# Atlantic Greater Amberjack Abundance Indices From Commercial Handline and Recreational Charter, Private, and Headboat Fisheries through fishing year 1997 

## By

N. J. Cummings, S. C. Turner, D. B. McClellan, and C. M. Legault

U.S. Department of Commerce

National Oceanic and Atmospheric Administration National Marine Fisheries Service
Southeast Fisheries Science Center
75 Virginia Beach Drive
Miami, Florida 33149 USA
July 1999
Sustainable Fisheries Division Contribution No. SFD-98/99-62

## Introduction

Parrack (1993a,b), Cummings and McClellan (1996, 1997) and Cummings (1998) presented catch per unit of effort (CPUE) statistics for the Atlantic greater amberjack, Seriola dumerili, stock. This report updates the CPUE analyses of earlier studies through the 1997 fishing year, incorporates new analysis procedures, and presents a new index of stock abundance developed from tagging data observations. Virtual Population Analysis (VPA) techniques have been used to evaluate the condition of this stock. Information external to the input fishery catch at age data are required for tuning (calibrating) the VPA estimates of stocksize and mortality trends. Standardized CPUE observations or abundance trends (independent and fishery dependent) have been shown effective in the standardization or calibration of VPA's. Therefore, providing updated CPUE trends and implementing improvements in analysis methods is important in the analysis of the Atlantic greater amberjack stock status.

## Materials and Methods

As in the previous 1996 evaluation of stock status, three separate sources of CPUE data existed for estimating temporal trends in abundance of the Atlantic greater amberjack stock. These included observations of recreational catch per angler and catch per angler hour from the Marine Recreational Fisheries Sampling Survey (MRFSS) data base of intercepted angler fishing trips from the shore, charter, and private vessel modes in the Atlantic ocean. In addition, for headboats operating from North Carolina through the Florida Keys, the National Marine Fisheries Service (NMFS), Beaufort Laboratory conducts a survey. This database is an archive of captain reports of landed catch in number per trip, and the number of anglers per trip collected since 1972 in the Carolinas, since 1976 in North Carolina through Daytona, Florida, and since 1977 in the Florida Keys. The third source of CPUE data was the NMFS Atlantic Snapper - Grouper and the Gulf of Mexico Reeffish Logbook database. This combined set of commercial data referred to as the "Logbook" data in this report, is an archival landings reports made by all commercial fishermen that have federal permits to fish in the Atlantic Snapper Grouper and also in the Gulf of Mexico fishery. The logbook survey was begun in January 1992 in the Atlantic in response to the mandate of Amendment 4 of the South Atlantic Snapper-Grouper Fishery Management Plan to collect information pertaining to the Atlantic reeffish fisheries. The logbook survey was begun in April 1990 in the Gulf of Mexico as part of the Gulf of Mexico Fishery Management Plan Data collection requirements (McClellan and Bohnsack 1991, Anon. 1987). During the early years only a sample of the permit holders (drawn randomly at a $25 \%$ rate) were required to report. Reporting became mandatory in 1993 in the Atlantic (Harris et al. 1994) and in 1991 for Florida vessels fishing in the Gulf of Mexico. It was necessary to select landings reports from the Gulf survey as well as the Atlantic since some permit holders who fish the Atlantic stock, especially in the Florida Keys region, may report to both systems. Reports were identified as duplicates and removed prior to analyses. A new abundance index for the Atlantic greater amberjack was developed from tagging observations from releases made by scientific and recreational anglers from the early 1960's through 1998.

## MRFSS CPUE Analysis Procedures

The MRFSS based CPUE indices included intercepts from recreational fishing trips having a positive catch of greater amberjack (whether targeting or not) and also intercepts with zero catches of greater amberjack that indicated they were targeting any of the four highly sought amberjack species (greater, banded rudder, almaco and lesser). As with previous MRFSS indices, observations from Monroe county Florida were considered in the analysis. Information recorded for each observation included the year, month, state and county of intercept, and whether targeting greater amberjack. In addition, a variable referred to as bag limit stanza corresponding to un-regulated or 3 fish per person was defined and each observation coded for this variable. Fishing year was calculated, as May 1-April 30, for each observation from the month and year of intercept. These five independent variables (fishing year, month, state-county, target id, and bag limit stanza) were used to adjust the fishing year CPUE indices for the estimated effects of these variables on catch rates. General linear modeling statistical methods (Robson 1966) as previously used in Parrack (1993a,b) and Cummings and McClellan (1996,1997), were used in calculating the abundance indices for the MRFSS observations of positive catch.

The Lo method (Lo et al. 1992) was explored to develop an abundance index for Atlantic greater amberjack from the combined set of charter and private CPUE observations of zero catch in addition to the positive catches. The Lo method assumes a delta-lognormal error structure in the raw data and generates an estimate of CPUE from the product of two models. One model estimates frequency (the proportion of intercepts in which greater amberjack was caught) assuming a binomial error distribution and the second model estimates the mean density of greater amberjack within positive trips assuming a lognormal error distribution.

Previous analyses of the MRFSS CPUE observations indicated the amount of observed variability in the raw catch per hour fished data was high, especially in the private boat $f$ fishery, therefore this year's analyses were made separately for the charter and the private boat data as well as the combined set of charter and private data. Few intercepts were from the shore-based fishery therefore these were not considered in the CPUE abundance analyses.

## Commercial Logbook CPUE Analysis Procedures

The NMFS Atlantic Snapper - Grouper and the Gulf of Mexico Logbooks were also used to develop an index of abundance for greater amberjack in the Atlantic. This combined set of commercial data referred to as 'logbook' below contains landings reports from commercial fishermen that have federal permits to fish in the Atlantic Snapper Grouper and the Gulf of Mexico reef fish fisheries. The logbook survey was initiated in the early 1990 's. During the early years only a sample of the permit holders (drawn randomly) were required to report, however, reporting became mandatory in about 1994 in the Atlantic. It was necessary to select landings reports from the Gulf survey as well as the Atlantic since some permit holders who fish the Atlantic stock, especially in the

Florida Keys region, could report to either or both systems. Duplicate reports in the two databases were identified and removed prior to analysis.

Logbook records existed from 1991 to 1998 by trip; no within trip information was available (if a trip involved fishing in multiple locations and/or on multiple days, the catch (i.e., landings) and effort data could not be identified for those separate events). Recorded data included landed weight by species (converted to whole weight as necessary), information on the primary fishing gear used and location of the effort and/or catch by $1^{\circ}$ latitude and longitude square. A preliminary definition of fishing year from April through March of the following year was used for initial examination of the data. A variety of effort measures were available depending on the fishing gear and at a minimum included days at sea and hours fished.

The number of trips, proportion of trips with greater amberjack and the proportion of greater amberjack in the total landings were summarized by fishing gear and fishing year to determine which gear(s) would be used for developing indices of abundance.

Multiple measures of fishing effort were evaluated to select one for use in developing indices of abundance; these reviews included both examination of frequency distributions of individual effort components (hours fished, sea days, hooks per line, lines, etc) as well as total annual effort measures. The data indicated that apparently extreme values occurred for some effort measures such as hooks per line. In an attempt to reduce the influence of extreme observations on catch rates, a policy of eliminating the upper and lower $1 \%$ of the observations for each effort measure was established.

Regression tree analysis (Venables and Ripley 1997) was used for preliminary data exploration primarily for the purpose of defining seasons and regions for use as explanatory variables. Separate analyses were conducted using as the dependent variable (1) proportion of positive caches and (2) catch per hook on trips on which caught greater amberjack (successful trips). Independent variables examined included fishing year, month or season ( 3 month seasons with January-March, etc.), latitude and longitude (based on the southeast corner of each one degree square) or region (defined from the results of earlier regression tree analyses).

General linear model (GLM) analyses were used to select the effort measure for trips on which greater amberjack were caught. The dependent variable was the log of the landed weight, and a lognormal error distribution was assumed. Multiple analyses were made each including a different measure of effort as an explanatory variable along with fishing year, season and region. The coefficient of determination $\left(\mathrm{R}^{2}\right)$ and the probability of the $F$ statistic for the effort effect were used to select the measure of effort for use in calculating catch rates.

Indices of abundance were derived using the Lo method (Lo et al 1992) which assumes a delta lognormal error structure, employs separate GLM analyses of the proportions of positive trips and the catch rates on trips which caught amberjack, and combines the results of the separate analyses to derive the index. A binomial error assumption was
used for the proportion positive analyses and a lognormal error assumption was used for the analyses of positive trips. Fishing year, season and region were included in the analyses along with the season $*$ region interaction if significant. In addition random effects year (fishing year) interactions with season and region were included where data were sufficient and if a significant effect was indicated. If possible three-way random effects year interactions (fishing year*season*region) were included; if the data would not support a three-way interaction one or more two-way random effects year interactions were included.

## NMFS Headboat CPUE Analysis Procedures

Data were available from the NMFS headboat survey from 1974-1998. The data consists of captains' reports of the number and estimated weight of landed catch by species, the number of anglers on board the boat, the type of trip. Trip type included information on the length of the trip: half day, $3 / 4$ day, full day, multi-day and the time of the trip: day or night for full as well as early or late in the day or night for partial day trips. The data usually included information on the geographical location of the catch by 10' square.

To eliminate effort clearly unrelated to greater amberjack, data were restricted to 10 , squares in which greater amberjack (Seriola dumerili) or other members of the genus were ever recorded. To ensure that catches came from the stock under consideration, it was assumed that greater amberjack from the other side of the Gulf Stream were from a different population and fishing effort in the Bahamas and similar areas were excluded.

To examine possible positive or negative associations of greater amberjack catch rates with other fishes; catches of other taxa were tabulated along with those of greater amberjack. Both relatively desegregated and highly aggregated taxonomic groupings were investigated. The desegregated classification included single species groups (such as black sea bass), genera (such as Myctoperca), families and groupings of similar species (such as large pelagics, or inshore Lujanids or offshore Lujanids); a total of 25 taxa were defined. Three classifications of aggregated taxa were defined; they were inshore species, offshore bottom oriented species and offshore pelagic species. Taxa were classified by their proportion of the total number of fish reported from a trip. Four levels of classification were used for preliminary analyses and two levels for subsequent analyses. The four levels were 0 (absent), and 1-3 indicating increasing proportions. The divisions between levels were taxa specific, because some taxa were frequently observed at all proportions while others were observed frequently up to some intermediate percentage of the catch. In final analyses the classifications for the aggregated taxa were summarized into 0 (absent or low) or 1 (moderate to high).

Preliminary analyses indicated that there was a negative relationship between greater amberjack catch and the number of anglers on the boat, and therefore the basic unit of effort was defined as the trip and catch rates were defined as catch per trip.
Regression tree analysis was used to explore the data set for possible aggregation into seasons and regions and to investigate the possible influences of the various taxonomic groupings.

Indices of abundance were derived using the Lo method (Lo et al. 1992), which assumes a delta lognormal error structure, employs separate GLM analyses of the proportions of positive trips and of the catch rates on trips which caught amberjack, and combines the results of the separate analyses to derive the index. A binomial error assumption was used for the proportion positive analyses and a lognormal error assumption was used for the analyses of positive trips. Fishing year, season and region were included in the analyses and in most cases information on catches of other species on the trip was also included. For the season, region and taxonomic factors two-way and three-way interactions were investigated to the extent permitted by the distribution of the data and the analysis. In addition random effects year (fishing year) interactions with the other factors were included if possible where data were sufficient and if a significant effect was indicated. The highest order random effect interactions, which could be estimated, were included.

## Tagging Data Analyses

Tag and recapture observations from 1960 through 1998 were available. McClellan and Cummings (1997) provided a description of the data for the years 1960-1994. The more recent data (1994-1998) were extracted from the NMFS, Cooperative Tagging Center tagging database (D. Rosenthal, pers. comm.) and the observations combined. For each observation the date of tagging and date of recapture was recorded thus allowing the corresponding fishing year period, May - April to be calculated. These data were used to generate a fishery independent estimates of the relative loss of the Atlantic greater amberjack stock. Several estimates were developed and the "best" estimate of relative loss used to create an index of relative abundance for the Atlantic greater amberjack stock.

Porch's (1998) method of estimating annual mortality rates from tagging data is based upon the probability of a tagged fish being recaptured assuming $M$ and reporting rates are constant over time:

$$
p(\text { recapture })=\frac{F_{w} e^{-\sum_{i=\alpha}^{w-1} Z_{i} \Delta_{i}-Z_{w}\left(t_{w}-d_{w}\right)}}{\sum_{i=\alpha}^{I} \frac{F_{i}}{Z_{i}}\left(1-e^{-Z_{i} \Delta_{i}}\right) e^{-\sum_{j=\alpha}^{i-1} Z_{j} \Delta_{j}}}
$$

## Where

$\mathrm{t}_{\alpha}=$ time (date) animal was tagged
$\mathrm{t}_{\mathrm{w}}=$ time (date) animal was recaptured
$\alpha=$ discrete time interval during which the animal was tagged
$\mathrm{w}=\quad$ discrete time interval during which the animal was recaptured
$d_{I}=$ time (date) at start of i'th interval, except $d_{\alpha}=t_{\alpha}$
$\Delta_{\mathrm{I}}=$ time spent in each interval $\left(\mathrm{d}_{\mathrm{i}+1}-\mathrm{d}_{\mathrm{i}}\right)$
$\mathrm{I}=$ last interval for which data are available
$F=\quad$ fishing mortality rate
$\mathrm{Z}=\quad$ all sources of tag loss, includes natural and fishing mortality, tag shedding, non-reported recaptures, emigration, etc.

This method in practice cannot distinguish between sources of tag loss. The method estimates the total tag losses $\left(Z^{*}\right)$ for each time period. These $Z^{*}$ values contain the effects of tag shedding, emigration from the study area, non reported recaptures, etc in addition to natural and fishing mortality rates. If all factors other than annual fishing mortality rates can be assumed constant during the study period, then the relative change in the annual $Z^{*}$ estimates from this method can be used as an index of the relative changes of annual F in the fishery. The loss rate reported here is formed by dividing each annual $Z^{*}$ estimate by the average of all estimates such that the index is centered about 1.0. This standardization allows for easy comparison of values estimated under different assumptions or from different subsets of the database.

The program TAP2 (C. Porch; NMFS, SEFSC, pers. comm.) allows for two methods of estimating these time period specific $Z^{*}$ rates. Each $Z^{*}$ can be estimated independently or through a Bayesian random walk. In the case of the random walk, the amount of change allowed in $Z^{*}$ between successive time intervals is determined by a user supplied value $(\sigma)$ which causes changes larger than $e^{\sigma}$ to be unlikely. The discrete time intervals used in the program were fishing years from May of one year to April of the next calendar year, with the exception of the first and last time periods. The first time period ranged from January 1960 to April 1961 and the last time period ranged from May 1997 to December 1998. Not enough data was available in these periods to use only single fishing years.

Four estimation procedures were considered for the time series; estimating annual loss rates for all years independently and fixing the random walk sigma at the values of 0.1 , 0.3 and 0.5 . The random walk sigmas constrain the model such that interannual variation in $Z^{*}$ 's are permitted to vary to lower degrees with lower sigmas. These estimates used only fish, which had at least one day at large between time of tagging and time of recapture. A number of sensitivity analyses were also performed, for example, estimating $Z^{*}$ values for periods of two years, changing the minimum number of days at large, or allowing changes in tag-recapture rates. The resulting series of $Z^{*}$ estimates were rescaled by the average for that time series and case to produce the relative loss values.

The number of greater amberjack caught each year in the private angler sector of the recreational fishery (Cummings and McClellan 1999) was used to transform the loss estimates into relative abundance estimates because the majority of tags were from this sector. The catches were either kept and released dead fish (A+B1) or all catch, including fish released alive, $(\mathrm{A}+\mathrm{B} 1+\mathrm{B} 2)$. These annual catches were used in the catch equation by assuming a constant level for the natural mortality rate ( M ) and a constant level for nonfishing related tag loss such that the $\mathrm{Z}^{*}$ could be separated into its components of $\mathrm{M}, \mathrm{F}$ and non-fishing related tag loss. The only remaining unknown in the catch equation is then the population abundance in numbers $\left[\mathrm{N}=\mathrm{CZ} /\left(\mathrm{F}\left(1-\mathrm{e}^{-\mathrm{Z}}\right)\right)\right]$. Each time series of values was re-scaled to average 1.0 by dividing annual values by the average of the time series.

These relative abundance values could be used as a tuning index for age structured assessment models, such as virtual population analysis.

## Greater Amberjack CPUE Analysis Results

## MRFSS CPUE Analyses

Summary statistics on the number of MRFSS intercepts and the nominal CPUE values of catch per angler hour (CPH) fished available for the charter and private boat fisheries for Atlantic greater amberjack are given in Tables 1a - 1e. Fishing year 1998/99 (May 1 1998 - April 30 1999) was incomplete as data were only available through December 1998 thus only fishing years 1981/82 - 1997/98 were used in developing abundance trends. The majority of the intercepts were from fishing years 1986 through 1995/96. Most anglers catching one or more of the amberjack species were intercepted on the East Coast of Florida ( $\mathrm{n}=767$ ) followed by North Carolina ( $\mathrm{n}=565$ ), and then off Monroe county, Florida ( $\mathrm{n}=319$ ) (Table 1b). Recreational fishing trips from Georgia and South Carolina contributed $4 \%$ and $12 \%$ respectively to the total database of charter and private boat intercepts. Intercepts from the charterboat recreational mode made up $69 \%$ of the data and $31 \%$ of the intercepts were from the private boat fishery. The nominal CPUE values suggest that the private boat fishery CPUE (CPH) was higher (across all years) than for the charterboat fishery for Atlantic greater amberjack (Table 1c). The exception to this was in Georgia. The monthly distribution of MRFSS intercept observations (Table 1 d ) indicates that amberjack were encountered during most months of most fishing years and also indicates temporal differences in the nominal CPUE values. Effort measures other than anglers or hours fished were not recorded thus catch per angler angler hour was calculated and used. The plotted CPH data showed skewed distributions suggesting a transformation was appropriate (Figure 1). A log transformation of the raw CPH fished observations resulted in a more symmetrical distribution of the charter and private boat fishery intercepts (Figure 1).

Relative abundance trends, referred to as standardized abundance trends, in this paper, were calculated using the GLM approach for the charter and private fisheries separately, the combined dataset of charter and private boat intercepts, and for the charter and private intercepts from Florida only. Summary results pertaining to each fit are included in Table 2. For each fit, the $y$-variate, was CPH fished and auxiliary information included in each fit as (X) variables to adjust CPH trends were the fishing year, month of intercept, state-county id, whether targeting greater amberjack or not, and bag limit stanza. In the case of the charter and private boat combined analysis, fishery type was also included as an independent variable in the mode.

As in previous year's analyses of the MRFSS CPUE observations for Atlantic greater amberjack, the analyses indicated the amount of un-explained variability in the data from these models was high. The proportional amount of the variance explained by the fits was not high, ranging from 21 to $44 \%$. The best fit in terms of $\mathrm{R}^{2}$ was obtained using the charterboat fishery observations, all states combined. For the majority of the fits, the inclusion of bag limit stanza ( 3 or none) or a term for whether the angler was targeting greater amberjack or was not targeting this species not was not significant in the model.

The calculated standardized abundance trends from the GLM fits for the charter and private boat fisheries off Florida were not in total agreement (Figure 2). The estimated fishing year abundance pattern from the charterboat fishery showed variability from 1981 through 1992 and stable abundance thereafter (Figure 2). The private boat fishery abundance pattern suggested increased abundance from 1990-1993 and followed by a decline through 1996 (Figure 2). The variability in these estimates was high.

Estimates of the number of fishing trips from these fisheries are available through the MRFSS survey; these were summarized by fishing year to obtain some feeling for the amount of total recreational effort expended by each fishery group (Figure 3). Estimated charterboat effort (number of fishing trips) was stable from about 1990 through 1997 while the estimated private boat effort (number of fishing trips) showed large increases between 1985 and 1988, a decline of about $30 \%$ in 1989, varied without trend through 1995, and declined again in 1996.

The GLM estimated abundance trends from the "all states combined fit" charter and private boat fisheries are shown by fishing year with the estimated total effort for each fishery (Figure 4). The all states combined abundance estimation results were used in subsequent analyses of the stock evaluation as the trends did not differ when only the Florida samples were used; using the observations from all states better represented the range of the Atlantic greater amberjack stock, increased the sample size and the $\mathrm{R}^{2}$ values were higher.

The Lo method was used to develop abundance trends of the Atlantic greater amberjack stock using the combined data from the charter and private boat fisheries, all states combined. For this analysis observations of positive catches of Atlantic greater amberjack and also observations of zero catches of greater amberjack that occurred when an angler was intercepted and reported catching one of the other commonly sought after amberjacks (e.g., Alamo, banded rudder, lesser) but did not catch greater amberjack. were used. These anglers could have reported targeting greater amberjack (or any of the other jacks listed here). The model included the same independent variables as used in the previous GLM runs (fishing year, month, state-county, fishery, bag_limit stanza, and whether targeting greater amberjack or not). The Lo method produced similar results, as did the GLM procedure for the charter and private boat data combined, however the variance estimates of the fishing year parameter were much higher. All of the independent variables were important in explaining CPUE except for bag limit as the case found using the GLM. In this application since very few observations of zero greater amberjack catches existed in this dataset, the analysis set was very similar to that of the input data used in the GLM. The proportion of positive catches was very high and similar results would be expected since this portion of the data dominated in the calculation of the expected CPUE. For this reason and also because the variability was lower for the fishing year estimates estimate via GLM, the GLM abundance estimates results were used in further analyses related to the status of the Atlantic greater amberjack stock. Future evaluation of this approach using the total intercept database may be warranted.

## Commercial Logbook CPUE Analyses

The number of trips, proportion with greater amberjack and the pounds per trip by fishing gear and fishing year (Tables 3-5), were reviewed to select which observations would be used for developing indices of abundance. Data from gears from which little data were available [cast net, gill net, trap, vertical longline (buoy) or other] were not included in those tables and were not considered for analysis. It was decided to develop an index from handline data because there were substantial numbers of trips recorded and greater amberjack generally represented a substantial fraction of the total landings (roughly $15-25 \%$ across all trips). Longlines were rejected because greater amberjack generally represented less than $1 \%$ of the total landings. Powerheads are a special kind of spear fishing gear. Development of a combined spear and powerhead index was initially considered, because greater amberjack were reported on a substantial fraction of the powerhead and spear trips (generally over $35 \%$ and $20 \%$, respectively) and represented substantial fractions of the total yield (roughly $20-40 \%$ ). It was thought necessary to distinguish the two gears in developing an index of abundance, because of probable differences in efficiency. It was noted that the two gears would not be distinguished in the earliest years of the data set and the consistency in the number of trips per year suggested that they were reliably distinguished only from 1995-1998. It was decided that an index would not be developed from these gears, because of the small number of years available. It was decided not to develop an index from the troll data, because of the relatively small proportion of amberjack in the total yield reported taken by trolling.

Regression tree analysis of handline catch rates and proportions of positive trips were used to define seasons and regions. Examination of proportion positive and catch rates by month (Appendices 1 and 2) indicated that lower positive catch rates and lower proportions positive occurred in the more northern latitudes. It was noted that latitude 29 had high positive catch rates in most months and high proportions positive. May and March (months 5 and 15) had high positive catch rates and high proportions positive, while April had lower statistics. Differences were noted in the relative importance of latitudes between the proportion positive and the positive catch rate analyses. In aggregating latitudes into regions, greater weight was placed on the proportion positive results (i.e., estimates of the probability of catching amberjack on a trip), because often proportion positive has a stronger influence on abundance index patterns than the positive catch rates. Therefore four regions ( $24-26 \mathrm{~N}, 27-28 \mathrm{~N}, 29-33 \mathrm{~N}$ and $34-35 \mathrm{~N}$ ) were defined. The analytical results with respect to month were more consistent between the proportion positive and positive catch rate analyses. Because May and March were the months of the highest positive catch rates and proportion positives, they were combined with April into one season and three other 3 month seasons were defined.

Units of effort considered for calculating catch rates from the handline fishery included days at sea, hours fished during the entire trip, number of lines fished per day, number of hooks fished per day, total number of line-days (lines per day * days at sea), total number of line-hours, total number of hook-days and total number of hook-hours. GLM analyses of greater amberjack landings per trip with the various effort measures (Table 6) had low
$\mathrm{R}^{2}(10-12 \%)$. Hooks per day had the lowest probability for the F statistic and highest $\mathrm{R}^{2}$ and was selected as the unit of effort.

## All Vessels

Fishing year, season and region were investigated as effects in analyses with data from all handline vessels. The season*region interactions were statistically significant in both the proportion positives and the yield per hook analyses and the random effects three way interaction was estimable and included (Tables 7 and 8). The estimated index of abundance (Table 9 and Figure 6) had a high fishing year 1991 value with a coefficient of variation of 0.51 which was roughly twice that of subsequent years. After 1991 the index shows an increase to 1995 and a decline thereafter. The confidence intervals (two standard errors above and below the index) suggest that differences in the annual index values were not highly significant.

## Analysis of Restricted Data Sets

There was concern that some handline vessels might not target amberjack or species with which amberjack might be associated. If the fraction of vessels not targeting amberjack or associated species changed over time then an index of abundance from the entire fishery might be influenced by changes not related to abundance. Therefore additional indices were derived using data from selected vessels which might have been targeting amberjack or associated species.

Restricted data sets were developed by limiting the observations to those vessels with either high proportions of trips with greater amberjack or high greater amberjack catch rates on successful trips. Additionally, to be included, vessels had to have reported at least 10 handline trips in each of at least 5 years in a region.
$90^{\text {th }}-99^{\text {th }}$ Percentile Vessels
Results of the analyses of data from vessels in the top $10 \%$ of vessels in proportion positive or catch rate on successful trips are presented in Tables 10 and 11. Data from the $27-28 \mathrm{~N}$ region were excluded to permit a more balanced distribution of data across fishing years, seasons and regions. The three-way random effects interaction (i.e., fishing year*season*region) appeared to be influential and was included in both models. The standardized index (Table 12 and Figure 7) indicated high catch rates of Atlantic greater amberjack in the 1994 and 1995 fishing years with catch rates in 1992-1993 and 19961997 at similar levels and the lowest catch rate in 1998.

## $75^{\text {th }}-99^{\text {th }}$ Percentile Vessels

Results of analyses of data from vessels in the top $25 \%$ of vessels in proportion positive or catch rate on successful trips are presented in Tables 13 and 14. Data from 1991 and from the $27-28 \mathrm{~N}$ region were excluded to permit a more balanced distribution of observations across fishing years, seasons and regions. The season*region interaction was significant in both models, and the three-way random effects interaction appeared to be influential in both. The standardized index (Table 15 and Figure 8) indicated high catch rates in 1992, 1994 and 1995 with catch rates in 1993 and 1996-1997 at similar levels and the lowest catch rate in 1998.

Comparison between the three indices (Figure 9) showed that the indices had similar patterns. The analyses using all handline vessels and using the vessels in the $90^{\text {th }}-99^{\text {th }}$ percentile included 1991; the divergence of those indices in 1991 may reflect sparse data distribution in that year. Because of the possibility that effects not related to greater amberjack abundance may have influenced the all vessel index trend, the restricted data set trends were preferred. Because the two restricted sets had generally similar patterns, the $75^{\text {th }}-99^{\text {th }}$ percentile index was selected for use in subsequent assessments, because this index was based on more data and the data were more evenly distributed because of the elimination of 1991 year data and the observations from 27-28N region.

## NMFS Headboat CPUE Analyses Results

Regression tree analysis of headboat catch (number of fish) per angler was used to gain insight into effects due to trip types, associated taxa as well as seasonal and geographic patterns in the data. Initial analyses included trip type and desegregated taxa (Appendix 3 ) and trip type with aggregated taxonomic information (Appendix 4). The analyses with trip type consistently indicated substantially higher catch rates for full day trips than for partial day trips. Because the partial day catch rates were so low, it was concluded that they might not provide much information on greater amberjack abundance patterns and they were excluded from subsequent analyses. Additional analysis was conducted using only full day trips with the aggregated taxonomic information (Appendix 5). Common components of all three of these analyses were that latitude 30 or 30 and 31 had higher catch rates than the other latitudes and that substantially higher average catch rates occurred in 1981-1982 or 1980-1982. When data splits were based on month, May was almost always classified with the higher catch rates and April was often in that grouping. It was also noted that in these analyses the first low catch rate node (node 2) accounted for $80-90 \%$ of the observations, and there was less discrimination provided by the explanatory variables for those low catch rates than for the higher catch rates classified below node 3. Thus the numerous low catch rate observations had less information content than the higher catch rate observations.

Examination of positive catch rates by latitude and month (Figures 10 and 11) provided additional detail for defining regions and seasons. The high catch rates in the area of 30 N were apparent. Examination of the associated raw data revealed that those catch rates were associated with multiple vessels in multiple months and that lower catch rates were observed in the same region at about the same time. Based on the regression tree results and Figure 10, four regions were defined: $24-25.4 \mathrm{~N}, 25.5-28.6 \mathrm{~N}, 28.7-30.6 \mathrm{~N}$, and $30.7-$ 35 N . It was decided to use four seasons of three months each. As noted from the regression tree results and in Figure 11, April and May often had higher catch rates and June had higher rates than March. Therefore April, May and June were combined and the remaining months assigned to appropriate seasons. The fishing year thus started in April and continued through March of the following year.

Initial analyses of proportions positive and positive catch rate revealed difficulties associated with the taxonomic variables. Those variables were often correlated and typically there were few observations for some levels of those effects which resulted in a
very sparse distribution of the data and associated difficulties in analytical fitting. Collapsing the three aggregated taxonomic groups to two levels from four did not eliminate the problems; therefore one of those variables (inshore species) was selected for inclusion in the final model.

The data available for the earliest years of the head boat survey were limited in number and geographic coverage. Because of very limited samples in those early years all data from 1974 and 1975 were excluded, the data for the $25.5-28.6 \mathrm{~N}$ region were restricted to 1978 and later, and the data for the $24-25.4 \mathrm{~N}$ region were restricted to 1979 and later. Final analyses included factors for fishing year, season, region and the relative importance of inshore species in the catch. A four-way random effect interaction term was included. Results of the analyses of proportions positive and of catch rates on trips with greater amberjack are shown in Tables 16 and 17 . The final index showed an increase from 1977 to a peak in 1982, a general decline through much of the 1980's, with fluctuating levels in the late 1980's through mid 1990's and lower levels since then (Table 18 and Figure 12).

One additional set of analyses was performed to investigate the influence of the inshore taxa factor on the estimated standardized catch rates by redoing the all final analytical steps without that factor. Figure 13 shows that the relative index was quite similar with or without the inshore factor.

## Tagging Data Loss Rate Results

The number of greater amberjack tagged, the number recaptured, and the relative loss values with coefficients of variation for each year are presented in Table 19 and relative loss is shown in Figure 14 The general pattern is the same in all four scenarios, high values in the 60's, decreasing to the lowest values in the 80's, followed by an increasing trend in the 90 's. The coefficients of variation (CV) for the random walk models reflect the constraint imposed by the random walk process while the CV for the independent estimation of each year depends upon the sample size for that year (Table 19 and Figure 15). Three types of sensitivity analyses were performed. Instead of estimating $Z^{*}$ for every year, one $Z^{*}$ was estimated for every two years in order to examine if the low number of observations in some years were causing artifacts in the $Z^{*}$ estimates. The relative loss pattern did not change (Figure 16). Secondly, the minimum number of days at large was changed from one to either zero or seven in order to examine whether sufficient mixing of tagged and untagged fish was occurring. Again, the relative loss pattern did not change (Figure 17). The final sensitivity analysis allowed the proportion of tags reported to vary through a random walk process instead of being held constant in order to examine whether changes in reporting rates could explain the changes in $\mathrm{Z}^{*}$. Under two levels of sigma for the reporting rate random walk, the relative loss pattern did not change (Figure 18) even though the fraction reported was estimated to decrease over time (Figure 19).

Since all sensitivity analyses showed the same relative loss pattern, the EstAll time series was used to form relative abundance time series using the two catch series (Figure 20) and bur assumptions of how $Z^{*}$ could be split into its components (Table 20 and

Figure 21). The eight relative abundance time series are similar in trend and have a similar pattern of uncertainty about the point estimates, as demonstrated by Case 1 (Figure 22). The trend in abundance is increasing, but highly variable especially at the two peaks, from 1981 through 1989. From 1989 to 1990, the relative abundance values drop suddenly both in value and in variability. The trend from 1990 through 1997 is decreasing, with lower levels of uncertainty than the 1981 through 1989 estimates. This change in level of variability is caused by a large change in sample size (see Table 19).
Given the high level of uncertainty in the 1981 through 1989 values, the relative abundance trends for the eight cases were computed for only years 1990 through 1997 (Table 21 and Figure 23). Note these computations only rescaled the estimates from 1990-1997, they did not require additional $Z^{*}$ estimation. These values grouped by the catch used in the calculations, with the four assumptions of how $\mathrm{Z}^{*}$ is split adding only a small additional amount of variability among the cases. The relative abundance trends are more dependent upon the choice of catch than the choice of $M$ and no $n$-fishing related tag loss, when the latter two are assumed constant over the time series. If either M or the non-fishing related tag loss varied within the time series, these variables would become much more important in determining the trend in relative abundance. Plotting the uncertainty about the point estimates for Case 1 shows that some of the values are significantly different from others at the $80 \%$ confidence level (Figure 24).

A summary of the indices considered for use in further evaluations of the Atlantic greater amberjack stock condition is given in Table 22. These results provide four measures of the abundance of this stock. Three of the indices overlap in some degree as they all derived from recreational fishery, MRFSS charter, MRFSS private, and the tagging index. Future CPUE examinations should further examine the disparity between the MRFSS charter and the private index. Further examinations of the logbook data might consider the effect of zero trips in the analysis as done for the headboat fishery. Analyses of the tagging data might consider the impact of time at liberty on the loss rate calculation. Sutherland and Scott (see McClellan and Cummings 1996) reported greater amberjack generally had limited movements with about one third of the tagged fish recaptured within 60 days of release. McClellan and Cummings (1996) reported that 41 \% were recovered within 90 days, $69 \%$ within one year and $84.8 \%$ within two years.

## Acknowledgments

Patty Phares provided all data pertaining to the recreational and headboat fishery. John Poffenberger provided commercial logbook data and Nelson Johnson answered many questions related to using these data. Clay Porch provided the analysis program, TAP, used in the tagging analysis. Larry Massey provided editorial assistance. Thanks are extended to these individuals for their timeliness.

## Literature Cited

Cummings, N. J. 1998. Fishery statistics for the Atlantic greater amberjack, Seriola dumerili, since 1995. U.S. Dept. of Comm., NOAA, NMFS, SEFSC, Sustainable Fisheries Division SFD/98/99-41 27p.

Cummings, N. J. and D. B. McClellan. 1999. Characteristics of the Atlantic greater amberjack fishery through fishing year 1997. U.S. Dept. of Comm., NOAA, NMFS, SEFSC, Sustainable Fisheries Division SFD/98/99-61p.

Cummings, N. J. and D. B. McClellan. 1997. Status of the Greater Amberjack, Seriola dumerili, in the southeastern United States through 1995. Proc. Gulf and Carib. Fish. Inst. 49: 246-272.

Harris, K. C., G. N. Johnson, C. W. Krouse, and A. J. Chester. 1994. The 1993 South Atlantic snapper- grouper logbook program. U.S. Dept. of Comm., NOAA, NMFS, SEFSC Miami Lab. Contr. MIA- 50p.

Lo, N.C., L.D. Jackson, and J.L. Squire. 1992. Indices of relative abundance From fish spotter data based on delta-lognormal models. Can. J. Fish. Aquat. Sci. 49: 2515-2526.

McClellan, D.B. and J. A. Bohnsack. 1991. Review of the Gulf of Mexico reef fish logbook reporting system. U.S. Dept. of Comm., NOAA, NMFS, SEFSC Miami Lab. Contr. MIA- 36p.

McClellan, D.B and N.J. Cummings. 1997. Preliminary analysis of tag and recapture data of the greater amberjack, Seriola dumerili, in the southeastern United States. Gulf and Caribbean Fisheries Institute Proceedings 49: 25-45.

Parrack, N. C. [1993] The exploitation status of the Atlantic amberjack fisheries through 1991. U.S. Dept. of Comm., NOAA, NMFS, SEFSC Miami Lab. Contr. MIA-92/93-77. 32 p.

Parrack, N. C. 1993b. Updated fisheries information for greater amberjack through 1992. U.S. Dept. of Comm., NOAA, NMFS, SEFSC, Miami Lab. Contr. No. MIA-92/93-77 32 p.

Porch, C. E. 1998. Estimating Atlantic bluefin tuna mortality from the release and recapture dates of recovered tags (preliminary results). ICCAT SCRS/98/65. 9 pp.

Robson, D. S. 1966. Estimation of the relative fishing power of individual ships. ICNAF Res. Bull. No. 3:5-15.

Venables, W. N. and B. D. Ripley. 1997. Modern applied statistics with S-Plus. Springer. New York. 548 p.

Table 1a. Summary of Atlantic Greater Amberjack MRFSS CPUE data on Catch per Hour )and Catch per Angler hour (CPA).

| Fishing <br> Year | Number of <br> Observations | Mean CPH | Mean CPA |
| :---: | :---: | :---: | :---: |
| $1981 / 82$ | 18 | 0.41 | 2 |
| $1982 / 83$ | 29 | 0.19 | 1.15 |
| $1983 / 84$ | 46 | 0.2 | 0.98 |
| $1984 / 85$ | 62 | 0.19 | 1.16 |
| $1985 / 86$ | 68 | 0.28 | 1.22 |
| $1986 / 87$ | 145 | 0.27 | 1.25 |
| $1987 / 88$ | 137 | 0.22 | 1.05 |
| $1988 / 89$ | 167 | 0.19 | 0.76 |
| $1989 / 90$ | 140 | 0.28 | 1.09 |
| $1990 / 91$ | 132 | 0.25 | 1.18 |
| $1991 / 92$ | 182 | 0.27 | 1.08 |
| $1992 / 93$ | 159 | 0.26 | 1.21 |
| $1993 / 94$ | 139 | 0.32 | 1.44 |
| $1994 / 95$ | 123 | 0.24 | 1.06 |
| $1995 / 96$ | 130 | 0.24 | 1.05 |
| $1996 / 97$ | 109 | 0.23 | 1.06 |
| $1997 / 98$ | 105 | 0.25 | 1.21 |
| $1998 / 99$ | 71 | 0.22 | 1.06 |

Table 1b. Summary of Atlantic Greater Amberjack MRFSS CPUE Data (Catch Per Hour) by fishing year and state of intercept.

|  | Florida (west) |  | Florida (east) |  | Georgia |  | South Carolina |  | North Carolina |  | All States |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\begin{aligned} & \text { Fishing } \\ & \text { Year } \end{aligned}$ | $\begin{array}{\|c\|} \hline \text { Number of } \\ \text { Trips } \end{array}$ | Mean CPH | Number of Trips | Mean CPH | Number of Trips | Mean CPH | Number of Trips | Mean CPH | $\begin{gathered} \text { Number of } \\ \text { Trips } \end{gathered}$ | Mean CPH | Nu Trips | er of Mean CPH |
| 1980/81 |  |  | 3 | 0.42 |  |  |  |  |  |  | 3 | 0.42 |
| 1981/82 | 10 | 0.51 | 8 | 0.29 |  |  |  |  |  |  | 18 | 0.41 |
| 1982/83 | 9 | 0.19 | 15 | 0.21 | 2 | 0.06 | 2 | 0.16 | 1 | 0.2 | 29 | 0.19 |
| 1983/84 | 14 | 0.26 | 27 | 0.19 |  |  | 3 | 0.07 | 2 | 0.13 | 46 | 0.2 |
| 1984/85 | 12 | 0.15 | 32 | 0.23 | 7 | 0.17 | 9 | 0.15 | 2 | 0.04 | 62 | 0.19 |
| 1985/86 | 8 | 0.55 | 34 | 0.27 | 8 | 0.09 | 14 | 0.23 | 4 | 0.34 | 68 | 0.28 |
| 1986/87 | 29 | 0.35 | 45 | 0.28 | 11 | 0.23 | 50 | 0.24 | 10 | 0.13 | 145 | 0.27 |
| 1987/88 | 10 | 0.28 | 33 | 0.41 | 10 | 0.15 | 39 | 0.15 | 45 | 0.16 | 137 | 0.22 |
| 1988/89 | 20 | 0.12 | 64 | 0.3 | 6 | 0.1 | 46 | 0.14 | 31 | 0.1 | 167 | 0.19 |
| 1989/90 | 11 | 0.23 | 58 | 0.41 | 1 | 0.4 | 26 | 0.21 | 44 | 0.17 | 140 | 0.28 |
| 1990/91 | 13 | 0.21 | 68 | 0.33 | 2 | 0.08 | 7 | 0.26 | 42 | 0.15 | 132 | 0.25 |
| 1991/92 | 34 | 0.33 | 91 | 0.3 | 1 | 0.12 | 5 | 0.17 | 51 | 0.21 | 182 | 0.27 |
| 1992/93 | 34 | 0.44 | 72 | 0.28 | 11 | 0.27 | 6 | 0.06 | 36 | 0.1 | 159 | 0.26 |
| 1993/94 | 24 | 0.26 | 52 | 0.34 |  |  | 10 | 0.2 | 53 | 0.34 | 139 | 0.32 |
| 1994/95 | 13 | 0.17 | 41 | 0.35 | 5 | 0.1 | 2 | 0.28 | 62 | 0.19 | 123 | 0.24 |
| 1995/96 | 12 | 0.17 | 41 | 0.35 | 2 | 0.06 | 1 | 0.06 | 74 | 0.19 | 130 | 0.24 |
| 1996/97 | 6 | 0.1 | 31 | 0.39 | 4 | 0.04 | 9 | 0.16 | 59 | 0.19 | 109 | 0.23 |
| 1997/98 | 45 | 0.23 | 23 | 0.24 | 4 | 0.04 | 8 | 0.37 | 25 | 0.31 | 105 | 0.25 |
| 1998/1999 | 15 | 0.1 | 29 | 0.23 | 2 | 0.07 | 1 | 0.08 | 24 | 0.29 | 71 | 0.22 |
| Total | 319 | 0.27 | 767 | 0.31 | 76 | 0.15 | 238 | 0.19 | 565 | 0.2 | 1965 | 0.25 |

Table 1c. Summary of Atlantic Greater amberjack MRFSS CPUE Data by fishing mode and state of intercept.

|  | Florida (west) |  | Florida (east) |  | Georgia |  | South Carolina |  | North Carolina |  | All States |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Fishing | Number of |  | Number of |  | Number of |  | Number of |  | Number of |  |  | er of |
| Mode | Trips | Mean CPH | Trips | Mean CPH | Trips | Mean CPH | Trips | Mean CPH | Trips | Mean CPH | Trips | Mean CPH |
| Charter | 268 | 0.24 | 501 | 0.3 | 47 | 0.17 | 192 | 0.17 | 351 | 0.16 | 1359 | 0.23 |
| Private | 51 | 0.45 | 266 | 0.32 | 29 | 0.12 | 46 | 0.28 | 214 | 0.26 | 606 | 0.3 |
| ALL | 319 | 0.27 | 767 | 0.31 | 76 | 0.15 | 238 | 0.19 | 565 | 0.2 | 1965 | 0.25 |

Table 1d. Summary of Atlantic Greater Amberjack MRFSS CPUE Data by fishing year and month of intercept.

|  | MONTH |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Fishing | Number of Trips | Mean CPH | Number of Trips | Mean CPH | Number of Trips | Mean CPH | Number of Trips | Mean CPH | Number of Trips | Mean CPH | Number of Trips | Mean CPH | Number of Tribs | Mean CPH |
| 1980/81 |  |  |  |  |  |  | 3 | 0.42 |  |  |  |  |  |  |
| 1981/82 |  |  |  |  |  |  | 4 | 0.78 | 4 | 0.24 | 6 | 0.47 |  |  |
| 1982183 |  |  | 1 | 0.15 | 2 | 0.11 | 4 | 0.27 | 6 | 0.12 | 2 | 0.11 | 4 | 0.17 |
| 1983/84 |  |  | 9 | 0.32 | 5 | 0.15 | 5 | 0.24 | 4 | 0.16 | 7 | 0.27 | 4 | 0.05 |
| 1984/85 |  |  | 2 | 0.06 | 9 | 0.26 | 1 | 0.4 | 21 | 0.25 | 3 | 0.12 | 4 | 0.11 |
| 1985/86 | 6 | 0.25 | 8 | 0.29 |  |  | 23 | 0.35 | 9 | 0.19 | 6 | 0.11 | 4 | 0.25 |
| 1986/87 | 1 | 0.17 | 2 | 0.49 | 8 | 0.25 | 16 | 0.42 | 7 | 0.17 | 19 | 0.18 | 16 | 0.25 |
| 1987/88 |  |  | 3 | 0.28 | 5 | 0.25 | 10 | 0.17 | 8 | 0.14 | 24 | 0.35 | 20 | 0.31 |
| 1988/89 | 3 | 0.18 |  |  | 9 | 0.14 | 40 | 0.27 | 17 | 0.17 | 17 | 0.32 | 19 | 0.1 |
| 1989/90 |  |  |  |  | 7 | 0.32 | 28 | 0.47 | 26 | 0.27 | 5 | 0.3 | 13 | 0.28 |
| 1990/91 | 1 | 0.25 | 2 | 0.15 | 13 | 0.38 | 27 | 0.26 | 34 | 0.29 | 9 | 0.21 | 11 | 0.19 |
| 199192 | 13 | 0.36 | 10 | 0.36 | 44 | 0.26 | 22 | 0.28 | 32 | 0.36 | 16 | 0.27 | 12 | 0.19 |
| 1992/93 |  |  | 15 | 0.37 | 9 | 0.45 | 9 | 0.44 | 28 | 0.36 | 20 | 0.19 | 23 | 0.16 |
| 1993/94 | 3 | 0.11 | 11 | 0.27 | 11 | 0.28 | 23 | 0.35 | 16 | 0.24 | 15 | 0.28 | 13 | 0.48 |
| $1994 / 95$ | 4 | 0.17 | 1 | 1.08 | 5 | 0.24 | 6 | 0.07 | 23 | 0.53 | 17 | 0.17 | 33 | 0.16 |
| 1995/96 | 2 | 0.17 | 8 | 0.2 | 27 | 0.23 | 19 | 0.41 | 16 | 0.38 | ${ }^{28}$ | 0.2 | 14 | 0.12 |
| 1996/97 | 4 | 0.18 | 5 | 0.56 | 13 | 0.27 | 5 | 0.22 | 15 | 0.14 | 11 | 0.18 | 18 | 0.31 |
| 1997/98 | 16 | 0.34 | 8 | 0.14 | 8 | 0.32 | 4 | 0.2 | 2 | 0.04 | 10 | 0.18 | 13 | 0.43 |
| 1998/99 |  |  |  |  |  |  |  |  | 13 | 0.24 | 8 | 0.31 | 11 | 0.34 |
| Total | 53 | 0.28 | 85 | 0.31 | 175 | 0.27 | 249 | 0.33 | 281 | 0.29 | 223 | 0.24 | 232 | 0.23 |


|  | August |  | September |  | October |  | November |  | December |  | $\frac{\text { All Months }}{\text { Number of }}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Fishing | Number of Trips | Mean CPH | Number of Trips | Mean CPH | $\begin{array}{\|l\|} \hline \text { Number of } \\ \text { Trivs } \end{array}$ | Mean CPH | $\begin{gathered} \text { Number of } \\ \text { Trips } \end{gathered}$ | Mean CPH | $\begin{array}{\|c\|} \hline \text { Number of } \\ \text { Trips } \end{array}$ | Mean CPH |  |  |
| 1980/81 |  |  |  |  |  |  |  |  |  |  |  | 0.42 |
| 1981/82 | 1 | 0.02 | 1 | 0.33 |  |  | 2 | 0.07 |  |  | 18 | 0.41 |
| 1982/83 | 3 | 0.43 | 4 | 0.18 | 2 | 0.18 |  |  | 1 | 0.05 | 29 | 0.19 |
| 1983/84 | 6 | 0.07 | 3 | 0.22 | 3 | 0.15 |  |  |  |  | 46 | 0.2 |
| 1984/85 | 9 | 0.18 | 5 | 0.15 | 6 | 0.09 | 2 | 0.07 |  |  | 62 | 0.19 |
| 1985/86 | 7 | 0.35 | 2 | 0.5 | 2 | 0.2 |  | 0 |  |  | 68 | 0.28 |
| 1986/87 | 16 | 0.25 | 33 | 0.19 | 7 | 0.18 | 10 | 0.38 | 10 | 0.48 | 145 | 0.27 |
| $1987 / 88$ | 27 | 0.21 | 24 | 0.15 | 9 | 0.1 | 6 | 0.15 | 1 | 0.25 | 137 | 0.22 |
| 1988/89 | 9 | 0.1 | 20 | 0.09 | 12 | 0.18 | 20 | 0.2 | 1 | 0.2 | 167 | 0.19 |
| 1989/90 | 11 | 0.19 | 15 | 0.21 | 26 | 0.18 | 8 | 0.29 | 1 | 0.11 | 140 | 0.28 |
| 1990/91 | 15 | 0.21 | 3 | 0.07 | 7 | 0.19 | 6 | 0.13 | 4 | 0.36 | 132 | 0.25 |
| 199192 | 13 | 0.16 | 12 | 0.2 | 4 | 0.07 | 4 | 0.2 |  |  | 182 | 0.27 |
| 1992/93 | 17 | 0.25 | 11 | 0.17 | 13 | 0.19 | 7 | 0.18 | 7 | 0.16 | 159 | 0.26 |
| $1993 / 94$ | 23 | 0.43 | 4 | 0.34 |  | 0.18 | 12 | 0.21 | 2 | 0.17 | 139 | 0.32 |
| 1994/95 | 8 | 0.12 | 8 | 0.21 | 9 | 0.27 |  | 0.1 | 3 | 0.04 | 123 | 0.24 |
| 1995/96 | 5 | 0.15 | 1 | 0.1 | ¢ | 0.11 | 3 | 0.04 | 1 | 0.13 | 130 | 0.24 |
| 1996/97 | 16 | 0.3 | 5 | 0.14 | 11 | 0.16 | 5 | 0.05 | 1 | 0.04 | 109 | 0.23 |
| 1997/98 | 12 | 0.32 | 6 | 0.15 | 11 | 0.18 | 6 | 0.21 | 9 | 0.13 | 105 | 0.25 |
| 1998/99 | 5 | 0.13 | 6 | 0.15 | 11 | 0.18 | 12 | 0.14 | 5 | 0.16 | 71 | 0.22 |
| Total | 203 | 0.24 | 163 | 0.17 | 145 | 0.17 | 110 | 0.19 | 46 | 0.23 | 1965 | 0.25 |

Table 1e. Summary of Atlantic Greater Amberiack MRFSS Data for observed Catch per Hour (CPH) by state and code county.

|  | Florida (west) |  | Florida (east) |  | Georgia |  | South Carolina |  | North Carolina |  | All States |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Fishing | Number of Trins | Mean CPH | Number of Trins | Mean CPH | Number of Trins | Mean CPH | Number of Trins | Mean CPH | Number of Trins | Mean CPH | $\xrightarrow{\text { N }}$ | er of Mean CPH |
| 5022 | 319 | 0.27 |  |  |  |  |  |  |  |  | 319 | 0.27 |
| 6023 |  |  | 110 | 0.29 |  |  |  |  |  |  | 110 | 0.29 |
| 6024 |  |  | 95 | 0.29 |  |  |  |  |  |  | 95 | 0.29 |
| 6025 |  |  | 397 | 0.33 |  |  |  |  |  |  | 397 | 0.33 |
| 6026 |  |  | 37 | 0.26 |  |  |  |  |  |  | 37 | 0.26 |
| 6027 |  |  | 9 | 0.43 |  |  |  |  |  |  | 9 | 0.43 |
| 6028 |  |  | 4 | 0.23 |  |  |  |  |  |  | 4 | 0.23 |
| 6029 |  |  | 24 | 0.18 |  |  |  |  |  |  | 24 | 0.18 |
| 6030 |  |  | 23 | 0.39 |  |  |  |  |  |  | 23 | 0.39 |
| 6032 |  |  | 12 | 0.72 |  |  |  |  |  |  | 12 | 0.72 |
| 6034 |  |  | 50 | 0.21 |  |  |  |  |  |  | 50 | 0.21 |
| 6035 |  |  | 6 | 0.24 |  |  |  |  |  |  | 6 | 0.24 |
| 7029 |  |  |  |  | 1 | 0.04 |  |  |  |  | 1 | 0.04 |
| 7039 |  |  |  |  | 10 | 0.2 |  |  |  |  | 10 | 0.2 |
| 7051 |  |  |  |  | 29 | 0.15 |  |  |  |  | 29 | 0.15 |
| 7127 |  |  |  |  | 34 | 0.14 |  |  |  |  | 34 | 0.14 |
| 7191 |  |  |  |  | 2 | 0.09 |  |  |  |  | 2 | 0.09 |
| 8013 |  |  |  |  |  |  | 118 | 0.21 |  |  | 118 | 0.21 |
| 8015 |  |  |  |  |  |  | 1 | 0.25 |  |  | 1 | 0.25 |
| 8019 |  |  |  |  |  |  | 33 | 0.21 |  |  | 33 | 0.21 |
| 8043 |  |  |  |  |  |  | 77 | 0.15 |  |  | 77 | 0.15 |
| 8051 |  |  |  |  |  |  | 9 | 0.16 |  |  | 9 | 0.16 |
| 9019 |  |  |  |  |  |  |  |  | 62 | 0.1 | 62 | 0.1 |
| 9031 |  |  |  |  |  |  |  |  | 98 | 0.11 | 98 | 0.11 |
| 9055 |  |  |  |  |  |  |  |  | 281 | 0.24 | 281 | 0.24 |
| 9095 |  |  |  |  |  |  |  |  | 48 | 0.32 | 48 | 0.32 0.13 |
| 9129 |  |  |  |  |  |  |  |  | 60 | 0.13 | 60 | 0.13 |
| 9133 |  |  |  |  |  |  |  |  | 7 | 0.19 | 7 | 0.19 |
| ALL | 319 | 0.27 | 767 | 0.31 | 76 | 0.15 | 238 | 0.19 | 565 | 0.2 | 1965 | 0.25 |

Table 2. Summary Results of the Atlantic Greater amberjack MRFSS CPUE Analyses

| RUN ID | TYPE MODEL FIT | \# Observations | Proportion of Total <br> Sum of Squares Explained | Comments On Model Fit |
| :---: | :---: | :---: | :---: | :---: |
| 1 Florida Charter | GLM ON Positive Catches | 735 | 21.56 | Bag Limit Term Not Significant |
| 2 Florida Private | GLM ON Positive Catches | 304 | 20.66 | Bag Limit Term Not Ssignificant |
| 3 Charter All States | GLM ON Positive Catches | 1195 | 44.16 | Bag Limit Term Not Significant |
| 4 Private All States | GLM ON Positive Catches | 525 | 27.71 | Bag Limit Term Not ignificant, Month not Significant |
| 5 Charter and Priva | GLM ON Positive Catches te | 1725 | 34.23 |  |

Table 3. Number of trips by gear and fishing year. $\mathrm{H}=$ handline, $\mathrm{L}=$ longline, $\mathrm{P}=$ powerhead, $\mathrm{S}=\mathrm{spear}$, and TR=troll.
FI SHYEAR
GEAR

| , H | L | P | , S |  |  | Tot al |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| fffffffff^ffffffff^ffffffff^ffffffff^ffffffff^ffffffff^ |  |  |  |  |  |  |
| 1990 | 91 | 6 | 0 |  | 0, | 101 |
| fffffffff^ffffffff^ffffffff^ffffffff^ffffffff^ffffffff^ |  |  |  |  |  |  |
| 1991 | 1467 | 77 | 0 | 83 | 156 | 1783 |
| fffffffff^ffffffff^ffffffff^ffffffff^ffffffff^fffffffe |  |  |  |  |  |  |
| 1992 | 6820 | 348 | 0, | 532 | 777 | 8477 |
| fffffffff^ffffffff^ffffffff^ffffffff^ffffffff ${ }^{\wedge} \mathrm{fffffff}{ }^{\wedge}$ |  |  |  |  |  |  |
| 1993 | 14216 | 773 | 5 | 1066 | 1705 | 17765 |
| fffffffff^ffffffff^ffffffff^ffffffff^ffffffff^ffffffff^ |  |  |  |  |  |  |
| 1994 | 15436 | 719 | 89 | 869 | 1972 | 19085 |
| fffffffff^ffffffff^ffffffff^ffffffff^ffffffff $\mathrm{ffffffff}{ }^{\wedge}$ |  |  |  |  |  |  |
| 1995 | 14613 | 886 | 245 | 708 | 2148 | 18600 |
| fffffffff^ffffffff^ffffffff^ffffffff^ffffffff^ffffffff^ |  |  |  |  |  |  |
| 1996 | 16338 | 653 | 275 | 763 | 1995 | 20024 |
| fffffffff^ffffffff^ffffffff^ffffffff^ffffffff^ffffffff |  |  |  |  |  |  |
| 1997 | 16140 | 584 | 286 | 833 | 4119 | 21962 |
| fffffffff^ffffffff^ffffffff^ffffffff^ffffffff^fffffff^ |  |  |  |  |  |  |
| 1998 | 13419 | 441 | 233 | 709 | 8050 | 22852 |
| fffffffff^ffffffff^ffffffff^ffffffff^ffffffff^ffffffff^ |  |  |  |  |  |  |
| Tot al | 98540 | 4487 | 1133 | 5567 | 20922 | 130649 |

Table 4. Proportion of trips with greater amberjack by gear and fishing year. $\mathrm{H}=$ handline, $\mathrm{L}=$ longline, $\mathrm{P}=$ powerhead, $\mathrm{S}=$ spear, and $T R=$ troll

FI SHYEAR
GEAR


Table 5. Proportion of greater amberjack in the total yield by gear and fishing year. $\mathrm{H}=$ handline, $\mathrm{L}=$ longline, $\mathrm{P}=$ powerhead, $\mathrm{S}=$ spear, and TR=troll

FI SHYEAR GEAR


Table 6. GLM analyses of greater amberjack landings per trip including fishing year, season, region and various measures of effort.
Class Level Information
Class Levels Values
FISHYEAR $8 \quad 19911992199319941995199619971998$ REGION 4 24-26N 27-28N 29-33N 34-35N SEASON 4 AprMayJun JanFebMar JulAugSep OctNovDec

Dependent Variable: LYIELD


Table 6. continued


Table 6. (Cont.)



Table 7. Anal ysis of proportions of all handine trips which caught greater anmerjackin the Atlantic Ocean in the United States EEZ.

| Cl ass | Level s | Val ues |  |
| :---: | :---: | :---: | :---: |
| FI SHYR2 | 8 | 1991199219931994 | 19951996 |
|  |  | 19971998 |  |
| SEASON2 | 4 | DecJ anFeb J unj ul Aug | Mar Apr May |
|  |  | SepOct Nov |  |
| REGI ON | 4 | 24-26N 27-28N 29-33N | N 34-35N |

Mdel Fitting Information for proportion positive
Description
Observations 122.0000
Res Log Li kel i hood -53.0508
Akai ke' s Inf or mation Criterion -56. 0508
Schwarz's Bayesian Criterion -59.9435
2 Res Log Li kel i hood 106. 101
Information Criteria

| Better | Parns | q | p | Al C | HOI C | BI C | CAI C |
| :--- | ---: | :--- | ---: | ---: | ---: | ---: | ---: |
|  |  |  |  |  |  |  |  |
| Larger | 3 | 3 | 0 | -56.1 | -57.6 | -59.9 | -61.4 |
| Larger | 26 | 3 | 23 | -79.1 | -92.7 | -112.8 | -125.8 |
| Snal I er | 3 | 3 | 0 | 112.1 | 115.3 | 119.9 | 122.9 |
| Smal I er | 26 | 3 | 23 | 158.1 | 185.4 | 225.6 | 251.6 |

                    Tests of Fixed Effects
    | Source | NDF | DDF | Type III Chi Sq | Type 111 F | Pr $>$ Chi Sq | Pr $>$ F |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
|  |  |  |  |  |  |  |
| FI SHR2 | 7 | 21 | 17.85 | 2.55 | 0.0127 | 0.0456 |
| SEASON2 | 3 | 21 | 12.22 | 4.07 | 0.0067 | 0.0199 |
| REG ON | 3 | 21 | 306.58 | 102.19 | 0.0001 | 0.0001 |
| SEASON2*REG ON | 9 | 57 | 171.46 | 19.05 | 0.0001 | 0.0001 |

Covariance Parameter Estimates (REML)
Cov Parm Estimate

FI SHYR2*SEASON2 0.01043981 FI SHYR2*REGI ON 0.05463983 Resi dual 2. 98853706

Sol ution for Fixed Effects


Table 7. (Cont.)

| FI SHYR2 | 1991 |  |  | 0. 56403445 | 0. 26525109 | 21 | 2. 13 | 0. 0455 | 0.05 | 0. 0124 | 1. 1157 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| FI SHYR2 | 1992 |  |  | 0. 32957568 | 0. 20967553 | 21 | 1. 57 | 0. 1309 | 0. 05 | -0. 1065 | 0. 7656 |
| FI SHYR2 | 1993 |  |  | 0. 54515533 | 0. 20017459 | 21 | 2. 72 | 0. 0127 | 0. 05 | 0. 1289 | 0. 9614 |
| FI SHYR2 | 1994 |  |  | 0. 53972636 | 0. 19943970 | 21 | 2. 71 | 0. 0132 | 0.05 | 0. 1250 | 0. 9545 |
| FI SHYR2 | 1995 |  |  | 0. 68474341 | 0. 19942989 | 21 | 3. 43 | 0. 0025 | 0. 05 | 0. 2700 | 1. 0995 |
| FI SHYR2 | 1996 |  |  | 0. 71064434 | 0. 19921340 | 21 | 3. 57 | 0. 0018 | 0. 05 | 0. 2964 | 1. 1249 |
| FI SHYR2 | 1997 |  |  | 0. 52088696 | 0. 19886114 | 21 | 2. 62 | 0. 0160 | 0.05 | 0. 1073 | 0. 9344 |
| FI SHYR2 | 1998 |  |  | 0. 00000000 |  |  |  |  |  |  |  |
| SEASON2 |  | DecJ anFeb |  | -0. 22143936 | 0. 19302191 | 21 | -1. 15 | 0. 2642 | 0. 05 | -0. 6229 | 0. 1800 |
| SEASON2 |  | $J$ unj ul Aug |  | 0. 47716142 | 0. 15463618 | 21 | 3. 09 | 0. 0056 | 0. 05 | 0. 1556 | 0. 7987 |
| SEASON2 |  | Mar Apr May |  | 0. 08318874 | 0. 17630589 | 21 | 0. 47 | 0. 6419 | 0. 05 | -0. 2835 | 0. 4498 |
| SEASON2 |  | Sepoct Nov |  | 0. 00000000 |  |  |  |  |  |  |  |
| REG ON |  |  | 24.26 N | -0. 42702914 | 0. 18418072 | 21 | - 2.32 | 0. 0306 | 0. 05 | -0. 8101 | - 0.0440 |
| REG ON |  |  | 27-28N | 0. 95113815 | 0. 24252795 | 21 | 3. 92 | 0. 0008 | 0.05 | 0. 4468 | 1. 4555 |
| REG ON |  |  | 29-33N | 1. 99511248 | 0. 17897000 | 21 | 11. 15 | 0. 0001 | 0. 05 | 1. 6229 | 2. 3673 |
| REGI ON |  |  | 34.35 N | 0. 00000000 |  |  |  |  |  |  |  |
| SEASON2*REG ON |  | DecJ anFeb | 24.26 N | 0. 56281743 | 0. 20408933 | 57 | 2. 76 | 0. 0078 | 0. 05 | 0. 1541 | 0. 9715 |
| SEASON2*REGI ON |  | DecJ anFeb | 27-28N | 0. 05965008 | 0. 29261423 | 57 | 0. 20 | 0. 8392 | 0. 05 | -0. 5263 | 0. 6456 |
| SEASON2*REG ON |  | DecJ anFeb | 29-33N | 0. 14603110 | 0. 20324331 | 57 | 0. 72 | 0. 4754 | 0. 05 | -0. 2610 | 0. 5530 |
| SEASON2*REG ON |  | DecJ anFeb | 3435 N | 0. 00000000 |  | . |  |  |  |  |  |
| SEASON2*REG ON |  | J unj ul Aug | 24.26 N | -0.45051447 | 0. 17425632 | 57 | -2. 59 | 0. 0123 | 0. 05 | -0.7995 | - 0.1016 |
| SEASON2*REG ON |  | J unj ul Aug | 27-28N | -0. 36412550 | 0. 25813015 | 57 | -1. 41 | 0. 1638 | 0. 05 | -0. 8810 | 0. 1528 |

Table 7. conti nued.

| SEASON2*REG ON | J unj ul Aug | 29-33N | -0. 46351812 | 0. 16529139 | 57 | -2. 80 | 0. 0069 | 0. 05 | -0. 7945 | -0.1325 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| SEASON2*REG ON | $J$ unj ul Aug | 3435 N | 0. 00000000 |  | . |  |  |  |  |  |
| SEASON2*REG ON | Mar Apr May | 24.26 N | 0. 83105120 | 0. 18641964 | 57 | 4. 46 | 0. 0001 | 0.05 | 0. 4578 | 1. 2044 |
| SEASON2*REGI ON | Mar Apr May | 27-28N | 0. 04561799 | 0. 26783357 | 57 | 0. 17 | 0. 8654 | 0. 05 | -0. 4907 | 0. 5819 |
| SEASON2*REG ON | Mar Apr May | 29-33N | - 0.25210842 | 0. 18684167 | 57 | -1. 35 | 0. 1826 | 0. 05 | -0. 6263 | 0. 1220 |
| SEASON2*REG ON | Mar Apr May | 3435 N | 0. 00000000 |  | . |  |  |  |  |  |
| SEASON2*REG ON | SepOct Nov | 2426 N | 0. 00000000 |  | . |  |  | . | . |  |
| SEASON2*REG ON | SepOct Nov | 27-28N | 0. 00000000 | . | . | . |  | . |  |  |
| SEASON2*REG ON | SepOct Nov | 29-33N | 0. 00000000 | . | . |  | . | . | . |  |
| SEASON2*REGI ON | SepOct Nov | 34 35N | 0. 00000000 | . | . | . | . | . | . |  |


| Effect | FI SHYR2 | LSMEAN | Std Error | DF | t | $\operatorname{Pr}>\|t\|$ | Al pha | Lower | Upper | COV1 | COV2 | cov3 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| FI SHYR2 | 1991 | - 1. 57938625 | 0. 22493361 | 21 | -7. 02 | 0. 0001 | 0.05 | - 2.0472 | -1. 1116 | 0. 05 | 0. 00 | 0. 00 |
| FI SHYR2 | 1992 | -1. 81384501 | 0. 15490599 | 21 | -11. 71 | 0. 0001 | 0.05 | -2. 1360 | -1. 4917 | 0. 00 | 0. 02 | 0. 00 |
| FI SHYR2 | 1993 | -1. 59826536 | 0. 14142199 | 21 | -11. 30 | 0. 0001 | 0.05 | -1. 8924 | -1. 3042 | 0. 00 | 0. 00 | 0.02 |
| FI SHYR2 | 1994 | -1. 60369434 | 0. 14040080 | 21 | -11.42 | 0. 0001 | 0.05 | -1.8957 | -1. 3117 | 0. 00 | 0. 00 | 0.00 |
| FI SHYR2 | 1995 | - 1. 45867729 | 0. 14041218 | 21 | -10. 39 | 0. 0001 | 0.05 | -1. 7507 | -1. 1667 | 0. 00 | 0. 00 | 0. 00 |
| FI SHYR2 | 1996 | -1. 43277636 | 0. 14019537 | 21 | -10. 22 | 0. 0001 | 0.05 | -1. 7243 | -1. 1412 | 0. 00 | 0. 00 | 0. 00 |
| FI SHYR2 | 1997 | -1. 62253374 | 0. 13957250 | 21 | -11.63 | 0. 0001 | 0.05 | -1. 9128 | -1. 3323 | 0. 00 | 0. 00 | 0. 00 |
| FI SHYR2 | 1998 | - 2. 14342070 | 0. 14236832 | 21 | -15. 06 | 0. 0001 | 0.05 | -2. 4395 | -1. 8473 | 0. 00 | 0. 00 | 0. 00 |

Table 8. Anal ysis of positive catch rates of greater anberjack fromall handine trips in the United States EEZ in the Atlantic.
Cl ass Levels Val ues
FI SHYR2 $\quad 8 \quad 199119921993199419951996$ 199119921993199419951996 19971998
SEASON2 4 DecJ anFeb J unJ ul Aug Mar Apr May SepOct Nov
REGI ON $\quad 4 \quad 24-26 \mathrm{~N}$ 27-28N 29-33N 34-35N
Mbdel Fitting Information for LCPHOOK
Description Val ue
Obser vati ons 13241.00
Res Log Li kelihood - 25454.6
Akai ke's Inf or mation Criterion -25456. 6
Schwarz's Bayesian Criterion -25464.1
-2 Res Log Li kel i hood 50909.17

Information Criteria

| Bet ter | Parns | q | $p$ | Al C | HQ C | BI C | CAI C |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Larger | 2 | 2 | 0 | -25457 | -25459 | -25464 | -25465 |
| Larger | 25 | 2 | 23 | -25480 | -25511 | -25573 | -25586 |
| Smal I er | 2 | 2 | 0 | 50913.2 | 50918.2 | 50928.1 | 50930.1 |
| Smal I er | 25 | 2 | 23 | 50959.2 | 51021.7 | 51146.4 | 51171.4 |


| Source | NDF | DDF | Type III F | Pr $>F$ |
| :--- | ---: | ---: | ---: | ---: |
|  |  |  |  |  |
| FI SHYR2 | 7 | 98 | 1.70 | 0.1183 |
| SEASONR | 3 | 98 | 4.87 | 0.0034 |
| REG ON | 3 | 98 | 229.95 | 0.0001 |
| SEASON2*REG ON | 9 | 98 | 6.46 | 0.0001 |

Covariance Parameter Estimates (REM)

| Cov Parm | Estimate |
| :--- | ---: |
| FI SHYR*SEASON*REGI ON | 0.06836473 |
| Resi dual | 2. 70392982 |

Sol ution for Fixed Effects

| Effect | FI SHYR2 | SEASON2 | REG ON | Estimate | Std Error | DF | t | Pr > $\|t\|$ | Al pha | Lower | Upper |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 NTERCEPT |  |  |  | 2. 08063743 | 0. 16768212 | 98 | 12. 41 | 0. 0001 | 0. 05 | 1. 7479 | 2. 4134 |
| FI SHYR2 | 1991 |  |  | 0. 38002686 | 0. 19554093 | 98 | 1. 94 | 0. 0548 | 0.05 | -0. 0080 | 0. 7681 |
| FI SHYR2 | 1992 |  |  | 0. 01035985 | 0. 13238102 | 98 | 0.08 | 0. 9378 | 0. 05 | - 0.2523 | 0. 2731 |
| FI SHYR2 | 1993 |  |  | -0.08193061 | 0. 11935143 | 98 | -0. 69 | 0. 4940 | 0. 05 | -0.3188 | 0. 1549 |
| FI SHYR2 | 1994 |  |  | 0. 07975599 | 0. 11808182 | 98 | 0.68 | 0. 5010 | 0. 05 | -0. 1546 | 0. 3141 |
| FI SHYR2 | 1995 |  |  | 0. 23178827 | 0. 11764791 | 98 | 1. 97 | 0. 0516 | 0. 05 | -0. 0017 | 0. 4653 |
| FI SHYR2 | 1996 |  |  | 0. 07485880 | 0. 11791557 | 98 | 0.63 | 0. 5270 | 0. 05 | -0. 1591 | 0. 3089 |
| FI SHYR2 | 1997 |  |  | 0. 02078504 | 0. 11797023 | 98 | 0. 18 | 0. 8605 | 0.05 | -0. 2133 | 0. 2549 |
| FI SHYR2 | 1998 |  |  | 0. 00000000 |  |  |  |  |  |  |  |
| SEASON2 |  | DecJ anFeb |  | 0. 12244925 | 0. 22329299 | 98 | 0. 55 | 0. 5847 | 0. 05 | - 0.3207 | 0. 5656 |
| SEASON2 |  | $J$ unj ul Aug |  | -0. 04830498 | 0. 19451186 | 98 | -0. 25 | 0. 8044 | 0. 05 | -0. 4343 | 0. 3377 |
| SEASON2 |  | Mar Apr May |  | -0. 16233190 | 0. 20873767 | 98 | -0.78 | 0. 4386 | 0. 05 | -0. 5766 | 0. 2519 |
| SEASON2 |  | SepOct Nov |  | 0. 00000000 |  |  |  |  |  |  |  |
| REGI ON |  |  | 24-26N | 1. 86023658 | 0. 19257579 | 98 | 9. 66 | 0. 0001 | 0. 05 | 1. 4781 | 2. 2424 |
| REG ON |  |  | 27-28N | 2. 22513015 | 0. 23462561 | 98 | 9. 48 | 0. 0001 | 0. 05 | 1. 7595 | 2. 6907 |
| REGI ON |  |  | 29-33N | 1. 06742680 | 0. 18294703 | 98 | 5. 83 | 0. 0001 | 0. 05 | 0. 7044 | 1. 4305 |
| REG ON |  |  | 34-35N | 0. 00000000 |  |  |  |  |  |  |  |
| SEASON2*REGI ON |  | DecJ anFeb | 24-26N | 0. 04425942 | 0. 27643427 | 98 | 0. 16 | 0. 8731 | 0. 05 | - 0.5043 | 0. 5928 |
| SEASON2*REGI ON |  | DecJ anFeb | 27-28N | -0. 61064583 | 0. 33065271 | 98 | -1. 85 | 0. 0678 | 0. 05 | -1. 2668 | 0. 0455 |
| SEASON2*REG ON |  | DecJ anFeb | 29-33N | -0. 15560263 | 0. 26983767 | 98 | -0. 58 | 0. 5655 | 0. 05 | -0. 6911 | 0. 3799 |
| SEASON2*REGI ON |  | DecJ anFeb | 34-35N | 0. 00000000 |  |  |  |  |  |  |  |
| SEASON2*REGI ON |  | $J$ unj ul Aug | 24-26N | 0. 16264767 | 0. 25968612 | 98 | 0. 63 | 0. 5326 | 0. 05 | - 0.3527 | 0. 6780 |
| SEASON2*REGI ON |  | $J$ unj ul Aug | 27-28N | -0. 38352089 | 0. 30681820 | 98 | -1. 25 | 0. 2143 | 0. 05 | -0. 9924 | 0. 2253 |
| SEASON2*REGI ON |  | $J$ unj ul Aug | 29-33N | -0.12648973 | 0. 24799538 | 98 | -0. 51 | 0. 6112 | 0. 05 | -0.6186 | 0. 3656 |
| SEASON2*REGI ON |  | $J$ unj ul Aug | 34-35N | 0. 00000000 |  |  |  |  |  |  |  |
| SEASON2*REGI ON |  | Mar Apr May | 24-26N | 1. 23843232 | 0. 26286674 | 98 | 4. 71 | 0. 0001 | 0. 05 | 0. 7168 | 1. 7601 |
| SEASON2*REGI ON |  | Mar Apr May | 27-28N | 0. 22341456 | 0. 31112261 | 98 | 0. 72 | 0. 4744 | 0. 05 | - 0.3940 | 0. 8408 |
| SEASON2*REGI ON |  | Mar Apr May | 29-33N | -0.03735220 | 0. 25835643 | 98 | -0. 14 | 0. 8853 | 0. 05 | -0. 5501 | 0. 4753 |
| SEASON2*REGI ON |  | Mar Apr May | 34-35N | 0. 00000000 | . | . | . |  |  |  |  |
| SEASON2*REGI ON |  | SepOct Nov | 24-26N | 0. 00000000 |  | . | . |  |  | . |  |
| SEASON2*REGI ON |  | SepOct Nov | 27-28N | 0. 00000000 | . |  | . |  |  |  |  |
| SEASON2*REGI ON |  | SepOct Nov | 29-33N | 0. 00000000 |  | . |  |  |  |  |  |
| SEASON2*REGI ON |  | SepOct Nov | 34-35N | 0. 00000000 |  |  |  |  |  |  |  |

Least Squares Means

| Effect | FI SHYR2 | LSMEAN | Std Error | DF | t | Pr $>\|\mathrm{t}\|$ | Al pha | Lower | Upper | COV1 | COV2 | COV3 |
| :--- | :--- | ---: | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
| FI SHYR2 | 1991 | 3.74901219 | 0.17673619 | 98 | 21.21 | 0.0001 | 0.05 | 3.3983 | 4.0997 | 0.03 | 0.00 | 0.00 |
| FI SHYR2 | 1992 | 3.37934518 | 0.10193083 | 98 | 33.15 | 0.0001 | 0.05 | 3.1771 | 3.5816 | 0.00 | 0.01 | 0.00 |
| FI SHYR2 | 1993 | 3.28705472 | 0.0835141 | 98 | 39.35 | 0.0001 | 0.05 | 3.1213 | 3.4528 | 0.00 | 0.00 | 0.01 |
| FI SHYR2 | 1994 | 3.44874132 | 0.08174061 | 98 | 42.19 | 0.0001 | 0.05 | 3.2865 | 3.610 | 0.00 | 0.00 | 0.00 |
| FI SHYR2 | 1995 | 3.60077359 | 0.08101959 | 98 | 44.44 | 0.0001 | 0.05 | 3.4400 | 3.7616 | 0.00 | 0.00 | 0.00 |
| FI SHYR2 | 1996 | 3.44384412 | 0.08188709 | 98 | 42.06 | 0.0001 | 0.05 | 3.2813 | 3.6063 | 0.00 | 0.00 | 0.00 |
| FI SHYR2 | 1997 | 3.38977036 | 0.08186413 | 98 | 41.41 | 0.0001 | 0.05 | 3.2273 | 3.522 | 0.00 | 0.00 | 0.00 |
| FI SHYR2 | 1998 | 3.36898532 | 0.08666967 | 98 | 38.87 | 0.0001 | 0.05 | 3.1970 | 3.5410 | 0.00 | 0.00 | 0.00 |

Table 9. Index of abundance of Atlantic greater anberjack fromall handline vessel s. CPUE is the standardized catch rates on trips with greater amberjack, PPOS is the proportion positive, INDEX is the standardized index and SE_l and CV_l are the standard error and coefficient of variation about the index.

| FI SHYR2 | CPUE | PPOS | BC_CPU | GC | BC_POS | I NDEX | SE_I | CV_I |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  |  |  |
| 1991 | 43.1477 | 0.17088 | 152.326 | 3.58591 | 0.17088 | 26.0298 | 13.4532 | 0.51684 |
| 1992 | 29.5044 | 0.14017 | 110.895 | 3.77818 | 0.14017 | 15.5446 | 3.8948 | 0.25055 |
| 1993 | 26.8575 | 0.16822 | 101.287 | 3.78446 | 0.16822 | 17.0389 | 3.9976 | 0.23462 |
| 1994 | 31.5660 | 0.16747 | 119.079 | 3.78501 | 0.16747 | 19.9417 | 4.6549 | 0.23343 |
| 1995 | 36.7470 | 0.18867 | 138.640 | 3.78522 | 0.18867 | 26.1571 | 6.0467 | 0.23117 |
| 1996 | 31.4122 | 0.19267 | 118.496 | 3.78496 | 0.19267 | 22.8302 | 5.2769 | 0.23114 |
| 1997 | 29.7587 | 0.16486 | 112.259 | 3.78497 | 0.16486 | 18.5065 | 4.3189 | 0.23337 |
| 1998 | 29.1583 | 0.10495 | 109.906 | 3.78348 | 0.10495 | 11.5344 | 2.7962 | 0.24242 |

Table 10. Anal ysis of proportions of trips which caught greater anberjack fromhandine vessel sin the $90^{\text {th }} 99^{\text {th }}$ percentile for positive catch rate or proportion of trips with greater amberjack in the Atlantic Ocean in the United States EEZ.

Cl ass Levels Val ues


Mbdel Fitting Information for proportion positive

| Description | Val ue |
| :--- | ---: |
| Observations | 88.0000 |
| Res Log Li kel i hood | -88.1390 |
| Akai ke's I nf ormation Criterion | -90.1390 |
| Schwarz's Bayesian Criterion | -92.3731 |
| - 2 Res Log Li kel i hood | 176.2779 |
| I nf ormation Criteria |  |


| Bet ter | Parms | q | p | Al C | HQI C | BI C | CAI C |
| :--- | ---: | :--- | ---: | ---: | ---: | ---: | ---: |
|  |  |  |  |  |  |  |  |
| Larger | 2 | 2 | 0 | -90.1 | -91.0 | -92.4 | -93.4 |
| Larger | 21 | 2 | 19 | -109.1 | -118.4 | -132.6 | -143.1 |
| Smal I er | 2 | 2 | 0 | 180.3 | 182.1 | 184.7 | 186.7 |
| Smal I er | 21 | 2 | 19 | 218.3 | 236.9 | 265.2 | 286.2 |

Tests of Fixed Effects

| Source | NDF | DDF | Type III Chi Sq | Type III F | Pr $>$ Chi Sq | Pr $>$ F |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
|  |  |  |  |  |  |  |
| FI SHYR2 | 7 | 69 | 18.66 | 2.67 | 0.0093 | 0.0168 |
| SEASON2 | 3 | 69 | 7.89 | 2.63 | 0.0483 | 0.0569 |
| REG ON | 2 | 69 | 259.82 | 129.91 | 0.0001 | 0.0001 |
| SEASON2*REG ON | 6 | 69 | 21.65 | 3.61 | 0.0014 | 0.0036 |

Table 10. conti nued.

## Covariance Parameter Estimates (REM)

Estimate
FI SHYR*SEASON*REG ON 0. 02716895
Resi dual 3.02890551

Sol ution for Fixed Effects

| Effect | FI SHYR2 | SEASON2 | REGI ON | Estimate | St d Error | DF | t | $\operatorname{Pr}>\|t\|$ | Al pha | Lower | Upper |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| I NTERCEPT |  |  |  | - 2. 68879147 | 0. 35750756 | 69 | -7. 52 | 0. 0001 | 0.05 | -3. 4020 | -1. 9756 |
| FI SHYR2 | 1991 |  |  | - 1. 19406655 | 0. 45073617 | 69 | -2. 65 | 0. 0100 | 0.05 | -2. 0933 | - 0. 2949 |

Table 10 (Cont.)

| FI SHYR2 | 1992 |  |  | 0. 23333536 | 0. 25551856 | 69 | 0. 91 | 0. 3643 | 0. 05 | - 0.2764 | 0. 7431 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| FI SHYR2 | 1993 |  |  | 0. 35453350 | 0. 22593506 | 69 | 1. 57 | 0. 1212 | 0. 05 | -0.0962 | 0. 8053 |
| FI SHYR2 | 1994 |  |  | 0. 29794656 | 0. 21307125 | 69 | 1. 40 | 0. 1665 | 0. 05 | -0. 1271 | 0. 7230 |
| FI SHYR2 | 1995 |  |  | 0. 34350400 | 0. 21874988 | 69 | 1.57 | 0. 1209 | 0. 05 | -0. 0929 | 0. 7799 |
| FI SHYR2 | 1996 |  |  | 0. 50978014 | 0. 21252487 | 69 | 2. 40 | 0. 0192 | 0. 05 | 0. 0858 | 0. 9338 |
| FI SHYR2 | 1997 |  |  | 0. 18608108 | 0. 21153549 | 69 | 0. 88 | 0. 3821 | 0. 05 | -0. 2359 | 0. 6081 |
| FI SHYR2 | 1998 |  |  | 0. 00000000 |  |  |  |  |  |  |  |
| SEASON2 |  | DecJ anFeb |  | -0. 01046292 | 0. 49706658 | 69 | -0. 02 | 0. 9833 | 0. 05 | -1. 0021 | 0. 9812 |
| SEASON2 |  | J unj ul Aug |  | 0. 73481476 | 0. 39910808 | 69 | 1. 84 | 0. 0699 | 0. 05 | -0.0614 | 1. 5310 |
| SEASON2 |  | Mar Apr May |  | 0. 29368077 | 0. 45734478 | 69 | 0. 64 | 0. 5229 | 0. 05 | -0. 6187 | 1. 2061 |
| SEASON2 |  | SepOct Nov |  | 0. 00000000 |  |  |  |  |  |  |  |
| REGI ON |  |  | 24.26 N | 1. 62405383 | 0. 35802196 | 69 | 4. 54 | 0. 0001 | 0. 05 | 0. 9098 | 2. 3383 |
| REG ON |  |  | 29-33N | 3. 63088962 | 0. 43772466 | 69 | 8. 29 | 0. 0001 | 0. 05 | 2. 7577 | 4. 5041 |
| REGI ON |  |  | 34.35 N | 0. 00000000 |  | . |  |  |  |  |  |
| SEASON2*REG ON |  | DecJ anFeb | 24.26 N | 0. 86410285 | 0. 53667362 | 69 | 1. 61 | 0. 1119 | 0. 05 | -0. 2065 | 1. 9347 |
| SEASON2*REGI ON |  | DecJ anFeb | 29-33N | 0. 31462152 | 0. 65314549 | 69 | 0. 48 | 0. 6315 | 0. 05 | -0. 9884 | 1. 6176 |
| SEASON2*REG ON |  | DecJ anFeb | 34.35 N | 0. 00000000 |  |  |  |  |  |  |  |
| SEASON2*REG ON |  | J unj ul Aug | 24.26 N | -0. 72502077 | 0. 45365002 | 69 | -1. 60 | 0. 1146 | 0. 05 | -1. 6300 | 0. 1800 |
| SEASON2*REG ON |  | J unJ ul Aug | 29-33N | -0. 73118867 | 0. 56943291 | 69 | -1. 28 | 0. 2034 | 0. 05 | -1.8672 | 0. 4048 |
| SEASON2*REGI ON |  | J unj ul Aug | 34-35N | 0. 00000000 |  |  |  |  |  |  |  |
| SEASON2*REG ON |  | Mar Apr May | 24.26 N | 0. 92120931 | 0. 49791794 | 69 | 1. 85 | 0. 0686 | 0. 05 | -0. 0721 | 1. 9145 |
| SEASON2*REG ON |  | Mar Apr May | 29-33N | -0. 01722804 | 0. 62119193 | 69 | -0.03 | 0. 9780 | 0. 05 | -1. 2565 | 1. 2220 |
| SEASON2*REG ON |  | Mar Apr May | 3435 N | 0. 00000000 |  | . |  |  |  |  |  |
| SEASON2*REG ON |  | SepOct Nov | 24.26 N | 0. 00000000 |  |  |  |  |  |  |  |
| SEASON2*REG ON |  | SepOct Nov | 29-33N | 0. 00000000 |  |  |  |  |  |  |  |
| SEASON2*REG ON |  | SepOct Nov | $34.35 N$ | 0. 00000000 | . | . | . | . | . | . |  |


| Effect | FI SHYR2 | LSMEAN | Std Error | DF | t | Pr $>\|\mathrm{t}\|$ | Al pha | Lower | Upper |
| :--- | :--- | ---: | :--- | :--- | :--- | :--- | :--- | :--- | ---: |
|  |  |  |  |  |  |  |  |  |  |
| FI SHYR2 | 1991 | -1.82449403 | 0.42672697 | 69 | -4.28 | 0.0001 | 0.05 | -2.6758 | -0.9732 |
| FI SHYR2 | 1992 | -0.39709212 | 0.20532531 | 69 | -1.93 | 0.0572 | 0.05 | -0.8067 | 0.0125 |
| FI SHYR2 | 1993 | -0.27589398 | 0.16753649 | 69 | -1.65 | 0.1042 | 0.05 | -0.6101 | 0.0583 |
| FI SHYR2 | 1994 | -0.33248092 | 0.15227606 | 69 | -2.18 | 0.0324 | 0.05 | -0.6363 | -0.0287 |
| FI SHYR2 | 1995 | -0.28692348 | 0.15924372 | 69 | -1.80 | 0.0759 | 0.05 | -0.6046 | 0.0308 |
| FI SHYR2 | 1996 | -0.12064734 | 0.15061298 | 69 | -0.80 | 0.4259 | 0.05 | -0.4211 | 0.1798 |
| FI SHYR2 | 1997 | -0.44434640 | 0.15002127 | 69 | -2.96 | 0.0042 | 0.05 | -0.7436 | -0.1451 |
| FI SHYR2 | 1998 | -0.63042748 | 0.16518984 | 69 | -3.82 | 0.0003 | 0.05 | -0.9600 | -0.3009 |

Table 11. Analysis of positive catch rates of greater anberjack fromhandine vessel sin the $90^{\text {th }}-99^{\text {th }}$ percentile for positive catch rate or proportion of trips with greater amberjack in the United States EEZ in the Atlantic.

Description Val ue
Observations 2516. 000 Res Log Likelihood -4833.58 Akai ke's Inf ormation Criterion -4835. 58 Schwarz's Bayesian Criterion -4841.41 - 2 Res Log Likelihood 9667. 169

Information Criteria

| Better | Parns | q | $p$ | Al C | HQIC | BI C | CAI C |
| :--- | ---: | :--- | ---: | ---: | ---: | ---: | ---: |
|  |  |  |  |  |  |  |  |
| Larger | 2 | 2 | 0 | -4835.6 | -4837.7 | -4841.4 | -4842.4 |
| Larger | 21 | 2 | 19 | -4854.6 | -4876.8 | -4915.7 | -4926.2 |
| Smal Ier | 2 | 2 | 0 | 9671.2 | 9675.4 | 9682.8 | 9684.8 |
| Smaller | 21 | 2 | 19 | 9709.2 | 9753.6 | 9831.4 | 9852.4 |

Tests of Fixed Effects

| Source | NDF | DDF | Type 111 F | Pr $>$ F |
| :--- | ---: | ---: | ---: | ---: |
| FI SHYR2 |  |  |  |  |
| SEASON2 | 7 | 66 | 0.93 | 0.4877 |
| REG ON | 3 | 66 | 2.25 | 0.0902 |
| SEASON2*REG ON | 2 | 66 | 86.78 | 0.0001 |
|  | 6 | 66 | 1.51 | 0.1892 |

## Covariance Parameter Estimates (REM) Cov Parm Estimate <br> FI SHYR*SEASON*REG ON 0. 17655247 <br> Resi dual 2.64402346

Sol ution for Fixed Effects

Table 11. conti nued

| 1 NTERCEP |  |  |  | 2. 39591751 |  | 0. 37487176 | 66 | 6. 39 | 0. 0001 | 0. 05 | 1. 6475 | 3. 1444 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| FI SHYR2 |  | 1991 |  | 0. 61 | 68191 | 0. 43669487 | 8766 | 1. 41 | 0. 1619 | 0.05 | -0. 2542 | 1. 4896 |
| FI SHYR2 |  | 1992 |  | -0.00 | 92735 | 0. 25141439 | 3966 | -0. 02 | 0. 9813 | 0.05 | -0. 5079 | 0. 4960 |
| FI SHYR2 |  | 1993 |  | -0.08 | 32597 | 0. 23753645 | 4566 | -0. 36 | 0. 7206 | 0.05 | -0. 5596 | 0. 3889 |
| FI SHYR2 |  | 1994 |  | -0.19 | 72427 | 0. 23821992 | 9266 | -0. 83 | 0. 4095 | 0.05 | -0.6733 | 0. 2779 |
| FI SHYR2 |  | 1995 |  | 0. 236 | 84258 | 0. 24053442 | 4266 | 0. 98 | 0. 3284 | 0. 05 | -0. 2434 | 0. 7171 |
| FI SHYR2 |  | 1996 |  | -0.05 | 81778 | 0. 23655368 | 6866 | -0. 22 | 0. 8273 | 0.05 | -0. 5241 | 0. 4205 |
| FI SHYR2 |  | 1997 |  | -0.0861 | 14604 | 0. 23862268 | 686 | -0. 36 | 0. 7192 | 0.05 | -0. 5626 | 0. 3903 |
| FI SHYR2 |  | 1998 |  | 0. 000 | 00000 |  |  |  |  |  |  |  |
| SEASON2 |  | DecJ anFeb |  | 0. 108 | 03930 | 0. 51908095 | 9566 | 0. 21 | 0. 8358 | 0.05 | -0.9283 | 1. 1444 |
| SEASONZ |  | J unJ ul Aug |  | 0. 18 | 49622 | 0. 41912004 | 0466 | 0. 45 | 0. 6527 | 0.05 | -0. 6473 | 1. 0263 |
| SEASON2 |  | Mar Apr May |  | 0. 20 | 93591 | 0. 47210371 | 7166 | 0. 43 | 0. 6656 | 0. 05 | -0. 7376 | 1. 1475 |
| SEASON2 |  | SepOct Nov |  | 0. 000 | 00000 |  |  |  |  |  |  |  |
| REG ON |  |  | 24.26 N | 2. 089 | 31846 | 0. 38800321 | 2166 | 5. 38 | 0. 0001 | 0.05 | 1. 3146 | 2. 8640 |
| REG ON |  |  | 29-33N | 1. 43 | 92352 | 0. 39191222 | 2266 | 3. 66 | 0. 0005 | 0. 05 | 0. 6524 | 2. 2174 |
| REG ON |  |  | 34.35 N | 0. 00 | 00000 |  |  |  |  |  |  |  |
| SEASON2*REG ON |  | DecJ anFeb | eb 24.26 N | 0. 302 | 98976 | 0. 58333961 | $61 \quad 66$ | 0. 52 | 0. 6052 | 0. 05 | -0.8617 | 1. 4677 |
| SEASON2*REG ON |  | DecJ anFeb | eb 29-33N | 0. 08 | 11037 | 0. 59184638 | 3866 | 0. 15 | 0. 8848 | 0. 05 | -1. 0955 | 1. 2678 |
| SEASON2*REG ON |  | DecJ anFeb | eb 3435 N | 0. 00 | 00000 |  |  |  |  |  |  |  |
| SEASON2*REG ON |  | J unJ ul Aug | gi 2426 N | 0. 102 | 48058 | 0. 50351230 | 3066 | 0. 20 | 0. 8393 | 0. 05 | -0.9028 | 1. 1078 |
| SEASON2*REG ON |  | $J$ unJ ul Aug | 29.33N | -0.66 | 43579 | 0. 50959332 | 3266 | - 1.30 | 0. 1968 | 0.05 | -1. 6819 | 0. 3530 |
| SEASON2*REG ON |  | J unJ ul Aug | ung 3435 N | 0. 00 | 00000 |  |  |  |  |  |  |  |
| SEASON2*REG ON |  | Mar Apr May | ay 2426 N | 0. 772 | 22870 | 0. 53963626 | $26 \quad 66$ | 1. 43 | 0. 1572 | 0. 05 | -0. 3054 | 1. 8494 |
| SEASON2*REGI ON |  | Mar Apr May | ay 29.33 N | -0.0780 | 94883 | 0. 55192403 | 0366 | -0. 14 | 0. 8867 | 0. 05 | -1. 1809 | 1. 0230 |
| SEASON2*REG ON |  | Mar Apr May | ay 3435 N | 0. 00 | 00000 |  |  |  |  |  |  |  |
| SEASON2*REGI ON |  | SepOct Nov | ov 2426 N | 0. 00 | 00000 |  | . |  |  |  |  |  |
| SEASON2*REG ONSEASON2*REGI ON |  | SepOct Nov | ov 29-33N | 0. 00 | 00000 |  | . |  |  |  |  |  |
|  |  | SepOct Nov | ov 3435 N | 0. 00 | 00000 |  | . . | . | . |  |  |  |
| Least Squares Means |  |  |  |  |  |  |  |  |  |  |  |  |
| Effect | FI SHYR2 | LSMEAN | St d Error | DF | t | Pr $>\|t\|$ A | Al pha | Lower | Upper |  |  |  |
| FI SHYR2 | 1991 | 4. 35731667 | 0. 40502632 | 66 | 10. 76 | 0. 0001 | 0. 05 | 3. 5487 | 5. 1660 |  |  |  |
| FI SHYR2 | 1992 | 3. 73370741 | 0. 18689290 | 66 | 19. 98 | 0. 0001 | 0. 05 | 3. 3606 | 4. 1069 |  |  |  |
| FI SHYR2 | 1993 | 3. 65430879 | 0. 16553883 | 66 | 22. 08 | 0. 0001 | 0.05 | 3. 3238 | 3. 9848 |  |  |  |
| FI SHYR2 | 1994 | 3. 54191049 | 0. 16427999 | 66 | 21. 56 | 0. 0001 | 0. 05 | 3. 2139 | 3. 8699 |  |  |  |
| FI SHYR2 | 1995 | 3. 97647734 | 0. 16672717 | 66 | 23. 85 | 0. 0001 | 0.05 | 3. 6436 | 4. 3094 |  |  |  |
| FI SHYR2 | 1996 | 3. 68781698 | 0. 16450706 | 66 | 22. 42 | 0. 0001 | 0.05 | 3. 3594 | 4. 0163 |  |  |  |
| FI SHYR2 | 1997 | 3. 65348872 | 0. 17075389 | 66 | 21. 40 | 0. 0001 | 0. 05 | 3. 3126 | 3. 9944 |  |  |  |
| FI SHYR2 | 1998 | 3. 73963476 | 0. 17944356 | 66 | 20. 84 | 0. 0001 | 0. 05 | 3. 3814 | 4. 0979 |  |  |  |

Table 12. Index of abundance of Atlantic greater amberjack from handline vessels in the $90^{\text {th }}$ - $99^{t h}$ percentile for positive catch rate or proportion of trips with greater amberjack. CPUE is the standardized catch rates on trips with greater amberjack, PPOS is the proportion positive, I NDEX is the standardized index and SE_I and CV_l are the standard error and coefficient of variation about the index

| FI SHYR2 | CPUE | PPOS | BC_CPU | GC | BC_POS | I NDEX | SE_I | CV_I |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  |  |  |  |  |  |  |  |  |
| 1991 | 84.7190 | 0.13890 | 255.758 | 3.27696 | 0.13890 | 35.5236 | 23.1021 | 0.65033 |
| 1992 | 42.5709 | 0.40201 | 150.565 | 3.59910 | 0.40201 | 60.5287 | 18.0386 | 0.29802 |
| 1993 | 39.1739 | 0.43146 | 139.578 | 3.61219 | 0.43146 | 60.2223 | 16.6506 | 0.27648 |
| 1994 | 35.0020 | 0.41764 | 124.764 | 3.61291 | 0.41764 | 52.1061 | 14.2605 | 0.27368 |
| 1995 | 54.0752 | 0.42876 | 192.597 | 3.61150 | 0.42876 | 82.5774 | 22.7552 | 0.27556 |
| 1996 | 40.5019 | 0.46987 | 144.358 | 3.61278 | 0.46987 | 67.8301 | 18.3756 | 0.27091 |
| 1997 | 39.1761 | 0.39071 | 139.346 | 3.60913 | 0.39071 | 54.4431 | 15.1309 | 0.27792 |
| 1998 | 42.7656 | 0.34741 | 151.659 | 3.60384 | 0.34741 | 52.6885 | 15.1951 | 0.28840 |

Table 13. Anal ysis of proportions of trips which caught greater anderjack fromhandline vessel s in the upper quartile for positive catch rate or proportion of trips with greater anberjack in the Atlantic Ocean in the United States EEZ.

## Cl ass Level I nfornmtion



Mbdel fitting Information for proportion positive

| Description | Val ue |
| :--- | ---: |
|  |  |
| Observations | 84.0000 |
| Res Log Li kel i hood | -29.5827 |
| Akai ke' s I nf or mation Criterion | -31.5827 |
| Schwarz' s Bayesian Criterion | -33.7723 |
| -2 Res Log Li keli hood | 59.1654 |

nformation Criteri

| Better | Parns | q | p | Al C | HQI C | BI C | CAI C |
| :--- | ---: | :--- | ---: | ---: | ---: | ---: | ---: |
|  |  |  |  |  |  |  |  |
| Larger | 2 | 2 | 0 | -31.6 | -32.4 | -33.8 | -34.8 |
| Larger | 20 | 2 | 18 | -49.6 | -58.2 | -71.5 | -81.5 |
| Smal I er | 2 | 2 | 0 | 63.2 | 64.9 | 67.5 | 69.5 |
| Snall er | 20 | 2 | 18 | 99.2 | 116.5 | 143.0 | 163.0 |

Tests of Fixed Effects

| Source | NDF | DDF | Type III Chi Sq | Type $\\|\\|$ Fr $>$ Chi Sq | Pr $>F$ |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| FI SHYR2 |  |  |  |  |  |  |  |
| SEASON2 | 6 | 66 | 37.05 | 6.17 | 0.0001 | 0.0001 |  |
| REG ON | 3 | 66 | 9.17 | 3.06 | 0.0271 | 0.0343 |  |
| SEASON2*REG ON | 2 | 66 | 696.37 | 348.18 | 0.0001 | 0.0001 |  |
| ( | 66 | 49.68 | 8.28 | 0.0001 | 0.0001 |  |  |

## Covariance Parameter Estimates (REM) <br> Cov Parm <br> SHYR*SEASON*REGI ON 0.02215622 <br> Resi dual 1. 36240272

Sol ution for Fixed Effects

Table 13. continued
I NTERCEPT
FI SHYR2 1992

FI SHYR2 1993
FI SHYR2 1995
FI SHYR2 1996

FI SHYR2 1997
SEASON2 1998

SEASO
SEASO
SEASONZ
SEASONZ
REG ON
REG ON
REG ON
SEASON2*REG ON
SEASON2*REG ON SEASON2*REG ON SEASON2*REG ON SEASON2*REG ON SEASON2*REG O SEASONZ*REG O SEASON2*REG ON SEASON2*REG ON

Table 13. conti nued.

# SEASON2*REG ON SEASON2*REG ON 

 SEASON2*REG ONSepOct Nov $24.26 \mathrm{~N} \quad 0.00000000$ SepOct Nov $29.33 \mathrm{~N} \quad 0.00000000$ $\begin{array}{lll}\text { SepOct Nov } & 34-35 \mathrm{~N} & 0.00000000\end{array}$

Least Squares Means

| Effect | FI SHYR2 | LSMEAN | Std Error | DF | t | Pr $>\|\mathrm{t}\|$ | Al pha | Lower | Upper |
| :--- | :--- | ---: | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  |  |  |  |  |  |  |  |  |  |
| FI SHYR2 | 1992 | -1.25251996 | 0.10211779 | 66 | -12.27 | 0.0001 | 0.05 | -1.4564 | -1.0486 |
| FI SHYR2 | 1993 | -1.05785922 | 0.08383715 | 66 | -12.62 | 0.0001 | 0.05 | -1.2252 | -0.8905 |
| FI SHYR2 | 1994 | -0.90529174 | 0.07370699 | 66 | -12.28 | 0.0001 | 0.05 | -1.0525 | -0.7581 |
| FI SHYR2 | 1995 | -0.87778505 | 0.07394219 | 66 | -11.87 | 0.0001 | 0.05 | -1.0254 | -0.7302 |
| FI SHYR2 | 1996 | -0.78872313 | 0.07199429 | 66 | -10.96 | 0.0001 | 0.05 | -0.9325 | -0.6450 |
| FI SHYR2 | 1997 | -1.03571427 | 0.07135239 | 66 | -14.52 | 0.0001 | 0.05 | -1.1782 | -0.8933 |
| FI SHYR2 | 1998 | -1.30650516 | 0.07751767 | 66 | -16.85 | 0.0001 | 0.05 | -1.4613 | -1.1517 |

Table 14. Anal ysis of positive catch rates of greater anberjack fromhandline vessels in the upper quartile for positive catch rate or proportion of trips with greater amberjack in the United States EEZ in the Atlantic.

Description Val ue
Observati ons 4737. 000 Res Log Li kel i hood -9359. 40
Akai ke' s I nf ormation Criterion -9361. 40 Schwarz's Bayesian Criterion -9367.86 - 2 Res Log Li kelihood 18718. 80

Information Criteria
Better Parns q p AIC HQIC BIC CAIC

| Lar ger | 2 | 2 | 0 | -9361.4 | -9363.7 | -9367.9 | -9368.9 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Larger | 20 | 2 | 18 | -9379.4 | -9402.1 | -9444.0 | -9454.0 |
| Smal I er | 2 | 2 | 0 | 18722.8 | 18727.3 | 18735.7 | 18737.7 |
| Small er | 20 | 2 | 18 | 18758.8 | 18804.2 | 18888.0 | 18908.0 |

Tests of Fixed Effects

| Source | NDF | DDF | Type 1111 F | $\mathrm{Pr}>\mathrm{F}$ |
| :--- | ---: | ---: | ---: | ---: |
|  |  |  |  |  |
| FI SHYR2 | 6 | 18 | 3.72 | 0.0139 |
| SEASON2 | 3 | 18 | 4.30 | 0.0188 |
| REGI ON | 2 | 4701 | 352.06 | 0.0001 |
| SEASON2*REG ON | 6 | 4701 | 6.93 | 0.0001 |

Covariance Parameter Estimates (REM)

| Cov Parm | Estimate |
| :--- | ---: |
|  |  |
| FI SHYR2*SEASON2 | 0.01317355 |
| Resi dual | 3.01818937 |

Sol ution for Fixed Effects

Table 14. conti nued.


Table 15. Index of abundance of Atlantic greater anberjack fromhandline vessels in the upper quartile for positive catch rate or proportion of trips with greater amberjack. CPUE is the standardized catch rates on trips with greater amberjack, PPOS is the proportion positive, INDEX is the standardized index and SE_I and CVI are the standard error and coefficient of variation about the index.

| FI SHYR2 | CPUE | PPOS | BC_CPU | GC | BC_POS | I NDEX | SE_I | CV_I |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1992 | 34.0089 | 0.22226 | 132.734 | 3.93201 | 0.22226 | 29.5021 | 15.5259 | 0.52626 |
| 1993 | 28.2382 | 0.25772 | 114.798 | 4.08616 | 0.25772 | 29.5856 | 11.3573 | 0.38388 |
| 1994 | 36.8301 | 0.28796 | 150.042 | 4.09019 | 0.28796 | 43.2067 | 16.4397 | 0.38049 |
| 1995 | 36.1333 | 0.29364 | 147.213 | 4.09032 | 0.29364 | 43.2273 | 16.4434 | 0.38040 |
| 1996 | 26.0123 | 0.31244 | 106.018 | 4.09103 | 0.31244 | 33.1247 | 12.5778 | 0.37971 |
| 1997 | 26.4804 | 0.26198 | 107.914 | 4.09082 | 0.26198 | 28.2711 | 10.7509 | 0.38028 |
| 1998 | 24.4842 | 0.21307 | 99.630 | 4.08796 | 0.21307 | 21.2284 | 8.1291 | 0.38293 |

Table 16. Anal ysis of the proportion of full day headboat trips catching greater anberjack in the Atlantic.

| Cl ass | Level s | Val ues |  |
| :---: | :---: | :---: | :---: |
| FI SHYEAR | 22 | 1976197719781979 | 19801981 |
|  |  | 1982198319841985 | 19861987 |
|  |  | 1988198919901991 | 19921993 |
|  |  | 1994199519961997 |  |
| SEASON | 4 | Apr MayJ un J anFebMar J | J ul AugSep |
|  |  | Oct NovDec |  |
| REGI ON | 4 | 24-25.4N 25. 5-28.6N |  |
|  |  | 28.7-30.6N 30.7-35N |  |
| I NSHORPL | 2 | 01 |  |

Covariance Parameter Estimates (REM)
Estimate

FI SH*SEAS*REG *I NSHO 0.59830579
Resi dual 1. 45801288
Mbdel Fitting Information for proportion positive
Description
Val ue
636. 0000

Res Log Li kel i hood -986. 820
Akai ke's Inf ormation Criterion -988. 820
Schwarz's Bayesian Criterion -993. 224
2 Res Log likelihood
1973.641

| Better | Parns | Information Criteria |  |  |  | BI C | CAI C |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | q | p | Al C | Hal C |  |  |
| Larger | 2 | 2 | 0 | -988. 8 | -990. 5 | -993. 2 | -994. 2 |
| Larger | 34 | 2 | 32 | -1020. 8 | -1050. 0 | -1095. 7 | -1112. 7 |
| Small er | 2 | 2 | 0 | 1977.6 | 1981. 1 | 1986.4 | 1988.4 |
| Small er | 34 | 2 | 32 | 2041.6 | 2099. 9 | 2191. 4 | 2225.4 |


| Source | NDF | DDF | Type III Chi Sq | Type III F | Pr $>$ Chi Sq | Pr > F |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
|  |  |  |  |  |  |  |
| FI SHYEAR | 21 | 604 | 145.31 | 6.92 | 0.0001 | 0.0001 |
| SEASON | 3 | 604 | 22.32 | 7.44 | 0.0001 | 0.0001 |
| REGI ON | 3 | 604 | 275.57 | 91.86 | 0.0001 | 0.0001 |
| I NSHORPL | 1 | 604 | 52.94 | 52.94 | 0.0001 | 0.0001 |
| SEASON* NSHORPL | 3 | 604 | 2.23 | 0.74 | 0.5265 | 0.5269 |

Sol ution for Fixed Effects

Table 16. conti nued

| Effect | FI SHYEAR | SEASON | REG ON | 1 NSHORPL | Estimate | St d Error | DF | t | $\operatorname{Pr}>\|t\|$ | Al pha | Lower | Upper |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| I NTERCEPT |  |  |  |  | -1. 12349358 | 0. 25284719 | 604 | -4. 44 | 0. 0001 | 0. 05 | - 1. 6201 | -0. 6269 |
| FI SHYEAR | 1976 |  |  |  | 0. 66521915 | 0. 32313991 | 604 | 2. 06 | 0. 0400 | 0.05 | 0. 0306 | 1. 2998 |
| FI SHYEAR | 1977 |  |  |  | 0. 59327828 | 0. 32317403 | 604 | 1. 84 | 0. 0669 | 0.05 | -0.0414 | 1. 2280 |
| FI SHYEAR | 1978 |  |  |  | 1. 37584146 | 0. 29692796 | 604 | 4.63 | 0. 0001 | 0.05 | 0. 7927 | 1. 9590 |
| FI SHYEAR | 1979 |  |  |  | 1. 44550375 | 0. 28907735 | 604 | 5. 00 | 0. 0001 | 0.05 | 0. 8778 | 2. 0132 |
| FI SHYEAR | 1980 |  |  |  | 1. 06458597 | 0. 28961701 | 604 | 3. 68 | 0. 0003 | 0.05 | 0. 4958 | 1. 6334 |
| FI SHYEAR | 1981 |  |  |  | 1. 20457285 | 0. 28875466 | 604 | 4. 17 | 0. 0001 | 0.05 | 0. 6375 | 1. 7717 |
| FI SHYEAR | 1982 |  |  |  | 1. 32576477 | 0. 29189779 | 604 | 4.54 | 0. 0001 | 0.05 | 0. 7525 | 1. 8990 |
| FI SHYEAR | 1983 |  |  |  | 1. 42876079 | 0. 29191615 | 604 | 4. 89 | 0. 0001 | 0.05 | 0. 8555 | 2. 0021 |
| FI SHYEAR | 1984 |  |  |  | 0. 98440360 | 0. 29065417 | 604 | 3. 39 | 0. 0008 | 0.05 | 0. 4136 | 1. 5552 |
| FI SHYEAR | 1985 |  |  |  | 0. 99361698 | 0. 28892729 | 604 | 3. 44 | 0. 0006 | 0.05 | 0. 4262 | 1. 5610 |
| FI SHYEAR | 1986 |  |  |  | 1. 11037145 | 0. 28662038 | 604 | 3. 87 | 0. 0001 | 0.05 | 0. 5475 | 1. 6733 |
| FI SHYEAR | 1987 |  |  |  | 1. 20482677 | 0. 28893582 | 604 | 4. 17 | 0. 0001 | 0.05 | 0. 6374 | 1. 7723 |
| FI SHYEAR | 1988 |  |  |  | 0. 78992575 | 0. 29525010 | 604 | 2. 68 | 0. 0077 | 0.05 | 0. 2101 | 1. 3698 |
| FI SHYEAR | 1989 |  |  |  | 0. 82251665 | 0. 29113025 | 604 | 2. 83 | 0. 0049 | 0.05 | 0. 2508 | 1. 3943 |
| FI SHYEAR | 1990 |  |  |  | 0. 16427913 | 0. 29793305 | 604 | 0. 55 | 0. 5816 | 0.05 | -0. 4208 | 0. 7494 |
| FI SHYEAR | 1991 |  |  |  | 0. 52598373 | 0. 29379480 | 604 | 1. 79 | 0. 0739 | 0.05 | -0.0510 | 1. 1030 |
| FI SHYEAR | 1992 |  |  |  | 0. 41189326 | 0. 29083629 | 604 | 1. 42 | 0. 1572 | 0.05 | -0.1593 | 0. 9831 |
| FI SHYEAR | 1993 |  |  |  | 0. 31344176 | 0. 29866424 | 604 | 1. 05 | 0. 2944 | 0.05 | -0.2731 | 0. 9000 |
| FI SHYEAR | 1994 |  |  |  | 0. 28904997 | 0. 29695369 | 604 | 0. 97 | 0. 3308 | 0.05 | -0. 2941 | 0. 8722 |
| FI SHYEAR | 1995 |  |  |  | -0. 24352314 | 0. 31209069 | 604 | -0.78 | 0. 4355 | 0.05 | -0.8564 | 0. 3694 |
| FI SHYEAR | 1996 |  |  |  | 0. 11839574 | 0. 30356622 | 604 | 0. 39 | 0. 6967 | 0.05 | -0.4778 | 0. 7146 |
| FI SHYEAR | 1997 |  |  |  | 0. 00000000 |  |  |  |  |  |  |  |
| SEASON |  | Apr MayJ un |  |  | -0.03097359 | 0. 13808804 | 604 | -0. 22 | 0. 8226 | 0.05 | - 0.3022 | 0. 2402 |
| SEASON |  | J anFebMar |  |  | -0. 57498712 | 0. 14644331 | 604 | -3. 93 | 0. 0001 | 0.05 | - 0.8626 | -0. 2874 |
| SEASON |  | J ul AugSep |  |  | -0.09376446 | 0. 14048917 | 604 | -0.67 | 0. 5048 | 0.05 | -0. 3697 | 0. 1821 |
| SEASON |  | Oct NovDec |  |  | 0. 00000000 |  |  |  |  |  |  |  |
| REGI ON |  |  | 24-25.4N |  | -1. 36325737 | 0. 13164354 | 604 | 10. 36 | 0.0001 | 0.05 | -1. 6218 | -1. 1047 |
| REGI ON |  |  | 25. 5-28.6N |  | -1. 38434368 | 0. 11760003 | 604 | 11. 77 | 0. 0001 | 0. 05 | 1. 6153 | 1. 1534 |
| REGI ON |  |  | 28.7-30.6N |  | 0. 14391753 | 0. 09574681 | 604 | 1. 50 | 0. 1333 | 0.05 | -0.0441 | 0. 3320 |
| REGI ON |  |  | 30.7-35N |  | 0. 00000000 |  |  |  |  |  |  |  |
| 1 NSHORPL |  |  |  | 0 | -0. 79748339 | 0. 17361097 | 604 | -4.59 | 0. 0001 | 0.05 | -1. 1384 | -0.4565 |
| I NSHORPL |  |  |  | 1 | 0. 00000000 |  |  |  |  |  |  |  |
| SEASON* NSHORPL |  | Apr MayJ un |  | 0 | 0. 29851697 | 0. 22817436 | 604 | 1. 31 | 0. 1913 | 0.05 | - 0.1496 | 0. 7466 |
| SEASON* ${ }^{\text {I NSHORPL }}$ |  | Apr MayJ un |  | 1 | 0. 00000000 |  |  |  |  |  |  |  |
| SEASON* 1 NSHORPL |  | J anFebMar |  | 0 | 0. 29883029 | 0. 24958985 | 604 | 1. 20 | 0. 2317 | 0. 05 | - 0. 1913 | 0. 7890 |
| SEASON* ${ }^{\text {I }}$ NSHOPPL |  | J anFebMar |  | 1 | 0. 00000000 |  |  |  |  |  |  |  |
| SEASON*I NSHORPL |  | J ul AugSep |  | 0 | 0. 13666297 | 0. 23015538 | 604 | 0. 59 | 0. 5529 | 0. 05 | -0.3153 | 0. 5887 |
| SEASON* ${ }^{\text {NSHORPL }}$ |  | J ul AugSep |  | 1 | 0. 00000000 |  |  |  |  |  | . |  |
| SEASON* I NSHORPL |  | Oct NovDec |  | 0 | 0. 00000000 |  |  |  |  |  |  |  |
| SEASON* ${ }^{\text {NSHORPL }}$ |  | Oct NovDec |  | 1 | 0. 00000000 | . |  | . | . | . | . |  |


| Effect | FI SHYEAR | LSMEAN | Std Error | DF | t | $\operatorname{Pr}>\|t\|$ | Al pha | Lower | Upper |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| FI SHYEAR | 1976 | - 1. 59111702 | 0. 22938800 | 604 | -6. 94 | 0. 0001 | 0.05 | -2. 0416 | -1. 1406 |
| FI SHYEAR | 1977 | - 1. 66305789 | 0. 22932360 | 604 | -7. 25 | 0.0001 | 0.05 | -2. 1134 | - 1. 2127 |
| FI SHYEAR | 1978 | -0. 88049472 | 0. 18752473 | 604 | -4. 70 | 0. 0001 | 0.05 | -1. 2488 | -0. 5122 |
| FI SHYEAR | 1979 | - 0.81083242 | 0. 17259415 | 604 | -4. 70 | 0. 0001 | 0.05 | -1. 1498 | -0. 4719 |
| FI SHYEAR | 1980 | - 1. 19175021 | 0. 17372769 | 604 | -6. 86 | 0. 0001 | 0.05 | -1. 5329 | -0. 8506 |
| FI SHYEAR | 1981 | - 1. 05176332 | 0. 17130183 | 604 | -6. 14 | 0. 0001 | 0.05 | -1. 3882 | -0. 7153 |
| FI SHYEAR | 1982 | -0.93057141 | 0. 17717425 | 604 | -5. 25 | 0. 0001 | 0.05 | -1. 2785 | -0. 5826 |
| FI SHYEAR | 1983 | -0. 82757539 | 0. 17832982 | 604 | -4. 64 | 0. 0001 | 0.05 | -1. 1778 | -0. 4774 |
| FI SHYEAR | 1984 | -1. 27193257 | 0. 17681728 | 604 | -7. 19 | 0. 0001 | 0.05 | -1. 6192 | -0. 9247 |
| FI SHYEAR | 1985 | -1. 26271919 | 0. 17349824 | 604 | -7. 28 | 0. 0001 | 0.05 | -1. 6035 | -0.9220 |
| FI SHYEAR | 1986 | - 1. 14596473 | 0. 16864509 | 604 | -6. 80 | 0. 0001 | 0. 05 | -1. 4772 | - 0.8148 |
| FI SHYEAR | 1987 | - 1. 05150940 | 0. 17397958 | 604 | -6. 04 | 0. 0001 | 0.05 | -1. 3932 | -0. 7098 |
| FI SHYEAR | 1988 | - 1. 46641043 | 0. 18420369 | 604 | -7. 96 | 0. 0001 | 0.05 | -1. 8282 | - 1. 1047 |
| FI SHYEAR | 1989 | - 1. 43381952 | 0. 17664831 | 604 | -8. 12 | 0. 0001 | 0.05 | -1. 7807 | -1. 0869 |
| FI SHYEAR | 1990 | -2. 09205705 | 0. 18845973 | 604 | -11. 10 | 0. 0001 | 0.05 | -2. 4622 | -1. 7219 |
| FI SHYEAR | 1991 | -1. 73035245 | 0. 18128319 | 604 | -9. 55 | 0. 0001 | 0.05 | -2. 0864 | -1. 3743 |
| FI SHYEAR | 1992 | - 1. 84444291 | 0. 17666911 | 604 | -10.44 | 0. 0001 | 0.05 | - 2. 1914 | -1. 4975 |
| FI SHYEAR | 1993 | - 1. 94289441 | 0. 18991864 | 604 | -10. 23 | 0. 0001 | 0. 05 | -2. 3159 | -1. 5699 |
| FI SHYEAR | 1994 | - 1. 96728621 | 0. 18721092 | 604 | -10. 51 | 0. 0001 | 0. 05 | -2. 3349 | -1. 5996 |
| FI SHYEAR | 1995 | - 2. 49985931 | 0. 21204049 | 604 | -11. 79 | 0. 0001 | 0. 05 | -2. 9163 | -2. 0834 |
| FI SHYEAR | 1996 | -2. 13794043 | 0. 19857801 | 604 | -10.77 | 0. 0001 | 0.05 | -2. 5279 | -1. 7480 |
| FI SHYEAR | 1997 | -2. 25633618 | 0. 23425202 | 604 | -9. 63 | 0. 0001 | 0.05 | -2. 7164 | -1. 7963 |

Table 17. Anal ysis greater anberjack catch per successful trip on full day headboats in the Atlantic


| FI SHYEAR | 1976 |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| FI SHYEAR | 1977 |  |  |  |
| FI SHYEAR | 1978 |  |  |  |
| FI SHYEAR | 1979 |  |  |  |
| FI SHYEAR | 1980 |  |  |  |
| FI SHYEAR | 1981 |  |  |  |
| FI SHYEAR | 1982 |  |  |  |
| FI SHYEAR | 1983 |  |  |  |
| FI SHYEAR | 1984 |  |  |  |
| FI SHYEAR | 1985 |  |  |  |
| FI SHYEAR | 1986 |  |  |  |
| FI SHYEAR | 1987 |  |  |  |
| FI SHYEAR | 1988 |  |  |  |
| FI SHYEAR | 1989 |  |  |  |
| FI SHYEAR | 1990 |  |  |  |
| FI SHYEAR | 1991 |  |  |  |
| FI SHYEAR | 1992 |  |  |  |
| FI SHYEAR | 1993 |  |  |  |
| FI SHYEAR | 1994 |  |  |  |
| FI SHYEAR | 1995 |  |  |  |
| FI SHYEAR | 1996 |  |  |  |
| FI SHYEAR | 1997 |  |  |  |
| SEASON |  | Apr MayJ un |  |  |
| SEASON |  | J anFebMar |  |  |
| SEASON |  | J ul AugSep |  |  |
| SEASON |  | Oct NovDec |  |  |
| REGI ON |  |  | 24-25. 4 N |  |
| REGI ON |  |  | 25. 5-28.6N |  |
| REGI ON |  |  | 28.7-30.6N |  |
| REGI ON |  |  | 30.7-35N |  |
| 1 NSHORPL |  |  |  | 0 |
| I NSHORPL |  |  |  | 1 |
| SEASON* ${ }^{\text {NSHORPL }}$ |  | Apr MayJ un |  | 0 |
| SEASON*I NSHORPL |  | Apr MayJ un |  | 1 |
| SEASON* I NSHORPL |  | J anFebMar |  | 0 |
| SEASON*I NSHORPL |  | J anFebMar |  | 1 |
| SEASON*I NSHORPL |  | J ul AugSep |  | 0 |
| SEASON*I NSHORPL |  | J ul AugSep |  | 1 |
| SEASON* ${ }^{\text {NSHORPL }}$ |  | Oct Nov Dec |  | 0 |
| SEASON*I NSHORPL |  | Oct NovDec |  | 1 |
| REGI ON*I NSHORPL |  |  | 24-25.4N | 0 |
| REGI ON*I NSHORPL |  |  | 24-25.4N | 1 |
| REGI ON*I NSHORPL |  |  | 25. 5-28.6N | 0 |
| REGI ON*I NSHORPL |  |  | 25. 5-28.6N | 1 |
| REGI ON*I NSHORPL |  |  | 28.7-30.6N | 0 |
| REGI ON*I NSHORPL |  |  | 28.7-30.6N | 1 |
| REGI ON*I NSHORPL |  |  | 30.7-35N | 0 |
| REGI ON*I NSHORPL |  |  | 30.7-35N | 1 |

$\begin{array}{lllllll}0.23907278 & 0.17497613 & 505 & 1.37 & 0.1724 & 0.05 & -0.1047 \\ 0.5828\end{array}$
0.23186154 $\begin{array}{llllllll}0.02186154 & 0.17569103 & 505 & 0.12 & 0.9010 & 0.05 & -0.3233 & 0.3670 \\ 0.00499205 & 0.16161827 & 505 & -0.03 & 0.9754 & 0.05 & -0.3225 & 0.3125\end{array}$ $\begin{array}{llllllll}0.00499205 & 0.16161827 & 505 & -0.03 & 0.9754 & 0.05 & -0.3225 & 0.3125\end{array}$ $\begin{array}{llllllll}0.02348899 & 0.15971175 & 505 & 0.15 & 0.8831 & 0.05 & -0.2903 & 0.3373 \\ 0.29548736 & 0.16161167 & 505 & 1.83 & 0.0681 & 0.05 & -0.0220 & 0.6130\end{array}$ $\begin{array}{llllllll}0.29548736 & 0.16161167 & 505 & 1.83 & 0.0681 & 0.05 & -0.0220 & 0.6130\end{array}$ $\begin{array}{llllllll}0.47978009 & 0.16021523 & 505 & 2.99 & 0.0029 & 0.05 & 0.1650 & 0.7946\end{array}$ $\begin{array}{llllllll}0.21980532 & 0.16114595 & 505 & 1.36 & 0.1732 & 0.05 & -0.0968 & 0.5364\end{array}$ $\begin{array}{llllllll}0.03203377 & 0.15950647 & 505 & 0.20 & 0.8409 & 0.05 & -0.2813 & 0.3454\end{array}$ $\begin{array}{lllllll}0.02513226 & 0.16065355 & 505-0.16 & 0.8758 & 0.05-0.3408 & 0.2905\end{array}$ $\begin{array}{llllllll}0.10528290 & 0.15965483 & 505 & 0.66 & 0.5099 & 0.05 & -0.2084 & 0.4190\end{array}$ $\begin{array}{lllllll}0.04772195 & 0.15830275 & 505 & 0.30 & 0.7632 & 0.05 & -0.2633 \\ 0.3587\end{array}$ $\begin{array}{lllllll}0.22612974 & 0.15906654 & 505 & 1.42 & 0.1558 & 0.05 & -0.0864 \\ 0.5386\end{array}$ $\begin{array}{llllllll}0.14496015 & 0.16294358 & 505 & 0.89 & 0.3741 & 0.05-0.1752 & 0.4651\end{array}$ $\begin{array}{llllllll}0.20941977 & 0.16195307 & 505 & 1.29 & 0.1966 & 0.05 & -0.1088 & 0.5276\end{array}$ $\begin{array}{llllllll}0.03317395 & 0.16919319 & 505 & 0.20 & 0.8446 & 0.05 & -0.2992 & 0.3656\end{array}$ $\begin{array}{lllllll}0.08114795 & 0.16408936 & 505 & 0.49 & 0.6211 & 0.05 & -0.2412\end{array} 0.4035$ $\begin{array}{llllllll}0.04262583 & 0.16219498 & 505 & 0.26 & 0.7928 & 0.05 & -0.2760 & 0.3613\end{array}$ $\begin{array}{llllllll}0.11099462 & 0.16736332 & 505 & 0.66 & 0.5075 & 0.05 & -0.2178 & 0.4398\end{array}$ 0.18593043 0. 16732754505 1.11 $\quad 0.2670 \quad 0.05-0.1428 \quad 0.5147$ $\begin{array}{llllllll}0.05517173 & 0.17952458 & 505 & 0.31 & 0.7587 & 0.05 & -0.2975 & 0.4079\end{array}$ $\begin{array}{llllllll}0.15501258 & 0.17267229 & 505 & 0.90 & 0.3698 & 0.05 & -0.1842 & 0.4943\end{array}$ 0. 00000000
$\begin{array}{llllllll}0.01260266 & 0.07314836 & 505 & 0.17 & 0.8633 & 0.05 & -0.1311 & 0.1563\end{array}$ 0.09185946 0.07997070 505-1.15 $0.2512 \quad 0.05-0.2490 \quad 0.0653$ 0.09185946 0.07997070 $505-1.15 \quad 0.2512$ 0.05-0.2490 0.0653 0. 0456000000
0. 00000000
$\begin{array}{lllll}0.37824293 & 0.09226893 & 505-4.10 & 0.0001 & 0.05-0.5595-0.1970\end{array}$
$0.54633774 \quad 0.07789239505-7.01 \quad 0.0001 \quad 0.05-0.6994-0.3933$
$\begin{array}{lllllll}0.00776815 & 0.06412718 & 505-0.12 & 0.9036 & 0.05-0.1338 & 0.1182\end{array}$ 0. 00000000
$\begin{array}{llllll}0.34837088 & 0.11487573 & 505-3.03 & 0.0025 & 0.05-0.5741-0.1227\end{array}$ 0. 00000000
$\begin{array}{llllllll}0.28644954 & 0.12379935 & 505 & 2.31 & 0.0211 & 0.05 & 0.0432 & 0.5297\end{array}$ o. 00000000
$\begin{array}{lllllllll}0.12770443 & 0.14146543 & 505 & 0.90 & 0.3671 & 0.05 & -0.1502 & 0.4056\end{array}$ 0. 00000000
$\begin{array}{lllllllll}0.09039852 & 0.12623129 & 505 & 0.72 & 0.4742 & 0.05 & -0.157 \dot{6} & 0.3384\end{array}$ 0. 00000000
0. 00000000
0. 00000000
$\begin{array}{llllllll}0.43541221 & 0.17185755 & 505 & 2.53 & 0.0116 & 0.05 & 0.0978 & 0.7731\end{array}$
0. 00000000
$\begin{array}{lllllll}0.39422457 & 0.14512734 & 505 & 2.72 & 0.0068 & 0.05 & 0.1091\end{array} 0.6794$ 0. 00000000
$\begin{array}{lllllll}0.34368319 & 0.10168813 & 505 & 3.38 & 0.0008 & 0.05 & 0.1439 \\ 0.5435\end{array}$ 0. 00000000
0. 00000000
0. 00000000

Table 17. conti nued
Least Squares Means

| Effect | FI SHYEAR | LSMEAN | Std Error | DF | t | Pr $>\|t\|$ | Al pha | Lower | Upper |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| FI SHYEAR | 1976 | 0. 90410826 | 0. 11998098 | 505 | 7. 54 | 0. 0001 | 0. 05 | 0. 6684 | 1. 1398 |
| FI SHYEAR | 1977 | 0. 68689702 | 0. 12102923 | 505 | 5. 68 | 0. 0001 | 0.05 | 0. 4491 | 0. 9247 |
| FI SHYEAR | 1978 | 0. 66004343 | 0. 09651112 | 505 | 6. 84 | 0. 0001 | 0. 05 | 0. 4704 | 0. 8497 |
| FI SHYEAR | 1979 | 0. 68852447 | 0. 09154141 | 505 | 7. 52 | 0. 0001 | 0.05 | 0. 5087 | 0. 8684 |
| FI SHYEAR | 1980 | 0. 96052284 | 0. 09647414 | 505 | 9. 96 | 0. 0001 | 0. 05 | 0. 7710 | 1. 1501 |
| FI SHYEAR | 1981 | 1. 14481557 | 0. 09165864 | 505 | 12. 49 | 0. 0001 | 0. 05 | 0. 9647 | 1. 3249 |
| FI SHYEAR | 1982 | 0. 88484081 | 0. 09470865 | 505 | 9. 34 | 0. 0001 | 0. 05 | 0. 6988 | 1. 0709 |
| FI SHYEAR | 1983 | 0. 69706925 | 0. 09210008 | 505 | 7. 57 | 0. 0001 | 0. 05 | 0. 5161 | 0. 8780 |
| FI SHYEAR | 1984 | 0. 63990322 | 0. 09529045 | 505 | 6. 72 | 0. 0001 | 0. 05 | 0. 4527 | 0. 8271 |
| FI SHYEAR | 1985 | 0. 77031838 | 0. 09343635 | 505 | 8. 24 | 0. 0001 | 0. 05 | 0. 5867 | 0. 9539 |
| FI SHYEAR | 1986 | 0. 71275743 | 0. 09073477 | 505 | 7. 86 | 0. 0001 | 0. 05 | 0. 5345 | 0. 8910 |
| FI SHYEAR | 1987 | 0. 89116523 | 0. 09296191 | 505 | 9. 59 | 0. 0001 | 0. 05 | 0. 7085 | 1. 0738 |
| FI SHYEAR | 1988 | 0. 80999563 | 0. 09866936 | 505 | 8. 21 | 0. 0001 | 0. 05 | 0. 6161 | 1. 0038 |
| FI SHYEAR | 1989 | 0. 87445526 | 0. 09659322 | 505 | 9. 05 | 0. 0001 | 0. 05 | 0. 6847 | 1. 0642 |
| FI SHYEAR | 1990 | 0. 69820943 | 0. 10887852 | 505 | 6. 41 | 0. 0001 | 0. 05 | 0. 4843 | 0. 9121 |
| FI SHYEAR | 1991 | 0. 74618343 | 0. 10023442 | 505 | 7. 44 | 0. 0001 | 0. 05 | 0. 5493 | 0. 9431 |
| FI SHYEAR | 1992 | 0.70766131 | 0. 09789412 | 505 | 7. 23 | 0. 0001 | 0.05 | 0. 5153 | 0. 9000 |
| FI SHYEAR | 1993 | 0. 77603010 | 0. 10656611 | 505 | 7. 28 | 0. 0001 | 0. 05 | 0. 5667 | 0. 9854 |
| FI SHYEAR | 1994 | 0. 85096591 | 0. 10680126 | 505 | 7. 97 | 0. 0001 | 0. 05 | 0. 6411 | 1. 0608 |
| FI SHYEAR | 1995 | 0. 72020722 | 0. 12614758 | 505 | 5. 71 | 0. 0001 | 0. 05 | 0. 4724 | 0. 9680 |
| FI SHYEAR | 1996 | 0. 82004806 | 0. 11574956 | 505 | 7. 08 | 0. 0001 | 0. 05 | 0. 5926 | 1. 0475 |
| FI SHYEAR | 1997 | 0. 66503548 | 0. 13330300 | 505 | 4. 99 | 0. 0001 | 0. 05 | 0. 4031 | 0. 9269 |

Table 18. Index of abundance of Atlantic greater anberjack fromheadboats. CPUE is the standardized catch rates on trips with greater anberjack, PPOS is the proportion positive, I NDEX is the standardized index and SEI and CVI are the standard error and coefficient of variation about the i ndex.

| FI SHYEAR | CPUE | PPOS | BC_CPU | GC | BC_POS | I NDEX | SE_I | CV_I |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  |  |  |
| 1976 | 2.48757 | 0.16923 | 3.70080 | 1.49846 | 0.16923 | 0.62627 | 0.15457 | 0.24681 |
| 1977 | 2.00290 | 0.15935 | 2.97850 | 1.49813 | 0.15935 | 0.47463 | 0.11961 | 0.25200 |
| 1978 | 1.95198 | 0.29308 | 2.92074 | 1.50335 | 0.29308 | 0.85600 | 0.14399 | 0.16821 |
| 1979 | 2.00432 | 0.30771 | 3.00192 | 1.50407 | 0.30771 | 0.92373 | 0.14282 | 0.15461 |
| 1980 | 2.65086 | 0.23295 | 3.9662 | 1.50338 | 0.23295 | 0.92401 | 0.15606 | 0.16889 |
| 1981 | 3.16758 | 0.25889 | 4.74408 | 1.50406 | 0.25889 | 1.22818 | 0.19686 | 0.16029 |
| 1982 | 2.44183 | 0.28281 | 3.65503 | 1.50362 | 0.28281 | 1.03367 | 0.16783 | 0.16237 |
| 1983 | 2.02340 | 0.30416 | 3.03018 | 1.50399 | 0.30416 | 0.92165 | 0.14616 | 0.15858 |
| 1984 | 1.91277 | 0.21893 | 2.86279 | 1.50354 | 0.21893 | 0.62674 | 0.10870 | 0.17344 |
| 1985 | 2.18760 | 0.22051 | 3.27532 | 1.50381 | 0.22051 | 0.72223 | 0.12241 | 0.16950 |
| 1986 | 2.05981 | 0.24123 | 3.08547 | 1.50418 | 0.24123 | 0.74430 | 0.12044 | 0.16182 |
| 1987 | 2.45873 | 0.25894 | 3.68156 | 1.50387 | 0.25894 | 0.95329 | 0.15551 | 0.16313 |
| 1988 | 2.26749 | 0.18749 | 3.39146 | 1.50304 | 0.18749 | 0.63586 | 0.11777 | 0.18521 |
| 1989 | 2.42726 | 0.19250 | 3.63193 | 1.50336 | 0.19250 | 0.69916 | 0.12428 | 0.17775 |
| 1990 | 2.03868 | 0.10987 | 3.04279 | 1.50146 | 0.10987 | 0.33432 | 0.07047 | 0.21078 |
| 1991 | 2.12165 | 0.15054 | 3.17232 | 1.50281 | 0.15054 | 0.47757 | 0.09138 | 0.19135 |
| 1992 | 2.05530 | 0.13653 | 3.07458 | 1.50316 | 0.13653 | 0.41976 | 0.07958 | 0.18958 |
| 1993 | 2.20624 | 0.12533 | 3.29452 | 1.50183 | 0.12533 | 0.41290 | 0.08523 | 0.20641 |
| 1994 | 2.37257 | 0.12268 | 3.54274 | 1.50180 | 0.12268 | 0.43463 | 0.08894 | 0.20463 |
| 1995 | 2.08874 | 0.07587 | 3.10472 | 1.49838 | 0.07587 | 0.23555 | 0.05887 | 0.24992 |
| 1996 | 2.31561 | 0.10546 | 3.45078 | 1.50030 | 0.10546 | 0.36393 | 0.08108 | 0.22279 |
| 1997 | 1.97239 | 0.09480 | 2.92625 | 1.49697 | 0.09480 | 0.27742 | 0.07390 | 0.26637 |

Table 19. Estimates of relative loss for the four cases with coefficients of variation and the number tagged (Tag) and recaptured (Rec). Year corresponds to May of the fishing year, May-April. EstAll=each year estimated independently, Sig01 = random walk sigma of 0.1, Sig03= random walk sigma of 0.3, Sig05= random walk sigma of 0.5.

|  |  |  | Relative Loss |  |  |  | Coefficient of Variation |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | Tag | Rec | EstAll | Sig01 | Sig03 | Sig05 | EstAll | Sig01 | Sig03 | Sig05 |
| 1960 | 9 | 4 | 2.040 | 1.875 | 2.576 | 2.467 | 53 | 15 | 19 | 22 |
| 1961 | 14 | 14 | 2.988 | 1.871 | 2.716 | 2.856 | 28 | 7 | 19 | 28 |
| 1962 | 37 | 33 | 3.376 | 1.825 | 2.708 | 3.014 | 19 | 7 | 17 | 24 |
| 1963 | 32 | 22 | 2.344 | 1.675 | 2.073 | 2.172 | 25 | 7 | 16 | 22 |
| 1964 | 36 | 28 | 1.203 | 1.506 | 1.457 | 1.348 | 21 | 7 | 16 | 21 |
| 1965 | 35 | 35 | 1.154 | 1.387 | 1.276 | 1.210 | 19 | 7 | 15 | 20 |
| 1966 | 30 | 34 | 1.197 | 1.300 | 1.221 | 1.203 | 19 | 7 | 15 | 20 |
| 1967 | 21 | 26 | 1.194 | 1.219 | 1.152 | 1.153 | 23 | 7 | 16 | 21 |
| 1968 | 28 | 19 | 0.838 | 1.135 | 1.023 | 0.962 | 25 | 7 | 16 | 22 |
| 1969 | 26 | 32 | 1.209 | 1.085 | 1.072 | 1.117 | 21 | 7 | 16 | 21 |
| 1970 | 31 | 27 | 1.038 | 1.014 | 0.961 | 0.986 | 23 | 7 | 16 | 21 |
| 1971 | 25 | 22 | 0.687 | 0.937 | 0.791 | 0.745 | 25 | 7 | 16 | 22 |
| 1972 | 34 | 23 | 0.639 | 0.887 | 0.728 | 0.683 | 24 | 7 | 16 | 22 |
| 1973 | 21 | 25 | 0.659 | 0.862 | 0.734 | 0.700 | 23 | 7 | 16 | 21 |
| 1974 | 20 | 27 | 0.832 | 0.854 | 0.807 | 0.820 | 23 | 7 | 16 | 21 |
| 1975 | 26 | 34 | 1.017 | 0.839 | 0.845 | 0.912 | 22 | 7 | 16 | 21 |
| 1976 | 32 | 24 | 0.688 | 0.796 | 0.714 | 0.706 | 26 | 7 | 16 | 22 |
| 1977 | 21 | 17 | 0.511 | 0.755 | 0.611 | 0.571 | 30 | 7 | 16 | 22 |
| 1978 | 22 | 20 | 0.526 | 0.731 | 0.577 | 0.549 | 29 | 7 | 17 | 23 |
| 1979 | 17 | 21 | 0.519 | 0.712 | 0.551 | 0.525 | 29 | 7 | 17 | 23 |
| 1980 | 5 | 9 | 0.390 | 0.696 | 0.507 | 0.461 | 39 | 7 | 17 | 24 |
| 1981 | 3 | 12 | 0.496 | 0.699 | 0.523 | 0.492 | 37 | 7 | 18 | 25 |
| 1982 | 0 | 5 | 0.410 | 0.702 | 0.521 | 0.481 | 51 | 7 | 18 | 27 |
| 1983 | 2 | 6 | 0.528 | 0.711 | 0.541 | 0.510 | 50 | 7 | 19 | 28 |
| 1984 | 2 | 2 | 0.390 | 0.720 | 0.550 | 0.509 | 77 | 7 | 19 | 29 |
| 1985 | 6 | 4 | 0.516 | 0.733 | 0.580 | 0.556 | 56 | 7 | 19 | 29 |
| 1986 | 9 | 10 | 0.751 | 0.748 | 0.617 | 0.622 | 39 | 7 | 19 | 29 |
| 1987 | 9 | 8 | 0.589 | 0.758 | 0.607 | 0.582 | 42 | 7 | 19 | 28 |
| 1988 | 11 | 3 | 0.382 | 0.769 | 0.590 | 0.526 | 61 | 7 | 18 | 27 |
| 1989 | 28 | 13 | 0.546 | 0.796 | 0.637 | 0.585 | 30 | 7 | 18 | 27 |
| 1990 | 43 | 26 | 0.655 | 0.842 | 0.738 | 0.699 | 21 | 7 | 17 | 23 |
| 1991 | 86 | 63 | 1.049 | 0.913 | 0.976 | 1.013 | 14 | 6 | 15 | 19 |
| 1992 | 62 | 85 | 1.158 | 0.948 | 1.062 | 1.111 | 13 | 6 | 13 | 16 |
| 1993 | 51 | 42 | 0.750 | 0.918 | 0.860 | 0.813 | 17 | 6 | 13 | 16 |
| 1994 | 37 | 50 | 0.861 | 0.926 | 0.900 | 0.880 | 16 | 7 | 14 | 17 |
| 1995 | 16 | 37 | 1.068 | 0.946 | 1.010 | 1.042 | 19 | 7 | 15 | 19 |
| 1996 | 2 | 18 | 1.191 | 0.953 | 1.074 | 1.149 | 28 | 7 | 18 | 23 |
| 1997 | 1 | 10 | 1.612 | 0.955 | 1.119 | 1.271 | 48 | 7 | 20 | 30 |

Table 20. Estimates of greater amberjack relative abundance from the $Z^{*}$ time series EstAll (see Table 1), two catch time series, and four assumptions of how $Z^{*}$ could be split into its components. Year corresponds to May of the fishing year, May-April.

|  | Case1 | Case2 | Case3 | Case4 | Case5 | Case6 | Case7 | Case8 |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Catch | A+B1 | A+B1 | A+B1 | A+B1 | A+B1+B2 | A+B1+B2 | A+B1+B2 | A+B1+B2 |
| M | 0.2 | 0.2 | 0.3 | 0.3 | 0.2 | 0.2 | 0.3 | 0.3 |
| Tag Loss | 0.05 | 0.1 | 0.05 | 0.1 | 0.05 | 0.1 | 0.05 | 0.1 |
| Year |  |  |  |  |  |  |  |  |
| 81 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 |
| 82 | 1.291 | 1.407 | 1.549 | 1.134 | 1.219 | 1.344 | 1.510 | 1.157 |
| 83 | 0.485 | 0.461 | 0.398 | 0.161 | 0.779 | 0.749 | 0.658 | 0.280 |
| 84 | 3.630 | 4.147 | 5.092 | 6.519 | 3.022 | 3.492 | 4.375 | 5.864 |
| 85 | 1.196 | 1.146 | 1.000 | 0.414 | 1.581 | 1.532 | 1.365 | 0.591 |
| 86 | 1.017 | 0.894 | 0.693 | 0.242 | 1.186 | 1.055 | 0.834 | 0.305 |
| 87 | 1.389 | 1.279 | 1.053 | 0.399 | 1.501 | 1.398 | 1.174 | 0.465 |
| 88 | 1.863 | 2.177 | 2.836 | 6.523 | 1.773 | 2.095 | 2.785 | 6.706 |
| 89 | 2.665 | 2.506 | 2.127 | 0.842 | 2.161 | 2.055 | 1.779 | 0.737 |
| 90 | 0.524 | 0.471 | 0.376 | 0.136 | 0.494 | 0.449 | 0.366 | 0.139 |
| 91 | 0.494 | 0.420 | 0.311 | 0.104 | 0.559 | 0.481 | 0.364 | 0.127 |
| 92 | 0.630 | 0.531 | 0.391 | 0.129 | 0.781 | 0.666 | 0.500 | 0.173 |
| 93 | 0.381 | 0.335 | 0.260 | 0.091 | 0.487 | 0.433 | 0.343 | 0.125 |
| 94 | 0.576 | 0.498 | 0.378 | 0.129 | 0.428 | 0.375 | 0.290 | 0.104 |
| 95 | 0.519 | 0.441 | 0.326 | 0.108 | 0.473 | 0.406 | 0.307 | 0.107 |
| 96 | 0.277 | 0.233 | 0.171 | 0.056 | 0.353 | 0.301 | 0.225 | 0.078 |
| 97 | 0.064 | 0.053 | 0.038 | 0.012 | 0.203 | 0.170 | 0.125 | 0.042 |

Table 21. Estimates of greater amberjack relative abundance from the $Z^{*}$ time series EstAll (see Table 1), two catch time series, and four assumptions of how $Z^{*}$ could be split into its components for years 1990-1997 only. Year corresponds to May of the fishing year, May-April.

|  | Case1 | Case2 | Case3 | Case4 | Case5 | Case6 | Case7 | Case8 |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Catch | A+B1 | A+B1 | A+B1 | A+B1 | A+B1+B2 | A+B1+B2 | A+B1+B2 | A+B1+B2 |
| M | 0.2 | 0.2 | 0.3 | 0.3 | 0.2 | 0.2 | 0.3 | 0.3 |
| Tag Loss | 0.05 | 0.1 | 0.05 | 0.1 | 0.05 | 0.1 | 0.05 | 0.1 |
| Year |  |  |  |  |  |  |  |  |
| $\quad 90$ | 1.209 | 1.263 | 1.335 | 1.423 | 1.045 | 1.095 | 1.161 | 1.242 |
| 91 | 1.141 | 1.126 | 1.106 | 1.083 | 1.184 | 1.172 | 1.155 | 1.135 |
| 92 | 1.454 | 1.425 | 1.389 | 1.348 | 1.653 | 1.625 | 1.588 | 1.546 |
| 93 | 0.881 | 0.900 | 0.924 | 0.950 | 1.032 | 1.057 | 1.088 | 1.123 |
| 94 | 1.330 | 1.336 | 1.343 | 1.346 | 0.907 | 0.914 | 0.921 | 0.926 |
| 95 | 1.199 | 1.182 | 1.160 | 1.133 | 1.001 | 0.989 | 0.974 | 0.955 |
| 96 | 0.640 | 0.626 | 0.608 | 0.589 | 0.748 | 0.733 | 0.715 | 0.695 |
| 97 | 0.147 | 0.142 | 0.135 | 0.129 | 0.429 | 0.415 | 0.398 | 0.380 |

Table 22. Indices considered for use in assessments.

\begin{tabular}{|c|c|c|c|c|c|c|c|c|c|c|c|}
\hline \& \begin{tabular}{l}
fishery \\
units \\
time of year age range
\end{tabular} \& hand
biom
mi
\(1-20\) \& \& mrfss

pri
numbe
mid

$1-20$ \& vate \& | mrfss char |
| :--- |
| number |
| mid |
| 1-20+ | \& \& headb

numb
mid

$1-20$ \& \& taggin $\mathrm{M}=0.3 \mathrm{Lo}$ numb mid 1-20 \& $$
\begin{aligned}
& \text { gg: } \\
& \text { ss= } \\
& \text { ser } \\
& \text { per }
\end{aligned}
$$ <br>

\hline fishing year \& \& index \& cV \& index \& cv \& index \& cv \& index \& cv \& index \& $\mathrm{cv}+0.1$ <br>
\hline 1977 \& \& \& \& \& \& \& \& 0.4746 \& 0.2520 \& \& <br>
\hline 1978 \& \& \& \& \& \& \& \& 0.8560 \& 0.1682 \& \& <br>
\hline 1979 \& \& \& \& \& \& \& \& 0.9237 \& 0.1546 \& \& <br>
\hline 1980 \& \& \& \& \& \& \& \& 0.9240 \& 0.1689 \& \& <br>
\hline 1981 \& \& \& \& 0.3158 \& 0.2324 \& 0.1786 \& 0.4076 \& 1.2282 \& 0.1603 \& \& <br>
\hline 1982 \& \& \& \& 0.1098 \& 0.2940 \& 0.1236 \& 0.2563 \& 1.0337 \& 0.1624 \& \& <br>
\hline 1983 \& \& \& \& 0.1436 \& 0.1652 \& 0.1851 \& 0.3452 \& 0.9217 \& 0.1586 \& \& <br>
\hline 1984 \& \& \& \& 0.1558 \& 0.1475 \& 0.1830 \& 0.2650 \& 0.6267 \& 0.1734 \& \& <br>
\hline 1985 \& \& \& \& 0.1627 \& 0.1680 \& 0.2107 \& 0.2309 \& 0.7222 \& 0.1695 \& \& <br>
\hline 1986 \& \& \& \& 0.1887 \& 0.1269 \& 0.2614 \& 0.2109 \& 0.7443 \& 0.1618 \& \& <br>
\hline 1987 \& \& \& \& 0.2059 \& 0.1178 \& 0.2104 \& 0.2205 \& 0.9533 \& 0.1631 \& \& <br>
\hline 1988 \& \& \& \& 0.1504 \& 0.1163 \& 0.1658 \& 0.2291 \& 0.6359 \& 0.1852 \& \& <br>
\hline 1989 \& \& \& \& 0.1762 \& 0.1183 \& 0.2367 \& 0.2110 \& 0.6992 \& 0.1778 \& \& <br>
\hline 1990 \& \& \& \& 0.1698 \& 0.1178 \& 0.1539 \& 0.2215 \& 0.3343 \& 0.2108 \& 1.2416 \& 0.31 <br>
\hline 1991 \& \& \& \& 0.1745 \& 0.0880 \& 0.2761 \& 0.1675 \& 0.4776 \& 0.1914 \& 1.1345 \& 0.24 <br>
\hline 1992 \& \& 29.5021 \& 0.5263 \& 0.1596 \& 0.1183 \& 0.4762 \& 0.2310 \& 0.4198 \& 0.1896 \& 1.5459 \& 0.23 <br>
\hline 1993 \& \& 29.5856 \& 0.3839 \& 0.1845 \& 0.1169 \& 0.5115 \& 0.2466 \& 0.4129 \& 0.2064 \& 1.1225 \& 0.27 <br>
\hline 1994 \& \& 43.2067 \& 0.3805 \& 0.1526 \& 0.1273 \& 0.3030 \& 0.2592 \& 0.4346 \& 0.2046 \& 0.9262 \& 0.26 <br>
\hline 1995 \& \& 43.2273 \& 0.3804 \& 0.1623 \& 0.1217 \& 0.3669 \& 0.2435 \& 0.2356 \& 0.2499 \& 0.9546 \& 0.29 <br>
\hline 1996 \& \& 33.1247 \& 0.3797 \& 0.1531 \& 0.1277 \& 0.2412 \& 0.2455 \& 0.3639 \& 0.2228 \& 0.6948 \& 0.38 <br>
\hline 1997 \& \& 28.2711 \& 0.3803 \& 0.1321 \& 0.1250 \& 0.4308 \& 0.2699 \& 0.2774 \& 0.2664 \& 0.3798 \& 0.58 <br>
\hline 1998 \& \& 21.2284 \& 0.3829 \& \& \& \& \& \& \& \& <br>
\hline
\end{tabular}

Figure 1. Observed CPUE observations (CPH) for the Atlantic Ocean greater amberjack for the private boat fishery (a) catch per angler hour, (b) log (catch per angler hour), for the charterboat fishery © catch per anglér hoiur ánd (d) log (catch per angler hour) based on the MRFSS data 1981-1998.
a)

b)


Figure 1. (Cont.)
c)



Figure 3. Estimated number of fishing trips by mode and fishing years in the Atlantic Ocean based on Summary of trips the MRFSS survey. The 1997 fishig year values are preliminary.


Figure 2. Standardized CPUE abundance trends (Catch per angler hour) and 95\% Confidence for the Atlantic Ocean greater amberjack by fishing year from the (a) Florida charter and private and b) all states combined charter and privatemodes, based on the MRFSS
a)

b)


Figure 4. Standardized catch per angler hour for the Atlantic Ocean greater amberiack by fishing year from the (a) charter and (b) private modes based on the MRFSS data. The estimated number of recreational fishing trips for the Atlantic Ocean by mode for all species is plotted on the second Y axis.
a.)

b).



Figure 6. Standardized yield per hook from all handline vessels with $80 \%$ confidence intervals.


Figure 7. Standardized yield per hook from handline vessels in the $90^{\text {th }}-99^{\text {th }}$ percentile for proportion positive or catch rate on successful trips with $80 \%$ confidence intervals.


Figure 8. Standardized yield per hook from handline vessels in the $75^{\text {th }}-99^{\text {th }}$ percentile for proportion positive or catch rate on successful trips. Confidence intervals


Figure 9. Comp arison of handline indices of abundance standardized to their means.

Figure 10. Greater amberjack catch per trip in the Atlantic by latitude.
Figure 11. Greater amberjack catch per trip in the Atlantic by month. Months 13 to 15 are January to March of the following calendar year.


Figure 12. Standardized catch of Atlantic greater amberjack per trip from the headboat fishery with confidence intervals of 2 standard errors.


Figure 13. Comparison of standardized indices from the headboat fishery with and without the inshore species factor.


Figure 14. Relative loss estimated from four procedures.


Figure 15. Relationship between sample size and uncertainty in relative loss estimation for the EstAll scenario.


Figure 16. Sensitivity analysis estimating one $Z^{*}$ for every two years.


Figure 17. Sensitivity analysis changing the minimum number of days at large.


Figure 18. Sensitivity analyses estimating reporting rates with a random walk


Figure 19. Sensitivity analyses estimating reporting rates with a random walk.


Figure 20. Two estimate types of the greater amberjack catch in the recreational fishery in the southeast Atlantic Ocean.


Figure 21. Relative abundance estimates (see Table 20).


Figure 22. Case I relative abundance estimates with $80 \%$ confidence intervals.


Figure 23. Relative abundance estimates (see Table 21).


Figure 24. Case I relative abundance with $80 \%$ confidence intervals.

Appendix 1. Regression tree examination of proportion of trips with greater amberjack with explanatory variables for fishing year, month, latitude and longitude (fishyrc, monthc, and latc). Months 11-16 indicate January-April in the calendar year after the associated months 5-12. For each node of the tree is given: node number (first node is 1 ), the levels of observations of the variable being split (latidude is split at node 2 and 3 with $24-28 \mathrm{~N}$ and $30-35 \mathrm{~N}$ falling in node 2 and latitude 29 N falling in node 3 ), the number of observations associated with that node ( 93121 observations for node 1), the deviance associated with the information in that node, and the average proportion positive associated with that node.

Tree formil a: PROPGA ~ FI SHYRC + MDNTHC + LATC,

```
1) root 93121 4326.00 0.06343
    2) LATC: 24, 25, 26, 27, 28, 30, 31, 32, 33, 34, 35 90471 3751.00 0. 05635
    4) MDNTHC: 6, 7, 8, 9, 10, 11, 12, 13, 14, 16 71428 2009.00 0. 04162
            8) LATC: 24, 25, 26, 32, 33, 34, 35 65649 1664.00 0. }0366
                16) MDNTHC: 6, 7, 8, 9, 10, 11, 12, 13, 16 56986 1256.00 0. }0330
                    32) FI SHYRC: 1991, 1992, 1993, 1994, 1998 28383 448. 80 0. }0246
                        64) MDNTHC: 7, 8, 9, 10, 11, 12, 16 20641 261.10 0.02130 *
                    65) MDNTHC: 6, 13 7742 187.00 0.03349 *
                    33) FI SHYRC: 1995, 1996, 1997 28603 803.40 0. }0414
                    66) LATC: 24, 25, 26, 34, 35 25114 704.70 0.03763
                        132) LATC: 34, 35 4101 29.34 0. 01755 *
                133) LATC: 24, 25, 26 21013 673.40 0.04155
                266) MONTHC: 6, 7, 8, 10, 11, 12, 16 16073 441.60 0. }0359
                        532) FI SHYRC: 1995, 1997 10358 247.90 0.03128 *
                        533) FI SHYRC: 1996 5715 193.10 0.04440 *
                267) MDNTHC: 9, 13 4940 229.70 0. 05979 *
                    67) LATC: 32, 33 3489 95.64 0.06922 *
            17) MDNTHC: }148663\mathrm{ 402.60 0.05981 *
            9) LATC: 27, 28, 30, 31 5779 324.40 0.09845 *
            5) MDNTHC: 5, 15 19043 1668.00 0. 11160
            10) LATC: 30, 31, 32, 33, 34, 35 3822 33.23 0. 02908 *
            11) LATC: 24, 25, 26, 27, 28 15221 1602.00 0. 13230
                22) LATC: 24, 27 12628 1236.00 0. 12120
                    44) MDNTHC: }156631\quad577.70 0. 1056
                    88) FI SHYRC: 1992, 1998 1620 100.30 0. 07075 *
                        89) FI SHYRC: 1991, 1993, 1994, 1995, 1996, 1997 5011 474. 80 0. }1169
                        178) FI SHYRC: 1994, 1995, 1996, 1997 3831 333.60 0. 10740 *
                        179) FI SHYRC: 1991, 1993 1180 139.80 0.14780 *
                45) MDNTHC: }55997\quad654.80 0.13840
                        90) FI SHYRC: 1991, 1993, 1994, 1995, 1996, 1997, 1998 5810 611.00 0. }1332
                        180) FI SHYRC: 1991, 1995, 1997, 1998 3202 311.00 0. 12310 *
                            181) FI SHYRC: 1993, 1994, 1996 2608 299. 20 0. 14570 *
                            91) FI SHYRC: 1992 187 38.70 0. 30000 *
            23) LATC: 25, 26, 28 2593 357.10 0. 18640 *
        3) LATC: 29 2650 415.90 0. 30520 *
```

Appendi $x$ 2. Regressi on tree examination of handline yi eld per hook on trips with greater anberj ack with expl anat ory variables for fishing year, month, latitude and Iongitude (fishyrc, monthc, I atc, and I ongc). Mbnths 11-16 indi cate January-April in the cal endar year after the associ at ed months 5-12. For each node of the tree is gi ven: node number (first node is 1), the levels of observations of the variable being split, the number of observations associ at ed with that node, the devi ance associ ated with the information in that node, and the average proportion positive associated with that node.

Tree for mol a: CPHOOK ~ FI SHYRC + MDNTHC + LATC + LONGC,

1) root 132411528000000182.40
2) LATC: 30, 31, 32, 33, 34, 35556664440000 38. 58 *
3) LATC: 24, 25, 26, 27, 28, 2976751265000000286.70
4) MONTHC: 6, $7, \quad 8, \quad 9,10,11,12,13,14,164976658300000228.90$
5) LATC: 24, 25, 26, 27, 283781427200000 203. 30
6) MDNTHC: 6, 7, 8, 9, 10, 11, 12, 13, 162993267000000 180. 90
7) LATC: $2530610610000 \quad 86.80$ *
8) LATC: 24, 26, 27, 282687253400000 191. 60
9) MDNTHC: 7, 11, 12, 16113268560000 149.50 *
10) MDNTHC: $6, \quad 8, \quad 9, \quad 10,131555181400000222.20$
11) FI SHYRC: 1992, 1993, 1996, 199871062710000 180. 30 *
12) FI SHYRC: 1991, 1994, 1995, 1997845116400000 257. 40 *
13) MDNTHC: $14788 \quad 153000000288.20$
14) FI SHYRC: 1995, 1996, 1997, 199846045430000233.70 *
15) FI SHYRC: 1991, 1992, 1993, 1994328104200000364.80 *
16) LATC: 291195220800000 309. 90
17) MONTHC: 6, 7, 8, 9, 11, 12675105800000264.00 *
18) MONTHC: 10, 13, 14, $16520 \quad 111700000$ 369. 40 *
19) MDNTHC: $5,152699559000000393.40$
20) FI SHYRC: 1993, 1994, 1995, 1996, 1997, 19982378426500000365.10
21) FI SHYRC: 1996, 1997, 19981173185900000 319. 50
22) FI SHYRC: 199824919100000 261. 10 *
23) FI SHYRC: 1996, 1997924165800000 335. 20
24) LONGC: 79, $81212 \quad 28020000$ 299. 60 *
25) LONGC: 80712137400000345.80 *
26) FI SHYRC: 1993, 1994, 19951205235700000 409. 50
27) LATC: 25, 26, 28, 2939463380000 303. 20 *
28) LATC: 24, 27811165700000461.10
29) LONGC: 8115625550000 326. 20 * 119) LONGC: 80655136700000 493. 30 *
30) FI SHYRC: 1991, 1992321116500000603.40 *

Appendix 3. Regression tree examination of headboat catch rates of greater amberjack. Explanatory variables included calendar year, month, triptype, latitude and longitude (yearc, monthc, triptypc, nlatmc, and nlongmc) as well as numerous variables related to the proportion of various disaggregated taxonomic groups. For each node of the tree is given: node number (first node is 1), the levels of observations of the variable being split, the number of observations associated with that node, the deviance associated with the information in that node, and the average catch rate for that node.

```
Tree formol a: CPUE ~ YEARC + MDNTHC + TRI PTYPC + NLATMC + NLONGMC + CARCHPC + HOLOCPC +
EPI NEPPC + MYCTERPC + OTHSERPC + BI GEYEPC +
    BLUELI PC + SANDTI PC + WAHCOBPC + OTSERI PC + I NPORGPC + LABRI DPC + OTHMACPC + BALI STPC
+ CENTROPC + CARANXPC + HMSPC + LUJ ANUPC +
    OCYURUPC + RHOMBOPC + GRUNTSPC + PAGARUPC + KATSUWPC + KI NGMAPC,
Variables actually used in tree construction:
[ 1] "NLATMC" "TRI PTYPC" "MDNTHC" "YEARC" "CENTROPC" "GRUNTSPC"
    1) root 153996 7874.00 0.017180
    2) NLATMC: -24, -25, -26, -27, - 28, - 29, - 32, -33, - 34, - 35 139103 1627.00 0.009383
        4) TRI PTYPC: 1/2 day am 1/2 day pm 1/2 nite pm 3/4 day 120719 1075.00 0.007068
            8) MDNTHC: 1, 2, 3, 6, 7, 8, 9, 10,12 91250 337.90 0.004600 *
            9) MDNTHC: 4, 5,11 29469 734.50 0.014710 *
        5) TRI PTYPC: full day 18384 547.30 0.024580 *
    3) NLATMC: - 30, - }3114893\mathrm{ 6159.00 0. }09003
        6) YEARC: 1977, 1978, 1979, 1980, 1983, 1984, 1985, 1986, 1987, 1988, 1989, 1990,
1991, 1992, 1993, 1994, 1995, 1996, 1997 13099 825.40 0. 051770
            12)
YEARC: 1978, 1979, 1983, 1984, 1985, 1986, 1987, 1988, 1989, 1990, 1991, 1992, 1993, 1994, 1995, 1996, 199
711540 223.40 0.034850 *
    13) YEARC: 1977, 1980 1559 574. 20 0. 177000 *
        7) YEARC: 1981, 1982 1794 5175.00 0. 369400
            14) MDNTHC: 3, 4, 6, 7, 8, 9, 10, 11, 12 1417 1238.00 0. 208900
                28) MDNTHC: 3, 7, 8, 9, 10,11 869 138.70 0.078460 *
                29) MDNTHC: 4, 6,12 548 1061.00 0.415800
                58) TRI PTYPC: day am, day pm 365 188.10 0. 184100 *
                59) TRI PTYPC: full day 183 814.50 0.878000
                        118) CENTROPC: 2, 3 109 28.62 0. 210500 *
                        119) CENTROPC: 0, 1 74 665.80 1.861000 *
            15) MDNTHC: 1, 2, 5 377 3763.00 0.972600
                30) CENTROPC: 0, 2, 3 254 665.90 0.428700 *
                31) CENTROPC: 1 123 2867.00 2. }09600
                    62) GRUNTSPC: 1, 2, 3 63 162. 10 0. 468700 *
                    63) GRUNTSPC: 0 60 2363.00 3. 804000 *
```

Appendix 4. Regression tree examination of headboat catch rates of greater amberjack. Explanatory variables included calendar year, month, triptype, latitude and longitude (yearc, monthc, triptypc, nlatmc, and nlongmc) as well as aggregated taxonomic groups indicators (inshorpl, offbotpl and offpelpl). For each node of the tree is given: node number (first node is 1 ), the levels of observations of the variable being split, the number of observatio ns associated with that node, the deviance associated with the information in that node, and the average catch rate for that node.

Tree formil a: CPUE ~ YEARC + MONTHC + TRI PTYPC + NLATMC + NLONGMC + I NSHORPL + OFFBOTPL + OFFPELPL,
dat $\mathrm{a}=$ Anal y6rev, na. action = na.fail, mincut $=25$, minsize $=50$, mindev $=0.1$

Variables act ually used in tree construction:
[ 1] " NLATMC" "TRI PTYPC" "MONTHC" "I NSHORPL" "OFFBOTPL" "YEARC"

1) root 1539967874.000 .017180
2) NLATMC: -24, -25, -26, -27, -28, - 29, -32, -33, -34, - 351391031627.000 .009383
3) TRI PTYPC: $1 / 2$ day am $1 / 2$ day pm $1 / 2$ nite $p m ~ 3 / 4$ day 120719 1075.00 0.007068
4) MONTHC: 1, 2, 3, 6, 7, 8, 9, 10, 1291250337.900 .004600 *
5) MONTHC: 4, 5, 1129469 734. 50 0. 014710 *
6) TRI PTYPC: full day 18384 547. 30 0. 024580 *
7) NLATMC: -30, - 31148936159.000 .090030
8) I NSHORPL: 2, $312816 \quad 639.20$ 0. 045620 *
9) I NSHORPL: 0, 12077 5339. 00 0. 364100
10) OFFBOTPL: $0,2,317931203.000 .177800$
11) OFFBOTPL: 31303 31.98 0. 042800 *
12) OFFBOTPL: 0, 2490 1084. 00 0. 536800
13) 

YEARC: 1977, 1978, 1979, 1983, 1984, 1985, 1986, 1987, 1988, 1989, 1990, 1991, 1992, 1993, 1994, 1995, 199
6, 1997294 59. 26 0. 113500 *
59) YEARC: 1980, 1981, 1982196 893. 00 1. 172000
118) MONTHC: 3, 4, 6, 7, 8, 9, 10, 11, 12140 214. 60 0. 755200 *
119) MDNTHC: 1, 2, 556 593. 40 2. 212000 *
15) OFFBOTPL: 12843681.00 1. 540000
30)

YEARC: 1978, 1979, 1981, 1983, 1984, 1985, 1986, 1987, 1988, 1989, 1990, 1992, 1993, 1994, 1995, 1996, 199
7204 262. 30 0. 345600 *
31) YEARC: 1980, 198280 2386. 00 4. 586000
62) TRI PTYPC: day am day pm 36 264. 50 1. 862000 *
63) TRI PTYPC: full day 44 1635. 00 6. 814000 *

Appendix 5. Regression tree examination of headboat catch rates of greater amberjack with data restricted to full day trips. Explanatory variables included calendar year, month, triptype, and latitude (yearc, monthc, and nlatmc) as well as aggregated taxonomic groups indicators (inshorpl, offbotpl and offpelpl). Year refers to fishing year with April as the first month of the fishing year. Months 13 to 15 refer to January to March. For each node of the tree is given: node number (first node is 1 ), the levels of observations of the variable being split, the number of observations associated with that node, the deviance associated with the information in that node, and the average catch rate for that node.

```
tree(formul a = CPUE ~ YEARC + MDNTHC + TRI PTYPC + NLATMC + I NSHORPL + OFFBOTPL +
OFFPELPL,
            data = Anal 6cfd, na. acti on = na. omit, mi ncut = 25, minsi ze = 50, mindev = 0.05)
Variables actually used in tree construction:
[ 1] "YEARC" "NLATMC" "OFFBOTPL" "I NSHORPL" "MDNTHC"
    1) root 25450 5623.00 0.05966
        2)
YEARC: 1975, 1976, 1977, 1978, 1979, 1983, 1984, 1985, 1986, 1987, 1988, 1989, 1990, 1991, 1992, 1993, 199
4, 1995, 1996, 1997 21924 476.70 0. }0338
            4) NLATMC: -24, -25, -26, -27, - 28, - 29, - 32, - 33, - 34, - 35 16214 235. 90 0.02185*
            5) NLATMC: - 30, -31 5710 231.90 0.06785 *
            3) YEARC: 1980, 1981, 1982 3526 5041.00 0. 22030
            6) NLATMC: - 24, -25, -26, -27, -28, -29, - 31, - 32, - 33, -34 2275 312. 30 0. 04411
                12) OFFBOTPL: 0, 2, 3 1975 55.03 0.02715 *
                13) OFFBOTPL: 1 300 253.00 0. 15580 *
            7) NLATMC: - }3012514530.00 0.5407
                14) I NSHORPL: 2, 3 843 339.20 0. }1748
                    28) MONTHC: 7, 8, 9, 10, 11, 12, 14, 15,16 657 83.97 0. 11380 *
                    29) MDNTHC: 5, 6, 13 186 244. 10 0.39050 *
                15) I NSHORPL: 0, 1 408 3845. }00\mathrm{ 1. }2960
                    30) OFFBOTPL: 0, 2, 3 344 642.10 0.56430
                    60) OFFBOTPL: 3 235 19.22 0. 13460 *
                    61) OFFBOTPL: 0, 2 109 485.90 1.49100
                        122) MDNTHC: 6, 7, 8, 9, 10, 11, 12, 15, 16 74 109.20 0. 95010 *
                123) MDNTHC: 5, 13,14 35 309.30 2.63400 *
            31) OFFBOTPL: }164\mathrm{ 2027.00 5. 23200
                    62) YEARC: 1980, 1981 39 91. 87 2. 45900 *
                    63) YEARC: 1982 25 1167.00 9.55700 *
```

