## Statistical Analysis for Genetic Epidemiology (S.A.G.E.) Version 5.1.1 User Reference Manual

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## **Chapter 1**

# Introduction

*Statistical Analysis for Genetic Epidemiology* (S.A.G.E.) is a collection of compiled C++ programs that perform a wide variety of genetic analyses. The range of functionality includes tools for

- extracting summary statistics describing the data and evaluating general data quality,
- estimating allele frequencies,
- estimating heritability and familial correlations,
- inferring mixture models for genetic transmission, and penetrance functions, including variable age of onset,
- estimating identity-by-descent (IBD) allele sharing probabilities between relative pairs,
- performing model-based linkage analysis,
- performing model-free linkage analysis,
- performing transmission/disequilibrium (TDT) analysis, and
- analyzing trait/allele associations.

S.A.G.E. runs on a variety of platforms, including Linux, Windows, Solaris, and Tru64. The programs may be run from a command line from a cross-platform graphical user interface (GUI) that is included as part of the complete package. The software is extremely flexible with respect to the structure of input pedigree data files and, unless otherwise stated, the dependent phenotypes and traits may be discrete (including dichotomous data) or continuous.

Please check our web page for the most up-to-date information on the S.A.G.E. 5.x programs at the following URL:

http://darwin.case.edu/

### **1.1 Program Descriptions**

#### **1.1.1 Summary Statistics**

#### PEDINFO

**PED***igree INFOrmation and statistics*: Provides many useful descriptive statistics on pedigree data including means, variances and histograms of family, sibship and pedigree sizes, and counts of each type of relative pair.

#### **1.1.2 Data Quality**

#### MARKERINFO

**MARKER INFO**rmation: Detects Mendelian inconsistencies of markers in pedigree data.

#### RELTEST

**RELationship TESTing:** Indicates pairs of relatives to be reclassified according to their true relationship using multi-point genome scan data. The method is based on a Markov process model of identity-by-descent (IBD) allele-sharing along chromosomes. This program currently analyzes four different types of putative pairs: full sib pairs, half sib pairs, parent offspring pairs and unrelated marital pairs. A summary file is produced that contains the pairs to be reclassified together with their Mean Allele-Sharing Statistic, Parent Offspring Statistic and, for each individual, the percentage of marker data that is missing.

#### **1.1.3** Allele Frequency Estimation

#### FREQ

Allele **FREQ**uency estimator: Estimates allele frequencies from marker data on related individuals with known pedigree structure and generates marker locus description files, needed by GENIBD, MLOD, and other S.A.G.E. programs. Future versions will also have the ability to estimate genotype and haplotype frequencies in the presence of allelic disequilibrium.

#### **1.1.4 Familial Aggregation**

#### ASSOC

*Marker-Trait* **ASSOC***iations in Pedigree Data*: Simultaneously analyzes from pedigree data the association between a trait and covariates, which can include marker phenotypes that have been transformed into quantitative covariates, and residual familial correlations/heritability.

#### FCOR

*Family CORrelations*: Calculates multivariate familial correlations with their asymptotic standard errors. Calculates familial correlations for all pair types available in the pedigrees without assuming multivariate normality of the traits across family members.

#### **1.1.5 Commingling Analysis**

#### SEGREG

**SEGREG***ation models*: This program can be used to fit mixtures of two or three normal distributions, simultaneously applying a power transformation to the data and also allowing for both ascertainment and residual familial correlations.

#### 1.1.6 Segregation Analysis

#### SEGREG

**SEGREG***ation models*: Fits and tests Mendelian segregation models in the presence of residual familial correlations. The trait analyzed can be continuous, binary, or a binary disease trait with variable age of onset. This program can also be used for commingling analysis, to predict the major genotype of any pedigree member, and to prepare penetrance files for model-based linkage analysis.

#### AGEON

*AGE of ONset*: Produces maximum likelihood estimates of the parameters of a mixed power-normal distribution for a binary trait with variable age of onset. The mean, variance and susceptibility parameters can be specified as dependent on covariates.

#### 1.1.7 IBD Allele Sharing Analysis

#### GENIBD

*GENerate IBD sharing probabilities*: Generates both single- and multi-point identityby-descent (IBD) distributions using a variety of algorithms tuned for different types of relative pairs in pedigrees. Exact methods can be used for small pedigrees with loops, and simulation methods are available for large extended pedigrees with loops. In the case of small pedigrees IBD sharing can also be interpolated between markers.

#### 1.1.8 Model-Based Linkage Analysis

#### LODLINK

*Single-point model-based* **LOD** *score* **LINK***age analysis*: LOD scores and recombination fractions are obtained between a marker and trait that follows any Mendelian model allowed by SEGREG (which can be used to generate the appropriate penetrance files). Test of linkage heterogeneity, and of linkage in the presence of linkage heterogeneity, are included.

#### MLOD

**M***ulti-point model-based* **LOD** *score analysis*: Performs multi-point model-based LODscore linkage analysis on small pedigrees. Analysis is greatly optimized for examining multiple one-locus trait models and will, in future versions, allow for meiosis specific (e.g., age and sex specific) recombination fractions.

#### 1.1.9 Model-Free Linkage Analysis

#### LODPAL

*LOD score Pair AnaLysis*: Performs analysis based on the LOD score formulation for affected-sib-pairs (ASP). The current implementation is of the general conditional logistic model, including the one-parameter model that allows for the inclusion of all affected-relative-pairs, covariates and epistatic interactions.

#### SIBPAL

*SIBling Pair AnaLysis*: Performs mean tests, proportion tests and linear regressionbased modeling of squared sib-pair differences and mean-corrected sums of a trait as a function of marker allele identity-by-descent sharing. Available analyses can use either single- or multi-point IBD information, and models allow for both binary and continuous traits due to multiple genetic loci, including epistatic interactions and covariate effects.

#### 1.1.10 Transmission Disequilibrium

#### TDTEX

*Transmission Disequilibrium Test (EXact)*: This program implements several asymptotic and exact versions of the transmission disequilibrium test (TDT) for testing linkage between marker and disease loci in the presence of allelic association. The exact tests are useful in cases where little data are available or there are many alleles at the marker locus. Different types of tests are available, including an exact test and a Markov chain Monte Carlo randomization test, as well as several exact marginal homogeneity tests.

#### 1.1.11 Allelic Association

#### ASSOC

*Marker-Trait* **ASSOC***iations in Pedigree Data*: Analyzes the association between a continuous trait and covariates, which can include marker phenotypes that have been transformed into quantitative covariates, from pedigree data in the presence of familial correlations. Together with the *Transmitted Allele Indicator* (available as a user-defined function), performs a pedigree transmission disequilibrium test (TDT).

#### 1.1.12 Haplotype Analysis

#### DECIPHER

Obtains maximum likelihood estimates of population haplotype frequencies for autosomal or X-linked markers, and determines all possible diplotypes and the most likely diplotypes for each individual. Estimates haplotype frequencies for different populations as specified by the user. Performs likelihood ratio tests and permutation tests to compare haplotype frequency distributions for dichotomous phenotypes.

## **1.2 Program Limitations**

All programs will cease to function when the user's current license expires.

All programs currently make the following assumptions in all of their analysis methods:

- each genetic marker has a known genotype-phenotype relation (which may be either deterministic or probabilistic),
- the founders of each constituent pedigree<sup>1</sup> are not inbred and are unrelated to one another, and in particular, the pedigrees do not comprise loops,
- the members of each constituent pedigree are unrelated to the members of any other constituent pedigree, and
- there are no selection, migration or mutation effects $^2$ .

<sup>&</sup>lt;sup>1</sup>The user defines pedigrees by giving each pedigree a unique identification number. A subset of pedigree members for whom there is no information on how they are related to other members of that pedigree is called a *constituent pedigree* and is treated as an independent pedigree in all analyses. With the exception of the program PEDINFO, the term *pedigree* will always refer to a *constituent pedigree* as defined here.

<sup>&</sup>lt;sup>2</sup>Several programs allow for non-Hardy-Weinberg equilibrium proportions and the SEGREG program allows for general transmission models.

## **1.3** Conventions Used in this Manual

This document uses the following typographical conventions to help clarify the correct specification of S.A.G.E. program commands and options:

- 1. All references to parameters and attributes are printed using a non-proportional font.
- 2. All references to *named constant* values (e.g., **true** and **false**) outside of a syntax table (see 2.2.4) are printed using a **bold** font.
- 3. Examples of parameter files are printed using a non-proportional font. Specific values within a given example are printed using a **bold non-proportional** font.
- 4. Examples of program outputs are printed using a non-proportional font.
- 5. Technical terms that have not been previously introduced in the manual are printed using an *italics* font. The term's definition will be explicitly given if its meaning is not evident from the context.
- 6. Text that needs to be otherwise EMPHASIZED is printed using an UPPER CASE font.

## **Chapter 2**

# **Program Input and Output**

Each S.A.G.E. program requires several input files in order to run. No program requires all of the possible input files. Refer to the individual program documentation for specific information on which files are required. The file types currently used for program input are:

Section	File Type	Description
2.2	Parameter file	Specifies the parameters and options with which
		to perform a particular analysis.
2.3.2	Pedigree data file	Contains delimited records for each individual,
		including fields for identifiers, sex, parents, trait
		and marker data.
2.5	Marker locus description file	Lists the alleles, allele frequencies and phenotype
		to genotype mapping for each marker locus.
2.5	Trait locus description file	Lists the genetic model for each of the traits to
		be analyzed for linkage using a specific genetic
		model
2.6	Genome description file	Contains a description of the linked marker re-
		gions, including distances between markers. This
		file is not required for single-point <sup><i>a</i></sup> analysis.
2.7	IBD sharing file	Stores identity-by-descent (IBD) distributions be-
		tween pairs of related individuals at one or more
		marker loci.
	LODPAL pairs file	Stores the pre-constructed pair-specific covariate
		and/or weight values to be used in the analysis

<sup>*a*</sup>Single-point in the sense that information is used from only one observed marker locus at a time. When performing linkage analysis this is often called "two-point" analysis.

Each program also produces one or more output files that contain results and diagnostic information. Refer to the individual program documentation for specific information on which files are produced and details of what they contain.

Section	File Type	Description			
2.8	Information output file	Contains informational diagnostic messages, warnings			
		and program errors. Each program generates one informa-			
		tion output file. No analysis results are stored in this file.			
		Information files are automatically named with a "INF"			
		extension (for example, "segreg_analysis.inf").			
2.9	Analysis output file	Each program may generate one or more analysis output			
		files. These files contain the results of each analysis or			
		may summarize the results of many analyses.			

The file types currently used for program output are:

## 2.1 Running S.A.G.E. Programs

S.A.G.E. programs may be executed from the provided Graphical User Interface (GUI) or, alternatively, by means of a command line directive that specifies the name of a selected program followed by a list of *arguments* to the program. The order of the arguments is important. If the user enters an incorrect or incomplete command line, then the software will display a usage statement that indicates the correct syntax for the desired program<sup>1</sup>. For example, if the PEDINFO program name were entered without any arguments, the result would be as follows:

#### >pedinfo

```
S.A.G.E. v4.x -- PEDINFO
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
usage: pedinfo <parameters> <pedigree>
Command line parameters:
parameters - parameter file
pedigree - pedigree data file
```

As indicated in this program usage statement, two input files need to be listed on the command line. A typical run of PEDINFO may look like the following, if run from the "example/output/pedinfo" directory:

```
>cd example/output/pedinfo
>pedinfo ../input/parameters ../input/pedigree
S.A.G.E. v4.x -- PEDINFO
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
Reading parameter file.....done.
Reading pedigree file.....
from ../../input/pedigree1..done.
```

<sup>&</sup>lt;sup>1</sup>The GUI is designed to generate syntactically correct lists of program arguments based on the user selections in the various screens and dialogs. The argument lists are in turn forwarded to the desired S.A.G.E. program(s) for processing. Although the GUI does perform validation of user selections before submitting them for processing, it is nevertheless possible that the user can generate an invalid S.A.G.E. command through the GUI. The user is therefore advised to check the S.A.G.E. outputs to the console as well as the information file (\*.INF) when the program does not seem to work as expected.

Sorting pedigrees.....done. Generating statistics.....done. Analysis complete!

### 2.2 The Parameter File

User options for analysis are specified to S.A.G.E. programs as a list of instructions within a *parameter file*<sup>2</sup>. When a particular S.A.G.E. program is executed it evaluates the contents of the specified parameter file to determine

- 1. how to interpret the contents of the given pedigree data file,
- 2. how many different analyses have been requested and
- 3. which options have been specified for each analysis.

A parameter file is simply a text file containing a list of S.A.G.E. program instructions written according to a specific syntax (see below). A single parameter file may be used to specify options for one or more S.A.G.E. programs in any combination. In other words, one parameter file could specify analysis options for several different S.A.G.E. programs, or different options for repeated calls to the same program, or both. And, of course, the user always has the option of creating a set of different parameter files if that makes it easier to manage a given project <sup>3</sup>. Since the parameter file also contains user-supplied specifications on how to interpret the pedigree data file, the ability to specify an arbitrary set of S.A.G.E. analyses within a single parameter file makes the software very flexible.

#### 2.2.1 Creating a Parameter File

One of the primary functions of the GUI is to create parameter files for S.A.G.E. programs. The GUI is designed to translate the user's selection on the various screens and dialogs into syntactically correct lists of program arguments, which are automatically passed into the appropriate S.A.G.E. program. This feature is intended to reduce the complexity associated with learning the syntax of S.A.G.E. parameter files, and is expected to be particularly beneficial to novice users of the software. Experienced users who prefer to create and edit their parameter files directly will continue to have the option of doing so.

A parameter file may be created and modified using a standard text editor on the local S.A.G.E. platform (i.e., the system on which S.A.G.E. has been installed), or the file may be produced on a different system and copied to the local S.A.G.E. platform. The user will normally want to copy the parameter file into the same directory that contains the pedigree data file for a given project, although this is not required.

<sup>&</sup>lt;sup>2</sup>The reader is cautioned that the word *parameter* will have two meanings in this document. In one context it will refer to the set of defined S.A.G.E. *keywords*, but in a statistical context it will refer to some distribution characteristic (e.g., the mean ( $\mu$ ) or variance ( $\sigma^2$ ) of a normal distribution). One goal of the typographical conventions (see 1.3) is to make the context of this word as clear as possible.

<sup>&</sup>lt;sup>3</sup>S.A.G.E. programs accept only one parameter file at a time (the one named as a program argument), regardless of the number available.

In computing environments that include both Unix workstations and Windows PCs, many individuals find the text editors available in Windows to be more user-friendly and convenient, and therefore would prefer to edit their parameter files with either Notepad or WordPad. Users who take this approach must remember to remove the spurious *carriage return* character (^M) which appears at the end of each line of the text file after it has been copied to the Unix target directory<sup>4</sup>.

<sup>&</sup>lt;sup>4</sup>Utility programs available under Unix, such as *dos2unix* make this task fairly easy.

#### 2.2.2 Parameter File Syntax and Structure

A parameter file consists of a list of S.A.G.E. program instructions known as *statements*. When a particular parameter file is passed to a S.A.G.E. program (as a command line argument), the specified program reads each line in the parameter file, from top to bottom, and configures itself to perform the analyses indicated by the listed statements.

All S.A.G.E. statements are formed according to the following format:

in which the square brackets ([]) indicate groupings of optional terms and are not to be entered by the user. The asterisk (\*) indicates that the preceding group or item may be repeated zero or more times. Note that the brackets ([]) and asterisk (\*) are artifacts of the above format definition, and are not to be entered by the user.

In words, the above format definition says:

"A *statement* is a parameter followed by an optional equal-sign-and-value pair<sup>5</sup>, followed by zero or more optional comma-and-attribute pairs (in which each attribute may be followed by an optional equal-sign-and-value pair). This totality, in turn, is optionally followed by a brace-enclosed list of zero or more *statements*."

The terms parameter and attribute represent S.A.G.E. *keywords* specified throughout this document, and the braces ({ }) are used to enclose an optional *block* of zero or more subsequent statements<sup>6</sup>. Further, the < and > symbols may be used instead if there are no { and } symbols on the user's keyboard.

The recursive manner by which a statement is defined in terms of itself is no accident and is, in fact, a common way to specify a formal language structure.

<sup>&</sup>lt;sup>5</sup>Parameters and attributes often do not require an explicit value to be assigned, allowing the user to run the selected S.A.G.E. program with its default values.

<sup>&</sup>lt;sup>6</sup>When a brace-enclosed block is nested within another, enclosing block, the nested blocks are referred to as *sub-blocks*.

An illustration of the overall structure of a parameter file is as follows:

```
parameter = value, attribute = value
parameter = value, attribute = value
parameter, attribute = value, attribute = value, attribute = value
parameter = value, attribute = value
parameter = value {
  parameter = value
   parameter, attribute = value
   parameter {
      parameter = value
      parameter = value
      parameter = value
   }
   parameter {
      parameter = value
      parameter = value
      parameter = value
      parameter = value, attribute = value
      parameter, attribute = value, attribute = value, attribute = value
      parameter = value, attribute = value
   }
   parameter = value, attribute = value
   parameter, attribute = value, attribute = value, attribute = value
   parameter = value, attribute = value
}
```

The S.A.G.E. statement grammar described above is complex, which can make the software difficult to learn. As noted previously the S.A.G.E. GUI is designed to eliminate the burden of learning the parameter file syntax; however, there may be times when an experienced user would prefer to manipulate the parameter files directly. The following section provides clarifying details and examples of parameter file syntax.

#### 2.2.2.1 Syntax Details

- 1. The specific parameters and attributes listed within this document are S.A.G.E. *reserved words*, meaning that they must be spelled exactly as shown in their corresponding syntax table (see 2.2.4).
- 2. The names of phenotypes, traits and covariates found in the pedigree data file may be the same as the names of parameters and attributes, although this practice is likely to cause confusion and is therefore not recommended.
- 3. White space, including blanks, tabs and newline characters, are required only to differentiate between successive parameters, attributes and values, and are otherwise ignored<sup>7</sup>

<sup>&</sup>lt;sup>7</sup>White space that occurs as part of a QUOTED character string (e.g., "Body Mass Index") is not ignored.

by S.A.G.E. programs. They may usually be inserted or omitted from statements at the user's discretion.<sup>8</sup>

- 4. Blank lines between successive statements are ignored and may be inserted as necessary to make the parameter file easier to read.
- 5. Due to the recursive nature of their definition, statements can be *nested* when listed in the parameter file. That is, a particular statement may be specified as containing another *sub-statement* which in turn contains one or more *sub-sub-statements*, etc. This manual refers to such sub-statements as *sub-blocks*.
- 6. A single statement may fit on a single line or may continue over several lines. Further, the placement of braces (for enclosed blocks and sub-blocks) is left entirely to the discretion of the user. The following example shows two ways to specify the same segreg statement:

```
segreg,out="my_analysis.out"{trait=BMI type_mean{option=three}}
```

```
segreg, out = "my_analysis.out"
{
    trait = BMI
    type_mean
    {
        option = three
    }
}
```

#### 2.2.2.2 Adding Comments to the Parameter File

The insertion of a pound sign (#) at any point of a line in a parameter file causes S.A.G.E. to ignore the remainder of that line. Thus, the user can *comment* on the contents of a parameter file that may need to be reviewed at some future time. The following example shows how the above-listed segreg block might be commented:

```
# Perform analysis on Body Mass Index
segreg, out = "my_analysis.out"
{
    trait = BMI
    type_mean
    {
        option = three # Run the 3-mean model
    }
}
```

<sup>&</sup>lt;sup>8</sup>Many users find that judicious use of blank spaces can make a parameter file easier to read, and therefore less prone to error.

Here is a more elaborate approach to commenting a S.A.G.E. parameter file:

```
# Begin File: Example_01.par
S.A.G.E. Parameter File
# *****
                     ********
          C. L. Dodgeson
# *****
                     ******
# ****
                     ************
            9 Sep 03
***** PEDIGREE BLOCK
                     **********
#
pedigree, character {
 # General specifications
 # ______
            _____
 delimiters = ","
delimiter_mode = single
 individual_missing_value = " "
 sex_code, male = M
sex_code, female = F
 _____
 # ------
 # Field specifications
 # _____
 binary,
  affected = 1,
  unaffected = 0
 phenotype = HEMATOCRIT, name = Hematrocrit
}
#
****** PEDINFO BLOCK ********** #
#
pedinfo, out = "assgn 01 pedinfo altered out.txt" {
 phenotype = "Aff Stat"
phenotype = Hematrocrit
 each_pedigree = false
}
# ------
# End File: Example_01.par
    _____
# ---
```

#### 2.2.3 Parameter and Attribute Values

#### 2.2.3.1 Character Strings

When a particular parameter or attribute takes a *character string*<sup>9</sup> for its value, the user should enter the desired *alphanumeric* character sequence<sup>10</sup> after the equal sign (=). Enclosing *double quotes*<sup>11</sup> are only required in the following cases:

- the string contains blank spaces (e.g., "Alice in Wonderland"),
- the string contains non-alphanumeric characters<sup>12</sup> (e.g., "Alice-in-Wonderland")
- the string contains no characters at all i.e., it has length of zero (e.g., ""). A zero-length string is sometimes referred to as a *null* string.

S.A.G.E. is *case insensitive* with respect to the names of traits, phenotypes and covariates, and therefore, the following statements are equivalent:

trait = HT, type = continuous trait = hT, type = continuous trait = Ht, type = continuous trait = ht, type = continuous

In all other cases, S.A.G.E. is case sensitive.

#### 2.2.3.2 Numeric Values

When a particular parameter or attribute takes a numeric quantity for its value, the user is required to enter a constant according to the normal conventions of decimal notation. Specific constraints on the value are as follows:

<sup>&</sup>lt;sup>9</sup>A character string is simply a contiguous sequence of zero or more letters, digits, or other typographic symbols, including spaces.

<sup>&</sup>lt;sup>10</sup>An alphanumeric string may contain only letters (upper or lower case) and decimal digits.

<sup>&</sup>lt;sup>11</sup>We distinguish between two kinds of quotation marks: double quotes ("") and single quotes (''). Unless otherwise stated, the double quotes should be used whenever the syntax rules call for a quoted string.

<sup>&</sup>lt;sup>12</sup>Typographic symbols OTHER than letters or digits: ~!@#\$%^&\*()+'-={}[[\:";'<>?,./

Numeric Value Constraint Notation					
Notational Form	Meaning				
$(-\infty,\infty)$	Any real number				
$(-\infty,\infty)$ - $\{a\}$	Any real number except for <i>a</i> .				
$(a,\infty)$	Any real number greater than <i>a</i> .				
$[a,\infty)$	Any real number greater than or equal to <i>a</i> .				
$(-\infty, b)$	Any real number less than <i>b</i> .				
$(-\infty, b]$	Any real number less or equal to <i>b</i> .				
(a,b)	Any real number greater than <i>a</i> and less than <i>b</i> .				
[a,b)	Any real number greater than or equal to <i>a</i> and less than <i>b</i> .				
(a, b]	Any real number greater than <i>a</i> and less than or equal to <i>b</i> .				
[ <i>a</i> , <i>b</i> ]	Any real number greater than or equal to <i>a</i> and less than or equal to <i>b</i> .				
$\{x_1, x_2, x_3,, x_n,\}$	Any one of a discrete list of given items.				
${i, i+1, i+2,, n}$	Any integer from <i>i</i> to <i>n</i> , inclusive.				
${i, i+1, i+2,}$	Any integer greater than or equal to <i>i</i> .				
{, <i>i</i> -2, <i>i</i> -1, <i>i</i> }	Any integer (positive or negative) less than or equal to <i>i</i> .				

In addition to decimal quantities, S.A.G.E. also accepts the following named constants:

- **pi**, designating the transcendental number  $\pi = 3.141592654...$
- e, designating the base of the natural logarithms = 2.718281828459...

The following example shows some ways in which numeric values may appear within S.A.G.E. statements:

```
segreg {
   composite_trait {
     covariate = BMI, val = 27.69
   }
   transmission {
     option = homog_general
     tau = A*, val = 0.5
   }
}
```

#### 2.2.4 Reading and Interpreting the Syntax Tables

As mentioned previously, S.A.G.E. requires the user to specify options to its various programs by entering a list of statements into a parameter file, in a manner similar to common programming languages such as C++ and Python. When S.A.G.E. is executed, it reads the parameter file to determine

1. how to interpret the given pedigree data file(s),

- 2. which analyses (programs) are to be run, and
- 3. how to configure itself for the requested analyses.

For every parameter defined within the S.A.G.E. family of programs, this user document provides the following information in tabular form:

- parameter designation
- list of attributes associated with a given parameter
- a brief explanation
- range of valid or possible values that the parameter or attribute can take (see 2.2.3.2)
- the default value
- whether or not a value is required
- a list of applicable notes when more detailed explanation is required; these notes will always be found immediately at the end of the table.

To understand	how to	o interpret	the s	syntax	tables	used	in this	document,	consider	the following	5
example:											

parameter	Explanation					
[, attribute]	•					
	Declare respective pedigree field names for pedigree					
	$ID^b$ , individual ID, parental ID and sex designator.					
		Character string representing the				
pedigree_id	Value Range	valid name of a field <sup>c</sup> in the				
individual id		pedigree data file.				
parent_id	Default Value	None <sup>d</sup>				
sex_field <sup>a</sup>	Required	Yes <sup>e</sup>				
Sex_IIEIu	Applicable Notes	$1, 2^{f}$				
	Declare codes used	to specify sex of individuals in the				
	pedigree.					
sex code	Value Range	N/A <sup>g</sup>				
sex_code	Default Value	None				
	Required	No				
	Applicable Notes	None				
	Specifies male sex code					
	Value Range	Character string. Typical values				
, male $^{h}$	value italiye	are: 1, 0, M, m				
, maie	Default Value	М				
	Required	No				
	Applicable Notes	None				
	Specifies female sex code.					
	Value Range	Character string. Typical values				
, female	value italiye	are: 0, 1, F, f				
, remare	Default Value	F				
	Required	No				
	Applicable Notes	None				

<sup>*a*</sup>The occurrence of multiple *parameter* names in a single cell means that the explanatory information at the right is applicable to all of them, and attributes listed within the cell are also applicable to all of them.

<sup>b</sup>An acronym for *identifier*.

<sup>c</sup>For users who are accustomed to spreadsheets, the database term *field* is analogous to *column*, and the term *record* is analogous to *row*.

<sup>*d*</sup>**None** means the default value is unspecified, inapplicable or both.

<sup>e</sup>If Yes, then the listed parameter or attribute is required, and the user must explicitly enter the listed parameter or attribute, optionally followed by an assignment expression (eg., sex\_code, male = "M"), into the parameter file. If No, then the listed parameter or attribute is *not* required, and the specified default value will be used in the analysis. Note: When relying on default values for a given analysis, the user should take care to ensure that they are appropriate for the intended model.

<sup>*f*</sup>The applicable notes will be found immediately below the table.

<sup>*g*</sup>N/A means *not applicable*, i.e., that the parameter or attribute in question is *self-defining* and does not take any values.

<sup>h</sup>Attributes are indented with respect to their associated parameters, but appear in the same cell. Relevant explanatory information appears to their immediate right.

## 2.3 The Pedigree Data File

For family data to be accurately analyzed they must be described and represented precisely. The following are the definitions of various non-obvious family structures and relationships that are used throughout this manual<sup>13</sup>.

Term	Definition				
pedigree <sup>a</sup>	A set of individuals identified as belonging to the same pedigree,				
	i.e., having the same pedigree $ID^b$ . These individuals may or may				
	not be related in any way, but those who are NOT members of the				
	same pedigree should NOT be related.				
individual	A member of a pedigree.				
founder <sup>c</sup>	An individual with at least one descendant who has neither parent				
	in the pedigree. Founders are assumed to be unrelated by ancestry				
	to any other founder.				
non-founder	An individual descended from at least one founder.				
mate relationship	Two individuals in a pedigree who have one or more offspring with				
	each other are related by a mate relationship. Each individual may				
	be a member of several mate relationships.				
nuclear family	A set of two individuals who have a mate relationship and their				
	natural children.				
constituent pedigree <sup>d</sup>	A <i>complete</i> set of individuals in the same pedigree who are related				
	by marriage, ancestry or descent and for whom there is enough				
	information to indicate that they are so related. By complete is				
	meant that all individuals in the pedigree who are so related must				
	be included in the constituent pedigree.				
singletons <sup>e</sup>	The set of individuals who have no relation to any other member				
	of the pedigree they belong to.				
marriage ring	A cycle of mate relationships in the undirected graph of individuals				
	in a constituent pedigree.				
non-marriage loop	A cycle containing at least one offspring and one mate relationship				
	in the undirected graph of individuals in a constituent pedigree.				
	(This includes consanguineous loops and non-consanguineous				
	loops that involve both mate and offspring relationships).				

<sup>*a*</sup>Other software packages refer to our definition of pedigrees as kindreds.

<sup>b</sup>"ID" is an acronym for *identifier*, and is used frequently throughout this document.

<sup>c</sup>Founders do not include singleton individuals.

 $^{d}$ A constituent pedigree is what is typically referred to as a pedigree in the literature. The distinction is made because of the prevalence of incomplete and fragmented datasets.

<sup>e</sup>Singletons are sometimes not differentiated from founders in the literature.

### 2.3.1 Pedigree Data File Specification

A pedigree data file is a text file composed of one or more records, each of which contains information about a single individual in a pedigree. Each record must end with a carriage return or linefeed

<sup>&</sup>lt;sup>13</sup>Some of these definitions are fairly technical but under most circumstances the conventional definitions will suffice.

Pedigree Data Field Requirements				
Field	Value Type	Description	Required	
Pedigree ID	character string or numeric	Uniquely identifies a particular pedi- gree within the file.	yes	
Individual ID	character string or numeric	Uniquely identifies a particular indi- vidual within a pedigree <sup><i>a</i></sup> .	yes	
Parent ID	character string or numeric	Identifier of the individual's parent (either father or mother).	yes	
Parent ID	character string or numeric	Identifier of the individual's other parent (either mother or father, depending on which was specified previously <sup><math>b</math></sup> ).	yes	
Sex	character string or numeric	Individual's sex <sup>c</sup> .	yes	
Continuous Traits, Phenotypes and Covariates	numeric	Observational trait and phenotype data with respect to the individual.	no	
Discrete Traits, Phenotypes and Covariates	character string or numeric	Observational trait and phenotype data with respect to the individual.	no	
Genotype Data	character string or numeric	Genotypic data with respect to the in- dividual.	no <sup>d</sup>	
Other Fields	character string or numeric	Other data related to the individual.	no	

character<sup>14</sup> and contains the following fields:

<sup>*a*</sup>Implicit in this is the possibility that the same Individual ID may appear more than once in a given pedigree data file, referring to a different individual at each occurrence.

<sup>b</sup>At the user's option, the pedigree data file may list the father's ID first, followed by the mother's ID.

<sup>c</sup>Incorrect use of the word *gender* is studiously avoided here. As the poet says, "Nouns have gender, whereas people have sex ... and enjoy it!"

<sup>d</sup>Genotypic data are, of course, required for programs that perform linkage analysis, allelic association analysis, etc.

There are two distinct types of record formats for pedigree data files: *character delimited* and *column delimited*. The two formats differ only in the method by which the pedigree fields are distinguished from one another within a given record, and are detailed in the following sections.

#### 2.3.1.1 Character Delimited records

In a *character delimited* record the individual fields are separated by one or more characters, known as *delimiters*, which are usually not present in any of the data elements themselves. Commonly used delimiters are the comma, the tab, and the space, but any non-alphanumeric character may be used. If your data are separated by a fixed known delimiter, then S.A.G.E. will read your pedigree file as character delimited records, and you will need to specify which delimiter is used along with some additional *metadata*<sup>15</sup> that specify the exact order, names and types of the fields in your pedigree

<sup>&</sup>lt;sup>14</sup>Any combination of carriage return and line feed characters is sufficient to terminate a record. This allows data files from most popular operating systems to be used without translation.

<sup>&</sup>lt;sup>15</sup>Database terminology that means "information about the data", i.e., field names, data types, value ranges, etc.

records.

Files that are formatted for LINKAGE, GENEHUNTER, PAP, GAS or similar computer programs may all be read as character delimited records with little or no modification<sup>16</sup>. Programs that readily generate data in a character delimited form are spreadsheet programs like Microsoft Excel, most pedigree drawing programs, and most database programs.

Character delimited pedigree data file records can also be read as *column delimited* records provided that:

- 1. each record is exactly one line long,
- 2. there is at least one delimiter character between fields,
- 3. all fields contain at least one non-delimiter character, i.e., no field is empty, and
- 4. the corresponding fields in different records are the same length.

Pedigrees stored in character-delimited format contain exactly one line for each individual with specific delimiter characters separating data fields.

THE FORMAT OF THE CHARACTER DELIMITED PEDIGREE DATA FILE IS DEFINED BY A CHARACTER-DELIMITED LIST OF DISTINCT NAMES THAT IDENTIFY EACH FIELD. THIS LIST OF NAMES MAY BE SPECIFIED AS THE FIRST LINE, OR HEADER, OF A CHARACTER DELIMITED PEDIGREE DATA FILE; OR ALTERNATELY, IT MAY BE GIVEN AS A SET OF PARAMETERS IN A PEDIGREE BLOCK WITHIN THE PARAMETER FILE.

The content of each field in this list has no default semantic meaning and the field it identifies may be used for any purpose once read in.

Associating a field with a meaning, such as a pedigree ID, individual ID, marker phenotype, trait, etc. is accomplished by specifying additional parameters to map the field names to data field types. It is not necessary to specify all fields, and the parameter file may even specify fields that do not appear in the file so long as all required fields are in the pedigree data file. Whitespace is stripped from the beginning and end of the content of each field.

Here is an example of a comma delimited pedigree data file that includes the name of each field in a header line, one of the many possible character delimited formats:

pedid,	indid,	mom,	dad,	sex,	trait1,	trait2,	markerl
1,	1,	,	,	М,	Affected,	10.3,	A/A
1,	2,	,	,	F,	Unaffected,	1.3,	a/a
1,	3,	1,	2,	Μ,	Affected,	7.9,	A/a

Several options are provided to let the user modify the way a character delimited pedigree is processed by S.A.G.E. programs. The sets of characters that represent whitespace and delimiter characters may be redefined. There is an option that alters the way multiple consecutive delimiter characters are interpreted, by treating them as a single delimiter. This is extremely useful when reading multiple space delimited, or other fixed column formats, that do not include empty fields. Empty

<sup>&</sup>lt;sup>16</sup>If necessary, column-delimited input files, such as those required for PAP can be imported into a spreadsheet program (Microsoft Excel, for example) and then exported in a character delimited format.

2.3. THE PEDIGREE DATA FILE

fields are a problem in this mode because it is not possible to detect them. For example, suppose each line in the following fixed column, space delimited, file is parsed into 6 fields using the delimiters and delimiter\_mode options to read multiple blanks as a single field delimiter and skip leading and trailing blanks. The following delimited pedigree file is correctly specified:

PEDID	INDID	MOM	DAD	SEX	TRAIT1
1	1	0	0	М	0
1	2	0	0	F	0
1	3	1	2	М	2

If 0, the missing value code for parents and traits in this example, were replaced with a space character as indicated below, the resulting fixed column records would be parsed inconsistently. The two parents would have the SEX field as their MOM field, as well as other errors due to missing values not being detected.

PEDID	INDID	MOM	DAD	SEX	TRAIT1
1	1			М	
1	2			F	
1	3	1	2	М	2

### 2.3.1.2 Column delimited records

In a *column delimited* record the individual fields are distinguished from one another by their respective locations within the record. Each character in the record occupies a single location, or *column* and, starting with the leftmost column, the locations are identified as column 1, column 2, column 3 and so on. In the following example, the word "Queen" is located *at* column five<sup>17</sup>, and the word "tarts" is located at column 36:

```
123456789_123456789_123456789_123456789_123456789_123456789_123456789_
The Queen of Hearts, she made some tarts, all on a summer day.
```

Once the appropriate pedigree record format has been determined, a single pedigree statement in the parameter file is used to specify the structure and content of the file; it has the following syntax:

```
pedigree[,column]
{
    [statement]*
}
```

Pedigree records are assumed to be character delimited by default. You can configure S.A.G.E. to accept column delimited records by including the column attribute, as in the following example:

<sup>&</sup>lt;sup>17</sup>Meaning that the first letter of the word is *in* column five, and extending to the right for as many additional columns as there are remaining letters.

```
pedigree, column
{
    format = "4A5, 1X, A1, ..."
    individual_missing_value=0
    sex_code, male=1, female=2, unknown=?
    pedigree_id
    individual_id
    mother_id
    father_id
    sex_field
    ...
}
```

Each record in a pedigree data file stored in column delimited form comprises one or more sequential lines that contain fields that are at a fixed offset from the beginning of the line. No separation is required between fields because each field is of fixed length and each record is a fixed number of columns long. Parameters must be specified to define the order of the fields, the locations and widths of each field, as well as information on how each field is encoded. A technique borrowed from the FORTRAN programming language, called a FORTRAN format statement, is used to specify the locations, widths and other information on how the fields are encoded. A tutorial on how to create FORTRAN format statements can be found on the S.A.G.E. web site at http://darwin.cwru.edu/sagegui/help/fortran.html, or in most FORTRAN reference texts.

### 2.3.1.3 Pedigree Data Quality

Users are always well-advised to ensure that their pedigree data files are as error-free as possible<sup>18</sup>, with particular attention paid to the correctness of family relationships within individual pedigrees. Nevertheless, S.A.G.E. programs are able to run in the presence of less-than-perfect data. With the exception of the five required fields mentioned previously, missing data will not prevent S.A.G.E. analyses from running to completion.

**Note**: If the pedigree block of the parameter file lists a variable that does not appear (or is spelled differently) within the corresponding pedigree data file, S.A.G.E. will issue an appropriate error message and halt immediately.

# 2.3.2 Pedigree Block Syntax

The following table shows the S.A.G.E. syntax for specifying the structure and content of a pedigree data file. Unless otherwise noted, all parameters and their corresponding attributes must be specified within a pedigree block of a parameter file.

The following table shows the syntax for a pedigree block statement:

<sup>&</sup>lt;sup>18</sup>A well-known software apothegm is "garbage in, garbage out", also expressed as the acronym GIGO.

parameter [, attribute]		Explanation
	Starts a pedigree sp	pecification block.
	Value Range	{column, character}
pedigree	Default Value	character
	Required	No
	Applicable Notes	1
	Specifies the name	of a pedigree data file.
		Character string representing a
641.	Value Range	valid file name.
, file	Default Value	None
	Required	No
	Applicable Notes	2,3
	Specifies that the	fields in the pedigree data file
	are character delir	nited, as opposed to the column-
	oriented format.	
		delimited
, delimited	Value Range	column
	Default Value	delimited
	Required	No
	Applicable Notes	2,4

1. In this case, a comma is used to indicate that a value is being given, as in the following examples:

```
pedigree, character #This is the default setting
{
    #Pedigree specifications follow
    ...
}
pedigree, column
{
    #Pedigree specifications follow
    format = "4A5, 1X, A1, ..."
}
```

2. S.A.G.E. programs are capable of processing multiple pedigree data files simultaneously. This feature is especially useful for analyzing marker data that span the entire genome, in which case each chromosome is normally allocated to its own pedigree data file. To analyze data across multiple pedigree files, create a separate pedigree block for each pedigree file, and use the file attribute to name a particular file, as in the following example:

```
pedigree, delimited, file = "Chr1.ped"{
   delimiters = "\t" # The '\t' indicates the tab key
delimiter_mode = multiple
   individual_missing_value = 0
   . . .
   pedigree_id = PID
   individual_id = ID
   parent_id = P1
   parent_id = P2
   sex_field = sex
   . . .
                              name = D1S2195
   allele = D1S2195a,
   allele = D1S2195b,
                               name = D1S2195
   allele = D1S1426a,
                               name = D1S1426
   allele = D1S1426b, name = D1S1426
   . . .
}
pedigree, delimited, file = "Chr2.ped"{
   delimiters
                               = "\t"
   delimiter_mode
                               = multiple
   individual_missing_value = 0
   . . .
   pedigree_id = PID
   individual_id = ID
   parent_id = P1
   parent_id = P2
   sex_field = sex
   . . .

      allele = D2S2195a,
      name = D2S2195

      allele = D2S2195b,
      name = D2S2195

      allele = D2S1426a,
      name = D2S1426

      allele = D2S1426b,
      name = D2S1426

   . . .
}
pedigree, delimited, file = "Chr3.ped"{
   delimiters = "\t"
   delimiter_mode
                              = multiple
   individual_missing_value = 0
   . . .
   pedigree id = PID
   individual_id = ID
   parent_id = P1
   parent_id
                 = P2
   sex_field = sex
   . . .
   allele = D3S2195a,
                              name = D3S2195
                               name = D3S2195
   allele = D3S2195b,
                               name = D3S1426
   allele = D3S1426a,
                               name = D3S1426
   allele = D3S1426b,
   . . .
}
```

3. Even if the user specifies a filename at the start of each pedigree block, as shown in the above example, it is still necessary to supply the name of a pedigree filename on the program command line when running the program.

4. When using the column delimited format, the data fields will automatically be converted into and stored as the most appropriate data type for the given data. Thus, the fields for pedigree\_id, individual\_id, parent\_id, and sex\_field will all be stored as string types, whereas trait, phenotype and covariate values will be stored as numeric types (integers or reals) if at all possible. Some types of categorical data may only be storable as strings (for example, "High", "Medium" and "Low"). Also the user can force a numeric quantity to be stored as a character string by using the string option when specifying the field's name in the pedigree block.

Finally, the pedigree data file name specified at the S.A.G.E. program *command line* will automatically be assigned to the first pedigree block that does not specify the file attribute. Suppose the statements listed in the above example were contained in a parameter file named *hypertension\_study.par*. Then the file attribute for the *first* pedigree block would be optional if the command line were:

>freq hypertension\_study.par Chr1.ped

parameter	Evaluation			
[, attribute]	Explanation			
	Designates either			
	<ol> <li>a FORTRAN-style format statement used to specify the record layout of a column delimited pedigree file, or</li> <li>a delimited sequential listing of each field of a</li> </ol>			
format	2. a delimited, sequential listing of each field of a character delimited pedigree data file.			
	Value Range Quoted character string.			
	Default Value None			
	Required Required for character delim- ited pedigree data files with a header record as the first entry. Required otherwise.			
	Applicable Notes 1, 2			
	Specifies delimiter characters			
	Value Range Quoted character string.			
delimiters	Default Value ", \t"			
	Required No			
	Applicable Notes 3			
	Specifies characters that must be treated as whites-			
	pace.			
whitespace	Value Range Quoted character string.			
wiiitespace	Default Value "" (blank space)			
	Required No			
	Applicable Notes 3			
	Specifies delimiter interpretation mode. If set to			
	multiple, then a set of successive delimiters in the			
	pedigree data file will be treated as a single delimiter.			
delimiter_mode	Value Range {single, multiple}			
	Default Value single			
	Required No			
	Applicable Notes 4			
	Specifies the number of individual records from the			
	pedigree file to be printed to the program information			
	output information file for visual verification of cor-			
	rectness.			
verbose	Value Range {0, 1, 2, 3,}			
	10, meaning that the first ten			
	Default Value pedigree records will be printed			
	to the information output file.			
	Required No			
	Applicable Notes None			

# 2.3.2.1 General Pedigree Formatting Options

	Specifies whether or not to omit automatically gener- ated "dummy" parent records in the pedigree data file for individuals with missing parent data. A value of		
	false means that the parent records will be added to		
require_record	the analysis as needed.		
	Value Range	{true, false}	
	Default Value	false	
	Required	No	
	Applicable Notes	5	

- 1. The format parameter is used to list the name of each field in the character delimited pedigree data file. Its value should be a delimited list of field names in the same order as those to be read from the file. The delimiter characters that separate each field name in this list are the same as those given in the delimiters parameter. If this parameter is not given, or is empty, then the first line of the character delimited pedigree file will be used to specify the format parameter.
- 2. In column delimited pedigree records, the fields are read into the program according to the order presented in the format parameter of the pedigree block.
- 3. The delimiters parameter specifies the characters that separate fields in each record. As a result, the delimiter characters should not be present in any fields. The default is that any comma (,) or tab (\t) character is interpreted as a delimiter character. Similarly, the whitespace parameter specifies characters that will be ignored when they appear at the beginning or end of fields.
- 4. The delimiter\_mode parameter is used to alter how records are parsed. When the value of delimiter\_mode is set to single each delimiter character found will terminate the current field. When the value of delimiter\_mode is set to multiple, consecutive delimiters are treated as a single delimiter and delimiters that occur at the beginning and end of the record are ignored. Typically, tab and comma delimited files should be set to the value single, while space delimited files should be set to the value multiple.
- 5. By default, each individual in a pedigree must have one record in the pedigree data file. However, data on sibships without parent data are not uncommon. Distinguishing parent IDs must still be assigned to all individuals, but empty records for the dummy parents can be omitted if the require\_record parameter is set to **false**.

parameter		Explanation		
[, attribute]	-			
	Specifies pedigree field names for pedigree ID, indi-			
	vidual ID, parental ID and Sex.			
		Character string representing the		
pedigree_id	Value Range	valid name of a field in the		
individual_id		pedigree data file.		
parent_id	Default Value	None		
sex_field	Required	Yes		
SCA_IICIU	Applicable Notes	1, 2		
	Specifies codes for	or missing individuals (typically		
	founders).			
		Character string. Typical values		
		are:		
		0,		
individual_missing_value	Value Range	"" (zero-length string),		
		"" (blank space),		
		999,		
	<b>-</b>	-1		
	Default Value	-1 "" (zero-length string)		
	Required	INO		
	Applicable Notes	3		
		to specify sex of individuals in the		
	pedigree.	NT / A		
sex_code	Value Range	N/A		
_	Default Value	N/A		
	Required	No		
	Applicable Notes	4		
	Specifies male sex code			
	Value Range	Character string. Typical values		
, male	-	are: 1, 0, M, m		
	Default Value	M		
	Required	No		
	Applicable Notes	None		
	Specifies female se			
, female	Value Range	Character string. Typical values		
	-	are: 0, 1, F, f		
	Default Value	F		
	Required	No		
	Applicable Notes	None		

# 2.3.2.2 Parameters for Individual and Family Identification Fields

	field. May be diff	value code specifically for the sex ferent from the missing value code e, trait and covariate fields.	
, missing	Value Range	Character string. Typical values are: 0, "" (zero-length string), " " (blank space), 999,	
	Default Value	-1 "" (zero-length string)	
	Required	No	
	Applicable Notes	None	
	Declares sex code	to be a binary trait, and subject to	
	statistical analysis as such. Automatically creates a		
	trait, sex_code,	with values as follows:	
, trait	<ul> <li>male = 0</li> <li>female = 1</li> </ul>		
	Value Range	N/A	
	Default Value	N/A	
	Required	No	
	Applicable Notes	5	

- 1. For character delimited pedigree records, the values assigned to the pedigree\_id, individual\_id, parent\_id, and sex\_field parameters are used to identify the order and location of the pedigree ID, individual ID, parent IDs and sex fields for each individual. Each of these values should match an element in the format parameter or the column header provided in the pedigree data file. The number of, and order in which, these fields are specified is arbitrary, and not all of these fields need appear in the pedigree data file, subject to several constraints. These constraints are that each record in the pedigree data file must include a pedigree ID and an individual ID, no more than one sex field and exactly two parent ID fields.
- 2. For column delimited pedigree records, the optional pedigree\_id, individual\_id, parent\_id, and sex\_field parameters should not be assigned any value. They are used to define the order and location of the pedigree ID, individual ID, parent IDs and the sex field for each individual. If any of these parameters are specified, then no default order is assumed and a parameter must be included for each field that is to be read in. If none are specified then it is assumed that the first five fields specified in the format statement correspond to the these fields in the following order:
  - pedigree ID
  - individual ID
  - Sex of individual

- First parent ID
- Second parent ID
- 3. This is the code that is used in the parent ID fields of founders.
- 4. Subject to the constraints described in note #2 (above), Sex codes may be specified separately or in the same parameter. e.g.:

```
# This is a valid sex_code parameter:
sex_code,male=M
sex_code,female=F
# ... and so also is this:
sex_code,male="M",female="F"
```

5. By including the attribute trait, the sex\_code can be used as a quantitative (0,1) variable.

	Specifies pedigree	Explanation
	Specifies pedigree 1	
	variate fields.	names for phenotype, trait and co-
phenotype	Value Range	Character string representing the valid name of a field in the pedigree data file.
trait	Default Value	None
covariate	Required	No
	Applicable Notes	1,2
		•
	Specifies missing va	
		Character string. Typical values
		are:
	Value Range	
	value Ralige	"" (zero-length string),
, missing		"" (blank space),
		999,
	Default Value	-1 "" (zero-length string)
	Required	No
	Applicable Notes	None
	Indicator for binary	
		N/A
, .	Value Range _ Default Value	N/A None
, binary	=	
	Required	No None
	Applicable Notes	
	_	affected status of binary pheno-
	types.	
		Character string. Typical values
, affected	Value Range	are: A, 1, AFFECTED,
		Affected, yes, true, pos
	Default Value	1
	Required	No
	Applicable Notes	None
	Specifies code for u	unaffected status of binary pheno-
	types.	
		Character string. Typical values
, unaffected	Value Range	are: U, 0, UNAFFECTED,
, unarrected		Unaffected, no, false, neg
	Default Value	0
	Required	No
	Applicable Notes	3

# 2.3.2.3 Parameters for Phenotype, Trait & Covariate Fields

	Specifies a field name different from the name given			
	in the pedigree header line.			
	Value Range Character string.			
, name	Default Value None			
	Required No			
	Applicable Notes 3			
	Designates a pedigree field that the user wishes to ma-			
	nipulate and/or report along with other data from an			
	individual's record (an assay bar code, for example).			
	Used for fields that do not fall into one of the previ-			
	ously mentioned types:			
	• phenotype			
	• trait			
string	• covariate			
	• marker			
	• allele			
	• trait-marker			
	Value Range Character string.			
	Default Value None			
	Required No			
	Applicable Notes None			

- 1. The phenotype, trait and covariate parameters all perform the same basic function; the values assigned to them identify fields in the pedigree data file that contain continuous or discrete phenotypic information The following guidelines may clarify when each different parameter should be used:
  - (a) phenotype fields are a generic designation and convey no suggestion of how the field is to be used.
  - (b) trait fields are typically selected by many analyses to be used as major variates.

Thus these parameters simply provide hints to S.A.G.E. on how to make reasonable use of phenotypic information. Refer to the program specific documentation for information on the specific behaviors of these parameters and how to override them.

Each trait field in a record may contain any character string that represents

- a missing value code,
- the affected or unaffected phenotype code (for binary phenotypes), or
- a numeric phenotype (for quantitative phenotypes).

- 2. The phenotype, trait and covariate parameters should be included for each field in the pedigree data file that contains quantitative or categorical phenotypic information. The value of each such parameter should be set to the name by which it will be referred to in the rest of the parameter file and in the program output. Like other parameters that specify fields in a column delimited pedigree file, the order of the parameters is important and determines how the fields specified in the format statement are interpreted. **Remember: any field specified as a phenotype will** *automatically* **be analyzed by S.A.G.E. programs**, whereas fields specified as either trait or covariate will be analyzed optionally, depending on whether or not they have been listed within the relevant analysis block. The choice between (dependent) trait and covariate is largely a device to help the user remember how the
- 3. A name attribute may optionally be attached to phenotype, trait and covariate parameters, effectively creating an internal *alias* for the field. This feature is useful when the original names listed in the pedigree data file are obscure or unclear (usually due to abbreviation), and the user would like to create analyses, models and reports with more informative names.

For example, if the pedigree data files contains four fields named Trait1, Trait2, Covariate1, and Affection, then the user may specify alternate names as in the following example:

```
phenotype = Trait1, name = "Generic phenotype", missing = "X"
trait = Trait2, missing = "-99"
covariate = Covariate1, name = "Covariate #1"
trait = Affection, binary, affected = 1, unaffected = 0, missing = "?"
```

As a result, the field originally designated as "Trait1" could be referenced as "Generic phenotype" within S.A.G.E. analyses, and similarly, the field originally designated as "Covariate1" could be referenced as "Covariate #1" within subsequent analyses.

parameter		Explanation
[, attribute]		-
		field name for a particular allele or
	marker.	
allele	Value Range	Character string
marker	Default Value	None
	Required	No
	Applicable Notes	1, 2, 3, 4, 5, 6
	Indicator for X-lin	ked marker.
	Value Range	N/A
, x_linked	Default Value	None
	Required	No
	Applicable Notes	None
	Indicator for Y-link	ked marker.
	Value Range	N/A
, y_linked	Default Value	None
	Required	No
	Applicable Notes	None
	Specifies missing v	value code
		Character string. Typical values
		are:
		0,
	Value Range	"" (zero-length string),
, missing		"" (blank space),
		999,
		-1
	Default Value	"" (zero-length string)
	Required	No
	Applicable Notes	None
	Specifies a marker	r or allele name different from the
	name given in the	pedigree header line.
	Value Range	Character string.
, name	Default Value	None
	Required	No
	Applicable Notes	6
	Character used to d	delimit alleles of codominant mark-
	ers in a pedigree of	data file. This is only necessary if
	markers are read in	n as a single field and are codomi-
<b>a a b b b b b b b b b b</b>	nant.	
, delimiter	Value Range	Quoted character string.
	Default Value	
	Required	No
	Applicable Notes	None

# 2.3.2.4 Parameters for Genotype Data Fields

	Specifies minimum	allele frequency for the marker.
	Value Range	[0, 1]
<pre>, minimum_allele_freq</pre>	Default Value	None
, minimum	Required	No
	Applicable Notes	7,8
Ĭ	Specifies maximum	n allele frequency for the marker.
	Value Range	[0, 1]
<pre>, maximum_allele_freq</pre>	Default Value	None
, maximum	Required	No
	Applicable Notes	7,8
Ň	Sets all allele frequ	iencies to be equal.
	Value Range	N/A
, equal	Default Value	None
, equal_allele_freq	Required	No
	Applicable Notes	7,9
	Sets allele frequence	cies proportional to complementary
	values.	
, complement	Value Range	N/A
, compl_allele_freq	Default Value	None
, compi_arrere_rreq	Required	No
	Applicable Notes	7, 10
	Designates a trait	for model-based linkage analysis
trait_marker	(e.g., for MLOD or	r LODLINK)
	Value Range	Character string
	Default Value	None
	Required	No
	Applicable Notes	11

- 1. The value is set to the name of the allele marker in the trait locus description file. The order of the parameters is important and determines how the fields specified in the format statement are interpreted.
- 2. Each allele field may be any character string that represents:
  - a missing value code, or
  - a single allele name.

Each marker field may be any character string that represents:

- a missing value code,
- an allele name, the allele delimiter character and another allele name, or
- a marker phenotype name.
- 3. A single marker parameter or two allele parameters should be included for each marker locus in the pedigree data file. Each marker locus field that is to be used should have a corresponding entry in the marker locus description file that defines its alleles, genotypes and

phenotype to genotype mapping. THOSE NOT FOUND IN THE MARKER LOCUS DESCRIP-TION FILE WILL NOT BE ANALYZED BY ANY APPLICATION THAT REQUIRES THE MARKER LOCUS DESCRIPTION FILE.

4. E.g., to specify three markers to be read, named D42S1, D42S3, D42S4; marker D42S1 is given by two allele fields, and the others are marker fields:

```
marker=D42S3  # Order is irrelevant
allele=D42S1a,name=D42S1 # First allele for marker D42S1
marker=D42S4
allele=D42S1b,name=D42S1 # Second allele for marker D42S1
```

5. E.g., to specify reading three markers named "D42S1", "D42S2", "D42S3", a trait named "Trait1", another marker "D42S4", a trait-marker called "MOD", and a binary covariate named "Cov" in sequential fields in each record of a column delimited pedigree file:

```
# Order is important
allele = D42S1 # First allele of D42S1
allele = D42S1 # Second allele of D42S1
marker = D42S2
marker = D42S3
trait = Trait1
marker = D42S4
trait_marker = MOD
covariate = Cov, binary, affected=1, unaffected=2, missing=3
```

6. A name attribute may optionally be specified for allele and marker parameters. It should be specified when the name of the field in the pedigree data file is not the same as the name that appears in the marker locus description file. If a name attribute is not specified, the marker name is assumed to be the field name. The order in which these fields are specified is arbitrary and not all the fields need appear in the pedigree data file.

A trait\_marker parameter should be included for each trait in the pedigree data file that is to undergo a model-based linkage analysis. The value is set to the name of the trait-marker in the trait locus description file. The order in which these fields are specified is arbitrary and not all fields need appear in the pedigree data file.

- 7. For each locus, the information can be modified by adding the proper attributes to marker/allele parameter within the pedigree block.
- 8. For frequency adjustment, add attributes to the marker/allele statement within the pedigree block. For example:

```
pedigree {
    marker = D1S111, minimum_allele_freq = p
}
```

This will replace with p all frequencies less than p, and then the frequencies will be normalized to sum to 1. The maximum\_allele\_freq parameter works in an analogous manner.

- 9. This will set all allele frequencies equal to 1/(number of alleles).
- 10. This will complement all allele frequencies and then normalize them to sum to 1.

11. A trait\_marker parameter should be included for each trait in the pedigree data file that is to undergo a model-based linkage analysis. Thus the trait becomes like a marker and has requirements similar to those of a marker parameter, and hence is called a trait-marker. Instead of mixing markers and trait-markers in the same locus description file, each trait-marker should have an entry in the trait locus description file. THOSE NOT FOUND IN THE TRAIT LOCUS DESCRIPTION FILE WILL NOT BE ANALYZED.

# 2.3.3 The marker Sub-Block

The following table show the correct syntax for the marker sub-block:

parameter	Explanation	
[, attribute]		
	Specifies allele frequency adjustment.	
	Value Range N/A	
allele_frequency	Default Value None	
	Required No	
	Applicable Notes None	
	Sets all allele frequencies to be equal.	
	Value Range N/A	
, equal	Default Value None	
	Required No	
	Applicable Notes 1, 2	
	Sets all allele frequencies proportional to complemen-	
	tary values	
a anna 1 an am t	Value Range N/A	
, complement	Default Value None	
	Required No	
	Applicable Notes 1, 2	
	Ensures that allele frequencies are not set below a	
	minimum value. Note that after normalization (to sum	
	to 1) some allele frequencies may be smaller than the	
, minimum	set minimum or larger than the set maximum	
, millinam	Value Range [0,1]	
	Default Value None	
	Required No	
	Applicable Notes 1, 2	
	Ensures that allele frequencies are not set above a	
	maximum value.	
. maximum	Value Range [0,1]	
, แลงาแนแ	Default Value None	
	Required No	
	Applicable Notes 1, 2	

	Specifies missing value code	
		Character string. Typical values
		are:
		0,
	Value Range	"" (zero-length string),
allele_missing		"" (blank space),
		999,
	_	-1
	Default Value	"" (zero-length string)
	Required	No
	Applicable Notes	4, 5, 7
	Character used to d	elimit alleles of codominant mark-
	ers in a pedigree of	lata file. This is only relevant if
	markers are read in as a single field and are codom-	
allele_delimiter	inant.	
	Value Range	Quoted character string.
	Default Value	·‹/››
	Required <sup>–</sup>	No
	Applicable Notes	2, 3, 6

- 1. See notes 7 through 10 of Section 2.3.2.4
- 2. equal has higher precedence than complement, complement has higher precedence than minimum and maximum, and minimum and maximum have the same precedence.
- 3. For setting delimiter other than the default value of /, add the parameter allele\_delimiter within the marker block.
- 4. The current global allele\_delimiter statement in the parameter file will be ignored when the value for allele\_delimiter is specified within the marker block.
- 5. For missing values other than the default, add the parameter allele\_missing within the marker block. For example:

```
marker {
    allele_frequency, minimum = p
    allele_delimiter = ":"
    allele_missing = "."
}
```

- 6. The current missing attribute in the locus description file is ignored when the value for allele\_missing in the marker block is specified.
- 7. For setting the delimiter to a value other than the default value of slash (/) add the allele\_delimiter attribute to the marker/allele parameter within the pedigree block. For example:

```
pedigree {
    marker = D1S111, allele_delimiter = ":"
}
```

8. To specify a missing value other than the default, add the parameter allele\_missing within the allele block. For example:

```
pedigree {
    marker = D1S111, allele_missing = "."
}
```

### 2.3.4 Character Delimited Pedigree Data File Examples

Here is a typical pedigree data file in comma delimited format:

Suppose each record in the above pedigree data file is one line long and you want to use the following fields:

Field	Field Name
Pedigree ID	PID
Individual ID	ID
Sex field	SEX
First Parent ID	P1
Second Parent ID	P2
Trait	TRAIT 1
Marker D42S1	D42S1
Marker D42S2	D42S2
Marker D42S3	D42S3
Marker D42S4	D42S4
Marker D42S6	D42S6

then the following pedigree parameter can be used to read the pedigree data file:

```
pedigree # Example character delimited parameter statement
{
    # The following format string could be used if the pedigree file did not
    # already include a header line. Do NOT include both!
    # format="PID,ID,P1,P2,SEX,JUNK,D42S1,D42S2,D42S3,D42S4,D42S5,D42S6,TRAIT 1"
    pedigree_id=PID # Pedigree Field Specification
    individual_id=ID
    sex_field=SEX
    parent_id=P1
    parent_id=P2
    trait="TRAIT 1",name="Trait",missing="XXXX" # order is irrelevant
    marker=D42S4
    marker=D42S1
    marker=D42S2
```

marker=D42S3
# Pedigree encoding information:
individual\_missing\_value="0"
sex\_code, male="m",female="f",missing="x"
}
allele\_delimiter=" " # Set the allele delimiter.

## 2.3.5 Column Delimited Pedigree File Examples

Here is a typical pedigree data file in column delimited format:

1018 1 m 1018 2 x 1018 3 1 2 f 2 5 3 1 3 6 7 8 1 2 2 7 2 7 1 23.1 1018 4 1 2 f 5 5 3 1 3 8 7 8 1 2 2 7 2 7 2 44.1 1018 5 1 2 f 2 5 3 1 3 8 7 8 2 2 2 7 2 7 1 1018 6 1 2 m 5 5 3 1 3 8 7 8 1 3 2 7 2 4 1 9.3 ....

Suppose each record is one line long with the following fields:

Field	Columns
Pedigree ID	1-4
Individual ID	6-7
Sex field	18
First Parent ID	9-10
Second Parent ID	12-13
Trait	65-68
Marker D42S1	29-33
Marker D42S2	35-39
Marker D42S3	41-45
Marker D42S4	47-51
Marker D42S5	53-57
Marker D42S6	59-63

Then the following pedigree parameter will read the pedigree data file:

```
# Set the marker delimiter to spaces
allele_delimiter=" "
# Example column delimited parameter
pedigree,column {
    # Pedigree Field Specification
    # FORTRAN Format Statement
    format="A4,1X,A2,T18,A1,T9,A2,1X,A2,T28,6(1x,A5),1X,A3"
    # Fields are listed in the order they are to be read
    marker="D42S1"
    marker="D42S2"
    marker="D42S3"
    marker="D42S4"
    marker="D42S5"
    marker="D42S6"
    trait="Trait 1"
```

```
# Pedigree encoding information
individual_missing_value=" "
sex_code, male="m",female="f",missing="x"
}
```

# **2.4 User-Defined Functions**

User-defined function parameters specify the creation of new traits, phenotypes, or covariates as a function of existing pedigree variables. Like other configuration parameters, function parameters may appear anywhere in the parameter file, but they are processed immediately after the pedigree data are read, IN THE ORDER IN WHICH THEY APPEAR. Thus, variables created by previous functions can be used in the specification of subsequent functions. Once created, a function variable may be used just like a trait, phenotype, or covariate read from a pedigree data file in all S.A.G.E programs except GENIBD, which does not currently support the use of function blocks.

# 2.4.1 The function Parameter

The following syntax table specifies the permissible parameter settings for the function parameter.

<pre>parameter [, attribute]</pre>	Explanation	
	Starts a function block.	
	Value Range	N/A
function	Default Value	None
	Required	N/A
	Applicable Notes	None

<pre>parameter [, attribute]</pre>		Explanation	
• • • • • • •	Names a constant. May appear multiple times if there		
		are multiple constants to be specified.	
	Value Range	Quoted character string.	
constant	Default Value	None	
	Required	No	
	Applicable Notes	None	
		ession used to calculate a value for	
	the constant.		
	Value Range	Quoted character string.	
, expression	Default Value	None	
	Required	No	
	Applicable Notes	See 2.4.2	
		riable. Only one function variable	
	per function statem	•	
	Value Range	Character string.	
phenotype	Default Value	None	
trait	Required	No	
covariate	Applicable Notes	1, 2	
	Specifies the algebraic expression used to calculate a value for the function variable.		
, expression	Value Range Default Value	Character string. None	
		Yes	
	Required		
	Applicable Notes	3, 4, 5	
	Specifies missing v		
		Character string. Typical values	
		are:	
	Value Range	0, "" ( have the state )	
	value Raliye	"" (zero-length string),	
, missing		"" (blank space),	
		999,	
	Default Value	-1 "" (zero-length string)	
	Required	No	
	Applicable Notes	None	
	Indicates a binary p		
	Value Range	N/A	
, binary	Default Value	N/A None	
, Dillary	Required	N/A	
		N/A None	
	Applicable Notes		

The following lists all parameters that may occur in a function block.

	Specifies code for	affected status of binary pheno-
	types.	
	Value Range	Character string
, affected	Default Value	1
	Required	No
	Applicable Notes	None
	Specifies code for	unaffected status of binary pheno-
	types.	
. unaffected	Value Range	Character string.
, unallected	Default Value	0
	Required	No
	Applicable Notes	None
	Specifies a time limit, in seconds, for evaluating con-	
	stants and expressions.	
time_limit	Value Range	{0, 1, 2, 3,}
	Default Value	30
	Required	No
	Applicable Notes	6

- 1. The three possible function variable types are phenotype, trait, and covariate. Phenotype and covariate fields are a generic designation and convey no suggestion of how the field is to be used. Trait fields are typically selected by many analyses to be used as major variates. Thus these parameters simply provide hints to make reasonable use of phenotypic information. Refer to the program-specific documentation for specific information on the behaviors of these parameters and how they may be overridden.
- 2. The value may be a character string representing the name of a new phenotype, trait or covariate; however, THE FIRST CHARACTER MAY NOT BE A DIGIT. The name may not be that of an existing pedigree variable. Note that S.A.G.E. is CASE INSENSITIVE with respect to the names of traits, phenotypes and covariates. Thus the name "mean\_bmi" is considered identical to "mean\_BMI".
- 3. A missing value for any of the variables in a function expression will result in a missing value for that function variable.
- 4. The value of expression should be an algebraic expression referring to one or more existing variables (traits, phenotypes, covariates or markers, either read from the pedigree data file or previously created as function variables) as well as allowable operators, elementary functions and constants. The variable name used in an expression may be specified by any character string, BUT THE FIRST CHARACTER OF THE STRING MAY NOT BE A DIGIT. Expressions should always be enclosed in double quotes (""), and MUST BE ALL ON ONE LINE.

Examples

• Derive a new trait from an existing trait

```
function {
    # Create trait x from traits HDL and LDL
    trait = x, expression="log(HDL) - log(LDL)"
}
```

• Derive a new trait from an existing trait

```
function {
    # Create trait x from traits HDL and LDL
    trait = x, expression="log(HDL / LDL)"
}
```

The above two functions (1 and 2) are equivalent. Note also that if LDL is 0, this trait is undefined (and hence a missing value is assigned to it).

• A more complex example

```
function {
    # Creates, from variables h1 and h2, the covariate "average" whose
    # value is 1 if (h1 + h2) / 2 is greater than .275, and 0 otherwise.
    # If the program cannot evaluate an expression in 2 seconds or less,
    # it will abort, giving a fatal error message.
    time_limit=2
    constant=gamma, expression = .275
    covariate = "average", expression = "(h1 + h2)/2 > gamma"
}
```

- 5. Variable names are not case sensitive, but elementary function and constant names are.
- 6. The time\_limit parameter is provided to avoid situations where the calculation of values takes an inordinate amount of time. In most cases it need not be changed.

# 2.4.2 Expression Elements

This section describes the constants, operators, and functions that may be used in function block expressions.

## 2.4.2.1 Constants

The following constants may be used<sup>19</sup>:

<b>Constant Type</b>	Example
Any rational number	-3.0, 5, 1.23,
e	2.71828
pi	3.14159

<sup>19</sup>The two names e and pi are reserved and may not be used as the names of phenotypes, traits or covariates.

Operator	Meaning	Example
	Unary Negation	expression = "-BMI"
_	Unary Negation	expression = "-X + 10"
* *	Exponentiation	expression = "DBH**2" (power of two)
	Exponentiation	expression = " $X^{**0.5''}$ (square root)
00	Modulus	expression = "Age % 10"
0	(remainder after	
	integer division)	
,	-	expression = "Age / 10"
/	Division	
*	Multiplication	expression = "Weight * 1.19"
	Multiplication	expression = "Weight * (Affected == 1)" <sup>a</sup>
	Subtraction	expression = "Height - AvgHeight "
_	Subtraction	
+	Addition	expression = "Var_X + Var_Y"
т		
!=, <>	Not equal to	expression = "Affected <> 0"
, </td <td>Not equal to</td> <td>expression = "Sex != M"</td>	Not equal to	expression = "Sex != M"
==	Equal to	expression = "Affected == 0"
	( <i>two</i> equal	expression = "Sex == M"
	signs!)	
		expression = "Age >= 65"
>=	Greater than or	expression = "Age >= (65 - AgeOnset)"
	equal to	
>	Greater than	expression = "Age > 65"
		<pre>expression = "Age &gt; (65 - AgeOnset)" expression = "Age &lt;= 65"</pre>
<=	Less than or	expression = "Age <= (65 - AgeOnset)"
	equal to	expression - Age <- (05 - Ageonset)
<	Less than	expression = "Age < 65"
		expression = "Age < (65 - AgeOnset)"
not	Logical nega-	expression = "not(Affected)"
	tion	expression = "not(Sex == M)"
		expression = "(Affected and Sex == M)"
and	Logical AND	expression = "(Age > 65 and not (Affected == 1))"
	(intersection)	
or	Logical OR (union)	expression = "(Affected or Sex == M)"
		expression = "(x + (y * z))"
()	Parentheses	
	(logical &	
	arithmetic	
	grouping) <sup>b</sup>	

# 2.4.2.2 Operators and Expressions

<sup>&</sup>lt;sup>*a*</sup>Parentheses may be used to group terms in the normal manner. In this example the use of the comparison operator, "==", creates a logical expression whose evaluation results in either 0 or 1.

<sup>&</sup>lt;sup>b</sup>Strictly speaking, the parentheses are not operators so much as punctuation, a subtle difference that does not prevent their being listed in this table.

Operators are evaluated in order of operator precedence from highest to lowest in the following list. Except when there are parentheses (see below), all operators of an equal precedence are evaluated before operators of lower precedence (from left to right, except for comparison operators which are evaluated from right to left). Operator precedence from highest to lowest is as follows (operators on the same line have equal precedence):

Operator	Precedence Level
- (Unary Negation)	(Highest) 8
**	7
*, /, %	6
+, -	5
<, <=, >, >=, ==, <>, !=	4
not	3
and	2
or	(Lowest) 1

Operator precedence may be overridden by parentheses. Expressions in parentheses are evaluated first (BRACKETS OR BRACES MAY NOT BE USED). Multiple parentheses are permissible; the computation starts within the innermost parentheses and works outwards.

The logical expression operators *or*, *and* & *not* may be used only in expressions that evaluate to either zero (false) or one (true). Otherwise the results are not defined.

### 2.4.2.3 Elementary Functions

The following elementary functions may be used:

<b>Function Syntax</b>	Mathematical Equivalent	Meaning
exp(x)	$e^x$	e to the power of x
log(x)	$\ln x$	natural log of x
log10(x)	$\log_{10} x$	log to the base 10 of x
pow(x,y)	$x^y$	x raised to the power of y
sqrt(x)	$\sqrt{x}$ , or $x^{\frac{1}{2}}$	positive square root of x
fabs(s)		absolute value of x
ceil(x)		smallest integer $>= x$ , for any x
floor(x)		largest integer <= x, for any x
$\min(x_1, x_2,, x_n)$		$x_i$ such that $x_i \le x_j$ , for $j = 1, 2,, n$
$\max(x_1, x_2,, x_n)$		$x_i$ such that $x_i \ll x_j$ , for $j = 1, 2,, n$

### 2.4.2.4 Marker Functions

The following functions are available for codominant markers only (In these functions the second argument (allele value) must be in single quotes as shown.):

```
• dominant(marker, 'Ai') or dom(marker, 'Ai')
```

returns the value 1 or 0 based on the alleles present at the specified marker locus, where  $A^*$  is any allele other than  $A_i$ , as follows<sup>20</sup>:

 $A_i/A_i$ ,  $A_i/A^*$  returns 1  $A^*/A^*$  returns 0,

• recessive(marker, 'Ai') or rec(marker, 'Ai')

returns the value 1 or 0 based on the alleles present at the specified marker locus, where  $A^*$  is any allele other than  $A_i$ , as follows:

 $A_i/A_i$  returns 1  $A_i/A^*$ ,  $A^*/A^*$  returns 0,

• additive(marker, 'Ai') or add(marker, 'Ai')

returns the value 2, 1, or 0 based on the alleles present at the specified marker locus, where  $A^*$  is any allele other than  $A_i$ , as follows:

 $A_i/A_i$  returns 2  $A_i/A^*$  returns 1  $A^*/A^*$  returns 0,

• genotype(marker, 'Ai', 'Aj') or gen(marker, 'Ai', 'Aj')

returns the value 1 or 0 based on the alleles present at the specified marker locus, where  $A^*$  is any allele other than  $A_i$  or  $A_j$ , as follows:

 $A_i/A_j, A_j/A_i$  returns 1  $A_i/A^*, A_j/A^*, A^*/A_i, A^*/A_j, A^*/A^*$  returns 0,

#### 1. An ABO example:

```
function {
    # Creates, from marker ABO, the covariate x whose value is 1 if marker
    # ABO genotypes are AB or BA, and 0 otherwise.
    covariate = x,
    expression = "dom(ABO, 'A') and dom(ABO, 'B')"
}
```

2. Another ABO example:

```
function {
    # Creates, from marker ABO, the covariate x whose value is 1 if marker
    # ABO genotypes are AB or BA, and 0 otherwise.
    covariate = x,
    expression = "gen(ABO, 'A', 'B')"
}
```

Note The above two functions (from examples 1 and 2) are equivalent. Also, if ABO is missing, the trait x will also be missing.

3. Another marker example:

<sup>&</sup>lt;sup>20</sup>The examples assume that / is the allele delimiter within genotypes. However, a different delimiter could be used.

```
function {
    # Creates, from marker D42S8 and trait z, a phenotype, y, whose value
    # is z if allele q1 is present at marker D42S8, and 0 otherwise.
    phenotype = y,
    expression = "dominant(D42S8, 'q1') * z"
}
```

### 2.4.3 Mean-Adjusted and Variance-Adjusted Data

S.A.G.E. provides the option of generating mean-adjusted, variance-adjusted or standardized values for each class of a stratification variable of a given trait, phenotype or covariate. There are two basic steps to creating an adjusted variable:

- 1. Specify the classes of the stratification variable<sup>21</sup>.
- 2. Define a new variable to be adjusted with respect to these classes.

The newly created variable can then be used in a S.AG.E. analysis.

### 2.4.3.1 The Binning Algorithm

The method by which data and their associated summary statistics are subdivided into multiple *bins* (i.e., classes) is herein referred to as a *binning algorithm*, and the primary goal of this algorithm is to ensure that each class has a sufficient level of membership to support valid statistical inference. First, we define the following set of variables:

- N the minimum number of individuals required for each bin
- $N_j$  the number of individuals occurring in class j
- *k* the number of ordered classes
- $x_j$  the statistic of interest for the j-the class, *before* the binning algorithm has been applied
- $x'_{i}$  the statistic of interest for the j-the class, *after* the binning algorithm has been applied

Then the algorithm for allocating data observations across the bins is

- 1. If  $\sum_{i=1}^{k} N_i < N$ , then the algorithm cannot be performed.
- 2. For each class  $i \in \{1, 2, ..., j\}$  let  $a_i$  be a starting point and  $b_i$  be the end point of class i on the number line, where

$$a_i = \begin{array}{cc} 0, & if \ j = 0\\ b_{i-1}, & if \ j > 0 \end{array}$$
 and  $b_i = a_j + N_i/N.$ 

<sup>&</sup>lt;sup>21</sup>A stratification variable is not strictly required, and the data adjustment procedure described here can be used to transform any given trait, phenotype or covariate into its mean- or variance-adjusted values, as explained in Sec. 2.4.3.6.

3. For any class *j* for which  $N_j < N$ , let *l* be the proportion of  $x'_j$  that comes from  $x_1, ..., x_{j-1}$ , where  $l = max \left[ min \left( \frac{N-N_j}{2N}, a_j \right), \frac{N-N_j}{2N} - min \left( \frac{N-N_j}{2N}, n_k - b_j \right) \right]$ . Then let  $A = a_i - l$  and  $x_j = \sum_{i=1}^k a_i x_i$ , where:  $a_i = \begin{cases} 0, if b_i < A \text{ or } a_i \ge A + 1 \\ b_i - A, if a_i < A \text{ and } b_i > A \\ N_i/N, if a_i \ge A \text{ and } b_i \le A + 1 \\ A + 1 - a_i, if a_i \ge A \text{ and } b_i > A + 1 \end{cases}$ 

#### 2.4.3.2 Specifying the Classes

Specify each class within a function block as the values of an expression attribute for the covariate parameter<sup>22</sup>. The following example shows how to create three classes of a covariate<sup>23</sup> named "Age":

```
function {
   covariate = class1, expression = "(Age <= 15)"
}
function {
   covariate = class2, expression = "(Age > 15 and Age <= 30)"
}
function {
   covariate = class3, expression = "(Age > 30)"
}
```

**IMPORTANT!** It is essential that the classification scheme partitions the data into exhaustive and mutually exclusive subsets with respect to the classification variable ("Age", in this case). If the data are not partitioned correctly, the resultant mean- and variance-adjusted variables will not be reliable.

#### 2.4.3.3 Creating a Mean-Adjusted Variable

Once the classes of the stratification variable have been defined, the mean-adjusted values of some other trait (or phenotype or covariate) can be calculated using the classes of the stratification variable to define classes of the trait. Specify the mean-adjustment within a function block as the value of an expression attribute for the trait parameter<sup>24</sup>. Assuming the pedigree data file lists a phenotype called "BP", the following example shows how to create a new mean-adjusted variable named "BP\_AgeAdjMean":

<sup>&</sup>lt;sup>22</sup>The classes could also be created from a trait or phenotype; however, it usually makes more sense to create them from a covariate.

<sup>&</sup>lt;sup>23</sup>In this example, "Age" is assumed to be a covariate; however the stratification variable could also be a trait or phenotype.

<sup>&</sup>lt;sup>24</sup>The variable could also be created as a phenotype or covariate.

```
function {
   trait = BP_AgeAdjMean,
   expression = "mean_adj(Age, BP, 10, class1, class2, class3)"
}
```

Note the use of the special keyword, **mean\_adj**. This is what tells S.A.G.E. to add a new set of information to the internal computer representation of the pedigree file.

The value of the third argument to **mean\_adj** (10 in this example) determines the minimum number of items required for the classes. If, after the data have been stratified, any of the resultant classes has less than the minimum number of entries, then a special algorithm is employed to "borrow" values from the ordered list of values in neighboring classes until the minimum number has been reached for the underrepresented class. Note that for the resulting mean-adjusted variable to be meaningful, the classes of the stratification variable must be in natural order. The user can disable this feature by entering a zero (0) as the minimum number.

The variable, *BP\_AgeAdjMean*, essentially becomes a new trait, and is surreptitiously added to the internal representation of the pedigree data file (the original file is left unchanged). In this case, there are three different means computed:

- $\bar{x}_1$ : the mean BP for individuals whose age is in the range  $\{0, 1, 2, ..., 15\}$ ,
- $\bar{x}_2$ : the mean BP for individuals whose age is in the range {16, 17, ..., 30} and
- $\bar{x}_3$ : the mean BP for individuals whose age is in the range {31, 32, 33, ...}.

If  $BP_i$  is the blood pressure value for individual *i*, then the value of "BP\_AgeAdjMean" for that individual will be

- $BP_i \bar{x}_1$ ; if the individual's age is in the range  $\{0, 1, 2, ..., 15\}$ ,
- $BP_i \bar{x}_2$ ; if the individual's age is in the range {16, 17, ..., 30} and
- $BP_i \bar{x}_3$ ; if the individual's age is in the range {31, 32, 33, ...}.

### 2.4.3.4 Creating a Variance-Adjusted Variable

The procedure for creating a variance-adjusted variable is analogous. The following example shows how to create a new variable named "BP\_AgeAdjVar":

```
function {
   trait = BP_AgeAdjVar,
   expression = "var_adj(Age, BP, 10, class1, class2, class3)"
}
```

Here, the required keyword is **var\_adj**, and the resultant values of BP\_AgeAdjVar (for an arbitrary individual *i*) will be:

- $BP_i/s_1$ : if the individual's age is in the range {0, 1, 2, ..., 15},
- $BP_i/s_2$ : if the individual's age is in the range {16, 17, ..., 30} and
- $BP_i/s_3$ : if the individual's age is in the range {31, 32, 33, ...},

where  $s_i$  (*i* = 1, 2, 3) is the sample standard deviation of the trait *BP* for age class *i*.

#### 2.4.3.5 Creating a Z-Score Variable

A standardized variable is obtained as follows:

```
function {
   trait = BP_AgeZScore
   expression = "z_score(Age, BP, 10, class1, class2, class3)"
}
```

In this last example, the required keyword is  $z\_score$ , and the values of BP\_AgeZScore (for an arbitrary individual *i*) will be:

- $(BP_i \bar{x}_1)/s_1$ : if the individual's age is in the range {0, 1, 2, ..., 15},
- $(BP_i \bar{x}_2)/s_2$ : if the individual's age is in the range {16, 17, ..., 30} and
- $(BP_i \bar{x}_3)/s_3$ : if the individual's age is in the range {31, 32, 33, ...}.

### 2.4.3.6 Creating Adjusted Variables Without Classes

It is also possible to create an adjusted variable that does not depend on the classes of a stratification variable. The result is simply the mean-adjusted, variance-adjusted or standardized value of a given variable with respect to the entire sample.

To create a mean-adjusted variable (BP\_AdjMean) from the variable BP, write:

```
function {
   trait = BP_AdjMean,
   expression = "mean_adj(BP)"
}
```

To create a variance-adjusted variable (BP\_AdjVar) from the variable BP, write:

```
function {
   trait = BP_AdjVar,
   expression = "var_adj(BP)"
}
```

To create a standardized variable (BP\_Normalized) from the variable BP, write:

```
function {
   trait = BP_Normalized,
   expression = "z_score(BP)"
}
```

## 2.4.4 Data Trimming and Winsorization

S.A.G.E. provides a way to minimize the adverse impact of outlier data by creating variables that either *trim* or *Winsorize* the tails of the distributions as shown in the following example:

	1212121	2	5 		i nemener i	0.10	1	1 (1) (1) (1) (1) (1) (1) (1) (1) (1) (1)	12332273
--	---------	---	-------	--	-------------	------	---	--	----------

After so	rting and	positioni	ng:						
0.10	0.12	0.13	0.15	0.19	0.27	0.38	0.51	0.58	0.86
1	2	3	4	5	6	7	8	9	10

After trimming:

19 I J	12 1. Contra	0.13	0.15	0.19	0.27	0.38	0.51	<u>a</u> s [	32
1	2	3	4	5	6	7	8	9	10

1 0	TTT	9.3	
After	W/ Inc.	0117:	ation
TTTPOT	AA 111.9.	01121	abivit.

0.13	0.13	0.13	0.15	0.19	0.27	0.38	0.51	0.51	0.51
1	2	3	4	5	6	7	8	9	10

Data that are subjected to the trim function are effectively thrown out of the analysis, whereas Winsorized data are revalued to a quantity that corresponds to some critical point along the distribution. Details of the method can be found in Armitage & Colton (1990).

### 2.4.4.1 Creating a trimmed variable

Create a trimmed variable using the **trim** S.A.G.E. keyword as in the following example:

function{trait=LNIGE\_trim, expression="trim(LNIGE,0.02)"}

The **trim** function takes two arguments:

- 1. the name of a trait, phenotype or covariate (here "LNIGE") previously specified in the pedigree block
- 2. a value  $\gamma \in (0, 1)$ , representing the "amount" of data to be trimmed (the value 0.02 will result in trimming 1% of the values in each tail of the distribution).

The newly created variable, *LNIGE\_trim*, can be used in the same manner as any other trait, covariate or phenotype within S.A.G.E. applications.

#### 2.4.4.2 Creating a Winsorized variable

Create a winsorized variable using the **winsor** S.A.G.E. keyword as in the following example:

function{trait=LNIGE\_wins, expression="winsor(LNIGE,0.02)"}

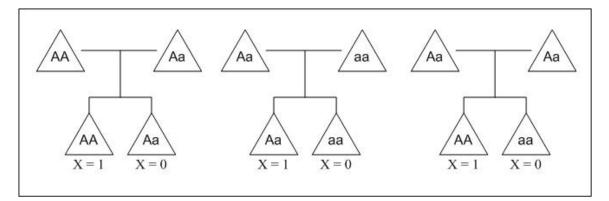
The **winsor** function takes two arguments:

- 1. the name of a trait, phenotype or covariate (here "LNIGE") previously specified in the pedigree block
- 2. a value  $\gamma \in (0, 1)$ , representing the "amount" of data to be winsorized (the value 0.02 will result in 1% of the values in each tail of the distribution being replaced by the corresponding 1 and 99 percentiles).

The newly created indicator variable, *LNIGE\_wins*, can be used in the same manner as any other trait, covariate or phenotype within S.A.G.E. applications.

#### 2.4.5 The Transmitted and Untransmitted Allele Indicators (TAI and UTAI)

The problem of performing a transmission disequilibrium test (TDT) to assess the linkage between a marker locus and a quantitative trait was addressed in a paper by George et al (1999), who proposed a linear-regression approach in which the disease trait (assumed to be continuous) is the dependent variable, *Y*. The primary independent variable in the model, *X*, is an indicator variable that reflects whether or not a given allele was transmitted to the individual from a heterozygous parent (see Figure 1). The authors refer to X as a *transmission status variable* which is referred to here by the slightly more accurate term: *transmitted allele indicator* (TAI).



**Figure 1**: Offspring who are informative for linkage, from relevant parental matings. *A* is the associated allele of interest, and *X* is the transmitted allele indicator variable such that X=1 if *A* was transmitted from a heterozygous parent, and X=0 otherwise.

For example, consider a diallelic locus  $\{A, a\}$  and suppose we wish to determine the TAI with respect to allele 'A'. Then the TAI values computed for a given individual would be as shown in the table below, which also indicates the UTAI as well<sup>25</sup>:

<sup>&</sup>lt;sup>25</sup>When the marker locus has more than two alleles, we appropriately extend this indicator to make use of the maximum amount of information available in an unbiased fashion. See the theory section of the TDTEX program in this manual.

	Parer Genot		Offspring Genotype	Informative?	TAI	UTAI Value
				Informacive:	varue	varue
1	AA x	AA	AA	Ν		
2	AA x	Aa	AA	Y	1	0
3	AA x	Aa	Aa	Y	0	1
4	AA x	aa	Аа	Ν		
5	Aa x	Aa	AA	Y	1	0
6	Аа х	Aa	Аа	Ν		
7	Aa x	Aa	aa	Y	0	1
8	Aa x	aa	Aa	Y	1	0
9	Aa x	aa	aa	Y	0	1
10	аа х	aa	aa	Ν		

The following table illustrates the way the TAI variable is created in S.A.G.E. (the table for UTAI variables is similar). The TAI user-defined function generates an internal table, similar to the one depicted below, that associates the value of the indicator status variable with respect to each individual in the pedigree, and for each allele. TAI variables for pedigree founders are simply interpreted as missing information.

PID	ID	M <sub>1</sub> A <sub>1(1)</sub>	M <sub>1</sub> A <sub>2(1)</sub>		M <sub>1</sub> A <sub>k(1)</sub>	M <sub>2</sub> A <sub>1(2)</sub>	M <sub>2</sub> A <sub>2(2)</sub>	0220	M2Ak(2)		M <sub>n</sub> A <sub>1(n)</sub>	M <sub>n</sub> A <sub>2(n)</sub>	2220	M <sub>n</sub> A <sub>k(r</sub>
001	001													
001	002			390				2020					2020	
001	003	0	0	394	1	0	0	2022	1		0	0	2332	0
001	004	1	0	3999	0	0	1	2352	0		0	1	-	1
001	005	0	0	30	0	1	0	3850	0	300	1	0	3332	1
002	001				()				0					0
002	002							0220						
002	003	0	0		1	1	0	0220	0		0	0	aa	0
002	004	1	0		0	0	1	0220	0		0	1	22.27	1
002	005	0	1		0	1	0	1220	1	an.	1	1	12.02	0
003	001							-					-	
003	002			300				2352					2000	
003	003	0	0	300	1	0	0	2022	0		0	0	2342	0
003	004	1	0		0	0	0	3352	1		0	0	1000	0
003	005	1	0		0	1	0	2022	0		1	0	22.0	1

#### 2.4.5.1 Creating TAI and UTAI Variables

To specify TAI and/or UTAI variables for a single marker, create a function block that defines the new variables using the **tai** and **utai** keywords as in the following example:

```
pedigree{
   .
   .
   .
   allele = "M1A", name = "M1" #marker specified in pedigree block
```

```
allele = "M1a", name = "M1" #marker specified in pedigree block
}
function{
   trait = M1A_tai, expression="tai(M1, A)"
}
function{
   trait = M1A_utai, expression="tai(M1, A)"
}
function{
   trait = M1A_utai, expression="utai(M1, A)"
}
function{
   trait = M1a_utai, expression="utai(M1, a)"
}
```

The newly created indicator variables, *M1A\_tai*, *M1a\_tai*, *M1A\_utai* and *M1a\_utai*, can be used in the same manner as any other trait, covariate or phenotype within S.A.G.E. applications<sup>26</sup>.

### **2.5 Locus Description Files**

The marker locus description file and the trait locus description file follow the same format as each other and contain records that define allele frequencies and phenotype to genotype mappings. The marker locus description file contains a record for each marker locus. The trait locus description file contains a record for each discrete trait, or "trait-marker", that is to undergo a model-based linkage analysis. A record must be included in the corresponding locus description file for each marker locus or trait-marker to be analyzed, and these records may appear in any order. All marker loci and trait-markers listed in the parameter file and/or the genome description file should be present in the marker locus description file: THOSE NOT PRESENT THERE ARE IGNORED. In the case of fully penetrant and codominant markers, the program FREQ can be used to prepare the marker locus description file.

Optionally, at the beginning of the file, a missing value code may be included. This code indicates which values, if any, indicate a missing marker value. For example, the user may specify one of the following lines as the first line of the marker locus description file:

```
Missing=CODE
missing=CODE
```

or

```
MISSING=CODE
```

<sup>&</sup>lt;sup>26</sup>S.A.G.E. currently requires the user to create these TAI and UTAI variables one at a time via the function block syntax as shown in the example. This would be an admittedly impractical solution when analyzing large numbers of markers (SNPs, for example), and the S.A.G.E. development team is working to provide shortcuts and improvements to this feature in a future release.

where CODE is the case-sensitive missing value code for ALL marker phenotypes or pairs of marker alleles. If, when identifying a marker phenotype by its two alleles, either allele is missing, the phenotype of the individual will be set to missing for that marker. If only one allele is missing, a non-fatal error message will be generated.

The locus description file should contain the following items for each locus to be analyzed:

- 1. The name of the locus.
- 2. A set of records that give the allele frequencies. The records should follow this format:

#### allele\_symbol = population allele frequency

The user should supply the information for the items on both sides of the "=" symbol. There can be any number of spaces before or after the equal sign as long as the allele symbol and its frequency remain on the same line. There should be only one allele symbol per line. The allele symbol can consist of up to 10 characters. It is also permissible to list just the alleles, leaving out " = allele frequency" from every line. When this is done, equal allele frequencies are substituted for each allele listed for that locus<sup>27</sup>. This option is useful when the marker locus description file is used in conjunction with a program that does not use allele frequencies (e.g., ASSOC).

- 3. A semicolon indicating the end of the alleles. This semicolon can be either on a line by itself or on the same line following the last population allele frequency of the set.
- 4. A set of records that defines the phenosets (i.e., the sets of genotypes compatible with each phenotype). The records should follow this format:

phenotype symbol =  $\{A1_1/A2_1[=P_1],...,A1_m/A2_m[=P_m]\}$  where

 $A1_1$  is the symbol for allele #1 in the first genotype of this phenoset;

A2<sub>1</sub> is the symbol for allele #2 in the first genotype of this phenoset;

••

 $A1_m$  is the symbol for allele #1 in the m-th (last) genotype of this phenoset;

 $A2_m$  is the symbol for allele #2 in the m-th (last) genotype of this phenoset,

and  $P_1...P_m$  are the penetrance values of the phenotype, i.e., the probabilities of the phenotype given the genotype. The penetrance values are strictly optional<sup>28</sup>.

There can be any number of spaces before or after the equal sign(s). The phenoset should begin with either a left curly brace ({) or less-than symbol (<), and end with a corresponding right curly brace (}) or greater-than symbol (>). The first and second allele of each genotype must be separated by a slash (/) or otherwise specified allele delimiter. Each genotype within the phenoset should be separated by a comma. This record may wrap onto as many lines as necessary. Complete the set by repeating this record for each phenotype at this locus. Any phenotype symbol that is not included here is interpreted as a missing phenotype value. The order of the alleles in a genotype has no effect.

Example:

<sup>&</sup>lt;sup>27</sup>This is a change from previous versions of S.A.G.E.

<sup>&</sup>lt;sup>28</sup>This is a change from previous versions of S.A.G.E. If no value is indicated, the phenotype is assumed to be fully penetrant and a value of 1 is assumed.

```
LOCA
A = 0.5
                           #(alleles/phenotype names are arbitrary
B = 0.25
                           #and need not be sequential)
C = 0.25
1 = \{A/A, A/B, A/C\}
                         #(A is dominant over B and C, and
2 = \{B/B, B/C\}
                           #B is dominant over C)
3 = \{C/C\}
ABO
A1 = 0.1904
A2 = 0.0612
B = 0.0728
0 = 0.6756
1 = { A1/A1, A1/A2, A1/O } \#(A1 is dominant over A2 and O)
2 = \{ A1/B \}
3 = \{ A2/A2, A2/O \}
                          #(A2 is dominant over O)
4 = \{ A2/B \}
5 = \{ B/B, B/O \}
                          #(B is dominant over O)
6 = \{ 0/0 \}
;
```

If a locus is fully penetrant and codominant it is not necessary to include the records for phenotypes mentioned in note 4 above. The program will generate the phenotype symbol by concatenating the two allele symbols of the genotype and putting a delimiter character between them (typically a /, but this can be modified in the parameter file). However, the semicolon indicating the end of the phenotypes still has to be included.

#### Example:

The following two locus descriptions are equivalent:

A	А		
1 = 0.645	1	=	0.6455
2 = 0.223	2	=	0.2230
3 = 0.1325;	3	=	0.1325;
$1/1 = \{1/1\}$	;		
$1/2 = \{1/2\}$			
$2/2 = \{2/2\}$			
$1/3 = \{1/3\}$			
$2/3 = \{2/3\}$			
$3/3 = \{3/3\}$			
i			

Trait-markers are specified similarly. As an example, suppose we have a trait "Disease", and an underlying model with two disease alleles (allele 1 has frequency 10% and allele 2 has frequency 90%) and two phenotypes (A = affected, U = unaffected). Suppose that we are assuming that allele 1 predisposes toward the expression of affection, and furthermore that it is recessive to allele 2.

Our penetrance table might look something like this:

	1/1	1/2	2/2
Α	0.6	0.01	0.01
U	0.4	0.99	0.99

i.e., 60% penetrance and a sporadic rate of 1%. The trait locus description file would then contain the following entry:

Disease 1 = 0.10 2 = 0.90 ; A = { 1/1 = 0.6, 1/2 = 0.01, 2/2 = 0.01 } U = { 1/1 = 0.4, 1/2 = 0.99, 2/2 = 0.99 } ;

Note that the trait need not be binary (any number of phenotypes may be specified), and the locus may have more than two alleles. For any particular phenotype, the sum of all (here two) penetrances must equal 1.

#### 2.6 Genome Description File

The genome description file describes the genomic region(s) used in analyses that require the order of, and distances between, linked marker loci. A genome is defined with at least one genomic region. This region contains the names of sequentially ordered marker loci and the distances or recombination fractions between pairs of adjacent markers. A map function is used to translate genetic map distances to and from recombination fractions. The general form of the file is as follows:

```
genome="genome name"[,map="map function"]
{
    [region1]
    [region2]
    [region3]
    .
    .
}
```

The genome name can be any name desired. The map attribute allows specification of a map function, which can be either the Haldane or Kosambi map functions. If no map function is supplied, Haldane is assumed. Map functions are not used during single-point analysis.

Each genomic region is described as follows:

```
region="region name"
{
    [marker and distance parameters]
}
```

The region name is used to identify the region being defined. If no name is specified, "region n" is used, where n is the number of the region within the genome. The attribute x\_linked is needed after region name to indicate the region as X-linked region as follows:

```
region="region name", x_linked
{
    [marker and distance parameters]
}
```

<pre>parameter [, attribute]</pre>	Explanation					
	Indicates a marker name. If none is specified, the					
	marker is ignored. There should be one marker pa-					
	rameter for each marker in the region.					
marker	Value Range Character string					
	Default Value None					
	Required Yes					
	Applicable Notes 1					
	Indicates a marker that is missing from the data, but is					
	included as a placeholder.					
missing	Value Range					
IIISSIIIG	Default Value None					
	Required No					
	Applicable Notes None					
	Specifies distance, in centimorgans, between adjacent					
	marker parameters.					
distance	Value Range $(0,\infty)$					
distance	Default Value None					
	Required Yes					
	Applicable Notes 2					
	Specifies distance between adjacent markers in terms					
	of the recombination fraction $\theta$ .					
theta	Value Range [0, 0.5]					
	Default Value None					
	Required Yes					
	Applicable Notes 2					

The following parameters are available within a region sub-block:

Notes

- 1. In the program output the first marker in each region is located at an absolute distance of 0.0 cM and all further markers are measured from this location in the map units specified by the map attribute.
- 2. There is a maximum of one distance or theta (i.e., recombination fraction) parameter between each pair of markers (marker and missing parameters.) When doing multi-point analysis, there must be either a distance or theta parameter between each pair of adjacent markers.

Here is an example of a typical Genome Description file:

	marker	=	"D4S2976"	#	at	155.1300	сМ
	distance	=	0.320000000				
	marker	=	"D4S2631"	#	at	155.4500	сМ
	distance	=	0.170000000				
	marker	=	"D4S3016"	#	at	155.6200	сМ
	distance	=	0.700000000				
	marker	=	"D4S1556"	#	at	156.3200	сМ
	distance	=	1.230000000				
	marker	=	"TSC0785934"	#	at	157.5500	сМ
	distance	=	0.000000001				
	marker	=	"TSC1312016"	#	at	157.5500	сМ
	distance	=	0.000000001				
	marker	=	"TSC0439917"	#	at	157.5500	сМ
	•						
	}						
ł	ſ						

# 2.7 IBD Sharing File

}

The IBD sharing file stores the probability distribution of allele-sharing identical-by-descent (IBD) between pairs of individuals at specific locations. The header of the file contains the n names (L1, L2, , Ln) of the locations at which IBD sharing information is stored for each pair of relatives. These locations are referred to as markers, even though they may not correspond to observed marker loci in a given dataset. The body of the file contains a line for each pair of individuals that includes the following fields:

- pedigree ID
- First individual ID
- Second individual ID
- $f_0$ : The probability that the pair shares 0 alleles IBD at marker  $L_1$
- $f_{1m-1p}$ : The probability that the pair shares 1 maternal allele minus the probability that it shares 1 paternal allele IBD at marker  $L_1$
- f<sub>2</sub>: The probability that the pair shares 2 alleles IBD at marker L<sub>1</sub>
- . . .
- f<sub>0</sub>: The probability that the pair shares 0 alleles IBD at marker L<sub>n</sub>
- $f_{1m-1p}$ : The probability that the pair shares 1 maternal allele minus the probability that it shares 1 paternal allele IBD at marker  $L_n$
- f<sub>2</sub>: The probability that the pair shares 2 alleles IBD at marker L<sub>n</sub>

The probability that a pair shares one allele IBD at a given marker is  $f_1 = 1 - f_0 - f_2$ , where  $f_0$  and  $f_2$  are the probabilities that the given pair shares 0 and 2 alleles IBD at the marker. Similarly, the estimated proportion of alleles shared IBD is  $f_2 + \frac{1}{2}f_1$ . These probabilities are conditional on the pedigree and marker information available and are usually denoted  $\hat{f}$  in the literature.

Notes

- 1. IBD sharing files are typically generated as a result of some prior analysis and will virtually never need to be constructed manually. IBD sharing files are generated by the program GENIBD and used as input to programs such as SIBPAL.
- 2. Packages other than S.A.G.E. may be able to use IBD sharing files as input, but the format in S.A.G.E. is subject to change.
- 3. The number of markers may be very large, so each line of the IBD sharing file can be extremely long. Loading these files into text-editors, especially those that wrap or truncate long lines, is not recommended.
- 4. IBD sharing files may be extremely large if there are many pairs and markers. When performing analyses on extremely large pedigrees and/or genome screens, IBD sharing files may consume disk space in excess of a gigabyte. Thankfully, IBD sharing files are amenable to many forms of data compression when not in use.

# 2.8 Information Output Files

An information output file is generated by all S.A.G.E. programs and contains diagnostic output generated during program execution. Typically, this includes information about how pedigree data files were read and diagnostic information on pedigree structure, phenotypes and marker loci. This file is named "program.inf", indicating the name of the specific program that was run<sup>29</sup>. No analysis results are stored in this file.

All S.A.G.E. programs that read trait or marker locus description files or genome description files generate the genome information File. This file contains diagnostic information on each marker or trait phenotype and genotype. This file is named "genome.inf". No analysis results are stored in this file, though errors relating to the markers and traits may be.

# 2.9 Analysis Output Files

All S.A.G.E. programs produce one or more analysis output files, which contain the results of the analyses. The number of analysis output files, their names and contents are program specific. Analysis output files may even correspond to other S.A.G.E. input file types. E.g., the analysis output file from GENIBD is an IBD sharing file that is an input file for SIBPAL.

<sup>&</sup>lt;sup>29</sup>Eg., fcor.inf, mlod.inf, segreg.inf, etc.

# Chapter 3

# **PEDINFO**

PEDINFO provides many useful descriptive statistics on pedigree structure including means, variances and histograms of family, sibship and pedigree sizes, and counts of each type of relative pair. Statistics based on trait phenotypic status (i.e., limited to traits not having missing values) are can also be requested.

# 3.1 Limitations

PEDINFO cannot correctly process a pedigree that contains loops; however, the program does indicate the presence of loops within the given pedigree data file.

# 3.2 Theory

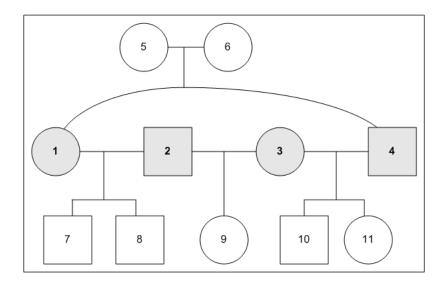
#### 3.2.1 Terminology

PEDINFO operates by iterating over the pedigree structures and keeps counts and distribution information of various elements. The following table defines some terms used in PEDINFO that are not defined elsewhere in this document:

Term	Definition
Brother Pair	A pair of individuals who share the same parents and are both
	male.
Sister Pair	A pair of individuals who share the same parents and are both fe-
	male.
Generations	In a pedigree without loops the number of generations is one more
	than the length of the longest chain of offspring relationships.
Inheritance Vector Bits	For a given pedigree, the maximum value of 2n-f maximized over
	its constituent pedigrees, where n is the number of non-founders
	and f is the number of founders in each constituent pedigree. This
	number represents the largest number of bits in an inheritance vec-
	tor that would be used in certain types of multi-point analysis al-
	gorithms. It is useful for evaluating whether it is feasible to run
	such algorithms on a given pedigree.

#### 3.2.2 Problematic Family Structures

A *marriage ring* is a chain of spouses who form a cycle, for example, the founders in the following pedigree (individuals #1, 2, 3 and 4):

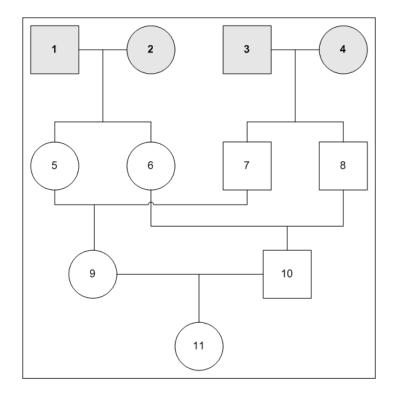


Individuals in a chain of spouses are listed in the output as individuals with multiple mates. Individuals in marriage ring are listed in both the individual column and the mates column.

	Individuals With Multiple Mates	
(Pedigree, Individual)		
(1, 1) (1, 2)	4, 2 1, 3	
(1, 3) (1, 4)	4, 2 3, 1	
=======================================		

These rings can cause computational difficulties for current programs using full pedigree structure information and are therefore enumerated so that users can break these rings as they see fit.

*Loops* indicate either consanguineous (marriage between relatives) or other marriage loops, eg., two brothers married to two sisters:



Consanguineous and other marriage loops can also cause computational difficulties for current programs using full pedigree structure information and may also need to be broken. To facilitate this process, consanguineous matings are listed by pedigree and by the pair of relatives who have mated.

======================================	Consanguineous Mating Pairs	
  Pedigree	Pair	
  1	10, 9	

When there are marriage rings or loops in the pedigree, the pairs are not independent and therefore the pair counts output by PEDINFO may not be accurate.

In the case of a consanguineous pedigree, the number of generations may be indeterminable and "undet" will appear in the generation statistics output by PEDINFO.

======================================	================		=========			=====	========
Generation Sta	tistics    Nu	ic Famil	ly Statis	stics	Inh Vecto	r Bit	Stats
# of Gens   #	of Peds    # c	of Nuc Fa	ams  # of	E Peds	# of Bits	#	of Peds
undet.	5	0 –	2	1	3 -	4	2
		3 -	4	4	5 -	8	3
======================================	=======================================		=========			=====	=======

Breaking loops can be done by duplicating individuals or by removing certain connecting individuals.

# 3.3 Program Input

File Type	Description
PEDINFO parameter file	Specifies the parameters and options with which
	to perform a particular analysis.
Pedigree data file	Contains delimited records for each individual in-
	cluding fields for identifiers, sex, parents, trait and
	marker data.

## 3.3.1 The pedinfo Parameter

The following syntax table specifies the permissible parameter and attribute settings for the main
PEDINFO parameter.

<pre>parameter [, attribute]</pre>		Explanation
	Starts a PEDINFO	analysis block.
pedinfo	Value Range	N/A
	Default Value	None
	Required	No
	Applicable Notes	None
	Specifies the root	name to be used for output files.
	Output file names v	will be formed by concatenating the
	root name and an a	ppropriate extension.
	Value Benge	Character string representing a
, out	Value Range	valid file name
	Default Value	pedinfo.out
	Required	No
	Applicable Notes	None

### 3.3.2 The pedinfo Sub-Block

The following syntax table specifies the permissible parameter and attribute settings for the pedinfo sub-block.

<pre>parameter [, attribute]</pre>	Explanation		
	Specifies a variable to be used in the analysis.		
		Character string representing the name of a trait, phenotype or	
trait	Value Range	covariate listed in the pedigree	
trait phenotype covariate		data file.	
	Default Value	None	
covariate	Required	No	
	Applicable Notes	1	
	Specifies option to by-pedigree basis	calculate statistics on a pedigree-	
	Value Range	{true, false}	
each_pedigree	Default Value	false	
	Required	No	
	Applicable Notes	2	
	Specifies suppressi	on of output for non-trait statistics.	
	Value Range	{true, false}	
suppress_general	Default Value	false	
	Required	No	
	Applicable Notes	None	

Notes

- 1. The trait, phenotype and covariate parameters are used to specify trait, covariate, or phenotype variables for which statistics are to be calculated. The value of a variable parameter should be set to the name of a variable field read from the pedigree data file or created using a function statement. To be included in the statistics an individual must not have a missing value for this variable. More than one variable may be specified in an analysis block, in which case an individual must have non-missing values for each of the specified variables to be included in the statistics. If a single binary variable is specified for analysis, counts of pairs that are concordant unaffected, discordant, concordant affected and uninformative will be displayed. If no variables are specified, only non-variable information (i.e., based on pedigree structure alone) will be used to determine counts.
- 2. The each\_pedigree parameter is used to specify whether results should be calculated for each pedigree separately in addition to a set of results for all the pedigrees taken as a whole.

The following are all valid pedinfo statements and could all occur within the same parameter file:

```
# A pedinfo statement that runs with all default values
pedinfo
# A pedinfo statement that runs with all default values
pedinfo
{
}
# A pedinfo statement that specifies the name of an output file
# and requests a separate report for each pedigree
pedinfo,out=allpeds.out
{
   each pedigree=true
}
# A pedinfo statement that specifies 2 traits,
# for each of which an individual must have no missing data
# to be included in the trait-specific pedigree statistics
pedinfo,out=analysis1 {
   trait=A
   trait=hematocrit
}
# A final example
pedinfo,out=output {
   phenotype=B
   each_pedigree=true
}
```

# **3.4 Program Execution**

PEDINFO is run via a command line interface on the supported UNIX and Windows platforms. This requires that the S.A.G.E. programs are properly installed and in the current execution path. Input files are specified on the command line and all output files are created in the current working directory.

Running PEDINFO from the command prompt with no arguments, or the wrong number of arguments, will result in the program printing its usage statement. This lists the input files the program requires on the command line:

```
>pedinfo
S.A.G.E. v5.x -- PEDINFO
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
usage: ./pedinfo <parameters> <pedigree>
Command line parameters:
parameters - parameter file
pedigree - pedigree data file
```

As indicated in the program usage statement, input files are listed on the command line. A typical run of PEDINFO may look like the following:

```
>pedinfo pedinfo.par example.ped
```

S.A.G.E. v5.x -- PEDINFO COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY. Reading parameter file.....done. Reading pedigree file.....done. from **example.ped**.....done. Sorting pedigrees.....done. Generating statistics.....done. Analysis complete!

# 3.5 Program Output

Filename	File Type	Description
pedinfo.inf	Information output file	Contains informational diagnostic messages, warnings and program er- rors. No analysis results are stored in this file.
pedinfo.out	Analysis output file	Contains a table of summary statis- tics for all pedigrees combined, and optionally a table for each individual pedigree.

Output files produced by PEDINFO containing results and diagnostic information are:

#### 3.5.1 Information Output File

The PEDINFO information output file contains a variety of useful information, including:

- Information on fields read from the pedigree data file. These tables, which provide information about what the program has read in, are included with all programs in S.A.G.E. and are very useful for debugging many common errors caused when reading the pedigree data. When first analyzing new data, it is recommended that these tables be carefully checked to make sure pedigree data are being correctly read.
- Information, warning and error messages generated throughout the program. It is recommended that you check this file for warning and error messages before examining the results of any run of the program. The program attempts to correct many common errors and this sometimes means analyses are not run as expected.

#### 3.5.2 Analysis Output File

The PEDINFO analysis output file may contain the following types of tables:

- Tables of statistics pertaining to the structure of all of the data as a whole.
- Tables of statistics pertaining to the structure of a single pedigree.
- Tables of statistics pertaining to a specific trait or set of traits for the data as a whole. For individuals to be counted in these tables they must be informative (not have a missing value) for each relevant trait<sup>1</sup>.
- Tables of statistics pertaining to a specific trait or set of traits for a single pedigree. For individuals to be counted in these tables they must be informative (not have a missing value) for each relevant trait.

<sup>&</sup>lt;sup>1</sup>In the counts of nuclear families per pedigree included in these tables, a nuclear family is defined as a family having at least one informative parent and one informative child.

# 3.6 Example Output File

Here are some typical examples of PEDINFO output:

	Count	Mean Size	+/- Std. Dev.	( Min.	, Max.)
Pedigrees	2	6.50	+/- 2.50	( 4	, 9)
Generation Stati # of Gens   # of					
2	2   	0 - 2  3 - 4	1   1	0 -	2
	Count	Mean Size	+/- Std. Dev.	( Min.	, Max.)
Sibships	4	1.25	+/- 0.43	( 1	, 2)
Constituent   Pedigrees		Marriage Rings	   0	   Loops	
Pairs		Count	Individuals		Cour
Parent/Off Sib/Sib Sis/Sis Bro/Bro Bro/Sis Grandp. Avunc. Half Sib Cousin		10   1   0   1   0   0   0   0	Male Female Unknown Total Founder Non-founder Singleton Total		
======================================		duals With D	Multiple Mate		
		anguineous 1	=================== Mating Pairs	==========	

=======================================		stics: All Pe	5			
	O Pa	rents w. Data	u  1 Parer	nt w. Data	2 Parents	w. Data
Sibships		(	)	0		4
	C	ount  Mean Si	.ze +/- Std.	Dev. (	Min.,	Max.)
Sibships		4  1.	00 +/-	(	1,	1)
Pedigrees			00 +/-	3.00 (	3,	9)
	ly Statistic ams  # of Pe					
0 - 3 -	2   4	1 1				
	======================================	Pair			Pain	 r
Pairs	Count	Correlatio	on  Pairs	Cour	nt Cori	relation
Parent/Off Sib/Sib Sis/Sis Bro/Bro Bro/Sis	8    0    0    0		Avunc.   Half Sik		0   0   0   0   0	  
Indiv.'s	Count	Mean +/- S	Std. Dev. (	Min.,	Max.)	
Male   Female   Unknown   All	4  8  0  12	40.00 +/- 40.00 +/- +/- 40.00 +/-	( ( (	40.00, 40.00, , 40.00,	40.00) 40.00) ) 40.00)	     
Founder   Nonfound.  Singleton  All	8  4  0  12	40.00 +/- 40.00 +/- +/- 40.00 +/-	( ( (	40.00, 40.00, , 40.00,	40.00) 40.00) ) 40.00)	

		stics: All Pedigrees		=======:
	Count  Mean	Size +/- Std. Dev. (	Min.,	Max.)
Pedigrees	1 1	6.00 +/ (	б,	6)
Generation Stat	istics   Nuc Fami	ly Statistics    Inh ams  # of Peds   # d	Vector Bit	Stats
3	1   0 -	2   1	3 - 4	
		Size +/- Std. Dev. (		Max.)
Sibships	1 1	1.50 +/- 0.50 (	1,	2)
Constituent Pedigrees	Marria 1   Rings		   oops	
Pairs		ount    Individuals		Coun
Parent/Off Sib/Sib Sis/Sis Bro/Bro Bro/Sis		6   Male 1   Female 0   Unknown 0   Total		
Grandp. Avunc. Half Sib Cousin		2   Founder 1   Non-founder 0   Singleton 0   Total		
======================================	Individuals	With Multiple Mates		
	Consanguin	eous Mating Pairs		======

			All Pedigree	es, Trait -		
		arents w. 1	=========== Data  1 Pa	arent w. Dat		ts w. Data
Sibships			0		0	
				======================================		Max.)
Sibships		2	1.50 +/-	0.50 (	1,	2)
Pedigrees		1	6.00 +/-	(	6 , 	6)
	ly Statisti ams  # of F 2					
======================================	2  =========== Concord.  Unaff.	Discord.		======================================	 Total  	Corr
Parent/Off	 1	3	2	0	6	0.000
Sib/Sib   Sis/Sis	0	1	0   0	0	1   0	
Bro/Bro	0	0	0		0	
Bro/Sis	0 j	1	0	0	1	
Grandp.	0	1	1	0	2	
Avunc.	0	1	0	0	1	
Half Sib   Cousin	0   0	0 0	0   0	0    0	0	
======== Indiv.'s	Affected	Unaff.	======================================	======================================		
 Male	0	3	0	3		
Female	3	0	0	3		
Unknown	0 j	0	0	0	İ	
Total	3	3	0	6		
Founder	1	2	0	3		
Nonfound.	2	1	0	3		
Singleton	0	0	0	0		
Total	3	3	0	6		

# **Chapter 4**

# FCOR

FCOR can estimate multivariate familial correlations, and their asymptotic standard errors, for all pair types available in a set of pedigrees. FCOR also estimates the equivalent count of independent pairs that could theoretically have been used to obtain the same standard error for each correlation. Familial correlations for both subtypes (sex-specific) and main types (pooled sex-specific correlations) are estimated, together with their corresponding asymptotic standard errors. The variance-covariance matrices of the estimated correlations are calculated and a test for homogeneity of correlations among subtypes can be performed.

## 4.1 Limitations

Further analysis, such as adjusting for covariates, is not supported. Standard errors are based on asymptotic theory and in some cases may not be estimable.

# 4.2 Theory

The theory underlying all the calculations performed by FCOR is given in Keen and Elston (2003).

#### 4.2.1 Relative Pairs and Treatment of Missing Data

For each type of familial correlation, FCOR uses all pairs of relatives where both members have data on at least one trait in common. All other pairs of that type are excluded from the calculations and output.

#### 4.2.2 Relative Pairs Naming Convention

We call relative pair types that depend on individuals' sexes *subtypes*, and those that do not *main types*. FCOR uses one of the following two naming conventions for each relative pair type, depending on the user's choice. The default is Sex Specific Name.

#### 4.2.2.1 Non-Sex Specific Name for a Pair Type

Each type of relative pair is described by a name for the (non-sex specific) relationship and, additionally, one or two lists of M's (for male) or F's (for female) within square brackets ([]) that describe their ancestry. These lists represent a sequence of sexes for the individuals that comprise a lineage connecting the individuals in the pair. Relationships that represent direct descent (that is, parent-offspring, grandparental, great grandparental, and so on) are displayed as a single list starting with the ancestor and ending with the descendant. Relationships that do not represent direct descent (for example, sibling, nephew-uncle, cousin, and so on) are displayed with two lists. The first list begins with the first individual in the pair and terminates at the common ancestral nuclear family. The second list begins with the common nuclear family and terminates at the other individual in the pair. If the common ancestor is a parent of two half siblings (that is, the second-to-last ancestors are half siblings), then the sex of the single common ancestor is displayed between the two lists.

#### 4.2.2.2 Sex Specific Name for a Pair Type

Each type of relative pair is described by the name of the relationship with the sequence of ancestors in the lineage.

#### 4.2.2.3 Examples

Non-Sex Specific Name		Sex Specific Name
parent:offspring	[MF]	father:daughter
grandparent	[MFM]	grandfather-through-mother:grandson
great-grandparent	[MMFM]	great-grandfather-through-mother's-father:great-
		grandson
sibling	[F,M]	sister:brother
half-sibling	[F,M,F]	paternal-half-sister:half-sister
cousin	[MF,FM]	male-cousin-through-mother:male-cousin-through-
		mother
half-cousin	[MF,F,MF]	maternal-male-half-cousin-through-mother:female-
		half-cousin-through-father
second-cousin	[MFF,MFF]	male-second-cousin-through-mother's-
		mother:female-second-cousin-through-mother's-
		father
avuncular	[M,FM]	uncle-through-mother:nephew
half-avuncular	[M,M,FM]	paternal-half-uncle-through-mother:half-nephew
great-avuncular	[M,FMF]	great-uncle-through-father's-mother:great-niece
first-cousin-once-	[MF,FMF]	male-cousin-through-mother:male-cousin-through-
removed		mother, removed-daughter
first-cousin-twice-	[MF,FFMM]	male-cousin-through-mother:female-cousin-through-
removed		mother, removed-son's-son
second-cousin-once-	[MFF,FMFF]	e e
removed		mother:female-second-cousin-through-father's-
		mother, removed-daughter

#### 4.2.3 Correlations

Consider the N pairs of the observations of a particular type in the sample as a set of random two-element vectors  $\{(x_i, y_i)\}_{i=1}^N$ . These vectors are not assumed to be independent or uncorrelated, but the structure of the pairwise correlations among them is known via the pedigree structure. The pedigree correlation between the two random variables  $x_i$  and  $y_i$  is consistently estimated from a random sample of pedigrees by

$$r_{xy} = \frac{\sum_{i=1}^{N} w_i (x_i - \overline{x}) (y_i - \overline{y})}{\sqrt{\sum_{i=1}^{N} w_i (x_i - \overline{x})^2 \sum_{i=1}^{N} w_i (y_i - \overline{y})^2}}$$
(4.1)

where  $\overline{x} = \sum_i w_i x_i / \sum_i w_i$  and  $\overline{y} = \sum_i w_i y_i / \sum_i w_i$  for arbitrary non-negative weights  $\{w_i\}$ .

The pedigree correlation  $r_{xy}$  can represent either an interclass correlation (if two classes of persons are involved) or an intraclass correlation (if only one class of persons is involved), for either the same trait or different traits. For example, suppose  $r_{xy}$  represents an interclass correlation between a trait measured on a woman and a trait measured on her daughter's son. Then we can let the random variable x denote the woman's trait and the random variable y denote the trait on one of her daughter's sons. In this way, grandmother is adopted as one class and daughters' sons as another class. Given a random sample of pedigrees, the pedigrees are scanned to produce N pairs from the two classes whereby for the *i*th pair,  $x_i$  equals the value of a woman's trait and  $y_i$  equals the value of a trait of one of the woman's daughter's sons. If a woman's daughter has more than one son, then there will be pairs that share the observation of the same grandmother – for which an accounting must be made when calculating the asymptotic standard error. Moreover, a sibling correlation will also need to be accommodated as well. If a woman has more than one daughter, and each has at least one son, then a cousin correlation will need to be accommodated for cousin pairs who share a common grandmother. The situation becomes even more complex when pedigrees contain, for example, pairs of grandmothers as sisters. Note that one or more of the correlations needed to calculate a standard error may not be estimable.

A special case in pedigrees is that of intraclass correlations. These correlations are defined and estimated with respect to, for example: siblings; cousins; brother/brother; and female-cousin/female-cousin. The intraclass correlations are not necessarily restricted to the same random variable, trait, or phenotype. In the situation of relating different random variables with members of the same class of individuals, the correlations are referred to here as intraclass cross-correlations. All possible pairs within a class of individuals are formed with the random variable x representing one trait and the random variable y representing the other trait measured on a different member of the same class.

The user can specify the largest number of generations to be considered when choosing the classes for which correlations are to be calculated. If this is not specified, FCOR can examine the pedigree structure and then decides for itself what pedigree correlations should be calculated for a given random sample of pedigrees (for large pedigrees this calculation can consume a lot of computer time). Thus correlations that are not calculated are those that cannot be adequately estimated from the sample (a minimum of three pairs must be available to estimate any correlation).

#### 4.2.4 Asymptotic Standard Errors of Correlations

The asymptotic standard error of a given correlation is estimated by using a second-order Taylor series expansion and replacing all correlation parameters with their respective estimates. If a required correlation is not estimable, it is replaced by zero or the user can suppress the calculation of such a standard error.

#### 4.2.5 Equivalent Pair Count

The equivalent pair count for a specific familial correlation coefficient estimate is the estimated number of independent pairs of observations that would have a standard error the same as the value estimated for the specific familial correlation. Letting r denote the value of the correlation and s the estimate of its standard error, the equivalent count is estimated by

$$EquivalentCount = \frac{1}{2} \left[ N_0 + \sqrt{N_0^2 + \frac{22(1-r^2)}{s^2}r^2} \right], \tag{4.2}$$

where  $N_0 = 1 + (1 - r^2)^2 / s^2$ .

#### 4.2.6 Test for Homogeneity of Correlations among Subtypes

This is a test of the hypothesis that all subtypes within a main type have the same correlation.

The main types are grouped by non-sex specific relationship type. For example, the SELF main type relationship contains two subtypes – male self and female self. As another example, the PARENT:OFFSPRING main type has four subtypes – father:son, father:daughter, mother:son, and mother:daughter.

Subtype correlations are always computed first, and then, if requested, main type correlations are calculated by appropriately pooling subtype correlations. After grouping subtypes into main types, chi-square statistics and p-values are calculated to test homogeneity of correlations among the subtypes within each main type. Under the null hypothesis of homogeneity, if only one dependent variable is being analyzed, the test statistic has an approximate chi-square distribution with degrees of freedom equal to the number of subtypes minus one. If multiple dependent variables are being analyzed, the test of homogeneity includes homogeneity of all possible subtype correlations. Thus, if there are k subtypes on p traits, then the number of degrees of freedom is $(k - 1) p^2$  for interclass correlations, and(k - 1) p (p - 1) / 2 for intraclass correlations.

# 4.3 Program Input

File Type	Description
FCOR Parameter File	Specifies the parameters and options with which
	to perform a particular analysis.
Pedigree Data File	Contains delimited records for each individual in-
	cluding fields for identifiers, sex, parents, and trait
	data.

## 4.3.1 The fcor Parameter

The following syntax table specifies the permissible parameter and attribute settings for the main FCOR parameter.

<pre>parameter [, attribute]</pre>		Explanation
	Starts an FCOR an	alysis block
	Value Range	N/A
fcor	Default Value	None
	Required	Yes
	Applicable Notes	None
	Specifies the root	name to be used for output files.
	Output file names v	will be formed by concatenating the
	root name and an a	ppropriate extension.
		Character string representing a
, out	Value Range	valid file name.
	Default Value	fcor
	Required	No
	Applicable Notes	None

## 4.3.2 The fcor Block

The following syntax table specifies the permissible parameter and attribute settings for the fcor block.

<pre>parameter [, attribute]</pre>	Explanation
	Specifies a trait to be used in the analysis.
	Value Range Character string
trait	Default Value None
	Required Yes
	Applicable Notes 1
	Specifies the weight to be used for interclass correla-
	tions.
	{pair_wise, pair, uniform,
interclass_weight	Value Range reduced_group, reduced}
	Default Value pair_wise
	Required No
	Applicable Notes 2
	Specifies the weight to be used for intraclass correla-
	tions.
	{pair_wise, pair, uniform,
intraclass_weight	Value Range reduced_group, reduced}
_	Default Value pair_wise
	Required No
	Applicable Notes 2
	Specifies calculation of correlations for subtypes only,
	or for main relative types only, or for both main rela-
	tive types and subtypes.
type	Value Range {subtypes, maintypes, both}
	Default Value subtypes
	Required No
	Applicable Notes 3
	Specifies option to produce an additional output file of
	alternate tabular structure for all correlation types.
	Value Range N/A
, tabular	Default Value None
	Required No
	Applicable Notes See 4.6.3
	Option to calculate asymptotic standard errors.
	Value Range {true, false}
standard_error	Default Value true
	Required No
	Applicable Notes 4

	Specifies calculation of asymptotic standard errors conservatively.		
	Value Range N/A		
, conservative	Default Value None		
	Required No		
	Applicable Notes 4		
	Specifies option to produce additional output file indi-		
	cating, for each standard error, the smallest number of		
	pairs used to calculate any of the required correlations.		
, pairs	Value Range N/A		
	Default Value None		
	Required No		
	Applicable Notes See 4.6.4		
	Option to print out relationship name with sex.		
	Value Range {true, false}		
sex_name	Default Value true		
	Required No		
	Applicable Notes 5		
	Specifies the largest number of steps permissible be-		
	tween a given pair of individuals and their closest		
	common ancestor. Relative pairs who exceed the		
	specified value will be excluded from analysis.		
generation_limit			
	Value Range {1, 2, 3,}		
	Default Value 2		
	Required No		
	Applicable Notes 6		
	Specifies the calculation of chi-square statistics and		
	associated p-values for homogeneity tests.		
homogeneity_test	Value Range {true, false}		
nomogenercy_test	Default Value false		
	Required No		
	Applicable Notes 7		
	Starts a parameter sub-block to specify the options to		
	print out the variance-covariance matrix of correlation		
	estimates.		
var_cov	Value Range N/A		
var_00v	Default Value None		
	Required No		
	••		
	Prints a single matrix for each trait specified.		
	Value Range N/A		
, single	Default Value None		
	Required No		
	Applicable Notes 8		

	Prints a joint matrix for each pair of traits specified.	
	Value Range	N/A
, joint	Default Value	None
	Required	No
	Applicable Notes	8

Notes

- 1. The value of a trait parameter should be set to the name of a trait, phenotype or covariate field either read from the pedigree data file or created by a function statement. If no valid trait parameters are listed, then all trait fields are used. Note that this can lead to long runs for highly multivariate data, and that the test for homogeneity among subtypes considers all specified traits jointly.
- 2. This parameter is used to specify the weight to be used to compute the correlations. The value of the interclass\_weight or intraclass\_weight parameter should be one of those from the following table. If no value has been selected by the user, or an invalid value has been selected, the default **pair\_wise** will be used. For the **pair\_wise** weighting scheme, the contributions to the sums of squares and cross-products for a given pedigree is proportional to the number of pairs of a main type or subtype in the pedigree; for the uniform weighting scheme, the contributions from each pedigree have the same weights, regardless of the number of pairs of the designated type in the pedigree.

In the following table,  $p_j$  denotes the number of pairs of a given type in the *j*th pedigree. For example, in the case of brother-sister correlations,  $p_j$  would denote the total number of brother-sister pairs in the *j*th pedigree.

Parameter Value	Description	Calculation	
pair_wise or pair	Each pair of the given type in the jth pedi- gree has the same weight regardless of $p_j$ .	$w_j = 1$	
uniform	Each pair has weight inversely propor- tional to the number of such pairs, $p_j$ , in the $j^{th}$ pedigree.	$w_i = \frac{1}{p_j}$ , if $p_j \ge 1$ ,	
reduced_group or reduced	Each pair has weight inversely propor- tional to the number of such pairs, $p_j$ , in the $j^{th}$ pedigree but pedigrees that con- tain only one pair of the type are ex- cluded.	$w_j = egin{array}{cc} rac{1}{p_j} &  ext{, if } p_j > 1 \ 0 &  ext{, if } p_j = 1 \end{array}$	

- 3. The type parameter is used to specify whether to calculate correlations for subtypes only, or for main relative types only, or for both main relative types and subtypes. If the value of type is set to **subtypes**, then correlations of subtypes will be computed. If the value of type is set to **both**, then both correlations of subtypes and main types will be computed. The default value is **subtypes**.
- 4. The standard\_error parameter is used to specify whether to calculate asymptotic standard errors of correlations. By default, any standard error for which a required

correlation is nonestimable is calculated by setting the value of that required correlation to a value of 0, and appears within [] in the output. This usually overestimates the standard error. The optional conservative attribute specifies that if any required correlation is nonestimable, then that standard error is not calculated. The default value for standard\_error is **true**.

- 5. If the value of sex\_name is set to **false**, then non-sex specific names will be printed in output tables. The default value is **true**.
- 6. The generation\_limit is the largest number of steps between the pair of individuals and their closest common ancestor. For example, a generation\_limit value of one would include only parent offspring, sibling and half sibling pair types. A generation\_limit value of 2 would include all first-and second-degree relationships, cousins and half avuncular pairs.
- 7. The homogeneity\_test parameter is used to specify calculation of chi-squares and p-values to test for homogeneity of subtypes within main types. The values of interclass\_weight and intraclass\_weight are use to calculate all required correlations. The default value of homogeneity\_test is **false**.
- 8. The var\_cov parameter block is used to specify options to print variance-covariance matrices of subsets of the correlations. The single attribute is used to print all matrices for all traits, one trait at a time, and the joint attribute is used to print a joint matrix for two traits. If no attribute is specified, then single is used as the default. The traits are specified in the var\_cov parameter block as trait parameters.

### 4.3.3 The var\_cov Sub-Block

The following lists al	l parameters that ma	y occur in a var_co	vv sub-block.

parameter	Explanation		
[, attribute]			
	Names a trait, phenotype or covariate for which a		
	variance-covariance matrix is to be printed.		
tracit	Value Range	Character string.	
trait	Default Value	None	
	Required	No	
	Applicable Notes	1	
	Specifies a set of correlations between two relative		
	types for a variance-covariance matrix.		
	Value Range	Any of the entries listed in the	
correlation		codes table below.	
	Default Value	None	
	Required	No	
	Applicable Notes	2	

Notes

1. The value of a trait parameter should be set to the name from a trait parameter that is used in the fcor block. If no valid trait parameters are listed, then all trait fields used in the fcor block are used.

2. The value of correlation should be set to one of following codes or names, and can be repeated.

Main Types		Subtypes	
Code	Name	Code	Name
0	self	m	male-self
		f	female-self
1	mother:father	m,f	mother:father
10	parent:offspring	mm	father:son
		fm	mother:son
		mf	father:daughter
		ff	mother:daughter
11	sibling	m,m	brother:brother
		f,m	sister:brother
		f,f	sister:sister
11h	half-sibling	m,m,m	paternal-half-brother:half-brother
		f,m,m	paternal-half-sister:half-brother
		f,m,f	paternal-half-sister:half-sister
		m,f,m	maternal-half-brother:half-brother
		f,f,m	maternal-half-sister:half-brother
		f,f,f	maternal-half-sister:half-sister
20	grandparental	mmm	grandfather-through-father:grandson
		fmm	grandmother-through-father:grandson
		m£m	grandfather-through-mother:grandson
		ffm	grandmother-through-mother:grandson
		mm£	grandfather-through-father:granddaughter
		fmf	grandmother-through-father:granddaughter
		mff	grandfather-through-mother:granddaughter
		fff	grandmother-through-mother:granddaughter
21	avuncular	m,mm	uncle-through-father:nephew
		f,mm	aunt-through-father:nephew
		m,fm	uncle-through-mother:nephew
		f,fm	aunt-through-mother:nephew
		m,mf	uncle-through-father:niece
		f,mf	aunt-through-father:niece
		m,ff	uncle-through-mother:niece
		f,ff	aunt-through-mother:niece
22	cousin	mm, mm	male-cousin-through-father:male-cousin-
			through-father
		mf,mm	male-cousin-through-mother:male-cousin-
			through-father
		mf,fm	male-cousin-through-mother:male-cousin-
			through-mother
		fm,mm	female-cousin-through-father:male-cousin-
			through-father

Main Types		Subtypes	
Code	Name	Code	Name
		fm,fm	female-cousin-through-father:male-cousin-
			through-mother
		ff,mm	female-cousin-through-mother:male-cousin-
			through-father
		ff,fm	female-cousin-through-mother:male-cousin-
			through-mother
		fm,mf	female-cousin-through-father:female-cousin-
			through-father
		ff,mf	female-cousin-through-mother:female-
			cousin-through-father
		ff,ff	female-cousin-through-mother:female-
			cousin-through-mother

#### 4.3.4 FCOR Examples

The following are all valid fcor statements:

```
fcor
{
}
fcor
{
   interclass_weight=uniform
}
fcor, out=test
ł
   trait=TRAIT1
   trait=TRAIT2
   trait=TRAIT3
   interclass_weight=uniform
   intraclass_weight=pair
   standard_error=true
   sex_name=false
   type=maintypes
   homogeneity_test=true
   generation_limit=3
   var_cov, single { # This will calculate separate variance-
                     # covariance matrices for TRAIT1 and TRAIT2
      trait=TRAIT1
                    # parent:offspring type correlations.
      trait=TRAIT2
      correlation=parent:offspring
   }
   var_cov, joint { # This will calculate the joint variance-
                    # covariance matrices for TRAIT1 and
      trait=TRAIT1 # TRAIT2, TRAIT1 and TRAIT3, and TRAIT2 and
      trait=TRAIT2 # TRAIT3 father:son and mother:son
      trait=TRAIT3 # correlations.
      correlation=mm # father:son
      correlation=fm # mother:son
   }
}
```

## 4.4 **Program Execution**

FCOR is run via a command line interface on the supported UNIX and Windows platforms. This requires the S.A.G.E. programs to be properly installed and in the current execution path. Input files are specified on the command line and all output files are created in the current working directory.

Running FCOR from the command prompt with no arguments, or the wrong number of arguments, will result in the program printing its usage statement. This lists the input files the program requires on the command line.

```
>fcor
S.A.G.E. v5.x -- FCOR
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
usage: fcor <parameters> <pedigree>
Command line parameters:
  parameters - parameter file
  pedigree - pedigree data file
```

As indicated in the program usage statement, input files are listed on the command line. A typical run of FCOR may look like the following:

>fcor data.par data.ped S.A.G.E. v5.x -- FCOR COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY. Reading Parameter File......done. Reading pedigree file......done. from data.ped......done. Sorting pedigrees.....done. No analyses specified. Performing FCOR default analysis... Computing subtypes correlations.....done. Computing standard errors.....done. Writing output files.....done. Analysis complete!

## 4.5 Program Output

FCOR produces six types of output files that contain results and diagnostic information:

File Name	Description
fcor.inf	Contains informational diagnostic messages, warnings and pro-
	gram errors. No calculation results are stored in this file.
fcor.sub	Contains tables of correlations and standard errors with used pair
	counts and equivalent pair counts for each pair of traits for each
	subtype of relative up to 2nd generation (by default) or the gen-
	eration specified by generation_limit. Generated when
	pair_type value is <b>subtypes</b> or <b>both</b> .
fcor.main	Contains tables of correlations and standard errors, with used pair
	counts and equivalent pair counts, for each pair of traits for each
	main type of relative up to 2nd generation (by default) or up to the
	generation specified by generation_limit. Generated when
	type value is <b>maintypes</b> or <b>both</b> .
fcor.htest	Contains chi-square values and p-values. Generated when
	homogeneity_test value is true.
fcor.alt	Contains tables of correlations and standard errors, with used pair
	counts and equivalent pair counts in the alternate tabular form.
	Generated when type has an attribute tabular.
fcor.pair	Contains tables of the smallest number of pairs used to calculate
	any of the required correlations for each standard error. Generated
	when standard_error has the attribute pairs.
fcor.cov	Contains the variance-covariance matrix of correlation estimates.
	Generated when there is a var_cov sub-block within the FCOR
	block.

## 4.5.1 Information Output File

The FCOR information file contains a variety of useful information, including:

- Information on fields read from the pedigree data file. These tables, which provide information about what the program has read from the pedigree data file, are included with all programs in S.A.G.E. and are very useful for debugging most common errors caused when reading the pedigree data. When first analyzing new data, it is recommended that these tables be checked carefully to make sure pedigree data are being correctly read.
- Information, warning and error messages generated throughout the program. It is recommended that you check this file for warning and error messages before examining the results of any run of the program. The program attempts to correct many common errors and this sometimes means analyses are not as expected. The file "fcor.inf" should be checked for errors and diagnostic information after each run of the program.

#### 4.5.2 Correlations and Standard Errors: Subtypes & Main Types

The FCOR subtypes and main types output file prints the tables of correlations and standard errors for interclass and intraclass relative types, within two generations by default. An additional output file contains the alternate tabular form for the tables.

#### 4.5.3 Smallest Pair Numbers

The FCOR pair number output prints the tables indicating, for each standard error, the smallest number of pairs used to calculate any of the required correlations.

.

## 4.6 Example Output Files

#### 4.6.1 Correlations and Standard Errors: Subtypes

Here is a typical example of an FCOR subtype output file for the father-son relationship when the sex\_name option is set equal to **true**:

```
Tables of Correlations +/- Asymptotic Standard Errors for Subtypes
_____
  Number of pedigrees : 207
  Number of traits
              : 2
  Weight method used for interclass : Equal Weight to Pairs
  Weight method used for intraclass : Uniform Weight to Pedigrees
  Legend :
    ***** : Value is not estimable.
   &&&&& : Value is greater than or equal to 100000.
   @@@@@@@ : Standard error is greater than or equal to 10.0.
    ####### : Equivalent pair count is greater than or equal to 10000.
    [StdErr]: Calculated by setting the value of an inestimable required
          correlation to a value of 0.
_____
Relationship Type : Row:Column
            father:son
Pairs Found
         : 241
_____
            TRAIT1
                         TRAIT2
 INTERCLASS -----
                      _____
         Count Correlation Count Correlation
        EqvCnt +/- StdErr EqvCnt +/- StdErr
_____
        241 -0.0011 241 -0.0265
 TRAIT1
         180.3 +/- [0.0747] 440.0 +/- [0.0477]
_____
 TRAIT2
         241 -0.0494 241 -0.0585
         440.8 +/- [0.0476] 168.4 +/- [0.0770]
_____
```

#### 4.6.2 Correlations and Standard Errors: Main Types

Here is a typical example of the FCOR main type output file for the parent-offspring relationship:

```
_____
 Tables of Correlations +/- Asymptotic Standard Errors for Maintypes
Number of pedigrees : 207
  Number of traits
              : 2
  Weight method used for interclass : Equal Weight to Pairs
  Weight method used for intraclass : Uniform Weight to Pedigrees
  Legend :
    * * * * * *
         : Value is not estimable.
    &&&&& : Value is greater than or equal to 100000.
    @@@@@@@ : Standard error is greater than or equal to 10.0.
    ####### : Equivalent pair count is greater than or equal to 10000.
    [StdErr]: Calculated by setting an inestimable constituent correlation
          to a value of 0.
_____
Relationship Type : parent:offspring
 Subtypes Pooled : father : son
             mother : son
             father : daughter
             mother : daughter
Total Pairs Found : 913
_____
             TRAIT1
                          TRAIT2
INTERCLASS -----
          Count Correlation Count Correlation
        EqvCnt +/- StdErr EqvCnt +/- StdErr
_____
         913 0.1825 913
                               0.1302
  TRAIT1
          620.7 +/- [0.0388] 609.3 +/- [0.0399]
_____
              _____
                                -----
         913
                  0.1136 913 0.0730]
  TRAIT2
          650.6 +/- [0.0387] 638.9 +/- [0.0394]
_____
```

#### 4.6.3 Output File of the Alternate Tabular Form

Here is a typical example of the alternate tabular form output tables:

```
      Relationship Type : father:son

      Pairs Found : 241

      Count Correlation EqvCnt StdError

      TRAIT1 - TRAIT1 241 0.0496812 190.9 0.072387

      TRAIT1 - TRAIT2 241 0.0443138 168.6 0.077100

      TRAIT2 - TRAIT1 241 0.0258354 190.8 0.072536

      TRAIT2 - TRAIT2 241 0.0295229 168.6 0.077172
```

#### 4.6.4 Output File of the Smallest Pair Numbers

Here is a typical example of the FCOR pair numbers output tables:

```
Tables of the Smallest Number of Pairs Used in
  Calculating Required Correlations for Subtypes
_____
  Number of pedigrees : 207
  Number of traits : 2
  Weight method used for interclass : Equal Weight to Pairs
  Weight method used for intraclass : Uniform Weight to Pedigrees
  ſ
      ] : Excluded the number of pairs for inestimable required
        correlations.
_____
Relationship Type : Row:Column
           father:son
Pairs Found
         : 241
_____
INTERCLASS
         TRAIT1
               TRAIT2
_____
           178
 TRAIT1
                  178
------
          178
                 178
 TRAIT2
-----
```

#### 4.6.5 Homogeneity Test Results Output File

The FCOR homogeneity test output file for main types of relatives has chi-square values and p-values. Here is a typical example of the FCOR homogeneity test output file for the parent-offspring relationship when the homogeneity\_test option is **true**:

```
Notes :
   ***** : Value is not estimable.
     ] : Calculated by setting nonestimable required correlations
   [
         to a value of 0.
_____
Relationship Type : parent:offspring
 Subtypes Pooled : father:son
              mother:son
              father:daughter
              mother:daughter
Total Pairs Found : 420
Chi-Square = [1.82247] with 3 degree(s) of freedom
      = [0.610058]
P-Value
_____
                 _____
```

#### 4.6.6 Variance-Covariance Matrix Output File

Here is a typical example of the variance-covariance matrices for TRAIT1 and TRAIT2 parent-offspring correlations:

```
_____
 Variance-Covariance Matrix for Correlations of
   PARENT: OFFSPRING
     with
   PARENT: OFFSPRING
   trait(s) : TRAIT1 TRAIT2 SINGLY
 **** : Value is not estimable.
 [ ] : Calculated by setting an inestimable
      required correlation to a value of 0.
-----
 Legend :
    [R1] PARENT: OFFSPRING TRAIT1: TRAIT1
    [C1] PARENT: OFFSPRING TRAIT1: TRAIT1
 _____
   \backslash
          [C1]
 _____
  [R1] 0.0052217
 _____
 The Smallest Number of Pairs Used in
 Calculating Required Correlations
 _____
   \backslash
          [C1]
 _____
  [R1]
             86
 _____
 Legend :
    [R1] PARENT: OFFSPRING TRAIT2: TRAIT2
    [C1] PARENT: OFFSPRING TRAIT2: TRAIT2
 _____
          [C1]
   \backslash
 _____
   [R1] [ 0.0062607]
 _____
 The Smallest Number of Pairs Used in
 Calculating Required Correlations
 _____
   \backslash
          [C1]
 ------
  [R1]
         86
 _____
```

Here is another typical example of the joint variance-covariance matrices for TRAIT1, TRAIT2, and TRAIT3 father:son and mother:son correlations.

FATHE wit MOTHE trait **** : Va [ ] : Ca r	R:SON h R:SON (s) : TRAIT] alue is not alculated by	Matrix for Corr	JOINTLY	
Legend :				
[R1]	FATHER: SON	- TRAIT1:TRAIT1		
[R2]	FATHER: SON	TRAIT1:TRAIT2		
[R3]	FATHER:SON	TRAIT2:TRAIT1		
[R4]	FATHER: SON	TRAIT2:TRAIT2		
[C1]	MOTHER: SON	- TRAIT1:TRAIT1		
[C2]	MOTHER: SON	- TRAIT1:TRAIT2	2	
[C3]	MOTHER: SON	TRAIT2:TRAIT1		
[C4]	MOTHER:SON	TRAIT2:TRAIT2		
\	[C1]	[C2]	[C3]	[C4]
	0 0004526	-0.0000822	0 0001700	
		-0.0004590		
[RZ]	-0.0000854			
	0 0002622	0 0001625		
[R3] [R4]  The Smal	0.0000577  lest Number	0.0001635 0.0002804 of Pairs Used i	0.0001000	
[R3] [R4]  The Smal	0.0000577  lest Number	0.0002804 of Pairs Used i d Correlations	0.0001000	
[R3] [R4]  The Smal Calculat 	0.0000577 lest Number ing Required [C1]	0.0002804 of Pairs Used i d Correlations [C2]	0.0001000 n [C3]	0.0001008 [C4]
[R3] [R4]  The Smal Calculat  \ [R1]	0.0000577 lest Number ing Required [C1] 178	0.0002804 of Pairs Used i d Correlations [C2] 178	0.0001000 n [C3] 178	0.0001008 [C4] 178
[R3] [R4]  The Small Calculat  [R1] [R2]	0.0000577 lest Number ing Required [C1] 178 178	0.0002804 of Pairs Used i d Correlations [C2] 178 178	0.0001000 n [C3] 178 178	0.0001008 [C4] 178 178
[R3] [R4]  The Small Calculat 	0.0000577 lest Number ing Required [C1] 178 178 178 178	0.0002804 of Pairs Used i d Correlations [C2] 178 178 178 178	0.0001000 n [C3] 178 178 178 178	0.0001008 [C4] 178 178 178 178
[R3] [R4]  The Small Calculat  [R1] [R2]	0.0000577 lest Number ing Required [C1] 178 178	0.0002804 of Pairs Used i d Correlations [C2] 178 178 178 178	0.0001000 n [C3] 178 178	0.0001008 [C4] 178 178
[R3] [R4]  The Smal Calculat (R1] [R2] [R3] [R4]	0.0000577 lest Number ing Required [C1] 178 178 178 178	0.0002804 of Pairs Used i d Correlations [C2] 178 178 178 178	0.0001000 n [C3] 178 178 178 178	0.0001008 [C4] 178 178 178 178
[R3] [R4]  Calculat  [R1] [R2] [R3] [R4] 	0.0000577 lest Number ing Required [C1] 178 178 178 178 178	0.0002804 of Pairs Used i d Correlations [C2] 178 178 178 178 178	0.0001000 n [C3] 178 178 178 178 178	0.0001008 [C4] 178 178 178 178
[R3] [R4] 	0.0000577 lest Number ing Required [C1] 178 178 178 178 178 178	0.0002804 of Pairs Used i d Correlations [C2] 	0.0001000 .n [C3] 178 178 178 178	0.0001008 [C4] 178 178 178 178
[R3] [R4] The Smal. Calculat. (R1] [R2] [R3] [R4] Legend : [R1] [R2]	0.0000577 lest Number [C1] 178 178 178 178 178 178 5ATHER:SON	0.0002804 of Pairs Used i d Correlations [C2] 178 178 178 178 178 178 178 178	0.0001000 .n [C3] 178 178 178 178	0.0001008 [C4] 178 178 178 178
[R3] [R4] The Smal Calculat (R1] [R2] [R3] [R4] Legend : [R1] [R2] [R2] [R3]	0.0000577 lest Number [C1] 178 178 178 178 178 178 178 5ATHER:SON FATHER:SON FATHER:SON	0.0002804 of Pairs Used i d Correlations [C2] 178 178 178 178 178 178 178 178 178 178	0.0001000 .n [C3] 178 178 178 178	0.0001008 [C4] 178 178 178 178
[R3] [R4] The Smal Calculat (R1] [R2] [R3] [R4] Legend : [R1] [R2] [R3] [R3] [R4]	0.0000577 lest Number [C1] 178 178 178 178 178 178 178 54THER:SON FATHER:SON FATHER:SON FATHER:SON	0.0002804 of Pairs Used i d Correlations [C2] 178 178 178 178 178 178 178 178 178 178	0.0001000 n [C3] 178 178 178 178	0.0001008 [C4] 178 178 178 178
[R3] [R4] The Smal Calculat Calculat [R1] [R2] [R3] [R4] Legend : [R1] [R2] [R3] [R4] [R3] [R4] [C1]	0.0000577 lest Number [C1] 178 178 178 178 178 178 178 5ATHER:SON FATHER:SON FATHER:SON FATHER:SON MOTHER:SON	0.0002804 of Pairs Used i d Correlations [C2] 178 178 178 178 178 178 178 178 178 178	0.0001000 n [C3] 178 178 178 178	0.0001008 [C4] 178 178 178 178
[R3] [R4]  The Smal Calculat  [R1] [R2] [R3] [R4] 	0.0000577 lest Number [C1] 178 178 178 178 178 178 178 178 178 54THER:SON FATHER:SON FATHER:SON FATHER:SON MOTHER:SON	0.0002804 of Pairs Used i d Correlations [C2] 178 178 178 178 178 178 178 178 178 178	0.0001000 n [C3] 178 178 178 178	0.0001008 [C4] 178 178 178 178
[R3] [R4] The Smal Calculat Calculat [R1] [R2] [R3] [R4] [R2] [R3] [R4] [R3] [R4] [C1] [C2] [C3]	0.0000577 lest Number [C1] 178 178 178 178 178 178 178 5ATHER:SON FATHER:SON FATHER:SON FATHER:SON MOTHER:SON	0.0002804 of Pairs Used i d Correlations [C2] - TRAIT1:TRAIT1 - TRAIT1:TRAIT3 TRAIT3:TRAIT3 - TRAIT1:TRAIT3 - TRAIT1:TRAIT3 - TRAIT1:TRAIT3 - TRAIT1:TRAIT3 - TRAIT1:TRAIT3	0.0001000 n [C3] 178 178 178 178	0.0001008 [C4] 178 178 178 178
[R3] [R4] The Smal Calculat Calculat [R1] [R2] [R3] [R4] [R2] [R3] [R4] [R3] [R4] [C1] [C2] [C3]	0.0000577 lest Number [C1] 178 178 178 178 178 178 178 178 178 54THER:SON FATHER:SON FATHER:SON FATHER:SON MOTHER:SON MOTHER:SON	0.0002804 of Pairs Used i d Correlations [C2] - TRAIT1:TRAIT1 - TRAIT1:TRAIT3 TRAIT3:TRAIT3 - TRAIT1:TRAIT3 - TRAIT1:TRAIT3 - TRAIT1:TRAIT3 - TRAIT1:TRAIT3 - TRAIT1:TRAIT3	0.0001000 n [C3] 178 178 178 178	0.0001008 [C4] 178 178 178 178
[R3] [R4] The Small Calculat (alculat [R1] [R2] [R3] [R4] Legend : [R1] [R2] [R3] [R4] [C1] [C2] [C3] [C4]	0.0000577 lest Number ing Required [C1] 178 178 178 178 178 178 178 178 178 178	0.0002804 of Pairs Used i d Correlations [C2] 178 178 178 178 178 178 178 178 178 178	0.0001000 n [C3] 178 178 178 178 178	0.0001008 [C4] 178 178 178 178 178 [C4]
[R3] [R4]  The Small Calculat  [R1] [R2] [R3] [R4]  Legend : [R1] [R2] [R3] [R4] [C1] [C2] [C3] [C4]  [C1]	0.0000577 lest Number ing Required [C1] 178 178 178 178 178 178 178 178	0.0002804 of Pairs Used i d Correlations [C2] 178 178 178 178 178 178 178 178 178 178	0.0001000 n [C3] 178 178 178 178 178 178 -0.0000811	0.0001008 [C4] 178 178 178 178 178 178 178 178 178
[R3] [R4] The Small Calculat (R1] [R2] [R3] [R4] Legend : [R1] [R2] [R3] [R4] [C1] [C2] [C3] [C4] 	0.0000577 lest Number ing Required [C1] 178 178 178 178 178 178 178 178	0.0002804 of Pairs Used i d Correlations [C2] 	0.0001000 n [C3] 178 178 178 178 178 178 178 178 178 178	0.0001008 [C4] 178 178 178 178 178 178 178 178 178 178
[R3] [R4]  The Small Calculat  [R1] [R2] [R3] [R4]  Legend : [R1] [R2] [R3] [R4] [C1] [C2] [C3] [C4]  [C1]	0.0000577 lest Number ing Required [C1] 178 178 178 178 178 178 178 178	0.0002804 of Pairs Used i d Correlations [C2] 	0.0001000 n [C3] 178 178 178 178 178 178 -0.0000811	0.0001008 [C4] 178 178 178 178 178 178 178 178

\	[C1]	[C2]	[C3]	[C4]
[R1]	 178	178	178	178
[R2]	178	178	178	178
[R3]	178	178	178	178
[R4]	178	178	178	178
egend :				
[R1]	FATHER: SON	TRAIT2:TRAIT2		
[R2]	FATHER: SON	TRAIT2:TRAIT3		
[R3]	FATHER: SON	TRAIT3:TRAIT2		
[R4]	FATHER: SON	TRAIT3:TRAIT3		
[C1]	MOTHER:SON	TRAIT2:TRAIT2		
[C2]	MOTHER:SON	TRAIT2:TRAIT3		
[C3]	MOTHER:SON	TRAIT3:TRAIT2		
[C4]	MOTHER: SON	TRAIT3:TRAIT3		
\	[C1]	[C2]	[C3]	[C4]
[R1]	0.0001008	-0.0003228	-0.0006320	0.0000243
[R2]	0.0001108	-0.0000552	0.0005476	-0.0005515
[R3]	0.0003562	0.0005819	-0.0003420	-0.0002770
[R4]	0.0000523	0.0005094	-0.0000148	-0.0002056
T		Number of Pairs any Required (		
\	[C1]	[C2]	[C3]	[C4]
[R1]	 178	178	 178	178
[R2]	178	178	178	178
[R3]	178	178	178	178
[R4]	178	178	178	178

The Smallest Number of Pairs Used in

## **Chapter 5**

# SEGREG

## 5.1 Introduction

SEGREG is a very general program that can be used for, among other things, commingling analysis, segregation analysis and to produce penetrance files for model-based linkage analysis (for use in the programs LODLINK and MLOD). The most significant improvements over the programs REGC, REGD and REGTL of the previous versions of S.A.G.E. are as follows:

- 1. It is no longer necessary to provide initial parameter estimates (but these can be provided if desired).
- 2. It is no longer necessary (or possible) to specify parameters that control the maximizing process.
- 3. Several related analyses can be automatically performed in a single run.
- 4. When a transformation of the data is performed, all location parameter estimates refer to the data on their original scale of measurement but parameter estimates of dispersion still refer to the transformed variable.
- 5. All covariates are initially centered, and the centering (average) values are given as part of the output.

## 5.2 Limitations

As with most S.A.G.E. programs, SEGREG cannot currently be used in the presence of pedigree loops.

Further, if the sample size is small relative to the number of parameters being estimated, the likelihood may have multiple maxima. There is no guarantee that in such a situation the maximum found and reported by the program is also the global maximum. Also, situations can occur in which it is not numerically possible to calculate the variance-covariance matrix of the estimates.

## 5.3 Theory

The segregation of a possible major locus is allowed for by letting one or more parameters depend on an unobserved (latent) qualitative factor u = AA, AB or BB. Following Go et al. (1978), we call u an individual's *type*. In this context, type is best defined in terms of the expected distribution of an individual's offspring. Two individuals have the same type if and only if the expected phenotypic distributions of their offspring by a mate of a given type are identical, and this is true for every type of mate. The same concept, but not with this definition, was denoted *ousiotype* by Cannings et al. (1978). Genotypes are the special case of types, or ousiotypes, that transmit to offspring in Mendelian fashion.

Thus we use the term *type* to allow for many kinds of discrete transmission, whether Mendelian or not. When there is no transmission from one generation to the next, the model can include the existence of only one type as defined above. In this situation, it will nevertheless be convenient to refer to several types, each with its own phenotypic distribution, but it must be understood that the model then essentially allows for only a single type, the corresponding phenotypic distribution being a mixture distribution. The incorporation of types introduces two sets of parameters, type frequencies<sup>1</sup> and transmission<sup>2</sup> parameters. The population frequencies of the types are designated  $\psi_u$ , for u = AA, AB, BB, and satisfy the condition:

$$\sum_{u} \psi_u = 1$$

If the type frequencies are in Hardy-Weinberg equilibrium proportions, then they are defined in terms of  $q_A$  = frequency of (component allele) A. Thus:  $\psi_{AA} = q_A^2$ ;  $\psi_{AB} = 2q_A(1 - q_A)$ ;  $\psi_{BB} = (1 - q_A)^2$ .

Each transmission parameter  $\tau_u$  is the probability that a parent of type u transmits allele (more generally, *component*) A to offspring, for u = AA, AB, BB. Mendelian transmission corresponds to the case in which  $\tau_{AA} = 1$ ,  $\tau_{AB} = 0.5$ , and  $\tau_{BB} = 0$ . These parameters give rise to transition<sup>3</sup> probabilities. The transition probability  $Pr(u|u_F, u_M)$  is the probability that parents of types  $u_F$  and  $u_M$  produce offspring of type u. Transition probabilities are determined by the transmission probabilities as follows:

$$\begin{aligned} Pr(AA|u_F, u_M) &= \tau_{u_F} \tau_{u_M}, \\ Pr(AB|u_F, u_M) &= \tau_{u_F} (1 - \tau_{u_M}) + \tau_{u_M} (1 - \tau_{u_F}), \\ Pr(BB|u_F, u_M) &= (1 - \tau_{u_F}) (1 - \tau_{u_M}). \end{aligned}$$

However, in the case that there is homogeneity of the phenotypic distributions between generations and no parent-offspring transmission of type, we define  $Pr(u|u_F, u_M) = \tau_u$ , for u = AA, AB, BB. In order to have homogeneity across generations when there is parent-offspring transmission of type,

<sup>1.</sup> the type frequencies must be in Hardy-Weinberg equilibrium proportions, and

<sup>&</sup>lt;sup>1</sup>We use the word *frequencies* in the sense used by geneticists, i.e., *relative frequencies* that sum to 1.

 $<sup>^{2}</sup>$ SEGREG uses the terms *transmission probability* and *transition probability* as defined by Elston and Stewart (1971).  $^{3}$ ditto

2.  $\tau_{AB}$  must be a function of  $\tau_{AA}$ ,  $\tau_{BB}$  and the allele frequency  $q_A$  (Demenais and Elston, 1981).

Details of the pedigree likelihoods that are calculated, on the assumption of random mating, are given below. It should be noted that singletons (unrelated individuals) may be included in the data. Although SEGREG counts and treats them separately for convenience, they are in fact simply one-person pedigrees and, as such, require no special treatment in the model. However, note that these singletons are not considered to be founders. Estimation is performed by maximum likelihood and standard errors are obtained by numerical double differentiation of the log likelihood surface. The output contains the overall ln(likelihood), -2ln(likelihood) and Akaike's *A* information criterion (AIC)<sup>4</sup> for each of the models that has been maximized in a run. When the model consists of two or three types, a table is produced indicating the respective likelihood ratio statistic for each type.

Transmission models are compared and p-values quoted according to the asymptotic distribution of the likelihood ratio as shown in the table below (Self and Liang, 1987).

Dist	ibutions of the Segre	gation Analysis Test S	statistic Used by SEG	REG
	Homo-No-Tran	Homo-Mendelian	Homo-General	Tau-AB-Free
Homo-Mendelian	—			
Homo-General	$\begin{array}{c} \chi_2^2 \\ (2t, 3t\text{-hwe}) \end{array}$	$\frac{\left(\frac{1}{4}\right) + \left(\frac{1}{2}\right)\chi_1^2 + \left(\frac{1}{4}\right)\chi_2^2}{(2t, 3t\text{-hwe})}$		
Tau-AB-Free	—	$\left(\frac{1}{2}\right) + \left(\frac{1}{2}\right)z_1^2$ (2t, 3t-hwe)	-	
General	$\chi_3^2$ (2t, 3t-hwe, 3t-nhwe)	$\frac{\left(\frac{1}{4}\right)z_1^2 + \left(\frac{1}{2}\right)z_2^2 + \left(\frac{1}{4}\right)z_3^2}{(2t, 3t\text{-hwe, }3t\text{-nhwe})}$	$\begin{array}{c} \chi_1^2 \\ (2t, 3t\text{-hwe}) \end{array}$	$\chi^2_2$ (2t, 3t-hwe, 3t-nhwe)
Legend	2t 3t-hwe 3t-nhwe	Not Applicable Two-type models Three-type models with HWE Three-type models without HWE		

<sup>4</sup>Contrary to popular belief, the acronym AIC stands for *A Information Criterion*, and not *Akaike's Information Criterion*.

## 5.3.1 Segregation Models

Certain aspects of the models available in SEGREG are common to all traits and models, and are described here. Later sections describe the aspects that are specific to regressive models for continuous traits, regressive multivariate logistic models for binary traits, the finite polygenic mixed model, and models for binary traits with variable age of onset.

#### 5.3.1.1 Single Ascertainment and/or Conditioning on a Subset

Instead of being sampled at random, a pedigree may be included in the analysis because one or more members of the pedigree have particular trait values or are in a certain *sampling frame*<sup>5</sup>. It may be desired to condition the likelihood on the phenotypes of these individuals or, more generally, on the phenotypes and/or structure of any prespecified subset C of the pedigree. This *conditioned subset* may be

- 1. the set of founders (members of the pedigree whose parents are not included in the pedigree),
- 2. the set of pedigree members in the pedigree proband sampling frame, or
- 3. the union of these two sets.

Currently, no model is assumed for the ascertainment and, for results to be correct, the observed pedigree must contain all members of the pedigree proband sampling frame. This subsumes both simplex and multiplex single ascertainment (see Elston and Bonney, 1986) as a special case. See Ginsberg et al (2003) for a discussion of what is meant by "pedigree", "correct results" and "pedigree proband sampling frame".

If no conditioned subset is indicated for a particular pedigree (either explicitly as a user-specified set or implicitly as the founders), random sampling is assumed for that pedigree. In general, the likelihood for a randomly sampled pedigree (L) is divided by a correction  $L_C$ , defined in one of three possible ways.

#### 1. Random Sampling

In this case, no correction is necessary, so C is empty and we define  $L_C = 1$ .

#### 2. Conditioning on Actual Phenotypes

In this case, the likelihood is conditioned on the available phenotype of each member of the conditioned subset. The correction  $L_C$  is then taken to be L computed as though all individuals not in C are missing.

#### 3. Conditioning on Phenotypes Being Above or Below a Threshold Value

In this case, the likelihood is conditioned, for each member of the conditioned subset for whom a phenotype is available, on that member's phenotype being at least as large as a threshold  $T_U$ , or at most as large as a threshold  $T_L$ .

<sup>&</sup>lt;sup>5</sup>The pedigree sampling frame can include pedigree members for whom the trait value is missing. (See Ginsberg et al (2003) for further discussion.)

#### 5.3.1.2 Type Probabilities and Penetrance Functions

Given a model with established parameter values, we can estimate the probability of each possible type for every individual conditional on all the sample data. We define the following terms:

- L(•) is a likelihood
- *S* is the set of all sampled data in the pedigree,
- $t_i$  is the analysis trait of individual i,
- $u_i$  is the type of individual *i*,
- $u_{i_M}$  is the type of individual *i*'s mother, and
- $u_{i_F}$  is the type of individual *i*'s father.

Then the posterior probability for a given individual is computed (using maximum likelihood estimates of unknown parameters) as:

$$L(u_i|S) = \frac{L(u_i, S)}{L(S)}.$$

Note that the denominator is the likelihood L computed for the whole pedigree to which *i* belongs.

If  $u_i$  is a genotype, SEGREG can prepare files of penetrance functions that can be used as input into LODLINK and MLOD using maximum likelihood estimates of all unknown parameters. These are of the form  $\Pr(t_i|u_i)$ .

#### 5.3.2 Regressive Models for Continuous Traits

Regressive models (Bonney, 1984; 1998) are those models in which distributions over pedigrees are specified by conditioning each individual's trait value on those of antecedent individuals. For continuous trait they assume (possibly after transformation) multivariate normality across pedigree members of the underlying individual residuals from the type means. Two classes of regressive models for continuous traits are implemented in SEGREG. Class A models assume that siblings are dependent only because of common parentage, while class D models assume that the sibling correlations are equal, but not necessarily due to common parentage alone. For a continuous trait, SEGREG assumes a model that is a close approximation to multivariate normal for the underlying individual residuals. The approximation used is a generalization of approximation 6 in Demenais et al (1990).

The following non-zero residual correlations are allowed in all the models:  $\rho_{FM}$  for father-mother (spouse),  $\rho_{MO}$  for mother-offspring,  $\rho_{FO}$  for father-offspring, and  $\rho_{SS}$  for the correlation between any two siblings (for a class D model). A class A model also includes, indirectly, a sibling correlation  $\rho_{SS}$  that satisfies the condition

$$\rho_{SS} = \frac{\rho_{MO}^2 + \rho_{FO}^2 - 2\rho_{FM}\rho_{MO}\rho_{FO}}{1 - \rho_{FM}^2}.$$

The residual correlations between half siblings are assumed to be zero, conditional on the common parent. Missing values are handled according to the formulas in Bonney (1984, 1998).

In the correlation structure indicated above, the means and variances of the underlying normal distribution can be dependent on covariates. All covariates are centered prior to inclusion in the likelihood.

When types are incorporated into the model, the correlation parameters ( $\rho$ s) are the correlations of the residual multivariate normal distribution. Thus the inference of a major gene can be made allowing for the cumulative effect, assumed to be multivariate normally distributed for the transformed trait, of various factors (such as polygenes, cultural, and other environmental factors) that are not separately distinguished.

#### 5.3.2.1 Composite Trait

The trait, or phenotype, to be analyzed may be a single variate, the *main phenotype*  $(y^* = y)$  or a linear function of the main phenotype (with coefficient 1) and p covariates (with coefficients  $\kappa_i$ ):

$$y^* = y + \kappa_1 x_1 + \kappa_2 x_2 + \dots + \kappa_{p_\kappa} x_p,$$

where the parameters  $\kappa_i$  may be estimated.

#### 5.3.2.2 Transformation of the Phenotype

The phenotype  $y^*$ , however composed, may be transformed by one of two transformations. For commingling analysis and segregation analysis, the first (Cox and Box) transformation is recommended.

The first possible transformation is:

$$t = h(y^*) = \begin{cases} \frac{(y^* + \lambda_1)^{\lambda_1} - 1}{\lambda_1(y^*_{G1})^{(\lambda_1 - 1)}} & if\lambda_1 \neq 0, \\ \\ y^*_{G1} \ln(y^* + \lambda_2) & if\lambda_1 = 0 \end{cases}$$

where

$$y_{G1}^* = \left[\prod_{i=1}^N (y_i^* + \lambda_2)\right]^{\frac{1}{N}},$$

and N = number of individuals in the full data set (possibly including more than one pedigree) with complete phenotype and covariate values (nothing missing). This is the standardized Box and Cox (1964) transformation with power parameter  $\lambda_1$  and shift parameter  $\lambda_2$ .

The second possible transformation is:

$$t = h(y*) = \begin{cases} \frac{sign(y*+\lambda_2)[(|y*+\lambda_2|+1)^{\lambda_1}-1]}{\lambda_1(y^*_{G2})^{(\lambda_1-1)}} & if \ \lambda_1 \neq 0\\\\ y^*_{G2}sign(y*+\lambda_2)\ln(|y*+\lambda_2|+1) & if \ \lambda_1 = 0 \end{cases}$$

where

$$y_{G2}^* = \left[\prod_{i=1}^N \left(|y_i^* + \lambda_2| + 1\right)\right]^{\frac{1}{N}}.$$

This is the standardized generalized modulus power transformation (George and Elston, 1988) with power parameter  $\lambda_1$  and shift parameter  $\lambda_2$ .

We call the transformed phenotype *t* the *analysis trait*. When a transformation is applied it is applied to *both sides* (Carroll and Ruppert, 1984), so that all location parameters are median unbiased on the original scale of measurement.

#### 5.3.2.3 Likelihood for a Randomly Sampled Pedigree

Let the pedigree contain n individuals (i = 1, ..., n) on each of whom we observe a value of the analysis trait. An individual is considered missing if the value of any variate for that individual, required for calculating the likelihood, is unknown. For individual i, let

 $t_i$  = analysis trait value of i

 $x_{ij} = j$ -th covariate value of i

 $u_i = type of i$ 

 $S_i$  = spouse of i

 $M_i$  = mother of i

 $F_i$  = father of i

 $B_{ij}$  = j'th observed elder sibling of i

 $n_{iB}$  = number of observed elder siblings of i.

We let the expected value of *t* conditional on type *u* be

$$\theta_u(i) = h(\beta_u + \xi_1 x_{i1} + \xi_2 x_{i2} + \dots + \xi_{p_{\mathcal{E}}} x_{ip_{\mathcal{E}}})$$

and the variance of t conditional on type u be

$$\eta_u^2(i) = \sigma_u^2 + \varsigma_1 x_{i1} + \varsigma_2 x_{i2} + \dots + \varsigma_{p_\varsigma} x_{ip_\varsigma}$$

Because the expected value of t conditional on type u undergoes the same transformation as is used to produce t ("transformation of both sides", see Carroll & Ruppert, 1984), the estimates of parameters in this conditional expectation are median unbiased on the same scale of measurement as the original untransformed data. However, the residual variance that is calculated, and all the covariate coefficients pertaining to it, are on the scale of the analysis trait. Further general quantities that apply to regressive models are defined as follows:

$$\alpha_{iS} = \begin{cases} \rho_{FM} & \text{if specific spouse of } i \text{ is observed}, \\ 0 & \text{otherwise}, \end{cases}$$

$$\alpha_{iM} = \begin{cases} \frac{\rho_{MO} - \rho_{FO} \rho_{FM}}{1 - \rho_{FM}^2} & \text{if both parents of $i$ are observed,} \\ \rho_{MO} & \text{if mother, but not father, of $i$ is observed,} \\ 0 & \text{if mother of $i$ is not observed,} \end{cases}$$

$$\alpha_{iF} = \begin{cases} \frac{\rho_{FO} - \rho_{MO}\rho_{FM}}{1 - \rho_{FM}^2} & \text{if both parents of } i \text{ are observed}, \\ \rho_{FO} & \text{if father, but not mother, of } i \text{ is observed}, \\ 0 & \text{if father of } i \text{ is not observed}, \end{cases}$$

$$\delta_{i} = \alpha_{iM}\rho_{MO} + \alpha_{iF}\rho_{FO} = \begin{cases} \rho_{SS} & \text{if both parents of } i \text{ are observed}, \\ \rho_{MO}^{2} & \text{if mother, but not father, of } i \text{ is observed}, \\ \rho_{FO}^{2} & \text{if father, but not mother, of } i \text{ is observed}, \\ 0 & \text{if neither parent of } i \text{ is observed}, \end{cases}$$

$$\phi(z_i, w_i) = \frac{1}{\sqrt{2\pi w_i}} \exp[-z_i^2/(2w_i)],$$

where the arguments  $z_i$  and  $w_i$  are defined differently for each of the model classes. For a class A model, the arguments of the normal density function are defined in SEGREG as

$$z_{i} = t_{i} - \theta_{u}(i) - b_{iS}V_{iS_{i}}(t_{S_{i}} - \theta_{u}(S_{i})) - b_{iM}V_{iM_{i}}(t_{M_{i}} - \theta_{u}(M_{i})) - b_{iF}V_{iF_{i}}(t_{F_{i}} - \theta_{u}(F_{i}))$$

and

$$w_{i} = \eta_{u}^{2}(i)(1 - b_{iS}\rho_{FM} - b_{iM}\rho_{MO} - b_{iF}\rho_{FO}),$$

where

$$V_{ij} = \eta_u(i)/\eta_u(j)$$
$$b_{iS} = \alpha_{iS},$$
$$b_{iM} = \alpha_{iM} \left( \frac{1 - \rho_{SS}}{1 - \rho_{SS} + n_{iB}(\rho_{SS} - \delta_i)} \right)$$

$$b_{iF} = \alpha_{iF} \left( \frac{1 - \rho_{SS}}{1 - \rho_{SS} + n_{iB}(\rho_{SS} - \delta_i)} \right),$$

with

$$\rho_{SS} = \frac{\rho_{MO}^2 + \rho_{FO}^2 - 2\rho_{MO}\rho_{FO}\rho_{FM}}{1 - \rho_{FM}^2}.$$

For a class D model, the arguments of the normal density function are defined as:

$$z_{i} = t_{i} - \theta_{u}(i) - b_{iS}V_{iS_{i}}(t_{S_{i}} - \theta_{u}(S_{i})) - b_{iM}V_{iM_{i}}(t_{M_{i}} - \theta_{u}(M_{i}))$$
$$-b_{iF}V_{iF_{i}}(t_{F_{i}} - \theta_{u}(F_{i})) - b_{iB}\sum_{j=1}^{n_{iB}}\hat{V}_{iB_{ij}}(t_{B_{ij}} - \hat{\mu}_{B_{ij}}),$$

and

$$w_{i} = \eta_{u}^{2}(i)(1 - b_{iS}\rho_{FM} - b_{iM}\rho_{MO} - b_{iF}\rho_{FO} - n_{iB}b_{iB}\rho_{SS}),$$

where

$$\begin{split} \hat{\mu}_{j} &= \sum_{u_{j}} \theta_{u}(j) f_{uj} / \sum_{u_{j}} f_{uj} \\ f_{uj} &= Pr(u_{j} | u_{F_{j}}, u_{M_{j}}) exp\{-(t_{j} - \theta_{u}(j))^{2} / (2\eta_{u}^{2}(j))\} / \eta_{u}(j), \\ \hat{\sigma}_{j}^{2} &= \sum_{u} f_{uj} \sigma_{u}^{2} / \sum_{u} f_{uj}, \\ \hat{V}_{ij} &= \sum_{u_{j}} f_{uj} / \sum_{u_{j}} f_{uj} \eta_{u}^{2}(j) \\ b_{iS} &= \alpha_{iS} \\ b_{iM} &= \alpha_{iM} \left( \frac{1 - \rho_{SS}}{1 - \rho_{SS} + n_{iB}(\rho_{SS} - \delta_{i})}, \right) \end{split}$$

To indicate all the potential variables in  $\phi(z_i, w_i)$ , except covariates, denote it

$$Pr(t_i|u_i, u_S, u_M, u_F, t_{S_i}, t_{M_i}, t_{F_i}, t_{B_{il}}, \dots, t_{B_{in_{iB}}})$$

(This quantity is a conditional phenotypic density function, sometimes referred to as a penetrance function.)

Using the components defined above, and letting

$$p_i(u_i, u_{M_i}, u_{F_i}) = \begin{cases} Pr(u_i | u_{F_i}, u_{M_i}) & \text{if the parents of } i \text{ are included in pedigree,} \\ \psi_{u_i} & \text{otherwise,} \end{cases}$$

 $H_i(u_i, u_{S_i}, u_{M_i}, u_{F_i}, t_i, t_{S_i}, t_{M_i}, t_{F_i}, t_{B_{il}}, \dots, t_{B_{in_iB}})$ 

$$= \begin{cases} p_i(u_i, u_{M_i}, u_{F_i}) & \text{if $i$ missing}, \\ \\ p_i(u_i, u_{M_i}, u_{F_i}) Pr(t_i | u_i, u_{S_i}, u_{M_i}, u_{F_i}, t_{S_i}, t_{M_i}, t_{F_i}, t_{B_{i1}}, \dots, t_{B_{in_{iB}}}) & \text{otherwise} \end{cases}$$

under random mating the likelihood for a randomly sampled pedigree is

$$L = \left[\sum_{u_1} \dots \sum_{u_n} \prod_{i=1}^n H_i(u_i, u_{S_i}, u_{M_i}, u_{F_i}, t_i, t_{S_i}, t_{M_i}, t_{F_i}, t_{B_{il}}, \dots, t_{B_{in_iB}})\right].$$

#### 5.3.2.4 Allowing for Ascertainment

Ascertainment is allowed for as indicated in 5.3.1.1. In order to condition on phenotypes being at least as large as  $T_U$  or at most as large as  $T_L$ , the correction  $L_C$  is taken to be the likelihood defined in 5.3.1.1 computed as though all individuals not in the prespecified subset C are missing, but with  $Pr(t_i|.)$ , for each individual i in C replaced by

$$\int_{T_U}^{\infty} Pr(t|.)dt = \Phi(-z_{iU}/\sqrt{w_i}), or \int_{-\infty}^{T_L} Pr(t|.)dt = \Phi(z_{iL}/\sqrt{w_i}),$$

where  $z_{iU}$  or  $z_{iL}$  is identical to  $z_i$  with  $h(T_U)$  or  $h(T_L)$ , respectively, substituted for  $t_i$ . However,  $z_i$  is always left unchanged for any founders not included in the proband sampling frame.

#### 5.3.3 Regressive Multivariate Logistic Models for Binary Traits

The multivariate logistic model for a binary trait was described by Karunaratne and Elston (1998) for nuclear family data. It is implemented in SEGREG for pedigree data by making the regressive model assumption that, conditional on the phenotype and major type of any individual who belongs to two nuclear families, the likelihoods for those two nuclear families are independent. In this model, unlike in Bonney's (1986) multiple logistic model, the marginal probability that any pedigree member has a particular phenotype is the same for all members who have the same values of any covariates in the model. This marginal probability, which we call susceptibility, is given by the cumulative logistic function

$$\gamma = \frac{e^{\theta(i)t_i}}{1 + e^{\theta(i)}},$$

where  $t_i$ , the analysis trait of the i-th individual, is 1 for an affected individual and 0 for an unaffected individual; and  $\theta(i)$ , the logit of the susceptibility for the i-th individual, can depend on both major type (u) and covariate values  $x_{i1}, x_{i2}, ..., x_{ip}$ :

$$\theta_u(i) = \beta_u + \xi_1 x_{i1} + \dots + \xi_p x_{ip}.$$

Composite traits and phenotype transformation are not relevant for a binary trait; nor is a Class A model possible.

Nuclear family residual association parameters, analogous to the correlation parameters in regressive models for continuous traits, are incorporated into the model. These are denoted in 5.3.4 below as  $\delta_{FM}$  for father-mother (spouse),  $\delta_{MO}$  for mother-offspring,  $\delta_{FO}$  for father-offspring, and  $\delta_{SS}$  for any two siblings. In the case of the multivariate logistic distribution these association parameters correspond to second-order correlations; it is assumed that all higher correlations are zero. The actual correlations are calculated from these associations measures for specific logit values [see Karunaratne and Elston (1988)].

For a binary trait, information about the population prevalence of the trait (for a binary trait with variable age of onset, the probability of having been affected since birth) can be incorporated into the likelihood as an independent factor. This is done by specifying that a sample of N independent individuals have been observed, of whom R have been affected for given values of the covariates (and/or up to a specified age), and this may be repeated for different sets of covariate values (The corresponding factor(s) in the likelihood are not shown in the next section). Similarly, the program can output the prevalence, for given sets of covariate values (and/or up to a specified age), estimated from the model using the maximum likelihood estimates of all parameters.

#### 5.3.3.1 Likelihood for a Randomly Sampled Nuclear Family

Let  $t_F$ ,  $t_M$  and  $t_i$  be the phenotypes of the father, mother and i-th child, i = 1, 2, ..., n and  $u_F$ ,  $u_M$  and  $u_i$  be the types of the father, mother and i-th child. Then the likelihood for a nuclear family is

$$\sum_{u_F} \sum_{u_M} \sum_{u_1} \dots \sum_{u_n} Pr(u_F) Pr(u_M | u_F) \prod_{i=1}^n Pr(u_i | u_F, u_M) L(t_F, t_M, t_1, \dots, t_n | u_1, \dots, u_n) + \sum_{i=1}^n Pr(u_i | u_F, u_M) L(t_F, t_M, t_1, \dots, t_n | u_1, \dots, u_n) + \sum_{i=1}^n Pr(u_i | u_F, u_M) L(t_F, t_M, t_1, \dots, t_n | u_1, \dots, u_n) + \sum_{i=1}^n Pr(u_i | u_F, u_M) L(t_F, t_M, t_1, \dots, t_n | u_1, \dots, u_n) + \sum_{i=1}^n Pr(u_i | u_F, u_M) L(t_F, t_M, t_1, \dots, t_n | u_1, \dots, u_n) + \sum_{i=1}^n Pr(u_i | u_F, u_M) L(t_F, t_M, t_1, \dots, t_n | u_1, \dots, u_n) + \sum_{i=1}^n Pr(u_i | u_F, u_M) L(t_F, t_M, t_1, \dots, t_n | u_1, \dots, u_n) + \sum_{i=1}^n Pr(u_i | u_F, u_M) L(t_F, t_M, t_1, \dots, t_n | u_1, \dots, u_n) + \sum_{i=1}^n Pr(u_i | u_F, u_M) L(t_F, t_M, t_1, \dots, t_n | u_1, \dots, u_n) + \sum_{i=1}^n Pr(u_i | u_F, u_M) L(t_F, t_M, t_1, \dots, t_n | u_1, \dots, u_n) + \sum_{i=1}^n Pr(u_i | u_F, u_M) L(t_F, t_M, t_1, \dots, t_n | u_1, \dots, u_n) + \sum_{i=1}^n Pr(u_i | u_F, u_M) L(t_F, t_M, t_1, \dots, t_n | u_1, \dots, u_n) + \sum_{i=1}^n Pr(u_i | u_F, u_M) L(t_F, t_M, t_1, \dots, t_n | u_1, \dots, u_n) + \sum_{i=1}^n Pr(u_i | u_F, u_M) L(t_F, t_M, t_1, \dots, t_n | u_1, \dots, u_n) + \sum_{i=1}^n Pr(u_i | u_F, u_M) L(t_F, t_M, t_1, \dots, t_n | u_1, \dots, u_n) + \sum_{i=1}^n Pr(u_i | u_F, u_M) L(t_F, t_M, t_1, \dots, t_n | u_1, \dots, u_n) + \sum_{i=1}^n Pr(u_i | u_F, u_M) L(t_F, t_M, t_1, \dots, t_n | u_1, \dots, u_n) + \sum_{i=1}^n Pr(u_i | u_F, u_M) L(t_F, t_M, t_1, \dots, t_n | u_1, \dots, u_n) + \sum_{i=1}^n Pr(u_i | u_F, u_M) L(t_F, t_M, t_1, \dots, t_n | u_1, \dots, u_n) + \sum_{i=1}^n Pr(u_i | u_F, u_M) L(t_F, t_M, t_1, \dots, t_n | u_1, \dots, u_n) + \sum_{i=1}^n Pr(u_i | u_F, u_M) L(t_F, t_M, t_1, \dots, t_n | u_1, \dots, u_n) + \sum_{i=1}^n Pr(u_i | u_F, u_M) L(t_F, t_M, t_M) L(t_F, t_M, t_M) + \sum_{i=1}^n Pr(u_i | u_F, u_M) L(t_F, t_M, t_M) + \sum_{i=1}^n Pr(u_i | u_H, u_M) + \sum_{i=1}^n Pr(u_i | u_F, u_M) L(t_F, t_M, t_M) + \sum_{i=1}^n Pr(u_i | u_H, u_M) + \sum_{i=1}^n Pr(u_i$$

where  $Pr(u_F)Pr(u_M|u_F)$  is the joint probability of types  $u_F$  and  $u_M$  in the population;  $Pr(u_i|u_F, u_M)$  is the probability that a sib has type  $u_i$ , given the parents' types are  $u_F$  and  $u_M$ ; and the penetrance function  $L(t_F, t_M, t_1, ..., t_n|u_F, u_M, u_1, ..., u_n)$  is given by

$$\begin{split} \prod_{i=F,M,1}^{n} \frac{e^{\theta_{u}(i)t_{i}}}{1+e^{\theta_{u}(i)}} \left\{ 1+\rho_{FO}\left(1-\frac{e^{\theta_{u}(F)t_{F}}}{1+e^{\theta_{u}(F)}}\right) \sum_{i=1}^{n} (-1)^{t_{F}+t_{i}} \left(1-\frac{e^{\theta_{u}(i)t_{i}}}{1+e^{\theta_{u}(i)}}\right) \right. \\ \left. +\rho_{MO}\left(1-\frac{e^{\theta_{u}(M)t_{M}}}{1+e^{\theta_{M}(u)}}\right) \sum_{i=1}^{n} (-1)^{t_{M}+t_{i}} \left(1-\frac{e^{\theta_{u}(i)t_{i}}}{1+e^{\theta_{u}(i)}}\right) \right. \\ \left. +\rho_{SS} \sum_{1\leq i< j\leq n} (-1)^{t_{i}+t_{j}} \left(1-\frac{e^{\theta_{u}(i)t_{i}}}{1+e^{\theta_{u}(i)}}\right) \left(1-\frac{e^{\theta_{u}(j)t_{j}}}{1+e^{\theta_{u}(j)}}\right) \right. \\ \left. +(-1)^{t_{M}+t_{F}} \rho_{MF} \left(1-\frac{e^{\theta_{u}(M)t_{M}}}{1+e^{\theta_{u}(M)}}\right) \left(1-\frac{e^{\theta_{u}(F)t_{F}}}{1+e^{\theta_{u}(F)}}\right) \right\}. \end{split}$$

#### 5.3.4 Finite Polygenic Mixed Model

The finite polygenic mixed model (Fernando et al, 1994; Lange, 1997) can be used for either continuous or binary traits, the only difference being in the particular penetrance function used. It can also be used for binary traits with variable age of onset.

In addition to type (AA, AB or BB), we assume the presence of k diallelic polygenic loci in the model. The alleles at each such locus are a and b, with effects  $\alpha$  and  $\beta$ , and frequencies p and 1-p (the default value of p is 0.5). The polygenic effect is the sum of the effects of alleles at all k loci. Thus, if a pedigree member has v a alleles and (2k - v) b alleles, then the polygenic effect is

$$\mu_v = v\alpha + n(2k - v)\beta,$$

where v is called the polygenic number, and  $\alpha$  and  $\beta$  are chosen to make the mean polygenic effect zero. It follows that

$$\mu_v = \frac{v - 2pk}{1 - p} \sqrt{\frac{\sigma_v^2(1 - p)}{2pk}},$$

where  $\sigma_v^2$  is the variance of the polygenic effect.

We assume that, conditional on the polygenic numbers of two parents, the polygenic number of any pedigree member is independent of the polygenic numbers of all other pedigree members. This allows us to use the Elston-Stewart (1971) algorithm summing over the 2k +1 possible genetic numbers times the three possible types. Although this is not strictly consistent with Mendelian inheritance, it leads to a conditional correlation of zero between the polygenic numbers of any two pedigree members.

It is possible to analyze a composite trait and to transform the phenotype in the case of a continuous trait. As for regressive models for continuous traits, the type mean and/or variance can depend on covariates. For a continuous trait, let  $t_i$  be the analysis trait for individual i, with expectation conditional on type u:

$$\theta_u(i) = h(\beta_u + \xi_1 x_{i1} + \xi_2 x_{i2} + \dots + \xi_{p_{\mathcal{E}}} x_{ip_{\mathcal{E}}})$$

and let the variance of t conditional on type u be

$$\eta_u^2(i) = \sigma_u^2 + \varsigma_1 x_{i1} + \varsigma_2 x_{i2} + \dots + \varsigma_{p_{\varsigma}} x_{ip_{\varsigma}}.$$

Then we define the penetrance function for a continuous trait to be

$$Pr(t_i|u_i, v_i) = \varphi(t_i - \theta_u(i) + \mu_{v_i}, \sigma_i^2),$$

with polygenic variance equal to the variance of  $\mu_{v_i}$ .

In the case of a binary trait, we define the penetrance function to be the cumulative logistic function

$$Pr(t_i|u_i, v_i) = \frac{e^{\theta_u(i)}}{1 + e^{\theta_u(i)}}$$

where, conditional on type *u*, we have the logit

$$\theta_u(i) = \beta_u + \mu_{v_i} + \xi_1 x_{i1} + \xi_2 x_{i2} + \dots + \xi_{p_{\mathcal{E}}} x_{ip_{\mathcal{E}}}.$$

#### 5.3.4.1 Likelihood for a Randomly Sampled Pedigree

Using the penetrance functions defined above, and letting

$$P_i(u_i, u_{Mi}, u_{Fi}, v_i, v_{Mi}, v_{Fi}) =$$

$$\begin{cases} Pr(u_i, u_{Mi}, u_{Fi}, v_i, v_{Mi}, v_{Fi}) & if the parents of i are included in the pedigree \\ \psi_i & otherwise \end{cases}$$

and

$$H_i(u_i, v_i, z_i) = \begin{cases} Pr(u_i, u_{M_i}, u_{F_i}, v_i, v_{M_i}, v_{F_i}) & \text{if $i$ is missing,} \\ Pr(u_i, u_{M_i}, u_{F_i}, v_i, v_{M_i}, v_{F_i}) Pr(t_i | u_i v_i) & \text{otherwise} \end{cases}$$

under random mating the likelihood for a randomly sampled pedigree is

$$L = \sum_{u_1} \dots \sum_{u_n} \sum_{v_1} \dots \sum_{v_n} \prod_{i=1}^n H_i(u_i, v_i, z_i).$$

#### 5.3.5 Binary Traits with Variable Age of Onset

SEGREG currently uses the finite polygenic mixed model for binary traits with variable age of onset.

In general terms, letting a be age of onset and a' the age at examination (for an unaffected person, the last age at which a person is known to be so), the penetrance functions are:

- $\gamma(f(a))$  for an affected person with known age of onset a,
- $\gamma(F(a'))$  for an affected person with unknown age of onset, age at examination a', and
- $1 \gamma(F(a'))$  for an unaffected person with age at examination a',

where  $\gamma$  is the susceptibility and f(a) is the age of onset density with cumulative distribution F(a').

The mean and variance of f, and susceptibility  $\gamma$ , can each be made dependent on covariates and/or type, in the same way as for a continuous analysis trait and a binary trait, respectively. However, age of onset is assumed to follow a logistic density function rather than a normal density function; because this causes the variance of the age of onset distribution to depend on the mean, it is not permissible for the mean and variance to depend on the same covariate. Letting  $\beta$  be a baseline parameter and  $\alpha$  the age coefficient, the density and cumulative distribution functions are:

$$f(a) = \frac{\alpha e^{\beta + \alpha a}}{(1 + e^{\beta + \alpha a})^2}$$

$$F(a') = \frac{e^{\beta + \alpha a'}}{1 + e^{\beta + \alpha a'}} = [1 + e^{-(\beta + \alpha a')}]^{-1}$$

For this distribution, the mean  $= -\frac{\beta}{\alpha}$ , and the variance  $= \frac{\pi^2}{3\alpha^2}$ .

As for any other continuous trait, the mean and variance can each depend linearly on covariates, and transformation of "both sides" is possible. Using the logistic distribution has the advantage that the parameters  $\alpha$  and  $\beta$  can be interpreted as increases in log odds in the susceptible population (in the whole population if  $\gamma = 1$ ).

The susceptibility  $\gamma$  is modeled by a cumulative logistic, in the same way as a binary trait is modeled. In order to avoid confounding among the parameters, there are restrictions on how age of onset and susceptibility depend on type and polygenic number. The following six possibilities are permissible:

1. Age of onset depends on major genotype alone, susceptibility depends on neither major genotype nor polygenic number

- 2. Age of onset depends on both major genotype and polygenic number, susceptibility depends on neither
- 3. Age of onset depends on major genotype alone, susceptibility depends on polygenic number alone
- 4. Susceptibility depends on major genotype alone, susceptibility depends on neither
- 5. Susceptibility depends on both major genotype and polygenic number, age of onset depends on neither
- 6. Susceptibility depends on major genotype, age of onset depends on polygenic number alone

As for a binary trait without variable age of onset, information about prevalence (probability of having been affected since birth) can be incorporated into the likelihood, or estimated from the model fitted.

## 5.4 Program Input

File Type	Description
Parameter File	Specifies the parameters and options with with to
	perform a particular analysis.
Pedigree Data File	Contains delimited records for each individual in-
	cluding fields for identifiers, sex, parents, trait and
	marker data.

## 5.4.1 The segreg Parameter

The following table shows the main SEGREG syntax:

<pre>parameter [, attribute]</pre>	Explanation	
	Starts a SEGREG a	analysis block.
	Value Range	N/A
segreg	Default Value	None
	Required	Yes
	Applicable Notes	None
	Specifies the "root	" name to be used for output files.
	Output file names v	will be formed by concatenating the
	root name and an appropriate extension.	
out		An alphanumeric constant
, out	Value Range	representing a valid filename
	Default Value	"segreg"
	Required	No
	Applicable Notes	None

## 5.4.2 The segreg Parameter Block

The following lists all parameters that may occur in a SEGREG block (see note 1).

parameter	Employedian		
[, attribute]	Explanation		
	Specifies a title for the analysis		
	Value Range Character string		
title	Default Value None		
	Required No		
	Applicable Notes 1		
	Specifies the name of a primary trait. Must be the		
	name of a trait, covariate, or phenotype in the pedigree		
	data file or created by means of a function block.		
trait	Value Range Character string		
	Default Value None		
	Required Yes		
	Applicable Notes 1		
	Primary trait type		
	Value Range {continuous, binary, age_onset}		
	continuous, if trait is continuous		
, type	Default Value binary, if trait is binary		
	Required No		
	Applicable Notes 2		
	Starts a sub-block for specifying composite trait co-		
	variates.		
••• •••	Value Range N/A		
composite_trait	Default Value None		
	Required No		
	Applicable Notes 3		
	Starts a sub-block for specifying type means.		
	Value Range N/A		
type_mean	Default Value None		
	Required No		
	Applicable Notes 4, 19		
	Starts a sub-block for specifying type variances.		
	Value Range N/A		
type_var	Default Value None		
	Required No		
	Applicable Notes 5, 19		
	Starts a sub-block for specifying type susceptibilities		
	or penetrances.		
	Value Range N/A		
type_suscept	Default Value None		
	Required No		
	Applicable Notes 6, 20		

	Starts a sub-block for specifying covariates for the
	mean.
	Value Range N/A
mean_cov	Default Value None
	Required No
	Applicable Notes 7, 19
	Starts a sub-block for specifying covariates for the
	variance.
	Value Range N/A
var_cov	Default Value None
	Required No
	Applicable Notes 8, 19
	Starts a sub-block for specifying covariates for the
	trait susceptibility or penetrance.
suscept_cov	Value Range N/A
	Default Value None
	Required No
	Applicable Notes 9, 20
	Specifies the model class
	Value Range {A, D, FPMM, MLM}
	D, for continuous traits
class	Default Value MLM, for binary traits
	FPMM, for age-of-onset traits
	Required No
	Applicable Notes 10
	Starts a sub-block for specifying an FPMM model.
	Value Range N/A
fpmm	Default Value None
	Required No
	Applicable Notes None
	Starts a sub-block for specifying residual correlations
	(or associations).
	Value Range N/A
resid	Default Value None
	Required No
	Applicable Notes 11
	Starts a sub-block for specifying transformation op-
	tions.
	Value Range N/A
transformation	Default Value None
	Required No
	Applicable Notes 12, 19
	$\frac{12}{12}$

	Starts a sub-block for specifying the founder genotype frequency model.		
geno_freq	Value Range N/A		
geno_rreq	Default Value None		
	Required No		
	Applicable Notes 13		
	Starts a sub-block for the specifying the transmission		
	model		
transmission	Value Range N/A		
	Default Value None		
	Required No		
	Applicable Notes 14		
	Starts a sub-block for specifying the pedigree ascer-		
	tainment options.		
ascertainment	Value Range N/A		
ascertainment	Default Value None		
	Required No		
	Applicable Notes 15		
	Starts a sub-block for specifying the constraints on the		
	population prevalence of a binary trait.		
	Value Range N/A		
prev_constraints	Default Value None		
	Required No		
	Applicable Notes 16		
	Starts a sub-block for specifying population preva-		
	lence model parameters for a binary trait.		
prove ogtimete	Value Range N/A		
prev_estimate	Default Value None		
	Required No		
	Applicable Notes 17		
	Starts a sub-block for specifying output options.		
	Value Range N/A		
output_options	Default Value None		
	Required No		
	Applicable Notes 18		

Notes

- 1. Each of the title and trait parameters is defined by its own block. Except when a binary trait with variable age of onset is being analyzed, no sub-blocks are required (a commingling analysis is automatically performed in this case). Whenever a sub-block is included, there can be required parameters.
- 2. Only necessary if a trait with variable age of onset is being analyzed, or a binary trait is to be analyzed as a continuous trait.
- 3. The trait analyzed can be a linear function of the primary trait (with coefficient 1) and other

covariates whose coefficients are fixed or estimated. This linear function is called a composite trait. Without this sub-block a composite trait is not formed. All covariates are centered, the centering (average) value being included as part of the output. The covariates can be any covariate, phenotype or trait (other than the primary trait) listed in the pedigree data file. Note: This sub-block is not applicable to binary traits.

- 4. This sub-block refers to means of continuous traits. Without this sub-block, one, two and three types are fitted successively (see notes 2 and 3 following the type\_mean sub-block for an interpretation of the type means).
- 5. This sub-block refers to variances of continuous traits conditional on type. Without this sub-block, one common variance is fitted.
- 6. This sub-block refers to logits of susceptibilities (or of penetrances). Without this sub-block, one, two and three types are fitted successively (see notes 2 and 3 following the type\_suscept sub-block for an interpretation of the type susceptibilities). Note that it is not possible to fit more than one type when either the **no\_trans** or **homog\_no\_trans** option for transmission is used unless the model includes non-zero residual correlations.
- 7. This sub-block indicates which covariates are to (linearly) modify the means indicated in the type\_mean sub-block. Without this sub-block, no such covariates are included in the analysis. All covariates are centered, the centering (average) value being included as part of the output.
- 8. This sub-block indicates which covariates are to (linearly) modify the variances in the type\_var sub-block. Without this sub-block, no such covariates are included in the analysis. All covariates are centered, the centering (average) value being included as part of the output.
- 9. This sub-block indicates which covariates are to (linearly) modify the logits of susceptibilities (or of penetrances) indicated in the type\_suscept sub-block. Without this sub-block, no such covariates are included in the analysis. All covariates are centered, the centering (average) value being included as part of the output.
- The values of A and D denote Bonney's class A and D regressive models, respectively.
   FPMM is the finite polygenic mixed model. Without this parameter, a class D regressive model is used for continuous traits and a multivariate logistic model is used for binary traits.
   FPMM is the only option currently available for binary traits with variable age of onset.
- 11. This sub-block is not relevant for the FPMM (finite polygenic mixed model). Residual correlations are relevant for continuous traits and residual associations are relevant for binary traits. We use the term "correlations" to cover both situations. Without this sub-block, the usual genetic mixed model assumption of no marital correlation and equal sib-sib and parent-offspring correlations is used.
- 12. Without this sub-block, the Box-Cox power parameter that provides the best fit to a normal distribution (logistic distribution for age of onset) conditional on type is estimated. An error message will be returned if any value of the analysis trait is at any time necessarily negative. When a composite trait is being analyzed this is avoided as much as possible.
- 13. Without this sub-block, it is assumed that there is no genotype correlation between spouses, and that there are Hardy-Weinberg equilibrium proportions when fitting three types.

- 14. Without this sub-block, homogeneity across generations, no transmission, and Hardy-Weinberg equilibrium proportions are assumed.
- 15. Without this sub-block, it is assumed that the pedigrees are randomly sampled.
- 16. Without this sub-block, the estimate of population prevalence (more correctly, for a binary trait with variable age of onset, the probability of having been affected since birth) is not constrained by data extraneous to the pedigree file.
- 17. Without this sub-block, the population prevalence (for a binary trait with variable age of onset, the probability of having been affected since birth) is not calculated.
- 18. Without this sub-block, the output contains the overall ln(likelihood), -2ln(likelihood) and Akaike's AIC criterion for each of the models that has been maximized in a run.
- 19. This sub-block is not applicable to binary traits, but does apply to the age-of-onset distribution of a binary trait with variable age of onset.
- 20. This sub-block is not applicable to continuous traits.

## 5.5 Sub-Block Syntax: composite\_trait

The following table shows the syntax for the composite_trait sub-block (see note 1):	
--	--

<pre>parameter [, attribute]</pre>	Explanation		
	Specifies the name of a covariate used to form a com- posite trait as a linear function of the primary trait and the covariate. This parameter may be specified mul- tiple times. A covariate that is specified in this sub- block may not be also specified in a mean_cov sub- block.		
covariate	Value Range	Character string representing the name of a trait, covariate or phenotype in the pedigree data file or created by means of a function block.	
	Default Value	None	
	Required	No	
	Applicable Notes	1, 3 e of the covariate coefficient.	
	Value Range	$(-\infty, \infty)$	
, val	Default Value	$\frac{(-\infty, \infty)}{\text{None}}$	
, vai	Required	No	
	Applicable Notes	2	
		fix the covariate coefficient.	
	Value Range	{true, false}	
, fixed	Default Value	false	
	Required	No	
	Applicable Notes	2	

Notes

- 1. This sub-block is not relevant for a binary trait (with or without variable age of onset).
- 2. If the fixed attribute is set to **true**, the attribute val must be included. If set to **false** and the attribute val is included, this determines the initial value of the variable to be used in the maximization process. If set to **false** and the attribute val is not included, then the program supplies various initial values for the maximization process.
- 3. A particular trait may not be specified as both a mean covariate and composite trait.

## 5.6 Sub-Block Syntax: type\_mean

parameter		Fundamation
[, attribute]		Explanation
	Specifies type_mean option	
option		one
	Value Range Default Value Required	two
		three
		two_dom
		two_rec
		three_add
		three_dec
		three_inc
		one
		No
	Applicable Notes	1, 2, 3
Specifies the effect of a type. This parameter may		
	be specified as many times as necessary to indi-	
	cate the values appropriate for the option chosen.	
	Value Range	AA (means $\beta_{AA}$ )
		AB (means $\beta_{AB}$ )
mean		BB (means $\beta_{BB}$ )
		A* (means $\beta_{AA}=\beta_{AB}$ )
		B* (means $\beta_{BB} = \beta_{AB}$ )
		** (means $\beta_{AA} = \beta_{AB} = \beta_{BB}$ )
	Default Value	None
	Required	No
	Applicable Notes	3
	Specifies value of a given mean.	
	Value Range	$(-\infty, +\infty)$
	Default Value	None
, val	Required	No
		See note 2 of the
	Applicable Notes	composite_trait sub-
		block.
, fixed	Specifies option to fix the given value.	
	Value Range	true
		false
	Default Value	false
	Required	No
	Applicable Notes	See note 2 of the
		composite_trait sub-
		block.

The following table shows the syntax for the type\_mean sub-block:

Notes

1. This option refers to the number of types fitted to a continuous trait. Note that if a

type\_mean sub-block is not included, the program successively fits one, two and three types (see note 4 of the segreg block). On the other hand, if a type\_mean sub-block is included without specifying an option, then only one type is fitted. This sub-block is only relevant for continuous traits. It is relevant for the age of onset distribution of a binary trait with variable age of onset, but is not otherwise relevant for a binary trait.

- 2. When specified in this sub-block, the type effects are means of continuous distributions. For a binary trait with variable age of onset, they are the mean values of age of onset.
- 3. Denoting the three type effects  $\beta_{AA}$ ,  $\beta_{AB}$ , and  $\beta_{BB}$ , the options correspond to:

Option	Estimated or Fixed	
one	$\beta = \beta_{AA} = \beta_{AB} = \beta_{BB}$	
two	$\beta_1 = \beta_{AA} = \beta_{AB},  \beta_2 = \beta_{BB}$	
	$\beta_1 = \beta_{AA},  \beta_2 = \beta_{AB} = \beta_{BB}$	
three	$\beta_{AA}, \beta_{AB}, \beta_{BB}$	
two_dom	$\beta_{AA} = \beta_{AB}, \beta_{BB}$	
two_rec	$\beta_{AA}, \beta_{AB} = \beta_{BB}$	
three_add	$\beta_{AA}, \beta_{AB} = (\beta_{AA} + \beta_{BB}) / 2, \beta_{BB}$	
three_dec	$\beta_{AA} \geq \beta_{AB} \geq \beta_{BB}$	
three_inc	$\beta_{AA} \leq \beta_{AB} \leq \beta_{BB}$	

For example,

```
type_mean
{
    option=three_inc
    mean="A*", val=5.0, fixed=false
    mean="BB", val=12.0, fixed=false
}
```

sets initial estimates  $\beta_{AA} = \beta_{AB} = 5.0$  and  $\beta_{BB} = 12.0$  when estimating  $\beta_{AA} \le \beta_{AB} \le \beta_{BB}$ .

### 5.7 Sub-Block Syntax: type\_var

The following table shows the syntax for the type\_var sub-block (see note 1):

parameter		Explanation
[, attribute]		Explanation
	Specifies type_v	ar option
		one
		two
	Value Range	three
option	value range	two_dom
0,000000		two_rec
	-	three_add
	Default Value	one
	Required	No
	Applicable Notes	1,2
	-	t of a type. This parameter may
	-	any times as necessary to indi-
	cate the values app	propriate for the option chosen.
		AA (means $\sigma_{AA}^2$ )
		AB (means $\sigma^2_{AB}$ )
var	Value Range	BB (means $\sigma_{BB}^2$ )
Val	value runge	A* (means $\sigma_{AA}^2 = \sigma_{AB}^2$ )
	–	B* (means $\sigma_{BB}^{2A} = \sigma_{AB}^{2}$ )
		** (means $\sigma_{AA}^2 = \sigma_{AB}^2 = \sigma_{BB}^2$ )
	Default Value	None
	Required	No
	Applicable Notes	2
	Specifies value of t	
	Value Range	$[0,\infty)$
	Default Value	None
, val	Required	No
		See note 2 of the
	Applicable Notes	composite_trait sub-
		block.
	<u> </u>	fix the given value.
	Value Range	{true, false}
	Default Value	false
, fixed	Required	No
		See note 2 of the
	Applicable Notes	composite_trait sub-
		block.

Notes

 This sub-block is only relevant for continuous traits. It is relevant for the age of onset distribution of a binary trait with variable age of onset, but is not otherwise relevant for a binary trait. There can be at most one variance for each type specified in the type\_mean sub-block. 2. Denoting the three variances  $\sigma_{AA}^2$ ,  $\sigma_{AB}^2$  and  $\sigma_{BB}^2$ , the six options are analogous to the first six options in the type\_mean sub-block (see note 3 of the type\_mean sub-block) with  $\sigma^2$  replacing  $\beta$ .

For example,

```
type_var
{
    option=three
    var="AA", val=5.0, fixed=false
    var="B*", val=30.0, fixed=true
}
```

sets the initial estimate  $\sigma_{AA}^2$ =5.0 and fixed values  $\sigma_{BB}^2 = \sigma_{AB}^2$ =30.0, when estimating only  $\sigma_{AA}^2$ .

#### Sub-Block Syntax: type\_suscept 5.8

The following table shows the synta	x for the type_suscept sub-block:

parameter		Explanation
[, attribute]		-
	Specifies type_s	uscept option
		one
		two
		three
	Value Range	two_dom
option	value Mariye	two_rec
0001011		three_add
		three_dec
		three_inc
	Default Value	one
	Required	No
	Applicable Notes	1, 2, 3
	Specifies the effect	t of a type. This parameter may
	be specified as ma	iny times as necessary to indi-
	cate the values app	propriate for the option chosen.
		AA (means $\beta_{AA}$ )
		AB (means $\beta_{AB}$ )
quagopt	Value Bongo	BB (means $\beta_{BB}$ )
suscept	Value Range	A* (means $\beta_{AA} = \beta_{AB}$ )
		<b>B</b> * (means $\beta_{BB} = \beta_{AB}$ )
		** (means $\beta_{AA} = \beta_{AB} = \beta_{BB}$ )
	Default Value	None
	Required	No
	Applicable Notes	3, 4
	Specifies value of a	a given mean.
	Value Range	$(-\infty, +\infty)$
	Default Value	None
, val	Required	No
	-	See note 2 of the
	Applicable Notes	composite_trait
		sub-block.
	Specifies option to	fix the given value.
	Value Range	{true, false}
	Default Value	false
, fixed	Required	No
	-	See note 2 of the
	Applicable Notes	composite_trait
		sub-block.

Notes

1. This option refers to the number of types fitted to a binary trait. Note that if a type\_suscept sub-block is not included, the program successively fits one, two and three types (see note 6 of the segreg block). On the other hand, if a type\_suscept sub-block is included without specifying an option, then only one type is fitted.

- 2. The type effects are mean logits of susceptibilities (of penetrances for a binary trait).
- 3. Denoting the three type effects  $\beta_{AA}$ ,  $\beta_{AB}$ , and  $\beta_{BB}$ , the options correspond to:

Option	Estimated or Fixed
one	$\beta = \beta_{AA} = \beta_{AB} = \beta_{BB}$
two	$\beta_1 = \beta_{AA} = \beta_{AB}, \ \beta_2 = \beta_{BB}$
	$\beta_1 = \beta_{AA}, \ \beta_2 = \beta_{AB} = \beta_{BB}$
three	$\beta_{AA}, \ \beta_{AB}, \ \beta_{BB}$
two_dom	$\beta_{AA} = \beta_{AB}, \ \beta_{BB}$
two_rec	$\beta_{AA}, \ \beta_{AB} = \beta_{BB}$
three_add	$\beta_{AA}, \ \beta_{AB} = \frac{1}{2}(\beta_{AA} + \beta_{BB}), \ \beta_{BB}$
three_dec	$\beta_{AA} \ge \beta_{AB} \ge \beta_{BB}$
three_inc	$\beta_{AA} \le \beta_{AB} \le \beta_{BB}$

4. For example,

```
type_suscept
```

```
{
    option=three_inc
    suscept="A*", val= -1.0, fixed=false
    suscept="BB", val= -2.0, fixed=false
}
```

sets initial estimates  $\beta_{AA} = \beta_{AB} = 5.0$  and  $\beta_{BB} = 12.0$  when estimating  $\beta_{AA} \leq \beta_{AB} \leq \beta_{BB}$ .

### 5.9 Sub-Block Syntax: mean\_cov

parameter		Explanation
[, attribute]	Compieto to mo	life the many of a contin
		dify the mean of a contin-
	uous trait. This parameter may be speci-	
	fied multiple times. A covariate that is spec-	
	ified in this sub-block may not be used in	
	either a composite_trait sub-block or a	
	suscept_cov sub-block.	
covariate		Character string representing the
		name of a trait, covariate or phe-
	Value Range	notype from the pedigree data
		file or a name created by means
		of a function block.
	Default Value	None
	Required	No
	Applicable Notes	1, 3
	Specifies the value	of the covariate coefficient.
	Value Range	$(-\infty, +\infty)$
	Default Value	None
, val	Required	No
		See note 2 of the
	Applicable Notes	composite_trait sub-
		block.
	Specifies option to	fix the given value.
	Value Range	{true, false}
	Default Value	false
, fixed	Required	No
, lixed	-	3
	Applicable Notes	See note 2 of the
	Applicable Notes	composite_trait sub-
		block.
ĺ	Specifies whether	interaction effects are to be es-
	timated.	
	Value Range	{true, false}
, interaction	Default Value	false
	Required	No
	Applicable Notes	2

The following table shows the syntax for the mean\_cov sub-block:

Notes:

 The default is to include no covariates in the analysis. The means indicated in the type\_mean sub-block are a linear function of this covariate. All covariates are centered, the centering (average) value being included as part of the output. This sub-block is only relevant for continuous traits. It is relevant for the age of onset distribution of a binary trait with variable age of onset, but is not otherwise relevant for a binary trait. In the case of an age of onset distribution, the same covariate cannot be specified to modify both the mean and the variance, or both the mean and the susceptibility. In the case of a continuous trait, the same trait cannot be specified as both the mean and composite trait.

- 2. The interaction attribute refers to an interaction with type; the default is to assume no interaction. If there is no interaction, we estimate  $\beta_{AA}$ ,  $\beta_{AB}$ ,  $\beta_{BB}$  (as many as are specified in the type\_mean sub-block) and one overall "mean" covariate coefficient for each covariate. If there is interaction, then for this covariate we estimate an additional two interaction effects that sum to 0 if two  $\beta$  parameters are being fitted; and an additional three interaction effects that sum to 0 if three  $\beta$  parameters are being fitted.
- 3. When analyzing covariates for an age of onset model the mean\_cov sub-block is allowed, regardless of whether the dependent trait is continuous or binary. In this case, the covariates affect the continuous component of the age of onset model.

### 5.10 Sub-Block Syntax: var\_cov

parameter	Explanation
[, attribute]	Explanation
	Covariate to modify the variance (conditional on
	type) of a continuous trait. Allowable values are
	the names of traits, covariates or phenotypes in the
	pedigree data file or created by means of a func-
covariate	tion block. This parameter may be specified mul-
Covariace	tiple times.
	Value Range N/A
	Default Value None
	Required No
	Applicable Notes 1
	Specifies value of the covariate coefficient
	Value Range $(-\infty, +\infty)$
	Default Value None
, val	Required No
	See note 2 of the
	Applicable Notes composite_trait sub-
	block.
	Specifies option to fix the given value.
	Value Range {true, false}
	Default Value false
, fixed	Required No
	See note 2 of the
	Applicable Notes composite_trait sub-
	block.
	Specifies whether interaction effects to be esti-
	mated.
interaction	Value Range {true, false}
, interaction	Default Value false
	Required No
	Applicable Notes 2

The following table shows the syntax for the var\_cov sub-block:

- The default is to include no covariates in the analysis. The variances indicated in the type\_var sub-block are a linear function of this covariate . All covariates are centered, the centering (average) value being included as part of the output. This sub-block is only relevant for continuous traits. It is relevant for the age of onset distribution of a binary trait with variable age of onset, but is not otherwise relevant for a binary trait. In the case of an age of onset distribution, the same covariate cannot be specified to modify both the mean and the variance.
- 2. The interaction attribute refers to an interaction with type; the default is to assume no interaction. If there is no interaction, we estimate  $\sigma_{AA}^2$ ,  $\sigma_{AB}^2$ ,  $\sigma_{BB}^2$  (as many as are specified

in the type\_var sub-block) and one overall "var" coefficient for each covariate. If there is interaction, then for this covariate we estimate an additional two interaction effects that sum to 0 if two  $\sigma^2$  parameters are being fitted; and an additional three interaction effects that sum to 0 if three  $\sigma^2$  parameters are being fitted.

### 5.11 Sub-Block Syntax: suscept\_cov

parameter		Explanation
[, attribute]		-
	•	of a covariate coefficient to be used
	-	n logit of susceptibility, or of pen-
	,	r function of the covariate. A co-
	•	this sub-block may not be used in
		ock. This parameter may be speci-
	fied multiple times	
covariate		Character string representing the
	Value Range	name of a trait, phenotype or
	value runge	covariate listed within the
	-	pedigree data file.
	Default Value	None
	Required	No
	Applicable Notes	1
		of the covariate coefficient.
	Value Range	$(-\infty,\infty)$
	Default Value	None
, val	Required	No
		See note 2 of the
	Applicable Notes	composite_trait sub-
		block.
	Specifies option to	
	Value Range	{true, false}
	Default Value	false
, fixed	Required	No
	Applicable Notes	See note 2 of the composite trait sub-
	Applicable Notes	composite_trait sub- block.
	Specifies option to	assume interaction with type.
	Value Range	{true, false}
, interaction	Default Value	false
, Interaction	Required	No
	Applicable Notes	2
	Applicable Notes	<i>L</i>

The following table shows the syntax for the suscept\_cov sub-block:

Notes

1. The default is to include no covariates in the analysis. The suscept\_cov sub-block indicates which covariates are to modify the logits of susceptibilities or penetrances indicated in the type\_suscept sub-block. All covariates are centered, the centering (average) value being included as part of the output. In the case of an age-of-onset distribution, the same covariate cannot be specified as both the mean and the type susceptibility.

2. The "interaction" attribute refers to an interaction with type; the default is to assume no interaction. If there is no interaction, we estimate  $\beta_{AA}$ ,  $\beta_{AB}$ ,  $\beta_{BB}$  (as many as are specified in the type\_suscept sub-block) and one overall "suscept" covariate coefficient for each covariate. If there is interaction, then for this covariate we estimate an additional two interaction effects that sum to 0 if two  $\beta$  parameters are being fitted; and an additional three interaction effects that sum to 0 if three  $\beta$  parameters are being fitted.

## 5.12 Sub-Block Syntax: fpmm

parameter		Explanation
[, attribute]		-
	Specifies number of	
	Value Range	{1, 2, 3,, 9}
loci	Default Value	3
	Required	No
	Applicable Notes	None
		quency of polygenic loci.
	Value Range	(0,1)
freq	Default Value	0.5
	Required	No
	Applicable Notes	None
		nce of the polygenic variable.
	Value Range	N/A
var	Default Value	N/A
	Required	No
	Applicable Notes	None
		of the polygenic variance
	Value Range	$(0, +\infty)$
	Default Value	None
, val	Required	No
		See note 2 below.
	Applicable Notes	See also note 2 of the
		composite_trait sub-
		block.
		fix the given value.
	Value Range	{true, false}
	Default Value	false
, fixed	Required	No
		See note 2 below.
	Applicable Notes	See also note 2 of the
		composite_trait sub-
		block.
	Starts a sub-block of onset.	for a binary trait with variable age
	Value Range	N/A
onset	Default Value	N/A
	Required	No
	Applicable Notes	3
	FT	

The following table shows the syntax for the fpmm sub-block:

#### Notes

1. The following is an example of an fpmm sub-block,

fpmm {

}

```
loci=6
freq=.4
var, val=10.3, fixed=false
onset { # See onset sub-block below.
    type_dependent=A
    multi_dependent=N
    status=disease
    age_onset=age
    age_exam=age
}
```

2. The attribute val is optional. Thus, the following two lines are equivalent.

```
var, val=0.2, fixed=false
var=0.2, fixed=false
```

3. This sub-block, nested within the fpmm sub-block, is required to analyze a disease trait with variable age of onset: a bivariate trait in which one trait is binary (affected versus unaffected) and the other is continuous (age of onset) censored for unaffected persons. The current version of S.A.G.E. can only analyze such disease traits under the finite polygenic mixed model.

### 5.13 Sub-Block Syntax: onset

parameter		Explanation
[, attribute]		-
	Specifies what is d	ependent on type.
	Value Range	{A, S}
type_dependent	Default Value	А
	Required	No
	Applicable Notes	1, 2
	Specifies what is d	ependent on a polygenic com-
	ponent.	
multi dependent	Value Range	$\{N, A, S\}$
multi_dependent	Default Value	N
	Required	No
	Applicable Notes	2, 3
	Specifies the age of	f onset.
		Name of a continuous trait, co-
		variate or phenotype that is either
	Value Range	listed in the pedigree data file or
age_onset		created by means of a function
		block.
	Default Value	None
	Required	No
	Applicable Notes	4
	Specifies the age at	t exam.
		Name of a continuous trait, co-
		variate or phenotype that is either
	Value Range	listed in the pedigree data file or
age_exam		created by means of a function
	_	block.
	Default Value	None
	Required	No
	Applicable Notes	4

The following table shows the syntax for the onset sub-block:

Notes

- 1. The type\_dependent values have the following meanings:
  - $\mathbf{A} Age$  of onset depends on type.
  - **S** Susceptibility depends on type.

The option chosen will not cause any dependence on type if A is specified and the type\_mean sub-block specifies an option value of **one**, or if S is specified and the type\_suscept sub-block specifies an option value of **one**.

- 2. See note 1 of the fpmm sub-block for an example.
- 3. The multi\_dependent values have the following meanings:
  - N There is no polygenic component.
  - A Age of onset has a polygenic component.
  - S Susceptibility has a polygenic component.
- 4. It is permissible for the age\_onset and age\_exam parameters to specify the same continuous trait, in which case the value of this trait is assumed to be age of onset for affected persons and age at exam for unaffected persons.

## 5.14 Sub-Block Syntax: resid

parameter		Explanation
[, attribute]		-
	Specifies residual f	amilial correlations.
	Value Range	equal_po_ss
		equal_po
option		arb
	Default Value	equal_po_ss
	Required	No
	Applicable Notes	2, 3
	·	lation between the residuals of
	father and mother	
fm	Value Range	N/A
Lui	Default Value	N/A
	Required	No
	Applicable Notes	None
	Specifies value of t	the residual correlation ( $\rho$ ).
		(-1, +1), for cont. traits
	Value Range	N/A, for binary traits
	Default Value	0
, val	Required	No
	-	4,
	Applicable Notes	See note 2 of the
	Applicable Notes	composite_trait
		sub-block.
	Option to fix the gi	ven value.
	Value Range	{true, false}
	Default Value	true
, fixed	Required	No
, IIXed	-	4,
	Applicable Notes	See note 2 of the
	Applicable Notes	composite_trait
		sub-block.
	Specifies the correl	lation between the residuals of
	mother and offsprin	ng.
	Value Range	N/A
mo	Default Value	N/A
	Required	No
	Applicable Notes	None

The following table shows the syntax for the resid sub-block (see note 1):

	Specifies value of	the residual correlation ( $\rho$ ).
	Value Range	(-1, +1), for cont. traits
	Ū	N/A, for binary traits
	Default Value	0
, val	Required	No
,		4,
		See note 2 of the
	Applicable Notes	composite_trait
		sub-block.
	Option to fix the g	
	Value Range	{true, false}
	Default Value	false
, fixed	Required	No
, rinca	rtoquirou	See note 2 of the
	Applicable Notes	composite_trait
		sub-block.
	Specifies initial co	prrelation between the residuals
	of father and offsp	
	Value Range	N/A
fo	Default Value	
	Required	No
	Applicable Notes	None
	Specifies value of	the residual correlation ( $\rho$ ).
	Value Range	(-1, +1), for cont. traits
		N/A, for binary traits
	Default Value	0
, val	Required	No
		4,
	Applicable Notes	See note 2 of the
		composite_trait
		sub-block.
	Option to fix the g	
	Value Range	{true, false}
	Default Value	false
, fixed	Required	No
		See note 2 of the
	Applicable Notes	composite_trait
		sub-block.
	Specifies the corre	elation between the residuals of
	siblings.	
	Value Range	N/A
	Default Value	N/A
SS	Delault value	1V/A
SS	Required	No
SS		

	Specifies value of t	the residual correlation ( $\rho$ ).
	Value Range	(-1, +1), for cont. traits
		N/A, for binary traits
	Default Value	0
, val	Required	No
	-	4,
	Appliaghle Notes	See note 2 of the
	Applicable Notes	composite_trait
		sub-block.
		Sub block.
	Option to fix the gi	
	Option to fix the gi	
	0	iven value.
, fixed	Value Range	iven value. {true, false}
, fixed	Value Range Default Value	iven value.       {true, false}       false
, fixed	Value Range Default Value	<pre>iven value.  {true, false}  false  No</pre>

- 1. This sub-block is not relevant for the FPMM (finite polygenic mixed model). Residual correlations are relevant for continuous traits and residual associations are relevant for binary traits. We use the term "correlations" to cover both situations.
- 2. The default option value, **equal\_po\_ss**, corresponds to the usual genetic mixed model assumption of no marital correlation and equal sib-sib and parent-offspring correlations (only one of the parameters from among mo, fo and ss may be specified)
- 3. With the second value of the option parameter, **equal\_po**, mother-offspring and father-offspring correlations are equal while the father-mother (marital) correlation and sibling-sibling correlation are functionally independent of the parent-offspring correlation and of each other (fm and ss may be specified as well as either mo or fo). With the option value of **arb**, all four correlations: father-mother, mother-offspring, father-offspring, and sibling-sibling are functionally independent of each other and any combination of these correlations may have their attributes specified.
- 4. The residual value range of (-1, +1) is valid only when modeling continuous data. In the multivariate logistic model used for non-continuous data with residuals, the range is a calculated value that changes based on parameter estimates.

## 5.15 Sub-Block Syntax: transformation

The following table show	a the aunter for the	tranaformation	sub block (see note 1).
	5 נווכ צעוונמג וטו נווכ	LIANSLUIMALIUN	SUD-DIOCK (SCC HOLC I).

parameter	Exploration		
[, attribute]	Explanation		
	Specifies transformation type.		
		none	
	Value Range	box_cox	
option		george_elston	
	Default Value	box_cox	
	Required	No	
	Applicable Notes	None	
	Specifies of the por	<u> </u>	
	Value Range	N/A	
lambdal	Default Value	N/A	
	Required	No	
	Applicable Notes	None	
	Specifies the value	for $\lambda_1$ .	
	Value Range	$(-\infty, +\infty)$	
	Default Value	1.0	
, val	Required	No	
		See note 2 of the	
	Applicable Notes	composite_trait	
		sub-block.	
	Specifies option to fix $\lambda_1$ at the given value.		
	Value Range	{true, false}	
	Default Value	false	
, fixed	Required	No	
		See note 2 of the	
	Applicable Notes	composite_trait	
		sub-block.	
	Specifies lower bou	und for $\lambda_1$ .	
	Value Range	$(-\infty, +\infty)$	
, lower_bound	Default Value	-1	
	Required	No	
	Applicable Notes	None	
	Specifies upper bo	und for $\lambda_1$ .	
	Value Range	$(-\infty, +\infty)$	
, upper_bound	Default Value	$\infty$	
	Required	No	
	Applicable Notes	None	
	Specifies the shift j		
	Value Range	N/A	
lambda2	Default Value	N/A	
	Required	No	
	Applicable Notes	None	

	Specifies the value for $\lambda_2$ .	
	Value Range	$(-\infty, +\infty)$
	Default Value	0
, val	Required	No
	-	See note 2 of the
	Applicable Notes	composite_trait
		sub-block.
	Option to fix $\lambda_2$ at t	the given value.
	Value Range	{true, false}
	Default Value	true
, fixed	Required	No
	-	See note 2 of the
	Applicable Notes	composite_trait
		sub-block.

Notes

1. This block is not relevant for a binary trait.

## 5.16 Sub-Block Syntax: geno\_freq

parameter		
[, attribute]	Explanation	
	Specifies whether Hardy Weinberg equilibrium	
	proportions are to be assumed.	
	Value Range {hwe, nhwe}	
option	Default Value hwe	
	Required No	
	Applicable Notes 1	
	Specifies probability of a given genotype. This	
	parameter should be specified at most twice and	
	is ignored if the option value is set to hwe	
prob	Value Range {AA, AB, BB}	
	Default Value None	
	Required No	
	Applicable Notes 2	
	Specifies value for given probability type.	
	Value Range [0, 1]	
, val	Default Value None	
	Required No	
	Applicable Notes 2	
	Option to fix or not fix genotype probabilities or	
	the probability (relative frequency) of allele (com-	
	ponent) A.	
probs_fixed	Value Range {true, false}	
	Default Value false	
	Required No	
	Applicable Notes 3	
	Specifies the relative frequency of allele A. Used	
	when option value is set to hwe, or when	
	option is set to <b>nhwe</b> and no prob parameters	
frog A	are specified.	
freq_A	Value Range N/A	
	Default Value N/A	
	Required No	
	Applicable Notes None	
	Specifies value for the allele frequency.	
	Value Range (0,1)	
, val	Default Value None	
	Required No	
	Applicable Notes	

The following table shows the syntax for the geno\_freq sub-block:

Notes

1. The hwe option imposes Hardy-Weinberg equilibrium proportions, nhwe does not.

- 2. If two prob parameters are specified, their sum must be less than 1.
- 3. If **true**, sufficient information (val attributes of probs\_fixed or allele\_freq\_A, depending on the option chosen) must be specified to fully cover all probabilities. If **false** and a sufficient number of vals are included to specify all probabilities, they determine initial values of the probabilities. If **false** and a sufficient number of vals are not included, the program supplies the necessary initial values for the maximization process.

## 5.17 Sub-Block Syntax: transmission

The fellering table above	Ale a an interest for a the a	+	a.h hlaal. ()	· · · · · · · · · · · · · · ·
The following table shows	the syntax for the	rrangmiggion	SUD-DIOCK (S	see note i r
	the syntax for the		Sub block (	

parameter	Explanation		
[, attribute]			
	Specifies transmission type.		
		homog_no_trans	
		homog_mendelian	
	Value Range	homog_general	
option		tau_ab_free	
		general	
	Default Value	no_trans	
		homog_no_trans	
	Required	<u>No</u>	
	Applicable Notes	2, 3, 4	
		ssion probability. The tau pa-	
	• •	becified as many times as nec-	
	model chosen.	the appropriate values for the	
	model chosen.		
		AA (means $\tau_{AA}$ )	
ton	Value Range	AB (means $\tau_{AB}$ )	
tau		BB (means $\tau_{BB}$ )	
		A* (means $\tau_{AA} = \tau_{AB}$ ) B* (means $\tau_{BB} = \tau_{AB}$ )	
		** (means $\tau_{BB} = \tau_{AB}$ )	
	Default Value	None	
	Required	No	
	Applicable Notes	None	
	Specifies a value for		
	Value Range	[0, 1]	
	Default Value	None	
, val	Required	No	
	· .	See note 2 of the	
	Applicable Notes	composite_trait	
		sub-block.	
	Option to fix the given value.		
	Value Range	{true, false}	
	Default Value	false	
, fixed	Required	No	
	-	See note 2 of the	
	Applicable Notes	composite_trait	
		sub-block.	

	Option to remove the range restriction on the transmission probabilities when the option value is set to either <b>homog_general</b> or <b>general</b> .	
no_bounds	Value Range	N/A
	Default Value	None
	Required	No
	Applicable Notes	6

- 1. This sub-block can only be used if two or three distinct types are specified in either the type\_mean or type\_suscept sub-block. If this sub-block is missing and a type\_mean or type\_suscept sub-block is included that specifies two or three types, then all of the option values of this sub-block, with the exception of **no\_trans**, are automatically performed.
- 2. Defining the transmission probability  $\tau_j$  to be the probability that a person of type j transmits A, and  $q_A$  to be the relative frequency of A, these options correspond to:

Option	Estimated or Fixed
homog_no_trans	$\tau_{AA} = \tau_{AB} = \tau_{BB} = q_A$
homog_mendelian	$\tau_{AA} = 1, \tau_{AB} = .5, \tau_{BB} = 0$
homog_general	$0 \leq  au_{AA},$
	$ au_{BB} \le 1$
	$\tau_{AB} = (q_A - q_A^2 \tau_{AA} - (1 - q_A)^2 \tau_{BB})/2q_A(1 - q_A)$
general	$0 \le \tau_{AA}, \tau_{AB}, \tau_{BB} \le 1$
tau_ab_free	$\tau_{AA} = 1, 0 \le \tau_{AB} \le 1, \tau_{BB} = 0$

- 3. For the 3 "homogeneous" options **hwe** must be specified in the geno\_freq sub-block (or, equivalently, a geno\_freq sub-block must not be included).
- 4. This default is appropriate for commingling analysis with the assumption of Hardy-Weinberg equilibrium proportions.
- 5. This option together with the option value of **hwe** in the geno\_freq sub-block will give the same result as the option value of **homog\_no\_trans**.
- 6. Does not apply to a tau parameter for which fixed = true or to user-specified initial values. The initial values of the val attribute, if specified, must always lie in the closed interval [0, 1].

## 5.18 Sub-Block Syntax: ascertainment

parameter	Explanation		
[, attribute]			
	Specifies the subset of persons on whom ascer-		
	tainment conditioning is performed.		
		none	
	Value Range	founders	
cond_set		psf	
	_	founders_plus_psf	
	Default Value	<b>psf</b> if psf_indic is given a	
		valid value, <b>none</b> otherwise.	
	Required	No	
	Applicable Notes	1	
	Specifies the proba	and sampling frame indicator.	
		of a trait, covariate or pheno-	
		ntinuous, in the pedigree data	
psf_indic	file or created by m	neans of a function block.	
psi_mare	Value Range	Character string.	
	Default Value	None	
	Required	No	
	Applicable Notes	1	
	Value of the proba	and sampling frame indicator	
	that is interpreted to mean an individual is in-		
	cluded in the proband sampling frame. May be re-		
	peated as many times as needed. Any other value		
	of the proband sampling frame indicator, includ-		
psf_indic_include	ing a missing value, means that the individual is		
	not part of the proband sampling frame.		
	Value Range	$(-\infty, +\infty)$	
	Default Value	1	
	Required	No	
	Applicable Notes	None	
	Specifies how a trai	t value is used to determine the	
	conditioning on a person's phenotype.		
		actual	
	Value Range	gte_thresh	
cond_val	raide rainge	lte_thresh	
	-	thresh_indic	
	Default Value	actual	
	Required	No	
	Applicable Notes	2, 3	

The following table shows the syntax for the ascertainment sub-block:

	Threshold value to be used if cond_val is <b>gte_thresh</b> or <b>lte_thresh</b> . If not specified, the value of thresh is estimated by the program.
, thresh	Value Range $(-\infty, +\infty)$
	Default Value None
	Required No
	Applicable Notes None
	Specifies the value for the greater-than-or-
	equal-to threshold if cond_val is set to
	thresh_indic. If not specified, the value of
	thresh_indic_high is estimated by the pro-
, thresh_indic_high	gram.
	Value Range $(-\infty, +\infty)$
	Default Value None
	Required No
	Applicable Notes 4
	Specifies the value for the less-than-or-
	equal-to threshold if cond_val is set to
	thresh_indic. If not specified, the value of
	thresh_indic_high is estimated by the
, thresh_indic_low	program.
	Value Range $(-\infty, +\infty)$
	Default Value None
	Required No
	Applicable Notes 4
	Specifies the threshold indicator variable. Must be
	the name of a continuous trait, covariate or pheno-
	type in the pedigree data file or created by means
thread india	of a function block.
thresh_indic	Value Range Character string
	Default Value None
	Required No
	Applicable Notes 5
	Specifies the cutoff value for using or not using
	thresh_indic
- la l-	Value Range $(-\infty, +\infty)$
, thresh	Default Value 0
	Required No
	Applicable Notes 5
	Specifies the type of conditioning when a binary
	trait with variable age of onset is being analyzed.
	Value Range {actual, by_onset}
onset_option	Default Value actual
	Required No
	•

Notes

- 1. This parameter determines whose phenotypes are conditioned on (the "conditioned subset") when calculating a conditional likelihood that allows for ascertainment as follows:
  - A value of **none** indicates that unconditional likelihoods are calculated (i.e. no correction for ascertainment).
  - A value of **psf** indicates the members of the pedigree proband sampling frame and is only permissible if a psf\_indic parameter is included in the sub-block.
  - A value of **founders** indicates all founder members of the pedigree.
  - A value of **founders\_plus\_psf** indicates all the founder members and the members of the pedigree proband sampling frame, and is only permissible if a psf\_indic parameter is included in the sub-block.
- 2. The cond\_val parameter is relevant for continuous traits only, and is ignored for binary traits, composite traits, and for age of onset models (for which the onset\_option parameter should be used). In the case of binary and composite traits, the default value of **actual** is always used. Also, **actual** is the value used for all founders not included in the proband sampling frame.

Value	Meaning
actual	Indicates that conditioning is on the actual trait value.
gte_thresh	Indicates that conditioning is on the trait value being greater than
	or equal to a threshold value.
lte_thresh	Indicates that conditioning is on the trait value being less than or
	equal to a threshold value.
thresh_indic	Indicates that for each person an indicator variable determines
	whether to apply the value <b>gte_thresh</b> or <b>lte_thresh</b> .

3. The meanings of the values of cond\_val are as follows:

- 4. If the value (specified or estimated) of thresh\_indic\_low is greater than the value of thresh\_indic\_high, a warning message is printed.
- 5. If cond\_val is set equal to the value **thresh\_indic**, then the value of the threshold indicator variable determines, separately for each individual, which cond\_val option to apply. The threshold indicator variable should:
  - be equal to thresh for those individuals for whom **actual** is to be applied.
  - be greater than or equal to thresh for those individuals for whom **gte** is to be applied.
  - be less than or equal to thresh for those individuals for whom **lte** is to be applied.
- 6. This parameter is required if a binary trait with variable age of onset is being analyzed (unless random sampling is to be assumed). If set equal to **actual**, the likelihood is conditioned on the binary phenotype and actual age of onset for each member of the conditioned subset, if available, otherwise by the age at exam. If the value **by\_onset** is specified, the likelihood is conditioned on the binary phenotype of each member of the conditioned subset and by the actual age of onset, if available, otherwise by the age at exam. However, **actual** is the value used for all founders not included in the proband sampling frame.

### 5.19 Sub-Block Syntax: prev\_constraints

The following table shows the syntax for the prev\_constraints sub-block:

<pre>parameter [, attribute]</pre>	Explanation	
	Starts a sub-block	k for specifying a particular
	prevalence constraint. May be repeated as many	
	times as needed.	
constraint	Value Range	N/A
	Default Value	None
	Required	No
	Applicable Notes	1

Notes:

1. Beginning with S.A.G.E. version 4.6, this represents a change from previous versions for the specification of the prev\_constraints parameter (see 5.19.1).

#### 5.19.1 Sub-Block Syntax: constraint

parameter	Explanation	
[, attribute]	-	
covariate	Specifies a covariate on which prevalence (prob-	
	ability of having been affected since birth) de-	
	pends. Allowable values are the names of traits,	
	covariates or phenotypes in the pedigree data file	
	or created by means of a function block. This pa-	
	rameter may be specified multiple times.	
	Value Range Character string	
	Default Value None	
	Required No	
	Applicable Notes 1, 2, 5	
	Specifies a value for the covariate.	
	Value Range $(-\infty, +\infty)$	
, val	Default Value None	
	Required Yes	
	Applicable Notes 2	
	Specifies the number of affected persons in a ran-	
	dom sample.	
<b>D</b>	Value Range (0, N)	
R	Default Value None	
	Required No	
	Applicable Notes 3, 5	
	Specifies the sample size.	
	Value Range $(R, +\infty)$	
N	Default Value None	
	Required No	
	Applicable Notes 3, 5	
	Specifies the age at which prevalence (probabil-	
	ity of having been affected since birth) should be	
	computed. Required for age of onset traits, and	
	disallowed otherwise.	
age	Value Range $(0, +\infty)$	
	Default Value None	
	Required No	
	Applicable Notes 2, 4, 5	

The following table shows the syntax for the constraint sub-block:

- 1. Any covariate in this sub-block must also appear in the mean\_cov, var\_cov or suscept\_cov sub-blocks.
- 2. Any covariate (including age) upon which prevalence depends and which is not specified as a covariate parameter, or for which a value is not specified, is set at its mean value.

- 3. It is assumed that, independent of the pedigree data, R of N persons are affected by the specified values of the covariates. R and N need not be integers.
- 4. The literal string **infinity** must be entered to indicate an "infinite" age.
- 5. The following example illustrates the constraint syntax:

```
segreg, out = myAnalysis {
   trait = BMI, type = continuous
   trait = aff, type = age_onset
   .
   prev_constraints {
      constraint {
         covariate = height
         covariate = weight
         age = infinity
         N = 1000
         R = 100
      }
      constraint {
         covariate = smoking
         covariate = drinking
         age = infinity
      }
      •
   }
}
```

### 5.20 Sub-Block Syntax: prev\_estimate

parameter [, attribute]	Explanation	
covariate	Specifies a covariate on which prevalence de-	
	pends. Allowable values are the names of traits,	
	covariates or phenotypes in the pedigree data file	
	or created by means of a function block. This pa-	
	rameter may be specified multiple times.	
	Value Range Character string	
	Default Value None	
	Required No	
	Applicable Notes 1, 2	
, val	Specifies a value for the covariate.	
	Value Range $(-\infty, +\infty)$	
	Default Value None	
	Required No	
	Applicable Notes 2	
	Specifies the age at which prevalence (probabil-	
	ity of having been affected since birth) should be	
	computed. Required for age of onset traits, and	
	disallowed otherwise.	
age	Value Range $(0, +\infty)$	
	Default Value None	
	Required No	
	Applicable Notes 2, 3	

The following table shows the syntax for the prev\_estimate sub-block:

- 1. Any covariate in this sub-block must also appear in the mean\_cov or suscept\_cov sub-blocks. Age of onset (or age at exam) may also be included as a covariate if an onset sub-block is included, and then prevalence is interpreted as the probability of having been affected since birth up to the specified age.
- 2. Any covariate upon which prevalence depends, but is not specified as a covariate parameter is set at its mean value.
- 3. The literal string **infinity** may be entered to indicate an "infinite" age.

### 5.21 Sub-Block Syntax: output\_options

<pre>parameter [, attribute]</pre>	Explanation	
type_prob	Specifies option to calculate type probabilities.	
	Value Range	{true, false}
	Default Value	false
	Required	No
	Applicable Notes	1
pen_func_out	Specifies option to create penetrance function out-	
	put file.	
	Value Range	{true, false}
	Default Value	false
	Required	No
	Applicable Notes	2

The following table shows the syntax for the output\_options sub-block:

- 1. Type probabilities can only be calculated if two or three types are specified in the type\_mean sub-block. In either case three probabilities (summing to 1) are output for an individual: the probabilities of being AA, AB or BB conditional on the model and all the pedigree information available, substituting maximum likelihood estimates for all unknown parameters.
- 2. This file can be used as input to LODLINK or MLOD. Because this only makes sense if the transmission option **homog\_mendelian** has been chosen, it will only be produced if that option is among those chosen in the transmission sub-block; otherwise, the option is ignored.

### 5.22 Program Execution

SEGREG is run via a command line interface on the supported UNIX and Windows platforms. This requires that the S.A.G.E. programs are properly installed and in the current execution path. Input files are specified on the command line and all output files are created in the current working directory.

Running SEGREG from the command prompt with no arguments, or the wrong number of arguments, will result in the program printing its usage statement. This lists the input files the program requires on the command line:

```
>segreg
S.A.G.E. v5.x -- SEGREG
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
usage: ./segreg <parameters> <pedigree>
Command line parameters:
parameters - parameter file
pedigree - pedigree data file
```

As indicated in the program usage statement, input files are listed on the command line. A typical run of SEGREG may look like the following:

```
>segreg segreg.par example.ped
S.A.G.E. v5.x -- SEGREG
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
Reading parameter file.....done.
Reading pedigree file.....done.
from example.ped.....done.
Sorting pedigrees.....done.
Generating statistics.....done.
Analysis complete!
```

### 5.23 Program Output

Output files produced by SEGREG containing results and diagnostic information are:

File Name	File Type	Description
segreg.inf	Information output file	Contains informational diagnostic messages,
		warnings and program errors. No analysis results
		are stored in this file.
segreg.sum	SEGREG summary output file	Contains the table of final estimates of the param-
		eters and their standard errors and other results.
segreg.det	SEGREG detailed output file	Contains the table of final estimates and variance-
		covariance matrix of the parameters.
segreg.typ	Trait genotype probability output file	Contains the individual specific type probabilities
		conditional on the model and all the pedigree in-
		formation available, and is suitable for input into
		model-based linkage programs such as MLOD
		and LODLINK.
segreg.pen	Trait penetrance function output file	Contains individual specific penetrance infor-
		mation. Will only be produced if the ho-
		mog_mendelian option of the transmission
		parameter has been enabled.

#### 5.23.1 Information Output File

The SEGREG Information file contains a variety of useful information, including:

- Information on fields read from the pedigree data file. These tables, which provide information about what the program has read from the pedigree data file, are included with all programs in S.A.G.E. Release and are very useful for debugging most common errors caused when reading the pedigree data. When first analyzing new data, it is recommended that these tables be checked carefully to make sure pedigree data are being correctly read.
- Information, warning and error messages generated throughout the program. It is recommended that this file be checked for warning and error messages before examining the results of any run of the program. The program attempts to correct many common errors and this sometimes means analyses are not as expected. The file "segreg.inf" should be checked for errors and diagnostic information after each run of the program.

#### 5.23.2 Summary Output File

The SEGREG summary output file stores the table of final estimates of the parameters with model information.

#### 5.23.3 Detailed Output File

The SEGREG Detailed output file stores the variance-covariance matrix as well as what the Summary File has.

### 5.24 Example Output File

#### 5.24.1 Summary Output File

Here is a typical example of a SEGREG summary output file.

```
_____
SEGREG Analysis for Trait : dbh
_____
  # Model Specification
      Model Class A
                                            : three means
      Type means
      Type variances
                                            : one variance
      Genotype frequency
                                            : Hardy-Weinberg equilibrium
      Residual correlations
                                           : no spouse correlation,
                                              parent-offspring and sib-sib correlations equal
      Transmission
                                            : homogeneous mendelian
      Transformation
                                            : Box-Cox
      Covariate means
          cov1 : (Mean = 0.16146)
          cov2
                     : (Mean = 0.68576)
                     : (Mean = 0.15278)
          cov3
  # Number of constituent pedigrees
                                        : 0
                                           : 4
  # Number of singletons
  # Final Estimates :
                     Parameter Est. Standard Err. First Deriv. Status
  Parameter
_____
                                                                              _____

        8.37728311
        1.58425845
        0.00001649
        IND, MAY VARY

        24.83059248
        2.25001207
        0.00000000
        IND, MAY VARY

        43.13638675
        1.93512951
        0.00001281
        IND, MAY VARY

        105.3057234
        14.03526432
        -0.00000393
        IND, MAY VARY

  mean_AA
  mean_AB
  mean_BB
  variance
  prob_AA
                      0.16311761
                                       0.05099851
                                                                       DEPENDENT
  prob_AB
                      0.48152120
                                       0.02427498
                                                                       DEPENDENT
                                      0.07527349 DEPENDENT
0.06313600 0.00027621 IND-FN, MAY VARY
0.00000000 FIXED EXTERNALLY
                      0.35536119
0.40387821
  prob_BB
  freq_A
  genotypic corr. 0.0000000
marital resid c 0.0000000
                                      0.0000000
                                                                       FIXED EXTERNALLY
                                      0.00000000
0.13603358 0.00027621 IND, MAY VARY
0.00000000 FIXED EXTERNAL
  po = ss resid c
                     0.07655435
  transm prob_AA
                     1.0000000
0.5000000
0.0000000
                                                                        FIXED EXTERNALLY
  transm prob_AB
                                        0.00000000
                                                                         FIXED EXTERNALLY
  transm prob_BB
lambda_one
lambda_two
                                       0.00000000
                                                                        FIXED EXTERNALLY
                                      0.05940720 -0.00025587 IND, MAY VARY
                      0.53976055
                      0.0000000
                                                                       FIXED EXTERNALLY
                                       0.00000000

        -3.71081358
        13.74889789
        0.00007443
        IND, MAY VARY

        1.13614517
        13.54930873
        -0.00012156
        IND, MAY VARY

        1.58174041
        13.87732080
        0.00000000
        IND, MAY VARY

                      -3.71081358
1.13614517
  covl
  cov2
  cov3
_____
        LN(Likelihood) : -1244.44
     -2 LN(Likelihood) : 2488.87
    Akaike's AIC score : 2526.87
_____
```

#### 5.24.2 Detailed Output File

Here is a typical example of a SEGREG detailed output file.

```
_____
SEGREG Analysis for Trait : dbh
_____
  # Model Specification
       Model Class A
       Type means
                                                : three means
       Type variances
                                                 : one variance
       Genotype frequency
                                                 : Hardy-Weinberg equilibrium
       Residual correlations
                                                : no spouse correlation,
                                                   parent-offspring and sib-sib correlations equal
       Transmission
                                                 : homogeneous mendelian
       Transformation
                                                 : Box-Cox
       Covariate means

      cov1
      : (Mean =
      0.16146)

      cov2
      : (Mean =
      0.68576)

      cov3
      : (Mean =
      0.15278)

  # Number of constituent pedigrees : 4
: 0
  # Number of singletons
  # Final Estimates :
•
  # Variance-Covariance Matrix :
      mean_AA mean_AB mean_BB
                                                                           variance . . . . . . . .
_____
 mean_AA2.509874832.167873160.821682467.45485406mean_AB2.167873165.062554331.2840008911.11488621mean_BB0.821682461.284000893.74472620-5.88210645variance7.4548540611.11488621-5.88210645196.9886444prob_AA0.040149030.070662680.043846160.10221457prob_AB0.019110690.033635010.020870500.04865352prob_BB-0.05925972-0.10429769-0.06471666-0.15086809freg_A0.049704380.087480190.054281410.12654133po = ss res0.063032920.076465950.007605230.82768163lambda_one0.038471060.032402270.036101660.03125240cov1-0.92646184-1.530581960.91183742-6.9737559cov20.147071610.20870282-0.245579580.17034505
                      0.14707161
                                       0.20870282 -0.24557958
  cov2
               0.17034505
  cov3
              -0.41230018
                                      -0.36967265
                                                          -0.55733522
                                                                            -2.43238363
          _____
        LN(Likelihood) : -1244.44
      -2 LN(Likelihood) : 2488.87
    Akaike's AIC score : 2526.87
_____
```

## **Chapter 6**

# MARKERINFO

MARKERINFO detects Mendelian inconsistencies in pedigree data. These inconsistencies are sorted by marker, by pedigree, and by whether one or more than one nuclear family is involved in the inconsistency.

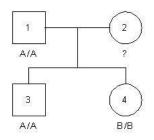
#### 6.1 Limitations

MARKERINFO analyses one marker at a time and is only guaranteed to detect all errors in the absence of loops. Mendelian inconsistencies cannot be localized beyond the nuclear family in which they are first detected (see theory).

#### 6.2 Theory

The phenoset of an individual is the set of all genotypes consistent with that individual's phenotype. Individuals labeled as missing are considered to be consistent with all possible phenotypes. MARKERINFO detects Mendelian inconsistencies in pedigree data by reducing the set of possible genotypes for each individual to the minimal possible subset on the basis of both the individual's phenoset and the phenosets of surrounding individuals. An empty minimal subset of genotypes for any individual indicates a Mendelian inconsistency.

Example 1

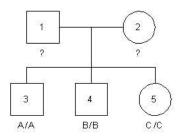


In this example pedigree, individuals 1 and 3 have a phenoset consisting of genotype A/A, while individual 4 has a phenoset consisting of genotype B/B. Individual 2 is unknown, so her phenoset includes all possible genotypes: {A/A, A/B, B/B, etc.}

These phenosets are reduced based on Mendelian inheritance from parent to child. Under Mendelian inheritance, a parent having marker genotype A/B can transmit either the A or the B allele to the child, but cannot transmit any other allele at that marker. Any genotype for which there is no valid transmission from a parent or to a child is removed from the phenoset. In this way, the subset of possible genotypes for individual 2 becomes A/B and that for individual 1 become empty.

MARKERINFO detects two sorts of inconsistencies, those involving one, and those involving more than one, nuclear family. In the above example, there is no valid transmission from individual 1 to individual 4 because 4 must receive a B allele from both parents and 1 has no B allele. In this and the next example it is sufficient to inspect a single nuclear family to detect an inconsistency.

Example 2:

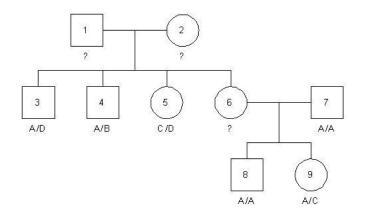


In this example, each of the children must receive a different set of alleles from each of their parents, but each parent has only two alleles. At least one child must be inconsistent with the parents, but it is impossible to determine which one.

Inconsistencies Involving More than one Nuclear Family

Often, a single nuclear family appears consistent until new information is added from surrounding nuclear families. Consider example 3.

Example 3:



Looking at only the nuclear family with parents 1 and 2, we see that this family is consistent, with 1 and 2 each having subset of possible genotypes A/C, B/D. Note here that if 1 is A/C, 2 must be B/D and vice versa. From this, we can deduce that the subset of possible genotypes for 6 is A/D, A/B, C/D, B/C.

Similarly, the nuclear family with parents 6 and 7 is consistent, with the subset of possible genotypes for 6 being A/C. However, A/C is not present in the subset of possible genotypes for 6 as derived from the first nuclear family. There is no genotype present in both subsets, so the minimal subset is empty. Because the sequence in which MARKERINFO traverses the pedigree depends on several factors, the inconsistency could be first detected in either of the nuclear families, and only one of them will be reported as being inconsistent.

# 6.3 Program Input

MARKERINFO requires the following input files in order to run:

File Type	Description	
MARKERINFO parameter file	Specifies the parameters and options with which	
	to perform a particular analysis.	
Pedigree data file	Contains delimited records for each individual in-	
	cluding fields for identifiers, parents, trait and	
	marker data.	

## 6.3.1 The markerinfo Parameter

The following syntax table specifies the permissible parameter and attribute settings for the main MARKERINFO parameter.

<pre>parameter [, attribute]</pre>		Explanation
	Starts MARKERIN	NFO block
	Value Range	N/A
markerinfo	Default Value	None
	Required	Yes
	Applicable Notes	None
	Specifies the root name to be used for output files.	
	Output file names v	will be formed by concatenating the
	root name and an a	ppropriate extension.
		Character string representing a
, out	Value Range	valid file name.
	Default Value	markerinfo
	Required	No
	Applicable Notes	None

### 6.3.2 The markerinfo Block

parameter	Explanation	
[, attribute]		
	Specifies an extra I	D field to be printed in the analysis
	output file.	
		Character string representing the
sample_id	Value Range	name of a field in the pedigree
Sampre_ra		data file.
	Default Value	None
	Required	No
	Applicable Notes	1
	Specifies that con	nsistent nuclear family members
	should be added to	the output.
	Value Range	true
consistent_out	value Mariye	false
	Default Value	false
	Required	No
	Applicable Notes	2

The following lists all parameters that may occur in a markerinfo block.

Notes

- 1. The value of sample\_id should be set equal to the name of a field read from the pedigree data file. This can be used to indicate the location where a sample is stored.
- 2. If consistent\_out is set to **true**, then the nuclear family members who are not inconsistent are added to the output with [] around them.

# 6.4 **Program Execution**

MARKERINFO is run via a command line interface on the supported UNIX and Windows platforms. This requires that the S.A.G.E. programs are properly installed and in the current execution path. Input files are specified on the command line and all output files are created in the current working directory.

Running MARKERINFO from the command prompt with no arguments, or the wrong number of arguments, will result in the program printing its usage statement.

```
>MARKERINFO
S.A.G.E. v5.x -- MARKERINFO
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
usage: markerinfo >parameters> <pedigree>
Command line parameters:
   parameters - parameter file
   pedigree - pedigree data file
```

As indicated in the program usage statement, input files are listed on the command line. A typical run of MARKERINFO may look like the following:

```
>markerinfo par example.ped
S.A.G.E. v5.x -- MARKERINFO
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
Reading Parameter File......done.
Reading pedigree file.......done.
Sorting pedigrees......done.
Markerinfo analysis.....
Processing pedigree '1'.....done.
Processing pedigree '102'.....done.
Processing pedigree '104'.....done.
Processing pedigree '105'.....done.
Processing pedigree '105'.....done.
Processing pedigree '106'.....done.
Processing pedigree '108'.....done.
```

# 6.5 Program Output

Output files produced by MARKERINFO containing results and diagnostic information are:

Filename	Filetype	Description
markerinfo.inf	Information output file	Contains informational diagnostic messages, warnings
		and program errors. No calculation results are stored in
		this file.
markerinfo.out	Analysis output file	Contains Mendelian inconsistency information on mark-
		ers (See note)

Note:

Two types of Mendelian inconsistencies are differentiated: those which occur within a single nuclear family, and those in which members of more than one nuclear family are involved - i.e. the inconsistency can only be detected if two or more nuclear families are simultaneously examined. In the latter case, only one of the nuclear families that could be involved is shown in the output, followed by \*.

### 6.5.1 Information Output File

The MARKERINFO information file contains a variety of useful information, including:

- Information on fields read from the pedigree data file. These tables, which provide information about what the program has read from the pedigree data file, are included with all programs in S.A.G.E. Release and are very useful for debugging most common errors caused when reading the pedigree data. When first analyzing new data, it is recommended that these tables be checked carefully to make sure pedigree data are being correctly read.
- Information, warning and error messages generated throughout the program. It is recommended that you check this file for warning and error messages before examining the results

of any run of the program. The program attempts to correct many common errors and this sometimes means analyses are not as expected. The file "markerinfo.inf" should be checked for errors and diagnostic information after each run of the program.

# 6.6 Example Output File

Here is a typical example of MARKERINFO output:

```
S.A.G.E. v5.x -- MARKERINFO
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
MARKERINFO Analysis Output
_____
_____
 Part 1: Number of Inconsistencies per pedigree/marker
_____
_____
                  Number of Markers
              Incon. Informative Total
Pedigree
_____
                     323
321
                 3
124
                            324
155
                  3
                             324
______
Number of Pedigrees
               Incon. Informative Total
Marker
_____
              _____
                    _____
                           _____
                    94 94
F13A1
                 61
1s225
                 3
                       93
                             94
D1S245
                 2
                        94
                             94
                       94
                             94
D14S608
                 1
                       94
D2S1328
                 1
                             94
D2S1334
                 1
                       91
                             94
_____
 Part 2: Inconsistencies
_____
 missing code = 0
 * More than one nuclear family must be examined to detect
  the inconsistency.
_____
Table 1 with Marker F13A1 1s225 D1S245
_____
Pedigree Individual F13A1 1s225 D1S245
             -----
-----
                           _____
  2 Mother 0
1 Father 1/1
2
2
2
     3
2
     4
               4/5
   2
        Mother
                    * 0
                           * 0
124
     1
                    * 0
                           * 0
124
          Father
                    * 2/9
                          * 4/6
124
     4
                    * 8/13
124
     5
                           * 1/4
124
    6
                    * 0
                           * 0
```

•

•

: Table 2 with Marker D14S608 D2S1328 D2S1334 Pedigree Individual D14S608 D2S1328 D2S1328 D2S1334 66 16 Mother \* 3/3 66 17 Father \* 0 66 27 \* 3/5 66 28 \* 3/4 66 29 \* 3/4 .

# **Chapter 7**

# FREQ

FREQ is a program that estimates allele frequencies from marker data among related individuals with known pedigree structure and generates marker locus description files, needed by GENIBD, MLOD, and other S.A.G.E. programs. Future versions will also have the ability to estimate genotype and haplotype frequencies in the presence of allelic disequilibrium.

# 7.1 Limitations

Maximum likelihood estimates of allele frequencies can only be calculated using information from pedigrees without mating rings or other loops. Any pedigrees with loops will automatically be skipped for maximum likelihood estimation. Sometimes numerical problems occur and standard errors of the frequency estimates cannot be calculated. Also, the computational time required to calculate maximum likelihood frequency estimates increases greatly with the number of alleles at any locus.

# 7.2 Theory

### 7.2.1 Initial Frequency Estimator

FREQ begins its analysis by computing allele frequencies using only founders and singletons (unrelated and unconnected individuals) from each pedigree for all codominant marker phenotypes (Boehnke 1991). These estimates are calculated by summing the number of times each allele appears and dividing by the total number of observed alleles. This estimator tends to be sub-optimal since much of the data are not used and, in many datasets, founders are not typed.

A second estimator is provided that attempts to use marker information from non-founders and nonsingletons by assuming that they are independent. Calculation is performed the same way as for the founders, by counting the number of times each allele appears and dividing by the total number of observed alleles. These estimates can be reported directly, or combined with the founder-only based estimates by giving the founder\_weight parameter a value. When the founder weight is not set, the founder and non-founder frequencies are combined by adding the number of times each allele appears in both founders and non-founders and dividing by the total number of observed alleles from both. When founder\_weight is set to a number between 0 and 1, say w, then a weighted average of the founder and non-founder frequencies is taken, with weights w and 1 - w, respectively. Setting founder\_weight to 1 generates founder-only frequency estimates, while setting founder\_weight to 0 results in non-founder-only frequency estimates.

These methods provide consistent but statistically inefficient frequency estimates which can be used for datasets that have many pedigrees with loops or markers with too many alleles for the frequencies to be computed efficiently, as well as automatically provide initial estimates for maximum likelihood estimation.

## 7.2.2 Maximum Likelihood Estimator

The likelihood formulation assumes that, with respect to the marker loci, the pedigrees are randomly ascertained from a single random mating population, and that genotypes occur with Hardy-Weinberg equilibrium frequencies. The likelihood for the data at each marker in the whole sample is numerically maximized over possible allele frequencies to obtain the maximum likelihood estimates for that marker. Standard errors are computed by double differentiation of the log likelihood. Those frequencies that maximize the likelihood are then reported. Non-codominant markers are fully supported, provided that the phenotype to genotype mapping is provided in a locus description file. It should be noted that singletons (unrelated and unconnected individuals) may be included in the data; they are simply one-person pedigrees with parent information missing and, as such, require no special treatment in the model.

# 7.3 Program Input

File Type	Description	
FREQ Parameter File	Specifies the parameters and options with which	
	to perform a particular analysis.	
Pedigree Data File	Contains delimited records for each individual in-	
	cluding fields for identifiers, sex, parents, trait and	
	marker data.	
Marker Locus Description File	Lists the alleles, allele frequencies and phenotype	
	to genotype mapping for each marker locus. Only	
	needed in FREQ for non-codominant markers.	

FREQ requires the following input files in order to run:

# 7.3.1 The freq Parameter

The following syntax table specifies the permissible parameter and attribute settings for the main FREQ parameter.

<pre>parameter [, attribute]</pre>		Explanation
	Starts FREQ analysis sub-block.	
	Value Range	N/A
freq	Default Value	None
	Required	Yes
	Applicable Notes	None
	Specifies the root name to be used for output files.	
	Output file names v	will be formed by concatenating the
	root name and an a	ppropriate extension.
		Character string representing a
, out	Value Range	valid file name.
	Default Value	freq.out
	Required	No
	Applicable Notes	None

## 7.3.2 The freq Sub-Block

The following syntax table specifies the permissible parameter and attribute settings for the freq sub-block.

<pre>parameter [, attribute]</pre>	Explanation	
	The weight used f	for founders to combine founder-
	only and approxir	nate non-founder frequency esti-
	mates.	
founder_weight	Value Range	[0, 1]
	Default Value	None
	Required	No
	Applicable Notes	1
	Specifies whether	to skip maximum likelihood esti-
	mate computation of allele frequencies.	
altin mla	Value Range	{true, false}
skip_mle	Default Value	false
	Required	No
	Applicable Notes	None
	Names a marker to be included in the current analysis.	
		Character string representing the
marker	Value Range	name of a marker listed in the
		pedigree data file.
	Default Value	None
	Required	No
	Applicable Notes	2

Notes

- 1. This parameter is useful when consistent (but inefficient) estimates are required from a dataset with many alleles. When not specified, the estimates labeled as "All Pedigree Members" are obtained on the assumption that all observed alleles are independent.
- 2. The value of a marker parameter should be set to the name of a marker for which allele frequencies are to be estimated. If no valid marker parameters are listed, then all markers are used.

# 7.4 Program Execution

FREQ is run via a command line interface on the supported UNIX and Windows platforms. This requires that the S.A.G.E. programs are properly installed and in the current execution path. Input files are specified on the command line and all output files are created in the current working directory.

Running FREQ from the command prompt with no arguments, or the wrong number of arguments, will result in the program printing its usage statement. This lists the input files the program requires on the command line:

>freq S.A.G.E. v5.x -- FREQ COPYRIGHT (C) 2002 CASE WESTERN RESERVE UNIVERSITY usage: freq <parameters> <pedigree> [locus] Command line parameters: parameters - Parameter File pedigree - Pedigree Data File locus - Locus Description File (optional)

As indicated in the program usage statement, input files are listed on the command line. A typical run of FREQ may look like the following:

```
> freq params ped
S.A.G.E. v5.x -- FREO
COPYRIGHT (C) 2002 CASE WESTERN RESERVE UNIVERSITY
Reading Parameter File.....done.
Reading Pedigree File.....
from ped.....done.
Sorting pedigrees.....done.
Estimating allele frequencies. (default analysis)
Detailed output file: freq.out
Locus description file: freq.loc
Processing marker: D5G1
Processing marker: D5G2
Processing marker: D5G3
Processing marker: D5G4
Processing marker: D5G5
Processing marker: D5G6
Processing marker: D5G7
Processing marker: D5G8
Processing marker: D5G9
Processing marker: D5G10
Processing marker: D5G11
Processing marker: D5G12
Processing marker: D5G13
Processing marker: D5G14
Processing marker: D5G15
Processing marker: D5G16
Processing marker: D5G17
Processing marker: D5G18
```

### 7.5 Program Output

Output files produced by FREQ containing results and diagnostic information are:

Filename	Filetype	Description
freq.inf	Information output file	Contains informational diagnostic messages,
		warnings and program errors. No analysis results
		are stored in this file.
freq.out	Analysis output file	Contains detailed tables of analysis options and
		results.
freq.loc	Locus Description	An output file that presents the allele frequencies
	File	estimated by FREQ in a form that may be read
		directly into any other S.A.G.E. program that re-
		quires it.

### 7.5.1 Information Output File

The FREQ information file contains a variety of useful information, including:

- Information on fields read from the pedigree data file. These tables, which provide information about what the program has read from the pedigree data file, are included with all programs in S.A.G.E. and are very useful for debugging most common errors caused when reading the pedigree data. When first analyzing new data, it is recommended that these tables be checked carefully to make sure pedigree data are being correctly read.
- Information, warning and error messages generated throughout the program. It is recommended that you check this file for warning and error messages before examining the results of any run of the program. The program attempts to correct many common errors and this sometimes means analyses are not as expected. The file "freq.inf" should be checked for errors and diagnostic information after each run of the program.

## 7.5.2 Locus Description File

Allele frequency estimates are output into the locus description file according to the following priority:

- 1. Maximum likelihood estimates
- 2. Weighted estimates
- 3. Non-weighted/naive estimates

For example, if skip\_mle = **true**, and founder\_weight is set to some nonzero value, then the locus description file will contain weighted estimates. On the other hand, if skip\_mle is set equal to **false**, then the locus description file will contain maximum likelihood estimates.

### 7.5.3 Example Output File

Here is a typical example of FREQ output:

S.A.G.E. v4.3 -- FREQ COPYRIGHT (C) 2002 CASE WESTERN RESERVE UNIVERSITY File generated on : Tue Oct 8 16:55:13 2002 Estimating allele frequencies. (default analysis) \_\_\_\_\_ Locus description file: freq.loc Allele frequency estimates for marker: D5G1 Frequency Estimates: Founders All Pedigree Maximum Standard Allele Members Likelihood Error Only \_\_\_\_\_ \_\_\_\_ 0.2918410 0.2912371 0.2918410 ( 0.0147032) В С 0.3033473 0.2993986 0.3033473 (0.0148680) D 0.3012552 0.3058419 0.3012552 (0.0148388) А 0.0575314 0.0640034 0.0575314 (0.0075311) 0.0460251 0.0395189 0.0460251 (0.0067770) E Allele frequency estimates for marker: ABO Marker was determined to be non-codominant, so only Maximum Likelihood estimates were computed. Frequency Estimates: Maximum Standard Likelihood Error Allele ----- -----0.1894737 ( 0.0290427) Α1 A2 0.0754243 ( 0.0212204) В 0.0428382 ( 0.0145766) 0 0.6922637 (0.0345212) Allele frequency estimates for marker: PGD Frequency Estimates: Founders All Pedigree Maximum Standard Members Likelihood Error Only Allele \_\_\_\_\_ \_\_\_\_ 0.98373980.96994990.9727053( 0.0089518)0.01626020.03005010.0272947( 0.0089518) А С Allele frequency estimates for marker: C3 Frequency Estimates: Founders All Pedigree Maximum Standard Allele Only Members Likelihood Error \_\_\_\_\_ \_\_\_\_ 0.20325200.21285480.1972428(0.0217400)0.79268290.78380630.7941248(0.0220807) F S 0.0040650 0.0033389 0.0086324 (0.0049628) F′

# **Chapter 8**

# GENIBD

GENIBD is a program for generating identity-by-descent (IBD) sharing distributions from family data and genetic marker loci by a variety of algorithms tuned for various types of pedigrees. Three methods of generating IBD sharing are provided: the Single Marker IBD Analysis (single-point only), the Exact IBD Analysis (single- or multi-point), and the Simulation IBD Analysis (single- and multi-point). Control of which algorithms are used in a given analysis is provided to the user through convenient automatic switching parameters. IBD sharing distributions are generated for five types of relative pairs: sibling, half sibling, avuncular, grandparental and cousin. In future versions there will be options to generate more types of pairs. The resulting output file(s) list at each location the probability of each pair sharing 0 or 2 alleles IBD, and the difference between the paternal and maternal probability of sharing 1 allele IBD, conditional on the marker data available. This file can then be read into other programs (e.g. SIBPAL) for analyses.

# 8.1 Limitations

IBD sharing for only five pair types can be generated:

- 1. full sib,
- 2. half sib,
- 3. grandparental,
- 4. avuncular and
- 5. first cousin.

There are three methods currently implemented that generate IBD sharing distributions. Each method has distinct capabilities and limitations:

#### 8.1.1 Single Marker IBD Analysis

The Single Marker IBD Analysis uses complete information at each single marker to generate the IBD distributions for each pair of relatives at that marker. It is strictly a single-point method, and does not support pedigrees with loops.

#### 8.1.2 Exact IBD Analysis

The Exact IBD Analysis computes the likelihood of each inheritance vector at one or several markers (including locations interpolated between markers) to generate IBD distributions for each pair of the five supported types of relative pairs at each marker. It can be used for either single- or multi-point analysis in pedigrees with or without loops and is not restricted in the type of relative pair for whom IBD is computed. It is, however, restricted to small pedigrees due to the exponential nature of the algorithm related to the number of individuals in the pedigree. The time and space complexity of the algorithm is largely characterized by the exponent 2n - f, the number of bits in an inheritance vector, where n is the number of non-founders and f is the number of founders in a pedigree. During parameter specification the maximum value of 2n - f may be set; any pedigree that has a value larger than the limit will use another of the analysis methods, if possible, or be skipped.

#### 8.1.3 Simulation IBD Analysis

The Simulation IBD analysis uses a Markov chain Monte Carlo (MCMC) simulation over the space of possible inheritance vectors for each pedigree to estimate the IBD distribution for each pair of the five supported pair types at each marker, without interpolation at locations between markers. Several batches are run to ensure coverage of the state space. Generation of IBD distributions at points between markers can be accomplished putting markers with no data at those locations.

Also note that since this is a simulation method, values differ between runs of the program. This method may be quite time consuming so it is only used when pedigrees are too large for the exact IBD analysis.

## 8.2 Theory

Let  $\hat{f}_{imj}$  be the probability, conditional on the marker data available, that relative pair j shares exactly i alleles IBD at marker m, where i=0, 1 or 2. GENIBD calculates  $\hat{f}_{imj}$  for each marker locus of interest for each of five types of relative pair in the data set as follows.

Given the marker data  $I_m$  for a single pedigree at marker m

$$\hat{f}_{imj} = \frac{P\left(I_m \mid pair \, j \, shares \, i \, alleles \, IBD\right) P\left(pair \, j \, shares \, i \, alleles \, IBD\right)}{L\left(I_m\right)} \tag{8.1}$$

or

$$\hat{f}_{imj} = \frac{P\left(I_m \mid pair \, j \, shares \, i \, alleles \, IBD\right) P\left(pair \, j \, shares \, i \, alleles \, IBD\right)}{L\left(I_m\right)} \tag{8.2}$$

where  $L(I_m)$  is the likelihood for the pedigree at marker m and Pr(pair j shares i alleles IBD) is the prior probability that depends on relationship alone.  $L(I_m)$  does not depend on the individual pair and is thus only calculated once for each pedigree at each marker locus.

In the case of full sib and half sibs, for i = 1 and pair j, the components  $\hat{f}_{1mj-maternal}$  and  $\hat{f}_{1mj-paternal}$  of  $\hat{f}_{1mj}$  are calculated separately, depending on the sex of the parent from whom

the sharing allele is descended, as follows:

$$\hat{f}_{1mj-maternal} = \frac{P(I_m | pair j shares 1 maternal allele IBD) P(pair j shares 1 maternal allele IBD)}{L(I_m)}$$
$$\hat{f}_{1mj-paternal} = \frac{P(I_m | pair j shares 1 paternal allele IBD) P(pair j shares 1 paternal allele IBD)}{L(I_m)}$$

or

$$\hat{f}_{1mj-maternal} = \frac{P(I_m, pair j shares 1 maternal allele IBD)}{L(I_m)}$$
$$\hat{f}_{1mj-paternal} = \frac{P(I_m, pair j shares 1 paternal allele IBD)}{L(I_m)}$$

The difference  $(\hat{f}_{1mj-maternal} - \hat{f}_{1mj-paternal})$  is reported in the GENIBD output for every marker location, denoted in the output as f1m-f1p.

The methods used to calculate these values depend on the type of analysis used.

#### 8.2.1 Single Marker Analysis

In the case of single marker analysis, only information at a single locus is used, with  $L(I_m)$  calculated using the recursive methods described in Fernando, Stricker and Elston (1993).

To calculate  $f_{imj}$  for sib pairs, we use equation 8.1, while for other pair types we use equation 8.2. For sib pairs, we use the *counting* method suggested by Amos, Dawson and Elston (1990). To evaluate equation 8.2 for other pair types, we condition upon a set of individuals in the pedigree that includes the pair and a chain of individuals connecting the pair genetically. This chain includes the parents of each member of the pair and the parents shared by any two individuals already in the chain [See Amos, Dawson and Elston (1990) for more detail.] We know that

$$P(I_m, pair j \ shares \ i \ alleles \ IBD) = \sum_{g \in G} P(I_m, pair \ j \ shares \ i \ alleles \ IBD, \ g),$$

where G is the set of all possible genotype configurations of the individuals in the conditioned set. We therefore calculate  $L(I_m, \text{ pair } j \text{ shares } i \text{ alleles i.b.d.},g)$  for each possible genotype configuration in G. We use the recursive methods of Fernando, Stricker and Elston (1993) to calculate the likelihood for the sections of the pedigree not in the conditioned set and reuse them for each likelihood calculation.

#### 8.2.2 Exact IBD Analysis

The exact IBD analysis is used for both single- and multi-point analysis. It uses the exact multipoint algorithm to generate likelihoods of inheritance vectors at target locations. These likelihoods are then summed separately for inheritance vectors corresponding to a given pair sharing 0, 1, and 2 alleles IBD.

#### 8.2.2.1 The Exact Multi-point Algorithm

The general algorithm used by MLOD and GENIBD to generate multi-point likelihoods and other statistics is called the exact multi-point algorithm. This algorithm takes a chromosomal region and generates likelihoods of all the possible inheritance patterns at each marker in the region. These likelihoods can then be combined to generate identity-by-descent statistics.

### 8.2.2.2 Single-point IBD Sharing

For single-point, a likelihood vector is generated for each marker of interest. For each inheritance pattern, the number of alleles shared by a given inheritance pattern can be determined by tracking which founder alleles each pair of individuals receives. By summing the likelihoods of all inheritance patterns that share a specific number of alleles IBD, and dividing by the total likelihood of the pedigree at that marker (equation 8.2 above), we obtain the probability of the pair sharing that number of alleles IBD.

### 8.2.2.3 Multi-Point IBD Sharing

The multi-point algorithm is essentially the same as single-point. For each location of interest along the chromosome, we generate a multi-point likelihood vector incorporating all the information provided by the markers. This vector can then be summed, as in the single-point case above, to give us the multi-point probability of sharing 0, 1 and 2 alleles IBD.

### 8.2.3 Simulation IBD Analysis

The simulation IBD analysis uses a modified Sobel and Lange (Sobel and Lange, 1996) algorithm to generate random inheritance patterns at each marker in the state space. A multi-point likelihood for all markers is generated, assuming no crossover interference. For each generated state, IBD values are noted. Heuristic methods are used to determine the number of states to be generated, as well as the number of batches and how much dememorization to perform.

#### 8.2.3.1 Calculating the Amount of Simulation

By default, GENIBD determines the amount of simulation to perform for each pedigree. It does this by multiplying the number of individuals in the pedigree by the number of markers in the region being simulated. This number is then multiplied by several factors, one each for the number of dememorization steps per batch, the number of simulation steps per batch, and the number of batches. The default factors have been set, based upon extensive in-house testing, to the following:

dememorization steps per batch	15
simulation steps per batch	150
batch factor	30

These values have been found to be sufficient in most cases, but may be changed.

# 8.3 Program Input

File Type	Description
GENIBD Parameter File	Specifies the parameters and options with which
	to perform a particular analysis.
Pedigree Data File	Contains delimited records for each individual in-
	cluding fields for identifiers, sex, parents, trait and
	marker data.
Marker Locus Description File	Lists the alleles, allele frequencies and phenotype
	to genotype mapping for each marker locus.
Genome Description File	Contains a description of the linked marker re-
	gions, including distances between markers. This
	file is not required for single-point analysis.

## 8.3.1 The genibd Parameter

<pre>parameter [, attribute]</pre>		Explanation
	Starts GENIBD an	alysis block.
	Value Range	N/A
genibd	Default Value	N/A
	Required	Yes
	Applicable Notes	None
	Specifies the root name to be used for output	
	Output file names v	vill be formed by concatenating the
	root name and an a	ppropriate extension.
		Character string representing a
, out	Value Range	valid file name.
	Default Value	None
	Required	No
	Applicable Notes	None

The following syntax table specifies the permissible parameter and attribute settings for the main GENIBD parameter.

Notes

The out attribute controls the filenames generated by the analysis. For each region in the analysis, a separate IBD file is generated. These filenames are in the format: "*out\_region.ibd*" where *out* is the value of the out attribute, and *region* is the region name. If no out attribute is specified, the analysis title is used instead.

## 8.3.2 The genibd Block

The following syntax table specifies the permissible parameter and attribute settings for the genibd block.

<pre>parameter [, attribute]</pre>		Explanation
	Specifies title of the run	
	Value Range	Character string.
title	Default Value	None
	Required	Yes
	Applicable Notes	None

region	A region to be analyzed. As many region parameters ters may be specified as required to specify which re- gions are to be analyzed. If no region parameters are specified, all regions in the genome descrip- tion file are analyzed. Regions with no valid markers (as may be the case when data for only one or two chromosomes are present in the pedigree data file) are skipped.Value Range RequiredCharacter string.Default Value RequiredNoneNoneNone
output_pair_types	Specifies pair types to be generated         siblings         Value Range       all_sibs         relatives         Default Value       all_sibs         Required       No         Applicable Notes       1
max_pedigree	The largest $2n - f$ value to be processed for a pedi- gree while performing exact multi-point or single- point analysis.Value Range Default Value $\{1, 2, 3,\}$ Default Value Required18NoNoApplicable NotesNone
scan_type	Indicates whether to compute IBD sharing at the ob- served markers or at the markers and intervals be- tween them.          Value Range       {markers, intervals}         Default Value       markers         Required       No         Applicable Notes       None
,distance	Sets the interval, in cM, to use as basis for computing         IBD sharing probabilities between observed markers.         Only applicable when value of scan_type parameter is set to intervals.         Value Range       [0, +INF]         Default Value       2.0         Required       No         Applicable Notes       None
allow_loops	Allows pedigrees with loops to be processed while performing single-point analysis.         Value Range       {true, false}         Default Value       false         Required       No         Applicable Notes       None

	Selects either single- or multi-point IBD generation.		
	Value Range	{singlepoint, multipoint}	
ibd_mode	Default Value	multipoint	
	Required	No	
	Applicable Notes	2	
	Option to allow p	edigrees that are too large for the	
split_pedigrees	exact analysis to b	e split into nuclear families before	
	processing. Setting	the value to <b>always</b> means that all	
	pedigrees will be split in this fashion.		
	Value Range	{yes, no, always}	
	Default Value	no	
	Required	No	
	Applicable Notes	None	
	Starts a sub-block for specifying simulation options.		
	Value Range	{yes, no, always}	
simulation	Default Value	yes	
use_simulation	Required	No	
	Applicable Notes	3	

Notes

- 1. output\_pair\_types may be set to any of three values: **siblings** if only full sibling pairs are desired, **all\_sibs** if both full and half sibling pairs, or **relatives** if all five relative pair types (sibs, half sibs, avuncular, grand-parental and cousin) are desired.
- 2. If **singlepoint** is selected, only data at each marker are used to calculate the IBD sharing at a marker. If **multipoint** is selected, all the marker data in a region is used to calculate the IBD sharing at each point, assuming no interference.
- 3. The simulation sub-block allows simulation on pedigrees that are too large for the exact methods. Setting the value of this parameter to **always** means that all pedigrees will use simulation. For example:

```
genibd, out = autism_study_01 {
   title = "Autism Study #1: IBD Results"
   region = "Chrom1"
   ibd_mode = multipoint
   scan_type = intervals, distance = 1.0
   simulation = always {
     use_factoring = true
     sim_steps = 100000
   }
}
```

### 8.3.3 Sub-Block Syntax: simulation

The following syntax table specifies the permissible parameter and attribute settings for the simulation sub-block.

parameter	Explanation	
[, attribute]	-	
sim_local_marker	The proportion of times during simulation that a marker adjacent to the current marker being simulated is chosen for simulation during the next simulation step. Value Range [0, 1]	
	Default Value     0.75       Required     No       Applicable Notes     None	
use_factoring	Controls whether the simulation scaling factors are used. If they are not used, simulation uses a constant number of steps regardless of pedigree size.         Value Range       {true, false}         Default Value       true         Required       No         Applicable Notes       1	
base_factor	The base scaling factor provides a method of adjusting all three scaling factors together. Will be ignored if use_factoring is set to false.Value Range $[0, \infty)$ Default ValueNoneRequiredNoApplicable NotesNone	
demem_factor	The dememorization scaling factor. This controls the number of dememorization steps done during each batch. Will be ignored if use_factoring is set to false. Will be set to 0.5 x base_factor if base_factor > 0.Value Range Default Value $\{1, 2, 3,\}$ Default Value $15$ RequiredNo 1	
sim_factor	The simulation step scaling factor. This controls the number of simulation steps during each batch. Will be ignored if use_factoring is set to false. Will be set to 10 x base_factor if base_factor > 0.Value Range Default Value150 Required Applicable Notes	

sim_batch_factor	The simulation batch count scaling factor. This controls the number of batches of simulation to perform.Will be ignored if use_factoring is set to false.Will be set to base_factor if base_factor > 0.Value Range $\{1, 2, 3,\}$ Default Value30RequiredNoApplicable Notes1
sim_steps	The number of simulation steps during each batch.         Value Range       {1, 2, 3,}         Default Value       200000         Required       No         Applicable Notes       1
demem_steps	The number of dememorization steps during each batch.         Value Range       {1, 2, 3,}         Default Value       50000         Required       No         Applicable Notes       1
batch_count	The number of batches of simulation to perform.         Value Range       {1, 2, 3,}         Default Value       100         Required       No         Applicable Notes       1

Notes

- 1. When calculating identity-by-descent values by simulation, it is usually unnecessary to specify the amount of simulation to be performed. GENIBD does this automatically for each pedigree being analyzed. However, an option to specify the amount of simulation is provided. There are two methods of doing this:
  - The first, called *factoring*, calculates the amount of dememorization, the amount of simulation, and the number of batches based upon pedigree size and number of markers in the region being simulated. It is selected by setting use\_factoring to **true** (default). The user may set the value of base\_factor (which automatically determines the values of demem\_factor, sim\_factor and sim\_batch\_factor as described in the syntax table) or may set the values of demem\_factor, sim\_factor and sim\_batch\_factor directly.
  - The second method uses the same number of steps and batches for every pedigree. It is used when use\_factoring is set to **false**. Setting demem\_steps, sim\_steps, and batch\_count parameters sets, respectively, the amount of dememorization per batch, the amount of simulation per batch, and the number of batches.

# 8.4 **Program Execution**

GENIBD is run via a command line interface on the supported UNIX and Windows platforms. This requires the S.A.G.E. programs to be properly installed and in the current execution path. Input files are specified on the command line and all output files are created in the current working directory.

Running GENIBD from the command prompt with no arguments, or the wrong number of arguments, will result in the program printing its usage statement. This lists the input files the program requires on the command line including any that are optional.

```
>GENIBD
S.A.G.E. v5.x -- GENIBD
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
usage: ./GENIBD <parameters> <pedigree> <locus> [map]
Command line parameters:
parameters - parameter file
pedigree - pedigree data file
locus - locus description file
map - Genome Map File (optional for single point analysis)
```

As indicated in the program usage statement, input files are listed on the command line. A typical run of GENIBD may look like the following:

```
>GENIBD data.par data.ped data.loc data.map
S.A.G.E. v5.x -- GENIBD
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
Loading Map file ...
Validating Analysis...
ANALYSES
_____
Single-point : Singlepoint analysis on region (CHR5):
* - Pedigrees with loops will not be processed.
Multi-point : Multipoint analysis on region (CHR5):
* - Maximum size for exact method is set to 18.
* - Pedigrees larger than 18 will be simulated.
* - Pedigrees will be simulated based on pedigree size and the number
of markers.
Processing Analyses....
_____
Single-point : Singlepoint analysis on region (CHR5):
* - Pedigrees with loops will not be processed.
Processing Region: CHR5
_____
Single-point: Pedigree 1
Generating Single Point Likelihoods.....Done.
Single-point: Pedigree 10
Generating Single Point Likelihoods.....Done.
Single-point: Pedigree 100
Generating Single Point Likelihoods.....Done.
```

# 8.5 Program Output

File Name	File Type	Description
GENIBD.inf	Information output file	Contains informational diagnostic messages,
		warnings and program errors. No analysis results
		are stored in this file.
genome.inf	Genome Information	Contains diagnostic information on the genetic
	File	map data and the marker loci that were provided
		for analysis. No analysis results are stored in this
		file.
analysis_region.ibd	IBD sharing files	There is one IBD sharing file for each region pro-
		cessed by each analysis performed by GENIBD.
		These files contain the IBD distribution of each
		pair of relatives for each marker in the analysis
		(see 2.7).

GENIBD produces several output files that contain results and diagnostic information:

### 8.5.1 Information Output File

The GENIBD information file contains a variety of useful information, including:

- Information on fields read from the pedigree data file. These tables, which provide information about what the program has read in, are included with all programs in S.A.G.E. and are very useful for debugging many common errors caused when reading the pedigree data. When first analyzing new data, it is recommended that these tables be checked carefully to make sure pedigree data are being correctly read.
- Information, warning and error messages generated throughout the program. It is recommended that you check this file for warning and error messages before examining the results of any run of the program. The program attempts to correct many common errors and this sometimes means analyses are not run as expected. The file "genibd.inf" should be checked for errors and diagnostic information after each run of the program.

### 8.5.2 Genome Information File

This file includes warnings and errors produced while parsing the marker locus description file, as well as a table for each marker listing allele and genotype population frequencies, assuming Hardy-Weinberg equilibrium. If allele frequencies do not sum to 1.0, they are standardized to 1.0, so these frequencies may not be identical to those in the marker locus description files.

### 8.5.3 IBD Sharing Files

The IBD sharing file stores the IBD probability distribution of allele-sharing identical-by-descent between pairs of individuals at specific locations.

# 8.6 Example Output File

The IBD sharing file is generated as output from GENIBD and is used as input to other programs, such as SIBPAL. It contains the following information (see 2.7):

- a list of the markers at which the IBD sharing distributions are generated.
- a table that contains a line for each relative pair and the probabilities of sharing 0 or 2 alleles at each marker (designated as  $f_0$  and  $f_2$ , respectively). Additionally, the table shows the value of the difference  $f_{1m} f_{1p}$  to support analysis of parent-of-origin effects. The table includes up to five types of relative pairs: sibling, half sibling, avuncular, grand-parental and cousin.

In our example, two IBD sharing files are generated, one using single- and one using multi-point analysis. Although the numerical results are generally different, the files are similar in structure. The following is a portion of the single-point file:

IBD File 1.9 : This File is automatically generated. Do NOT edit! #\_\_\_\_\_ # ANALYSIS #----= Analysis 1 scan\_type = intervals
allow\_loops = off
ibd\_mode = multipoint, exact split\_pedigrees = no use\_simulation = no # MARKERS #----11s1984 0.0 11\_2.0 2.0 11\_4.0 4.0 11s2362 6.0 11\_8.0 8.0 11\_10.0 10.0 11\_12.0 12.0 11s1999 14.0 #Pedigree Ind 1 Ind 2 11s1984 f0, 11s1984 f1m-f1p, 11s1984 f2, ... #----- ---- ----- ------ ------. . . . . . . . • • • • • • • . . . . ... . . . . .

# **Chapter 9**

# RELTEST

RELTEST helps classify pairs in a (sib pair) linkage study according to their true relationship using genome scan data. It is based on a Markov process model of allele-sharing along chromosomes. The program currently performs analyses to classify putative sib pairs, putative half-sib pairs, putative parent-offspring pairs, and putative marital pairs into five different types of pairs: MZ twin pairs, full sib pairs, half sib pairs, parent-offspring pairs, and unrelated pairs. A summary file is produced that contains the identifiers of the putative full-sib pairs to be reclassified and their sibling classification statistics; for each pair, missing data rates over the genome; and histograms of the sibling classification statistic and parent offspring classification statistic. An optional output file contains the same pair-specific statistics, but for all putative pairs other than MZ twins (i.e., including putative half-sib pairs, parent-offspring pairs, parent-offspring pairs).

## 9.1 Limitations

The probability of misclassification depends on the total length of the genotyped genome provided and overall marker informativeness. The misclassification rates are minimal when at least half the genome is genotyped using microsatellite markers at most 20 cM apart. Individual pairs may be misclassified if one or both members have a high proportion of missing genotypes, as the classification cut points are based on the length of the genotyped genome and marker informativeness calculated for the entire sample. It should also be noted that the proportion of missing genotypes is calculated using as the denominator the number of markers listed in the genome file.

## 9.2 Theory

This program is intended primarily for late-onset diseases, for which parents are not typed and the number of typed sibs is often two. In this case, one cannot detect errors in relationship by looking for inconsistencies, and one must use the entire genome (or as much of it as possible) to examine the overall allele-sharing between the sibs. In practice, this program can be used for other types of data sets, and even pairs with late-onset disease will sometimes have typed parents or additional sibs. However, we do not use all the marker information to construct the relationship statistics. For each pair, only the marker information for that pair is used, and none from the other relatives, including other sibs and parents. Pair-wise allele-sharing is computed using multipoint algorithms.

#### 9.2.1 Full Sib Pairs

Let  $\hat{f}_{jis}$  be the estimated probability that sib pair *j* shares *i* marker alleles identical-by-descent (IBD) at location *s* on a chromosome. We assume throughout that these IBD probabilities are obtained using multi-point methods. Feingold et al. (1993) proposed a Gaussian process model to describe the ideal (i.e., infinitely dense, fully informative) process for the estimated mean number of alleles shared IBD by a sample of N sib pairs at location *s*:

$$X_s = \sum_{j=1}^{N} (\hat{f}_{j1s} + 2\hat{f}_{j2s}).$$

If the marker is fully informative,  $X_s$  is the total number of alleles shared IBD in the sample at location s.

For the ideal process and a large sample of randomly sampled sib pairs, the mean-sharing statistic

$$Z_s = (X_s - N) / (N/2)^{1/2}$$

has mean equal to 0, variance equal to 1, and approximate Gaussian process covariance function  $\exp(-\beta |t|)$ , where t is the distance between markers and  $\beta$ =0.04 for sib pairs (Feingold et al., 1993). The parameter  $\beta$  is a function of the recombination process and assumes that crossovers are independent, i.e., that there is no crossover interference.

Here we consider a single random sib pair j, and let  $Z_{js}$  be the mean-sharing statistic for a single pair (N=1). We obtain a measure of the average number of alleles shared by this pair over the entire genome. Let  $k=1,2,\ldots,22$  index the human autosomes and  $L_k$  be the length of chromosome k in cM. The statistic

$$Y_{jk} = \frac{1}{L_k} \int_0^{L_k} Z_{js} ds$$

has expectation

$$E(Y_{jk}) = \frac{1}{L_k} \int_0^{L_k} E(Z_{js}) ds = 0$$

and variance

$$Var(Y_{jk}) = \frac{1}{L_k^2} \int_0^{L_k} \int_0^{L_k} Cov(Z_{js}, Z_{jr}) dr ds = \frac{2}{\beta L_k} - \frac{2}{(\beta L_k)^2} (1 - e^{-\beta L_k})$$
(9.1)

(Parzen, 1962; Olson, 1999). In the ideal case of fully informative, infinitely dense markers, the statistic  $Y_{jk}$  is the difference between the proportions of the chromosomes sharing 2 and 0 alleles IBD. More generally, it is the difference between the absolute areas above and below the null mean (sharing 1 allele IBD), divided by the length of the chromosome.

If putative sib pair j is a true sib pair, then  $Y_{jk}/[Var(Y_{jk})]^{1/2}$  has a standard normal distribution as  $L_k \to \infty$ . In practice, the normal approximation is somewhat inadequate for single chromosomes of modest length. A genome-wide measure, the sibling classification statistic given by

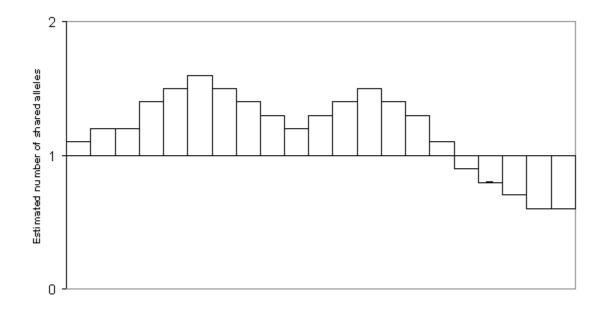


Figure 9.1: Approximate Mean-Corrected Allele Sharing

$$Y_j = \left(\sum_{k=1}^{22} Y_{jk}\right) / \left[\sum_{k=1}^{22} Var(Y_{jk})\right]^{1/2},$$

is well approximated by a standard normal distribution in the fully informative, infinitely dense case. Similarly, for any number of chromosomes K,

$$Y_j = \left(\sum_{k=1}^{K} Y_{jk}\right) / \left[\sum_{k=1}^{K} Var(Y_{jk})\right]^{1/2}.$$

Relationship estimation for each pair j in the sample is based on estimating genome-wide  $Y_j$  for each of the sib pairs. These statistics can be obtained in practice using a standard algorithm to calculate multipoint IBD at equally spaced points throughout the genome. For each chromosome, the absolute areas above and below the estimated mean-corrected allele-sharing curve is approximated using rectangles (see Figure 9.1),

which is equivalent to computing:

$$\hat{Y}_{jk} = [c\sqrt{2}\sum_{s=1}^{P} (X_{sj} - 1)]/P,$$

where P is the number of points at which allele-sharing is computed and c is the distance (cM) between points.

#### 9.2.2 Parent/Offspring Pairs

Parent/offspring pairs are always expected to share exactly one allele IBD, and so  $\hat{Y}_j$  cannot be used to discriminate between sib pairs and parent/offspring pairs. Therefore, a second Markov process statistic is used to classify sibs vs. parent/offspring pairs. At location *s*, the estimated number of alleles shared IBD by a parent/offspring pair is obtained using

$$X_s^* = (\hat{f}_{j2s} + \hat{f}_{j0s} - \hat{f}_{j1s}).$$

For a fully informative location s, the Gaussian process statistic

$$Z_s^* = \sum_{j=1}^N \frac{X_s^*}{N^{1/2}}$$

has a standard normal distribution in a large sample of sib pairs, with covariance function  $\exp(-\beta |t|)$ , where now  $\beta = 0.08$ . The new statistic  $Y_j^*$ , the parent offspring classification statistic, is calculated in the same manner as before, i.e.,

$$Y_j^* = \left(\sum_{k=1}^K Y_{jk}^*\right) / \left[\sum_{k=1}^K Var(Y_{jk}^*)\right]^{1/2},$$

with

$$\hat{Y}_{jk} = [c \sum_{s=1}^{P} X_{sj}^*] / P,$$

and the variance is calculated using equation 9.1 with  $\beta = 0.08$ .

#### 9.2.3 Incomplete Marker Information

When markers are not infinitely dense and fully informative, the variance of the Sibling and Parent-Offspring Classification Statistics are less than one. Classification criteria (cut points) may be determined using the overall marker informativity and the length of the genotyped genome. The *Average Marker Information Content* (AMIC) (Kruglyak and Lander 1995)

is defined as follows,

$$AMIC = \sum_{p=1}^{M} r^2(s)/M,$$

where M is the total number of points over which the genome IBD probabilities are calculated and

$$r^{2}(s) = 1 - \frac{\sum_{i=1}^{N} \sigma_{i, residual}^{2}(s)}{\sum_{i=1}^{N} \sigma_{i, initial}^{2}} = 1 - \frac{2\sum_{i=1}^{N} \sigma_{i, residual}^{2}(s)}{N},$$

N is total number of sib pairs in the sample and  $\sigma_{i,residual}^2(s)$  is the variance of the IBD distribution at point s for sib pair i.

The best-fit regression equations for obtaining classification values, the cut points, are:

- $\log_{10}(-C_u) = 0.421 + 0.506 \log_{10}(T) + 1.162 \log_{10}(AMIC) + 0.472 (\log_{10}(AMIC))^2$ ,
- $\log_{10}(-C_h) = 0.141 + 0.524 \log_{10}(T) + 0.237 \log_{10}(AMIC) 0.861 (\log_{10}(AMIC))^2$ ,
- $\log_{10}(-C_p) = 0.2 + 0.518 \log_{10}(T) + 2.220 \log_{10}(AMIC),$
- $C_m = 3.27$ ,

where T is the total length of the genotyped genome in cM divided by 150, and  $C_u$ ,  $C_h$ ,  $C_m$ , and  $C_p$  are the classification cut points for unrelated pairs, half sib pairs, MZ twins, and parent offspring pairs respectively.  $C_u$ ,  $C_h$ ,  $C_m$  are used to classify pairs on the basis of the sibling classification statistic into unrelated, half sibs, full sibs, and MZ twins.  $C_p$  is used to classify pairs into full sib and parent-offspring pairs on the basis of the parent-offspring classification statistic.

#### 9.2.4 Strategy for Classifying Putative Full-Sib and Non-Full-Sib Pairs

There are two steps to classify each pair:

- 1. Using  $Y_j$  and the cut points defined above, we classify as follows: Unrelated  $< C_u <$  Half sib  $< C_h <$  Sib  $< C_m <$  MZtwin
- 2. If the pair is classified as a sib pair in step 1, we use  $Y_j^*$  and the parent\_offspring cut point: Parent/offspring  $< C_p < Sib$

#### 9.2.5 Nonparametric Estimation Procedure

After calculating the  $Y_j$  and  $Y_j^*$ , a nonparametric estimation procedure is used to obtain the mean and variance of the sib-pair distributions of these two sets of statistics.

1. Estimating the shift:

We use the  $L_2$ -error procedure (Scott, 2000) to maximize the function

$$\frac{2}{n}\sum_{j=1}^{n}\phi(Y_j|\mu,\sigma^2) - \frac{1}{2\sqrt{\pi\sigma^2}},$$

where  $\mu$  and  $\sigma^2$  are parameters, n is the total number of sib pairs (all putative full sib pairs), and  $\phi$ (.) is the normal density function

$$\phi(Y_j|\mu,\sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(Y_j-\mu)^2}$$

2. We then adjust the cut points:

New Cut point = Old Cut point +  $\mu$  from step 1.

- 3. We repeat the same step 1 and 2 for the  $Y_i^*$  obtained from all putative full sib pairs.
- 4. We perform the classification as described in 9.2.4 using the new cut points.

To test the deviation of the sib pair mean from zero, we use the  $Y_j$  from putative full sib pairs now classified as true sib pairs to compute the mean

$$\bar{Y} = \frac{\sum_{j=1}^{n} Y_j}{n}$$

and the standard error of the mean

$$S.E.(\bar{Y}) = \frac{1}{\sqrt{n}} \sqrt{(\sum_{j=1}^{n} Y_j^2 - \frac{(\sum_{j=1}^{n} Y_j)^2}{n})/n} \quad .$$

Then a confidence interval is constructed as

$$\bar{Y} \pm 2S.E.(\bar{Y}).$$

If zero is not included in this interval, a warning is printed in the output. The user should at this point note that the sib-pair histogram is shifted significantly (in the statistical sense) away from its null hypothesis mean value of zero. If such significant deviation is substantial, there may be large-scale error in the data or specification of parameters. Our previous experience with real data sets has shown that such error may be due to

- 1. Gross misspecification of marker allele frequencies,
- 2. Misalignment of marker description information between the parameter file, the data file and/or the genome file, and
- 3. Large-scale genotype errors.

Examples of large-scale genotype errors that have caused large "shifts" in the sib-pair histogram have included:

- 1. Errors in programs translating data from the genotyping lab to the pedigree data file and
- 2. Extensive binning errors in the assignment of genotypes.

The above list includes only errors we have been alerted to by RELTEST; other sources of error detectable by RELTEST are clearly possible. We suggest using RELTEST not only to classify pairs according to relationships, but also as a general test of the overall accuracy of the data and parameter specifications (Olson et al., 2004).

# 9.3 Program Input

File Type	Description
Parameter file	Specifies the parameters and options with which
	to perform an analysis.
Pedigree data file	Contains delimited records for each individual in-
	cluding fields for identifiers, sex, parents, and
	marker data.
Marker locus description file	Lists the alleles, allele frequencies and phenotype
	to genotype mapping for each marker locus.
Genome description file	Contains a description of the linked marker re-
	gions, including distances between markers.

### 9.3.1 Parameter File Syntax

RELTEST can read multiple pedigree data files in cases where each pedigree file contains the markers for a single chromosome. For each pedigree file, there has to be a corresponding pedigree block with file name specified. All other fields should be the same, except for the marker fields, for all pedigree files used.

Example:

```
pedigree, file=ped1
{
    .
    .
    marker="chlml"
    .
    .
    pedigree, file=ped2
    {
    .
    marker="ch2ml"
    .
    .
    pedigree, file=ped3
    {
    .
    .
    .
}
pedigree, file=ped3
{
    .
.
```

marker="ch3ml"
.
.
.
}

The specific syntax for RELTEST parameters, attributes and values is described in the following sections.

### 9.3.1.1 The reltest Parameter

The following syntax table specifies the permissible parameter and attribute settings for the main RELTEST parameter.

<pre>parameter [, attribute]</pre>	Explanation				
	Starts a RELTEST	parameter block.			
	Value Range	N/A			
reltest	Default Value	N/A			
	Required	Yes			
	Applicable Notes	None			
	Specifies the root	name to be used for output files.			
	Output file names will be formed by concatenating the				
	root name and an appropriate extension.				
		Character string representing a			
, out	Value Range	valid file name.			
	Default Value	reltest			
	Required	No			
	Applicable Notes	None			

### 9.3.1.2 The reltest Block

The following syntax table specifies the permissible parameter and attribute settings for the reltest block.

<pre>parameter [, attribute]</pre>	Explanation				
[,]	Specifies the putative pair to be analyzed.				
	sib				
		hsib			
	Value Range	parent_offspring			
pair_type		marital			
	Default Value	None			
	Required	No			
	Applicable Notes	1			
		mic regions that will be used in the			
	analysis.	the regions that will be used in the			
		Character string representing the			
		name of a region in the genome			
		description file. If no region is			
region	Value Range	specified then analysis will take			
1091011		place with respect to all			
		available marker data.			
	Default Value	None			
	Required	No			
	Applicable Notes	None			
	Specifies pre-calculated cut points to be used to clas				
	sify pairs.				
	Value Range	N/A			
cut_points	Default Value	None			
	Required	No			
	Applicable Notes	2			
	Cut point for siblin	ng pairs.			
	Value Range	(-∞, ∞)			
, sib	Default Value	None			
	Required	No			
	Applicable Notes	2			
	Cut point for half-s	sib pairs.			
	Value Range	$(-\infty,\infty)$			
, hsib	Default Value	None			
	Required	No			
	Applicable Notes	2			
	Cut point for MZtv	vins			
	Value Range	$(-\infty,\infty)$			
, MZtwin	Default Value	None			
	Required	No			
	Applicable Notes	2			

	Cut point for paren	t-offspring pairs.	
	Value Range	$(-\infty,\infty)$	
, parent_offspring	Default Value	None	
	Required	No	
	Applicable Notes	2	
	Specifies option to	print out the sibling in the nuclear	
	family file.		
nuctor filo	Value Range	{true, false}	
nucfam_file	Default Value	false	
	Required	No	
	Applicable Notes	3	
	Specifies option to print detailed output file.		
	Value Range	{true, true}	
detailed_file	Default Value	false	
	Required	No	
	Applicable Notes	None	

Notes

- 1. By default, all four pair types will be analyzed.
- 2. Normally, cut points are automatically generated based on the pedigree data, as given in the theory section. The program will use the generated cut points if the cut\_points parameter is not specified here.

Example: All of the following are valid RELTEST analysis statements:

3. See 9.5.3

## 9.4 Program Execution

RELTEST is run via a command line interface on the supported UNIX and Windows platforms. This requires the S.A.G.E. programs to be properly installed and in the current execution path. Input files are specified on the command line and all output files are created in the current working directory.

Running RELTEST from the command prompt with no arguments, or the wrong number of arguments, will result in the program printing its usage statement. This lists the input files the program requires on the command line.

```
>reltest
```

```
S.A.G.E. v5.x -- RELTEST
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
usage: ./reltest <parameters> <pedigree> <locus>
<map>
Command line parameters:
parameters - parameter file
pedigree - pedigree data file
locus - locus description file
map - Genome Map File
```

As indicated in the program usage statement, input files are listed on the command line. A typical run of RELTEST may look like the following:

```
>reltest reltest.par example.ped example.loc example.map
S.A.G.E. v5.x -- RELTEST
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
Reading parameter file.....done.
Reading locus description file.....done.
Reading pedigree file.....done.
from example.ped.....done.
Reading genome file.....done.
Reading genome file.....done
RELTEST analysis.....1
Testing sib pairs......done
Testing parent_offspring pairs.....done
Testing mother_father pairs.....done
Analysis complete!
```

## 9.5 Program Output

Filename	File Type	Description
reltest.inf	Information output file	Contains informational diagnostic
		messages, warnings and program
		errors.
reltest.sum	Reclassification summary	Contains the values of all cut points
	file	and pairs to be reclassified, together
		with related statistics. Contains
		histograms and classification
		statistics for all putative pairs.
reltest.fam	Sibling in nuclear family in-	Contains information about all sib
	formation file	pairs of the nuclear families in which
		at least one sib pair should be
		reclassified.
reltest.det	Detailed pair information file	Contains statistics for all pairs used
		in the analysis.

RELTEST produces several output files that contain results and diagnostic information:

### 9.5.1 Information Output File

The RELTEST information file contains a variety of useful information, including:

- Information on fields read from the pedigree data file. These tables provide information about what the program has read in, and are included with all programs in S.A.G.E. They are very useful for debugging many common errors caused when reading the pedigree data. When first analyzing new data, it is recommended that these tables be checked carefully to make sure pedigree data are being correctly read.
- Information, warning and error messages generated throughout the program. The program attempts to correct many common errors by making reasonable assumptions that are usually sufficient, but may not be for your data set. It is strongly recommended that you check this file for warning and error messages before examining the results of any run of the program. The file "reltest.inf" should be checked for errors and diagnostic information after each run of the program.

### 9.5.2 Reclassification Summary File

The reclassification summary file contains the cut point values to classify pairs and the total length of genome used in the analysis. It also provides a separate table for each putative pair type, listing pairs to be reclassified with their individual IDs and pedigree IDs from the original pedigree data file, new class, sibling classification statistic, parent offspring classification statistic and missing genotype rate. Note: misclassification may occur if one or both members of the pair have a high rate of missing genotypes. For each putative pair type, the total number of original pairs and the total number of pairs to be reclassified are also included.

This file also provides text-based histograms of the sibling classification statistic and the parent offspring classification statistic for each putative pair type included in the analyses. The minimum and maximum values of these statistics are also included.

### 9.5.3 Sibling in Nuclear Family Information File

The sibling-in-nuclear-family information file contains information about all sib pairs in nuclear families in which at least one sib pair should be reclassified. This file provides heuristic information intended to aid understanding the statistical distribution related to pairs that should be reclassified.

### 9.5.4 Detailed Pair Information File

This file provides a table of the  $Y_j$  and  $Y_j^*$  values for all pairs used in the analysis for each putative pair type.

## 9.6 Example Output Files

### 9.6.1 Reclassification Summary File

Here is a typical example of a RELTEST summary output file:

```
S.A.G.E. v5.x -- RELTEST
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
RELATIONSHIP TEST PROGRAM SUMMARY OUTPUT
Analysis Name
                                         : default_analysis
    Average Marker Information Content : 0.61057
                               : 3539 (cM)
    Total Length of Genome
    Cut-points
                                                   original adjusted
                                        ------

        Sibling Classification
        unrelated
        -7.73259
        -6.56278

        Statistics(Yj)
        half sib
        -3.07633
        -1.90652

        MZtwins
        3.27000
        4.43981

                                                     3.27000
      Parent/Offspring Classification |parent/
        Statistics(Yj*) | offspring | -2.72544 -3.22814
      _____
    Sibling Classification Statistics(Yj)
       robust (L2) mean : 1.16981
       robust (L2) variance : 0.5
    Parent/Offspring Classification Statistics(Yj*)
       robust (L2) mean : -0.502698
       robust (L2) variance : 0.5
    Average Yj of Pairs
      Reclassified as Full Sibs : 1.18377
    Standard Error : 0.0477349
    95% Confidence Interval : 1.0883 to 1.27924
  ! WARNING : THE MEAN OF THE SIB-PAIR DISTRIBUTION DIFFERS SIGNIFICANTLY FROM
               ZERO. YOU MAY HAVE SUBSTANTIAL DATA ERROR OR MISSPECIFICATION OF
              PARAMETERS SUCH AS ALLELE FREQUENCIES.
_____
PUTATIVE FULL SIB PAIRS TO BE RECLASSIFIED :
                    reclassified
          pair pair type
                                       Υj
                                                    Yj*
  pid
                                                                 missing data
 _____

      118
      3/4
      HSIB
      -2.2236
      -1.6296
      4% / 5%

      159
      3/5
      HSIB
      -4.2079
      0.0942
      8% / 8%

      4
      3/4
      HSIB
      -2.8302
      -1.6862
      2% / 1%

      45
      3/4
      HSIB
      -3.0415
      -0.6532
      3% / 5%

      58
      5/6
      HSIB
      -3.1239
      -1.6834
      3% / 4%

      60
      3/5
      HSIB
      -2.7622
      -1.0231
      3% / 4%

      66
      30/31
      HSIB
      -2.7980
      -1.1299
      5% / 5%

      66
      12/16
      MZTWINS
      8.3388
      7.1082
      30% / 4%

  118
  159
  _____
 total putative pairs : 342
 total pairs to be reclassified : 8
```

====		====
==		==
==	HISTOGRAM OF SIBLING CLASSIFICATION STATISTIC (Yj)	==
==	FOR PUTATIVE PAIRS	==
==		==
==	putative pair type : FULL SIB	==
==	maximum Yj : 8.33881	==
==	minimum Yj : -4.20787	==
==	bin size : 0.25	==
==		==

-----

Inter			nt (one * is equal to 1 or 2 pairs.)
-4.22 to		1	
-3.97 to	-3.72	0	
-3.72 to	-3.47	0	
-3.47 to	-3.22	0	
-3.22 to	-2.97	2	*
-2.97 to	-2.72	3	**
-2.72 to	-2.47	0	
-2.47 to	-2.22	1	*
-2.22 to	-1.97	0	
-1.97 to	-1.72	0	
-1.72 to	-1.47	1	*
-1.47 to	-1.22	1	*
-1.22 to		1	
-0.97 to	-0.72	2	
-0.72 to		2	
-0.47 to			* * * *
-0.22 to			* * * * * *
0.03 to			* * * * * * * * *
0.28 to			* * * * * * * * * * * * *
0.53 to		= -	* * * * * * * * * * * *
0.78 to			* * * * * * * * * * * * * * * * * * * *
1.03 to			* * * * * * * * * * * * * * * * * * * *
1.28 to			* * * * * * * * * * * * * * *
1.53 to			* * * * * * * * * * * * * * * * *
1.78 to		==	* * * * * * * * *
2.03 to			* * * * * * *
2.28 to			* * * * * * * * *
2.53 to			* * *
2.78 to			**
3.03 to 3.28 to		3	^ ^
3.28 to 3.53 to		0	
		1	*
3.78 to 4.03 to		1	
4.03 to 4.28 to		1 0	
4.28 to 4.53 to		0	
4.55 LO	7./0	0	

. .

•

### 9.6.2 Sibling in Nuclear Family Information File

Here is a typical example of a sibling in nuclear family information file:

```
S.A.G.E. v5.x -- RELTEST
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
RELATIONSHIP TEST PROGRAM NUCLEAR FAMILY INFORMATION
Note : This file contains information about all sib pairs of the
nuclear families in which at least one sib pair has been
reclassified.
Analysis Name : analysis_1
```

\_\_\_\_\_

pid	pair	reclassified pair type	Yj	Yj*	missing data
118	3/4	HSIB	-2.2236	-1.6296	4% / 5%
118	7/8	SIB	0.4039	-1.0336	4% / 3%
159	3/5	HSIB	-4.2079	0.0942	8% / 8%
159	7/8	SIB	1.5395	-0.2638	7% / 7%
4	3/4	HSIB	-2.8302	-1.6862	2% / 1%
45	3/4	HSIB	-3.0415	-0.6532	3% / 5%
58	5/6	HSIB	-3.1239	-1.6834	3% / 4%
58	14/15	SIB	2.0909	-0.6428	2% / 2%
60	3/5	HSIB	-2.7622	-1.0231	3% / 4%
=======					

## Chapter 10

# SIBPAL

This is a model-free linkage program that models trait data from full-sib pairs as a function of marker allele sharing identity-by-descent (IBD). Available analyses can use both single- and multi-point IBD information, and models allow for both binary and continuous traits due to multiple genetic loci, including epistatic interactions, and covariate effects. Like the original SIBPAL, it uses linear regression and hence is extremely fast.

## **10.1** Limitations

The Haseman-Elston linkage test in this release only includes support for univariate analysis of sibling pairs for autosomal regions. Full support for multivariate analysis and using all relative pairs will be available in future releases.

Unlike earlier versions of SIBPAL, this program does not generate IBD sharing estimates itself. That must be done using GENIBD, which outputs an IBD sharing file as input for SIBPAL.

### 10.2 Theory

### **10.2.1** Basic notation

Let the number of individuals be N with trait values:  $x_1, x_2, ..., x_N$ .

Let the number of covariates be *c* with values for sib i:  $z_{1i}, z_{2i}, \ldots, z_{ci}$ .

Let  $\hat{f}_{11}, \hat{f}_{12}, ..., \hat{f}_{1j}, ...$  be the probability of sharing 1 allele IBD for the jth sib pair.

Let  $\hat{f}_{21}, \hat{f}_{22}, ..., \hat{f}_{2j}, ...$  be the probability of sharing 2 alleles IBD for the jth sib pair.

Let  $\hat{\pi}_1$ ,  $\hat{\pi}_2$ , ...,  $\hat{\pi}_j$ , ... be the average allele sharing IBD (proportion of alleles shared) for the jth sib pair, where  $\hat{\pi}_j = \hat{f}_{2j} + w_1 \hat{f}_{1j}$  and  $w_1 \in [0, 0.5]$ .<sup>1</sup>

These probabilities are conditional on the marker<sup>2</sup> information available.

<sup>&</sup>lt;sup>1</sup>The default value of  $w_1$  is 0.5.

<sup>&</sup>lt;sup>2</sup>In this context, marker  $\equiv$  marker location, and need not be a measured marker. This is mainly an issue dealt with in the IBD generation phase.

Let the number of sibships be P.

Let the number of full sibling pairs in the i'th sibship be  $n_i: n_1, n_2, \ldots, n_P$ .

Let j index the sib pair: j=1, 2, ...,  $\sum_i n_i = n$ .

### 10.2.2 Test of Mean Allele Sharing

We want estimates of the means of the  $\hat{\pi}_j$  and  $\hat{f}_{ij}$ , which we will denote  $\bar{\pi}$  and  $\hat{f}_i$ , and test the hypothesis that their values agree with expectation under random sampling. These tests are that  $E(\bar{\pi}) = \pi$  and  $E(\bar{f}_i) = f_i$ , where  $\pi = f_2 + w_1 f_1$  and  $(f_0, f_1, f_2) = (\frac{1}{4}, \frac{1}{2}, \frac{1}{4}, )$  for a random sample of full sib pairs if there is no meiotic drive or selection. The means and their variances are estimated by calculating:

$$\bar{\hat{\pi}} = \frac{1}{n} \sum_{j} \hat{\pi}_{j} \qquad s_{\hat{\pi}}^{2} = \frac{1}{n(n-1)} \sum_{j} (\hat{\pi}_{j} - \bar{\hat{\pi}})^{2} 
\bar{\hat{f}}_{i} = \frac{1}{n} \sum_{j} \hat{f}_{ij} \qquad s_{\hat{f}_{i}}^{2} = \frac{1}{n(n-1)} \sum_{j} (\hat{f}_{ij} - \bar{\hat{f}}_{i})^{2}.$$
(10.1)

From each mean, a t-statistic is computed and referred to the t-distribution with n-1 d.f. for a twosided test, i.e., the p-values are

$$P\left(t_{n-1} \ge \frac{\left|\bar{\hat{\pi}} - \pi\right|}{s_{\hat{\pi}}}\right) \text{ and } P\left(t_{n-1} \ge \frac{\left|\bar{\hat{f}}_i - f_i\right|}{s_{\hat{f}_i}}\right),$$

where  $t_{n-1}$  is a random variable that is distributed as t with n-1 d.f.

### 10.2.3 Test of Mean Allele Sharing for Binary Traits in Selected Pairs

The above tests can be performed separately for pairs with 0, 1, and 2 affected members as tests for linkage. However, all tests are then one-sided and the p-values are

$$P\left(t_{n-1} \ge \frac{\delta\left[\bar{\hat{\pi}} - \pi\right]}{s_{\hat{\pi}}}\right) \text{ and } P\left(t_{n-1} \ge \frac{\delta\left[\bar{\hat{f}}_i - f_i\right]}{s_{\hat{f}_i}}\right)$$

where  $\delta=1$  for affected pairs (2 affected members) if i = 2 and for unaffected pairs (0 affected members) if i = 0, and  $\delta = -1$  for discordant pairs (1 affected member). No tests are performed if i = 1.

#### 10.2.4 Generalized Haseman and Elston Linkage Test

### **10.2.4.1** Dependent variables

Denote the j-th sib-pair with the subscript ii', and let

$$\bar{x} = \frac{1}{2N} \sum_{j=ii'=1}^{N} (x_i + x_{i'}).$$

Then the dependent variable for the ii'-th pair can be

$$y_{j} = \begin{cases} (x_{i} - \overline{x})(x_{i'} - \overline{x}) & mean-corrected cross - product (default) \\ -\frac{1}{2}[(x_{i} - \overline{x}) - (x_{i'} - \overline{x})]^{2} = -\frac{1}{2}(x_{i} - x_{i'})^{2} & -\frac{1}{2}(squared pair trait difference) \\ \frac{1}{2}[(x_{i} - \overline{x}) + (x_{i'} - \overline{x})]^{2} & \frac{1}{2}(squared mean - corrected trait sum) \\ w_{i}[(x_{i} - \overline{x}) + (x_{i'} - \overline{x})]^{2} - (1 - w_{i})(x_{i} - x_{i'})^{2} & weighted combination of the squared trait difference and squared mean-corrected sum \end{cases}$$

Any individual who has a missing trait, marker or covariate value is not used, and not included in any pair that requires it - i.e., those pairs are not included in n.

Note that in the case of a binary trait,  $x_i = 1$  for an affected individual and 0 for an unaffected individual, and  $x_i$  is then treated the same as for any other quantitative trait to obtain  $y_i$ .

#### 10.2.4.2 Regression Model

The basic model we fit is of the form

$$y = \alpha + \sum_{h} a_h \hat{\pi}_h + \sum_{h} d_h \hat{f}_{2h} + \sum_{k} c_k f(z_k) + \varepsilon$$
(10.2)

where  $\alpha$  is the intercept, and in a random sample  $a_i$  is the additive genetic variance due to the h-th marker when  $w_1 = 0.5$ ,  $\hat{\pi} = \hat{f}_2 + w_1 \hat{f}_1$ ,  $d_h$  is the dominant genetic variance due to the h-th marker,  $c_k$  is a nuisance parameter accounting for the effect of some function f of the k'th covariate, and  $\varepsilon$  is the residual error. The variances  $a_h$  and  $d_h$  are the trait locus-specific variances attenuated by the recombination fraction between the trait and marker loci, when  $w_1$  is 0.5. An iterative method using generalized estimating equations (GEE) is used to fit this model to allow for the non-independence of sibling pairs.

A problem may occur when performing multiple regression using multi-point IBD estimates. The IBD sharing between closely linked markers can be almost totally linearly dependent, resulting in a singular design matrix during Trait Regression. If multiple regression is to be performed using multiple markers with multi-point IBD sharing information, it is recommended that the loci used should have significantly different information content (i.e. be on different chromosomes or have at least one informative marker between them).

#### **10.2.4.3** Correlation Matrices

In a nuclear family, if there are s sibs in the sibship, there are s(s - 1)/2 sib pairs and  $s^2(s - 1)^2/4$  entries in the correlation matrix for that sibship, of which

s(s - 1)/2 are 1 (down the main diagonal)

s(s - 1)(s - 2) are  $r_1$  (when the pair of pairs has one sib in common)

s(s - 1)(s - 2)(s - 3)/4 are  $r_o$  (when the pair of pairs has no sibs in common)

where  $r_1$  and  $r_o$  are correlations that differ depending on the dependant variable.

Below are the correlation matrices for sib pairs in sibships of size 3 (top left  $3 \times 3$  matrix), 4 (top left  $6 \times 6$  matrix) and 5 (whole matrix).

Sib1,Sib2	1,2	1,3	2,3	1,4	2,4	3,4	1,5	2,5	3,5	4,5
1,2	1	$r_1$	$r_1$	$r_1$	$r_1$	$r_0$	$r_1$	$r_1$	$r_0$	$r_0$
1,3	$r_1$	1	$r_1$	$r_1$	$r_0$	$r_1$	$r_1$	$r_0$	$r_1$	$r_0$
2,3	$r_1$	$r_1$	1	$r_1$	$r_0$	$r_0$	$r_0$	$r_0$	$r_0$	$r_0$
1,4	$r_1$	$r_1$	$r_1$	1	$r_1$	$r_0$	$r_0$	$r_1$	$r_0$	$r_1$
2,4	$r_1$	$r_0$	$r_1$	$r_1$	1	$r_0$	$r_1$	$r_1$	$r_0$	$r_0$
3,4	$r_0$	$r_1$	$r_1$	$r_1$	$r_1$	1	$r_1$	$r_0$	$r_1$	$r_1$
1,5	$r_1$	$r_1$	$r_0$	$r_1$	$r_0$	$r_0$	1	$r_1$	$r_1$	$r_1$
2,5	$r_1$	$r_0$	$r_1$	$r_0$	$r_1$	$r_0$	$r_1$	1	$r_1$	$r_1$
3,5	$r_0$	$r_1$	$r_1$	$r_0$	$r_0$	$r_1$	$r_1$	$r_1$	1	$r_1$
4,5	$r_0$	$r_0$	$r_0$	$r_1$	$r_1$	$r_1$	$r_1$	$r_1$	$r_1$	1

Let  $\mathbf{r} = (r_1, r_0)$  be a vector of correlations, where  $r_i$  is the correlation between pairs sharing *i* individuals in common.  $\mathbf{r}$  is either estimated from the data for the chosen dependent variable, with the restriction that the correlations are constrained to be greater than 0 to avoid numerical instability, or set to 0 by the user.

Let  $R_s$  be the correlation matrix for a sibship of size s. There are  $n_i$  (i = 1, 2, ..., P) pairs in the i-th family,  $n = \sum_{i=1}^{P} n_i$  pairs all told,  $n_i = s_i(s_i - 1)/2$ . Let W be a block diagonal matrix of the  $R_i$ :

	$R_{n1}$	0	•••	0
117	0	$R_{n2}$		0
W =			·	0
	0		0	$R_{nP}$

### 10.2.4.4 Univariate Test of Linkage Using Full Sib Pairs

We want to calculate the vector of m estimates<sup>3</sup>:

$$b = (A^T W^{-1} A)^{-1} A^T W^{-1} y, (10.3)$$

where y is an n × 1 vector of dependent variates with transpose  $y^T = (y_1, y_2...y_n)$ . A is an n × m design matrix, where m is the number of parameters estimated - each parameter corresponds to a particular column of A.

Columns of A:

- 1. The first column is a column of 1's
- 2. Following this come one or two columns for each marker locus entered in the model. The first of each of these is a column whose elements are  $\hat{\pi}_j$  and the second of each (if present) is a column whose elements are  $\hat{f}_{2j}$ . For each marker, the user can choose whether or not to include dominance  $(f_2)$  in the model. If it is not included,  $a_i$  is an attenuated measure of the total (additive and dominant) genetic variance when  $w_1$  is 0.5.

<sup>&</sup>lt;sup>3</sup>The equations are not actually computed by SIBPAL as listed. A significantly more complex method is implemented that is efficient and numerically stable.

- 3. Following this may come one or more columns each element of which is the product of elements of two (or more) of the previous columns.
- 4. Following this come one or more columns for each covariate entered in the model. Each of these is a column whose elements are the mean-corrected sum, absolute difference, or meancorrected cross product between the sib pair covariate values. Additional columns may be entered that are powers of these covariate sums, differences or products.

Let the k'th covariate term included in the model for the ii'-th pair be

$$z_{kii'} = \begin{vmatrix} (z_{ki} - \bar{z}_k) + (z_{ki'} - \bar{z}_k) & \text{mean-corrected covariate sum} \\ |(z_{ki} - \bar{z}_k) - (z_{ki'} - \bar{z}_k)| & \text{mean-corrected covariate absolute difference} \\ (z_{ki} - \bar{z}_k)(z_{ki'} - \bar{z}_k) & \text{mean-corrected covariate product (default)} \end{vmatrix}$$

for covariate k and individuals i and i'.

For each covariate, the user can choose whether or not to include any combination of the sum, difference and product terms in the model, as well as powers of them. Including too many covariate terms may cause A to be singular due to linear dependencies in the data. Covariate means may be specified in the parameter file or estimated from the set of relative pairs used in the analysis.

1. Following this may come one or more columns, each element of which is the product of elements of two (or more) of the covariate terms in previous columns.

Thus, A will be of the form:

1	$\hat{\pi}_{11}$ $f_{211}$	$\hat{\pi}_{21}$ $f_{221}$	$\cdots \hat{\pi}_{11}\hat{\pi}_{21}$	• • •	$[z_{11}]^p$ $[z_{21}]^p$	$\cdots \qquad [z_{11}z_{21}]^p$	• • •
1	$\hat{\pi}_{12}$ $f_{212}$	$\hat{\pi}_{22}$ $f_{222}$	$\cdots \qquad \hat{\pi}_{12}\hat{\pi}_{22}$		$[z_{12}]^p$ $[z_{22}]^p$	$\cdots \qquad [z_{12}z_{22}]^p$	
1	$\hat{\pi}_{13}$ $f_{223}$	$\hat{\pi}_{23}$ $f_{223}$	$\cdots \qquad \hat{\pi}_{13}\hat{\pi}_{23}$		$[z_{13}]^p$ $[z_{23}]^p$	$\cdots \qquad [z_{13}z_{23}]^p$	
÷	:	÷	÷		÷	÷	
(a)	(b)	(b)	(c)		(d)	(e)	
	first	second	marker		covariate sum,	covariate	
	marker	marker	interactions		difference and	interactions	
					product terms		

The first  $n_1$  rows of A are for family 1,

the next  $n_2$  are for family 2,

the last  $n_P$  are for family *P*.

### 10.2.4.5 Output of estimates and t-statistics

We want the estimates  $b_i$ , the three elements of b indicated by (10.3), and the test statistics  $\frac{b_i}{s_i}$ , for i = 2, 3, ..., m, where  $s_i^2$  is

$$\frac{y^T W^{-1} (y - Ab) d_{ii}}{n - m}$$
(10.4)

in which  $d_{ii}$  is the i-th diagonal element of  $(A^T W^{-1} A)^{-1}$ . For each test statistic we calculate a p-value which is either

$$\mathbf{P}_t = P\left(t_{n-m} \ge \frac{b_i}{s_i}\right) \tag{10.5}$$

or

$$\mathbf{P}_t = P\left(t_{n-m} \ge \frac{|b_i|}{s_i}\right). \tag{10.6}$$

Estimates  $b_i$  corresponding to a column of  $\hat{\pi}s$  or  $\hat{f}_2s$  and other columns of marker terms (i.e., products of  $\hat{\pi}s$  and  $\hat{f}_2s$ ) use 10.5. A two-sided test (10.6) is used for all remaining columns that contain any covariate terms.

Alternatively, the above tests can be performed using variances estimated using an estimator that is robust to misspecification of the model and the correlation matrices. When this option is specified, the covariance matrix of the parameter estimates is computed using the *sandwich* variance estimator

$$(A^{T}W^{-1}A)^{-1} \left[ y^{T}W^{-1}(y-Ab_{j}) \right] \left[ y^{T}W^{-1}(y-Ab_{j}) \right]^{T} (A^{T}W^{-1}A)^{-1}.$$
 (10.7)

These variance estimates can be extremely conservative and caution should be exercised when using this option.

#### **10.2.4.6** Empirical estimates of significance

We can also estimate an empirical p-value of the test statistic using a Monte Carlo permutation procedure with N replicate permutations. For each replicate, we permute the allele sharing among the pairs (both within sibships and across sibships of the same size), recalculate the test statistic, and determine the proportion of the replicates that are equal to or greater than the statistic calculated from the original observations. We choose N, the number of replicates, such that the estimated empirical p-value,  $\hat{p}$ , is within a proportion w (the width parameter) of its true p-value, p, with predetermined confidence probability  $\gamma$  (the confidence parameter). That is, we want the standard deviation  $s_{\hat{p}}$  of  $\hat{p}$  to be proportional to  $\hat{p}$ . This permutation process can be viewed as a set of N independent Bernoulli trials each with success probability p. The sample variance,  $s_{\hat{p}}^2$ , of  $\hat{p}$  is  $s_{\hat{p}}^2 = \frac{\hat{p}(1-\hat{p})}{N}$ . So we choose N such that  $Pr(|\hat{p} - p| \le w\hat{p}) = \gamma$ . Using a normal approximation for the distribution of  $\hat{p}$ , we obtain

$$N = \left(\frac{1-\hat{p}}{w^2\hat{p}} \left[\Phi^{-1}\left(\frac{\gamma+1}{2}\right)\right]^2\right),\,$$

where  $\Phi$  is the standard normal cumulative distribution function. We estimate N by substituting for  $\hat{p}$  the p-value obtained on assuming the test statistic follows a t-distribution, and use this number of replicates to obtain an empirical p-value within any prespecified proportion of its true value with a known confidence coefficient. For example, if we wish to estimate an empirical p-value within 20% of its true value with 95% confidence, then N should be approximately  $\frac{100(1-\hat{p})}{\hat{p}}$ . The number of replicates, *N*, can be limited to avoid excessive computing time.

## **10.3 Program Input**

File Type	Description
Parameter file	Specifies the parameters and options with which
	to perform a particular analysis.
Pedigree data file	Contains delimited records for each individual in-
	cluding fields for identifiers, sex, parents, trait and
	marker data.
IBD sharing file	Stores identity-by-descent (IBD) distributions be-
	tween pairs of related individuals at one or more
	marker loci.

## 10.3.1 The sibpal Parameter

The following syntax table specifies the permissible parameter and attribute settings for the main
SIBPAL parameter.

parameter [, attribute]		Explanation
	Starts a SIBPAL pa	arameter block.
	Value Range	N/A
sibpal	Default Value	None
	Required	Yes
	Applicable Notes	None
	Specifies the root name to be used for output files.	
	Output file names v	will be formed by concatenating the
	root name and an appropriate extension.	
out	Value Range	Character string representing a
, out	value Ralige	valid file name.
	Default Value	traits
	Required	No
	Applicable Notes	None

### 10.3.2 The sibpal Block

The following syntax table specifies the permissible parameter and attribute settings for the sibpal block.

parameter [, attribute]	Explanation	
[, attribute]	Starts a sub-block for specifying a test of mean IBD sharing.	
mean_test	Value Range     N/A       Default Value     None       Required     Yes	
	Applicable Notes None	
	Starts a sub-block for specifying a regression of traits on one or more markers, covariates, and interactions	
trait_regression	Value Range N/A Default Value None	
	Specifies that regression will be performed on covari- ate(s) only, and any listed markers will be disregarded.	
	Value Range N/A	
, zero_marker	Default Value None	
, zero	Required No	
	Applicable Notes None	
	Selects single regression on one marker at a time.	
	Value Range N/A	
, single	Default Value None	
, single_marker	Required No	
	Applicable Notes None	
	Selects multiple regression on all chosen markers at	
	once.	
	Value Range N/A	
, multiple	Default Value None	
, multiple_marker	Required No	
	Applicable Notes None	
	Specifies default type of regression for listed trait re-	
	gression blocks.	
trait regregation default	Value Range {zero, single, multiple}	
trait_regression_default	Default Value single	
	Required Yes	
	Applicable Notes 1	

Notes

1. If a trait\_regression statement does not have either the single or multiple attributes, then the trait\_regression\_default statement will determine whether the given marker or interval estimates will be regressed one at a time (single) or all at once (multiple).

Each trait\_regression statement performs a test of linkage of a trait to one or more markers. The analysis may consist of several regression tests each using a single marker, if either the single attribute is included or the value of the trait\_regression\_default parameter is set to single. Similarly, a single multiple-regression test is performed if either the multiple attribute is included or the value of the trait\_regression\_default parameter is set to multiple. The traits, covariates, markers and other options to be used may be listed in a sub-block of the trait\_regression statement. All options changed in a sub-block are local to the analysis being performed, and do not affect further analyses. If no sub-blocks are listed, then analysis will be performed using all traits and all markers. All parameters that may be included in the sub-block are optional and all values are case-insensitive.

2. A single regression is performed by default.

## 10.3.3 The mean\_test Sub-Block

The following syntax table specifies the permissible parameter and attribute settings for the mean\_test sub-block.

parameter	Explanation	
[, attribute]		
	Specifies the name of a marker for which to test mean	
	IBD sharing	
	Character string representing the	
marker	Value Range name of a marker listed in the	
	pedigree data file.	
	Default Value None	
	Required No	
	Applicable Notes 1	
	Names a trait denoting affection status. Analysis is	
	performed separately on concordantly affected, unaf-	
	fected and discordant pairs.	
	Character string representing the	
trait	Value Range name of a trait listed in the	
	pedigree data file.	
	Default Value None	
	Required Yes	
	Applicable Notes None	
	Specifies a trait used as an indicator variable to select	
	subsets of pairs to analyze.	
	Character string representing the	
_	Value Range name of a trait listed in the	
subset	pedigree data file.	
	Default Value None	
	Required No	
	Applicable Notes 2	
	Prints more verbose output information. This causes	
	some output tables to be more than 80 columns wide.	
	Value Range {true, false}	
wide_out	Default Value false	
	Required No	
	Applicable Notes 3	
	Specifies option to produce tab-delimited output that	
	can easily be imported to other programs such as Ex-	
	cel, SAS and SPlus.	
export_output	Value Range {true, false}	
ever c_ouchac	Default Value false	
	Required No	
	Applicable Notes None	

	Specifies option to print p-values using scientific no- tation as opposed to the default of fixed decimal nota- tion.	
pval_scientific_notation	Value Range	{true, false}
	Default Value	false
	Required	No
	Applicable Notes	None

Notes

1. The value of a marker parameter should be set to the name of a marker for which IBD sharing information was generated and stored in the IBD sharing file. If no valid marker parameters are listed, then all markers are used. The following are all valid mean\_test statements:

```
mean_test # Test each marker
mean_test { # Equivalent to the previous statement.
}
mean_test {
    marker=M1
    marker="region 1 MRK"
    marker=M3
}
```

- 2. The subset parameter specifies a trait to be used as an indicator variable to limit the individuals that may be used in an analysis; individuals for whom this indicator is zero are assumed to have missing trait values. It may be included more than once, in which case the only individuals included in the analysis are those for which all the indicated binary traits are coded 1. The trait being analyzed for linkage should not be used as a subset variable. If the trait specified is a binary trait, it should be coded as 0 for individuals to be excluded from analysis and 1 for individuals to be included. Only those individuals that are affected will be considered. If the trait is continuous, only individuals with trait values greater than 0 will be included. This option does not alter the direction of any of the test statistics as would the trait parameter, so it is usually not appropriate to specify subsets based on phenotypes that are useful for testing linkage.
- 3. If the wide\_out parameter is set to **true**, then additional columns are added to the output from Trait Regression analyses, including a column of t-values corresponding to each parameter estimate.

## 10.3.4 The trait\_regression Sub-Block

The following syntax table specifies the permissible parameter and attribute settings for the trait\_regression sub-block.

parameter	Explanation	
[, attribute]	-	
	Specifies a trait to be used as the dependant variable	
	in the current test.	
	Character string representing the	
trait	Value Range name of a trait listed in the	
	pedigree data file.	
	Default Value None	
	Required No	
	Applicable Notes 1	
	Fixes the trait mean to a value other than the sample	
	mean.	
	Value Range N/A	
, mean	Default Value None	
	Required No	
	Applicable Notes 1	
	Specifies a marker to be included in the current test.	
	Character string representing the	
	Value Range name of a marker listed in the	
marker	pedigree data file.	
	Default Value None	
	Required No	
	Applicable Notes 2	
	Specifies option to test the additive and dominance	
	variances linked to the marker separately instead of	
	the total variance.	
, dominance	Value Range N/A	
, dom	Default Value None	
	Required No	
	Applicable Notes 3	
	Names a covariate.	
	Character string representing the	
	Value Range name of a covariate listed in the	
covariate	pedigree data file.	
	Default Value None	
	Required No	
	Applicable Notes 4	
	Include the covariate mean-corrected product.	
	Value Range true	
, prod	Value Range false	
	Default Value true	
	Required No	
	Applicable Notes None	

	Include the covariate mean-corrected sum.		
, sum		true	
	Value Range	false	
	Default Value	false	
	Required	No	
	Applicable Notes	None	
	Include the covaria	te difference.	
	Value Denge	true	
2455	Value Range	false	
, diff	Default Value	false	
	Required	No	
	Applicable Notes	None	
	Include all covariat	te terms (sum, difference and prod-	
	uct).		
		true	
, all	Value Range	false	
	Default Value	false	
	Required	No	
	Applicable Notes	None	
	Raise covariate terr	ms to specified power.	
	Value Range	$(-\infty,\infty)$	
, power	Default Value	1.0	
	Required	No	
	Applicable Notes	None	
	Starts a parameter sub-block that contains marker		
	and covariate parameters that represent a multiplica-		
	tive interaction terr	m to be included in the regression	
interaction	model.		
Interaction	Value Range	N/A	
	Default Value	None	
	Required	No	
	Applicable Notes	5	
	Specifies which of	the following dependent variables	
	to use in the curren		
		diff	
		sum	
regression_method	Volue Dense	prod	
	Value Range	W2	
		W3	
		W4	
	Default Value	prod	
	Required	No	
	Applicable Notes	6	

add_sum_covariate	will automatically	eter is set to <b>true</b> and ethod = <b>sum</b> , then SIBPAL include the trait sum as a regres- ereby increasing power to detect {true, false} false No None
	· ·	use only a subset of the data. The
subset	Value Range	Id be an indicator variable. Character string representing the name of a trait, phenotype or covariate listed in the pedigree file, or the name of a variable specified within a function block.
	Default Value	None
	Required	No
	Applicable Notes	7
	Names a trait to use as a regression weight for pairs. Weights are computed as the product of each individ- uals trait value.	
weight	Value Range	Character string representing the name of a trait, phenotype or covariate listed in the pedigree file, or the name of a variable specified within a function block.
	Default Value	None
	Required	No
	Applicable Notes	None
		assume that all sib pairs are inde-
		he identity working matrix.
	Value Range	{true, false}
identity_weights	Default Value	false
	Required	No
	Applicable Notes	None
robust_variance	the robust, or <i>sand</i> lead to very conser	ance of parameter estimates using wich, variance estimator. This can vative tests for larger samples con- ndent sibling pairs.
- Coupe_var rance	Value Range	{true, false}
	Default Value	false
	Required	No
	Applicable Notes	None

	Use the sibship mean when mean-correcting trait values.
·	Value Range {true, false}
sibship_mean	Default Value false
	Required No
	Applicable Notes None
	Use the grand mean for all sibships with fewer than
	the specified number of siblings.
	Value Range {1, 2, 3,}
, threshold	Default Value 3
	Required No
	Applicable Notes None
	Prints more verbose output information. This causes
	some output tables to be $> 80$ columns wide.
	Value Range {true, false}
wide_out	Default Value false
	Required No
	Applicable Notes 8
	Compute empirical p-values by permutation.
	Value Range {true, false}
compute_empirical_pvalues	Default Value false
compute_empiricar_pvarues	Required No
	Applicable Notes None
	Only compute empirical p-values for asymptotic p-
	values less than this value.
	Value Range [0,1]
, threshold	Default Value $0.05$
	Required No
	Applicable Notes None
	Specifies an exact number of permutations that should
	always be performed if the asymptotic p-value is less
	than threshold. Use of this option effectively overrides
	all of the following attributes.
, permutations	All of the following attributes:Value Range $\{0, 1, 2, 3,\}$
	Default Value None
	Required No
	Applicable Notes None
	Specifies the maximum number of permutations that
	should be performed.
, max_permutations	î
	Value Range $\{0, 1, 2, 3,\}$ Default Value 10000
	Required No
	•
	Applicable Notes None

, width	Specifies the relative precision of the empirical p- value. E.g., if width=0.2, p-values will be estimated to be within 20% of their true value with a given con- fidence level. This value is used to choose the number of replicates necessary. Note that the number of repli- cates required varies quadratically with the inverse of the width.Value Range[0,1]Default Value0.2RequiredNoApplicable NotesNone
, confidence	Applicable Notes       Note         Specifies the confidence with which an empirical p-value is required to be within a relative interval (i.e., the width) of its true value.         Value Range       [0,1]         Default Value       0.95         Required       No         Applicable Notes       None
skip_uninformative_pairs	Option to skip pairs of individuals whose prior and observed IBD sharing probabilities are numerically identical given the machine precision.         Value Range       {true, false}         Default Value       false         Required       No         Applicable Notes       None
export_output	Specifies option to produce tab-delimited output that can easily be imported to other programs such as Excel, SAS and SPlus.         Value Range       {true, false}         Default Value       false         Required       No         Applicable Notes       None
pval_scientific_notation	Specifies option to print p-values using scientific no- tation as opposed to the default of fixed decimal nota- tion.         Value Range       {true, false}         Default Value       false         Required       No         Applicable Notes       None
print_design_matrix	Specifies option to print the design matrix A.         Character string representing the name of a marker or location listed in the IBD sharing file.         Default Value       None         Required       No         Applicable Notes       9

	Specifies option to print the sibship specific correla- tion matrices for dependent variable.	
print_correlation_matrix	Value Range	Character string representing the name of a marker or location listed in the IBD sharing file.
	Default Value	None
	Required	No
	Applicable Notes	10

Notes

- 1. The value of a trait parameter should be set to the name of a trait, phenotype or covariate field read from the pedigree data file. If no valid trait parameters are listed, then all trait fields read in are used. If more than one trait is specified, then multiple univariate regressions are performed using each trait with all markers and covariates listed. The *population* mean of the trait may be used in computing the mean-corrected trait values. This is specified by including an attribute, mean, with value set to the desired trait mean. Otherwise, the trait mean is estimated from the sample of individuals used in the regression.
- 2. The value of a marker parameter should be set to the name of a marker for which IBD sharing information was generated and stored in the IBD sharing file. If no valid marker parameters are listed then all markers are used.
- 3. If a marker parameter has the dom or dominance attribute, then the additive and dominance variances due to that marker will be tested separately (i.e. there will be regression on both  $\hat{\pi}$  and  $\hat{f}_2$ ); and a marker parameter without this attribute will test total genetic variance due to that marker (i.e. there will be regression on  $\hat{\pi}$  only).
- 4. The value of a covariate parameter should be set to the name of a trait, phenotype or covariate field read from the pedigree data file. If no valid covariate parameters are listed, then by default no covariates are included.
- 5. The interaction parameter should contain a sub-block of marker and covariate parameters that specify a multiplicative interaction term in the regression model. e.g., the following interaction sub-block specifies a gene-environment interaction term between the dominance component of D1S344 and the squared BMI difference:

```
interaction {
    marker = D1S344, dom
    covariate = BMI, diff, power = 2
}
```

6. The values for the regression\_method are explained as follows:

Value	Meaning
diff	$-\frac{1}{2}$ squared trait difference ( $-\frac{1}{2} \times$ traditional Haseman-Elston).
sum	$\frac{1}{2}$ squared mean-corrected trait sum.
prod	Mean-corrected cross-product
W2	Weighted combination of squared trait difference and squared mean-corrected trait sum. Weights are chosen proportional to the inverses of the residual variances of the squared differences and sums.
W3	Weighted combination of squared trait difference and squared mean-corrected trait sum, as above but further adjusted for the non-independence of sib-pairs. <sup><i>a</i></sup>
W4	Weighted combination of squared trait difference and squared mean-corrected trait sum, as above but further adjusted for the non-independence of sib-pairs and the non-independence of squared trait sums and differences. <sup><math>b</math></sup>

<sup>a</sup>This method should be more powerful asymptotically (see Shete, et al., 2003)

<sup>b</sup>This method should be the most powerful asymptotically (see Shete, et al., 2003)

- 7. The subset parameter specifies a trait to be used as an indicator variable to limit the individuals that may be used in an analysis; individuals for whom this indicator is zero are assumed to have missing trait values. It may be included more than once, in which case the only individuals included in the analysis are those for which all the indicated binary traits are coded 1. The trait being analyzed for linkage should not be used as a subset variable. If the trait specified is a binary trait, it should be coded as 0 for individuals to be excluded from analysis and 1 for individuals to be included. Only those individuals that are affected will be considered. If the trait is continuous, only individuals with trait values greater than 0 will be included. This option does not alter the direction of any of the test statistics as would the trait parameter, so it is usually not appropriate to specify subsets based on phenotypes that are useful for testing linkage.
- 8. If the wide\_out parameter is set to **true**, then additional columns are added to the output from Trait Regression analyses, including a column of t-values corresponding to each parameter estimate.
- 9. If either the zero\_marker or multiple\_marker attribute is specified for the sibpal parameter, then no value is required to specify the location. If the single\_marker attribute is specified, then a character string representing the name of a marker or location listed in the IBD sharing file may be used to specify the location to print. If no value is specified for single marker regression, then the first n rows of the design matrix for all locations will be printed.
- 10. If either the zero\_marker or multiple\_marker attribute is specified for the sibpal parameter, then no value is required to specify the location. If the single\_marker attribute is specified, then a character string representing the name of a marker or location listed in the IBD sharing file must be used to specify the location to print.

## **10.4 Program Execution**

SIBPAL is run via a command line interface on the supported UNIX and Windows platforms. This requires the S.A.G.E. programs to be properly installed and in the current execution path. Input files are specified on the command line and all output files are created in the current working directory.

Running SIBPAL from the command prompt with no arguments, or the wrong number of arguments, will result in the program printing its usage statement. This lists the input files the program requires on the command line.

```
>sibpal
S.A.G.E. v5.x -- SIBPAL
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
usage: sibpal <parameters> <pedigree> <IBD...>
Command line parameters:
parameters - parameter file
pedigree - pedigree data file
IBD - IBD sharing file(s)
```

As indicated in the program usage statement, input files are listed on the command line. A typical run of SIBPAL may look like the following:

```
>sibpal sibpal.par example.ped example.ibd
S.A.G.E. v5.x -- SIBPAL
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
Loading parameters.....done.
Reading pedigree data.....
from example.ped.....done.
Sorting pedigrees.....done.
Sorting pairs.....done.
Computing mean test......done.
Computing mean test......
Computing mean test......
Computing trait regression 1..done.
Analysis complete!
```

### **10.5 Program Output**

SIBPAL produces several output files that contain results and diagnostic information:

Filename	File Type	Description
sibpal.inf	Information output file	Contains informational diagnostic
		messages, warnings and program er-
		rors. No analysis results are stored in
		this file.
means.out	Mean analysis output file	Contains the results of each test of
		mean allele sharing IBD.
	Trait Regression analysis	
traits.out	output file	Contains the results of each linkage
		test.

### **10.5.1** Information Output File

The SIBPAL Information Output file contains a variety of useful information, including:

- Information on fields read from the pedigree data file. These tables, which provide information about what the program has read in, are included with all programs in the S.A.G.E. and are very useful for debugging most common errors caused when reading the pedigree data. When first analyzing new data, it is recommended that these tables be checked carefully to make sure pedigree data are being correctly read.
- Information, warning and error messages generated throughout the program. It is recommended that you check this file for warning and error messages before examining the results of any run of the program. The program attempts to correct many common errors and this sometimes means analyses are not run as expected. The file "sibpal.inf" should be checked for errors and diagnostic information after each run of the program.

### 10.5.2 Mean Analysis Output File

One Mean Analysis output file, named "means.out", is generated per run of SIBPAL. It includes the results of all non-trait specific mean analyses: average allele sharing, as well as 0, 1 and 2 alleles IBD are output in a table with standard errors and p-values for each estimate.

### 10.5.3 Trait Regression Analysis Output File

One Trait Regression analysis output file, named "traits.out", is generated per run of SIBPAL. It contains the results of all Trait Regression linkage tests. Each coefficient estimated is printed in a table with its standard error and p-value.

## **10.6 Example Output Files**

## 10.6.1 Mean Analysis Output File

	ean Allele Shar				
Estimates pi - Ave fi - Est	erage proportio	on of alle	les shared l b pairs shar	IBD. ring i all	eles IBD.
======================================	Pairs	Estir		d Error	P-value
D5G1	f	1     0.29       1     0.50	464286 0. 000000 0.	04820449 05310089 02571722 04645782	0.36259122038 0.40789006347 0.34511298675
D5G2	f f	E00.30E10.51	357143 0. 785714 0.	05696380 07426574 07919128 06410910	0.28224850647 0.47689606098  0.27502563114
D5G3	f f	0.17 0.55	857143 0. 357143 0.	05310089 .05729972 .06967083 .06916042	0.40789006347 0.22325916300  0.79821143299
D5G4	f f	0         0.17           1         0.53	857143 0. 571429 0.	04699492 04786642 05709323 06128351	0.26432288020 0.14722754816 0.56488301140
 D5G5	f f	0.19 0.55	642857 0. 357143 0.	06212092 06967083 09033772 08333333	0.66975583933 0.44861400075 1.000000000000

. .

### **10.6.2** Trait Regression Analysis Output File

SIBPAL Output -- 2 Aug 2005 15:43:12 -- [S.A.G.E. v5.0.3; bld 01 Aug 2005] COPYRIGHT (C) 2005 CASE WESTERN RESERVE UNIVERSITY \_\_\_\_\_ Haseman-Elston Regression Analysis of Full Sibs - single\_marker regression \_\_\_\_\_ Binary trait : affection, affected = 'A' , unaffected = 'U' Number of full sib pairs = 28 Sample mean = 0.0370 = 0.0370 Sample variance Sample skewness = 4,9029 Sample kurtosis = 22.0385 Sibling correlation 0.0610 = Dependent variate : Squared trait difference Correlation between pairs with no sibs in common = 0.0000 Correlation between pairs with one sib in common = 0.0000 Correlation between squared difference and squared mean corrected sum = -1.0000 -0.0121 Intercept = Total variance 0.0086 = Residual variance = 0.0093 Residual skewness -4.9942 = Residual kurtosis 22.9761 = Other options used : Identity weights = no Robust variance = no Use sibship mean = no Legend : Note: kurtosis = coefficient of kurtosis - 3 - significance .05 level; \*\* - significance .01 level; \*\*\* - significance .001 level; \_\_\_\_\_ Independent Nominal variable Pairs Parameter Estimate Std Error P-value \_\_\_\_\_ \_\_\_\_ D5G1 28 (A+D)GenVar -0.0127 0.0726 0.5688081 -0.0127 0.0614 0.5813646 D5G2 28 (A+D)GenVar 0.0107 D5G3 28 (A+D)GenVar 0.0659 0.4361386 0.0163 D5G4 28 (A+D)GenVar 0.0748 0.4147840 28 (A+D)GenVar D5G5 0.0466 0.0557 0.2047478 

 28
 (A+D)GenVar
 0.0466
 0.0557
 0.2047478

 28
 (A+D)GenVar
 0.0076
 0.0629
 0.4525494

 28
 (A+D)GenVar
 0.0667
 0.0691
 0.1715900

 28
 (A+D)GenVar
 0.0000
 0.0755
 0.5000000

 28
 (A+D)GenVar
 -0.0512
 0.0587
 0.8046652

 28
 (A+D)GenVar
 0.0000
 0.0681
 0.5000000

 28
 (A+D)GenVar
 0.0042
 0.0662
 0.4748203

 28
 (A+D)GenVar
 0.0085
 0.0756
 0.4553801

 D5G6 D5G7 D5G8D5G9 D5G10 D5G11 D5G12 D5G13 . D5G24 28 (A+D)GenVar 0.0141 0.0612 0.4100034 D5G25 28 (A+D)GenVar 0.0942 0.0554 0.0502490 \_\_\_\_\_

## **Chapter 11**

# LODPAL

LODPAL performs a linkage analysis based on the LOD score formulation for affected-sib-pairs (ASPs) (Risch, 1990). The current implementation is of the general conditional logistic model proposed by Olson (1999) modified to give the one-parameter model of Goddard et al. (2001). The model allows for the inclusion of all affected-relative-pairs (ARPs) and covariates or discordant sibling pairs, with the possibility of pooling unaffected relative pairs together with ARPs in the analysis.

## 11.1 Limitations

The current release only includes support for a single disease locus and assumes all pairs of relatives are independent.

## 11.2 Theory

### 11.2.1 Basic notation

Let the number of relative pairs be N.

Let *i* index the relative pair: i = 1, 2, ..., N.

Let  $f_{r0}$ ,  $f_{r1}$ , and  $f_{r2}$  be the prior probabilities of sharing 0, 1, or 2 alleles IBD given a relative pair of type r.

Let  $w_i$  be a weight corresponding to the *i*th pair.

Let

 $\hat{f}_{0i}$  be the probability of sharing 0 alleles IBD at a given marker location, for the *i*th pair,

 $\hat{f}_{1i}$  be the probability of sharing 1 allele IBD at a given marker location, for the *i*th pair, and

 $\hat{f}_{2i}$  be the probability of sharing 2 alleles IBD at a given marker location, for the *i*th pair.

These three IBD-sharing probabilities are estimated by GENIBD given the available marker data and given the pedigree relationship (i.e., type of relative pair). They may be multipoint or single-marker estimates. Marker is equivalent to marker location, and need not be a measured marker. This is mainly an issue dealt with in the IBD generation phase.

# Alleles Shared IBD	Probabilities			
0	$Z_0$	$\frac{1}{\lambda_0+2\lambda_1+\lambda_2}$	$\frac{1}{1+2e^{\beta_1}+e^{\beta_2}}$	
1	$Z_1$	$\frac{2\lambda_1}{\lambda_0 + 2\lambda_1 + \lambda_2}$	$\frac{2e^{\beta_1}}{1+2e^{\beta_1}+e^{\beta_2}}$	
2	$Z_2$	$rac{\lambda_2}{\lambda_0+2\lambda_1+\lambda_2}$	$\frac{e^{\beta_2}}{1+2e^{\beta_1}+e^{\beta_2}}$	

The following table summarizes the various notations that have been used for the probability of sharing *i* alleles IBD between affected sib pairs at a particular locus:

The sibling locus-specific relative recurrence risk is given by

$$\lambda_s = \frac{1}{4} \left[ \lambda_0 + 2\lambda_1 + \lambda_2 \right] = \frac{1}{4} \left[ 1 + 2\lambda_1 + \lambda_2 \right] = \frac{1}{4} + \frac{1}{2}\lambda_1 + \frac{1}{4}\lambda_2$$

### 11.2.2 Affected Relative Pair Linkage Analysis

### 11.2.2.1 Two-parameter Model (Olson 1999)

The LOD score for a set of N ARPs is

$$z = \sum_{i=1}^{N} \log_{10} \left\{ w_i \frac{\hat{f}_{0i} + \hat{f}_{1i} e^{\beta_1} + \hat{f}_{2i} e^{\beta_2}}{f_{r0} + f_{r1} e^{\beta_1} + f_{r2} e^{\beta_2}} + (1 - w_i) \right\}$$

$$=\sum_{i=1}^{N}\log_{10}\left\{w_{i}\frac{\sum\limits_{k=0,1,2}^{\infty}\hat{f}_{ki}e^{\beta_{k}}}{\sum\limits_{k=0,1,2}^{\infty}f_{rk}e^{\beta_{k}}}+(1-w_{i})\right\}=\sum_{i=1}^{N}\log_{10}\left\{w_{i}\frac{\sum\limits_{k=0,1,2}^{\infty}\hat{f}_{ki}\lambda_{k}}{\sum\limits_{k=0,1,2}^{\infty}f_{rk}\lambda_{k}}+(1-w_{i})\right\},$$

where  $\lambda_k$  is the (locus-specific) relative recurrence risk for an individual sharing k alleles IBD with an affected person and  $w_i$  is a user-specified weight to be given to the i-th relative pair. Here,  $\beta_0 =$ 0, and  $\beta_1$ ,  $\beta_2$  are estimated by maximizing the LOD score with the constraints  $\beta_1 \ge 0$  and  $\beta_2 \ge \log_e (2e^{\beta_1} - 1)$  (i.e.,  $\lambda_1 > 1$  and  $\lambda_2 > 2\lambda_1 - 1$ ).

For full sibs,  $f_{S_0} = \frac{1}{4}, f_{S_1} = \frac{1}{2}, f_{S_2} = \frac{1}{4}$ , giving for the *i*th sib pair

$$\log_{10} \left\{ w_i 4 \frac{\hat{f}_{0i} + \hat{f}_{1i} e^{\beta_1} + \hat{f}_{2i} e^{\beta_2}}{1 + 2e^{\beta_1} + e^{\beta_2}} + (1 - w_i) \right\}.$$

For half sibs,  $f_{h0} = \frac{1}{2}$ ,  $f_{h1} = \frac{1}{2}$ ,  $f_{h2} = 0$ , giving for the *i*th half sib pair

$$\log_{10} \left\{ w_i 2 \frac{\hat{f}_{0i} + \hat{f}_{1i} e^{\beta_1}}{1 + e^{\beta_1}} + (1 - w_i) \right\}.$$

In summary,

r	$\mathbf{f_{r0}}$	$\mathbf{f_{r1}}$	$\mathbf{f_{r2}}$
Sibs	$\frac{1}{4}$	$\frac{1}{2}$	$\frac{1}{4}$
Half-sibs	$\frac{1}{2}$	$\frac{1}{2}$	0
Grandparents	$\frac{1}{2}$	$\frac{1}{2}$	0
Avuncular	$\frac{1}{2}$	$\frac{1}{2}$	0
Cousins	$\frac{3}{4}$	$\frac{1}{4}$	0

In the next sections, the subscript *i* indexing the pair and the summation over *i* will be suppressed.

### 11.2.2.2 One Parameter Model

Under the optimal one parameter model, the LOD score contribution of a single pair is

$$\log_{10} \left\{ w \frac{\hat{f}_0 + \hat{f}_1 e^{\beta_1} + \hat{f}_2 (3.634 e^{\beta_1} - 2.634)}{f_{r0} + f_{r1} e^{\beta_1} + f_{r2} (3.634 e^{\beta_1} - 2.634)} + (1 - w) \right\}.$$

The constants in the above expression fix the mode of inheritance to a value approximately halfway between a dominant and a recessive model and correspond to the Whittemore and Tu (1998) minmax model mode of inheritance parameter (they defined two parameters,  $w_1$  and a, and the minmax values of these are  $w_1 = 0.275$  and  $\alpha = (2 - 3a)/a$ . To allow more flexibility, the user may specify a different "mode of inheritance" parameter. Thus, under a generalization of this model, the LOD score contribution of a single pair is

$$\log_{10} \left\{ w \frac{\hat{f}_0 + \hat{f}_1 e^{\beta_1} + \hat{f}_2[(\alpha + 1)e^{\beta_1} - \alpha]}{f_{r0} + f_{r1}e^{\beta_1} + f_{r2}[(\alpha + 1)e^{\beta_1} - \alpha]} + (1 - w) \right\},\$$

where  $\alpha \ge 1$  is a mode of inheritance parameter:  $\alpha = 1$  (corresponding to Whittemore & Tu's  $w_1 = a = 0.5$ ) gives a dominant model and  $\alpha \to \infty$  (corresponding to Whittemore & Tu's  $w_1 = a = 0$ ) gives a recessive model. (In practice,  $\alpha \approx 10$  gives a pretty good recessive model.) Compared to the two-parameter model, this model has the constraints  $\lambda_2 = (\alpha + 1) \lambda_1 - \alpha$  in terms of relative recurrence risks and we estimate  $\beta_1 \ge 0$  (default).

### **11.2.2.3** Covariates<sup>1</sup>

Inclusion of a single covariate (z) gives

<sup>&</sup>lt;sup>1</sup>Covariates are pair-specific and are allowed only in the one-parameter model.

$$\log_{10} \left\{ w \frac{\hat{f}_0 + \hat{f}_1 e^{\beta_1 + y\delta_1} + \hat{f}_2[(\alpha + 1)e^{\beta_1 + z\delta_1} - \alpha]}{f_{r0} + f_{r1}e^{\beta_1 + y\delta_1} + f_{r2}[(\alpha + 1)e^{\beta_1 + z\delta_1} - \alpha]} + (1 - w) \right\},\$$

where  $\delta_1$  is an additional parameter to be estimated and z is the adjusted (see below) value of the covariate for that pair. The model extends easily to include more than one covariate; there is one additional parameter for each covariate.

• Constraints on  $\delta_1$ :

Let the original (unadjusted) covariate value be denoted x. Two options are allowed:

- 1. Genetic constraint on  $\beta_1$  holds at the average value,  $x = \bar{x}$ , but not necessarily for all x. The covariate value is centered to give  $z = x - \bar{x}$  before inclusion in the likelihood, so that the mean of the centered covariate = 0. Then  $\delta_1$  is unconstrained.
- 2. Genetic constraint on  $\beta_1$  holds at all values of x. The minimum value of a covariate is subtracted (i.e.,  $z = x \min x$ ), so that the smallest value of the covariate equals zero. Then for a set of covariates, indexed by j, the following constraint is applied:

$$\min_{y>0}\sum_j z_j\delta_{1j} \ge -\beta_1.$$

# 11.2.3 Adding Discordant Sib Pairs (DSP) to an ARP Analysis (one-parameter model only)

This model is the same as the ARP one parameter model with one covariate that indicates nonconcordance status:

$$\lambda_1 = e^{\beta_1 + y\delta_1},$$

where the covariate y is set to 0 if the pair is concordant and 1 if the pair is discordant for affection status. A related model sets the covariate y to

- 0 if the pair is concordantly affected and to
- 1 if the pair is discordant for affection status (concordantly unaffected pairs are not used in the analysis).

When either option is chosen,  $\beta_1$  and  $\delta_1$  are estimated subject to the constraints:  $\beta_1 \ge 0$ ,  $\delta_1 \le -\beta_1$ . No additional covariates may be included when discordant sib pairs are included in the analysis this way.

### 11.2.4 X-linked Models

Models for X-linkage are similar to those for autosomal inheritance. Recall the autosomal model: the LOD score contribution for a particular affected relative pair (ARP) of type r is

$$\log_{10} \left\{ w \frac{\sum_{k=0,1,2} \hat{f}_k \lambda_k}{\sum_{k=0,1,2} f_{rk} \lambda_k} + (1-w) \right\}.$$

For X-linked models, the LOD score is

$$\log_{10} \left\{ w \frac{\sum\limits_{k=0,1,2} \hat{f}_{kuv} \lambda_{kuv}}{\sum\limits_{k=0,1,2} f_{rkuv} \lambda_{kuv}} + (1-w) \right\},\$$

where u, v denote the sex (m=male, f=female) of the members of the pair. For male-female ARPs, m and f are interchangeable, i.e.  $\lambda_{kmf} = \lambda_{kfm}$ .

There are four possible relative risk parameters:  $\lambda_{1ff}$ ,  $\lambda_{2ff}$ ,  $\lambda_{1mm}$ , and  $\lambda_{1mf}$  (= $\lambda_{1fm}$ ). All others equal 1 (e.g.,  $\lambda_{0ff} = \lambda_{2mm} = 1$ , etc.). The following table gives the  $\lambda$  parameters for each type of ARP.

		$\lambda$ Corresponding to IBD-sharing equal to		
Type of ARP		0	1	2
Male-Male		1	$\lambda_{1mm}$	1
Male - Female		1	$\lambda_{1mf}$	1
Female - Female	ASP	1	$\lambda_{1ff}$	$\lambda_{2ff}$
Female - Female	Other types	1	$\lambda_{1ff}$	1

- Constraints on  $\lambda_{1ff}$ ,  $\lambda_{1mm}$ ,  $\lambda_{1mf}$ :
  - DEFAULT VALUE : all  $\lambda_1$  constrained to be equal:

$$\lambda_{1ff} = \lambda_{1mm} = \lambda_{1mj}$$

– OPTIONAL : all  $\lambda_1$  not constrained to be equal:

 $\lambda_{1ff}$ ,  $\lambda_{1mm}$ , and  $\lambda_{1mf}$  are estimated separately.

- Constraints on  $\lambda_{2ff}$ :
  - DEFAULT VALUE :  $\lambda_{2ff} = (\alpha+1) \lambda_{1ff} \alpha$ The default value of  $\alpha$  is 2.634.
  - OPTIONAL :  $\lambda_{2ff}$  is not constrained to be dependent on  $\lambda_{1ff}$ .

 $\lambda_{1ff}$  and  $\lambda_{2ff}$  are estimated separately. Since both parameters are not estimable if the data contains only ASPs or no ASPs, unless  $\lambda_{1ff}$  is estimated in part using male-male and/or male-female ASPs, this option will be carried out only if either the data contains at least 15 male-male and male-female sib pairs (ASPs) under the default constraints on  $\lambda_1$ , or if the data set contains at least 15 sister-sister ASPs and at least 15 female-female ARPs other than ASPs under the optional constraints on  $\lambda_1$ .

Under this model, the additional constraint  $\beta_{2ff} \ge \log_e (2e^{\beta_{1ff}} - 1)$  is used<sup>2</sup>.

### 11.2.4.1 Covariates

Inclusion of a single covariate (z) gives

$$\log_{10} \left\{ w \frac{\hat{f}_{0uv} + \hat{f}_{1uv} e^{\beta_{1uv} + \delta_{1uv}z} + \hat{f}_{2uv}[(\alpha + 1)e^{\beta_{1uv} + \delta_{uv1}z} - \alpha]}{f_{r0uv} + f_{r1uv} e^{\beta_{1uv} + \delta_{1uv}z} + f_{r2uv}[(\alpha + 1)e^{\beta_{1uv} + \delta_{1uv}z} - \alpha]} + (1 - w) \right\},$$

where  $\delta_{1uv}$  is an additional parameter to be estimated and y is the adjusted value of the covariate for that pair as with the autosomal models. The model extends easily to include more that one covariate; there is one additional parameter for each covariate.

Under a generalization of this model, the LOD score is

$$\log_{10} \left\{ w \frac{\hat{f}_{0uv} + \hat{f}_{1uv} e^{\beta_{1uv} + \sum_{l} \delta_{luv} z_{l}} + \hat{f}_{2uv}[(\alpha + 1)e^{\beta_{1uv} + \sum_{l} \delta_{luv} z_{l}} - \alpha]}{f_{r0uv} + f_{r1uv} e^{\beta_{1uv} + \sum_{l} \delta_{luv} z_{l}} + f_{r2uv}[(\alpha + 1)e^{\beta_{1uv} + \sum_{l} \delta_{luv} z_{l}} - \alpha]} + (1 - w) \right\},$$

where *l* indexes the covariate. Note that covariates can only be included in the one-parameter model, with the constraint  $\lambda_{2ff} = (\alpha+1) \lambda_{1ff} - \alpha$ .

Constraints on the  $\delta s$  are the same as with the autosomal models.

### 11.2.5 Parent-of-Origin Models

The expression of an allele may depend on the sex of the parent from whom the allele was inherited; this phenomenon is known as a parent-of-origin effect (alternatively, genetic imprinting). For example, individuals affected with the autosomal dominant condition Beckwith-Wiedemann syndrome almost always inherited the defective allele from their mother. Individuals who inherit the defective allele from their father are rarely affected with this disorder.

For the model that includes a parent-of-origin effect, the ARP lod score model fits separate parameters for the maternal and paternal effects. The test of parent-of-origin effect is obtained by comparing the likelihood-ratio statistics (i.e., 4.6 times the lod score) for the models with and without the parent-of-origin effect. The parent-of-origin model can only be applied to autosomal loci.

For the parent-of-origin model, the LOD score for a particular affected relative pair is

$$\log_{10} \left\{ w \frac{\sum\limits_{k=0,1m,1p,2} \hat{f}_k \lambda_k}{\sum\limits_{k=0,1m,1p,2} f_{rk} \lambda_k} + (1-w) \right\},$$

where *m* denotes maternal and *p* denotes paternal, so that the sum is over k = 0, *1m*, *1p*, *2* rather than k = 0, 1, 2. As in previous models,  $\lambda_0 = 1$  and  $\lambda_k = \exp(\beta_k)$ , where  $\beta_k$  is the parameter estimated.

<sup>&</sup>lt;sup>2</sup>As with the autosomal models,  $\beta$  in  $\lambda_{kuv} = \exp(\beta_{kuv})$  will be estimated instead of estimating  $\lambda$  itself.

### 11.2.5.1 One Parameter Model

First note that  $\lambda_1 = \frac{\lambda_{1m} + \lambda_{1p}}{2}$ . The one-parameter model employs the same mode-of-inheritance constraint, i.e.,  $\lambda_2 = (\alpha + 1) \lambda_1 - \alpha$ .

### 11.2.5.2 Covariates

Covariates may be included only in the one-parameter model. Inclusion of a single covariate (z) gives  $\lambda_{1m} = e^{\beta_{1m} + \delta_{1m}z}$  and  $\lambda_{1p} = e^{\beta_{1p} + \delta_{1p}z}$  where  $\delta_{1m}$  and  $\delta_{1p}$  are the additional parameters to be estimated and z is the adjusted value of the covariate for that pair, as with the autosomal models. The model extends easily to include more that one covariate; there are two additional parameters for each covariate, so that  $\lambda_{1m} = e^{\beta_{1m} + \sum_l \delta_{lm}z_l}$  and  $\lambda_{1p} = e^{\beta_{1p} + \sum_l \delta_{lp}z_l}$ , where l indexes the covariate under the generalization of this model. We include an option that fixes either  $\lambda_{1m}$  or  $\lambda_{2m}$  to be equal to 1. In such situations, only one covariate parameter is fitted for each covariate.

Constraints on the  $\delta s$  are the same as with the other autosomal models.

These models only apply to ASPs and affected half-sib pairs because the problem of computing the right IBD probabilities for other types of ARPs is daunting. By default, other types of ARPs are excluded from the analysis, but an option to include other types into the analysis is provided. (It should be recognized that, for other types of ARPs, parent-of-origin effects may be highly confounded with ascertainment.) When other types of ARPs are included in an analysis with a parent-of-origin effect  $\lambda_1$  is replaced with  $\{(\lambda_{1m} + \lambda_{1p})/2\}$  in other ARPs to avoid fitting an extra parameter. The only information about parent-of-origin effect in the models comes from the ASPs, and so parent-of-origin models will not be allowed if the number of ASPs in which  $\hat{f}_{1m} \neq \hat{f}_{1p}$  is less than 10, even if many other ARPs are available.

File Type	Description
LODPAL parameter file	Specifies the parameters and options with which
	to perform a particular analysis.
Pedigree data file	Contains delimited records for each individual
	including fields for identifiers, sex, parents, trait
	(including individual-specific covariate values)
	and marker data.
IBD sharing file	Stores identity-by-descent (IBD) distributions
	between pairs of related individuals at one or
	more marker loci.
Pair information file (Optional)	Contains character delimited records for each
	known relative pair including fields for
	identifiers, weights, and pair-specific covariate
	values.

## 11.3 Program Input

Notes

To use the X-linked model in LODPAL, an IBD sharing file has to include "x\_linked" after the name of the marker in the file header.

## **11.3.1** Parameter File Syntax

## 11.3.1.1 The lodpal Parameter

The following syntax table specifies the permissible parameter and attribute settings for the main LODPAL parameter.

<pre>parameter [, attribute]</pre>	Explanation		
	1	ar LOD score analysis for affected	
	relative pairs.		
lodpal	Value Range	N/A	
lodpal	Default Value	None	
	Required	Yes	
	Applicable Notes	None	
	Specifies the root name to be used for output files.		
	Output file names will be formed by concatenating the		
	root name and an appropriate extension.		
out		Character string representing a	
, out	Value Range	valid file name.	
	Default Value	lodpal	
	Required	No	
	Applicable Notes	None	

## 11.3.1.2 The lodpal Block

The following syntax table specifies the permissible parameter and attribute settings for a LODPAL block.

parameter	Explanation		
[, attribute]	Explanation		
	Specifies a binary trait to be used in the current analysis.		
trait	Value Range	Character string representing the name of a binary trait, phenotype or covariate listed in	
		the pedigree data file.	
	Default Value	None	
	Required	Yes	
	Applicable Notes	1	
	Traits that are not value.	t binary are dichotomized at this	
	Value Range	$(-\infty,\infty)$	
, cutpoint	Default Value	0	
	Required	No	
	Applicable Notes	1	
	Specifies option to analyze affected relative pairs only.		
	Value Range	N/A	
, conaff	Default Value	None	
	Required	No	
	Applicable Notes	1	
	Specifies option to	pool concordantly affected relative	
	pairs with concordantly unaffected sib pairs, and in-		
	clude discordant si	b pairs in the analysis.	
, condisc	Value Range	N/A	
	Default Value	None	
	Required	No	
	Applicable Notes	1	
	Specifies option to	analyze concordantly affected rel-	
	ative pairs and disc	•	
, noconunaff	Value Range	N/A	
, noconunari	Default Value	None	
	Required	No	
	Applicable Notes	1	

		use only a subset of the data. The be a binary trait used as an indica-		
subset	Value Range	Character string representing the name of a binary trait listed in the pedigree data file.		
	Default Value	None		
	Required	No		
	Applicable Notes	2		
	Specifies a marker ysis.	to be included in the current anal-		
marker	Value Range	Character string representing the name of a marker listed in the pedigree data file.		
	Default Value	None		
	Required	No		
	Applicable Notes	3		
	Specifies a covari only.	ate for the one-parameter model		
		Character string representing the		
	Value Range	name of a covariate listed in the		
covariate	Ŭ	pedigree data file.		
	Default Value	None		
	Required	No		
	Applicable Notes	4		
	Covariate terms are taken to the power specified.			
	Value Range	$(-\infty,\infty)$		
, power	Default Value	1		
	Required	No		
	Applicable Notes	None		
		include the covariate sum.		
	Value Range	N/A		
, sum	Default Value	None		
	Required	No		
	Applicable Notes	4		
		include the covariate difference.		
	Value Range	N/A		
, diff	Default Value	None		
	Required	No		
	Applicable Notes	4		
		o include covariate terms: sum &		
	difference.			
	Value Range	N/A		
, both	Value Range Default Value	N/A None		
, both				

	Specifies option to	include the covariate product.
	Value Range	N/A
, prod	Default Value	None
	Required	No
	Applicable Notes	4
		include the covariate average.
	Value Range	N/A
, avg	Default Value	None
	Required	No
	Applicable Notes	4
	Specifies option to	include the covariate value for only
	the first member of	
	Value Range	N/A
, single	Default Value	None
	Required	No
	Applicable Notes	4
	Specifies option to	center covariates around the ob-
	served mean or the	user-supplied value.
	Value Range	$(-\infty,\infty)$
, mean	Default Value	observed mean
	Required	No
	Applicable Notes	4
	Specifies option to	set covariate values as offsets from
	the smallest observ	ved value.
	Value Range	N/A
, minimum	Default Value	None
	Required	No
	Applicable Notes	4
	Specifies option to	print diagnostic information to a
	separate file.	
		Character string representing the
diagnastic	Value Range	name of a marker or location
diagnostic		listed in the IBD sharing file.
	Default Value	None
	Required	No
	Applicable Notes	5
	Specifies option to	b disable the default maximization
	process.	
turn off default	Value Range	N/A
turn_off_default	Default Value	None
	Required	No
	Applicable Notes	6

	Specifies option to ysis.	use only full sib-pairs in the anal-		
., . ,	Value Range	N/A		
sib_pairs_only	Default Value	None		
sib	Required	No		
	Applicable Notes	None		
	Specifies option to	print more verbose output informa-		
	tion. This causes some output tables to be more than			
	80 columns wide.			
wide_out		true		
wide_out	Value Range	false		
	Default Value	false		
	Required	No		
	Applicable Notes	7		
	Specifies a variable	e name to be used as weight in the		
	current analysis.			
		Character string representing the		
		name of a trait, phenotype or		
weight	Value Range	covariate listed in the pedigree		
		data file.		
	Default Value	None		
	Required	No		
	Applicable Notes	8		
	Starts a sub-block for specification of pair-specific co-			
	variate(s) and/or w	reight values to be used in the cur-		
	rent analysis.			
pair_info_file	Value Range	Character string representing a		
parr_rmo_rme	value Italiye	valid file name.		
	Default Value	None		
	Required	No		
	Applicable Notes	9		
		for specification of an autosomal		
	model on an existin	ng autosomal marker.		
autosomal	Value Range	N/A		
autosomal_model	Default Value	None		
aucosonal_model	Required	No		
	Applicable Notes	10		
		for specification of an X-linked		
	model on existing			
x_linkage	Value Range	N/A		
x_linkage_model	Default Value	None		
A_TIIKAYE_MOUET	Required	No		
	Applicable Notes	11		

Notes

1. The value of a trait parameter should be set to the name of a binary trait, phenotype or covariate field read from the pedigree data file.

- (a) If no valid trait parameters are listed, then all trait fields read in from the pedigree data file are used.
- (b) If more than one trait is specified, then each will be used in a separate analysis.
- (c) If a trait is not a binary trait, then it will be dichotomized at 0 (trait values <= 0 will be treated as unaffected and values > 0 will be treated as affected) or at the value of the cutpoint attribute. When dichotomizing a trait using a cutpoint, all values less than or equal to the cutpoint are considered unaffected and all values strictly greater than the cutpoint are considered to be affected.
- (d) If a trait parameter has the conaff attribute, then the program selects only concordantly affected relative pairs (ARPs) and performs an analysis on these pairs.
- (e) If a trait parameter has the condisc attribute, then the program pools the concordantly affected relative pairs with the concordantly unaffected sib pairs and performs a one-parameter model analysis in which these are analyzed together with the discordant sib pairs (DSP) by creating a covariate to indicate concordance status of the pairs.
- (f) If a trait parameter has the noconunaff attribute, the program performs the same analysis as with the condisc attribute, but without including the concordantly unaffected sib pairs.
- (g) When either the condisc attribute or the noconunaff attribute is used, no covariates can be included. If the user specifies any covariates, they are ignored by the program.
- (h) If no attributes are listed, then by default an ARP analysis is performed.
- 2. The trait specified by a subset parameter should be a binary trait coded as 0 for individuals to be excluded from, and 1 for individuals to be included in, the analysis. The subset parameter may be included more than once, in which case the only individuals included in the analysis are those for which all the indicated binary traits are coded 1.
- 3. The value of a marker parameter should be set to the name of a marker (or marker location) for which IBD sharing information was generated and stored in the IBD sharing file. The marker parameter may be included more than once. If no valid marker parameters are listed then all markers are used.
- 4. The value of a covariate parameter should be set to the name of a trait, phenotype or covariate field read from the pedigree data file. The covariate parameter may be included more than once. A covariate parameter may have two attributes. One is to specify the function to compute a pair-specific value from two individual-specific values (sum, diff, single, avg or prod), and the other is to adjust the covariate value to impose genetic constraints on them (mean or minimum).
  - (a) If sum, diff, avg or prod attributes are specified, then a single covariate sum, difference, average or product term of two individual-specific values is included as a pair-specific covariate value.
  - (b) If the both attribute is specified, then both sum and difference terms are included.
  - (c) If the single attribute is specified, then the covariate value for the first member of the pair is included as a pair-specific value.
  - (d) If no attribute of this kind is specified, then the sum is included by default.

- (e) If the mean attribute is specified (the default), then the program automatically centers each pair-specific covariate value before inclusion in the likelihood, using the sample mean or a user-supplied value (for example, mean = 0.5).
- (f) If the minimum attribute is included, then the program automatically puts the offset from the smallest observed covariate value as the pair-specific covariate value into the likelihood, so that the smallest value of the pair-specific covariate equals zero.
- 5. The value of a diagnostic parameter should be set to the name of a marker or a location (in centiMorgans) for which IBD sharing information was generated and stored in the IBD sharing file. If the diagnostic parameter has a valid value, then an additional output file, "LODPAL.lod", will be generated that contains the individual pair LOD score contributions for the final model at the particular location specified by the diagnostic parameter value.
- 6. If the program finds the turn\_off\_default parameter, then the program maximizes the LOD score in a somewhat simpler way than the default way. By default, the program uses a method that avoids, as much as possible, spuriously high LOD scores. However, because there may be multiple true maxima, the result obtained using the turn\_off\_default parameter may also be of interest.
- 7. If the wide\_out parameter is set to **true**, then additional columns are added to the output of the LOD score Analysis of Affected Relative Pairs. The information contained in the additional columns are :
  - detailed info on the number of pairs
  - first derivatives of parameter estimates
  - the number of iterations it took for maximization
  - the return flag from the MAXFUN library function
- 8. The value of a weight parameter should be set to the name of a trait, phenotype or covariate field read from the pedigree data file. The weight value for the first member of the pair is included as a pair-specific value. Weights must be between 0 and 1, inclusive (i.e., in [0,1]). Pairs with weights outside this interval will not be included in the analysis. If a pair-specific weight from the pair information file is to be specified, it must instead be specified within the pair\_info\_file sub-block using the pair\_weight parameter (see note 9). Only one weight, specified either by the weight parameter or by the pair\_weight parameter in pair\_info\_file sub-block (see note 9) may be included; and weights are given values of 1 for all pairs in the analysis by default, if neither a weight parameter nor a pair\_weight parameter is found.
- 9. If the program finds the pair\_info\_file parameter with a valid file name, then the program uses the pre-constructed pair-specific covariate and/or weight values from the file name specified. The pair\_info\_file parameter may have its own sub-block to specify the name of the pair-specific covariate(s) and/or weight to be used in the current analysis.
- 10. If the program finds the autosomal or autosomal\_model parameter, then the program uses the specified autosomal model for the autosomal locations. The autosomal parameter may have its own sub-block to specify the model to be used in the current analysis. If no sub-block is found, the default autosomal model will be used, i.e., the one-parameter model with the default alpha value, and without parent-of-origin effect.

11. If the program finds the x\_linkage or x\_linkage\_model parameter, then the program uses an X-linked model for the X-linked markers. The x\_linkage parameter may have its own sub-block to specify the model to be used in the current analysis. If no sub-block is found, the default X-linked model will be used, in which all three  $\lambda_1$  parameters are constrained to be equal, and  $\lambda_{2ff}$  is fixed.

### 11.3.1.3 The pair\_info\_file Sub-Block

The following syntax table specifies the permissible parameter and attribute settings for a pair\_info\_file sub-block.

parameter	Explanation			
[, attribute]	-			
	Specifies a variable name to be used as a covariate in			
	the current test.			
		Character string representing the		
	Value Range	name of a trait, phenotype or		
pair_covariate		covariate listed in the pair		
		information file.		
	Default Value	None		
	Required	No		
	Applicable Notes	1		
	Specifies option to	center covariates around the ob-		
	served mean or the	user-supplied value.		
	Value Range	$(-\infty,\infty)$		
, mean	Default Value	observed mean		
	Required	No		
	Applicable Notes	3		
	Specifies option to	set covariate values as offsets from		
	the smallest observed value.			
, minimum	Value Range	N/A		
, miriimum	Default Value	None		
	Required	No		
	Applicable Notes	3		
	Specifies a variable	e name to be used as a weight in the		
	current test.			
		Character string representing the		
	Value Denge	name of a trait, phenotype or		
pair_weight	Value Range	covariate listed in the pair		
		information file.		
	Default Value	None		
	Required	No		
	Applicable Notes	2		

Notes

- 1. The value of a pair\_covariate parameter should be set to the name of a covariate field read from the Pair Information File. The pair\_covariate parameter may be included more than once.
- 2. The value of a pair\_weight parameter should be set to the name of a weight field read from the Pair Information File. Weights must be in the interval [0, 1]. Pairs with weights outside this interval will not be included in the analysis. Only one pair\_weight parameter may be included, and by default weights are given values of 1 for all pairs if no

pair\_weight parameter is specified. If the pair\_info\_file parameter does not have a sub-block, then the program uses every field listed in the pedigree block, with the exception of ID fields, as a mean-centered covariate in the analysis. The pair\_info\_file parameter is ignored for the two-parameter model unless pair\_weight is found in this sub-block.

- 3. The value of a pair\_covariate parameter should be set to the name of a trait, phenotype or covariate field read from the pedigree data file. The pair\_covariate parameter may be included more than once. A pair\_covariate parameter may have an attribute to adjust the covariate value to impose genetic constraints on them (mean or minimum).
  - (a) If the mean attribute is specified (the default), then the program automatically centers each pair-specific covariate value before inclusion in the likelihood, using the sample mean or a user-supplied value (for example, mean = 0.5).
  - (b) If the minimum attribute is included, then the program automatically puts the offset from the smallest observed covariate value as the pair-specific covariate value into the likelihood, so that the smallest value of the pair-specific covariate equals zero.

### 11.3.1.4 The autosomal Sub-Block

The following syntax table specifies the permissible parameter and attribute settings for a autosomal sub-block.

parameter	Explanation		
[, attribute]	-		
	Specifies the type of model to use in the analysis.		
	Value Range	{one_parameter,	
model	value range	two_parameter}	
model	Default Value	one_parameter	
	Required	No	
	Applicable Notes	1	
	Specifies that no ge	enetic constraints on the parameters	
	are to be estimated		
ungon	Value Range	N/A	
, uncon , unconstrained	Default Value	None	
, unconstrained	Required	No	
	Applicable Notes	None	
	Specifies the Whitt	emore and Tu one-parameter mod-	
	els alpha value. A	In alpha value of $= 1$ specifies a	
	model with no don	ninant genetic variance.	
, alpha	Value Range	$(1,\infty)$	
	Default Value	2.634	
	Required	No	
	Applicable Notes	None	
		n to test for the parent-of-origin ef-	
	fect. By default, or	nly sib-pairs are used in the analy-	
	sis.		
parent_of_origin	Value Range	{true, false}	
	Default Value	false	
	Required	No	
	Applicable Notes	None	
	Specifies the option	n to fix either $\lambda_{1m}$ or $\lambda_{1p}$ to 1.	
	Value Range	{maternal, paternal}	
, fixed	Default Value	None	
	Required	No	
	Applicable Notes	2	
		n to include non-sibs in the analysis	
	Value Range	N/A	
, all_pairs	Default Value	None	
	Required	No	
	Applicable Notes	3	

Notes

1. The value **one\_parameter** specifies the Whittemore and Tu one-parameter model. The value **two\_parameter** specifies the two-parameter model. The two-parameter model does not allow

the inclusion of covariate data. If **two\_parameter** is specified, any covariate parameter is ignored and no covariates are included in the analysis.

- 2. A fixed value of **maternal** sets  $\lambda_{1m}$  equal to 1, and a fixed value of **paternal** sets  $\lambda_{1p}$  equal to 1.
- 3. By default, other types of affected relative pairs are excluded from the analysis because the problem of computing the right IBD sharing probabilities for other types of affected relative pairs is daunting. If the all\_pairs attribute is specified, then other types of affected relative pairs are included in an analysis with parent-of-origin effect test for affected sib pairs. When other types of affected relative pairs are included,  $\lambda_1$  will be replaced with  $\{(\lambda_{1m} + \lambda_{1p})/2\}$  in other affected relative pairs to avoid fitting an extra parameter.

## 11.3.1.5 The x\_linkage Sub-Block

The following syntax table specifies the permissible parameter and attribute settings for a

x_	_1	i:	nka	age	sut	o-bl	lock.
----	----	----	-----	-----	-----	------	-------

parameter	Explanation		
[, attribute]	Explanation		
	Specifies the sex-specific pair types to be included in		
	the analysis. Multiple pair_type statements can be		
	specified to include more than one pair type.		
pair_type	Value Range {M-M, M-F, F-F, all}		
	Default Valueall		
	Required No		
	Applicable Notes None		
	Specifies $\lambda_{1mm} = \lambda_{1mf} = \lambda_{1ff}$ regardless of sexes of		
	the pair in the current test. When set to false, all three		
	$\lambda_1$ s are estimated separately.		
lambda1_equal	Value Range {true, false}		
	Default Value true		
	Required No		
	Applicable Notes None		
	Specifies $\lambda_{2ff}$ to be dependent on $\lambda_{1ff}$ . When set to		
	<b>false</b> , $\lambda_{1ff}$ and $\lambda_{2ff}$ are estimated separately.		
lambda2_fixed	Value Range {true, false}		
Tambdaz_TIxed	Default Value true		
	Required No		
	Applicable Notes None		
	The alpha value to compute $\lambda_{2ff}$ when it is dependent		
	on $\lambda_{1ff}$ . This is ignored when lambda2_fixed is		
	set to <b>false</b> .		
, alpha	Value Range $(-\infty, \infty)$		
	Default Value 2.634		
	Required No		
	Applicable Notes None		

The following are all valid LODPAL statements:

```
# Test all markers
lodpal, multipoint
lodpal, multipoint
                        # Equivalent to the previous statement.
{
}
lodpal, singlepoint {
   marker
          = M1
   covariate = ageexam, minimum, diff
}
lodpal, multipoint {
   trait = T1
   autosomal_model {
      model = two_parameter
   }
   diagnostic = "20 44.0" # The additional output file for location
                          # "20 44.0" is generated.
}
lodpal, multipoint,out="tlcondisc.out" {
   trait = T1, condisc
   turn_off_default
                           # Turn off default maximization process
}
                                # The covariate is ignored.
lodpal, multipoint {
                                # Analysis is done with the one-
   trait = T1, noconunaff
   covariate = ageexam
                                # parameter model using default alpha.
}
lodpal,multipoint {
   trait = T1
   pair_info_file = "cov.in" {
                  = probability
      pair_weight
      pair_covariate = covariate1, mean
   }
}
lodpal,multipoint {
   trait = T1
   x_linkage {
                    = "M-M"
      pair_type
      lambdal_equal = false # All three lambdas are estimated separately.
      lambda2_fixed = false # The data set has to have at least 15
                             # sister-sister pairs and at least 15
                             # female-female pairs other then sister
                             # -sister pairs to use this model
   }
}
lodpal,multipoint {
   trait = T1
   x_linkage {
```

```
lambda2_fixed = true, alpha = 3.5 # The same as default model,
                                           # but the different alpha value
                                           # is used.
   }
}
lodpal, multipoint {
   trait = T1
   autosomal {
      model
                        = one_parameter, alpha=2.634
      parent_of_origin = true, fixed=maternal
   }
}
lodpal,multipoint {
   trait=T1
   autosomal {
      model
                        = one_parameter, alpha=2.634
      parent_of_origin =true, all_pairs
   3
}
lodpal,multipoint {
   trait = T1
   autosomal {
      model = one_parameter, uncon
   }
}
lodpal,multipoint {
   trait = T1
   autosomal {
      model = two_parameter, unconstrained
}
```

### 11.3.2 Pair Information File

The pair information file is a character delimited file that stores the pre-constructed pair-specific covariate and/or weight values for the pairs to be used in the analysis. The first line of the file is the header that contains the name of each field, and the rest of the file contains the line for each pair with the required IDs, covariate(s) and/or weight fields. The pedigree ID (PEDID in the example below), first individual ID (ID1 in the example), and second individual ID (ID2 in example) fields are required in that order, and the weight and covariate fields can be in any order. Each individual is expected to be found in the pedigree data file, and the pairs are expected to be found in the IBD sharing file, for the analysis to proceed. Any individual or pair that is not in both of these files will be ignored. The weight and covariate values should be numerics, and no missing values are allowed.

A pair information file may look like the following:

 PEDID ID1 ID2 probability covariate1

 1
 3
 4
 0.0033619
 0.0033619

 102
 3
 6
 0.0114638
 0.000000

102	6	7	0.0022620	0.3283151
102	3	7	0.0162358	0.0000000
104	5	б	0.9802018	0.0000000
105	б	7	0.0135131	0.9079691
106	3	4	0.8125513	0.0334500
107	7	8	0.9497964	0.0006405
•				
•				

Another Pair Information File may look like:

```
PEDID, ID1, ID2, probability, covariate1
1, 3, 4, 0.0033619, 0.0033619
102, 3, 6, 0.0114638, 0.0000000
102, 6, 7, 0.0022620, 0.3283151
102, 3, 7, 0.0162358, 0.0000000
104, 5, 6, 0.9802018, 0.0000000
105, 6, 7, 0.0135131, 0.9079691
106, 3, 4, 0.8125513, 0.0334500
107, 7, 8, 0.9497964, 0.0006405
.
.
```

## 11.4 Program Execution

LODPAL is run via a command line interface on the supported UNIX and Windows platforms. This requires the S.A.G.E. programs to be properly installed and in the current execution path. Input files are specified on the command line and all output files are created in the current working directory.

Running LODPAL from the command prompt with no arguments, or the wrong number of arguments, will result in the program printing its usage statement. This lists the input files the program requires on the command line.

```
>LODPAL
S.A.G.E. v5.x -- LODPAL
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
usage: ./LODPAL <parameters> <pedigree> <IBD>
Command line parameters:
   parameters - parameter file
   pedigree - pedigree data file
   IBD - IBD sharing file
```

As indicated in the program usage statement, input files are listed on the command line. A typical run of LODPAL may look like the following:

## 11.5 Program Output

Filename	File Type	Description	
lodpal.inf	Information output file	Contains informational diagnostic messages,	
		warnings and program errors. No analysis results	
		are stored in this file.	
lodpal.out lodpal.xln	Pair analysis output file	Contains tables of LOD scores and parameter es-	
		timates. For autosomal markers the extension is	
		'out', for X-linked markers it is 'xln'.	
lodpal.lod	Diagnostic output file	Contains tables of individual LOD score contribu-	
		tions and other variables at a particular location.	

LODPAL produces several output files that contain results and diagnostic information:

## **11.5.1 Information Output File**

The LODPAL information output file contains a variety of useful information, including:

- Information on fields read from the pedigree data file. These tables, which provide information about what the program has read in, are included with all programs in the S.A.G.E. Release and are very useful for debugging most common errors caused when reading the pedigree data. When first analyzing new data, it is recommended that these tables be checked carefully to make sure pedigree data are being correctly read.
- Information, warning and error messages generated throughout the program. It is recommended that you check this file for warning and error messages before examining the results of any run of the program. The program attempts to correct many common errors and this sometimes means analyses are not run as expected. The file "lodpal.inf" should be checked for errors and diagnostic information after each run of the program.

## 11.5.2 Pair Analysis Output File

One pair analysis output file, named either "LODPAL.out" or "LODPAL.xln", is generated per run of LODPAL. It contains tables of LOD scores and parameters estimates for each marker location tested.

## 11.5.3 Diagnostic Output File

One diagnostic output file, named "LODPAL.lod" by default, is generated per run of LODPAL when a valid diagnostic location (i.e., marker) has been specified by the user. It contains a table of individual LOD score contributions, covariates, and allele-sharing probabilities at the specified location, along with a variance-covariance matrix of the parameter estimates (assuming independence of all pairs) and a histogram of the individual LOD score contributions.

## **11.6 Example Output Files**

### **11.6.1** Pair Analysis Output File

Here is a typical example of a LODPAL output table.

```
S.A.G.E. v5.x -- LODPAL
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
Wed Oct 10 10:40:33 2001
Conditional Logistic Analysis of
    Affected Relative Pairs - multipoint
_____
    Trait : affection
                concordantly affected relative pairs
    Covariate: cov1
                sum of two individual covariate values
                mean centered
                mean before adjusting = 0.232673
                mean after adjusting = 0.000000
                std. deviation
                                  = 0.423585
    Method : default analysis method
    Model : one-parameter model, constrained, alpha = 2.634
_____
                           Full
                                          Parameter Estimates
                          Sib All
                                          _____
                 LOD
         cM SCORE Pairs Pairs Betal
MARKER
                                                    cov1
20s103
          ----- 0.080913 117 202 0.053795 -0.053795
            2.0 0.064902 117 202 0.051271 -0.051271
 20_2.0

      2.0
      0.064902
      117
      202
      0.051271
      0.051271

      4.0
      0.048328
      117
      202
      0.045973
      -0.045973

      6.0
      0.033579
      117
      202
      0.038439
      -0.038439

      8.0
      0.022750
      117
      202
      0.030557
      -0.030557

      -----
      0.019024
      117
      202
      0.027139
      -0.027139

      -----
      0.029545
      -0.029545
      -0.029545
      -0.029545

 20_4.0
 20_6.0
 20_8.0
 20s482
 20_10.0 10.0 0.020959 117 202 0.029545 -0.029545
 20_12.0 12.0 0.025838 117 202 0.035060 -0.035060
 20_14.0 14.0 0.031819 117 202 0.041004 -0.041004
 20_16.0 16.0 0.037990 117 202 0.046230 -0.046230
 20_18.0 18.0 0.042800 117 202 0.049278 -0.049278
 20 20.0
           20.0 0.044706 117 202 0.049128 -0.049128
 20_22.0
           22.0 0.043029 117 202 0.045823 -0.045823
 20_24.0 24.0 0.038242 117 202 0.040305 -0.040305
          ----- 0.035044 117 202 0.037096 -0.037096
 20s851
 20_26.026.00.0617691172020.052506-0.05250620_28.028.00.1302341172020.081308-0.081308
_____
```

### **11.6.2** Diagnostic Output File

Here is a typical example of a LODPAL diagnostic output table.

```
S.A.G.E. v5.x -- LODPAL
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
Wed Aug 29 13:30:52 2001
Conditional Logistic Analysis of Affected Relative Pairs - multipoint
_____
   Trait : affection
             concordantly affected relative pairs
   Covariate: cov1
             sum of two individual covariate values
             mean centered
             mean before adjusting = 0.232673
             mean after adjusting = 0.000000
             std. deviation = 0.423585
   Method
          : default analysis method
   Model : one-parameter model, constrained, alpha = 2.634
   Location : 20_36.0
_____
# Final Result Summary
   Parameter Estimates:
     1. beta 1 = 0.121141
     2. cov1(delta 1) = 0.132373
   Variance-Covariance Matrix(assuming independent pairs):
     1 0.017046 0.003377
     _____
        2 | 0.003377 | 0.001245 |
     ------
# Histogram of Individual LOD Score Contributions
   Maximum LOD Score = 0.5936
   Minimum LOD Score = -0.4504
   Bin Size= 0.1044IntervalCount (one * is equal up to 2 pair(s).)
_____
                                                     _____
                           ------
0.0716 to 0.1760 35 **************
                    13 ******
 0.1760 to 0.2805
0.2805 to 0.3849
                    0
                      4 **
 0.2805 to
 0.3849 to 0.4893
 0.4893 to 0.5937 1 *
_____
                          _____
            Total : 202
# Individual LOD Score Contribution
                                       Cov. prin2
FAMID IDSIB1 IDSIB2 F0
                              F2
                                                     LOD SCORE CONTRIBUTION

        3
        4
        0.0033619
        0.0033619
        3.5349345
        -0.0715001694

        3
        6
        0.0114638
        0.000000
        0.9441527
        0.0490763128

        6
        7
        0.0022620
        0.3283151
        1.1912290
        0.0415850027

        3
        7
        0.0162358
        0.000000
        1.7190322
        0.0670693326

1
102
102
102
. . .

        4
        5
        0.0051893
        0.0271982
        -4.2598469

        3
        4
        0.0028969
        0.1407563
        -1.0821854

                                                    0.0957374556
0.0003981387
108
109
      Total Pair Count = 202
                                         Total LOD Score = 3.3944467415
_____
```

## Chapter 12

# LODLINK

LODLINK performs model-based lod score calculations for two-point linkage between a main trait and each of the other markers in the pedigree file. The main trait may be a marker or a trait that follows Hardy Weinberg proportions, Mendelian transmission and has either two or three types. In the latter case, output from SEGREG can be used as input. LODLINK uses the genotype/phase elimination algorithms proposed by Lange and Boehnke (1983) and Lange and Goradia (1987), together with other enhancements, to perform fast linkage calculations

## **12.1** Limitations

Pedigrees may not contain loops or marriage rings.

## 12.2 Theory

### 12.2.1 Computation of the Likelihood and Lod Scores

Let  $T_1, ..., T_k$  be the alleles at the trait locus,  $q_{T_1}, ..., q_{T_k}$  be the corresponding frequencies,  $M_1, ..., M_m$  be the alleles at the marker locus, and  $q_M, ..., q_{M_m}$  be the corresponding allele frequencies.

Define

- 1. a phased joint genotype to be:  $\frac{T_b M_c}{T_d M_e}$  where b, d = 1, ..., k; c, e = 1, ..., m;
- 2. the probability of a joint genotype in the population to be:

$$\psi\left(\frac{T_b M_c}{T_d M_e}\right) = C\psi(T_b T_d)\psi(M_c M_e),$$

where

$$C = \begin{cases} 1 & , if T_b = T_d \text{ or } M_c = M_e \\ \\ \frac{1}{2} & , otherwise \end{cases}$$

and

 $\psi(\mathbf{T}_b\mathbf{T}_d)$  = probability of trait genotype in the population,

 $\psi(\mathbf{M}_c\mathbf{M}_e)$  = probability of marker genotype in the population;

3. the transmission probability to be:

$$\begin{aligned} \tau_S \left( \frac{T_b M_c}{T_d M_e} \to T_f M_g \right) &= Pr \left( \begin{array}{c} parent \ of \ sex \ s \ and \ genotype \frac{T_b M_c}{T_d M_e} \\ transmits \ haplotype \ T_f M_g \ to \ child \end{array} \right) \\ &= \frac{(1 - \theta_S)(\delta_{T_b T_f} \delta_{M_c M_g} + \delta_{T_d T_f} \delta_{M_e M_g})}{2} + \frac{\theta_S(\delta_{T_b T_f} \delta_{M_e M_g} + \delta_{T_d T_f} \delta_{M_c M_g})}{2}, \end{aligned}$$

where  $\theta_S$  is the sex-dependent recombination fraction between the trait and marker loci ( $\theta_S = \theta_{male}$  or  $\theta_{female}$ ) and

$$\delta_{xy} = \begin{cases} 1 & , if x = y \\ \\ 0 & , if x \neq y \end{cases}$$

4. the transition probability to be:

$$\Pr\left(\begin{array}{c} Mother \ with \ genotype \ \frac{T_bM_c}{T_dM_e} \ and \ Father \ with \ genotype \ \frac{T_rM_s}{T_uM_v} \\ have \ a \ child \ with \ genotype \ \frac{T_fM_g}{T_hM_j} \end{array}\right)$$

$$= \tau_{female} \left(\frac{T_bM_c}{T_dM_e} \to T_fM_g\right) \tau_{male} \left(\frac{T_rM_s}{T_uM_v} \to T_hM_j\right), \ if \ T_f = T_h \ and \ M_g = M_j$$
or
$$= \tau_{female} \left(\frac{T_bM_c}{T_dM_e} \to T_fM_g\right) \tau_{male} \left(\frac{T_rM_s}{T_uM_v} \to T_hM_j\right) + \tau_{female} \left(\frac{T_bM_c}{T_dM_e} \to T_hM_j\right) \tau_{male} \left(\frac{T_rM_s}{T_uM_v} \to T_fM_g\right), \ otherwise.$$

For a joint genotype  $T_bM_c/T_dM_e$  of pedigree member *i*, let the separate one-locus genotypes be denoted  $u_i = T_bT_d$  for the trait and  $v_i = M_cM_e$  for the marker. Let  $y_i$  be the trait phenotype and  $m_i$  be the marker phenotype (discrete). Let  $w_{male_i}$ ,  $w_{female_i}$  and  $w_i$  be the joint genotypes of the father of individual *i*, mother of individual *i*, and individual *i* respectively. The likelihood for a pedigree of *n* persons is

$$L(\theta) = \sum_{w_1} \dots \sum_{w_n} \prod_{i=1}^n H_i,$$

where

$$H_{i} = \begin{cases} p_{i}(w_{female_{i}}, w_{male_{i}}, w_{i}) &, if \ i \ is \ missing \\ p_{i}(w_{female_{i}}, w_{male_{i}}, w_{i})g_{u_{i}}(y_{i})g_{v_{i}}(m_{i}) &, otherwise \end{cases}$$

in which

$$p_i(w_{female_i}, w_{male_i}, w_i) = \begin{cases} Tr(w_{female_i}, w_{male_i}, w_i), & if the parents of i are in the pedigree \\ \psi(w_i), & otherwise, \end{cases}$$

 $g_{v_i}(m_i) =$  probability of marker phenotype  $m_i$  given marker genotype  $v_i$  (assumed to be always 0 or 1).

 $g_{u_i}(y_i) =$  probability (density) of trait phenotype  $y_i$  conditional on genotypes  $u_i$  and possibly other factors. These can be obtained as output from SEGREG by specifying type\_prob = true in the SEGREG output\_options sub-block.

Lod scores are defined as

 $Z(\theta) = Log_{10}L(\theta) - Log_{10}L(0.5).$ 

### 12.2.2 Estimation of Parameters

When estimating the recombination fraction  $\theta$ , maximum likelihood estimates of  $\theta$  are obtained as the values that make the likelihood largest in the parameter space [0, 0.5]. If a larger likelihood exists for  $\theta$  in the parameter space [0, 1], the corresponding estimate(s) are also given.

When estimating both the recombination fraction  $\theta$  and the proportion of linked families,  $\alpha$ , maximum likelihood estimates are obtained over the range of parameter values indicated in the output.

### 12.2.3 Hypothesis Tests

### 12.2.3.1 Maximum Lod Score Test for Linkage

If we are estimating recombination fractions with  $\theta_{male} = \theta_{female}$ , then the asymptotic chi-square statistic calculated is

$$\chi_1^2 = 2[\log_e L(\hat{\theta}) - \log_e L(0.5)]$$

and the corresponding p-value quoted is

$$1 - \Phi\left(\sqrt{\chi_1^2}\right),$$

where  $\Phi$  is the standard cumulative normal distribution. The upper bound of the p-value is calculated as  $\frac{1}{10^{z(\hat{\theta})}}$ . The p-value and upper bound are quoted only if  $0 \le \hat{\theta} < 0.5$ .

If we are calculating the recombination fractions for males and females separately, the chi-square statistic calculated is

$$\chi_2^2 = 2[\log_e L(\hat{\theta}_{male}, \hat{\theta}_{female}) - \log_e L(0.5, 0.5)]$$

The corresponding p-value quoted as corresponding to this lod score is calculated on the assumption that the estimates  $\hat{\theta}_{male}$  and  $\hat{\theta}_{female}$  are independent, i.e. assuming that, under the null hypothesis  $\hat{\theta}_{male} = \hat{\theta}_{female} = 0.5$ ,  $2 \log_e 10 \times$  (maximum lod) is distributed as  $\frac{1}{4} + \frac{1}{2}\chi_1^2 + \frac{1}{4}\chi_2^2$ . The upper bound of the p-value is calculated as

$$\frac{1}{10^{z(\hat{\theta}_{male},\hat{\theta}_{female})}}.$$

The p-value and upper bound are quoted only if  $0 \leq \hat{\theta}_{male}, \hat{\theta}_{female} < 0.5$ .

#### 12.2.3.2 Cleves and Elston's (1997) Likelihood Ratio Test for Linkage

Let  $L(\hat{\theta}_{male}, \hat{\theta}_{female})$  be the likelihood evaluated at the maximum likelihood estimates  $\hat{\theta}_{male}, \hat{\theta}_{female}$ and  $L(\tilde{\theta}_{male}, \tilde{\theta}_{female})$  be the likelihood estimated at the values  $\tilde{\theta}_{male}, \tilde{\theta}_{female}$  that maximize the likelihood under the constraint  $\tilde{\theta}_{male} + \tilde{\theta}_{female} = 1$ . Then the asymptotic chi-square statistic calculated is

$$\chi_1^2 = 2[\log_e L(\hat{\theta}_{male}, \hat{\theta}_{female}) - \log_e L(\hat{\theta}_{male}, \hat{\theta}_{female})]$$

and the corresponding p-value quoted is

$$1 - \Phi\left(\sqrt{\chi_1^2}\right),$$

where  $\Phi$  is the standard cumulative normal distribution. If both  $\hat{\theta}_{male}$  and  $\hat{\theta}_{female}$  are > 0.5, no p-value is calculated.

### 12.2.3.3 Morton's (1956) Likelihood Ratio Test for Homogeneity of the Recombination Fraction

Let  $\sum_{i=1}^{n} \log_e L_i(\hat{\theta}_i)$  be the maximum log likelihood over n groups of pedigrees with  $\hat{\theta}_i$  estimated separately for each group, and let  $\sum_{i=1}^{n} \log_e L_i(\hat{\theta})$  be the maximum log likelihood over the n groups with a common  $\hat{\theta}$  estimated; then the asymptotic chi-square statistic is

$$2[\sum_{i=1}^{n} \log_{e} L_{i}(\hat{\theta}_{i}) - \sum_{i=1}^{n} \log_{e} L_{i}(\hat{\theta})], \quad with n - 1 \text{ degrees of freedom if } \theta_{male} = \theta_{female}$$
$$2(n-1) \text{ degrees of freedom if } \theta_{male} \neq \theta_{female}$$

The "asymptotic p-value" is the p-value based on the statistic following a chi-square distribution.

### 12.2.3.4 Smith's (1963) Test for Homogeneity of the Recombination Fraction

Let  $\theta < 0.5$  be the recombination fraction in a proportion  $\alpha$  of the families, and suppose there is no linkage in the remaining 1 -  $\alpha$  of the families. Define the log likelihood of the i-th family as  $\log_e L_i(\alpha, \theta) = \log_e [\alpha L_i(\theta) + (1 - \alpha)L_i(0.5)]$ . Under the model  $0 \le \alpha \le 1$ , and  $0 \le \theta \le 0.5$ , we test the null hypothesis  $\alpha = 1$ .

Let  $\sum_{i=1}^{n} \log_e L_i(\hat{\alpha}, \hat{\theta})$  be the maximum log likelihood over *n* constituent pedigrees with  $\alpha$  and  $\theta$  estimated, and  $\sum_{i=1}^{n} \log_e L_i(1, \hat{\theta})$  be the maximum log likelihood over n constituent pedigrees with  $\alpha$  = 1 and  $\theta$  estimated.

If  $\hat{\theta}$  is scalar (i.e., we assume  $\theta_{male} = \theta_{female}$ ) then the asymptotic chi-square statistic for heterogeneity versus homogeneity is

$$\chi_1^2 = 2\left[\sum_{i=1}^n \log_e L_i(\hat{\alpha}, \hat{\theta}) - \sum_{i=1}^n \log_e L_i(1, \hat{\theta})\right], and the one sided p-value is 1 - \Phi(\sqrt{\chi_1^2}),$$

where  $\Phi$  is the standard cumulative normal distribution.

If  $\theta_{male} = \theta_{female}$  is not assumed, so that both  $\hat{\theta}_{male}$  and  $\hat{\theta}_{female}$  are estimated, the chi-square statistic is compared to the chi-square distribution with 2 degrees of freedom and the asymptotic p-value is "two-sided".

### 12.2.3.5 Faraway's (1993) Test for Linkage Under Smith's (1963) Heterogeneity Model.

The asymptotic "chi-square" for linkage in the presence of heterogeneity is

$$2[\sum_{i=1}^{n} \log_e L_i(\hat{\alpha}, \hat{\theta}) - \sum_{i=1}^{n} \log_e L_i(0.5)],$$

for which the p-value is obtained on the assumption that this statistic is distributed as the maximum of two independent chi-square variables, each with one degree of freedom.

If  $\theta_{male} = \theta_{female}$  is not assumed, the "chi-square" statistic is assumed to be distributed as the maximum of two independent chi-square variables, each with 2 degrees of freedom, and the asymptotic p-value quoted is "two-sided".

The posterior probability that the i-th family belongs to the linked type, given the observations, is computed as

$$w_i(\hat{\alpha}, \hat{\theta}) = \frac{\hat{\alpha} L_i^*(\hat{\theta})}{\hat{\alpha} L_i^*(\hat{\theta}) + 1 - \hat{\alpha}}$$

where

$$L_i^* = \frac{L_i}{L_i(0.5)}$$

 $w_i(\widehat{\alpha},\widehat{\theta}) > \widehat{\alpha}$  indicates that the i-th family contains evidence for linkage.

### **12.2.4** Conditional Trait Genotype Probabilities

The table in the detail file headed "Individual Genotype Probabilities" gives, for each pedigree member, the probabilities of having genotypes bd conditional on that member's output marker phenotype and assuming maximum likelihood estimates of the recombination fraction (or fractions, sex specific), assuming homogeneity across pedigrees, i.e., expressing  $L(\theta)$  as a function of the two locus genotypes bc/de (bd for the trait and de for the marker), L(bc/de),

$$P_{bd} = \frac{\sum_{all ce} L(bc/de)}{\sum_{all bd} \sum_{all ce} L(bc/de)}$$

where, by default, bd = AA, AB or BB as in SEGREG.

## **12.3** Program Input

File Type	Description	
LODLINK Parameter File	Specifies the parameters and options with which	
	to perform a particular analysis.	
Pedigree Data File	Contains delimited records for each individual in-	
	cluding fields for identifiers, sex, parents, and trait	
	data.	
Marker Locus Description File <sup>a</sup>	<sup><i>i</i></sup> Lists the alleles, allele frequencies and phenotype	
	to genotype mapping for each marker locus.	
Trait Locus Description File <sup>b</sup>	Lists the alleles, allele frequencies and phenotype	
	to genotype mapping for each trait marker.	
Trait Genotype Probability File	Produced by SEGREG and has a ".typ" exten-	
	sion. Lists the trait-marker penetrance functions	
	for each individual.	

<sup>*a*</sup>If an allele frequency for a particular individual is zero, then the likelihood for that individual's pedigree will be zero, and the pedigree will effectively be skipped during analysis.

<sup>b</sup>Both the Trait Locus Description File and the Type File are optional. One, but not both, may be used for LODLINK input.

### 12.3.1 Parameter File Syntax

### 12.3.2 The lodlink Parameter

The following syntax table specifies the permissible parameter and attribute settings for the main LODLINK parameter.

<pre>parameter [, attribute]</pre>	Explanation		
	Starts a LODLINK	arameter block.	
	Value Range	N/A	
lodlink	Default Value	None	
	Required	Yes	
	Applicable Notes	None	
Specifies the root name to be		name to be used for output files.	
	Output file names will be formed by concatenating the		
	root name and an appropriate extension.		
	Value Range	Character string representing a	
, out		valid file name.	
	Default Value	lodlink_analysis	
	Required	No	
	Applicable Notes	None	

## 12.3.2.1 The lodlink Block

The following syntax table specifies the permissible parameter and attribute settings for the lodlink block.

parameter	Evalenation	
[, attribute]	Explanation	
	Specifies a title for the analysis defined within the pa-	
	rameter block.	
title	Value Range Quoted character string.	
	Default Value "LODLINK Analysis"	
	Required No	
	Applicable Notes None	
	Specifies model for linkage calculations.	
	Value Range N/A	
model	Default Value None	
	Required No	
	Applicable Notes None	
	Model name in ".typ" file generated by SEG-	
	REG when type_prob = true in the SEGREG	
	output_options sub-block, or the name of a valid	
	trait marker. Linkage is calculated between each	
	marker and this locus.	
, trait	Character string representing a	
	Value Range valid model name, or the name	
	of a valid trait marker.	
	Default Value None	
	Required No	
	Applicable Notes 1	
	Marker against which all others are tested for linkage.	
	Value Range Character string representing a	
, marker	valid marker name.	
, marker	Default Value None	
	Required No	
	Applicable Notes 1	
	Specifies option to perform linkage tests.	
	Value Range {true, false}	
linkage_tests	Default Value true	
	Required No	
	Applicable Notes 2	
	Specifies option to use sex-specific recombination	
	fractions.	
, sex_specific	Value Range {true, false}	
, BCX_BPECITIC	Default Value false	
	Required No	
	Applicable Notes 2	

	Specifies option to assume linkage homogeneity.
	Value Range {true, false}
, homog	Default Value true
	Required No
	Applicable Notes 2
	Starts a sub-block that specifies tests for linkage ho-
	mogeneity.
howen to sta	Value Range N/A
homog_tests	Default Value None
	Required No
	Applicable Notes 3
	Starts a sub-block that specifies lod score calculations.
	Value Range N/A
lods	Default Value None
	Required No
	Applicable Notes 4
	Specifies option to calculate genotype probabilities.
	Value Range {true, false}
genotypes	Default Value false
	Required No
	Applicable Notes None
	Specifies option to use sex-specific recombination
	fractions.
	Value Range {true, false}
, sex_specific	Default Value false
	Required No
	Applicable Notes None

### Notes:

- 1. A value for either a trait or marker must be specified, but both should not be specified.
- 2. Linkage tests are performed according to the following table, depending on the values assigned to the sex\_specific and homog attributes of the linkage\_tests parameter.

	Sex-specific Recombination F	
Homogeneity Assumed	true	false
true	Lod Score test Cleves-Elston test	Lod Score test
false	Faraway's test	Faraway's test

- 3. The default is to perform no linkage homogeneity tests. Otherwise a homog\_tests subblock must be included.
- 4. The default is to calculate lod scores for the following non-sex-specific recombination fractions: 0, .01, .05, 0.1, 0.2, 0.3 and 0.4. Otherwise a lods sub-block must be included.

## 12.3.2.2 The homog\_tests Sub-Block

The following syntax table specifies the permissible parameter and attribute settings for the homog\_tests sub-block.

parameter	Explanation		
[, attribute]	Explanation		
	Specifies option to	p perform Smith's test for linkage	
	homogeneity.		
smiths_test	Value Range	{true, false}	
Sur CHS_CCSC	Default Value	false	
	Required	No	
	Applicable Notes	None	
	Use sex-specific re	combination fractions.	
	Value Range	{true, false}	
, sex_specific	Default Value	false	
	Required	No	
	Applicable Notes	None	
	Starts a sub-block t	that specifies Morton's test for link-	
	age homogeneity.		
mortons test	Value Range	{true, false}	
	Default Value	false	
	Required	No	
	Applicable Notes	None	
	Use sex-specific recombination fractions.		
	Value Range	{true, false}	
, sex_specific	Default Value	false	
	Required	No	
	Applicable Notes	None	

### 12.3.2.3 The mortons\_test Sub-Block

The following syntax table specifies the permissible parameter and attribute settings for the mortons\_test sub-block.

<pre>parameter [, attribute]</pre>	Explanation	
	This sub-block specifies groups of pedigree IDs to be	
	used for Morton's test. The value of the group param-	
	eter is the name of the group. This parameter may be	
	specified as many times as necessary.	
group	Value Range	Character string that uniquely
		names a pedigree group.
	Default Value	None
	Required	No
	Applicable Notes	1, 2

Notes

- 1. If no groups are specified, each pedigree is its own group.
- 2. Each pedigree must be listed in one, and only one group in the group sub-block described below.

### 12.3.2.4 The group Sub-Block

The following syntax table specifies the permissible parameter and attribute settings for the group sub-block.

<pre>parameter [, attribute]</pre>	Explanation	
	This sub-block specifies groups of pedigree IDs to be	
	used for Morton's test. This parameter may be speci-	
	fied as many times as necessary to describe the group.	
pedigree_id	Value Range	Character string representing a
pedigree_id		valid pedigree ID.
	Default Value	None
	Required	No
	Applicable Notes	1, 2

Notes:

- 1. Required if group parameter is specified.
- 2. Example:

```
lodlink {
   model, trait = T1
   linkage_tests = false
   homog_tests {
      smiths_test = false #explicitly set to the default value
      mortons_test = true, sex_specific = false {
         group = 1 {
            pedigree_id = 1
            pedigree_id = 2
            pedigree_id = 3
            pedigree_id = 4
            pedigree_id = 5
         }
         group = 2 {
            pedigree_id = 6
            pedigree_id = 7
            pedigree_id = 8
         }
      }
   }
   lods {
      option = none
   }
}
```

#### 12.3.2.5 The lods Sub-Block

The following syntax table specifies the permissible parameter and attribute settings for the lods sub-block.

<pre>parameter [, attribute]</pre>	Explanation		
	Specifies calculation option.		
	Value Range {none, standard, specified}		
option	Default Value standard		
	Required No		
	Applicable Notes 1		
	Specifies option to use sex-specific recombination		
	fractions.		
sex_specific	Value Range {true, false}		
Sex_specific	Default Value false		
	Required No		
	Applicable Notes None		
	Starts a sub-block for specifying sex-specific recom-		
	bination fractions at which lods will be calculated.		
male female	Value Range N/A		
	Default Value None		
	Required No		
	Applicable Notes 2		
	Starts a sub-block for specifying sex-averaged recom-		
	bination fractions at which lods will be calculated.		
average	Value Range N/A		
average	Default Value None		
	Required No		
	Applicable Notes 3		

Notes

- 1. If **none** is specified, no lod scores will be calculated. If **standard** is specified, lod scores will be calculated for the following non-sex-specific recombination fractions: 0, .01, .05, 0.1, 0.2, 0.3 and 0.4. If **specified** is specified, the desired recombination fractions must be specified using male\_female or average sub-blocks for sex-specific or sex-averaged recombination fractions, respectively.
- 2. Required if the option parameter is set to **specified** and the sex\_specific parameter is set to **true**.
- 3. Required if the option parameter is set to **specified** and the sex\_specific parameter is set to **false**.

#### 12.3.2.6 The male\_female Sub-Block

The following syntax table specifies the permissible parameter and attribute settings for the male\_female sub-block.

<pre>parameter [, attribute]</pre>	Explanation		
	Specifies sex-specific recombination fractions for		
	which a lod score is to be calculated. This parame-		
	ter may be repeated	d as often as desired.	
theta	Value Range	N/A	
	Default Value	None	
	Required	No	
	Applicable Notes	None	
	Specifies the male	recombination fraction.	
	Value Range	[0, 1]	
, male	Default Value	None	
	Required	No	
	Applicable Notes	1	
	Specifies the femal	le recombination fraction.	
	Value Range	[0, 1]	
, female	Default Value	None	
	Required	No	
	Applicable Notes	1	

Notes:

1. Required if the theta parameter is specified.

#### 12.3.2.7 The average Sub-Block

The following syntax table specifies the permissible parameter and attribute settings for the average sub-block.

<pre>parameter [, attribute]</pre>	Explanation		
	Specifies a sex-averaged recombination fraction for		
	which a lod score is to be calculated. This parame-		
	ter may be repeated as often as desired.		
theta	Value Range	[0, 1]	
	Default Value	None	
	Required	No	
	Applicable Notes	1	

Notes:

1. Required if the average parameter is specified.

#### Example 1

Do Morton's test for linkage homogeneity between model T1 (produced by SEGREG) and each marker in the pedigree file, estimating non-sex-specific recombination fractions. For these tests the group designated "1" consists of pedigrees 1-5 and group "2" consists of pedigrees 6-7.

```
lodlink {
   model, trait = T1
   linkage_tests = false
   homog_tests {
      smiths_test = false #explicitly set to the default value
      mortons_test = true, sex_specific = false {
         group = 1 {
            pedigree_id = 1
            pedigree_id = 2
            pedigree_id = 3
            pedigree_id = 4
            pedigree_id = 5
         }
         group = 2 {
            pedigree_id = 6
            pedigree_id = 7
            pedigree_id = 8
         }
      }
   }
   lods {
      option = none
   }
}
```

#### Example 2

Test for linkage between marker "Mfd154" and each of the other markers in the pedigree file estimating sex-specific recombination fractions assuming linkage homogeneity. Also calculate lod scores for the following pairs of recombination fractions: male .4, female 0; male .4, female .1; male .3, female .2.

Use the title "linkage test" in the output files. Name the summary and detail output files "example2.sum" and "example2.det", respectively.

```
lodlink, out = "example2" {
   title = "linkage test"
   model, marker = Mfd154
   linkage_tests = true, sex_specific = true, homog = true
   homog_tests {
      smiths_test = false #explicitly set to the default value
      mortons_test = false #explicitly set to the default value
   }
   lods {
      option
                   = specified
      sex_specific = true
      male female {
         theta, male = .4, female = 0
         theta, male = .4, female = .1
         theta, male = .3, female = .2
      }
   }
}
```

#### **12.4 Program Execution**

LODLINK is run via a command line interface on the supported UNIX and Windows platforms. This requires the S.A.G.E. programs to be properly installed and in the current execution path. Input files are specified on the command line and all output files are created in the current working directory.

Running LODLINK from the command prompt with no arguments, or the wrong number of arguments, will result in the program displaying its usage statement, which lists the input files names required on the command line.

```
S.A.G.E. v4.4 -- LODLINK
COPYRIGHT (C) 2003 CASE WESTERN RESERVE UNIVERSITY
usage: ./lodlink <parameters> <pedigree> <locus> [trait/type]
Command line parameters:
  parameters - Parameter File
  pedigree - Pedigree Data File
  locus - Marker Locus Description File
  trait/type - Trait Locus Description File or SEGREG Generated Type File (optional)
```

A typical run of LODLINK may look like the following:

```
>lodlink par ped mld
S.A.G.E. v5.x -- LODLINK
COPYRIGHT (C) 2002 CASE WESTERN RESERVE UNIVERSITY
Reading Parameter File.....done.
Reading Marker Locus Description File....done.
Reading Pedigree File.....
from ped.....done.
Sorting pedigrees.....done.
Generating statistics.....
Parsing LODLINK analyses ...
Parsing new analysis block ...
Parsing of LODLINK Analysis 1 complete. Analysis valid.
_____
Performing lod score calculations ...
Performing lod ratio test for linkage ...
Parsing new analysis block ...
Parsing of LODLINK Analysis 2 complete. Analysis valid.
_____
Performing test for linkage under Smith's model ...
Performing Smith's homogeneity test ...
done.
Analysis complete!
```

## 12.5 Program Output

File Name	File Type	Description
lodlink.inf	Diagnostic information output	Contains informational diagnostic messages,
	file	warnings and program errors. No calculation re-
		sults are stored in this file.
genome.inf	Genome information output file	Contains marker allele and genotype frequencies.
lodlink_analysis1.sum	Summary output	Contains lod scores and results of linkage and
		linkage homogeneity tests. Results in this file are
		based on calculations done on the pedigree data
		file as a whole.
lodlink_analysis1.det	Detailed output	Contains lod scores by family, the posterior prob-
		ability that each family belongs to the linked pedi-
		gree, recombination fraction estimates by group
		for Morton's test, and individual trait genotype
		probabilities, as appropriate

LODLINK produces four types of output files that contain results and diagnostic information:

#### 12.5.1 Information Output File

The Information Output File contains a variety of useful information, including:

• Information on fields read from the Pedigree Data File. These tables, which provide information about what the program has read from the Pedigree Data File, are included with all programs in S.A.G.E. and are very useful for debugging most common errors caused when reading the pedigree data. When first analyzing new data, it is recommended that these tables be checked carefully to make sure pedigree data are being correctly read.

• Information, warning and error messages generated throughout the program. It is recommended that this file be checked for warning and error messages before examining the results of any run of the program. The program attempts to correct many common errors and this sometimes means analyses are not as expected. The file "lodlink.inf" should be checked for errors and diagnostic information after each run of the program.

#### 12.5.2 Genome Information Output File

Contains marker allele and genotype frequencies.

#### 12.5.3 Summary Output File

Contains results pertaining to whole data set. See section 12.2 for details regarding interpretation of these results.

#### 12.5.4 Detailed Output File

Contains results on a per individual, per family or per group basis. See section 12.2 for details regarding interpretation of these results.

## **12.6 Example Output Files**

# 12.6.1 Summary Output File

LODLINK Analysi	s 1 SUMMARY FI	LE				
Options Selecte		===				
	=					
Main locus type			trait			
Main locus name	2		T1			
Lod scores			yes			
Linkage tests			yes			
Sex-specific	recombination	fractions	no			
Assume homoge	-		yes			
Smith's test fo			no			
-	recombination		no			
Morton's test f			no			
-	recombination	fractions	no			
Genotype probab			no			
	recombination		no			
Recombination F						
			0 0500	0 1000	0 0000	0 2000
Sex averaged	0.0000	0.0100	0.0500	0.1000	0.2000	0.3000
Results						
	Lod Scores Non-Sex-Sp	ecific Recor	nbination Fra	ctions		
Locus	0.0000	0.0100	0.0500	0.1000	0.2000	0.3000
 M1			-24.2			4 31
M2			6.76			
112	MLE	51		nd Linkage T		20
				m. Fract. in		
			002119 11000		[0] 10]	
Locus	Recom	Recom	Lod	Chi	P-Value	P-Value
	Fract in	Fract in	Score	Square		Upper
	[0, .5]			Stat		Bound
Ml	0.2737		4.42	20.4	3.196e-06	3.779e-05

### 12.6.2 Detailed Output File

Results						
		s By Family pecific Recom	bination Fr	actions		
Locus		Pedigree 1 C 0.0100	0.0500			0.3000
 м1	0					
M2	0			0	0	(
Locus		Pedigree 2 C 0.0100	0.0500	0.1000	0.2000	0.3000
 м1			0.259	0.215	0 124	0.064
M2	0.501	0.292 0	0.250	0.215	0.151	0.001
Locus	0.0000	0.0100	0.0500	0.1000	0.2000	
M1 M2	-INFINITY	-1.11 -1.11	-0.442	-0.188	0.0103 0.0103	
	Family in	Pedigree 4 C	ontaining M	ember 1		
Locus	0.0000	0.0100		0.1000	0.2000	0.3000
M1	0	0	0	0	0	
м2	1.2 Family in	1.19 Pedigree 5 C	1.12 Containing M	1.02 ember 1	0.816	0.58
Locus	0.0000	0.0100	0.0500	0.1000	0.2000	0.3000
 м1	INFINITY	0.809	-0.164	0.0668		
M2	-INFINITY	-0.809	-0.164	0.0668		
Locus		Pedigree 6 C 0.0100			0.2000	0.3000
M1 M2	0		0 0.986		0 0.369	0.098
•	1.2	1.10	0.980	0.77	0.309	0.090
•	Family in	Pedigree 195	Containing	Member 1		
Locus	0.0000	0 0100				
		0.0100	0.0500	0.1000	0.2000	0.3000
MI						
	0.602 -INFINITY	0.593 -1.4	0.558 -0.721	0.511 -0.444		0.29
M2	0.602 -INFINITY Family in	0.593 -1.4 Pedigree 196	0.558 -0.721 Containing	0.511 -0.444 Member 1	0.408 -0.194	0.29 -0.075
M2	0.602 -INFINITY Family in 0.0000	0.593 -1.4 Pedigree 196 0.0100	0.558 -0.721 Containing 0.0500	0.511 -0.444 Member 1 0.1000	0.408 -0.194 0.2000	0.29 -0.075 0.3000
M2 Locus 	0.602 -INFINITY Family in 0.0000	0.593 -1.4 Pedigree 196 0.0100	0.558 -0.721 Containing 0.0500	0.511 -0.444 Member 1 0.1000	0.408 -0.194 0.2000	0.29 -0.075 0.3000 
M2 Locus 	0.602 -INFINITY Family in 0.0000 -INFINITY -INFINITY	0.593 -1.4 Pedigree 196 0.0100	0.558 -0.721 Containing 0.0500  -0.721 -1.2	0.511 -0.444 Member 1 0.1000  -0.444 -0.708	0.408 -0.194 0.2000	0.29 -0.075 0.3000 
M2 Locus  M1 M2	0.602 -INFINITY Family in 0.0000 -INFINITY -INFINITY Family in 0.0000	0.593 -1.4 Pedigree 196 0.0100  -1.4 -2.52 Pedigree 197 0.0100	0.558 -0.721 Containing 0.0500 	0.511 -0.444 Member 1 0.1000  -0.444 -0.708 Member 1 0.1000	0.408 -0.194 0.2000  -0.194 -0.303 0.2000	0.29 -0.075 0.3000  -0.075 -0.12 0.3000
M2 Locus M1 M2 Locus	0.602 -INFINITY Family in 0.0000 -INFINITY -INFINITY Family in 0.0000	0.593 -1.4 Pedigree 196 0.0100  -1.4 -2.52 Pedigree 197 0.0100	0.558 -0.721 Containing 0.0500 	0.511 -0.444 Member 1 0.1000  -0.444 -0.708 Member 1 0.1000	0.408 -0.194 0.2000 	0.29 -0.075 0.3000  -0.12 0.3000  0.300
M2 Locus M1 M2 Locus M1	0.602 -INFINITY Family in 0.0000 	0.593 -1.4 Pedigree 196 0.0100 -1.4 -2.52 Pedigree 197 0.0100  0.93 -0.168	0.558 -0.721 Containing 0.0500  -0.721 -1.2 Containing 0.0500 	0.511 -0.444 Member 1 0.1000 -0.444 -0.708 Member 1 0.1000 	0.408 -0.194 0.2000 -0.194 -0.303 0.2000	0.29 -0.075 0.3000  -0.12 0.3000  0.300
M2 Locus M1 M2 Locus M1 M2	0.602 -INFINITY Family in 0.0000 	0.593 -1.4 Pedigree 196 0.0100  -1.4 -2.52 Pedigree 197 0.0100  0.93 -0.168 Pedigree 198	0.558 -0.721 Containing 0.0500  -0.721 -1.2 Containing 0.0500  0.851 -0.137 Containing		0.408 -0.194 0.2000 	0.29 -0.075 0.3000 -0.075 -0.12 0.3000 
M2 Locus M1 M2 Locus M1 M2	0.602 -INFINITY Family in 0.0000 -INFINITY -INFINITY Family in 0.000 	0.593 -1.4 Pedigree 196 0.0100  Pedigree 197 0.0100  0.93 -0.168 Pedigree 198 0.0100 	0.558 -0.721 Containing 0.0500 	0.511 -0.444 Member 1 0.1000  0.444 -0.708 Member 1 0.1000 	0.408 -0.194 0.2000 -0.194 -0.303 0.2000 	0.29 -0.075 0.3000 -0.075 -0.12 0.3000 -0.023 0.3000 -3000
M2 Locus M1 M2 Locus M1 M2 Locus M1	0.602 -INFINITY Family in 0.0000 -INFINITY -INFINITY Family in 0.000 	0.593 -1.4 Pedigree 196 0.0100  Pedigree 197 0.0100  0.93 -0.168 Pedigree 198 0.0100 	0.558 -0.721 Containing 0.0500 	0.511 -0.444 Member 1 0.1000  0.444 -0.708 Member 1 0.1000 	0.408 -0.194 0.2000 -0.194 -0.303 0.2000 	0.29 -0.075 0.3000 -0.075 -0.12 0.3000 -0.023 0.3000 -3000
M2 Locus M1 M2 Locus M1 M2 Locus M1	0.602 -INFINITY Family in 0.0000 		0.558 -0.721 Containing 0.0500  -0.721 -1.2 Containing 0.0500  0.851 -0.137 Containing 0.0500  0 -1.44 Containining	0.511 -0.444 Member 1 0.1000 -0.444 -0.708 Member 1 0.1000 -0.75 -0.104 Member 1 0.1000 -0.887 Member 1	0.408 -0.194 0.2000 -0.194 -0.303 0.2000 -0.534 -0.0555 0.2000 -0.388	0.29 -0.075 0.3000 -0.075 -0.12 0.3000 -0.023 0.3000 -0.15
M2 Locus M1 M2 Locus M1 M2 Locus M1 M2 Locus M1 M2	0.602 -INFINITY Family in 0.0000 	 Pedigree 196 0.0100  -1.4 -2.52 Pedigree 197 0.0100  0.93 -0.168 Pedigree 198 0.0100  0.2.8	0.558 -0.721 Containing 0.0500  -0.721 -1.2 Containing 0.0500  0.851 -0.137 Containing 0.0500  0 -1.44 Containining	0.511 -0.444 Member 1 0.1000 -0.444 -0.708 Member 1 0.1000 -0.75 -0.104 Member 1 0.1000 -0.887 Member 1	0.408 -0.194 0.2000 -0.194 -0.303 0.2000 	0.29 -0.075 0.3000 -0.075 -0.12 0.3000 -0.023 0.3000 -0.15
M2 Locus M1 M2 Locus M1 M2 Locus M1 M2 Locus Locus	0.602 -INFINITY Family in 0.0000 		0.558 -0.721 Containing 0.0500  -0.721 -1.2 Containing 0.0500  0.851 -0.137 Containing 0.0500  0 -1.44 Containing 0.0500	0.511 -0.444 Member 1 0.1000 -0.444 -0.708 Member 1 0.1000 -0.75 -0.104 Member 1 0.1000 -0.887 Member 1 0.1000 -0.245	0.408 -0.194 0.2000 -0.194 -0.303 0.2000 -0.534 -0.0555 0.2000 -0.388 0.2000 -0.388	0.29 -0.075 0.3000 -0.075 -0.12 0.3000 -0.023 0.3000 -0.15 0.3000
M2 Locus M1 M2 Locus M1 M2 Locus M1 M2 Locus M1 M2 Locus M1	0.602 -INFINITY Family in 0.0000 -INFINITY -INFINITY Family in 0.0000 -0.176 Family in 0.0000 -INFINITY Family in 0.0000 -0.1NFINITY Family in 0.0000			0.511 -0.444 Member 1 0.1000 -0.444 -0.708 Member 1 0.1000 -0.75 -0.104 Member 1 0.1000 -0.887 Member 1 0.1000 -0.887 Member 1 0.1000	0.408 -0.194 0.2000 -0.194 -0.303 0.2000 -0.534 -0.0555 0.2000 -0.388 0.2000 -0.388	0.29 -0.075 0.3000 -0.075 -0.12 0.3000 -0.023 0.3000 -0.15 0.3000 -0.015
M2 Locus M1 M2 Locus M1 M2 Locus M1 M2 Locus M1 M2 Locus M1 M2	0.602 -INFINITY Family in 0.0000 -INFINITY -INFINITY Family in 0.0000 -0.176 Family in 0.0000 -0.176 Family in 0.0000 -0.176 Family in 0.0000 -0.175INITY Family in 0.0000 -0.347 -INFINITY Family in Family in			0.511 -0.444 Member 1 0.1000  -0.444 -0.708 Member 1 0.1000 	0.408 -0.194 0.2000 -0.194 -0.303 0.2000 -0.534 -0.0555 0.2000 -0.388 0.2000 -0.388	0.29 -0.075 0.3000 -0.075 -0.12 0.3000 -0.023 0.3000 -0.15 0.3000 -0.015
M1 M2 Locus M1 M2 Locus M1 M2 Locus M1 M2 Locus M1 M2 Locus M1 M2 Locus	0.602 -INFINITY Family in 0.0000 -INFINITY -INFINITY Family in 0.0000 -0.176 Family in 0.0000 -INFINITY Family in 0.0000 		0.558 -0.721 Containing 0.0500 -0.721 -1.2 Containing 0.0500 	0.511 -0.444 Member 1 0.1000 -0.444 -0.708 Member 1 0.1000 -0.75 -0.104 Member 1 0.1000 -0.887 Member 1 0.1000 -0.245 -0.396 Member 1 0.1000	0.408 -0.194 0.2000 -0.194 -0.303 0.2000 -0.534 -0.0555 0.2000 -0.388 0.2000 -0.151 -0.167 0.2000	0.29 -0.075 0.3000 -0.075 -0.12 0.3000 -0.023 0.3000 -0.15 0.3000 -0.15 0.3000 -0.072 -0.063 0.3000
M2 Locus M1 M2 Locus M1 M2 Locus M1 M2 Locus M1 M2 Locus M1 M2 Locus M1 M2	0.602 -INFINITY Family in 0.0000 -INFINITY -INFINITY Family in 0.0000 -0.176 Family in 0.0000 -0.176 Family in 0.0000 -0.175INITY Family in 0.0000 -0.347 -INFINITY Family in 0.0000 -0.347 -INFINITY Family in 0.0000 -0.347 -0.362	 Pedigree 196 0.0100      Pedigree 197 0.0100  0.93 -0.168 Pedigree 198 0.0100  0.336 Pedigree 199 0.0100  0.336 Pedigree 200 0.0100 		0.511 -0.444 Member 1 0.1000  0.444 -0.708 Member 1 0.1000 	0.408 -0.194 0.2000 -0.194 -0.303 0.2000 -0.534 -0.0555 0.2000 -0.388 0.2000 -0.151 -0.167 0.2000 -0.318	0.29 -0.075 0.3000 -0.075 -0.12 0.3000 -0.023 0.3000 -0.15 0.3000 -0.072 -0.063 0.3000 -0.072
M2 Locus M1 M2 Locus M1 M2 Locus M1 M2 Locus M1 M2 Locus M1 M2 Locus Locus	0.602 -INFINITY Family in 0.0000 	 Pedigree 196 0.0100  -1.4 -2.52 Pedigree 197 0.0100  0.93 -0.168 Pedigree 198 0.0100  0.336 Pedigree 199 0.0100  0.336 Pedigree 200 0.0100  0.338 Pedigree 200 0.0100		0.511 -0.444 Member 1 0.1000  0.444 -0.708 Member 1 0.1000 	0.408 -0.194 0.2000 -0.194 -0.303 0.2000 -0.534 -0.0555 0.2000 -0.388 0.2000 -0.151 -0.167 0.2000	0.29 -0.075 0.3000 -0.075 -0.12 0.3000 -0.023 0.3000 -0.15 0.3000 -0.072 -0.063 0.3000 -0.072
M2 Locus M1 M2 Locus M1 M2 Locus M1 M2 Locus M1 M2 Locus M1 M2 Locus M1 M2	0.602 -INFINITY Family in 0.0000 -INFINITY -INFINITY Family in 0.0000 -0.176 Family in 0.0000 -0.176 Family in 0.0000 -0.176 Family in 0.0000 -0.347 -INFINITY Family in 0.0000 -0.347 -INFINITY Family in 0.0000 -0.347 -0.362 0.301 Lod Score Variance	0.593 -1.4 Pedigree 196 0.0100  -2.52 Pedigree 197 0.0100  0.93 -0.168 Pedigree 198 0.0100  0.336 Pedigree 199 0.0100  0.336 Pedigree 200 0.0100  0.336 Pedigree 200 0.0100  0.589 0.288 Linkage Test Covariance Ma	 0.558 -0.721 Containing 0.0500  0.721 -1.2 Containing 0.0500  0.851 -0.137 Containing 0.0500  0.44 Containing 0.0500  0.295 -0.662 Containing 0.0500  0.235 0.238 trices	0.511 -0.444 Member 1 0.1000  0.444 -0.708 Member 1 0.1000 	0.408 -0.194 0.2000 -0.194 -0.303 0.2000 -0.534 -0.0555 0.2000 -0.388 0.2000 -0.151 -0.167 0.2000 -0.318	0.29 -0.075 0.3000 -0.075 -0.12 0.3000 -0.023 0.3000 -0.15 0.3000 -0.072 -0.063 0.3000 -0.072 -0.063
M2 Locus M1 M2 Locus M1 M2 Locus M1 M2 Locus M1 M2 Locus M1 M2 Locus	0.602 -INFINITY Family in 0.0000 - INFINITY -INFINITY Family in 0.0000 - 0.949 -0.176 Family in 0.0000 - INFINITY Family in 0.0000 - 0.347 -INFINITY Family in 0.0000 - 0.347 -INFINITY -INFINITY - 0.347 - 0.301 Lod Score Variance- Parameter	0.593 -1.4 Pedigree 196 0.0100  -1.4 -2.52 Pedigree 197 0.0100  0.93 -0.168 Pedigree 199 0.0100  0.336 -1.33 Pedigree 200 0.0100  0.589 0.288 Linkage Test	 0.558 -0.721 Containing 0.0500  0.721 -1.2 Containing 0.0500  0.851 -0.137 Containing 0.0500  0.44 Containing 0.0500  0.295 -0.662 Containing 0.0500  0.295 -0.623 Containing 0.0535 0.238 trices tecomb	0.511 -0.444 Member 1 0.1000  0.444 -0.708 Member 1 0.1000 	0.408 -0.194 0.2000 -0.194 -0.303 0.2000 -0.534 -0.0555 0.2000 -0.388 0.2000 -0.151 -0.167 0.2000 -0.318 0.2000	
M2 Locus M1 M2 Locus M1 M2 Locus M1 M2 Locus M1 M2 Locus M1 M2 Locus M1 M2	0.602 -INFINITY Family in 0.0000 -INFINITY -INFINITY Family in 0.0000 -0.949 -0.176 Family in 0.0000 -INFINITY Family in 0.0000 -INFINITY Family in 0.0000 -INFINITY Family in 0.0000 -0.347 -INFINITY Family in 0.0000 -0.347 -0.301 Lod Score Variance	0.593 -1.4 Pedigree 196 0.0100  -1.4 -2.52 Pedigree 197 0.0100  0.93 -0.168 Pedigree 199 0.0100  0.336 -1.33 Pedigree 200 0.0100  0.336 2.589 0.0100  0.589 0.288 Linkage Test Covariance Ma Order (Avg R in [0, .5]		0.511 -0.444 Member 1 0.1000  0.444 -0.708 Member 1 0.1000  0.75 -0.104 Member 1 0.1000  0.245 -0.396 Member 1 0.1000  0.245 -0.396 Member 1 0.1000	0.408 -0.194 0.2000 -0.194 -0.303 0.2000 -0.534 -0.0555 0.2000 -0.388 0.2000 -0.151 -0.167 0.2000 -0.318 0.0849	

# Chapter 13

# MLOD

MLOD Performs multi-point model-based LOD-score linkage analysis on small constituent pedigrees. Analysis is optimized for examining one-locus trait models and will, in future versions, allow for meiosis specific recombination fractions.

#### 13.1 Limitations

MLOD calculates the likelihood of each possible inheritance pattern (i.e., ancestral origin of each allele) at each marker location for each constituent pedigree, using all marker data and assuming no crossover interference. It is restricted to small pedigrees due to the exponential nature of the algorithm related to the number of individuals in the pedigree. Only discrete traits may be analyzed, but there is no limit on the number of discrete categories allowed (this effectively allows the analysis of continuous traits). The time and space complexity of the algorithm is largely characterized by the exponent 2n - f, the number of bits in an inheritance vector, where n is the number of nonfounders and f is the number of founders in a constituent pedigree. During parameter specification the maximum value of 2n - f may be set, so that any constituent pedigree that has a value larger than this maximum will be skipped.

#### 13.2 Theory

Given trait data, T, marker data, M, for a chromosomal region, and a point of interest in that region, p, MLOD computes a multi-point LOD score, defined as:

$$Z(p) = \log_{10} \left( \frac{P(M \mid T \text{ at } p)}{P(M)P(T)} \right),$$

where P(T) can be a probability mass or density function.

Given a chromosomal region, a trait, and several pedigrees, MLOD calculates multi-point LOD scores for each point of interest along the chromosome by first generating exact multi-point likelihoods at each marker using a modified Lander-Green approach (Idury and Elston, 1996), and then computing the likelihood for the trait of each inheritance pattern (which is proportional to the probability of the trait for each inheritance pattern). These likelihoods are combined to generate the final LOD score at each location specified by the user.

#### 13.2.1 The Exact Multi-point Algorithm

The general algorithm used by MLOD to generate multi-point likelihoods and other related statistics is called the exact multi-point algorithm. This algorithm takes a chromosomal region and generates likelihoods of all the possible inheritance patterns at each marker location in the region. These likelihoods are then combined at each marker location to generate multi-point LOD scores.

Given a pedigree with f founders and n non-founders and a pattern of segregation at a particular locus for this pedigree, we may represent this segregation as a vector of binary (0 or 1) digits of length 2n where each element represents one of the 2n meioses in the pedigree. The value of each binary element is determined by that meiosis having either a grandpaternal or grandmaternal allele from the parent. This "inheritance vector" is the basis for the Lander-Green multi-point algorithm (Lander and Green, 1987).

Because each meiosis is a separate event at a given locus, there are  $2^{2n}$  possible patterns of locus segregation in the pedigree for each marker. However, because founder phase is unknown, it is impossible to determine the true state of the meioses from the founders. This means that, for the founder meioses, we do not know the binary values to be used in the inheritance vectors for a given inheritance pattern. Each inheritance pattern can therefore be represented by  $2^f$  different inheritance vectors that represent the same inheritance pattern and share the same likelihood. These "equivalence classes" of inheritance vectors reduce the number of vectors that we must consider to  $2^{2n-f}$ .

For a given set M of i markers  $m_1 \dots m_i$  (including a trait-marker, i.e. a trait considered in the same manner as any other marker but with more general penetrance functions), we calculate the joint probability of each inheritance vector and the pedigree data at each marker. The set of  $2^{2n-f}$  joint probabilities at a particular marker is called the *likelihood vector* for that marker. The sum of these  $2^{2n-f}$  joint probabilities is proportional to the likelihood for the pedigree data.

#### 13.2.2 Combining Likelihood Vector Elements to Obtain a Multi-point Likelihood

Given two likelihood vectors,  $v_1$  and  $v_2$  at markers  $m_1$  and  $m_2$ , and a recombination fraction  $\theta_1$  between them, we wish to calculate the joint likelihood.

To do this, we form a transition matrix  $T_1$ . This is a  $2^{2n-f} \times 2^{2n-f}$  matrix with elements  $t_{\alpha\beta} = \theta_1^q (1-\theta_1)^{2n-f-q}$  where  $\alpha$ ,  $\beta$  are inheritance vectors of the two markers and q is the Hamming Distance between them (the number of elements of  $\alpha$ ,  $\beta$  that differ). Then,

$$L(v_1, v_2) = v_1' T_1 v_2.$$

To add a third likelihood vector  $v_3$  at marker  $m_3$ , with recombination fraction  $\theta_2$  between  $m_2$  and  $m_3$ , we form a transition matrix  $T_2$  analogous to  $T_1$ . Then

 $L(v_1, v_2, v_3) = v'_1 T_1 V_2 T_2 v_3$ , where  $V_2$  is a  $2^{2n-f} \times 2^{2n-f}$  diagonal matrix containing the elements of  $v_2$ .

In general,

$$L(v_1, v_2, \dots, v_{i-1}, v_i) = v'_1 T_1 V_2 T_2 \dots V_{i-1} T_{i-1} v_i.$$

Idury and Elston (1996) suggested methods of calculating these likelihoods that are efficient, given the underlying structure of the transition matrices. S.A.G.E. extends these methods to include additional optimizations that use the genetic information at the markers to reduce the time complexity of these algorithms. Even so, the algorithm takes time and space that increases exponentially with the size of the pedigree. It is for this reason that these algorithms are restricted to small-to-medium sized pedigrees.

#### 13.2.3 Using Genetic Information to Improve Algorithm Performance

There are  $2^{2n-f}$  inheritance vectors that we must consider at each marker. However, when most individuals are typed, the joint probability of the data and many of these inheritance patterns will be zero, because the inheritance pattern indicated by the vector is not consistent with the observed phenotypes at the marker in question.

A *fixed point* is any meiosis where the transmission is known with certainty. Given a fixed point in our likelihood vector, all inheritance vectors that do not match the transmission of the fixed point have a joint probability of 0. This information is used to speed up the computation. For each fixed point, we can reduce the time required for calculation by a factor of 2. These reductions are cumulative, so that for *n* fixed points, the time is reduced by a factor of  $2^n$ .

#### 13.2.4 Calculating Multi-point Likelihood Vectors

It is often necessary to calculate the multi-point likelihood vector at a specific location p along a chromosome. Assume we have a chromosome containing markers  $m_1, \ldots, m_i$  with distances  $d_1, \ldots, d_{i-1}$  between them. We have two adjacent markers,  $m_j$  and  $m_{j+1}$  between which is a point p for which we wish to calculate a multi-point likelihood vector v, with p some known distances  $d_{j1}$  and  $d_{j2}$  (where  $d_{j1} + d_{j2} = d_j$ ) from  $m_j$  and  $m_{j+1}$ , respectively. Distances are expressed as recombination fractions and may be computed from genetic distance using either the Kosambi or Haldane map function.

First, we calculate  $v_1, \ldots, v_i$ , the single-point likelihood vectors for each marker. Then we calculate the following:

$$P_{j1} = v'_1 T_1 V_2 T_2 \dots V_j \text{ and } P_{j2} = v'_i T_{i-1} V_{i-1} \dots V_{j+1}.$$

 $P_{j1}$  is the multi-point information contributed to point p by all markers before point p, while  $P_{j2}$  is the multi-point information contributed by all markers after p. Each is a  $1 \times 2^{2n-f}$  vector representing the combined multi-point information contributing to v. Calculating v is now trivial:

$$v = P_{j1}T_{j1}T_{j2}P'_{j2}$$

where  $P'_{j2}$  is a diagonal matrix consisting of elements of  $P_{j2}$ .

#### 13.2.5 Computing LOD Scores

For a given point p on a chromosome, we calculate the multi-point LOD score given that a trait locus T (the trait-marker), is at that location by first calculating P(M|T at p), the multi-point likelihood for the chromosome given that T is present at that location and follows the model specified. We then calculate, P(M), the multi-point likelihood for the chromosomal region without T, and P(T), the probability of the trait given the underlying model. Then the LOD score for T being at point p is

$$Z(p) = \log_{10} \left( \frac{P(M \mid T \text{ at } p)}{P(M)P(T)} \right)$$

At each location p we generate a LOD score for each pedigree. The combined LOD score at p is the sum of each constituent pedigree's individual LOD score at p.

#### 13.2.6 Computing Information Content

Information content at a location is determined based on the probabilities of each inheritance pattern within the likelihood vector at that location. If we have n possible inheritance patterns,  $i_1 \dots i_n$ , each with b bits and probability  $p_i$  such that

$$\sum_{i} p_i = 1,$$

then, the Information I is defined by [Kruglyak and Lander, 1995b]

$$I = 1 + \frac{\sum_{i} p_i \frac{\log(p_i)}{\log(2)}}{h}.$$

## 13.3 Program Input

MLOD requires the following input files in order to run:

File Type	Description
Parameter file	Specifies the parameters and options with which
	to perform a particular analysis.
Pedigree data file	Contains delimited records for each individual,
	including fields for identifiers, sex, parents, trait,
	and marker data.
Marker locus Description File	Lists the alleles, allele frequencies and phenotype
	to genotype mapping for each marker locus.
Trait locus description file	Lists the genetic model for each trait being ana-
	lyzed.
Genome description file	Contains a description of the linked marker re-
	gions, including distances between markers.

### 13.3.1 The mlod Parameter

parameter	Explanation		
[, attribute]			
	Starts a MLOD and	alysis block.	
	Value Range	N/A	
mlod	Default Value	None	
	Required	Yes	
	Applicable Notes	None	
	Specifies the "root	" name to be used for output files.	
	Output file names will be formed by concatenating the		
	root name and an appropriate extension.		
out	Value Range	An alphanumeric constant	
, out	value Raliye	representing a valid filename	
	Default Value	"mlod"	
	Required	No	
	Applicable Notes	1	

Notes

1. An analysis output file is generated for each analysis performed. The name of this file may be provided in the out attribute of the mlod parameter. If no filename is provided, the filename defaults to the name of the region with the extension ".lod" appended to it.

### 13.3.2 The mlod Parameter Block

The following lists all parameters that may occur in a MLOD block.

parameter	Explanation			
[, attribute]	Explanation			
	Specifies option to computes LOD scores at the ob- served markers or at the markers and intervals be- tween them.			
scan_type	Value Range marker interval			
	Default Value marker			
	Required No			
	Applicable Notes 3			
	Specifies title of the run.			
	Value Range Character string			
title	Default Value None			
	Required No			
	Applicable Notes None			
	Specifies the name of the region to be analyzed. Must			
	be a name listed in the genome description file.			
	Value Range Character string			
region	Default Value None			
	Required No			
	Applicable Notes 5			
	Character string representing the name of a trait- marker to be analyzed. MLOD requires at least one trait-marker to be specified, but the user may list as many as desired.			
trait_marker	Value Range Character string			
	Default Value None			
	Required Yes			
	Applicable Notes None			
	Maximum size (2n - f) of pedigree to analyze.			
	Value Range {0, 1, 2,}			
may gize	Default Value 18			
max size				
max_size				
max_size	Required Yes			
max_size	Required     Yes       Applicable Notes     None			
max_size	Required     Yes       Applicable Notes     None       Sets the interval used to compute LOD scores between			
	Required       Yes         Applicable Notes       None         Sets the interval used to compute LOD scores between observed markers in centiMorgans.			
max_size	$\begin{tabular}{ c c c c c } \hline Required & Yes & & & & & & & & & & & & & & & & & & &$			
	Required       Yes         Applicable Notes       None         Sets the interval used to compute LOD scores between observed markers in centiMorgans.			

	Print verbose resu pedigree.	lts of LOD score results for each	
	Value Range	true	
pedigree_lod_out		false	
	Default Value	false	
	Required	No	
	Applicable Notes	2	
	Controls the amour	nt of output generated on a per pedi-	
	gree basis.		
output_pedigrees	Value Range	{none, markers, all}	
output_pedigrees	Default Value	none	
	Required	No	
	Applicable Notes	3	
	Controls the amount of detail provided about the use-		
	able pedigree data sample.		
appropriate detail	Value Range	{none, removed, all}	
sample_detail	Default Value	removed	
	Required	No	
	Applicable Notes	4	

Notes

- 1. The scan\_type parameter defines the locations where LOD scores are to be computed. If the value of scan\_type is set to **markers**, then LOD scores are computed only at observed marker loci. If set to **intervals**, then LOD scores are computed both at the marker locations and at intervals between markers defined by the distance parameter.
- 2. The value of the pedigree\_lod\_out parameter affects the verbosity of the results in the LOD analysis output file. If pedigree\_lod\_out is set to **true**, LOD scores will be printed at marker loci, and each of the locations computed in between markers for each individual pedigree. Otherwise LOD scores are printed only at marker loci. This option does not affect how the overall LOD score results are output in the LOD analysis summary output file.
- 3. If output\_pedigrees is set equal to **markers**, pedigree tables are printed, but only for the markers. If set equal to **all**, all points in the region are produced.
- 4. If sample\_detail is set equal to **removed**, the table only includes those individuals removed from analysis (with reasons for removal), if set equal to **all**, all individuals are included in the table with reason for removal or being kept.
- 5. This causes the region to be analyzed using the current parameter settings and the corresponding output to be generated. If the value of the region parameter is not the name of a valid region, then the analysis is skipped. If no region parameters are specified in the parameter file, then multi-point LOD scores will be computed for each region in the genome description file. If no genome description file is given then single-point LOD scores will be generated at each marker in the marker locus description file.

#### **13.4 Program Execution**

MLOD is run via a command line interface on the supported UNIX and Windows platforms. This requires the S.A.G.E. programs to be properly installed and in the current execution path. Input files are specified on the command line and all output files are created in the current working directory.

Running MLOD from the command prompt with no arguments, or with the wrong number of arguments, will result in the program printing its usage statement. This lists the input files the program requires on the command line.

```
>mlod
S.A.G.E. v5.x -- MLOD
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
usage: mlod <params> <pedigree> <traits> <locus>
<map>
Command line parameters:
params - Parameters File
pedigree - pedigree data file
traits - trait-marker description file
locus - locus description file
map - Genome Map File
```

As indicated in the program usage statement, input files are listed on the command line. A typical run of MLOD may look like the following:

```
>mlod data.par data.ped data.trt data.loc data.map
S.A.G.E. v5.x -- MLOD
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
Loading Map file ...
Validating Analysis ...
ANALYSES
_____
LOD Score Analysis : LOD score analysis on region 'chr5'. Maximum connected
pedigree size is set to 18.
Processing Analyses....
LOD Score Analysis : LOD score analysis on region 'chr5'. Maximum connected
pedigree size is set to 18.
LOD Score Analysis: Pedigree 1
Generating Marker Likelihoods .....Done.
Generating Multipoint Combined Info.....Done.
Generating Marker Likelihoods .....Done.
LOD Score Analysis: Pedigree 2
Generating Marker Likelihoods .....Done.
Generating Multipoint Combined Info.....Done.
Generating Marker Likelihoods .....Done.
LOD Score Analysis: Pedigree 3
Generating Marker Likelihoods .....Done.
Generating Multipoint Combined Info.....Done.
Generating Marker Likelihoods .....Done.
```

# **13.5 Program Output**

Filename	File Type	Description
mlod.inf	Information output file	Contains informational diagnostic
		messages, warnings and program er-
		rors. No analysis results are stored in
		this file.
genome.inf	Genome Information File	Contains diagnostic information on
		the genetic map data and the marker
		loci that were provided for analysis.
		No analysis results are stored in this
		file.
region.lod	LOD analysis output file(s)	There is one LOD analysis output
		file for each analysis performed by
		MLOD. This file contains a table for
		each pedigree analyzed, listing LOD
		scores and information content for
		each trait analyzed at each marker lo-
		cation.
	LOD Analysis summary out-	
summary.out	put file	Contains a table for each analy-
		sis performed by MLOD. This table
		sums LOD scores and pools infor-
		mation content over all pedigrees for
		each point considered in the analysis.

MLOD produces several output files that contain results and diagnostic information:

#### **13.5.1** Information Output File

The MLOD Information file contains a variety of useful information, including:

- Information on fields read from the pedigree data file. These tables, which provide information about what the program has read in, are included with all programs in S.A.G.E. and are very useful for debugging many common errors caused when reading the pedigree data. When first analyzing new data, it is recommended that these tables be carefully checked to make sure pedigree data are being correctly read.
- Information, warning and error messages generated throughout the program. It is recommended that you check this file for warning and error messages before examining the results of any run of the program. The program attempts to correct many common errors and this sometimes means analyses are not run as expected. The file "mlod.inf" should be checked for errors and diagnostic information after each run of the program.

#### 13.5.2 Genome Output File

This file includes warnings and errors produced while parsing the locus description files, as well as a table for each marker listing allele and genotype population frequencies, assuming HardyWeinberg equilibrium. If allele frequencies do not sum to 1.0, they are standardized to 1.0, so these frequencies may not be as described in the locus description files. Note that a table is produced only for markers and not for trait-markers (traits with models.)

### 13.5.3 LOD Analysis Output File

A separate LOD analysis output file is created for each analysis performed by MLOD. This file contains a table for each pedigree analyzed, listing LOD scores and information content at each marker for each trait analyzed. Points between markers are also be listed if the pedigree\_lod\_out parameter has been set to **true**.

#### 13.5.4 LOD Summary Output File

The LOD summary output file contains a table for each analysis performed by MLOD. These tables summarize LOD scores and information content for each point considered in the analysis by summing LOD scores from all pedigrees in the data set into a single LOD statistic. Information content is similarly summarized.

# 13.6 Example Output Files

```
LOD Table File
```

Analysis	'Analys	====== sis 1'	on	marker	region	'chr5'	=====	=====			
Pedigree	1 1 1	LOD sc	ore	====== S							
=======			===:	====== D	ominant			=====	====== Reces:	====== sive	=====

		DOILL	nant	Reces	sive
	Marker	LOD score	Information	LOD score	Information
0.0	D5G1	-0.00001	0.94751443	-0.00001	0.94746010
0.7	D5G2	-0.00000	1.00000000	-0.00001	1.00000000
3.0	D5G3	-0.00000	0.97313461	-0.00001	0.97312913
5.4	D5G4	-0.00000	0.98934838	-0.00001	0.98933651
6.5	D5G5	-0.00000	1.00000000	-0.00001	1.00000000
9.3	D5G6	-0.00000	0.96847998	-0.00001	0.96847998
11.6	D5G7	-0.00056	0.96911857	-0.00001	0.96915157
14.8	D5G8	-0.00056	1.00000000	-0.00001	1.00000000

# Summary LOD score Output file

Analysis: Analysis 1 (chr5

-----

Pos	Trait	Marker	LOD score	Information	# Ped.
0.0	Dominant	D5G1	-5.519175122	0.96972173829	239
0.7	Dominant	D5G2	-5.555971669	0.98266560809	239
3.0	Dominant	D5G3	-5.979345816	0.98893249560	239
5.4	Dominant	D5G4	-5.724044828	0.98947251776	239
6.5	Dominant	D5G5	-5.771942142	0.98631478376	239
9.3	Dominant	D5G6	-6.028611158	0.97542619041	239
11.6	Dominant	D5G7	-6.498739837	0.99294067379	239
14.8	Dominant	D5G8	-6.815240677	0.96763531437	239
17.2	Dominant	D5G9	-6.298380200	0.98134489280	239
19.6	Dominant	D5G10	-5.971049479	0.98499045128	239
22.6	Dominant	D5G11	-5.115293921	0.97504538569	239
23.5	Dominant	D5G12	-5.219382005	0.98451863326	239
26.1	Dominant	D5G13	-6.053329330	0.99199116082	239
27.9	Dominant	D5G14	-5.906801214	0.99470950295	239
30.3	Dominant	D5G15	-6.435072155	0.99356584224	239

# **Chapter 14**

# ASSOC

ASSOC analyzes the association between a continuous trait and one or more covariates from pedigree data in the presence of familial correlations, simultaneously estimating familial variance components (and hence familial correlations and heritability). Covariates may include marker phenotypes that have been transformed into quantitative covariates by using a function block in the parameter file. Given data on one or more independent pedigrees sampled at random, this program estimates (by maximum likelihood assuming a generalization of multivariate normality) the parameters of the model with and without inclusion of a specified set of covariates. It provides the corresponding values of the ln(likelihood), as well as twice the difference between these values, so that the significance of the set of covariates can be determined. It also calculates numerically the standard errors of the estimates of all individual parameters in the model and performs an appropriate Wald test on each.

## 14.1 Limitations

ASSOC does not support pedigrees with loops, mating clusters of more than three total individuals (one individual with two spouses), or mating chains. Any constituent pedigree data that contains these structures will be excluded from the analysis. Also, ASSOC does not infer marker genotypes from their relatives' genotypes.

Further, if the sample size is small relative to the number of parameters being estimated, the likelihood may have multiple maxima. There is no guarantee that in such a situation the maximum found and reported by the program is also the global maximum. Also, situations can occur in which it is not numerically possible to calculate the variance-covariance matrix of the estimates.

### 14.2 Theory

#### 14.2.1 Description of the Model

To incorporate familial correlations and arbitrary covariates into a likelihood, we assume the correlation structure described in Elston, George and Severtson (1992) and the regression model described in George and Elston (1987). For individual *i*, let:

$Y_i$	=	a continuous trait
$x_i$	=	a vector of covariates
$G_i$	=	a polygenic effect
$F_i$	=	family effect
$F_i^{\prime}$	=	family effect
$M_i$	=	a marital effect
$S_i$	=	a sibship effect
$E_i$	=	a random environmental effect

Then for a continuous trait the model is of the form

$$h(Y_{i}) = h(\beta^{T} x_{i}) + G_{i} + F_{i} + F_{i}' + M_{i} + S_{i} + E_{i},$$

where h is a transformation<sup>1</sup>, and the polygenic effect  $(G_i)$  and all the environmental effects  $(F_i, M_i, S_i, E_i)$ are assumed to be normally distributed random effects with zero means. All covariates are centered prior to inclusion in the likelihood (see below). Because the transformation h is applied to both sides of the equation, the estimates of the parameter values in  $\beta$  are on the original scale on which  $Y_i$  is measured.  $F_i$  is a random effect that a person shares with his/her spouse and children;  $F'_i$ is an effect that person shares with his/her parents, full sibs, half sibs and half sibs' parents;  $M_i$ is an effect that spouses share with each other;  $S_i$  is an effect that full sibs share with each other; and  $E_i$  is a person-specific random effect. These random effects are assumed to have variances  $\sigma_G^2$ ,  $\sigma_F^2 = \sigma_{F'}^2$ ,  $\sigma_M^2$ ,  $\sigma_S^2$  and  $\sigma_E^2$ , respectively, and these variances are on the transformed scale. Thus,

$$V[h(Y)] = \sigma_G^2 + 2\sigma_F^2 + \sigma_M^2 + \sigma_S^2 + \sigma_E^2,$$
(14.1)

where any of the variances other than  $\sigma_E^2$  can be set to zero. However, if only 2-generation families are present in the sample,

$$V\left[h\left(Y\right)\right] = \sigma_{G}^{2} + \sigma_{F}^{2} + \sigma_{M}^{2} + \sigma_{S}^{2} + \sigma_{E}^{2}.$$

For  $\sigma_F^2$  to be estimable, it is often necessary to have large pedigrees or large numbers of pedigrees, or both, and therefore  $\sigma_F^2$  is set equal to zero by default. Variance components divided by the total variance can be interpreted as intraclass correlations (interclass in case of the marital correlation); it is not possible to estimate any variances to be less than zero.

#### 14.2.2 Likelihood for a Randomly Sampled Pedigree

The likelihood formulation is based on the assumption of normality of the residuals and on the assumed correlational structure of the  $y_j$ . An algorithm for computing this likelihood is described in Elston et al. (1992).

It should be noted that singletons (unrelated individuals) may be included in the data. Although ASSOC counts and treats them separately for convenience, they are in fact simply one-person pedigrees with parent information missing and, as such, require no special treatment in the model.

<sup>&</sup>lt;sup>1</sup>The dependent variable y may be transformed by one of two transformations: Box-Cox or George-Elston. See section 5.3.2.2 for details on the transformation theory implemented in this program.

#### 14.2.3 Estimation of Parameters

Estimation is performed by maximizing the likelihood numerically. For computational reasons, however, the likelihood is not considered directly. Instead of the likelihood L based on the above model description, we maximize its natural logarithm ln(L). If several independent pedigrees are analyzed jointly, the logarithms of the likelihoods are summed over all pedigrees.

The program itself determines initial estimates for the maximizing process. The user, however, may override the initial estimate of any of the variance components or covariate coefficients.

#### 14.2.4 Tests

ASSOC can be executed in one of two ways. If the user does not specify any test covariates, the likelihood maximization will be performed *once*, and there will be no joint-test output.

If, however, the user does specify test covariates, the maximum ln(likelihood) of the model is determined under two hypotheses:  $H_1$  assumes the general model, including all test covariates specified by the user;  $H_0$ , the null hypothesis, excludes from the regression model the test covariates specified by the user. If  $L_1$  and  $L_0$  are the maximum likelihoods under  $H_1$  and  $H_0$  respectively, then the likelihood ratio statistic is  $2[ln(L_1) - ln(L_0)]$ . Under the assumption of normality of the transformed variable and the null hypothesis that the test covariates have no effect, this statistic is asymptotically distributed as chi-square with the number of degrees of freedom being equal to the number of test covariates. In addition, p-values are calculated for each individual parameter in the model using its standard error obtained by double differentiation of the ln likelihood. These p-values are two-sided for the covariate coefficients  $\beta$ ,  $\lambda_1$ , and  $\lambda_2$ ; they are one-sided for all variances. In each case the test is for the null hypothesis that the parameter is 0, except for  $\lambda_1$ , where the null hypothesis is  $\lambda_1 = 1$ .

# 14.3 Program Input

File Type	Description		
Parameter file	Specifies the parameters and options with which		
	to perform a particular analysis.		
Pedigree data File	Contains delimited records for each individual in-		
	cluding fields for identifiers, parents, trait and co-		
	variates.		
Marker locus description file	Lists the alleles at each marker locus. This file		
	will be used by ASSOC only if a marker, trans-		
	formed to be quantitative, is used as a covariate $^{a}$ .		

<sup>*a*</sup>ASSOC does not use any information on allele frequencies or phenotype to genotype mapping that may be in the Marker Locus Description File.

#### 14.3.1 Parameter File Syntax

The specific syntax for ASSOC parameters, attributes and values is described in the following sections.

#### 14.3.1.1 The assoc Parameter

The following syntax table specifies the permissible parameter and attribute settings for the main ASSOC parameter.

<pre>parameter [, attribute]</pre>		Explanation		
	Starts an ASSOC a	nalysis block.		
	Value Range	N/A		
assoc	Default Value	None		
assoc_analysis,	Required	Yes		
	Applicable Notes	None		
Ì	Specifies the root name to be used for output files.			
	Output file names will be formed by concatenating the			
	root name and an appropriate extension.			
	Value Range	Valid file name in the form of a		
		quoted character string.		
, out	-	analysis_n, where $n = 1, 2,, k$		
	Default Value	for a given set of k specified		
		ASSOC analyses.		
	Required	No		
	Applicable Notes	None		

#### 14.3.1.2 The assoc Block

The following syntax table specifies the permissible parameter and attribute settings for the assoc block.

parameter		Explanation		
[, attribute]		_		
	Specifies the title o			
	Value Range	Quoted character string.		
		analysis_n, where $n = 1, 2,, k$		
title	Default Value	for a given set of k specified		
		ASSOC analyses.		
	Required	No		
	Applicable Notes	1		
	Specifies a depend	lent variable as the trait in the re-		
	gression model.			
		Character string representing the		
		name of a trait, phenotype or		
trait	Value Range	covariate from the pedigree data		
primary_trait		file.		
	Default Value	None		
	Required	Yes		
	Applicable Notes	None		
	Specifies a variable in the regression model. It can be			
	a trait name, a covariate name, or a phenotype name.			
	Value Range	The name of a trait, phenotype,		
covariate		covariate.		
COV	Default Value	None		
	Required	No		
	Applicable Notes	2, 3		
	Specifies a covaria	ate to be included in the general		
	model $H_1$ , but excluded from the null hypothesis $H_0$ .			
	Value Range	N/A		
, test	Default Value	None		
	Required	No		
	Applicable Notes	3		
	Specifies the initia	l estimate for the covariate coeffi-		
	cient.			
	Value Range	$(-\infty,\infty)$		
, val	Default Value	None		
	Required	No		
	Applicable Notes	None		
		oefficient for this covariate is fixed.		
	Value Range	{true, false}		
, fixed	Default Value	false		
	Required	No		
	Applicable Notes	4		
		•		

	Specifies the inclusion of a polygenic variance component in the model.
	Value Range {true, false}
polygenic_effect	Default Value true
ре	Required No
	Applicable Notes 5
	Specifies the initial estimate for this variance compo-
	nent.
, val	Value Range $[0,\infty)$
, vai	Default Value None
	Required No
	Applicable Notes 5
	Specifies that the effect is fixed.
	Value Range {true, false}
, fixed	Default Value false
,	Required No
	Applicable Notes 4
	Specifies the inclusion of a nuclear family variance
	component in the model.
	Value Range {true, false}
family_effect	Default Value false
fe	
	Required No
	Applicable Notes 5
	Specifies the initial estimate for this variance compo-
	nent.
, val	Value Range $[0,\infty)$
, vai	Default Value None
	Required No
	Applicable Notes None
	Specifies that this effect is fixed.
	Value Range {true, false}
, fixed	Default Value false
	Required No
	Applicable Notes 4
	Specifies the inclusion of a marital (i.e., spousal) vari-
	ance component in the model.
	Value Range {true, false}
marital_effect	Default Value true
me	Required No
	Applicable Notes 5
	Specifies the initial estimate for this variance compo-
	nent.
, val	Value Range $[0,\infty)$
,	
	Default Value None
	Default value     None       Required     No       Applicable Notes     None

	Specifies that the e	ffect is fixed.		
	Value Range	true		
, fixed		false		
, TIXEU	Default Value	false		
	Required	No		
	Applicable Notes	4		
	-	sion of a sibling variance compo-		
	nent in the model.			
sibship_effect	Value Range	{true, false}		
se	Default Value	true		
	Required	No		
	Applicable Notes	5		
	Specifies the initia	l estimate for this variance compo-		
	nent.			
, val	Value Range	$[0,\infty)$		
, var	Default Value	None		
	Required	No		
	Applicable Notes	None		
	Specifies that the effect is fixed.			
	Value Range	{true, false}		
, fixed	Default Value	false		
	Required	No		
	Applicable Notes	4		
	Starts a transforma			
	Value Range	N/A		
transformation	Default Value	None		
transform	Required	No		
trans	Applicable Notes	7		
		ub-block to specify diagnostics op-		
	tions for the likelih	nood maximization process.		
maxfun	Value Range	N/A		
	Default Value	None		
	Required	No		
	Applicable Notes	8		
	-	s the user to substitute covariates'		
		or missing covariate data.		
allow_averaging	Value Range	{mean, none}		
allow_averaging aa	Default Value	none		
	Required	No		
	Applicable Notes	6		

Notes

- 1. The title parameter also specifies the naming convention for ASSOC output files. However, the value of the out attribute (of the assoc parameter) will override the title parameter as the name of output files.
- 2. If a sex\_code covariate is specified, the estimated effect will be for that of a female (i.e.

males are coded 0, females are coded 1). This requires that the sex\_code has been specified as available to be used as a trait.

- 3. The trait analyzed can be a linear function of the primary trait (with coefficient 1) and other covariates whose coefficients are fixed or estimated. This linear function is called a composite trait. Without this sub-block a composite trait is not formed. All covariates are centered, the centering (average) value being included as part of the output. The covariates can be any covariate, phenotype or trait (other than the primary trait) listed in the pedigree data file. Note: This sub-block is not applicable to binary traits.
- 4. If the fixed attribute is set to **true**, the attribute val must be included. If set to **false** and the attribute val is included, this determines the initial value of the variable to be used in the maximization process. If set to **false** and the attribute val is not included, then the program supplies various initial values for the maximization process.
- 5. If val is set to 0 and fixed is set to **true**, the relevant effect (polygenic, family, marital, etc.) will be excluded from the model. This is equivalent to setting the effect to **false** (i.e., pe = false, fe = false, or me = false).
- 6. If the value **none** is specified, and any individual's value is missing for any particular covariate, then that individual will be treated as uninformative for the purpose of the analysis. If **mean** is specified, the individual's missing covariate value will be replaced with the sample mean of that covariate (calculated on the basis of all fully informative individuals).
- 7. By default, ASSOC will estimate  $\lambda_1$  (and fix  $\lambda_2$  at 0), using the George-Elston transformation. See section 5.3.2.2 for details on the transformation theory implemented in this program.
- 8. MAXFUN (*function maximization*) is the name of the S.A.G.E. library component designed to find the global maximum of a given function with respect to a vector of parameters. When a S.A.G.E. program reports a failure of convergence to a global maximum, the maxfun subblock provides the ability to generate diagnostic information about the maximization process. The diagnostic information is generated to a text file with a ".max" extension on the filename.

#### 14.3.1.3 The transformation sub-block

The following syntax table specifies the permissible parameter and attribute settings for the transformation sub-block.

parameter [, attribute]	Explanation		
[, actibute]	Specifies a particular transformation option.		
	Value Range	{none, box_cox, george_elston}	
option	Default Value		
option		george_elston	
	Required	1	
	Applicable Notes	1	
	Specifies the power	*	
	Value Range	N/A	
lambda1	Default Value	None	
	Required	No	
	Applicable Notes	None	
	Value of the parame	eter	
	Value Range	$(-\infty,\infty)$	
, val	Default Value	1	
	Required	No	
	Applicable Notes	None	
	Specifies option to t	fix the value.	
	Value Range	{true, false}	
, fixed	Default Value	true	
	Required	No	
	Applicable Notes	None	
	Specifies inclusive	lower bound for power parameter.	
	Value Range	$(-\infty,\infty)$	
, lower_bound	Default Value	-1	
	Required	No	
	Applicable Notes	None	
		upper bound for power parameter.	
	Value Range	$\frac{11}{(-\infty,\infty)}$	
, upper_bound	Default Value	+∞	
,	Required	No	
	Applicable Notes	None	
	Specifies the shift p		
	Value Range	N/A	
lambda2	Default Value	None	
Tambadz	Required	No	
	Applicable Notes	None	
	Specifies the value		
	Value Range	$\frac{1}{(-\infty,\infty)}$	
	Default Value	$\frac{(-\infty,\infty)}{0}$	
, val	Required	No	
	Applicable Notes	None	
	Applicable Notes	110110	

	Option to fix this v	alue.
	Value Range	{true, false}
, fixed	Default Value	None
	Required	No
	Applicable Notes	None

Notes

1. An option value of **none** disables transformation calculations for the analysis, and an option value of either **george\_elston** or **box\_cox** means that both the  $\lambda_1$  and  $\lambda_2$  transformation parameters are to be estimated.

#### 14.3.1.4 The maxfun sub-block

The following syntax table specifies the permissible parameter and attribute settings for the maxfun sub-block.

<pre>parameter [, attribute]</pre>	Explanation		
	Specifies the quantity and type of diagnostic informa-		
	tion desired from the	he MAXFUN component.	
		no_debug_info	
	Value Benge	basic	
level	Value Range	per_run	
		complete	
	Default Value	None	
	Required	No	
	Applicable Notes	1	

Notes

The following are all valid assoc statements:

```
assoc_analysis {
  trait = TRAIT1
  cov = TRAIT2, test
}
assoc_analysis, out = my_test {
  trait = trait1
  cov = a_covariate, test
}
assoc_analysis {
  title = "Analysis, Oct. 8, 2001"
   trait = TRAIT3
   cov
       = COV1, test
}
assoc_analysis, out=Assoc_res {
  trait = TRAIT3
   cov = COV1, test
   cov = cov2, test
   cov
       = Cov3
   cov
       = TRAIT1
}
assoc_analysis, out=test_analysis {
  title = "Test Ignore"
  COV
               = a_trait1
  cov
               = a_trait2
  cov
              = a_test_cov_1, test
  COV
               = a_test_cov_2, test
  primary_trait = the_trait
}
```

#### 14.3.2 Exclusion Criteria for Individuals and Pedigrees

Under some conditions ASSOC will exclude individuals and/or pedigrees from the analysis, and the user is advised to take note of program outputs that indicate the numbers of valid and invalid individuals being counted. The exclusion criteria are:

- 1. Any structurally invalid families will be excluded from the analysis (see 16.1).
- 2. Any individual whose primary trait is missing or who is missing at least one covariate value (if the allow\_averaging option is disabled the default behavior) will be retained to provide relationship information for the analysis, but all of the individual's trait information will be treated as missing.

## 14.4 Program Execution

ASSOC is run via a command line interface on the supported UNIX and Windows platforms. This requires the S.A.G.E. programs to be properly installed and in the current execution path. Input files are specified on the command line and all output files are created in the current working directory.

Running ASSOC from the command prompt with no arguments, or the wrong number of arguments, will result in the program printing its usage statement. This lists the input files the program requires on the command line.

```
>Assoc
S.A.G.E. v5.x -- ASSOC
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
usage: Assoc <parameters> <pedigree> [locus]
Command line parameters:
parameters - Parameter File
pedigree - Pedigree Data File
locus - Locus Description File (optional)
```

As indicated in the program usage statement, input files are listed on the command line. A typical run of ASSOC may look like the following:

```
>ASSOC data.par data.ped
S.A.G.E. v5.x -- ASSOC
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
Reading Parameter File......done.
Reading Pedigree File......done.
from data.ped.....done.
Sorting pedigrees.....done.
Verifying Analyses:
Analysis 1.....done.
Analysis 2......done.
```

# 14.5 Program Output

ASSOC produces three output files for each analysis, an Information Output File, a Summary Output File and a Detailed Output File:

Filename	Filetype	Description
analysis.inf	Information output file	Contains informational diagnostic messages,
		warnings and program errors. No analysis results
		are stored in this file.
analysis.sum	ASSOC calculation	Contains the final estimates and standard errors of
	summary output	the parameters used in the model
analysis.det	ASSOC calculation	This file contains the variance-covariance matrix
	detailed output	of the estimates and the partial derivatives of the
		log likelihood with respect to the parameters.

#### 14.5.1 Information Output File

The ASSOC Information file contains a variety of useful information, including:

- Information on fields read from the Pedigree Data File. These tables, which provide information about what the program has read from the Pedigree Data File, are included with all programs in S.A.G.E., and are very useful for debugging most common errors caused when reading the pedigree data. When first analyzing new data, it is recommended that these tables be checked carefully to make sure pedigree data are being correctly read.
- Information, warning and error messages generated throughout the program. It is recommended that this file be checked before examining the results of any run of the program. The program attempts to correct many common errors and this sometimes means analyses are not performed as expected. The file "assoc.inf" should be checked for errors and diagnostic information after each run of the program.
- A table indicating the structural validity of each constituent subpedigree in the pedigree data file used for the analysis. For each constituent subpedigree, the table lists whether or not it has loops, mating chains, or mating clusters larger than 3 total individuals.

#### 14.5.2 Summary Output File

The Summary Output File contains the final estimates, standard error, and p-value of the parameters used in the model:

- 1. Transformation parameters
  - $\lambda_1$ -Lambda 1
  - $\lambda_2$ -Lambda 2
- 2. Variance components on the transformed scale.
  - $\sigma_G^2$ -Polygenic variance
  - $\sigma_E^2$ -Random variance
  - $\sigma_F^2$ -Familial variance
  - $\sigma_M^2$ -Marital variance
  - $\sigma_S^2$ -Sibship variance
- 3. Coefficients
  - $\beta_0$ -Intercept
  - $\beta_j$ -Covariate coefficients, j > 0
- 4. Total variance:  $V[h(Y_i)]$
- 5. "Heritability":  $\sigma_G^2/V[h(Y_i)]$

 $(\sigma_F^2 + \sigma_S^2 + \frac{1}{2}\sigma_G^2) / \{V[h(Y_i)]\}$ 

 $(\sigma_F^2 + \frac{1}{4}\sigma_G^2) / \{V[h(Y_i)]\}$ 

 $(\sigma_F^2 + \frac{1}{2}\sigma_G^2) / \{V[h(Y_i)]\}$ 

 $\sigma_{F}^{2}/\{V[h(Y_{i})]\}$ 

- 6. Environmental intraclass correlations
  - Nuclear Family:  $\sigma_F^2 / \{V[h(Y_i)] \sigma_G^2\}$
  - Marital:  $(\sigma_F^2 + \sigma_F^2) / \{V[h(Y_i)] \sigma_G^2\}$
  - Sibship:  $(\sigma_F^2 + \sigma_S^2) / \{V [h(Y_i)] \sigma_G^2\}$
- 7. Residual familial correlations
  - Full Sib
  - Half Sib

•

- Parent-offspring
- Step-parent-step-offspring
  - Spouses:  $(\sigma_F^2 + \sigma_M^2) / \{V[h(Y_i)]\}$

In addition, several p-values are quoted based on the asymptotic distribution of the test statistics (likelihood ratio, Wald). p-values quoted for  $\sigma_G^2$ ,  $\sigma_E^2$ ,  $\sigma_F^2$ ,  $\sigma_M^2$  and  $\sigma_S^2$  use a 1-sided test. All other p-values use 2-sided tests.

#### 14.5.3 Detailed Output File

The detailed output includes:

- 1. The estimated variance-covariance matrix of all the estimated parameters.
- 2. The partial first derivative of the natural logarithm of the likelihood with respect to each of the parameters estimated.

# 14.6 Example Output Files

#### 14.6.1 ASSOC Summary Output File

The following is an example of an ASSOC Summary Output File:

-----Results -----------Sample description -Number of pedigrees in dataset Number of analyzable pedigrees 4 Number of individuals in dataset 942 Number of analyzable individuals 923 Number of analyzable invalid individuals 424 Number of analyzable valid individuals 499 Model description Primary Trait gqrtdbb Test covariate: Test covariate: Mean = 0.512146 Std. dev. = 0.499852 Min. = 0.000000 Max. = 1.000000 Note: No transformation is applied. -----MAXIMIZATION RESULTS without test covariates ------\_\_\_\_\_ ------Estimate S.E. P-value Parameter Variance components Polygenic 2.251733 0.335338 < 1e-07 Family 0.000000 Ind. func. fixed @ bnd 0.000000 Ind. func. fixed @ bnd Marital Sibling 0.179865 0.098181 0.033478 Random 0.300521 0.191533 0.058320 Correlations 0.000000 0.000000 0.000000 0.000000 Nuclear family 0.004495 Marital 0.004495 0.374417 0.237980 Sibling 0.115646 Full sibs 0.477919 0.047367 < 1e-07 Half sibs 0.206043 0.017349 < 1e-07 Parent-offspring 0.412085 0.034699 < 1e-07 Step parent-offspring 0.000000 0.000000 < 1e-07 Spouses or spouses of common spouse 0 000000 0 000000 < 1e-07 Intercept 4.941494 0.137505 < 1e-07 Other parameters Total variance 2.732120 0.225741 Heritability 0.824171 0.069397 < 1e-07 < 1e-07 -----Final ln likelihood: -881 373404 MAXIMIZATION RESULTS with test covariates Estimate S.E. P-value Parameter Variance components 2 223050 0 335194 Polvgenic < 1e-07 0.000000 Ind. func. fixed @ bnd Family 0.000000 Ind. func. fixed @ bnd Marital 0.031135 Sibling 0.185202 0.099337 0.311680 0.191454 0.051766 Random Correlations Nuclear family 0.000000 0.000000 0.003702 Marital 0.000000 0.000000 0.003702 Sibling 0.372727 0.229468 0.104309 Full sibs Half sibs 0.476750 0.047508 < 1e-07 0.204330 0.017667 < 1e-07 Parent-offspring 0.408659 0.035334 < 1e-07 0.000000 0.000000 Step parent-offspring < 1e-07 0.000000 0.000000 < 1e-07 Spouses or spouses of common spouse 4.948390 0.137428 Intercept < 1e-07 Test covariates apres 0.094199 0.142815 0.509517 Other parameters Total variance 2.719933 0.224600 < 1e-07 Heritability 0.817318 0.070668 < 1e-07 \_\_\_\_\_ \_ \_ \_ \_ \_ \_ \_ \_ \_ \_ \_ \_ \_ \_ \_ Final ln likelihood: -881.155989 \_\_\_\_\_ JOINT TEST ------HO ln likelihood without test covariates -881.373404 H1 ln likelihood with test covariates -881.155989 0.434830 2 \* |HO - H1| Degrees of freedom P-value 0.509628

# 14.6.2 ASSOC Detailed Output File

The following is an example of an ASSOC Detailed Output File:

-----Results \_\_\_\_\_ -----Sample description Number of pedigrees in dataset Number of analyzable pedigrees 4 Number of individuals in dataset 942 Number of analyzable individuals 923 Number of analyzable invalid individuals 424 Number of analyzable valid individuals 499 Model description Primary Trait pe\_se\_fe\_me Primary Trait sqrtdbh Test covariate: Mean = 0.512146 Std. dev. = 0.499852 Min. = 0.000000 Max. = 1.000000 apres Note: No transformation is applied. --MAXIMIZATION RESULTS without test covariates \_\_\_\_\_ Estimate S.E. P-value Deriv Parameter ----------Variance components Polygenic 2.251733 0.335338 < 1e-07 0.000003594 0.000000 Ind. func. fixed @ bnd 0.000000 Ind. func. fixed @ bnd Family Marital Ind. func. fixed @ bnd Sibling 0.179865 0.098181 0.033478 0.0000002697 Random 0.300521 0.191533 0.058320 0.0000004046 Correlations 0.000000 0.00000 0.000000 0.000000 Nuclear family 0.004495 0.000000000 Marital 0.004495 0 0000000000 0.374417 0.237980 0.115646 Sibling 0.0000000000 Full sibs 0.477919 0.047367 < 1e-07 0.000000000 Half sibs 0.206043 0.017349 < 1e-07 0 0000000000 Parent-offspring 0.412085 0.034699 < 1e-07 0.000000000 Step parent-offspring 0.000000 0.000000 < 1e-07 0.000000000 Spouses or spouses of common spouse 0 000000 0.000000 < 1e-070 0000000000 Intercept 4.941494 0.137505 < 1e-07 0.0000000000 Other parameters Total variance 2.732120 0.225741 < 1e-07 Heritability 0.824171 0.069397 < 1e-07 0 0000000000 0.0000000000 -----Final ln likelihood: -881.373404 MAXIMIZATION RESULTS with test covariates Estimate S.E. P-value Deriv Parameter Variance components 2.223050 0.335194 < 1e-07 Polvgenic 0 0000001517 0.000000 Ind. func. fixed @ bnd 0.000000 Ind. func. fixed @ bnd Family Marital 0.185202 0.099337 0.311680 0.191454 Sibling 0.031135 -0.0000001349 0.051766 0.000001349 Random Correlations Nuclear family 0.000000 0.000000 0.003702 0.0000000000 Marital 0.000000 0.000000 0.003702 0.0000000000 Sibling 0.372727 0.229468 0.104309 0.0000000000 Full sibs Half sibs 0.476750 0.047508 < 1e-07 0.0000000000 0.204330 0.017667 < 1e-07 0.000000000 Parent-offspring 0.408659 0.035334 < 1e-07 0.000000 < 1e-07 0.0000000000 0.000000 Step parent-offspring 0.000000000 0.000000 0.000000 < 1e-07 0.000000000 Spouses or spouses of common spouse Intercept 4.948390 0.137428 < 1e-07 0.0000000000 Test covariates apres 0.094199 0.142815 0.509517 0.000000000 Other parameters Total variance 2.719933 0.224600 < 1e-07 0.000000000 Heritability 0.817318 0.070668 < 1e-07 0.0000000000 Final ln likelihood: -881.155989 \_\_\_\_\_ JOINT TEST ------HO ln likelihood without test covariates -881.373404 H1 ln likelihood with test covariates -881.155989 2 \* |HO - H1| 0.434830 Degrees of freedom 0.509628 P-value \_\_\_\_\_ VARIANCE-COVARIANCE MATRIX without test covariates -----Polygenic Family Marital Sibling Random Nuclear family Marital ... \_\_\_\_\_

	Polygenic	0.112451	0.000000	0.000000	0.007029	-0.052069	0.00000	0.00000	
	Family	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	
	Marital	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	
	Sibling	0.007029	0.00000	0.000000	0.009640	-0.008868	-0.000000	-0.000000	
	Random	-0.052069	0.00000	0.000000	-0.008868	0.036685	-0.000000	-0.000000	
	Nuclear family	0.000000	0.00000	0.00000	-0.000000	-0.000000	0.00000	0.00000	
	Marital	0.000000	0.00000	0.00000	-0.000000	-0.000000	0.00000	0.00000	
	Sibling	0.049737	0.00000	0.000000	0.019465	-0.040141	0.000000	0.000000	
	Full sibs	0.011360	0.00000	0.00000	0.003450	-0.008533	0.00000	0.00000	
	Half sibs	0.005206	0.00000	0.00000	0.000055	-0.002936	0.00000	0.00000	
	Parent-offspring	0.010412	0.00000	0.000000	0.000110	-0.005871	0.00000	0.00000	
	Step parent-offspring	-0.000000	0.00000	0.000000	-0.000000	0.000000	-0.000000	-0.000000	
	Spouses or spouses of common spouse	-0.000000	0.00000	0.00000	-0.000000	0.000000	-0.000000	-0.000000	
	Intercept	-0.003622	0.00000	0.000000	-0.000332	0.002230	-0.000000	-0.000000	
	Total variance	0.067411	0.00000	0.00000	0.007801	-0.024253	0.00000	0.00000	
	Heritability	0.020824	0.00000	0.000000	0.000220	-0.011742	0.00000	0.00000	
==			======						
	VARIANCE-COVARIANCE MATRIX with t	est covaria	tes						
==									
		Polygenic	Family	Marital	Sibling	Random	Nuclear family	Marital	
	Polygenic	0.112355	0.000000	0.000000	0.006410	-0.052018	0.000000	0.000000	
	Family	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	
	Marital	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	
	all'	0.000000			0.000000			0.000000	

rorygenic	0.1123333	0.000000	0.000000	0.000110	0.052010	0.000000	0.000000	
Family	0.000000	0.000000	0.000000	0.00000	0.00000	0.00000	0.00000	
Marital	0.000000	0.000000	0.00000	0.00000	0.00000	0.00000	0.00000	
Sibling	0.006410	0.000000	0.000000	0.009868	-0.008608	-0.000000	-0.00000	
Random	-0.052018	0.000000	0.00000	-0.008608	0.036654	-0.000000	-0.000000	
Nuclear family	0.000000	0.000000	0.000000	-0.000000	-0.000000	0.000000	0.00000	
Marital	0.000000	0.000000	0.000000	-0.00000	-0.000000	0.00000	0.00000	
Sibling	0.047113	0.000000	0.000000	0.018915	-0.038363	0.00000	0.00000	
Full sibs	0.011311	0.000000	0.000000	0.003462	-0.008525	0.00000	0.00000	
Half sibs	0.005313	0.000000	0.000000	0.000013	-0.002980	0.00000	0.00000	
Parent-offspring	0.010626	0.000000	0.000000	0.000026	-0.005961	0.00000	0.00000	
Step parent-offspring	-0.000000	0.000000	0.000000	-0.00000	0.00000	-0.000000	-0.00000	
Spouses or spouses of common spouse	-0.000000	0.000000	0.000000	-0.00000	0.00000	-0.000000	-0.00000	
Intercept	-0.003931	0.000000	0.000000	-0.000244	0.002354	-0.000000	-0.00000	
Total variance	0.066747	0.000000	0.000000	0.007670	-0.023972	0.00000	0.00000	
Heritability	0.021251	0.000000	0.000000	0.000052	-0.011921	0.00000	0.00000	
apres	-0.005830	0.000000	0.000000	0.001160	0.002525	-0.000000	-0.00000	

# **Chapter 15**

# TDTEX

The transmission-disequilibrium test (TDT) introduced by Spielman et al. (1993) is a method for detecting linkage between a marker locus and a disease susceptibility locus when linkage disequilibrium or any other type of allelic association is present. The basic TDT test for binary traits has been generalized by Bickeböller and Clerget-Darpoux (1995), Rice et al. (1995), Curtis and Sham (1995), Olson et al. (1997). TDTEX is a computer program based on this work, and implements a very general system for detecting linkage in the presence of linkage disequilibrium between a marker locus and a disease locus with a binary phenotype.

# 15.1 Limitations

The TDTEX program makes the following assumptions:

- 1. Each marker has a known genotype-phenotype relation.
- 2. Only autosomal loci are considered.
- 3. Only binary phenotypes are considered.

This program is limited by the program execution time of the computer on which it runs. As the transmission table size and number of marker alleles increase, processing time becomes slower. The major computational limitation is the exact permutation algorithm. This becomes prohibitively slow for transmission tables with greater than around 300 observations, or with more than about 8 alleles. In such cases, the asymptotic or Monte Carlo test statistics are recommended instead.

# 15.2 Theory

TDTEX consists of four main components:

1. A scoring algorithm to identify which alleles or genotypes are transmitted to affected offspring.

- 2. Transmission tables (i.e., contingency tables) to summarize the number of transmitted vs. non-transmitted alleles or genotypes.
- 3. A pedigree sampler to identify and collect informative transmissions from pedigree data. The sampler collects transmission information in transmission tables, conditional on the types of relatives to be sampled (individual affected offspring or affected sibling pairs), the availability of marker data, and optionally on parental traits such as sex or affection status.
- 4. A suite of statistical tests to evaluate significance of the computed transmission tables under the null hypothesis of complete symmetry or marginal homogeneity. These tests include the standard asymptotic TDT tests which rely on large sample theory for validity. Exact tests that do not rely on asymptotic approximations are also provided at the expense of greater computational requirements.

### 15.2.1 Allele and Genotype Transmissions

Consider a sample of affected individuals and their parents typed for a genetic marker. The basis of the transmission-disequilibrium test is a case/control study, matching alleles found in an affected individual with internal family-based control alleles. The "case" alleles are those that were transmitted to an affected individual, and "control" alleles are the alleles not transmitted from the parents of the individual. By scoring these transmitted and non-transmitted alleles from pedigree data, it is possible to estimate the distribution of these transmissions. If the marker and trait loci are unlinked or are unassociated (in equilibrium), then the distribution of parental alleles transmitted to the affected offspring will not differ in expectation from that of alleles that were not transmitted to the affected offspring. Otherwise, if *both linkage and disequilibrium* (or, more generally, linkage and allelic association, whatever the cause of that association) are present between marker and trait loci, then the distribution of alleles transmitted to the affected offspring will differ from that of the non-transmitted alleles. This scheme has the advantage of being robust to the presence of population stratification, a situation caused by admixture of populations with distinct marker allele and disease frequencies. For more details see Spielman et al. (1993).

We define an *allele transmission* from a single parent to a child to be an ordered pair of alleles, where the first allele is transmitted from the parent to the child and the second allele is the other parental allele, i.e., the one that is not transmitted to the child. In other words, an allele transmission is the ordered pair  $(A_1, A_2)$  where  $A_1$  is the transmitted allele, and  $A_2$  is the non-transmitted allele.

It is possible to combine the information from the allele transmissions from each of the two parents to a child. Since two allele transmissions involve two transmitted alleles (and two non-transmitted alleles), we can group the transmitted (and non-transmitted) alleles together to form a genotype. Thus a *genotype transmission* is defined as an ordered pair of genotypes, where the first genotype is formed by the two alleles transmitted from the parents to the child, i.e., the genotype of the child. Similarly, the second genotype includes the two alleles not-transmitted from the parents to the child. Consider a pair of allele transmissions from the two parents,  $(A_1, A_2)$  and  $(A_3, A_4)$ . We denote a genotype transmission from these parents as  $(A_1/A_3, A_2/A_4)$ , where  $A_1/A_3$  is the transmitted genotype and  $A_2/A_4$  is the non-transmitted genotype (see Figure 15.1).

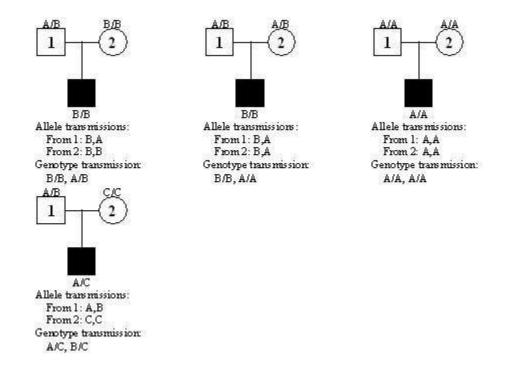


Figure 15.1: Allele and Genotype Transmission Examples

### 15.2.2 Scoring affected offspring

Scoring affected offspring requires computing the allele or genotype transmissions from the parents of an affected individual. However, not all such transmissions are informative and, in the presence of missing parental data, some transmissions cannot be used due to potential bias introduced by population stratification (Curtis and Sham, 1995). Table 1 represents the scoring function used by TDTEX for affected offspring. The basic distinct patterns of allele configurations for parents and children are shown, together with the resulting allele and genotype transmissions. All possible configurations can be obtained from these by relabeling alleles, permuting the two parents, or permuting the alleles within individuals.

Empty cells in the table represent uninformative or unusable transmissions. Notice that some information on allele transmission can be obtained from affected individuals with only one typed parent.

### 15.2.3 Scoring affected sibling pairs

In some situations it is advantageous to test for linkage disequilibrium in data sets consisting of pairs of affected offspring and their parents (Spielman et al., 1993; Olson et al., 1997). This variant of the TDT scores only the same allele transmissions to both affected offspring. This is a narrower sampling scheme than the standard affected offspring version, because transmissions from heterozygous parents that transmit a different allele to each offspring are ignored. In some situations, sampling affected sib pairs rather than affected individuals greatly improves the power of the TDT (see Figure 15.2).

Parent		Parent		Child	Parent 1	Parent 2	Genotype
1		2			transmis-	transmis-	transmis-
					sion	sion	sion
A/A	X	A/A	$\rightarrow$	A/A			
A/A	X	A/B	$\rightarrow$	A/A	A,A	A,B	A/A, A/B
A/A	X	A/B	$\rightarrow$	A/B	A,A	B,A	A/B, A/A
A/A	X	B/B	$\rightarrow$	A/B	A,A	B,B	A/B, A/B
A/A	Х	B/C	$\rightarrow$	A/B	A,A	B,C	A/B, A/C
A/B	Х	A/B	$\rightarrow$	A/A	A,B	A,B	A/A, B/B
A/B	Х	A/B	$\rightarrow$	A/B			
A/B	Х	A/C	$\rightarrow$	A/A	A,B	A,C	A/A, B/C
A/B	X	A/C	$\rightarrow$	A/B	B,A	A,C	A/B, A/C
A/B	X	A/C	$\rightarrow$	B/C	B,A	C,A	B/C, A/A
A/B	X	B/C	$\rightarrow$	A/B	A,B	B,C	A/B, B/C
A/B	X	C/C	$\rightarrow$	A/C	A,B	C,C	A/C, B/C
?/?	X	A/A	$\rightarrow$	A/A			
?/?	X	A/A	$\rightarrow$	A/B		A,A	
?/?	Х	A/B	$\rightarrow$	A/A			
?/?	Х	A/B	$\rightarrow$	A/B			
?/?	Х	A/B	$\rightarrow$	A/C		A,B	
?/?	Х	?/?	$\rightarrow$	A/A			
?/?	X	?/?	$\rightarrow$	A/B			

Table 15.1: Transmission scores for all possible distinct configurations of parents and offspring.

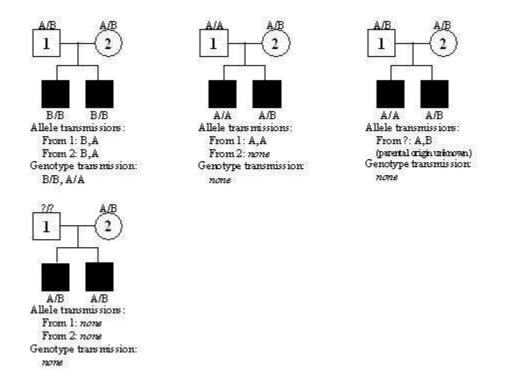


Figure 15.2: Allele and Genotype Transmission to Individual Offspring

Table 2 represents the scoring function used by TDTEX for affected sibling pairs. The basic possible allele configurations for parents and children are shown, together with the resulting allele and genotype transmissions. Empty cells in the table represent uninformative or unusable transmissions. Notice that some information on allele transmission can be obtained from affected pairs with only one typed parent.

Parent		Parent		Child 1	Child 2	Parent	Parent	Genotype
1		2				1 allele	2 allele	transmis-
						transmis-	transmis-	sion
						sion	sion	
A/A	X	A/A	$\rightarrow$	A/A	A/A			
A/A	Х	A/B	$\rightarrow$	A/A	A/A	A,A	A,B	A/A, A/B
A/A	х	A/B	$\rightarrow$	A/A	A/B	A,A		
A/A	х	A/B	$\rightarrow$	A/B	A/B	A,A	B,A	A/B, A/A
A/A	х	B/B	$\rightarrow$	A/B	A/B	A,A	B,B	A/B, B/B
A/A	Х	B/C	$\rightarrow$	A/B	A/B	A,A	B,C	A/B, A/C
A/A	х	B/C	$\rightarrow$	A/B	A/C	A,A		
A/A	X	B/C	$\rightarrow$	A/C	A/C	A,A	C,B	A/C, A/B
A/B	X	A/B	$\rightarrow$	A/A	A/A	A,B	A,B	A/A, B/B
A/B	X	A/B	$\rightarrow$	A/A	A/B	А,	B*	
A/B	X	A/B	$\rightarrow$	A/A	B/B			
A/B	х	A/B	$\rightarrow$	A/B	A/B			
A/B	Х	A/C	$\rightarrow$	A/A	A/A	A,B	A,C	A/A, B/C
A/B	X	A/C	$\rightarrow$	A/A	A/B		A,C	
A/B	х	A/C	$\rightarrow$	A/A	B/C			
A/B	Х	A/C	$\rightarrow$	A/B	A/B	B,A	A,C	A/B, A/C
A/B	Х	A/C	$\rightarrow$	A/B	A/C			
A/B	x	A/C	$\rightarrow$	A/B	B/C	B,A		
A/B	X	A/C	$\rightarrow$	A/C	A/C	A,B	C,A	A/C, A/B
A/B	Х	A/C	$\rightarrow$	A/C	B/C		C,A	
A/B	Х	A/C	$\rightarrow$	B/C	B/C	B,A	C,A	B/C, A/A
A/B	Х	C/D	$\rightarrow$	A/C	A/C	A,B	C,D	A/C, B/D
A/B	X	C/D	$\rightarrow$	A/C	A/D	A,B		
A/B	X	C/D	$\rightarrow$	A/C	B/C		C,D	
A/B	X	C/D	$\rightarrow$	A/C	B/D			
?/?	X	A/A	$\rightarrow$	A/A	A/A			
?/?	X	A/A	$\rightarrow$	A/B	A/B		A,A	
?/?	X	A/B	$\rightarrow$	A/A	A/A			
?/?	X	A/B	$\rightarrow$	A/A	A/B			
?/?	x	A/B	$\rightarrow$	A/B	A/B			
?/?	x	A/B	$\rightarrow$	A/B	B/B			
?/?	X	A/B	$\rightarrow$	A/C	A/C		A,B	
?/?	X	A/B	$\rightarrow$	A/C	B/C			

Table 15.2: Transmission scores for all possible distinct configurations of parents and two offspring
receiving the same transmission

\* - parental origin is unknown

### 15.2.4 Transmission Tables

To test for differences between the distribution of transmitted alleles and genotypes and non-transmitted alleles and genotypes, TDTEX tabulates all the pairs of transmissions and non-transmissions into contingency tables, henceforth called "transmission tables".

Let  $M_1..M_K$  represent the K alleles or genotypes at a given marker locus. Transmission tables are defined to be K x K tables of counts, where the rows represent transmitted alleles or genotypes, and columns are the non-transmitted alleles or genotypes (Table 15.3). The entries  $n_{ij}$  are the number of times  $M_i$  was transmitted and  $M_j$  was not transmitted to an affected individual/pair.

The diagonal elements of the table, (when scoring allele transmissions, those from homozygous parents) contain no information and are ignored in the analysis.

Non-transmitted					
Transmitted	$\mathbf{M}_1$	$M_2$		$\mathbf{M}_{\mathbf{K}}$	Total
$\mathbf{M}_{1}$	$n_{12}$	$n_{12}$		$n_{1K}$	$n_{1\bullet}$
$\mathbf{M_2}$	$n_{21}$	$n_{22}$		$n_{2K}$	$n_{2\bullet}$
:	÷	÷	·	÷	÷
$\mathbf{M}_{\mathbf{K}}$	$n_{K1}$	$n_{K2}$		$n_{KK}$	$n_{K\bullet}$
Total	$n_{\bullet 1}$	$n_{\bullet 2}$	•••	$n_{\bullet k}$	$n_{\bullet\bullet}$

Table 15.3: The structure of a transmission table

### 15.2.5 Pedigree sampler

The pedigree sampler is the component of TDTEX that controls the construction of transmission tables. It traverses the pedigree data, identifies potentially informative individuals and pairs based on trait and marker data, scores them, and tabulates the results into a transmission table. For each nuclear family considered, the sampler first attempts to find any *informative* affected sibling pairs, up to a user-specified maximum number. This maximum can be set to zero to disable the sampling of affected sibling pairs, or to an unlimited value to select as many as possible. The sampler will only allow each child to participate in at most one transmission, so there is no problem with overlapping affected sibling pairs. The remaining offspring not already used in a sibling pair are then scored, up to a separate user-specified maximum number. This maximum can also be set to zero to disable the sampling of affected sibling pairs, or to an unlimited value to select as many as possible.

The traditional TDT test corresponds to setting the maximum number of affected children per nuclear family to 1 and the maximum number of affected sibling pairs to none. The sampler will then score the first informative allele or genotype transmission to an affected offspring, and then move on to score the next nuclear family. This will result in a valid test of allelic association in the presence of linkage.

Some other implementations of the TDT test work by setting the maximum number of affected children per nuclear family to unlimited, and the number of affected sibling pairs to none. This allows the sampler to score all informative affected offspring in each nuclear family. Similarly, basic TDT tests utilizing only sibling pairs are possible by setting the maximum number of affected

offspring to none, and the maximum number of affected sibling pairs to 1 or unlimited. This will result in valid tests of linkage in the presence of allelic association.

An interesting option exists to enable the sampling of both affected sibling offspring *and* affected sibling pairs. This very general variation gives preference to informative affected sibling pairs over affected offspring. Overall, this configuration provides a way to take advantage of more information from data sets that include a mixture of family types, not all of which have two affected offspring. Equal weight is given to all transmissions, so power may not be optimal in spite of the larger sample size.

### 15.2.6 Testing significance of transmission tables

Two null hypotheses have been proposed to test transmission tables for deviations from the expected pattern of allele and genotype transmissions. The first hypothesis is that of complete symmetry between the transmitted and non-transmitted alleles. This states that the expected number of any transmission type is equal to the expected number of observed transmission of the opposite pattern, i.e.,  $E(n_{ij}) = E(n_{ji})$ . The second hypothesis is the hypothesis of marginal homogeneity: in this case, the number of alleles or genotypes transmitted is compared to the number not transmitted, i.e.,  $E(n_{i\bullet}) = E(n_{\bullet j})$ . Which null hypothesis is optimal depends on the sample size, number and distribution of alleles, and the structure of the disequilibrium present in the sample. TDTEX provides tests based on both hypotheses for maximum flexibility.

TDTEX also includes both exact and asymptotic tests. Exact tests, as the name suggests, provide exact significance levels at the expense of being computationally intensive. Asymptotic tests are based on distributional theory and approximations that are only precise for very large sample sizes. They tend to be very quick to compute, but there are situations when asymptotic tests are significantly less powerful than exact versions. Typically, this occurs when sample sizes are small, transmission tables are sparse, and cells have less than 5 observations.

Statistics based on both the hypotheses of complete symmetry and marginal homogeneity may be applied to tables of allele transmissions as well as genotype transmissions. Genotype transmission tables may be preferred because the transmission patterns of the two parents, which include transmission from the homozygous parents, are not independent in the multiallelic case, except when linkage is complete (Bickeböller and Clerget-Darpoux, 1995). However, because of the larger size and increased sparseness of genotype transmission tables for markers with multiple alleles, the marginal homogeneity test is less prone than the complete symmetry test to problems arising from table sparseness.

### 15.2.6.1 Asymptotic Tests

Under the hypothesis of complete symmetry, the NcNemar test statistic

$$T_{mc} = \sum_{i < j} \frac{(n_{ij} - n_{ji})^2}{n_{ij} + n_{ji}} \sim \chi^2_{K(K-1)/2}$$

has an asymptotically  $\chi^2$  distribution with K(K - 1)/2 degrees of freedom (Bickeböller and Clerget-Darpoux, 1995). In practice, the number of degrees of freedom equals the number of types of parental heterozygotes in the sample. A continuity corrected version of the NcNemar test statistic

$$T_{mcc} = \sum_{i < j} \frac{(|n_{ij} - n_{ji}| - 1)^2}{n_{ij} + n_{ji}} \sim \chi^2_{K(K-1)/2}$$

is also provided, since it tends to be more robust to small sample sizes.

Under the hypothesis of marginal homogeneity, the test statistic

$$T_{mh} = \frac{K-1}{K} \sum_{i} \frac{(n_{i.} - n_{.i})^2}{n_{i.} + n_{.i} - 2n_{ii}} \sim \chi^2_{K-1}$$

has an asymptotically  $\chi^2$  distribution with K-1 degrees of freedom, provided the table margins are independent of each other (Spielman and Ewens, 1996).

#### 15.2.6.2 Exact tests

The exact test of complete symmetry or marginal homogeneity is generally a more powerful test than the asymptotic tests in the presence of table sparseness and/or a small sample size. To obtain the null permutation distribution for the exact test, we write the distribution of the  $n_{ij}$ , conditional on the sums of complementary off-diagonal cells, as the product of K(K-1)/2 binomial random variables with equal probability of transmission vs. non-transmission:

$$Pr(n) = \prod_{i < j} \begin{pmatrix} n_{ij} + n_{ji} \\ n_{ij} \end{pmatrix} \begin{pmatrix} \frac{1}{2} \end{pmatrix}^{n_{ij} + n_{ji}}$$

An exact significance level is determined by calculating the probability of finding a permutation of the observed data, conditional on the sums of complementary off-diagonal cells, which is as extreme as, or more extreme than, the observed transmission table. Let  $N = \{n' : n'_{ij} + n'_{ji} = n_{ij} + n_{ji}\}$  be the set of all permutations of the observed data, conditional on the sums of complementary off-diagonal cells. Let  $N' = \{n' : Pr(n') \leq Pr(n), n' \in N\}$ , be the set of all permutations with probability less than or equal to that of the observed data. Then the significance level, or p-value, is  $P_{cs} = \sum_{n \in N'} Pr(n')$ .

Since enumerating all possible permutations of the observed transmission table is infeasible for larger tables, the exact permutation algorithm relies upon methods of ordering permutations of the observed table, and by avoiding the evaluation of many equivalent tables. The algorithm uses the fact that the probability after permuting a pair of symmetric odd-diagonal cells in a transmission table does not involve the remaining cells. The null probability distribution is also independent of the direction of asymmetry. For example, a configuration in which  $n_{12} = 4$  and  $n_{21} = 0$  has the same probability as that of  $n_{12} = 0$  and  $n_{22} = 4$ .

### 15.2.6.3 Monte Carlo Approximations

As the transmission table size and number of marker alleles increases, program execution time of the exact permutation test becomes prohibitively slow. For transmission tables with greater than about 300 observations, or with more than about 8 alleles, the Monte Carlo approximation is recommended. Instead of considering every possible permutation, a random sample from the set of all possible permutations, conditional on the observed transmission table, is taken.

The proportion of permutations with significance equal to or greater than the observed table is computed. This proportion is an estimate of the exact pvalue of the observed table. The standard error of the estimated p-value is obtained by computing the variance among several batches of permutations. The total number of permutations considered is chosen to estimate the resulting p-value within 20% of its true value with 95% confidence.

# 15.3 Program Input

File Type	Description
Parameter file	Specifies the parameters and options with which
	to perform a particular analysis.
Pedigree data file	Contains delimited records for each individual in-
	cluding fields for identifiers, sex, parents, trait and
	marker data.

## 15.3.1 The tdtex Parameter

The following syntax table specifies the permissible parameter and attribute settings for the main
TDTEX parameter.

<pre>parameter [, attribute]</pre>	Explanation				
	Starts a TDTEX sp	ecification block.			
	Value Range	N/A			
tdtex	Default Value	None			
	Required	Yes			
	Applicable Notes	None			
	Specifies the root name to be used for				
	Output file names will be formed by concatenating the				
	root name and an appropriate extension.				
		Character string representing a			
, out	Value Range	valid file name.			
	Default Value	tdtex			
	Required	No			
	Applicable Notes	None			

# 15.3.2 The tdtex Block

The following syntax table lists the permissible parameter and attribute settings for the tdtex block.

parameter		Explanation			
[, attribute]	Specifies a marker for which transmissions are scored.				
marker	Value Range	Character string representing the name of a valid marker listed in the pedigree data file.			
	Default Value Required Applicable Notes	None       Yes       4			
	Specifies a trait der and sibling pairs.	noting affection status for offspring			
trait	Value Range	Character string representing the name of a valid trait, phenotype or covariate listed in the pedigree data file.			
	Default Value Required	None       Yes       None			
	Applicable Notes				
	Specifies a trait used as an indicator variable to select subsets of pairs to analyze.				
		Character string representing the			
parental_trait	Value Range	name of a valid trait, phenotype or covariate listed in the			
	Default Value Required	pedigree data file. None No			
	Applicable Notes	None			
sample	Value Range Default Value Required Applicable Notes	e of transmission is to be scored. {alleles, genotypes} alleles No None			
	Specifies the max	imum number of informative af- ansmissions per nuclear family that se.			
max_children	Value Range	none unlimited {0, 1, 2, 3,}			
	Default Value Required Applicable Notes	None           No           1, 3			
	, , , , , , , , , , , , , , , , , , , ,	-, -			

	Specifies the maximum number of informative af- fected sibling pair transmissions per nuclear family that the sampler may use.
max_sib_pairs	None Value Range unlimited
	Default Value $\frac{\{0, 1, 2, 3,\}}{\text{None}}$
	Required No
	Applicable Notes 2, 3
	Causes three tests to be performed:
	1. one scoring transmissions from all parents,
	2. one that scores only paternal transmissions, and
sex_differential	3. one that scores only maternal transmissions.
	Value Range {true, false}
	Default Value false
	Required No
	Applicable Notes None
	Specifies that no exact tests are to be performed.
	· ·
	-
skip_exact_tests	
	Default Value true
	Required No
	Value Range {true, false}
skip_permutation_test	Default Value true
	Required No
	Applicable Notes None
	Specifies that the exact Monte Carlo McNemar test
	should not be performed.
	Value Range {true, false}
skip_mc_test	Required       No         Applicable Notes       None         Specifies that no exact tests are to be performed         This option is shorthand for setting all three         of the parameters:       skip_permutation_test         skip_mc_test and skip_mcmh_test.         Value Range       {true, false}         Default Value       true         Required       No         Applicable Notes       None         Specifies that the exact permutation McNemar test         should not be performed.         Value Range       {true, false}         Default Value       true         Required       No         Applicable Notes       None         Specifies that the exact permutation McNemar test         should not be performed.         Value Range       {true, false}         Default Value       true         Required       No         Applicable Notes       None         Specifies that the exact Monte Carlo McNemar test         should not be performed.       Value Range         Value Range       {true, false}         Default Value       true         Required       No
	Required No
	Applicable Notes None
	Specifies that the exact Monte Carlo marginal homo-
	geneity test should not be performed.
	Value Range {true, false}
skip_mcmh_test	Default Value true
	Required No
	Applicable Notes None

- 1. A classic TDT can be performed by setting max\_children to 1 and max\_sib\_pairs to **none**. All affected children in a pedigree can be used as if they are independent by setting max\_children to **unlimited** and max\_sib\_pairs to **none**.
- 2. A sibling TDT using only one sib-pair per pedigree can be performed by setting max\_children to **none** and max\_sib\_pairs to 1. A sibling TDT using all sibling-pairs in a pedigree as if they are independent can be performed by setting max\_children to **none** and max\_sib\_pairs to **unlimited**.
- 3. Regardless of the values of max\_children and max\_sib\_pairs, pedigrees must have at least one typed parent.
- 4. The user may list as many different marker parameters as desired. If no marker parameters are specified, then the default TDTEX behavior is to score transmissions for all markers found in the pedigree data file.

# **15.4 Program Execution**

TDTEX is run via a command line interface on the supported UNIX and Windows platforms. This requires the S.A.G.E. programs to be properly installed and in the current execution path. Input files are specified on the command line and all output files are created in the current working directory.

Running TDTEX from the command prompt with no arguments, or the wrong number of arguments, will result in the program printing its usage statement. This lists the input files the program requires on the command line.

```
>tdtex
S.A.G.E. v5.x -- TDTEX
COPYRIGHT (C) 2002 CASE WESTERN RESERVE UNIVERSITY.
usage: tdtex <parameters> <pedigree>
Command line parameters:
parameters - parameter file
pedigree - pedigree data file
```

As indicated in the program usage statement, input files are listed on the command line. A typical run of TDTEX may look like the following:

```
>tdtex tdtex.par example.ped
S.A.G.E. v5.x -- TDTEX
COPYRIGHT (C) 2002 CASE WESTERN RESERVE UNIVERSITY
Reading Parameter File.....done.
Reading Pedigree File.....done.
from example.ped.....done.
Sorting pedigrees.....done.
```

# **15.5 Program Output**

Filename	File Type	Description
tdtex.inf	Information output file	Contains informational diagnostic messages, warnings and program errors. No analysis results are stored in this file.
tdtex.out	TDTEX analysis output file	Contains the results of each TDT analysis.

TDTEX produces several output files that contain results and diagnostic information:

### **15.5.1** Information Output File

The TDTEX Information Output file contains a variety of useful information, including:

- Information on fields read from the pedigree data file. These tables, which provide information about what the program has read in, are included with all programs in the S.A.G.E. and are very useful for debugging most common errors caused when reading the pedigree data. When first analyzing new data, it is recommended that these tables be checked carefully to make sure pedigree data are being correctly read.
- Information, warning and error messages generated throughout the program. It is recommended that you check this file for warning and error messages before examining the results of any run of the program. The program attempts to correct many common errors and this sometimes means analyses are not run as expected. The file "tdtex.inf" should be checked for errors and diagnostic information after each run of the program.

### 15.5.2 TDTEX Analysis Output File

One analysis output file, named "tdtex.out", is generated per run of TDTEX. It contains the results of all tests.

## **15.6 Example Output Files**

```
S.A.G.E. v5.x -- TDTEX
COPYRIGHT (C) 2002 CASE WESTERN RESERVE UNIVERSITY
  Allele transmissions to affected children:
   _____
  NOTICE: 15 error(s) were found in your sample! See the Information Output
            File for a detailed description of each problem.
  Marker: mrk1
  Trait: Trait1
  Max. children/family: 1
  Max. sib pairs/family: none
  [13 empty rows/columns not shown]
  T \setminus NT | A | C | F | G |
   ----|----
              ____
      A 0 1 1 1
      C 0 0 0
F 0 0 0
                        0
                       0
      G 0 0 0 0
                Informative / Total = % Informative

        Pedigrees
        :
        1 /
        2 =
        50.00%

        Families
        :
        3 /
        10 =
        30.00%

        Affected children
        :
        3 /
        3 =
        100.00%

  Affected sib pairs : 0 / 0 = 0.00%
  Sample size : 3 / 3 = 100.00%
                                               e-value (std. error)
      p-value
  Exact test statistics

        Exact McNemar test
        1.00000000

        Monte Carlo McNemar test
        1.00000000 (0.0000000)

        Monte Carlo Marginal Homogeneity
        0.25012750 (0.00111201)

                                           p-value
  Asymptotic test statistics
   -----
                                              _____
                                             0.39162518
  McNemar test
  Continuity corrected McNemar test 1.00000000
  Marginal homogeneity test
                                             0.21229020
```

# **Chapter 16**

# AGEON

This program produces maximum likelihood estimates of the parameters of a mixed power-normal distribution for a binary trait (affect versus unaffected) with variable age of onset. These estimates can be used in a special function to produce either of two new variables, for a binary trait with variable age of onset, to be used in a SIBPAL analysis. The mean, variance, and susceptibility can each depend linearly on covariates. By default a class susceptibility covariate is generated according to the value of the parental binary trait(s).

# 16.1 Limitations

No account is taken of ascertainment or familial correlations; i.e. all individuals are assumed to be randomly sampled. This does not affect the validity or robustness of any SIBPAL analysis. Genetic susceptibilities are not estimated for classes with fewer than 5 informative members. If for any reason the power parameter  $\lambda_1$  is fixed by the user at 0, then the initial value of the shift parameter  $\lambda_2$  *must* be greater than the inverse of the minimum value for age-of-onset or age-at-exam, whichever is smaller.

# 16.2 Theory

The purpose of the AGEON program is to estimate the parameters needed to calculate either of two new quantitative variable that can be used in SIBPAL. These variables are to detect linkage to

- 1. genes that affect susceptibility to disease, and
- 2. genes that affect age of onset of disease.

### 16.2.1 Susceptibility

Susceptibility to disease conditional on whether the individual is affected or not by age  $a'_j$  is given by

$$x_{j} = \begin{cases} 1 & , \text{ if affected} \\ \\ \frac{\gamma_{j} - \gamma_{j} \Phi \left[ \frac{\left( \frac{a_{j}' + \lambda_{2} \right)^{\lambda_{1}}}{\lambda_{1}} - \mu_{j}}{\frac{\sigma_{j}}{\sigma_{j}}} \right]}{1 - \gamma_{j} \Phi \left[ \frac{\left( \frac{a_{j}' + \lambda_{2} \right)^{\lambda_{1}}}{\lambda_{1}} - \mu_{j}}{\sigma_{j}} \right]} & , \text{ if not affected by age } a_{j}'$$

where  $\Phi$  is the standard cumulative normal distribution function.

The disease age of onset is given by the survival analysis residual

$$x_{j} = \begin{cases} 1 - \gamma_{j} \Phi \left[ \frac{\left(a_{j} + \lambda_{2}\right)^{\lambda_{1}}}{\sigma_{j}} - \mu_{j} \right] &, \text{ if affected at age } a_{j} \\ \\ -\gamma_{j} \Phi \left[ \frac{\left(a_{j}' + \lambda_{2}\right)^{\lambda_{1}}}{\sigma_{j}} - \mu_{j} \right] &, \text{ if not affected by age } a_{j}' \end{cases}$$

where again  $\Phi$  is the standard cumulative normal distribution function.

The program AGEON estimates  $\lambda_1$ ,  $\lambda_2$ ,  $\mu$  and  $\gamma$  wherein the parameters are defined, possibly as functions of specified covariates, as follows:

$$\mu_{j} = \mu_{0} + \xi_{1}x_{1j} + \xi_{2}x_{2j} + \cdots,$$

$$\gamma_{j} = \frac{e^{\theta_{j}}}{1 + e^{\theta_{j}}}, \text{ where } \theta_{j} = \gamma_{0} + \xi_{1}x_{1j} + \xi_{2}x_{2j} + \cdots, \text{ and}$$

$$\sigma^{2} = \sigma_{0}^{2} + \xi_{1}x_{1j} + \xi_{2}x_{2j} + \cdots$$

In particular,  $\gamma_0$  can be a function of parental affection status if sufficient parental data are available.

Using one of these values of x as a quantitative trait can be more powerful in the usual Haseman-Elston test for linkage than using disease status as a simple binary trait (Zhu, et al., 1997; Hanson and Knowler, 1998).

The log likelihood maximized is  $\sum_{i}^{n} \ln L(i)$ , where n is the number of sibs in the sample and L(i) is the likelihood for the i-th sib. Let  $a_i$  denote age of onset and  $a'_i$  the age at examination, the latter being available for all unaffected persons to be included in the analysis. Define

$$\begin{split} \varphi\left(x,\sigma\right) &= \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{-\frac{1}{2}\left(\frac{x}{\sigma}\right)^2\right\}\\ \Phi\left(x\right) &= \frac{1}{\sqrt{2\pi}}\int_{-\infty}^x \exp\left\{-\frac{1}{2}u^2\right\} du\\ h\left(x\right) &= \left\{\begin{array}{ll} \frac{(x+\lambda_2)^{\lambda_1}-1}{\lambda_1}, & \text{if } \lambda_1 \neq 0\\ \ln(x+\lambda_2), & \text{if } \lambda_1 = 0 \end{array}\right. \end{split}$$

$$sign(x) = \begin{cases} 1, \text{ if } x \ge 0\\ -1, \text{ if } x < 0 \end{cases}$$

Then the likelihood L(i) is given by:

Category	$L\left(i ight)$
Affected individuals with known age of onset	$L(i) = \gamma_i \varphi \left[ h(a_i) - h(\mu_i), \sigma_i^2 \right] (a_i + \lambda_2)^{\lambda_1 - 1}$
Affected individuals with un- known age of onset	$L^*(i) = \gamma_i \Phi\left[\frac{h(a_i') - h(\mu_i)}{\sigma_i}\right]$
Unaffected individuals	$L(i) = 1 - L^*(i)$

As mentioned above, the mean  $\mu_i$ , variance  $\sigma_i^2$ , and susceptibility  $\gamma_i$ , may depend on covariates. In the case of susceptibility, the logit is assumed to be a linear function of the covariates.

Because  $a_i + \lambda_2$  must be positive to apply the Box and Cox (1964) transformation, so that prior to transformation  $a_i + \lambda_2$  cannot strictly follow a normal distribution, there the maximization is also performed using the following likelihoods, which allow for the truncation (see Pericak-Vance et al, 1983):

Category	$L\left(i ight)$
Affected individuals with known age of onset	$L(i) = \frac{\gamma_i \varphi[h(a_i) - h(\mu_i), \sigma_i^2](a_i + \lambda_2)^{\lambda_1 - 1}}{\Phi\left[sign(\lambda_1)\left(\frac{\mu_i - h(0)}{\sigma_i}\right)\right]}$
Affected individuals with un- known age of onset	$L^*(i) = \frac{\gamma_i sign(\lambda_1) \left[ \Phi\left(\frac{h(a_i') - h(\mu_i)}{\sigma_i}\right) - \Phi\left(\frac{-\frac{1}{\lambda_1} + h(\mu_i)}{\sigma_i}\right) \right]}{\Phi\left[sign(\lambda_1) \left(\frac{\mu_i - h(0)}{\sigma_i}\right)\right]}$
Unaffected individuals	$L\left(i\right) = 1 - L^{*}\left(i\right)$

# 16.3 Program Input

File Type	Description	
Parameter file	Specifies the parameters and options with which	
	to perform a particular analysis.	
Pedigree data File	Contains delimited records for each individual in-	
	cluding fields for identifiers, parents, trait and co-	
	variates.	

### 16.3.1 The ageon Parameter

The following syntax table specifies the permissible parameter and attribute settings for the main ageon parameter.

<pre>parameter [, attribute]</pre>	Explanation	
	Starts an AGEON analysis block	
	Value Range	N/A
ageon	Default Value	None
	Required	Yes
	Applicable Notes	None
	Specifies the root name to be used for output files.	
	Output file names will be formed by concatenating the root name and an appropriate extension.	
out		Character string representing a
, out	Value Range	valid file name.
	Default Value	ageon
	Required	No
	Applicable Notes	None

# 16.3.2 The ageon Parameter Block

The following table s	shows the syntax f	or the ageon block:
		J

parameter	Evaluation		
[, attribute]	Explanation		
	Specifies a title for the analysis defined within the block.		
title	Value Range Quoted character string		
LILIE	Default Value "AGEON Analysis"		
	Required No		
	Applicable Notes None		
	Specifies name of a binary trait containing affection status. Must be the name of a trait, covariate or phe- notype in the pedigree data file or created by means of		
	a function block.		
affected, affectedness	Value Range       Character string representing a valid file name.		
	Default Value None		
	Required Yes		
	Applicable Notes None		
	Specifies name of a trait containing age of onset for		
	affected individuals. Must be the name of a trait, co-		
	variate or phenotype in the pedigree data file or cre-		
	ated by means of a function block.		
age_of_onset	Value Range Character string		
	Default Value None		
	Required Yes		
	Applicable Notes 1		
	Specifies name of a trait containing age of examina-		
	tion for unaffected individuals. Must be the name of		
	a trait, covariate or phenotype in the pedigree data file		
<u> </u>	or created by means of a function block.		
age_of_exam	Value Range Character string		
	Default Value None		
	Required Yes		
	Applicable Notes 1		
	Specifies option to substitute covariate mean values		
	for missing covariate data		
	Value Range {mean, none}		
allow_averaging	Default Value none		
	Required No		
	Applicable Notes 2		

	Starts a sub-block for specifying covariates for the		
	mean.		
	Value Range N/A		
mean_cov	Default Value None		
	Required No		
	Applicable Notes None		
	Starts a sub-block for specifying covariates for the		
	variance.		
var cov	Value Range N/A		
val_cov	Default Value None		
	Required No		
	Applicable Notes None		
	Starts a sub-block for specifying covariates for the		
	trait susceptibility.		
suscept_cov	Value Range N/A		
suscept_cov	Default Value None		
	Required No		
	Applicable Notes None		
	Starts a sub-block for specifying transformation op-		
transformation	tions.		
	Value Range N/A		
	Default Value None		
	Required No		
	Applicable Notes 3, 4		

Notes:

- 1. The age\_of\_onset parameter and the age\_of\_exam parameter can both specify the same pedigree data field.
- 2. If the value of **none** is specified and any single individual's covariate value is missing, then that individual will be treated as uninformative for the purpose of the analysis. If **mean** is specified, missing covariate values will be replaced with the covariate's mean value as calculated from the sample.
- 3. See section 5.3.2.2 for details on the transformation theory implemented in this program.

### 16.3.3 Sub-Block Syntax: mean\_cov

parameter	Explanation		
[, attribute]			
	Covariate to modify the mean value of the age of		
	onset. This param	eter may be specified multiple	
	times.		
		Character string representing the	
		name of a trait, covariate or phe-	
covariate	Value Range	notype from the pedigree data	
		file or a name created by means	
		of a function block.	
	Default Value	None	
	Required	No	
	Applicable Notes	1	
	Specifies the value	of the covariate coefficient.	
	Value Range	$(-\infty, +\infty)$	
, val	Default Value	None	
	Required	No	
	Applicable Notes	None	
	Specifies option to	fix the given value.	
	Value Range	{true, false}	
, fixed	Default Value	false	
	Required	No	
	Applicable Notes	2	

The following table shows the syntax for the mean\_cov sub-block:

Notes:

- 1. The default is to include no mean covariates in the analysis. The means indicated in the type\_mean sub-block are a linear function of this covariate. All covariates are centered, the centering (average) value being included as part of the output.
- 2. The fixed attribute, used in conjunction with the val attribute, specifies that the covariate's value will remain fixed at the given value and will not be estimated by the program.

### 16.3.4 Sub-Block Syntax: var\_cov

parameter	Explanation		
[, attribute]	-		
	Covariate to modify the variance of the trans-		
	formed age of onse	et. This parameter may be spec-	
	ified multiple time	s.	
		Character string representing the	
		name of a trait, covariate or phe-	
covariate	Value Range	notype from the pedigree data	
		file or a name created by means	
		of a function block.	
	Default Value	None	
	Required	No	
	Applicable Notes	1	
	Specifies value of	the covariate coefficient	
	Value Range	$(-\infty, +\infty)$	
, val	Default Value	None	
	Required	No	
	Applicable Notes	None	
	Specifies option to	fix the given value.	
	Value Range	{true, false}	
, fixed	Default Value	false	
	Required	No	
	Applicable Notes	2	

The following table shows the syntax for the var\_cov sub-block:

Notes

- 1. The default is to include no covariates in the analysis. The variances indicated in the type\_var sub-block are a linear function of this covariate . All covariates are centered, the centering (average) value being included as part of the output.
- 2. The fixed attribute, used in conjunction with the val attribute, specifies that the covariate's value will remain fixed at the given value and will not be estimated by the program.

### 16.3.5 Sub-Block Syntax: suscept\_cov

The following table shows	the syntax	for the suscept	cov sub-block:
The following there blows	cite by fitteri	Ior the babeepe_	_000 010010

parameter	Explanation		
[, attribute]			
	Covariate to modify the mean of a continuous trait. This parameter may be specified multiple		
	times.		
covariate	Value Range	Character string representing the name of a trait, covariate or phe- notype from the pedigree data file or a name created by means of a function block.	
	Default Value	None	
	Required	No	
	Applicable Notes	1	
ĺ	Specifies the value	of the covariate coefficient.	
	Value Range	$(-\infty,\infty)$	
, val	Default Value	None	
	Required	No	
	Applicable Notes	None	
	Specifies option to	fix this value.	
	Value Range	{true, false}	
, fixed	Default Value	false	
	Required	No	
	Applicable Notes	2	
		c for specifying classification op-	
	tions.		
class	Value Range	N/A	
	Default Value	None	
	Required	No	
	Applicable Notes	None	

Notes

- 1. The default is to include no susceptibility covariates in the analysis. The suscept\_cov sub-block indicates which covariates are to modify the logits of susceptibilities. All covariates are centered, the centering (average) value being included as part of the output.
- 2. The fixed attribute, used in conjunction with the val attribute, specifies that the covariate's value will remain fixed at the given value and will not be estimated by the program.

# 16.3.6 Sub-Block Syntax: transformation

The following table shows the syntax for the transformation sub-block:

parameter		
[, attribute]	Explanation	
	Specifies of the po	wer parameter, $\lambda_1$
	Value Range	N/A
lambda1	Default Value	N/A
	Required	No
	Applicable Notes	None
	Specifies the value	e for $\lambda_1$ .
	Value Range	$(-\infty, +\infty)$
, val	Default Value	1.0
	Required	No
	Applicable Notes	None
	Specifies option to	fix $\lambda_1$ at the given value.
	Value Range	{true, false}
, fixed	Default Value	false
	Required	No
	Applicable Notes	None
	Specifies the shift	parameter, $\lambda_2$
	Value Range	N/A
lambda2	Default Value	N/A
	Required	No
	Applicable Notes	None
	Specifies the value	e for $\lambda_2$ .
	Value Range	$(-\infty, +\infty)$
	Default Value	0.05
, val	Required	No
	Applicable Notes	2
		None
	Option to fix $\lambda_2$ at	the given value.
	Value Range	{true, false}
, fixed	Default Value	true
	Required	No
	Applicable Notes	None

## 16.3.7 class sub-block

The following lists all parameters that may occur in a class sub-block.

parameter [, attribute]	Explanation	
trait	Specifies the name of an alternate trait on which	
	to base the individual's classification code.	
	Value Range	Character string
	Default Value	None
	Required	No
	Applicable Notes	None
num_of_classes	Specifies the number of classes to be considered	
	during the analysis	
	Value Range	{1, 2, 3,}
	Default Value	6
	Required	No
	Applicable Notes	None

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### 16.3.8 Sample parameter file

```
pedigree {
   delimiter_mode = multiple
   delimiters = " "
   individual_missing_value = "0"
   sex_code, male = "1", female = "2", unknown = "?"
   pedigree_id = fam
   individual id = id
   parent_id = mom
   parent id = dad
   sex field = sex
   trait = aff, binary, affected = 1, unaffected = 0, missing = -1
   trait = ao, missing = -1
   trait = ae, missing = -1
   trait = cov1, missing = -999
   trait = cov2, missing = -999
   trait = classx
}
ageon {
   title = "analysis"
   affectedness = aff
   age_of_onset = ao
   age_of_exam = ae
   mean cov {
      covariate = cov2
   }
   suscept_cov {
      covariate = cov1
      class {
         trait = classx
         num_of_classes = 10
      }
   }
}
```

### 16.3.9 Exclusion Criteria for Individuals and Pedigrees

Under some conditions AGEON will exclude individuals and/or pedigrees from analysis, and the user is advised to take note of program outputs that indicate the numbers of valid and invalid individuals being counted. The exclusion criteria are:

- 1. Any structurally invalid families will be excluded from analysis (see 16.1).
- 2. Any individual whose primary trait is missing or who is missing at least one covariate value (if the allow\_averaging option is disabled the default behavior) will be retained for

analysis, but all of its trait information will be treated as missing.

### **16.3.10 Program Execution**

AGEON is run via a command line interface on the supported UNIX and Windows platforms. This requires that the S.A.G.E. programs are properly installed and in the current execution path. Input files are specified on the command line and all output files are created in the current working directory.

Running AGEON from the command prompt with no arguments, or the wrong number of arguments, will result in the program printing its usage statement. This lists the input files the program requires on the command line:

```
>ageon
S.A.G.E. v5.x -- AGEON
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
usage: ./ageon <parameters> <pedigree>
Command line parameters:
parameters - Parameter File
pedigree - Pedigree Data File
As indicated in the program usage statement, input files are
listed on the command line. A typical run of AGEON may look like
the following:
>ageon ageon.par example.ped
S.A.G.E. v5.x -- AGEON
COPYRIGHT (C) CASE WESTERN RESERVE UNIVERSITY.
Reading parameter file.....done.
Reading pedigree file.....
from example.ped.....done.
Sorting pedigrees.....done.
Parsing Ageonset Analyses...
Beginning new analysis block....
Analysis parsing complete. Analysis valid.
_____
Importing individual data into AGEON...
Import complete.
_____
```

### 16.3.11 Program Output

File Name	File Type	Description	
ageon.inf	Information output file	Contains informational diagnostic messages, warnings and program errors. No analysis results are stored in this file.	
ageon.out	AGEON summary output file	Contains the table of final estimates of the param- eters and their standard errors and other results.	

Output files produced by AGEON containing results and diagnostic information are:

### 16.3.11.1 Information Output File

The AGEON Information file contains a variety of useful information, including:

- Information on fields read from the Pedigree Data File. These tables, which provide information about what the program has read from the Pedigree Data File, are included with all programs in S.A.G.E. Release and are very useful for debugging most common errors caused when reading the pedigree data. When first analyzing new data, it is recommended that these tables be checked carefully to make sure pedigree data are being correctly read.
- Information, warning and error messages generated throughout the program. It is recommended that this file be checked for warning and error messages before examining the results of any run of the program. The program attempts to correct many common errors and this sometimes means analyses are not as expected. The file "ageon.inf" should be checked for errors and diagnostic information after each run of the program.

### 16.3.12 Output Files

The AGEON Output Files present the tables of final estimates of the parameters, along with relevant model information. By default a class susceptibility covariate is generated according to the value of the parental binary trait(s), as shown in the following table:

Class	Description
0	Both parents are unknown
1	One of the parents is unknown, the other is affected
2	One of the parents is unknown, the other is unaffected
3	Both parents are affected
4	One of the parents is affected, the other is unaffected
5	Both parents are unaffected

#### 16.3.12.1 Example Summary Output File

The summary output file contains descriptive information about each of the six classifications, as well as final estimates, standard errors, and p-values of the parameters estimated in the model, including

- mean intercept and covariates
- susceptibility intercept(s) and covariates
- variance intercept and covariates
- transformation parameters ( $\lambda_1$  and  $\lambda_2$ ).

The file also includes a likelihood ratio test statistic for the comparison of separate susceptibilities for each category of the classification variable against a common susceptibility for all categories.

Here is an example of an AGEON summary output file:

```
_____
   Sample description
-----
 Number of pedigrees in dataset
                                    200
 Number of analyzable pedigrees
                                    200
                                    787
 Number of individuals in dataset
 Number of analyzable individuals
                                    787
 Number of analyzable invalid individuals
                                    149
 Number of analyzable valid individuals
                                    638
------
   MODEL DESCRIPTION
Title
        AGEONSET Analysis 1
 Affectedness trait AFF
 Age-of-onset trait AO
 Age-at-exam trait AE
_____
    CLASSIFICATION SYSTEM
_____
 Using default classification system:
  0 Both parental traits are unknown.
  1 One of the parental traits is unknown, the other is affected.
  2 One of the parental traits is unknown, the other is unaffected.
  3 Both parental traits are affected.
  4 One of the parental traits is affected, the other is unaffected.
  5 Both parental traits are unaffected.
_____
   CLASS STATISTICS
_____
_____
    CLASS 0
_____
 TOTAL NUMBER OF INDIVIDUALS USED IN ANALYSIS
                                             60
 NUMBER OF INDIVIDUALS WITH AN AGE OF ONSET
                                             25
 MEAN OF AGE OF ONSET
                                       75.458537
```

VARIANCE OF AGE OF ONSET 4.872671 NUMBER OF INDIVIDUALS AFFECTED 9 PROPORTION OF INDIVIDUALS AFFECTED 0.150000 MEAN OF AGE AT EXAM OF THE UNAFFECTED 74.816667 VARIANCE OF AGE AT EXAM OF THE UNAFFECTED 6.949722 \_\_\_\_\_ CLASS 1 \_\_\_\_\_ TOTAL NUMBER OF INDIVIDUALS USED IN ANALYSIS 80 NUMBER OF INDIVIDUALS WITH AN AGE OF ONSET 33 MEAN OF AGE OF ONSET 74.696970 VARIANCE OF AGE OF ONSET 6.514233 NUMBER OF INDIVIDUALS AFFECTED 20 PROPORTION OF INDIVIDUALS AFFECTED 0.250000 MEAN OF AGE AT EXAM OF THE UNAFFECTED75.350000VARIANCE OF AGE AT EXAM OF THE UNAFFECTED4.377500 \_\_\_\_\_ CLASS 5 TOTAL NUMBER OF INDIVIDUALS USED IN ANALYSIS 43 NUMBER OF INDIVIDUALS WITH AN AGE OF ONSET 15 MEAN OF AGE OF ONSET 76.266667 VARIANCE OF AGE OF ONSET 4.595556 NUMBER OF INDIVIDUALS AFFECTED 9 PROPORTION OF INDIVIDUALS AFFECTED 0.209302 MEAN OF AGE AT EXAM OF THE UNAFFECTED 75.744186 VARIANCE OF AGE AT EXAM OF THE UNAFFECTED 5.446187 MAXIMIZATION RESULTS susceptibilities equal, no truncation \_\_\_\_\_ \_\_\_\_\_ P-value Estimate S.E. Parameter \_\_\_\_\_ Susceptibility intercepts Class 0 -0.489288 0.162019 0.002528 -0.489288 0.162019 0.002528 Class 1 -0.489288 0.162019 0.002528 Class 2 -0.489288 0.162019 Class 3 0.002528 Class 4 -0.489288 0.162019 0.002528 Class 5 -0.489288 0.162019 0.002528 Mean intercept 73.682004 0.182158 < 1e-07 Mean covariates cov1 0.343027 0.172083 0.046219 Variance intercept 6.308833 1.561902 0.000054 Variance covariates 1.886914 1.201458 0.116294 cov2 Transformation Lambda1 1.000000 Fixed Lambda2 0.050000 Fixed \_\_\_\_\_ Final ln likelihood: -383.779624 \_\_\_\_\_ MAXIMIZATION RESULTS susceptibilities free, no truncation \_\_\_\_\_ \_\_\_\_\_

Parameter	Estimate	S.E.	P-value	
Susceptibility inte	rcepts			
Class 0	-0.893122	0.422414	0.034487	
Class 1	-0.295137	0.330400	0.371711	
Class 2	-0.437648	0.322448	0.174696	
Class 3	-0.761468	0.647488	0.239582	
Class 4	-0.396851	0.305525	0.193974	
Class 5	-0.564788	0.451150	0.210612	
Mean intercept	73.683039	0.182786	< 1e-07	
Mean covariates				
covl	0.346353	0.172818	0.045054	
Variance intercept	6.358667	1.578147	0.000056	
Variance covariates				
	1.907274	1.214693	0.116376	
Transformation				
Lambdal	1.000000	Fix	ed	
Lambda2	0.050000	Fix	ed	
Final ln likelihood: -382.959654 JOINT TEST				
HO ln likelihood su H1 ln likelihood su	-			
2 *  H0 - H1  Degrees of freedom P-value				1.639942 5 0.896377

#### 16.3.12.2 Example Detailed Output File

The detailed output file contains all information present in the summary output file, and has the following additional information:

- first partial derivatives for all parameters
- estimates for all four models (with/without truncation)
- variance-covariance matrices for all four models
- additional likelihood ratio test statistics for the models not listed in the summary output file.

Here is an example of an AGEON detailed output file:

```
------
     Sample description
-
  Number of pedigrees in dataset
                                             200
  Number of analyzable pedigrees
                                            200
  Number of individuals in dataset
                                             787
  Number of analyzable individuals
Number of analyzable invalid individuals
                                             787
                                            149
  Number of analyzable valid individuals
                                             638
-----
    MODEL DESCRIPTION
-----
  Title
                     AGEONSET Analysis 1
  Affectedness trait AFF
  Age-of-onset trait AO
  Age-at-exam trait
                     AE
-----
    CLASSIFICATION SYSTEM
-----
  Using default classification system:
   0 Both parental traits are unknown.
      One of the parental traits is unknown, the other is affected.
   2 One of the parental traits is unknown, the other is unaffected.
    3 Both parental traits are affected.
      One of the parental traits is affected, the other is unaffected.
   5 Both parental traits are unaffected.
CLASS STATISTICS
 -----
     CLASS 0
-----
  TOTAL NUMBER OF INDIVIDUALS USED IN ANALYSIS
                                                       60
  NUMBER OF INDIVIDUALS WITH AN AGE OF ONSET
                                                       25
                                                75.458537
  MEAN OF AGE OF ONSET
  VARIANCE OF AGE OF ONSET
NUMBER OF INDIVIDUALS AFFECTED
                                                 4.872671
                                                        9
                                                 0.150000
  PROPORTION OF INDIVIDUALS AFFECTED
  MEAN OF AGE AT EXAM OF THE UNAFFECTED
VARIANCE OF AGE AT EXAM OF THE UNAFFECTED
                                                74.816667
                                                 6.949722
     CLASS 1
_____
  TOTAL NUMBER OF INDIVIDUALS USED IN ANALYSIS
                                                       80
  NUMBER OF INDIVIDUALS WITH AN AGE OF ONSET
                                                       33
  MEAN OF AGE OF ONSET
                                                74.696970
  VARIANCE OF AGE OF ONSET
NUMBER OF INDIVIDUALS AFFECTED
                                                 6.514233
                                                       20
                                                 0.250000
  PROPORTION OF INDIVIDUALS AFFECTED
  MEAN OF AGE AT EXAM OF THE UNAFFECTED
VARIANCE OF AGE AT EXAM OF THE UNAFFECTED
                                                75.350000
                                                 4.377500
-----
     CLASS 5
------
  TOTAL NUMBER OF INDIVIDUALS USED IN ANALYSIS
                                                       43
  NUMBER OF INDIVIDUALS WITH AN AGE OF ONSET
                                                       15
                                                76 266667
  MEAN OF AGE OF ONSET
```

VARIANCE OF AGE OF NUMBER OF INDIVIDUA			4.595	556 9
PROPORTION OF INDIV		TED	0.209	
MEAN OF AGE AT EXAM VARIANCE OF AGE AT			75.744 5.446	
VARIANCE OF AGE AI				
MAXIMIZATION RES				
Parameter	Estimate			
Susceptibility inte				
Class 0 Class 1	-0.489288	0.162019	0.002528	-0.0000002360
Class 2	-0.489288	0.162019	0.002528	
Class 3	-0.489288	0.162019	0.002528	
Class 4	-0.489288 -0.489288 -0.489288	0.162019	0.002528	
Class 5	-0.489288	0.162019	0.002528	0.000000000
Mean intercept	73.682004	0.182158	< 1e-07	0.000006171
Mean covariates cov1	0.343027	0.172083	0.046219	-0.0000017870
Variance intercept				-0.0000001/8/0
Variance covariates		1.501902	0.000054	-0.0000000000000
	1.886914	1.201458	0.116294	0.0000008041
Transformation				
Lambda1	1.000000		Fixed	
Lambda2	0.050000		Fixed	
nal ln likelihood: -				
MAXIMIZATION RES				
MAXIMIZATION RES				
Parameter				
Parameter	Estimate			
Susceptibility inte				
Class 0	-0.893122	0.422414	0.034487	-0.000000189
Class 1	-0.295137		0.371711	0.000000000
Class 2	-0.437648	0.322448	0.174696	-0.0000002023
Class 3	-0.761468 -0.396851 -0.564788	0.647488	0.239582	0.000000000
Class 4	-0.396851	0.305525	0.193974	
Class 5	-0.564788 73.683039	0.451150	0.210612	0.0000000597
Mean intercept Mean covariates	/3.083039	0.182/80	< 1e-07	-0.000000016
cov1	0.346353	0.172818	0.045054	-0.0000001012
Variance intercept				-0.000000080
Variance covariates				
cov2	1.907274	1.214693	0.116376	-0.000000088
Transformation				
Lambda1 Lambda2			Fixed Fixed	
	202 050654			
nal ln likelihood: -				
MAXIMIZATION RES	ULTS suscept	ibilities e	equal, using	g truncation
	ULTS suscept	ibilities e	qual, using	f truncation
MAXIMIZATION RES	ULTS suscept	ibilities e S.E.	equal, using P-value	truncation Deriv
MAXIMIZATION RES	ULTS suscept Estimate	ibilities e S.E.	equal, using P-value	truncation Deriv
MAXIMIZATION RES Parameter Susceptibility inte	ULTS suscept Estimate Prcepts	ibilities e S.E.	equal, using P-value	g truncation Deriv
MAXIMIZATION RES Parameter Susceptibility inte Class 0 Class 1	ULTS suscept Estimate 	ibilities e S.E. 0.162019 0.162019	equal, using P-value 0.002528 0.002528	g truncation Deriv -0.0000002360 0.0000000000
MAXIMIZATION RES Parameter Susceptibility inte Class 0 Class 1 Class 2	ULTS suscept Estimate -0.489288 -0.489288 -0.489288	S.E. 0.162019 0.162019 0.162019 0.162019	equal, using P-value 0.002528 0.002528 0.002528	g truncation Deriv -0.0000002360 0.000000000 0.000000000
MAXIMIZATION RES Parameter Susceptibility inte Class 0 Class 1 Class 2 Class 3	ULTS suscept Estimate -0.489288 -0.489288 -0.489288 -0.489288 -0.489288	ibilities e S.E. 0.162019 0.162019 0.162019 0.162019 0.162019	equal, using P-value 0.002528 0.002528 0.002528 0.002528 0.002528	g truncation Deriv -0.0000002360 0.000000000 0.000000000 0.000000000
MAXIMIZATION RES Parameter Susceptibility inte Class 0 Class 1 Class 2 Class 3	ULTS suscept Estimate -0.489288 -0.489288 -0.489288 -0.489288 -0.489288	ibilities e S.E. 0.162019 0.162019 0.162019 0.162019 0.162019	equal, using P-value 0.002528 0.002528 0.002528 0.002528 0.002528	g truncation Deriv -0.0000002360 0.000000000 0.000000000 0.000000000
MAXIMIZATION RES Parameter Susceptibility inte Class 0 Class 1 Class 2 Class 3	ULTS suscept Estimate -0.489288 -0.489288 -0.489288 -0.489288 -0.489288	ibilities e S.E. 0.162019 0.162019 0.162019 0.162019 0.162019	equal, using P-value 0.002528 0.002528 0.002528 0.002528 0.002528	g truncation Deriv -0.0000002360 0.000000000 0.000000000 0.000000000
MAXIMIZATION RES Parameter Susceptibility inte Class 0 Class 1 Class 2 Class 3 Class 4 Class 4 Class 5 Mean intercept	ULTS suscept Estimate -0.489288 -0.489288 -0.489288 -0.489288 -0.489288	ibilities e S.E. 0.162019 0.162019 0.162019 0.162019 0.162019	equal, using P-value 0.002528 0.002528 0.002528 0.002528 0.002528	g truncation Deriv -0.0000002360 0.000000000 0.000000000 0.000000000
MAXIMIZATION RES Parameter Susceptibility inte Class 0 Class 1 Class 2 Class 3 Class 4 Class 5 Mean intercept Mean covariates	ULTS suscept Estimate -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 73.682004	ibilities e S.E. 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019	equal, using P-value 0.002528 0.002528 0.002528 0.002528 0.002528 0.002528 0.002528 <1e-07	-0.000002360 0.000000000 0.00000000 0.000000000 0.000000
MAXIMIZATION RES Parameter Susceptibility inte Class 0 Class 1 Class 2 Class 3 Class 4 Class 5 Mean intercept Mean covariates Covl	ULTS suscept Estimate -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 73.682004 0.343027	ibilities e 5.E. 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019	P-value 0.002528 0.00258 0.	<pre>g truncation Deriv -0.000002366 0.000000000 0.000000000 0.000000000 0.000000</pre>
MAXIMIZATION RES Parameter Susceptibility inte Class 0 Class 1 Class 2 Class 3 Class 4 Class 4 Class 5 Mean intercept Mean covariates cov1 Variance intercept Variance covariates	ULTS suscept Estimate -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 73.682004 0.343027 6.308833	ibilities e S.E. 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019	equal, using P-value 0.002528 0.002528 0.002528 0.002528 0.002528 0.002528 0.002528 	-0.000002360 0.000000000 0.00000000 0.000000000 0.000000
MAXIMIZATION RES Parameter Susceptibility inte Class 0 Class 1 Class 2 Class 3 Class 3 Class 4 Class 5 Mean intercept Mean covariates cov1 Variance intercept Variance covariates cov2	ULTS suscept Estimate -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 73.682004 0.343027 6.308833	ibilities e S.E. 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019	equal, using P-value 0.002528 0.002528 0.002528 0.002528 0.002528 0.002528 0.002528 	-0.000002360 0.000000000 0.00000000 0.000000000 0.000000
MAXIMIZATION RES Parameter Susceptibility inte Class 0 Class 1 Class 2 Class 3 Class 4 Class 5 Mean intercept Mean covariates cov1 Variance intercept Variance intercept Variance covariates cov2 Transformation	ULTS suscept Estimate -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 73.682004 0.343027 6.308833 1.886914	ibilities e S.E. 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.182158 0.172083 1.561902 1.201458	P-value 0.002528 0.002528 0.002528 0.002528 0.002528 0.002528 0.002528 0.002528 (-07) 0.046219 0.000054 0.116294	-0.000002360 0.000000000 0.000000000 0.000000000 0.000000
MAXIMIZATION RES Parameter Susceptibility inte Class 0 Class 1 Class 2 Class 3 Class 4 Class 5 Mean intercept Mean covariates cov1 Variance intercept Variance covariates cov2 Transformation Lambdal	ULTS suscept Estimate -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 73.682004 0.343027 6.308833 1.886914 1.000000	ibilities e S.E. 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.182158 0.172083 1.561902 1.201458	P-value 0.002528 0.002528 0.002528 0.002528 0.002528 0.002528 0.002528 0.002528 (-07) 0.046219 0.000054 0.116294 Fixed	-0.000002360 0.000000000 0.000000000 0.000000000 0.000000
MAXIMIZATION RES Parameter Susceptibility inte Class 0 Class 1 Class 2 Class 3 Class 4 Class 5 Mean intercept Mean covariates cov1 Variance intercept Variance intercept Variance covariates cov2 Transformation	ULTS suscept Estimate -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 73.682004 0.343027 6.308833 1.886914 1.000000 0.050000	ibilities e S.E. 0.162019 0.152019 0.150019 0.15200000000000000000000000000000000000	equal, using P-value 0.002528 0.002528 0.002528 0.002528 0.002528 0.002528 0.002528 < 1e-07 0.046219 0.000054 0.116294 Fixed Fixed	-0.000002360 0.00000000 0.00000000 0.00000000 0.000000
MAXIMIZATION RES Parameter Susceptibility inte Class 0 Class 1 Class 2 Class 3 Class 4 Class 4 Class 5 Mean intercept Mean covariates cov1 Variance intercept Variance covariates cov2 Transformation Lambda1 Lambda2 nal ln likelihood: -	ULTS suscept Estimate -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 73.682004 0.343027 6.308833 1.886914 1.000000 0.050000	ibilities e S.E. 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.182158 0.172083 1.561902 1.201458	equal, using P-value 0.002528 0.002528 0.002528 0.002528 0.002528 0.002528 0.002528 0.002528 (-07) 0.046219 0.000054 0.116294 Fixed Fixed	-0.000002360 0.000000000 0.000000000 0.000000000 0.000000
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MAXIMIZATION RES Parameter Susceptibility inte Class 0 Class 1 Class 2 Class 3 Class 4 Class 4 Class 5 Mean intercept Mean covariates cov1 Variance intercept Variance covariates cov2 Transformation Lambda1 Lambda2 nal ln likelihood: -	ULTS suscept Estimate -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 73.682004 0.343027 6.308833 1.886914 1.000000 0.050000 -383.779624 ULTS suscept	ibilities e S.E. 0.162019 0.1621900000000000000000000000000000000000	P-value 0.002528 0.002528 0.002528 0.002528 0.002528 0.002528 0.002528 (-07) 0.046219 0.000054 0.116294 Fixed Fixed Fixed	<pre>g truncation Deriv -0.0000002360 0.000000000 0.000000000 0.000000000 0.000000</pre>
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MAXIMIZATION RES Parameter Susceptibility inte Class 0 Class 1 Class 2 Class 3 Class 4 Class 5 Mean intercept Wean covariates cov1 Variance intercept Variance covariates cov2 Transformation Lambda1 Lambda1 nal ln likelihood:	ULTS suscept Estimate -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 73.682004 0.343027 6.308833 1.886914 1.000000 0.050000 -383.779624 ULTS suscept Estimate -0.893122	ibilities e S.E. 0.16201900000000000000000000000000000000	P-value P-value 0.002528 0.002528 0.002528 0.002528 0.002528 0.002528 (-07) 0.046219 0.000054 0.116294 Fixed Fixed Fixed Fixed 0.016294 0.000054 0.016294 0.000054 0.016294 0.000054 0.016294 0.000054 0.016294 0.000054 0.016294 0.000054 0.0	g truncation Deriv -0.0000002360 0.000000000 0.000000000 0.000000000
MAXIMIZATION RES Parameter Susceptibility inte Class 0 Class 1 Class 2 Class 3 Class 4 Class 5 Mean intercept Mean covariates cov1 Variance intercept Variance covariates cov2 Transformation Lambda1 Lambda1 Lambda2 nal ln likelihood: Parameter Susceptibility inte Class 1	ULTS suscept Estimate -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 73.682004 0.343027 6.308833 1.886914 1.000000 0.050000 -383.779624 Estimate -0.893122 -0.295137	ibilities e S.E. 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.182158 0.172083 1.561902 1.201458 ibilities f	equal, using P-value 0.002528 0.00054 0.116294 Fixed Fixed P-value 0.034487 0.37711	g truncation Deriv -0.000002360 0.000000000 0.000000000 0.000000000
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MAXIMIZATION RES Parameter Susceptibility inte Class 0 Class 1 Class 2 Class 3 Class 4 Class 5 Mean intercept Mean covariates cov1 Variance intercept Variance covariates cov2 Transformation Lambda1 Lambda1 Lambda2 nal ln likelihood: Parameter Susceptibility inte Class 1	ULTS suscept Estimate -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 73.682004 0.343027 6.308833 1.886914 1.000000 0.050000 -383.779624 Estimate -0.893122 -0.295137	ibilities e S.E. 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.182158 0.172083 1.561902 1.201458 ibilities f	equal, using P-value 0.002528 0.00054 0.116294 Fixed Fixed P-value 0.034487 0.377711	g truncation Deriv -0.000002360 0.000000000 0.000000000 0.000000000
MAXIMIZATION RES Parameter Susceptibility inte Class 0 Class 1 Class 2 Class 3 Class 4 Class 5 Mean intercept Mean covariates cov1 Variance intercept Variance covariates cov2 Transformation Lambda1 Lambda1 Lambda2 nal ln likelihood: Parameter Susceptibility inte Class 1	UULTS suscept Estimate -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 -0.489288 73.682004 0.343027 6.308833 1.886914 1.000000 0.050000 -383.779624 Estimate -0.893122 -0.295137	ibilities e S.E. 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.162019 0.182158 0.172083 1.561902 1.201458 ibilities f	equal, using P-value 0.002528 0.00054 0.116294 Fixed Fixed P-value 0.034487 0.377711	g truncation Deriv -0.000002360 0.000000000 0.000000000 0.000000000
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cov1		D.172818 0.045054 L.578147 0.000056				
Variance covariates	1					
cov2 Transformation	1.907274	1.214693 0.116376	-0.00000008	8		
	1.000000 0.050000	Fixed Fixed				
Final ln likelihood: -						
JOINT TEST						
		free, no truncation equal, no truncation				
2 *  HO - H1			1.63994	2		
Degrees of freedom				5		
P-value			0.89637	7		
JOINT TEST						
		free, using truncati				
Hl ln likelihood su	sceptibilities	equal, using truncat	ion -383.77	9624		
2 *  HO - H1  Degrees of freedom			1.63	9942 5		
P-value			0.89	5		
		ceptibilities equal,				
		Variance intercept				Class 5
Mean intercept	0.033182	0.057390 2.439538 -0.000076 1.257227 0.041093 0.041093	-0.003310	0.032536	0.010836	0.010836
Variance intercept	0.057390	2.439538	-0.000076	1.257227	0.041093	0.041093
COVI	-0.003310	-0.000076	0.029612	0.006981	-0.000758	0.000758
COV2	0.032536	1.25/22/	0.006981	1.443502	0.021532	0.021532
Class U	0.010836	0.041093 0.041093 0.041093 0.041093 0.041093	-0.000758	0.021532	0.026250	0.026250
Class 2	0.010836	0.041093	-0.000758	0.021532	0.026250	0.026250
	0.010836	0.041093	-0.000758	0.021532	0.026250	0.026250
Clagg 4	0 010836	0 041093	-0.000758	0.021532	0.026250	0.026250 0.026250
Class 5	0.010836	0.041093	-0.000758	0.021532	0.026250	0.026250
				====		
VARIANCE-COVARIA		ceptibilities free, n				
Error: Matrix is not a						
	NCE MATRIX sus	ceptibilities equal,	using trunca	tion		
Mean intercept	0.033182	0.057390 2.439538	-0.003310	0.032536	0.010836	0.010836
Variance intercept	0.057390	2.439538	-0.000076	1.257227	0.041093	0.041093
covl	-0.003310	-0.000076	0.029612	0.006981	-0.000758	0.000758
cov2	0.032536	1.257227	0.006981	1.443502	0.021532	0.021532
Class 0	0.010836	0.041093 0.041093 0.041093 0.041093 0.041093	-0.000758	0.021532	0.026250	0.026250 0.026250 0.026250 0.026250 0.026250 0.026250
Class 1 Class 2	0.010030	0.041093	-0.000/58	0.021532	0.020250	0 026250
Class 3	0.010836 0.010836 0.010836	0.041093	-0.000758	0.021532	0.020250	0.026250
	0 010936	0.041093	-0.000758	0.021532	0.026250	0.026250
Class 5	0.010836	0.041093	-0.000758	0.021532	0.026250	0.026250
VARIANCE-COVARIA		ceptibilities free, u				
Error: Matrix is not						

Error: Matrix is not available.

# Chapter 17

# DECIPHER

DECIPHER obtains maximum likelihood estimates of population haplotype frequencies for autosomal or X-linked markers, and determines all possible diplotypes and the most likely diplotypes for each individual. Genotypes of other pedigree members can be used to infer phase for ambiguous individuals, which improves the population haplotype frequency estimates over those obtained using unrelated individuals only. Haplotype frequencies can be estimated separately for different populations that are specified by the user. A likelihood ratio test and a permutation test are provided to compare haplotype frequency distributions for dichotomous phenotypes.

#### 17.1 Limitations

Genotypes of other pedigree members can be used to infer phase for ambiguous individuals only for non-recombinant regions (i.e., no recombination is observed in the pedigree between those markers). Memory constraints may be encountered in situations where a large fraction of markers are missing, or when a large number of markers (more than 25) are haplotyped. Finally, markers in the haplotyping region must be codominant, and family information may not be used with X-linked markers. The current version of DECIPHER limits the number of individuals per pool, k, to at most 1; this limitation will eventually be removed in a future version of the program.

## 17.2 Theory

Maximum likelihood estimates of haplotype frequencies can be obtained from pooled DNA using a form of the expectation-maximization (EM) algorithm developed expressly for that purpose (Quade et al. 2005; Ito et al. 2003; Wang, Kidd, and Zhao 2003). The approach incorporates a variety of data types, including unrelated individuals, sets of related individuals (i.e., families), and pooled samples, or combinations of these data types. The key modification was the recognition that each of the other types of data can be considered a special case of pooled data. For example, unrelated individuals can be considered a pool with one individual. Groups of founders in a pedigree can be considered a pool of f individuals, where f is the number of founders. To allow combinations of the EM algorithm to allow different numbers of individuals in each unit.

To estimate population haplotype frequencies, a random set of unrelated individuals or pedigrees must be used. For pedigree data, the user can specify a single representative from each pedigree. The family representative is either selected by the user or, if no individual is indicated for a particular pedigree, the program will randomly select one individual out of those individuals in the pedigree with the most marker genotypes (i.e., we are assuming the genotypes are missing at random).

The form of the EM algorithm for pooled data is as follows. Suppose we are given n pools and each pool contains k individuals. The total number of markers is m. In this description, we primarily focus on single nucleotide polymorphisms (SNPs) which are diallelic markers with alleles encoded as 0 or 1; however, DECIPHER allows more than two alleles per locus. For each pool, at each marker position, we are given the number of 0s and the number of 1s. The summation of these two numbers is 2k since each individual provides 2 alleles and there are k individuals in each pool. The input data can be represented by a nonnegative integer matrix M of size  $n \ge m$ , where the *i*-th row,  $M_{i,i}$ , represents one pool and the *j*-th column,  $M_{i,j}$ , represents one SNP, where  $1 \le i \le n$  and  $1 \le j \le m$ . Each entry,  $M_{ij}$ , is an integer representing the number of allele 1s in pool *i* at SNP *j*. The value of each entry thus should be  $\le 2k$  and  $\ge 0$ . For *m* markers, there are total of

 $T = 2^m$  possible haplotypes. Let  $h_t$  denote the *t*-th haplotype and let  $f_t$  denote its population frequency for  $0 \le t \le T$ . Let  $H = \{h_t : 0 \le t \le T\}$  and  $F = \{f_t : 0 \le t \le T\}$  be the set of all haplotypes and the set of haplotype frequencies, respectively. For a given pool  $M_i$ , let  $H_i$ denote the set of all possible haplotype assignments for  $M_{i.}$ , i.e., each element  $\Delta$  of  $H_i$  contains 2k haplotypes for the *k* individuals in pool *i*. Under the assumption of Hardy-Weinberg equilibrium and random mating, and assuming that all the individuals are independent, the likelihood for the given data can be expressed as

$$P(M,F) = \prod_{i=1}^{n} \sum_{\Delta \in H_i} P(\Delta)$$
(17.1)

The standard EM algorithm starts with an initial assignment of the haplotype frequencies for F. During the E step, the expected number of each haplotype is calculated under the assumption that the haplotype frequencies are known, and during the M step the haplotype frequencies are updated according to the haplotype counts calculated in the previous E step. The two steps are iterated until convergence, defined as the minimum difference between haplotype frequencies in successive iterations being less than a small number,  $\varepsilon$ , which is specified by the user. To ensure that a global maximum is reached rather than a local maximum, the user can specify the number of starting points that will be used. DECIPHER will obtain maximum likelihood estimates for each of these randomly selected starting points, and the set of estimates corresponding to the maximum likelihood will be displayed. We have modified this algorithm so that the value of k can differ for each pool. Note that for a pool that consists of a single male with X-linked data, k equals 1/2; however, in this instance the haplotype is always known with certainty.

For pedigree data, we use descent graphs to identify compatible haplotypes for a particular individual in the pedigree consistent with the observed data in the pedigree. We assume all markers are in a region with no observed recombination within the pedigrees. Using the method of descent graphs described by (Sobel and Lange 1996), we can identify all possible allele states at each locus for each individual. A complete list of all possible haplotype states for each individual can then be obtained by taking the Cartesian product of the possible allele states at each locus. The possible founder haplotypes are linked through the descent graphs, such that sets of founder haplotypes that are simultaneously consistent with the observed data can be obtained. These sets of possible haplotypes,  $H_i$ , are then used in equation 17.1 above. There are several types of information that can be obtained from this procedure. First, haplotype frequencies can be estimated for sets of individuals. The user has the option of partitioning the individuals into subpopulations (e.g., case-control status, ethnic groups, etc) and obtaining haplo-type frequencies separately for each subpopulation. Second, we can obtain a list of all possible non-recombinant diplotypes for each individual or pool (with the constraint of < 30 markers). In addition, we can obtain the posterior probability of each of these possibilities. Third, we can obtain a list of the most likely pairs of haplotypes (i.e., diplotypes) for each individual, together with the relative likelihood of each, based on population data. The list of haplotypes or diplotypes can be quite large, particularly when there is a large number of markers and/or alleles. Therefore, an option is provided to specify a threshold, such that only haplotypes (diplotypes) with a frequency (posterior probability) greater than this threshold will be displayed. In the case where the most likely diplotypes are requested, more than one diplotype will be returned if they have the same (maximum) posterior probability.

A likelihood ratio test is available to compare the distribution of haplotypes in populations (e.g., cases versus controls). Assume we have N groups, and we have estimated haplotype frequencies separately for each group and for the whole sample combined. Assume there are  $h_i$  haplotypes with frequency  $p_{ij}$  for haplotype *i* in group *j*. For the likelihood ratio test, the null hypothesis is  $H_0$ :  $p_{i1} = p_{i2} = \dots = p_{in}$ , versus the alternative hypothesis,  $H_A$ :  $p_{ij} \neq p_{ik}$  for at least one haplotype *i*, and at least one pair of groups *j* and *k*. The likelihood is maximized under these two conditions (i.e., forcing  $p_{ij}$  to be the same for all j versus allowing them to be different). The likelihood ratio (LR) is then formed, and -2ln(LR) asymptotically follows a chi-square distribution with  $(N-1)(h_T-1)$  degrees of freedom, where  $h_T$  is the number of haplotypes for the whole sample. This asymptotic distribution is conservative when there are rare haplotypes, and is not recommended under those circumstances. Therefore, we also provide a method for obtaining an empirical p-value for the LR test statistic. This is obtained by sampling permutations of the category assignment (e.g., case-control status), and recomputing the LR test statistic for each permutation. The empirical p-value is determined from the sample permutations as the number of permutations where the LR test statistic exceeds the observed LR test statistic, divided by the total number of permutations.

File Type	Description
Parameter file	Specifies the parameters and options with which
	to perform a particular analysis.
Pedigree data file	Contains delimited records for each individual,
	including fields for identifiers, sex, parents, trait
	and marker data.
Genome description file	Contains a description of the linked marker re-
	gions, including distances between consecutive
	markers.

## 17.3 Program Input

Note: A marker locus description file is not accepted by DECIPHER. Missing allele and allele delimiter symbols may be specified using the marker sub-block in the parameter file. See section 2.3.3 for details.

#### 17.3.1 Parameter File

The following syntax table specifies the permissible parameter and attribute settings for the main DECIPHER parameter.

<pre>parameter [, attribute]</pre>		Explanation	
	Starts a DECIPHE	R analysis block.	
	Value Range	N/A	
decipher	Default Value	N/A	
	Required	Yes	
	Applicable Notes	None	
	Specifies the root	name to be used for output files.	
	Output file names will be formed by concatenating the		
	root name and an a	ppropriate extension.	
	Value Range	Character string representing a	
, out	value Raliye	valid file name	
	Default Value	decipher.out	
	Required	No	
	Applicable Notes	None	

### 17.3.2 The decipher Parameter Block

The following syntax table specifies the permissible parameter and attribute settings for the decipher parameter block.

parameter [, attribute]		Explanation	
	Specifies the title of	of the analysis	
	Value Range	Character string	
title	Default Value	Analysis 1	
01010	Required	No	
	Applicable Notes	None	
		e of the chromosomal region to be	
	analyzed.		
		Character string naming a region	
	Value Range	listed in the genome description	
region	-	file.	
	Default Value	None	
	Required	Yes	
	Applicable Notes	1	
	Specifies the minir	num difference between haplotype	
	frequencies in successive iterations as a convergence		
	criterion for the EM	A algorithm.	
epsilon	Value Range	(0, 1)	
	Default Value	0.00001	
	Required	No	
	Applicable Notes	None	
		of randomly chosen starting points	
	against which the I	EM algorithm is to run.	
starting_points	Value Range	{1, 2, 3,}	
starting_points	Default Value	10	
	Required	No	
	Applicable Notes	None	
		write haplotype frequencies (and	
	their log-likelihoods) for each set of EM algorithm		
	starting points to a	_	
dump	Value Range	{true, false}	
	Default Value	false	
	Required	No	
	Applicable Notes	2, 3	

, cutoff	Specifies minimum haplotype frequency threshold value for display. If none of the estimated haplotype frequencies meet or exceed the specified value, then the haplotype with the greatest estimated frequency is displayed.Value Range[0, 1]Default Value0.001RequiredNoApplicable NotesNone
	Starts a sub-block to specify
data	<ol> <li>how to treat relatedness of individuals,</li> <li>individuals to represent families, and</li> <li>individuals to represent subpopulations.</li> </ol>
	Value Range N/A
	Default Value N/A
	Required No
	Applicable Notes None
	Starts a sub-block to specify analysis tasks to be per-
	formed.
tasks	Value Range N/A
Cabib	Default Value N/A
	Required No
	Applicable Notes None

Notes

- 1. Used for establishing marker order (no marker order is implied by parameter or pedigree files). Distances between the markers are ignored by DECIPHER.
- 2. Applicable only if pop\_freq or most\_likely\_diplotypes in tasks sub-block is set to **true**.
- 3. Starting points shown in the dump file output are generated by choosing random phase probabilities for each pool and then calculating the haplotype frequencies.

#### 17.3.3 The data Sub-Block

The following syntax table specifies the permissible parameter and attribute settings for the data sub-block.

parameter	Explanation		
[, attribute]			
related	sidered in determin rameter is set to <b>fa</b> be independent and tion of haplotype fr typed person per of Familial informatio ing possible diploty be treated as if the Value Range Default Value Required	Tamilial relationships are to be con- ting possible diplotypes. If this pa- lse, individuals will be assumed to a they will all be used in the estima- equencies. If set to <b>true</b> , one geno- constituent pedigree will be used. on will be considered in determin- types. Unconnected individuals will parameter where set to <b>false</b> .         {true, false}         true         No	
	Applicable Notes	1	
	Variable used to specify one genotyped individual per constituent pedigree when related equals <b>true</b> .		
family_rep	Value Range	Character string representing the name of a trait, phenotype, covariate or string field listed in the pedigree data file.	
	Default Value	None	
	Required	No	
	Applicable Notes	2,8	
	Specifies the value	of the family_rep variable that bed individual for haplotype analy-	
, family_rep_value	Value Range	Character string representing the (literal) value to be matched in the designated family_rep field.	
	Default Value	None	
	Required	Yes, if family_rep is specified	
	Applicable Notes	3	

	Starts a sub-block for specifying subpopulations of in- dividuals from the original data set.		
partition	Value Range	Character string representing the name of a trait, phenotype, covariate or string field listed in the pedigree data file. The named field will be used as the basis of classification for the created partition (see notes).	
	Default Value	None	
	Required	No	
	Applicable Notes	4, 5, 6, 7, 8	

Notes:

- 1. If this option is selected and a Mendelian inconsistency is detected in a constituent pedigree at a particular locus, all members of the constituent pedigree are treated as if they had missing values for that locus.
- 2. If no variable is specified, the program will arbitrarily pick a genotyped individual in each constituent pedigree to be used in haplotype calculations.
- 3. If more than one individual in a pedigree has this value, the one with the most genotyped loci in the haplotype region is chosen. In the case of a tie, the selection is made at random.
- 4. This sub-block may appear no more than twice per analysis block and each partition sub-block in an analysis block must have a unique value.
- 5. If this sub-block is not specified, all individuals will be treated as a single population.
- 6. All individuals having the same value for this variable belong to the same subpopulation.
- 7. The order in which the partitions are listed is significant. See note #3 of the tasks sub-block for details.
- 8. The same trait, phenotype or covariate may not be used as a value for both the family\_rep and partition parameters.

#### 17.3.4 The partition Sub-Block

The following syntax table specifies the permissible parameter and attribute settings for the partition
sub-block.

<pre>parameter [, attribute]</pre>		Explanation	
	Specifies name of subpopulation.		
	Value Range	Character string	
sub_pop	Default Value	None	
	Required	No	
	Applicable Notes	1, 2	
	Specifies value of p	partition variable common to all in-	
	dividuals in this subpopulation.		
		Character string representing the	
	Value Range	(literal) value to be matched in	
, sub_pop_value		the designated sub_pop field.	
	Default Value	None	
	Required	No	
	Applicable Notes	3, 4, 5	

Notes:

- 1. If not specified, subpopulation name is the same as sub\_pop\_value.
- 2. This parameter may be repeated as needed but sub\_pop and sub\_pop\_value must be unique within a partition.
- 3. Required if sub\_pop is specified.
- 4. If no valid values are specified for the sub\_pop\_value option, then every distinct value of the partition variable found in the pedigree data file (except the missing value), will designate a subpopulation.
- 5. Missing value code may not be specified as a sub\_pop\_value.

#### 17.3.5 The tasks Sub-Block

The following syntax table specifies the permissible parameter and attribute settings for the tasks sub-block.

parameter	Explanation
[, attribute]	-
	Specifies option to estimate population haplotype fre-
pop_freq	quencies.
	Value Range {true, false}
	Default Value
	Required No
	Applicable Notes None
	Specifies minimum haplotype frequency threshold
	value for display. If none of the estimated haplotype
	frequencies meet or exceed the specified value, then
	the haplotype with the greatest estimated frequency is
, cutoff	displayed.
	Value Range [0, 1]
	Default Value 0.001
	Required No
	Applicable Notes 1
	Specifies option to display diplotypes for each indi-
	vidual in tabular form.
	Value Range {true, false}
all_possible_diplotypes_table	Default Value false
	Required No
	Applicable Notes None
	Specifies option to display the most likely diplotypes
	for each individual.
most_likely_diplotypes	Value Range {true, false}
most_invery_dipiotypes	Default Value false
	Required No
	Applicable Notes None
	Specifies minimum haplotype frequency threshold
	value for display. If none of the estimated haplotype
	frequencies meet or exceed the specified value, then
	the haplotype with the greatest estimated frequency is
, cutoff	displayed.
	Value Range [0, 1]
	Default Value 0.05
	Required No
	Applicable Notes 2
	Specifies option to perform likelihood ratio test.
	Value Range {true, false}
likelihood_ratio_test	Default Value false
	Required No
	Applicable Notes 3

	Specifies option to estimate p-values by permutation methods.
compute_empirical_pvalue	Value Range {true, false}
	Default Value false
	Required No
	Applicable Notes None
	Specifies an exact number of permutations to be per-
	formed. Use of this option effectively overrides all of
	the following attributes.
, permutations	Value Range {1, 2, 3,}
	Default Value None
	Required No
	Applicable Notes None
	Specifies the maximum number of permutations that
	should be performed.
	Value Range {1, 2, 3,}
, max_permutations	Default Value 10,000
	Required No
	Applicable Notes None
	Specifies the relative precision of the empirical p-
	value. For example, if width $= 0.2$ , then p-values
	will be estimated to be within 20% of their true
	value with a given confidence level. This value is
	used to choose the number of permutations necessary.
, width	Note that the number of permutations required varies
	quadratically with the inverse of the width.
	Value Range [0, 1]
	Default Value 0.2
	Required No
	Applicable Notes None
	Specifies the confidence with which an empirical p-
	value is required to be within a relative interval (i.e.,
	the width) of its true value.
, confidence	Value Range [0, 1]
	Default Value 0.95
	Required No
	Applicable Notes None

Notes:

- 1. To display all *haplotype* frequency estimates, specify a cutoff of 0.
- 2. To display all *diplotype* probabilities, specify a cutoff of 0.
- 3. If two partition sub-blocks are specified, the first partition listed is *Partition 1*, and the second partition listed is *Partition 2*. Likelihood ratio tests are performed across the subpopulations of Partition 1 for each of the subpopulations in Partition 2. At least two subpopulations must be specified in a partition sub-block to do this test.

The following are all valid decipher analysis blocks:

```
decipher, region = "chrom 1" {
}
decipher, region = Chrl2 { # Quotes not required since the region name does not contain spaces
   related = false
                           # Do not use family information in determining
                           # possible haplotypes.
}
decipher, out="run 1" {
   title = "1st run"
   region = "chrom 14"
   epsilon = .0001
                         # End EM algorithm when differences in frequency estimates for
                         # successive iterations are less than .0001 for all haplotypes.
   starting_points = 3 # Run EM algorithm 3 times with a different set of
                         # starting points each time.
   data {
      related = true
      family_rep = "T1", # Values for this trait designate a
                         # genotyped individual in each family whose haplotypes
                         # are to be determined.
         family_rep_value = 1 # Haplotypes to be determined for genotyped individuals
                               # whose value for trait, T1, equals 1.
      partition = "T2" {
                               # Values for this trait will determine membership
                               # in subpopulations.
         sub_pop = "popl", sub_pop_value = 1 # If individual has value of 1 for T2, he
                                              # belongs to pop1.
         sub_pop = "pop2", sub_pop_value = 2 # if individual has value of 2 for T2, he
                                              # belongs to pop2.
      }
   }
  tasks {
     pop_freq=true, cutoff = .1 # Show only haplotype frequency estimates greater
                                 # than .1.
     likelihood_ratio_test = true
     compute_empirical_pvalue = true, permutations = 1000
  }
}
```

## **17.4** Program Execution

DECIPHER is run via a command line interface on the supported UNIX and Windows platforms. This requires the S.A.G.E. programs to be properly installed and in the current execution path. Input files are specified on the command line and all output files are created in the current working directory.

Running DECIPHER from the command prompt with no arguments, or the wrong number of arguments, will result in the program displaying its usage statement, which lists the input files names required on the command line.

```
>decipher
DECIPHER Output -- 13 Jul 2005 10:43:16 -- [S.A.G.E. v5.0.3; bld 27 Jun 2005]
COPYRIGHT (C) 2005 CASE WESTERN RESERVE UNIVERSITY
usage: ./decipher <parameters> <pedigree> <map>
Command line parameters:
  parameters - Parameter File
  pedigree - Pedigree Data File
  map - Genome Description File
```

A typical run of DECIPHER may look like the following:

```
>decipher par ped mld gen
DECIPHER Output -- 13 Jul 2005 10:52:36 -- [S.A.G.E. v5.0.3; bld 27 Jun 2005]
COPYRIGHT (C) 2005 CASE WESTERN RESERVE UNIVERSITY
Reading Parameter File.....done.
Reading Pedigree File.....
        from ped.....done.
Sorting Pedigrees.....done.
Reading Genome Description File.....done.
Parsing DECIPHER analyses ...
Parsing new analysis block ...
Parsing of Analysis 1 complete. Analysis valid.
_____
 Performing analysis: Analysis 1
 Total population.
   Determining possible diplotypes ..... done.
   Maximizing likelihood ..... done.
 _____
```

## **17.5 Program Output**

DECIPHER produces four types of output files that contain results and diagnostic information:

File Name	File Type	Description
decipher.inf	DECIPHER Information output file	Contains informational diagnostic messages, warnings and program errors. No analysis results are stored in this file.
decipher.sum	DECIPHER summary output file	Contains population haplotype frequency esti- mates.
decipher.det	DECIPHER detailed output file	Contains possible diplotypes and most likely diplotypes for individuals.
decipher.dmp	DECIPHER data dump file	Contains haplotype frequency estimates and ln likelihoods for each set of starting points for which the EM algorithm is run.

#### **17.5.1** Information Output File

The Information Output File contains a variety of useful information, including:

- Information on fields read from the Pedigree Data File. These tables, which provide information about what the program has read from the Pedigree Data File, are included with all programs in S.A.G.E. and are very useful for debugging most common errors caused when reading the pedigree data. When first analyzing new data, it is recommended that these tables be checked carefully to make sure pedigree data are being correctly read.
- Information, warning and error messages generated throughout the program. It is recommended that this file be checked for warning and error messages before examining the results of any run of the program. The program attempts to correct many common errors and this sometimes means analyses are not as expected. The file "decipher.inf" should be checked for errors and diagnostic information after each run of the program.

#### 17.5.2 Summary Output File

Contains results pertaining to whole data set specifically haplotype frequency estimates, likelihood ratio test results and empirical p-values.

#### 17.5.3 Detail Output File

Contains results on an individual basis, specifically possible and most likely diplotypes.

## **17.6 Example Output Files**

#### 17.6.1 Example Summary Output File

```
==========
Analysis 1
==========
Options Selected
_____
Haplotype region
                                       one
EM algorithm convergence criterion
                                      1e-05
Number of EM algorithm starting states 10
  Dump
                                       no
Subpopulations specified
                                       no
Use family information
                                       yes
  Family representatives specified
                                       yes
Estimate haplotype frequencies
                                       yes
  Cutoff
                                       1e-06
List all possible diplotypes
                                       no
Show all possible diplotypes table
                                       no
List most likely diplotypes
                                       no
Do likelihood ratio test
                                       no
Compute empirical p-value
                                       no
Results
======
Markers in order:
  M1 M2 M3 M4 M5 M6 M7 M8 M9
Missing allele symbol (first marker):
  1 ? 1
                        Haplotype Frequency Estimates
Note: Haplotypes listed have estimated frequencies greater than or equal to
      the cutoff or have the greatest frequency estimate.
Haplotype
                         Frequency
_____
                          _____
1-1-2-2-2-2-1-1-1
                         0.190474
1-1-2-2-1-1-2-2-1
                         0.187969
2-1-2-2-1-2-1-1-1
                         0.186462
2-2-2-1-1-2-2-1-2
                        0.172941
1-1-1-1-1-2-1-1
                        0.163969
1-2-2-2-1-1-2-2-1
                         0.00303146
1-2-1-1-1-1-1-1
                        0.00301807
2-2-1-2-1-1-2-1-2
                         0.00300129
1-1-2-2-1-2-2-1-2
                         0.00300000
1-1-2-1-2-2-2-2-1
                         0.00300000
1-1-2-1-2-1-1-2-2
                          0.00300000
•
1-1-1-1-2-1-2-1
                          0.00100000
2-2-2-2-2-2-1-2-1
                         0.000992059
1-2-2-1-1-1-2-1
                         0.000753648
1-2-1-1-2-1-1-2
                         0.000702381
1-2-1-2-1-2-1-1-2
                         0.000297619
1-2-2-1-1-2-1-2-1
                         0.000246352
1-2-2-2-1-1-2-2-2
                          7.94147e-06
2-2-2-1-2-2-1-2-2
                         1.28784e-06
1-2-2-1-2-2-1-1-2
                         1.27373e-06
                          _____
                          0.999999
Total
Ln likelihood -1860.13
```

#### 17.6.2 Example Detail Output File

```
==========
Analysis 1
===========
Markers in order:
 M1 M2
Missing allele symbol (first marker):
 '?'
                   Most Likely Diplotypes
Note: Diplotypes listed have estimated probabilities greater than or equal
     to the cutoff or have the greatest probability estimate.
Pedigree
         Member
                  Diplotype Probability
           ____
                    _____
                                 _____
_____
1
           1
                    a-A a-a
                                 1.00000
1
           10
                    A-a a-a
                                 1.00000
1
           2
                                 1.00000
                    А-А А-а
          3
1
                                1.00000
                    А-а А-а
          4
                                0.934722
1
                    А-а а-А
                                0.0652776
                    A-A a-a
      5
1
                                0.934722
                    А-а а-А
                    А-А а-а
                                0.0652776
1
           6
                    а-А а-А
                                1.00000
1
           7
                                0.934722
                    A-a a-A
                    A-A a-a
                                0.0652776
                    A-a a-A
1
           8
                                0.934722
                    А-А а-а
                                0.0652776
1
           9
                    А-а а-а
                                 1.00000
                   All Possible Diplotypes
                 pedigree:member
         1:1 1:10 1:2 1:3 1:4 1:5 1:6 1:7 1:8 1:9
A-A A-a
                  х
         -
             _
                      -
                           -
                               _
                                    _
                                        -
                                            _
                                                 -
                      _
A-A a-a
              _
                                    _
                                                 _
         _
                  _
                           х
                               х
                                        х
                                            х
         _
             _
                          _
                                    _
А-а А-а
                  _
                      х
                               _
                                        _
                                            _
                                                 _
А-а а-А -
              -
                  _
                      -
                          х
                                    _
                                        х
                                                 _
                               х
                                            х
A-a a-a - x
                  _
                      -
                           _
                               -
                                    -
                                        _
                                            -
                                                х
а-А а-А -
                - -
                           _
                               _
                                        _
                                            _
             -
                                   x
                                                 _
а-А а-а х
              _
                  _
                      _
                           _
                               _
                                    _
                                        _
                                             _
                                                 _
```

# 17.6.3 Example Dump File

DECIPHER Output	13 Jul 2005 11:14:37 [S.A.G.E. v5.0.3; bld 27 Jun	1 200
	05 CASE WESTERN RESERVE UNIVERSITY	
Maximizing tota		_
	s listed have estimated frequencies greater than or equ	ial t
	f or have the greatest frequency estimate.	
run 1		
start point		
Haplotype	Frequency	
A-a	0.401786 0.301786	
a-A a-a	0.198214	
A-A	0.0982143	
A-A		
Total	1.000000	
end point	1.00000	
Haplotype	Frequency	
A-a	0.436945	
a-A	0.336945	
a-a	0.163055	
A-A	0.0630555	
Total	1.000000	
Ln likelihood	-17.45593957	
run 2		
start point		
Haplotype	Frequency	
A-a	0.370000	
a-A	0.270000	
a-a	0.230000	
A-A	0.130000	
Totol	1.000000	
Total	1.00000	
end point Haplotype	Fromionau	
партосуре	Frequency	
 А-а	0.436945	
a-A	0.336945	
a-a	0.163055	
A-A	0.0630555	
Total	1.000000	
Ln likelihood	-17.45593957	
•		
•		
•		
run 9		
start point		
Haplotype	Frequency	
A-a	0.313158	
a-a	0.286842	
a-A	0.213158	

A-A	0.186842
Total	1.000000
end point	
Haplotype	Frequency
A-a	0.436944
a-A	0.336944
a-a	0.163056
A-A	0.0630555
Total	1.000000
Ln likelihood	-17.45593957
run 10	
start point	
Haplotype	Frequency
A-a	0.344915
a-a	0.255085
a-A	0.244915
A-A	0.155085
Total	1.000000
end point	
Haplotype	Frequency
A-a	0.436944
a-A	0.336944
a-a	0.163056
A-A	0.0630557
Total	1.000000
Ln likelihood	-17.45593957
THI TIVETIHOOD	T1. IJJJJJJ1

# **Chapter 18**

# DESPAIR

DESPAIR is a program to help in designing linkage studies for searching the whole autosomal genome. Originally created for a study comprising affected pairs of relatives of a particular type, the latest version of DESPAIR has been modified to further incorporate discordant relative pairs into the study. The program can be used to determine, for specified power and significance level, the optimal two-stage study design – i.e., how many pairs of relatives should be studied, how many equally spaced markers should be used initially, and what criterion should be used to specify the markers around which further searching should be done. Alternatively, the program will calculate either the number of relative pairs required for a given number of first-stage markers, or the number of markers required for a given number of relative pairs.

#### 18.1 Limitations

#### **18.1.1** Theoretical Limitations

The method used assumes that independent pairs of relatives of a single particular type (full sibling, half-sibling, grandparent-grandchild, avuncular, or first cousin) are being sampled. Only three levels of interference are considered, corresponding to Haldane's mapping function (no interference), Kosambi's mapping function (moderate interference), and Morgan's linear mapping function (extreme interference). The spacing between markers is not allowed to be less than one tenth of a centimorgan, nor as much as one morgan, and markers are assumed to be in linkage equilibrium. Two test statistics are allowed for in the cases of sibling pairs, but only one (that based on the mean test) is implemented for designs that use both affected and discordant pairs.

#### 18.2 Theory

It is well understood that linkage of a putative disease locus to a polymorphic marker can be conducted through a study design of affected pairs of relatives, and this is usually the most powerful sampling strategy for binary traits (Blackwelder and Elston, 1985; Risch, 1990). However, recent research shows that, under certain situations, using discordant relative pairs can be as powerful as, or even more powerful than, using affected relative pairs. Moreover, combining discordant with affected relative pairs provides a more valid and reasonable study from both a theoretical and practical point of view (Guo and Elston, 2000). Specifically, linkage can be studied by typing pairs of relatives and examining the proportions of the pairs sharing 0, 1, or 2 alleles identical by descent (IBD) at the marker locus. The test for linkage in DESPAIR is based on either the proportion of pairs sharing 0 alleles IBD or the mean proportion of marker alleles shared IBD, which depend on the type of relative pair.

Denote the expected values of either of these proportions under the null hypothesis of free linkage  $\pi_0$ . If there is linkage, the expected values are  $\pi_0 + \delta_c$  and  $\pi_0 - \delta_d$ , corresponding to a design using affected relative pairs alone and a design using discordant relative pairs alone, respectively;  $\delta_c$  and  $\delta_d$  are the expected deviations respectively for affected pairs and discordant pairs due to linkage. Both these measures depend not only on the type of relative pair, but also on the recombination fraction  $\theta$  between the marker and disease loci. In addition,  $\delta_c$  depends on the relative recurrence risk of disease, due to the disease locus, to first degree relatives of affected persons:

 $\lambda = \frac{Pr(\text{first degree relative of affected person is affected})}{Pr(\text{random member of population is affected})}$ 

and  $\delta_d$  depends on the corresponding relative non-recurrence risk ratio for an affected-unaffected first degree relative pair:

$$\lambda^{-} = \frac{Pr(\text{first degree relative of affected person is unaffected})}{Pr(\text{random member of population is unaffected})}.$$

Each of these relative risks, often called risk ratios, can be to either a parent/offspring  $(\lambda_o, \lambda_o^-)$  or to a full sibling  $(\lambda_s, \lambda_s^-)$ .

If several disease loci act multiplicatively, the relative risk is the product of  $\lambda$ 's, one for each locus. For a study design that combines affected relative pairs with discordant relative pairs, the test statistic is based on the notion that, in the presence of linkage, affected relative pairs are expected to share a larger proportion of marker alleles IBD, whereas discordant relative pairs are expected to share a smaller proportion of alleles IBD. The difference in the proportion of alleles shared IBD between affected pairs and discordant pairs is quantified by  $\Delta$ , a weighted difference in the deviations of the mean proportions from  $\pi_0$ .  $\Delta$  equals zero under the null hypothesis of no linkage, and is greater than zero when linkage is present. The values of  $\Delta$  can be expressed as a function of  $\theta$ ,  $\lambda$ ,  $\lambda^-$ , and the ratio  $(r_p)$  of the number of affected relative pairs to the number of discordant relative pairs that are sampled. Values of  $\pi_0 + \delta_c$  were given by Risch (1990), and values of  $\pi_0 - \delta_d$ , and  $\Delta$ were given by Guo and Elston (2000), for five relative pairs: full sibling, half sibling, avuncular, grandparent-grandchild, and first cousin.

The test based on the proportion sharing 0 alleles IBD and the mean test give identical results except in the case of full sib pairs. The test based on the proportion sharing 0 alleles IBD is not implemented for designs using both concordant affected and discordant full sib pairs.

Assume that at a first stage, *m* fully informative markers, equally spaced along an autosomal genome *M* morgans long, are determined on *n* pairs of relatives of a particular type. For each marker, a onesided test is performed at the  $\alpha^*$  significance level to decide whether the sample proportion of alleles shared IBD deviates significantly from  $\pi_0$ , suggesting linkage. Around each marker suggesting linkage at the first stage, a further 2*k* fully informative markers are tested for linkage at a second stage, assuming that these are placed, *k* on either side of the first stage marker, to span in an optimal manner the interval of interest suggested by the significant first-stage marker (see Figure 18.1).

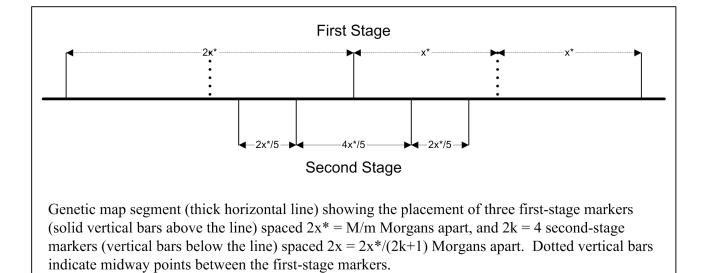


Figure 18.1: Stage-1 and Stage-2 Marker Placement

Assume that the study is designed to have power  $1 - \beta$  of detecting a disease locus with relative risk ratio  $\lambda$  at a significance level  $\alpha$  at the second stage, and that there are actually *d* such disease loci present. Finally, assume that the cost of recruiting a person into the study is *R* times the cost of determining one marker on one person. Under these assumptions, if at most one first stage marker is linked to any disease locus, the expected cost of the study is proportional to

$$2n\{R + m + 2k[\alpha^*m + (1 - \beta)d]\}$$
(18.1)

However, because there may be more than one first-stage marker linked to the disease locus, the total expected cost is more appropriately reflected by

$$C = 2n\{R + m + 2k[\alpha^*(m - \sum_{i=1}^d l_i) + \sum_{i=1}^d \sum_{j=1}^{l_i} (1 - \beta_{ij})]\}$$
(18.2)

where  $l_i$  is the number of first stage markers linked to disease locus *i*, and  $1 - \beta_{ij}$  is the probability that 2k second stage markers are typed around marker *j* that is linked to disease locus *i* (Ziegler et al. 2001). In this revised version of DESPAIR, which implements cost function (18.2), users have the option to input a maximum distance (g) between any disease locus and a "linked" marker. Then significant results obtained within *g* morgans from any disease locus are considered to be successes, and any outside that range are considered to be false positives. By making the distance *g* small in comparison to the distance between first stage markers, for a large number of markers cost function (18.2) approaches cost function (18.1), which was the function used in the original version of DESPAIR.

Given  $\alpha$ ,  $\beta$ ,  $\lambda$ , R, d, g, M, the type of relative pairs, and the type of data (affected relative pairs, discordant relative pairs, or both discordant and affected pairs: for the latter two cases,  $\lambda^-$  must also be specified; and for the last one case, the ratio  $(r_p)$  of the number of affected to the number

of discordant relative pairs to be sampled must also be specified), DESPAIR finds the values of m, n, and  $\alpha^*$  that minimize this expected cost for different mapping functions (linear, Kosambi's, and Haldane's), and for values of k from 0 (a one-stage design) to a specified maximum value of k, subject to the limitation M < m < 1000 (i.e., the markers must be spaced less than one morgan apart, and must be no closer than one tenth of a centimorgan apart). There is an option (c) to include the cost of screening the population to find the desired sample (the cost of screening is taken to be the same as the cost of recruiting), in which case the user must also enter the proportion of the screened population ( $r_s$ ) that becomes the final sample.

It is assumed that *n* is large enough, in determining the test criterion corresponding to  $\alpha^*$  and  $\beta$ , that the distribution of the proportion of pairs sharing 0 alleles IBD or the mean proportion of marker alleles shared IBD is normally distributed. However, in the case of  $\alpha$ , which is typically much closer to zero, there is the option of using either this same approximation assumption (the approximate method), or exact binomial distribution probabilities (the exact method, not implemented for the case where the sample includes both affected and discordant pairs).

To allow for less than fully informative markers, a value of the polymorphism information content (PIC), which measures the markers informativeness (assumed to be the same for all markers), can be specified. This is converted by the program to the corresponding type-of-pair-specific LIC value (Guo and Elston, 1999; Guo et al. 2002). Similarly, a fraction *h*, heterogeneity, can be specified that represents the proportion of the sample pairs affected due to causes other than segregation at the linked locus (in this case one would typically specify a large value for  $\lambda$  and/or a small value of  $\lambda^{-}$ ).

Further details of the method are given in the references.

# **18.3** Running the Program

DESPAIR can be run by clicking on the DESPAIR GUI link on the S.A.G.E. website

http://darwin.cwru.edu/despair/

and inputting one or more sets of parameters for which the sample size (numbers of affected sib pairs and/or number of markers) is desired. The parameters may be specified as follows:

parameter		Explanation
	Specifies relative p	air type.
		<b>S</b> (full siblings),
		G (grand-parental),
	Value Range	A (avuncular),
relative_pair_type		<b>H</b> (half siblings),
		C (first cousins)
	Default Value	S
	Required	Yes
	Applicable Notes	None
	Specifies the phene	otypic concordance status (adb) of
	the observations.	
		A (affected relative pairs)
annanrdanga tuma	Value Range	<b>D</b> (discordant relative pairs)
concordance_type		<b>B</b> (both)
	Default Value	A
	Required	Yes
	Applicable Notes	None
	Specifies the analysis method to be used.	
		A (approximate)
method	Value Range	E (exact)
meenod	Default Value	А
	Required	Yes
	Applicable Notes 1	1
	Specifies the statistical significance level $\alpha$ .	
	Value Range	(0, 1)
significance	Default Value	0.000101
	Required	Yes
	Applicable Notes	2
	Specifies the statist	tical power level $1 - \beta$ .
power	Value Range	(α, 1)
	Default Value	None
	Required	Yes
	Applicable Notes	None

	Specifies the type of test statistic to be employed (mp).
test_statistic	Value Range M (mean statistic)
	<b>P</b> (proportion statistic)
	Default Value M
	Required Yes
	Applicable Notes 3
	Specifies $\lambda_o$ , the locus-specific relative recurrence risk
	ratio of disease for an offspring of the affected person.
	Value Range $(1, +\infty)$ for ARPs
offspr_recurrence_risk	(0, 1) for DRPs
	Default Value None
	Required Yes
	Applicable Notes 3
	Specifies $\lambda_o^-$ , the locus-specific relative nonrecur-
	rence risk ratio of disease for an offspring of the af-
	fected person.
offspr_nonrecurrence_risk	Value Range (0, 1) for ARPs
	$(1, +\infty)$ for DRPs
	Default Value None
	Required Yes
	Applicable Notes 3
	Specifies $\lambda_s$ , the locus-specific relative recurrence risk
	ratio of disease for a sibling of the affected person.
	Value Range $(1, +\infty)$ for ASPs
sib_recurrence_risk	Value Range $(1, +\infty)$ for ASPs $(0, 1)$ for DSPs
sib_recurrence_risk	Value Range $(1, +\infty)$ for ASPs $(0, 1)$ for DSPsDefault ValueNone
sib_recurrence_risk	Value Range $(1, +\infty)$ for ASPs $(0, 1)$ for DSPsDefault ValueNoneRequiredYes
sib_recurrence_risk	Value Range $(1, +\infty)$ for ASPs $(0, 1)$ for DSPsDefault Value RequiredNoneApplicable Notes3
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	Value Range $(1, +\infty)$ for ASPs $(0, 1)$ for DSPsDefault Value RequiredNonePapelicable Notes3Specifies $\lambda_s^-$ , the locus-specific relative nonrecurrence risk ratio of disease for a sibling of the affected person.Value Range $(0, 1)$ for ASPs $(1, +\infty)$ for DSPsDefault Value RequiredNoneValue Range $(0, 1)$ for ASPs $(1, +\infty)$ for DSPsDefault Value RequiredNoneYes 33Specifies the ratio (R) of the cost of recruiting a person
	Value Range $(1, +\infty)$ for ASPs $(0, 1)$ for DSPsDefault Value RequiredNoneApplicable Notes3Specifies $\lambda_s^-$ , the locus-specific relative nonrecurrence risk ratio of disease for a sibling of the affected person.Value Range $(0, 1)$ for ASPs $(1, +\infty)$ for DSPsDefault Value Required Applicable NotesNoneValue Range $(0, 1)$ for ASPs $(1, +\infty)$ for DSPsDefault Value Required Applicable NotesNoneSpecifies the ratio (R) of the cost of recruiting a person to the cost of performing one marker assay.
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	Specifies the number (d) of disease loci being ana- lyzed.
num_loci	Value Range {1, 2, 3,}
	Default Value 1
	Required Yes
	Applicable Notes None
	Specifies the length (M), in morgans, of the underly-
	ing genome.
genome_length	$\frac{1}{\text{Value Range}} \left\{ 1, 2, 3, \right\}$
	Default Value $36$
	Required Yes
	Applicable Notes None
	Specifies the maximum distance (g), in morgans, be-
	tween any disease locus and a "linked" marker.
linked_distance	Value Range $(0, +\infty)$
	Default Value 0.4
	Required Yes
	Applicable Notes None
	Specifies the value of the polymorphism information
	content (PIC) used to constrain marker selection.
pic	Value Range (0, 1]
pic	Default Value 1
	Required Yes
	Applicable Notes None
	Specifies the heterogeneity proportion (h) of sample
	pairs affected due to causes other than segregation at
	the linked locus.
heterogeneity	Value Range [0, 1)
	Default Value 0
	Required Yes
	Applicable Notes None
	Specifies option (c) to include the cost of screening
	the population to obtain desired pairs.
	<b>V</b> (include the cost)
screening_cost	Value Range $\mathbf{N}$ (do not include the cost)
	Default Value N
	Required Yes
	Applicable Notes None
	Specifies proportion of collected samples in the
	screened population $(r_s)$
screened_proportion	$\frac{   }{   } \frac{   }{   } \frac{   }{   } \frac{   }{   }$
	Default Value 1
	•
	Applicable Notes 4

	Specifies the ratio $(r_p)$ of concordantly affected to dis-
conc_disc_ratio	cordant relative pairs to be sampled.
	Value Range $(0, +\infty)$
	Default Value None
	Required Yes
	Applicable Notes 5
	Specifies the number (m) of first-stage markers to be
	used.
	Value Range $\{M + 1, M + 2, 1000M\}$
num_stage_one_markers	Default Value None
	Required No
	Applicable Notes 6
	Specifies the maximum value for the number of mark-
	ers (k) to be typed, during the second stage, on each
	side of the markers found to be significant during the
num stage two membrans	first stage.
num_stage_two_markers	Value Range {0, 1, 2,}
	Default Value None
	Required Yes
	Applicable Notes None
	Specifies the number of relative pairs (n) to be ana-
num_pairs	lyzed.
	Value Range {1, 2, 3,}
	Default Value None
	Required No
	Applicable Notes 6

Notes

- 1. The method parameter is not applicable for sample data comprising both affected pairs and discordant pairs; only the approximate method (A) is implemented for such data.
- 2. The default value for  $\alpha$  corresponds to a lod score of 3 if the method parameter is set to A (approximate).
- 3. The parameters offspr\_recurrence\_risk and offspr\_nonrecurrence\_risk are used by the proportion test for linkage, while the parameters sib\_recurrence\_risk and sib\_nonrecurrence\_risk are used by the mean test.
- 4. When the value of the screening\_cost parameter is set to **N**, the screened\_proportion parameter will be ignored.
- 5. When the value of the screening\_cost parameter is set to **N**, or the concordance\_type parameter is set to either A or D, the conc\_disc\_ratio parameter will be ignored. In other words, the conc\_disc\_ratio parameter is applicable only when the concordance\_type parameter is set to **B**.
- 6. The user may specify a value for either num\_stage\_one\_markers or num\_pairs, but not both. If a value for either one of the parameters is specified, the other will be determined by the program. If neither parameter is specified, the program will determine both.

## 18.4 Output

DESPAIR produces a Standard Output File that includes:

- Title, version, and date of the program for each problem
- Control values specified by user
- For each k = 0, ..., max k, and for each mapping function, tabulation of optimal values of m and n with corresponding  $\alpha^*$ , cost (in units of the cost of typing one marker on one person), and the first and second stage marker spacings in centimorgans

For an example output, see the end of this document.

#### **18.4.1** Error Messages

DESPAIR has an error checking routine. Values of any parameter that are out of bounds are not allowed. When an error is detected during the analysis, DESPAIR will identify the error and display the error message associated with it. The error messages that may be displayed are as follows:

- The following fields were set to values out of bounds: <FIELD LIST>
- The exact test is not implemented for the case in which both concordant and discordant pairs are available.
- The test based on the proportion sharing 0 alleles i.b.d. is not available. The above results are for the mean test.

# **Chapter 19**

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