# A Stock Assessment for Gray Triggerfish, Balistes capriscus, in the Gulf of Mexico 

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## SUMMARY

Standardized indices of abundance were estimated for gray triggerfish (Balistes capriscus) in the Gulf of Mexico from five recreational and commercial fisheries data sets: the Marine Recreational Fishery Statistics Survey (MRFSS), the Southeast Fisheries Science Center, National Marine Fisheries Service (SEFSC-NMFS) Headboat Survey, the Alabama Charterboat Survey, the Panama City Charterboat Survey, and the commercial Florida Logbook System Program. A sixth data set from the Texas Park and Wildlife Department (TPWD) Recreational Creel Survey was examined but the indices developed were not considered for subsequent analyses. The standardized indices were estimated using Generalized Linear Mixed Models under a delta lognormal model approach.

Catch-effort statistics from the recreational and commercial sectors for years 1986 to 1998 were used for stock assessment. The standardized catch rates developed here were used to tune a non-equilibrium production model (ASPIC).

## INTRODUCTION

The gray triggerfish, Balistes capriscus (Gmelin, 1788), is an important component of the Gulf of Mexico reef fishery, particularly for the recreational fishing sector (Goodyear and Thompson, 1993). The species is widely distributed in tropical and temperate waters throughout the Atlantic; in the Western Atlantic it ranges from Nova Scotia through Bermuda and the Gulf of Mexico to Argentina (Harper and McClellan 1997).

Until recently, gray triggerfish were not considered a desirable catch by most fishers, however, the decline in other reef fish stocks (e.g., red snapper and groupers) has probably caused an increased targeting of this and other "under-utilized" species. This has resulted in an initial increase in average annual landings from 1.46 million pounds (1986) to 2.88 million pounds (1990) followed by a steady decline since then ( 0.85 million pounds in 1998). The cause of this decline has not been determined, but it could be attributed to a consistent increase in fishing effort and a possible consequent decrease in stock size. A thorough stock assessment is required to test this hypothesis.

The gray triggerfish (Balistes capriscus) in the Gulf of Mexico EEZ is managed under the 1981 Reef Fish Fishery Management Plan (FMP) and subsequent amendments. It was first added to the list of species included in the FMP in Amendment 1 (8/1989). Amendment 12 to the Reef Fish FMP (12/1995) established an EEZ aggregate daily bag (possession) limit for all reef fish species not having a bag limit. The aggregate bag limit was established to improve enforceability of commercial reef fish harvest regulations by preventing non-permitted fishermen from harvesting commercial quantities of those species under pretense of recreational fishing, which might subsequently be sold. It also served as a pro-active conservation measure to prevent uncontrolled increase in harvest of species for which no regulations or stock assessment existed. This aggregate bag limit applied to reef fish, including gray triggerfish. Species not in the reef fish fishery which did not have a bag limit could continue to be caught in unlimited quantities. Amendment

15 (6/1997) to the Reef Fish FMP continued to include gray triggerfish in the 20 fish aggregate bag limit.

Amendment 16b (1/1999) established compatible bag limits and size limits for several species of reef fish regulated under Florida statutes, for which there previously were either no corresponding limits in federal waters, or for which federal limits differed from the state limits. For consistency of regulations and improvement of enforceability, Florida requested that compatible limits be adopted in federal waters. As part of these changes, a minimum size limit of 12 inches (TL) was adopted for gray triggerfish.

As a result of these amendments, current regulations for gray triggerfish in the Gulf of Mexico are:

1) Recreational regulations: Minimum Size $=12 \mathrm{in}$. TL, no closed season, with a 5 fish/person bag limit, included in the 20 reef fish aggregate limit.
2) Commercial regulations: Minimum Size $=12$ in. TL, no closed season, no trip limit.

The increase in economic value and the steady decline in total landings since 1990 have raised concern regarding the status of gray triggerfish stocks and the effectiveness of the existing management regulations in the Gulf of Mexico. Due to this concern, the SEFSC initiated a thorough examination of the existing information for the species in 1993, with an evaluation of data on size and catch limits conducted by Goodyear and Thompson (1993). They showed catch trends by sector, state, mode, area, and depth strata for the period 1986 to 1991, and estimated various morphometric relationships for the Gulf of Mexico stock. They concluded that there could be significant reductions in landings by size if length and trip limits were implemented.

As a continuation of these efforts to evaluate the Gulf of Mexico fishery for gray triggerfish, Harper and McClellan (1997) conducted a thorough review of the biology and the fishery, and updated the estimates from the previous study. They suggested that several factors could be involved in the initial increase and recent decline in gray triggerfish landings, such as an increased targeting of this species by both recreational and commercial fishers, the reduction in other reef fish stocks, and more restrictive regulations on other reef fish stocks.

Based on this background information, and given that more complete information on landings statistics, CPUE and size-weight relationships is available since 1986, it is now possible to conduct a formal stock assessment. The objective of this study is to evaluate the current status of the fishery and the gray triggerfish population in the Gulf of Mexico.

## DATA SOURCES AND METHODS

Commercial landings statistics for gray triggerfish in the Gulf of Mexico exist since 1962. For the recreational sector, landings statistics date back to 1981. For the purpose of this study, only data for the period 1986 to 1998 for both sectors was considered complete and useful for stock assessment. Survey data on catch rates and biological information exists for the recreational sector since 1979 from various sources, including NMFS/SEFSC surveys and state-based fishery statistical programs. Additional
information on landings by size from both sectors is available from the SEFSC Trip Interview Program (TIP) for the period 1989-1999, but this last year is still incomplete. Only size information for the period 1986-1998 was included in the present analysis.

Five recreational fisheries survey data were included in the analyses: 1) the NMFS, Marine Recreational Fishery Statistics Survey (MRFSS) (1981-1999) for landings estimates from charterboat, shore, and private/rental modes, for CPUE information, and for samples of landings at size; 2) the NMFS/SEFSC Beaufort Laboratory Headboat Survey (1986-1998) for landings estimates, CPUE and size samples; 3) the Texas Parks and Wildlife Department Recreational Creel Survey (19831998) for landings estimates from all modes for Texas, for CPUE, and for size samples; 4) the Alabama Department of Conservation, Marine Resources Division, Charterboat Logbook Survey (1991-1995) for CPUE and landings by size; and 5) the SEFSC Panama City Charterboat Survey (1989-1996) for CPUE estimates.

For analyses of the commercial sector, data were obtained from the SEFSC General Canvass Program for landings in weight. Trip specific landings information from the SEFSC Logbook Program was used for commercial CPUE standardization.

## CPUE ANALYSES

## Recreational Sector

All the recreational survey data-bases were used to estimate relative indices of abundance for gray triggerfish in the Gulf of Mexico. Recreational logbooks generally record the numbers of fish caught (kept and released) by fishing trip, the number of anglers on the trip, the hours spent fishing, the fishing mode(s), the gear(s) used, the area fished, the target species, and occasionally other, more specific information, such as the number of hooks, the number of trips in the day, and finer categories of the hours spent fishing and the catch. Each data set was analyzed separately, but the estimation of nominal fishing effort, total catch, and nominal catch rates was performed in a similar manner. The fishing effort unit considered was angler hour, estimated as the total hours spent fishing times the total number of anglers in the trip. Catch was summed over all types (kept and released, dead or alive, caught while trolling or not). Nominal catch rates (CPUEs) were estimated as the total catch per angler hour, and were used for abundance index standardization. The peculiarities of each data set will not be described, but the main features, useful for the analyses, will be outlined.

MRFSS. For this data set, CPUE analysis used data from 1981 through 1999. All trips with successful and unsuccessful gray triggerfish catch were considered, whether this species was the primary target or not. The index is the standardized number of fish caught (landed + discarded) per angler hour adjusted to non-interviewed anglers assuming similar catch to those anglers interviewed for a given trip. Only hook and line catches were included, as they accounted for over $99 \%$ of the data. The other fishing gear reported in this data set (dip net, gill net, seine, trawl, spear, other) were therefore excluded. Texas data was also removed from the analysis, since this state is covered by the Texas Parks and Wildlife Department Survey, and the catches reported for Texas in MRFSS are negligible. The explanatory variables considered for the MRFSS Gulf of

Mexico analysis included: year, state, fishing mode (shore, headboats, charter, private/rental boats), area (distance from shore: ocean<3 miles, ocean>3 miles, ocean <10 miles, ocean $>10$ miles), season (Jan-Apr, May-August, September-December), and fishing target, where target 1 specifically included gray triggerfish as a target species.

Headboats. Data for years 1986 through 1998 were available for CPUE analysis. All trips were considered and the index is the standardized number of fish caught (landed + discarded) per angler hour. The explanatory variables analyzed included year, state, and season. Fishing areas (defined here $\boldsymbol{\infty}$ a subdivisions within each state) were not included in the analysis since their effect was nested within the state variable.

Texas Parks and Wildlife. For the Texas Parks and Wildlife Department Recreational Angler Creel Survey (TPWD) data set, CPUE analysis used data from 1983 through 1990. The index is the standardized number of fish caught (landed + discarded) per angler hour. The explanatory variables included were year, area (distance from shore: ocean $<10$ miles, ocean>10 miles, and inshore bays and passes), and season of the year.

Alabama Logbook Survey. Data for years 1991 to 1995 was used for CPUE analysis. Fishing effort was estimated as the total number of hours spent trolling and not trolling in estuarine and ocean waters (targeting or not targeting gray triggerfish), times the number of anglers. Catch was calculated as the sum of gray triggerfish kept and released while trolling and not trolling. Therefore the index is defined as the standardized number of fish (landed + discarded) per angler hour. This data set contains catch information for other species, so positive catches of these and/or zero gray triggerfish catch were defined as unsuccessful gray triggerfish trips, and were included in the analysis. Unreported fishing modes were excluded from the analyses, as they accounted for a very small proportion of the catches. The explanatory variables included were only year and season.

Panama City Charterboat Survey. This data base spans years 1989 to 1996 and includes fishing area (state subdivisions) and target species information. Only years 1989 through 1995 were included in the analysis because 1996 was incomplete. The explanatory variables considered were only year, season, and fishing area. State was not included since its effect is confounded with the area effect. The index is defined as the standardized number of fish caught (landed + discarded) per angler hour.

## Commercial Sector

The Florida Logbook System (FLS) data-base (1990-1999) was used to estimate relative indices of abundance for the commercial sector of the gray triggerfish fishery in the Gulf of Mexico. The commercial logbook program for the Gulf of Mexico records trip-specific information for various fisheries (reef fish, swordfish, tuna/bluefin, sharks, king mackerel, dolphin, etc.). Trip-specific data include landings in weight by species, information about the vessel, the crew, the location fished (state, county, area), the type of gear used (traps, longlines, gill nets, handlines, trolls, divers), and the amount of fishing effort exerted (days/hours fishing, number of lines/hooks/traps/divers/nets, size of
lines/nets). In order to perform the CPUE standardization for gray triggerfish, this dataset was filtered according to the following criteria (see Figures 3, 4, 5, 6, and 7):

1. Use only Gulf of Mexico data-base.
2. Extract only Gulf of Mexico reef fish (based on Gulf of Mexico Fisheries Management Plan).
3. Select vessels that caught at least one pound of gray triggerfish in their catch history.
4. Extract handlines only (they make up for $86 \%$ of gray triggerfish catch).
5. Select vessels that caught gray triggerfish for five years or more (5-10) during the period 1990-1999.

Four fishing effort units were considered for analysis: Effortl= angler hours (estimated as the number of crew imes hours fished), Effort2= number of hooks (number of lines times number of hooks per line), Effort3=hooks hour (number of hooks times hours fished), and Effort $4=$ hours fished. Nominal catch rates were estimated as the total catch in pounds divided by each effort unit.

CPUE analyses only used data from 1993 to 1999, as there were insufficient data in years 1990-1992. This explains the large standard errors observed for those three years (Figure 5). Nominal CPUE trajectories showed highly fluctuating and dissimilar trends for vessels that caught gray triggerfish for less than 5 years (Figures 5, 6, 7). Conversely, nominal catch rates were comparable among vessels that harvested this species for 5 years or more, and standard errors were smaller as the as the number of years with gray triggerfish catch decreased (Figure 5). Hence, these vessels (>=5 years) were selected for analyses, even when they constituted only $30 \%$ of the total number harvesting the species.

All trips with successful and unsuccessful gray triggerfish catch were considered; catch of all other species was aggregated into the "not successful" catch. Only handline catches were included as the other fishing gear reported in the FLS data-base accounted for a small proportion of the catch (less than $14 \%$ ) (Figure 3). Crew size information was insufficient to use angler hour as an effort unit for CPUE standardization, therefore it was dropped from subsequent analyses.

The commercial indices developed are: CPUE2= standardized catch in weight (pounds) per hook; CPUE3= pounds per hook hour; CPUE4= pounds per hour. The explanatory variables considered were year, state, county, area (Gulf of Mexico grids), and season (Jan-Apr, May-August, September-December).

## THE DELTA LOGNORMAL MODEL FOR CPUE STANDARDIZATION

Relative indices of abundance for gray triggerfish were estimated by a Generalized Linear Mixed Model Approach (GLMM) assuming a delta lognormal model distribution. The delta model estimates separately the proportion of positive trips/stratum (in the GLM matrix), assuming a binomial error distribution, and the mean catch rate of trips where at least one fish was caught assuming a lognormal error distribution. The log-
transformed frequency distributions of catch rates in numbers for gray triggerfish are shown in Figure 1. The estimated proportion of successful trips per stratum is assumed to be the result of $r$ positive trips of a total $n$ number of trips, and each one is an independent Bernoulli-type realization. The estimated proportion is a linear function of fixed effects and interactions. The logit function was used as a link between the linear factor component and the binomial error. For trip/days that caught at least one fish (positive observations), estimated catch rates were assumed to follow a lognormal error distribution of a linear function of fixed factors and random effect interactions, particularly when the year effect was within the interaction. In some cases other interactions were tested.

A step-wise regression procedure was used to determine the set of systematic factors and interactions that significantly explained the observed variability. The difference of deviance between two consecutive models follows a $\mathbf{P}^{2}$ (Chi-square) distribution; this statistic was used to test for the significance of an additional factor in the model. The number of additional parameters associated with the added factor minus one corresponds to the number of degrees of freedom in the $\mathbf{P}^{2}$ test (McCullagh and Nelder, 1989). Deviance analysis tables for catch rates in numbers are presented for each index developed. Each table contains the deviance for the proportion of positive observations (i.e. positive trips/total trips), and the deviance for the positive catch rates. Final selection of explanatory factors was conditional to: a) the relative percent of deviance explained by adding the factor in evaluation; normally factors that explained more than $5 \%$ were selected. The year term was always included regardless of statistical significance because a time series is desired. b) The $\mathbf{P}^{2}$ test significance, and c) the type III test significance within the final specified model.

Once a set of fixed factors was specified, possible interactions were evaluated, in particular interactions between the year effect and other factors. Selection of the final mixed model was based on the Akaike's Information Criterion (AIC), the Schwarz's Bayesian Criterion (SBC), and a likelihood-ratio test for successive model formulations, based on a chi-square test (Littell et al. 1996). Relative indices for the delta model formulation were calculated as the product of the year effect least square means (LSMeans) from the binomial and the lognormal model components. The LSMeans estimates use a weighted factor of the proportional observed margins in the input data to account for the unbalanced characteristics of the data. LSMeans of lognormal positive trips were bias corrected using Lo et al. (1992) algorithms. Analyses were done using a computer program developed by Ortiz et al. that incorporates the GLIMMIX and MIXED procedures from the SAS statistical computer software (SAS Institute Inc. 1997). This methodology has been applied and refined by Legault and Ortiz (1998), by Ortiz et al. (2000), and Ortiz and Farber (2000), to standardize catch rates of Spanish mackerel, king mackerel, and marlins, respectively.

## HARVEST

Recreational landings in numbers of fish by state and fishing mode were estimated for the period 1986-1998 (Table 9, Figure 12). Recreational landings peaked in 1990, followed by a steady decline ever since. The majority of annual landings since

1986 have been reported from the West Coast of Florida, followed by Louisiana. The other states account for a very small proportion of the catch.

Landings from the charterboat mode have dominated recreational landings for most of the period studied. Private and rental boats have also accounted for a significant proportion of the landings, whereas the headboat mode has generally contributed with a small proportion. The shore mode has scarcely been represented during this period.

Landings in weight by year and state were estimated for the commercial sector for the period 1986-1998 (Table 10, Figure 13). The trend throughout this period is similar to that of the recreational sector, but with a peak in 1993 and a steady decline since then. The greatest proportion of commercial landings has been reported for the west coast of Florida, followed only by Louisiana. The other states have generally reported very small proportions of the total commercial catch.

Total landings in weight were estimated for both the recreational and the commercial sectors (Table 11, Figure 14). To evaluate landings in weight from the recreational sector, landings in numbers of fish were converted to total weight. To accomplish this conversion, size and weight samples from each recreational survey (MRFSS, Headboat, and TPWD) were analyzed separately using the information presented in Tables 7 and Figures 10 and 11. Within each data set, when fish weight was not provided, it was estimated from fork length or total length using the morphometric relationships given in Goodyear and Thompson (1993):

Fork length to whole weight: $\mathrm{Wt}=(8.975 \mathrm{E}-4) \mathrm{FL}^{2.96}$
Total length to whole weight: $\mathrm{Wt}=(9.953 \mathrm{E}-4) \mathrm{TL}^{2.773}$
These weight samples were used to estimate mean fish weight by year, state and fishing mode strata. In cases where the sample size by stratum exceeded 25 individuals, the mean weight estimate corresponded to the average by stratum; if the sample size by stratum was less than 25 , the state annual mean was substituted (if $\mathrm{n}>25$ individuals), else the gulfwide annual mean was used (Table 8). The three data sets were combined and mean weights were multiplied by numbers of fish stratified in the same manner (year, state, mode) to derive total recreational landings in weight (Table 9). Comparison of these estimates with those from previous studies (Goodyear and Thompson 1993, Harper and McClellan 1997) was made. The stratified landings in weight could not be matched up. The source of he discrepancy was associated with the estimate of mean fish weight per stratum, rather than with the estimate of landings in numbers per stratum. The differences among the estimates from the three studies did not indicate bias in any direction, and, unfortunately, the detailed procedure and assumptions used in those other studies could not be established. Thus, the exact reason of the mismatch could not be determined. However, after a careful review of our method, we believe that our estimates are reasonable as they fall within the range of the previous studies.

Estimated total landings over the period 1986 through 1998 have been dominated by the recreational sector. Total landings increased each year for the first few years, reaching a peak in 1990, and then declined steadily through 1998, to an estimated 854,000 pounds. Both sectors have showed a proportional decline in landings throughout the period.

## STOCK ASSESSMENT MODEL AND APPLICATION TO DATA

No previous assessments have been made of the gray triggerfish fishery of the Gulf of Mexico. Given the characteristics of the data available (annual yields in weight and standardized catch rates) and that a simple, straight-forward, and flexible method may be desired as a first approach, a non-equilibrium surplus-production model was selected to conduct this stock assessment. The ASPIC computer program of Prager (1994, 2000) was used for model fitting. This method incorporates various extensions to classical stock-production models, such as the possibility of including several simultaneous or sequential fisheries on the same stock, "tuning" the model to a biomass index, estimating missing values of fishing effort, and constructing confidence intervals of parameter values via bootstrapping (Prager 1994).

Data needed for parameter estimation under ASPIC are a series of observations on catch (yield in biomass) and corresponding effort or CPUE. The program can fit data from up to 10 data series. In addition to data, ASPIC requires starting guesses and the ranges for its estimated parameters: $r$, the intrinsic rate of increase; $M S Y$, maximum sustainable yield; the ratio $B_{l} / B_{M S Y}$, the ratio of the biomass at the beginning of the first year to the biomass at which MSY can be attained; and $q$, the catchability coefficient. A separate estimate of $q$ is made for each data series (Prager 2000). Initial parameter estimates and their ranges were based on biological knowledge of the species and of the fishery in question.

Initial runs of the model used all the recreational standardized catch rates obtained in this study and the commercial CPUE 4 ( $\mathrm{lb} / \mathrm{hr}$ ) as tuning indices (Figure 15 (B)), parameters were not constrained, and the program was allowed to estimate all the parameters. Under these circumstances, the minimization routine wandered off to unrealistic scenarios and rarely attained convergence. It was thus necessary to select fewer and more representative catch rates, along with the yield corresponding to each user group. For the recreational sector, the MRFSS and Headboat indices were selected, and the Logbook-Handline index (in lb/hour) was used to represent the commercial sector (Table 12, Fig. 15). In addition, it was necessary to constrain parameters within reasonable bounds, by fixing some parameters to estimate the others.

Trials to narrow down the search for an absolute minimum included: 1) eliminating some tuning indices; 2) setting $r$ to fixed values to estimate $B_{l} / B_{M S Y}, M S Y$, and $q$; 3)fixing the $B_{l} / B_{M S Y}$ ratio and $r$ at different levels to estimate MSY and $q ; 4$ ) fixing MSY, $r$, and $B_{1} / B_{M S Y}$ to estimate $q ; 5$ ) fixing $B_{1} / B_{M S Y}$ and MSY to estimate $r$ and $q ; 6$ ) further limiting the bounds for MSY and $r$ and fixing only $B_{l} / B_{M S Y}$ at different levels to estimate MSY, $r$, and $q$, and 7) using a first run's results (with constrained parameters) as starting guesses for subsequent runs. The specifics of these sensitivity trials are provided in the Stock Assessment Results section below. Each of these tests resulted in a number of combinations of fixed parameter values and estimates of others, which guided subsequent searches for more reasonable parameter bounds.

## RESULTS AND DISCUSSION

## STANDARDIZED CATCH RATES

## Recreational Sector

Figure 1 shows the frequency distributions of $\log$ CPUE of successful trips, where an approximate normal trend is observed in most cases. Table 1 shows the deviance analysis for each index developed. In each case, the main factors and interactions that exceeded $5 \%$ of the total deviance were considered significant and were selected as the explanatory variables for the positive catch rates and the proportion of positive catch. These variables are highlighted in the tables.

Table 2 shows the results from the random test analyses for each index, and the three criteria statistics used for model selection. The selected model is highlighted.

Standardized CPUE series for each index are presented in Table 3 and Figure 2.

MRFSS Index. The mean catch rate for positive observations was explained by the year*state, year*mode, and year*area fixed factor interactions, even when the area factor by itself was not significant. The major fixed factors determining the proportion of positive catches were year, state, mode, and area, with no significant interactions among them.

Once these sets of fixed factors were selected, we evaluated the first level random interactions between the year and other effects, only for the positive catch rates, since no fixed interactions were observed for the proportion of positives. All the random interactions between year and state, area, mode, and season proved significant and were included in the final run of the model.

The standardized catch rate series follows the same general trend of the nominal series, particularly from 1987 on. The variability observed and the occasional lack of agreement between the standardized and nominal indices may be partially attributed to the very low proportion ( $\sim 2 \%$ ) of positive catches observed in the database.

Headboat Index. Both the mean catch rate for positive observations and the proportion of positive observations were explained by the year, state, season fixed factors and by the year*state interaction. The significant random interactions for positive observations were year*state and year*season. In this case, the area factor was completely eliminated, since its effect was confounded with the state effect and did not provide any additional information.

Positive observations accounted for a higher proportion in the data (~ $40 \%$ ) compared to MRFSS. This is reflected in the good agreement between the observed and standardized indices depicted in the second panel of Figure 2. The peak observed in 1990 corresponds to that observed for the MRFSS catch rates. A large variability, not explained by the model, was observed.

Panama City Index. The mean catch rate for positive observations was explained by year, area, season and year*area interaction. The significant factors and interactions for the proportion of positive catch values were year, state, area, year*state and year*area. The random tests showed significant interactions between year*state, year*area, year* season and area*season for the positive catch.

The standardized index shows a somewhat flat trend and a lack of correspondence with the observed catch rates. This is a portrait of the year factor not being the most important one. In this case, state and area are by far ( $\sim 60 \%$ ) the most influential, as well as the interactions state*year and area*year, so these factors determine the yearly predicted trend. The proportion of positive catches was generally large ( $\sim 40 \%$ ), except for 1989 and 1995, where lower values were observed. The amount of variability not explained by the model was relatively small; the coefficients of variation of the standardized index ranged around $30 \%$.


#### Abstract

Alabama Index. Both the mean catch rate for positive observations and the proportion of positive observations were explained by the year and season fixed factors and by the year*season interaction. This interaction was also significant in the random test for positive observations. A good agreement between the nominal and standard catch rates was observed, even when only four years of data were used. The proportion of positive catches constituted approximately $16 \%$ of the data. The proportion of unexplained variability was fairly reasonable (CVs $\sim 30 \%$ ).


Texas Index. As a result of a highly unbalanced design, it was not possible to standardize catch rates for this data set. The proportion of positive catches accounted for a very small fraction of the data ( $\sim 0.4 \%$ ), which is problematic for CPUE standardization. Even when fixed and random factors and interactions were carefully evaluated, the model fits were generally poor and the amount of unexplained variability remained extremely high for most model configurations (CV>200\%). Therefore the results presented here for TPWD were excluded from all further analyses.

## Commercial Sector

Figure 8 shows the frequency distributions of log CPUE of successful commercial trips (gray triggerfish catch present). Approximate normal distributions are observed in all cases. Table 4 shows the deviance analysis for each commercial index developed. In each case, the main factors and interactions that explained the positive catch rates and the proportion of positive catch are highlighted.

Table 5 shows the results from the random test analyses for each index, and the three criteria statistics used for model selection. The selected model is highlighted.

Standardized CPUE series for each index are presented in Table 6 and Figure 9.
The area factor was found to be nested within state, so the latter was removed as an explanatory variable because area provides more detailed information. The county factor was also removed because the same county identifiers are used in different states in the FLS database, which created confounding effects.

In all CPUEs, the mean catch rate for positive observations was explained by the year and area fixed factors, and the year*area interaction. The same factors and interaction were significant in the proportion of positive catches. It is important to note that deviance values for positive catches vary across indices because different effort units are used and thus the number of observations may also vary if all the effort information is not present for a particular fishing trip. In cases with missing values (effort estimates), observations are omitted from the delta lognormal analysis. On the other hand, the deviance tables for the proportion of positive observations are equal for all the indices because this proportion is a constant in the data base, regardless of the units used to measure effort.

Once these sets of fixed factors were selected, the first level random interaction between the year and area effect were evaluated, both for the positive and the proportion of positive catch rates for each index. The season effect was not considered because it was not significant in the fixed factor evaluation. The random year*area interaction proved significant in all cases and was included in the final run of each model. The random effects evaluation is presented in Table 5.

The nominal and standardized commercial CPUEs are presented in Table 6 and Figure 9. All catch rate estimates follow closely the trend of the nominal series, which may be partially attributed to the large proportion of positive catches observed ( $\sim 65 \%$ ) once the FLS database was filtered (see page 5). In CPUE2 (lb/hooks), there is a certain mismatch between observed and expected values in the first three years (1993-1995) and the variability is relatively low ( $18 \%$ ). Model fit is good for CPUE3 (lb/hook*hr), but the coefficients of variation were the largest observed ( $27 \%$ - $33 \%$ ). The best fit and smallest variability (CVs $\sim 16 \%$ ) was observed for CPUE4 ( $\mathrm{lb} / \mathrm{hr}$ ). This standardized index was thus selected for use in the production model analysis.

## STOCK ASSESSMENT

Initial runs of the production model analysis (ASPIC) failed to converge when no constraints were placed on parameters using the original data set (Table 12). After a number of trials and sensitivity analyses, model fits were slightly improved by fixing parameters at different levels $\left(B_{1} / B_{M S Y}=0.2 ; r=0.75-1.0-1.2\right)$, and constraining the bounds for $r$ and $M S Y$ ( $r=0.5-1.5, M S Y=1-5$ million pounds). Despite the number and variety of trials attempted, ASPIC was still unable to provide reasonable parameter estimates with this data set (1986-1998). It was thus necessary to make additional assumptions regarding the catch-rate time-series. The discrepancies and fluctuations in the catch rates between 1986-1989 may have introduced extra noise to the assessment, so those years were truncated from all further analyses (Figure 15). Year 1993 in the commercial index was also dropped because its opposing trend with the rest of the data made model convergence difficult, so this CPUE included only years 1994-1998.

A similar procedure to that described before (fixing parameters, constraining bounds and conducting various sensitivity runs) was needed to fit the model with the reduced (1990-1998) data set, as ASPIC was again unable, without constraints, to provide reasonable estimates of all parameters. The best fits were obtained by: fixing $r$, fixing
$B_{I} / B_{M S Y}$, fixing $B_{I} / B_{M S Y}$ and $r$, and fixing $B_{I} / B_{M S Y}$ and $r$ at different levels. Results of these sensitivity trials are presented in Figure 16.

## Fixed $r$.

The range of $r$ values examined was $r=0.5-2.0$. Convergence was limited to a reduced $r$ - range ( $r=0.95-1.4$ ) and the best fits were obtained with $r=1.0-1.2$. Over this $r$ range, MSY estimates were between 2.65 and 2.91 million pounds (Figure 16 (A)). These estimates were used as initial guesses for the final bootstrap runs, where no parameters were held constant. Fishing mortality rates were high, between $F=0.8-1.6$, with low stock biomass values.

## Fixed $\boldsymbol{B}_{l} / \boldsymbol{B}_{M S Y}$

Initial biomass ratio levels tested ranged from $B_{I} / B_{M S Y}=0.5$ to 2.0 . Convergence was only attained with $B_{1} / B_{M S Y}=0.5-0.65$, as shown in Figure 16 (B). MSY and $r$ estimates decreased with increasing biomass ratio, and the objective function values increased with increasing $B_{l} / B_{M S Y}$. MSY and $r$ estimates are relatively low ( $M S Y=1.36-1.86$ million pounds; $r=0.14-0.16$ ). Fishing mortality estimates were also relatively low ( $\mathrm{F}=0.2-0.3$ ) compared to all other sensitivity trials.

## Fixed $\boldsymbol{B}_{I} / \boldsymbol{B}_{M S Y}$ and $r$

ASPIC converged with most combinations of $r=0.1-2.0$ and $B_{l} / B_{M S Y}=0.5-1.0$ as seen in Figure 16 (C). For all initial biomass ratio levels, MSY and fishing mortality estimates ( F ) increased with increasing $r$ levels. However, the largest $M S Y$ estimates were obtained with combinations of the smallest $B_{l} / B_{M S Y}=0.5$ and the largest $r$ values. As biomass increased, F estimates decreased, which demonstrates the expected opposing trend between current stock size and fishing mortality. Objective function values show that the model fits best to the low $B$, low $r$ combination. The wide range of parameter values obtained (as seen clearly in the plots presented here) may indicate the range of uncertainty present in the data.

## Fixed $B_{l} / \boldsymbol{B}_{M S Y}$ and $M S Y$

The range of biomass ratio levels examined was from $B_{l} / B_{M S Y}=0.5-1.5$, and $M S Y=1.5$ - 3.5 million pounds. Model convergence was attained with only a few combinations of these fixed parameters, as shown in Figure 16 (D). As biomass values increased, $M S Y$ declined. For all $B$ levels, $r$ estimates declined as $M S Y$ increased. In general, $r$ also increased with increasing levels of $B$ (as seen in Figure 16 (A)). Objective function values were smallest at large $M S Y$, small $B$, and small $r$ values, which is similar to the previous model runs.

## Sensitivity Trials

All the sensitivity trial results discussed above show a range of parameter estimates that could describe the status of the stock. This range might also represent the range of uncertainty present in the data and thus the range of uncertainty in parameter
values. Different parameter combinations give different stock size and fishing mortality rates, however, all sensitivities demonstrate that estimates of current stock level are inversely correlated with fishing mortality rate.

ASPIC was very sensitive to starting values and constraints placed on parameters, particularly of MSY. Model convergence or the lack thereof often depended on the initial values used, even when the rest of the input data remained unchanged. It is possible that local minima were often encountered because the response surface may be too flat, resulting from the limited time-series and the tendency in the catch and effort data used. A single reasonable solution does not appear to exist for this data set, so several scenarios need to be tested and explored.

## Model Projections

In order to show possible scenarios, that include population trajectories and confidence intervals for the parameters, two additional ASPIC runs were performed. These final runs were based on one fit from the fixed $r$ sensitivity analyses $(r=1.0)$. This intrinsic rate of population growth was based on individual growth rate and longevity considerations. One thousand bootstrap trials were used in each case in order to characterize the error associated with population parameter estimates.

The first model (Model 1) assumed fixed $r=1.0$. Resulting parameter estimates from this model were used as initial guess values in the second model (Model 2), where the program was allowed to freely estimate all parameters. The two models thus differ in initial guess values and the number of parameters estimated by ASPIC. Results for both model runs are presented in Table 13 and Figures 17, 18, and 19.

The ASPIC model fits for the observed catch to the indices of abundance resulted in a relatively high R -squared values for both models and the three user groups (Figure 18). In Model 1, R-squared in CPUEs were: $R^{2}=0.97$ (MRFSS), $R^{2}=0.58$ (HEADBOAT), and $\mathrm{R}^{2}=0.75$ (Commercial-Handlines). In Model 2, R -squared values were: $R^{2}=0.96, R^{2}=0.62$, and $R^{2}=0.7$, respectively. Both models appear to capture the major dynamics in CPUEs by fishery.

Population and fishing mortality trajectories for both models (Figure 17) follow similar trends, but a difference in scales is observed, attributed to fixing $r=1.0$ in Model 1. The estimates of the biomass-ratio obtained with both models $\left(\mathrm{B} / \mathrm{B}_{\mathrm{MSY}}(1)=0.208\right.$, $\mathrm{B} / \mathrm{B}_{\mathrm{MSY}}(2)=0.164$ ) denote that the initial biomass in 1999 is estimated to be approximately 20 percent of the biomass the stock would be at if fished at MSY. The Fratio estimates $\left(\mathrm{F} / \mathrm{F}_{\mathrm{MSY}}(1)=1.62, \mathrm{~F} / \mathrm{F}_{\mathrm{MSY}}(2)=1.65\right)$ indicate that the 1998 fishing mortality is about 65 percent higher than that estimated for $\mathrm{F}_{\text {MSY }}$.

Results from these analyses suggest that throughout the period 1990-1998, biomass levels have been below $\mathrm{B}_{\mathrm{MSY}}$, being the lowest in recent years. Accordingly, fishing mortality rates have exceeded $\mathrm{F}_{\text {MSY }}$ throughout the whole period. Therefore, the declining stock size seems consistent with the pattern of exploitation. The MSY estimate obtained with Model 1 ( $\mathrm{MSY}(1)=2.65$ million pounds) is very similar to the largest recorded catch of 2.88 million pounds in 1990. The MSY estimate from Model 2 was somewhat larger ( $\mathrm{MSY}(2)=3.37$ million pounds). This estimate is above the maximum
observed catch since 1986 and the results would imply that historical removals were somewhat larger than this.

Confidence limits around these estimates were constructed by running a bootstrap analysis with the same model inputs. A total of 1,000 trials were run in each case. Associated with the ordinary ASPIC parameter estimates, are bias-corrected estimates, percent bias, and upper and lower $80 \%$ confidence intervals. These estimates are also given in Table 13. Time series of relative biomass and fishing mortality with 80 percent confidence intervals are shown in Figure 19. Diagrams of generic default limit control rules with M assumed equal to 0.2 are included. For both models, the stock appears as overfished since 1990. Fishing mortality rates indicate that overfishing is still occurring. Biomass levels have declined steadily, attaining the lowest levels in 1998. Fishing mortality was estimated as greater that $\mathrm{F}_{\text {MSY }}$ throughout the whole period.

Up to this point in the analysis, there is reasonable evidence that the current rate of removal is not sustainable: a steady decline in landings since the peak in 1990, current landings ( 850,000 pounds in 1998) are below the MSY range, estimated biomass levels are low, and estimated exploitation rates are high. This evidence suggests that the stock is overfished and that catches should at least be held constant if not reduced to bring the population back to sustainable levels. In order to test this hypothesis, projections of the possible future condition of the stock under different fishing scenarios were made using the parameter outputs from the ASPIC models 1 and 2 . Even when some data are available for years 1999 and 2000, it has not been processed for this study, so all projections were made using 1998 as the last year in the assessment and 1999 as the first year of management. To determine whether or not the sock could be rebuilt $\left(\mathrm{B} / \mathrm{B}_{\mathrm{MSY}}=\right.$ 1.0) within a ten year time frame, projections were carried out to the years 1999-2008.

Three fishing scenarios were projected for Models 1 and 2: 1) no fishing for a tenyear period, 2) the 1998 catch repeated for ten years, and 3) the 1998 F value repeated for 10 years. In each case, diagrams of a generic default limit control rule with M assumed equal to 0.2 were constructed.

The first projection assumed that all fishing would completely cease for ten years. Under these circumstances, the stock was estimated to be rebuilt to a level of $\mathrm{B} / \mathrm{B}_{\mathrm{MSY}}=1.0$ in approximately 3 years (2001) in both Models 1 and 2 (Figure 20).

The second projection assumed a constant catch scenario, whereby the 1998 catch value ( 854,000 pounds) is repeated over the ten-year management period. In Model 1, the stock was estimated to be rebuilt to a level of $\mathrm{B} / \mathrm{B}_{\mathrm{MSY}}=1.0$ in approximately 6 years (2004) (Figure 21, panels A, B, C). There is a large uncertainty around this estimate (Figure 21, panel A), which doesn't stabilize until approximately 2006, meaning that it could take the stock up to 8 years to recuperate. Under this scenario, fishing mortality rates are reduced to sustainable levels $\left(\mathrm{F} / \mathrm{F}_{\mathrm{MSY}}=1.0\right)$ within a 3 year period (2001), but uncertainty around these estimates is also rather high (Figure 21, panel B). In Model 2, the stock was estimated to be rebuilt in approximately 4 years (2003), which is faster than in Model 1 (Figure $21 \mathrm{D}, \mathrm{E}, \mathrm{F}$ ). Fishing mortality rate is reduced to $\mathrm{F} / \mathrm{F}_{\mathrm{MSY}}=1.0$ in approximately 2 years (2000).

The third projection assumed the 1998 fishing mortality rates repeated for ten years ( $\mathrm{F}=0.81$ in Model $1, \mathrm{~F}=1.052$ in Model 2). In this projection, estimates of $\mathrm{B} / \mathrm{B}_{\mathrm{MSY}}$
increased slowly during the management period, but did not attain sustainable levels within the ten-year frame (Figure 22). In Model 1, the estimated mean $\mathrm{B} / \mathrm{B}_{\mathrm{MSY}}$ in ten years was 0.59 , and in Model 2, $\mathrm{B} / \mathrm{B}_{\mathrm{MSY}}=0.22$, which indicates that if the fishery continued to operate at either of these fishing mortality rates, the stock would not recover (ie. $\mathrm{B} / \mathrm{B}_{\mathrm{MSY}} \geq 1$ ) within the ten-year modeling projection. With Model 1, yield would increase even at this low biomass level, and with Model 2, yield would remain relatively low, near the current (1998) level, or even decline.

These projections indicate that, indeed, as was suggested before, the stock may be overfished, that overfishing is still occurring, and that catches should at least be held constant or preferably reduced to bring the stock back to healthy levels.

In conclusion, the production model analyses utilized here to project the stock trajectory indicate that current fishing mortality rates are not sustainable.

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TEXAS


Figure 1. Frequency distributions of logarithm CPUE of successful trips of gray triggerfish from recreational survey data. The plots show density on $\operatorname{lnCPUE}$. Catch rates are given in numbers of fish (fish/angler hour). The line represents the estimated normal distribution of the data.

Table 1. Deviance analysis tables for recreational gray triggerfish catch rates using the delta lognormal model. Proportion positive/total observations assumed a binomial error distribution. The dependent variable is the total number of fish caught per hour per angler. P refers to the Chi-square test probability (alpha=5\%) test between two consecutive model specifications. Factors and interactions with total deviance $\$ 5 \%$ were selected and are shown in shaded areas.

## RECREATIONAL SECTOR

| MRFSS |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Model factors positive catch rates values | d. f. | Residual deviance | Change in deviance | \% of total deviance | $p$ |
| NULL | 1 | 6362.71 |  |  |  |
| YEAR | 18 | 5840.91 | 521.8 | 33.7\% | < 0.001 |
| ... + STATE | 4 | 5650.55 | 190.4 | 12.3\% | < 0.001 |
| $\ldots+$ MODE | 3 | 5631.02 | 19.5 | 1.3\% | < 0.001 |
| $\ldots+$ AREA | 4 | 5588.24 | 42.8 | 2.8\% | < 0.001 |
| $\ldots+$ TARGET | 1 | 5518.19 | 70.0 | 4.5\% | < 0.001 |
| $\ldots+$ SEASON | 2 | 5465.22 | 53.0 | 3.4\% | < 0.001 |
| $\ldots+$ YEAR:STATE | 50 | 5157.96 | 307.3 | 19.8\% | < 0.001 |
| $\ldots+$ YEAR:MODE | 35 | 5015.91 | 142.0 | 9.2\% | < 0.001 |
| $\ldots+$ YEAR:AREA | 57 | 4908.29 | 107.6 | 7.0\% | < 0.001 |
| $\ldots+$ YEAR:TARGET | 14 | 4883.85 | 24.4 | 1.6\% | 0.040 |
| $\ldots+$ YEAR:SEASON | 36 | 4814.54 | 69.3 | 4.5\% | < 0.001 |
| Model factors proportion positive catch rates values | d. f. | Residual deviance | Change in deviance | \% of total deviance | $p$ |
| NULL | 1 | 19132.19 |  |  |  |
| YEAR | 18 | 18557.51 | 574.7 | 3.6\% | < 0.001 |
| $\ldots+$ STATE | 22 | 16491.44 | 2066.1 | 12.8\% | < 0.001 |
| $\ldots+$ SEASON | 24 | 16281.05 | 210.4 | 1.3\% | < 0.001 |
| $\ldots+$ MODE | 27 | 6114.02 | 10167.0 | 63.1\% | < 0.001 |
| $\ldots+$ AREA | 31 | 3834.53 | 2279.5 | 14.1\% | < 0.001 |
| $\ldots+$ TARGET | 32 | 3212.73 | 621.8 | 3.9\% | < 0.001 |
| $\ldots+$ YEAR*SEASON | 68 | 3010.56 | 202.2 | 1.3\% | <0.001 |


| HEADBOAT |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Model factors positive catch rates values | d. f. | Residual deviance | Change in deviance | \% of total deviance | $p$ |
| NULL | 0 | 86811.59 |  |  |  |
| YEAR | 12 | 85135.20 | 1676.4 | 6.3\% | $<0.001$ |
| $\ldots+$ STATE | 3 | 64349.30 | 20785.9 | 78.7\% | < 0.001 |
| $\ldots+$ SEASON | 2 | 62959.73 | 1389.6 | 5.3\% | < 0.001 |
| $\ldots+$ YEAR:STATE | 36 | 60763.72 | 2196.0 | 8.3\% | < 0.001 |
| $\ldots+$ YEAR:SEASON | 24 | 60396.07 | 367.7 | 1.4\% | < 0.001 |
| Model factors proportion positive catch rates values | d. f. | Residual deviance | Change in deviance | \% of total deviance | $p$ |
| NULL | 1 | 46.52381 |  |  |  |
| YEAR | 12 | 44.83039 | 1.69342 | 4.0\% | 1.000 |
| $\ldots$ + STATE | 3 | 7.9974 | 36.83299 | 87.6\% | < 0.001 |
| $\ldots+$ SEASON | 2 | 7.42834 | 0.56906 | 1.4\% | 0.752 |
| $\ldots+$ YEAR:STATE | 36 | 5.27057 | 2.15777 | 5.1\% | 1.000 |
| $\ldots+$ YEAR:SEASON | 24 | 4.47764 | 0.79293 | 1.9\% | 1.000 |

Table 1. (Continued).

| PANAMA CITY |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Model factors positive catch rates values | d. f. | Residual deviance | Change in deviance | \% of total deviance | $p$ |
| NULL | 0 | 8113.768 |  |  |  |
| YEAR | 7 | 7690.60 | 423.172 | 22.7\% | $<0.001$ |
| $\ldots+$ AREA | 8 | 6736.15 | 954.4 | 51.2\% | < 0.001 |
| $\ldots+$ SEASON | 2 | 6597.41 | 138.7 | 7.4\% | $<0.001$ |
| $\ldots$ + YEAR:AREA | 39 | 6346.85 | 250.6 | 13.4\% | < 0.001 |
| $\ldots+$ YEAR:SEASON | 13 | 6285.42 | 61.4 | 3.3\% | < 0.001 |
| $\ldots+$ AREA:SEASON | 15 | 6249.67 | 35.8 | 1.9\% | 0.002 |
| Model factors proportion positive catch rates values | d. f. | Residual deviance | Change in deviance | \% of total deviance | $p$ |
| NULL | 0 | 86.75046 |  |  |  |
| YEAR | 7 | 82.24468 | 4.50578 | 5.7\% | 0.7200 |
| ... + STATE | 4 | 46.66724 | 35.57744 | 44.9\% | < 0.001 |
| $\ldots+$ AREA | 4 | 28.02606 | 18.64118 | 23.5\% | < 0.001 |
| $\ldots+$ SEASON | 2 | 26.95972 | 1.06634 | 1.3\% | 0.5867 |
| ... + YEAR:STATE | 25 | 17.8025 | 9.15722 | 11.6\% | 0.9984 |
| $\ldots+$ YEAR:AREA | 22 | 11.74364 | 6.05886 | 7.6\% | 0.9997 |
| $\ldots+$ YEAR:SEASON | 13 | 10.75838 | 0.98526 | 1.2\% | 1.0000 |
| $\ldots+$ SEASON:AREA | 16 | 7.4894 | 3.26898 | 4.1\% | 0.9997 |



| TEXAS |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Model factors positive catch rates values | d. f. | Residual deviance | Change in deviance | \% of total deviance | $p$ |
| NULL | 1 | 562.3247 |  |  |  |
| YEAR | 15 | 552.5328 | 9.7919 | 8.0\% | 0.8326 |
| $\ldots+$ AREA | 2 | 545.1547 | 7.3781 | 6.0\% | 0.0250 |
| $\ldots+$ SEASON | 2 | 540.3293 | 4.8254 | 3.9\% | 0.0896 |
| $\ldots+$ YEAR:AREA | 30 | 480.4296 | 59.8997 | 48.9\% | < 0.001 |
| $\ldots+$ YEAR:SEASON | 26 | 449.408 | 31.0216 | 25.3\% | 0.2275 |
| $\ldots+$ AREA:SEASON | 4 | 439.731 | 9.677 | 7.9\% | 0.0462 |
| Model factors proportion positive catch rates values | d. f. | Residual deviance | Change in deviance | \% of total deviance | $p$ |
| NULL | 1 | 7757.3524 |  |  |  |
| YEAR | 15 | 7679.8294 | 77.523 | 4.2\% | $<0.001$ |
| $\ldots+$ AREA | 17 | 5955.0071 | 1724.8223 | 92.7\% | < 0.001 |
| $\ldots+$ SEASON | 19 | 5895.9803 | 59.0268 | 3.2\% | <0.001 |

Table 2. Recreational sector. Random effects evaluation for delta lognormal mixed model specifications. Highlighted rows refer to the final model.

## RECREATIONAL SECTOR

| MRFSS |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| RANDOM TESTS | -2 RES Log likelihood | Akaike's Information Criterion | Schwartz's Bayesian Criterion | Likelihood Ratio Test |  |
| Positive Catch |  |  |  |  |  |
| Year State Season Mode Area | 13283.58 | -6642.79 | -6645.96 |  |  |
| Year State Season Mode Area Year*Area | 13170.65 | -6587.33 | -6593.67 | 112.93 | $2.2353 \mathrm{E}-26$ |
| Year State Season Mode Area Year*Area Year*State | 13149.52 | -6577.76 | -6587.28 | 21.13 | $4.2916 \mathrm{E}-06$ |
| Year State Season Mode Area Year*Area Year*State Year*Season | 13137.03 | -6572.51 | -6585.21 | 12.49 | $4.0914 \mathrm{E}-04$ |
| Year State Season Mode Area Year*Area Year*State Year*Season Year*Mod | 13084.87 | -6547.43 | -6563.3 | 52.16 | $5.1158 \mathrm{E}-13$ |


| HEADBOAT |  |  |  |
| :--- | :--- | :--- | :--- | :--- |
| RANDOM TESTS | A2 RES Log <br> likelihood | Akaike's <br> Information <br> Criterion | Schwartz's <br> Bayesian <br> Criterion |
| Positive Catch | 159299.1 | -79650.6 | Likelihood Ratio Test |


| PANAMA CITY |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| RANDOM TESTS | -2 RES Log <br> likelihood | Akaike's Information Criterion | Schwartz's Bayesian Criterion | Likelihood Ratio Test |  |
| Positive Catch |  |  |  |  |  |
| Year State Area Season | 17785.39 | -8893.69 | -8897.05 |  |  |
| Year State Area Season Year*State | 17695.65 | -8849.82 | -8856.53 | 89.74 | $2.7161 \mathrm{E}-21$ |
| Year State Area Season Year*State Year*Area | 17650.46 | -8828.23 | -8838.29 | 45.19 | 1.7882E-11 |
| Year State Area Season Year*State Year*Area Year*Season | 17622.96 | -8815.48 | -8828.89 | 27.5 | $1.5709 \mathrm{E}-07$ |
| Year State Area Season Year*State Year*Area Year*Season Area*Season | 17612.86 | -8811.43 | -8828.2 | 10.1 | $1.4827 \mathrm{E}-03$ |


| ALABAMA |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| RANDOM TESTS | -2 RES Log likelihood | Akaike's Information Criterion | Schwartz's Bayesian Criterion | Likelihood | o Test |
| Positive Catch |  |  |  |  |  |
| Year Season | 4730.526 | -2366.26 | -2368.93 |  |  |
| Year Season Year*Seasor. | 4721.647 | -2362.82 | -2368.16 | 8.879 | $2.8847 \mathrm{E}-03$ |


| TEXAS |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| RANDOM TESTS | -2 RES Log likelihood | Akaike's Information Criterion | Schwartz's Bayesian Criterion | Likelihood Ratio Test |  |
| Positive Catch |  |  |  |  |  |
| Year Area Season | 1511.413 | -756.707 | -758.805 |  |  |
| Year Area Season Year*Season | 1510.967 | -757.483 | -761.68 | 0.446 | $5.04 \mathrm{E}-01$ |
| Year Area Season Year*Season Year*Area | 1501.881 | -753.941 | -760.235 | 9.086 | $2.58 \mathrm{E}-03$ |
| Year Area Season Year*Season Year*Area Area* Season | 1500.565 | -754.283 | -762.676 | 1.316 | $2.51 \mathrm{E}-01$ |
| Proportion Positives |  |  |  |  |  |
| Year Area Season | 383.5551 | -192.778 | -194.167 |  |  |
| Year Area Season Year*Area | 371.0854 | -187.543 | -190.322 | 12.4697 | $4.14 \mathrm{E}-04$ |
| Year Area Season Year*Area Year*Season | 371.0854 | -188.543 | -192.711 | 0 | $1.00 \mathrm{E}+00$ |
| Year Area Season Year*Area Area*Season | 361.2482 | -183.624 | -187.793 | 9.8372 | $1.71 \mathrm{E}-03$ |
| Year Area Season Year*Season Year*Area Area* Season | 360.9467 | -184.473 | -190.032 | 0.3015 | 5.83E-01 |

Table 3. Nominal and delta lognormal standardized CPUE indices for gray triggerfish from recreational survey data. CPUE units are number of fish/angler hour.


| Year | Nominal | Standard | CV |
| ---: | ---: | ---: | :--- |
| 1981 | 0.01454 | 0.02923 | $47.10 \%$ |
| 1982 | 0.00386 | 0.02641 | $42.87 \%$ |
| 1983 | 0.00698 | 0.01899 | $44.68 \%$ |
| 1984 | 0.00357 | 0.00847 | $56.61 \%$ |
| 1985 | 0.00272 | 0.00985 | $57.15 \%$ |
| 1986 | 0.02873 | 0.04165 | $36.44 \%$ |
| 1987 | 0.00781 | 0.01809 | $38.44 \%$ |
| 1988 | 0.01407 | 0.05353 | $35.17 \%$ |
| 1989 | 0.01885 | 0.08643 | $34.80 \%$ |
| 1990 | 0.02307 | 0.10552 | $36.41 \%$ |
| 1991 | 0.02542 | 0.08428 | $33.97 \%$ |
| 1992 | 0.01564 | 0.07310 | $31.18 \%$ |
| 1993 | 0.00864 | 0.05082 | $34.13 \%$ |
| 1994 | 0.00874 | 0.05032 | $32.73 \%$ |
| 1995 | 0.00575 | 0.03895 | $35.97 \%$ |
| 1996 | 0.00707 | 0.03325 | $36.22 \%$ |
| 1997 | 0.00675 | 0.03270 | $33.70 \%$ |
| 1998 | 0.00900 | 0.03524 | $32.89 \%$ |
| 1999 | 0.00768 | 0.03021 | $31.53 \%$ |

PANAMA CITY

| Year | Nominal | Standard | CV |
| ---: | ---: | ---: | :--- |
| 1989 | 0.09282 | 0.28874 | $30.60 \%$ |
| 1990 | 0.26740 | 0.37156 | $32.15 \%$ |
| 1991 | 0.48079 | 0.31247 | $29.33 \%$ |
| 1992 | 0.35190 | 0.38472 | $28.49 \%$ |
| 1993 | 0.51404 | 0.33463 | $34.49 \%$ |
| 1994 | 0.56815 | 0.44076 | $28.52 \%$ |
| 1995 | 0.22909 | 0.44667 | $27.73 \%$ |

HEADBOAT

| Year | Nominal | Standard | CV |
| ---: | ---: | ---: | ---: |
| 1986 | 0.00851 | 0.00828 | $147.78 \%$ |
| 1987 | 0.00894 | 0.00833 | $144.09 \%$ |
| 1988 | 0.01762 | 0.01180 | $114.25 \%$ |
| 1989 | 0.02219 | 0.01730 | $88.38 \%$ |
| 1990 | 0.03801 | 0.02705 | $61.15 \%$ |
| 1991 | 0.02595 | 0.02371 | $70.35 \%$ |
| 1992 | 0.03010 | 0.02425 | $65.64 \%$ |
| 1993 | 0.02459 | 0.02351 | $67.24 \%$ |
| 1994 | 0.02582 | 0.01938 | $77.01 \%$ |
| 1995 | 0.02384 | 0.01372 | $97.60 \%$ |
| 1996 | 0.02357 | 0.01477 | $94.91 \%$ |
| 1997 | 0.01990 | 0.01118 | $106.56 \%$ |
| 1998 | 0.01768 | 0.01026 | $117.38 \%$ |

## ALABAMA

| Year | Nominal | Standard | CV |
| ---: | ---: | ---: | ---: |
| 1991 | 0.84120 | 0.14492 | $28.46 \%$ |
| 1992 | 0.92154 | 0.13550 | $30.19 \%$ |
| 1993 | 0.66591 | 0.07176 | $36.06 \%$ |
| 1994 | 0.66605 | 0.08390 | $34.02 \%$ |

## TEXAS

| Year | Nominal | Standard | CV |
| :---: | ---: | ---: | :---: |
| 1983 | 0.00037 | 0.00300 | $364.21 \%$ |
| 1984 | 0.00065 | 0.00306 | $357.50 \%$ |
| 1985 | 0.00093 | 0.00335 | $312.86 \%$ |
| 1986 | 0.00050 | 0.00184 | $502.57 \%$ |
| 1987 | 0.00042 | 0.00245 | $384.43 \%$ |
| 1988 | 0.00102 | 0.00604 | $230.39 \%$ |
| 1989 | 0.00125 | 0.00505 | $266.15 \%$ |
| 1990 | 0.00134 | 0.00406 | $287.77 \%$ |
| 1991 | 0.00065 | 0.00277 | $369.28 \%$ |
| 1992 | 0.00148 | 0.00329 | $332.67 \%$ |
| 1993 | 0.00072 | 0.00364 | $315.74 \%$ |
| 1994 | 0.00079 | 0.00456 | $265.87 \%$ |
| 1995 | 0.00067 | 0.00252 | $370.78 \%$ |
| 1996 | 0.00066 | 0.00236 | $409.34 \%$ |
| 1997 | 0.00063 | 0.00199 | $425.23 \%$ |
| 1998 | 0.00012 | 0.00100 | $788.95 \%$ |



Figure 2. Scaled nominal and delta lognormal standardized catch rates (CPUE) of gray triggerfish from recreational survey data. CPUE units are number of fish/angler hour. The solid line represents the average of the standardized catch rates ( $\pm 95 \% \mathrm{CI}$ ); the dotted line represents the nominal average CPUE.


| GEAR | \% LANDINGS |
| :--- | :---: |
| Handlines | $85.96 \%$ |
| Traps | $8.46 \%$ |
| Longlines | $3.32 \%$ |
| Trolling | $1.19 \%$ |
| Other + Unknown | $0.74 \%$ |
| Diving | $0.33 \%$ |

Figure 3. Commercial logbook data. Percent gray triggerfish landings by gear.


Figure 4. Cumulative percentage of vessels that caught gray triggerfish in the Gulf of Mexico during the period 1990-1999 from the Florida Logbook System (FLS) database.


Figure 5. Nominal commercial CPUEs by number of years where individual vessels caught gray triggerfish in the period 1990-1999. NYRSOBS = number of years where gray triggerfish was observed in the catch. Units are: CPUE1= pounds/angler*hour, CPUE2= pounds/hook, CPUE3= pounds/hook*hour, CPUE4=pounds/hour.


Figure 6. Nominal commercial catch rates for Gulf of Mexico gray triggerfish by number of years the species was caught by individual vessels.


Figure 7. Nominal commercial CPUE4 with gray triggerfish present in the catch for 5 years or more, selected for standardization and use in production model analysis.


LGCPUE3



LGCPUE4


Figure 8. Frequency distributions of logCPUE of successful trips of gray triggerfish from commercial logbook data (handlines only). The plots show density on $\operatorname{lnCPUE}$. The line represents the estimated normal distribution of the data. Catch rates are given in: CPUE1=pounds/angler*hour, CPUE2=pounds/hook, CPUE3= pounds/hook*hour, CPUE4=pounds/hour.

Table 4. Deviance analysis tables for commercial gray triggerfish catch rates using the delta lognormal model. Separate analyses were conducted for different units of fishing effort and CPUE. Units for each index are given in parenthesis. Proportion positive/total observations assumed a binomial error distribution. $P$ refers to the Chi-square test probability (alpha=5\%) test between two consecutive model specifications. Factors and interactions with total deviance $\$ 5 \%$ were selected and are shown in shaded areas.

## COMMERCIAL SECTOR (HANDLINES)

| CPUE2 (lb/hook) |  | Residual deviance | Change in deviance | \% of total deviance |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Model factors positive catch rates values | d-f |  |  |  | 0 |
| NULL | 1 | 45769.14 |  |  |  |
| YEAR | 6 | 44600.79 | 1168.4 | 9.6\% | $<0.001$ |
| $\ldots$ + AREA | 22 | 35563.09 | 9037.7 | 74.1\% | < 0.001 |
| $\ldots$ + SEASON | 2 | 35030.55 | 532.5 | 4.4\% | < 0.001 |
| ... + YEAR:AREA | 124 | 33773.07 | 1257.5 | 10.3\% | $<0.001$ |
| $\ldots+$ YEAR:SEASON | 12 | 33573.92 | 199.2 | 1.6\% | < 0.001 |
| Model factors proportion positive catch rates values | d.f. | Residual deviance | Change in deviance | \% of total deviance | $p$ |
| NULL | 1 | 129.53 |  |  |  |
| YEAR | 6 | 128.57 | 1.0 | 0.9\% | 0.987 |
| ... + AREA | 22 | 44.28 | 84.3 | 81.7\% | $<0.001$ |
| ... + SEASON | 2 | 43.14 | 1.1 | 1.1\% | 0.565 |
| ... + YEAR:AREA | 130 | 26.92 | 16.2 | 15.7\% | 1.000 |
| $\ldots+$ YEAR:SEASON | 12 | 26.38 | 0.5 | 0.5\% | 1.000 |


| CPUE3 (lbs/hook*hr) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Model factors positive catch rates values | d.f. | Residual deviance | Change in deviance | \% of total deviance | $p$ |
| NULL | 1 | 50687.23 |  |  |  |
| YEAR | 6 | 49284.99 | 1402.2 | 16.4\% | < 0.001 |
| ... + AREA | 22 | 44065.93 | 5219.1 | 61.2\% | < 0.001 |
| . + SEASON | 2 | 44045.84 | 20.1 | 0.2\% | < 0.001 |
| $\ldots$ + YEAR:AREA | 124 | 42421.13 | 1624.7 | 19.1\% | < 0.001 |
| ... + YEAR:SEASON | 12 | 42159.77 | 261.4 | 3.1\% | < 0.001 |
| Model factors proportion positive catch rates values | d. f. | Residual deviance | Change in deviance | \% of total deviance | $p$ |
| NULL | 1 | 129.53 |  |  |  |
| YEAR | 6 | 128.57 | 1.0 | 0.9\% | 0.987 |
| ... + AREA | 22 | 44.28 | 84.3 | 81.7\% | < 0.001 |
| ... + SEASON | 2 | 43.14 | 1.1 | 1.1\% | 0.565 |
| $\ldots+$ YEAR:AREA | 130 | 26.92 | 16.2 | 15.7\% | 1.000 |
| $\ldots+$ YEAR:SEASON | 12 | 26.38 | 0.5 | 0.5\% | 1.000 |


| CPUE4 (lbs/hr) |  | Residual deviance | Change in deviance | \% of total deviance |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Model factors positive catch rates values | d.f |  |  |  | 0 |
| NULL | 1 | 43472.74 |  |  |  |
| YEAR | 6 | 42983.21 | 489.5 | 6.8\% | $<0.001$ |
| $\ldots$. + AREA | 23 | 37597.25 | 5386.0 | 75.1\% | < 0.001 |
| ... + SEASON | 2 | 37566.67 | 30.6 | 0.4\% | < 0.001 |
| ... + YEAR:AREA | 130 | 36425.63 | 1141.0 | 15.9\% | < 0.001 |
| $\ldots+$ YEAR:SEASON | 12 | 36302.77 | 122.9 | 1.7\% | $<0.001$ |
| Model factors proportion positive catch rates values | d.f. | Residual deviance | Change in deviance | \% of total deviance | 0 |
| NULL | 1 | 129.53 |  |  |  |
| YEAR | 6 | 128.57 | 1.0 | 0.9\% | 0.987 |
| $\ldots$. AREA | 22 | 44.28 | 84.3 | 81.7\% | < 0.001 |
| ... + SEASON | 2 | 43.14 | 1.1 | 1.1\% | 0.565 |
| $\ldots+$ YEAR:AREA | 130 | 26.92 | 16.2 | 15.7\% | 1.000 |
| $\ldots+$ YEAR:SEASON | 12 | 26.38 | 0.5 | 0.5\% | 1.000 |

Table 5. Commercial sector. Random effects evaluation for delta lognormal mixed model specifications. Separate analyses were conducted for different units of fishing effort and CPUE. Units for each index are given in parenthesis. Highlighted rows refer to the final model.

## COMMERCIAL SECTOR (HANDLINES)

| CPUE2 (lb/hook) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| RANDOM TESTS | -2 RES Log likelihood | Akaike's Information Criterion | Schwartz's Bayesian Criterion | Likelihood Ratio Test |  |
| Positive Catch |  |  |  |  |  |
| Year Area | 54088.81 | -27045.4 | -27049.2 |  |  |
| Year Area Year*Area | 53858.31 | -26931.2 | -26938.7 | 230.5 | 0.0000 |
| Proportion Positive |  |  |  |  |  |
| Year Area | 233.7685 | -117.884 | -119.318 |  |  |
| Year Area Year*Area | 228.8015 | -116.401 | -119.268 | 4.967 | 0.0258 |


| CPUE3 (lbs/hook*hr) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| RANDOM TESTS | -2 RES Log likelihood | Akaike's Information Criterion | Schwartz's <br> Bayesian <br> Criterion | Likelihood Ratio Test |  |
| Positive Catch Year Area | 56882.11 | -28442.1 | -28445.8 |  |  |
| Year Area Year*Area | 56638.56 | -28321.3 | -28328.8 | 243.55 | 0.0000 |
| Proportion Positive Year Area | 233.7685 | -117.884 | -119.318 |  |  |
| Year Area Year*Area | 228.8015 | -116.401 | -119.268 | 4.967 | 0.0258 |


| CPUE4 (lbs/hr) |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | :--- |
|  | -2 RES Log <br> likelihood | Akaike's <br> Information <br> Criterion | Schwartz's <br> Bayesian <br> Criterion | Likelihood Ratio Test |  |
| RANDOM TESTS |  |  |  |  |  |
| Positive Catch | 51836.7 | -25919.4 | -25923.1 |  |  |
| Year Area | 51579.01 | -25791.5 | -25799.1 | 257.69 | 0.0000 |
| Year Area Year*Area |  |  |  |  |  |
| Proportion Positive | 233.7685 | -117.884 | -119.318 |  |  |
| Year Area | 228.8015 | -116.401 | -119.268 | 4.967 | 0.0258 |
| Year Area Year*Area |  |  |  |  |  |

Table 6. Nominal and delta lognormal standardized CPUE indices for gray triggerfish from commercial logbook survey (handlines). Separate analyses were conducted with different CPUE units. Units for each index are given in parenthesis.

CPUE2 (lb/hook)

| Year | Nominal | Standard | CV |
| :---: | ---: | ---: | :---: |
| 1993 | 1.000 | 0.836 | $18.25 \%$ |
| 1994 | 0.865 | 1.000 | $17.84 \%$ |
| 1995 | 0.907 | 0.711 | $18.57 \%$ |
| 1996 | 0.631 | 0.687 | $18.02 \%$ |
| 1997 | 0.535 | 0.651 | $18.02 \%$ |
| 1998 | 0.463 | 0.545 | $18.19 \%$ |
| 1999 | 0.422 | 0.560 | $18.24 \%$ |

CPUE3 (lb/hook*hr)

| Year | Nominal | Standard | CV |
| :---: | ---: | ---: | :---: |
| 1993 | 0.835 | 0.807 | $29.51 \%$ |
| 1994 | 1.000 | 1.000 | $27.23 \%$ |
| 1995 | 0.901 | 0.843 | $29.75 \%$ |
| 1996 | 0.596 | 0.664 | $30.00 \%$ |
| 1997 | 0.480 | 0.611 | $30.54 \%$ |
| 1998 | 0.437 | 0.526 | $32.26 \%$ |
| 1999 | 0.332 | 0.495 | $32.82 \%$ |

CPUE4 (lb/hr)

| Year | Nominal | Standard | CV |
| ---: | ---: | ---: | ---: |
| 1993 | 0.924 | 0.804 | $15.99 \%$ |
| 1994 | 1.000 | 1.000 | $15.60 \%$ |
| 1995 | 0.958 | 0.928 | $16.18 \%$ |
| 1996 | 0.752 | 0.711 | $15.70 \%$ |
| 1997 | 0.680 | 0.658 | $15.68 \%$ |
| 1998 | 0.623 | 0.610 | $15.78 \%$ |
| 1999 | 0.532 | 0.594 | $15.85 \%$ |

Figure 9. Scaled nominal and delta lognormal standardized catch rates (CPUE) of gray triggerfish from commercial logbook data. Units are: CPUE2=lbs/hook, CPUE3= lbs/hook*hr, CPUE4=lbs/hour. The solid line represents the average of the standardized catch rates $( \pm 95 \% \mathrm{CI})$; the dotted line represents the nominal average CPUE.





Fork Length (in)


Figure 10. Size and weight frequency distributions by year for gray triggerfish from all recreational surveys combined (MRFSS, HEADBOAT, TPWD) for years 1979-1999.

Table 7. Mean size (fork length) and mean weight (whole) of gray triggerfish measured from the MRFSS (1981-1999), HEADBOAT (1986-1998) and TPWD (1983-1998) recreational surveys.

| MRFSS |  |  |  |  |  |  |  | N obs | Mean weight (lb) | SD |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | N obs | Mean FL (in) | SD | N |  |  |  |  |  |  |
| 1981 | 80 | 12.99 | 2.46 | 81 | 1.84 | 1.45 |  |  |  |  |
| 1982 | 188 | 12.67 | 2.39 | 188 | 1.91 | 1.04 |  |  |  |  |
| 1983 | 139 | 13.72 | 2.75 | 139 | 2.46 | 1.45 |  |  |  |  |
| 1984 | 68 | 14.57 | 3.70 | 69 | 3.23 | 3.61 |  |  |  |  |
| 1985 | 46 | 14.22 | 2.64 | 49 | 2.88 | 1.61 |  |  |  |  |
| 1986 | 123 | 14.71 | 2.83 | 128 | 2.97 | 1.92 |  |  |  |  |
| 1987 | 422 | 13.92 | 2.77 | 424 | 2.42 | 1.47 |  |  |  |  |
| 1988 | 397 | 14.00 | 2.54 | 420 | 2.49 | 1.78 |  |  |  |  |
| 1989 | 210 | 13.37 | 2.40 | 230 | 2.00 | 1.15 |  |  |  |  |
| 1990 | 313 | 13.15 | 2.14 | 323 | 2.26 | 1.35 |  |  |  |  |
| 1991 | 658 | 13.83 | 2.33 | 667 | 2.46 | 1.30 |  |  |  |  |
| 1992 | 1412 | 13.24 | 2.35 | 1436 | 2.18 | 1.30 |  |  |  |  |
| 1993 | 400 | 13.29 | 2.31 | 401 | 2.17 | 1.30 |  |  |  |  |
| 1994 | 381 | 13.28 | 2.40 | 392 | 2.18 | 1.32 |  |  |  |  |
| 1995 | 325 | 12.59 | 1.94 | 340 | 1.84 | 0.98 |  |  |  |  |
| 1996 | 187 | 12.93 | 2.11 | 195 | 1.96 | 1.20 |  |  |  |  |
| 1997 | 501 | 13.69 | 2.47 | 515 | 2.30 | 1.34 |  |  |  |  |
| 1998 | 1374 | 12.90 | 2.09 | 1379 | 1.88 | 1.05 |  |  |  |  |
| 1999 | 2128 | 12.59 | 1.90 | 2128 | 1.78 | 0.92 |  |  |  |  |
| Total | 9352 |  |  | 9504 |  |  |  |  |  |  |

HEADBOAT

| Year | N obs | Mean FL (in) | SD | N obs | Mean weight (Ib) | SD |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1986 | 469 | 12.67 | 2.59 | 546 | 1.99 | 1.37 |
| 1987 | 552 | 12.32 | 2.61 | 607 | 1.81 | 1.87 |
| 1988 | 597 | 12.42 | 2.48 | 676 | 1.74 | 1.48 |
| 1989 | 1352 | 12.21 | 2.40 | 1458 | 1.64 | 1.16 |
| 1990 | 2071 | 12.13 | 2.27 | 2161 | 1.64 | 2.36 |
| 1991 | 1638 | 12.46 | 2.09 | 1666 | 1.68 | 0.93 |
| 1992 | 2499 | 12.11 | 2.06 | 2510 | 1.58 | 0.91 |
| 1993 | 1373 | 12.92 | 2.13 | 1375 | 1.91 | 1.07 |
| 1994 | 2137 | 12.54 | 2.02 | 2167 | 1.77 | 1.00 |
| 1995 | 1735 | 12.70 | 1.90 | 1760 | 1.81 | 0.89 |
| 1996 | 1501 | 12.68 | 2.08 | 1564 | 1.84 | 1.61 |
| 1997 | 1149 | 12.30 | 2.04 | 1218 | 1.65 | 0.94 |
| 1998 | 1486 | 12.30 | 2.01 | 1586 | 1.66 | 1.00 |
| Total | 18559 |  |  | 19294 |  |  |


| TEXAS |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | N obs | Mean FL (in) | SD | N obs | Mean weight (lb) | SD |  |
| 1983 | 153 | 10.98 | 2.15 | 153 | 1.24 | 0.84 |  |
| 1984 | 175 | 11.26 | 2.41 | 175 | 1.35 | 0.86 |  |
| 1985 | 93 | 9.02 | 1.64 | 93 | 0.68 | 0.39 |  |
| 1986 | 49 | 10.19 | 1.75 | 49 | 0.97 | 0.54 |  |
| 1987 | 80 | 10.51 | 1.90 | 80 | 1.07 | 0.64 |  |
| 1988 | 137 | 10.59 | 1.82 | 137 | 1.09 | 0.67 |  |
| 1989 | 92 | 10.33 | 1.73 | 92 | 1.00 | 0.49 |  |
| 1990 | 115 | 11.14 | 1.46 | 115 | 1.22 | 0.46 |  |
| 1991 | 80 | 10.72 | 1.93 | 80 | 1.13 | 0.61 |  |
| 1992 | 93 | 11.43 | 1.90 | 93 | 1.35 | 0.88 |  |
| 1993 | 95 | 11.95 | 1.96 | 95 | 1.53 | 0.73 |  |
| 1994 | 149 | 12.39 | 1.71 | 149 | 1.67 | 0.63 |  |
| 1995 | 134 | 12.59 | 1.43 | 134 | 1.72 | 0.56 |  |
| 1996 | 83 | 12.05 | 1.82 | 83 | 1.57 | 0.72 |  |
| 1997 | 100 | 12.66 | 2.43 | 100 | 1.87 | 1.30 |  |
| 1998 | 24 | 12.73 | 1.42 | 24 | 1.77 | 0.55 |  |
| Total | 1652 |  |  | 1652 |  |  |  |



Figure 11. Mean size FL (in) and mean whole weight (lbs) of gray triggerfish measured from recreational surveys.

Table 8. Mean weight estimates for recreationally harvested gray triggerfish by year, state and fishing mode. Conversions from fork length used $\mathrm{Wt}=(8.975 \mathrm{E}-4) \mathrm{FL}^{2.96}$ and from total length used $\mathrm{Wt}=(9.953 \mathrm{E}-4) \mathrm{Len}^{2.773}$ (equations taken from Goodyear and Thompson, 1993). Mean weight estimates by year, state, and mode correspond to the mean estimate where the sample size exceeded 25 individuals; where the sample size was less than 25 , the state or gulfwide annual mean was substituted following the same convention. Units are in pounds.

| Year | TX | LA | MS | AL | FLW |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1986 |  |  |  |  |  |
| 1987 |  |  |  |  |  |
| 1988 |  |  |  |  |  |
| 1989 |  |  |  |  |  |
| 1990 |  |  |  | 2.25 | 2.29 |
| 1991 |  |  |  | 2.47 | 2.61 |
| 1992 |  |  |  |  | 1.90 |
| 1993 |  |  |  |  | 2.06 |
| 1994 |  |  |  | 2.38 |  |
| 1995 |  |  |  |  | 1.53 |
| 1996 |  |  |  |  |  |
| 1997 |  |  |  |  | 2.19 |
| 1998 |  |  |  |  |  |


|  |  | CHARTER |  |  |  |
| ---: | ---: | ---: | ---: | ---: | ---: |
| Year | TX | LA | MS | AL | FLW |
| 1986 |  | 2.92 |  | 3.27 | 2.77 |
| 1987 | 1.07 | 2.37 | 2.42 | 2.49 | 2.48 |
| 1988 | 1.09 | 2.49 | 2.49 | 2.63 | 2.38 |
| 1989 | 1.00 |  | 2.00 | 2.49 | 1.32 |
| 1990 | 1.22 | 2.35 | 2.26 | 2.33 | 2.36 |
| 1991 | 1.13 | 2.36 |  | 2.53 | 2.63 |
| 1992 | 1.35 | 1.92 | 2.18 | 2.32 | 2.06 |
| 1993 |  | 2.17 | 2.17 | 2.32 | 2.23 |
| 1994 | 1.67 | 1.76 | 2.18 | 2.42 | 1.95 |
| 1995 |  | 2.15 | 1.84 | 2.00 | 1.54 |
| 1996 | 1.57 | 1.96 | 1.96 | 2.11 | 1.87 |
| 1997 | 1.87 | 2.30 | 2.30 | 2.52 | 2.27 |
| 1998 | 1.77 | 1.98 |  | 1.93 | 1.80 |

HEADBOAT

| Year | TX | LA | MS | AL | FLW |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 1986 | 1.66 | 1.99 |  | 2.22 | 2.54 |
| 1987 | 1.85 | 1.76 |  | 1.89 | 1.41 |
| 1988 | 1.87 | 1.45 |  | 1.70 | 2.12 |
| 1989 | 1.65 | 1.76 |  | 1.72 | 1.25 |
| 1990 | 1.76 | 3.36 |  | 1.51 | 1.45 |
| 1991 | 1.95 | 1.93 |  | 1.62 | 1.68 |
| 1992 | 1.42 | 1.82 |  | 1.62 | 1.27 |
| 1993 | 2.00 | 2.33 |  | 1.70 | 1.91 |
| 1994 | 1.88 | 2.07 |  | 1.53 | 1.84 |
| 1995 | 1.99 | 2.05 |  | 1.61 | 1.41 |
| 1996 | 1.95 | 1.90 |  | 1.62 | 3.01 |
| 1997 | 2.08 | 1.73 |  | 1.60 | 1.19 |
| 1998 | 2.15 | 2.01 |  | 1.51 | 1.35 |


| Year | TX | LA | MS | AL | FLW |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 1986 | 0.96 | 2.85 |  | 3.17 | 2.77 |
| 1987 | 1.08 | 2.34 | 2.42 | 2.46 | 2.33 |
| 1988 | 1.08 | 2.49 |  | 2.63 | 2.36 |
| 1989 | 1.00 | 2.00 |  | 2.36 | 1.28 |
| 1990 | 1.23 | 2.32 | 2.26 | 2.25 | 2.29 |
| 1991 | 1.13 | 2.36 | 2.46 | 2.47 | 2.61 |
| 1992 | 1.38 | 1.99 | 2.18 | 1.96 | 1.73 |
| 1993 | 1.53 | 2.17 | 2.17 | 2.31 | 1.62 |
| 1994 | 1.66 | 1.87 | 2.18 | 2.38 | 2.03 |
| 1995 | 1.73 | 2.15 | 1.84 | 1.90 | 1.53 |
| 1996 | 1.54 | 1.96 | 1.96 | 2.11 | 1.49 |
| 1997 | 1.87 | 2.30 | 2.30 | 2.53 | 1.65 |
| 1998 | 1.77 | 1.98 | 1.88 | 2.09 | 1.82 |

Table 9. Recreational harvest estimates for Gulf of Mexico gray triggerfish by year, state and fishing mode for the period 1986-1998. The estimates are based on the MRFSS, the NMFS Headboat Survey, and the Texas Parks and Wildlife size-frequency samples and catch estimates. The weight estimates are the sums of products of the annual harvest and mean weight estimates for each state by mode. Units are in number of fish and pounds.

|  | TX |  | LA |  | MS |  | AL |  | FLW |  | TOTAL GULF |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | N | Wt(lbs) | N | Wt(lbs) | N | Wt(lbs) | N | Wt(lbs) | N | Wt(lbs) | N | Wt(lbs) |
| $\begin{aligned} & 1986 \\ & 1987 \\ & 1988 \\ & 1989 \end{aligned}$ |  |  |  |  |  |  |  |  |  |  |  |  |
| 1990 1991 |  |  |  |  |  |  | $\begin{gathered} 30765 \\ 5664 \end{gathered}$ | 69110.80 <br> 14005.10 | $\begin{aligned} & 27485 \\ & 41830 \end{aligned}$ | $\begin{gathered} 62822.86 \\ 109051.24 \end{gathered}$ | $\begin{aligned} & 58250 \\ & 47494 \end{aligned}$ | $\begin{aligned} & 131933.66 \\ & 123056.35 \end{aligned}$ |
| 1992 |  |  |  |  |  |  |  |  | 27981 4193 | 53217.52 8640.34 | $\begin{gathered} 27981 \\ 1102 \end{gathered}$ | 53217.52 8640.34 |
| 1994 |  |  |  |  |  |  | 1265 | 3005.03 | 4193 | 8640.34 | 1265 |  |
| 1995 |  |  |  |  |  |  |  |  | 2782 | 4246.00 | 2782 | 4246.00 |
| 1996 |  |  |  |  |  |  |  |  | 1161 | 2541.34 | 1161 | 2541.34 |
| 1998 |  |  |  |  |  |  |  |  |  |  |  |  |


|  |  |  |  | A |  | S |  |  | FLW |  | TOTAL GULF |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | N | Wt(lbs) | N | Wt(lbs) | N | Wt(lbs) | N | Wt(lbs) | N | Wt(lbs) | N | Wt(lbs) |
| 1986 |  |  | 1725 | 5034.83 |  |  | 13958 | 45672.05 | 394155 | 1090586.11 | 409838 | 1141292.98 |
| 1987 | 1388 | 1486.39 | 1803 | 4280.67 | 13 | 31.45 | 10267 | 25524.13 | 463119 | 1148154.63 | 476590 | 1179477.28 |
| 1988 | 203 | 220.51 | 1341 | 3343.96 | 909 | 2266.71 | 85830 | 226103.59 | 320627 | 762023.00 | 408910 | 993957.77 |
| 1989 | 102 | 102.15 |  |  | 4655 | 9332.47 | 129322 | 322245.14 | 247969 | 328516.17 | 382048 | 660195.93 |
| 1990 | 315 | 382.97 | 5093 | 11963.28 | 82 | 185.34 | 319420 | 743223.57 | 278075 | 655178.75 | 602985 | 1410933.92 |
| 1991 | 137 | 154.83 | 56613 | 133506.24 |  |  | 94231 | 237936.32 | 552407 | 1455492.61 | 703388 | 1827090.00 |
| 1992 | 1870 | 2531.50 | 14410 | 27736.13 | 72 | 157.22 | 91477 | 212366.30 | 245723 | 507377.86 | 353552 | 750169.01 |
| 1993 |  |  | 16834 | 36469.64 | 930 | 2014.78 | 95899 | 222753.25 | 269815 | 601475.04 | 383478 | 862712.71 |
| 1994 | 30 | 49.97 | 22272 | 39167.95 | 1360 | 2965.27 | 64069 | 155193.51 | 420498 | 821712.07 | 508229 | 1019088.77 |
| 1995 |  |  | 28497 | 61294.13 | 1148 | 2116.05 | 114976 | 229686.07 | 258845 | 397585.92 | 403466 | 690682.18 |
| 1996 | 26 | 40.80 | 4913 | 9628.90 | 4443 | 8707.75 | 76716 | 162087.58 | 105903 | 197686.62 | 192001 | 378151.65 |
| 1997 | 815 | 1523.79 | 2250 | 5177.36 | 1733 | 3987.72 | 72837 | 183561.00 | 102112 | 231509.58 | 179747 | 425759.45 |
| 1998 | 7902 | 14013.55 | 5148 | 10171.74 |  |  | 58608 | 113154.07 | 123962 | 223276.85 | 195620 | 360616.21 |

HEADBOAT

|  | TX |  | LA |  | MS |  | AL |  | FLW |  | TOTAL GULF |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | N | Wt(lbs) | N | Wt(lbs) | N | Wt(lbs) | N | Wt(lbs) | N | Wt(lbs) | N | Wt(lbs) |
| 1986 | 15642 | 25965.57 | 407 | 809.09 |  |  | 23209 | 51452.66 | 6797 | 17243.84 | 46055 | 95471.16 |
| 1987 | 16085 | 29728.32 | 612 | 1076.82 |  |  | 16602 | 31299.06 | 7206 | 10155.44 | 40505 | 72259.63 |
| 1988 | 39569 | 74052.96 | 1927 | 2802.05 |  |  | 22609 | 38412.74 | 5846 | 12375.41 | 69951 | 127643.16 |
| 1989 | 23589 | 38896.87 | 1355 | 2383.57 |  |  | 39033 | 67303.66 | 18820 | 23592.46 | 82797 | 132176.56 |
| 1990 | 21762 | 38181.08 | 3915 | 13163.86 |  |  | 93659 | 141552.51 | 14043 | 20407.97 | 133379 | 213305.42 |
| 1991 | 24100 | 46936.58 | 7028 | 13599.08 |  |  | 53014 | 85968.51 | 6038 | 10150.66 | 90180 | 156654.83 |
| 1992 | 35890 | 50928.09 | 5862 | 10677.64 |  |  | 62408 | 101187.20 | 7965 | 10147.91 | 112125 | 172940.84 |
| 1993 | 38226 | 76559.80 | 5958 | 13863.58 |  |  | 53022 | 90198.04 | 6823 | 13065.41 | 104029 | 193686.82 |
| 1994 | 50034 | 94116.14 | 6678 | 13793.39 |  |  | 49259 | 75291.49 | 5624 | 10370.61 | 111595 | 193571.63 |
| 1995 | 47925 | 95567.33 | 3916 | 8035.06 |  |  | 42187 | 67969.12 | 4493 | 6326.81 | 98521 | 177898.32 |
| 1996 | 37501 | 73181.71 | 2828 | 5369.38 |  |  | 33016 | 53588.75 | 4400 | 13239.48 | 77745 | 145379.33 |
| 1997 | 28731 | 59740.14 | 496 | 858.23 |  |  | 27295 | 43583.15 | 8227 | 9814.52 | 64749 | 113996.04 |
| 1998 | 15222 | 32756.41 | 881 | 1767.42 |  |  | 29324 | 44217.83 | 8357 | 11295.01 | 53784 | 90036.68 |

PRIVATE/RENTAL

|  | TX |  | LA |  | MS |  | AL |  | FLW |  | TOTAL GULF |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | N | Wt(lbs) | N | Wt(lbs) | N | Wt(lbs) | N | Wt(lbs) | N | Wt(lbs) | N | Wt(lbs) |
| 1986 | 4394 | 4204.94 | 8643 | 24599.49 |  |  | 2222 | 7045.59 | 34769 | 96202.23 | 50028 | 132052.25 |
| 1987 | 5134 | 5522.10 | 2029 | 4742.69 | 1429 | 3457.07 | 4224 | 10376.61 | 144248 | 335685.70 | 157064 | 359784.17 |
| 1988 | 13797 | 14908.93 | 7449 | 18575.08 |  |  | 941 | 2478.89 | 272253 | 642455.96 | 294440 | 678418.86 |
| 1989 | 32589 | 32628.79 | 49453 | 99144.66 |  |  | 38941 | 91771.82 | 395901 | 505027.56 | 516884 | 728572.83 |
| 1990 | 8763 | 10766.60 | 89754 | 208453.67 | 9291 | 20999.94 | 75263 | 169071.55 | 110495 | 252560.00 | 293566 | 661851.75 |
| 1991 | 8793 | 9951.52 | 1055 | 2486.68 | 1399 | 3447.63 | 10177 | 25164.19 | 47553 | 123971.16 | 68977 | 165021.17 |
| 1992 | 70559 | 97241.46 | 13435 | 26772.10 | 3607 | 7876.10 | 57701 | 113162.18 | 209148 | 362556.31 | 354450 | 607608.15 |
| 1993 | 39204 | 60086.21 | 1619 | 3507.45 | 983 | 2129.60 | 52531 | 121338.75 | 110030 | 178259.22 | 204367 | 365321.22 |
| 1994 | 6272 | 10410.83 | 18788 | 35134.28 | 3022 | 6589.00 | 24761 | 58820.25 | 50259 | 101877.64 | 103102 | 212831.99 |
| 1995 | 4439 | 7662.17 | 38499 | 82807.41 | 7968 | 14687.04 | 73409 | 139527.67 | 15504 | 23662.81 | 139819 | 268347.11 |
| 1996 | 2291 | 3525.15 | 2068 | 4053.03 | 1876 | 3676.74 | 32087 | 67794.26 | 52559 | 78261.38 | 90881 | 157310.55 |
| 1997 | 4150 | 7759.19 | 13233 | 30449.77 | 1629 | 3748.41 | 18315 | 46347.05 | 61472 | 101184.71 | 98799 | 189489.12 |
| 1998 | 2950 | 5231.58 | 2961 | 5850.53 | 8505 | 15948.70 | 16192 | 33835.05 | 92527 | 168117.54 | 123135 | 228983.40 |

Table 9. Recreational harvest estimates for Gulf of Mexico gray triggerfish by year, state and fishing mode for the period 1986-1998 (continued).

|  | TX |  | LA |  | MS |  | AL |  | FLW |  | TOTAL GULF |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | N | Wt(lbs) | N | Wt(lbs) | N | Wt(lbs) | N | Wt(lbs) | N | Wt(lbs) | N | Wt(lbs) |
| 1986 | 20036 | 30170.5114 | 10775 | 30443.40604 | 0 | 0 | 39389 | 104170.3009 | 435721 | 1204032.177 | 505921 | 1368816.396 |
| 1987 | 22607 | 36736.814 | 4444 | 10100.18244 | 1442 | 3488.5152 | 31093 | 67199.80025 | 614573 | 1493995.77 | 674159 | 1611521.081 |
| 1988 | 53569 | 89182.4009 | 10717 | 24721.08781 | 909 | 2266.713 | 109380 | 266995.2268 | 598726 | 1416854.366 | 773301 | 1800019.795 |
| 1989 | 56280 | 71627.8132 | 50808 | 101528.2296 | 4655 | 9332.4654 | 207296 | 481320.6165 | 662690 | 857136.1922 | 981729 | 1520945.317 |
| 1990 | 30840 | 49330.6577 | 98762 | 233580.8051 | 9373 | 21185.276 | 519107 | 1122958.427 | 430098 | 990969.5815 | 1088180 | 2418024.748 |
| 1991 | 33030 | 57042.9308 | 64696 | 149591.9984 | 1399 | 3447.6277 | 163086 | 363074.1161 | 647828 | 1698665.68 | 910039 | 2271822.353 |
| 1992 | 108319 | 150701.048 | 33707 | 65185.87294 | 3679 | 8033.3129 | 211586 | 426715.6829 | 490817 | 933299.6027 | 848108 | 1583935.519 |
| 1993 | 77430 | 136646 | 24411 | 53840.66415 | 1913 | 4144.3755 | 201452 | 434290.0353 | 390861 | 801440.0174 | 696067 | 1430361.093 |
| 1994 | 56336 | 104576.936 | 47738 | 88095.62378 | 4382 | 9554.2666 | 139354 | 292310.2838 | 476381 | 933960.3166 | 724191 | 1428497.427 |
| 1995 | 52364 | 103229.503 | 70912 | 152136.602 | 9116 | 16803.094 | 230572 | 437182.8555 | 281624 | 431821.5418 | 644588 | 1141173.596 |
| 1996 | 39818 | 76747.6482 | 9809 | 19051.31279 | 6319 | 12384.488 | 141819 | 283470.5936 | 162862 | 289187.4803 | 360627 | 680841.523 |
| 1997 | 33696 | 69023.1289 | 15979 | 36485.3493 | 3362 | 7736.1226 | 118447 | 273491.1992 | 172972 | 345050.1509 | 344456 | 731785.951 |
| 1998 | 26074 | 52001.5468 | 8990 | 17789.68684 | 8505 | 15948.703 | 104124 | 191206.9544 | 224846 | 402689.3977 | 372539 | 679636.29 |

Figure 12. Recreational harvest in numbers of fish for Gulf of Mexico gray triggerfish by state and fishing mode for the period 1986-1998.



Table 10. Commercial harvest estimates in weight for Gulf of Mexico gray triggerfish by year and state for the period 1986-1998. The estimates based on the SEFSC General Canvass Program.

| Year | TX | LA | MS | AL | FLW | TOTAL_GULF |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1986 | 572 | 14493 | 4008 | 5881 | 70978 | 95932 |
| 1987 | 289 | 21941 | 5550 | 3778 | 92742 | 124300 |
| 1988 | 1885 | 36980 | 8242 | 7641 | 140790 | 195538 |
| 1989 | 429 | 60856 | 7682 | 10389 | 238974 | 318330 |
| 1990 | 6951 | 69798 | 9027 | 16613 | 359553 | 461942 |
| 1991 | 6242 | 90572 | 7991 | 6993 | 332674 | 444472 |
| 1992 | 7941 | 101495 | 12433 | 6551 | 321883 | 450303 |
| 1993 | 11287 | 128947 | 38273 | 10413 | 374260 | 563180 |
| 1994 | 15428 | 119758 | 15382 | 8389 | 247156 | 406113 |
| 1995 | 26168 | 75744 | 22681 | 5268 | 208449 | 338310 |
| 1996 | 17226 | 79331 | 12644 | 2867 | 152502 | 264570 |
| 1997 | 15022 | 50583 | 8813 | 2534 | 109682 | 186634 |
| 1998 | 20944 | 34378 | 10120 | 1288 | 107651 | 174381 |
| 1999 | 12452 | 50030 | 5613 | 1709 | 118248 | 188052 |

Figure 13. Estimated Gulf of Mexico gray triggerfish annual landings by weight for the commercial sector for the period 1986-1998.


Table 11. Estimated Gulf of Mexico gray triggerfish annual landings by weight for the commercial and recreational sectors for the period 1986-1998.

| Year | Commercial | Recreational | Total Gulf |
| :---: | :---: | :---: | :---: |
| 1986 | 95932 | 1368816 | 1464748 |
| 1987 | 124300 | 1611521 | 1735821 |
| 1988 | 195538 | 1800020 | 1995558 |
| 1989 | 318330 | 1520945 | 1839275 |
| 1990 | 461942 | 2418025 | 2879967 |
| 1991 | 444472 | 2271822 | 2716294 |
| 1992 | 450303 | 1583936 | 2034239 |
| 1993 | 563180 | 1430361 | 1993541 |
| 1994 | 406113 | 1428497 | 1834610 |
| 1995 | 338310 | 1141174 | 1479484 |
| 1996 | 270593 | 680842 | 951435 |
| 1997 | 186634 | 731786 | 918420 |
| 1998 | 174381 | 679636 | 854017 |

Figure 14. Estimated Gulf of Mexico gray triggerfish annual landings by weight for the commercial and recreational sectors for the period 1986-1998.


Table 12. Annual yield and CPUE index data used to fit ASPIC production model.

|  |  | RECREATIONAL_INDICES |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| MRFSS | CPUE(num/angler hour) |  | HEADBOAT |  |  |
| Year |  | Yield (wt) | Year | CPUE(num/angler hour) | Yield (wt) |
| 1986 | 0.39 | 1273345 | 1986 | 0.31 | 95471 |
| 1987 | 0.17 | 1539261 | 1987 | 0.31 | 72260 |
| 1988 | 0.51 | 1672377 | 1988 | 0.44 | 127643 |
| 1989 | 0.82 | 1388769 | 1989 | 0.64 | 132177 |
| 1990 | 1.00 | 2204719 | 1990 | 1.00 | 213305 |
| 1991 | 0.80 | 2115168 | 1991 | 0.88 | 156655 |
| 1992 | 0.69 | 1410995 | 1992 | 0.90 | 172941 |
| 1993 | 0.48 | 1236674 | 1993 | 0.87 | 193687 |
| 1994 | 0.48 | 1234926 | 1994 | 0.72 | 193572 |
| 1995 | 0.37 | 963275 | 1995 | 0.51 | 177898 |
| 1996 | 0.32 | 535462 | 1996 | 0.55 | 145379 |
| 1997 | 0.31 | 617790 | 1997 | 0.41 | 113996 |
| 1998 | 0.33 | 589600 | 1998 | 0.38 | 90037 |
|  | COMMERCIAL INDEX |  |  |  |  |
| LOGBOOKHEADBOAT Year | Std CPUE (pounds/hour) | Yield (wt) |  |  |  |
| 1986 | -- | 95932 |  |  |  |
| 1987 | -- | 124300 |  |  |  |
| 1988 | -- | 195538 |  |  |  |
| 1989 | -- | 318330 |  |  |  |
| 1990 | -- | 461942 |  |  |  |
| 1991 | -- | 444472 |  |  |  |
| 1992 | -- | 450303 |  |  |  |
| 1993 | 0.80 | 563180 |  |  |  |
| 1994 | 1.00 | 406113 |  |  |  |
| 1995 | 0.93 | 338310 |  |  |  |
| 1996 | 0.71 | 270593 |  |  |  |
| 1997 | 0.66 | 186634 |  |  |  |
| 1998 | 0.61 | 174381 |  |  |  |

Figure 15. Gray triggerfish data from the Gulf of Mexico used to fit production model (ASPIC). (A) Total yield. (B) Overlay of all standardized recreational and commercial CPUEs constructed in this study. (C) Standardized CPUE trajectories selected for use in ASPIC. Recreational CPUE units are in number of fish per angler hour, commercial CPUE units are in pounds per hour.



Figure 16. Results of ASPIC production model analyses. (A) Fixing the intrinsic rate of increase from $r=$ 0.5 to 2.0. (B) Fixing the starting biomass ratio from $B_{1} / B_{M S Y}=0.5$ to 2.0. (C) Fixing $B_{1} / B_{M S Y}=0.5$ to 1.0 and $r=0.1$ to 2.0 . (D) Fixing $B_{l} / B_{M S Y}=0.5$ to 1.5 and $\mathrm{MSY}=1.5 \mathrm{E}+06$ to $3.5 \mathrm{E}+06$. Only parameter combinations that allowed the model to converge are shown.

## (A) Fixed $r$





Figure 16. (Continued).
(B) Fixed B1/B мsү




## (C) Fixed B1/B msy and $r$




Figure 16. (Continued).

## (D) Fixed B1/B msy $_{\text {and }}$ MSY




Table 13 . Results of two bootstapped production model analyses for gray triggerfish in the Gulf of Mexico. In Model (2), constraints were placed on parameters and $r$ was held constant at $r=1.0$. Resulting parameters from Model (2) were used as initial guess values for Model (1), and no parameters were fixed. The two models thus differ in initial guess values and the number of parameters estimated by ASPIC. Each analysis included a bootstrap with 1,000 trials. Nonparametric bias-corrected $80 \%$ confidence intervals are derived from the bootstrap within ASPIC.

| Parameter | Bias-corrected estimate | MODEL 1 <br> Ordinary estimate | (r=1.0) | $\begin{gathered} 80 \% \\ \text { Upper CL } \\ \hline \end{gathered}$ | Relative IQ range |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | $\begin{gathered} 80 \% \\ \text { Lower CL } \end{gathered}$ |  |  |
| Model Parameters |  |  |  |  |  |
| B1ratio | 0.672 | 0.697 | 0.531 | 1.070 | 30.2\% |
| K | $1.09 \mathrm{E}+07$ | $1.06 \mathrm{E}+07$ | $8.14 \mathrm{E}+06$ | $1.30 \mathrm{E}+07$ | 20.4\% |
| r | 1 | 1 | 1 | 1 | 0.0\% |
| q(1) | 2.97E-07 | 2.92E-07 | $2.70 \mathrm{E}-07$ | $3.20 \mathrm{E}-07$ | 8.7\% |
| $\mathrm{q}(2)$ | $4.00 \mathrm{E}-07$ | 3.93E-07 | $3.63 \mathrm{E}-07$ | $4.28 \mathrm{E}-07$ | 8.5\% |
| q(3) | 6.68E-07 | $6.57 \mathrm{E}-07$ | $6.01 \mathrm{E}-07$ | 7.33E-07 | 10.1\% |
| Management Benchmarks |  |  |  |  |  |
| MSY | $2.72 \mathrm{E}+06$ | $2.65 \mathrm{E}+06$ | $2.03 \mathrm{E}+06$ | $3.24 \mathrm{E}+06$ | 20.4\% |
| Bmsy | $5.43 \mathrm{E}+06$ | $5.31 \mathrm{E}+06$ | $4.07 \mathrm{E}+06$ | $6.48 \mathrm{E}+06$ | 20.4\% |
| Fmsy | 0.5 | 0.5 | 0.5 | 0.5 | 0.0\% |
| fmsy(1) | $1.68 \mathrm{E}+06$ | $1.71 \mathrm{E}+06$ | $1.56 \mathrm{E}+06$ | $1.85 \mathrm{E}+06$ | 8.8\% |
| fmsy(2) | $1.25 \mathrm{E}+06$ | $1.27 \mathrm{E}+06$ | $1.17 \mathrm{E}+06$ | $1.38 \mathrm{E}+06$ | 8.7\% |
| fmsy(3) | $7.48 \mathrm{E}+05$ | 7.61E+05 | $6.82 \mathrm{E}+05$ | $8.32 \mathrm{E}+05$ | 10.2\% |
| $\mathrm{B}_{1999} / \mathrm{B}_{\text {MsY }}$ | 0.196 | 0.208 | 0.152 | 0.294 | 33.3\% |
| $\mathrm{F}_{1998} / \mathrm{Fmsy}$ | 1.651 | 1.620 | 1.437 | 1.879 | 13.7\% |

Table 13. (Continued.).

| Parameter | MODEL 2 | (No fixed parameters) |  | $\begin{gathered} 80 \% \\ \text { Upper CL } \\ \hline \end{gathered}$ | Relative IQ range |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Bias-corrected estimate | Ordinary estimate | $\begin{gathered} 80 \% \\ \text { Lower CL } \end{gathered}$ |  |  |
| Model Parameters |  |  |  |  |  |
| B1ratio | 0.525 | 0.532 | 0.506 | 0.554 | 3.8\% |
| K | $1.06 \mathrm{E}+07$ | $1.06 \mathrm{E}+07$ | $1.01 \mathrm{E}+07$ | 1.23E+07 | 3.5\% |
| r | 1.287 | 1.277 | 1.224 | 1.346 | 3.9\% |
| q(1) | 3.83E-07 | $3.75 \mathrm{E}-07$ | $3.41 \mathrm{E}-07$ | $4.06 \mathrm{E}-07$ | 7.0\% |
| $\mathrm{q}(2)$ | 5.18E-07 | 5.04E-07 | $4.66 \mathrm{E}-07$ | 5.55E-07 | 8.3\% |
| $\mathrm{q}(3)$ | 8.63E-07 | 8.43E-07 | 7.48E-07 | $9.48 \mathrm{E}-07$ | 10.7\% |
| Management Benchmarks |  |  |  |  |  |
| MSY | $3.41 \mathrm{E}+06$ | 3.37E+06 | $3.28 \mathrm{E}+06$ | $3.48 \mathrm{E}+06$ | 3.0\% |
| Bmsy | $5.29 \mathrm{E}+06$ | 5.27E+06 | $5.05 \mathrm{E}+06$ | $6.14 \mathrm{E}+06$ | 3.5\% |
| Fmsy | 0.644 | 0.638 | 0.612 | 0.673 | 3.9\% |
| fmsy(1) | $1.69 \mathrm{E}+06$ | 1.70E+06 | $1.60 \mathrm{E}+06$ | $1.82 \mathrm{E}+06$ | 5.7\% |
| fmsy(2) | $1.26 \mathrm{E}+06$ | $1.27 \mathrm{E}+06$ | $1.18 \mathrm{E}+06$ | $1.36 \mathrm{E}+06$ | 7.1\% |
| fmsy(3) | $7.56 \mathrm{E}+05$ | 7.57E+05 | $6.91 \mathrm{E}+05$ | $8.35 \mathrm{E}+05$ | 9.3\% |
| $\mathrm{B}_{1999} / \mathrm{B}_{\text {MSY }}$ | 0.155 | 0.164 | 0.128 | 0.192 | 21.7\% |
| $\mathrm{F}_{1998} / \mathrm{Fmsy}$ | 1.666 | 1.648 | 1.481 | 1.898 | 12.9\% |







Figure 17. Estimated population trajectories (non-bootstrapped) of two production model analyses (ASPIC) of gray triggerfish in the Gulf of Mexico. In Model (1) $r=1.0$. The dashed line and squares represent the results from Model (1); the solid line and triangles represent the results from Model (2).


Figure 18. Goodness-of-fit of two production model analyses (ASPIC) by data series (user group). In Model (1) $r$ was held fixed at $r=1.0$. In Model (2) all parameters were estimated by ASPIC. The solid line represents the observed values and dashed lines represent the estimated values for each model (1 and 2).







Figure 19. Estimated trajectories of relative biomass and relative fishing mortality, and control rule plots from two production model (ASPIC) runs. Panels (A), (B), and (C) correspond to Model 1 (constant $r=1.0$ ) and panels (D), (E), and (F) to Model 2 (no fixed parameters).

## NO FISHING

Model 1






Figure 20. Projected trajectories of $\mathrm{B} / \mathrm{B}_{\mathrm{MSY}}, \mathrm{F} / \mathrm{F}_{\mathrm{MSY}}$, and control rule plots under a no fishing scenario. Panels (A), (B), and (C) correspond to ASPIC Model 1 (constant $r=1.0$ ) and panels (D), (E), and (F) to Model 2 (no fixed parameters).

## Model 1



Model 2




Figure 21. Projected trajectories of $\mathrm{B} / \mathrm{B}_{\mathrm{MSY}}, \mathrm{F} / \mathrm{F}_{\mathrm{MSY}}$, and control rule plots assuming a constant (current) catch scenario. The 1998 catch value ( 854,000 pounds) is repeated for ten years. Panels (A), (B), and (C) correspond to ASPIC Model 1 and panels (D), (E), and (F) to Model 2.

## 1998 Fishing Mortality Rate

Model 1





## Model 2






Figure 22. Projected trajectories of $\mathrm{B} / \mathrm{B}_{\mathrm{MSY}}, \mathrm{F} / \mathrm{F}_{\mathrm{MSY}}$, yield, and control rule plots assuming status quo fishing mortality. The 1998 F value ( $\mathrm{F}=0.81$ in Model $1, \mathrm{~F}=1.052$ in Model 2) is repeated for ten years. Panels (A), (B), (C), and (D) correspond to ASPIC Model 1and panels (E), (F), (G), and (H) to Model 2.

