Source and Accuracy Statement for the October 2004 CPS Microdata File on School Enrollment

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Source and Accuracy of Estimates for the October 2004 CPS Microdata File on School Enrollment

SOURCE OF DATA

The data in this microdata file are from the October 2004 Current Population Survey (CPS). The Census Bureau conducts the CPS every month, although this file has only October data. The October survey uses two sets of questions, the basic CPS and a set of supplemental questions. The CPS, sponsored jointly by the Census Bureau and the U.S. Bureau of Labor Statistics, is the country's primary source of labor force statistics for the entire population. The supplemental questions for October are sponsored jointly by the Census Bureau and the National Center for Education Statistics.

<u>Basic CPS</u>. The monthly CPS collects primarily labor force data about the civilian noninstitutional population living in the United States. The institutionalized population, which is excluded from the population universe, is composed primarily of the population in correctional institutions and nursing homes (91 percent of the 4.1 million institutionalized people in Census 2000). Interviewers ask questions concerning labor force participation about each member 15 years old and over in sample households.

The CPS uses a multistage probability sample based on the results of the decennial census, with coverage in all 50 states and the District of Columbia. When files from the most recent decennial census become available, the Census Bureau gradually introduces a new sample design for the CPS¹.

In April 2004, the Census Bureau began phasing out the 1990 sample and replacing it with the 2000 sample, creating a mixed sampling frame. Two simultaneous changes will occur during this phase-in period. First, primary sampling units (PSUs)² selected for only the 2000 design will gradually replace those selected for the 1990 design. This will involve 10 percent of the sample. Second, within PSUs selected for both the 1990 and 2000 designs, sample households from the 2000 design will gradually replace sample households from the 1990 design. This will involve about 90 percent of the entire sample. By July 2005, the new sample design will be completely implemented, and the sample will come entirely from Census 2000 files.

In the first stage of the sampling process, PSUs are selected for sample. In the 1990 design, the United States was divided into 2,007 PSUs. These were grouped into 754 strata, and one PSU was selected for sample from each stratum. In the 2000 sample design, the United States is divided into 2,025 PSUs. These PSUs are grouped into 824 strata. Within each stratum, a single PSU is chosen for the sample, with its probability of selection proportional to its population as of the most recent decennial census. This PSU represents the entire stratum from which it was selected. In the case of strata consisting of only one PSU, the PSU is chosen with certainty.

For detailed information on the 1990 sample redesign, see the Department of Labor, Bureau of Labor Statistics report, *Employment and Earnings*, Volume 41 Number 5, May 1994.

² The PSUs correspond to substate areas, counties or groups of counties that are geographically contiguous.

The 1990 design and 2000 design stratum numbers are not directly comparable since the 1990 design contained some PSUs in New England and Hawaii that were based on minor civil divisions instead of counties, while the PSUs in the 2000 design are strictly county-based. The PSUs have also been redefined to correspond to the new Office of Management and Budget (OMB) definitions of Core-Based Statistical Area definitions and to improve efficiency in field operations.

Approximately 72,000 housing units were selected for sample from the mixed sampling frame in October. Based on eligibility criteria, 11 percent of these housing units were sent directly to Computer-Assisted Telephone Interviewing (CATI). The remaining units were assigned to interviewers for Computer-Assisted Personal Interviewing (CAPI).³ Of all housing units in sample, about 60,000 were determined to be eligible for interview. Interviewers obtained interviews at about 55,000 of these units. Noninterviews occur when the occupants are not found at home after repeated calls or are unavailable for some other reason.

<u>October 2004 Supplement</u>. In October 2004, in addition to the basic CPS questions, interviewers asked supplementary questions of household members three years old and over on school enrollment.

Estimation Procedure. This survey's estimation procedure adjusts weighted sample results to agree with independently derived population estimates of the civilian noninstitutional population of the United States and states (including the District of Columbia). These population estimates, used as controls for the CPS, are prepared annually to agree with the most current set of population estimates that are released as part of the Census Bureau's population estimates and projections program.

The population controls for the nation are distributed by demographic characteristics in two ways:

- Age, sex, and race (White alone, Black alone, Asian alone, and all other groups combined), and
- Age, sex, and Hispanic origin.

The projections for the states are distributed by race (Black alone and all other race groups combined), age (0-15, 16-44, and 45 and over), and sex.

The independent estimates by age, sex, race, and Hispanic origin and for states by selected age groups and broad race categories are developed using the basic demographic accounting formula whereby the population from the latest decennial data is updated using data on the components of population change (births, deaths, and net international migration) with net internal migration as an additional component in the state population estimates.

For further information on CATI and CAPI and the eligibility criteria, please see: Technical Paper 63RV, *The Current Population Survey: Design and Methodology*, U.S. Census Bureau, U.S. Department of Commerce, 2002. (http://www.census.gov/prod/2002pubs/tp63rv.pdf)

The net international migration component in the population estimates includes a combination of:

- Legal migration to the United States,
- Emigration of foreign born and native people from the United States,
- Net movement between the United States and Puerto Rico,
- Estimates of temporary migration, and
- Estimates of net residual foreign-born population, which include unauthorized migration.

Because the latest available information on these components lag the survey date, it is necessary to make short-term projections of these components to develop the estimate for the survey date.

ACCURACY OF THE ESTIMATES

A sample survey estimate has two types of error: sampling and nonsampling. The accuracy of an estimate depends on both types of error. The nature of the sampling error is known given the survey design; the full extent of the nonsampling error is unknown.

<u>Sampling Error</u>. Since the CPS estimates come from a sample, they may differ from figures from an enumeration of the entire population using the same questionnaires, instructions, and enumerators. For a given estimator, the difference between an estimate based on a sample and the estimate that would result if the sample were to include the entire population is known as sampling error. Standard errors, as calculated by methods described in "Standard Errors and Their Use," are primarily measures of the magnitude of sampling error. However, they may include some nonsampling error.

Nonsampling Error. For a given estimator, the difference between the estimate that would result if the sample were to include the entire population and the true population value being estimated is known as nonsampling error. Sources of nonsampling errors include the following:

- Inability to get information about all sample cases (nonresponse).
- Definitional difficulties.
- Differences in the interpretation of questions.
- Respondent inability or unwillingness to provide correct information.
- Respondent inability to recall information.
- Errors made in data collection, such as recording and coding data.
- Errors made in processing the data.
- Errors made in estimating values for missing data.
- Failure to represent all units with the sample (undercoverage).

The Census Bureau employs quality control procedures throughout the production process including the overall design of surveys, the wording of questions, the review of the work of interviewers and coders, and the statistical review of reports to minimize these errors.

Two types of nonsampling error that can be examined to a limited extent are nonresponse and undercoverage.

<u>Nonresponse</u>. The effect of nonresponse cannot be measured directly, but one indication of its potential effect is the nonresponse rate. For the October 2004 basic CPS, the nonresponse rate was 7.7 percent. The nonresponse rate for the School Enrollment supplement was an additional 4.0 percent. These two nonresponse rates lead to a combined supplement nonresponse rate of 11.4 percent.

<u>Coverage</u>. The concept of coverage in the survey sampling process is the extent to which the total population that could be selected for sample "covers" the survey's target population. CPS undercoverage results from missed housing units and missed people within sample households. Overall CPS undercoverage for October 2004 is estimated to be about 11 percent. CPS undercoverage varies with age, sex, and race. Generally, undercoverage is larger for males than for females and larger for Blacks than for non-Blacks.

The CPS weighting procedure partially corrects for bias due to undercoverage, but biases may still be present when people who are missed by the survey differ from those interviewed in ways other than age, race, sex, Hispanic ancestry, and state of residence. How this weighting procedure affects other variables in the survey is not precisely known. All of these considerations affect comparisons across different surveys or data sources.

A common measure of survey coverage is the coverage ratio, calculated as the estimated population before post-stratification divided by the independent population control. Table 1 shows October 2004 CPS coverage ratios by age and sex for certain race and Hispanic groups. The CPS coverage ratios can exhibit some variability from month to month.

	Table 1. CPS Coverage Ratios: October 2004										
<u>Totals</u>			Whit	e Only	Black	k Only	Residu	ıal Race	His	<u>panic</u>	
Age Group	All People	Male	Female	Male	Female	Male	Female	Male	Female	Male	Female
0-15	0.93	0.92	0.93	0.94	0.95	0.83	0.80	0.94	0.98	0.93	0.96
16-19	0.88	0.88	0.88	0.88	0.90	0.91	0.72	0.82	0.95	0.93	0.98
20-24	0.78	0.77	0.80	0.79	0.80	0.61	0.78	0.79	0.78	0.74	0.84
25-34	0.87	0.84	0.89	0.86	0.91	0.70	0.82	0.80	0.85	0.82	0.91
35-44	0.90	0.87	0.93	0.88	0.95	0.77	0.88	0.87	0.88	0.82	0.91
45-54	0.91	0.90	0.93	0.92	0.94	0.78	0.86	0.93	0.90	0.86	0.90
55-64	0.95	0.96	0.94	0.95	0.94	0.90	0.90	1.09	1.00	0.89	0.88
65+	0.93	0.93	0.93	0.93	0.93	0.94	0.91	0.88	0.95	0.81	0.86
15+	0.90	0.88	0.91	0.89	0.92	0.79	0.85	0.88	0.90	0.83	0.90
0+	0.90	0.89	0.91	0.90	0.93	0.80	0.83	0.89	0.92	0.86	0.92

Notes: (1) The Residual Race group includes cases indicating a single race other than White or Black, and cases indicating two or more races.

(2) Hispanics may be of any race.

<u>Comparability of Data</u>. Data obtained from the CPS and other sources are not entirely comparable. This results from differences in interviewer training and experience and in differing survey processes. This is an example of nonsampling variability not reflected in the standard errors. Therefore, caution should be used when comparing results from different sources.

Caution should also be used when comparing the data from this microdata file, which reflects Census 2000-based controls, with microdata files from March 1994 through December 2001, which reflect 1990 census-based controls. Caution should also be used when comparing the data from this microdata file to certain microdata files from 2002, namely June, October, and November, which contain both Census 2000-based estimates and 1990 census-based estimates. When comparing estimates, the same controls should be used when possible. Microdata files from previous years reflect the latest available census-based controls. Although this change in population controls had relatively little impact on summary measures such as averages, medians, and percentage distributions, it did have a significant impact on levels. For example, use of Census 2000-based controls results in about a one percent increase from the 1990 census-based controls in the civilian noninstitutional population and in the number of families and households. Thus, estimates of levels for data collected 2002 and later years will differ from those for earlier years by more than what could be attributed to actual changes in the population. These differences could be disproportionately greater for certain subpopulation groups than for the total population.

Users should also exercise caution due to changes caused by the phase-in of the Census 2000 files. During this time period, CPS data are collected from sample designs based on different censuses. Three features of the new CPS design have the potential of affecting published estimates: (1) the temporary disruption of the rotation pattern from August 2004 through June 2005 for a comparatively small portion of the sample, (2) the change in sample areas, and (3) the introduction of the new Core-Based Statistical Areas (formerly called metropolitan areas). Most of the known effect on estimates during and after the sample redesign will be the result of changing from 1990 to 2000 geographic definitions. Research has shown that the national-level estimates of the metropolitan and nonmetropolitan populations should not change appreciably because of the new sample design. However, users should still exercise caution when comparing metropolitan and nonmetropolitan estimates across years with a design change, especially at the state level.

Caution should also be used when comparing Hispanic estimates over time. No independent population control totals for people of Hispanic ancestry were used before 1985.

A Nonsampling Error Warning. Since the full extent of the nonsampling error is unknown, one should be particularly careful when interpreting results based on small differences between estimates. Even a small amount of nonsampling error can cause a borderline difference to appear significant or not, thus distorting a seemingly valid hypothesis test. Caution should also be used when interpreting results based on a relatively small number of cases. Summary measures probably do not reveal useful information when computed on a subpopulation smaller than 75,000.

For additional information on nonsampling error including the possible impact on CPS data when known, refer to

- Statistical Policy Working Paper 3, *An Error Profile: Employment as Measured by the Current Population Survey*, Office of Federal Statistical Policy and Standards, U.S. Department of Commerce, 1978. (http://www.fcsm.gov/working-papers/spp.html)
- Technical Paper 63RV, The Current Population Survey: Design and Methodology, U.S. Census Bureau, U.S. Department of Commerce, 2002. (http://www.census.gov/prod/2002pubs/tp63rv.pdf)

Standard Errors and Their Use. A number of approximations are required to derive, at a moderate cost, standard errors applicable to all the estimates in this microdata file. Instead of providing an individual standard error for each estimate, parameters are provided to calculate standard errors for various types of characteristics. These parameters are listed in Tables 2 and 3. Also, tables are provided that allow the calculation of parameters for prior years and parameters for U.S. regions. Tables 4 and 5 provide factors to derive prior year parameters. Table 6 provides factors to derive U.S. regional parameters.

The sample estimate and its standard error enable one to construct a confidence interval. A confidence interval is a range that would include the average result of all possible samples with a known probability. For example, if all possible samples were surveyed under essentially the same general conditions and using the same sample design, and if an estimate and its standard error were calculated from each sample, then approximately 90 percent of the intervals from 1.645 standard errors below the estimate to 1.645 standard errors above the estimate would include the average result of all possible samples.

A particular confidence interval may or may not contain the average estimate derived from all possible samples. However, one can say with specified confidence that the interval includes the average estimate calculated from all possible samples.

Standard errors may also be used to perform hypothesis testing, a procedure for distinguishing between population parameters using sample estimates. The most common type of hypothesis is that the population parameters are different. An example of this would be comparing the percentage of men who were part-time workers to the percentage of women who were part-time workers.

Tests may be performed at various levels of significance. A significance level is the probability of concluding that the characteristics are different when, in fact, they are the same. For example, to conclude that two characteristics are different at the 0.10 level of significance, the absolute value of the estimated difference between characteristics must be greater than or equal to 1.645 times the standard error of the difference.

The Census Bureau uses 90-percent confidence intervals and 0.10 levels of significance to determine statistical validity. Consult standard statistical textbooks for alternative criteria.

<u>Estimating Standard Errors</u>. The Census Bureau uses replication methods to estimate the standard errors of CPS estimates. These methods primarily measure the magnitude of sampling error. However, they do measure some effects of nonsampling error as well. They do not measure systematic biases in the data due to nonsampling error. Bias is the average over all possible samples of the differences between the sample estimates and the true value.

Generalized Variance Parameters. While it is possible to compute and present an estimate of the standard error based on the survey data for each estimate in a report, there are a number of reasons why this is not done. A presentation of the individual standard errors would be of limited use, since one could not possibly predict all of the combinations of results that may be of interest to data users. Additionally, variance estimates are based on sample data and have variances of their own. Therefore, some methods of stabilizing these estimates of variance, for example, by generalizing or averaging over time, may be used to improve their reliability.

Experience has shown that certain groups of estimates have similar relationships between their variances and expected values. Modeling or generalizing may provide more stable variance estimates by taking advantage of these similarities. The generalized variance function is a simple model that expresses the variance as a function of the expected value of the survey estimate. The parameters of the generalized variance function are estimated using direct replicate variances. These generalized variance parameters provide a relatively easy method to obtain approximate standard errors for numerous characteristics. In this source and accuracy statement, Table 2 provides the generalized variance parameters for labor force estimates, Table 3 provides the parameters for school enrollment estimates, and Tables 4, 5, and 6 provide factors for use with the parameters.

<u>Standard Errors of Estimated Numbers</u>. The approximate standard error, s_x , of an estimated number, with the exception of school enrollment estimates, from this microdata file can be obtained by using the formula:

$$s_x = \sqrt{ax^2 + bx} \tag{1}$$

Here x is the size of the estimate and a and b are the parameters in Table 2 associated with the particular type of characteristic. When calculating standard errors from cross-tabulations involving different characteristics, use the set of parameters for the characteristic that will give the largest standard error.

Illustration No. 1

Suppose there were 4,200,000 unemployed men (ages 16 and up) in the civilian labor force (CLF). Use the appropriate parameters from Table 2 and Formula (1) to get

Illustration 1				
Number of unemployed males in the CLF (x)	4,200,000			
A parameter (a)	-0.000032			
B parameter (b)	2,971			
Standard error	109,000			
90% confidence interval	4,021,000 to 4,379,000			

The standard error is calculated as

$$s_x = \sqrt{-0.000032 \times 4,200,000^2 + 2,971 \times 4,200,000} = 109,000$$

The 90-percent confidence interval is calculated as $4,200,000 \pm 1.645 \times 109,000$.

A conclusion that the average estimate derived from all possible samples lies within a range computed in this way would be correct for roughly 90 percent of all possible samples.

<u>Standard Errors of Estimated School Enrollment Numbers</u>. The approximate standard error, s_x , of an estimated school enrollment number from this microdata file can be obtained by using the formula:

$$s_{x} = \sqrt{-\left(\frac{b}{T}\right)x^{2} + bx} \tag{2}$$

Here x is the size of the estimate, T is the total number of persons in a specific age group and b is the parameter in Table 3 associated with the particular type of characteristic. If T is not known, for Total or White use 100,000,000; for Blacks and Hispanic use 10,000,000. When calculating standard errors for numbers from cross-tabulations involving different characteristics, use the set of parameters for the characteristic that will give the largest standard error.

Illustration No. 2

Suppose there were 4,600,000 three and four year olds enrolled in school and 8,400,000 total children in that age group. Use the appropriate b parameter from Table 3 and Formula (2) to get

Illustration 2	
Number of three and four year olds enrolled in school (x)	4,600,000
Total (T)	8,400,000
B parameter (b)	2,453
Standard error	71,000
90% confidence interval	4,483,000 to 4,717,000

The standard error is calculated as

$$s_x = \sqrt{-\left(\frac{2,453}{8,400,000}\right) \times 4,600,000^2 + 2,453 \times 4,600,000} = 71,000$$

The 90-percent confidence interval for this estimate is approximately 4,483,000 to 4,717,000 (i.e., $4,600,000 \pm 1.645 \times 71,000$). Therefore, a conclusion that the average estimate derived from all possible samples lies within a range computed in this way would be correct for roughly 90 percent of all possible samples.

Standard Errors of Estimated Percentages. The reliability of an estimated percentage, computed using sample data for both numerator and denominator, depends on both the size of the percentage and its base. Estimated percentages are relatively more reliable than the corresponding estimates of the numerators of the percentages, particularly if the percentages are 50 percent or more. When the numerator and denominator of the percentage are in different categories, use the parameter from Table 2 or 3 as indicated by the numerator.

The approximate standard error, $s_{x,p}$, of an estimated percentage can be obtained by using the formula:

$$s_{x,p} = \sqrt{\frac{b}{x} p(100 - p)}$$
 (3)

Here x is the total number of people, families, households, or unrelated individuals in the base of the percentage, p is the percentage ($0 \le p \le 100$), and b is the parameter in Table 2 or 3 associated with the characteristic in the numerator of the percentage.

Illustration No. 3

Suppose there were 16,000,000 people aged 18 to 21, and 47.0 percent were enrolled in college. Use the appropriate parameter from Table 3 and Formula (3) to get

Illustration 3	
Percentage of people aged 18-21 enrolled in college (p)	47.0
Base (x)	16,000,000
b parameter (b)	2,131
Standard error	0.58
90% confidence interval	46.05 to 47.95

The standard error is calculated as

$$s_{x,p} = \sqrt{\frac{2,131}{16,000,000} \times 47.0 \times (100.0 - 47.0)} = 0.58$$

The 90 percent confidence interval for the estimated percentage of people aged 18 to 21 enrolled in college is from 46.05 to 47.95 percent (i.e., $47.0 + 1.645 \times 0.58$).

Standard Errors of Estimated Differences. The standard error of the difference between two sample estimates is approximately equal to

$$s_{x-y} = \sqrt{s_x^2 + s_y^2} \tag{4}$$

where s_x and s_y are the standard errors of the estimates, x and y. The estimates can be numbers, percentages, ratios, etc. This will result in accurate estimates of the standard error of the same characteristic in two different areas, or for the difference between separate and uncorrelated characteristics in the same area. However, if there is a high positive (negative) correlation

between the two characteristics, the formula will overestimate (underestimate) the true standard error.

Illustration No. 4

Suppose that of the 7,300,000 employed men between 20-24 years of age, 24 percent were part-time workers, and of the 6,500,000 employed women between 20-24 years of age, 39 percent were part-time workers. Use the appropriate parameters from Table 2 and Formulas (3) and (4) to get

Illustration 4						
	Male (x)	Female (y)	Difference			
Percentage working part-time (p)	24.0	39.0	15.0			
Number	7,300,000	6,500,000	-			
b parameter (b)	2,971	2,782	-			
Standard error	0.86	1.01	1.33			
90% confidence interval	22.58 to 25.42	37.34 to 40.66	12.82 to 17.18			

The standard error of the difference is calculated as

$$s_{x-y} = \sqrt{0.86^2 + 1.01^2} = 1.33$$

The 90-percent confidence interval around the difference is calculated as $15.0 \pm 1.645 \times 1.33$. Since this interval does not include zero, we can conclude with 90 percent confidence that the percentage of part-time women workers between 20-24 years of age is greater than the percentage of part-time men workers between 20-24 years of age.

<u>Standard Errors of Quarterly or Yearly Averages</u>. For information on calculating standard errors for labor force data from the CPS which involve quarterly or yearly averages, see "Explanatory Notes and Estimates of Error: Household Data" in *Employment and Earnings*, a monthly report published by the U.S. Bureau of Labor Statistics.

Table 2. Parameters for Computation of Standard Errors for Labor Force Characteristics: October 2004					
Characteristic	a	b			
Total or White					
Civilian Labor Force, Employed	-0.000016	3,068			
Not in Labor Force	-0.000009	1,833			
Unemployed	-0.000016	3,096			
Total or White					
Civilian Labor Force, Employed, Not in L					
Men	-0.000032	2,971			
Women	-0.000031	2,782			
Both sexes, 16 to 19 years	-0.000022	3,096			
Black or African American					
Civilian Labor Force, Employed, Not in L	abor Force, and Unemplo				
Total	-0.000151	3,455			
Men	-0.000311	3,357			
Women	-0.000252	3,062			
Both sexes, 16 to 19 years	-0.001632	3,455			
Hispanic or Latino ethnicity					
Civilian Labor Force, Employed, Not in Labor Force, and Unemployed					
Total	-0.000141	3,455			
Men	-0.000253	3,357			
Women	-0.000266	3,062			
Both sexes, 16 to 19 years	-0.001528	3,455			
Asian Civilian Labor Force, Employed, Not in Labor Force, and Unemployed					
Total	-0.000346	3,198			
Men	-0.000729	3,198			
Women	-0.000659	3,198			
Both sexes, 16 to 19 years	-0.004146	3,198			

Notes: (1) These parameters are to be applied to basic CPS monthly labor force estimates. (2) For foreign-born and noncitizen characteristics for Total and White, the a and b parameters should be multiplied by 1.3. No adjustment is necessary for foreign-born and noncitizen characteristics for Blacks, Hispanics, and Asians.

Table 3. Parameters for Computation of Standard Errors for School Enrollment Characteristics: October 2004					
		b			
Characteristics	Total or White	Black	API, AIAN, NH, and OPI	Hispanic	
People					
Persons Enrolled in School:					
Total	2,131	2,410	2,410	2,744	
Children 13 and under	2,453	2,775	2,775	3,159	
Marital Status, Household and Family Characteristics, Health Insurance					
Some household members	4,687	6,733	6,733	11,347	
All household members	5,695	9,929	9,929	16,733	
Families, Households, or Unrelated Individuals					
Income, Earnings	2,016	2,201	2,201	3,709	
Marital Status, Household and Family Characteristics, Educational Attainment, Population by Age and/or Sex	1,860	1,683	1,683	2,836	

Notes: (1) These parameters are to be applied to the October 2004 School Enrollment Supplement data.

- (2) API, AIAN, NH, and OPI are Asian and Pacific Islander, American Indian and Alaska Native, Native Hawaiian, and Other Pacific Islander, respectively.
- (3) Hispanics may be of any race.
- (4) The Total or White, Black, and API parameters are to be used for both "alone" and "in combination" race group estimates.
- (5) These b parameters should be multiplied by 1.5 for nonmetropolitan residence categories.
- (6) These b parameters should be multiplied by the factors in Table 6 for regional data.

In 1994, we calculated school enrollment parameters directly from the 1994 CPS data. Since that time, the school enrollment parameters have been based on these updated parameters. Therefore, when calculating past school enrollment parameters, a separate set of year factors should be used.

Table 4 shows the prior year factors to apply to the **Non-School Enrollment** parameters while Table 5 shows prior year factors to apply to **School Enrollment** parameters.

Table 4. Year Factors for Non-School Enrollment Characteristics (1942-2003)				
Time Period	Total or White	Black		Hispanic
Time Period	a and b parameters	a parameter	b parameter	a and b parameters
January 2003 - Present	1.00	1.00	1.00	1.00
July 2001 - December 2002 (SCHIP)	1.00	1.20	1.00	1.00
January 1996 - June 2001 (Non-SCHIP)	1.11	1.33	1.11	1.11
April 1989 - December 1995	1.03	1.23	1.03	1.03
April 1988 - March 1989	1.14	1.37	1.14	1.20
January 1985 – March 1988	0.96	1.15	0.96	0.96
January 1982 - December 1984	0.96	1.15	0.96	1.35
March 1973 - December 1981	0.86	1.03	0.86	1.20
January 1967 - February 1973	0.86	1.03	0.86	1.20
May 1956 - December 1966	1.29	1.55	1.29	1.81
August 1942 - April 1956	1.93	2.32	1.93	2.71

Note: These factors are for use with the 2004 non-School Enrollment a and b parameters.

Table 5. Year Factors for School Enrollment Characteristics (1945-2003)					
Time Period	Total or White	Black		Hispanic	
Time Period	a and b	a	b	a and b	
	parameters	parameter	parameter	parameters	
January 2003 - Present	1.00	1.00	1.00	1.00	
July 2001 - December 2002 (SCHIP)	1.00	1.20	1.00	1.00	
January 1996 - June 2001 (Non-SCHIP)	1.11	1.33	1.11	1.11	
March 1995 – December 1995	1.03	1.23	1.03	1.03	
April 1989 – February 1995	1.19	1.70	1.42	2.10	
April 1988 - March 1989	1.32	1.89	1.58	2.45	
January 1985 - March 1988	1.11	1.60	1.33	1.97	
January 1982 - December 1984	1.11	1.60	1.33	2.76	
March 1973 - December 1981	0.99	1.43	1.19	2.46	
January 1967 - February 1973	0.99	1.43	1.19	2.46	
May 1956 - December 1966	1.49	2.14	1.78	3.69	
October 1945 - April 1956	2.24	3.21	2.67	5.54	

Note: These factors are for use with the 2004 School Enrollment 'b' parameters.

Table 6 provides the U.S. regional factors to apply to parameters in order to calculate standard errors for U.S. regional estimates.

Table 6. Regional Factors to Apply To 2004 Parameters				
Type of Characteristic Factor				
U. S. Totals	1.00			
Regions:				
Northeast	0.90			
Midwest	0.93			
South	1.14			
West	1.14			