## Sample Size, Power and Effect Size

Below we provide estimates of power and effect size for various genetic tests and subgroups within the Framingham Heart Study. Estimates are provided for the following situations:

1) linkage analyses with the 330 pedigrees included in the genome scan with microsatellite markers, conducted by the Mammalian Genotyping Service,
2) association analyses for unrelated individuals with varying sample sizes
3) association analyses for family-based association tests

## Linkage Analyses in 330 Pedigrees

All power calculations for linkage were performed using SOLAR v1.6.6 (Almasy and Blangero, 1998). We evaluated the power of a likelihood-ratio test for linkage for a range of QTL heritabilities ( $5 \%$ to $40 \%$ ) and various sample sizes. The different sample sizes will reflect the variation in the availability of data for different phenotypes. Three sample sizes, ranging from the largest possible size (e.g. for a trait measured in all Framingham participants such as blood pressure) to smaller sizes (e.g. traits measured only in the Offspring Cohort or a subset of Offspring) were considered within the existing 330 pedigrees. For these samples we computed power for three critical values (LOD scores 1.0, 2.0 and 3.0). These critical values encompass LOD scores that are suggestive of linkage ( $\mathrm{LOD} \geq 1.0$ or 1.5 ) or establish significant linkage $(\operatorname{LOD} \geq 3)$. Given a selected critical value, an approximate value of the power of the likelihoodratio test is computed using the non-central $\chi^{2}$ distribution whose non-centrality parameter is twice the difference between the expected log-likelihoods of the null and alternative hypotheses of linkage (Sham et al., 2000). In this analysis the non-centrality parameters were estimated by means of simulation using our pedigrees.

For each QTL heritability, 100 data sets were simulated using the 330 pedigree structure and a genetic model containing a marker with 10 alleles of equal frequency ( 0.90 heterozygosity) that was linked to a diallelic QTL with $10 \%$ allelic frequency. We evaluated other allele frequencies, but found that all were essentially equivalent since the primary determinant of power here is the proportion of variance that the QTL explains. Simulation was performed using the "simqtl" command in SOLAR. First, the marker and the phenotypic data (under the assumed QTL variance) were generated and the IBD probabilities were computed for all observations in the pedigrees. For a given sample the non-centrality parameter of the $\chi^{2}$ distribution was estimated from a twopoint linkage analysis of that sample.

Tables 1 presents the expected lod score (ELOD, the average twopoint LOD score from 100 simulated data sets) and power computed for the various samples and levels of QTL heritability in the 330 families. In these tables the three columns for power represent different critical values. For example, the estimated power for sample size $\mathrm{N}=2885$ in 330 pedigrees and $20 \%$ QTL heritability was $99 \%, 92 \%$ and $76 \%$ if linkage is deemed significant for LOD $\geq 1.0$, $\mathrm{LOD} \geq 2.0$ and $\mathrm{LOD} \geq 3.0$, respectively.

Table 1: Power for Linkage Analyses in 330 Pedigrees

| QTL <br> Heritabilit | Original Cohort and Offspring [ $\mathrm{N}=2885$, $\mathrm{Ped}=330$ ] |  |  |  | $\begin{gathered} \text { Offspring } \\ {[\mathrm{N}=1672, \text { Ped }=330]} \end{gathered}$ |  |  |  | $\begin{gathered} \text { Offspring Subset } \\ {[\mathrm{N}=1228, \text { Ped=326] }} \end{gathered}$ |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | ELOD LOD $\geq 1$ LOD $\geq 2$ LOD $\geq 3$ |  |  |  | ELOD LOD $\geq 1$ LOD $\geq 2$ LOD $\geq 3$ |  |  |  | ELOD LOD $\geq 1$ LOD $\geq 2$ LOD $\geq 3$ |  |  |  |
| y |  |  |  |  |  |  |  |  |  |  |  |  |
| (\%) |  |  |  |  |  |  |  |  |  |  |  |  |
| 5 | 0.39 | 21 | 4 | 1 | 0.30 | 16 | 3 | 1 | 0.21 | 12 | 2 | 0 |
| 10 | 1.20 | 58 | 25 | 9 | 0.83 | 43 | 14 | 4 | 0.51 | 27 | 7 | 1 |
| 15 | 2.47 | 89 | 63 | 36 | 1.68 | 74 | 40 | 17 | 0.98 | 49 | 18 | 6 |
| 20 | 4.22 | 99 | 92 | 76 | 2.84 | 93 | 72 | 46 | 1.63 | 72 | 38 | 16 |
| 25 | 6.50 | 100 | 99 | 96 | 4.36 | 99 | 93 | 78 | 2.47 | 89 | 63 | 37 |
| 30 | 9.35 | 100 | 100 | 100 | 6.26 | 100 | 99 | 95 | 3.53 | 97 | 84 | 62 |
| 35 | 12.87 | 100 | 100 | 100 | 8.60 | 100 | 100 | 99 | 4.84 | 99 | 95 | 84 |

*For 40\% QTL heritability, power greater than 95\%

## Association Studies with Biologically Unrelated Subjects

Table 2 displays detectable effect sizes ( $\Delta \mathrm{R}^{2}$ or percent variation explained by a QTL SNP) for a quantitative trait measured on unrelated subjects for power of $80 \%$ or $90 \%$ and alpha levels of 0.010 .001 and 0.0001 . These calculations are performed for varying samples sizes, from 1800 to 800 . The numbers for 1800 correspond to genotyping the unrelated plate set and all subjects having the trait. Thus, for $1 \%$ significance level, if background (baseline) covariates explain $30 \%$ of the variation in the trait ( $\mathrm{R}^{2}$ base), then a sample of size 1800 will have $80 \%$ power to detect a QTL variance of 0.0054 and $90 \%$ power to detect a QTL variance of 0.0068 .

Table 2: QTL variance sufficient for power $\mathbf{0 . 8 0}(\mathbf{0 . 9 0})$ in association analyses of unrelated subjects Analysis with General Linear Model having unrestricted genotype means [u(AA), u(Aa), u(aa)]

Entries are values of QTL variance for stated power and significance level given fixed sample size, and non-QTL baseline $\mathbf{R}$-squared, with total variance=1.0

| Sample | Significance Level |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Size | Power | $\mathbf{R}^{2}$ base | $\mathbf{0 . 0 1}$ | $\mathbf{0 . 0 0 1}$ | $\mathbf{0 . 0 0 0 1}$ |
| 1800 | 0.8 | 0.5 | 0.0039 | 0.0055 | 0.0070 |
| 1800 | 0.8 | 0.3 | 0.0054 | 0.0076 | 0.0098 |
| 1800 | 0.8 | 0.1 | 0.0069 | 0.0098 | 0.0126 |
| 1800 | 0.9 | 0.5 | 0.0048 | 0.0066 | 0.0083 |
| 1800 | 0.9 | 0.3 | 0.0068 | 0.0092 | 0.0116 |
| 1800 | 0.9 | 0.1 | 0.0087 | 0.0119 | 0.0149 |
| 1700 | 0.8 | 0.5 | 0.0041 | 0.0058 | 0.0074 |
| 1700 | 0.8 | 0.3 | 0.0057 | 0.0081 | 0.0104 |
| 1700 | 0.8 | 0.1 | 0.0074 | 0.0104 | 0.0133 |
| 1700 | 0.9 | 0.5 | 0.0051 | 0.0070 | 0.0087 |
| 1700 | 0.9 | 0.3 | 0.0072 | 0.0098 | 0.0122 |


| 1700 | 0.9 | 0.1 | 0.0092 | 0.0126 | 0.0157 |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 1600 | 0.8 | 0.5 | 0.0043 | 0.0061 | 0.0079 |
| 1600 | 0.8 | 0.3 | 0.0061 | 0.0086 | 0.0110 |
| 1600 | 0.8 | 0.1 | 0.0078 | 0.0110 | 0.0142 |
| 1600 | 0.9 | 0.5 | 0.0054 | 0.0074 | 0.0093 |
| 1600 | 0.9 | 0.3 | 0.0076 | 0.0104 | 0.0130 |
| 1600 | 0.9 | 0.1 | 0.0098 | 0.0133 | 0.0167 |
| 1500 | 0.8 | 0.5 | 0.0046 | 0.0065 | 0.0084 |
| 1500 | 0.8 | 0.3 | 0.0065 | 0.0092 | 0.0117 |
| 1500 | 0.8 | 0.1 | 0.0083 | 0.0118 | 0.0151 |
| 1500 | 0.9 | 0.5 | 0.0058 | 0.0079 | 0.0099 |
| 1500 | 0.9 | 0.3 | 0.0081 | 0.0111 | 0.0139 |
| 1500 | 0.9 | 0.1 | 0.0104 | 0.0142 | 0.0178 |
| 1400 | 0.8 | 0.5 | 0.0050 | 0.0070 | 0.0090 |
| 1400 | 0.8 | 0.3 | 0.0069 | 0.0098 | 0.0126 |
| 1400 | 0.8 | 0.1 | 0.0089 | 0.0126 | 0.0162 |
| 1400 | 0.9 | 0.5 | 0.0062 | 0.0085 | 0.0106 |
| 1400 | 0.9 | 0.3 | 0.0087 | 0.0119 | 0.0148 |
| 1400 | 0.9 | 0.1 | 0.0112 | 0.0152 | 0.0191 |
| 1300 | 0.8 | 0.5 | 0.0053 | 0.0075 | 0.0097 |
| 1300 | 0.8 | 0.3 | 0.0075 | 0.0106 | 0.0135 |
| 1300 | 0.8 | 0.1 | 0.0096 | 0.0136 | 0.0174 |
| 1300 | 0.9 | 0.5 | 0.0067 | 0.0091 | 0.0114 |
| 1300 | 0.9 | 0.3 | 0.0094 | 0.0128 | 0.0160 |
| 1300 | 0.9 | 0.1 | 0.0120 | 0.0164 | 0.0205 |
| 1200 | 0.8 | 0.5 | 0.0058 | 0.0082 | 0.0105 |
| 1200 | 0.8 | 0.3 | 0.0081 | 0.0114 | 0.0147 |
| 1200 | 0.8 | 0.1 | 0.0104 | 0.0147 | 0.0188 |
| 1200 | 0.9 | 0.5 | 0.0072 | 0.0099 | 0.0124 |
| 1200 | 0.9 | 0.3 | 0.0101 | 0.0138 | 0.0173 |
| 1200 | 0.9 | 0.1 | 0.0130 | 0.0178 | 0.0222 |
| 1100 | 0.8 | 0.5 | 0.0063 | 0.0089 | 0.0114 |
| 1100 | 0.8 | 0.3 | 0.0088 | 0.0125 | 0.0160 |
| 1100 | 0.8 | 0.1 | 0.0114 | 0.0160 | 0.0206 |
| 1100 | 0.9 | 0.5 | 0.0079 | 0.0108 | 0.0135 |
| 1100 | 0.9 | 0.3 | 0.0111 | 0.0151 | 0.0189 |
| 1100 | 0.9 | 0.1 | 0.0142 | 0.0194 | 0.0242 |
| 1000 | 0.8 | 0.5 | 0.0069 | 0.0098 | 0.0126 |
| 1000 | 0.8 | 0.3 | 0.0097 | 0.0137 | 0.0176 |
| 1000 | 0.8 | 0.1 | 0.0125 | 0.0177 | 0.0226 |
| 1000 | 0.9 | 0.5 | 0.0087 | 0.0118 | 0.0148 |
| 1000 | 0.9 | 0.3 | 0.0122 | 0.0166 | 0.0207 |
| 1000 | 0.9 | 0.1 | 0.0156 | 0.0213 | 0.0266 |
| 900 | 0.8 | 0.5 | 0.0077 | 0.0109 | 0.0139 |
| 900 | 0.8 | 0.3 | 0.0108 | 0.0152 | 0.0195 |
| 900 | 0.8 | 0.1 | 0.0139 | 0.0196 | 0.0251 |
| 900 | 0.9 | 0.5 | 0.0097 | 0.0131 | 0.0164 |
| 900 | 0.9 | 0.3 | 0.0135 | 0.0184 | 0.0230 |
|  |  |  |  |  |  |
|  | 0 | 0 | 0 |  |  |


| 900 | 0.9 | 0.1 | 0.0174 | 0.0236 | 0.0296 |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 800 | 0.8 | 0.5 | 0.0087 | 0.0122 | 0.0157 |
| 800 | 0.8 | 0.3 | 0.0122 | 0.0171 | 0.0219 |
| 800 | 0.8 | 0.1 | 0.0156 | 0.0220 | 0.0282 |
| 800 | 0.9 | 0.5 | 0.0109 | 0.0148 | 0.0185 |
| 800 | 0.9 | 0.3 | 0.0152 | 0.0207 | 0.0258 |
| 800 | 0.9 | 0.1 | 0.0195 | 0.0266 | 0.0332 |

Differences between SNP genotypes relate to $\Delta R^{2}$ through allele frequencies (let f denote minor allele frequency) - which determine expected genotype frequencies ( $g_{1}, g_{2}$ and $g_{3}$ ) - and through covariate-adjusted genotype-specific means ( $\mu_{1}, \mu_{2}$ and $\mu_{3}$ ). Without loss of generality, we assume trait values are standardized to variance 1 . Then $\Delta R^{2}=\Sigma g_{i}\left(\mu_{i}-\mu\right)^{2}$ where $\mu=\Sigma g_{i} \mu_{i}$ is the overall mean. With an additive genetic model, $\mu_{2}-\mu_{1}=\theta$ and $\mu_{3}-\mu_{1}=2 \theta$. Figure 1 displays the relation between the increment $\Delta \mathrm{R}^{2}$, the minor allele frequency, f , and the additive difference between means, $\theta$. For example, suppose we determine that we have power 0.80 to detect a SNP that contributes $2 \%$ of the variance $(\Delta=0.02)$. If the minor allele frequency is $\mathrm{f}=0.10$ then $\theta \approx 0.33$; note that $\theta$ is in standard deviation units. If the minor allele frequency is $\mathrm{f}=0.30$, then $\theta \approx 0.22$. In conjunction with the power table, this Figure helps us to determine which combinations of allele frequencies and genotype mean differences will be compatible with the value of $\Delta R^{2}$ from power calculations for specified power and significance level.

Figure 1. Delta R-square related to Allele Frequency Theta $=$ difference between genotype means


## Association Studies using Family-Based Association Test

Table 3 displays power calculated using PBAT for the family-based association test with the family plate set. Two sample sizes were evaluated: $\mathrm{n}=1398$ from 400 nuclear families using essentially all subjects on the plate set and $\mathrm{n}=1061$ from 291 nuclear families, a subset of the total. Power was computed for a range of QTL and marker minor allele frequencies. The QTL and the marker are assumed to be in linkage disequilibrium ( $\mathrm{D}^{\prime} \sim 1$ ). Thus, when the minor allele for the QTL is equal to that of the marker, it is assumed that the marker and the QTL are the same and power is at its maximum. Although the D' remains constant, as the differences between the minor alleles for the QTL and the marker increase, the LD correlation $\left(\mathrm{R}^{2}\right)$ and the study power decrease.

For example, if a study used 1400 subjects on the family plate set for a trait with allele frequency for the QTL of $5 \%$ and allele frequency of the nearby marker of $10 \%$, the power to detect a QTL variance (heritability) of $10 \%$ would be $52 \%$ for an alpha of 0.001 and $93 \%$ for an alpha of 0.05 . When the allele frequencies of the marker and the underlying functional variant of the QTL are the same, power in these examples increase to a maximum of $98 \%$ and $100 \%$.

Table 3: Power for FBAT analyses

| Table 3: Power for FBAT analyses <br> Allele <br> Freq for <br> Allele <br> Fample for <br> Fize <br> Fanctional <br> Variant |  |  |  |  | Typed <br> Marker |
| :---: | :---: | :---: | :---: | :---: | :---: |
| QTL H ${ }^{2}$ | Alpha | Power |  |  |  |
| 1400 | 5 | 5 | 5 | 0.001 | 66 |
|  |  |  | 5 | 0.05 | 98 |
|  |  |  | 10 | 0.001 | 98 |
|  |  |  | 10 | 0.05 | 100 |
| 1400 | 5 | 10 | 5 | 0.001 | 21 |
|  |  |  | 5 | 0.05 | 72 |
|  |  |  | 10 | 0.001 | 52 |
|  |  |  | 10 | 0.05 | 93 |
| 1400 | 5 | 20 | 5 | 0.001 | 6 |
|  |  |  | 5 | 0.05 | 41 |
|  |  |  | 10 | 0.001 | 16 |
|  |  |  | 10 | 0.05 | 65 |
| 1400 | 5 | 30 | 5 | 0.001 | 2 |
|  |  |  | 5 | 0.05 | 26 |
|  |  |  | 10 | 0.001 | 7 |
|  |  |  | 10 | 0.05 | 44 |
| 1400 | 10 | 5 | 5 | 0.001 | 24 |
|  |  |  | 5 | 0.05 | 77 |



| 1400 | 30 | 20 | 5 | 0.001 | 39 |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | 5 | 0.05 | 87 |
|  |  |  | 10 | 0.001 | 86 |
|  |  |  | 10 | 0.05 | 99 |
| 1400 | 30 | 30 | 5 | 0.001 | 75 |
|  |  |  | 5 | 0.05 | 98 |
|  |  |  | 10 | 0.001 | 99 |
|  |  |  | 10 | 0.05 | 100 |
| 1061 | 5 | 5 | 5 | 0.001 | 29 |
|  |  |  | 5 | 0.05 | 85 |
|  |  |  | 10 | 0.001 | 72 |
|  |  |  | 10 | 0.05 | 99 |
| 1061 | 5 | 10 | 5 | 0.001 | 8 |
|  |  |  | 5 | 0.05 | 50 |
|  |  |  | 10 | 0.001 | 22 |
|  |  |  | 10 | 0.05 | 76 |
| 1061 | 5 | 20 | 5 | 0.001 | 2 |
|  |  |  | 5 | 0.05 | 27 |
|  |  |  | 10 | 0.001 | 7 |
|  |  |  | 10 | 0.05 | 45 |
| 1061 | 5 | 30 | 5 | 0.001 | 1 |
|  |  |  | 5 | 0.05 | 18 |
|  |  |  | 10 | 0.001 | 3 |
|  |  |  | 10 | 0.05 | 30 |
| 1061 | 10 | 5 | 5 | 0.001 | 9 |
|  |  |  | 5 | 0.05 | 54 |
|  |  |  | 10 | 0.001 | 28 |
|  |  |  | 10 | 0.05 | 84 |
| 1061 | 10 | 10 | 5 | 0.001 | 36 |
|  |  |  | 5 | 0.05 | 87 |
|  |  |  | 10 | 0.001 | 82 |
|  |  |  | 10 | 0.05 | 100 |
| 1061 | 10 | 20 | 5 | 0.001 | 9 |
|  |  |  | 5 | 0.05 | 51 |
|  |  |  | 10 | 0.001 | 27 |
|  |  |  | 10 | 0.05 | 78 |
| 1061 | 10 | 30 | 5 | 0.001 | 4 |
|  |  |  | 5 | 0.05 | 33 |
|  |  |  | 10 | 0.001 | 11 |


|  |  |  | 10 | 0.05 | 55 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1061 | 20 | 5 | 5 | 0.001 | 2 |
|  |  |  | 5 | 0.05 | 28 |
|  |  |  | 10 | 0.001 | 7 |
|  |  |  | 10 | 0.05 | 50 |
| 1061 | 20 | 10 | 5 | 0.001 | 9 |
|  |  |  | 5 | 0.05 | 54 |
|  |  |  | 10 | 0.001 | 31 |
|  |  |  | 10 | 0.05 | 84 |
| 1061 | 20 | 20 | 5 | 0.001 | 40 |
|  |  |  | 5 | 0.05 | 88 |
|  |  |  | 10 | 0.001 | 86 |
|  |  |  | 10 | 0.05 | 100 |
| 1061 | 20 | 30 | 5 | 0.001 | 15 |
|  |  |  | 5 | 0.05 | 64 |
|  |  |  | 10 | 0.001 | 46 |
|  |  |  | 10 | 0.05 | 91 |
| 1061 | 30 | 5 | 5 | 0.001 | 1 |
|  |  |  | 5 | 0.05 | 19 |
|  |  |  | 10 | 0.001 | 3 |
|  |  |  | 10 | 0.05 | 33 |
| 1061 | 30 | 10 | 5 | 0.001 | 4 |
|  |  |  | 5 | 0.05 | 35 |
|  |  |  | 10 | 0.001 | 13 |
|  |  |  | 10 | 0.05 | 62 |
| 1061 | 30 | 20 | 5 | 0.001 | 17 |
|  |  |  | 5 | 0.05 | 66 |
|  |  |  | 10 | 0.001 | 53 |
|  |  |  | 10 | 0.05 | 93 |
| 1061 | 30 | 30 | 5 | 0.001 | 41 |
|  |  |  | 5 | 0.05 | 88 |
|  |  |  | 10 | 0.001 | 88 |
|  |  |  | 10 | 0.05 | 100 |

