

Continuation of NLS Discussion Paper 96-33  
Part 2 of 2

This version of the paper was split for web delivery.

**Table 1**  
**Definition of Variables**

<b>Variable</b>	<b>Definition</b>
hinsur	Current Employer Offers Health Insurance
lnwage	Log Of Real Hourly Wage Rate
experien	Years of Employment Experience
expersqr	Years of Employment Experience Squared
grade	Highest Grade Completed
unemrate	State Unemployment Rate
hi_pct	% of Persons in R's State Covered by Private HI
pct_ch	% Change in Coverage Proportion Between Years
lnhosp	Log Real Ave. Daily Hospital Room Charge in R's State
pctchghp	% Change in Hospital Room Charge Between Years
illspell	Away From Work At Least 1 Week For Illness
child	At Least One Child Living At Home
leave	Left Job Between Interviews
sep_inv	Fired, Dismissed, or Layed Off From Job
sep_quit	Quit Job for New Job or for Personal Reasons
age	In Years, as of May 1 in Each Survey Year
male	Equals One If Respondent is Male
married	Equals One If Respondent is Married

Table 2  
Descriptive Statistics

Variable	Mean	Std. Dev.
hinsur	0.804	0.397
lnwage	1.936	0.446
experien	7.696	3.209
grade	13.121	2.172
unemrate	2.920	1.045
hi_pct	0.775	0.059
pct_ch	-0.001	0.038
lnhosp	5.410	0.334
pctchghp	0.093	0.079
illspell	0.032	0.176
child	0.339	0.474
leave	0.284	0.451
sep_inv	0.068	0.252
sep_quit	0.179	0.384
age	26.237	3.229
male	0.579	0.494
married	0.477	0.499

Notes:

1) Observations are pooled across years and weighted.

2)  $N = 24325$

Table 3  
Sample of Respondents Selected from the 1979-1992 NLSY Record

Description of Selection Rule	# of Obs. Deleted	Sample Size
Was in the 1979 NLSY respondent sample.	0	12686
Was a NLSY cross-section respondent.	6575	6111
Interviewed at least 8 times after reaching the age of 21.	1156	4955
Main activity during the week preceding an interview was 'keeping house' no more than once.	1040	3915
In the labor force during the week preceding an interview at least 6 times.	199	3716
Had no missing information in any year for wages, employment experience, health insurance, or job mobility.	57	3659
Did not report being 'disabled' at two or more consecutive interviews.	24	3635
Was employed at all consecutive interviews	5	3630

Final sample contains **3630** respondents and **25174** total observations from 1979-1992.

Table 4  
Health Insurance (HI) and Job Transitions Between Consecutive Observations at Times  $t$   
and  $t + 1$ , By Type of Job Transition

Weighted Cell Percentages  
[Weighted Column Percentages]  
(Unweighted Cell Frequencies)

Period $t+1$	Period $t$		
	Job Does Not Offer HI	Job Offers HI	Total
<b>Health Insurance Status, Same Job Status</b>			
w/o HI, Not At Same Job - Quit	2.39 [12.63] (637)	3.00 [3.70] (757)	5.39 (1394)
w/ HI, Not At Same Job - Quit	3.50 [18.54] (895)	10.31 [12.72] (2489)	13.81 (3384)
w/o HI, Not At Same Job - LFD <sup>1</sup>	1.18 [6.24] (337)	1.35 [1.67] (351)	2.53 (688)
w/ HI, Not At Same Job - LFD	1.34 [7.06] (353)	3.02 [3.73] (772)	4.36 (1125)
w/o HI, At Same Job	7.25 [38.35] (1812)	4.20 [5.17] (1057)	11.45 (2869)
w/ HI, At Same Job	3.25 [17.18] (832)	59.21 [73.02] (14061)	62.46 (14893)
<b>Total</b>	18.91 [100.0] (4866)	81.09 [100.0] (19487)	100.0 (24353)

(1) LFD = Layed off, Fired, or Dismissed

**Table 5a**

**Probability of a Transition From a Period  $t$  Job that Doesn't Offer Health Insurance to a Period  $t + 1$  Job that Does Offer Health Insurance Coverage, by Type of Job Transition and by State Health Insurance Availability**

	<b>Quit Old Job</b>	<b>Involuntary Separation</b>	<b>Stay At Old Job</b>
<b>State HI Coverage Rates:</b>			
<b>Quartiles (% Covered)</b>			
<b>Min (54%)</b>	.616	.486	.372
<b>1st (72%)</b>	.618	.540	.345
<b>2nd (77%)</b>	.618	.555	.337
<b>3rd (81%)</b>	.618	.567	.331
<b>Max (90%)</b>	.619	.594	.318

Note: Predicted probabilities calculated using estimates from table 8a.

**Table 5b**

**Probability of a Transition From a Period  $t$  Job that Doesn't Offer Health Insurance to a Period  $t + 1$  Job that Does Offer Health Insurance Coverage, by Type of Job Transition and by Annual Percentage Change In State Health Insurance Coverage**

	<b>Quit Old Job</b>	<b>Involuntary Separation</b>	<b>Stay At Old Job</b>
<b>% Change In Rate:</b>			
<b>Quartiles (% Change)</b>			
<b>1st (-2.61%)</b>	.557	.431	.225
<b>2nd (.20%)</b>	.564	.461	.246
<b>3rd (2.31%)</b>	.570	.484	.263

Note: Predicted probabilities calculated using estimates from table 8a.

Table 5c

Probability of a Transition From a Period  $t$  Job that Doesn't Offer Health Insurance to a Period  $t + 1$  Job that Does Offer Health Insurance Coverage, by Type of Job Transition and by State Health Care Prices

	Quit Old Job	Involuntary Separation	Stay At Old Job
<b>State Hospital Prices:</b>			
Quartiles (ln Price)			
Min (4.09)	.738	.723	.183
1st (4.98)	.658	.606	.276
2nd (5.26)	.630	.565	.312
3rd (5.53)	.601	.524	.349
Max (6.41)	.508	.394	.477

Note: Predicted probabilities calculated using estimates from table 8a.

Table 5d

Probability of a Transition From a Period  $t$  Job that Doesn't Offer Health Insurance to a Period  $t + 1$  Job that Does Offer Health Insurance Coverage, by Type of Job Transition and by Annual Percentage Change In State Health Care Prices

	Quit Old Job	Involuntary Separation	Stay At Old Job
<b>% Change In Price:</b>			
Quartiles (%)			
1st (5.95%)	.620	.555	.337
2nd (9.78%)	.617	.551	.336
3rd (13.40%)	.615	.548	.335

Note: Predicted probabilities calculated using estimates from table 8a.

Table 6a

Probability of a Transition From a Period  $t$  Job that Offers Health Insurance to a Period  $t + 1$  Job that Does Not Offer Health Insurance Coverage, by Type of Job Transition and by State Health Insurance Availability

	Quit Old Job	Involuntary Separation	Stay At Old Job
<b>State HI Coverage Rates:</b>			
Quartiles (% Covered)			
Min (54%)	.344	.289	.113
1st (72%)	.264	.314	.078
2nd (77%)	.245	.321	.071
3rd (81%)	.230	.326	.065
Max (90%)	.197	.340	.053

Note: Predicted probabilities calculated using estimates from table 8b.

Table 6b

Probability of a Transition From a Period  $t$  Job that Offers Health Insurance to a Period  $t + 1$  Job that Does Not Offer Health Insurance Coverage, by Type of Job Transition and by Annual Percentage Change In State Health Insurance Coverage

	Quit Old Job	Involuntary Separation	Stay At Old Job
<b>% Change In Rate:</b>			
Quartiles (% Change)			
1st (-2.61%)	.245	.329	.074
2nd (.20%)	.246	.321	.069
3rd (2.31%)	.247	.315	.066

Note: Predicted probabilities calculated using estimates from table 8b.



Table 6c

Probability of a Transition From a Period  $t$  Job that Offers Health Insurance to a Period  $t + 1$  Job that Does Not Offer Health Insurance Coverage, by Type of Job Transition and by State Health Care Prices

	Quit Old Job	Involuntary Separation	Stay At Old Job
<b>State Hospital Prices:</b>			
Quartiles (ln Price)			
Min (4.09)	.175	.325	.012
1st (4.98)	.223	.325	.038
2nd (5.26)	.240	.325	.053
3rd (5.53)	.257	.325	.074
Max (6.41)	.317	.325	.198

Note: Predicted probabilities calculated using estimates from table 8b.

Table 6d

Probability of a Transition From a Period  $t$  Job that Offers Health Insurance to a Period  $t + 1$  Job that Does Not Offer Health Insurance Coverage, by Type of Job Transition and by Annual Percentage Change In State Health Care Prices

	Quit Old Job	Involuntary Separation	Stay At Old Job
<b>% Change In Price:</b>			
Quartiles (%)			
1st (5.95%)	.243	.319	.067
2nd (9.78%)	.246	.321	.070
3rd (13.40%)	.249	.323	.073

Note: Predicted probabilities calculated using estimates from table 8b.

Table 7  
Determinants of Real Log Wages

	Coefficient	t-Statistic
experien	.056	17.90
expersqr	-.001	-6.06
grade	.061	54.80
unemrate	-.032	-12.96
constant	.840	39.89

Notes:

(1)  $R^2 = .413$ ,  $N = 29015$

(2) Dependent variable is natural logarithm of the hourly wage rate in 1983 dollars.

Table 8a  
 Multinomial Logit Estimates of Health Insurance Coverage and Job Mobility Hazard  
 Function - Individuals Observed At Time  $t$  At Jobs Not Offering Health Insurance Coverage<sup>1</sup>

Transition	Independent Variable	Coefficient	z-Statistic
[ $H(t + 1) = 0, T(t) = 1$ ]	experien	-.152	-2.21
	expersqr	.001	0.30
	grade	-.046	-2.00
	hi_pct	-2.284	-2.59
	pct_ch	.144	0.12
	lnhosp	.572	3.30
	pctchghp	-.822	-1.30
	unemrate	-.146	-2.99
	illspell	.662	3.00
	lnwage	-.453	-3.94
	constant	.305	0.27
[ $H(t + 1) = 1, T(t) = 1$ ]	experien	-.112	-1.84
	expersqr	.003	0.67
	grade	.089	4.35
	hi_pct	-2.234	-2.89
	pct_ch	.004	0.00
	lnhosp	.134	0.86
	pctchghp	-1.093	-1.97
	unemrate	-.158	-3.63
	illspell	.561	2.70
	lnwage	-.448	-4.49
	constant	1.047	1.07
[ $H(t + 1) = 0, T(t) = 2$ ]	experien	-.298	-3.50
	expersqr	.006	0.95
	grade	-.076	-2.60
	hi_pct	-3.993	-3.55
	pct_ch	-.840	-0.56
	lnhosp	.860	4.00
	pctchghp	-.405	-0.50
	unemrate	-.016	-0.28
	illspell	.632	2.28
	lnwage	.051	0.35
	constant	-.786	-0.56

Table 8a (continued)  
 Multinomial Logit Estimates of Health Insurance Coverage and Job Mobility Hazard  
 Function – Individuals Observed At Time  $t$  At Jobs Not Offering Health Insurance Coverage<sup>1</sup>

Transition	Independent Variable	Coefficient	z-Statistic
[ $H(t + 1) = 1, T(t) = 2$ ]	experien	-.098	-1.14
	expersqr	-.000	-0.07
	grade	.035	1.24
	hi_pct	-2.700	-2.49
	pct_ch	.982	0.65
	lnhosp	.253	1.17
	pctchgph	-.742	-0.96
	unemrate	-.067	-1.13
	illspell	.669	2.48
	lnwage	.075	0.54
	constant	-.570	-0.42
[ $H(t + 1) = 1, T(t) = 3$ ]	experien	-.142	-2.36
	expersqr	.011	2.95
	grade	.093	4.43
	hi_pct	-.572	-0.74
	pct_ch	3.191	2.57
	lnhosp	.593	3.69
	pctchgph	-.095	-0.18
	unemrate	.024	0.55
	illspell	-.280	-1.01
	lnwage	.060	0.60
	constant	-4.546	-4.42

Notes:

(1) At time of initial observation (time  $t$ ) each individual is working at a job which does not offer health insurance coverage.

(2) In column 1  $H(t + 1)$  represents health insurance status in the year following the initial observation, and  $T(t)$  represents the job transition that took place between the two observations.  $T(t) = 1$  represents a job quit;  $T(t) = 2$  represents an involuntary job separation; and  $T(t) = 3$  represents no job change.

(3) For identification the category [ $H(t + 1) = 0, T(t) = 3$ ] is excluded.

(4) Pseudo  $R^2 = .027$ ,  $N = 4623$

Table 8b  
 Multinomial Logit Estimates of Health Insurance Coverage and Job Mobility Hazard  
 Function – Individuals Observed At Time  $t$  At Jobs Offering Health Insurance Coverage<sup>1</sup>

Transition	Independent Variable	Coefficient	z-Statistic
[ $H(t + 1) = 0, T(t) = 1$ ]	experien	-.125	-2.18
	expersqr	.000	0.02
	grade	-.063	-3.20
	hi_pct	-2.477	-3.72
	pct_ch	-.425	-0.43
	lnhosp	.673	4.98
	pctchgph	.559	1.10
	unemrate	-.154	-3.92
	illspell	1.073	6.71
	lnwage	-1.229	-11.87
	constant	-.241	-.28
[ $H(t + 1) = 1, T(t) = 1$ ]	experien	.006	0.16
	expersqr	-.006	-2.73
	grade	.092	8.36
	hi_pct	-.211	-0.53
	pct_ch	-.481	-0.79
	lnhosp	.279	3.38
	pctchgph	.023	0.08
	unemrate	-.218	-8.89
	illspell	.522	4.08
	lnwage	-.851	-13.87
	constant	-1.704	-3.32
[ $H(t + 1) = 0, T(t) = 2$ ]	experien	-.293	-3.84
	expersqr	.009	1.80
	grade	-.245	-8.33
	hi_pct	-1.445	-1.49
	pct_ch	-2.740	-1.99
	lnhosp	.897	4.69
	pctchgph	1.529	2.26
	unemrate	.057	1.11
	illspell	.829	4.84
	lnwage	-.139	-1.33
	constant	-2.481	-2.92

Table 8b (continued)  
**Multinomial Logit Estimates of Health Insurance Coverage and Job Mobility Hazard  
 Function – Individuals Observed At Time  $t$  At Jobs Offering Health Insurance Coverage<sup>1</sup>**

Transition	Independent Variable	Coefficient	z-Statistic
[ $H(t + 1) = 1, T(t) = 2$ ]	experien	-.116	-2.00
	expersqr	-.002	-0.50
	grade	-.154	-7.83
	hi_pct	-2.015	-3.12
	pct_ch	-1.301	-1.30
	lnhosp	.818	6.22
	pctchghp	1.123	2.37
	unemrate	-.058	-1.54
	illspell	.829	4.84
	lnwage	-.139	-1.33
	constant	-2.481	-2.92
[ $H(t + 1) = 0, T(t) = 3$ ]	experien	-.234	-4.90
	expersqr	.018	7.06
	grade	-.009	-0.52
	hi_pct	-2.359	-4.37
	pct_ch	-2.402	-2.20
	lnhosp	1.486	12.67
	pctchghp	1.399	3.12
	unemrate	.110	3.35
	illspell	-.226	-0.95
	lnwage	-.856	-9.69
	constant	-7.116	-9.31

Notes:

(1) At time of initial observation (time  $t$ ) each individual is working at a job which offers health insurance coverage.

(2) In column 1  $H(t + 1)$  represents health insurance status in the year following the initial observation, and  $T(t)$  represents the job transition that took place between the two observations.  $T(t) = 1$  represents a job quit;  $T(t) = 2$  represents an involuntary job separation; and  $T(t) = 3$  represents no job change.

(3) For identification the category [ $H(t + 1) = 1, T(t) = 3$ ] is excluded.

(4) Pseudo  $R^2 = .046$ ,  $N = 18690$

Table 9a  
 Probit Estimates of the Likelihood of a Job Change

Independent Variable	(1)		(2)		(3)		(4)	
	No Fixed Effects	Fixed Effects	No Fixed Effects	Fixed Effects	No Fixed Effects	Fixed Effects	No Fixed Effects	Fixed Effects
experien	-.052 (.028)	-.055 (.041)	-.052 (.028)	-.055 (.041)	-.052 (.028)	-.055 (.041)	-.052 (.028)	-.055 (.041)
expersqr	-.002 (.002)	.003 (.002)	-.002 (.002)	.003 (.002)	-.002 (.002)	.003 (.002)	-.002 (.002)	.003 (.002)
grade	-.027* (.009)	-.029 (.075)	-.027* (.009)	-.032 (.075)	-.027* (.009)	-.029 (.075)	-.027* (.009)	-.031 (.075)
illspell	.475* (.106)	.563* (.140)	3.970* (1.515)	3.168 (1.953)	-.756 (1.801)	2.925 (2.402)	2.926 (2.424)	6.013 (3.281)
hi_pct	-1.233* (.326)	1.20 (.898)	-1.100* (.331)	1.319 (.903)	-1.227* (.326)	1.192 (.898)	-1.082* (.331)	1.327 (.904)
pct_ch	-1.119* (.500)	.741 (.728)	-1.071* (.501)	.795 (.729)	-1.122* (.500)	.758 (.729)	-.962 (.508)	.906 (.738)
lnhosp	.346* (.067)	-.458* (.236)	.342* (.067)	-.466* (.236)	.339* (.068)	-.442 (.237)	.338* (.068)	-.442 (.237)
pctchghp	.169 (.243)	-.243 (.340)	.169 (.243)	-.239 (.340)	.166 (.243)	-.232 (.340)	.129 (.246)	-.280 (.344)
illhipct	--	--	-4.556 (1.969)	-3.387 (2.529)	--	--	.148 (.351)	-.579 (.475)
illhipc	--	--	--	--	--	--	1.622 (1.654)	2.222 (2.252)
illhp	--	--	--	--	.230 (.335)	-.440 (.446)	-4.415* (1.984)	-3.312 (2.552)
illhpc	--	--	--	--	--	--	-3.887 (3.112)	-3.252 (4.150)
constant	-.636 (.419)	-3.77 --	-.722 (.421)	-3.798 --	-.601 (.422)	-3.85 --	-.706 (.424)	-3.940 --

Notes:

\* Indicates significance at the 5% level.

(1) Standard errors in parentheses.

(2) Probit estimates based on a subset of 800 respondents randomly drawn from the sample of respondents represented in table 3.

(3) Eight-hundred individual level fixed effects are not reported.

(4) Sample includes 5216 observations.

Table 9b  
 Probit Estimates of the Likelihood of a Job Change – MALES ONLY

Independent Variable	No Fixed Effects	Fixed Effects
experien	-.072* (.027)	-.087* (.040)
expersqr	-.001 (.002)	.005* (.002)
grade	-.057* (.008)	-.055 (.078)
illspell	.571* (.104)	.337* (.136)
hi_pct	-.988* (.336)	.925 (.921)
pct_ch	-1.748* (.508)	-.708 (.921)
lnhosp	.268* (.067)	-.608* (.238)
pctchghp	-.058 (.234)	-.869* (.313)
constant	.072 (.418)	-1.61 --

Notes:

\* Indicates significance at the 5% level.

(1) Standard errors in parentheses.

(2) Probit estimates based on a subset of 800 respondents randomly drawn from the sample of respondents represented in table 3.

(3) Eight-hundred individual level fixed effects are not reported.

(4) Sample includes 5275 observations.



Table 9c  
 Probit Estimates of the Likelihood of a Job Change – FEMALES ONLY

Independent Variable	No Fixed Effects	Fixed Effects
experien	-.030 (.027)	-.013 (.040)
expersqr	-.003 (.002)	.002 (.002)
grade	.003 (.009)	-.050 (.060)
illspell	.543* (.097)	.536* (.129)
hi_pct	-.505 (.320)	1.359 (.883)
pct_ch	-.181 (.497)	1.475* (.697)
lnhosp	.174* (.066)	-.864* (.226)
pctchgph	.015 (.229)	-.471 (.309)
constant	-.770 (.415)	4.704 --

Notes:

\* Indicates significance at the 5% level.

(1) Standard errors in parentheses.

(2) Probit estimates based on a subset of 800 respondents randomly drawn from the sample of respondents represented in table 3.

(3) Eight-hundred individual level fixed effects are not reported.

(4) Sample includes 5280 observations.

Table 9d  
 Probit Estimates of the Likelihood of a Job Change - MARRIED ONLY

Independent Variable	No Fixed Effects	Fixed Effects
experien	-.054 (.028)	-.112* (.042)
expersqr	-.002 (.002)	.005* (.002)
grade	-.047* (.008)	-.061 (.083)
illspell	.500* (.093)	.505* (.124)
hi_pct	-1.325* (.340)	.473 (.905)
pct_ch	-.959 (.495)	.753 (.699)
lnhosp	.196* (.069)	-.475 (.255)
pctchghp	-.157 (.255)	-.784* (.349)
constant	.453 (.426)	-1.709 --

Notes:

\* Indicates significance at the 5% level.

(1) Standard errors in parentheses.

(2) Probit estimates based on a subset of 800 respondents randomly drawn from the sample of respondents represented in table 3.

(3) Eight-hundred individual level fixed effects are not reported.

(4) Sample includes 5336 observations.

Table 9e  
 Probit Estimates of the Likelihood of a Job Change - UNMARRIED ONLY

Independent Variable	No Fixed Effects	Fixed Effects
experien	-.085 (.027)	-.069 (.042)
expersqr	-.001 (.002)	.003 (.002)
grade	-.022* (.009)	-.207* (.076)
illspell	.320* (.104)	.180 (.138)
hi_pct	-1.148* (.325)	.726 (.934)
pct_ch	-1.074 (.494)	-.124 (.726)
lnhosp	.168* (.068)	-.601* (.235)
pctchghp	-.081 (.247)	-.268 (.329)
constant	.381 (.421)	-.968 --

Notes:

\* Indicates significance at the 5% level.

(1) Standard errors in parentheses.

(2) Probit estimates based on a subset of 800 respondents randomly drawn from the sample of respondents represented in table 3.

(3) Eight-hundred individual level fixed effects are not reported.

(4) Sample includes 5182 observations.

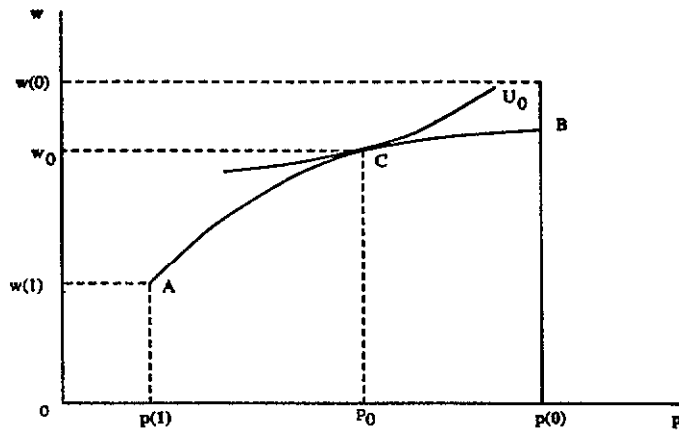


Figure 1: Individual Preferences For Wages and Health Insurance

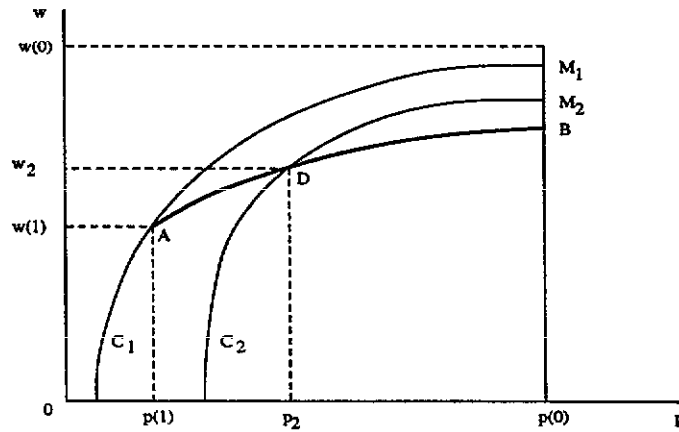


Figure 2: Derivation of the Employer Utility Offer Curve

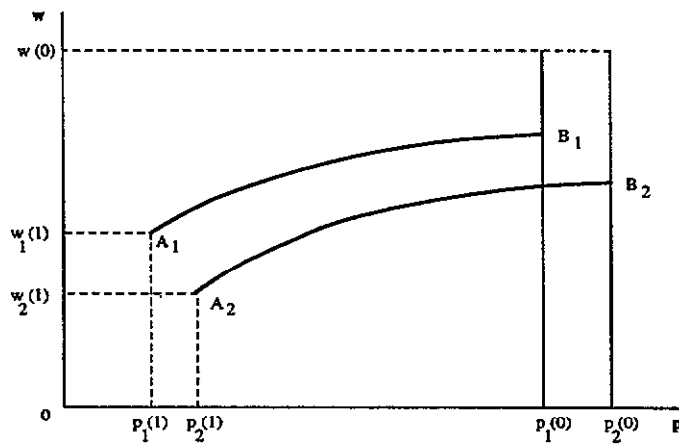


Figure 3a: Differences In State Health Insurance Costs

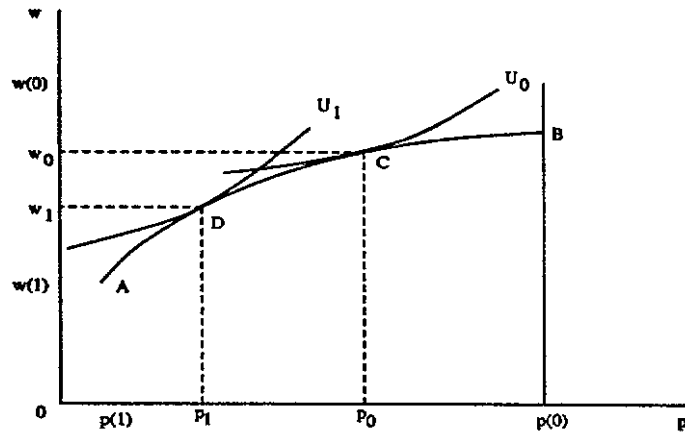


Figure 3b: Differences In Health Insurance Preferences

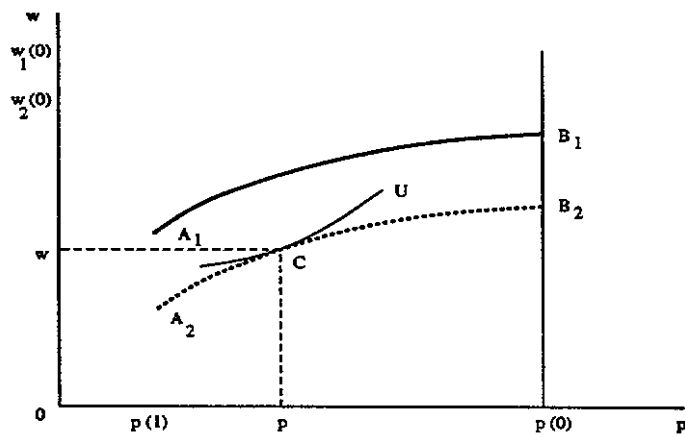


Figure 3c: A Low Employer-Employee Match

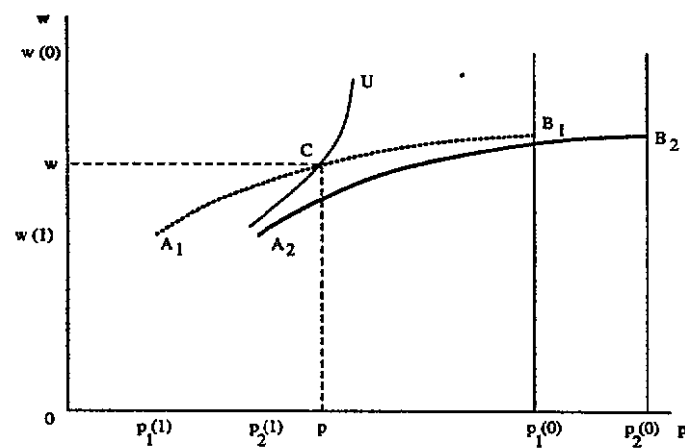


Figure 3d: The Effect of Pre-existing Conditions On Job Mobility

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## Appendix

### Solution of the Search Maximization Problem<sup>61</sup>

Let time be represented by a sequence of discrete periods of length  $h$ . Let  $\beta(h)$  be the discount factor which is a function of  $h$ . Let  $c$  be the out of pocket cost of search per unit of time. Let  $\bar{u} = \bar{u}(\bar{w}, \bar{p})$  be the instantaneous utility flow from a current job that offers wage  $\bar{w}$  and insurance price  $\bar{p}$ . Let  $u = u(w, p)$  be the instantaneous utility flow of an employment offer of  $(w, p)$  from the market. Let the probability distribution function of utility offers from the market be  $F(u)$ . A new employment offer from the current employer, given by  $y$ , is drawn at the end of every period  $h$ . The probability distribution function of  $y$ ,  $\tilde{F}(y)$ , depends upon all currently available information, so that the expectation of  $y$  is time dependent. Offers from both the market and the current employer distributions are drawn independently of one another, and are independently and identically distributed over time.

Let  $q(n, h)$ ,  $n = 0, 1, \dots$ , be a probability distribution of the number of employment offers from the market,  $n$ , in time interval  $h$ . The offer to be considered in each period is the best of those that were offered in the period, i.e.,

$$u = \max [u_1, \dots, u_n] \quad (22)$$

Let  $H(u, n)$  represent the probability that the best of  $n$  offers is less than  $u$  given that  $n \geq 1$ . As in Mortensen I specify the the density  $q$  to be Poisson,

$$q(n, h) = e^{-\lambda h} (\lambda h)^n / n!, \quad (23)$$

where  $\lambda$  is the offer arrival rate. Let  $V^G(u)$  denote the value of discontinuing search, accepting the offer from the market,  $u$ , and working forever at the same level of compensation.  $V^G(u)$  is continuous and strictly increasing, and  $V^G(0) = 0$  by assumption.  $V^S(\bar{u})$  is the value of the current utility offer  $\bar{u}$ , and  $V^S(y)$  is the value of the new offer from the current employer. Then the individual's problem is to maximize  $V^S$ , where  $V^S$  solves:

$$V^S(\bar{u}) = (\bar{u} - c)h + \beta(h) \int_0^\infty \left[ \sum_{n=1}^\infty q(n, h) \right]$$

---

<sup>61</sup>The solution in this section is adapted from Mortensen (1986). A few unique phrases are taken directly from that text.

$$\int_0^\infty \max(V^S(y), V^G(u)) dH(u, n) + q(0, h)V^S] d\tilde{F}(y) \quad (24)$$

$$= (\tilde{u} - c)h + \beta(h)E[V^S(y)] +$$

$$\beta(h) \int_0^\infty \left[ \sum_{n=1}^\infty q(n, h) \int_0^\infty \max(0, V^G(u) - V^S(y)) dH(u, n) \right] d\tilde{F}(y) \quad (25)$$

Assuming that the mean utility of the market offer distribution and that of the current employer offer distribution are finite we know that (25) has a unique solution given by a reservation utility level,  $u^*$ .  $u^*$  is defined as the unique solution to

$$V^G(u^*) = V^S(\tilde{u}). \quad (26)$$

Equation (25) is greatly simplified by converting the problem to continuous time, which is the limiting case when the period length,  $h$ , becomes very small. The following results are useful for the transformation to continuous time:

$$\lim_{h \rightarrow 0} q(1, h)/h = \lambda \quad \text{and} \quad \lim_{h \rightarrow 0} q(n, h)/h = 0, \quad \text{for } n \geq 1,$$

and if the discount factor is specified as

$$\beta(h) = e^{-rh}, \quad \text{then} \quad \lim_{h \rightarrow 0} [1 - \beta(h)]/h = r,$$

where  $r$  is the rate of interest. With the above results handy we can divide both sides of (25) by  $h$ , and re-express it as

$$\frac{1 - \beta(h)}{h} V^S(\tilde{u}) = (\tilde{u} - c) + \beta(h) \frac{E[V^S(y)] - V^S(\tilde{u})}{h} + \beta(h) \int_0^\infty \left[ \frac{1}{h} \sum_{n=1}^\infty q(n, h) \int_0^\infty \max(0, V^G(u) - V^S(y)) dH(u, n) \right] d\tilde{F}(y) \quad (27)$$

Taking the limit of (27) as  $h \rightarrow 0$  we get

$$rV^S(\tilde{u}) = \tilde{u} - c + \frac{dE[V^S(y)]}{dt} + \int_0^\infty \int_0^\infty \max(0, V^G(u) - V^S(y)) dF(u) d\tilde{F}(y), \quad (28)$$

where  $dt$  represents an infinitesimal change in time. Also, since  $V^G(u) = u/r$  we know from (26) that

$$rV^S(\tilde{u}) = rV^G(u^*) = u^*. \quad (29)$$

Using (29) to substitute for  $V^S(\tilde{u})$ ,  $V^G$  in (28) gives

$$u^* = \tilde{u} - c + \frac{dE[V^S(y)]}{dt} + \frac{\lambda}{r} \int_0^\infty \int_{u^*}^\infty (u - rV^S(y)) dF(u) d\tilde{F}(y) \quad (30)$$

Equation (30) represents the first order condition for the solution to the optimal stopping problem.

It follows from (30) that an individual will search if and only if

$$u^* > \tilde{u} \Leftrightarrow \frac{\lambda}{r} \int_0^\infty \int_{rV^S(y)}^\infty (u - rV^S(y)) dF(u) d\tilde{F}(y) + \frac{dE[V^S(y)]}{dt} > c. \quad (31)$$

That is, search will occur if and only if the expected increase in discounted utility from search is greater than the cost of search.

Replacement of the utility flow on the current job with its indirect utility equivalent,  $\tilde{u} = \tilde{w} e^{-\alpha\tilde{p}}$ , into (30) gives

$$u^* = \tilde{w} e^{-\alpha\tilde{p}} - c + \frac{dE[V^S(y)]}{dt} + \frac{\lambda}{r} \int_0^\infty \int_{rV^S(y)}^\infty (u - rV^S(y)) dF(u) d\tilde{F}(y) \quad (32)$$

where  $u^*$  is now explicitly a function of current utility flow parameter  $\alpha$  and prices  $w$  and  $p$ .

### Comparative Static Results from the Search First Order Condition

As was noted in the text comparative static effects are derived under the assumption of stationarity. Stationarity is imposed on the model by assuming that the current job utility offer is a fixed constant. When stationarity is imposed, equation (32) reduces to:

$$u^* = \tilde{w} e^{-\alpha\tilde{p}} - c + \frac{\lambda}{r} \int_{u^*}^\infty (u - u^*) dF(u) \quad (33)$$

Equation (33) can be rewritten in a useful form as

$$u^* \left( \frac{r + \lambda}{r} \right) = \tilde{w} e^{-\alpha\tilde{p}} - c + \frac{\lambda}{r} \int_0^{u^*} F(u) du + \frac{\lambda}{r} E_F(u) \quad (34)$$

Equation (34) allows us to derive the following comparative static results:

$$\frac{\partial u^*}{\partial \alpha} = \frac{-\tilde{w} \tilde{p} e^{-\alpha\tilde{p}} r}{r + \lambda(1 - F(u^*))} < 0 \quad (35)$$

$$\frac{\partial u^*}{\partial c} = \frac{-r}{r + \lambda(1 - F(u^*))} < 0 \quad \text{and} \quad > -1 \quad (36)$$

$$\frac{\partial u^*}{\partial \tilde{w}} = \frac{r e^{-\alpha\tilde{p}}}{r + \lambda(1 - F(u^*))} > 0 \quad (37)$$

$$\frac{\partial u^*}{\partial \tilde{p}} = \frac{-\alpha r \tilde{w} e^{-\alpha\tilde{p}}}{r + \lambda(1 - F(u^*))} < 0 \quad (38)$$

$$\frac{\partial u^*}{\partial \lambda} = \frac{\int_{u^*}^\infty (u - u^*) dF(u)}{r + \lambda(1 - F(u^*))} > 0 \quad (39)$$

$$\frac{\partial u^*}{\partial r} = \frac{-\frac{\lambda}{r} [\int_{u^*}^\infty (u - u^*) dF(u)]}{r + \lambda(1 - F(u^*))} < 0 \quad (40)$$

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