holes remained (Table C9). Remaining holes were filled for swept-area biomass calculations but not for trend analysis using a model described below. As pointed out earlier (i.e., see Assessment A. Ocean quahog, in this Report NEFSC), NEFSC survey data are used only from surveys during 1986-2002 because of limited sampling during other years.

A cooperative surfclam survey was conducted in SVA, DMV and NJ during 2004 (Weinberg et al. 2005). It is used in calculation of swept area biomass but not for trend analysis.

## Tows with poor survey dredge performance

NEFSC developed a set of objective criteria based on sensor data used to identify NEFSC clam survey tows with poor dredge performance (see Assessment A. of this Report). These criteria were used in this assessment to identify tows in the 2005 survey with poor dredge performance.

## Dredge performance during the 2002 survey

Sensor data from the 2002 survey review were reviewed to see if dredge performance problems during 2005 also occurred during 2002. If so, the dredge performance issues might occur during most surveys.

Because of time constraints the review for 2002 was limited to a visual inspection of sensor data plots for a sample ( 213 out of 556) of stations. Details are available in Appendix C2 but the visual criteria used to judge dredge performance were the same as used in a preliminary analysis of the 2005 SSP data. In particular, manifold pressure and angle of attach were reviewed for significant deviations from "normal" values.

In general, results showed that poor dredge performance problems are likely to arise due to a number of factors that affect either manifold pressure or the angle of attack for the dredge while in operation on the bottom. The main reason for a poor dredge performance differed during 2002 and 2005 (Appendix C3). Compared to the survey during 2002, the 2005 survey had a high number of poor incidents due to manifold blockage that occurred when a screen over the pumps water intake failed and allowed small stones to lodge in the manifold nozzles. In 2002 the main problem was the dredge pump being shut off early.

It is important to realize that most of the tows with poor dredge performance would have been excluded from stock assessment analyses anyway due to haul and tow data routinely collected by the survey watch chief or chief scientist at each station. After tows with haul or gear problems were omitted, many of the remaining tows with poor dredge performance would be excluded from trend and swept area biomass calculations because they were nonrandom (Figures C19-C20).

Based on rates of occurrence during the 2002 and 2005 surveys, it was hypothesized that poor dredge performance occurs regularly during NEFSC clam surveys. Random stations during the 2002 and 2005 surveys with poor dredge performance and not otherwise were therefore used in estimation of survey trends for surfclam. In practical terms, it would have been impossible to exclude such tows consistently in all years because sensors were not used prior to 1997. As shown below, tows with poor dredge performance during 2002 and 2005 had an imperceptible effect on survey trend indices and swept area biomass estimates with the exception of the LI area during 2005.

## Survey dredge performance during depletion studies

Based on data for 2002 and 2005 surveys, the frequency of tows with poor dredge performance ${ }^{29}$ was relatively high during depletion experiments by the $R / V$ Delaware $I I$, probably because repeated tows in the same area loosened sediments which obstructed

[^26]the intake and exhaust nozzles on the survey dredge. Surfclam depletion experiments by the $R / V$ Delaware II during the 1997, 1999 and 2002 surveys were therefore not used in this stock assessment.

Based on the sampled tows and visual analysis, the frequency of tows with poor dredge performance (Table C10) during 2002 was about $15 \%$, almost twice as high as in $2005(8 \%)$. In both cases, roughly $30 \%$ of the tows with poor dredge performance were made during depletion experiments.

In contrast to trend analysis, 2005 survey stations with poor dredge performance and not otherwise were excluded from swept-area biomass calculations. The goal of swept-area biomass calculations was to obtain the best biomass estimate possible and consistency from year to year was not as important. No stations with poor dredge performance were omitted from the 2002 survey because not all stations were examined and the determination was subjective.

## Imputed survey data for remaining holes

Negative binomial GLM models were fit to survey catch data for surfclam to impute survey data for remaining holes (Table C9). Imputed data were used only in swept area biomass calculations and were not used in trend analysis due to lack of time and because the approach was experimental. Effects of imputed values on survey trends and swept-area biomass were minor because most holes had already been filled by borrowing (Table C12). Residual plots for SVA, GBK, and SNE (Figures C21-C23 suggest that the model was a reasonable approach that performed acceptably. Pending further evaluation, imputed survey data might be used in place of borrowing for future surfclam assessments.

Models used to impute missing survey data were fit in Splus using the glm.negbin() function available in the MASS library of functions for Splus and R statistical analysis software (Venables and Ripley 1997). The linear predictor had categorical year and stratum effects and the log link was employed so that year and stratum effects were multiplicative. Parameters were estimated by maximum likelihood assuming that the observed survey data were drawn from a negative binomial distribution with mean estimated by the model and a variance parameter common to all observations. The primary advantage of the negative binomial model was that it accommodated noisy data and tows with zero catch in a natural manner without adding constants and taking logs or otherwise changing the data.

A separate model was fit to tow by tow mean kg/tow (standardized using Doppler tow distances) for surfclam $120+\mathrm{mm}$ SL in each stock assessment region. All data for successful random tows during 1982-2002 were used. The imputed values used to fill remaining holes were predictions from the model for year and strata combinations missing in the original survey data.

## 2005 survey results

Based on CVs for means in stratified random sampling, the 2005 NEFSC clam survey was reasonably precise for well sampled regions (Table C11). Of particular interest, small recruit surfclams ( 50 to 119 mm SL ) were taken from near shore strata in southern DMV (Figure C4.6) where warm water probably caused extensive mortality during 1999-2004 (Weinberg 2005; Weinberg et al. 2005). However, no large fishable surfclams ( $120+\mathrm{mm}$ ) were found in near shore strata off southern DMV (Figures C24-

C25). See NEFSC (2005) for a summary of survey station locations and catches during the 2005 NEFSC clam survey.

## Survey trends

Survey trend data (Figures C26-C28) were more variable for small surfclams than for large surfclams. Based on survey trend data, fishable biomass ( $120+\mathrm{SL}$ ) declined in southern regions SVA, DMV and NJ. The decline in SVA was gradual beginning in the mid-1980s. The declines in DMV and NJ were relatively rapid beginning in the mid1990s. Fishable biomass in LI may have increased gradually after 1982 but the survey data are variable and difficult to interpret.

Recruitment indices 2005 were at or near record lows for all regions surveyed with the exception of LI and GBK which was not surveyed in 2005 (Figures C26-C27). During the 2002 survey, recruitment in GBK was relatively high.

With the exception of LI during 2005, tows with poor dredge performance during 2002 and 2005 had an imperceptible effect on estimated trends in fishable biomass (Figure C28).

Year effects and the 1994 survey
Trends in NEFSC survey data (Table C11) for small recruit surfclams (mean n tow-1, $50-119 \mathrm{~mm} \mathrm{SL}$ ) and large fishable surfclams (mean kg tow, $120+\mathrm{mm} \mathrm{SL}$ ) showed some evidence of year effects when estimates for the same year and region increased or decreased together (Figure C26). Year effects in NEFSC clam survey may be due to changes in survey dredge equipment or protocols between surveys (NEFSC 2003).

Based on survey trend data, it was decided to include the 1994 survey in all analyses for surfclam. In contrast, previous surfclam assessments (NEFSC 1998; 2000; 2003) included 1994 survey data in graphics but excluded the data from swept area biomass and other analyses because of hypothesized year effects that may have increased catch rates. In particular, the voltage supplied to the pump on the dredge was reportedly set at 480 V , rather than 460 V as specified and higher voltage during the 1994 survey may have increased catch rates (NEFSC 2003). However, based on additional survey data there is insufficient evidence of a year effect during the 1994 survey for surfclam. Moreover, field tests with the survey dredge operating with 460 and 480 V were inconclusive (J. Weinberg, pers. comm.). Additionally, a comparison of tows during the 2002 and 2005 survey with good and poor dredge efficiency suggested that surfclam catches were not sensitive to dredge performance (Appendix C3).

The decision to use 1994 survey data for surfclams in stock assessment analyses does not apply to ocean quahogs. Evidence for a strong year effect due to high voltage appears stronger for ocean quahogs (see Assessment A. in this Report).

## Survey length and age data

Survey length composition data show a wide range of lengths for surfclam in SNE, LI, and NJ with relatively few large surfclam in DMV and a relatively narrow range of lengths in GBK (Figures C30-C34). Survey length data for LI during 2005 was too variable to be interpreted. It may be possible to track a recruitment event in the survey length data for LI beginning in 1983. Length data for SVA are scant.

Survey age composition data for NJ and DMV show the strong 1992 and 1998 year classes relatively consistently and clearly (Figure C34b). During 2005 these two
year classes dominated the population as 7 and 13 year-olds. There is some evidence of a recruitment event in the age composition data for age 2 surfclams in DMV during 2005.

### 5.0 STOCK BIOMASS AND FISHING MORTALITY (TOR-2)

Efficiency corrected swept area biomass estimates were based on NEFSC and cooperative clam survey data for 1997, 1999, 2002, 2004 and 2005 and cooperative depletion experiments. They are a key source of information about the scale (magnitude, thousand mt ) of surfclam biomass during recent years in this assessment.

Efficiency corrected swept area estimates are relatively direct, model-free and independent estimates of biomass and fishing mortality. Surfclams have proven difficult to model in some cases (e.g. NEFSC 2003) and it is useful to have another method available for estimating recent biomass and fishing mortality. Fishing mortality, in particular, can be estimated on a regional basis as the ratio of catch and efficiency corrected swept area biomass. Fishing mortality rates are low for surfclams and the June survey occurs when the stock is near the average annual level so that the ratio of catch and biomass gives nearly the same result as solving the catch equation exactly. Swept area biomass and fishing mortality estimates were not made for years with surveys prior to 1997 because no sensor-based tow distance data were available.

NEFSC clam survey trend data are the main source of information about trends in fishable biomass and recruitment since 1982. Survey data (mean $\mathrm{kg} /$ tow, based on sensor tow distances) for trend and swept area analyses were from random stations with no problems recorded on standard survey logs. Some survey stations with poor dredge performance identified using sensors during 2005 were omitted from swept area biomass calculations. As described above, negative binomial GLM models were used to impute missing survey data used to fill remaining holes in NEFSC data.

The KLAMZ delay-difference stock assessment model was used to make estimates for surfclams in DMV, NNJ and for the entire stock. The assessment model is advantageous because it estimates long term biomass and fishing mortality levels during 1982-2005, "balances the books" to ensure that all assumptions can be reconciled, and smoothes out measurement errors in swept area biomass and survey trend data. The KLAMZ model was not applied to SNE, LI and GBK in this assessment because the survey data are difficult to interpret and very little fishing has occurred in northern regions.

In the previous assessment (NEFSC 2003), the KLAMZ model was used only for DMV because it did not give reasonable results for southern and northern New Jersey (which were modeled separately). The KLAMZ model and data used in this assessment involve improvements that enhance model performance. In particular, the southern and northern New Jersey regions are combined in this assessment to form the NJ region with relatively precise survey data. Additional survey data for 2004 and 2005 are available and show clear trends over the last decade.

All of the methods for estimating surfclam biomass and fishing mortality levels and calculating variances are described in Assessment A. Ocean Quahogs, in this same Report. A few differences in methodology for surfclams are described below where relevant.

## Survey and commercial dredge efficiency

As for ocean quahogs (in Assessment A. Ocean Quahogs of this Report), the best estimate of survey dredge efficiency for surfclams in this assessment was the median of estimates from all available depletion studies (Table C13). In particular, the best estimate of efficiency for commercial dredges was the median $E=0.765$ (mean $0.704, \mathrm{CV}=0.081$, $\mathrm{n}=19$ ) and the best estimate for the NEFSC survey dredge was $e=0.226$ (mean=0.262, $\mathrm{CV}=0.17, \mathrm{n}=16$ ).

All commercial efficiency estimates for surfclam in this assessment were from Rago et al.'s (2006) "Patch" model fit to data from depletion studies by commercial vessels. Survey dredge efficiencies were estimated for depletion experiments with setup tows by $R / V$ Delaware II during NEFSC clam surveys. In contrast to ocean quahog and as described above, depletion studies carried out entirely by the $R / V$ Delaware II were not used because of problems with survey dredge performance during repeated tows in the same location. A variety of ad-hoc estimators for survey dredge efficiency used by NEFSC (2003) for surfclams were not used in this assessment because they have unknown statistical characteristics and were not necessary.

Eight new depletion studies have been carried out since the last assessment, three during 2004 and five during 2005 (Table C14). Additionally, it was necessary to reanalyze depletion experiment data from fourteen depletion experiments during 19971999 so that consistent methodology and corrected estimators were used in all cases.

## Assumed length at full recruitment

The most important difference in estimating dredge efficiencies for surfclam in this assessment and in the previous assessment was the assumed length at full recruitment to the commercial gear used in each depletion experiment. Surfclams were assumed in this assessment to be fully recruited to commercial gear used in depletion experiments at $150 \mathrm{~mm} \mathrm{SL} .{ }^{30}$ Elsewhere, in mortality and biomass calculations for this assessment, surfclams are assumed to recruit to the commercial fishery and become fishable at about 120 mm SL. However, full recruitment is likely to occur at some larger size.

Depletion experiments for surfclams included vessels that specialize in surfclam (e.g. F/V Jersey Girl in Table C14) and vessels that specialize in ocean quahog (e.g. F/V Lisa Kim). Gear on quahog vessels is designed to catch relatively small ocean quahog efficiently. Thus, surfclams likely recruit to gear on ocean quahog vessels at a smaller size than gear used on surfclam vessels. However, it was important too choose an assumed length at full recruitment that was high enough to assure full recruitment to both types of gear in all experiments. A single length criterion was important for the sake of efficient data processing and consistency of surfclam density estimates.

NEFSC (2005) used 90 mm SL as the assumed size at full recruitment for ocean quahog because commercial selectivity at that size was at least $85 \%$ at 90 mm SL based on a commercial fishery selectivity curve. No directly estimated selectivity curves are available for surfclams. However, a "relative" selectivity curve that relates catches in commercial surfclam gear to catches in the NEFSC survey dredge indicates that $85 \%$ relative selectivity occurs at $145-150 \mathrm{~mm}$ SL (Figure C30 in NEFSC 2004). A review of

[^27]length data from surfclam depletion experiments with setup tows indicated that 150 mm SL would suffice as the assumed size of full selectivity in all experiments.

The disadvantage in choosing a relatively large assumed size at full recruitment was that data from the SC2002-4, SC2004-3 and SC2005-6 depletion experiments were not useable. In these experiments, catches of surfclams $150+\mathrm{mm}$ SL were either zero or too low and variable.

## Relationships between efficiency and other variables

There were no clear relationships between Patch model estimates and environmental variables such as depth and sediment size (Figure C35 and C36). With one exception, there were no clear relationships among Patch model estimates themselves (Figure C35 and C36).

The apparent negative relationship between estimates of efficiency and initial surfclam density from the Patch is potentially important (Figure C36). However, the pattern is readily explained as an artifact of the natural statistical correlation between the two parameters in the Patch model. Sites for depletion experiments are chosen to have relatively high surfclam densities. If efficiency decreases at high surfclam densities and experiments are conducted at sites with high density, then mean efficiency for the stock as a whole (in areas of high and low density) might be underestimated. If efficiency is underestimated, then stock biomass might be overestimated and fishing mortality under estimated.

As described in Rago et al. (2006) and illustrated by a typical bivariate likelihood profile for density and efficiency estimates from the Patch model (Figure C37), uncertainty in initial density and efficiency estimates take the form of an elongated "banana" shaped region so that lower estimates of initial density are associated with higher estimate of efficiency and vice-versa. In other words, sets of parameters with density low and efficiency high tend to fit the data from a depletion experiment as well as sets with density high and efficiency low. This type of statistical correlation is common in nonlinear parameter estimation (Bard 1974). In linear regression modeling, it takes the form familiar statistical correlation between estimates of the slope and intercept of the regression line.

A simple simulation analysis using linear regression and a simulated Leslie-Davis depletion experiment showed the same relationship between efficiency and density estimates, although no relationship was included in the simulation scenario. The Patch model is quite similar to a linear regression problem because, in effect, it is the result of applying Leslie-Davis depletion models to a number of depletion experiments sites simultaneously (Rago et al. 2006). Leslie-Davis depletion models were fit originally by simple linear regression (Ricker 1975).

## Sensitivity of Patch model estimates to smoothing position data

As described in Assessment A. Ocean quahogs, in this Report, position data from depletion experiments was smoothed and interpolated prior to use in the Patch model. NEFSC (2006) carried out a number of analyses to determine the sensitivity of Patch estimates to assumptions and procedures but did not consider smoothing.

Procedures and equipment improved steadily in each survey. Precision of position data was relatively low for 1997 depletion experiments because Loran was used to measure location (accuracy 30-40 ft) and position data were recorded at relatively long
time intervals (e.g. 1 minute). In later years, more precise differential GPS was used to measure location to a precision of about 6-9 ft and at shorter intervals of 1-6 seconds.

To accommodate differences in precision of location data among depletion experiments, the Patch model was fit with and without smoothing to data from one surfclam depletion experiment in each survey year. Results (Table C15) show that smoothed data produces higher estimates of initial density and lower estimates of dredge efficiency than unsmoothed data. Area swept during each depletion tow decreased by 1$20 \%$ when using smoothed data (Table C15).

## Building a bridge

Assessment A. Ocean quahogs, of this Report (see Tables A14-A15) evaluated effects of the many changes made in estimation of dredge efficiency for ocean quahog. Results from those analyses for ocean quahog are probably also applicable to surfclam.

As with ocean quahog and with the exception of experiments in 2002, revised efficiency estimates for surfclam were lower and more precise (lower CVs) than estimates previous estimates (Table C16). However, care is required in making comparisons with efficiency estimates in NEFSC (2003) because previous estimates were from a variety of estimation procedures. In addition, previous estimates from the Patch model were usually made under different assumptions, data for different sizes of surfclam were included and less accurate formulas may have been used.

## Efficiency corrected swept area biomass

The best estimate of survey dredge efficiency ( $e=0.226$ ) was used to estimate efficiency corrected swept area biomass (Table C17) and fishing mortality (Table C18) for surfclams 120 mm SL is 1997, 1999, 2002 and 2005.

## 2004 Cooperative Survey

Additional information was available from a cooperative survey carried out during 2004 by the F/V Lisa Kim in SVA, DMV and NJ (Weinberg et al. 2005). Sweptarea biomass estimates in Weinberg et al. (2005) were recalculated using the median commercial dredge efficiency ( $E=0.714$, Table C19) from six depletion experiments by the FV Lisa Kim during 2004-2005 (Table C14). The updated calculations excluded some nonrandom tows that may have been used inadvertently by NEFSC (2003).

Cooperative 2004 survey analyses in this assessment used catch data for surfclams $120+\mathrm{mm}$ SL (all sizes in the fishable biomass) because the $F / V$ Lisa Kim normally targets ocean quahog and is equipped to catch relatively small commercial size ocean quahog, which are smaller than commercial size surfclam. As described above, the assumed size at full recruitment was 150 mm SL in other analyses because commercial vessels were used in some experiments that target surfclams use gear that retains larger clams. Survey length composition data from the depletion experiments indicated that surfclams probably recruited to the dredge on the F/V Lisa Kim at about 120 mm SL.

Results from the 2004 survey (Table C20) confirmed downward trends in biomass evident in biomass estimates for DMV and NJ based on NEFSC surveys during 19972005 (Table C21; Figure C38). In particular, the 2004 estimates from the cooperative survey were nearly intermediate between biomass estimates from the 2002 and 2005 NEFSC surveys. The 2004 survey did not cover all strata in SVA and catch rates for SVA were too variable to be used in estimating biomass (Figure C38).

## KLAMZ modeling

KLAMZ delay-difference models for surfclam biomass dynamics were similar to those used by in the Ocean quahog Assessment (see Assessment A. of this Report) for ocean quahog. ${ }^{31}$ A few changes were made to model surfclams more realistically. These changes involved configuration of survey trend data, assumptions about recruitment, growth patterns that changed over time, and application to the stock as whole as well as to individual regions. Surfclams require slightly different modeling approaches because more data are available, surfclams are inherently more productive and their population dynamics are more variable, surfclams grow relatively quickly, growth varies over time, surfclams have a higher assumed natural mortality rate ( $M=0.15 \mathrm{y}^{-1}$ instead of $0.02 \mathrm{y}^{-1}$ ), and recruitment patterns are substantially different. Many of these factors appear to be influenced by density dependent factors (Weinberg 1998), oceanographic conditions and bottom temperatures in particular (Weinberg 2005).

The most important challenges in modeling surfclams stem from variability in NEFSC clam survey data for recruits and fishable sizes, and lack of survey data between triennial NEFSC clam surveys. In a nutshell, recruitment trend data change too rapidly to be readily tracked by the triennial survey data. LPUE trend data are available and can be compared to model results but were not used in fitting KLAMZ models for surfclams due to well known problems relating commercial catch rates and trends in stock biomass (Hilborn and Walters 1992). Catch data used in KLAMZ models for surfclams included discards that occurred prior to 1993 when size limits were used to manage the fishery (Table C2).

Despite problems, a number of factors enhance the utility of the KLAMZ model for surfclam. Most importantly, direct estimates of stock biomass based on depletion studies and swept area estimates are easily incorporated in the assessment model. The KLAMZ model is flexible and has a number of features that can be used to take advantage of various aspects of surfclam biology. Landings data for surfclams are relatively accurate because of accounting procedures inherent in the ITQ fishery management program. Survey data for surfclams include CVs that characterize sampling variability and that can be used to determine when the model fits the survey data "too well" (i.e. better than could be expected based on the inherent precision of the data). Auxiliary information is available for many important parameters (e.g. survey dredge efficiency and swept area biomass and growth). Surfclams are relatively long lived ( $\sim 35$ y) and expected rates of change in fishable stock biomass are lower for relatively longlived organisms.

Year effects and correlated measurement errors (the same year effect in survey data for recruits and fishable size groups in the same year) are a concern in using survey data for surfclams. Simulation analyses have not been carried out using the KLAMZ model, but detailed simulation analyses with the abundance-based Collie Sissenwine model (ASMFC 2006) which is similar to KLAMZ showed that model performance (mean squared error, bias and variance) actually improved when survey data for recruits and fishable size groups had strong correlated year effects.

[^28]
## Growth curves

Growth is a key part of biomass dynamics in the KLAMZ delay-difference model. Survey data for surfclams in KLAMZ models (particularly for new recruits) are calculated based on assumptions about growth.

The Schnute-Deriso delay difference equation in the KLAMZ model (Schnute 1985) uses a version of the von Bertalanffy model for growth in weight with two parameters. In particular, $\rho=e^{K}$ where $K$ is from a von Bertalanffy model for weight, and $J_{t}=W_{k-1, t} / W_{k, t}$, where $W_{k, t}$ is predicted weight at age $k$ when recruitment occurs based on the growth curve for year $t$. The von Bertlanffy parameters $W_{\max }$ and $t_{0}$ are implicit in $J_{t}$. In delay-difference model calculations (Schnute 1985), the parameters $J_{t}$ may change over time but $K$ is constant in all years.

Survey mean length at age data for NJ and DMV in each survey (Figure C2) were converted to mean weights at age in each survey by applying region specific lengthweight relationships (Table C1). The growth curves used different $W_{\max }$ and $t_{0}$ parameters for 1982-1992 and 1994-2005, but used the same $K$ parameter in all years (Table C22). Growth parameters for NJ were used also in modeling the whole stock.

## Survey indices

NEFSC clam survey data in the KLAMZ model were for recruit (Table C23) and fishable size groups (Table C11). The recruit index was mean $\mathrm{kg} /$ tow for surfclam in the survey that were 120 to $L_{k+1} \mathrm{~mm} \mathrm{SL}$, where $L_{k+1}$ is the predicted size at age $k+1$ and $k$ is the predicted age at recruitment $\left(L_{k}=120 \mathrm{~mm} \mathrm{SL}\right)$ based on a growth curve. The fishable index was survey mean kg/tow for surfclams $120+\mathrm{mm}$ SL. Recruit trend data were assumed to track trends in the biomass of new recruits. Trend data for fishable surfclams were assumed to track trends in total fishable biomass (new recruits plus survivors from the previous year). Surveys were assumed to occur in the middle of the year because the NEFSC clam survey is carried out during late May-early July.

As described above, survey data for surfclams 120 to $L_{k+1} \mathrm{~mm}$ SL were used in both the recruit and fishable biomass trend indices. This strategy was intentional and meant to link the relatively noisy recruit and more stable fishable survey data indices in the model, to reduce potential problems stemming from uncertainty about where to split the index for fishable biomass, and to help insure that the survey scaling factor for both recruit and fishable indices would be about the same. In practical terms, it had little effect on the survey data themselves because recruit $\mathrm{kg} /$ tow was small relative to $\mathrm{kg} /$ tow for the remaining fishable size groups.

NEFSC (2003) used a more complicated system of survey trend data for prerecruits, recruits and remaining fishable size groups. Fishable sizes were 100+ or $120+\mathrm{mm}$ SL, depending on area. Prerecruit size groups were $L_{k-1}$ to either 100 or 120 mm SL based on region specific von Bertalanffy growth curves. The prerecruit index was lagged in the model by one year so that data collected in year $t$ would be used in the model to estimate recruitment in year $t+1$. The prerecruit index was not used in this assessment because it is highly variable for surfclams with noisy trends that are difficult to resolve given the rest of the survey and catch data in the model.

For convenience in interpreting model results, survey mean $\mathrm{kg} /$ tow data for fishable surfclams in the entire stock were scaled up to approximate efficiency corrected swept area biomass before use in the KLAMZ model. The scaling factor was the average ratio of the survey data and efficiency corrected swept area biomass during 1995-2005
surveys (see below and Table C25). With this adjustment, the survey scaling factors for fishable biomass trends estimated in the KLAMZ model are expected to be close to one. The adjustment to the survey data did not affect biomass or fishing mortality estimates.

## Survey dredge efficiency and swept-area biomass

Following NEFSC (2003), efficiency corrected swept area biomass estimates were included in the assessment model as a measure of scale but not as measures of trend. In fitting the model, the likelihood of the estimated scaling parameter for swept area biomass was calculated based on a lognormal prior distribution with mean 1.0 and arithmetic $\mathrm{CV}=0.5$. The relatively large CV means that the prior information about the scaling parameter was relatively "weak". However, experience shows that the prior information tends to have a strong impact when survey data are limited and there is little other information in the model data about biomass scale.

## Recruitment assumptions

Following NEFSC (2003) surfclam recruits were estimated in the KLAMZ model as a random walk with steps constrained by a variance parameter. A smooth, random walk process is probably not ideal from a biological perspective because of the possibility of strong year classes in surfclams but the approach was necessary because of the lack of annual recruitment data. The random walk approach keeps the recruitment estimate in year $t$ at the same level as in year $t-1$, unless there is a good reason in terms of goodness of fit to change it. For surfclams in the KLAMZ model, the random walk approach was used primarily to fill gaps in information due to not having a recruit index for each year, to avoid excessive variation in recruitment and to ensure that some recruitment was estimated for each year.

In modeling surfclam population dynamics with random walk recruitment, it is important to control the "random walk recruitment variance" $\sigma_{R}^{2}$ (NEFSC 2003) which measures variability in the size of successive steps taken during the random walk (i.e. variance in $\left[\ln \left(R_{1} / R_{2}\right), \ln \left(R_{2} / R_{3}\right), \ln \left(R_{3} / R_{4}\right)\right.$, etc.], where $R_{t}$ is the recruitment estimate for year $t$ ). As $\sigma_{R}^{2}$ approaches zero, recruitment estimates become smooth and tend towards a constant value with no changes from year to year. As $\sigma_{R}^{2}$ becomes large, estimated recruitments will change randomly and more widely from one year to next.

Following NEFSC (2003), initial KLAMZ model runs assumed a $20 \%$ CV for steps in the random walk so that $\ln \left(\sigma_{R}^{2}\right)=\ln \left(0.2^{2}\right)$. The constraint was relaxed gradually in subsequent runs until the model was just able to fit the survey data without pattern in residuals. In final runs, $\ln \left(\sigma_{R}^{2}\right)=\ln \left(0.3^{2}\right)$ for NJ and the entire stock, and $\ln \left(\sigma_{R}^{2}\right)=$ $\ln \left(0.35^{2}\right)$ for DMV. In each case, the CV for fit to the survey data (residual CV) was compared to CVs for the actual survey data to determine if $\sigma_{R}^{2}$ was too large and the model was fitting the survey data more closely than could be expected based on the precision of the survey data. The goal was basically to find the simplest model (fewest effective recruitment parameters) that would adequately explain the survey data for surfclam. Choices were subjective but had only modest effects on biomass and fishing mortality estimates for surfclam, because many different recruitment patterns imply similar biomass and recruitment levels.

## Results-whole stock

Survey data for the entire stock in the KLAMZ model were filled as described above. However, no provision was made for filling remaining holes that could not be filled by borrowing (Table C9). Mean surfclam densities for strata with data (original or filled) were used to compute the weighted mean density for the stock as a whole (i.e. strata with no data were ignored in computing the mean density for the stock as a whole). However, the mean density for the stock as a whole was applied to the entire stock area, which included the area of strata with no data. The effects of remaining wholes were reduced in whole stock runs because remaining wholes were a relatively small proportion of the total number of strata and total area of the stock.

The KLAMZ model fit survey biomass trend data reasonably well although the fishable biomass trend datum for 1994 was not completely reconciled in the model fit (Figure C40). The model fit the recruit index better than the fishable biomass index, although the latter was more precise based on survey CVs. LPUE and swept area biomass trends did not affect model estimates, but estimated biomass trends from the model were similar to trends in LPUE after 1999 and to trends in swept area biomass for in all years.

The survey scaling parameter for the scaled fishable biomass index was $Q=1.26$ and reasonably close to one. The survey scaling parameter for efficiency corrected swept area biomass was $Q=0.99$ indicating that the trend data, landings and model estimates were compatible with the prior information about $Q$ for efficiency corrected swept area biomass estimates.

Model results suggest that surfclam biomass increased from 1981-1997 to record high levels due to high surplus production (relatively good recruitment and fast growth rates) which occurred during the mid 1980s and early 1990s (Table C24 and Figure C41). Surplus production declined steadily after 1993 as recruitment declined, the stock aged and growth rates slowed. Surplus production was negative after 1997 while stock biomass declined steadily. By 2005, stock biomass had declined to about the same level as in 1986-1992 but was still relatively high in historical terms. Fishing mortality rates were much lower than natural mortality and probably inconsequential during 1981-2005.

Bootstrap analysis (2000 iterations) indicated a tendency towards negative bias in biomass and fishing mortality estimates during peak recruitment years, but good model performance and little bias overall. CVs and confidence intervals from bootstrapping indicate that biomass and fishing mortality estimates were reasonable precise, particularly for recent years (Table C24; Figures C42-C43), probably due to the swept area biomass data for 1997-2005. Recruitment was estimated less precisely than biomass and fishing mortality (Table C24; Figure C44). The model did not completely converge during a substantial fraction of bootstrap runs (roughly $50 \%$ ), due to uncertainty in estimated recruitments (Table C24). In other words, a range of recruitment patterns probably explained the survey data equal well.

## Results-DMV and NJ

The KLAMZ model for DMV fit survey index data quite well (Figure C45). The model for NJ fit reasonably well although the fishable biomass indices for NJ during 1994 and 1997 were not reconciled (Figure C46). Survey scaling factors for scaled fishable biomass trends and efficiency corrected swept area biomass were reasonably close to one in all cases.

Model results for DMV indicate that biomass declined continuously from relatively high levels during the early 1980s due to declining recruitment, slow growth, and surplus production levels that were usually negative (Figure C47). Model results for NJ were similar to results for the whole stock but biomass declined more steeply during recent years to lower levels during 2005 (Figure C48). Fishing mortality appears to have been a minor factor in both areas during 1981-2005 (Figures C47-C48).

## Stock biomass by region

Average ratios for survey data (Doppler standardized) and efficiency corrected swept area biomass were calculated for each region (Table C25) and used to rescale survey trend data to approximate swept area biomass levels (Table C23). The proportions of swept area biomass in each region were used to prorate fishable biomass estimates from the KLAMZ model for the entire stock during years with NEFSC clam surveys into regional components. Results clearly show the shift over time in biomass from southern to northern regions (Figures C49 to C50).

## Recruitment parameters

Recruitment estimates for surfclam from the KLAMZ model were made with limited survey data and are complicated to interpret. Under these conditions, recruitment estimates for surfclam should probably be regarded as "nuisance" parameters of less interest than biomass and fishing mortality estimates. As nuisance parameters, recruitment estimates basically amount to adjustments in the KLAMZ model that implicitly account for model misspecification, survey noise, survey year effects, changes in recruitment, natural mortality and variability in growth not explicitly included in the modeling framework.

## Proportions of total fishable biomass at various density levels

As described in the first assessment in this Report (A. Ocean quahogs), best biomass estimates and survey data were combined to partition best biomass estimates into components found in areas with relatively high and low biomass density levels. Biomass density is important to profitability of the ocean quahog fishery because it determines commercial catch rates. Biomass density was measured as survey catch per tow (fishable $\mathrm{kg} / \mathrm{tow}$ ) because commercial catch rate data for random locations and the entire stock area were not available.

Results (Table C26) show reductions in stock within high density areas in the southern DMV and SVA regions. During 2005 (Table C27), the largest component (29\% or 47 thousand mt meats) of total fishable stock biomass was on GBK in the highest (25+ $\mathrm{kg} /$ tow) biomass density category. In contrast, stock biomass levels in density categories larger than 10 kg /tow were low for other regions.

### 6.0 BIOLOGICAL REFERENCE POINTS (TOR-3)

According to the Surfclam and Ocean Quahog FMP, overfishing occurs whenever the fishing mortality rate on the entire stock is larger than $F_{M S Y}$. The stock is overfished if total biomass falls below $B_{\text {Threshold }}$ (estimated as $1 / 2 B_{M S Y}$ ). When stock biomass is less than the biomass threshold, the fishing mortality rate threshold is reduced from $F_{M S Y}$ in a linear fashion to zero.

The current best proxy for $F_{M S Y}$ is $F=M=0.15 \mathrm{y}^{-1}$. The proxy for $B_{M S Y}$ is onehalf of the estimated fishable biomass during 1999 which was estimated to be 1,460 thousand mt in this assessment based on KLAMZ model results for the entire stock. Revised biomass reference points are higher than previous values (see table below) because of new information about the efficiency of the dredge used in NEFSC clam surveys.

| Reference Point | Last assessment | Revised |
| :---: | :---: | :---: |
| $F_{M S Y}$ | $M=0.15 \mathrm{y}^{-1}$ | Same |
| $B_{1999}$ | 1,460 thousand mt <br> meats | 1,799 thousand mt <br> meats |
| $B_{M S Y}=1 / 2 B_{1999}$ (target) | 730 thousand mt <br> meats | 900 thousand mt <br> meats |
| $B_{\text {Threshold }}=1 / 2 B_{M S Y}$ | 365 thousand mt <br> meats | 490 thousand mt <br> meats |

Status determinations by comparisons of biomass estimates and biomass reference points are almost unaffected by new information about dredge efficiency because the changes in biomass estimates and the $B_{M S Y}$ proxy "cancel out" when current biomass is compared to or divided by the $B_{M S Y}$ proxy (Figure C51). Comparison of fishing mortality estimates and the $F_{M S Y}$ proxy are more sensitive because fishing mortality estimates depends on dredge efficiency but the $F_{M S Y}$ proxy does not (Figure C52).

Fortunately, conclusions in this assessment about fishing mortality and reference points are robust because fishing mortality rates for the stock are relatively low. In particular, conclusions about stock status would not change unless either the mortality estimate or threshold was changed by 7 fold (Figure C52).

## Critique

Current reference points for surfclams suffice for use in this assessment because surfclam biomass is relatively high (at near average levels) and fishing mortality is low. However, biomass referenced points should be reconsidered the next time the stock is assessed.

Use of $1 / 2 B_{1999}$ as a proxy for $B_{M S Y}$ implicitly assumes that the stock was at carrying capacity during 1999. The carrying capacity assumption should be reevaluated based on the longer time series of data that are currently available. In addition, it may be useful to consider possible climate change effects on $B_{M S Y}$ and $F_{M S Y}$ proxies as evidenced by loss of surfclams in the south near the coast of the Delmarva Peninsula (Weinberg 2005).

### 7.0 STOCK STATUS (TOR-4)

The Atlantic surfclam stock is not overfished and overfishing is not occurring. Estimated fishable stock biomass during 2005 ( $120+\mathrm{mm}$ shell length, SL) was 1,170 thousand mt meats, which is above the management target of $1 / 21999$ biomass $=900$ thousand mt meats (Figure C51). Estimated fishing mortality during 2005 was $F=0.0192$ $\mathrm{y}^{-1}$, which is below the management threshold $F_{M S Y} \cong M=0.15 \mathrm{y}^{-1}$ (Figure C52).

### 8.0 PROJECTION METHODS (TOR-5)

For the first time, a fully integrated assessment model, variance estimation and stochastic projection approach was used to provide example projections for surfclam stock biomass and fishing mortality. In particular, simulation runs for projection analysis were carried out using the same delay difference equation as used in the KLAMZ model and were initialized exactly as in the last year of each bootstrap run.

Projections can be made for assumed levels of constant fishing mortality or assumed constant catch levels, and can be carried out for time periods of any length. In projections for surfclams with assumed levels of catch, likely levels of incidental mortality should be considered and probably included. For example, constant quota levels can be increased by $12 \%$ to accommodate incidental mortality and to obtain a more realistic estimate of fishery impacts. A large number of individual stochastic simulation runs (e.g. 1000) should be carried out in projection analysis. Normally, the number of simulation runs is the same as the number of bootstrap runs because bootstrap results are saved for later use by the projection software. It is possible, however, to make more than one projection from each bootstrap run.

Each simulation run in the projection analysis starts with the terminal conditions estimated in one bootstrap run. Thus, uncertainty about current stock biomass, age structure, recent recruitments and other factors is included in the projection analysis.

Uncertainty in future conditions is included by simulating random future recruitments. For surfclams, random recruitments $\left(R_{t}\right)$ were chosen to mimic a random walk with user specified mean and lag-1 autocorrelation. Projected recruitments were modeled as a random walk to match assumptions in the stock assessment model. As described above, the random walk recruitment assumption in the stock assessment model was pragmatic and may not be ideal from a biological perspective. The algorithm for surfclams in this assessment was:

$$
\begin{aligned}
\sigma & =\sqrt{\ln \left(C V^{2}+1\right)} \\
b & =\frac{\sigma^{2}}{2} \\
s & =\sqrt{1-\rho^{2}} \\
j_{t} & \sim N(0,1) \\
\delta_{t} & =s j_{t} \\
\gamma_{t} & =\rho \gamma_{t-1}+\delta \\
R_{t} & =\overline{\mathrm{R}} \mathrm{e}^{\gamma_{t} \sigma-b}
\end{aligned}
$$

where $j_{t}$ is drawn from the standard normal distribution, $\rho$ is the lag-1 autocorrelation for successive $\log$ scale recruitments [i.e. the correlation of $\ln \left(R_{t}\right)$ and $\ln \left(R_{t+1}\right)$, specified by the user], $\sigma$ is the standard deviation of log scale recruitments based on an arithmetic scale CV (specified by the user), $\bar{R}$ is the mean arithmetic recruitment (specified by the user), and $b$ is a bias correction factor. The term $\gamma_{\mathrm{t}}$ is normally distributed with mean zero, standard deviation 1.0 and lag-1 autocorrelation $\rho$. At the end of the projection
analysis, the model calculates the means and CVs for biomass, recruitment, catch and fishing mortality at the beginning of each year.

Based on the KLAMZ model run for the entire stock, $\rho=0.72, \mathrm{CV}=0.53$, and $\bar{R}=$ 121 thousand mt in example projection calculations. The simulation runs were for 20052015 ( 10 y beyond the last year in the KLAMZ model).

## Procedures for setting TAL and TAC levels

It is not necessary to describe approaches for setting TAC or TAL levels in the surfclam fishery because it is managed using constant quota levels.

### 9.0 EXAMPLE PROJECTIONS (TOR-6)

Example projections were carried out assuming the following conditions during 2006-2015: i) constant fishing mortality $=0.15$; ii) constant landings at the minimum quota level $=1.85$ million bu; iii) constant landings at mean level during 2003-2005; and iv) constant landings at the maximum quota level $=3.4$ million bu. In each case, landings in bushels were converted to meat weights and increased by $12 \%$ to account for potential incidental mortality during fishing.

Results (Table C28 and Figure C53) indicate that current downward trends in biomass will persist during the next few years because of the tendency for runs of good and bad recruitment in surfclams. Declines are largest for the $\mathrm{F}=0.15$ scenario. Results for the status quo and maximum quota scenarios are very similar.

Projected biomass levels out by about 2015 in all scenarios. However, CVs are very large in all years and, in particular, larger than $250 \%$ after 2008. The high CV levels indicate very high uncertainty in projected results, particularly after 2008.

### 10.0 RESEARCH RECOMMENDATIONS (TOR-7)

Research recommendations from the previous assessments are listed below (not in priority order).
i) Consider using year-, region- or episodic natural mortality rates. This was discussed in the working group but deferred until a later assessment when the necessity for incorporating this feature might be more pressing.
ii) Develop a forward casting age-structured, numbers-based stock assessment model. This work is in progress for sea scallop, ocean quahog and surfclam. In the interim, the KLAMZ model is implicitly age structured and numbers based, although it does not make full use of survey and fishery age or length data. NEFSC convened an age readers workshop during 2006 (Jacobson et al. 2006) to address questions about age data and results will be useful in formulating the new model. NEFSC has begun to characterize variability in survey length data and the results are expected to be useful in modeling as well.
iii) Reconcile survey trends for pre- and new- recruits relative to trends in survey data for older recruits. Pre-recruit survey indices were not used for modeling in this assessment because they are too variable. Survey data procedures for modeling were redesigned to ease interpretation.
iv) Reconcile survey data with consistently declining trends in LPUE during the last decade. Recent trends in survey and LPUE data were similar in this assessment for southern regions, where fishing is heaviest, and for the stock as a whole.
v) Focus on analysis of declining LPUE trends and examine new approached for describing fishing power among commercial clam vessels. This issue was addressed by standardizing LPUE data in models that included individual vessel effects. Thus, it was not necessary to characterize fishing power based on GRT, horsepower, etc.
vi) Collect commercial age and length data to monitor and predict recruitment and for use in length and age structured models. Length data but no age data are currently being collected from port samples. Sampling rates for length data should be increased particularly for new northern fishing grounds. All available survey age, length and commercial length data were used at least qualitatively in this assessment to characterize and predict recruitment.
vii) Reexamine coefficients used to convert commercial catches in bushels to meat weights. No progress.
viii) Consider using a sensor that tracks dredge position, rather than the ships position, during surveys and depletion studies. New acoustic sensor equipment was tried experimentally during the 2005 survey but with poor results.
ix) Conduct surveys more frequently than every three years in critical areas. $A$ cooperative survey in the SVA, DMV and NJ areas was carried out during 2004, in the interim between the 2002 and 2005 NEFSC clam surveys.
x) Select a new set of fixed stations in unfished areas to monitor dredge efficiency changes between surveys. Fixed station analysis was abandoned in this assessment due to variable environmental conditions that may affect density in unfished areas.
xi) Consider new technological methods that rely less heavily on estimating dredge efficiency. No progress.
xii) Consider new methods to estimate variability in the spatial distribution of biomass. All depletion studies were reanalyzed for this assessment producing estimates of the negative binomial parameter $k$, which measures spatial patchiness in the density of surfclams within depletion study areas. However, this topic is of relatively low importance.
xiii) Continue to bring outside experts to Invertebrate Working Group meetings. One outside expert was included in each of the meetings for this assessment.

The following are new research recommendations (not in priority order).
a) Refine logbook data collection, focusing on spatial details. Resolve apparent problems with locations for some records. Can recent data show patterns on finer spatial scales (e.g. for 1 -minute rather than 10 -minute squares)?
b) Improve collection and use of port sample data form the commercial fishery.
c) Characterize relationships between shell height, width and length for potential use in understanding the size selectivity of commercial and survey dredges and commercial sorting gear.
d) Test the Patch model for depletion experiments with simulations focusing on potential effects of uncertainty about position data and including all effects of cell size and smoothing.
e) Determine the size selectivity of survey and commercial fishing equipment experimentally.
f) Improve procedures for filling holes in the survey data using statistical models with year and spatial effects. Determine if filling holes is preferable to borrowing data from previous and subsequent surveys.
g) Review survey age data carefully to determine if strong year classes can be used to estimate mortality rates outside of a stock assessment model (e.g. "empirical" $Z$ estimates).
h) Further investigate spatial trends in survey data.
i) Devote sufficient time and resources to fully develop and improve dynamic population models.
j) Review the technical basis of the current $B_{M S Y}$ proxy given new data and possible climate effects.
k) Utilize New Jersey and New York inshore clam survey data more fully in the EEZ surfclam assessment.

### 11.0 ACKNOWLEDGMENTS

The captains, ships' crew, and scientific crews aboard the F/V Lisa Kim and $R / V$ Delaware II made substantial contributions to the report during 2004-2005 that are not mentioned elsewhere in this report. J. Womack (a contributing member of the Invertebrate Subcommittee) made very substantial contributions by analyzing NEFSC survey sensor data. J. Weinberg, C. Pickett and A. Chute (NEFSC) and E. Powell (Rutgers Univeristy) made important contributions. SARC reviewers (C. Jones, V. Haist and P . Cordue) made helpful suggestions and comments.

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[^29]
## SURFCLAM TABLES

Table C1. Length-weight parameters for Atlantic surfclam, by region. Parameters are for the relationship $W=e^{a} L^{b}$, where $W$ is meat weight in grams, $L$ is shell length in mm , and $a$ and $b$ are parameters in the table.

| Region | a | b |
| :---: | :---: | :---: |
| SVA | -7.05830 | 2.30330 |
| DMV | -9.48913 | 2.86018 |
| NJ | -9.31210 | 2.86371 |
| LI | -7.98370 | 2.58020 |
| SNE | -7.98370 | 2.58020 |
| GBK | -8.27443 | 2.65422 |

Table C2. Discard estimates for surfclam in the commercial fishery during 1982-1994 from Table D4 in NEFSC (1995).

| Year | Discard (mt meats) |  |  |  |  |  | Landings <br> $(\mathrm{mt}$ | NNJ | SNJ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | NJ | DMV | Total | Discards / <br> meats) <br> Landings | Catch | Size <br> limit <br> $(\mathrm{mm})$ |  |  |  |
| 1982 | 3,684 | 215 | 3,899 | 2,295 | 6,194 | 16,688 | $37 \%$ | 22,882 | 140 |
| 1983 | 2,122 | 385 | 2,507 | 2,127 | 4,634 | 18,592 | $25 \%$ | 23,226 | 140 |
| 1984 | 2,266 | 458 | 2,724 | 2,015 | 4,739 | 22,888 | $21 \%$ | 27,627 | 133 |
| 1985 | 1,938 | 248 | 2,186 | 1,725 | 3,911 | 22,480 | $17 \%$ | 26,391 | 127 |
| 1986 | 2,328 | 233 | 2,561 | 239 | 2,800 | 24,520 | $11 \%$ | 27,320 | 127 |
| 1987 | 1,414 | 61 | 1,475 | 415 | 1,890 | 21,744 | $9 \%$ | 23,634 | 127 |
| 1988 | 1,317 | 13 | 1,330 | 106 | 1,436 | 23,377 | $6 \%$ | 24,813 | 127 |
| 1989 | 1,048 | 6 | 1,054 | 258 | 1,312 | 21,887 | $6 \%$ | 23,199 | 127 |
| 1990 | 1,089 | 57 | 1,146 | 123 | 1,269 | 24,018 | $5 \%$ | 25,287 | 127 |
| 1991 | 495 | 36 | 531 | 5 | 536 | 20,615 | $3 \%$ | 21,151 | -- |
| 1992 | 918 | 102 | 1,020 | 4 | 1,024 | 21,685 | $5 \%$ | 22,709 | - |
| 1993 | 0 | 0 | 0 | 0 | 0 | 21,859 | $0 \%$ | 21,859 | -- |
| 1994 | 0 | 0 | 0 | 0 | 0 | 21,942 | $0 \%$ | 21,942 | -- |

Table C3. Atlantic surfclam landings in state waters and the EEZ with EEZ surfclam quotas ( mt meat weights). Total landings for 2002-2005 from dealer records. EEZ landings for 2002-2005 from MAFMC (2006). Other figures from logbooks or NEFSC (2003). Landings for state waters + unknown areas were estimated as total landings - EEZ landings.
$\left.\begin{array}{cccccc}\hline & & & \begin{array}{c}\text { State Waters } \\ \text { Total } \\ \text { Year } \\ \end{array} & \begin{array}{c}\text { Enknown } \\ \text { Landings }\end{array} & \begin{array}{c}\text { Percent } \\ \text { Landings } \\ \text { from EEZ }\end{array}\end{array} \begin{array}{c}\text { EEZ } \\ \text { Quota } \\ \text { Landings }\end{array}\right]$

Table C4. EEZ surfclam landings (mt meats) by stock assessment area and year based on NEFSC (2003) for 1979 and logbook data for 1980-2005. Logbook landings from unknown areas in each year were prorated to known areas based on proportions of landings in known areas.

| Year | SVA | DMV | NJ | LI | SNE | Other | Total EEZ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1979 | 0 | 11,836 | 1,350 | 0 | 0 | 0 | 13,186 |
| 1980 | 64 | 12,788 | 2,878 | 17 | 0 | 0 | 15,748 |
| 1981 | 568 | 7,472 | 8,820 | 88 | 0 | 0 | 16,947 |
| 1982 | 1,705 | 6,679 | 8,086 | 94 | 125 | 0 | 16,688 |
| 1983 | 2,225 | 7,173 | 8,095 | 264 | 836 | 0 | 18,592 |
| 1984 | 1,796 | 5,978 | 11,904 | 7 | 382 | 2,819 | 22,888 |
| 1985 | 741 | 7,856 | 11,246 | 0 | 452 | 2,185 | 22,480 |
| 1986 | 529 | 2,853 | 17,730 | 17 | 1,223 | 2,168 | 24,520 |
| 1987 | 378 | 1,302 | 18,017 | 0 | 1,140 | 907 | 21,744 |
| 1988 | 557 | 1,149 | 19,420 | 0 | 1,512 | 739 | 23,377 |
| 1989 | 439 | 3,123 | 16,531 | 0 | 1,361 | 433 | 21,887 |
| 1990 | 1,502 | 3,546 | 17,887 | 0 | 998 | 86 | 24,018 |
| 1991 | 0 | 1,634 | 18,913 | 15 | 33 | 21 | 20,615 |
| 1992 | 0 | 1,221 | 20,398 | 61 | 5 | 0 | 21,685 |
| 1993 | 0 | 3,414 | 18,365 | 62 | 3 | 14 | 21,859 |
| 1994 | 0 | 3,454 | 18,417 | 71 | 0 | 0 | 21,942 |
| 1995 | 0 | 2,752 | 16,497 | 0 | 378 | 0 | 19,627 |
| 1996 | 0 | 2,233 | 17,430 | 26 | 82 | 0 | 19,771 |
| 1997 | 0 | 1,540 | 16,998 | 73 | 0 | 0 | 18,611 |
| 1998 | 0 | 484 | 17,517 | 117 | 121 | 0 | 18,240 |
| 1999 | 0 | 648 | 18,749 | 157 | 16 | 0 | 19,570 |
| 2000 | 0 | 2,039 | 17,487 | 121 | 102 | 0 | 19,749 |
| 2001 | 0 | 3,282 | 17,719 | 935 | 81 | 0 | 22,017 |
| 2002 | 64 | 4,489 | 18,271 | 1,130 | 52 | 0 | 24,006 |
| 2003 | 0 | 1,432 | 21,693 | 1,625 | 267 | 0 | 25,017 |
| 2004 | 0 | 1,482 | 19,197 | 906 | 2,612 | 0 | 24,197 |
| 2005 | 0 | 1,668 | 16,850 | 759 | 1,885 | 0 | 21,163 |
| Min | 0 | 484 | 1,350 | 0 | 0 | 0 | 13,186 |
| Max | 2,225 | 12,788 | 21,693 | 1,625 | 2,612 | 2,819 | 25,017 |
| Mean | 391 | 3,834 | 15,425 | 242 | 506 | 347 | 20,746 |

Table C5. EEZ fishing effort (all vessels, hours fished) for surfclam by stock assessment area and year based on logbook data. The fraction of logbook effort from unknown areas in each year was prorated to known areas based on fishing effort in known areas.

| Year | SVA | DMV | NJ | LI | SNE | Other | Total EEZ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1991 | 0 | 1,254 | 17,243 | 21 | 107 | 293 | 18,917 |
| 1992 | 0 | 797 | 21,379 | 67 | 0 | 0 | 22,243 |
| 1993 | 0 | 2,423 | 18,232 | 57 | 15 | 5 | 20,732 |
| 1994 | 0 | 1,930 | 21,494 | 70 | 0 | 0 | 23,494 |
| 1995 | 0 | 1,560 | 18,625 | 0 | 1,059 | 0 | 21,244 |
| 1996 | 0 | 1,577 | 20,995 | 40 | 287 | 0 | 22,899 |
| 1997 | 0 | 1,098 | 20,383 | 77 | 0 | 0 | 21,558 |
| 1998 | 0 | 289 | 19,609 | 134 | 518 | 0 | 20,550 |
| 1999 | 0 | 734 | 18,146 | 151 | 149 | 0 | 19,179 |
| 2000 | 0 | 1,859 | 16,787 | 115 | 368 | 0 | 19,128 |
| 2001 | 0 | 2,536 | 18,462 | 962 | 148 | 0 | 22,108 |
| 2002 | 12 | 5,505 | 19,825 | 1,241 | 62 | 0 | 26,746 |
| 2003 | 0 | 2,367 | 25,071 | 1,827 | 176 | 0 | 29,441 |
| 2004 | 0 | 3,161 | 26,453 | 1,267 | 1,108 | 0 | 31,989 |
| 2005 | 0 | 2,654 | 24,335 | 1,206 | 1,340 | 0 | 29,534 |
| Min | 0 | 289 | 16,787 | 0 | 0 | 0 | 1,917 |
| Max | 112 | 5,505 | 26,453 | 1,827 | 1,340 | 293 | 31,989 |
| Mean | 7 | 1,983 | 20,469 | 482 | 356 | 20 | 23,317 |

Table C6. Nominal landings per unit effort (LPUE, bushels $h^{-1}$ ) for surfclam fishing (all vessels) in the US EEZ based on logbooks. Nominal LPUE is the ratio of total reported landings and total hours fished. Landings and fishing effort from unknown areas were prorated to area before LPUE was calculated.

| Year | SVA | DMV | NJ | LI | SNE | Other | All areas |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1991 |  |  | 142 | 95 | 40 | 9 | 141 |
| 1992 |  | 199 | 124 | 119 |  |  | 126 |
| 1993 |  | 183 | 131 | 143 | 28 | 390 | 137 |
| 1994 |  | 232 | 111 | 132 |  |  | 121 |
| 1995 |  | 229 | 115 |  | 46 |  | 120 |
| 1996 |  | 184 | 108 | 85 | 37 |  | 112 |
| 1997 |  | 182 | 108 | 122 |  |  | 112 |
| 1998 |  | 217 | 116 | 114 | 30 |  | 115 |
| 1999 |  | 115 | 134 | 135 | 14 |  | 132 |
| 2000 |  | 142 | 135 | 137 | 36 |  | 134 |
| 2001 |  | 168 | 124 | 126 | 71 |  | 129 |
| 2002 | 74 | 106 | 120 | 118 | 108 |  | 116 |
| 2003 |  | 78 | 112 | 115 | 197 |  | 110 |
| 2004 |  | 61 | 94 | 93 | 306 |  | 98 |
| 2005 |  | 82 | 90 | 82 | 183 |  | 93 |
| Min | 74 | 61 | 90 | 82 | 14 | 9 | 93 |
| Max | 74 | 232 | 142 | 143 | 306 | 390 | 141 |
| Mean | 74 | 155 | 118 | 115 | 91 | 199 | 120 |

Table C7. Standardized annual LPUE (bushels per hour) based on log-linear GLM models. Results are scaled to LPUE during June for an arbitrary vessel that fished in all areas.

|  | DMV |  | NJ |  | LI |  | SNE |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | LPUE | CV | LPUE | CV | LPUE | CV | LPUE | CV |
| 1990 | 241 | 0.69 | 138 | 0.05 |  |  |  |  |
| 1991 | 206 | 0.69 | 107 | 0.05 |  |  |  |  |
| 1992 | 232 | 0.69 | 101 | 0.05 |  |  |  |  |
| 1993 | 237 | 0.69 | 110 | 0.05 |  |  |  |  |
| 1994 | 322 | 0.69 | 98 | 0.05 |  |  | 0.59 |  |
| 1995 | 287 | 0.69 | 96 | 0.05 |  |  |  |  |
| 1996 | 215 | 0.69 | 91 | 0.05 |  |  |  |  |
| 1997 | 202 | 0.69 | 88 | 0.05 | 157 | 0.49 |  | 0.66 |
| 1998 | 210 | 0.70 | 97 | 0.05 | 105 | 0.50 | 24 | 0.99 |
| 1999 | 185 | 0.69 | 101 | 0.05 | 119 | 0.48 | 39 | 0.97 |
| 2000 | 185 | 0.69 | 93 | 0.05 | 130 | 0.49 | 28 | 0.62 |
| 2001 | 200 | 0.69 | 78 | 0.05 | 116 | 0.47 | 44 | 0.64 |
| 2002 | 119 | 0.69 | 85 | 0.05 | 104 | 0.47 | 83 | 0.56 |
| 2003 | 86 | 0.69 | 75 | 0.05 | 91 | 0.47 | 109 | 0.54 |
| 2004 | 69 | 0.69 | 63 | 0.05 | 71 | 0.47 | 72 | 0.53 |
| 2005 | 85 | 0.69 | 54 | 0.04 | 60 | 0.46 | 81 | 0.53 |
| Min | 69 | 0.69 | 54 | 0.04 | 60 | 0.46 | 6 |  |
| Max | 322 | 0.70 | 138 | 0.05 | 157 | 0.50 | 109 | 0.99 |
| Average | 193 | 0.69 | 92 | 0.05 | 106 | 0.48 | 50 | 0.69 |

Table C8. Numbers of commercial trips sampled and numbers of surfclam measured in port samples from landings during 1982-2005, by region. Numbers of measurements for 1982-1999 are from NEFSC (2003, Table C5) and numbers of trips during 1982-1999 were estimated assuming 30 individuals sampled per trip, as specified in port sample instructions.

| Year | DMV |  | NJ |  | LI |  | SNE |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Trips | Lengths | Trips | Lengths | Trips | Lengths | Trips | Lengths |
| 1982 | 259 | 7,756 | 249 | 7,477 |  |  | 1 | 30 |
| 1983 | 197 | 5,923 | 375 | 11,253 |  |  | 1 | 30 |
| 1984 | 102 | 3,066 | 425 | 12,751 |  |  | 3 | 90 |
| 1985 | 61 | 1,832 | 256 | 7,674 |  |  | 5 | 150 |
| 1986 | 42 | 1,260 | 171 | 5,130 |  |  | 11 | 330 |
| 1987 | 24 | 730 | 30 | 900 |  |  | 19 | 569 |
| 1988 | 14 | 420 | 30 | 900 |  |  | 27 | 810 |
| 1989 | 29 | 866 | 31 | 919 |  |  | 15 | 449 |
| 1990 | 30 | 892 | 30 | 901 |  |  | 7 | 209 |
| 1991 | 36 | 1,080 | 76 | 2,272 |  |  |  |  |
| 1992 | 39 | 1,170 | 57 | 1,710 |  |  |  |  |
| 1993 | 46 | 1,392 | 31 | 928 |  |  |  |  |
| 1994 | 4 | 119 | 30 | 900 |  |  |  |  |
| 1995 | 24 | 720 | 17 | 510 |  |  |  |  |
| 1996 | 38 | 1,154 | 37 | 1,117 |  |  |  |  |
| 1997 | 54 | 1,622 | 32 | 957 |  |  |  |  |
| 1998 | 52 | 1,560 | 23 | 690 |  |  |  |  |
| 1999 | 57 | 1,720 | 29 | 856 |  |  |  |  |
| 2000 | 20 | 600 | 111 | 3,315 | 1 | 30 |  |  |
| 2001 | 33 | 970 | 42 | 1,260 |  |  |  |  |
| 2002 | 7 | 210 | 37 | 1,111 |  |  |  |  |
| 2003 | 2 | 60 | 80 | 2,455 | 5 | 150 |  |  |
| 2004 |  |  | 36 | 1,080 | 2 | 60 |  |  |
| 2005 | 19 | 581 | 61 | 1,834 | 11 | 330 |  |  |
| Min | 2 | 60 | 17 | 510 | 1 | 30 | 1 | 30 |
| Max | 259 | 7,756 | 425 | 12,751 | 11 | 330 | 27 | 810 |
| Mean | 52 | 1,552 | 96 | 2,871 | 5 | 143 | 10 | 296 |

Table C9. Numbers of random survey stations in NEFSC and cooperative clam surveys by stratum, region and survey year. The 2004 survey was cooperative and carried out on a commercial vessel. All others were NEFSC clam surveys carried out on the $R / V$ Delaware II. Numbers of NEFSC clam survey stations for 2005 include a few tows with poor dredge performance used to trends but not for swept area biomass. For NEFSC surveys, figures in plain text are the number of original random tows (without borrowing). Bold and outlined figures are for NEFSC survey data are "holes" (strata in with no stations), which where filled by borrowing data from the same stratum during previous and/or subsequent cruises. Black cells are remaining zeroes for NEFSC survey data that could not be filled by borrowing. Only SVA, DMV and NJ were sampled during 2004 (cells for strata not sampled are crosshatched). Survey data for GBK during 1982-1984 and 2005 (stippled) should not be used in most analyses due to limited sampling.

| Region | Stratum | Survey Year |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1982 | 1983 | 1984 | 1986 | 1989 | 1992 | 1994 | 1997 | 1999 | 2002 | 2004 | 2005 |
| SVA | 1 | 10 | 10 | 14 | 7 | 10 | 10 | 11 | 10 | 10 |  |  |  |
|  | 2 |  |  |  | 1 | 1 | 2 | 1 | 1 | 1 |  |  |  |
|  | 5 | 4 | 9 | 13 | 8 | 8 | 8 | 7 | 8 | 16 | 8 | 8 | 8 |
|  | 6 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 3 | 2 | 2 | 1 |
|  | 80 | 6 | 6 | 9 | 3 | 7 | 7 | 8 | 7 | 7 |  |  |  |
|  | 81 | 4 | 4 | 7 | 3 | 5 | 5 | 5 | 5 | 5 | 5 |  | 5 |
| DMV | 9 | 30 | 26 | 35 | 29 | 37 | 37 | 39 | 39 | 38 | 39 | 37 | 36 |
|  | 10 | 2 | 2 | 3 | 3 | 3 | 3 | 3 | 3 | 3 | 3 | 3 | 3 |
|  | 13 | 19 | 18 | 25 | 20 | 20 | 20 | 21 | 22 | 19 | 20 | 20 | 18 |
|  | 14 | 2 | 2 | 3 | 3 | 3 | 3 | 5 | 3 | 3 | 3 | 3 | 3 |
|  | 82 | 1 | 1 | 1 | 1 |  | 1 | 1 | 1 | 2 | 2 | 2 | 2 |
|  | 83 | 2 | 2 | 2 | 2 | 2 | 2 | 2 | 2 | 2 | 2 | 2 | 2 |
|  | 84 | 4 | 3 | 3 | 4 | 4 | 4 | 4 | 4 | 3 | 4 | 4 | 4 |
|  | 85 | 6 | 5 | 4 | 5 | 5 | 5 | 5 | 5 | 5 | 5 | 3 | 3 |
|  | 86 | 2 | 2 | 3 | 3 | 3 | 2 | 3 | 3 | 3 | 3 | 3 | 3 |
| NJ | 17 | 11 | 11 | 18 | 12 | 12 | 12 | 12 | 14 | 12 | 12 | 12 | 12 |
|  | 18 | 3 | 3 | 6 | 3 | 3 | 3 | 3 | 3 | 3 | 3 | 3 | 3 |
|  | 21 | 18 | 18 | 22 | 19 | 20 | 20 | 23 | 26 | 39 | 29 | 27 | 20 |
|  | 22 | 3 | 3 | 6 | 3 | 3 | 3 | 5 | 3 | 3 | 3 | 3 | 3 |
|  | 25 | 9 | 9 | 13 | 8 | 9 | 9 | 9 | 12 | 8 | 9 | 9 | 9 |
|  | 26 | 2 | 2 | 5 | 3 | 3 | 3 | 3 | 3 | 3 | 3 | 3 | 3 |
|  | 87 | 8 | 7 | 10 | 9 | 9 | 9 | 9 | 9 | 9 | 16 | 14 | 8 |
|  | 88 | 15 | 15 | 24 | 17 | 20 | 20 | 20 | 21 | 23 | 20 | 20 | 17 |
|  | 89 | 15 | 15 | 21 | 15 | 18 | 17 | 17 | 19 | 18 | 18 | 17 | 15 |
|  | 90 | 2 | 2 | 3 | 2 | 2 | 2 | 2 | 2 | 2 | 2 | 2 | 2 |
| LI | 29 | 11 | 10 | 20 | 10 | 10 | 10 | 10 | 10 | 11 | 10 |  | 10 |
|  | 30 | 7 | 8 | 14 | 6 | 6 | 6 | 6 | 6 | 7 | 6 |  | 7 |
|  | 33 | 4 | 4 | 8 | 4 | 4 | 4 | 5 | 4 | 4 | 4 |  | 4 |
|  | 34 | 2 | 2 | 4 | 2 | 2 | 2 | 5 | 2 | 2 | 2 |  | 2 |
|  | 91 | 3 | 2 | 4 | 4 | 3 | 3 | 3 | 3 | 3 | 3 |  | 3 |
|  | 92 | 2 | 2 | 3 | 2 | 2 | 2 | 2 | 2 | 2 | 2 |  | 2 |
|  | 93 | 1 | 1 | 2 | 1 | 1 | 1 | 1 | 1 | 1 | 2 |  | 1 |

Table C9. (continued)

| Survey Year |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Region | Stratum | 1982 | 1983 | 1984 | 1986 | 1989 | 1992 | 1994 | 1997 | 1999 | 2002 | 2004 | 2005 |
| SNE | 37 | 7 | 4 | 7 | 3 | 6 | 3 | 5 | 4 | 4 | 3 |  | 3 |
|  | 38 | 3 | 2 | 5 | 3 | 3 | 3 | 5 | 3 | 3 | 3 |  | 2 |
|  | 41 | 6 | 5 | 7 | 5 | 6 | 6 | 6 | 6 | 5 | 6 |  | 6 |
|  | 45 | 3 | 7 | 9 | 4 | 4 | 4 | 4 | 4 | 4 | 3 |  | 3 |
|  | 46 | 2 | 5 | 5 | 3 | 2 | 3 | 5 | 3 | 3 | 2 |  | 3 |
|  | 47 | 4 | 3 | 4 | 2 | 2 | 4 | 5 | 4 | 3 | 1 |  | 7 |
|  | 94 | 1 | 2 | 2 |  | 1 | 1 | 2 | 2 | 4 | 2 |  | 2 |
|  | 95 | 4 | 14 | 11 | 4 | 4 | 4 | 4 | 4 | 4 | 4 |  | 4 |
|  | 96 | 12 | 12 | 13 | 1 | 1 | 3 | 2 | 4 | 4 |  |  |  |
| GBK | 54 |  | 3 | 3 | 3 | 6 | 3 | 3 | 3 | 3 |  |  |  |
|  | 55 | 3 | 3 | 3 | 3 | 1 | 3 | 3 | 3 | 2 | 2 |  | 2 |
|  | 57 |  |  | 2 | 2 | 1 | 2 | 5 | 2 | 2 | 2 |  | 2 |
|  | 59 | 1 | 4 | 5 | 1 | 2 | 6 | 5 | 5 | 4 | 5 |  | 5 |
|  | 61 | 8 | 1 | 6 | 5 | 12 | 7 | 6 | 6 | 6 | 6 |  | 6 |
|  | 65 |  |  | 3 | 3 | 5 | 2 | 4 | 3 | 4 | 1 |  | 1 |
|  | 67 |  | 5 | 5 | 5 | 7 | 7 | 7 | 7 | 7 |  |  |  |
|  | 68 | 1 | 8 | 7 | 3 | 6 | 6 | 5 | 5 | 5 |  |  |  |
|  | 69 | 2 | 5 | 11 | 6 | 6 | 6 | 7 | 6 | 8 | 8 |  |  |
|  | 70 | 1 | 2 | 6 | 4 | 8 | 4 | 4 | 4 | 3 | 2 |  | 2 |
|  | 71 |  | 2 | 2 | 3 | 1 | 2 | 3 | 3 | 1 | 2 |  | 2 |
|  | 72 | 2 | 10 | 8 | 1 | 8 | 8 | 8 | 8 | 6 | 6 |  |  |
|  | 73 | 1 | 1 | 4 | 3 | 6 | 6 | 6 | 6 | 5 | 6 |  | 6 |
|  | 74 | 3 | 4 | 1 | 3 | 7 | 4 | 4 | 4 | 3 | 3 |  | 3 |

Table C10. Bad tows identified using objective criteria in the 2005 survey and by eye in the 2002 survey using sensor data.

| Statistic | 2005 | 2002 |
| :---: | :---: | :---: |
| Total |  |  |
| N examined tows | 433 | 556 |
| \% examined | 399 | 213 |
| Number w/poor dredge performance | $92 \%$ | $38 \%$ |
| Proportion w/poor dredge perfomance | 0.08 | 32 |
| Depletion tows only |  |  |
| Total | 30 | 75 |
| N examined (estimate) | 28 | 29 |
| Number bad | 8 | 10 |
| Proportion w/poor dredge perfomance |  |  |
| Assuming 100\% examined* | $27 \%$ | $13 \%$ |
| Expanded based on \% reviewed | $29 \%$ | $35 \%$ |

[^30]Table C11. NEFSC clam survey data for surfclam abundance (mean N/tow) and biomass (mean $\mathrm{KG} /$ tow). Data are for two size groups: small recruits ( $50-119 \mathrm{~mm} \mathrm{SL}$ ) and large fishable ( $120+\mathrm{mm} \mathrm{SL}$ ). Survey holes (strata with no sampling) were filled by borrowing but no imputed survey data were used.

| Small recruits (50-119 mm SL) |  |  |  |  |  | Large fishable (120+ mm SL) |  |  |  | N Tows | N Positive Tows | N Strata Sampled |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Region | Year | N/ Tow | CV | KG / Tow | CV | N/Tow | CV | KG / Tow | CV |  |  |  |
| SVA | 1982 | 3.529 | 0.88 | 0.134 | 0.91 | 0.920 | 1.00 | 0.257 | 0.87 | 25 | 5 | 5 |
| SVA | 1983 | 6.600 | 0.62 | 0.249 | 0.64 | 0.620 | 1.00 | 0.405 | 0.60 | 30 | 9 | 5 |
| SVA | 1984 | 7.849 | 0.37 | 0.303 | 0.40 | 0.310 | 1.00 | 1.609 | 0.30 | 44 | 16 | 5 |
| SVA | 1986 | 1.498 | 0.35 | 0.058 | 0.41 | 0.750 | 1.00 | 1.553 | 0.74 | 23 | 11 | 6 |
| SVA | 1989 | 3.109 | 0.75 | 0.083 | 0.71 | 0.830 | 1.00 | 0.758 | 0.82 | 32 | 10 | 6 |
| SVA | 1992 | 18.151 | 0.86 | 0.760 | 0.90 | 0.770 | 1.00 | 0.812 | 0.79 | 33 | 17 | 6 |
| SVA | 1994 | 43.379 | 0.46 | 0.784 | 0.31 | 0.440 | 1.00 | 0.427 | 0.38 | 33 | 19 | 6 |
| SVA | 1997 | 10.309 | 0.44 | 0.294 | 0.46 | 0.460 | 1.00 | 0.030 | 0.44 | 32 | 14 | 6 |
| SVA | 1999 | 9.317 | 0.41 | 0.234 | 0.35 | 0.460 | 1.00 | 0.084 | 0.47 | 47 | 19 | 6 |
| SVA | 2002 | 13.693 | 0.61 | 0.360 | 0.62 | 0.550 | 1.00 | 0.399 | 0.55 | 15 | 5 | 3 |
| SVA | 2005 | 3.646 | 0.66 | 0.051 | 0.57 | . | 0.00 | 0.000 | . | 14 | 4 | 3 |
| DMV | 1982 | 157.134 | 0.46 | 6.621 | 0.44 | 21.360 | 0.23 | 2.687 | 0.29 | 68 | 37 | 9 |
| DMV | 1983 | 30.679 | 0.54 | 1.534 | 0.61 | 31.205 | 0.46 | 3.168 | 0.35 | 61 | 30 | 9 |
| DMV | 1984 | 184.102 | 0.74 | 5.247 | 0.61 | 34.911 | 0.28 | 3.555 | 0.28 | 79 | 47 | 9 |
| DMV | 1986 | 58.771 | 0.43 | 3.120 | 0.46 | 74.792 | 0.38 | 6.703 | 0.32 | 70 | 44 | 9 |
| DMV | 1989 | 16.705 | 0.54 | 0.813 | 0.55 | 31.237 | 0.26 | 3.065 | 0.24 | 78 | 37 | 9 |
| DMV | 1992 | 13.494 | 0.28 | 0.580 | 0.38 | 28.855 | 0.29 | 2.918 | 0.24 | 77 | 52 | 9 |
| DMV | 1994 | 68.704 | 0.33 | 2.787 | 0.43 | 60.964 | 0.21 | 5.958 | 0.20 | 83 | 63 | 9 |
| DMV | 1997 | 77.184 | 0.17 | 3.346 | 0.20 | 54.528 | 0.24 | 4.928 | 0.22 | 82 | 61 | 9 |
| DMV | 1999 | 29.612 | 0.28 | 1.543 | 0.28 | 26.363 | 0.22 | 2.406 | 0.20 | 78 | 44 | 9 |
| DMV | 2002 | 16.467 | 0.28 | 0.594 | 0.28 | 20.698 | 0.21 | 2.235 | 0.19 | 81 | 50 | 9 |
| DMV | 2005 | 6.437 | 0.42 | 0.252 | 0.43 | 4.757 | 0.26 | 0.508 | 0.28 | 74 | 40 | 9 |
| NJ | 1982 | 33.102 | 0.30 | 1.787 | 0.31 | 32.777 | 0.22 | 4.084 | 0.20 | 85 | 50 | 10 |
| NJ | 1983 | 27.780 | 0.51 | 1.627 | 0.55 | 25.382 | 0.22 | 3.147 | 0.20 | 85 | 54 | 10 |
| NJ | 1984 | 15.932 | 0.23 | 0.714 | 0.22 | 29.970 | 0.20 | 3.731 | 0.18 | 126 | 68 | 10 |
| NJ | 1986 | 10.335 | 0.21 | 0.493 | 0.20 | 29.677 | 0.18 | 4.172 | 0.18 | 91 | 59 | 10 |
| NJ | 1989 | 9.877 | 0.29 | 0.489 | 0.31 | 31.527 | 0.15 | 4.160 | 0.13 | 99 | 60 | 10 |
| NJ | 1992 | 16.462 | 0.33 | 0.849 | 0.42 | 23.221 | 0.16 | 3.193 | 0.15 | 98 | 62 | 10 |
| NJ | 1994 | 67.394 | 0.20 | 2.664 | 0.18 | 82.766 | 0.17 | 11.014 | 0.16 | 103 | 84 | 10 |
| NJ | 1997 | 17.910 | 0.16 | 1.012 | 0.17 | 83.720 | 0.13 | 11.442 | 0.12 | 112 | 83 | 10 |
| NJ | 1999 | 8.021 | 0.25 | 0.389 | 0.28 | 50.578 | 0.21 | 6.903 | 0.17 | 120 | 77 | 10 |
| NJ | 2002 | 10.678 | 0.16 | 0.464 | 0.16 | 35.035 | 0.17 | 5.503 | 0.17 | 115 | 94 | 10 |
| NJ | 2005 | 7.808 | 0.20 | 0.397 | 0.22 | 19.090 | 0.18 | 2.818 | 0.17 | 92 | 60 | 10 |
| LI | 1982 | 0.032 | 1.00 | 0.002 | 1.00 | 3.994 | 0.61 | 0.641 | 0.62 | 29 | 1 | 7 |
| LI | 1983 | 0.175 | 0.61 | 0.005 | 0.60 | 0.407 | 0.72 | 0.055 | 0.72 | 29 | 3 | 7 |
| LI | 1984 | 0.561 | 0.30 | 0.021 | 0.36 | 1.635 | 0.34 | 0.248 | 0.34 | 55 | 12 | 7 |
| LI | 1986 | 0.581 | 0.39 | 0.022 | 0.40 | 1.715 | 0.61 | 0.285 | 0.61 | 29 | 7 | 7 |
| LI | 1989 | 2.237 | 0.87 | 0.089 | 0.88 | 3.484 | 0.72 | 0.475 | 0.74 | 28 | 4 | 7 |
| LI | 1992 | 5.733 | 0.44 | 0.301 | 0.47 | 2.544 | 0.33 | 0.275 | 0.32 | 28 | 9 | 7 |
| LI | 1994 | 4.232 | 0.17 | 0.213 | 0.20 | 7.243 | 0.19 | 0.901 | 0.21 | 32 | 11 | 7 |
| LI | 1997 | 1.444 | 0.49 | 0.082 | 0.53 | 4.171 | 0.64 | 0.563 | 0.63 | 28 | 6 | 7 |
| LI | 1999 | 1.608 | 0.64 | 0.047 | 0.50 | 10.710 | 0.65 | 1.433 | 0.61 | 30 | 8 | 7 |
| LI | 2002 | 0.854 | 0.45 | 0.034 | 0.44 | 1.944 | 0.67 | 0.304 | 0.67 | 29 | 7 | 7 |
| LI | 2005 | 1.415 | 0.34 | 0.060 | 0.38 | 12.624 | 0.50 | 1.658 | 0.47 | 29 | 7 | 7 |
| SNE | 1982 | 2.584 | 0.29 | 0.112 | 0.35 | 12.402 | 0.41 | 1.776 | 0.42 | 42 | 14 | 9 |
| SNE | 1983 | 0.839 | 0.40 | 0.040 | 0.44 | 7.883 | 0.39 | 1.267 | 0.39 | 54 | 18 | 9 |
| SNE | 1984 | 0.810 | 0.36 | 0.034 | 0.43 | 10.838 | 0.34 | 1.676 | 0.34 | 63 | 18 | 9 |
| SNE | 1986 | 1.115 | 0.14 | 0.027 | 0.26 | 4.125 | 0.68 | 0.644 | 0.69 | 25 | 8 | 8 |
| SNE | 1989 | 1.178 | 0.43 | 0.044 | 0.44 | 4.569 | 0.33 | 0.687 | 0.33 | 29 | 10 | 9 |
| SNE | 1992 | 1.147 | 0.56 | 0.032 | 0.51 | 2.491 | 0.58 | 0.399 | 0.58 | 31 | 9 | 9 |
| SNE | 1994 | 1.265 | 0.52 | 0.061 | 0.58 | 1.693 | 0.53 | 0.265 | 0.54 | 38 | 10 | 9 |
| SNE | 1997 | 2.947 | 0.31 | 0.120 | 0.35 | 12.279 | 0.30 | 1.913 | 0.30 | 34 | 13 | 9 |
| SNE | 1999 | 2.601 | 0.42 | 0.089 | 0.47 | 4.296 | 0.66 | 0.725 | 0.66 | 34 | 15 | 9 |
| SNE | 2002 | 1.006 | 0.69 | 0.057 | 0.72 | 3.852 | 0.27 | 0.601 | 0.22 | 24 | 5 | 8 |
| SNE | 2005 | 0.261 | 0.49 | 0.008 | 0.51 | 1.986 | 0.19 | 0.355 | 0.19 | 30 | 6 | 8 |
| GBK | 1986 | 19.998 | 0.79 | 0.719 | 0.78 | 4.967 | 0.52 | 0.708 | 0.55 | 44 | 20 | 14 |
| GBK | 1989 | 5.214 | 0.34 | 0.285 | 0.42 | 24.858 | 0.73 | 3.004 | 0.73 | 75 | 36 | 14 |
| GBK | 1992 | 15.535 | 0.40 | 0.706 | 0.46 | 7.894 | 0.33 | 0.956 | 0.34 | 66 | 43 | 14 |
| GBK | 1994 | 30.010 | 0.33 | 1.610 | 0.34 | 45.843 | 0.39 | 5.853 | 0.41 | 70 | 47 | 14 |
| GBK | 1997 | 58.550 | 0.31 | 3.002 | 0.33 | 23.517 | 0.25 | 2.730 | 0.25 | 65 | 45 | 14 |
| GBK | 1999 | 24.014 | 0.41 | 1.340 | 0.41 | 29.590 | 0.31 | 3.385 | 0.30 | 59 | 34 | 14 |
| GBK | 2002 | 22.093 | 0.52 | 1.163 | 0.54 | 27.052 | 0.43 | 3.250 | 0.41 | 43 | 21 | 11 |

C:\Assessments\Surfclam2006ISurveys\Trends\[SurveyTrends-20.x|s]Table 1.

Table C12. Original mean $\mathrm{kg} /$ tow for surfclam in regions that had strata with remaining holes and mean $\mathrm{kg} /$ tow with remaining holes filled. Remaining holes were filled with imputed values from a negative binomial GLM model. Estimates of mean kg /tow for swept area biomass were computed from estimates for trends using the mean ratio of doppler and sensor distances during 1997-2005 for each region.

|  |  |  |  | Mean kg/tow for trends |  | Mean kg/tow for <br> swept-area biomass |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  |  | CV for trends |  |  |  |  |  |
|  |  | Imputed | Original | Imputed | Original | Imputed | Original |
| 198204 | GBK | 0.059 |  |  |  | 0.219 |  |
| 198204 | SVA | 0.243 | 0.257 |  |  | 0.874 | 0.870 |
| 198305 | GBK | 0.485 |  |  |  | 0.678 |  |
| 198305 | SVA | 0.383 | 0.405 |  |  | 0.597 | 0.600 |
| 198403 | SVA | 1.522 | 1.609 |  |  | 0.296 | 0.300 |
| 198604 | SNE | 0.609 | 0.680 |  |  | 0.688 | 0.690 |
| 200206 | GBK | 3.411 | 3.250 | 1.847 | 1.890 | 0.349 | 0.410 |
| 200206 | SNE | 0.715 | 0.601 | 0.418 | 0.396 | 0.264 | 0.220 |
| 200206 | SVA | 0.263 | 0.399 | 0.157 | 0.268 | 0.517 | 0.550 |
| 200507 | SNE | 0.317 | 0.355 | 0.185 | 0.224 | 0.190 | 0.190 |
| 200507 | SVA | 0.000 | 0.000 | 0.000 | 0.000 | 310 |  |

Table C13. Summary of commercial dredge efficiency, population density and negative binomial parameter $k$ estimates from the Patch model, setup tow densities and NEFSC survey dredge efficiency estimates from setup tows, by year. All estimates are for surfclam 150+ mm SL.

| Statistic | N successful experiments | Population Density ( $\mathrm{Nft}{ }^{-2}$ ) | Depletion Vessel Efficiency | $k$ | Setup Density ( $\mathrm{Nft}{ }^{-2}$ ) | NEFSC Dredge Efficiency |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1997 |  |  |  |  |  |  |
| Mean |  | 0.023 | 0.79 | 4.758 | 0.0061 | 0.317 |
| Median |  | 0.017 | 0.89 | 3.261 | 0.0069 | 0.27 |
| Lower 80\% bound |  | 0.012 | 0.613 | 3.134 | 0.0049 | 0.225 |
| Upper 80\% bound |  | 0.033 | 0.967 | 6.382 | 0.0072 | 0.409 |
| SE |  | 0.007 | 0.115 | 1.059 | 0.0008 | 0.06 |
| CV (SE / Mean) | 5 | 0.296 | 0.146 | 0.223 | 0.1281 | 0.189 |
| 1999 |  |  |  |  |  |  |
| Mean |  | 0.035 | 0.652 | 20.534 | 0.0061 | 0.189 |
| Median |  | 0.025 | 0.726 | 12.841 | 0.0058 | 0.199 |
| Lower 80\% bound | commercial | 0.024 | 0.469 | 10.137 | 0.0051 | 0.13 |
| Upper 80\% bound | depletion, 5 | 0.046 | 0.835 | 30.93 | 0.007 | 0.248 |
| SE | with setup | 0.007 | 0.124 | 7.044 | 0.0006 | 0.039 |
| CV (SE / Mean) | tows | 0.211 | 0.19 | 0.343 | 0.1012 | 0.205 |
| 2002 |  |  |  |  |  |  |
| Mean |  | 0.014 | 0.584 | 16.792 | 0.007 | 0.516 |
| Median |  | 0.014 | 0.584 | 16.792 | 0.007 | 0.516 |
| Lower 80\% bound |  | 0.012 | -0.268 | -26.157 | -0.0032 | -0.282 |
| Upper 80\% bound |  | 0.016 | 1.437 | 59.74 | 0.0173 | 1.313 |
| SE |  | 0.001 | 0.277 | 13.955 | 0.0033 | 0.259 |
| CV (SE / Mean) | 2 | 0.038 | 0.474 | 0.831 | 0.474 | 0.503 |
| 2004 |  |  |  |  |  |  |
| Mean |  | 0.024 | 0.736 | 5.939 | NA | NA |
| Median |  | 0.024 | 0.736 | 5.939 | NA | NA |
| Lower 80\% bound | 2 | 0.004 | 0.517 | 0.22 | NA | NA |
| Upper 80\% bound | commercial | 0.043 | 0.955 | 11.658 | NA | NA |
| SE | depletion | 0.006 | 0.071 | 1.858 | NA | NA |
| CV (SE / Mean) | experiments | 0.268 | 0.097 | 0.313 | NA | NA |
| 2005 |  |  |  |  |  |  |
| Mean |  | 0.037 | 0.717 | 4.078 | 0.005 | 0.158 |
| Median |  | 0.034 | 0.676 | 4.593 | 0.005 | 0.158 |
| Lower 80\% bound |  | 0.023 | 0.551 | 3.121 | 0.004 | 0.105 |
| Upper 80\% bound |  | 0.051 | 0.882 | 5.035 | 0.006 | 0.21 |
| SE |  | 0.008 | 0.101 | 0.584 | 0 | 0.032 |
| CV (SE / Mean) | 4 | 0.229 | 0.141 | 0.143 | 0.084 | 0.203 |
| All years |  |  |  |  |  |  |
| Mean |  | 0.029 | 0.704 | 10.988 | 0.006 | 0.262 |
| Median |  | 0.025 | 0.765 | 5.676 | 0.006 | 0.226 |
| Lower 80\% bound | 19 | 0.024 | 0.628 | 7.073 | 0.005 | 0.203 |
| Upper 80\% bound | commercial depletion | 0.033 | 0.779 | 14.903 | 0.007 | 0.32 |
| SE | 16 with | 0.004 | 0.057 | 2.943 | 0 | 0.044 |
| CV (SE / Mean) | setup tows | 0.128 | 0.081 | 0.268 | 0.076 | 0.168 |

Table C14. Summary of depletion experiments, setup tows, Patch model estimates, and survey dredge efficiency estimates for surfclam. All depletion results are for surfclam $150+\mathrm{mm}$ SL. Depletion experiments by $R / V$ Delaware $I I$ are not shown.

|  |  | Experiment | nd Study | Area |  |  |  |  | Depletion Tow |  |  |  |  | Patch | Model |  |  | Survey | y setu p tows |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Experiment | Original Name | Region | Approx. latitude degrees) | Approx. longitude (decimal degrees) | $\begin{aligned} & \text { Depth } \\ & (\mathrm{m}) \end{aligned}$ | $\left.\begin{gathered} \text { Mean } \\ \text { Sediment } \\ \text { Size } \\ \text { (microns) } \end{gathered} \right\rvert\, \text { S }$ | Depletion Study Vessel | Depletion | Ship Position Data (source / nominal accuracy / time interval) | Depletion tows: N used, [ N with bushel count and length samples] | Depletion vessel blade width (ft) | $\underset{(f t)}{\text { Cell size }}$ | Population Density ( $D,>=150$ mm SL, Nft ) | Depletion vessel efficiency ( $E$, fully recruited, $>=150 \mathrm{~mm}$ SL ) | Negative parameter (K) | Gamma (indirect effects, $\gamma$ ) | Goodness of fit (-log likelihood) | Survey id, [station id] (N tows) $\{\mathrm{N}$ tows with length data\} | $\begin{gathered} \text { Catch } \\ \text { density } \\ (d,>=150 \\ m m \text { SL, } \\ \mathrm{Nfft}) \end{gathered}$ | $\begin{gathered} \text { CV for } \\ \text { catch } \\ \text { density } \\ \text { (se / } \\ \text { mean) } \end{gathered}$ | NEFSC survey dredge efficiency ( $e$, fully recruited | Notes |
| SC 1997-2 | PP-1 | $\begin{array}{\|c} \text { NNJ } \\ \text { (Pt. Pleasant) } \end{array}$ | 40.05317 | 73.83917 | 26 |  | Sherri Ann | 6/9/1997 | Loran / 9-12 M/ 1 Minute | 39 [9] | 8.33 | 16.67 | 0.0492 | 0.3540 | 7.5313 | 0.5 | 210.3 |  | 0.0081 | 0.1498 | 0.1645 | Forty depletion tows total but tow 1 (and samples) omitted. Setup tows during calibration survey 199703 prior to 199704 clam survey. |
| SC 1997-3 | AC2-1 | $\underset{\text { (Atlantic City) }}{\text { NNJ }}$ | 39.39317 | 73.91033 | 30 |  | Jersey Girl | 6/10/1997 | Loran / 9-12 M / 1 Minute | 13 [4] | 10.83 | 21.67 | 0.0172 | 0.7646 | 2.6272 | 0.5 | 66.1 | $\begin{gathered} 199703[169, \\ 175-181](8) \\ \{8\} \end{gathered}$ | 0.0042 | 0.1011 | 0.2463 | Setup tows during calibration survey 199703 prior to 199704 clam survey. |
| SC 1997-4 | AC2-2 | $\begin{gathered} \text { NNJ } \\ \text { (Atlantic City) } \end{gathered}$ | 39.39317 | 73.91033 | 30 |  | Jersey Girl | 6/10/1997 | Loran / 9-12 <br> M/1 Minute | 31 [4] | 10.83 | 21.67 | 0.0157 | 0.9900 | 3.2368 | 0.5 | 95.8 | Same as SC1997-2 | 0.0042 | 0.1011 | 0.2698 | Setup tows during calibration survey 199703 prior to 199704 clam survey. |
| SC 1997-5 | AC1-1 | $\underset{\text { (Atlantic City) }}{\text { NNJ }}$ | 31.36500 | 73.89833 | 30 |  | Judy Marie | 6/11/1997 | Loran / 9-12 M/ 1 Minute | 17 [4] | 8.33 | 16.67 | 0.0137 | 0.9500 | 3.2606 | 0.5055 | 86.9 | $\begin{gathered} \hline 199703[166- \\ 168,170- \\ 174](8)\{8\} \\ \hline \end{gathered}$ | 0.0069 | 0.1173 | 0.5016 | Same as above plus -> Previous analyses at SAW26 (NEFSC 1998) omitted depletion tow 10, which was included here |
| SC 1997-6 | AC1-2 | $\underset{\text { (Atlantic City) }}{\text { NNJ }}$ | 39.36500 | 73.89833 | 30 |  | Judy Marie | 6/11/1997 | Loran / 9-12 <br> M/ 1 Minute | 19 [4] | 8.33 | 16.67 | 0.0171 | 0.8902 | 7.1339 | 0.5 | 99.2 | $\begin{gathered} \text { Same as } \\ \text { SC } 1997-5\} \\ \hline \end{gathered}$ | 0.0069 | 0.1173 | 0.4022 | Same as above plus -> Previous analyses at SAW26 (NEFSC 1998) omitted depletion tows 17 and 19, which were included here |
| SC 1999-2 | $\begin{gathered} \hline \mathrm{JG}-1 \text { (S99 } \\ 5 \text { ) } \\ \hline \end{gathered}$ | NNJ | 39.68133 | 73.74667 | 24 | 0.88 | Jersey Girl | 9/14/1999 | $\begin{aligned} & \text { Loran } / 9-12 \\ & \mathrm{M} / 1 \text { Minute } \\ & \hline \end{aligned}$ | 4 [1] | 10.83 | 21.67 | 0.0249 | 0.8453 | 10.2855 | 0.5 | 21.5 | $\begin{array}{\|c\|} \hline 199903 \text { [105- } \\ 108](4)\{4\} \\ \hline \end{array}$ | 0.0075 | 0.2273 | 0.3004 |  |
| SC 1999-3 | $\begin{gathered} \mathrm{JG}-2 \text { (s99-1 } \\ 5 \text { ) } \end{gathered}$ | NNJ | 39.68133 | 73.74667 | 24 | 0.88 | Jersey Girl | 9/14/1999 | $\begin{aligned} & \text { Loran / } 9-12 \\ & \text { M/1 Minute } \end{aligned}$ | 5 [2] | 10.83 | 21.67 | 0.0631 | 0.4625 | 9.3468 | 0.5 | 30.0 | $\begin{gathered} \text { Same as } \\ \text { SC1999-2 } \end{gathered}$ | 0.0075 | 0.2273 | 0.1186 |  |
| SC 1999-4 | $\begin{gathered} \hline \text { JG-3 (S99 } \\ 6 \text { ) } \end{gathered}$ | NNJ | 39.52133 | 73.77867 | 26 | 0.67 | Jersey Girl | 9/14/1999 | $\begin{aligned} & \text { Loran } / 9-12 \\ & \text { M / } 1 \text { Minute } \\ & \hline \end{aligned}$ | 6 [2] | 10.83 | 21.67 | 0.0251 | 0.9900 | 15.3974 | 0.5 | 31.5 | $\begin{array}{\|l\|} \hline 199903(112- \\ 115](4)\{4\} \\ \hline \end{array}$ | 0.0050 | 0.1398 | 0.1990 |  |
| SC 1999-5 | $\begin{gathered} \mathrm{CH}-1 \text { (S999 } \\ \mathrm{DEII}) \end{gathered}$ | DMV | 36.90200 | 74.97583 | 35 | 1.13 | Christy | 9/25/1999 | Loran /9-12 <br> M/1 Minute | 28 [6] | 10.83 | 21.67 | 0.0193 | 0.1641 | 5.6765 | 0.5 | 92.8 | $\begin{aligned} & 19903[367- \\ & 370](4)\{0\} \end{aligned}$ |  |  |  | No length data for setup tows |
| SC 1999-6 | $\begin{gathered} \hline \mathrm{MJ}-1 \text { ( } \mathrm{s99-} \\ 3, \mathrm{NJ} \\ \text { Inshore } \\ \text { Site 1) } \\ \hline \end{gathered}$ | NNJ | 39.56333 | 73.91167 | 26 | 1.08 | Melissa J | 9/28/1999 | Loran / 9-12 M / 1 Minute | 4 [1] | 10.83 | 21.67 | 0.0245 | 0.8357 | 32.4987 | 0.5 | 18.7 | $\begin{array}{\|c\|c\|} \hline 199903[82- \\ 85](4)\{4\} \end{array}$ | 0.0058 | 0.4363 | 0.2374 | Sarc31 list Blade at 13 |
| SC 1999-7 | $\begin{gathered} \hline \mathrm{MJ-1} \text { (s999-1 } \\ 3, \mathrm{NJ} \\ \text { Inshore } \\ \text { Sito } 2 \text { ) } \end{gathered}$ | NNJ | 39.76800 | 73.91633 | 24 | 3.85 | Melissa J | 9/28/1999 | Loran /9-12 M/1 Minute | 10 [2] | 10.83 | 21.67 | 0.0513 | 0.6164 | 49.9988 | 0.5 | 52.2 | $\begin{array}{\|c\|c\|} \hline 199903[88- \\ 90](3)\{3\} \end{array}$ | 0.0046 | 0.1742 | 0.0888 | Sarc31 list Blade at 13 |
| SC2002-2 | SC02-2 | NNJ | 40.10908 | 73.84423 | 38 | 0.43 | Jersey Girl | 8/20/2002 | GPS-DT3M1/2 sec. | 16 [3] | 10.83 | 21.67 | 0.0144 | 0.8610 | 30.7464 | 0.5 | 74.1 | $\begin{gathered} 200206[87- \\ 91](5)\{1\} \\ \hline \end{gathered}$ | 0.0037 | 0.2774 | 0.2563 |  |
| SC2002-3 | SC02-3 | SNJ | 39.26923 | 73.78116 | 31 | 1.12 | Jersey Girl | 8/19/2002 | GPS-D/3M/2 sec. | 19 [see footnote] | 10.83 | 21.67 | 0.0134 | 0.3071 | 2.8366 | 0.5 | 88.3 | $\begin{gathered} 200206[202- \\ 206](5)\{4\} \end{gathered}$ | 0.0104 | 0.6116 | 0.7748 | Depletion tows: 1) bushel count for tow 1 only (tows 2-19 had catch < 1 bu ); 2) clams counted for all tows because catches were low; and 3 ) lengths measured for 10 for tows. Setup tows: zero clams caught at setup tow (station 206); only two clams measured at station 205 |
| SC2002-4 | SC02-4 | DMV | 38.85791 | 74.40888 | 31 | 0.48 | Jersey Girl | 8/20/2002 | $\begin{aligned} & \text { GPS-D/3M/2 } \\ & \text { sec. } \end{aligned}$ | 18 [4] | 10.83 | 21.67 |  |  |  |  |  | $\begin{array}{\|c\|} \hline 200206[335- \\ 339](5)\{1\} \\ \hline \end{array}$ |  |  |  | Zero clams $>=150 \mathrm{~mm}$ in tows 1-3, very low and variable catches in other tows. |
| SC2004-1 | SC04-1 | NJ | 39.28611 | 73.87778 | 35 |  | Lisa Kim | 4/8/2004 | $\begin{gathered} \begin{array}{c} \text { GPS-D/3M/2 } \\ \text { sec. } \end{array} \\ \hline \end{gathered}$ | $\begin{gathered} 24 \text { [5] (see } \\ \text { note) } \end{gathered}$ | 10 | 20.00 | 0.0301241 | 0.8072 | 4.0810 | 0.5000 | 130.753 |  |  |  |  | 200416 Cooperative Survey (shakedown leg) stations 15-38 |
| SC2004-2 | SC04-2 | NJ | 39.58278 | 74.02778 | 21 |  | Lisa Kim | 4/8/2004 | $\begin{gathered} \substack{\text { GPS-C. } / 3 \mathrm{M} / 2 \\ \mathrm{sec} .} \\ \hline \end{gathered}$ | $\begin{gathered} 20 \text { [4] (see } \\ \text { note) } \\ \hline \end{gathered}$ | 10 | 20.00 | 0.017376 | 0.6646 | 7.7973 | 0.5000 | 102.320 |  |  |  |  | $\begin{aligned} & \text { latin } \\ & \text { stations Cooperative Survey (shakedown leg) } \end{aligned}$ |
| SC2004-3 | SC04-3 | DMV | 38.27075 | 74.37920 | 38 |  | Lisa Kim | 7/3/2004 | $\begin{gathered} \text { GPS-D/3M/10 } \\ \text { sec. } \end{gathered}$ | $\begin{gathered} 20[4] \text { (see } \\ \text { note) } \end{gathered}$ | 10 | 20.00 |  |  |  |  |  |  |  |  |  | 200416 Cooperative Survey stations 146-165; zero clams $>=150 \mathrm{~mm}$ in tow 1 , very low and variable catches in other tows. |
| SC2005-2 | SC05-02 | NNJ | 39.56383 | 73.90364 | 24 | 0.29 | Lisa Kim | 97/12005 | $\underset{\mathrm{sec}}{\mathrm{GPS} / 6 \mathrm{ft} / 6}$ | 17 [3] | 10.00 | 20.00 | 0.0407 | 0.7633 | 4.7110 | 0.5 | 98.5 |  | 0.004 | 0.3635 | 0.1035 |  |
| SC2005-3 | SC05-03 | NNJ | 39.89733 | 73.90591 | 38 | 0.24 | Lisa Kim | 9/8/2005 | $\underset{\text { sec }}{\operatorname{GPS} / 6 \mathrm{ft} / 6}$ | 20 [4] | 10.00 | 20.00 | 0.0590 | 0.5879 | 4.7883 | 0.5 | 120.6 | $\begin{gathered} \hline 200507[21, \\ 384-387](5) \\ \{3 ?\} \\ \hline \end{gathered}$ | 0.006 | 0.2999 | 0.1008 |  |
| SC2005-4 | SC05-04 | DMV | 39.56972 | 73.54946 | 41 | 0.20 | Lisa Kim | 9/9/2005 | $\underset{\text { sec }}{\mathrm{GPS} / 6 \mathrm{ft} / 6}$ | 20 [4] | 10.00 | 20.00 | 0.0264 | 0.5341 | 4.4756 | 0.5 | 104.5 | $\begin{array}{\|c} 200507[41, \\ 391-393, \\ 395](5)\{3\} \\ \hline \end{array}$ | 0.006 | 0.2597 | 0.2143 |  |
| SC2005-5 | SC05-05 | NNJ | 39.43615 | 73.37320 | 33 | 0.28 | Lisa Kim | 9/10/2005 | $\underset{\text { sec }}{\substack{\text { GPS } / 6 \mathrm{ft} / 6 \\ \hline}}$ | 17 [4] | 10.00 | 20.00 | 0.0212 | 0.9823 | 2.3360 | 0.5 | 96.1 | $\begin{array}{\|c} 200507[143, \\ 397-402](7) \\ \{5\} \end{array}$ | 0.004 | 0.1809 | 0.2119 |  |
| SC2005-6 | SC05-01 | NNJ | 38.26530 | 74.37947 | 26 | 0.19 | Lisa Kim | 9/7/2005 | $\underset{\text { sec }}{\mathrm{GPS} / 6 \mathrm{ft} / 6}$ | 20 [4] | 10.00 | 20.00 |  |  |  |  |  | $\begin{aligned} & 200507(123- \\ & 127,354](6) \\ & \{6\} \end{aligned}$ |  |  |  | Low catches >= 150 mm SL in setup and depletion tows (less than 6\% of total). |

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Table C15. Sensitivity analysis to determine effects of smoothing position data on Patch model estimates for surfclam 150+ mm SL. Model runs were preliminary.

| Smoothing | Density | Efficiency | K | Gamm | Goodness <br> na to catch d log-likelih | of fit ata (ood) | Area swept ( $\mathrm{ft}^{2}$ ) | Effective area swept (ft ${ }^{2}$ ) | Number stations | Number position observations | Mean Observations per station | Shape of tow tracks based on original position data | Type Smooth |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| SC1997-4 |  |  |  |  |  |  |  |  |  |  |  |  |  |
| No | 0.0147 | 0.99 | 3.42 | 0.50 | 95.38 |  | 296,196 | 134,886 | 18 | 107 | 5.9 | Curved, stair step | Linear or cubic polynomial |
| Yes | 0.0157 | 0.99 | 3.24 | 0.50 | 95.82 |  | 288,202 | 133,535 |  |  |  |  |  |
| \% Difference | -7\% | 0\% | 6\% | 0\% | 0\% |  | 3\% | 1\% |  |  |  |  |  |
|  | SC1999-2 |  |  |  |  |  |  |  |  |  |  |  |  |
| No | 0.0206 | 0.95 | 21.53 | 0.50 | 20.22 |  | 53,854 | 35,480 | 4 | 46 | 11.5 | Linear, stair step | Quadratic polynomial |
| Yes | 0.0249 | 0.85 | 10.29 | 0.50 | 21.54 |  | 52,268 | 30,236 |  |  |  |  |  |
| \% Difference | -17\% | 12\% | 109\% | 0\% | -6\% |  | 3\% | 17\% |  |  |  |  |  |
| SC2002-2 |  |  |  |  |  |  |  |  |  |  |  |  |  |
| No | 0.0134 | 0.94 | 30.21 | 0.50 | 74.42 |  | 279,668 | 162,865 | 16 | 2664 | 166.5 | Linear or curved, | Spline |
| Yes | 0.0144 | 0.86 | 30.75 | 0.50 | 74.13 |  | 276,021 | 152,182 |  |  |  |  |  |
| \% Difference | -7\% | 9\% | -2\% | 0\% | 0\% |  | 1\% | 7\% |  |  |  |  |  |
| SC2005-2 |  |  |  |  |  |  |  |  |  |  |  |  |  |
| No | 0.0350 | 0.93 | 5.35 | 0.50 | 97.55 |  | 287,369 | 117,195 | 17 | 905 | 53.2 | Curved, wavy | Spline |
| Yes | 0.0407 | 0.76 | 4.71 | 0.50 | 98.55 |  | 283,682 | 100,759 |  |  |  |  |  |
| \% Difference | -14\% | 22\% | 14\% | 0\% | -1\% |  | 1\% | 16\% |  |  |  |  |  |
| Table C16. NEFSC survey dredge efficiency estimates for surfclam in the 1997-2005 NEFSC clam surveys (revised and values used in the last |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  | NEFSC survey dredge efficiency estimates for surfclam in the 1997-2005 NEFSC clam surveys (revised and values used in the last assessment). CVs are the standard error / mean. "NA" means not available. Efficiency estimates shown in the table are averages, not |  |  |  |  |  |  |  |  |  |  |  |  |
| Survey year | Revised (Patch model \& setup tows) |  |  |  | From previous assessment (various types of estimates, NEFSC 2003) |  |  | ent tes, |  |  |  |  |  |
|  | Efficiency | y CV | $N$ |  | Efficiency | CV |  | N |  |  |  |  |  |
| 1997 | 0.317 | 0.19 |  | 5 | 0.460 | 0.47 |  | 4 |  |  |  |  |  |
| 1999 | 0.189 | 0.20 |  | 5 | 0.276 | 0.349 |  | 5 |  |  |  |  |  |
| 2002 | 0.516 | 0.50 |  | 3 | 0.389 | 0.52 |  | 6 |  |  |  |  |  |
| 2005 | 0.158 | 0.20 |  | 5 | NA | NA |  | NA |  |  |  |  |  |
| All | 0.262 | 0.17 |  | 18 | 0.370 | 0.492 |  | 15 |  |  |  |  |  |

Table C17. Efficiency corrected swept-area biomass estimates ( 1000 mt ) and CVs for the fishable stock of surfclam during 1997-2005 by stock assessment region. Figures for SVA and GBK during 2005 were taken from 2003 because no data were available for 2005.

| INPUT: Nominal tow distance ( $\mathrm{d}_{n}, n m$ ) INPUT: Dredge width (nm) | Estimate | CV |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 0.15 |  |  |  |  |  |  |  |
|  | 0.0008225 |  |  |  |  |  |  |  |
| Area swept per standard tow ( $a, \mathrm{~nm}^{2}$ ) | $1.23375 \mathrm{E}-04$ | 10\% |  |  |  |  |  |  |
| Area of assessment region ( $A, \mathrm{~nm}^{\mathbf{2}}$ ) - no correction for stations with unsuitable clam habitat |  |  |  |  |  |  |  |  |
| S. Virginia and N. Carolina (SVA) | 3,119 | 10\% |  |  |  |  |  |  |
| Delmarva (DMV) | 4,660 | 10\% |  |  |  |  |  |  |
| New Jersey ( NJ ) | 5,078 | 10\% |  |  |  |  |  |  |
| Long Island (LI) | 2,917 | 10\% |  |  |  |  |  |  |
| Southern New England (SNE) | 4,321 | 10\% |  |  |  |  |  |  |
| Georges Bank (GBK) | 5,772 | 10\% |  |  |  |  |  |  |
| Total | 25,867 |  |  |  |  |  |  |  |
| INPUT: Fraction suitable habitat ( $u$ ) |  |  |  |  |  |  |  |  |
| S. Virginia and N. Carolina (SVA) | 100\% | 10\% |  |  |  |  |  |  |
| Delmarva (DMV) New Jersey (NJ) | 100\% | 10\% |  |  |  |  |  |  |
| Long Island (LI) | 100\% | 10\% |  |  |  |  |  |  |
| Southern New England (SNE) | 100\% | 10\% |  |  |  |  |  |  |
| Georges Bank (GBK) | 88\% | 10\% |  |  |  |  |  |  |
| Habitat area in assessment region ( $\left.A^{\prime}, \mathrm{nm} 2\right)$ |  |  | INPUT: Biomass fraction in unsurveyd deep water |  |  |  |  |  |
| S. Virginia and N. Carolina (SVA) | 3,119 | 14\% |  | S. Virginia and | Carolina (SVA) | 0\% | 10\% |  |
| Delmarva (DMV) | 4,660 | 14\% |  |  | Delmarva (DMV) | 0\% | 10\% |  |
| New Jersey ( NJ ) | 5,078 | 14\% |  |  | New Jersey (NJ) | 0\% | 10\% |  |
| Long Island (LI) | 2,917 | 14\% |  |  | Long Island (LI) | 0\% | 10\% |  |
| Southern New England (SNE) | 4,321 | 14\% |  | Southern New | England (SNE) | 0\% | 10\% |  |
| Georges Bank (GBK) | 5,079 | 14\% |  |  | ges Bank (GBK) | 0\% | 10\% |  |
| INPUT: Original survey mean catch from fishable stock (kg/tow, for tows adjusted to nominal tow distance using sensors) |  |  |  |  |  |  |  |  |
|  | Estimates for 1997 | cv | Estimates for $1999$ | CV | $\begin{array}{\|c\|} \hline \text { Estimates for } \\ 2002 \end{array}$ | CV | $\begin{aligned} & \text { Estimates for } \\ & 2005 \end{aligned}$ | CV |
| S. Virginia and N. Carolina (SVA) $120+\mathrm{mm}$ | 0.0142 | 43\% | 0.0532 | 52\% | 0.2676 | 58\% | 0.2676 | 58\% |
| Delmarva (DMV) $120+\mathrm{mm}$ | 2.3751 | 22\% | 1.4130 | 20\% | 2.2406 | 20\% | 0.4038 | 30\% |
| New Jersey (NJ) $120+\mathrm{mm}$ | 5.8453 | 12\% | 4.0036 | 17\% | 3.5823 | 16\% | 2.1776 | 17\% |
| Long Island (LI) $120+\mathrm{mm}$ | 0.3179 | 66\% | 0.7895 | 53\% | 0.1849 | 64\% | 1.9644 | 37\% |
| Southern New England (SNE) $120+\mathrm{mm}$ | 0.8868 | 32\% | 0.4839 | 67\% | 0.4180 | 26\% | 0.1851 | 19\% |
| Georges Bank (GBK) 120+mm | 1.5228 | 25\% | 2.0445 | 31\% | 1.8469 | 35\% | 1.8469 | 35\% |
| Swept-area biomass without efficiency correction (B', 1000 mt ): |  |  |  |  |  |  |  |  |
| S. Virginia and N. Carolina (SVA) $120+\mathrm{mm}$ | 0.3597 | 47\% | 1.3447 | 56\% | 6.7651 | 61\% | 6.7651 | 61\% |
| Delmarva (DMV) $120+\mathrm{mm}$ | 89.7081 | 30\% | 53.3720 | 28\% | 84.6301 | 28\% | 15.2519 | 36\% |
| New Jersey (NJ) $120+\mathrm{mm}$ | 240.5850 | 23\% | 164.7861 | 26\% | 147.4441 | 26\% | 89.6280 | 26\% |
| Long Island (LI) $120+\mathrm{mm}$ | 7.5155 | 69\% | 18.6664 | 57\% | 4.3707 | 67\% | 46.4441 | 42\% |
| Southern New England (SNE) $120+\mathrm{mm}$ | 31.0590 | 38\% | 16.9471 | 70\% | 14.6411 | 33\% | 6.4817 | 28\% |
| Georges Bank (GBK) $120+\mathrm{mm}$ | 62.6950 | 32\% | 84.1714 | 37\% | 76.0380 | 40\% | 76.0380 | 40\% |
| Total fishable biomass less GBK | 369 | 17\% | 255 | 19\% | 258 | 18\% | 165 | 19\% |
| Total fishable biomass | 432 | 15\% | 339 | 17\% | 334 | 16\% | 241 | 18\% |
| INPUT: Survey dredge efficiency (e) | 0.226 | 17\% | 0.226 | 17\% | 0.226 | 17\% | 0.226 | 17\% |
| Efficiency adjusted swept area fishable biomass (B, 1000 mt ) |  |  |  |  |  |  |  |  |
| S. Virginia and N. Carolina (SVA) $120+\mathrm{mm}$ | 1.593 | 50\% | 5.955 | 58\% | 29.961 | 64\% | 29.961 | 64\% |
| Delmarva (DMV) $120+\mathrm{mm}$ | 397 | 34\% | 236 | 33\% | 375 | 33\% | 68 | 40\% |
| New Jersey (NJ) $120+\mathrm{mm}$ | 1,065 | 29\% | 730 | 31\% | 653 | 31\% | 397 | 31\% |
| Long Island (LI) $120+\mathrm{mm}$ | 33 | 71\% | 83 | 59\% | 19 | 69\% | 206 | 45\% |
| Southern New England (SNE) $120+\mathrm{mm}$ | 138 | 41\% | 75 | 72\% | 65 | 37\% | 29 | 32\% |
| Georges Bank (GBK) $120+\mathrm{mm}$ | 278 | 36\% | 373 | 41\% | 337 | 44\% | 337 | 44\% |
| Total fishable biomass less GBK | 1,635 | 24\% | 1,130 | 25\% | 1,142 | 24\% | 729 | 25\% |
| Total fishable biomass | 1,913 | 23\% | 1,503 | 24\% | 1,479 | 23\% | 1,066 | 25\% |
| Lower bound for $\mathbf{8 0 \%}$ confidence intervals on fishable biomass ( 1000 mt , for lognormal distribution with no bias correction) |  |  |  |  |  |  |  |  |
|  | Estimates for 1997 | Estimates for 1999 | Estimates for 2002 | $\begin{array}{\|c} \hline \text { Estimates for } \\ 2005 \end{array}$ |  |  |  |  |
| S. Virginia and N. Carolina (SVA) $120+\mathrm{mm}$ | 0.867 | 2.983 | 14.208 | 14.208 |  |  |  |  |
| Delmarva (DMV) $120+\mathrm{mm}$ | 260 | 157 | 249 | 41 |  |  |  |  |
| New Jersey (NJ) $120+\mathrm{mm}$ | 743 | 494 | 445 | 269 |  |  |  |  |
| Long Island (LI) $120+\mathrm{mm}$ | 15 | 41 | 9 | 118 |  |  |  |  |
| Southern New England (SNE) $120+\mathrm{mm}$ | 83 | 33 | 41 | 19 |  |  |  |  |
| Georges Bank (GBK) $120+\mathrm{mm}$ | 177 | 226 | 197 | 197 |  |  |  |  |
| Total fishable biomass less GBK | 1,207 | 821 | 840 | 529 |  |  |  |  |
| Total fishable biomass | 1,434 | 1,112 | 1,100 | 780 |  |  |  |  |
| Upperbound for $\mathbf{8 0 \%}$ confidence intervals on fishable biomass (1000 mt, for lognormal distribution with no bias correction) |  |  |  |  |  |  |  |  |
| S. Virginia and N. Carolina (SVA) $120+\mathrm{mm}$ | 2.926 | 11.888 | 63.180 | 63.180 |  |  |  |  |
| Delmarva (DMV) $120+\mathrm{mm}$ | 608 | 356 | 565 | 110 |  |  |  |  |
| New Jersey (NJ) $120+\mathrm{mm}$ | 1,528 | 1,078 | 958 | 586 |  |  |  |  |
| Long Island (LI) $120+\mathrm{mm}$ | 75 | 167 | 43 | 358 |  |  |  |  |
| Southern New England (SNE) $120+\mathrm{mm}$ | 229 | 172 | 103 | 43 |  |  |  |  |
| Georges Bank (GBK) 120+mm | 435 | 614 | 574 | 574 |  |  |  |  |
| Total fishable biomass less GBK | 2,215 | 1,555 | 1,552 | 1,004 |  |  |  |  |
| Total fishable biomass | 2,551 | 2,031 | 1,988 | 1,456 |  |  |  |  |

Table C18. Fishing mortality estimates for surfclams based on catch and efficiency corrected swept-area biomass for fishable surfclams during 1997, 1999, 2002 and 2005.

| INPUT: Upper bound incidental mortality allowance | 12\% |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| INPUT: Assumed CV for catch | 10\% |  |  |  |  |  |  |  |
| INPUT: Landings (1000 mt, discard ~ 0) | Estimates for 1997 | Estimates for 1999 | Estimates for 2002 | Estimates for 2005 |  |  |  |  |
| S. Virginia and N. Carolina (SVA) | 0.000 | 0.000 | 0.064 | 0.000 |  |  |  |  |
| Delmarva (DMV) | 1.540 | 0.648 | 4.489 | 1.668 |  |  |  |  |
| New Jersey ( NJ ) | 16.998 | 18.749 | 18.271 | 16.850 |  |  |  |  |
| Long Island (LI) | 0.073 | 0.157 | 1.130 | 0.759 |  |  |  |  |
| Southern New England (SNE) | 0.000 | 0.016 | 0.052 | 1.885 |  |  |  |  |
| Georges Bank (GBK) | 0.000 | 0.000 | 0.000 | 0.000 |  |  |  |  |
| Total | 18.611 | 19.570 | 24.006 | 21.163 |  |  |  |  |
| Catch (1000 mt, landings + upper bound incidental mortality allowance) |  |  |  |  |  |  |  |  |
| S. Virginia and N. Carolina (SVA) | 0.000 | 0.000 | 0.072 | 0.000 |  |  |  |  |
| Delmarva (DMV) | 1.725 | 0.726 | 5.028 | 1.868 |  |  |  |  |
| New Jersey (NJ) | 19.038 | 20.999 | 20.463 | 18.872 |  |  |  |  |
| Long Island (LI) | 0.081 | 0.176 | 1.265 | 0.850 |  |  |  |  |
| Southern New England (SNE) | 0.000 | 0.018 | 0.058 | 2.112 |  |  |  |  |
| Georges Bank (GBK) | 0.000 | 0.000 | 0.000 | 0.000 |  |  |  |  |
| Total | 20.844 | 21.919 | 26.886 | 23.702 |  |  |  |  |
| INPUT: Efficiency Corrected Swept Area Biomass for Fishable Stock | Estimates for |  | Estimates for |  | Estimates for |  | Estimates for |  |
| ( 1000 mt ) | 1997 | CV | 1999 | CV | 2002 | CV | 2005 | CV |
| S. Virginia and N. Carolina (SVA) $120+\mathrm{mm}$ | 2 | 50\% | 6 | 58\% | 30 | 64\% | 30 | 64\% |
| Delmarva (DMV) $120+\mathrm{mm}$ | 397 | 34\% | 236 | 33\% | 375 | 33\% | 68 | 40\% |
| New Jersey (NJ) 120+ mm | 1,065 | 29\% | 730 | 31\% | 653 | 31\% | 397 | 31\% |
| Long Island (LI) $120+\mathrm{mm}$ | 33 | 71\% | 83 | 59\% | 19 | 69\% | 206 | 45\% |
| Southern New England (SNE) 120+mm | 138 | 41\% | 75 | 72\% | 65 | 37\% | 29 | 32\% |
| Georges Bank (GBK) 120+ mm | 278 | 36\% | 373 | 41\% | 337 | 44\% | 337 | 44\% |
| Total fishable biomass less GBK | 1,635 | 24\% | 1,130 | 25\% | 1,142 | 24\% | 729 | 25\% |
| Total fishable biomass | 1,913 | 23\% | 1,503 | 24\% | 1,479 | 23\% | 1,066 | 25\% |
| Fishing mortality ( $\mathrm{y}^{-1}$ ) |  |  |  |  |  |  |  |  |
| S. Virginia and N. Carolina (SVA) $120+\mathrm{mm}$ | 0.0000 | 51\% | 0.0000 | 59\% | 0.0024 | 64\% | 0.0000 | 64\% |
| Delmarva (DMV) $120+\mathrm{mm}$ | 0.0043 | 36\% | 0.0031 | 34\% | 0.0134 | 34\% | 0.0277 | 41\% |
| New Jersey (NJ) 120+ mm | 0.0179 | 30\% | 0.0288 | 33\% | 0.0313 | 32\% | 0.0475 | 33\% |
| Long Island (LI) $120+\mathrm{mm}$ | 0.0024 | 72\% | 0.0021 | 60\% | 0.0654 | 70\% | 0.0041 | 46\% |
| Southern New England (SNE) 120+mm | 0.0000 | 42\% | 0.0002 | 73\% | 0.0009 | 38\% | 0.0736 | 34\% |
| Georges Bank (GBK) 120+ mm | 0.0000 | NA | 0.0000 | NA | 0.0000 | NA | 0.0000 | NA |
| Total fishable biomass less GBK | 0.0127 | 26\% | 0.0194 | 27\% | 0.0235 | 26\% | 0.0325 | 27\% |
| Total fishable biomass | 0.0109 | 25\% | 0.0146 | 26\% | 0.0182 | 25\% | 0.0222 | 27\% |
| Lower bound for $\mathbf{8 0 \%}$ confidence intervals for fishing mortality ( $\mathbf{y}^{-1}$, for lognormal distribution with no bias correction) | Estimates for 1997 | Estimates for 1999 | Estimates for 2002 | $\begin{aligned} & \text { Estimates for } \\ & 2005 \end{aligned}$ |  |  |  |  |
| S. Virginia and N. Carolina (SVA) $120+\mathrm{mm}$ | NA | NA | 0.0011 | NA |  |  |  |  |
| Delmarva (DMV) $120+\mathrm{mm}$ | 0.0028 | 0.0020 | 0.0087 | 0.0167 |  |  |  |  |
| New Jersey (NJ) 120+ mm | 0.0122 | 0.0191 | 0.0210 | 0.0316 |  |  |  |  |
| Long Island (LI) $120+\mathrm{mm}$ | 0.0011 | 0.0010 | 0.0292 | 0.0023 |  |  |  |  |
| Southern New England (SNE) $120+\mathrm{mm}$ | NA | 0.0001 | 0.0006 | 0.0483 |  |  |  |  |
| Georges Bank (GBK) 120+ mm | NA | NA | NA | NA |  |  |  |  |
| Total fishable biomass less GBK | 0.0092 | 0.0138 | 0.0169 | 0.0231 |  |  |  |  |
| Total fishable biomass | 0.0080 | 0.0105 | 0.0132 | 0.0159 |  |  |  |  |
| Upper bound for $\mathbf{8 0 \%}$ confidence intervals for fishing mortality $\left(\boldsymbol{y}^{-1}\right.$, for lognormal distribution with no bias correction) |  |  |  |  |  |  |  |  |
| S. Virginia and N. Carolina (SVA) $120+\mathrm{mm}$ | NA | NA | 0.0051 | NA |  |  |  |  |
| Delmarva (DMV) $120+\mathrm{mm}$ | 0.0068 | 0.0047 | 0.0206 | 0.0458 |  |  |  |  |
| New Jersey (NJ) 120+ mm | 0.0262 | 0.0433 | 0.0469 | 0.0715 |  |  |  |  |
| Long Island (LI) $120+\mathrm{mm}$ | 0.0056 | 0.0043 | 0.1465 | 0.0073 |  |  |  |  |
| Southern New England (SNE) $120+\mathrm{mm}$ | NA | 0.0006 | 0.0014 | 0.1121 |  |  |  |  |
| Georges Bank (GBK) 120+mm | NA | NA | NA | NA |  |  |  |  |
| Total fishable biomass less GBK | 0.0177 | 0.0273 | 0.0328 | 0.0458 |  |  |  |  |
| Total fishable biomass | 0.0149 | 0.0202 | 0.0251 | 0.0311 |  |  |  |  |

Table C19. Patch model estimates for surfclam depletion experiments carried out by the F/V Lisa Kim during 2004-2005.

| Statistic | N successful <br> experiments | Population <br> Density <br> $\left(\mathrm{N} \mathrm{ft}^{-2}\right)$ | Depletion <br> Vessel <br> Efficiency | $k$ | Setup <br> Density <br> $\left(\mathrm{N} \mathrm{ft}^{-2}\right)$ | NEFSC <br> Dredge <br> Efficiency |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Mean | F/V Lisa Kim (2004-2005) |  |  |  |  |  |
| Median | 6 commercial | 0.032 | 0.028 | 0.723 | 4.698 | 0.0051 |
| Lower 80\% bound | depletion, 4 | 0.023 | 0.625 | 4.593 | 0.0051 | 0.158 |
| Upper 80\% bound | with setup | 0.042 | 0.822 | 5.763 | 0.0044 | 0.1051 |
| SE | tows | 0.006 | 0.067 | 0.722 | 0.0004 | 0.2101 |
| CV (SE / Mean) |  | 0.192 | 0.092 | 0.154 | 0.0839 | 0.203 |

Table C20. Revised surfclam efficiency corrected swept-area abundance and biomass estimates ( $120+$ SL) from the cooperative 2004 clam survey and assuming dredge efficiency $E=0.714$. Estimates from Weinberg et al. (2005) assuming $E=0.792$ are shown for comparison.

| Statistics | Revised |  | SVA |
| :---: | :---: | :---: | :---: |
|  |  | DMV | NJ |
| N tows | 15 | 77 | 110 |
| Mean n/tow | 0.143 | 23.253 | 71.079 |
| Var | 0.012 | 35.412 | 82.763 |
| CV | 0.78 | 0.26 | 0.13 |
| Mean kg/tow | 0.011 | 2.365 | 10.863 |
| Var | 0.000 | 0.348 | 1.907 |
| CV | 0.81 | 0.25 | 0.13 |
| Area (sq nm) | 1,074 | 4,660 | 5,078 |
| Efficiency | 0.714 | 0.714 | 0.714 |
| Swept area abundance (10 ${ }^{6}$ clams) | 1.7 | 1230.3 | 4098.1 |
| Swept area biomass (mt) | 128 | 125,139 | 626,302 |
| Weinberg et al. (2005) |  |  |  |
| Swept area biomass (mt) | 300 | 143,000 | 535,000 |

Table C21. Efficiency corrected swept-area biomass estimates (1,000 mt) for SVA, DMV and NJ , which were covered during the 2004 cooperative surfclam survey.

| Region | 1997 | 1999 | 2002 | 2004 | 2005 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| SVA | 1.59 | 5.96 | 29.96 | 0.13 | 29.96 |
| DMV | 397 | 236 | 375 | 125 | 68 |
| NJ | 1,065 | 730 | 653 | 626 | 397 |
| Total | 1,464 | 972 | 1,058 | 752 | 494 |

Table C22. Von Bertalanffy growth model parameters for surfclam weight at age in the DMV and NJ regions during 1982-1992 and 1994-2005 based on NEFSC survey data with estimates of meat weight ( $W$, grams) and shell length $(L, \mathrm{~mm}$ ) at the age of recruitment $(k)$, one year before recruitment $(k-1)$ and one year after recruitment $(k+1)$. The parameters for NJ were also used for the whole stock.

| Parameter | NJ |  | DMV |  |
| :---: | :---: | :---: | :---: | :---: |
|  | $1982-1992$ | $1994-20051982-19921994-2005$ |  |  |
| $K\left(y^{-1}\right)$ | 0.1449 |  | 0.1258 |  |
| $r$ | 0.8651 | 0.8818 |  |  |
| $t 0(y)$ | 1.5365 | 1.6919 | 1.5176 | 1.6026 |
| $W_{\max }$ | 240.5 | 206.8 | 197.4 | 138.0 |
| Age at recruitment $k(\mathrm{y})$ | 4.4 | 5.1 | 4.8 | 6.9 |
| $W_{k-1}$ | 56.4 | 61.7 | 49.5 | 57.4 |
| $W_{k}$ | 81.3 | 81.3 | 66.9 | 66.9 |
| $W_{k+1}$ | 102.8 | 98.2 | 82.4 | 75.3 |
| $L_{k-1}$ | 105.7 | 109.0 | 107.9 | 113.7 |
| $L_{k}$ | 120.0 | 120.0 | 120.0 | 120.0 |
| $L_{k+1}$ | 130.2 | 128.2 | 129.0 | 125.1 |
| $J$ | 0.6945 | 0.7592 | 0.7388 | 0.8578 |

Table C23. NEFSC survey index trend data (doppler tow distance measurements) used in KLAMZ models for surfclam recruits and

|  | DMV recruit index |  |  | $N J$ recruit index |  |  | Whole stock recruit index |  |  | Whole stock fishable biomass index |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Kg/Tow | CV | $\begin{gathered} \text { Size } \\ \text { groups } \\ \text { (mm SL) } \end{gathered}$ | Kg/Tow | CV | Size (mm SL ) | Kg/Tow | CV |  | Mean kg/tow | Approx. swept area biomass (mt) | CV | Size groups ( mm SL ) |
| 1982 | 0.598 | 0.34 | 120-129 | 0.976 | 0.24 | 120-129 | 0.350 | 0.19 | 120-129 | 1.756 | 1,069 | 0.15 | 120+ |
| 1983 | 1.177 | 0.78 | 120-129 | 0.882 | 0.34 | 120-129 | 0.448 | 0.45 | 120-129 | 1.615 | 983 | 0.18 | 120+ |
| 1984 | 0.846 | 0.3 | 120-129 | 0.776 | 0.27 | 120-129 | 0.396 | 0.18 | 120-129 | 2.149 | 1,308 | 0.13 | 120+ |
| 1986 | 3.165 | 0.53 | 120-129 | 0.493 | 0.29 | 120-129 | 0.815 | 0.43 | 120-129 | 2.695 | 1,641 | 0.19 | 120+ |
| 1989 | 0.745 | 0.31 | 120-129 | 0.508 | 0.2 | 120-129 | 0.352 | 0.22 | 120-129 | 2.290 | 1,395 | 0.22 | $120+$ |
| 1992 | 0.730 | 0.57 | 120-129 | 0.399 | 0.28 | 120-129 | 0.313 | 0.3 | 120-129 | 1.601 | 975 | 0.14 | 120+ |
| 1994 | 1.328 | 0.33 | 120-125 | 1.536 | 0.23 | 120-129 | 0.792 | 0.16 | 120-129 | 4.613 | 2,809 | 0.14 | $120+$ |
| 1997 | 1.933 | 0.34 | 120-125 | 1.060 | 0.22 | 120-129 | 0.790 | 0.19 | 120-129 | 3.989 | 2,429 | 0.09 | 120+ |
| 1999 | 0.989 | 0.29 | 120-125 | 0.707 | 0.58 | 120-129 | 0.550 | 0.21 | 120-129 | 2.717 | 1,654 | 0.13 | 120+ |
| 2002 | 0.380 | 0.34 | 120-125 | 0.242 | 0.24 | 120-129 | 0.296 | 0.28 | 120-129 | 2.481 | 1,510 | 0.14 | 120+ |
| 2005 | 0.075 | 0.26 | 120-125 | 0.193 | 0.24 | 120-129 | 0.101 | 0.24 | 120-129 | 1.339 | 815 | 0.17 | 120+ |

Table C24. Estimated biomass, recruitment biomass and fishing mortality for the entire surfclam stock from the KLAMZ model. CVs are from 1000 bootstrap iterations.

| Year | Biomass <br> $(1000$ <br> $\mathrm{mt})$ | CV | Recruitment <br> $(1000 \mathrm{mt})$ | CV | Fishing <br> mortality <br> $1)$ $\mathrm{y}^{-}$ | CV |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1981 | 1,020 | 0.26 | NA |  | 0.0173 | 0.25 |
| 1982 | 1,036 | 0.23 | 96 | 0.17 | 0.0231 | 0.22 |
| 1983 | 1,059 | 0.20 | 110 | 0.20 | 0.0229 | 0.19 |
| 1984 | 1,083 | 0.18 | 109 | 0.17 | 0.0266 | 0.17 |
| 1985 | 1,141 | 0.16 | 147 | 0.26 | 0.0241 | 0.15 |
| 1986 | 1,225 | 0.15 | 170 | 0.20 | 0.0231 | 0.15 |
| 1987 | 1,271 | 0.14 | 130 | 0.30 | 0.0193 | 0.15 |
| 1988 | 1,290 | 0.15 | 106 | 0.28 | 0.0200 | 0.15 |
| 1989 | 1,289 | 0.14 | 93 | 0.15 | 0.0187 | 0.15 |
| 1990 | 1,285 | 0.15 | 96 | 0.31 | 0.0205 | 0.15 |
| 1991 | 1,283 | 0.15 | 102 | 0.32 | 0.0172 | 0.15 |
| 1992 | 1,290 | 0.15 | 109 | 0.15 | 0.0184 | 0.15 |
| 1993 | 1,476 | 0.13 | 289 | 0.30 | 0.0153 | 0.14 |
| 1994 | 1,613 | 0.12 | 231 | 0.13 | 0.0141 | 0.13 |
| 1995 | 1,709 | 0.09 | 201 | 0.33 | 0.0119 | 0.09 |
| 1996 | 1,780 | 0.07 | 185 | 0.32 | 0.0115 | 0.08 |
| 1997 | 1,842 | 0.07 | 189 | 0.14 | 0.0105 | 0.07 |
| 1998 | 1,824 | 0.05 | 116 | 0.35 | 0.0104 | 0.05 |
| 1999 | 1,799 | 0.04 | 121 | 0.17 | 0.0114 | 0.04 |
| 2000 | 1,723 | 0.04 | 76 | 0.36 | 0.0120 | 0.04 |
| 2001 | 1,628 | 0.04 | 62 | 0.36 | 0.0142 | 0.04 |
| 2002 | 1,531 | 0.04 | 63 | 0.18 | 0.0166 | 0.04 |
| 2003 | 1,415 | 0.05 | 43 | 0.24 | 0.0187 | 0.05 |
| 2004 | 1,292 | 0.05 | 32 | 0.22 | 0.0199 | 0.05 |
| 2005 | 1,170 | 0.06 | 27 | 0.16 | 0.0192 | 0.06 |

Table C25. Factors used to scale NEFSC survey trend data (mean $\mathrm{kg} / \mathrm{tow}, 120+\mathrm{mm}$, doppler tow distances) to approximate efficiency corrected swept-area biomass (based on sensor distance data and efficiency estimates.

| SVA | DMV | NJ |
| :---: | :---: | :---: |
| 68.462 | 119.917 | 114.584 |
| $L I$ | SNE | $G B K$ |
| 76.107 | 89.164 | 103.414 |

Table C26. Proportions of total fishable surfclam biomass during 1980-2005 at a range of survey biomass density levels, by region.

| Years | Fishable biomass density levels (kg/tow) from survey data |  |  |  |  |  | Sum of Proportions (check) | Total Total Number of Number of Tows Surveys |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 0 to 4 | 5 to 9 | 10 to 14 | 15 to 19 | 20 to 24 | 25+ |  |  |  |
| Proportions of tows (and stock area) at each survey catch rate level: |  |  |  |  |  |  |  |  |  |
| Southern Virgina (SVA) |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 0.94 | 0.03 | 0.01 | 0.02 | 0.01 |  | 1.00 | 154 | 5 |
| 1990-1999 | 0.99 |  |  | 0.01 |  |  | 1.00 | 107 | 3 |
| 2000-2005 | 1.00 |  |  |  |  |  | 1.00 | 29 | 2 |
| Delmarva (DMV) |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 0.81 | 0.10 | 0.03 | 0.01 | 0.01 | 0.04 | 1.00 | 355 | 5 |
| 1990-1999 | 0.78 | 0.11 | 0.05 | 0.03 | 0.02 | 0.02 | 1.00 | 237 | 3 |
| 2000-2005 | 0.90 | 0.05 | 0.03 | 0.01 |  | 0.01 | 1.00 | 152 | 2 |
| New Jersy (NJ) |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 0.71 | 0.12 | 0.06 | 0.04 | 0.02 | 0.05 | 1.00 | 484 | 5 |
| 1990-1999 | 0.56 | 0.13 | 0.10 | 0.05 | 0.05 | 0.11 | 1.00 | 330 | 3 |
| 2000-2005 | 0.69 | 0.11 | 0.11 | 0.03 | 0.01 | 0.04 | 1.00 | 206 | 2 |
| Long Island (LI) |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 0.97 | 0.02 | 0.01 |  |  |  | 1.00 | 170 | 5 |
| 1990-1999 | 0.95 | 0.02 | 0.01 |  | 0.01 |  | 1.00 | 86 | 3 |
| 2000-2005 | 0.93 | 0.02 | 0.02 | 0.04 |  |  | 1.00 | 57 | 2 |
| Southern New England (SNE) |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 0.87 | 0.03 | 0.05 | 0.02 | 0.01 | 0.02 | 1.00 | 202 | 5 |
| 1990-1999 | 0.90 | 0.02 | 0.04 |  | 0.01 | 0.02 | 1.00 | 90 | 3 |
| 2000-2005 | 0.96 | 0.04 |  |  |  |  | 1.00 | 48 | 2 |
| Georges Bank (GBK) |  |  |  |  |  |  |  |  |  |
| 1986-1992 | 0.87 | 0.05 | 0.03 | 0.02 | 0.01 | 0.02 | 1.00 | 126 | 3 |
| 1997-2002 | 0.79 | 0.04 | 0.04 | 0.02 | 0.01 | 0.10 | 1.00 | 119 | 3 |
| All years | 0.83 | 0.04 | 0.04 | 0.02 | 0.01 | 0.06 | 1.00 | 245 | 6 |
| Mean survey catch rate (kg/tow) at each survey catch rate level ( $p_{L}$ ): <br> Southern Virgina (SVA) |  |  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 0.267 | 6.206 | 11.779 | 16.929 | 21.086 |  |  |  |  |
| 1990-1999 | 0.119 |  |  | 17.468 |  |  |  |  |  |
| 2000-2005 | 0.171 |  |  |  |  |  |  |  |  |
| Delmarva (DMV) |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 0.853 | 7.528 | 11.999 | 16.412 | 21.738 | 50.956 |  |  |  |
| 1990-1999 | 0.820 | 7.348 | 12.039 | 17.431 | 22.697 | 50.709 |  |  |  |
| 2000-2005 | 0.518 | 6.800 | 11.471 | 17.350 |  | 25.869 |  |  |  |
| New Jersy (NJ) |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 1.018 | 7.559 | 12.270 | 17.662 | 22.426 | 52.603 |  |  |  |
| 1990-1999 | 0.939 | 7.343 | 12.017 | 17.518 | 22.016 | 45.320 |  |  |  |
| 2000-2005 | 1.216 | 7.215 | 12.195 | 15.867 | 22.468 | 32.093 |  |  |  |
| Long Island (LI) |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 0.095 | 6.554 | 13.132 |  |  |  |  |  |  |
| 1990-1999 | 0.240 | 6.216 | 11.010 |  | 23.237 |  |  |  |  |
| 2000-2005 | 0.121 | 7.404 | 10.151 | 17.446 |  |  |  |  |  |
| Southern New England (SNE) |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 0.311 | 8.573 | 11.768 | 18.272 | 22.628 | 43.811 |  |  |  |
| 1990-1999 | 0.118 | 7.898 | 12.033 |  | 20.543 | 30.708 |  |  |  |
| 2000-2005 | 0.640 | 6.301 |  |  |  |  |  |  |  |
| Georges Bank (GBK) |  |  |  |  |  |  |  |  |  |
| 1986-1992 | 0.223 | 8.360 | 10.987 | 17.530 | 21.017 | 85.534 |  |  |  |
| 1997-2002 | 0.500 | 7.110 | 10.928 | 17.167 | 22.838 | 40.544 |  |  |  |
| All years | 0.351 | 7.792 | 10.954 | 17.385 | 21.927 | 46.971 |  |  |  |
| Proportions of stock biomass at each survey catch rate level ( $\boldsymbol{X}_{L}$ ): <br> Southern Virgina (SVA) |  |  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 0.242 | 0.156 | 0.148 | 0.320 | 0.133 |  | 1.00 |  |  |
| 1990-1999 | 0.431 |  |  | 0.569 |  |  | 1.00 |  |  |
| 2000-2005 | 1.000 |  |  |  |  |  | 1.00 |  |  |
| Delmarva (DMV) |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 0.164 | 0.187 | 0.072 | 0.055 | 0.044 | 0.478 | 1.00 |  |  |
| 1990-1999 | 0.162 | 0.197 | 0.142 | 0.131 | 0.097 | 0.271 | 1.00 |  |  |
| 2000-2005 | 0.311 | 0.233 | 0.197 | 0.149 |  | 0.111 | 1.00 |  |  |
| New Jersy ( NJ ) |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 0.118 | 0.146 | 0.119 | 0.100 | 0.060 | 0.457 | 1.00 |  |  |
| 1990-1999 | 0.055 | 0.102 | 0.129 | 0.094 | 0.104 | 0.516 | 1.00 |  |  |
| 2000-2005 | 0.168 | 0.161 | 0.272 | 0.108 | 0.044 | 0.249 | 1.00 |  |  |
| Long Island (LI) |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 0.255 | 0.319 | 0.426 |  |  |  | 1.00 |  |  |
| 1990-1999 | 0.296 | 0.187 | 0.166 |  | 0.350 |  | 1.00 |  |  |
| 2000-2005 | 0.109 | 0.126 | 0.172 | 0.593 |  |  | 1.00 |  |  |
| Southern New England (SNE) |  |  |  |  |  |  |  |  |  |
| 1980-1989 | 0.101 | 0.095 | 0.218 | 0.135 | 0.126 | 0.325 | 1.00 |  |  |
| 1990-1999 | 0.062 | 0.102 | 0.310 |  | 0.132 | 0.395 | 1.00 |  |  |
| 2000-2005 | 0.700 | 0.300 |  |  |  |  | 1.00 |  |  |
| Georges Bank (GBK) |  |  |  |  |  |  |  |  |  |
| 1986-1992 | 0.067 | 0.138 | 0.121 | 0.145 | 0.058 | 0.471 | 1.00 |  |  |
| 1997-2002 | 0.069 | 0.052 | 0.080 | 0.050 | 0.034 | 0.715 | 1.00 |  |  |
| All years | 0.068 | 0.082 | 0.094 | 0.083 | 0.042 | 0.630 | 1.00 |  |  |

Table C27. Proportions of total 2005 stock biomass at a range of survey density levels, by region.

| Survey catch rate level (kg/tow) |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Region | 0 to 4 | 5 to 9 | 10 to 14 | 15 to 19 | 20 to 24 | 25+ | Total |
| Total 2005 biomass (mt meats) |  |  |  |  |  |  |  |
| Southern Virginia (SVA) | 36 | 0 | 0 | 0 | 0 | 0 | 36 |
| Delmarva (DMV) | 29 | 21 | 18 | 14 | 0 | 10 | 92 |
| New Jersey (NJ) | 61 | 59 | 99 | 39 | 16 | 91 | 365 |
| Long Island (LI) | 18 | 21 | 29 | 99 | 0 | 0 | 167 |
| Southern New England (SNE) | 28 | 12 | 0 | 0 | 0 | 0 | 40 |
| Georges Bank (GBK) | 33 | 25 | 38 | 24 | 16 | 336 | 471 |
| Total | 205 | 138 | 184 | 176 | 32 | 437 | 1,170 |
| Total 2005 biomass (bushels) |  |  |  |  |  |  |  |
| Southern Virginia (SVA) | 4,678 | 0 | 0 | 0 | 0 | 0 | 4,678 |
| Delmarva (DMV) | 3,713 | 2,786 | 2,350 | 1,777 | 0 | 1,325 | 11,951 |
| New Jersey (NJ) | 7,959 | 7,598 | 12,843 | 5,085 | 2,058 | 11,755 | 47,299 |
| Long Island (LI) | 2,354 | 2,721 | 3,731 | 12,823 | 0 | 0 | 21,628 |
| Southern New England (SNE) | 3,615 | 1,548 | 0 | 0 | 0 | 0 | 5,162 |
| Georges Bank (GBK) | 4,218 | 3,188 | 4,900 | 3,079 | 2,048 | 43,632 | 61,065 |
| Total | 26,537 | 17,841 <br> Percen | $\begin{gathered} 23,823 \\ \text { of total } 20 \end{gathered}$ | $22,764$ <br> 5 biomass | 4,106 | 56,712 | 151,783 |
| Southern Virginia (SVA) | 3.082\% | 0.000\% | 0.000\% | 0.000\% | 0.000\% | 0.000\% | 3.082\% |
| Delmarva (DMV) | 2.45\% | 1.84\% | 1.55\% | 1.17\% | 0.00\% | 0.87\% | 7.87\% |
| New Jersey (NJ) | 5.24\% | 5.01\% | 8.46\% | 3.35\% | 1.36\% | 7.74\% | 31.16\% |
| Long Island (LI) | 1.55\% | 1.79\% | 2.46\% | 8.45\% | 0.00\% | 0.00\% | 14.25\% |
| Southern New England (SNE) | 2.38\% | 1.02\% | 0.00\% | 0.00\% | 0.00\% | 0.00\% | 3.40\% |
| Georges Bank (GBK) | 2.78\% | 2.10\% | 3.23\% | 2.03\% | 1.35\% | 28.75\% | 40.23\% |
| Total | 17.48\% | 11.75\% | 15.70\% | 15.00\% | 2.70\% | 37.36\% | 100.00\% |

Table C28. Example projection results for surfclam showing projected average biomass, and fishing mortality during 2006-2015 under three possible scenarios: i) constant landings at the minimum quota; ii) status-quo landings (i.e. mean landings during 2003 to 2005); iii) constant landings at the maximum quota; and iv) constant fishing mortality at the $F_{M S Y}$ proxy $=M=0.15 \mathrm{y}^{-1}$. CVs measure variability between simulation runs in the projection analysis for a scenario. CVs were similar for each scenario in the same year and the CVs shown in the table are averages for simplicity in presentation.

| Year | $\begin{gathered} \text { Landings }=\mathrm{min} \\ \text { quota } \\ =1.85 \text { million bu } \end{gathered}$ | $\begin{aligned} & \begin{array}{c} \text { Status quo } \\ \text { landings } \end{array} \\ &= \text { mean } 2003- \\ & 2005 \\ &= 3.042 \text { million } \\ & \text { bu } \end{aligned}$ | $\begin{aligned} & \text { Landings = max } \\ & \text { quota } \\ & =3.4 \text { million bu } \end{aligned}$ | $\begin{gathered} F=F_{M S Y} \\ =M= \\ 0.15 \end{gathered}$ | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Catch (landings + 12\%, 1000 mt ) |  |  |  |  |  |
| All | 16.0 | 49.7 | 49.7 | variable | NA |
| Biomass (1000 mt) |  |  |  |  |  |
| 2005 | 1,198 | 1,198 | 1,198 | 1,198 | 251\% |
| 2006 | 1,093 | 1,093 | 1,093 | 1,093 | 275\% |
| 2007 | 1,010 | 1,001 | 998 | 889 | 322\% |
| 2008 | 944 | 925 | 920 | 739 | 417\% |
| 2009 | 892 | 866 | 858 | 632 | 560\% |
| 2010 | 856 | 823 | 813 | 559 | 744\% |
| 2011 | 832 | 793 | 781 | 512 | 944\% |
| 2012 | 820 | 776 | 762 | 485 | 1150\% |
| 2013 | 819 | 769 | 754 | 472 | 1350\% |
| 2014 | 826 | 772 | 755 | 470 | 1532\% |
| 2015 | 839 | 781 | 763 | 474 | 1679\% |
| Fishing mortality (annual rate) |  |  |  |  |  |
| 2005 | 0.0188 | 0.0188 | 0.0188 | 0.0188 | 255\% |
| 2006 | 0.0156 | 0.0258 | 0.0288 | 0.1500 | 279\% |
| 2007 | 0.0169 | 0.0282 | 0.0317 | 0.1500 | 327\% |
| 2008 | 0.0181 | 0.0306 | 0.0345 | 0.1500 | 412\% |
| 2009 | 0.0193 | 0.0329 | 0.0372 | 0.1500 | 531\% |
| 2010 | 0.0202 | 0.0349 | 0.0396 | 0.1500 | 676\% |
| 2011 | 0.0210 | 0.0367 | 0.0418 | 0.1500 | 836\% |
| 2012 | 0.0216 | 0.0381 | 0.0435 | 0.1500 | 1009\% |
| 2013 | 0.0220 | 0.0392 | 0.0449 | 0.1500 | 1187\% |
| 2014 | 0.0222 | 0.0399 | 0.0458 | 0.1500 | 1369\% |
| 2015 | 0.0223 | 0.0403 | 0.0465 | 0.1500 | 1551\% |

## SURFCLAM FIGURES



Figure C1. Surfclam stock assessment regions and NEFSC clam survey strata. Northern and southern New Jersey is combined to form the larger New Jersey (NJ) assessment region.


Figure C2. Size at age data for surfclam in DMV and NJ from NEFSC clam surveys during 1982-1992 and 1994-2005. The dark line shows average size at age in all years.


Figure C3. Atlantic surfclam landings and EEZ surfclam quotas (all converted to mt meats).


Figure C4. Surfclam landings from the US EEZ during 1979-2005 by stock assessment region.


Figure C5. Total fishing effort (hours fished, all trips and all vessels) in the US EEZ during 1991-2005 by stock assessment region.


Figure C6. Real and nominal exvessel prices (US\$ per bushel) for surfclam landed (EEZ and state waters) during 1982-2005. Real prices use 1980-1982 as the base year. The current "break-even" price (to meet variable costs) is about 3-4 \$ bu-1 (nominal, in 2005 dollars) and shown for comparison.


Figure C7. Nominal and standardized LPUE for surfclam in the EEZ, by region. Regions with similar trends are plotted together.


Figure C8. Spatial distribution of surfclam landings (annual means, 1 kilobushel $=1000$ bu y-1) during 1981-1990, 1991-1995, 1996-2000 and 2001-2005 based on logbook data and ten-minute squares. Categories correspond approximately with quartiles during 1981-1990.


Figure C9. Spatial distribution of surfclam fishing effort (annual means, $\mathrm{hy}^{-1}$ ) during 1981-1990, 1991-1995, 1996-2000 and 2001-2005 based on logbook data and ten-minute squares. Categories correspond approximately with quartiles during 1981-1990.


Figure C10. Spatial distribution of surfclam LPUE (annual means, bu h ${ }^{-1} \mathrm{y}^{-1}$ ) during 1981-1990, 1991-1995, 1996-2000 and 2001-2005 based on logbook data and ten-minute squares. Categories correspond approximately with quartiles during 1981-1990.


Year
Figure C11. Annual surfclam landings (1000 bushels per year) for important ten-minute squares during 1980-2005 based on logbook data. The smooth dark line is a spline intended to show general trends.


Figure C12. Annual surfclam fishing effort (hours of fishing per year) for important tenminute squares during 1980-2005 based on logbook data. The smooth dark line is a spline intended to show general trends.


Figure C13. Annual surfclam landings per unit of fishing effort (LPUE, mean $\mathrm{h}^{-1}$ ) for important ten-minute squares during 1980-2005 based on logbook data. The smooth dark line is a spline intended to show general trends.


Figure C14. Commercial length composition data for surfclam caught in the DMV area, based on port samples. The dashed vertical line is at 120 mm SL.

## DMV



Figure C14. (continued)

## NJ



Figure C14. (continued)


Figure C15. Commercial length composition data for surfclam caught in the NJ area, based on port samples. The dashed vertical line is at 120 mm SL.


Figure C16. Commercial length composition data for surfclam caught in the LI area, based on port samples. The dashed vertical line is at 120 mm SL.

## SNE



Figure C17. Commercial length composition data for surfclam caught in the SNE area, based on port samples. The dashed vertical line is at 120 mm SL.


Figure C18. Commercial age composition data for surfclam in the NJ and DMV areas during 2005. There is uncertainty about timing of ring formation. Assuming rings form during the fall after the NEFSC clam survey, dark circles identify the 1992 ( 14 rings in 2005) year class and dark triangles identify the 1999 year class ( 7 rings in 2005).


Figure C19. Locations of 2002 survey stations with poor dredge performance that would not have been excluded from trend and swept area trend analyses based on haul or gear damage codes, with station numbers. Codes 1 and 2 (dark squares and open circles) are random stations. Stations in close-up 1 are all from a depletion experiment.


Figure C19 (continued)


Figure C19 (continued)


Figure C20. Locations of 2005 survey stations with poor dredge performance that would not have been excluded from trend and swept area trend analyses based on haul or gear damage codes, with station numbers. Codes 1 and 2 (dark squares and open circles) are random stations. Stations in close-up 2 are all from a depletion experiment.


Figure C20 (continued)


Figure C20 (continued)

> SVA


Predicted

SVA


SVA


Cruise

SVA


SVA


SVA


Figure C21. Residuals and diagnostics for negative binomial GLM model used to impute missing survey data for surfclam in SVA.

SNE


Figure C22. Residuals and diagnostics for negative binomial GLM model used to impute missing survey data for surfclam in SNE.

GBK


GBK


GBK


GBK


GBK


GBK


Figure C23. Residuals and diagnostics for negative binomial GLM model used to impute missing survey data for surfclam in GBK.


Figure C24. Location of successful random survey stations during the 2005 NEFSC clam survey with catches for small recruit surfclam $80-119 \mathrm{~mm}$ SL. Catches are numbers per tow, standardized by doppler distance with no borrowing.


Figure C25. Location of successful random survey stations during the 2005 NEFSC clam survey with catches for large fishable surfclam $120+\mathrm{mm} \mathrm{SL}$. Catches are numbers per tow, standardized by doppler distance with no borrowing.


Figure C26. Trends in abundance (mean $n$ tow $^{-1}$ ) for small recruit surfclam (50-119 mm) and trends in biomass (mean $\mathrm{kg} \mathrm{tow}^{-1}$ ) for large fishable ( $120+\mathrm{mm}$ ) surfclam based on NEFSC clam surveys, by region.


Figure C27. Trends in abundance indices (mean n tow ${ }^{-1}$ ) for small recruit surfclam (50119 mm SL ) in NEFSC clam surveys, with $80 \%$ confidence intervals assuming lognormal measurement errors and arithmetic CVs for stratified random sampling based on Students- $t$ distribution with the number of tows as degrees of freedom.


Figure C28. Trends in biomass indices (mean $\mathrm{kg} \mathrm{tow}^{-1}$ ) for large fishable surfclam (120+ mm SL) in NEFSC clam surveys, with $80 \%$ confidence intervals assuming lognormal measurement errors and arithmetic CVs for stratified random sampling based on Students- $t$ distribution with the number of tows as degrees of freedom. Different symbols show effects of omitting tows with poor gear performance during 2005.


Figure C29. Trends in biomass indices for large fully large surfclam (mean kg tow ${ }^{-1}$, $120+\mathrm{mm}$ SL) in NEFSC clam surveys and standardized LPUE in the commercial fishery.

Survey Length Data for Surfclam in DMV


Figure C30. Survey length composition data for surfclam in the DMV region.

Survey Length Data for Surfclam in NJ


Figure C31. Survey length composition data for surfclam in the NJ region.

## Survey Length Data for Surfclam in LI



Figure C32. Survey length composition data for surfclam in the LI region.

Survey Length Data for Surfclam in SNE


Figure C33. Survey length composition data for surfclam in the SNE region.

## Survey Length Data for Surfclam in GBK



Figure C34a. Survey length composition data for surfclam in the GBK region.


Figure 34b. NEFSC clam survey age composition data for surfclam in the NJ and DMV areas during 1997-2005. There is uncertainty about the timing of annual ring formation. Assuming rings form during the fall after the NEFSC clam survey, dark circles identify the 1992 year class ( 14 rings in 2005) and dark triangles identify the 1999 year class ( 7 rings in 2005).


Figure C35. Relationships between depletion study variables and sediment grain size based on depletion studies during 1999, 2002 and 2005. Sediment data were not collected during 1997 and 2004.


Figure C36. Relationships between depletion study variables based on all depletion studies during 1997-2005.


Joint 50\%, 90\%, 95\% and 99\% profile confidence intervals

Figure C37. Likelihood profile analysis for efficiency and density estimates from the Patch model for the SC1999-7 surfclam depletion experiment. The joint $50 \%$ confidence interval for efficiency and density is the area within the outermost contour. The joint $99 \%$ confidence interval is the area inside the innermost contour lines. Contour lines for the joint $90 \%$ and $95 \%$ confidence intervals lie between.


Figure C38. Efficiency corrected swept-area biomass estimates for surfclam $120+\mathrm{mm}$ SL, including estimates from the 2004 cooperative survey and NEFSC clam surveys during 1997, 1999, 2002 and 2005.


Figure C39. Von Bertlanffy curves for size (meat weight) at age of surfclam during 1982-1992 and 1994-2005 in the NNJ and DMV regions, based on NEFSC clam survey data for 1982-2005.


Figure C40. Model diagnostics for the KLAMZ model for the entire stock of surfclam. Trends in nominal LPUE and efficiency corrected swept area biomass are shown with predicted trends from the model for comparison, but trends in these indices did not affect model estimates. Survey scaling parameter ( $Q$ ) estimates are shown below plots for each set of data.




Figure C41. Population dynamics estimates from the KLAMZ model for the entire surfclam stock.


Figure C42. Fishable biomass estimates with $80 \%$ empirical confidence intervals from bootstrapping for the entire surfclam stock. Nominal LPUE from logbooks (total reported landings / total reported hours fished, all vessels and all trips) for the entire fishery is shown for comparison. LPUE data were not used in estimating biomass.


Figure C43. Fishing mortality estimates for the entire surfclam stock with $80 \%$ confidence intervals from bootstrapping.


Figure C44. Recruitment for the entire surfclam stock with $80 \%$ empirical confidence intervals from bootstrapping.


Q $=0.9961$
Figure C45. Model diagnostics for the KLAMZ model surfclam in DMV. Trends in nominal LPUE and efficiency corrected swept area biomass are shown with predicted trends from the model for comparison, but trends in these indices did not affect model estimates. Survey scaling parameter $(Q)$ estimates are shown below plots for each set of data.


Figure C46. Model diagnostics for the KLAMZ model surfclam in NJ. Trends in nominal LPUE and efficiency corrected swept area biomass are shown with predicted trends from the model for comparison, but trends in these indices did not affect model estimates. Survey scaling parameter $(Q)$ estimates are shown below plots for each set of data.


Figure C47. Population dynamics estimates from the KLAMZ model for surfclam in DMV region.


Figure C48. Population dynamics estimates from the KLAMZ model for surfclam in NJ region.


Figure C49. Surfclam biomass for the whole stock prorated into regional components based on rescaled regional survey trend data.


Figure C50. Proportions of total surfclam biomass by region during 1982 and 2005.


Figure C51. Fishable surfclam biomass during 2005 with probability distributions to characterize uncertainty. The long dash vertical line on the left is the biomass threshold. The short dash vertical line on the right is the biomass target.


Fishing Mortality
Figure C52. Fishing mortality for surfclam during 2005 with probability distributions to characterize uncertainty. The dash vertical line on the right is the fishing mortality threshold reference point.


Figure C53. Average biomass and fishing mortality during 2005-2015 based on stochastic projection analysis under four assumed scenarios for constant landings of constant fishing mortality. CVs are for the variability between simulation runs in the same scenario.

## APPENDIX C1. Invertebrate Subcommittee

Persons who attended Invertebrate Subcommittee meetings (September 25-26, October 16-17, and October 30-November 1, 2006) and contributed to this report are:
T. Alspach (Sea Watch International, Ltd.)
M. Bell (Invited external participant, Lowestoft, Suffolk, UK)
A. Chute (NEFSC)
H. Dobby (Invited external participant, FRS Marine Laboratory, Aberdeen, Scotland)
C. Heaton (Mid-Atlantic Fishery Management Council, MAFMC)
J. Heifitz (Invited external participant, NMFS, AKFSC)
T. Hoff (MAFMC)
L. Jacobson (Northeast Fisheries Science Center, NEFSC) - assessment lead
C. Pickett (NEFSC)
E. Powell (Haskin Shellfish Laboratory, Rutgers University)
D. Wallace (Wallace \& Associates, Inc.)
J. Womack (Wallace \& Associates, Inc.)
R. Mayo (NEFSC) - Subcommittee Chair
J. Weinberg (NEFSC)

## APPENDIX C2. Analyses of tows with poor dredge performance in the 2002 NEFSC clam survey.

The review of the Survey Sensor Pack (SSP) data from the 2005 clam survey showed a significant number survey tows with anomalies that would likely affect the performance of the survey dredge. These anomalies in 2005 were mostly with problems in the manifold pressure in addition to several tows that had erratic towing angles. The number of 2005 survey tows deemed to have poor dredge performance by the proposed evaluation criteria (see Appendix C3) was approximately $8 \%$ of the total number of survey stations reviewed.

To see if the anomalies present in the 2005 survey were a unique situation or a continuation of an inherit inconsistency with the NMFS survey dredge, a review of the SSP data from the 2002 clam survey was undertaken. Because of time constraints and the limited number of survey station data plots available, this review was limited to a visual inspection of the data plots. The visual criteria used to judge a tow to have either "good" or "poor" dredge performance is the same as was used to perform a preliminary grading of the 2005 SSP data. In general the manifold pressure and fore/aft tilt angle plots were the parameters reviewed for significant deviations from normal values. Sample plots are shown below.

Manifold Pressure Visual Inspection Criteria Samples


Dredge Angle Visual Inspection Criteria Samples


## Summary of Results (for APPENDIX C2.)

The review of 2002 survey SSP data showed that similar anomalies found in 2005 survey were also found in the 2002 survey in addition to a problem with early shutoff of the dredge pump before the completion of the tow. The summary of the anomalies is shown below for both the 2002 and 2005 surveys.

|  | 2002 | 2005 |
| :--- | ---: | ---: |
| Description | Survey | Survey |
| Total \# of DE2 Survey Stations | 556 | 433 |
| Total \# of Stations Tows Reviewed | 213 | 399 |
| Total \# of Stations Labeled Good | 181 | 366 |
| \% of Total Stations Reviewed | $85.0 \%$ | $91.7 \%$ |
| Total \# of Stations Labeled Poor for Any Reason | 32 | 33 |
| \% of Total Stations Reviewed | $15.0 \%$ | $8.3 \%$ |
| Total \# of Stations Labeled for Intake Blockage | 11 | 22 |
| \% of total Stations Reviewed | $5.2 \%$ | $5.5 \%$ |
| Total \# of Stations Labeled Poor for Manifold Blockage | 1 | 10 |
| \% of total Stations Reviewed | $0.5 \%$ | $2.5 \%$ |
| Total \# of Stations Labeled Poor for Dredge Angle | 0 | 2 |
| \% of total Stations Reviewed | $0.0 \%$ | $0.5 \%$ |
| Total \# of Stations Labeled Poor for Early Pump Shutoff | 20 | 0 |
| \% of total Stations Reviewed | $9.4 \%$ | $0.0 \%$ |

In general the results show that the NMFS survey dredge is likely to experience a significant number of poor tows during any given survey from a number of possible reasons that affect either manifold pressure or fore and aft dredge running angle. From survey to survey, however, the predominate reason for a poor tow can vary. For example, the 2005 survey had a high number of poor tows due to manifold blockage compared to the 2002 survey. This was from an intake screen failure in 2005 on the dredge pump which allowed small stones to lodge in the manifold nozzles. In 2002, the predominate problem was the dredge pump being shutoff early which did not happen in 2005.

The list of poor tows for the 2002 tows from the tows reviewed is below. As pointed out elsewhere, many of the tows with poor gear performance would have been omitted from use in the stock assessment due to standard haul or gear condition criteria or were nonrandom tows used for special purposes.


APPENDIX C3. Comparison of surfclam and ocean quahog catches in tows with poor dredge performance during the 2002 and 2005 NEFSC clam surveys and 2002 cooperative survey tows (prepared by John Womack, Wallace and Associates, Ltd.)

```
2002 Stock Assessment Survey Results
    Total # of DE2 Survey Stations = 556
    Total # of Stations Reviewed = 213
    Total # of Stations Labeled Good = 181
        % of total Stations Reviewed = 85.0%
    Total # of Stations Labeled Poor = 32 (Any Reason, Visual Inspection of Plots)
            % of total Stations Reviewed = 15.0%
            Total # of Stations Labeled Poor = 11 (Intake Blockage)
            % of total Stations Reviewed = 5.2%
    Total # of Stations Labeled Poor = 1 (Manifold Blockage)
            % of total Stations Reviewed = 0.5%
    Total # of Stations Labeled Poor = 0 (Dredge Angle)
            % of total Stations Reviewed = 0.0%
    Total # of Stations Labeled Poor = 20 (Early Pump Shutoff)
            % of total Stations Reviewed = 9.4%
    Average # of Surfclam per Good Tow - 24.2
    Average # of Surfclam per Poor Tow - 28.5
    Average # of Quahogs per Good Tow - 69.3
    Average # of Quahogs per Poor Tow - 64.3
```

Poor Stations, Intake Blockage - 4, 52, 76, 218, 250, 386, 394, 458
Poor Stations, Manifold Blockage - 382
Poor Stations, Early Pump Shutoff - 32, 42, 44, 45, 82, 90, 101, 103, 104, 106, 111, 118,
$125,137,140,141,254,278,360,368,496,498,506$

## Comments on Review of Pump Manifold Pressure (See Figure 1)

For initial portion of the cruise, station 0-230, the pump voltage was about 388 VAC. During this part of the cruise the pump manifold pressure followed a similar value and decrease in pressure pattern, i.e. normal wear, as was seen in the 2005 survey.
After about station 230 the pump voltage suddenly rises to about 400 VAC till about station 300. The pump manifold pressure also showed a small increase over the first portion of the cruise from about 34 PSI to about 35-36 PSI.
After station 300 this rise can not be tracked as voltage data is lost from around station 300 till around station 400.
At around station 400 the pump voltage suddenly rises to about 417 VAC This voltage rise lasted till the survey end. The pump manifold pressure also showed a significant increase over the first portion of the cruise from about 34 PSI to about 40 PSI.
The total voltage rise from cruise start to end is about $7.5 \%$. The power the pump was drawing also showed a similar increase from 11.87 to 12.79 .


Figure 1 (Appendix C3)

| Station Number | All | $0-230$ | $231-409$ | $410-546$ |
| :--- | :---: | :---: | :---: | :---: |
| Avg \# of Surfclam per Good Tow - | 24.2 | 30.1 | 15.8 | 12.3 |
| Avg \# of Surfclam per Poor Tow - | 28.5 | 33.6 | 30.1 | 0.0 |
| Avg \# of Quahogs per Good Tow - | 69.3 | 34.3 | 45.0 | 232.5 |
| Avg \# of Quahogs per Poor Tow - | 64.3 | 4.1 | 14.3 | 465.8 |
| Total \# Of Good Tows | 181 | 114 | 37 | 30 |
| Total \# Of Poor Tows | 32 | 20 | 8 | 4 |

For all stations and 0-230 and 231-409 groups, the NMFS dredge appears to fish surfclam better during a poor tow, generally which was a loss of manifold pressure, than a good tow. The last group, 410-546, did not show this pattern but this could be due to the fact that it appears to be primarily composed of quahog habitat stations.
The manifold may have seen some blockage in the stations around 375 to 400 as the pressure is higher but the amps draw has dropped.

For all groups as the manifold pressure rises, the surfclam catch per tow falls significantly, over $50 \%$. See Figure 2. Caveat, limited number of stations in last two groups and last group was likely in quahog habitat.
For all groups as the manifold pressure rises, the quahog catch per tow increases significantly. See Figure 3. Caveat, limited number of stations in last two groups and last group was likely in quahog habitat.



Figure 2 and Figure 3 (Appendix C3)

2005 Stock Assessment Survey Results
Total \# of DE2 Survey Stations $=433 \quad 82=556$
Total \# of Stations Reviewed $=399 \quad 82=213$
Total \# of Stations Labeled Good $=366 \quad 82=181$
$\%$ of total Stations Reviewed $=91.7 \% \quad 82=85.0 \%$
Total \# of Stations Labeled Poor $=33$ (Any Reason) $82=32$
$\%$ of total Stations Reviewed $=8.3 \% \quad 82=15.0 \%$
Total \# of Stations Labeled Poor = 22 (Intake Blockage) $82=11$

$$
\% \text { of total Stations Reviewed }=5.5 \% \quad 82=5.2 \%
$$

Total \# of Stations Labeled Poor $=10$ (Manifold Blockage)
$\%$ of total Stations Reviewed $=2.5 \%$
$82=1$
Total \# of Stations Labeled Poor = 2 (Dredge Angle)
$82=0.5 \%$
$\%$ of total Stations Reviewed $=0.5 \%$
$82=0$
Total \# of Stations Labeled Poor $=0$ (Early Pump Shutoff)
$\%$ of total Stations Reviewed $=0.0 \%$
$82=0.0 \%$
$82=20$
Average \# of Surfclam per Good Tow - 18.20
$82=9.4 \%$
Average \# of Surfclam per Poor Tow - 28.68
$82=24.2$
Average \# of Quahogs per Good Tow - 42.91
$82=28.5$
Average \# of Quahogs per Poor Tow - 1.19
$82=69.3$
$82=64.3$

## General Comments on 2002/2005 Survey Tows

2002 Speed fairly smooth and consistent as opposed to 2205 survey which had more variation and steeper spikes and dips.
2002 Dredge angle relatively smooth even when pump intake was blocked or pump was shutoff early. (i.e. may have continued to fish effectively)
2002 Survey had significant changes in the dredge pump voltage and thus a significant increase in manifold pressure during the survey cruise.
NMFS Dredge fished surfclam better on poor tows then good tows for both 2002 and 2005 surveys.

F/V Lisa Kim Poor Tows
Station 12 - Dredge angle high. Odd as angle is about 5 degrees above normal and basically smooth throughout the tow.
Station 72 - Dredge angle very erratic varying from 0 to 25 degrees.
F/V Lisa Kim Tows with Blips, Not severe enough for a poor tow.
Station 95 - Very brief bump up in dredge angle.
F/V Jersey Girl had no Poor tows or tows with blips.

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[^0]:    ${ }^{1}$ This assessment was prepared by the Invertebrate Subcommittee. Contributing members are listed in INTRODUCTION TO SAW-44 ASSESSMENT REPORT.

[^1]:    ${ }^{2}$ See Cargnelli et al. (1999) for additional information.

[^2]:    3 "Efficiency" of a clam dredge is the probability that an ocean quahog in the path of the dredge will be caught. Efficiency of capture may differ between quahog of difference size and the definition used here applies to quahog large enough to be fully available to the sampling gear. Efficiency estimates for the survey dredge are used with a variety of other information to estimate the "catchability" coefficients for NEFSC clam surveys that relate survey catches to stock abundance and biomass.

[^3]:    ${ }^{4}$ Standard survey procedures omit stations with database Station Type-Haul Type-Gear Condition (SHG) codes greater than 136.

[^4]:    5 Steps 1-2 were done in SAS (note that interpolation precedes smoothing).
    proc expand data=sdata1 out=sdata2 to=second;
    by station;
    ID TowTime;
    convert TiltY=SmoothAngle / transform=(cmovave 7);
    convert GPS1_SOG=SmoothSOG / transform=(cmovave 7);
    run;

[^5]:    ${ }^{6}$ Contact Alan Seaver (Alan.Seaver@noaa.gov), Northeast Fisheries Science Center, Woods Hole, MA, USA for information and access to the Stock Assessment Toolbox.

[^6]:    Parametric statistics.
    ${ }^{2}$ Bootstrap statistics (15,000 iterations).

[^7]:    ${ }^{7}$ Prepared by John Womack, Wallace and Associates, Ltd.

[^8]:    ${ }^{8}$ Prepared by John Womack, Wallace and Associates, Inc.

[^9]:    ${ }^{9}$ In applications, assumptions about knife-edge selectivity can be relaxed by assuming the model tracks "fishable", rather that total, biomass (NEFSC 2000a; 2000b). An analogous approach assigns pseudo-ages based on recruitment to the fishery so that new recruits in the model are all pseudo-age $k$. The synthetic cohort of fish pseudo-age $k$ may consist of more than one biological cohort. The first pseudo-age ( $k$ ) can be the predicted age at first, $50 \%$ or full recruitment based a von Bertalanffy curve and size composition data (Butler et al. 2002). The "incomplete recruitment" approach (Deriso 1980) calculates recruitment to the model in each year $R_{t}$ as the weighted sum of contributions from two or more biological cohorts (yearclasses) from spawning during successive years (i.e. $R_{t}=\sum_{a=1}^{k} r_{a} \Pi_{t-a}$ where k is the age at full recruitment to the fishery, $r_{a}$ is the contribution of fish age $k-a$ to the fishable stock, and $\Pi_{t-a}$ is the number or biomass of fish age $k$ - $a$ during year $t$ ).
    ${ }^{10}$ In some applications, and more generally, new recruits might be defined as individuals recruiting at the beginning or at any time during the current time step (e.g. NEFSC 1996).

[^10]:    ${ }^{11}$ Otter Research Ltd., Box 2040, Sydney, BC, Canada V8L 3S3 (otter@otter-rsch.com).

[^11]:    ${ }^{12}$ The constraint is implemented by adding $L=\lambda \bar{\sigma}^{2}$ (where $\bar{\sigma}$ is the average deviation) to the objective function, generally with a high weighting factor $(\lambda=1000)$ so that the constraint is binding.

[^12]:    ${ }^{13}$ Another approach to using time dependent natural mortality rates is to treat estimates of predator consumption as discarded catch (see "Predator consumption as discard data"). In addition, estimates of predator abundance can be used in fishing effort calculations (see "Predator data as fishing effort").

[^13]:    ${ }^{14}$ By convention, the instantaneous rates $G_{t}, F_{t}$ and $M_{t}$ are always expressed as numbers $\geq 0$.
    ${ }^{15}$ The traditional catch equation $C_{t}=F_{t}\left(1-e^{-Z_{t}}\right) B_{t} / Z_{t}$ where $Z_{t}=F_{t}+M_{t}$ underestimates catch biomass for a given level of fishing mortality $F_{t}$ and overestimates $F_{t}$ for a given level of catch biomass. The errors can be substantial for fast growing fish, particularly if recent recruitments were strong.

[^14]:    ${ }^{16}$ Unfortunately, concentrated likelihood calculations cannot be used with MCMC and other Bayesian approaches to characterizing posterior distributions. Therefore, in the near future, concentrated NLL calculations will be replaced by calculations for the entire NLL. At present, MCMC calculations in KLAMZ are not useful.

[^15]:    ${ }^{17}$ Quinn and Deriso (1999) describe another approach attributed to a manuscript by C. Walters.
    ${ }^{18}$ Normally, $n_{G} \leq 2$.

[^16]:    ${ }^{19}$ Future versions of the KLAMZ model will allow equilibrium initial biomass to be calculated based on other recruitment values and for a user-specified level of F (Butler et al. 2003).

[^17]:    ${ }^{20}$ Abundance indices with statistical distributions other than log normal may be used as well, but are not currently programmed in the KLAMZ model. For example, Butler et al. (2003) used abundance indices with binomial distributions in a delay-difference model for cowcod rockfish. The next version of KLAMZ will accommodate presence-absence data with binomial distributions.
    ${ }^{21}$ The relationship between data and fish populations is affected by factors (process errors) that are not accounted for in CV calculations.

[^18]:    ${ }^{22}$ It may be important to project biomass forward if an absolute estimate of biomass is available (e.g. from a hydroacoustic or daily egg production survey), if fishing mortality rates or high or if the timing of the survey varies considerably from year to year.

[^19]:    ${ }^{23}$ If $x$ has a standardized beta distribution with parameters $a$ and $b$, then the probability of $x$ is $P(x)=\frac{x^{a-1}(1-x)^{b-1}}{\Gamma(a, b)}$.

[^20]:    ${ }^{24}$ Variances in NLL for surplus production-biomass models are a subject of ongoing research. The advantage in assuming normal errors is that negative production values (which occur in many stocks, e.g. Jacobson et al. 2001) are accommodated. In addition, production models can be fit easily by linear regression of $P_{t}$ on $B_{t}$ and $B_{t}{ }^{2}$ with no intercept term. However, variance of production estimate residuals increases with predicted surplus production. Therefore, the current approach to fitting production curves in KLAMZ is not completely satisfactory.

[^21]:    ${ }^{25}$ MCMC calculations are not available in the current version because objective function calculations use concentrated likelihood formulas. However, the C++ version of KLAMZ is programmed in other respects to accommodate Bayesian estimation.

[^22]:    ${ }^{26}$ At present, only Beverton-Holt recruitment calculations are available in SPROJDDF.

[^23]:    ${ }^{27}$ Available with assessment for reviewer's convenience.

[^24]:    ${ }^{28}$ Contact Alan.Seaver@noaa.gov for information about the NMFS Stock Assessment Toolbox.

[^25]:    Table B6.7. Diet composition of Clearnose Skate. All values are expressed as whole numbers rather than percentages. Relmsw = relative mean stomach weight, on average for the size class and time period given. $\mathrm{AR}=$ animal remains, a welldigested, highly unresolved category.

[^26]:    ${ }^{29}$ During the 2005 survey, tows with poor dredge performance occurred at survey stations: $1,2,14,17,20-$ $26,28,29-34,45,48,56,58,67,75,76,108,218,225,262,282,405,411,413,414,417$, and 422-424. Based on a sample from the 2002 survey, tows with poor dredge performance occurred at survey stations: $4,32,42,44,45,52,76,82,90,101,103,105,106,111,118,125,137,140,141,218,250,254,278,360$, $368,382,386,394,458,496,498$, and 506.

[^27]:    ${ }^{30}$ Surfclam appear to recruit to the NEFSC survey dredge by about 120 mm SL. Surfclam recruit to the NEFSC survey dredge at smaller sizes that to commercial dredges because the survey dredge is made with closely spaced bars and a wire mesh liner. Moreover, survey catches are not sorted mechanically on a shaker table to remove trash and undersized objects.

[^28]:    ${ }^{31}$ See Appendix A5 of the ocean quahog assessment (NEFSC 2007) for a complete technical description of the KLAMZ model.

[^29]:    ${ }^{32}$ Available at: http://www.nefsc.noaa.gov/nefsc/publications/crd/crd0316/surfclam.pdf.
    ${ }_{34}^{33}$ Available at: http://www.nefsc.noaa.gov/esb/survey_reports/Clam\%202005/all.pdf.
    ${ }^{34}$ Available at: http://www.nefsc.noaa.gov/nefsc/publications/crd/crd0501/.

[^30]:    * Minimal estimate assuming that all depletion tows were examined

