# Traffic-Related Air Pollution and Perinatal Mortality: A Case-Control Study

Andréa Paula Peneluppi de Medeiros,<sup>1</sup> Nelson Gouveia,<sup>1</sup> Reinaldo Paul Pérez Machado,<sup>2</sup> Miriam Regina de Souza,<sup>1</sup> Gizelton Pereira Alencar,<sup>3</sup> Hillegonda Maria Dutilh Novaes,<sup>1</sup> and Márcia Furquim de Almeida<sup>3</sup>

<sup>1</sup>Department of Preventive Medicine, Faculty of Medicine, <sup>2</sup>Department of Geography, and <sup>3</sup>Department of Epidemiology, Faculty of Public Health, University São Paulo, São Paulo, Brazil

BACKGROUND: Ambient levels of air pollution may affect the health of children, as indicated by studies of infant and perinatal mortality. Scientific evidence has also correlated low birth weight and preterm birth, which are important determinants of perinatal death, with air pollution. However, most of these studies used ambient concentrations measured at monitoring sites, which may not consider differential exposure to pollutants found at elevated concentrations near heavy-traffic roadways.

OBJECTIVES: Our goal was to examine the association between traffic-related pollution and perinatal mortality.

METHODS: We used the information collected for a case–control study conducted in 14 districts in the City of São Paulo, Brazil, regarding risk factors for perinatal deaths. We geocoded the residential addresses of cases (fetal and early neonatal deaths) and controls (children who survived the 28th day of life) and calculated a distance-weighted traffic density (DWTD) measure considering all roads contained in a buffer surrounding these homes.

RESULTS: Logistic regression revealed a gradient of increasing risk of early neonatal death with higher exposure to traffic-related air pollution. Mothers exposed to the highest quartile of the DWTD compared with those less exposed exhibited approximately 50% increased risk (adjusted odds ratio = 1.47; 95% confidence interval, 0.67–3.19). Associations for fetal mortality were less consistent.

CONCLUSIONS: These results suggest that motor vehicle exhaust exposures may be a risk factor for perinatal mortality.

KEY WORDS: air pollution, epidemiology, geographic information systems, perinatal mortality, traffic density. *Environ Health Perspect* 117:127–132 (2009). doi:10.1289/ehp.11679 available via http://dx.doi.org/ [Online 22 September 2008]

Child mortality has decreased substantially in Brazil in past decades; nonetheless, it still has higher values than those found in developed countries (Vermelho et al. 2002). Perinatal deaths—those that occur in utero (≥ 22 weeks of gestation or fetal deaths) and in the first 6 days of life (early neonatal deaths)—contribute most to this situation. Perinatal mortality is an indicator of mother and child health and may reflect the conditions of reproductive health, which is related to socioeconomic position, quality of antenatal care, and delivery characteristics (Jackson et al. 1999). Some of the several risk factors identified for perinatal mortality include mother-related factors and complications in pregnancy, labor, and delivery, which may affect both the newborn and the fetus (e.g., maternal hypertension, placenta previa, premature placental detachment, and other morphologic and functional abnormalities of the placenta); respiratory and cardiovascular disorders specific to the perinatal period; infections; and disorders related to the length of pregnancy and fetal growth [e.g., intrauterine growth restriction (IUGR), preterm birth, and low birth weight] (Almeida et al. 2007; Conde-Agudelo et al. 2000; Kramer et al. 2002; Schoeps et al. 2007). More recently, studies have indicated that exposure to air pollution may be associated with low birth weight and preterm birth (Ritz and Yu 1999; Ritz et al. 2000), both of which are important risk factors for perinatal death. In addition, urban air pollution has also been directly implicated in perinatal mortality (Lin et al. 2004; Nishioka et al. 2000; Pereira et al. 1998). The association between deaths in the neonatal period (0–28 days of life) and air pollution was also identified in the studies by Lipfert et al. (2000) and Hajat et al. (2007).

However, these studies assessed the exposure to air pollution based on the average concentration of pollutants obtained from air quality monitoring stations. This direct measuring system reflects the levels of pollutants in a relatively large area. Although this procedure may adequately reflect average exposure of pregnant women to background levels of air pollution in their neighborhood, it does not take into account differences in exposure within neighborhoods because of proximity to heavy-traffic roadways and freeways. It is likely that those living closer to these sources may experience greater exposure to toxic compounds released directly in vehicle exhaust or formed in the atmosphere adjacent to roadways (Wilhelm and Ritz 2003). In addition, measurement studies have shown that concentrations of traffic-related air pollutants are elevated near roadways (Kinney et al. 2000; Lena et al. 2002).

Many studies have investigated the effects of air pollution on health using traffic-related exposure (Langholz et al. 2002; Pearson et al. 2000; Wilhelm and Ritz 2003) and found higher health effects among those living closer to streets with higher traffic volumes. Therefore, in the present study we aimed to investigate the association between perinatal mortality and maternal exposure to traffic-related air pollutants, assessed indirectly by calculating the distance-weighted traffic density (DWTD) in the vicinities of mothers' homes, which we based on the subjects' residential address, the distance of their homes to surrounding streets, and the traffic flow in these streets.

## **Materials and Methods**

We used information collected from a case-control study conducted in 14 districts located in the south region of São Paulo, Brazil, regarding risk factors for early fetal and neonatal deaths (Almeida et al. 2007; Schoeps et al. 2007). These districts account for approximately 23% of the population and 44% of the city's geographic area (Almeida et al. 2004). This study was approved by the ethical committees of all institutions involved, and all participants gave informed consent before interviews

Subjects. The population studied comprised all births and fetal deaths from women living in this area in the period between 1 August 2000 and 31 January 2001. Cases were all perinatal deaths that occurred in the study period in the area identified through a linkage of the Brazilian Live Birth Information System (SINASC) and the Brazilian Mortality Information System (SIM) databases (both available at www.datasus.gov.br). Both systems have very high population coverage for São Paulo (Almeida et al. 2006). Controls

Address correspondence to N. Gouveia, Department of Preventive Medicine, Faculty of Medicine, University of São Paulo, Av. Dr. Arnaldo, 455, São Paulo, SP 01246-903, Brazil. Telephone: (55 11) 3061-7075. Fax: (55 11) 3062-6018. E-mail: ngouveia@usp.br

This study is part of the project "Perinatal Mortality in the South Region of the City of São Paulo: A Case—Control Population-Based Study" approved by the Committee of Ethics in Research of the Faculty of Public Health, University São Paulo; it is funded by the State of São Paulo's Research Support Foundation (Fapesp Process 99/11985-9), and by the National Scientific and Technological Development Council (CNPq Process 479007/2001-2).

The authors declare they have no competing financial interests.

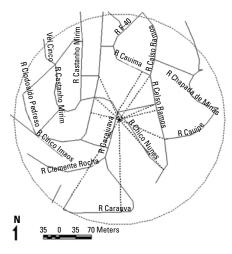
Received 8 May 2008; accepted 22 September 2008.

were a random sample of all children who survived the 28th day of life, and were obtained through the same databases. We excluded all nonhospital deliveries. The initial sample was defined for a study on perinatal mortality; however, preliminary analysis of results indicated differences between risk factors for early neonatal and fetal deaths. Thus, we decided to analyze these components separately, maintaining the number of controls previously established.

Primary data were collected by home interviews up to 6 months after the event using a form containing variables regarding families' socioeconomic characteristics, maternal reproductive history and risk factors, outcome of previous and present pregnancies and deliveries, and use of and access to health care. A protocol for collecting data from patient charts in the hospitals where deliveries took place was used for both cases and controls. For deaths that occurred in a different hospital (other than that of the delivery), data were also obtained from patient charts of those services.

Exposure assessment. We geocoded the residential addresses of case and control mothers based on their postal code and street number using MapInfo (professional version 8.5; MapInfo Corporation, New York, NY, USA).

A digital cartographic base of the studied area was provided by the Traffic Engineering Company (CET) of the Department of Transport of the city of São Paulo. This contained a simulation of traffic flows on every segment of all freeways, arterials, and collectors attained from systematic counts conducted in the peak morning hours, from 0730 to 0830 hours. This data set also contained an estimate of the total traffic flow from all local



Streets
--- Buffer 750 feet

Geocoded home/event

····· Shortest distance

**Figure 1.** Representation of the calculation of shortest distances between location of a case or control and all roadways in a radius of 750 feet.

roads inside predefined microareas (polygons). Therefore, for local roads without information on flows, we followed a procedure developed by CET that divides the estimated total flow for each polygon proportionally for each local road according to its length inside the area. We performed this procedure using Maptitude software (version 4.6; Calliper Corp., Newton, MA, USA), which enabled us to distribute and attribute traffic flow values for all local roads without counts.

Using a method similar to that of Pearson et al. (2000), we calculated the DWTD for each subject, which we used as an indicator of exposure to traffic-related air pollutants close to their homes. More specifically, we delineated a radius of 750 feet (228.6 m) around the home of each case and control and identified all streets contained in this buffer for later calculation of the perpendicular distance between each street and home. We applied a model also developed by Pearson et al. (2000) to estimate the dispersion of pollutants emitted by cars inside this circumference. This model assumes that 96% of the total pollutants emitted disperse up to 500 feet (152.4 m) from the road according to the equation

$$Y = 1/[0.4(2\Pi)^{1/2}] \times \exp\{[(0.5)(D/500)^2]/0.4^2\},$$
 [1]

where *D* is the shortest distance from the home to each street within the delimited area (Figure 1) and *Y* is the value used to weight the traffic flow for each street. We then summed the weighted values of the traffic flow for all streets surrounding the subject's home.

Analysis. We fit separate models for fetal and neonatal deaths to explore in detail the effect of our indicator of exposure to traffic pollution (DWTD) on each outcome. We grouped the DWTD values into quartiles derived from its distribution for all subjects and examined their association with each outcome using logistic regression analysis.

Based on previous analyses of these data (Almeida et al. 2007; Schoeps et al. 2007), we adjusted our models for covariates representing four different groups of potential risk factors for perinatal mortality according to a hierarchical conceptual framework. This included *a*) socioeconomic and demographic characteristics of mothers and families (family income, head of the household education, maternal occupation, maternal marital status and length of union, housing type, and crowding), *b*) characteristics

of mothers before pregnancy (age, parity, previous low-birth-weight infant), c) conditions during pregnancy (vaginal bleeding, hypertension, diabetes, cigarette and alcohol consumption, planning of the pregnancy, and prenatal care), and d) characteristics of delivery and the fetus (sex, birth weight, gestational age, IUGR, presence of fetal malformation, and problems during delivery).

We combined marital status and length of union into one variable. We grouped maternal occupation during the pregnancy in a dichotomous variable "working" and "not working" (unemployed, retired, housewife, student, or without activity). Housing type refers to the construction material (brickwork or other materials, e.g., wood, usually found in more precarious housing). We based problems during delivery on maternal report (yes/no), and they refer mostly to hemorrhage and eclampsia.

We used univariate regression to obtain crude odds ratios (OR) and to select the most important covariates (p < 0.20) from each of the four groups to be included in the multivariate logistic models. We kept variables in the final models if they attained statistical significance (p < 0.05) or if we considered them important potential confounders of the association between our indicator of exposure and perinatal mortality, regardless of their significance. We did not include birth weight, gestational age, and IUGR in our models because these conditions may lie in the causal pathway between traffic-related air pollution and perinatal mortality, so it would not be adequate to control for them.

#### Results

Of the total 749 eligible events (378 cases and 371 controls), we omitted 60 cases (15.9%) and 58 controls (15.6%) because of incorrect residential address (16 cases and 8 controls), mother refusal (12 cases and 9 controls), event without hospital record (21 cases and 19 controls), or hospital refused access to records (11 cases and 22 controls).

Therefore, we obtained a sample of 631 events in the period between August 2000 and January 2001, corresponding to 318 perinatal deaths (146 early neonatal and 172 fetal deaths) and 313 unpaired controls. Of the total 172 fetal deaths, we excluded 10 from the study because eight were intrapartum deaths (based on obstetric records), one event did not discriminate sex, and another could not be geocoded, leaving 162 fetal deaths for analysis.

Table 1. Descriptive statistics of the DWTD values according to case and control categories.

Category	Mean ± SD	Median	25th percentile	75th percentile
All events	504 ± 1225.8	45.3	6.1	370.2
Cases	565.3 ± 1361.8	52.6	7.4	434.1
Fetal	530.9 ± 1364.3	45.1	6.8	426.5
Early neonatal	626.8 ± 1398.6	72.5	8.5	586.6
Controls	441.9 ± 1069.3	35.8	5.4	340.9

The distribution of the DWTD for this population was highly skewed, with a minimum of 6 vehicles/hr and maximum of > 10,000 vehicles/hr (median of 45.3 vehicles/hr) (Table 1). On average, cases had higher values than did controls for this indicator, although the difference was not statistically significant. In addition, DWTD values were higher among early neonatal deaths than among fetal deaths.

Table 2 summarizes the demographic characteristics of cases and controls as well as maternal conditions during pregnancy. As expected, proportions of low birth weight and prematurity were much higher among cases than among controls. Mothers of early neonatal

deaths were younger than other groups, lived in more crowded environments, and had worse social conditions than did mothers of controls. There were also a greater proportion of single mothers among cases, especially for fetal deaths. In addition, mothers of cases had higher parity and greater proportions of unplanned pregnancies and previous low-birth-weight children. They also had worse antenatal care, smoked substantially more, and exhibited higher rates of morbidities during pregnancy than did mothers of controls (Table 2).

As expected, low birth weight and prematurity exhibited the strongest association with perinatal mortality, followed by morbidities developed during pregnancy such as diabetes, vaginal bleeding (possible proxy for placenta previa), and hypertension. The presence of congenital anomalies yielded a much higher risk for neonatal mortality (OR = 13.6) than for fetal mortality (OR = 3.2).

The univariate analysis of the association between DWTD and perinatal mortality showed that the OR for neonatal deaths increased with DWTD quartiles, and in the highest quartile we observed a 78% increase in risk for this outcome [OR = 1.78; 95% confidence interval (CI), 1.01-3.16]. For fetal deaths we did not observe such a consistent increase in risk (OR in the highest quartile = 1.35; 95% CI, 0.79-2.30).

Table 2. Subjects [no. (%)] and univariate ORs according to demographic and socioeconomic characteristics, maternal conditions before and during pregnancy, and characteristics of delivery and fetus.

			l deaths		al deaths				deaths		ıl deaths
Variable	Controls [no. (%	6)] no. (%)	OR	no. (%)	OR	Variable	Controls [no. (%	)] no. (%)	OR	no. (%)	OR
Socioeconomic a	and demographic	characteristic	S			Conditions during	pregnancy				
Per capita family	/ income at Brazil	ian minimum	wage			Diabetes					
> 10	38 (12.6)	14 (9.1)	1	13 (9.4)	1	No	309 (99.7)	150 (95.5)	1	140 (97.9)	1
3–9	143 (47.5)	74 (48.4)	1.4	61 (44.2)	1.25	Yes	1 (0.3)	7 (4.5)	14.42	3 (2.1)	6.62
≤ 2	120 (39.9)	65 (42.5)	1.47	64 (46.4)	1.56				p = 0.013		p = 0.103
			p = 0.353		p = 0.157	Maternal smoking					
Head-of-family e	education (years)					No	254 (81.2)	118 (72.8)	1	98 (67.1)	1
< 12	299 (95.5)	159 (98.1)	1	144 (98.6)	3.37	Yes	59 (18.8)	44 (27.2)	1.6	48 (32.9)	2.11
≥ 12	14 (4.5)	3 (1.9)	2.48	2 (1.4)	1				p = 0.038		p = 0.001
			p = 0.158		p = 0.111	Maternal alcohol					
Maternal occupa	ation					No	257 (82.1)	172 (75.3)	1	120 (82.2)	1
Working	121 (39.7)	76 (51.7)	1	61 (46.2)	1	Yes	56 (17.9)	40 (24.7)	1.5	26 (17.8)	0.99
Not working	184 (60.3)	71 (48.3)	0.66	71 (53.8)	0.76				p = 0.081		p = 0.983
			p = 0.042		p = 0.203	Planned pregnand	СУ				
Marital status ar	nd length of unio	n (years)				Yes	116 (37.1)	46 (28.4)	1	39 (26.7)	1
≥ 1	243 (77.6)	94 (58.0)	1	96 (65.8)	1	No	197 (62.9)	116 (71.6)	1.48	107 (73.3)	1.61
< 1	22 (7.0)	20 (12.4)	2.35	17 (11.6)	1.96				p = 0.06		p = 0.03
Single mother	48 (15.4)	48 (29.6)	2.58	33 (22.6)	1.74	Antenatal care					
			<i>p</i> < 0.001		p = 0.014	Adequate	232 (74.1)	83 (51.2)	1	73 (50.0)	1
Housing type						Inadequate	77 (24.6)	70 (43.2)	2.54	52 (35.6)	2.15
Brickwork	245 (78.3)	133 (82.0)	1	93 (63.7)	1	None	4 (1.3)	9 (5.6)	6.29	21 (14.4)	16.68
Other materials	s 68 (21.7)	29 (18.0)	0.78	53 (36.3)	2.05				<i>p</i> < 0.001		p < 0.001
			p = 0.328		p = 0.001	Characteristics of	f delivery and the	efetus			
No. occupants p						Sex					
< 1	92 (29.3)	42 (25.9)	1	31 (21.2)	1	Female	156 (49.8)	72 (44.4)	1	58 (39.7)	1
1–2	203 (64.9)	104 (64.2)	1.12	94 (64.4)	1.37	Male	157 (50.2)	90 (55.6)	1.14	88 (60.3)	1.51
3–4	18 (5.8)	16 (9.9)	1.92	21 (14.4)	3.46				p = 0.265		p = 0.044
			p = 0.159		p = 0.003	Birth weight (g)					
	teristics before p	regnancy				≥ 2,500	288 (92.0)	38 (24.5)	1	27 (18.5)	1
Age (years)						< 2,500	25 (8.0)	117 (75.5)	37.6	119 (81.5)	50.77
> 20	261 (83.4)	132 (81.5)	1	102 (69.9)	1				p < 0.001		p < 0.001
≤ 20	52 (16.6)	30 (18.5)	1.14	44 (30.1)	2.16	Length of gestation					
			p = 0.603		p = 0.001	≥ 37	269 (85.9)	51 (32.7)	1	28 (19.2)	1
Parity						< 37	44 (14.1)	105 (67.3)	12.59	118 (80.8)	25.76
2nd or 3rd	155 (49.5)	79 (48.8)	1	53 (36.3)	1				<i>p</i> < 0.001		p < 0.001
1st	128 (40.9)	61 (37.6)	0.93	69 (47.3)	0.93	Congenital anoma			_		
≥ 4th	30 (9.6)	22 (13.6)	1.44	24 (16.4)	1.8	No	304 (97.1)	148 (91.4)	1	104 (71.2)	1
			p = 0.467		p = 0.185	Yes	9 (2.9)	14 (8.6)	3.19	42 (28.8)	13.64
Previous low-bir		100 /00 41		444 (70 0)		B			p = 0.008		<i>p</i> < 0.001
No	281 (89.8)	133 (82.1)	1	111 (76.0)	1	Problems during of		400 (75.0)		444/70 61	
Yes	32 (10.2)	29 (17.9)	1.91	35 (24.0)	2.77	No	289 (92.3)	123 (75.9)	1	114 (78.6)	1
0 1111			p = 0.019		<i>p</i> < 0.001	Yes	24 (7.7)	39 (24.1)	3.82	31 (21.4)	3.27
Conditions durin						DWITD			<i>p</i> < 0.001		<i>p</i> < 0.001
Vaginal bleeding		1.40 /00 41		110 (01 =)		DWTD	00 (00 4)	00 /00 5		00 (40 4)	4
No	307 (98.1)	140 (86.4)	1	119 (81.5)	1	≤ 6.0	89 (28.4)	38 (23.5)	1	28 (19.1)	1
Yes	6 (1.9)	22 (13.6)	8.04	27 (18.5)	11.61	> 6.0 to ≤ 45.3	75 (24.0)	43 (26.5)	1.34	38 (26.1)	1.61
			<i>p</i> < 0.001		<i>p</i> < 0.001	> 45.3 to ≤ 370.		39 (24.1)	1.2	39 (26.7)	1.63
Hypertension	004 (00.0)	400 (07.0)	4	404 (04.6)		> 370.2	73 (23.3)	42 (25.9)	1.35	41 (28.1)	1.78
No	291 (93.0)	109 (67.3)	1	124 (84.9)	1				p = 0.36		p = 0.06
Yes	22 (7.0)	53 (32.7)	6.43	22 (15.1)	2.35						
			<i>p</i> < 0.001		p = 0.008						

The multivariate analysis showed that length of union, previous low-birth-weight infant, vaginal bleeding, hypertension, antenatal care, problems during delivery, and congenital anomalies were all associated with fetal deaths. The OR for fetal deaths associated with the traffic-related air pollution indicator decreased after adjustment. The adjusted OR for the highest category of DWTD was 1.20 (95% CI, 0.65–2.24) (Table 3).

For neonatal deaths, length of union, housing type, maternal age, previous low-birth-weight infant, vaginal bleeding, antenatal care, sex, congenital anomalies, and presence of problems during delivery remained in the multivariate analysis as significant risk factors. The association of these deaths with DWTD adjusted for covariates exhibited a pattern similar to that in the univariate analysis—that is, an increase in risk of neonatal deaths for mothers in the higher exposure categories. However, the trend observed in the univariate analysis was no longer clear. The adjusted OR in the higher category of exposure was 1.47 (95% CI, 0.67–3.19) (Table 4).

### **Discussion**

To our knowledge, this is the first study to evaluate the association between perinatal mortality and exposure of pregnant women and newborns to pollutants from heavy-traffic roadways in the vicinity of their homes. We divided perinatal deaths into two components, fetal and early neonatal deaths, given that marked differences in the etiology of their determinants have been observed (Kramer et al. 2002).

We observed an association of fetal and early neonatal deaths with our indicator of traffic-related air pollution. Mothers in the highest categories of exposure exhibited increased risk, although this association was stronger for neonatal than for fetal deaths because risks were higher for neonatal deaths in all categories of exposure. In addition, adjusted analysis for fetal deaths exhibited an increase in risk only for the higher category of exposure to traffic-related air pollution.

For early neonatal deaths, results suggest that women with a higher load of exposure to traffic have a nearly 50% increase in risk compared with less exposed mothers. We also observed an increasing trend in the univariate analysis, but it disappeared after adjustment for other covariates.

The case-control study that provided the data for our analysis was not originally planned to evaluate the role in perinatal mortality of living close to heavy-traffic roadways, and thus potentially being exposed to higher levels of motor vehicle exhaust. Therefore, the sample size obtained may not have provided sufficient power to explore this relationship.

Although birth weight, gestational age, and IUGR are important risk factors for perinatal mortality, we did not include them in our final models. Studies have shown that these conditions are also associated with exposure to air pollution (Ritz and Yu 1999; Ritz et al. 2000), so it is very likely that they lie in the causal pathway between traffic-related air pollution and perinatal deaths. Other pregnancy complications such as hypertension and diabetes may also make the fetus more susceptible to the additional insult of air pollution, but it is less likely that they are in the causal pathway because, in our study, these are conditions developed during pregnancy and thus essentially are associated with the pregnancy itself. In all these cases, these variables might modify the effect of exposure. However, we did not have sufficient statistical power to evaluate this hypothesis.

Another limitation of this study is the potential for misclassification of the exposure of pregnant women to traffic-related air pollution. Because we estimated this exposure based on the home address, it is possible that we incorrectly classified those who spent most of their time during pregnancy in another location (e.g., work). It

 $\begin{tabular}{ll} \textbf{Table 3.} & \textbf{Adjusted OR (95\% CI)} & \textbf{for fetal deaths for each covariate included in the adjusted model.} \end{tabular}$ 

Covariate	Adjusted OR (95% CI)	<i>p</i> -Value
Marital status and length of union (yea	irs)	
≥ 1	1.00	0.001
<1	2.80 (1.32-5.96)	
Single mother	2.24 (1.30-3.85)	
Previous low-birth-weight infant		
No	1.00	0.021
Yes	2.19 (1.14–4.19)	
Conditions during pregnancy		
Vaginal bleeding		
No	1.00	0.001
Yes	6.14 (2.11–17.85)	
Hypertension		
No	1.00	< 0.001
Yes	6.61 (3.59–12.16)	
Antenatal care	4.00	
Adequate	1.00	< 0.001
Inadequate	2.33 (1.46–3.73)	
None	3.57 (0.83–15.37)	
Problems during delivery	4.00	0.004
No	1.00	< 0.001
Yes	3.31 (1.75–6.29)	
Congenital malformation	1.00	0.004
No	1.00	0.004
Yes DWTD	4.17 (1.58–11.05)	
≤ 6.0	1.00	0.709
≤ 6.0 > 6.0 and ≤ 45.3	1.06 (0.57–1.96)	0.709
> 6.0 and ≤ 45.5 > 45.3 and ≤ 370.2		
> 45.3 and ≤ 370.2 > 370.2 and ≤ 10,810.9	0.92 (0.48–1.77) 1.20 (0.65–2.24)	
> 3/U.Z dIIU ≤ 1U,01U.3	1.20 (0.00–2.24)	

Table 4. Adjusted OR (95% CI) for neonatal deaths for each covariate included in the adjusted model.

Covariate	Adjusted OR (95% CI)	<i>p</i> -Value
Length of union (years)		
≥ 1 < 1 Single mother	1.00 3.94 (1.61–9.65) 1.89 (0.95–3.76)	0.024
Housing type Brickwork Other materials	1.00 1.64 (0.93–2.91)	0.049
Maternal age > 20 ≤ 20	1.00 1.88 (1.01–3.51)	0.022
Previous low-birth-weight infant No Yes	1.00 4.82 (2.44–9.53)	< 0.001
Vaginal bleeding No Yes	1.00 12.71 (4.28–37.68)	< 0.001
Antenatal care Adequate Inadequate None	1.00 2.12 (1.20–3.75) 36.09 (10.30–126.48)	< 0.001
Sex Female Male	1 1.91 (1.12–3.27)	0.023
Congenital malformation No Yes	1.00 27.85 (11.34–68.41)	< 0.001
Problems during delivery No Yes	1.00 5.17 (2.41–11.07)	< 0.001
DWTD $\leq 6.0$ > 6.0 and $\leq 45.3$ > 45.3 and $\leq 370.2$ > 370.2 and $\leq 10810.9$	1.00 1.46 (0.67–3.18) 2.82 (1.32–6.03) 1.47 (0.67–3.19)	0.215

is also possible that some of these women might have moved during pregnancy. We do not have data on mobility, but according to the information on maternal occupation during pregnancy, 58% of mothers of controls, 63% of mothers of neonatal deaths, and 51% of mothers of fetal deaths were housewives, unemployed, retired, or students, which means they might have stayed at home most of the time during pregnancy. Even for those who worked, it is likely that they might have stayed mostly at home during the final months of pregnancy, thus enhancing their exposure to the local traffic-related air pollutants. It should also be noted that São Paulo has a mild climate, and people keep windows open throughout the year. Therefore, a significant portion of outdoor pollution from traffic exhaust penetrates indoors.

In addition, we estimated the DWTD measures without knowledge of the case-control status, so any errors in the DWTD measurement are likely to be nondifferential, resulting in an underestimation of the risk associated with traffic-related air pollution.

Other sources of imprecision in our estimate of exposure include the lack of information on wind patterns and local topography, which may alter the spatial distribution of air pollutants (e.g., traffic-related air pollution may be different for those living upwind or downwind) and the different types of vehicles passing in a given street, such as gasoline-, ethanol-, or diesel-fueled vehicles, which have different emission profiles. These are all potential sources of variability in our estimates of exposure, but again, one does not expect them to be differentially distributed between cases and controls.

Despite these limitations, this indicator of exposure to traffic-related pollutants has been applied in different studies (Bayer-Oglesby et al. 2006; Ciccone et al. 1998; Kramer et al. 2000; Pearson et al. 2000; Wilhelm and Ritz 2003) because it is relatively easy to interpret and provides an account of the importance of living close to busy roads, and therefore exposure to vehicular emissions. Other studies have shown that concentrations of pollutants near roadways are well correlated with traffic counts (Kinney et al. 2000; Lena et al. 2002), so the DWTD model can be used as an indicator of population exposure to urban air pollutants.

The original study was carefully designed to avoid selection bias, because all cases of perinatal deaths during the study period were included and controls were randomly selected among survivors of the neonatal period identified from live-born birth certificates. We compared the data on sex, birth weight, and gestational age available in the birth certificates for those whom we did not include in the study (~16% of cases and controls) with the included infants and found no statistical significant differences among them, confirming that there was no selection bias.

We used extensive collected data in this study, including variables related to socioeconomic conditions, maternal reproductive history and risk factors, and outcome of previous and present pregnancy and delivery. We assessed important risk factors for perinatal deaths and identified and controlled for potential confounding variables. Nevertheless, we cannot rule out some residual confounding by other factors for which we did not have data, such as meteorologic aspects. Background air pollution is correlated with temperature and/ or other meteorologic parameters, and its effect on health might vary according to season. However, we could not examine this in detail because we selected the events of this study from a period of only 6 months.

Few other studies have investigated the association between deaths in the perinatal period and air pollution (Hajat et al. 2007; Lin et al. 2004; Nishioka et al. 2000; Pereira et al. 1998). Most studies have used a timeseries approach and therefore examined the effect of air pollution on day(s) before the event (temporal scale). In addition, they relied on air quality monitoring sites for assessment of exposure to urban air pollution. In general, they all found statistically significant associations between daily counts of deaths (late fetal, neonatal, infant) and the average concentration of pollutants in the atmosphere. Pereira et al. (1998), in a study conducted in São Paulo, also showed that carboxyhemoglobin levels of blood collected from the umbilical cord of nonsmoking mothers were correlated with environmental carbon monoxide. These studies, although not directly comparable with the present study, which we based on spatial comparisons, suggest that fetuses and newborns may suffer from the consequences of a contaminated environment.

Potential biologic mechanisms behind this association may involve an effect of air pollution on the placenta, embryo, maternal immunologic system, ovarian-hypothalamic axis, and the induction of IUGR, which can lead to a more vulnerable fetus. It is possible that certain toxics emitted in motor vehicle exhaust, such as polycyclic aromatic hydrocarbons (PAHs), may be responsible for these adverse birth outcomes. Exposure to PAHs during pregnancy has been demonstrated to alter levels of serum progesterone and estrogens, decrease survival of rats, and relate to endocrine disorders in rodents (Takeda et al. 2004). Air particles may be associated with changes in plasma viscosity (Peters et al. 1997), inflammatory process (Dubowsky et al. 2006), and elevation in blood pressure in susceptible populations (Brook 2005). Carbon monoxide easily crosses the placental barrier, coming into contact with the fetus, leading to a rapid accumulation of carboxyhemoglobin, with subsequent reduction in oxygen transportation by blood (Aleixo Neto 1990).

The results of the present study suggest that the early neonatal component of perinatal mortality may be associated with mothers' exposure to air pollution from traffic near their homes. Although we could not provide strong evidence of such association, the consistent literature and the biologic plausibility indicate that motor vehicle exhaust exposures may be important for this outcome.

#### REFERENCES

- Aleixo Neto A. 1990. Smoking effects on pregnancy [in Portuguese]. Rev Saude Publica 24:420–424.
- Almeida MF, Alencar GP, Novaes HMD, França Júnior I, Siqueira AAF, Campbell OMR, et al. 2007. Risk-factors for antepartum fetal deaths in the city of São Paulo, Brazil. Rev Saude Publica 41:35–43.
- Almeida MF, Alencar GP, Novaes HMD, Ortiz LP. 2006. Information systems and perinatal mortality: concepts and conditions for the utilization of data in epidemiological studies. Rev Bras Epidemiol 9:56–68.
- Almeida MF, Ortiz LP, Novaes HMD, Yamasaki LC, Alencar GP, Schoeps D. 2004. Diferenciais intra-urbanos de mortalidade perinatal em São Paulo (MSP) [in Portuguese]. In: Anais do VI Congresso Brasileiro de Epidemiologia, 19-24 June 2004, Recife, Brasil. Rio de Janeiro:ABRASCO (Brazilian Association of Collective Health), CD-ROM.
- Bayer-Oglesby L, Schindler C, Arx MEH, Braun-Fahrlander C, Keidel D, Rapp R, et al. 2006. Living near main streets and respiratory symptoms in adults. Am J Epidemiol 164:1190–1198.
- Brook RD. 2005. You are what you breathe: evidence linking air pollution and blood pressure. Curr Hypertens Rep 7:427–434.
- Ciccone G, Forastiere F, Agabiti N, Biggeri A, Bisanti L, Chellini E, et al. 1998. Road traffic and adverse respiratory effects in children. Occup Environ Med 55:771–778.
- Conde-Agudelo A, Belizan JM, Diaz-Rossello JL. 2000. Epidemiology of fetal death in Latin America. Acta Obstet Gynecol Scand 79:371–378.
- Dubowsky SD, Suh H, Schwartz J, Coull BA, Gold DR. 2006. Diabetes, obesity and hypertension may enhance associations between air pollution and markers of systemic inflammation. Environ Health Perspect 114:992–998.
- Hajat S, Armstrong B, Wilkinson P, Busby A, Dolk H. 2007. Outdoor air pollution and infant mortality: analysis of daily time-series data in 10 English cities. J Epidemiol Community Health 61:719–722.
- Jackson DJ, Lang JM, Ganiats TG. 1999. Epidemiological issues in perinatal outcomes research. Paediatr Perinat Epidemiol 13:392–404.
- Kinney PL, Aggarwal M, Northridge ME, Janssen N, Shepard P. 2000. Airborne concentrations of PM<sub>2.5</sub> and diesel exhaust particles on Harlem sidewalks: a community-based pilot study. Environ Health Perspect 108:213–218.
- Kramer MS, Liu S, Luo Z, Yuan H, Platt RW, Joseph KS. 2002. Analysis of perinatal mortality and its components: time for change? Am J Epidemiol 156:493–497.
- Kramer U, Koch T, Ranft U, Ring J, Behrendt H. 2000. Trafficrelated air pollution is associated with atopy in children living in urban areas. Epidemiology 11:64–70.
- Langholz B, Ebi K, Thomas DC, Peters JM, London SJ. 2002. Traffic density and the risk of childhood leukemia in a Los Angeles case-control study. Ann Epidemiol 12:482–487.
- Lena TS, Ochieng V, Carter M, Holguin-Veras J, Kinney PL. 2002. Elemental carbon and  $\mathrm{PM}_{25}$  levels in an urban community heavily impacted by truck traffic. Environ Health Perspect 110:1009–1015.
- Lin CA, Pereira LAA, Nishioka DC, Conceição GMS, Braga ALF, Saldiva PHN. 2004. Air pollution and neonatal deaths in São Paulo, Brazil. Braz J Med Biol Res 37:765–770.
- Lipfert FW, Zhang J, Wyzga RE. 2000. Infant mortality and air pollution: a comprehensive analysis of U.S. data for 1990. J Air Waste Manag Assoc 50:1350–1366.
- Nishioka DC, Coura FLD, Pereira LAA, Conceicao GMS. 2000. Estudo dos efeitos da poluição atmosferica na mortalidade neonatal e fetal na cidade de São Paulo, Brasil [in Portuguesel. Rev Med (São Paulo) 79:81-89.
- Pearson RL, Wachtel H, Ebi KL. 2000. Distance-weighted traffic

- density in proximity to a home is a risk factor for leukemia and other childhood cancers. J Air Waste Manag Assoc 50:175–180.
- Pereira LA, Loomis D, Conceição GM, Braga AL, Arcas RM, Kishi HS, et al. 1998. Association between air pollution and intrauterine mortality in São Paulo, Brazil. Environ Health Perspect 106:325–329.
- Peters A, Doring A, Wichman HE, Koening W. 1997. Increased plasma viscosity during an air pollution episode: a link to mortality? Lancet 349:1582–1587.
- Ritz B, Yu F. 1999. The effect of ambient carbon monoxide on low birth weight among children born in southern California between 1989 and 1993. Environ Health Perspect 107:17–25.
- Ritz B, Yu F, Chapa G, Fruin S. 2000. Effect of air pollution on preterm birth among children born in southern California between 1989 and 1993. Epidemiology 11:502–511.
- Schoeps D, Almeida MF, Alencar GP, França Júnior I, Novaes HMD, Siqueira AAF, et al. 2007. Risk factors for early neonatal mortality [in Portuguese]. Rev Saude Publica 419:1013–1022.
- Takeda K, Tsukue N, Yoshida S. 2004. Endocrine-disrupting activity of chemicals in diesel exhaust and diesel exhaust particles. Environ Sci 11:33–45.
- Vermelho LL, Costa AJL, Kale PL. 2002. Indicadores de saúde [in Portuguese]. In: Epidemiologia (Medronho RA, Carvalho DM, Block KV, Luiz RR, Werneck GL, eds). São Paulo:Editora Atheneu, 33–55.
- Wilhelm M, Ritz B. 2003. Residential proximity to traffic and adverse birth outcomes in Los Angeles County, California, 1994–1996. Environ Health Perspect 111:207–216.